

# Forecasting secular variation using core flows

Ciarán Beggan<sup>1</sup> and Kathy Whaler<sup>2</sup>

<sup>1</sup>British Geological Survey, Murchison House, West Mains Road, Edinburgh, EH9 3LA, United Kingdom

<sup>2</sup>School of GeoSciences, University of Edinburgh, West Mains Road, Edinburgh, EH9 3JW, United Kingdom

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Over the past ten years satellite measurements in combination with data from ground-based observatories have allowed very detailed models of the secular variation (SV) of the Earth's magnetic field to be constructed. However, forecasting the change of the main field still remains a challenge, primarily because the core processes controlling SV are not sufficiently well understood. Hence, most forecasts do not appeal to any physical modelling constraints but use, for example, polynomial extrapolation from previous measurements. We attempt to apply a physical model to forecast the average SV during 2010–2015 by developing a core flow model. This steady flow model, derived from SV data during 2004.5 to 2009.5, generates a set of Gauss SV coefficients which are used to advect the large scale magnetic field forwards in time. Although this model has not been submitted as a candidate for IGRF-11, we present our SV prediction model and compare it to other candidate IGRF-11 SV models. In addition, we examine the use of the Ensemble Kalman filter to optimally assimilate field models derived from (1) forecast methods and (2) noisy data measurements. Such a scenario might conceivably arise if high quality satellite data with global coverage are not available for a significant period of time. We show that the overall misfit of the assimilated model to the actual field can be lower than the individual misfits of the input models, provided the uncertainties of each model are reasonably well known.

**Key words:** Magnetic field, secular variation, forecasting, core flows.

## 1. Introduction

Forecasting of magnetic field change has been attempted in many forms over the past 300 years since Halley's observation of westward motion of the agonic line in the Atlantic hemisphere (Halley, 1692). Currently, a widely used forecast of the average secular variation (SV) is produced for the International Geomagnetic Reference Field (IGRF) model every five years. The IGRF models are an agreed set of field coefficients representing snapshots of the magnetic field at defined times and a set of coefficients forecasting the average SV for a period of five years into the future. The tenth generation of the International Geomagnetic Reference Field model (IGRF-10) covered the period from 2005.0 to 2010.0 (Macmillan and Maus, 2005). For the IGRF-10 SV model, the methods to estimate the average SV over its five year lifetime used a combination of polynomial extrapolation of satellite data and linear prediction filters applied to observatory data. However, these approaches do not invoke any particular physical arguments to support the assumption that the field coefficients will continue to change linearly (which, of course, they do not). Lowes (2000) succinctly outlines the issues faced in forecasting of SV, noting that an understanding of the fluid processes within the core might be useful in producing more accurate predictions. Since 2005, a number of forecasting techniques have been developed to incorporate physical approximations. For example, Sun *et al.* (2007) and Fournier *et al.* (2007) both outline frameworks in which ob-

served magnetic field data can be assimilated into physical magnetohydrodynamic models. More recently, Kuang *et al.* (2009) have assimilated historical field data into numerical dynamo models to investigate if improvements can be made to forecast field models.

We present a forecast based upon a steady core flow model generated from satellite magnetic data measured over the period 2004.5–2009.5. The SV coefficients derived from a flow model are used to advect the main field Gauss coefficients forwards in time. In Section 2 we examine how well the IGRF-10 SV model forecast has performed while in Section 3 we compare the forecast from a steady flow model over a five year period starting in 2004.5 to the IGRF-10 estimate for 2009.5.

In Section 4, we investigate a method known as the Ensemble Kalman Filter to assimilate field models derived from magnetic measurements (from observatories or satellite) and field models computed from a forecast. Ensemble techniques to examine variability and error in geomagnetic studies have previously been applied to core flow modelling by Gillet *et al.* (2009). We outline our implementation scheme and use error estimates from comparisons to past IGRF and Definitive Geomagnetic Reference Field (DGRF) models to determine reasonable variances for the filter and to illustrate the potential benefits for forecasting in the absence of high-quality satellite magnetic data.

## 2. Forecasting Ability

To ascertain how accurate previous forecasts have been, it is useful to examine how the IGRF-10 field model and its SV prediction compares against the DGRF and IGRF-11

models (Finlay *et al.*, 2010). The models are compared using a root mean square (RMS) difference (or misfit) metric ( $\sqrt{dP}$ ) calculated by Maus *et al.* (2008):

$$dP = \sum_{n=1}^{n_{\max}} (n+1) \sum_{m=0}^n [(\mathbf{g}_n^m)_{\text{Model A}} - (\mathbf{g}_n^m)_{\text{Model B}}]^2 \quad (1)$$

where  $\{\mathbf{g}_n^m\}$  are the Gauss coefficients of the field and are used to represent both  $g_n^m$  and  $h_n^m$ . The maximum degree of the models is  $n = m = 13$  for the main field and  $n = m = 8$  (giving  $d\dot{P}$  from  $\dot{g}_n^m$ ) for the SV.

The IGRF-10 model coefficients for 2005.0 were generated by forward extrapolation of data measured up to mid-2004 and are a combination of three candidate field models (Maus *et al.*, 2005). Hence a revision is required to produce a retrospective DGRF model for 2005.0. The RMS difference between the IGRF-10 model and the DGRF model for 2005 is approximately 13 nT. The misfit of the IGRF-10 model coefficients to the best available estimate of the internal magnetic field is thus slightly larger than the 5 nT suggested in Maus *et al.* (2005).

The change in the magnetic field over five years can be estimated by examining the RMS difference between the IGRF-11 model for 2010 and the DGRF model for 2005; it is 399 nT. By comparison, the difference between the IGRF-11 model for 2010 and the IGRF-10 model for 2005 is 401 nT. However, although the IGRF-10 field model benefitted from greatly improved availability of satellite data compared to previous IGRF models, the predicted SV coefficients were still relatively poor and did not accurately forecast the actual SV from 2005 to 2010. The RMS difference between the field for 2010 as estimated from the IGRF-10 model (i.e. IGRF-10 in 2005 plus the sum of the predicted annual SV for five years) and the IGRF-11 model is 119 nT. This is an average RMS misfit of about 24 nT/yr, which is slightly larger than the error estimate of 20 nT/yr given in Maus *et al.* (2005). Figure 1 summarises the RMS difference relationships.

It can be concluded that adequate techniques to accurately estimate the global SV over five years have not yet been developed. This is due, primarily, to the unknown changes of fluid flow within the core (e.g. Holme, 2007). In the next section, we attempt to improve upon the SV forecast by use of a steady core flow model to advect the magnetic field model forward in time.

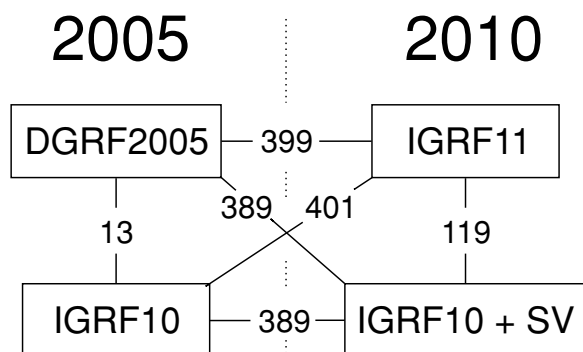


Fig. 1. Root Mean Square differences between field models (in nT) for 2005 and 2010.

### 3. Flow Modelling

We investigate whether using a core flow model to predict SV can reduce the RMS difference between the actual and predicted field at the end of five years. A study by Maus *et al.* (2008) examined how well hindcasting of the magnetic field using core flow models reproduced the observed field over 13 years. They employed the ‘frozen flux’ approximation in which diffusion of the magnetic field on large scales is assumed to be negligible on short timescales (Roberts and Scott, 1965). Maus *et al.* (2008) compared the prediction of the field back in time from the SV generated by a number of core flow modelling assumptions including steady, toroidal only and accelerated flows. They concluded that a steady flow produces the best average fit over a ten-year period. Following on from this result, Beggan and Whaler (2009) investigated how well a steady flow generated from satellite data could predict secular variation. Using the same techniques, here we test how well a steady core flow model developed from CHAMP satellite data from 2001.4–2004.5 can predict the change in the magnetic main field between 2004.5 and 2009.5. These time periods are covered by current satellite field models and allow consistent comparisons to be made.

For this study, we prepared a series of 27 monthly SV data sets, over the period 2001.4–2004.5, generated from CHAMP satellite data using the ‘Virtual Observatory’ (VO) method of Manda and Olsen (2006). The SV data were inverted for toroidal and poloidal flow using the linear relationship between SV and flow spherical harmonic coefficients. The relation is through the Gaunt/Elsasser matrix ( $\mathbf{H}$ ) whose elements depend on the main field coefficients which change with time (Whaler, 1986). The main field, SV and flow coefficients are truncated at degree and order  $n_{\max} = 14$ , and thus we have assumed that only large scale flows are responsible for the large scale SV. We assumed a steady flow model, with tangential geostrophy (Hills, 1979; Le Mouél, 1984) imposed as a weak penalty norm constraint. The Gauss coefficients (in the vector  $\mathbf{g}$ ) from the POMME3 main field model (Maus *et al.*, 2006) were used in the flow inversion.

Employing the method outlined in Beggan and Whaler (2009), the steady flow model coefficients ( $\hat{\mathbf{m}}_{\text{SF}}$ ) were used to forecast the change in the magnetic field over the five year period from 2004.5 to 2009.5. This period was chosen to allow direct comparison between the forecast field model coefficients and the CHAOS-2s model (Olsen *et al.*, 2009). Thus the Gauss coefficients from CHAOS-2 (up to degree 12) for 2004.5 were used as the starting field model. The field was advected forward (forecast) over successive months ( $t$ ) for five years (until 2009.5) using the equation:

$$\mathbf{g}_{t+1} = \mathbf{g}_t + (\mathbf{H}_t \hat{\mathbf{m}}_{\text{SF}})/12 \quad (2)$$

with the  $\mathbf{H}_t$  matrix updated at every timestep using the main field coefficients forecast from the previous timestep, making the system slightly non-linear.

After five years, the RMS difference between the forecast model and the satellite field model CHAOS-2 (at 2009.5) was 85 nT, or approximately 17 nT/yr. As a comparison, the IGRF-10 SV coefficients were used to compute the field in 2009.5 using the CHAOS-2 coefficients for 2004.5 as

Table 1. Predicted average SV coefficients from the steady flow model, up to degree and order 8.

$n$	$m$	$g_n^m$	$h_n^m$
1	0	12.5	0
1	1	16.9	-28.2
2	0	-11.4	0
2	1	-5.4	-21.3
2	2	2.0	-12.1
3	0	0.9	0
3	1	-3.2	7.9
3	2	-3.4	-3.4
3	3	-7.5	-2.2
4	0	-1.6	0
4	1	1.8	0.4
4	2	-8.2	3.6
4	3	5.0	3.0
4	4	-2.5	-0.9
5	0	-0.6	0
5	1	0.5	-0.1
5	2	-1.7	1.7
5	3	-0.9	1.1
5	4	1.4	3.7
5	5	0.8	-0.8
6	0	0.2	0
6	1	-0.3	0.3
6	2	-0.3	-1.6
6	3	2.0	-0.2
6	4	-1.5	-0.2
6	5	-0.3	0.6
6	6	2.0	0.4
7	0	0.2	0
7	1	-0.1	0.6
7	2	-0.7	0.2
7	3	1.1	-0.2
7	4	0.1	-0.2
7	5	0.1	-0.7
7	6	-0.8	0.2
7	7	0.8	0.2
8	0	0.3	0
8	1	0.0	0.0
8	2	-0.3	0.3
8	3	0.3	0.0
8	4	-0.3	0.3
8	5	0.3	-0.1
8	6	0.3	-0.2
8	7	-0.6	0.6
8	8	0.1	0.3

the starting point. The misfit of the IGRF-10 SV model at 2009.5 to the CHAOS-2 model was 102 nT.

We verified that the flow model generated directly from the SV derived from the VO method is insensitive to the main field model used in the inversion by comparing the prediction of a steady flow derived from the CHAOS main field model (Olsen *et al.*, 2006) instead of POMME3. The resulting flow model was only slightly different from that derived from the POMME3 model, and the RMS difference between the SV forecast model using the CHAOS-derived steady flow after five years was also approximately 85 nT. Thus, SV estimates from steady core flows give a

lower RMS misfit over five years and suggest a forecasting method employing core flow modelling may be beneficial.

We applied the same methodology to forecast the average SV over the period 2010–2015, now based on a steady flow derived from a longer time-series of fifty monthly VO SV models using CHAMP data from 2004.5–2009.5. Due to changes in core flows associated with geomagnetic jerks (e.g. Gubbins, 1984; Wardinski *et al.*, 2008), we use satellite data after the jerk in about 2003.4 (Olsen and Manda, 2007). The field was computed using Eq. (2), starting with the main field coefficients for 2009.5 from CHAOS-2. The coefficients of the average SV (given in Table 1) were calculated by dividing the difference between the main field coefficients predicted at 2010 and those forecast at 2015 by five years.

Figure 2 shows the Lowes-Mauersberger spectrum of the SV derived from the steady flow model for 2009.5–2014.5, compared to those of IGRF-11 candidate models A (DTU), E (EOST/ LPGN/ LATMOS/ IPGP), G (GFZ) and H (GFSC/JCET) illustrating that the SV prediction of the steady flow model is within the range of these proposed models. It also lies within the range of the other four candidates which are not shown to avoid cluttering the figure. The prediction of candidate Model A has more power in the lower degrees ( $n = 1-4$ ), while Model E and Model H predictions are very similar for degrees  $n = 2-6$ .

#### 4. Data Assimilation

We wish to investigate whether it is possible to combine a relatively good forecast field model with a lower quality ‘measured’ field model to improve the estimate of the actual field. For example, if the existing set of satellites fail before Swarm is fully operational, there could be a gap in which our present high-quality global coverage is diminished and we may become mostly reliant upon ground-based observatories and repeat stations to produce global magnetic field models. Due to the uneven geographic distribution of ground-based observations these models have much higher uncertainties than models employing satellite data. This section looks at a potential method for mitigating the impact of such an event by employing an optimal data assimilation algorithm to make best use of all available information. We investigate whether a sufficiently accurate forecast can be obtained using an initial high-resolution satellite field model, combined with a flow model for advection of the field and intermittent updates from a simulated lower quality ground-based field model.

Beggan and Whaler (2009) established a method for applying the Ensemble Kalman Filter (EnKF) technique of Evensen (1994) to combine SV forecasts from a steady flow model with intermittent updates from a relatively ‘noisy’ field model. As they noted, the results are highly dependent on the assumptions about the errors of the flow and field models. Indeed, if the field model is of high quality, then the EnKF technique is redundant. However, if the field model is relatively poor, then a combination of a forecast field model using a steady core flow model and a field model computed from ground-based data can reduce the overall misfit to the ‘true’ field.

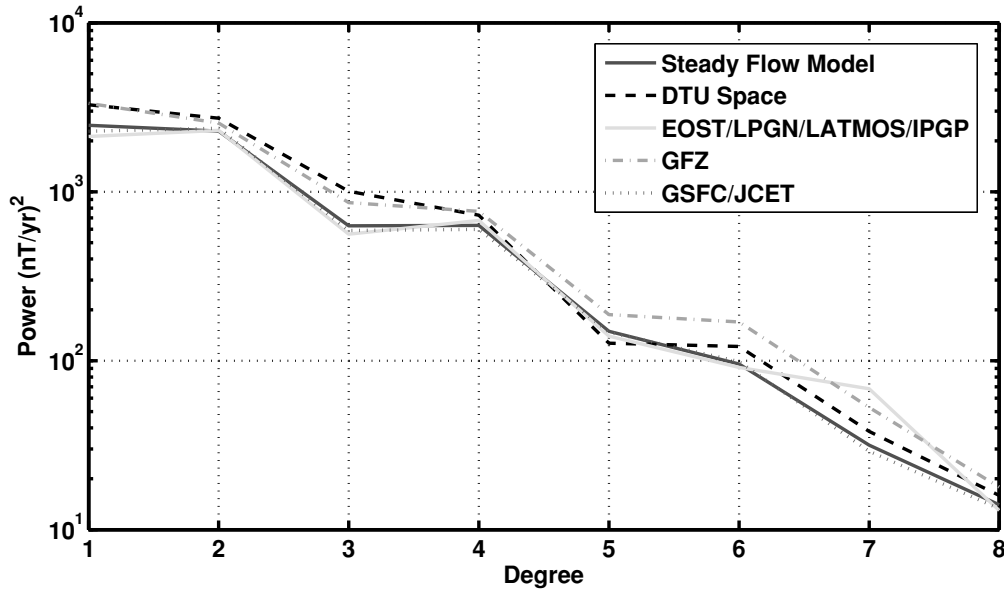


Fig. 2. Spectrum of the Steady Flow model SV prediction for 2010–2015 compared to that of the IGRF-11 SV candidate Models A (DTU Space), E (EOST/LPGN/LATMOS/IPGP), G (GFZ) and H (GFSC/JCET).

#### 4.1 Ensemble Kalman filtering

A traditional single-state Kalman Filter is implemented in two steps: (1) prediction of the evolution of the model state by equations believed to adequately represent the system and (2) assimilation of a measurement to correct any accumulated error from the model (Kalman, 1960). The advantage of the Kalman filter is that measurements can be assimilated whenever they are available. When no data are available, the process is modelled by forecasting. Even relatively poor data can be used to constrain the forecast, as they are optimally included into the filter. At a time  $t$ , the optimal blending of a forecast ( $\mathbf{x}_t^f$ ) and measurement ( $\mathbf{z}_t$ ) to generate the assimilated state vector,  $\mathbf{x}_t^a$ , is through the so-called Kalman gain matrix ( $\mathbf{K}_t$ ):

$$\mathbf{x}_t^a = \mathbf{x}_t^f + \mathbf{K}_t (\mathbf{z}_t - \mathbf{x}_t^f) \quad (3)$$

with

$$\mathbf{K}_t = \mathbf{P}_t^f (\mathbf{P}_t^f + \mathbf{Q})^{-1}, \quad (4)$$

where  $\mathbf{P}_t^f$  is the error covariance matrix of the model equations and  $\mathbf{Q}$  is the error covariance matrix for the measurement.

In our application of this method, we set  $\mathbf{x}_t^f$  to be a vector of Gauss coefficients produced from a forecast as defined in Eq. (2). If a field model, derived from observed data, becomes available it can be represented as  $\mathbf{z}_t$ , another vector of Gauss coefficients. Combining the measurement and the forecast depends on how the errors for each vector are defined, which we will discuss in Section 4.2.

Examination of Eqs. (3) and (4) reveals some important aspects of the filter. If the measurement  $\mathbf{z}_t$  and model prediction  $\mathbf{x}_t^f$  are equal, then the update has no effect. If the elements of the measurement covariance error matrix ( $\mathbf{Q}$ ) are small, then  $(\mathbf{z}_t - \mathbf{x}_t^f)$  is more heavily weighted in the update. Equally, if the model covariance error matrix elements ( $\mathbf{P}^f$ ) are small then the forecast state ( $\mathbf{x}_t^f$ ) is more important. The balance between these matrices (i.e. the estimated un-

certainties of the model and measurement) controls the assimilation step.

The Kalman Filter was designed as a linear filter. In a non-linear regime, it can become unwieldy and unreliable. To overcome such limitations, Evensen (1994) suggested an algorithm for data assimilation employing Monte Carlo methods as an alternative. The state of a process at any particular time is represented as a vector in  $r$ -dimensional space, where  $r$  is the number of parameters in the system. In an Ensemble Kalman Filter the uncertainty of our knowledge of the process is represented by perturbing the inputs to the model forecast equations randomly by a known variance (with zero-mean) to produce an ‘ensemble’ of states—conceptually imagined as a ‘cloud’ of points in  $r$ -dimensional space. The evolution of the states through time is controlled by propagating the entire ensemble forward using model equations of the system behaviour. In essence, this is nominally equivalent to running many instances of the Kalman Filter in parallel.

When an ‘observation’ is available, it can be optimally assimilated into the ensemble by applying the standard Kalman Filter equations. With a sufficiently large ensemble (determined through experimentation), the mean state should represent the maximum likelihood value for the process at the time. The evolution of the ensemble can be explored by examining the ‘spread’ of the states about the mean.

There are three stages required to implement the EnKF for this problem: (1) generation of the initial ensemble, (2) forecasting the change of the field by driving the field model with SV predicted by core flow models and (3) assimilation of measurements e.g. from a ‘true’ field model. In our implementation (full details of the algorithm are given in Beggan and Whaler (2009)), the state ensemble ( $\mathbf{x}_k$ ) is composed of a matrix of Gauss coefficients up to degree and order 12. The number of ensemble states was set to 1000 after experimentation, though it was found that any more

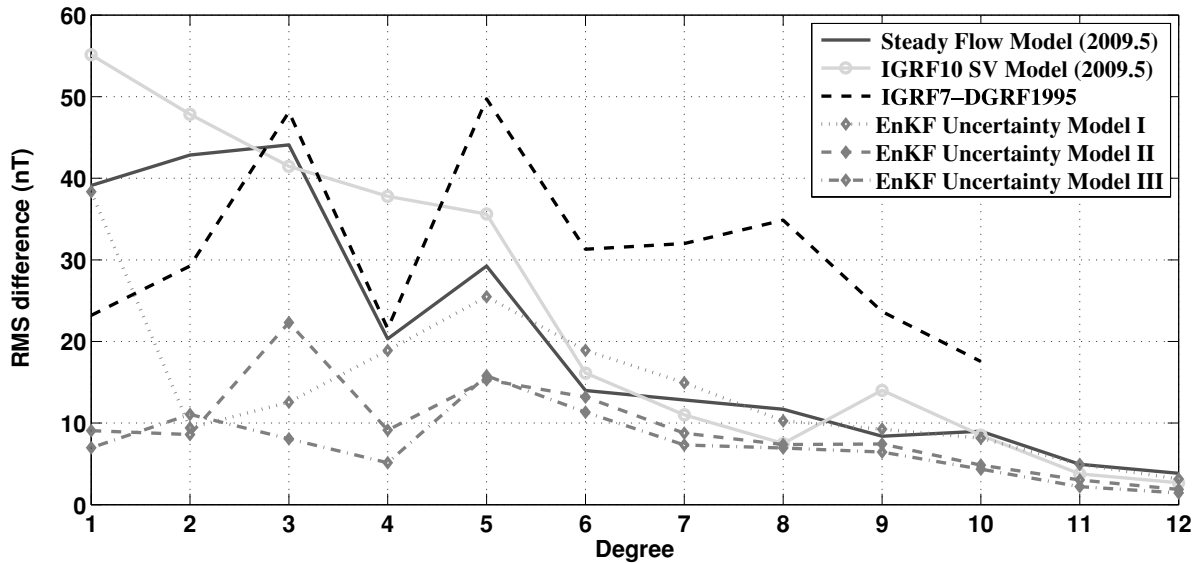


Fig. 3. Root Mean Square difference per degree (defined in Eq. (5)) to the CHAOS-2 model (at 2009.5) of the field predicted by the IGRF-10 SV model, the Steady Flow SV model and three EnKF simulations with different errors assumptions as discussed in Section 4.3. Also shown is the RMS difference per degree between the IGRF-7 and DGRF field models for 1995.0. See text for details.

than 500 is adequate.

The ensemble is initiated by perturbing the Gauss coefficients of the CHAOS-2 field model in 2004.5. The initial perturbation to the  $\mathbf{g}_n^m$  field coefficients is based on an estimate of the standard deviation for each coefficient, discussed in detail in the next section. A matrix of normally distributed random numbers  $N(0, 1)$  is generated, which is then multiplied by the standard deviation of the flow coefficients to give a perturbed flow coefficient matrix. This perturbed flow coefficient matrix is pre-multiplied by the  $\mathbf{H}$  matrix to produce a matrix of perturbed SV coefficients, correctly scaled to reflect the uncertainty in the flow models. The perturbed SV coefficient matrix is then added to the initial state vector to produce the initial ensemble matrix. Once this initial ensemble has been created, forecasting and assimilation can take place.

The forecast (prediction) of the field is driven forwards by the summation of (1) the field coefficients and (2) the monthly SV coefficients from the flow model. In addition, at each timestep, model noise is added to simulate the variance of the ensemble, forcing it to grow at each forecast iteration. These steps are repeated until a measurement becomes available for assimilation into the ensemble.

Over time, the forecast field will begin to diverge from the actual field. To improve the forecast, data can be input into the ensemble to update (correct) it. The data have associated errors which are used to generate a perturbed data ensemble. Data, for example a set of Gauss coefficients ( $\mathbf{z}_k$ ) with a certain (estimated or known) error are available. A matrix of zero-mean Gaussian random numbers is generated and scaled with the data error. The data are then added to the matrix of scaled random numbers to produce a matrix of ‘perturbed data’. Using Eq. (3) this data perturbation matrix and the perturbed SV coefficients are optimally assimilated into the ensemble at this timestep. The covariance matrices ( $\mathbf{P}$  and  $\mathbf{Q}$ ) can be estimated from the ensemble and measurement errors (Evensen, 1994). In this study

we assimilate low quality (i.e. large assumed variance; see Section 4.3) Gauss coefficients of the CHAOS-2 field model into the ensemble via the  $\mathbf{z}_k$  term.

#### 4.2 Uncertainty estimation

The determination of realistic uncertainties to apply to the model coefficients of the field and the SV from the flow is difficult. For a Kalman Filter to be optimal it is important that the uncertainty estimate of the forecast field generated by the flow model and the uncertainty of the ‘measured’ field model should be similar in magnitude. If one has a much lower uncertainty estimate than the other, then it tends to dominate the filter.

One approach to estimating the uncertainty is to examine which field coefficients have in the past been most poorly predicted at the end of a five year period. As noted in Section 3, the field model prediction from 2004.5 to 2009.5 of the IGRF-10 SV model has a RMS difference of approximately 102 nT, while the RMS difference of the steady flow model prediction is 85 nT. We can investigate the difference in more detail by defining the RMS difference per degree ( $\sqrt{dP_n}$ ) as:

$$dP_n = \sum_{m=0}^n (n+1) [(g_n^m)_{\text{Model A}} - (g_n^m)_{\text{Model B}}]^2 \quad (5)$$

Figure 3 shows the RMS difference per degree to the CHAOS-2 model (in 2009.5) of the IGRF-10 SV prediction for five years and the steady flow model SV prediction. The RMS difference per degree shows that the field model forecast by the IGRF-10 SV prediction (grey solid line, circle marker) differs strongly from the CHAOS-2 model at degrees 1–5. In contrast, the field model forecast by the steady flow model (black solid line) is poorest at degree 3 (coefficients  $h_3^3$  and  $h_3^1$ ), but better at degrees 1, 2 and 4–6.

Prior to the launch of Ørsted in 1999, main field IGRF models were computed primarily from ground-based observatory data. To determine where the largest errors

between satellite field models and models derived from ground-based measurements arise, we examined the difference between IGRF-7 model (valid 1995–2000) coefficients as defined for 1995 and the DGRF model for 1995 from Chambodut *et al.* (2005). The DGRF model was calculated from the backward projection of high quality satellite data from 2000 onwards, though it should be noted that the back-propagation is imperfect. The RMS difference between the IGRF-7 model and the revised DGRF is 104 nT (up to degree 10) and we have plotted the RMS difference per degree between models in Fig. 3 (dashed black line), again up to degree 10. The IGRF-7 model is relatively good at degrees 1, 2 and 4 but poorer at degrees 3 and 5. Whether this pattern is peculiar to the IGRF-7 model or applies in general to field models derived mainly from ground-based data is not known.

We can now use the absolute differences of the Gauss coefficients between the IGRF-7 and DGRF models (dashed black line) as reasonable estimates for the noise perturbations to apply to an assimilated ‘measurement’ within the EnKF. Similarly, the differences in the Gauss coefficients between the field forecast by the SV from the steady flow model and CHAOS-2 (solid black line) can be used as estimates for the variance of the forecast. The absolute differences are relatively similar in magnitude and somewhat complementary—the errors of the IGRF-7 field model are lower at degree 1 and 2 and similar for degree 3 and 4, while the errors of the flow model SV forecast are lower at degrees 5–12. As the IGRF-7 errors only extend to degree 10, the errors for coefficients of degrees 11 and 12 were set to a value of 0.5 nT. (This is the median of the absolute value of the degree 11 and 12 coefficients of IGRF-10).

In the EnKF, random perturbations to the ensemble states can now be controlled using the absolute differences between the coefficients of IGRF-7 and DGRF for 1995 as the expected standard deviations. Each state coefficient (or Gauss coefficient for this study i.e.  $x_i^r$ ) perturbation is thus drawn from a normal distribution  $N(0, 1)$  and scaled by the respective error. We explore the effect of using three different uncertainty assumptions for the ‘measured’ field model input to the filter.

### 4.3 EnKF simulation results

A steady flow model (with a tangential geostrophic constraint) was generated from SV data over the period 2001.4–2004.5. The CHAOS-2 model for 2004.5 was used as the starting point for the EnKF field ensemble with each ensemble state propagated forward with a timestep equal to one-twelfth of a year using a set of perturbed SV coefficients from the steady flow model. The ensemble states were forecast forward in time using:

$$\mathbf{g}_{t+1}^r = \mathbf{g}_t^r + ((\mathbf{H}_t \hat{\mathbf{m}}_{\text{SF}}) + N(0, 1)_t^r * \sigma_n^m) / 12 \quad (6)$$

where  $\mathbf{g}_t^r$  are the Gauss coefficients of ensemble state  $r$  at timestep  $t$ . The values of  $\sigma_n^m$  are the estimated uncertainties, up to degree and order  $n = m = 12$ , computed from the difference between the steady flow model prediction and the CHAOS-2 model at 2009.5.

At every twelfth timestep (i.e. annually), a set of noisy Gauss coefficients was assimilated into the system using Eqs. (3) and (4). The Gauss coefficients, initially generated

from the CHAOS-2 model, have Gaussian noise added with a standard deviation derived from the RMS difference of the IGRF-7 and DGRF1995 models to simulate the expected uncertainty of a ground-based field model. The results of three different uncertainty assumptions for the assimilated field model were tested. (Note, the EnKF uncertainty models are given roman numerals to avoid confusion with the IGRF-11 candidate models).

Figure 4 shows the RMS difference between the CHAOS-2 model and three Ensemble Kalman Filter noise simulations. In each panel, the thin green lines represent the evolution of 1000 individual field model ensemble states over 61 months. The solid black line represents the mean of the ensemble states, the dashed line is the  $+1\sigma$  field model and the dot-dashed line is the  $-1\sigma$  field model (computed from the mean of the ensemble). In the upper panel, uncertainty model I, the forecast model and ‘measurement’ field model errors are similar in magnitude. The mean RMS difference at 2009.5 is 58 nT. Figure 4 (upper panel, I) illustrates the spread of the ensemble states but shows that the mean of the ensemble is the best fit model overall (though two of the ensemble states do have a slightly lower misfit at 2009.5). The RMS difference per degree of the mean of the ensemble to the CHAOS-2 model is shown in Fig. 3 (dotted red line, diamond marker). The overall difference per degree is, on average, smaller at 2009.5 than either of the input errors (from the IGRF-7 minus DGRF1995 model or the CHAOS-2 model minus the Steady Flow model at 2009.5). This illustrates one of the strengths of a Kalman filter—the assimilation of ‘forecast’ and ‘measurement’ can improve upon both, by optimally combining the best parts of each, provided the uncertainties are reasonably complementary.

We investigated the outcome of using better field models in the assimilation by reducing the noise added to the CHAOS-2 field coefficients. In uncertainty model II, the ‘measured’ field model errors were sampled from a distribution with spread approximately half as large as the field model errors in I (i.e. the standard deviations were divided by 2). The RMS difference to the CHAOS-2 model in 2009.5 is 47 nT. In uncertainty model III the field model errors are approximately a quarter the size of the model errors in I, with RMS difference to CHAOS-2 in 2009.5 of 34 nT. Thus, as the field model improves, the RMS difference becomes smaller, as would be expected. The RMS difference per degree for models II and III are also shown in Fig. 3 (dashed red line and dash-dot red line with diamond markers, respectively). Due to the random nature of the added noise, some degrees are better fit than others. For example, degree 3 is better fit in model I than model II, despite the average uncertainty being smaller for the latter. At higher degrees ( $n = 11$  and  $12$ ), the difference of all the models to the CHAOS-2 model is roughly equal.

Reducing the expected error of the forecast model is not really meaningful as we cannot control the goodness of fit of the forecast in this study. Note also that as each simulation is a random realisation of the errors, the results are, of course, indicative of the potential improvements possible.

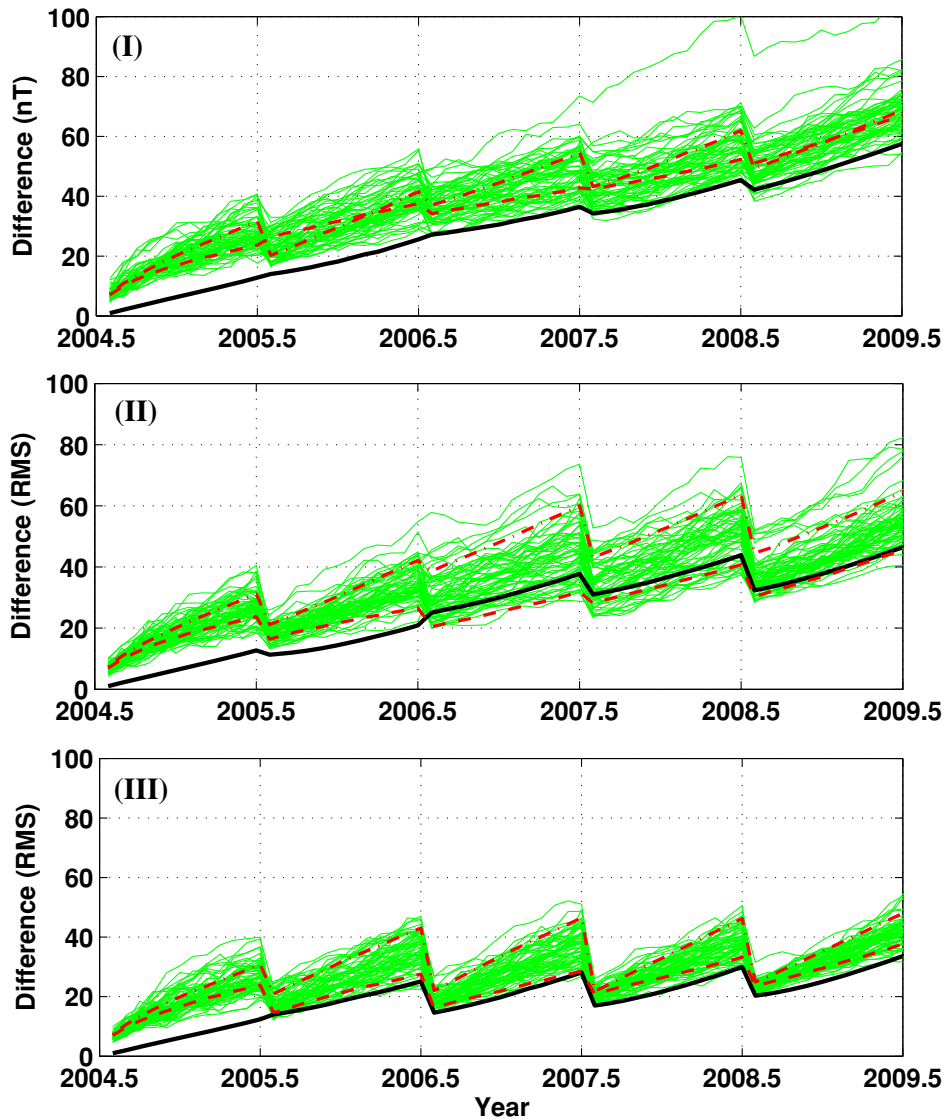


Fig. 4. Root Mean Square difference to the CHAOS-2 model of three Ensemble Kalman Filter simulations. The thin green lines represent the evolution of 1000 individual field model ensemble states over 61 months. The solid black line represents the mean of the ensemble states, the (lower red) dashed line is the  $+1\sigma$  field model and the (upper red) dot-dashed line is the  $-1\sigma$  field model as calculated from the mean of the ensemble. Model I assumes that the forecast model uncertainties and the assimilated field model uncertainties are set to be approximately similar in magnitude. Model II has assimilated field model uncertainties that are approximately one-half the size of model I. The assimilated field model uncertainties for model III are approximately one-quarter the size of those in model I.

## 5. Conclusions

The RMS difference between the IGRF-10 SV prediction for 2009.5 starting from a model of the field for 2004.5 is estimated to be approximately 102 nT. Thus, the average annual RMS misfit between the IGRF-10 model and the ‘true’ magnetic field was about 21 nT. Using a steady flow model generated from SV data prior to 2004.5 to predict the SV for a similar period of time (2004.5–2009.5) resulted in an average RMS difference of approximately 17 nT/yr. This suggests improvement in the SV prediction may be possible. While there is a large variation between proposed IGRF-11 SV candidate models, the coefficients derived from the steady flow model are within the range of most of the candidate models, indicating it is not an unreasonable estimate.

The RMS difference between the DGRF1995 field model (derived from the backward projection of satellite data) and

the IGRF-7 model for 1995 (derived primarily from ground-based observatories) is 104 nT. In a future scenario where only a relatively poor field model is available, we show that it is possible to improve the prediction using data assimilation with a field model derived from a steady core flow forecast. If a high-quality starting field model is available with only lower-quality updates, then the combination of forecasts from a core flow model and lower quality field models using an Ensemble Kalman filter can reduce the RMS errors of the resulting forecast. The potential improvements are strongly dependent on the relative uncertainties of each model.

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## References

- Beggan, C. D. and K. Whaler, Forecasting change of the magnetic field using core surface flows and ensemble Kalman filtering, *Geophys. Res. Lett.*, **36**, L18303, 2009.
- Chambodut, A., B. Langlais, and M. Manda, Candidate main-field models for the Definitive Geomagnetic Reference Field 1995.0 and 2000.0, *Earth Planets Space*, **57**, 1197–2002, 2005.
- Evensen, G., Sequential data assimilation with a nonlinear quasi-geostrophic model using Monte Carlo methods to forecast error statistics, *J. Geophys. Res.*, **99**, 10,143–10,162, 1994.
- Finlay, C. C., S. Maus, C. D. Beggan, M. Hamoudi, F. J. Lowes, N. Olsen, and E. Thébault, Evaluation of candidate geomagnetic field models for IGRF-11, *Earth Planets Space*, **62**, this issue, 787–804, 2010.
- Fournier, A., C. Eymin, and T. Alboussiere, A case for variational geomagnetic data assimilation: insights from a one-dimensional, nonlinear, and sparsely observed MHD system, *Nonlin. Proc. Geophys.*, **14**, 163–180, 2007.
- Gillet, N., A. Pais, and D. Jault, Ensemble inversion of time-dependent core flow models, *Geochem. Geophys. Geosyst.*, **10**, Q06004, 2009.
- Gubbins, D., Geomagnetic field analysis—II. Secular variation consistent with a perfectly conducting core, *Geophys. J. R. Astron. Soc.*, **77**, 753–766, 1984.
- Halley, E., On the cause of the change in the variation of the magnetic needle; with an hypothesis of the structure of the internal parts of the Earth, *Phil. Trans. R. Soc. Lond.*, **17**, 470–478, 1692.
- Hills, R., Convection in the Earth's mantle due to viscous shear at the core-mantle interface and due to large-scale buoyancy, Ph.D. thesis, N. M. State Univ., Las Cruces, 1979.
- Holme, R., Large scale flow in the core, in *Treatise on Geophysics*, Vol. 8, 107–130, Elsevier, 2007.
- Kalman, R., A new approach to linear filtering and prediction problems, *Trans. ASME J. Basic Eng.*, **82**, 35–45, 1960.
- Kuang, W., A. Tangborn, Z. Wei, and T. Sabaka, Constraining a numerical geodynamo model with 100 years of surface observations, *Geophys. J. Int.*, **179**, 1458–1468, 2009.
- Le Mouél, J.-L., Outer-core geostrophic flow and secular variation of Earth's geomagnetic field, *Nature*, **311**, 734–735, 1984.
- Lowes, F. J., An estimate of the errors of the IGRF/DGRF fields 1945–2000, *Earth Planets Space*, **52**, 1207–1211, 2000.
- Macmillan, S. and S. Maus, International Geomagnetic Reference Field—the tenth generation, *Earth Planets Space*, **57**, 1135–1140, 2005.
- Manda, M. and N. Olsen, A new approach to directly determine the secular variation from magnetic satellite observations, *Geophys. Res. Lett.*, **33**, L15306, 2006.
- Maus, S., S. Macmillan, F. Lowes, and T. Bondar, Evaluation of candidate geomagnetic field models for the 10th generation of IGRF, *Earth Planets Space*, **57**, 1173–1181, 2005.
- Maus, S., M. Rother, C. Stolle, W. Mai, S. Choi, H. Lühr, D. Cooke, and C. Roth, Third generation of the Potsdam Magnetic Model of the Earth (POMME), *Geochem. Geophys. Geosyst.*, **7**, Q07008, 2006.
- Maus, S., L. Silva, and G. Hulot, Can core-surface flow models be used to improve the forecast of the Earth's main magnetic field?, *J. Geophys. Res.*, **113**, B08102, 2008.
- Olsen, N. and M. Manda, Investigation of a secular variation impulse using satellite data: the 2003 geomagnetic jerk, *Earth Planet. Sci. Lett.*, **255**, 94–105, 2007.
- Olsen, N., H. Lühr, T. Sabaka, M. Manda, M. Rother, and L. Toffner-Clausen, CHAOS: a model of the Earth's magnetic field derived from CHAMP, Oersted, and SAC-C magnetic satellite data, *Geophys. J. Int.*, **166**, 67–75, 2006.
- Olsen, N., M. Manda, T. J. Sabaka, and L. Toffner-Clausen, CHAOS-2—a geomagnetic field model derived from one decade of continuous satellite data, *Geophys. J. Int.*, **179**, 1477–1487, 2009.
- Roberts, P. and S. Scott, On the analysis of the secular variation. 1. A hydromagnetic constraint: Theory, *J. Geomag. Geoelectr.*, **17**, 137–151, 1965.
- Sun, Z., A. Tangborn, and W. Kuang, Data assimilation in a sparsely observed one-dimensional modeled MHD system, *Nonlin. Proc. Geophys.*, **14**, 181–192, 2007.
- Wardinski, I., R. Holme, S. Asari, and M. Manda, The 2003 geomagnetic jerk and its relation to the core surface flows, *Earth Planet. Sci. Lett.*, **267**, 468–481, 2008.
- Whaler, K., Geomagnetic evidence for fluid upwelling at the core-mantle boundary, *Geophys. J. R. Astron. Soc.*, **86**, 563–588, 1986.

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C. Beggan (e-mail: ciar@bgs.ac.uk) and K. Whaler