The Random Version of Dvoretzky's Theorem in ℓ_{∞}^{n}

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Summary. We show that with "high probability" a section of the ℓ_{∞}^n ball of dimension $k \leq c\varepsilon \log n$ (c > 0 a universal constant) is ε close to a multiple of the Euclidean ball in this section. We also show that, up to an absolute constant the estimate on k cannot be improved.

1 Introduction

Milman's version of Dvoretzky's theorem states that:

There is a function $c(\varepsilon) > 0$ such that for all $k \le c(\varepsilon) \log n$, $\ell_2^k (1 + \varepsilon)$ -embeds into any normed space of dimension n.

See [Dv] for the original theorem of Dvoretzky (in which the dependence of k on n is weaker), [Mi] for Milman's original work, and [MS] and [Pi] for expository outlets of the subject (there are many others). It would be important for us to notice that the proof(s) of the theorem above actually give more: The vast majority of subspaces of the stated dimension are $(1+\varepsilon)$ -isomorphic to ℓ_2^k .

The dependence of k on n in the theorem above is known to be best possible (for ℓ_{∞}^n) but the dependence on ε is far from being understood. The best known estimate is $c(\varepsilon) \geq c\varepsilon/(\log\frac{1}{\varepsilon})^2$ given in [Sc] (here and elsewhere in this paper c and C denote positive universal constants). However, the proof in [Sc] does not give the additional information that most subspaces are $(1+\varepsilon)$ -isomorphic to ℓ_2^k . If one also want this requirement then the best estimate for $c(\varepsilon)$ that was known was $c(\varepsilon) \geq c\varepsilon^2$ ([Go]).

As an upper bound for $c(\varepsilon)$ one gets $C/\log\frac{1}{\varepsilon}$ for some universal C. Indeed, if ℓ_2^k $(1+\varepsilon)$ embed into ℓ_∞^n then $k \leq C\log n/\log\frac{1}{\varepsilon}$. This is also the right order of k in the ℓ_∞ case: If $k \leq c\log n/\log\frac{1}{\varepsilon}$ then ℓ_2^k $(1+\varepsilon)$ embed into ℓ_∞^n .

We show here that, in the ℓ_{∞} case, if one is interested in the probabilistic statement of Dvoretzky theorem (i.e, that the vast majority of subspaces of

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 ℓ_{∞}^{n} of a certain dimension are $(1+\varepsilon)$ -isomorphic to Euclidean spaces) then the right estimate for $c(\varepsilon)$ is $c\varepsilon$.

Theorem 1. For $k < c\varepsilon \log n$, with probability $> 1 - e^{-ck}$, the ℓ_{∞}^n norm and a multiple of the ℓ_2^n norm are $1 + \varepsilon$ equivalent on a k dimensional subspace. Moreover, this doesn't hold anymore for k of higher order. i.e., For every a there is an A such that if, with probability larger than $1 - e^{-ak}$, a k dimensional subspace satisfies that the ratio between the ℓ_{∞}^n norm and a multiple of the ℓ_2^n norm are $1 + \varepsilon$ equivalent for all vectors in the subspace, then $k \leq A\varepsilon \log n$.

2 Computation of the Concentration of the Max Norm

Let g_1, g_2, \ldots be a sequence of standard independent Gaussian variables. fix n and let M be the median of $\|(g_1, g_2, \ldots, g_n)\|_{\infty}$. In this section we compute some fine estimates on the probability of deviation of $\|(g_1, g_2, \ldots, g_n)\|_{\infty}$ from M.

Claim 1.

$$(1 - 2^{-1/n}) \frac{\sqrt{\pi}M}{\sqrt{2}} \le e^{-M^2/2} \le (1 - 2^{-1/n}) \frac{\sqrt{\pi}(M+1)}{\sqrt{2}(1 - e^{-\frac{1}{2}}e^{-M})}.$$
 (1)

Proof.

$$\frac{1}{2} = P\left(\max_{1 \le i \le n} |g_i| < M\right) = \left(1 - \sqrt{\frac{2}{\pi}} \int_M^\infty e^{-s^2/2} ds\right)^n.$$

Consequently,

$$1 - 2^{-1/n} = \sqrt{\frac{2}{\pi}} \int_{M}^{\infty} e^{-s^{2}/2} ds \ge \sqrt{\frac{2}{\pi}} \frac{1}{M+1} \int_{M}^{M+1} s e^{-s^{2}/2} ds$$

$$\ge \sqrt{\frac{2}{\pi}} \frac{1}{M+1} e^{-M^{2}/2} (1 - e^{-\frac{1}{2}} e^{-M}),$$
(2)

or

$$e^{-M^2/2} \le (1 - 2^{-1/n}) \frac{\sqrt{\pi}(M+1)}{\sqrt{2}(1 - e^{-\frac{1}{2}}e^{-M})}.$$
 (3)

Similarly,

$$1 - 2^{-1/n} = \sqrt{\frac{2}{\pi}} \int_{M}^{\infty} e^{-s^2/2} ds \le \sqrt{\frac{2}{\pi}} \frac{1}{M} \int_{M}^{\infty} s e^{-s^2/2} ds \le \sqrt{\frac{2}{\pi}} \frac{e^{-M^2/2}}{M},$$

or

$$e^{-M^2/2} \ge (1 - 2^{-1/n}) \frac{\sqrt{\pi}M}{\sqrt{2}}.$$
 (4)

Claim 2.

$$\frac{\log 2}{4 + \log 2} e^{-3\varepsilon M^2/2} \le P\left(\max_{1 \le i \le n} |g_i| > (1 + \varepsilon)M\right) \le \log 2(1 + o(1)) e^{-\varepsilon M^2}$$
 (5)

where o(1) means a(n) with $a(n) \to 0$ as $n \to \infty$ independently of ε .

Proof. (3) implies

$$P\left(\max_{1\leq i\leq n}|g_{i}| > (1+\varepsilon)M\right)$$

$$\leq \sqrt{\frac{2}{\pi}} \frac{n}{(1+\varepsilon)M} e^{-(1+\varepsilon)^{2}M^{2}/2}$$

$$\leq \frac{n}{(1+\varepsilon)M} (1-2^{-1/n}) \frac{M+1}{1-e^{-\frac{1}{2}}e^{-M}} e^{-\varepsilon M^{2}} e^{-\varepsilon^{2}M^{2}/2}$$
(6)

and, since M is of order $\sqrt{\log n}$, we get from this that

$$P\left(\max_{1 \le i \le n} |g_i| > (1+\varepsilon)M\right) \le \log 2(1+o(1))e^{-\varepsilon M^2}.$$
 (7)

(For a fixed ε one can replace $\log 2(1+o(1))$ with a quantity tending to 0 with n.)

We now look for a lower bound on $P(\max_{1 \le i \le n} |g_i| > (1+\varepsilon)M)$. Since for iid X_i -s,

$$P\left(\max_{1 \le i \le n} (X_i > t) = 1 - (1 - P(X_1 > t))^n \ge 1 - e^{-nP(X_1 > t)}$$

$$\ge 1 - \frac{1}{1 + nP(X_1 > t)} = \frac{nP(X_1 > t)}{1 + nP(X_1 > t)},$$
(8)

$$P\left(\max_{1\leq i\leq n}|g_i| > (1+\varepsilon)M\right) \geq \frac{nP(|g_1| > (1+\varepsilon)M)}{1+nP(|g_1| > (1+\varepsilon)M)}.$$
 (9)

The right hand side is an increasing function of $P(|g_1| > (1 + \varepsilon)M)$ and, by (4),

$$P(|g_{1}| > (1+\varepsilon)M)$$

$$\geq \sqrt{\frac{2}{\pi}} \frac{1}{(1+\varepsilon)M+1} e^{-(1+\varepsilon)^{2}M^{2}/2} \left(1 - e^{-\frac{1}{2}} e^{-(1+\varepsilon)M}\right)$$

$$= \sqrt{\frac{2}{\pi}} \frac{1}{(1+\varepsilon)M+1} e^{-M^{2}/2} e^{-\varepsilon M^{2} - \varepsilon^{2}M^{2}/2} \left(1 - e^{-\frac{1}{2}} e^{-(1+\varepsilon)M}\right)$$

$$\geq \frac{M(1-2^{-1/n})}{(1+\varepsilon)M+1} e^{-\varepsilon M^{2} - \varepsilon^{2}M^{2}/2} \left(1 - e^{-\frac{1}{2}} e^{-(1+\varepsilon)M}\right)$$

$$\geq \frac{\log 2}{4n} e^{-\varepsilon M^{2} - \varepsilon^{2}M^{2}/2} \geq \frac{\log 2}{4n} e^{-3\varepsilon M^{2}/2}, \tag{10}$$

for $\varepsilon \leq 1$ and n large enough (independently of ε). Using (9), we get

$$P\left(\max_{1 \le i \le n} |g_i| > (1+\varepsilon)M\right) \ge \frac{\frac{\log 2}{4}e^{-3\varepsilon M^2/2}}{1 + \frac{\log 2}{4}e^{-3\varepsilon M^2/2}} \ge \frac{\log 2}{4 + \log 2}e^{-3\varepsilon M^2/2}.$$
(11)

Claim 3. For some absolute positive constants c, C and for all $0 < \varepsilon < 1/2$,

$$\exp\left(-Ce^{\varepsilon M^2}\right) \le P\left(\max_{1 \le i \le n} |g_i| < (1-\varepsilon)M\right) \le C\exp\left(-ce^{3\varepsilon M^2/4}\right). \tag{12}$$

Proof.

$$P\left(\max_{1 \le i \le n} |g_i| < (1 - \varepsilon)M\right)$$

$$= \left(1 - \sqrt{\frac{2}{\pi}} \int_{(1 - \varepsilon)M}^{\infty} e^{-s^2/2}\right)^n \le \left(1 - \sqrt{\frac{2}{\pi}} \int_{(1 - \varepsilon)M}^{M} e^{-s^2/2}\right)^n$$

$$\le \left(1 - \sqrt{\frac{2}{\pi}} \frac{1}{M} \int_{(1 - \varepsilon)M}^{M} se^{-s^2/2}\right)^n$$

$$= \left(1 - \sqrt{\frac{2}{\pi}} \frac{1}{M} \left(e^{-(1 - \varepsilon)^2 M^2/2} - e^{-M^2/2}\right)\right)^n$$

$$= \left(1 - \sqrt{\frac{2}{\pi}} \frac{1}{M} e^{-M^2/2} \left(e^{\varepsilon M^2 - \varepsilon^2 M^2/2} - 1\right)\right)^n$$

$$\le \left(1 - (1 - 2^{-1/n}) \left(e^{\varepsilon M^2 - \varepsilon^2 M^2/2} - 1\right)\right)^n \text{ by (4)}$$

$$\le \exp\left(-n(1 - 2^{-1/n}) \left(e^{\varepsilon M^2 - \varepsilon^2 M^2/2} - 1\right)\right)$$

$$\le 2(1 + o(1)) \exp\left(-\log 2(1 + o(1)) e^{3\varepsilon M^2/4}\right)$$

which proves the right hand side inequality in (12). As for the left hand side,

$$P\left(\max_{1 \le i \le n} |g_{i}| < (1 - \varepsilon)M\right)$$

$$= \left(1 - \sqrt{\frac{2}{\pi}} \int_{(1 - \varepsilon)M}^{\infty} e^{-s^{2}/2}\right)^{n} \ge \left(1 - \sqrt{\frac{2}{\pi}} \frac{1}{(1 - \varepsilon)M} \int_{(1 - \varepsilon)M}^{\infty} se^{-s^{2}/2}\right)^{n}$$

$$= \left(1 - \sqrt{\frac{2}{\pi}} \frac{1}{(1 - \varepsilon)M} (e^{-(1 - \varepsilon)^{2}M^{2}/2}\right)^{n} = \left(1 - \sqrt{\frac{2}{\pi}} \frac{e^{-M^{2}/2}}{(1 - \varepsilon)M} e^{\varepsilon M^{2} - \varepsilon^{2}M^{2}/2}\right)^{n}$$

$$\ge \exp\left(-\frac{2n(1 - 2^{-1/n})}{1 - e^{-\frac{1}{2} - M}} \frac{M + 1}{M} e^{\varepsilon M^{2} - \varepsilon^{2}M^{2}/2}\right) \text{ by (3)}$$

$$\ge \exp\left(-2\log 2(1 + o(1))e^{\varepsilon M^{2}}\right).$$

We summarize Claims 2 and 3 in a form that will be useful for us later in the following Proposition.

Proposition 1. For some positive absolute constants c, C and for all $0 < \varepsilon < 1$ and $n \in \mathbb{N}$, denoting $g = (g_1, g_2, \dots, g_n)$,

$$ce^{-C\varepsilon\log n} \le P\left(\|g\|_{\infty} < \frac{(1-\varepsilon)M}{\sqrt{n}}\|g\|_2 \text{ or } \|g\|_{\infty} > \frac{(1+\varepsilon)M}{\sqrt{n}}\|g\|_2\right)$$

$$\le Ce^{-c\varepsilon\log n}.$$

Proof. This follows easily from Claims 2 and 3 and the facts that $e^x > x$ for all x, M is of order $\sqrt{\log n}$ and

$$P(\|g\|_2 < (1-\varepsilon)\sqrt{n} \text{ or } \|g\|_2 > (1+\varepsilon)\sqrt{n}) < Ce^{-\varepsilon^2 n}.$$

3 Proof of the Theorem

The first part of the Theorem follows easily from the, by now well exposed, proof of Milman's version of Dvoretzky's theorem (see e.g, [MS] or [Pi]) with the improved concentration estimate in (the right hand side of the inequality in) Proposition 1 replacing the classical estimates. For the proof of the second part we need:

Lemma 1. Let \mathcal{A} be a subset of $G_{n,k}$ of $\mu_{n,k}$ measure a. Put $U_{\mathcal{A}} = \bigcup_{E \in \mathcal{A}} E$, then

$$P((g_1, g_2, \dots, g_n) \in U_{\mathcal{A}}) \ge a^{1/k}.$$

Proof. Let X_1, X_2, \ldots, X_k be k independent random vectors distributed according to P, the canonical Gaussian measure on \mathbb{R}^n . Note that, since $\mu_{n,k}$ is the unique rotational invariant probability measure on $G_{n,k}$, the distribution of span $\{X_1, \ldots, X_k\}$ is $\mu_{n,k}$. Accordingly,

$$P(U_{\mathcal{A}})^k = P(X_1, X_2, \dots, X_k \in U_{\mathcal{A}})$$

$$\geq P(\operatorname{span}\{X_1, X_2, \dots, X_k\} \in \mathcal{A})$$

$$= \mu_{n,k}(\mathcal{A}).$$

Remark 1. As we'll see below we use only a weak form of Lemma 1. We actually believe there is a much stronger form of it.

Proof of the moreover part in Theorem 1. Let $A \subset G_{n,k}$ be such that every $E \in A$ there is an M_E such that

$$M_E ||x||_2 \le ||x||_\infty \le (1+\varepsilon)M_E ||x||_2$$

for all $x \in E$. Let \mathcal{B} be the subset of \mathcal{A} of all E for which $\frac{(1-3\varepsilon)M}{\sqrt{n}} \leq M_E \leq \frac{(1+\varepsilon)M}{\sqrt{n}}$, and let $\mathcal{C} = \mathcal{A} \setminus \mathcal{B}$. By Lemma 1,

$$\mu_{n,k}(\mathcal{C})^{1/k} \le P\left(\left\{x; \ \|x\|_{\infty} < \frac{(1+\varepsilon)(1-3\varepsilon)M}{\sqrt{n}} \|x\|_{2} \text{ or } \|x\|_{\infty} > \frac{(1+\varepsilon)M}{\sqrt{n}} \|x\|_{2}\right\}\right)$$

and, by Proposition 1, this last quantity is smaller than $Ce^{-c\varepsilon \log n}$. It follows that

$$\mu_{n,k}(\mathcal{B}) > 1 - e^{-ak} - Ce^{-c\varepsilon k \log n}$$
.

We may assume that $\varepsilon \log n$ is much larger than a so that the last term above is dominated by e^{-ak} . Applying Lemma 1 once more we get

$$P\left(\left\{x; \ (\frac{(1-3\varepsilon)M}{\sqrt{n}} \|x\|_2 \le \|x\|_{\infty} \le \frac{(1+\varepsilon)^2 M}{\sqrt{n}} \|x\|_2\right\}\right) \ge \mu_{n,k}(\mathcal{B}) > 1 - 2e^{-ak}.$$

Using now the other part of Proposition 1 we get that

$$C\varepsilon \log n > ak$$
.

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