

A

Appendix

A.1 Summary of Probability Distributions

Here we briefly summarize all probability distributions used in this book. For an exhaustive review of probability distributions see Johnson et al. (1993, 1994, 1995). The parameterization of the densities closely follows Bernardo and Smith (1994).

Notation

Y denotes a univariate random variable; y refers to a realization of Y . $f(y)$ refers to the density of the probability distribution of Y with respect to an appropriate measure (Lebesgue measure or counting measure, depending on the context). $E(Y)$ denotes the expectation of Y , whereas $\text{Var}(Y)$ denotes the variance of Y . For multivariate random variables \mathbf{Y} and \mathbf{y} are used, whereas Y_j and y_j refer to a certain element of \mathbf{Y} and \mathbf{y} , respectively. $E(\mathbf{Y})$ denotes the mean vector; $\text{Var}(Y)$ denotes the variance–covariance matrix of \mathbf{Y} .

A.1.1 The Beta Distribution

The Beta distribution $Y \sim \mathcal{B}(\alpha, \beta)$ with $\alpha, \beta \in \mathbb{R}^+$, is a univariate distribution defined on the unit interval $y \in [0, 1]$. For mixture models the Beta distribution appears mainly as a posterior distribution of an unknown probability. Density, mean, and variance are given by

$$f_B(y; \alpha, \beta) = \frac{1}{B(\alpha, \beta)} y^{\alpha-1} (1-y)^{\beta-1}, \quad (\text{A.1})$$
$$E(Y) = \frac{\alpha}{\alpha + \beta}, \quad \text{Var}(Y) = \frac{\alpha\beta}{(\alpha + \beta)^2(\alpha + \beta + 1)},$$

where $B(\alpha, \beta)$ is the Beta function:

$$B(\alpha, \beta) = \frac{\Gamma(\alpha)\Gamma(\beta)}{\Gamma(\alpha + \beta)}. \quad (\text{A.2})$$

For a proper density both α and β need to be positive. If $0 < \alpha < 1$, the density is unbounded at 0; if $0 < \beta < 1$, the density is unbounded at 1. If $\alpha = 1$ and $\beta = 1$, the density is equal to the density of a uniform distribution. If $\alpha = 1$ and $\beta > 1$, the mode of the density lies at 0; if $\alpha > 1$ and $\beta = 1$, the mode of the density lies at 1. For $\alpha, \beta > 1$ the mode of the density lies in the interior of $[0, 1]$ at $(\alpha - 1)/(\alpha + \beta - 2)$.

A.1.2 The Binomial Distribution

The binomial distribution, $Y \sim \text{BiNom}(n, p)$, with $n = 1, 2, \dots$ and $p \in [0, 1]$, is frequently chosen to model the outcome of repeated measurements. The density is defined for $y \in \{0, 1, 2, \dots, n\}$,

$$f_{BN}(y; n, p) = \binom{n}{y} p^y (1-p)^{n-y}, \quad (\text{A.3})$$

and mean and variance are given by

$$E(Y) = np, \quad \text{Var}(Y) = np(1-p). \quad (\text{A.4})$$

A.1.3 The Dirichlet Distribution

The Dirichlet distribution, $\mathbf{Y} \sim \mathcal{D}(\alpha_1, \dots, \alpha_K)$, is a standard choice in the context of modeling an unknown discrete probability distribution $\mathbf{Y} = (Y_1, \dots, Y_K)$ where $\sum_{j=1}^K Y_j = 1$ and therefore is of great importance for mixture and switching models. The Dirichlet distribution is a distribution on the unit simplex $\mathcal{E}_K \subset (\mathfrak{R}^+)^K$, defined by the following constraint,

$$\mathcal{E}_K = \left\{ \mathbf{y} = (y_1, \dots, y_K) \in (\mathfrak{R}^+)^K : \sum_{j=1}^K y_j = 1 \right\}.$$

The density is given by

$$f_D(\mathbf{y}; \alpha_1, \dots, \alpha_K) = y_1^{\alpha_1-1} \dots y_K^{\alpha_K-1} c, \quad c = \frac{\Gamma(\Sigma_\alpha)}{\prod_{i=1}^K \Gamma(\alpha_i)}, \quad (\text{A.5})$$

where $\Sigma_\alpha = \sum_{j=1}^K \alpha_j$. As each y_j is completely determined given the remaining elements of \mathbf{y} , one element y_j has to be substituted in (A.5) by $1 - \sum_{k \neq j} y_k$ to obtain a proper density. The density is proper if all α_j are positive; the density is bounded if all α_j are greater than or equal to 1. The density of a $\mathcal{D}(\alpha_1, \dots, \alpha_K)$ -distribution is improper whenever any α_j is equal to 0.

For $\alpha_1 = \dots = \alpha_K = 1$ the uniform distribution over the unit simplex results. For $K = 2$ the Dirichlet distribution is equal to $Y_1 \sim \mathcal{B}(\alpha_1, \alpha_2)$.

The marginal distribution of Y_j is a $\mathcal{B}(\alpha_j, \Sigma_\alpha - \alpha_j)$ -distribution, therefore mean and variance of Y_j are given by

$$E(Y_j) = \frac{\alpha_j}{\Sigma_\alpha}, \quad \text{Var}(Y_j) = \frac{\alpha_j(\Sigma_\alpha - \alpha_j)}{\Sigma_\alpha^2(\Sigma_\alpha + 1)}. \quad (\text{A.6})$$

The mode of the marginal density of Y_j is given by $(\alpha_j - 1)/(\Sigma_\alpha - K)$.

The easiest way to sample a random variable $\mathbf{Y} = (Y_1, \dots, Y_K)$ from the $\mathcal{D}(\alpha_1, \dots, \alpha_K)$ -distribution is to sample K independent random variables Y_1^*, \dots, Y_K^* from the following Gamma distributions, $Y_j^* \sim \mathcal{G}(\alpha_j, 1)$, $j = 1, \dots, K$, and to normalize: $Y_j = Y_j^* / \sum_{k=1}^K Y_k^*$.

A.1.4 The Exponential Distribution

The exponential distribution, $Y \sim \mathcal{E}(\beta)$ with $\beta > 0$, is a univariate distribution defined on the positive real line $y \geq 0$ and is mainly used as a sampling distribution in the context of finite mixture models. There exist different ways to parameterize this distribution and we follow Bernardo and Smith (1994), by defining the density as

$$f_E(y; \beta) = \beta e^{-\beta y}. \quad (\text{A.7})$$

Mean and variance are given by

$$E(Y) = \text{Var}(Y) = 1/\beta. \quad (\text{A.8})$$

In this parameterization, the exponential distribution is equal to the $\mathcal{G}(1, \beta)$ -distribution.

A.1.5 The F-Distribution

The F-distribution, $Y \sim F(\alpha_1, \alpha_2)$, is a univariate distribution defined on the positive real line $y \geq 0$. For finite mixture models based on the normal distribution, it appears as part of modeling the prior of heterogeneous variances. For $\alpha_1 > 0, \alpha_2 > 0$ the density is given by

$$f_F(y; \alpha_1, \alpha_2) = \frac{\Gamma((\alpha_1 + \alpha_2)/2)\alpha_1}{\Gamma(\alpha_1/2)\Gamma(\alpha_2/2)\alpha_2} \left(\frac{\alpha_1}{\alpha_2}y\right)^{\alpha_1/2-1} \left(1 + \frac{\alpha_1}{\alpha_2}y\right)^{-(\alpha_1+\alpha_2)/2} \quad (\text{A.9})$$

For $y \rightarrow \infty$, the density behaves as $y^{-\alpha_2/2-1}$, whereas for $y \rightarrow 0$, the density behaves as $y^{\alpha_1/2-1}$. The density is an improper density, whenever α_1 or α_2 are 0. If $\alpha_1 > 0$ and $\alpha_2 > 0$, the density is proper, but unbounded at $y = 0$ for $0 < \alpha_1 < 2$.

Relation to the χ^2 -Distribution

Assume that Y_1 and Y_2 are independent random variables, each following a χ^2 -distribution: $Y_1 \sim \chi_{\alpha_1}^2$ and $Y_2 \sim \chi_{\alpha_2}^2$. Then

$$Y = \frac{Y_1/\alpha_1}{Y_2/\alpha_2} \sim F(\alpha_1, \alpha_2). \quad (\text{A.10})$$

A.1.6 The Gamma Distribution

The Gamma distribution, $Y \sim \mathcal{G}(\alpha, \beta)$, is a univariate distribution defined on the positive real line $y \geq 0$. It is encountered in finite mixture models as a posterior density within a Bayesian analysis for certain nonnormal models, in particular for observations from the Poisson and the exponential distribution. There exist various ways to parameterize this distribution and we follow Bernardo and Smith (1994), where the density is given by

$$f_G(y; \alpha, \beta) = \frac{\beta^\alpha}{\Gamma(\alpha)} y^{\alpha-1} e^{-\beta y}. \quad (\text{A.11})$$

The density is an improper density whenever $\beta = 0$. If $\beta > 0$, the density is proper, but unbounded at $y = 0$ for $0 < \alpha < 1$. For $\alpha > 0, \beta > 0$, mean and variance are given by

$$E(Y) = \frac{\alpha}{\beta}, \quad \text{Var}(Y) = \frac{\alpha}{\beta^2}. \quad (\text{A.12})$$

The mode is given by $(\alpha - 1)/\beta$. If $Y \sim \mathcal{G}(\alpha, \beta)$, then $\omega Y \sim \mathcal{G}(\alpha, \beta/\omega)$.

The χ^2 -Distribution

The $\mathcal{G}(\nu/2, 1/2)$ -distribution with $\nu = 1, 2, \dots$ is called the χ_ν^2 -distribution, which is the distribution of the sum of squares of ν independent standard normal random variables, $Y = \sum_{j=1}^{\nu} Y_j^2$, $Y_j \sim \mathcal{N}(0, 1)$. Therefore $E(Y) = \nu$, $\text{Var}(Y) = 2\nu$.

Inverted Gamma Distribution

A random variable Y follows an inverted Gamma distribution, $Y \sim \mathcal{G}^{-1}(\alpha, \beta)$, if Y^{-1} has a Gamma distribution: $Y^{-1} \sim \mathcal{G}(\alpha, \beta)$. The inverted Gamma distribution often appears as a posterior distribution of an unknown variance within a Bayesian analysis of finite mixture models based on the normal distribution. The density is given by

$$f_{IG}(y; \alpha, \beta) = \frac{\beta^\alpha}{\Gamma(\alpha)} \left(\frac{1}{y}\right)^{\alpha+1} e^{-\beta/y}. \quad (\text{A.13})$$

Y has finite expectation, iff $\alpha > 1$, and finite variance, iff $\alpha > 2$:

$$E(Y) = \frac{\beta}{\alpha - 1}, \quad \text{Var}(Y) = \frac{\beta^2}{(\alpha - 1)^2(\alpha - 2)}. \quad (\text{A.14})$$

The mode is given by $\beta/(\alpha + 1)$. $\mathcal{G}^{-1}(0, 0)$ has an improper density proportional to $1/y$. If $Y \sim \mathcal{G}^{-1}(\alpha, \beta)$, then $\omega Y \sim \mathcal{G}^{-1}(\alpha, \omega\beta)$.

A.1.7 The Geometric Distribution

The geometric distribution is defined for $y \in \{1, 2, \dots\}$. Density, mean and variance are given by

$$f(y; p) = p(1 - p)^{y-1}, \quad (\text{A.15})$$

$$E(Y) = \frac{1}{p}, \quad \text{Var}(Y) = \frac{1 - p}{p^2}.$$

A.1.8 The Multinomial Distribution

Let \mathbf{Y} be a categorical random variable with D categories coded as 1 to D . $\mathbf{Y} \sim \text{MulNom}(n, p_1, \dots, p_D)$, iff for all $\mathbf{y} = (y_1, \dots, y_D)$, where $\sum_{j=1}^D y_j = n$ and $y_j \geq 0$,

$$\Pr(\mathbf{Y} = \mathbf{y}) = f_{MN}(\mathbf{y}; n, p_1, \dots, p_D) = \binom{n}{y_1 \dots y_D} p_1^{y_1} \dots p_D^{y_D}, \quad (\text{A.16})$$

with the general binomial coefficient being defined as:

$$\binom{n}{y_1 \dots y_D} = \frac{n!}{y_1! \dots y_D!}$$

The mean and variance of Y_j are given by

$$E(Y_j) = np_j, \quad \text{Var}(Y_j) = np_j(1 - p_j). \quad (\text{A.17})$$

The binomial distribution is that special case of the multinomial distribution where $D = 2$.

A.1.9 The Negative Binomial Distribution

The negative binomial distribution is defined for $y \in \{0, 1, 2, \dots\}$: $Y \sim \text{NegBin}(\alpha, \beta)$. For $\alpha > 0, \beta > 0$ density, mean, and variance of Y are given by

$$f_{NB}(y; \alpha, \beta) = \binom{\alpha + y - 1}{\alpha - 1} \left(\frac{\beta}{\beta + 1}\right)^\alpha \left(\frac{1}{\beta + 1}\right)^y,$$

$$E(Y) = \frac{\alpha}{\beta}, \quad \text{Var}(Y) = \frac{\alpha}{\beta^2}(\beta + 1). \quad (\text{A.18})$$

The $\text{NegBin}(\alpha, \beta)$ -distribution is an infinite mixture of $\mathcal{P}(\mu)$ -distributions, where $\mu \sim \mathcal{G}(\alpha, \beta)$:

$$f_{NB}(y; \alpha, \beta) = \int f_P(y; \mu) f_G(\mu; \alpha, \beta) d\mu.$$

A.1.10 The Normal Distribution

The normal distribution is the most important density in finite mixture modeling, both as a common choice as sampling density as well as posterior density in a Bayesian analysis of such models.

The Univariate Normal Distribution

The univariate normal distribution, $Y \sim \mathcal{N}(\mu, \sigma^2)$, with $\mu \in \Re$ and $\sigma > 0$, is defined on \Re . Density, mean, and variance of Y are given by

$$f_N(y; \mu, \sigma^2) = \frac{1}{\sqrt{2\pi\sigma^2}} e^{-(y-\mu)^2/(2\sigma^2)},$$

$$E(Y) = \mu, \quad \text{Var}(Y) = \sigma^2. \quad (\text{A.19})$$

$1/\sigma^2$ is often called precision. A normal distribution with zero mean and zero precision has an improper density proportional to a constant. Higher-order moments $E((Y - \mu)^m)$ are zero for m odd, otherwise:

$$E((Y - \mu)^m) = \sigma^{2m} \prod_{n=1}^{m/2} (2n - 1). \quad (\text{A.20})$$

The Multivariate Normal Distribution

The multivariate normal distribution, $\mathbf{Y} \sim \mathcal{N}_d(\boldsymbol{\mu}, \boldsymbol{\Sigma})$, where $\boldsymbol{\mu} \in \Re^d$ and $\boldsymbol{\Sigma} \in \Re^{d \times d}$ is a symmetric, positive definite matrix, is defined for $\mathbf{y} \in \Re^d$. Density, mean, and covariance matrix of \mathbf{Y} are given by

$$f_N(\mathbf{y}; \boldsymbol{\mu}, \boldsymbol{\Sigma}) = (2\pi)^{-d/2} |\boldsymbol{\Sigma}|^{-1/2} \exp\left(-\frac{1}{2}(\mathbf{y} - \boldsymbol{\mu})' \boldsymbol{\Sigma}^{-1}(\mathbf{y} - \boldsymbol{\mu})\right),$$

$$E(\mathbf{Y}) = \boldsymbol{\mu}, \quad \text{Var}(\mathbf{Y}) = \boldsymbol{\Sigma}. \quad (\text{A.21})$$

$\boldsymbol{\Sigma}^{-1}$ is the information matrix. A normal distribution with $\boldsymbol{\mu} = \mathbf{0}$ and $\boldsymbol{\Sigma}^{-1} = \mathbf{0}$ has an improper density proportional to a constant.

To sample from a multivariate normal distribution the following result is useful. Let $\mathbf{Z} = (Z_1 \dots Z_d)'$ be d independent copies of an $\mathcal{N}(0, 1)$ -distribution and let $\boldsymbol{\Sigma} = \mathbf{A}\mathbf{A}'$ be the Cholesky decomposition of $\boldsymbol{\Sigma}$. Then $\mathbf{Y} = \boldsymbol{\mu} + \mathbf{A}\mathbf{Z} \sim \mathcal{N}_d(\boldsymbol{\mu}, \boldsymbol{\Sigma})$.

Marginal and Conditional Distributions

The quadratic form $(\mathbf{Y} - \boldsymbol{\mu})' \boldsymbol{\Sigma}^{-1}(\mathbf{Y} - \boldsymbol{\mu}) \sim \chi_d^2$. Assume that \mathbf{Y} is divided into two blocks, \mathbf{Y}_1 containing the first d_1 components of \mathbf{Y} , and \mathbf{Y}_2 containing the remaining $d_2 = d - d_1$ components. Apply a similar partition on $\boldsymbol{\mu}$ and $\boldsymbol{\Sigma}$:

$$\boldsymbol{\mu} = \begin{pmatrix} \mu_1 \\ \mu_2 \end{pmatrix}, \quad \boldsymbol{\Sigma} = \begin{pmatrix} \boldsymbol{\Sigma}_{11} & \boldsymbol{\Sigma}_{12} \\ \boldsymbol{\Sigma}_{21} & \boldsymbol{\Sigma}_{22} \end{pmatrix}.$$

Then, marginally, $\mathbf{Y}_2 \sim \mathcal{N}_{d_2}(\boldsymbol{\mu}_2, \boldsymbol{\Sigma}_{22})$ and the conditional distribution of $\mathbf{Y}_1|\mathbf{Y}_2$ is $\mathcal{N}_{d_1}(\boldsymbol{\mu}_{1|2}, \boldsymbol{\Sigma}_{11|2})$, where

$$\begin{aligned} \boldsymbol{\mu}_{1|2} &= \boldsymbol{\mu}_1 + \boldsymbol{\Sigma}_{12}\boldsymbol{\Sigma}_{22}^{-1}(\mathbf{Y}_2 - \boldsymbol{\mu}_2), \\ \boldsymbol{\Sigma}_{11|2} &= \boldsymbol{\Sigma}_{11} - \boldsymbol{\Sigma}_{12}\boldsymbol{\Sigma}_{22}^{-1}\boldsymbol{\Sigma}'_{12}. \end{aligned}$$

The Normal Gamma Distribution

A pair (\mathbf{Y}_1, Y_2) of random variables \mathbf{Y}_1 and Y_2 with $\mathbf{Y}_1|Y_2 \sim \mathcal{N}_d(\boldsymbol{\mu}, Y_2\boldsymbol{\Sigma})$ and $1/Y_2 \sim \mathcal{G}(\nu/2, \nu/2)$ follows the so-called normal Gamma distribution. The marginal distribution of \mathbf{Y}_1 is the d -variate $t_\nu(\boldsymbol{\mu}, \boldsymbol{\Sigma})$ -distribution, whereas the marginal distribution of Y_2 is $\mathcal{G}^{-1}(\nu/2, \nu/2)$.

A.1.11 The Poisson Distribution

The Poisson distribution, $Y \sim \mathcal{P}(\mu)$ with $\mu > 0$, is defined for $y \in \{0, 1, 2, \dots\}$ making it a standard distribution to model a random count variable. Density, mean, and variance are given by

$$\begin{aligned} f_P(y; \mu) &= \frac{\mu^y}{y!} e^{-\mu}, \\ E(Y) &= \text{Var}(Y) = \mu. \end{aligned} \tag{A.22}$$

A.1.12 The Student- t Distribution

The Student- t distribution is useful as a sampling distribution for robust modeling.

The Univariate Student- t Distribution

The univariate Student- t distribution, $Y \sim t_\nu(\mu, \sigma^2)$ with $\mu \in \Re$, $\sigma > 0$, and $\nu > 0$, is defined for $y \in \Re$. The density is given by

$$f_{t_\nu}(y; \mu, \sigma^2) = \frac{\Gamma((\nu + 1)/2)}{\Gamma(\nu/2)\sqrt{\nu\pi\sigma^2}} \left(1 + \frac{(y - \mu)^2}{\nu\sigma^2}\right)^{-(\nu+1)/2}.$$

If $\nu > 1$, then $E(Y) = \mu$; if $\nu > 2$, then

$$\text{Var}(Y) = \sigma^2 \frac{\nu}{\nu - 2}.$$

The univariate $t_\nu(\mu, \sigma^2)$ -distribution is an infinite mixture of $\mathcal{N}(\mu, \sigma^2/\omega)$ -distributions, where $\omega \sim \mathcal{G}(\nu/2, \nu/2)$:

$$f_{t_\nu}(y; \mu, \sigma^2) = \int f_N(y; \mu, \sigma^2/\omega) f_G(\omega; \nu/2, \nu/2) d\omega.$$

The Multivariate Student-*t* Distribution

The multivariate Student-*t* distribution, $\mathbf{Y} \sim t_\nu(\boldsymbol{\mu}, \boldsymbol{\Sigma})$, where $\boldsymbol{\mu} \in \mathbb{R}^d$ and $\boldsymbol{\Sigma} \in \mathbb{R}^{d \times d}$ is a symmetric, positive definite matrix, is defined for $\mathbf{y} \in \mathbb{R}^d$. The density is given by

$$f_{t_\nu}(\mathbf{y}; \boldsymbol{\mu}, \boldsymbol{\Sigma}) = \frac{\Gamma((\nu + d)/2)}{\Gamma(\nu/2)} (\nu\pi)^{-d/2} |\boldsymbol{\Sigma}|^{-1/2} \left(1 + \frac{1}{\nu} (\mathbf{y} - \boldsymbol{\mu})' \boldsymbol{\Sigma}^{-1} (\mathbf{y} - \boldsymbol{\mu}) \right)^{-(\nu+d)/2}. \tag{A.23}$$

If $\nu > 1$, then $E(\mathbf{Y}) = \boldsymbol{\mu}$; if $\nu > 2$, then

$$\text{Var}(\mathbf{Y}) = \frac{\nu}{\nu - 2} \boldsymbol{\Sigma}.$$

The multivariate $t_\nu(\boldsymbol{\mu}, \boldsymbol{\Sigma})$ -distribution is an infinite mixture of $\mathcal{N}_d(\boldsymbol{\mu}, \boldsymbol{\Sigma}/\omega)$ -distributions, where $\omega \sim \mathcal{G}(\nu/2, \nu/2)$:

$$f_{t_\nu}(\mathbf{y}; \boldsymbol{\mu}, \boldsymbol{\Sigma}) = \int f_N(\mathbf{y}; \boldsymbol{\mu}, \boldsymbol{\Sigma}/\omega) f_G(\omega; \nu/2, \nu/2) d\omega.$$

A.1.13 The Uniform Distribution

The uniform distribution, $Y \sim \mathcal{U}[0, 1]$, is defined over the unit interval $y \in [0, 1]$. Density, mean, and variance are given by

$$f_U(y) = 1, \quad E(Y) = 1/2, \quad \text{Var}(Y) = 1/12. \tag{A.24}$$

A.1.14 The Wishart Distribution

The Wishart distribution, $\mathbf{Y} \sim \mathcal{W}_d(\alpha, \mathbf{S})$, with \mathbf{S} being a $(d \times d)$ symmetric, nonsingular matrix ($|\mathbf{S}| > 0$), is a standard distribution law used for a random $(d \times d)$ symmetric, positive-definite matrix \mathbf{Y} . Many parameterizations are in use for this density; the following one, introduced by Bernardo and Smith (1994), is used in this book, because in this parameterization the $\mathcal{W}_1(\alpha, S)$ distribution reduces the $\mathcal{G}(\alpha, S)$ distribution.

For $\alpha > (d - 1)/2$ the density of the $d(d + 1)/2$ -dimensional random vector of distinct elements of \mathbf{Y} is given by

$$f_W(\mathbf{Y}; \alpha, \mathbf{S}) = \frac{|\mathbf{S}|^\alpha}{\Gamma_d(\alpha)} |\mathbf{Y}|^{\alpha - (d+1)/2} \exp\{-\text{tr}(\mathbf{S}\mathbf{Y})\}, \tag{A.25}$$

where

$$\Gamma_d(\alpha) = \pi^{d(d-1)/4} \prod_{j=1}^d \Gamma\left(\frac{2\alpha + 1 - j}{2}\right) \tag{A.26}$$

is the generalized Gamma function. The mean of \mathbf{Y} is given by

$$\mathbf{E}(\mathbf{Y}) = \alpha \mathbf{S}^{-1};$$

the mode reads:

$$\left(\alpha - \frac{d+1}{2}\right) \mathbf{S}^{-1}.$$

The variance of the elements Y_{ij} of \mathbf{Y} is equal to

$$\text{Var}(Y_{ij}) = \alpha \left((S_{ij}^{-1})^2 + S_{ii}^{-1} S_{jj}^{-1} \right),$$

where S_{ij} is the (i, j) th element of \mathbf{S} . If $\mathbf{Y} \sim \mathcal{W}_d(\alpha, \mathbf{S})$ and \mathbf{A} is an $(m \times d)$ matrix with $m \leq d$, then

$$\mathbf{A} \mathbf{Y} \mathbf{A}' \sim \mathcal{W}_m \left(\alpha, (\mathbf{A} \mathbf{S}^{-1} \mathbf{A}')^{-1} \right),$$

provided that the scale matrix exists.

Inverted Wishart Distribution

A random $(d \times d)$ symmetric, positive-definite matrix \mathbf{Y} follows an inverted Wishart distribution, $\mathbf{Y} \sim \mathcal{W}_d^{-1}(\alpha, \mathbf{S})$, if $\mathbf{Y}^{-1} \sim \mathcal{W}_d(\alpha, \mathbf{S})$. In the parameterization used for the Wishart density in (A.25), the $\mathcal{W}_1^{-1}(\alpha, S)$ distribution reduces to the $\mathcal{G}^{-1}(\alpha, S)$ distribution. The density of the inverted Wishart distribution is given by

$$f_{IW}(\mathbf{Y}; \alpha, \mathbf{S}) = \frac{|\mathbf{S}|^\alpha}{\Gamma_d(\alpha)} |\mathbf{Y}^{-1}|^{\alpha+(d+1)/2} \exp\{-\text{tr}(\mathbf{S} \mathbf{Y}^{-1})\}, \quad (\text{A.27})$$

with $\Gamma_d(\alpha)$ being the same as in (A.26). \mathbf{Y} has finite expectation, iff $\alpha > (d+1)/2$:

$$\mathbf{E}(\mathbf{Y}) = \mathbf{S}/(\alpha - (d+1)/2);$$

the mode is given by

$$\mathbf{S}/(\alpha + (d+1)/2).$$

A.2 Software

A toolbox of MATLAB version 6 scripts and functions has been written by the author with the purpose of carrying out practical statistical modeling based on finite mixture and Markov switching models. The package is available at <http://www.ifas.jku.at/personal/fruehwirth/fruehwirthspringer06>.

You will need a valid MATLAB license including the STATISTICS toolbox to run functions from this package. The MATLAB software is available from MathWorks, Inc. (<http://www.mathworks.com/>).

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Index

- χ^2 -distribution, 434
- Adaptive radial-based direction sampling, 188
- AIC, 116–117
 - choosing number of clusters, 218
 - choosing number of components, 117, 238
 - Markov switching model, 346
 - mixture GLM, 294
 - normal mixtures, 200
 - Poisson mixtures, 285
 - state space model, 422
 - switching ARCH model, 382
- Akaike's criterion, *see* AIC
- AR model, 358, 360, 374, 424
 - finite mixture, 361
 - random coefficient, 377
 - self-exciting threshold, 361
 - smooth transition, 385
 - switching, *see* Markov Switching autoregressive model
 - threshold, 357, 361
- ARCH model, 374, 377, 378, 398
 - finite mixture, 379
 - switching, *see* Switching ARCH model
- ARIMA model, 390
- ARMA model, 311, 363
 - observed with noise, 399
 - state space representation, 396–397
 - switching, *see* Switching ARMA model
- Auxiliary mixture sampling, 238, 422
- Basic Markov switching model, *see* Markov switching model
- Basic structural model, 397, 401
- Bayes p -value, 113
- Bayes factor, 119–122, 143, 255, 387
 - asymptotic behavior, 120–121
 - sensitivity to prior, 122, 123, 142
- Bayes' classifier, *see* Classification
- Bayes' rule, 26, 28, 32, 118
- Bayes' theorem, 32
- Bayesian clustering, 65, 68, 220–224
 - loss function
 - 0/1, 220
 - misclassification, 221
 - similarity matrix, 223, 224
- Bayesian estimation of finite mixtures
 - choosing the prior, 58–63, 104
 - overfitting mixtures, 103–105
 - posterior density, *see* Mixture posterior density
 - simulation study, 54–56
 - using posterior draws, 87–89
- Bayesian estimation of Markov mixture posterior density, *see* Markov mixture posterior density
- Bayesian interval estimation, 35
- Bayesian model selection, 117, 120–121, 125
 - choosing priors, 122–123, 141
- Bayesian point estimation, 34, 93
- Beta distribution, 431, 432
- Beta function, 431
- Beta-binomial model, 287, 288

- Between-group heterogeneity, 170, 203, 211
- BIC, 116–117
 - choosing number of clusters, 212, 215, 218, 220
 - choosing number of components, 117, 238
 - Markov switching model, 346
 - mixture GLM, 294
 - normal mixtures, 200
 - Poisson mixtures, 285
 - random-effects models, 269
 - relation to Bayes factor, 120
 - relation to Schwarz's criterion, 116
 - state space model, 422
 - switching ARCH model, 382
- Binary data
 - auxiliary mixture sampling, 239
 - finite mixture modeling, 286–288, 294–297
 - mixture regression model, 292–293
 - time series, 315
- Binomial
 - distribution, 432, 435
 - finite mixtures, 286–288, 296
 - Bayesian estimation, 288
 - identifiability, 21, 23, 286
 - MCMC, 288
 - overdispersion, 288
 - reversible jump MCMC, 288
 - Markov mixtures, 315
- Birth and death MCMC, 137–139
 - Markov switching model, 348
 - mixture regression model, 255
 - mixtures of GLMs, 294
 - Student t mixtures, 233
- Breakpoint, *see* Structural break
- Bridge sampling, 150–151
 - MSE, 152
- Burn-in, 69
- Business cycle analysis, 359, 365–369, 385, 386, 401
- Candidate's formula, 159
- Categorical data
 - auxiliary mixture sampling, 239
 - finite mixture modeling, 288–289, 293, 297–298
- Characteristic function, 22
- Chi-squared distribution, *see* χ^2 -distribution
- Chib's estimator, 160–164
 - finite mixtures, 161–164
- Choosing the number of clusters, 210–216
 - AIC, 220
 - BIC, 212, 215, 218, 220
 - Calinski–Harabasz statistic, 211
 - classification likelihood criterion, 213
 - classification-based criteria, 213–215
 - CLC, 213
 - distortion curve, 212
 - gap statistic, 211
 - ICL, 213–215
 - integrated classification likelihood criterion, 213–215
 - Krzanowski–Lai statistic, 211
 - marginal likelihood, 212, 217, 219
 - overlapping clusters, 215
 - well-separated clusters, 215
- Classification
 - Bayes' rule, 26–29, 236, 237
 - Bayesian MAP, 210
 - combining classified and unclassified data, 236
 - discriminant analysis, 235–236
 - naïve Bayes' classifier, 27
 - supervised, 235
 - unsupervised, *see* Model-based clustering
 - within MCMC, 76
 - without parameter estimation, 68–70
 - multivariate normal mixtures, 195
 - normal mixtures, 181
 - Poisson mixtures, 70
- Classification likelihood approach, *see* Model-based clustering
- CLC, *see* Choosing the number of clusters
- Clustering, 68, 203
 - Bayesian, *see* Bayesian clustering
 - criteria, 203–204, 207–211
 - finite mixtures, 6, 204–206
 - for model identification, 96–97
 - model-based, *see* Model-based clustering
 - of the point process representation, 97

- Coefficient of determination, 170, 204
- Complete-data Bayesian estimation, 31–41, 74
 - of the weights, 35–41
 - posterior density, 32
- Complete-data ML estimation, 30–31
 - normal mixtures, 30, 31
 - Poisson mixtures, 31
- Count data
 - auxiliary mixture sampling, 239
 - mixture regression model, 290–292
 - overdispersion, *see* Overdispersion
 - time series, 315, 348–356
- Covariance selection, *see* Finite mixture of multivariate normal distributions
- Credibility interval, 35
- Data
 - DARWIN'S DATA, 225–227
 - EYE TRACKING DATA, 280, 283–285
 - FISHER'S IRIS DATA, 195, 218–220, 224
 - FISHERY DATA, 1, 2, 5, 109, 188–190, 216–218
 - GDP DATA, 358, 365–369
 - LAMB DATA, 348, 351–356
 - MARKETING DATA, 270–273
 - NEW YORK STOCK EXCHANGE DATA, 375, 382–383
 - U.S./U.K. REAL EXCHANGE RATE DATA, 423–430
- Data augmentation
 - finite mixture, 76
- Density estimation
 - bandwidth selection, 238
 - kernel methods, 237
 - using finite mixture, 237
- Diagnosing finite mixtures
 - Bayes p -value, 113
 - method of moments, 110–112
 - predictive methods, 112–114
 - residual diagnostics, 114
- Diagnosing Markov switching models, 346
- Dickey–Fuller regression, 364
- Dirichlet distribution, 432, 433
- Dirichlet process prior, 14
- Dirichlet-multinomial distribution, 289
- Discrete-valued time series, 315
- Disease mapping, 291
- Distance-based estimation, 54, 94, 112, 117, 123, 229
 - Bayesian approach, 54, 94, 123, 255
 - Hellinger distance, 54
 - Kullback–Leibler, 54, 94, 255
- Dynamic factor model
 - switching, 400, 416
 - variable selection, 422
- Eigenvalue decomposition, 198–199, 234, 236
- Elevation plot, 6
- EM algorithm
 - classification, 205
 - finite mixture, 49–53
 - multivariate normal, 172
 - normal, 52, 172
 - overfitting, 172
 - Poisson, 51
 - Markov switching model, 333
 - model-based clustering, 205
 - standard errors, 53
 - stochastic, 205
- Empty components
 - Bayesian estimation, 38–41, 141
 - ML estimation, 38
- Entropy, 28, 29, 213
- Excess kurtosis, 377
 - finite mixtures, 11
- Excess zeros, 282, 292
- Exponential
 - distribution, 433
 - finite mixtures, 4, 6, 277–279
 - applications, 277
 - Bayesian estimation, 278–279
 - definition, 4, 277
 - EM algorithm, 277
 - identifiability, 22
 - MCMC, 278
 - method of moments, 277
 - partially proper prior, 278
 - prior, 278
 - reparametrization, 278
 - reversible jump MCMC, 278
 - unimodality, 6
- Exponential family

- finite mixture of, *see* Finite mixture distribution
- Exposures, 290, 291
- Extra-binomial variation, 287, 292
- F-distribution, 433, 434
- Filtering
 - general linear regression model, 402–404
 - Kalman filter, *see* Linear Gaussian state space model
 - Kim’s algorithm, 409
 - Markov switching model, 320–324
 - infinite memory, 327
 - long memory, 326
 - numerical stabilization, 322
 - multiprocess model, 406
 - switching linear Gaussian state space model, 406–410
- Financial time series, 375, 378
- Finite Markov mixture distribution, *see* Markov mixture distribution
- Finite Markov mixture model, *see* Markov switching model
- Finite mixture distribution, 3–11, 312
 - definition, 3–5
 - descriptive features, 5–8
 - general exponential family, 5, 120
 - identifiability, 14–23
 - moments, 10–11
 - multimodality, 6–8
 - point process representation, 10, 16, 18, 20, 87, 96, 105, 109
- Finite mixture model, 12–13
 - including covariates, 13, 243
 - longitudinal data, *see* Random-effects models
 - panel data, *see* Random-effects models
 - random-effects, *see* Random-effects models
 - repeated measurements, *see* Random-effects models
- Finite mixture of regression models, 241–275, 289–294
 - allocations known, 250–252, 257–258
 - Bayesian estimation, 252–255, 258–259
 - birth and death MCMC, 255
 - EM algorithm, 250
 - false convergence of Gibbs sampler, 253
 - high-dimensional parameters, 258
 - identifiability, 243–246, 256, 290, 292
 - label switching, 254
 - latent class regression model, 246, 263
 - marginal likelihood, 255
 - MCMC, 253–254, 258–259
 - method of moments, 249
 - mixed effects, 256–259, 263, 264
 - ML estimation, 249–250
 - outliers, 249
 - prior distribution, 252, 256
 - hierarchical, 252, 253
 - unbounded likelihood, 250
- Fisher information, 30, 31, 35, 52
- Forward-filtering-backward-sampling
 - Markov switching model, 342
 - state space model, 419
- Forward-filtering-backward-smoothing, 324, 325
- Frailty, *see* Unobserved heterogeneity
- Gamma distribution, 434, 435
 - inverted, *see* Inverted Gamma distribution
- GARCH model, 377, 379
 - structural break, 380
- General linear regression model, 402
- Generalized Gamma function, 439
- Generalized linear model, 289
 - finite mixture, 289–294
 - AIC, 294
 - application, 290
 - Bayesian estimation, 294
 - BIC, 294
 - birth and death MCMC, 294
 - EM algorithm, 293
 - random effects, 299
 - reversible jump, 294, 299
- Geometric distribution, 435
 - state duration, 308
- Gibbs sampling, 69–71, 74
 - finite mixture models, 76–77, 81–83
 - hierarchical prior, 77
 - improper priors, 78, 182–185

- linear Gaussian state space model, 416
 - Markov switching model, 338–345
 - mixture regression models, 253–254, 258
 - multivariate normal mixtures, 193–194
 - non-Gaussian state space model, 423
 - normal mixtures, 180–185
 - Poisson mixture, 75
 - random-effects model, 267–269
 - single move marginal sampling of the allocations, 69–71, 73
 - Student- t mixtures, 231
- Hellinger distance, 54
- Heterogeneity
 - unobserved, *see* Unobserved heterogeneity
- Heterogeneity model, *see* Random-effects model
- Hidden indicator, 1, 12
- Hidden Markov chain
 - ergodic, 314
 - inhomogeneous, 318
 - initial distribution, 318
 - nonergodic, 316
 - nonstationary, 318
 - reducible, 318
 - sampling, 1, 342–345
- Hidden Markov chain model, 301, 303, 310, 314, 320, 333
- Hidden Markov random field model, 14
- Hidden multinomial model, 12, 14
- Hierarchical Bayes model, *see* Random-effects model
- Hit-and-run sampling, 188
- Homogeneity
 - testing, 99, 115, 255
- Hyperparameter, 61, 92, 185, 192, 193, 280, 335
- ICL, *see* Choosing the number of clusters
- Identifiability
 - constraint, 19–21, 47–49, 95, 255
 - formal, 19–21, 47–49, 95
 - finite mixture distribution, 14–23
 - generic, 21–23
 - mixture regression model, 243–246
- Identification, 94–98, 222, 255
- Importance density, 145, 150
- Importance sampling, 146–148, 296
 - allocations, 84
 - marginal likelihood, 146
 - reciprocal, 147–148
 - sequential, 188
 - state space models, 413
- Individual parameter inference, 91
- Information filter, 266, 403
- Integrated classification likelihood
 - Poisson mixtures, 66
- Intervention analysis, 396, 398, 400, 421
- Invariance
 - component specific functionals, 64
 - Markov mixture posterior, 336
 - mixture posterior density, 63
 - posterior mean, 65
 - posterior of the allocations, 65, 67
- Inverted Gamma distribution, 434
- Inverted Wishart distribution, 439
- Kalman filter, 402, 404, 411
 - derivation, 405
- Kiefer–Wolfowitz example, 174–176
 - MCMC, 183–185
- Kim’s algorithm, 409
- Kullback–Leibler distance, 9, 54, 112, 117, 123
 - symmetrized, 9, 94, 255
- Label switching, 84, 87, 89, 107, 221
 - finite mixture models, 78–83
 - mixture regression models, 254
 - unbalanced, 81
- Labeling, 46–47
 - unique, 47–49, 80, 92–95, 222
- Laplace approximation, 164, 165
- Laplace regularity, 120
- Laplace–Metropolis estimator, 165, 199, 212
- Latent class model, 294–298
 - Bayesian estimation, 296, 297
 - identifiability, 296
 - linear logistic, 298
 - mixed-mode data, 298–299
- Latent class regression model, 246, 263–264, 270

- Latent structure analysis, *see* Latent class model
- Level shift, 391, 398, 400, 416
- Leverage effect, 377, 379, 382
- Likelihood function
 - complete-data, 29–30
 - finite mixture model, *see* Mixture likelihood function
 - Markov mixture model, *see* Markov mixture likelihood function
 - penalized, 115
- Likelihood ratio statistic, 114–115
 - modified, 115
- Lindley's paradox, 122
- Linear Gaussian state space model, 391–393, 395
 - application, 393
 - filtering, 404–406
 - finite mixture, *see* Switching linear Gaussian state space model
 - identification, 392, 397
 - local level model, *see* Local level model
 - sampling state process, 418–420
 - switching, *see* Switching linear Gaussian state space model
- Local level model, 389–391, 414, 415
 - outlier, 390
- Local linear trend model, *see* Trend
- Logit regression model
 - finite mixture, 292–294
 - applications, 293
 - identifiability, 292
- Longitudinal data
 - mixed-effects, *see* Random-effects models
 - random-effects, *see* Random-effects models
- Loss function
 - clustering, 220–223
 - model selection, 121
 - parameter estimation, 93–94
 - state estimation, 345
- Mahalanobis distance, 7–9, 232
- Marginal likelihood, 118–119
 - Poisson distribution, 160
- Marginal likelihood for finite mixture, 139–141, 165–167
- bridge sampling, 150–159, 218
- computation, 143
- harmonic mean estimator, 148, 150, 154–159
- hierarchical priors, 159
- importance sampling, 154–159, 218
- Laplace approximation, 164, 165
- Laplace–Metropolis estimator, 165, 199, 212
- Monte Carlo integration, 146
- reciprocal importance sampling, 147–148, 154–159, 218
- Markov chain, 302, 304–308
 - aperiodicity, 307
 - ergodicity, 305, 307
 - invariance property, 304
 - invariant distribution, 304–306
 - irreducibility, 305
 - nonergodic, 316
 - periodic, 307
 - persistence, 307, 308
 - reducible, 305
- Markov chain Monte Carlo, 68, 74, 88
 - Gibbs sampling, *see* Gibbs sampling
 - Metropolis–Hastings algorithm, *see* Metropolis–Hastings algorithm
 - reversible jump, *see* Reversible jump MCMC
 - trans-dimensional, *see* Trans-dimensional MCMC
- Markov mixture distribution, 301–318
 - autocorrelation, 310–311
 - autocorrelation squared process, 311–312
 - excess kurtosis, 309
 - identifiability, 313–314
 - moments, 308–309
 - multimodality, 309
 - relation to ARMA, 311
 - skewness, 309
- Markov mixture likelihood function, 330–333
 - computation, 331, 332
 - conditional, 332
 - multimodality, 333
 - unboundedness, 333
 - unconditional, 332
- Markov mixture posterior density
 - hidden Markov chain, 336

- invariance, 336
- marginal densities, 336
- single state, 337
- Markov switching autoregressive model, 358–369
 - application, 365
 - ARMA representation, 363
 - asymmetry, 361, 366
 - autocorrelation, 362
 - Bayesian estimation, 365
 - EM algorithm, 365
 - endogenous selection, 384
 - identification, 368
 - marginal likelihood, 366
 - ML estimation, 365
 - moments, 362
 - nonstationary time series, 363
 - order selection, 365
 - prior, 366
- Markov switching conditional heteroscedasticity, 373–384, 398
- Markov switching dynamic regression model, 317, 371–372
 - Bayesian estimation, 371–372
 - mixed-effects, 371
 - prior, 372
- Markov switching model, 314–315, 319–348
 - AIC, 346
 - Bayesian estimation, 334–346
 - BIC, 346
 - birth and death MCMC, 348
 - Chib’s estimator, 355
 - classification likelihood approach, 334
 - complete-data Bayesian estimation, 329
 - transition matrix, 329–330
 - complete-data likelihood function, 327–328
 - conditional heteroscedasticity, *see* Markov switching conditional heteroscedasticity
 - diagnostics, 346
 - EM algorithm, 333
 - forecasting, 374–375
 - identification, 354
 - label switching, 355
 - likelihood function, *see* Markov mixture likelihood function
 - marginal likelihood, 347–348, 352
 - MCMC, 338–339
 - method of moments, 334
 - ML estimation, 333–334
 - multivariate time series, 386–388
 - nonstationary series, 363
 - panel data, 385–386
 - posterior, *see* Markov mixture posterior density
 - predictive density, 372–374
 - prior, 335
 - reparameterization, 345
 - reversible jump MCMC, 348
 - state estimation, *see* State estimation
- Markov switching regression model, 315–316
- Markov switching trend model, 364
 - multivariate time series, 387
- Method of moments, 42–43
 - for diagnosing mixtures, 110–112
 - Poisson mixtures, 43
 - simulation study, 54–56
- Metropolis–Hastings algorithm, 72–73
 - binomial mixtures, 288
 - mixture regression models, 253
 - mixtures of GLMs, 294
 - sampling allocations, 72–73
 - sampling parameters in a finite mixture model, 83–84
 - stationary hidden Markov chain, 341
 - Student t mixture, 233
 - switching ARCH model, 380
- Misclassification rate, 27, 221–222, 346
- Missing data problem, 54, 73, 236, 253, 294, 334, 337, 396
- Mixed-effects model, *see* Random-effects models
- Mixed-mode data, 298
- Mixture density, *see* Finite mixture distribution
- Mixture likelihood function, 43–49, 74, 333
 - definition, 43
 - maximization, 49
 - multimodality, 44–49
 - overfitting mixture, 100–103
 - potential overfitting, 105
 - relation to complete-data likelihood, 44

- surface plot, 45, 46, 48, 50, 102, 106, 107, 174, 175
- Mixture posterior density
 - improper, 59–60
 - invariance, 63, 78–83
 - visualization, 85–87
- Mixtures of distributions, *see* Finite mixture distribution
- Mixtures of factor analyzer models, 234
- Mixtures of mixtures, 236
- Mixtures of probabilistic component analyzer, 234
- Mixtures-of-experts models, 246, 274
- ML estimation
 - EM algorithm, *see* EM algorithm
 - finite mixtures
 - asymptotic properties, 52, 100–101
 - gradient method, 49
 - Newton's method, 49, 412
 - practical difficulties, 52–53
 - simulation study, 54–56
 - under model uncertainty, 100–103
 - Markov switching autoregressive model, 365
 - Markov switching model, 333–334
 - mixture regression model, 249–250
- Mode hunting
 - mixture posterior, 108–109
 - sample histogram, 109–110
- Model checking, *see* Diagnosing mixtures
- Model selection, *see* Bayesian model selection
- Model-based clustering, 203–224
 - applications, 205, 234
 - Bayesian approach, 205–206, 210, 220–223
 - classification likelihood approach, 207–210
 - dimension reducing techniques, 233–235
 - EM algorithm, 205
 - high-dimensional data, 233–235
 - large data sets, 206
 - mixed-mode data, 298
 - mixtures of Student- t , 229, 230
 - nagging problems, 205
 - noise component, 229
 - outliers, 227
 - robustness, 227, 230
 - variable selection, 234
- Modes of a mixture density, 6–8
- Moment-generating function, 22, 42, 112, 249
- Monte Carlo Integration, 143
- MSAR model, *see* Markov switching autoregressive model
- Multimodality
 - finite mixture distribution, *see* Finite mixture distribution
 - mixture likelihood, *see* Mixture likelihood function
- Multinomial
 - distribution, 435
 - mixtures, 288–289
 - applications, 289
 - Bayesian estimation, 289
- Multiprocess model, 393, 394, 406
 - filtering, 406
- Multivariate normal
 - distribution, 436
 - mixtures
 - AIC, 200
 - allocations known, 190
 - application, 195, 205, 236, 237
 - Bayesian parameter estimation, 173, 190–195
 - BIC, 200
 - classification, 181, 194, 195
 - conditionally conjugate prior, 192
 - constrained ML estimation, 173
 - definition, 4, 169
 - eigenvalue decomposition, 198, 234, 236
 - EM algorithm, 172
 - hierarchical prior, 193, 194, 200
 - homoscedastic, 197, 208, 236
 - identifiability, 23
 - isotropic, 197, 208, 234
 - marginal likelihoods, 199, 219
 - MCMC, 193–195, 197
 - method of moments, 172
 - ML estimation, 49, 172, 173
 - multimodality, 6–8
 - prior, 192–193, 200
 - reversible jump, 201, 202
 - spherical, 197, 209
 - unboundedness of likelihood, 173

- variance decomposition, 170
- Multivariate time series
 - cointegration model, 387
 - dynamic factor model, *see* Dynamic factor model
 - Markov switching cointegration model, 387
 - Markov switching model, 386–388
 - Markov switching VAR model, 386
 - Markov switching VEC model, 387
 - Markov trend model, 387
- Negative binomial
 - distribution, 435
 - mixture regression, 292
 - modeling count data, 281, 283
- Newton's method, 49, 52, 412
- Nonidentifiability
 - invariance to relabeling, 15–16
 - of a finite mixture distribution, 15–19
 - potential overfitting, 17–19
- Nonidentifiability set, 15–19, 80, 100, 101, 103–105, 107, 109
- Nonparametric mixture modeling, 14
- Nonregular likelihood, 38, 53
- Nonstationary time series
 - local level model, 389
 - Markov switching models, 363
 - smoothing, 401
- Normal
 - distribution, 436, 437
 - Gamma distribution, 437
 - Markov mixture, 303
 - mixtures
 - AIC, 200
 - allocations known, 177
 - application, 176, 188–190, 237
 - Bayesian parameter estimation, 173, 177–190
 - BIC, 200
 - complete-data ML estimation, 30, 31
 - conditionally conjugate prior, 179
 - constrained ML estimation, 173
 - definition, 4, 169
 - dependence prior, 185–187
 - EM algorithm, 172
 - excess kurtosis, 11
 - hierarchical prior, 185–186, 200
 - identifiability, 22
 - independence prior, 179
 - marginal likelihoods, 199, 217
 - Markov prior, 187
 - MCMC, 180–185
 - method of moments, 42, 172
 - ML estimation, 49, 172, 173
 - multimodality, 6–8
 - multivariate, *see* Multivariate normal
 - prior, 179, 185–187, 200
 - prior variance ratio, 179
 - reversible jump, 200
 - skewness, 5, 11
 - unboundedness of likelihood, 173, 176
- Normalizing constant, 34, 44, 59, 66, 140, 150, 159, 160, 321
- Numerical stabilization, 322
- Omitted regressors, 1, 3, 248, 261, 287, 292
- Outliers, 224–230
 - Bayesian approach, 225–226
 - prior, 226
 - clustering, 227–230
 - finite mixture, 224–226
 - local level model, 390
 - location shift model, 225, 226
 - regression model, 249
 - time series, 398–400
 - variance inflation model, 225
- Overdispersion, 280–282
 - Poisson mixtures, 281
 - Poisson mixture regression, 290
 - time series, 348
- Overfitting
 - mixtures, 97, 100
 - potential, 105–107
- Panel data
 - clustering, 234, 385
 - Markov switching model, 385–386
 - random-effects, *see* Random-effects models
- Particle filtering, 188, 345
- Perfect sampling, 85
- Permutation sampling
 - finite mixtures, 81–83

- Markov switching models, 340
- Point process, 10, 137
- Poisson
 - distribution, 435, 437
 - finite mixtures, 279–285
 - AIC, 285
 - application, 279, 280, 282–285, 291
 - Bayesian estimation, 280
 - BIC, 285
 - clustering, 70
 - complete-data Bayesian estimation, 33, 34
 - complete-data ML estimation, 31
 - data augmentation, 74–76
 - definition, 4, 279
 - excess zeros, 282–283
 - hierarchical prior, 76
 - identifiability, 23
 - inflated zeros, 282
 - marginal likelihoods, 283
 - MCMC, 74–76
 - method of moments, 43
 - prior distribution, 60, 61, 76, 280
 - reversible jump MCMC, 135–136, 285
 - Markov mixture, 315, 328
 - autocorrelation, 350, 351
 - marginal likelihoods, 352
 - MCMC, 337–338
 - overdispersion, 350, 351
- Poisson regression model
 - excess zeros, 292
 - finite mixture, 290–292
 - application, 291
 - estimation, 294
 - identifiability, 290
 - overdispersion, 281, 290
 - inflated zeros, 292
- Poisson–Gamma model, 281, 282
- Posterior density ratio, 159–160
- Posterior mean, 65, 93
- Posterior mode, 94
- Posterior predictive density, *see* Predictive density
- Predictive density
 - diagnosing mixtures, 112–114
 - estimation, 89
 - Markov switching model, 372
 - posterior, 89, 90
 - sampling from, 90, 374
- Principal component analysis, 235
- Prior
 - conditionally conjugate, 60, 66
 - hierarchical, 61–62
 - improper, 58–60
 - invariant, 62, 63
 - Jeffrey’s, 36, 62
 - natural conjugate, 33
 - objective, 58
 - partially proper, 61–62
 - reference prior, 62
 - subjective, 58
- Probability distributions, 431–439
- Probit regression model
 - finite mixture, 292–294
 - identifiability, 292
- Product space MCMC, *see* Trans-dimensional MCMC
- Pseudo-prior, 127, 128
- Random-effects models, 14, 259–275
 - Bayesian estimation, 265–269
 - BIC, 269
 - heterogeneity model, 264–265, 270
 - hierarchical Bayes, 262–263
 - marginal likelihood, 269, 271
 - MCMC, 267–269
 - misspecification, 260, 263, 299
 - mixed-effects, 270
 - nonnormal data, 299
 - pooling, 260
 - prior distribution, 265
 - random coefficient model, 261
 - reversible jump MCMC, 269
 - shrinkage, 260, 262
 - variance heterogeneity, 261, 268
- Ratio-of-uniform method, 188
- Relabeling, *see* Labeling
- Reparameterization, 186, 278, 337, 345
- Repeated measurements
 - clustering, 234
 - finite mixture modeling, *see* Random-effects models
 - random-effects, *see* Random-effects models
- Residual heterogeneity, *see* Unobserved heterogeneity

- Reversible jump MCMC, 129–137, 139, 165–167
 - binomial mixtures, 288
 - birth moves, 136–137
 - death moves, 136–137
 - exponential mixtures, 129, 278
 - finite mixture models, 129, 131–137
 - finite mixture of GLMs
 - with random effects, 299
 - Markov switching model, 129, 348
 - merge moves, 133–135
 - mixtures of GLMs, 294
 - multivariate normal mixtures, 201, 202
 - normal mixtures, 200
 - Poisson mixtures, 129, 135–136, 285
 - random-effects model, 269
 - split moves, 133–135
- Ridgeline surface, 6
- Sampling allocations
 - Gibbs sampling, *see* Gibbs sampling
 - importance sampling, *see* Importance sampling
 - Metropolis–Hastings algorithm, *see* Metropolis–Hastings algorithm
- Savage–Dickey density ratio, 122
- Schwarz criterion, *see* BIC
- SETAR model, *see* AR model
- Similarity matrix, 223, 224, 346
- Similarity of mixture components, 9
- Skewness, 5, 113, 377
 - finite mixtures, 11, 113
- Smoothing
 - Markov switching model, 324–326
 - infinite memory, 327
 - long memory, 326
- Software, 439
- Spurious modes, 173, 176, 179, 183
- Standard finite mixture model, *see* Finite mixture model
- State
 - duration, 361
 - estimation, 319–327
 - loss function, 345
 - without parameter estimation, 344
 - probabilities
 - filtered, 320–324, 326–327
 - predictive, 320
 - smoothed, 320, 324–327
- State space model, 389–393
 - AIC, 422
 - auxiliary mixture sampling, 422
 - BIC, 422
 - complete-data Bayesian estimation, 414
 - EM algorithm, 412
 - filtering, 402
 - general form, 394
 - likelihood function, 411
 - linear Gaussian, *see* Linear Gaussian state space model
 - Metropolis–Hastings algorithm, 418, 420
 - non-Gaussian
 - Gibbs sampling, 423
 - outlier, 398–400
 - regime switching, *see* Switching state space model
 - robust, 399
 - Student-*t* errors, 399
- Stochastic volatility model, 377, 423
- Structural break, 316, 318, 363, 380, 398
 - regression model, 246–247
- Student-*t*
 - distribution, 437
 - multivariate, 438
 - mixtures, 229–233
 - Bayesian parameter estimation, 230–233
 - EM algorithm, 230
 - known degrees of freedom, 231
 - marginal likelihood, 233
 - MCMC, 231–233
 - Metropolis–Hastings, 233
 - prior degrees of freedom, 232
 - robustness, 231
- Switching ARCH model, 326, 378–383
 - AIC, 382
 - Bayesian estimation, 380
 - BIC, 382
 - leverage, 377, 379, 382
 - marginal likelihood, 382
 - MCMC, 380
 - Metropolis–Hastings algorithm, 380
 - ML estimation, 380
 - second-order stationarity, 379

- Switching ARMA model, 327, 396
- Switching dynamic factor model, *see*
 - Dynamic factor model
- Switching GARCH model, 327, 383–384
 - applications, 384
 - stationarity, 383
- Switching linear Gaussian state space model, 393–394
 - application, 423–430
 - filtering, 406–410
 - identification, 428
 - marginal likelihood, 422, 427
 - random permutation sampler, 427
- Switching regression model, 246, 315
- Switching state space model, 389–430
 - Bayesian estimation, 412–421
 - EM algorithm, 412
 - likelihood function, 411
 - linear Gaussian, *see* Switching linear Gaussian state space model
 - MCMC, 415–421
 - prior, 413
 - sampling
 - continuous states, 417–420
 - discrete states, 420–421
- TAR model, *see* AR model
- Testing homogeneity, *see* Homogeneity
- Time-varying transition matrix, *see*
 - Transition matrix
- Trans-dimensional MCMC, 118, 125–139
 - birth and death MCMC, *see* Birth and Death MCMC
 - product space MCMC, 126–129
 - Metropolis-Hasting, 129
 - reversible jump, *see* Reversible jump MCMC
- Transition matrix, 302, 394
 - eigenvalues, 304, 307
 - irreducible aperiodic chains, 304–308
 - posterior sampling, 340–341
 - prior, 335
 - time-varying, 318, 384–385, 401
- Trapping states, 183
- Trend
 - local linear trend model, 397, 398
 - Markov switching, 364
 - stochastic, 363
- Uniform distribution, 438
 - finite mixtures
 - identifiability, 21
- Unit root test
 - structural break, 363–364
- Unit simplex, 4, 37, 41, 302, 432
- Unobserved component time series model, 397–398, 416
 - ARCH disturbances, 398
 - local level model, *see* Local level model
- Unobserved heterogeneity, 1–3, 281, 287
 - distribution of heterogeneity, 260, 262–264
 - occurrence probabilities, 287–288
 - omitted regressors, 1, 3, 248, 261, 287, 292
 - pooling across units, 260
 - unknown segments in the population, 1, 3, 248, 260
- Variable selection, 234, 422
 - and clustering, 234
- Variance decomposition, 170
- Viterbi algorithm, 346
- Volatility clustering, 377
- Weight distribution
 - finite mixture, 4
 - choosing prior, 36–41
 - complete-data Bayesian estimation, 35–41
 - including covariates, 274–275
 - mixtures-of-experts, 274
- Wishart distribution, 438, 439
 - inverted, *see* Inverted Wishart distribution
- Within-group heterogeneity, 170, 177, 203, 211

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