

A survey of time consistency of dynamic risk measures and dynamic performance measures in discrete time: LM-measure perspective



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Abstract In this work we give a comprehensive overview of the time consistency property of dynamic risk and performance measures, focusing on a the discrete time setup. The two key operational concepts used throughout are the notion of the LM-measure and the notion of the update rule that, we believe, are the key tools for studying time consistency in a unified framework.

Keywords Time consistency · Update rule · Dynamic LM-measure · Dynamic risk measure · Dynamic acceptability index · Measure of performance

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“The dynamic consistency axiom turns out to be the heart of the matter.”

A. Jobert and L. C. G. Rogers

Valuations and dynamic convex risk measures, *Math Fin* 18(1), 2008, 1–22.

Introduction

The goal of this work is to give a comprehensive overview of the time consistency property of dynamic risk and performance measures. We focus on discrete time setup, since most of the existing literature on this topic is dedicated to this case.

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The time consistency surveyed in this paper is related to dynamic decision making subject to various uncertainties that evolve in time. Typically, decisions are made subject to the decision maker's preferences, which may change in time and thus they need to be progressively assessed as an integral part of the decision making process. Naturally, the assessment of preferences should be done in such a way that the future preferences are assessed consistently with the present ones. This survey is focusing on this aspect of time consistency of a dynamic decision making process.

Traditionally, in finance and economics, the preferences are aimed at ordering cash and/or consumption streams. A convenient way to study preferences is to study them via numerical representations, such as (dynamic) risk measures, (dynamic) performance measures, or, more generally, dynamic LM-measures¹ (Bielecki et al. 2014a). Consequently, the study of the time consistency of preferences is also conveniently done in terms of their numerical representations. This work is meant to survey various approaches to modelling and analysis of the time consistency of numerical representations of preferences.

As stated above, the objects of our survey—the dynamic LM-measures—are meant to “put a preference order” on the sets of underlying entities. There exists a vast literature on the subject of preference ordering, with various approaches towards establishing an order of choices, such as the decision theory or the expected utility theory, that trace their origins to the mid 20th century. We focus our attention, essentially, on the axiomatic approach to defining risk or performance measures.

The axiomatic approach to measuring risk of a financial position was initiated in the seminal paper by Artzner et al. (1999), and has been going through a flourishing development since then. The measures of risk introduced in (Artzner et al. 1999), called coherent risk measures, were meant to determine the regulatory capital requirement by providing a numerical representation of the riskiness of a portfolio of financial assets. In this framework, from mathematical point of view, the financial positions are understood as either discounted terminal values (payoffs) of portfolios, that are modeled in terms of random variables, or they are understood as discounted dividend processes, cumulative or bullet, that are modeled as stochastic processes. Although stochastic processes can be viewed as random variables (on appropriate spaces), and vice versa - random variables can be treated as particular cases of processes—it is convenient, and in some instances necessary, to treat these two cases separately—the road we are taking in this paper.

In the paper (Artzner et al. 1999), the authors considered the case of random variables, and the risk measurement was done at time zero only. This amounts to considering a one period time model in the sense that the measurement is done today of the cash flow that is paid at some fixed future time (tomorrow). Accordingly, the related risk measures are referred to as static measures. Since then, two natural paths were followed: generalizing the notion of risk measure by relaxing or changing the set of axioms, or/and considering a dynamic setup. By dynamic setup we mean that the measurements are done throughout time and are adapted to the flow of available

¹An LM-measure is a function that is *local* and *monotone*; see Definition 1. These two properties have to be satisfied by any reasonable dynamic measure of performance and/or measure of risk, and are shared by most such measures in the existing literature.

information. In the dynamic setup, both discrete and continuous time evolutions have been studied, for both random variables and stochastic processes as the inputs. In the present work, we focus our attention on the discrete time setup, although we briefly review the literature devoted to continuous time.

This survey is organized as follows. We start with the literature review relevant to the dynamic risk and performance measures focusing on the time consistency property in the discrete time setup. In Section “Mathematical preliminaries”, we set the mathematical scene; in particular, we introduce the main notations used in this paper and the notion of LM-measures. Section “Approaches to time consistent assessment of preferences” is devoted to the time consistency property. There we discuss two generic approaches to time consistent assessment of preferences and point out several idiosyncratic approaches. We put forth in this section the notion of an update rule that, we believe, is the key tool for studying time consistency in a unified framework. Sections “Time consistency for random variables” and “Time consistency for stochastic processes” survey some concepts and results regarding time consistency in the case of random variables and in the case of stochastic processes, respectively. Our survey is illustrated by numerous examples that are presented in Section “Examples”. We end the survey with two appendices. In “Appendix” we provide a brief exposition of the three fundamental concepts used in the paper: the dynamic LM-measures, the conditional essential suprema/infima, and LM-extensions. Finally, in Appendix “Proofs” we collect proofs of several results stated throughout our survey.

Literature review

The aim of this section is to give a chronological survey of the developments of the theory of dynamic risk and performance measures. Although it is not an obvious task to establish the exact lineup, we tried our best to account for the most relevant works according to adequate chronological order.

We trace back the origins of the research regarding time consistency to Koopmans (1960) who put on the precise mathematical footing, in terms of the utility function, the notion of persistency over time of the structure of preferences.

Subsequently, in the seminal paper, Kreps and Porteus (1978) treat the time consistency at a general level by axiomatising the “choice behavior” of an agent by taking into account how choices at different times are related to each other; in the same work, the authors discuss the motivations for studying the dynamic aspect of choice theory.

Before we move on to reviewing the works on dynamic risk and performance measures, it is worth mentioning that the robust expected utility theory proposed by Gilboa and Schmeidler (1989) can be viewed as a more comprehensive theory than the one discussed in (Artzner et al. 1999); we refer to (Roorda et al. 2005) for the relevant discussion.

Starting with (Artzner et al. 1999), the axiomatic theory of risk measures, understood as functions mapping random variables into real numbers, was developing around the following main goals: a) to define a set of properties (or axioms) that a risk measure should satisfy; b) to characterize all functions that satisfy these

properties; c) provide particular examples of such functions. Each of the imposed axioms should have a meaningful financial or actuarial interpretation. For example, in (Artzner et al. 1999), a static coherent risk measure is defined as a function $\rho : L^\infty \rightarrow [-\infty, \infty]$ that is monotone decreasing (larger losses imply larger risk), cash-additive (the risk is reduced by the amount of cash added to the portfolio today), sub-additive (a diversified portfolio has a smaller risk) and positive homogenous (the risk of a rescaled portfolio is rescaled correspondingly), where L^∞ is the space of (essentially) bounded random variables on some probability space $(\Omega, \mathcal{F}, P)^2$. The descriptions or the representations of these functions, also called robust representations, usually are derived via duality theory in convex analysis, and are necessary and sufficient in their nature. Traditionally, among such representations we find: representations in terms of the level or the acceptance sets; numerical representations in terms of the dual pairings (e.g., expectations). For example, the coherent risk measure ρ mentioned above can be described in terms of its acceptance set $\mathcal{A}_\rho = \{X \in L^\infty \mid \rho(X) \leq 0\}$. As it turns out, the acceptance set \mathcal{A}_ρ satisfies certain characteristic properties, and any set \mathcal{A} with these properties generates a coherent risk measure via the representation $\rho(X) = \inf\{m \in \mathbb{R} \mid m + X \in \mathcal{A}\}$. Alternatively, the function ρ is a coherent risk measure if and only if there exists a nonempty set \mathcal{Q} of probability measures, absolutely continuous with respect to P , such that

$$\rho(X) = - \inf_{Q \in \mathcal{Q}} E^Q[X]. \quad (1)$$

The set \mathcal{Q} can be viewed as a set of generalized scenarios, and a coherent risk measure is equal to the worst expected loss under various scenarios. By relaxing the set of axioms, the static coherent risk measures were generalized to static convex risk measures and to an even more general class called monetary risk measures. See, for instance, (Szegö 2002) for a survey of static risk measures, as well as (Cheridito and Li 2009, 2008). On the other hand, axiomatic theory of performance measures was originated in (Cherny and Madan 2009). A general theory of risk preferences and their robust representations, based on only two generic axioms, was studied in (Drapeau 2010, Drapeau and Kupper 2013).

Moving to the dynamic setup, we first introduce an underlying filtered probability space $(\Omega, \mathcal{F}, \{\mathcal{F}_t\}_{t \geq 0}, P)$, where the increasing collection of σ -algebras \mathcal{F}_t , $t \geq 0$, models the flow of information that is accumulated through time.

Artzner et al. (2002b) and (Artzner et al. 2002a) study an extension of the static models examined in (Artzner et al. 1999) to the multiperiod case, assuming discrete time and discrete probability space. The authors proposed a method of constructing dynamic risk measures $\{\rho_t : L^\infty(\mathcal{F}_T) \rightarrow \bar{L}^0(\mathcal{F}_t), t = 0, 1, \dots, T\}$, by a backward recursion, starting with $\rho_T(X) = -X$, and letting

$$\rho_t(X) = - \inf_{Q \in \mathcal{Q}} E^Q[-\rho_{t+1}(X) \mid \mathcal{F}_t], \quad 0 \leq t < T, \quad (2)$$

where, as before, \mathcal{Q} is a set of probability measures. If, additionally, \mathcal{Q} satisfies a property called recursivity or consistency (cf. (Riedel 2004)), namely

²In the original paper (Artzner et al. 1999), the authors considered finite probability spaces, but later the theory was elevated to a general probability space (Delbaen 2000, 2002).

$$\inf_{Q \in \mathcal{Q}} E^Q[Z | \mathcal{F}_t] = \inf_{Q \in \mathcal{Q}} E^Q[\inf_{Q_1 \in \mathcal{Q}} E^{Q_1}[Z | \mathcal{F}_{t+1}] | \mathcal{F}_t], \quad t = 0, 1, \dots, T-1, Z \in L^\infty, \tag{3}$$

then one can show that (2) is equivalent to

$$\rho_t(X) = \rho_t(-\rho_{t+1}(X)), \quad 0 \leq t < T, X \in L^\infty(\mathcal{F}_T). \tag{4}$$

The property (4) represents what has become known in the literature as the *strong time consistency* property. For example, if $\mathcal{Q} = \{P\}$, then the strong time consistency reduces to the tower property for conditional expectations. From a practical point of view, this property essentially means that assessment of risks propagates in a consistent way over time: assessing at time t future risk, represented by random variable X , is the same as assessing at time t a risky assessment of X done at time $t + 1$ and represented by $-\rho_{t+1}(X)$. Additionally, the property (4) is closely related to the Bellman principle of optimality or to the dynamic programming principle (see, for instance, (Bellman and Dreyfus 1962; Carpentier et al. 2012).

Delbaen (2006) studies the recursivity property in terms of m -stable sets of probability measures, and also describes the time consistency of dynamic coherent risk measures in the context of martingale theory. The recursivity property is equivalent to properties known as time consistency and the rectangularity in the multi-prior Bayesian decision theory. Epstein and Schneider (2003) study time consistency and rectangularity property in the framework of “decision under ambiguity.”

It needs to be said that several authors refer to (Wang 2002) for an alternative axiomatic approach to time consistency of dynamic risk measures.

The first study of dynamic risk measures for stochastic processes (finite probability space and discrete time) is attributed to Riedel (2004), where the author introduced the (strong) time consistency as one of the axioms. If $\rho_t, t = 0, \dots, T$, is a dynamic coherent risk measure, acting on the set of discounted terminal cash flows³, then ρ is *strongly time consistent* if the following implication holds true:

$$\rho_{t+1}(X) = \rho_{t+1}(Y) \Rightarrow \rho_t(X) = \rho_t(Y). \tag{5}$$

This means that if tomorrow we assess the riskiness of X and Y at the same level, then today X and Y must have the same level of riskiness. It can be shown that for dynamic coherent risk measures, or more generally for dynamic monetary risk measures, property (5) is equivalent to (4).

Motivated by results regarding the pricing procedure in incomplete markets, based on use of risk measures, Roorda et al. (2005) study dynamic coherent risk measures (for the case of random variables on finite probability space and discrete time) and introduce the notion of (strong) time consistency; note that their work was similar and contemporaneous to (Riedel 2004). They show that strong time consistency entails recursive computation of the corresponding optimal hedging strategies. Moreover, time consistency is also described in terms of the collection of probability measures that satisfy the “product property,” similar to the rectangularity property mentioned above.

³In (Riedel 2004), the author considered discounted dividend processes, but for simplicity here we write the time consistency for random variables.

Similarly, as in the static case, the dynamic coherent risk measures were extended to dynamic convex risk measures by replacing sub-additivity and positive homogeneity properties with convexity. In the continuous time setup, Rosazza Gianin (2002) links dynamic convex risk measures to nonlinear expectations or g -expectations, and to Backward Stochastic Differential Equations (BSDEs). Strong time consistency plays a crucial role and, in view of (4), it is equivalent to the tower property for conditional g -expectations. These results are further studied in a sequel of papers (Frittelli and Rosazza Gianin 2004; Peng 2004; Rosazza Gianin 2006), as well as in Coquet et al. (2002).

A representation similar to (1) holds true for dynamic convex risk measure

$$\rho_t(X) = - \inf_{Q \in \mathcal{M}(P)} \left(E^Q[X | \mathcal{F}_t] + \alpha_t^{\min}(Q) \right), \quad t = 0, 1, \dots, T, \quad (6)$$

where $\mathcal{M}(P)$ is the set of all probability measures absolutely continuous with respect to P , and α^{\min} is the minimal penalty function.⁴ The natural question of describing (strong) time consistency in terms of properties of the minimal penalty functions was studied by Scandolo (2003). Also in (Scandolo 2003), the author discusses the importance in the dynamic setup of the special property called locality. It should be mentioned that locality property was part of the earlier developments in the theory of dynamic risk measures. For example, it was called dynamic relevance axiom in (Riedel 2004), and zero-one law in (Peng 2004). Similarly to previous studies, (Scandolo 2003) finds a relationship between time consistency, the recursive construction of dynamic risk measures, and the supermartingale property. These results are further investigated in Detlefsen and Scandolo (2005). Also in these works, it was shown that the dynamic entropic risk measure is a strongly time consistent convex risk measure.

Weber (2006) continues the study of dynamic convex risk measures for random variables in a discrete time setup and introduces weaker notions of time consistency *acceptance and rejection time consistency*. Mainly, the author studies the law invariant risk measures, and characterizes time consistency in terms of the acceptance indicator $a_t(X) = \mathbb{1}_{\rho_t(X) \leq 0}$ and in terms of the acceptance sets of the form $\mathcal{N}_t = \{X \mid \rho_t(X) \leq 0\}$. Along the same lines, Föllmer and Penner (2006) investigate the dynamic convex risk measures, representation of strong time consistency as a recursivity property, and they relate it to the Bellman principle of optimality. They also prove that the supermartingale property of the penalty function corresponds to the weak or acceptance/rejection time consistency. Moreover, the authors study the co-cycle property of the penalty function for the dynamic convex risk measures that admit robust representation (see Definition 50).

Artzner et al. (2007) continue to study the strong time consistency for dynamic risk measures, its equivalence with the stability property of test probabilities and with the optimality principle.

⁴See Appendix “Robust representations for dynamic monetary utility measures” for the definition of minimal penalty functions, up to a sign, and for the corresponding robust representations.

It is worth mentioning that Bion-Nadal (2004) studies dynamic monetary risk measures in a continuous time setting and their time consistency property in the context of model uncertainty when the class of probability measures is not specified.

Motivated by optimization subject to risk criterion, Ruszczyński and Shapiro (2006a) elevate the concepts from (Ruszczyński and Shapiro 2006b) to the dynamic setting, with the main goal to establish conditions under which the dynamic programming principle holds.

Cheridito and Kupper (2011) introduce the notion of aggregators and generators for dynamic convex risk measures and give a thorough discussion about the composition of time-consistent convex risk measures in the discrete time setup, for both random variables and stochastic processes. They link time consistency to one step dynamic penalty functions. In this regard, we also refer to (Cheridito et al. 2006; Cheridito and Kupper 2009).

Jobert and Rogers (2008) take the valuation concept as the starting point, rather than the dynamics of acceptance sets, with the valuation functional being the negative of a risk measure. To quote the authors (strong) “time consistency is the heart of the matter.” Kloeppel and Schweizer (2007) use dynamic convex risk measures for valuation in incomplete markets, where the time consistency plays a key role. Cherny (2007) uses dynamic coherent risk measure for pricing and hedging European options; see also (Cherny and Madan 2006).

Roorda and Schumacher (2007) study the weak form of time consistency for dynamic convex risk measure.

Bion-Nadal (2006) continues to study various properties of dynamic risk measures, both in discrete and in continuous time, mainly focusing on the composition property mentioned above, and thus on the strong time consistency. The composition property is characterized in terms of stability of probability sets. The author defines the *co-cycle condition for the penalty function* and shows its equivalence to strong time consistency. In the followup paper, (Bion-Nadal 2008), the author continues to study the characterization of time consistency in terms of the co-cycle condition for minimal penalty function. For further related developments in the continuous time framework see (Bion-Nadal 2009b).

Observing that Value at Risk ($V@R$) is not strongly time consistent, Boda and Filar (2006), and Cheridito and Stadje (2009) construct a strongly time consistent alternative to $V@R$ by using a recursive composition procedure.

Tutsch (2008) gives a different perspective on time consistency of convex risk measures by introducing the update rules⁵ and generalizes the strong and weak form of time consistency via test sets.

The theory of dynamic risk measures finds its application in areas beyond the regulatory capital requirements. For example, Cherny (2010) applies dynamic coherent risk measures to risk-reward optimization problems and in (Cherny 2009) to capital allocation; Bion-Nadal (2009a) uses dynamic risk measures for time consistent pricing; Barrieu and El Karoui (2004; 2005; 2007) study optimal derivatives design under dynamic risk measures; Geman and Ohana (2008) explore the time consistency

⁵In the present manuscript, we also use the name ‘update rules’, although the concept used here is different from that introduced in (Tutsch 2008).

in managing a commodity portfolio via dynamic risk measures; Zariphopoulou and Zitkovic (2010) investigate the maturity independent dynamic convex risk measures.

In Delbaen et al. (2010), the authors establish a representation of the penalty function of dynamic convex risk measure using g -expectation and its relation to the strong time consistency.

There exists a significant literature on a special class of risk measures that satisfy the law invariance property. Kupper and Schachermayer (2009) prove that the only relevant, law invariant, strongly time consistent risk measure is the entropic risk measure.

For a fairly general study of dynamic convex risk measures and their time consistency we refer to (Bion-Nadal and Kervarec 2012) and (Bion-Nadal and Kervarec 2010). Acciaio et al. (2012) give a comprehensive study of various forms of time consistency for dynamic convex risk measures in a discrete time setup. This includes strong and weak time consistency, representations of time consistency in terms of acceptance sets, and the supermartingale property of the penalty function. We would like to point out the survey by Acciaio and Penner (2011) of discrete time dynamic convex risk measures. This work deals with (essentially bounded) random variables and examines most of the papers mentioned above from the perspective of the robust representation framework.

Although the connection between BSDEs and the dynamic convex risk measures in a continuous time setting had been established for some time, it appears that Stadje (2010) was the first author to create a theoretical framework for studying dynamic risk measures in discrete time via the Backward Stochastic Difference Equations (BS Δ Es). Due to the backward nature of BS Δ Es, the strong time consistency of risk measures played a critical role in characterizing the dynamic convex risk measures as solutions of BS Δ Es. In a series of papers, Cohen and Elliott further studied the connection between dynamic risk measures and BS Δ Es (Cohen and Elliott 2010, 2011; Elliott et al. 2015).

Föllmer and Penner (2011) developed the theory of dynamic monetary risk measures under Knightian uncertainty, where the corresponding probability measures are not necessarily absolutely continuous with respect to the reference measure. See also Nutz and Soner (2012) for a study of dynamic risk measures under volatility uncertainty and their connection to G -expectations.

From a slightly different point of view, Ruszczyński (2010) studies Markov risk measures, that enjoy strong time consistency, in the framework of risk-averse preferences; see also (Shapiro 2009, 2011, 2012; Fan and Ruszczyński 2014). Some concepts from the theory of dynamic risk measures are adopted to the study of the dynamic programming for Markov decision processes.

In the recent paper, Mastrogiacomo and Rosazza Gianin (2015) provide several forms of time consistency for sub-additive dynamic risk measures and their dual representations.

Finally, we want to mention that during the last decade significant advances were made towards developing a general theory of set-valued risk measures (Hamel and Rudloff 2008; Hamel et al. 2011; Feinstein and Rudloff 2013; Hamel et al. 2013; Feinstein and Rudloff 2015), including the dynamic version of them, where mostly the corresponding form of strong time consistency is considered.

We recall that the main objective of use of risk measures for financial applications is mapping the level of risk of a financial position to a regulatory monetary amount expressed in units of the relevant currency. Accordingly, the key property of any risk measure is cash-additivity $\rho(X - m) = \rho(X) + m$. Clearly, one can think of the risk measures as generalizations of $V@R$.

A concept that is, in a sense, complementary to the concept of risk measures, is that of performance measures, which can be thought as generalizations of the well known Sharpe ratio. In similarity with the theory of risk measures, the development of the theory of performance measures followed an axiomatic approach. This was initiated by Cherny and Madan (2009), where the authors introduced the (static) notion of the coherent acceptability index—a function on L^∞ with values in \mathbb{R}_+ that is monotone, quasi-concave, and scale invariant. As a matter of fact, scale invariance is the key property of acceptability indices that distinguishes them from risk measures, and, typically, acceptability indices are not cash-additive. The dynamic version of coherent acceptability indices was introduced by Bielecki et al. (2014b), for the case of stochastic processes, finite probability space, and discrete time. From now on, we will use as synonyms the terms *measures of performance*, *performance measures*, and *acceptability indices*.

As it turns out, the time consistency for measures of performance is a delicate issue. None of the forms of time consistency, which had been coined for dynamic risk measures, are appropriate for dynamic performance measures. In (Bielecki et al. 2014b), the authors introduce a new form of time consistency that is suitable for dynamic coherent acceptability indices. Let α_t , $t = 0, 1, \dots, T$, be a dynamic coherent acceptability index acting on L^∞ (i.e., discounted terminal cash flows). We say that α is time consistent if the following implications hold true:

$$\begin{aligned} \alpha_{t+1}(X) \geq m_t &\Rightarrow \alpha_t(X) \geq m_t, \\ \alpha_{t+1}(X) \leq n_t &\Rightarrow \alpha_t(X) \leq n_t, \end{aligned} \quad (7)$$

where $X \in L^\infty$, and m_t, n_t are \mathcal{F}_t -measurable random variables. Biagini and Bion-Nadal (2014) study dynamic performance measures in a fairly general setup that generalize the results of (Bielecki et al. 2014b). Later, using the theory of dynamic coherent acceptability indices developed in (Bielecki et al. 2014b), Bielecki et al. (2013) propose a pricing framework, called dynamic conic finance, for dividend paying securities in discrete time. The time consistency property was at the core of establishing the connection between dynamic conic finance and classical arbitrage pricing theory. The static conic finance, that served as motivation for (Bielecki et al. 2013), was introduced in (Cherny and Madan 2010). Finally, in recent papers (Bielecki et al. 2015b, Rosazza Gianin and Sgarra 2013), the authors elevate the notion of dynamic coherent acceptability indices to the case of sub-scale invariant performance measures. For that, BSDEs are used in (Rosazza Gianin and Sgarra 2013) and BSDEs are used in (Bielecki et al. 2015b).

For a general theory of robust representations of quasi-concave maps that covers both dynamic risk measures and dynamic acceptability indices, see (Frittelli and Maggis 2011; Bielecki et al. 2016; Frittelli and Maggis 2014;

Bion-Nadal 2016). Also in (Bielecki et al. 2016), the authors study the strong time consistency of quasi-concave maps via the concept of certainty equivalence; see also (Frittelli and Maggis 2010).

To our best knowledge, (Bielecki et al. 2014a) is the only paper that combines into a unified framework the time consistency for dynamic risk measures and dynamic performance measure. It uses the concept of update rules that serve as a vehicle for connecting preferences at different times. We take the update rules perspective as the main tool for surveying the existing forms of time consistency.

We conclude this literature review by listing works, which in our opinion, are most relevant to this survey (not all of which are mentioned above though).

Dynamic coherent risk measures

- random variables, strong time consistency: Artzner et al. 2002a, 2002b; Roorda 2005.
- stochastic processes, strong time consistency: Riedel 2004; Artzner et al. 2007.

Dynamic convex risk measures,

- random variables, strong time consistency: (discrete time) (Bion-Nadal 2006; Boda and Filar 2006; Cheridito and Stadje 2009; Detlefsen and Scandolo 2005; Föllmer and Penner 2006; Frittelli and Scandolo 2006; Geman and Ohana 2008; Ruzyczyński and Shapiro 2006a; Scandolo 2003), (Acciaio and Penner 2011; Acciaio et al. 2012; Bielecki et al. 2014a; Bielecki et al. 2016; Bion-Nadal 2008; Cheridito and Kupper 2009; Cohen and Elliott 2010, 2011; Elliott et al. 2015; Fasen and Svejda 2012; Iancu et al. 2015; Kupper and Schachermayer 2009; Mastrogiacomo and Rosazza Gianin 2015; Roorda and Schumacher 2015; Stadje 2010); (continuous time) (Barrieu and El Karoui 2004, 2007; Bion-Nadal 2006, 2008, 2009b; Bion-Nadal and Kervarec 2012; Delbaen 2012; Delbaen et al. 2010; Frittelli and Rosazza Gianin 2004; Jiang 2008; Klöppel and Schweizer 2007; Nutz and Soner 2012; Penner and Réveillac 2014; Rosazza Gianin 2002, 2006; Sircar and Sturm 2015).
- random variables, supermartingale time consistency: (Scandolo 2003; Detlefsen and Scandolo 2005).
- random variables, acceptance/rejection time consistency: (Acciaio et al. 2012; Bielecki et al. 2014a; Föllmer and Penner 2006; Roorda and Schumacher 2007, 2015; Tutsch 2008; Weber 2006).
- stochastic processes, strong and supermartingale time consistency: (discrete time) (Bielecki et al. 2014a; Scandolo 2003), (continuous time) (Jobert and Rogers 2008)

Dynamic monetary risk measures, strong time consistency:

(discrete time) (Cheridito and Kupper (2011); Cheridito et al. (2006)); (continuous time) (Bion-Nadal 2004; Föllmer and Penner 2011).

Dynamic acceptability indices: (Bielecki et al. (2014b); Biagini and Bion-Nadal (2014); Bielecki et al. (2013); Rosazza Gianin and Sgarra (2013); Bielecki et al. (2016); Frittelli and Maggis (2014); Bielecki et al. (2014a); Bielecki et al. (2015b)).

Mathematical preliminaries

Let $(\Omega, \mathcal{F}, \mathbb{F} = \{\mathcal{F}_t\}_{t \in \mathbb{T}}, P)$ be a filtered probability space, with $\mathcal{F}_0 = \{\Omega, \emptyset\}$, and $\mathbb{T} = \{0, 1, \dots, T\}$, where $T \in \mathbb{N}$ is a fixed and finite time horizon. We will also use the notation $\mathbb{T}' = \{0, 1, \dots, T - 1\}$.

For $\mathcal{G} \subseteq \mathcal{F}$ we denote by $L^0(\Omega, \mathcal{G}, P)$ and $\bar{L}^0(\Omega, \mathcal{G}, P)$ the sets of all \mathcal{G} -measurable random variables with values in $(-\infty, \infty)$ and $[-\infty, \infty]$, respectively. In addition, we use the notation $L^p(\mathcal{G}) := L^p(\Omega, \mathcal{G}, P)$, $L_t^p := L^p(\mathcal{F}_t)$, and $L^p := L_T^p$, for $p \in \{0, 1, \infty\}$; analogously we define \bar{L}_t^0 . We also use the notation $\mathbb{V}^p := \{(V_t)_{t \in \mathbb{T}} : V_t \in L_t^p\}$, for $p \in \{0, 1, \infty\}$.⁶ Moreover, we use $\mathcal{M}(P)$ to denote the set of all probability measures on (Ω, \mathcal{F}) that are absolutely continuous with respect to P , and we set $\mathcal{M}_t(P) := \{Q \in \mathcal{M}(P) : Q|_{\mathcal{F}_t} = P|_{\mathcal{F}_t}\}$.

Throughout this paper, \mathcal{X} relates to either the space of random variables L^p , or the space of adapted processes \mathbb{V}^p . If $\mathcal{X} = L^p$, then the elements $X \in \mathcal{X}$ are interpreted as discounted terminal cash flow. On the other hand, if $\mathcal{X} = \mathbb{V}^p$, then the elements of \mathcal{X} are interpreted as discounted dividend processes. All concepts developed for $\mathcal{X} = \mathbb{V}^p$ can be easily adapted to the case of the cumulative discounted value processes. The case of random variables can be viewed as a particular case of stochastic processes by considering cash flow with only the terminal payoff, i.e., stochastic processes such that $V = (0, \dots, 0, V_T)$. Nevertheless, we treat this case separately for transparency. In both cases, we consider the standard pointwise order, understood in the almost sure sense. In what follows, we also make use of the multiplication operator denoted as \cdot_t and defined by:

$$\begin{aligned} m \cdot_t V &:= (V_0, \dots, V_{t-1}, mV_t, mV_{t+1}, \dots), \\ m \cdot_t X &:= mX, \end{aligned} \tag{8}$$

for $V \in \{(V_t)_{t \in \mathbb{T}} \mid V_t \in L_t^0\}$, $X \in L^0$, $m \in L_t^\infty$, and $t \in \mathbb{T}$. In order to ease the notation, if no confusion arises, we drop \cdot_t from the above product, and we simply write mV and mX instead of $m \cdot_t V$ and $m \cdot_t X$, respectively. For any $t \in \mathbb{T}$ we set

$$1_{\{t\}} := \begin{cases} (\underbrace{0, 0, \dots, 0}_t, 1, 0, 0, \dots, 0), & \text{if } \mathcal{X} = \mathbb{V}^p, \\ 1 & \text{if } \mathcal{X} = L^p. \end{cases}$$

For any $m \in \bar{L}_t^0$, the value $m1_{\{t\}}$ corresponds to a cash flow of size m received at time t . We use this notation for the case of random variables to present more unified definitions (see Appendix ‘‘Dynamic LM-measures’’).

⁶Unless otherwise specified, it will be understood in the rest of the paper that $p \in \{0, 1, \infty\}$.

Remark 1 We note that the space \mathbb{V}^p , endowed with the multiplication \cdot_t , does not define a proper L^0 -module (Filipovic et al. 2009; Vogelpoth 2009) (e.g., in general, $0 \cdot_t V \neq 0$). However, in what follows, we will adopt some concepts from L^0 -module theory, which naturally fit into our study. We refer the reader to (Bielecki et al. 2015a, 2016) for a thorough discussion on this matter.

We use the convention $\infty - \infty = -\infty + \infty = -\infty$ and $0 \cdot \pm\infty = 0$. Note that the distributive law does not hold true in general: $(-1)(\infty - \infty) = \infty \neq -\infty + \infty = -\infty$. For $t \in \mathbb{T}$ and $X \in \bar{L}^0$ define the (generalized) \mathcal{F}_t -conditional expectation of X by

$$E[X|\mathcal{F}_t] := E[X^+|\mathcal{F}_t] - E[X^-|\mathcal{F}_t],$$

where $X^+ = (X \vee 0)$ and $X^- = (-X \vee 0)$. See Appendix “Conditional expectation and conditional essential supremum/infimum” for some relevant properties of the generalized expectation.

For $X \in \bar{L}^0$ and $t \in \mathbb{T}$, we will denote by $\text{ess inf}_t X$ the unique (up to a set of probability zero), \mathcal{F}_t -measurable random variable, such that

$$\text{ess inf}_{\omega \in A} X = \text{ess inf}_{\omega \in A} (\text{ess inf}_t X), \tag{9}$$

for any $A \in \mathcal{F}_t$. We call this random variable the \mathcal{F}_t -conditional essential infimum of X . Similarly, we define $\text{ess sup}_t(X) := -\text{ess inf}_t(-X)$, and we call it the \mathcal{F}_t -conditional essential supremum of X . Again, see Appendix “Conditional expectation and conditional essential supremum/infimum” for more details and some elementary properties of conditional essential infimum and supremum.

The next definition introduces the main object of this work.

Definition 1 A family $\varphi = \{\varphi_t\}_{t \in \mathbb{T}}$ of maps $\varphi_t : \mathcal{X} \rightarrow \bar{L}_t^0$ is a Dynamic LM-measure if φ satisfies

- 1) (Locality) $\mathbb{1}_A \varphi_t(X) = \mathbb{1}_A \varphi_t(\mathbb{1}_A \cdot_t X)$;
- 2) (Monotonicity) $X \leq Y \Rightarrow \varphi_t(X) \leq \varphi_t(Y)$;

for any $t \in \mathbb{T}$, $X, Y \in \mathcal{X}$ and $A \in \mathcal{F}_t$.

It is well recognized that locality and monotonicity are two properties that must be satisfied by any reasonable dynamic measure of performance and/or measure of risk, and in fact are shared by most, if not all, of such measures studied in the literature. The monotonicity property is natural for any numerical representation of an order between the elements of \mathcal{X} . The locality property (also referred to as regularity, or zero-one law, or relevance) essentially means that the values of the LM-measure restricted to a set $A \in \mathcal{F}$ remain invariant with respect to the values of the arguments outside of the same set $A \in \mathcal{F}$; in particular, the events that will not happen in the future do not affect the value of the measure today.

Remark 2 While in most of the literature the axiom of locality is not stated directly, it is very often implied by other assumptions. For example, if $\mathcal{X} = L^\infty$, then monotonicity and cash-additivity imply locality (cf. (Pitera 2014), Proposition 2.2.4).

Similarly, any convex (or concave) map is also local (cf. (Detlefsen and Scandolo 2005)). It is also worth mentioning that locality is strongly related to time consistency. In fact, in some papers locality is considered as a part of the time consistency property discussed below (see e.g. (Detlefsen and Scandolo 2009)).

In this paper, we only consider dynamic LM-measures φ , such that

$$0 \in \varphi_t[\mathcal{X}], \quad (10)$$

for any $t \in \mathbb{T}$. We impose this (technical) assumption to ensure that the maps φ_t that we consider are not degenerate in the sense that they are not taking infinite values for all $X \in \mathcal{X}$ on some set $A_t \in \mathcal{F}_t$ of positive probability, for any $t \in \mathbb{T}$; in the literature, sometimes such maps are referred to as *proper* (Kaina and Rüschemdorf 2009). If this is the case, then there exists a family $\{Y_t\}_{t \in \mathbb{T}}$, where $Y_t \in \mathcal{X}$, such that $\varphi_t(Y_t) \in L_t^0$ for any $t \in \mathbb{T}$, and so we can consider maps $\tilde{\varphi}$ given by $\tilde{\varphi}_t(\cdot) := \varphi_t(\cdot) - \varphi_t(Y_t)$, that satisfy assumption (10) and preserve the same order as the maps φ_t do. Typically, in the risk measure framework, one assumes that $\varphi_t(0) = 0$, which implies (10). However, here we cannot assume that $\varphi_t(0) = 0$, as we will also deal with dynamic performance measures for which $\varphi_t(0) = \infty$.

Finally, let us note that in the literature, traditionally the dynamic risk measures are monotone decreasing. On the other hand, the measures of performance are monotone increasing. In view of condition 2) in Definition 1, whenever our LM-measure corresponds to a dynamic risk measure, it needs to be understood as the negative of that risk measure. In such cases, in order to avoid confusion, we refer to the respective LM-measure as to *dynamic (monetary) utility measure* rather than as *dynamic (monetary) risk measure*. See Appendix “Dynamic LM-measures” for details.

Approaches to time consistent assessment of preferences

In this section, we present a brief survey of approaches to time consistent assessment of preferences, or to time consistency—for short, that were studied in the literature. As discussed in the Introduction, time consistency is studied via numerical representations of preferences. Various numerical representations will be surveyed below, and discussed in the context of dynamic LM-measures.

To streamline the presentation, we focus our attention on the case of random variables, that is $\mathcal{X} = L^p$, for $p \in \{0, 1, \infty\}$.⁷ Usually, the risk measures and the performance measures are studied on spaces smaller than L^0 , such as L^p , $p \in [1, \infty]$. This is motivated by the aim to obtain so called robust representation of such measures (see Appendix “Dynamic LM-measures”), since a certain topological structure is required for that (cf. Remark 13). On the other hand, *time consistency* refers only to consistency of measurements in time, where no particular topological structure is needed, and thus most of the results obtained here hold true for $p = 0$.

⁷Most of the concepts discussed in this Section can be modified to deal with the case of stochastic processes, as we will do in Section “Time consistency for stochastic processes”.

In Section “Generic approaches”, we outline two generic approaches to time consistent assessment of preferences: an approach based on update rules and an approach based on benchmark families. These two approaches are generic in the sense that nearly all types of time consistency can be represented within these two approaches. On the contrary, the approaches outlined in Section “Idiosyncratic approaches” are specific. That is to say, those approaches are suited only for specific types of time consistency, specific classes of dynamic LM-measures, specific spaces, etc.

Generic approaches

In this section, we outline two concepts that underlie the generic approaches to time consistent assessment of preferences: the update rules and the benchmark families. It will be seen that different types of time consistency can be characterized in terms of these concepts.

Update rules

The approach to time consistency using update rules was developed in Bielecki et al. (2014a). An update rule is a tool that is applied to preference levels, and used for relating assessments of preferences done using a dynamic LM-measure at different times.

Definition 2 A family $\mu = \{\mu_{t,s} : t, s \in \mathbb{T}, t < s\}$ of maps $\mu_{t,s} : \bar{L}_s^0 \rightarrow \bar{L}_t^0$ is called an update rule if μ satisfies the following conditions:

- 1) (Locality) $\mathbb{1}_A \mu_{t,s}(m) = \mathbb{1}_A \mu_{t,s}(\mathbb{1}_A m)$;
- 2) (Monotonicity) if $m \geq m'$, then $\mu_{t,s}(m) \geq \mu_{t,s}(m')$;

for any $s > t$, $A \in \mathcal{F}_t$, and $m, m' \in \bar{L}_s^0$.

Next, we give a definition of time consistency in terms of update rules.

Definition 3 Let μ be an update rule. We say that the dynamic LM-measure φ is μ -acceptance (resp. μ -rejection) time consistent if

$$\varphi_s(X) \geq m_s \quad (\text{resp. } \leq) \implies \varphi_t(X) \geq \mu_{t,s}(m_s) \quad (\text{resp. } \leq), \quad (11)$$

for all $s > t$, $s, t \in \mathbb{T}$, $X \in \mathcal{X}$, and $m_s \in \bar{L}_s^0$. If property (11) is satisfied for $s = t + 1$, $t \in \mathbb{T}'$, then we say that φ is one-step μ -acceptance (resp. one-step μ -rejection) time consistent.

We see that m_s and $\mu_{t,s}(m_s)$ serve as benchmarks to which the measurements of $\varphi_s(X)$ and $\varphi_t(X)$ are compared, respectively. Thus, the interpretation of acceptance time consistency is straightforward: if $X \in \mathcal{X}$ is accepted at some future time $s \in \mathbb{T}$, at least at level m_s , then today, at time $t \in \mathbb{T}$, it is accepted at least at level $\mu_{t,s}(m_s)$. Similar reasoning holds for the rejection time consistency. Essentially, the update rule μ converts the preference levels at time s to the preference levels at time t .

We started our survey of time consistency with Definition 3 since, as we will demonstrate below, this concept of time consistency covers various cases of time consistency for risk and performance measures that can be found in the existing literature. In particular, it allows to establish important connections between different types of time consistency. The time consistency property of an LM-measure, in general, depends on the choice of the updated rule; we refer to Section “Time consistency for random variables” for an in-depth discussion.

It is useful to observe that the time consistency property given in terms of update rules can be equivalently formulated as a version of the dynamic programming principle (see (Bielecki et al. 2014a), Proposition 3.6): φ is μ -acceptance (resp. μ -rejection) time consistent if and only if

$$\varphi_t(X) \geq \mu_{t,s}(\varphi_s(X)) \quad (\text{resp. } \leq), \tag{12}$$

for any $X \in \mathcal{X}$ and $s, t \in \mathbb{T}$, such that $s > t$. The interpretation of (12) is as follows: if the numerical assessment of preferences about X is given in terms of a dynamic LM-measure φ , then this measure is μ -acceptance time consistent if and only if the numerical assessment of preferences about X done at time t is greater than the value of the measurement done at any future time $s > t$ and updated at time t via $\mu_{t,s}$. The analogous interpretation applies to the ejection time consistency.

Next, we define two interesting and important classes of update rules.

Definition 4 *Let μ be an update rule. We say that μ is*

- 1) *s-invariant, if there exists a family $\{\mu_t\}_{t \in \mathbb{T}}$ of maps $\mu_t : \bar{L}^0 \rightarrow \bar{L}_t^0$, such that $\mu_{t,s}(m_s) = \mu_t(m_s)$ for any $s, t \in \mathbb{T}, s > t$, and $m_s \in \bar{L}_s^0$;*
- 2) *projective, if it is s-invariant and $\mu_t(m_t) = m_t$, for any $t \in \mathbb{T}$, and $m_t \in \bar{L}_t^0$.*

Remark 3 *If an update rule μ is s-invariant, then it is enough to consider only the corresponding family $\{\mu_t\}_{t \in \mathbb{T}}$. Hence, with slight abuse of notation, we write $\mu = \{\mu_t\}_{t \in \mathbb{T}}$ and call it an update rule as well.*

Example 1 *The families $\mu^1 = \{\mu_t^1\}_{t \in \mathbb{T}}$ and $\mu^2 = \{\mu_t^2\}_{t \in \mathbb{T}}$ given by*

$$\mu_t^1(m) = E[m|\mathcal{F}_t], \quad \text{and} \quad \mu_t^2(m) = \text{ess inf}_t m, \quad m \in \bar{L}^0,$$

are projective update rules. It will be shown in Example 5 that there is a dynamic LM-measure that is μ^2 -time consistent but not μ^1 -time consistent.

Benchmark families

The approach to time consistency based on families of benchmark sets was initiated by (Tutsch 2008), where the author applied this approach in the context of dynamic risk measures. Essentially, a benchmark family is a collection of subsets of \mathcal{X} that contain reference or test objects. The idea of time consistency in this context, is that the preferences about objects of interest must compare in a consistent way to the preferences about the reference objects.

Definition 5

(i) A family $\mathcal{Y} = \{\mathcal{Y}_t\}_{t \in \mathbb{T}}$ of sets $\mathcal{Y}_t \subseteq \mathcal{X}$ is a benchmark family if

$$0 \in \mathcal{Y}_t \quad \text{and} \quad \mathcal{Y}_t + \mathbb{R} = \mathcal{Y}_t,$$

for any $t \in \mathbb{T}$.

(ii) A dynamic LM-measure φ is acceptance (resp. rejection) time consistent with respect to the benchmark family \mathcal{Y} , if

$$\varphi_s(X) \geq \varphi_s(Y) \quad (\text{resp. } \leq) \quad \implies \quad \varphi_t(X) \geq \varphi_t(Y) \quad (\text{resp. } \leq), \quad (13)$$

for all $s \geq t$, $X \in \mathcal{X}$, and $Y \in \mathcal{Y}_s$.

Informally, the “degree” of time consistency with respect to \mathcal{Y} is measured by the size of \mathcal{Y} . Thus, the larger the sets \mathcal{Y}_s are, for each $s \in \mathbb{T}$, the stronger the degree of time consistency of φ .

Example 2 The families of sets $\mathcal{Y}^1 = \{\mathcal{Y}_t^1\}_{t \in \mathbb{T}}$ and $\mathcal{Y}^2 = \{\mathcal{Y}_t^2\}_{t \in \mathbb{T}}$ given by

$$\mathcal{Y}_t^1 = \mathbb{R} \quad \text{and} \quad \mathcal{Y}_t^2 = \mathcal{X},$$

are benchmark families. They relate to weak and strong types of time consistency, as will be discussed later on.

For future reference, we recall from (Bielecki et al. 2014a, Proof of Proposition 3.9) that φ is acceptance (resp. rejection) time consistent with respect to \mathcal{Y} , if and only if φ is acceptance (resp. rejection) time consistent with respect to the benchmark family $\widehat{\mathcal{Y}}$ given by

$$\widehat{\mathcal{Y}}_t := \{Y \in \mathcal{X} : Y = \mathbb{1}_A Y_1 + \mathbb{1}_{A^c} Y_2, \text{ for some } Y_1, Y_2 \in \mathcal{Y}_t \text{ and } A \in \mathcal{F}_t\}. \quad (14)$$

Relation between update rule approach and the benchmark approach

The difference between the update rule approach and the benchmark family approach is that the preference levels are chosen differently. Specifically, in the former approach, the preference level at time s is chosen as any $m_s \in \bar{L}_s^0$, and then updated to the preference level at time t , using an update rule. In the latter approach, the preference levels at both times s and t are taken as $\varphi_s(Y)$ and $\varphi_t(Y)$, respectively, for any reference object $Y \in \mathcal{Y}_s$, where \mathcal{Y}_s is an element of the benchmark family \mathcal{Y} .

These two approaches are strongly related to each other. Indeed, for any LM-measure φ and for any benchmark family \mathcal{Y} , one can construct an update rule μ such that φ is time consistent with respect to \mathcal{Y} if and only if it is μ -time consistent.

For example, in case of acceptance time consistency of φ with respect to \mathcal{Y} , using the locality of φ , it is easy to note that (13) is equivalent to

$$\varphi_t(X) \geq \text{ess sup}_{A \in \mathcal{F}_t} \left[\mathbb{1}_A \text{ess sup}_{Y \in \mathcal{Y}_{A,s}^-(\varphi_s(X))} \varphi_t(Y) + \mathbb{1}_{A^c} (-\infty) \right],$$

where $\mathcal{Y}_{A,s}^-(m_s) := \{Y \in \widehat{\mathcal{Y}}_s : \mathbb{1}_A m_s \geq \mathbb{1}_A \varphi_s(Y)\}$ and $\widehat{\mathcal{Y}} = \{\widehat{\mathcal{Y}}_s\}_{s \in \mathbb{T}}$ is defined in (14). Consequently, setting

$$\widetilde{\mu}_{t,s}(m_s) := \operatorname{ess\,sup}_{A \in \mathcal{F}_t} \left[\mathbb{1}_A \operatorname{ess\,sup}_{Y \in \mathcal{Y}_{A,s}^-(m_s)} \varphi_t(Y) + \mathbb{1}_{A^c}(-\infty) \right],$$

and using (12), we deduce that φ satisfies (13) if and only if φ is time consistent with respect to the update rule $\widetilde{\mu}_{t,s}$ (see (Bielecki et al. 2014a, Proposition 3.9) for details). The analogous argument works for rejection time consistency.

Generally speaking, the converse implication does not hold true; the notion of time consistency given in terms of update rules is more general. For example, time consistency of a dynamic coherent acceptability index cannot be expressed in terms of a single benchmark family.

Idiosyncratic approaches

Each such approach to time consistency of a given LM-measure exploits the idiosyncratic properties of this LM-measure, which are not necessarily shared by other LM measures, and typically is suited only for a specific subclass of dynamic LM-measures. For example, in case of dynamic convex or monetary risk measures the time consistency can be characterized in terms of the relevant properties of associated acceptance sets and/or the dynamics of the penalty functions and/or the rectangular property of the families of probability measures. These idiosyncratic approaches, and the relevant references, were mentioned and briefly discussed in Section “Literature review”. Detailed analysis of each of these approaches is beyond the scope of this survey.

Time consistency for random variables

In this section, we survey the time consistency of LM-measures applied to random variables. Accordingly, we assume that $\mathcal{X} = L^p$, for a fixed $p \in \{0, 1, \infty\}$. We proceed with the discussion of various related types of time consistency, without much reference to the existing literature. Such references are provided in Section “Literature review”.

Weak time consistency

The main idea behind this type of time consistency is that if “tomorrow”, say at time s , we accept $X \in L^p$ at level $\varphi_s(X)$, then “today”, say at time t , we would accept X at any level less than or equal to $\varphi_s(X)$, adjusted by the information \mathcal{F}_t available at time t . Similarly, if tomorrow we reject X at level $\varphi_s(X)$, then today, we should also reject X at any level greater than or equal to $\varphi_s(X)$, adapted to the information \mathcal{F}_t .

Definition 6 A dynamic LM-measure φ is weakly acceptance (resp. weakly rejection) time consistent if

$$\varphi_t(X) \geq \text{ess inf}_t \varphi_s(X), \quad (\text{resp. } \varphi_t(X) \leq \text{ess sup}_t \varphi_s(X))$$

for any $X \in L^p$ and $s, t \in \mathbb{T}$, such that $s > t$.

Propositions 1 and 2 provide some characterizations of weak acceptance time consistency.

Proposition 1 Let φ be a dynamic LM-measure on L^p . The following properties are equivalent:

- 1) φ is weakly acceptance time consistent.
- 2) φ is μ -acceptance time consistent, where μ is a projective update rule, given by

$$\mu_t(m) = \text{ess inf}_t m.$$

- 3) The following inequality is satisfied

$$\varphi_t(X) \geq \text{ess inf}_{Q \in \mathcal{M}_t(P)} E_Q[\varphi_s(X)|\mathcal{F}_t], \tag{15}$$

for any $X \in L^p$, $s, t \in \mathbb{T}$, $s > t$.

- 4) For any $X \in L^p$, $s, t \in \mathbb{T}$, $s > t$, and $m_t \in \bar{L}_t^0$, it holds that

$$\varphi_s(X) \geq m_t \Rightarrow \varphi_t(X) \geq m_t.$$

Similar results hold true for weak rejection time consistency.

For the proof of the equivalence between 1), 2), and 4), see (Bielecki et al. 2014a, Proposition 4.3). Regarding 3), note that any measure $Q \in \mathcal{M}_t(P)$ may be expressed in terms of a Radon-Nikodym derivative with respect to measure P . In other words, instead of (26), we may write

$$\varphi_t(X) \geq \text{ess inf}_{Z \in P_t} E[Z\varphi_s(X)|\mathcal{F}_t],$$

where $P_t := \{Z \in L^1 \mid Z \geq 0, E[Z|\mathcal{F}_t] = 1\}$. Thus, one can show equivalence between 1) and 3) noting that for any $m \in \bar{L}_t^0$ we get $\text{ess inf}_t m = \text{ess inf}_{Z \in P_t} E[Zm|\mathcal{F}_t]$. See (Bielecki et al. 2014a, Proposition 4.4) for the proof.

It is worth mentioning that Property 4) in Proposition 1 was suggested as the notion of (weak) acceptance and (weak) rejection time consistency in the context of scale invariant measures, called acceptability indices (cf. (Biagini and Bion-Nadal 2014; Bielecki et al. 2014b)).

Usually, the weak time consistency is considered for dynamic monetary risk measures on L^∞ (cf. (Acciaio and Penner 2011) and references therein). This case lends itself to even more characterizations of this property.

Proposition 2 Let φ be a representable dynamic monetary utility measure⁸ on L^∞ . The following properties are equivalent:

⁸See Section “Dynamic LM-measures” for details.

- 1) φ is weakly acceptance time consistent.
- 2) φ is acceptance time consistent with respect to $\{\mathcal{Y}_t\}_{t \in \mathbb{T}}$, where $\mathcal{Y}_t = \mathbb{R}$.
- 3) For any $X \in L^P$ and $s, t \in \mathbb{T}, s > t$,

$$\varphi_s(X) \geq 0 \Rightarrow \varphi_t(X) \geq 0. \tag{16}$$

- 4) $\mathcal{A}_{t+1} \subseteq \mathcal{A}_t$, for any $t \in \mathbb{T}$, such that $t < T$.
- 5) For any $Q \in \mathcal{M}(P)$ and $t \in \mathbb{T}$, such that $t < T$,

$$\alpha_t^{\min}(Q) \geq E_Q[\alpha_{t+1}^{\min}(Q) | \mathcal{F}_t],$$

where α^{\min} is the minimal penalty function in the robust representation of φ .

Analogous results are obtained for weak rejection time consistency.

We note that equivalence of properties 1), 2), and 3) also holds true in the case of $\mathcal{X} = L^0$, and not only for representable, but for any dynamic monetary utility measure; for the proof, see (Bielecki et al. 2014a, Proposition 4.3). Property 4) is a characterisation of weak time consistency in terms of acceptance sets, and property 5) gives a characterisation in terms of the supermartingale property of the penalty function. For the proof of the equivalence of 3), 4), and 5), see (Acciaio and Penner 2011, Proposition 33).

The next result shows that weak time consistency is indeed one of the weakest forms of time consistency, in the sense that the weak time consistency is implied by any time consistency generated by a projective update rule; we refer to (Bielecki et al. 2014a, Proposition 4.5) for the proof.

Proposition 3 *Let φ be a dynamic LM-measure on L^P , and let μ be a projective update rule. If φ is μ -acceptance (resp. μ -rejection) time consistent, then φ is weakly acceptance (resp. weakly rejection) time consistent.*

Remark 4 *An important feature of the weak time consistency is its invariance with respect to monotone transformations. Specifically, let $g : \mathbb{R} \rightarrow \mathbb{R}$ be a strictly increasing function and let φ be a weakly acceptance/rejection time consistent dynamic LM-measure. Then, $\{g \circ \varphi_t\}_{t \in \mathbb{T}}$ is also a weakly acceptance/rejection time consistent dynamic LM-measure.*

Remark 5 *In the case of general LM-measures, the weak time consistency may not be characterized as in 2) of Proposition 2. For example, if φ is a (normalized) acceptability index, then $\varphi_t(\mathbb{R}) = \{0, \infty\}$, for $t \in \mathbb{T}$, which does not agree with 4) in Proposition 1.*

Strong time consistency

As already stated in the Introduction, the origins of the strong form of time consistency can be traced to (Koopmans 1960). Historically, this is the first and the most extensively studied form of time consistency in dynamic risk measures literature.

It is fair to mention, that this form of time consistency also appears in the insurance literature, as the iterative property, and it is related to the mean value principle (Gerber 1974; Goovaerts and De Vylder 1979).

We start with the definition of strong time consistency.

Definition 7 *Let φ be a dynamic LM-measure on L^p . Then, φ is said to be strongly time consistent if*

$$\varphi_s(X) = \varphi_s(Y) \implies \varphi_t(X) = \varphi_t(Y), \quad (17)$$

for any $X, Y \in L^p$ and $s, t \in \mathbb{T}$, such that $s > t$.

Strong time consistency gains its popularity and importance due to its equivalence to the dynamic programming principle. This equivalence, as well as other characterisations of strong time consistency, are the subject of the following two propositions.

Proposition 4 *Let φ be a dynamic LM-measure on L^p . The following properties are equivalent:*

- 1) φ is strongly time consistent.
- 2) There exists an update rule μ such that φ is both μ -acceptance and μ -rejection time consistent.
- 3) φ is acceptance time consistent with respect to $\{\mathcal{Y}_t\}_{t \in \mathbb{T}}$, where $\mathcal{Y}_t = L^p$.
- 4) There exists an update rule μ such that for any $X \in L^p$, $s, t \in \mathbb{T}$, $s > t$,

$$\mu_{t,s}(\varphi_s(X)) = \varphi_t(X). \quad (18)$$

- 5) There exists a one-step update rule μ such that for any $X \in L^p$, $t \in \mathbb{T}$, $t < T$,

$$\mu_{t,t+1}(\varphi_{t+1}(X)) = \varphi_t(X).$$

See Appendix “Proofs” for the proof of Proposition 4. Property 4) in this proposition is referred to as Bellman’s principle or the dynamic programming principle. Also, note that 5) implies that any strongly time consistent dynamic LM-measure can be constructed using a backward recursion starting from $\varphi_T := \varrho$, where ϱ is an LM-measure. See (Cheridito and Kupper 2011) where the recursive construction for dynamic risk measures is discussed in details.

An important, and frequently studied, type of strong time consistency is the strong time consistency for dynamic monetary risk measures on L^∞ (cf. (Acciaio and Penner 2011) and references therein). As the next result shows, there are more equivalences that are valid in this case.

Proposition 5 *Let φ be a representable dynamic monetary utility measure on L^∞ . The following properties are equivalent:*

- 1) φ is strongly time consistent.
- 2) φ is recursive, i.e., for any $X \in L^p$, $s, t \in \mathbb{T}$, $s > t$,

$$\varphi_t(X) = \varphi_t(\varphi_s(X)).$$

- 3) $\mathcal{A}_t = \mathcal{A}_{t,s} + \mathcal{A}_s$, for all $t, s \in \mathbb{T}, s > t$.
- 4) For any $Q \in \mathcal{M}(P), t, s \in \mathbb{T}, s > t$,

$$\alpha_t^{\min}(Q) = \alpha_{t,s}^{\min}(Q) + E_Q[\alpha_s^{\min}(Q) | \mathcal{F}_t].$$

- 5) For any $X \in L^p, Q \in \mathcal{M}(P), s, t \in \mathbb{T}, s > t$,

$$\varphi_t(X) - \alpha_t^{\min}(Q) \leq E_Q[\varphi_s(X) - \alpha_s^{\min}(Q) | \mathcal{F}_t].$$

For the proof see, for instance, (Acciaio and Penner 2011, Proposition 14).

Remark 6 (i) *In general, for dynamic LM-measures, the strong time consistency does not imply either the weak acceptance or weak rejection time consistency. Indeed, let us consider $\varphi = \{\varphi_t\}_{t \in \mathbb{T}}$, such that $\varphi_t(X) = t$ (resp. $\varphi_t(X) = -t$) for all $X \in L^0$. Since $\varphi_t(0) = t \not\geq \text{ess inf}_t \varphi_s(0) = s$ (resp. $-t \not\leq -s$), for $s > t$, we conclude that φ is not weakly acceptance (resp. weakly rejection) time consistent. However, since $\varphi_t(X) = \varphi_t(\varphi_s(X))$ for any $X \in L^0$, then φ is strongly time consistent. We note, that if the update rule in Definition 7 is projective, as it is usually the case for dynamic monetary risk measures, then, due to Proposition 3, the strong time consistency implies the weak time consistency.*

(ii) *It is worth mentioning that, in principle, strong time consistency is not suited for acceptability indices (Bielecki et al. 2014b, 2015b; Cherny and Madan 2009). Let φ be a scale invariant dynamic LM-measure, and let $A \in \mathcal{F}_s$ be such that $P[A] = 1/2$, for some $s > 0, s \in \mathbb{T}$. Additionally, assume that \mathcal{F}_0 is trivial. We consider the sequence of random variables $X_n = n\mathbb{1}_A - \mathbb{1}_{A^c}, n \in \mathbb{N}$. By locality and scale invariance of φ , we have that $\varphi_s(X_n) = \varphi_s(X_1)$, for $n \in \mathbb{N}$. If φ is strongly time consistent, then we also have that $\varphi_0(X_n) = \varphi_0(X_1), n \in \mathbb{N}$. On the other hand, any reasonable measure of performance should assess X_n at the higher level as n increases, which contradicts the fact that $\varphi_0(X_n)$ is a constant sequence.*

Robust expectations, submartingales, and supermartingales

The concept of a projective update rule is connected with the concept of the (conditional) nonlinear expectation (see, for instance, (Peng 1997) for the definition and properties of nonlinear expectation). In (Peng 2004; Rosazza Gianin 2006), the authors established a link between nonlinear expectations and dynamic risk measures. One particularly important example of an projective update rule is the standard conditional expectation operator. Time consistency in L^∞ framework, defined in terms of conditional expectation, was studied in (Detlefsen and Scandolo 2005, Section 5) and associated with the super(sub)martingale property.

The next result introduces a general class of updates rules that are generated by conditional expectations and determining families of sets. First, we recall the concept of the determining family of sets (see, for instance, (Cherny 2006) for more details).

For each $t \in \mathbb{T}$ define

$$\mathcal{P}_t := \{Z \in L^1 \mid Z \geq 0, E[Z | \mathcal{F}_t] = 1\}.$$

A family of sets $\mathcal{D} = \{\mathcal{D}_t\}_{t \in \mathbb{T}}$ is a *determining family* if for any $t \in \mathbb{T}$, the set \mathcal{D}_t satisfies the following properties: $\mathcal{D}_t \neq \emptyset$, $\mathcal{D}_t \subseteq \mathcal{P}_t$, it is L^1 -closed, \mathcal{F}_t -convex⁹, and uniformly integrable.

Proposition 6 *Let \mathcal{D} be a determining family of sets, and let φ be a dynamic LM-measure. Consider the family of maps $\phi = \{\phi_t\}_{t \in \mathbb{T}}$, $\phi_t : \bar{L}^0 \rightarrow \bar{L}_t^0$, given by the following robust expectations¹⁰*

$$\phi_t(m) = \text{ess inf}_{Z \in \mathcal{D}_t} E[Zm | \mathcal{F}_t]. \tag{19}$$

Then,

- 1) the family ϕ is a projective update rule;
- 2) if φ is ϕ -acceptance time consistent, then $\{g \circ \phi_t\}_{t \in \mathbb{T}}$ is also ϕ -acceptance time consistent, for any increasing and concave function $g : \bar{\mathbb{R}} \rightarrow \mathbb{R}$.

Remark 7 *Classical (static) coherent risk measures defined on L^∞ admit robust representation of the form (1) for some set of probability measures \mathcal{Q} . It is known that the set \mathcal{Q} might not be unique. Consequently, there may exist multiple extensions of ρ to a map defined on \bar{L}^0 (see Appendix “LM-extensions” for the concept of the extension). Nevertheless, as in (Cherny 2006), one can consider the maximal set \mathcal{D} called determining set of a risk measure, which guarantees the uniqueness of such extension. The family of maps defined in (19) is an example of a family of such extensions. Consequently, we see that the coherent risk measures constitute a good starting point for generation of update rules.*

For the proof of Proposition 6, see Appendix “Proofs”. The counterpart of Proposition 6 for rejection time consistency is obtained by taking ess sup instead of ess inf in (19), and assuming that g is convex.

In the particular case of determining family with $\mathcal{D}_t = \{1\}$, for any $t \in \mathbb{T}$, the projective update rule takes the form $\mu_t(m) = E[m | \mathcal{F}_t]$, $m \in \bar{L}^0$. This is an important case, as it produces the concept of supermartingale and submartingale time consistency.

Definition 8 *Let φ be a dynamic LM-measure on L^p . We say that φ is supermartingale (resp. submartingale) time consistent if*

$$\varphi_t(X) \geq E[\varphi_s(X) | \mathcal{F}_t], \quad (\text{resp. } \leq)$$

for any $X \in L^p$ and $t, s \in \mathbb{T}$, $s > t$.

Remark 8 (i) *Note that any dynamic LM-measure that is ϕ -acceptance time consistent, where ϕ is given in (19), is also weakly acceptance time consistent, as ϕ is projective. In particular, any supermartingale time consistent LM-measure*

⁹By \mathcal{F}_t -convex we mean that for any $Z_1, Z_2 \in \mathcal{D}_t$, and $\lambda \in L_t^0$ such that $0 \leq \lambda \leq 1$ we get $\lambda Z_1 + (1 - \lambda)Z_2 \in \mathcal{D}_t$.

¹⁰The term robust is inspired by robust representations of risk measures.

is also weakly acceptance time consistent. A similar statement holds true for rejection time consistency.

(ii) As mentioned in (Bielecki et al. 2014a), the idea of update rules might be used to weight the preferences. Intuitively speaking, the risk of loss in the far future might be more preferred than the imminent risk of loss. This idea was used in (Cherny 2010). For example, the update rule μ of the form

$$\mu_{t,s}(m, X) = \begin{cases} \alpha^{s-t} E[m|\mathcal{F}_t] & \text{on } \{E[m|\mathcal{F}_t] \geq 0\}, \\ \alpha^{t-s} E[m|\mathcal{F}_t] & \text{on } \{E[m|\mathcal{F}_t] < 0\}. \end{cases} \tag{20}$$

for a fixed $\alpha \in (0, 1)$ would achieve this goal.

Other types of time consistency

The weak, strong, and super/sub-martingale forms of time consistency have attracted the most attention in the existing literature. In this section, we present other forms of time consistency that have been studied.

Middle time consistency

The notion of middle time consistency was originally formulated for dynamic monetary risk measures on L^∞ (cf. (Acciaio and Penner 2011)). The main idea is to replace the equality in (17) by an inequality. The term *middle acceptance* or *middle rejection* is used depending on the direction of the inequality.

Definition 9 A dynamic LM-measure φ on L^p is middle acceptance (resp. middle rejection) time consistent if

$$\varphi_s(X) \geq \varphi_s(Y) \quad (\text{resp. } \leq) \implies \varphi_t(X) \geq \varphi_t(Y) \quad (\text{resp. } \leq),$$

for any $X \in L^p$, $s, t \in \mathbb{T}$, $s > t$, and $Y \in L^p \cap L_s^0$.

The middle acceptance (resp. middle rejection) time consistency is equivalent to the acceptance (resp. rejection) time consistency with respect to the benchmark family $\mathcal{Y} = \{\mathcal{Y}_t\}_{t \in \mathbb{T}}$, given by $\mathcal{Y}_t = L^p \cap L_t^0$. In the case of dynamic convex risk measures, other characterizations of middle acceptance time consistency are available, as the following proposition shows.

Proposition 7 Let φ be a representable dynamic monetary utility measure on L^∞ , which is continuous from above. The following properties are equivalent:

- 1) φ is middle acceptance time consistent.
- 2) φ is φ^- -acceptance time consistent.¹¹
- 3) For any $X \in L^p$, $s, t \in \mathbb{T}$, $s > t$,

$$\varphi_t(X) \geq \varphi_t(\varphi_s(X)). \tag{21}$$

¹¹See Appendix “LM-extensions” for the definition of φ^-

4) For any $X \in L^P$ and $t \in \mathbb{T}$, such that $t < T$,

$$\varphi_{t+1}(X) - \varphi_t(X) \in \mathcal{R}_{t,t+1}.$$

5) For any $X \in \mathcal{R}_t$ and $t \in \mathbb{T}$, such that $t < T$,

$$\varphi_{t+1}(X) \in \mathcal{R}_t.$$

6) For any $t \in \mathbb{T}$, such that $t < T$, $\mathcal{A}_t \supseteq \mathcal{A}_{t,t+1} + \mathcal{A}_{t+1}$.

7) For any $Q \in \mathcal{M}(P)$ and $t \in \mathbb{T}$, such that $t < T$,

$$\alpha_t^{\min}(Q) \geq \alpha_{t,t+1}^{\min}(Q) + E_Q[\alpha_{t+1}^{\min}(Q) | \mathcal{F}_t].$$

7) For any $Q \in \mathcal{M}(P)$ and $t \in \mathbb{T}$, such that $t < T$,

$$\varphi_t(X) \geq E_Q[\varphi_{t+1}(X) | \mathcal{F}_t] + \alpha_{t,t+1}^{\min}(Q).$$

Since φ^- is an LM-extension of φ , and $\varphi_s(Y) = Y$, for any $Y \in L^P \cap \bar{L}_s^0$, the equivalence between 1) and 2) is immediate. For all other equivalences see (Acciaio and Penner 2011, Section 4.2) and references therein. Property 1) in Proposition 7 is sometimes called *prudence* (see (Penner 2007)).

Time consistency induced by LM-measure

It turns out that any dynamic LM-measure generates an update rule. Indeed, as the next result shows, any LM-extension of an LM-measure (see Appendix “LM-extensions” for the definition of LM-extension) is an s-invariant update rule.

Proposition 8 Any LM-extension $\hat{\varphi}$ of a dynamic LM-measure φ is an s-invariant update rule. Moreover, $\hat{\varphi}$ is projective if and only if $\varphi_t(X) = X$, for $t \in \mathbb{T}$ and $X \in L^P \cap \bar{L}_t^0$.

The proof is deferred to Appendix “Proofs”.

LM-extensions may be used to give stronger forms of strong and middle time consistency, that are especially well suited in the case of dynamic monetary risk measures.

Recall that for dynamic monetary risk measure φ on L^∞ , strong time consistency is equivalent to the property that

$$\varphi_t(X) = \varphi_t(\varphi_s(X)),$$

for any $X \in \mathcal{X}$, $s, t \in \mathbb{T}$, $s > t$.

However, if \mathcal{X} is larger than L^∞ , then this characterisation is problematic, as we might get $\varphi_s(X) \notin \mathcal{X}$. In this case, LM-extensions come in handy and one defines strong time consistency via the following equality,

$$\varphi_t(X) = \hat{\varphi}_t(\varphi_s(X)), \quad X \in \mathcal{X}, \quad s, t \in \mathbb{T}, \quad s > t, \tag{22}$$

where $\hat{\varphi}$ is an extension of φ from \mathcal{X} to \bar{L}^0 . Accordingly, we say that φ is *strongly* time consistent*, if there exists an LM-extension $\hat{\varphi}$, of φ , such that φ is both $\hat{\varphi}$ -acceptance and $\hat{\varphi}$ -rejection time consistent.

Note that since $\hat{\varphi}$ is an update rule, the strong* time consistency implies strong time consistency in the sense of Definition 7. In general, the converse implication is not true; to see this, it is enough to consider strong time consistency for an update rule that is not s-invariant.

In the same fashion, we say that φ is *middle* acceptance time consistent*, if there exists an LM-extension of φ , say $\hat{\varphi}$, such that φ is $\hat{\varphi}$ -acceptance time consistent. In view of Proposition 17, this is equivalent to saying that φ is middle* acceptance time consistent if it is φ^- -acceptance time consistent. Likewise, to define *middle* rejection time consistency* we use the mapping φ^+ .

Taxonomy of results

For the convenience of the reader, in Fig. 1 below, we summarize the results surveyed in Section “Time consistency for random variables”. For transparency, we label (by circled numbers) each arrow (implication or equivalence) in the flowchart, and we relate the labels to the relevant results, also providing comments on converse implications whenever appropriate.

- ① Proposition 2, 2)
- ② Proposition 1, 4)
- ③ Remark 8 and Proposition 3. The converse implication is not true in general, see Example 8.

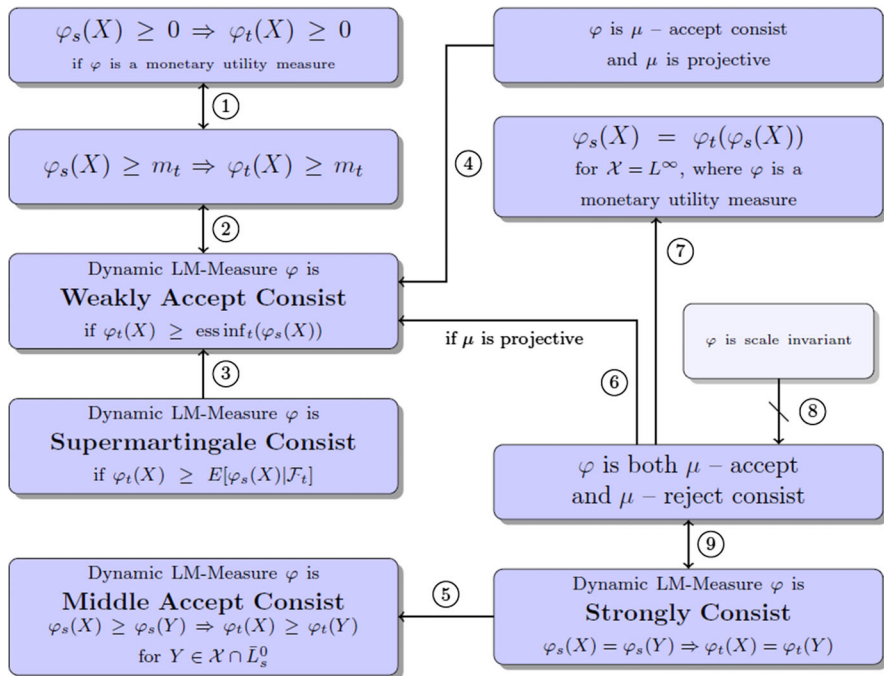


Fig. 1 Summary of results for acceptance time consistency for random variables

- ④ Proposition 3. Generally speaking, the converse implication is not true. See Example 8: the negative of Dynamic Entropic Risk Measure with $\gamma < 0$ is weakly acceptance time consistent, but it is not supermartingale time consistent, i.e., it is not acceptance time consistent with respect to the projective update rule $\mu_t = E_t[m|\mathcal{F}_t]$.
- ⑤ Proposition 4, 4). The converse implication is not true in general. For the counterexample, see (Acciaio and Penner 2011, Proposition 37).
- ⑥ Proposition 3, and see also ④. In general, strong time consistency does not imply weak acceptance time consistency, see Remark 6.
- ⑦ Proposition 7, 3)
- ⑧ This is heuristic statement. See Remark 6.(ii).
- ⑨ Proposition 4, 5)

Time consistency for stochastic processes

We preserve the same names for various types of time consistency for both the random variables and the stochastic processes. However, we stress that the nature of time consistency for stochastic processes is usually much more intricate. If φ is an LM-measure, and $V \in \mathbb{V}^p$, then in order to compare $\varphi_t(V)$ and $\varphi_s(V)$, for $s > t$, one also needs to take into account the cash flows between times t and s .

In order to account for the intermediate cash flows, we modify appropriately the concept of the update rule.

Definition 10 *The family $\mu = \{\mu_{t,s} : t, s \in \mathbb{T}, t < s\}$ of maps $\mu_{t,s} : \bar{L}_s^0 \times \mathcal{X} \rightarrow \bar{L}_t^0$ is called a generalized update rule if for any $X \in \mathcal{X}$ the family $\mu(\cdot, X) = \{\mu_{t,s}(\cdot, X) : t, s \in \mathbb{T}, t < s\}$ is an update rule.*

Note that the update rule introduced in Definition 3 may be considered as the generalized update rule, which is constant with respect to X , i.e., $\mu(\cdot, X) = \mu(\cdot, Y)$ for any $X, Y \in \mathcal{X}$. In what follows, if there is no ambiguity, we drop the term *generalized*.

As before, we say that the update rule μ is *s-invariant*, if there exists a family $\{\mu_t\}_{t \in \mathbb{T}}$ of maps $\mu_t : \bar{L}^0 \times \mathcal{X} \rightarrow \bar{L}_t^0$, such that $\mu_{t,s}(m_s, X) = \mu_t(m_s, X)$ for any $s, t \in \mathbb{T}, s > t, X \in \mathcal{X}$, and $m_s \in \bar{L}_s^0$.

We now arrive at the corresponding definition of time-consistency.

Definition 11 *Let μ be a generalized update rule. We say that the dynamic LM-measure φ is μ -acceptance (resp. μ -rejection) time consistent if*

$$\varphi_s(X) \geq m_s \quad (\text{resp. } \leq) \implies \varphi_t(X) \geq \mu_{t,s}(m_s, X) \quad (\text{resp. } \leq), \tag{23}$$

for all $s, t \in \mathbb{T}, s > t, X \in \mathcal{X}$, and $m_s \in \bar{L}_s^0$. In particular, if property (23) is satisfied for $s = t + 1, t = 0, \dots, T$, then we say that φ is one-step μ -acceptance (resp. one-step μ -rejection) time consistent.

Throughout this section, we assume that $\mathcal{X} = \mathbb{V}^P$.¹² We will focus our attention on one-step update rules μ such that

$$\mu_{t,t+1}(m, V) = \tilde{\mu}_{t,t+1}(m) + f(V_t), \quad t = 0, \dots, T - 1, \tag{24}$$

where $\tilde{\mu}$ is the one-step update rule for random variables, and $f : \bar{\mathbb{R}} \rightarrow \bar{\mathbb{R}}$ is a Borel measurable function such that $f(0) = 0$. Property (24) is postulated primarily to allow establishing a direct connection between our results and the existing literature. Moreover, when using one-step update rules of form (24), the one-step time consistency for random variables is a particular case of one-step time consistency for stochastic processes by considering cash flows with only terminal payoff, namely stochastic processes such that $V = (0, \dots, 0, V_T)$.

Finally, we note that for update rules, which admit the so called nested composition property (cf. (Ruszczyński 2010; Ruszczyński and Shapiro 2006b)),

$$\mu_{t,s}(m, V) = \mu_{t,t+1}(\mu_{t+1,t+2}(\dots \mu_{s-2,s-1}(\mu_{s-1,s}(m, V), V) \dots V), V), \tag{25}$$

we have that μ -acceptance (resp. μ -rejection) time consistency is equivalent to one step μ -acceptance (resp. μ -rejection) time consistency. This is another reason why we consider only one step update rules for stochastic processes.

Weak time consistency

We start with the following definition.

Definition 12 *A dynamic LM-measure φ on \mathbb{V}^P is weakly acceptance (resp. weakly rejection) time consistent if*

$$\varphi_t(V) \geq \text{ess inf}_t \varphi_{t+1}(V) + V_t, \quad (\text{resp. } \varphi_t(V) \leq \text{ess sup}_t \varphi_{t+1}(V) + V_t)$$

for any $V \in \mathbb{V}^P$ and $t \in \mathbb{T}$, such that $t < T$.

The next result is the counterpart of Proposition 1 and Proposition 2.

Proposition 9 *Let φ be a dynamic LM-measure on \mathbb{V}^P . The following properties are equivalent:*

- 1) φ is weakly acceptance time consistent.
- 2) φ is μ -acceptance time consistent, where μ is an s -invariant update rule, given by

$$\mu_t(m, V) = \text{ess inf}_t m + V_t.$$

- 3) For any $V \in \mathbb{V}^P$ and $t < T$

$$\varphi_t(V) \geq \text{ess inf}_{Q \in \mathcal{M}_t(P)} E_Q[\varphi_{t+1}(V) | \mathcal{F}_t] + V_t. \tag{26}$$

- 4) For any $V \in \mathbb{V}^P$, $t < T$, and $m_t \in \bar{L}_t^0$,

$$\varphi_{t+1}(V) \geq m_t \Rightarrow \varphi_t(V) \geq m_t + V_t.$$

¹²We recall that the elements of \mathbb{V}^P are interpreted as discounted dividend processes.

Additionally, if φ is a dynamic monetary risk measure, then the above properties are equivalent to

5) For any $V \in \mathbb{V}^P$ and $t < T$,

$$\varphi_{t+1}(V) \geq 0 \Rightarrow \varphi_t(V) \geq V_t.$$

Analogous equivalences are true for weak rejection time consistency.

The proof of Proposition 9 is analogous to the proofs of Proposition 1 and Proposition 2, and we omit it.

As mentioned earlier, the update rule, and consequently time consistency for stochastic processes, depends also on the value of the process (the dividend paid) at time t . In the case of weak time consistency this feature is interpreted as follows: if tomorrow, at time $t + 1$, we accept $V \in \mathbb{V}^P$ at the level greater than $m_{t+1} \in \mathcal{F}_{t+1}$, then today at time t , we will accept V at least at the level $\text{ess inf}_t m_{t+1}$ (i.e., the worst level of m_{t+1} adapted to the information \mathcal{F}_t) plus the dividend V_t received today.

Finally, we present the counterpart of Proposition 3 for the case of stochastic processes.

Proposition 10 *Let ϕ be a projective update rule for random variables and let the update rule μ for stochastic processes be given by*

$$\mu_{t,t+1}(m, V) = \phi_t(m) + V_t, \quad m \in \bar{L}_{t+1}^0, \quad V \in \mathbb{V}^P. \tag{27}$$

If φ is a dynamic one-step LM-measure on \mathbb{V}^P , which is μ -acceptance (resp. μ -rejection) time consistent, then φ is weakly acceptance (resp. weakly rejection) time consistent.

Proposition 10 can be proved in a way analogous to the proof of Proposition 3.

Remark 9 *The statement of Proposition 10 remains true if we replace (27) with*

$$\mu_{t,t+1}(m, V) = \phi_t(m + V_t), \quad m \in \bar{L}_{t+1}^0, \quad V \in \mathbb{V}^P.$$

Indeed, it is enough to note that, for any $V \in \mathbb{V}^P$ and $t < T$,

$$\begin{aligned} \varphi_t(V) &\geq \mu_{t,t+1}(\varphi_{t+1}(V), V) = \phi_t(\varphi_{t+1}(V) + V_t) \\ &\geq \phi_t(\text{ess inf}_t[\varphi_{t+1}(V) + V_t]) = \text{ess inf}_t[\varphi_{t+1}(V) + V_t] \geq \text{ess inf}_t \varphi_{t+1}(V) + V_t. \end{aligned}$$

Semi-weak time consistency

In this section, we introduce the concept of semi-weak time consistency for stochastic processes. We have not discussed semi-weak time consistency in the case of random variables, since, in that case, semi-weak time consistency coincides with the weak time consistency.

As it was shown, (Bielecki et al. 2014b), none of the forms of time consistency existing in the literature at the time when that paper was written were suitable for scale-invariant maps such as acceptability indices. In fact, even the weak acceptance and the weak rejection time consistency for stochastic processes (as defined

in the present paper) are too strong in the case of scale invariant maps. This is a reason why we introduce yet a weaker notion of time consistency, which we will refer to as *semi-weak acceptance* and *semi-weak rejection* time consistency. The notion of semi-weak time consistency for stochastic processes, introduced next, is well suited for scale-invariant maps; we refer the reader to (Bielecki et al. 2014b) for a detailed discussion on time consistency for such maps and their dual representations.¹³

Definition 13 *Let φ be a dynamic LM-measure on \mathbb{V}^P . Then, φ is semi-weakly acceptance time consistent if*

$\varphi_t(V) \geq \mathbb{1}_{\{V_t \geq 0\}} \text{ess inf}_t(\varphi_{t+1}(V)) + \mathbb{1}_{\{V_t < 0\}}(-\infty)$, for all $V \in \mathbb{V}^P$, $t \in \mathbb{T}$, $t < T$, and it is semi-weakly rejection time consistent if

$\varphi_t(V) \leq \mathbb{1}_{\{V_t \geq 0\}} \text{ess sup}_t(\varphi_{t+1}(V)) + \mathbb{1}_{\{V_t > 0\}}(+\infty)$, for all $V \in \mathbb{V}^P$, $t \in \mathbb{T}$, $t < T$.

Clearly, weak acceptance/rejection time consistency for stochastic processes implies semi-weak acceptance/rejection time consistency.

Next, we will show that the definition of semi-weak time consistency is indeed equivalent to the time consistency introduced in (Bielecki et al. 2014b).

Proposition 11 *Let φ be a dynamic LM-measure on \mathbb{V}^P . The following properties are equivalent*

- 1) φ is semi-weakly acceptance time consistent.
- 2) φ is one step μ -acceptance time consistent, where the (generalized) update rule is given by

$$\mu_{t,t+1}(m, V) = \mathbb{1}_{\{V_t \geq 0\}} \text{ess inf}_t m + \mathbb{1}_{\{V_t < 0\}}(-\infty).$$

- 3) For all $V \in \mathbb{V}^P$, $t \in \mathbb{T}$, $t < T$, and $m_t \in \bar{L}_t^0$, such that $V_t \geq 0$

$$\varphi_{t+1}(V) \geq m_t \implies \varphi_t(V) \geq m_t.$$

A similar result is true for semi-weak rejection time consistency.

For the proof, see (Bielecki et al. 2014a, Proposition 4.8).

Property 3) in Proposition 11, which is the definition of the (acceptance) time consistency given in (Bielecki et al. 2014b), best illustrates the financial meaning of semi-weak acceptance time consistency: if tomorrow we accept the dividend stream $V \in \mathbb{V}^P$ at level m_t , and if we get a positive dividend V_t paid today at time t , then today we accept the cash flow V at least at level m_t as well. A similar interpretation is valid for semi-weak rejection time consistency.

The next two results are important. In particular, they generalize the work done in (Bielecki et al. 2014b) regarding duality between cash-additive risk measures and acceptability indices.

¹³In (Bielecki et al. 2014b), the authors combine both semi-weak acceptance and rejection time consistency into one single definition and call it time consistency.

Proposition 12 Let $\{\varphi^x\}_{x \in \mathbb{R}_+}$ be a decreasing family¹⁴ of dynamic LM-measures on \mathbb{V}^P . Assume that for each $x \in \mathbb{R}_+$, φ^x is weakly acceptance (resp. weakly rejection) time consistent. Then, the family $\{\alpha_t\}_{t \in \mathbb{T}}$ of maps $\alpha_t : \mathbb{V}^P \rightarrow \bar{L}_t^0$ defined by

$$\alpha_t(V) := \operatorname{ess\,sup}_{x \in \mathbb{R}_+} \{x \mathbb{1}_{\{\varphi_t^x(V) \geq 0\}}\}, \tag{28}$$

is a semi-weakly acceptance (resp. semi-weakly rejection) time consistent dynamic LM-measure.

For the proof, see (Bielecki et al. 2014a, Proposition 4.9). It will be useful to note that $\alpha_t(V)$ defined in (28) can also be written as

$$\alpha_t(V) = \sup\{x \in \mathbb{R}^+ \mid \varphi_t^x(V) \geq 0\}. \tag{29}$$

Proposition 13 Let $\{\alpha_t\}_{t \in \mathbb{T}}$ be a dynamic LM-measure, which is independent of the past and translation invariant.¹⁵ Assume that $\{\alpha_t\}_{t \in \mathbb{T}}$ is semi-weakly acceptance (resp. semi-weakly rejection) time consistent. Then, for any $x \in \mathbb{R}_+$, the family $\varphi^x = \{\varphi_t^x\}_{t \in \mathbb{T}}$ of maps $\varphi_t^x : \mathbb{V}^P \rightarrow \bar{L}_t^0$ defined by

$$\varphi_t^x(V) := \operatorname{ess\,inf}_{c \in \mathbb{R}} \{c \mathbb{1}_{\{\alpha_t(V - c \mathbb{1}_{\{t\}}) \leq x\}}\}, \tag{30}$$

is a weakly acceptance (resp. weakly rejection) time consistent dynamic LM-measure.

For the proof, see (Bielecki et al. 2014a, Proposition 4.10). In what follows, we will use the fact that $\varphi_t^x(V)$ defined in (30) can also be written as

$$\varphi_t^x(V) = \inf\{c \in \mathbb{R} \mid \alpha_t(V - c \mathbb{1}_{\{t\}}) \leq x\}. \tag{31}$$

This type of dual representation, i.e., (28) and (30), or, equivalently, (29) and (31), first appeared in (Cherny and Madan 2009) where the authors studied the static (one period of time) case. Subsequently, in (Bielecki et al. 2014b), the authors extended these results to the case of stochastic processes with special emphasis on the time consistency property. In contrast to the results of (Bielecki et al. 2014b), Propositions 12 and 13 consider an arbitrary probability space, not just a finite one.

Strong time consistency

Let us start with the definition of strong time consistency.

Definition 14 Let φ be a dynamic LM-measure on \mathbb{V}^P . Then φ is said to be strongly time consistent if

$$V_t = V'_t \text{ and } \varphi_{t+1}(V) = \varphi_{t+1}(V') \implies \varphi_t(V) = \varphi_t(V'),$$

for any $V, V' \in \mathbb{V}^P$ and $t \in \mathbb{T}$, such that $t < T$.

¹⁴A family, indexed by $x \in \mathbb{R}_+$, of maps $\{\varphi_t^x\}_{t \in \mathbb{T}}$, is called *decreasing*, if $\varphi_t^x(X) \leq \varphi_t^y(X)$ for all $X \in \mathcal{X}$, $t \in \mathbb{T}$ and $x, y \in \mathbb{R}_+$, such that $x \geq y$.

¹⁵See Appendix “Dynamic LM-measures” for details.

Now, let us present the counterpart of Proposition 4.

Proposition 14 *Let φ be a dynamic LM-measure on \mathbb{V}^P , which is independent of the past. The following properties are equivalent:*

- 1) φ is strongly time consistent.
- 2) There exists an update rule μ such that: for any $t \in \mathbb{T}'$, $m \in \bar{L}_t^0$, and $V, V' \in \mathbb{V}^P$, satisfying $V_t = V'_t$, we have $\mu_{t,t+1}(m, V) = \mu_{t,t+1}(m, V')$; the family φ is both one-step μ -acceptance and one-step μ -rejection time consistent.
- 3) There exists an update rule μ such that for any $t < T$ and $V \in \mathbb{V}^P$

$$\varphi_t(V) = \mu_{t,t+1}(\varphi_{t+1}(V), 1_{\{t\}}V_t).$$

As in the case of random variables, strong time consistency is usually considered for dynamic monetary risk measures on \mathbb{V}^∞ . In this case, additional equivalent properties can be established. For brevity, we skip the details, and only show the general idea for deriving a litany of equivalent properties. This idea is rooted in a specific construction of strongly time consistent dynamic LM-measures.

Corollary 1 *Let μ be a update rule for random variables. Let $\tilde{\varphi}$ be a dynamic LM-measure on \mathbb{V}^∞ given by*

$$\begin{cases} \tilde{\varphi}_T(V) = V_T \\ \tilde{\varphi}_t(V) = \mu_{t,t+1}(\tilde{\varphi}_{t+1}(V)) + V_t, \end{cases}$$

Then, $\tilde{\varphi}$ is a strongly time consistent dynamic LM-measure on \mathbb{V}^∞ .

For a more detailed explanation of this idea and other equivalent properties see, e.g., (Cheridito and Kupper 2011) or (Ruszczyński and Shapiro 2006b).

Other types of time consistency

Other types of time consistency for stochastic processes may be defined in analogy to what is done in Section “Other types of time consistency” for the case of random variables. For brevity, we limit our discussion here to the update rules derived from dynamic LM-measures.

First, given a dynamic LM-measure φ on \mathbb{V}^P , we denote by $\tilde{\varphi}$ the family of maps $\tilde{\varphi}_t : L_{t+1}^P \rightarrow \bar{L}_t^0$ given by

$$\tilde{\varphi}_t(X) := \varphi_t(1_{\{t+1\}}X), \text{ for } t \in \mathbb{T}'. \tag{32}$$

Since φ is monotone and local on \mathbb{V}^P , then, clearly, $\tilde{\varphi}_t$ is local and monotone on L_{t+1}^P .

Next, for any $t \in \mathbb{T}'$, we extend $\tilde{\varphi}_t$ to \bar{L}_{t+1}^0 , preserving locality and monotonicity (see Remark 12), and this extension produces a one-step update rule.

For example, the middle acceptance time consistency is obtained by taking the update rule μ given as

$$\mu_{t,t+1}(m, V) = \tilde{\varphi}_t^-(m + V_t), \quad t \in \mathbb{T}',$$

where $\tilde{\varphi}_t^- : \tilde{L}_{t+1}^0 \rightarrow \tilde{L}_t^0$ is defined as in (55), with the sets $\mathcal{Y}_A^-(X)$ replaced by

$$\mathcal{Y}_{t,A}^-(X) := \{Y \in L_{t+1}^p \mid \mathbb{1}_A Y \leq \mathbb{1}_A X\}, \quad X \in \tilde{L}_{t+1}^0.$$

Taxonomy of results

In Fig. 2, we summarize the results surveyed in Section “Time consistency for stochastic processes”. We label each arrow (implication or equivalence) in the flowchart with numbers in squares and we relate the labels to the relevant results. Additionally, we provide comments on converse implications whenever appropriate.

- Ⓛ Proposition 9, 5)
- Ⓜ Proposition 9, 4)
- Ⓝ Proposition 11, 3)
- Ⓞ Proposition 10
- Ⓟ Proposition 10, and see also Ⓞ
- Ⓠ Proposition 14.

Remark 10 *The converse of implications Ⓞ and Ⓟ in Flowchart 2 do not hold true in general; one can use the same counterexamples as in the case of random variables. For a counterexample showing that the converse of Ⓝ does not hold true in general, see Example 5.*

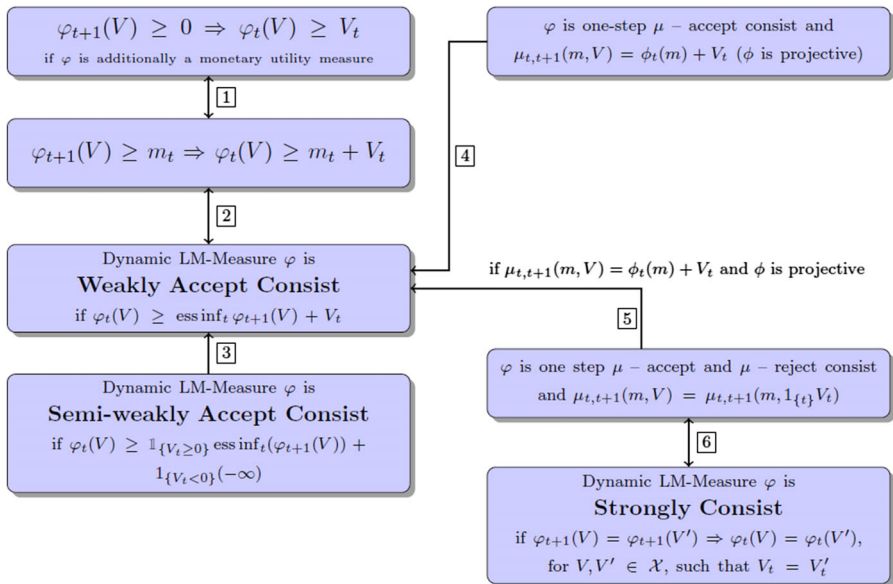


Fig. 2 Summary of results for acceptance time consistency for stochastic processes

Examples

In this section, we present examples that illustrate the different types of time consistency for dynamic risk measures and dynamic performance measures, as well as the relationships between them.

We recall that according to the convention adopted in this paper, the dynamic LM-measures representing risk measures are the negatives of their classical counterparts. With this understanding, in the titles of the examples representing risk measures below we will skip the term “negative”.

Example 3 (Value at Risk (V@R)) *Let $\mathcal{X} = L^0$ and $\alpha \in (0, 1)$. We denote by $\varphi_t^\alpha(X)$ an \mathcal{F}_t -conditional α -quantile of X*

$$\varphi_t^\alpha(X) := \text{ess sup}\{Y \in L_t^0 \mid P[X \leq Y \mid \mathcal{F}_t] \leq \alpha\}. \tag{33}$$

According to our convention, the conditional V@R is defined by $V@R_t^\alpha(X) := -\varphi_t^\alpha(X)$.

The family of maps $\{\varphi_t^\alpha\}_{t \in \mathbb{T}}$ is a dynamic monetary utility measure. It is well known that $\{\varphi_t^\alpha\}_{t \in \mathbb{T}}$ is not strongly time consistent; see (Cheridito and Stadje 2009) for details. However, it is both weakly acceptance and weakly rejection time consistent. Indeed, if $\varphi_s^\alpha(X) \geq 0$, for some $s > t$, and $X \in L^0$, then for any $\epsilon < 0$ we get $P[X \leq \epsilon \mid \mathcal{F}_s] = E[\mathbb{1}_{\{X \leq \epsilon\}} \mid \mathcal{F}_s] \leq \alpha$, and

$$E[\mathbb{1}_{\{X \leq \epsilon\}} \mid \mathcal{F}_t] = E[E[\mathbb{1}_{\{X \leq \epsilon\}} \mid \mathcal{F}_s] \mid \mathcal{F}_t] \leq \alpha.$$

Since $\epsilon < 0$ was chosen arbitrarily, we get $\varphi_t^\alpha(X) \geq 0$, and thus, in view of Proposition 2, $\{\varphi_t^\alpha\}_{t \in \mathbb{T}}$ is weakly acceptance time consistent.

Now, let us assume that $\varphi_s^\alpha(X) \leq 0$. Then, due to the locality of the conditional expectation, we have that $E[\mathbb{1}_{\{X \leq \epsilon\}} \mid \mathcal{F}_s] > \alpha$, for any $\epsilon > 0$. In fact, if $P[E[\mathbb{1}_{\{X \leq \epsilon\}} \mid \mathcal{F}_s] > \alpha] < 1$, then there exists an \mathcal{F}_s -measurable set A with positive measure on which

$$E[\mathbb{1}_{\{X \leq \epsilon\}} \mid \mathcal{F}_s] \leq \alpha.$$

Taking any $Y' \in L_s^0$ such that $E[\mathbb{1}_{\{X \leq Y'\}} \mid \mathcal{F}_s] \leq \alpha$, we know that for \mathcal{F}_s -measurable random variable $Z := \mathbb{1}_A \epsilon + \mathbb{1}_{A^c} Y'$ we get

$$\begin{aligned} E[\mathbb{1}_{\{X \leq Z\}} \mid \mathcal{F}_s] &= \mathbb{1}_A E[\mathbb{1}_{\{X \leq Z\}} \mid \mathcal{F}_s] + \mathbb{1}_{A^c} E[\mathbb{1}_{\{X \leq Z\}} \mid \mathcal{F}_s] = \mathbb{1}_A E[\mathbb{1}_{\{X \leq \epsilon\}} \mid \mathcal{F}_s] \\ &\quad + \mathbb{1}_{A^c} E[\mathbb{1}_{\{X \leq Y'\}} \mid \mathcal{F}_s] \leq \alpha. \end{aligned}$$

Thus,

$$0 \geq \text{ess sup}\{Y \in L_s^0 \mid P[X \leq Y \mid \mathcal{F}_s] \leq \alpha\} \geq Z,$$

which leads to the contradiction.

Consequently, for any $Y \in L_t^0$ and $\epsilon > 0$, we get

$$\begin{aligned} E[\mathbb{1}_{\{X \leq Y\}} \mid \mathcal{F}_t] &\geq E[\mathbb{1}_{\{X \leq \epsilon < Y\}} \mid \mathcal{F}_t] = E[\mathbb{1}_{\{X \leq \epsilon\}} \mathbb{1}_{\{Y > \epsilon\}} \mid \mathcal{F}_t] \\ &= \mathbb{1}_{\{Y > \epsilon\}} E[E[\mathbb{1}_{\{X \leq \epsilon\}} \mid \mathcal{F}_s] \mid \mathcal{F}_t], \end{aligned}$$

and, consequently, $E[\mathbb{1}_{\{X \leq Y\}} \mid \mathcal{F}_t] > \alpha$ on \mathcal{F}_t -measurable set $\{Y > \epsilon\}$. Hence,

$$\varphi_t^\alpha(X) = \text{ess sup}\{Y \in L_t^0 \mid E[\mathbb{1}_{\{X \leq Y\}} \mid \mathcal{F}_t] \leq \alpha\} \leq \text{ess sup}\{Y \in L_t^0 \mid Y \leq \epsilon\} = \epsilon.$$

Since $\epsilon > 0$ was chosen arbitrary, we conclude that $\varphi_t^\alpha(X) \leq 0$, thus $\{\varphi_t^\alpha\}_{t \in \mathbb{T}}$ is weakly rejection time consistent.

Example 4 (Conditional Weighted Value at Risk) Let $\mathcal{X} = L^0$. For a fixed $\alpha \in (0, 1)$, we consider the family of sets $\{D_t^\alpha\}_{t \in \mathbb{T}}$ defined by

$$D_t^\alpha := \{Z \in L^1 : 0 \leq Z \leq \alpha^{-1}, E[Z|\mathcal{F}_t] = 1\}, \tag{34}$$

and we set

$$\varphi_t^\alpha(X) := \operatorname{ess\,inf}_{Z \in D_t^\alpha} E[ZX|\mathcal{F}_t], \quad t \in \mathbb{T}, X \in L^0. \tag{35}$$

The family of maps $\{\varphi_t^\alpha\}_{t \in \mathbb{T}}$ is a dynamic coherent utility measure (see, e.g., (Cherny 2006) for details). Moreover, it is submartingale time consistent. Indeed, let $t, s \in \mathbb{T}, s > t$. Clearly, $D_s^\alpha \subseteq D_t^\alpha$, thus

$$\begin{aligned} \varphi_t^\alpha(X) &= \operatorname{ess\,inf}_{Z \in D_t^\alpha} E[ZX|\mathcal{F}_t] \leq \operatorname{ess\,inf}_{Z \in D_s^\alpha} E[ZX|\mathcal{F}_t] \\ &= \operatorname{ess\,inf}_{Z \in D_s^\alpha} E[E[ZX|\mathcal{F}_s]|\mathcal{F}_t]. \end{aligned} \tag{36}$$

Now, using the fact that D_s^α is L^1 -closed (see (Cherny 2006) for details), for any $X \in L^0$, there exist $Z_X^* \in D_s^\alpha$ such that $\varphi_s^\alpha(X) = E[Z_X^*X|\mathcal{F}_s]$. This implies that

$$\begin{aligned} \operatorname{ess\,inf}_{Z \in D_s^\alpha} E[E[ZX|\mathcal{F}_s]|\mathcal{F}_t] &\leq E[E[Z_X^*X|\mathcal{F}_s]|\mathcal{F}_t] = E[\operatorname{ess\,inf}_{Z \in D_s^\alpha} E[ZX|\mathcal{F}_s]|\mathcal{F}_t] \\ &= E[\varphi_s^\alpha(X)|\mathcal{F}_t]. \end{aligned} \tag{37}$$

Combining (36) and (37), we conclude that φ^α is submartingale time-consistent. In particular, by Remark 8, φ^α is also weakly rejection time consistent.

On the other hand, as shown in (Artzner et al. 2007), φ^α is neither middle rejection time consistent nor weakly acceptance time consistent.

Example 5 (Dynamic TV@R Acceptability Index for Processes) Tail Value at Risk Acceptability Index was introduced in (Cherny and Madan 2009), as an example of static scale invariant performance measure for the case of random variables. Here, along the lines of (Bielecki et al. (Bielecki et al. 2014b)), we extend this notion to the dynamic setup and apply it to the case of stochastic processes. Let $\mathcal{X} = \mathbb{V}^0$, and for a fixed $\alpha \in (0, 1]$, we consider the sets $\{D_t^\alpha\}_{t \in \mathbb{T}}$ defined in (34). We consider the distortion function $g(x) = \frac{1}{1+x}, x \in \mathbb{R}^+$, and we define $\rho^x = \{\rho_t^x\}_{t \in \mathbb{T}}, x \in \mathbb{R}_+$, as follows

$$\rho_t^x(V) = \operatorname{ess\,inf}_{Z \in D_t^{g(x)}} E \left[Z \sum_{i=t}^T V_i \middle| \mathcal{F}_t \right], \quad V \in \mathbb{X}, t \in \mathbb{T}. \tag{38}$$

Then, ρ^x is an increasing (with respect to x) family of dynamic coherent utility measures for processes, and the map $\alpha = \{\alpha_t\}_{t \in \mathbb{T}}$ given by

$$\alpha_t(V) = \sup\{x \in \mathbb{R}_+ \mid \rho_t^x(V) \leq 0\}, \tag{39}$$

is a dynamic acceptability index for processes (see (Cherny and Madan 2009) and (Bielecki et al. (Bielecki et al. 2014b))). Moreover,

$$\rho_t^x(V) = \inf\{c \in \mathbb{R} \mid \alpha_t(V + c\mathbb{1}_{\{t\}}) \geq x\}. \tag{40}$$

Clearly, (39) and (40) are the counterparts of (29) and (31), respectively.

Considering the above, then, similarly to Example 4, one can show that ρ^x is weakly rejection time consistent, but it is not weakly acceptance time consistent, for any fixed $x \in \mathbb{R}_+$, and hence, by Proposition 12 and Proposition 13, α is semi-weakly rejection time consistent but not semi-weakly acceptance time consistent.

Example 6 (Dynamic RAROC) *The Risk Adjusted Return On Capital (RAROC) is a popular scale invariant measure of performance; see (Cherny and Madan (Cherny and Madan 2009)) for static RAROC and (Bielecki et al. (Bielecki et al. 2014b)) for its extension to the dynamic setup. We consider the space $\mathcal{X} = \mathbb{V}^1$, and for a fixed $\alpha \in (0, 1)$ the dynamic RAROC is defined as follows*

$$\varphi_t(V) := \begin{cases} \frac{E[\sum_{i=t}^T V_i | \mathcal{F}_t]}{-\rho_t^\alpha(V)} & \text{if } E\left[\sum_{i=t}^T V_i | \mathcal{F}_t\right] > 0, \\ 0 & \text{otherwise,} \end{cases} \tag{41}$$

when $\rho_t^\alpha(V) < 0$, where $\rho_t^\alpha(V) = \text{ess inf}_{Z \in D_t^\alpha} E[Z \sum_{i=t}^T V_i | \mathcal{F}_t]$, and $\{D_t^\alpha\}_{t \in \mathbb{T}}$ given in (34), and $\varphi_t(V) = +\infty$, if $\rho_t(V) \geq 0$. It was shown in (Bielecki et al. 2014b) that φ is a dynamic acceptability index for processes. Moreover, for any fixed $t \in \mathbb{T}$, we have that (cf. (Bielecki et al. 2015a))

$$\varphi_t(V) = \sup\{x \in \mathbb{R}_+ : \phi_t^x(V) \geq 0\},$$

where $\phi_t^x(V) = \text{ess inf}_{Z \in \mathcal{B}_t^x} E[Z(\sum_{i=t}^T V_i) | \mathcal{F}_t]$, with $\mathcal{B}_t^x = \{Z \in L^1 : Z = \frac{1}{1+x} + \frac{x}{1+x} Z_1, \text{ for some } Z_1 \in D_t^\alpha\}$. It is easy to check that the family $\{\varphi_t^x\}_{t \in \mathbb{T}}$ is a dynamic coherent utility measure for processes, and by similar arguments as in Example 4, we get that for any fixed $x \in \mathbb{R}_+$, φ_t^x is weakly rejection time consistent, but not weakly acceptance time consistent. Since $1 \in D_t^\alpha$, it follows that $\{\varphi_t^x\}_{t \in \mathbb{T}}$ is increasing in $x \in \mathbb{R}_+$, and by similar arguments as in Example 5, we conclude that φ is semi-weakly rejection time consistent, but not semi-weakly acceptance time consistent.

Example 7 (Dynamic Gain Loss Ratio) *Dynamic Gain Loss Ratio (dGLR) is another popular measure of performance, which essentially improves on some drawbacks of Sharpe Ratio (such as penalizing for positive returns), and it is given by the ratio of expected return over expected losses. Formally, for $\mathcal{X} = \mathbb{V}^1$, dGLR is defined as*

$$\varphi_t(V) := \begin{cases} \frac{E\left[\sum_{i=t}^T V_i | \mathcal{F}_t\right]}{E\left[(\sum_{i=t}^T V_i)^- | \mathcal{F}_t\right]}, & \text{if } E\left[\sum_{i=t}^T V_i | \mathcal{F}_t\right] > 0, \\ 0, & \text{otherwise.} \end{cases} \tag{42}$$

For various properties and dual representations of dGLR see (Bielecki et al. 2014b, 2016). In (Bielecki et al. (Bielecki et al. 2014b)), assuming that Ω is finite, the authors showed that dGLR is both semi-weakly acceptance and semi-weakly rejection time consistent. For the sake of completeness, we will show here that dGLR is semi-weakly acceptance time consistent.

Assume that $t \in \mathbb{T}'$, and $V \in \mathcal{X}$. In view of Definition 13, it is enough to show that

$$\varphi_t(V) \geq \mathbb{1}_{\{V_t \geq 0\}} \text{ess inf}_t(\varphi_{t+1}(V)) + \mathbb{1}_{\{V_t < 0\}}(-\infty). \tag{43}$$

On the set $\{V_t < 0\}$, the inequality (43) is trivial. Since φ_t is non-negative and local, without loss of generality, we may assume that $\text{ess inf}_t(\varphi_{t+1}(V)) > 0$. Since, $\varphi_{t+1}(V) \geq \text{ess inf}_t(\varphi_{t+1}(V))$, we have that

$$E \left[\sum_{i=t+1}^T V_i \middle| \mathcal{F}_{t+1} \right] \geq \text{ess inf}_t(\varphi_{t+1}(V)) \cdot E \left[\left(\sum_{i=t+1}^T V_i \right)^- \middle| \mathcal{F}_{t+1} \right]. \tag{44}$$

Using (44) we obtain

$$\begin{aligned} \mathbb{1}_{\{V_t \geq 0\}} E \left[\sum_{i=t}^T V_i \middle| \mathcal{F}_t \right] &\geq \mathbb{1}_{\{V_t \geq 0\}} E \left[E \left[\sum_{i=t+1}^T V_i \middle| \mathcal{F}_{t+1} \right] \middle| \mathcal{F}_t \right] \\ &\geq \mathbb{1}_{\{V_t \geq 0\}} \text{ess inf}_t(\varphi_{t+1}(V)) \cdot E \left[\mathbb{1}_{\{V_t \geq 0\}} E \left[\left(\sum_{i=t+1}^T V_i \right)^- \middle| \mathcal{F}_{t+1} \right] \middle| \mathcal{F}_t \right] \\ &\geq \mathbb{1}_{\{V_t \geq 0\}} \text{ess inf}_t(\varphi_{t+1}(V)) \cdot E \left[\left(\sum_{i=t}^T V_i \right)^- \middle| \mathcal{F}_t \right]. \end{aligned} \tag{45}$$

Note that $\text{ess inf}_t(\varphi_{t+1}(V)) > 0$ implies that $\varphi_{t+1}(V) > 0$, thus $E \left[\sum_{i=t+1}^T V_i \middle| \mathcal{F}_{t+1} \right] > 0$. Hence, on the set $\{V_t \geq 0\}$, we have

$$E \left[\sum_{i=t}^T V_i \middle| \mathcal{F}_t \right] \geq E \left[E \left[\sum_{i=t+1}^T V_i \middle| \mathcal{F}_{t+1} \right] \middle| \mathcal{F}_t \right] > 0.$$

We conclude the proof by combining the last inequality with (45).

Example 8 (Dynamic Entropic Risk Measure) *Entropic Risk Measure is a classical convex risk measure. The dynamic version of it (up to the negative sign) is defined as follows*

$$\varphi_t^\gamma(X) = \begin{cases} \frac{1}{\gamma} \ln E[\exp(\gamma X) | \mathcal{F}_t] & \text{if } \gamma \neq 0, \\ E[X | \mathcal{F}_t] & \text{if } \gamma = 0, \end{cases} \tag{46}$$

where $X \in \mathcal{X} = L^\infty$, $t \in \mathbb{T}$. The parameter $\theta = -\gamma$ is commonly known as the risk-aversion parameter. It can be proved that for $\gamma \leq 0$, the map φ_t^γ is a dynamic concave utility measure, and that for any $\gamma \in \mathbb{R}$, the map φ^γ is strongly time consistent (cf. (Kupper and Schachermayer 2009)). Since it is also cash-additive, strong time consistency implies both weak rejection and weak acceptance time consistency. Moreover (see (Bielecki et al. 2015a; Kupper and Schachermayer 2009) for details), $\{\varphi_t^\gamma\}_{t \in \mathbb{T}}$ is supermartingale time consistent if and only if $\gamma \geq 0$, and submartingale time consistent if and only if $\gamma \leq 0$.

Example 9 (Dynamic Entropic Risk Measure with non-constant risk aversion) *One can generalize the Dynamic Entropic Risk Measure (46) by taking time dependent risk aversion parameters. Let*

$$\varphi_t^{\gamma_t}(X) = \begin{cases} \frac{1}{\gamma_t} \ln E[\exp(\gamma_t X)|\mathcal{F}_t] & \text{if } \gamma_t \neq 0, \\ E[X|\mathcal{F}_t] & \text{if } \gamma_t = 0, \end{cases} \tag{47}$$

where $\{\gamma_t\}_{t \in \mathbb{T}}$ is such that $\gamma_t \in L_t^\infty$, $t \in \mathbb{T}$. It has been shown in (Acciaio and Penner 2011) that $\{\varphi_t^{\gamma_t}\}_{t \in \mathbb{T}}$ is strongly time consistent if and only if $\{\gamma_t\}_{t \in \mathbb{T}}$ is a constant process, and that it is middle acceptance time consistent if and only if $\{\gamma_t\}_{t \in \mathbb{T}}$ is a non-increasing process, and that it is middle rejection time consistent if and only if $\{\gamma_t\}_{t \in \mathbb{T}}$ is non-decreasing.

Example 10 (Dynamic Certainty Equivalent) *Dynamic Certainty Equivalents form a large class of dynamic risk measures, with Dynamic Entropic Risk Measure being a particular case. In this example, following Kupper and Schachermayer (2009), we consider an infinite time horizon, and take $\mathbb{T} = \mathbb{N}$ and $\mathcal{X} = L^\infty$. We let $U : \bar{\mathbb{R}} \rightarrow \bar{\mathbb{R}}$ be a strictly increasing and continuous function on $\bar{\mathbb{R}}$, i.e., strictly increasing and continuous on \mathbb{R} , with $U(\pm\infty) = \lim_{n \rightarrow \pm\infty} U(n)$. Let $\varphi = \{\varphi_t\}_{t \in \mathbb{T}}$ be defined by*

$$\varphi_t(X) = U^{-1}(E[U(X)|\mathcal{F}_t]), \quad X \in \mathcal{X}, \quad t \in \mathbb{T}. \tag{48}$$

It is easy to check that φ is a strongly time consistent dynamic LM-measure. It belongs to the class of so called dynamic certainty equivalents Kupper and Schachermayer (2009). In Kupper and Schachermayer (2009), the authors showed that every dynamic LM-measure, which is finite, normalized, strictly monotone, continuous, law invariant, admits The Fatou property, and is strongly time consistent, can be represented as (48) for some U . We also refer to Biagini and Bion-Nadal (2014) for a more general approach to dynamic certainty equivalents (e.g., by using stochastic utility functions U), and to Bielecki et al. (2016) for the definition of certainty equivalents for processes.

Example 11 (Dynamic Risk Sensitive Criterion) *In Bielecki et al. (2015a) the authors introduced the notion of the Dynamic Limit Growth Index (dLGI) that is designed to measure the long-term performance of a financial portfolio in discrete time. The dynamic analog of Risk Sensitive Criterion (cf. (Bielecki and Pliska 2003; Davis and Lleo 2014; Whittle 1990) and references therein) is a particular case of dLGI. We consider an infinite time horizon setup, $\mathbb{T} = \mathbb{N}$, and the following space suitable for our needs $\mathbb{V}_{\ln}^P := \{(W_t)_{t \in \mathbb{T}} : W_t > 0, \ln W_t \in L_t^P\}$. To be consistent with Bielecki et al. (2015a), we view the elements of \mathcal{X} as cumulative value processes of portfolios of some financial securities, which have integrable growth expressed as cumulative log-return (note that everywhere else in the present paper, the stochastic processes represent dividend streams). Let $\varphi^\gamma = \{\varphi_t^\gamma\}_{t \in \mathbb{T}}$ be defined by*

$$\varphi_t^\gamma(W) = \begin{cases} \liminf_{T \rightarrow \infty} \frac{1}{T} \frac{1}{\gamma} \ln E[W_T^\gamma|\mathcal{F}_t], & \text{if } \gamma \neq 0, \\ \liminf_{T \rightarrow \infty} \frac{1}{T} E[\ln W_T|\mathcal{F}_t], & \text{if } \gamma = 0, \end{cases} \tag{49}$$

where γ is a fixed real number. It was proved in Bielecki et al. (2015a) that φ^γ is a dynamic measure of performance, and it is μ -acceptance time consistent with respect to $\mu_t(m) = E[m|\mathcal{F}_t]$, $t \in \mathbb{T}$, if and only if $\gamma > 0$, and μ -rejection time consistent, with respect to μ if and only if $\gamma < 0$.

Taxonomy of examples

The following table is meant to help the reader to navigate through the examples presented above relative to various types of time consistency studied in this paper. We will use the following abbreviations for time consistency: WA - weak acceptance; WR - weak rejection; sWA - semi-weak acceptance; sWR - semi-weak rejection; MA - middle acceptance; MR - middle rejection; STR - strong; Sub - submartinagle; Sup - supermartinagle.

If a cell is marked with the check mark, that means that the corresponding property of time consistency is satisfied; otherwise the property is not satisfied in general.

We note that Example 11 is not represented in the table due to the distinct nature of the example. The dGLI evaluates a process V , but it does it through a limiting procedure, which really amounts to evaluating the process through its “values at $T = \infty$.” We refer the reader to (Bielecki et al. 2015a) for a detailed discussion on various properties of this measure.

	\mathcal{X}	WA	WR	sWA	sWR	MA	MR	STR	Sub	Sup
Example 3	L^p	✓	✓	✓	✓					
Example 4	L^p		✓		✓				✓	
Example 5	\mathbb{V}^p				✓					
Example 6	\mathbb{V}^p				✓					
Example 7	\mathbb{V}^p			✓	✓					
Example 8	$\gamma \geq 0$	L^p	✓	✓	✓	✓	✓	✓		✓
	$\gamma \leq 0$	L^p	✓	✓	✓	✓	✓	✓	✓	
Example 9	$\gamma_t \downarrow$	L^p	✓		✓	✓				✓*
	$\gamma_t \uparrow$	L^p		✓		✓	✓		✓**	
Example 10	L^p	✓	✓	✓	✓	✓	✓	✓		

*if $\gamma_t \geq 0$

**if $\gamma_t \leq 0$

Appendix

Here we provide a brief exposition of the three fundamental concepts used in the paper: the dynamic LM-measures, the conditional essential suprema/infima, and the LM-extensions.

Dynamic LM-measures

Let \mathcal{X} denote the space of random variables or adapted stochastic processes as described in Section “Mathematical preliminaries”.

We start with listing additional properties that may be enjoyed by a dynamic LM-measure. Let φ be a dynamic LM-measure. We say that φ is

- *Super-additive* if $\varphi_t(X + Y) \geq \varphi_t(X) + \varphi_t(Y)$;
- *Normalized* if $\varphi_t(0) = 0$;
- *Cash-additive* if $\varphi_t(X + m1_{\{t\}}) = \varphi_t(X) + m$;
- *Quasi-concave* if $\varphi_t(\lambda \cdot_t X + (1 - \lambda) \cdot_t Y) \geq \varphi_t(X) \wedge \varphi_t(Y)$;
- *Concave* if $\varphi_t(\lambda \cdot_t X + (1 - \lambda) \cdot_t Y) \geq \lambda\varphi_t(X) + (1 - \lambda)\varphi_t(Y)$;
- *Scale invariant* if $\varphi_t(\beta \cdot_t X) = \varphi_t(X)$;
- *Positively homogeneous* if $\varphi_t(\beta \cdot_t X) = \beta\varphi_t(X)$;
- *Lower semi-continuous with respect to the topology η* , if $\{Z \in \bar{L}_t^0 \mid \varphi_t(X) \leq Z\}$ is η -closed;¹⁶
- *Upper semi-continuous with respect to the topology η* , if $\{Z \in \bar{L}_t^0 \mid \varphi_t(X) \geq Z\}$ is η -closed,

for any $X, Y \in \mathcal{X}$, $t, s \in \mathbb{T}$, such that $s > t$, and $m, \lambda, \beta \in L_t^p$, such that $0 \leq \lambda \leq 1$, $\beta > 0$ and $\|\beta\|_\infty < \infty$. Moreover, if $\mathcal{X} = \mathbb{V}^p$, then we say that φ is

- *Independent of the past* if $\varphi_t(X) = \varphi_t(X - 0 \cdot_t X)$;
- *Translation invariant* if $\varphi_t(X + m1_{\{t\}}) = \varphi_t(X + m1_{\{s\}})$.

These last two properties are automatically satisfied for $\mathcal{X} = L^p$.

Most of the above properties have a natural financial interpretation. For example, *quasi-concavity*, *concavity*, or *super-additivity* correspond to the positive effect of portfolio diversification. See (Cherny and Madan 2009; Föllmer and Schied 2010) for more details and for a financial interpretation of other properties listed above.

Next, we recall the *Fatou property*, the *Lebesgue property*, as well as the *law-invariance property*. For simplicity, we present them only for the case of random variables. We say that a dynamic LM-measure φ admits

- *Fatou property*, if $\varphi_t(X) \geq \limsup_{n \rightarrow \infty} \varphi_t(X_n)$;
- *Lebesgue property*, if $\varphi_t(X) = \lim_{n \rightarrow \infty} \varphi_t(X_n)$;
- *Law-invariant property* if $\varphi_t(X) = \varphi_t(Y)$, whenever $Law(X) = Law(Y)$;

for any $t \in \mathbb{T}$, $X, Y \in \mathcal{X}$ and any dominated sequence¹⁷ $\{X_n\}_{n \in \mathbb{N}}$ such that $X_n \in L^p$ and $X_n \xrightarrow{a.s.} X$.

Classes of dynamic LM-measures

We say that a dynamic LM-measure φ is a

- *Dynamic monetary utility measure*, or just dynamic utility measure for short, if φ is translation invariant, independent of the past, normalized, monotone, and cash-additive;
- *Dynamic concave utility measure*, if φ is a dynamic utility measure and concave;

¹⁶That is closed with respect to topology η ; if η will be clear from the context, we will simply write that f is *lower semi-continuous*. If $\mathcal{X} = L^p$, then we use the topology induced by $\|\cdot\|_p$ norm (see (Föllmer and Schied 2004, Appendix A.7) for details).

¹⁷This means that there exist $Y \in \mathcal{X}$ such that for all $n \in \mathbb{N}$ we have $|X_n| \leq |Y|$.

- *Dynamic coherent utility measure*, if φ is a dynamic utility measure, is positive homogeneous, and super-additive;
- *Dynamic performance measure*, if φ is adapted, translation invariant, independent of the past, monotone increasing, and scale invariant;
- *Dynamic acceptability index*, if φ is a dynamic performance measure, and it is quasi-concave.

It needs to be stressed that in the literature, typically, the negative of the dynamic (monetary, concave, or coherent) utility measure is used and referred to as *dynamic (monetary, convex, or coherent) risk measure*.

Robust representations for dynamic monetary utility measures

Robust representations have been studied for general dynamic LM-measures, not only for dynamic monetary utility measures. However, in this paper we only use robust representations for dynamic monetary utility measures for random variables, and that is why our discussion here is limited to this case. Consequently, we take $\mathcal{X} = L^p$ for a fixed $p \in \{0, 1, \infty\}$.

Let φ be a dynamic monetary utility measure. We associate with φ the following families of objects:

- *acceptance and rejection sets* denoted by $\mathcal{A} = \{\mathcal{A}_t\}_{t \in \mathbb{T}}$ and $\mathcal{R} = \{\mathcal{R}_t\}_{t \in \mathbb{T}}$, respectively, where

$$\begin{aligned} \mathcal{A}_t &:= \{X \in L^p : \varphi_t(X) \geq 0\}, \\ \mathcal{R}_t &:= \{X \in L^p : \varphi_t(X) \leq 0\}. \end{aligned}$$

- *conditional acceptance and conditional rejection sets* denoted by $\{\mathcal{A}_{t,s} : t, s \in \mathbb{T}, s > t\}$ and $\{\mathcal{R}_{t,s} : t, s \in \mathbb{T}, s > t\}$, respectively, where

$$\begin{aligned} \mathcal{A}_{t,s} &:= \{X \in L^p \cap \bar{L}_s^0 : \varphi_t(X) \geq 0\}, \\ \mathcal{R}_{t,s} &:= \{X \in L^p \cap \bar{L}_s^0 : \varphi_t(X) \leq 0\}. \end{aligned}$$

- *minimal penalty functions* denoted by $\alpha^{\min} = \{\alpha_t^{\min}\}_{t \in \mathbb{T}}$, where $\alpha_t^{\min} : \mathcal{M}(P) \rightarrow \mathbb{R}$ is given by

$$\alpha_t^{\min}(Q) := -\text{ess inf}_{X \in \mathcal{A}_t} E_Q[X | \mathcal{F}_t].$$

- *conditional minimal penalty functions* denoted by $\{\alpha_{t,s}^{\min} : t, s \in \mathbb{T}, s > t\}$, where $\alpha_{t,s}^{\min} : \mathcal{M}(P) \rightarrow \bar{L}_t^0$ is given by

$$\alpha_{t,s}^{\min}(Q) := -\text{ess inf}_{X \in \mathcal{A}_{t,s}} E_Q[X | \mathcal{F}_t].$$

The following important definition is frequently used in this paper.

Definition 15 *Let φ be a dynamic monetary utility measure. We call φ representable, if*

$$\varphi_t(X) = \text{ess inf}_{Q \in \mathcal{M}(P)} (E_Q[X | \mathcal{F}_t] + \alpha_t^{\min}(Q)), \tag{50}$$

for any $X \in \mathcal{X}$.

This type of representation is called robust or numerical representations. Moreover, such representation characterizes dynamic concave utility measures that admit the Fatou property.

Conditional expectation and conditional essential supremum/infimum

We present here some relevant properties of the generalized conditional expectation and conditional essential superemum and infimum, in the context of \bar{L}^0 .

Proposition 15 *For any $X, Y \in \bar{L}^0$ and $s, t \in \mathbb{T}, s > t$, it holds that*

- 1) $E[\lambda X|\mathcal{F}_t] \leq \lambda E[X|\mathcal{F}_t]$ for $\lambda \in L_t^0$, and $E[\lambda X|\mathcal{F}_t] = \lambda E[X|\mathcal{F}_t]$ for $\lambda \in L_t^0, \lambda \geq 0$;
- 2) $E[X|\mathcal{F}_t] \leq E[E[X|\mathcal{F}_s]|\mathcal{F}_t]$, and $E[X|\mathcal{F}_t] = E[E[X|\mathcal{F}_s]|\mathcal{F}_t]$ for $X \geq 0$;
- 3) $E[X|\mathcal{F}_t] + E[Y|\mathcal{F}_t] \leq E[X + Y|\mathcal{F}_t]$, and $E[X|\mathcal{F}_t] + E[Y|\mathcal{F}_t] = E[X + Y|\mathcal{F}_t]$, if $X, Y \geq 0$.

For the proof, see (Bielecki et al. 2014a, Proposition A.1).

Remark 11 *All inequalities in Proposition 15 can be strict. Assume that $t = 0$ and $k, s \in \mathbb{T}, k > s > 0$, and let $\xi \in L_k^0$ be such that $\xi = \pm 1$, ξ is independent of \mathcal{F}_s , and $P(\xi = 1) = P(\xi = -1) = 1/2$. We consider $Z \in L_s^0$ such that $Z \geq 0$, and $E[Z] = \infty$. By taking $\lambda = -1, X = \xi Z$ and $Y = -X$, we get strict inequalities in 1), 2), and 3).*

We proceed with presenting a definition of the conditional essential infimum and supremum that is equivalent to the one presented in Section ‘‘Mathematical preliminaries’’, cf. (51). We start by recalling the definition of the conditional essential infimum for bounded random variables. For $X \in L^\infty$ and $t \in \mathbb{T}$, we denote by $\text{ess inf}_t X$ the unique (up to a set of measure zero), the \mathcal{F}_t -measurable random variable, such that for any $A \in \mathcal{F}_t$, the following equality holds true

$$\text{ess inf}_{\omega \in A} X = \text{ess inf}_{\omega \in A} (\text{ess inf}_t X). \tag{51}$$

We call this random variable the \mathcal{F}_t -conditional essential infimum of X . Accordingly, we define $\text{ess sup}_t(X) := -\text{ess inf}_t(-X)$, the \mathcal{F}_t -conditional essential supremum of $X \in L^\infty$. The reader is referred to (Barron et al. 2003) for a proof of the existence and uniqueness of the conditional essential supremum/infimum.

Consequently, for any $t \in \mathbb{T}$ and $X \in \bar{L}^0$, we define the \mathcal{F}_t -conditional essential infimum by

$$\text{ess inf}_t X := \lim_{n \rightarrow \infty} \left[\text{ess inf}_t (X^+ \wedge n) \right] - \lim_{n \rightarrow \infty} \left[\text{ess sup}_t (X^- \wedge n) \right]. \tag{52}$$

Respectively, we put $\text{ess sup}_t(X) := -\text{ess inf}_t(-X)$.

Proposition 16 *For any $X, Y \in \bar{L}^0, s, t \in \mathbb{T}, s \geq t$, and $A \in \mathcal{F}_t$ the following properties hold,*

- 1) $\text{ess inf}_{\omega \in A} X = \text{ess inf}_{\omega \in A} (\text{ess inf}_t X)$;
- 2) If $\text{ess inf}_{\omega \in A} X = \text{ess inf}_{\omega \in A} U$ for some $U \in \bar{L}_t^0$, then $U = \text{ess inf}_t X$;
- 3) $X \geq \text{ess inf}_t X$;
- 4) If $Z \in \bar{L}_t^0$, is such that $X \geq Z$, then $\text{ess inf}_t X \geq Z$;
- 5) If $X \geq Y$, then $\text{ess inf}_t X \geq \text{ess inf}_t Y$;
- 6) $\mathbb{1}_A \text{ess inf}_t X = \mathbb{1}_A \text{ess inf}_t (\mathbb{1}_A X)$;
- 7) $\text{ess inf}_s X \geq \text{ess inf}_t X$;

Analogous results are true for $\{\text{ess sup}_t\}_{t \in \mathbb{T}}$.

The proof for the case $X, Y \in L^\infty$ can be found in (Barron et al. 2003). Since for any $n \in \mathbb{N}$ and $X, Y \in \bar{L}^0$, we get $X^+ \wedge n \in L^\infty$, $X^- \wedge n \in L^\infty$, and $X^+ \wedge X^- = 0$, the extension of the proof to the case $X, Y \in \bar{L}^0$ is straightforward.

It is worth mentioning that properties 3) and 4) from Proposition 16 imply that the conditional essential infimum $\text{ess inf}_t(X)$ can be defined as the largest \mathcal{F}_t -measurable random variable, which is smaller than X (cf. (Barron et al. 2003)).

Next, we define the generalized versions of ess inf and ess sup of a (possibly uncountable) family of random variables. For $\{X_i\}_{i \in I}$, where $X_i \in \bar{L}^0$, we let

$$\text{ess inf}_{i \in I} X_i := \lim_{n \rightarrow \infty} [\text{ess inf}_{i \in I} (X_i^+ \wedge n)] - \lim_{n \rightarrow \infty} [\text{ess sup}_{i \in I} (X_i^- \wedge n)]. \quad (53)$$

Note that, in view of (Karatzas and Shreve 1998, Appendix A), $\text{ess inf}_{i \in I} X_i \wedge n$ and $\text{ess sup}_{i \in I} X_i \wedge n$ are well defined, so that $\text{ess inf}_{i \in I} X_i$ is well defined. It needs to be observed that the operations of the right-hand side of (53) preserve measurability. In particular, if $X_i \in \mathcal{F}_t$ for all $i \in I$, then $\text{ess inf}_{i \in I} X_i \in \mathcal{F}_t$.

Furthermore, if for any $i, j \in I$, there exists $k \in I$, such that $X_k \leq X_i \wedge X_j$, then there exists a sequence $i_n \in I, n \in \mathbb{N}$, such that $\{X_{i_n}\}_{n \in \mathbb{N}}$ is non-increasing and $\text{ess inf}_{i \in I} X_i = \inf_{n \in \mathbb{N}} X_{i_n} = \lim_{n \rightarrow \infty} X_{i_n}$. Analogous results hold true for $\text{ess sup}_{i \in I} X_i$.

LM-extensions

In this part of the appendix, we introduce the concept of an LM-extension of a dynamic LM-measure for random variables.

Definition 16 Let φ be a dynamic LM-measure on L^p . We call a family $\widehat{\varphi} = \{\widehat{\varphi}_t\}_{t \in \mathbb{T}}$ of maps $\widehat{\varphi}_t : \bar{L}_t^0 \rightarrow \bar{L}_t^0$ an LM-extension of φ , if for any $t \in \mathbb{T}$, $\widehat{\varphi}_t|_{\mathcal{X}} \equiv \varphi_t$, and $\widehat{\varphi}_t$ is local and monotone on \bar{L}_0 .¹⁸

We will show below that such extension exists, for which we will make use of the following auxiliary sets:

$$\mathcal{Y}_A^+(X) := \{Y \in \mathcal{X} \mid \mathbb{1}_A Y \geq \mathbb{1}_A X\}, \quad \mathcal{Y}_A^-(X) := \{Y \in \mathcal{X} \mid \mathbb{1}_A Y \leq \mathbb{1}_A X\},$$

defined for any $X \in \bar{L}^0$ and $A \in \mathcal{F}$.

¹⁸That is, it satisfies monotonicity and locality on \bar{L}_0 , as in 5) and 6) in Proposition 16.

Definition 17 Let φ be a dynamic LM-measure. The collection of functions $\varphi^+ = \{\varphi_t^+\}_{t \in \mathbb{T}}$, where $\varphi_t^+ : \bar{L}^0 \rightarrow \bar{L}_t^0$ is defined as¹⁹

$$\varphi_t^+(X) := \operatorname{ess\,inf}_{A \in \mathcal{F}_t} \left[\mathbb{1}_A \operatorname{ess\,inf}_{Y \in \mathcal{Y}_A^+(X)} \varphi_t(Y) + \mathbb{1}_{A^c} (+\infty) \right], \tag{54}$$

is called the upper LM-extension of φ . Respectively, the collection of functions $\varphi^- = \{\varphi_t^-\}_{t \in \mathbb{T}}$, where $\varphi_t^- : \bar{L}^0 \rightarrow \bar{L}_t^0$, and

$$\varphi_t^-(X) := \operatorname{ess\,sup}_{A \in \mathcal{F}_t} \left[\mathbb{1}_A \operatorname{ess\,sup}_{Y \in \mathcal{Y}_A^-(X)} \varphi_t(Y) + \mathbb{1}_{A^c} (-\infty) \right], \tag{55}$$

is called the lower LM-extension of φ .

The next result shows that φ^\pm are two “extreme” extensions, and any other extension is sandwiched between them.

Proposition 17 Let φ be a dynamic LM-measure. Then, φ^- and φ^+ are LM-extensions of φ . Moreover, let $\widehat{\varphi}$ be an LM-extension of φ . Then, for any $X \in \bar{L}^0$ and $t \in \mathbb{T}$,

$$\varphi_t^-(X) \leq \widehat{\varphi}_t(X) \leq \varphi_t^+(X). \tag{56}$$

Clearly, in general, the maps (54) and (55) are not equal, and thus the extensions of an LM-measure are not unique.

Remark 12 Let $t \in \mathbb{T}$ and $\mathcal{B} \subseteq \bar{L}^0$ be such that, for any $A \in \mathcal{F}_t$, $\mathbb{1}_A \mathcal{B} \subseteq \mathcal{B}$, and $\mathbb{1}_{A^c} \mathcal{B} + \mathbb{1}_A \mathcal{B} \subseteq \mathcal{B}$. As a generalization of Proposition 17, one can show that for any \mathcal{F}_t -local and monotone mapping²⁰ $f : \mathcal{B} \rightarrow \bar{L}_t^0$, the maps f^\pm defined analogously as in (54) and (55) are extensions of f to \bar{L}^0 , preserving locality and monotonicity.

Remark 13 For a large class of LM-measures, as mentioned earlier, there exists a “robust representation” type theorem—essentially a representation, via convex duality, as a function of conditional expectation. We refer the reader to (Bielecki et al. 2016) and references therein, where the authors present a general robust representation for dynamic quasi-concave upper semi-continuous LM-measures. Hence, an alternative construction of extensions can be obtained through the robust representations of LM-measures, by considering conditional expectations defined on the extended real number line, etc.

¹⁹We will use the convention $\operatorname{ess\,sup} \emptyset = -\infty$ and $\operatorname{ess\,inf} \emptyset = \infty$.

²⁰That is, \mathcal{F}_t -local and monotone on \mathcal{B} .

Proofs

Proof of proposition 4

Proof 1) \Rightarrow 2). Let $t, s \in \mathbb{T}$ be such that $s > t$, and consider the following set

$$\mathcal{X}_{\varphi_s} = \{X \in \bar{L}^0 \mid X = \varphi_s(Y) \text{ for some } Y \in \mathcal{X}\},$$

where $\mathcal{X} = L^p$. From 1), for any $X, Y \in \mathcal{X}$, such that $\varphi_s(X) = \varphi_s(Y)$, we get $\varphi_t(X) = \varphi_t(Y)$. Next, we define the map $\phi_{t,s} : \mathcal{X}_{\varphi_s} \rightarrow \bar{L}_t^0$ as follows: for any $X' \in \mathcal{X}_{\varphi_s}$

$$\phi_{t,s}(X') = \varphi_t(X), \quad X \in \mathcal{X}, \tag{57}$$

where $X \in \mathcal{X}$ is such that $X' = \varphi_s(X)$. In view of the definition of \mathcal{X}_{φ_s} and strong time consistency of φ , the map $\phi_{t,s}$ is well-defined.

Since there exists $Z \in \mathcal{X}$, such that $\varphi_s(Z) = 0$ (see property (10)), using locality of φ , we get that for any $X \in \mathcal{X}_{\varphi_s}$, $A \in \mathcal{F}_t$, there exists $Y \in \mathcal{X}$, so that

$$\mathbb{1}_A X = \mathbb{1}_A \varphi_s(Y) = \mathbb{1}_A \varphi_s(\mathbb{1}_A Y) + \mathbb{1}_{A^c} \varphi_s(\mathbb{1}_{A^c} Z) = \varphi_s(\mathbb{1}_A Y + \mathbb{1}_{A^c} Z).$$

Thus, $\mathbb{1}_A X \in \mathcal{X}_{\varphi_s}$, for any $A \in \mathcal{F}_t$, $X \in \mathcal{X}_{\varphi_s}$. Hence, from 1) and the locality of φ , for any $X, Y \in \mathcal{X}_{\varphi_s}$, $A \in \mathcal{F}_t$, we get

- (A) $X \geq Y \Rightarrow \phi_{t,s}(X) \geq \phi_{t,s}(Y)$;
- (B) $\mathbb{1}_A \phi_{t,s}(X) = \mathbb{1}_A \phi_{t,s}(\mathbb{1}_A X)$.

In other words, $\phi_{t,s}$ is local and monotone on $\mathcal{X}_{\varphi_s} \subseteq \bar{L}_s^0$. By Remark 12), there exists an extension of $\phi_{t,s}$, say $\widehat{\phi}_{t,s} : \bar{L}_s^0 \rightarrow \bar{L}_t^0$, which is local and monotone on \bar{L}_s^0 . Finally, we take $\mu_{t,s} : \bar{L}_s^0 \rightarrow \bar{L}_t^0$ defined by

$$\mu_{t,s}(m) := \widehat{\phi}_{t,s}(m), \quad m \in \bar{L}_s^0.$$

Clearly, the family $\mu = \{\mu_{t,s} : t, s \in \mathbb{T}, s > t\}$ is an update rule, and using (57), we get that φ is both μ -acceptance and μ -rejection time consistent.

2) \Rightarrow 3). Let $s, t \in \mathbb{T}$ and $X, Y \in \mathcal{X}$ be such that $s > t$ and $\varphi_s(X) \geq \varphi_s(Y)$. From 2),(12), and by the monotonicity of μ , we have $\varphi_t(X) = \mu_{t,s}(\varphi_s(X)) \geq \mu_{t,s}(\varphi_s(Y)) = \varphi_t(Y)$.

3) \Rightarrow 1), 4) \Leftrightarrow 2), and 4) \Rightarrow 5) are obvious.

5) \Rightarrow 4). Let a family $\tilde{\mu} = \{\tilde{\mu}_{t,s} : t, s \in \mathbb{T}, t < s\}$ of maps $\tilde{\mu}_{t,s} : \bar{L}_s^0 \rightarrow \bar{L}_t^0$ be given by

$$\begin{cases} \tilde{\mu}_{t,s}(\cdot) := \mu_{t,t+1}(\cdot) & \text{if } s = t + 1, \\ \tilde{\mu}_{t,s}(\cdot) := \mu_{t,t+1} \circ \dots \circ \mu_{s-1,s}(\cdot) & \text{if } s > t + 1, \end{cases}$$

where μ is an update rule from 5). It is straightforward to check that $\tilde{\mu}$ is an update rule, and that φ is both $\tilde{\mu}$ -acceptance and $\tilde{\mu}$ -rejection time consistent, which proves that 4) holds.

The proof is complete. □

Proof of proposition 6

Proof Let us consider $\{\phi_t\}_{t \in \mathbb{T}}$ as given in (19).

1) The proof of monotonicity and locality is similar to the one for the conditional essential infimum and supremum, Proposition 16. Finally, for any $t \in \mathbb{T}$, $Z \in D_t$, and $m \in \bar{L}_t^0$, since $E[Z|\mathcal{F}_t] = 1$, we immediately get

$$E[Zm|\mathcal{F}_t] = 1_{\{m \geq 0\}}mE[Z|\mathcal{F}_t] + 1_{\{m < 0\}}(-m)E[-Z|\mathcal{F}_t] = m, \tag{58}$$

and thus, $\phi_t(m) = m$, for any $m \in \bar{L}_t^0$. Hence, $\{\phi_t\}_{t \in \mathbb{T}}$ is projective.

2) Let φ be a dynamic LM-measure which is ϕ -rejection time consistent, and $g : \mathbb{R} \rightarrow \mathbb{R}$ be an increasing, concave function. Then, for any $X \in \mathcal{X}$, we get

$$g(\varphi_t(X)) \geq g(\phi_t(\varphi_s(X))) = g(\operatorname{ess\,inf}_{Z \in D_t} E[Z\varphi_s(X)|\mathcal{F}_t]) = \operatorname{ess\,inf}_{Z \in D_t} g(E[Z\varphi_s(X)|\mathcal{F}_t]). \tag{59}$$

Recall that any $Z \in D_t$ is a Radon-Nikodym derivative of some measure Q with respect to P , and thus we have $E[Z\varphi_s(X)|\mathcal{F}_t] = E_Q[\varphi_s(X)|\mathcal{F}_t]$. Hence, by Jensen's inequality, we deduce

$$\operatorname{ess\,inf}_{Z \in D_t} g(E[Z\varphi_t(X)|\mathcal{F}_t]) \geq \operatorname{ess\,inf}_{Z \in D_t} E[Zg(\varphi_t(X))|\mathcal{F}_t] = \phi_t(g(\varphi_s(X))). \tag{60}$$

Combining (59) and (60), ϕ -acceptance time consistency of $\{g \circ \varphi_t\}_{t \in \mathbb{T}}$ follows. \square

Proof of proposition 8

Proof The first part follows immediately from the definition of LM-extension. Clearly, projectivity of $\widehat{\varphi}$ implies that $\varphi_t(X) = X$, for $X \in \mathcal{X} \cap \bar{L}_t^0$. To prove the opposite implication, it is enough to prove that φ^+ and φ^- are projective. Assume that φ is such that $\varphi_t(X) = X$, for $t \in \mathbb{T}$ and $X \in L^p \cap \bar{L}_t^0$. Let $X \in \bar{L}_t^0$. For any $n \in \mathbb{N}$, we get

$$\begin{aligned} \mathbb{1}_{\{n \geq X \geq -n\}}\varphi_t^+ &= \mathbb{1}_{\{n \geq X \geq -n\}}\varphi_t^+(\mathbb{1}_{\{n \geq X \geq -n\}}X) = \mathbb{1}_{\{n \geq X \geq -n\}}\varphi_t(\mathbb{1}_{\{n \geq X \geq -n\}}X) \\ &= \mathbb{1}_{\{n \geq X \geq -n\}}X. \end{aligned}$$

Thus, on set $\bigcup_{n \in \mathbb{N}}\{-n \leq X \leq n\} = \{-\infty < X < \infty\}$, we have

$$\varphi_t^+(X) = X, \quad \text{for } X \in \bar{L}_t^0. \tag{61}$$

Next, for any $A \in \mathcal{F}_t$, such that $A \subseteq \{X = \infty\}$, we get $\mathcal{Y}_A^+(X) = \emptyset$, which implies $\mathbb{1}_{\{X = \infty\}}\varphi^+(X) = \infty$. Finally, for any $n \in \mathbb{R}$, using locality of φ_t^+ and the fact that $n \in \mathcal{X} \cap \bar{L}_t^0$, we get

$$\mathbb{1}_{\{X = -\infty\}}\varphi_t^+(X) \leq \mathbb{1}_{\{X = -\infty\}}\varphi_t^+(\mathbb{1}_{\{X = -\infty\}}n) = \mathbb{1}_{\{X = -\infty\}}\varphi_t(n) = \mathbb{1}_{\{X = -\infty\}}n,$$

which implies $\mathbb{1}_{\{X = -\infty\}}\varphi^+(X) = -\infty$. Hence, (61) holds true on entire space. The proof for φ^- is analogous. \square

Proof of proposition 14

Proof Let φ be a dynamic LM-measure, which is independent of the past.

1) \Rightarrow 2). Let $t \in \mathbb{T}'$ and consider the following set

$$\mathcal{X}_{\varphi_{t+1}} = \{X \in \bar{L}^0 \mid X = \varphi_{t+1}(V) \text{ for some } V \in \mathbb{V}^p\}.$$

From 1), for any $V, V' \in \mathcal{X}$, such that $\varphi_{t+1}(V) = \varphi_{t+1}(V')$ and $V_t = V'_t$, we get $\varphi_t(X) = \varphi_t(Y)$. Thus, using the independence of the past of φ , there exists a map $\phi_{t,t+1} : \mathcal{X}_{\varphi_{t+1}} \times L_t^p \rightarrow \bar{L}_t^0$ such that

$$\phi_{t,t+1}(\varphi_{t+1}(X), Y_t) = \varphi_t(X - 1_{\{t\}}(X_t - Y_t)), \quad X \in \mathcal{X}.$$

Next, since there exists $Z \in \mathcal{X}$, such that $\varphi_{t+1}(Z) = 0$, using the locality of φ , we get that for any $X \in \mathcal{X}_{\varphi_{t+1}}$, $A \in \mathcal{F}_t$, there exist $Y \in \mathcal{X}$, so that

$$\begin{aligned} \mathbb{1}_A X &= \mathbb{1}_A \varphi_{t+1}(Y) = \mathbb{1}_A \varphi_{t+1}(\mathbb{1}_A \cdot_{t+1} Y) + \mathbb{1}_{A^c} \varphi_{t+1}(\mathbb{1}_{A^c} \cdot_{t+1} Z) \\ &= \varphi_{t+1}(\mathbb{1}_A \cdot_{t+1} Y + \mathbb{1}_{A^c} \cdot_{t+1} Z). \end{aligned}$$

Thus, $\mathbb{1}_A X \in \mathcal{X}_{\varphi_{t+1}}$, for any $A \in \mathcal{F}_t$, $X \in \mathcal{X}_{\varphi_{t+1}}$. Hence, from 2) and the locality of φ , for any $X, X' \in \mathcal{X}_{\varphi_{t+1}}$, $Y_t \in L_t^p$ and $A \in \mathcal{F}_t$, we get

- (A) $X \geq X' \Rightarrow \phi_{t,t+1}(X, Y_t) \geq \phi_{t,t+1}(X', Y_t)$;
- (B) $\mathbb{1}_A \phi_{t,t+1}(X, Y_t) = \mathbb{1}_A \phi_{t,t+1}(\mathbb{1}_A X, Y_t)$.

In other words, for any fixed $Y_t \in L_t^p$, $\phi_{t,t+1}(\cdot, Y_t)$ is local and monotone on $\mathcal{X}_{\varphi_{t+1}} \subseteq \bar{L}_{t+1}^0$. In view of Remark 12, for any fixed $Y_t \in L_t^p$ there exists an extension (to \bar{L}_{t+1}^0) of $\phi_{t,t+1}(\cdot, Y_t)$, say $\widehat{\phi}_{t,t+1}(\cdot, Y_t)$, which is local and monotone on \bar{L}_{t+1}^0 . Finally, we take $\mu_{t,t+1} : \bar{L}_{t+1}^0 \times \mathcal{X} \rightarrow \bar{L}_t^0$ defined by

$$\mu_{t,t+1}(m, X) := \widehat{\phi}_{t,t+1}(m, X_t), \quad X \in \mathcal{X}, m \in \bar{L}_{t+1}^0.$$

Clearly, the family $\mu_{t,t+1}$ is a (one step) update rule. Moreover, we get

$$\mu_{t,t+1}(m, X) = \mu_{t,t+1}(m, X'),$$

for $m \in \bar{L}_{t+1}^0$ and $X, X' \in \mathcal{X}$, such that $X_t = X'_t$. Finally, φ is both μ -acceptance and μ -rejection time consistent, as

$$\varphi_t(X) = \varphi_t(X - 1_{\{t\}}(X_t - X_t)) = \phi_{t,t+1}(\varphi_{t+1}(X), X_t) = \mu_{t,t+1}(\varphi_{t+1}(X), X).$$

2) \Rightarrow 1). Assume that μ is an update rule, fulfilling 2), such that φ is both μ -acceptance and μ -rejection time consistent. Then, we get $\varphi_t(X) = \mu_{t,t+1}(\varphi_{t+1}(X), Y)$, for any $t \in \mathbb{T}'$, $X \in \mathcal{X}$, and $Y \in \mathcal{X}$, such that $X_t = Y_t$. Let $t \in \mathbb{T}'$ and $X, Y \in \mathcal{X}$ be such that $X_t = Y_t$ and $\varphi_{t+1}(X) \geq \varphi_{t+1}(Y)$. From the above, and by monotonicity of μ , we have

$$\varphi_t(X) = \mu_{t,t+1}(\varphi_{t+1}(X), X) = \mu_{t,t+1}(\varphi_{t+1}(X), Y) \geq \mu_{t,t+1}(\varphi_{t+1}(Y), Y) = \varphi_t(Y).$$

The proof of the equivalence between 2) and 3) is straightforward and hence omitted here. □

Proof of Proposition 17

Proof We show the proof for φ^+ only; the proof for φ^+ is similar. Consider a fixed $t \in \mathbb{T}$.

(Adaptivity) It is easy to note that for any $X \in \bar{L}^0$, and $A \in \mathcal{F}_t$, we get

$$\left[\mathbb{1}_A \operatorname{ess\,inf}_{Y \in \mathcal{Y}_A^+(X)} \varphi_t(Y) + \mathbb{1}_{A^c}(\infty) \right] \in \bar{L}_t^0. \tag{62}$$

Indeed, for any $X \in \bar{L}^0$, ess inf of the set of \mathcal{F}_t -measurable random variables $\{\varphi_t(Y)\}_{Y \in \mathcal{Y}_A^+(X)}$ is \mathcal{F}_t -measurable (see (Karatzas and Shreve 1998), Appendix A), which implies (62) for any $A \in \mathcal{F}_t$. Thus, $\varphi_t^+(X) \in \bar{L}_t^0$.

(Monotonicity) If $X \geq X'$ then for any $A \in \mathcal{F}_t$ we get $\mathcal{Y}_A^+(X) \subseteq \mathcal{Y}_A^+(X')$, and consequently, for any $A \in \mathcal{F}_t$,

$$\mathbb{1}_A \text{ess inf}_{Y \in \mathcal{Y}_A^+(X)} \varphi_t(Y) \geq \mathbb{1}_A \text{ess inf}_{Y \in \mathcal{Y}_A^+(X')} \varphi_t(Y),$$

which implies $\varphi_t^+(X) \geq \varphi_t^+(X')$.

(Locality) Let $B \in \mathcal{F}_t$ and $X \in \bar{L}^0$. It is enough to consider $A \in \mathcal{F}_t$, such that $\mathcal{Y}_A^+(X) \neq \emptyset$, as otherwise we get $\varphi_t^+(X) \equiv \infty$. For any such $A \in \mathcal{F}_t$, we get

$$\mathbb{1}_{A \cap B} \text{ess inf}_{Y \in \mathcal{Y}_A^+(X)} \varphi_t(Y) = \mathbb{1}_{A \cap B} \text{ess inf}_{Y \in \mathcal{Y}_{A \cap B}^+(X)} \varphi_t(Y). \tag{63}$$

Indeed, let us assume that $\mathcal{Y}_A^+(X) \neq \emptyset$. As $\mathcal{Y}_A^+(X) \subseteq \mathcal{Y}_{A \cap B}^+(X)$, we have

$$\mathbb{1}_{A \cap B} \text{ess inf}_{Y \in \mathcal{Y}_A^+(X)} \varphi_t(Y) \geq \mathbb{1}_{A \cap B} \text{ess inf}_{Y \in \mathcal{Y}_{A \cap B}^+(X)} \varphi_t(Y).$$

On the other hand, for any $Y \in \mathcal{Y}_{A \cap B}^+(X)$, and any fixed $Z \in \mathcal{Y}_A^+(X)$ (note that $\mathcal{Y}_A^+(X) \neq \emptyset$), we get

$$\mathbb{1}_B Y + \mathbb{1}_{B^c} Z \in \mathcal{Y}_A^+(X).$$

Thus, using the locality of φ_t , we deduce

$$\mathbb{1}_{A \cap B} \text{ess inf}_{Y \in \mathcal{Y}_{A \cap B}^+(X)} \varphi_t(Y) = \mathbb{1}_{A \cap B} \text{ess inf}_{Y \in \mathcal{Y}_{A \cap B}^+(X)} \mathbb{1}_B \varphi_t(\mathbb{1}_B Y + \mathbb{1}_{B^c} Z) \geq \mathbb{1}_{A \cap B} \text{ess inf}_{Y \in \mathcal{Y}_A^+(X)} \varphi_t(Y),$$

which proves (63). It is easy to see that $\mathcal{Y}_{A \cap B}^+(X) = \mathcal{Y}_{A \cap B}^+(\mathbb{1}_B X)$, and thus

$$\mathbb{1}_A \text{ess inf}_{Y \in \mathcal{Y}_{A \cap B}^+(X)} \varphi_t(Y) = \mathbb{1}_A \text{ess inf}_{Y \in \mathcal{Y}_{A \cap B}^+(\mathbb{1}_B X)} \varphi_t(Y). \tag{64}$$

Combining (63), (64), and the fact that $\mathcal{Y}_A^+(X) \neq \emptyset$ implies $\mathcal{Y}_A^+(\mathbb{1}_B X) \neq \emptyset$, we continue

$$\begin{aligned} \mathbb{1}_B \varphi_t^+(X) &= \mathbb{1}_B \text{ess inf}_{A \in \mathcal{F}_t} \left[\mathbb{1}_A \text{ess inf}_{Y \in \mathcal{Y}_A^+(X)} \varphi_t(Y) + \mathbb{1}_{A^c}(\infty) \right] \\ &= \mathbb{1}_B \text{ess inf}_{A \in \mathcal{F}_t} \left[\mathbb{1}_{A \cap B} \text{ess inf}_{Y \in \mathcal{Y}_A^+(X)} \varphi_t(Y) + \mathbb{1}_{A^c \cap B}(\infty) \right] \\ &= \mathbb{1}_B \text{ess inf}_{A \in \mathcal{F}_t} \left[\mathbb{1}_{A \cap B} \text{ess inf}_{Y \in \mathcal{Y}_{A \cap B}^+(X)} \varphi_t(Y) + \mathbb{1}_{A^c \cap B}(\infty) \right] \\ &= \mathbb{1}_B \text{ess inf}_{A \in \mathcal{F}_t} \left[\mathbb{1}_{A \cap B} \text{ess inf}_{Y \in \mathcal{Y}_{A \cap B}^+(\mathbb{1}_B X)} \varphi_t(Y) + \mathbb{1}_{A^c \cap B}(\infty) \right] \\ &= \mathbb{1}_B \text{ess inf}_{A \in \mathcal{F}_t} \left[\mathbb{1}_A \text{ess inf}_{Y \in \mathcal{Y}_A^+(\mathbb{1}_B X)} \varphi_t(Y) + \mathbb{1}_{A^c}(\infty) \right] \\ &= \mathbb{1}_B \varphi_t^+(\mathbb{1}_B X). \end{aligned}$$

(Extension) If $X \in \mathcal{X}$, then for any $A \in \mathcal{F}_t$, we get $X \in \mathcal{Y}_A^+(X)$. Thus,

$$\varphi_t^+(X) = \text{ess inf}_{A \in \mathcal{F}_t} \left[\mathbb{1}_A \text{ess inf}_{Y \in \mathcal{Y}_A^+(X)} \varphi_t(Y) + \mathbb{1}_{A^c}(\infty) \right] = \text{ess inf}_{A \in \mathcal{F}_t} \left[\mathbb{1}_A \varphi_t(X) + \mathbb{1}_{A^c}(\infty) \right] = \varphi_t(X).$$

As the above results are true for any $t \in \mathbb{T}$, thus we have proved that φ^+ is an extension of φ . Let us now show (56) for φ^+ .

Let $\widehat{\varphi}$ be an extension of φ , and let $X \in \bar{L}^0$ and $t \in \mathbb{T}$. Due to monotonicity and locality of $\widehat{\varphi}_t$, for any $A \in \mathcal{F}_t$ and $Y \in \mathcal{Y}_A^+(X)$, we get $\mathbb{1}_A \widehat{\varphi}_t(X) \leq \mathbb{1}_A \widehat{\varphi}_t(Y)$. Thus, recalling that $\text{ess inf } \emptyset = \infty$, we have

$$\widehat{\varphi}_t(X) \leq \mathbb{1}_A \text{ess inf}_{Y \in \mathcal{Y}_A^+(X)} \widehat{\varphi}_t(Y) + \mathbb{1}_{A^c}(\infty) = \mathbb{1}_A \text{ess inf}_{Y \in \mathcal{Y}_A^+(X)} \varphi_t(Y) + \mathbb{1}_{A^c}(\infty). \quad (65)$$

Since (65) holds true for any $A \in \mathcal{F}_t$, we conclude that

$$\widehat{\varphi}_t(X) \leq \text{ess inf}_{A \in \mathcal{F}_t} \left[\mathbb{1}_A \text{ess inf}_{Y \in \mathcal{Y}_A^+(X)} \varphi_t(Y) + \mathbb{1}_{A^c}(\infty) \right] = \varphi_t^+(X).$$

The proof of the second inequality is analogous. \square

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Authors' contributions

All authors read and approved the final manuscript.

Competing interests

The authors declare that they have no competing interests.

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