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Validation of triple pass 24-hour dietary recall in Ugandan children by simultaneous weighed food assessment

Helen Nightingale^{1†}, Kevin J. Walsh^{1†}, Peter Olupot-Olupot^{2,3}, Charles Engoru⁴, Tonny Ssenyondo³, Julius Nteziyaremye³, Denis Amorut⁴, Margaret Nakuya⁴, Margaret Arimi⁴, Gary Frost^{1*} and Kathryn Maitland^{5,6*} 

Abstract

Background: Undernutrition remains highly prevalent in African children, highlighting the need for accurately assessing dietary intake. In order to do so, the assessment method must be validated in the target population. A triple pass 24 h dietary recall with volumetric portion size estimation has been described but not previously validated in African children. This study aimed to establish the relative validity of 24-h dietary recalls of daily food consumption in healthy African children living in Mbale and Soroti, eastern Uganda compared to simultaneous weighed food records.

Methods: Quantitative assessment of daily food consumption by weighed food records followed by two independent assessments using triple pass 24-h dietary recall on the following day. In conjunction with household measures and standard food sizes, volumes of liquid, dry rice, or play dough were used to aid portion size estimation. Inter-assessor agreement, and agreement with weighed food records was conducted primarily by Bland-Altman analysis and secondly by intraclass correlation coefficients and quartile cross-classification.

Results: Nineteen healthy children aged 6 months to 12 years were included in the study. Bland-Altman analysis showed 24-h recall only marginally under-estimated energy (mean difference of 149 kJ or 2.8 %; limits of agreement -1618 to 1321 kJ), protein (2.9 g or 9.4 %; -12.6 to 6.7 g), and iron (0.43 mg or 8.3 %; -3.1 to 2.3 mg). Quartile cross-classification was correct in 79 % of cases for energy intake, and 89 % for both protein and iron. The intraclass correlation coefficient between the separate dietary recalls for energy was 0.801 (95 % CI, 0.429–0.933), indicating acceptable inter-observer agreement.

Conclusions: Dietary assessment using 24-h dietary recall with volumetric portion size estimation resulted in similar and acceptable estimates of dietary intake compared with weighed food records and thus is considered a valid method for daily dietary intake assessment of children in communities with similar diets. The method will be utilised in a sub-study of a large randomised controlled trial addressing treatment in severe childhood anaemia.

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* Correspondence: g.frost@imperial.ac.uk; k.maitland@imperial.ac.uk

†Equal contributors

¹Faculty of Medicine, Nutrition and Dietetic Research Group, Division of Diabetes, Endocrinology and Metabolism, Department of Investigative Medicine, Imperial College London, Hammersmith Campus, London W12 0NN, UK

⁵Kilifi Clinical Trials Facility, KEMRI-Wellcome Trust Research Programme, PO Box 230, Kilifi, Kenya

Full list of author information is available at the end of the article



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Trial registration: This study was approved by the Mbale Research Ethics committee (Reference: 2013–050). Transfusion and Treatment of severe Anaemia in African Children: a randomized controlled Trial (TRACT) registration: ISRCTN84086586.

Keywords: Dietary assessment, Validation, Children, Portion size estimation, Uganda, Undernutrition

Abbreviations: 24hDR, 24 h dietary recall; CDC, Centres for disease control and prevention; DR, Dietary recall; FAO, Food and agricultural organisation of the united nations; ICC, Intraclass correlation coefficient; IQR, Interquartile range; LOA, Level of agreement; SD, Standard deviation; TRACT, Transfusion and treatment of severe anaemia in African children: a randomised controlled trial; UDHS, Ugandan demographic and health survey; UFT, Ugandan food tables; UNU, United nations university; WAZ, Weight-for-age z-score; WFR, Weighed food record; WHO, World health organisation

Background

Undernutrition, estimated to affect 100,000,000 children in the developing world, is implicated in approximately 45 % of childhood mortality globally [1, 2] and its reduction has been one of the United Nations Millennium Development Goals since 2000 s [3]. Aside from affecting mortality, poor nutrition in the first 1000 days of life is also associated with impaired cognitive ability, and reduced school and work performance [4]. Nutritional intake is fundamentally important to the health of the child and there is an intimate relationship between nutritional intake, nutritional status and infection. In order to develop and assess nutritional strategies and policies aimed at reducing childhood undernutrition, evaluation and validation of reliable methods of quantifying an individual's macro- and micronutrient intakes are therefore of critical importance.

Several studies in African countries have used single methods for assessing diet including household consumption surveys [5], weighed food records [6], food frequency questionnaires [7, 8], and 24-h dietary recall (24hDR) [5–7, 9] with variable success. Common methodologies, such as food frequency questionnaires and retrospective information on dietary history, are largely qualitative and considered as poor barometers of daily intake due to their imprecision [10–12]. Quantitative methods, measuring individual foods consumed (weighed food records, WFR) are the most precise methods for providing quantitative dietary data [13]. These are, however, time-consuming to conduct that results often in a small sample size, as they have been found to be burdensome and disruptive to the respondents. Interactive dietary recall is a potential substitute for a weighed food record. This has been investigated in Ghanaian children [14] and in Malawian children [15] in studies using a single 24hDR the day following independent weighed food assessment. This method of dietary recall could only be considered partially validated in the study groups due to some biases and imprecision. The Ghanaian study reported that averaged 24hDR

assessments tended to underestimate energy and nutrient intake compared with WFR, while the Malawian study reported the opposite. The over- and under-estimation of energy and nutrients may be reduced by the modifying the triple pass method for 24hDR, which has been shown to maximise recall accuracy for quantitation [16–18] by including volumetric portion size estimation, but this has yet to be evaluated in African children.

The current pilot study sought to establish the relative validity of an interactive 24hDR method with volumetric portion size estimation, compared to concurrent WFR in children in rural Uganda. The tool is intended for future use to assess the impact of daily dietary intake on outcome for a controlled trial of children hospitalised with severe anaemia (Transfusion and Treatment of severe anaemia in African children: a randomised controlled Trial (TRACT), ISRCTN84086586) [19].

Methods

Aim

The study's aims were first, to establish the relative validity of a 24hDR method compared to a weighed food record in estimating intakes of macro- and selected micro-nutrients in children in rural Uganda. Second, to ensure the recall method is feasible and culturally acceptable in this population.

Design

Dietary data from a weighed food record carried out by an independent researcher in the home of the subject was compared to estimated intakes from 24hDR assessments carried out by two other independent researchers the following day, to assess the relative validity of 24hDR. These researchers (clinicians and nurses) were not aware of the outcome of either the weighed food record or the other dietary recall. We opportunistically recruited 24 well children aged 6 months to 12 years attending Mbale and Soroti Regional Referral Hospitals over a two-week period in May 2014. We excluded infants who were entirely breastfed and children currently

unwell. Prospective consent was sought from parents or guardians.

Pre-study training involved role-play simulations of data collection including recall and weighed food interviews with non-study child–parent pairs attending hospital to consolidate clinician and nurse training.

Portion size estimation

In developing the 24hDR method for this population, issues specific to East African diets emerged such as the estimation of portion sizes for semi-solid foods since much of the diet is a semi-solid consistency (such as a maize flour-based paste known as ‘posho’ or ‘ugali’) and eaten by hand, often from one communal family bowl [20]. Thus, it was problematic to estimate by volume using standard household measures (bowfuls, spoonfuls). We developed a number of novel approaches to estimate portion sizes (see Table 1). We considered an alternative method of estimating portions of semi-solid foods by utilising play dough and volume displacement, previously proposed [16] but not yet validated. Estimated volumes or number of items eaten were then converted into grams. For this a database of local foodstuffs was generated with weight per 100mls or weight of a whole food item. Local reference sizes were used where appropriate (for example small/medium/large mango) or for certain foods including cassava chips or sugar cane three using representative lengths to which they were closest. Consensus approaches were agreed for other items, for example loaves of bread were classified by price, since these are consistently sized in this community.

Dietary data collection

Dietary data collection occurred in three stages: weighed food record (WFR) and two dietary recalls (DR) each carried out by a separate member of the research team following published protocols [14, 15]. Each researcher completed only one stage with each child and guardian in the home of the child and were blinded to details recorded by other observers. The details of each stage are summarised in Table 1.

For all measures the specific time frame was from the time the child awoke in the morning to the time they

slept at night. Any food taken after this time was not included in either WFR or DR since it was not realistic to expect researchers to remain in participants’ houses overnight.

The triple pass 24-h recall, shown to maximise recall accuracy for quantitation [18], used the following algorithm. The first pass encourages the respondent (guardian/parent) to freely report all food and drink intake for the prior day uninterrupted; in the second pass the interviewer probes for greater details on the exact time, type and quantity of food or drink taken; in the third and final pass the interviewer reviews all food reported in order, prompting for omissions and clarifying ambiguities. Completion of both DR used the same methodology and the same guardian and child to provide information about inter-assessor reliability and reproducibility. Interviews and assessments were carried out in English or local languages to ensure accuracy.

Calculation of estimated requirements

Total daily energy and protein requirements were estimated using the methods recommended by the relevant World Health Organization (WHO), Food and Agriculture Organisation of the United Nations (FAO), United Nations University (UNU) or joint publications [21, 22]. Iron requirements were based on the age and gender specific recommended daily allowances presented by Food and Nutrition Board of the US Institute of Medicine [23].

Data entry and analysis

Data from WFR, DR1 and DR2 were entered into Dietplan 6 (Forestfield Software Limited), and energy, macro- and micronutrient intakes were automatically computed for most foods using McCance and Widdowstone’s ‘The Composition of Foods (Food Standards Agency)’ [24]. These were supplemented, when recipes or foods were not available, by the Ugandan Food Tables (UFT) [25] which are derived from the United States Department of Agriculture National Nutrient Database for Standard Reference. For food items, such as milk, meat and flour, where composition may vary geographically, both UFT and The Composition of Foods values were

Table 1 Methodology of dietary data collection and portion size estimation

Stage	Methodology	Person conducting	Portion Size Estimation
1	Weighed food record	First researcher	Weighing
2	24-h dietary recall	Second researcher Third Researcher	<ul style="list-style-type: none"> Volume of play dough^a Household measures^b Standardised food item size^c

^afor foods eaten by hand

^bcups, bowls, table- and teaspoons of water or dry uncooked rice

^cfor example 1 egg, half of 1 medium onion

compared, and generally the lower of the two values used. Some foods such as oil, and maize and wheat flours are fortified in Uganda with vitamin A, and iron respectively, however this does not appear to be consistent [26]. Since the current study is concerned with method validation only and as such, unfortified values have been used.

We could find no data of direct nutrient analysis of food in Uganda or East Africa therefore some uncertainty remains regarding the accuracy of food composition data in this setting. It is recognised that neither US based UFT values [25], nor the UK Composition of Foods [24] may reflect actual nutrient composition of Ugandan foods.

Statistical analysis

Weight-for-age z-scores (WAZ) were calculated with WHO Anthro using the WHO reference population [27] and compared to the Uganda Demographic and Health Survey (UDHS), which use the median of the National Centre for Health Statistics [28], Centres for Disease Control and Prevention (CDC) [29], and WHO reference populations [27]. All other statistical analysis was completed using IBM SPSS Statistics for Windows v22 (IBM). Prior to statistical tests, Kolmogorov-Smirnov statistic and Q-Q plots were used to assess data distribution. Only estimated energy requirements were non-normally distributed, therefore Wilcoxon signed-rank test was used when comparing estimated energy requirements and estimated intakes and variability was assessed using inter-quartile range (IQR, 25-75th centiles). Bland-Altman analysis was conducted for a range of macro- and micronutrients, to compare each individual assessment of 24hDR (DR1 and DR2) and then to compare these with WFR [30]. Mean difference and standard deviation of the difference between each DR, and DR and reference method were generated for energy, protein and iron consumption, and reported as mean difference and limits of agreement (i.e. ± 1.96 *standard deviation of mean difference).

The relationship between estimated intakes of energy, protein and iron were explored using intraclass correlation coefficients (ICC) and by quartile cross-classification. ICCs compared absolute agreement of average measures, using a two-way random model. Classification was defined as correct (same quartile), adjacent (± 1 quartile), or grossly misclassified by 2 or more quartiles. Differences between estimated requirements and estimated intakes by WFR, DR1, and DR2 were analysed using paired t-tests. Initial analysis was completed between WFR, DR1 and DR2 in pairs. Statistical significance was defined as $p < 0.05$.

Results

Demographics and anthropometry

Of 24 children recruited (14 in Mbale and 10 in Soroti), two did not complete the dietary assessment and three were excluded due to recurring or new illness. Of the remaining 19, 12 were female (61.9 %), mean age (\pm SD) was 3.4 years (± 2.6), and mean weight (\pm SD) was 14.0 kg (± 5.6). The mean WAZ score (\pm SD) was -0.19 (± 1.75). Three children were moderately or severely underweight defined as WAZ scores ≤ -2.0 . The majority ($n = 13$) had WAZ scores between -2 and 2 . Three children had high WAZ scores ≥ 2 . Four children were partially breastfed therefore were not included in comparisons with estimated requirements as determining a reliable 'portion size' was impossible. A post hoc power analysis showed that with 19 participants, this study has 80 % power to detect a difference of 16.7 % or 1097 kJ in energy intake at a significance level of 0.05, using the mean energy consumption of 6563 kJ and SD of 1706 kJ.

Inter-assessor variation

Figure 1 shows Bland-Altman analysis with mean difference, absolute limits of agreement and percentage (%) between DR1 and DR2 for energy 289.4 kJ, -2111.9 to 2690.6 kJ (-40.0 to 51.0 %); protein 1.3 g, -9.93 to 12.6 g (-32.8 to 41.7 %); and iron 0.2 mg, -2.5 to 2.8 mg (-48.3 to 55.1 %). The intraclass correlation coefficient for the two 24-h dietary recalls for energy was 0.802 (95 % CI, 0.429–0.933), for protein 0.925 (95 % CI, 0.779–0.975), and for iron 0.868 (95 % CI, 0.618–0.955) suggesting high inter-assessor reliability. Since the estimates by DR1 and DR2 for each of these parameters were comparable as assessed by cross-validation and Bland-Altman analyses, we therefore used the global mean of these estimates to compare with WFR data for conciseness.

Comparability of WFR and 24-h dietary recall methods

Figure 2 shows the mean difference for energy was -149.1 kJ with limits of agreement of -1619 to 1321 kJ (-30.4 to 24.8 %), mean difference for protein was -2.9 g with limits of agreement of -12.6 to 6.7 g (-40.4 to 21.6 %) and mean difference for iron was -0.4 mg with limits of agreement of -3.1 to 2.3 mg (-60.2 to 43.7 %). Mean differences with associated upper and lower limits of agreement comparing WFR and combined DR1 and DR2 are displayed for all nutrients included in Additional file 1: Table S1 and the associated dataset (Excel format) is provided to enable validation of results and statistical interpretation (Additional file 2: Table S2).

Intraclass correlation coefficients for WFR and combined 24-h dietary recall estimates of nutritional intake were 0.979 (95 % CI, 0.899–0.984) for energy, 0.972 (95 % CI, 0.903–0.990) for protein, and 0.936 (95 % CI, 0.837–0.975) for iron, summarized in Table 2.

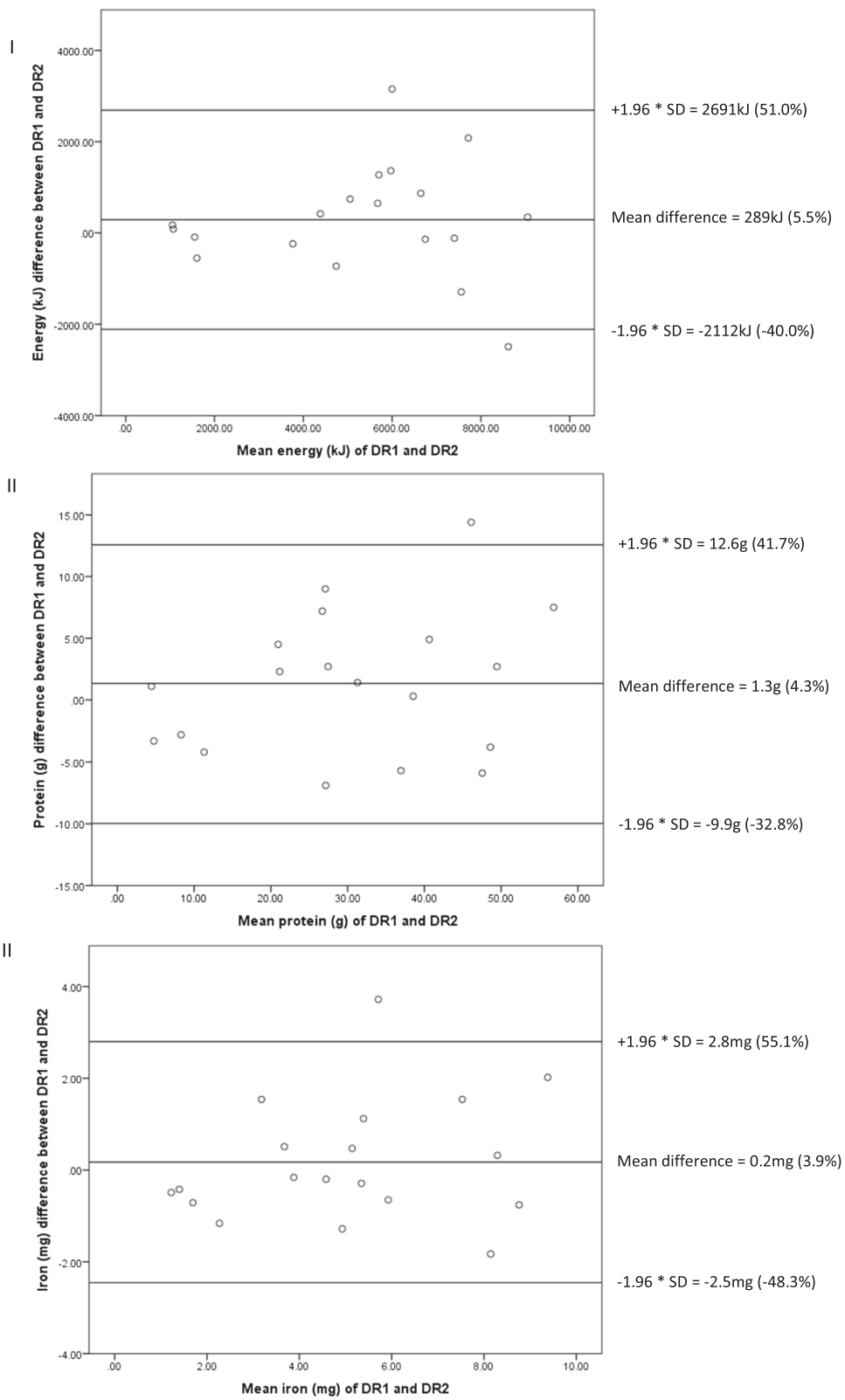


Fig. 1 Bland-Altman plots of first and secondary dietary recalls: (I) energy, (II) protein, (III) iron. DR, dietary recall; SD, standard deviation

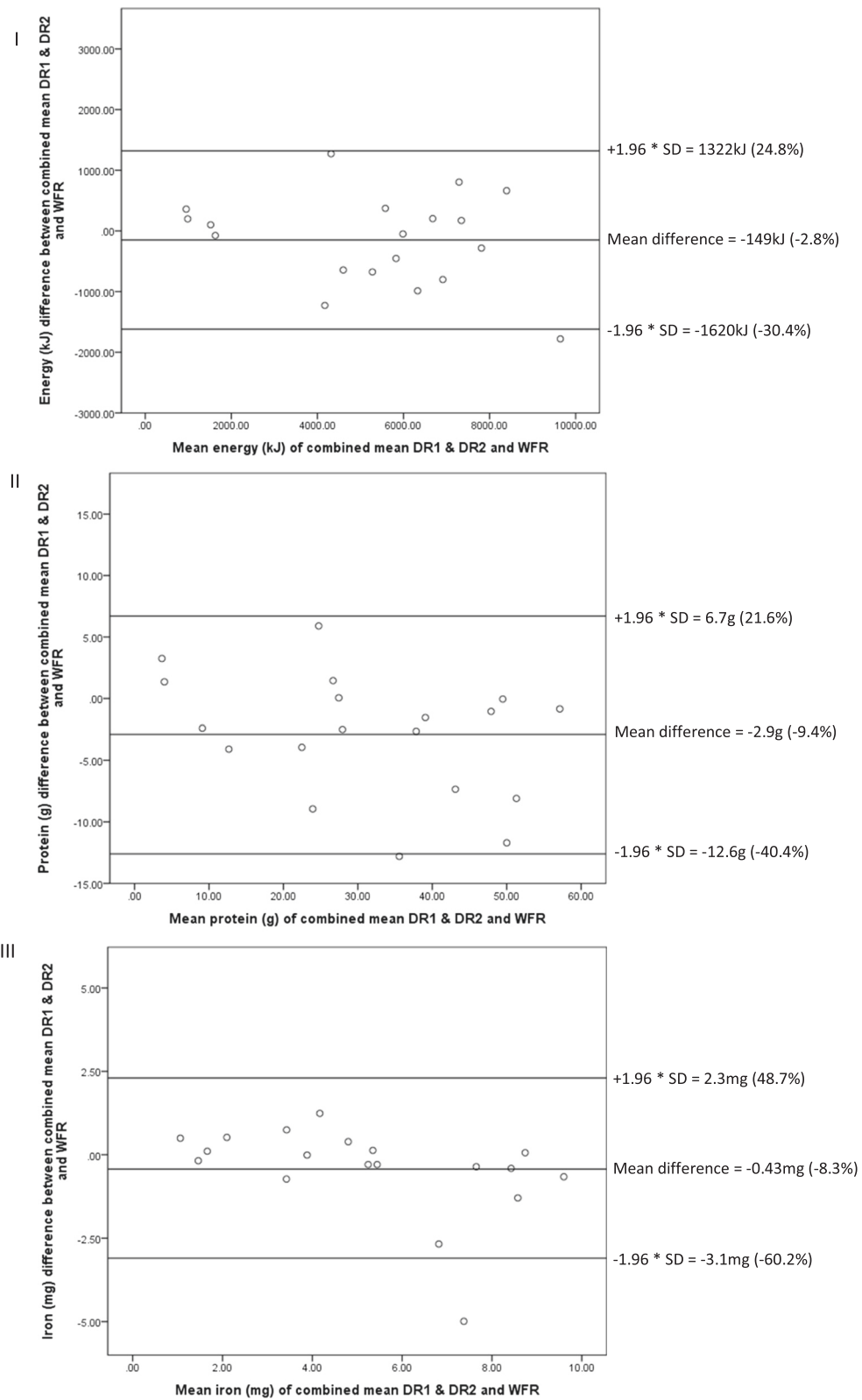


Fig. 2 Bland-Altman plots of combined dietary recalls and weighed food records: (I) energy, (II) protein, and (III) iron. DR, dietary recall; SD, standard deviation; WFR, weighed food record

Table 2 Intraclass and bivariate correlation coefficients comparing estimated intakes by weighed food records and 24-h dietary recalls^a

Nutrient	Weighed Food Record		Combined Dietary Recalls			
	Mean	SD	Mean	SD	ICC (95 % CI)	<i>r</i> (<i>p</i> -value)
Energy (kJ)	6563	1706	6335	1537	0.98 (0.90–0.98)	0.96*
Protein (g)	40.0*	12.9	36.4*	11.4	0.97 (0.90–0.99)	0.985*
Iron (mg)	6.6	2.6	6.0	2.0	0.94 (0.84–0.98)	0.91*

SD standard deviation, ICC intraclass correlation coefficient, CI confidence interval, *r* correlation coefficient

^aICCs compared absolute agreement of average measures, using a two-way random model

**p* < 0.001

[†]*p* = 0.02

Classification into quartiles of intake and assessment of this agreement by Cohen's Kappa (*k*) statistic is displayed in Table 3. This showed that in the majority of cases WFR and dietary recalls agreed on classification, in 79 % of cases for energy and 89 % for protein and iron. The remainder were classified adjacently, with none being grossly misclassified. Agreement of classification in quartiles was substantial (*k* 0.61–0.80) or almost perfect (*k* 0.81–1.00) for all nutrients tested [31].

Estimated requirements and intake

The median estimated requirement for energy was 4602 kJ/day (IQR 25-75th centile = 3836–5208 kJ), and intake was estimated at 6544 kJ (IQR 25-75th centile = 5330–7448 kJ) by the WFR, showing a significant surplus of 1942 kJ (*p* = 0.001). Mean estimated protein requirement was 14.2 g (\pm 5.1), while WFR-estimated intake was 40.0 g (\pm 12.9)/day, 26.8 g in excess of requirements (*p* < 0.001). Mean iron requirement was 8.3 mg (\pm 1.6)/day, while WFR-estimated intake was 6.6 mg (\pm 2.6) (*p* = 0.004).

Discussion

The 24-h multi-pass recall method described compared favourably to a weighed food records, with regards to energy, protein and iron intakes. Bland-Altman analyses showed overall agreement for energy, protein and iron intakes between two separate interviewers, suggesting high inter-assessor reproducibility, which is further reinforced by high intraclass correlation coefficients. Classification of energy intake into quartiles showed substantial agreement for energy and almost perfect agreement for protein and iron intakes.

High intraclass correlation coefficients, and low mean differences for energy, protein and iron with weighed food records suggest the triple-pass 24 h recalls are comparable for assessing daily intakes. The method suggested by Gibson & Ferguson [16] was adapted to the local setting and validated in this pilot. Using play-dough and volume displacement generally worked well, and was intuitive for both researchers and subjects. The estimated nutrient intakes must be interpreted with caution owing to wide limits of agreement; in the case of iron particularly, only gross differences in intake can be inferred. For iron one extreme outlier was noted with 4.99 mg lower estimated intake by recalls compared to WFR. The cause of this large discrepancy was found to be due to inaccuracy in the portion size estimation of a ready-to-use nutrient-dense nutritional supplement, which contributed over 6 mg of iron alone to intake, the only instance in this pilot where this supplement was noted. Studies involving severely malnourished children are likely to encounter ready-to-use feed or calorie enhanced milks, and particular care in estimating the portion size is advised due to nutrient density, for the future study (TRACT), where it is intended to be used the numbers of children with severe malnutrition are anticipated to be few.

Although inter-assessor variability was assessed, this study did not address intra-observer repeatability, which must be borne in mind when the method is used. One limitation of only assessing the preceding 24-h period is that a habitual identical intake cannot be assumed. Both dietary recalls were undertaken on the same day, which may have introduced bias in parental recall, for example memory of information provided during the first recall may have been reinforced for the second recall, whether accurate or not, thus artificially reducing the inter-

Table 3 Cross-classification of children to quartiles according to intake estimates

Nutrient	Classified correctly (%)	Classified adjacently (%)	Grossly misclassified (by \geq 2 quartiles) (%)	Cohen's Kappa <i>k</i> (<i>p</i> -value)
Energy	15 (79)	4 (21)	0	0.719 (<0.001)
Protein	17 (89)	2 (11)	0	0.859 (<0.001)
Iron	17 (89)	2 (11)	0	0.859 (<0.001)

assessor variability. Noteworthy, is that whilst the results presented using the mean of two 24-h dietary recalls will technically reduce the observed variability, the inter-assessor variability was low, therefore conducting a single recall should not have a substantive effect.

Although method validation was the main aim of this study, it is prudent to comment on the intakes observed. Energy intake was higher than in previous reported studies, at 6563 kJ compared to 5606 kJ [32]. Intakes of energy and protein were also in excess of requirements by a factor of 1.39 for energy and 2.97 for protein. Similar high protein intakes of 41.0 g/day have been reported in children in other regions of Uganda [32]. Another potential reason for the difference is variation over the week in energy consumption, which is not reflected on a single day recall assessment. Two reasons are suggested for this observation. Firstly, although it was explained to participants that the priority was to observe the children's intake unbiased, the effect of the researchers' presence is difficult to estimate. Secondly, while these were healthy children, all had had recent contact with healthcare services, and as such may be experiencing catch-up growth and provided with additional food for recuperation. Indeed, WAZ scores observed showed that 16 % were severely or moderately underweight, and is similar to the most recent UDHS 2011 census for the Eastern Uganda region [33], where prevalence was 15.4 %. In contrast to the UDHS results which showed only 0.1 % had WAZ scores >2, compared to 15.8 % ($n = 3$) of subjects in this pilot.

Conclusions

The methods we have described and validated in children in Uganda appear consistent and correlate satisfactorily with quantitative assessment of dietary intake. A study comparing a single pass 24DR to assess dietary intake with a subsequent 7-day weighed food record in Sri Lankan adults found that 24DR tended to underestimate mean energy levels and macronutrients however the difference in the energy percentages were not statistically different [34]. Underestimation using single pass 24DR has been previously reported and is improved by triple-pass 24DR [17, 18]. We consider that the method we have assessed to be valid for an on going a factorial treatment trial of African children presenting to hospital with severe anaemia (TRACT trial) [19]. The method will be used to assess nutritional intake as a surrogate marker of general wellbeing and the association of acute nutritional intake with severity of anaemia, impaired gut barrier function and susceptibility to infection. The TRACT study combines sequential dietary intake assessment using the multi-pass method at each follow-up visit to estimate macro and micro nutrient intake and will be subsequently linked to biomarkers of gut barrier function, gut microbiome, immunity and hormonal appetite control.

Additional files

Additional file 1: Table S1. Individual nutrient and nutrient mean differences, standard deviation, and upper and lower limits of agreement (LOA) comparing WFR and combined DR1 and DR1; Supplementary table summarising Bland-Altman analysis results for all nutrients. Results presented in absolute and percentage values. (DOCX 16 kb)

Additional file 2: Dataset for validation of triple pass 24-h dietary recall in Ugandan children by simultaneous weighed food assessment; Associated dataset in Excel format to enable validation of results and statistical interpretation. Gender has been removed to limit indirect identifiable information, but can be made available on request. (XLSX 26 kb)

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Availability of data and materials

The dataset supporting the conclusions of this article is included within the article and its additional files.

Authors' contributions

HN, GF, KM, POO, and CE designed and planned the study. HN, TS, JN, CM, MN, MA, DA developed the method and carried out data collection. KJW analysed and interpreted the data. HN and KJW drafted the manuscript, and all authors commented on the manuscript. All authors read and approved the final manuscript.

Competing interests

The authors declare that they have no competing interests.

Consent for publication of individual data

Not applicable.

Ethics approval and consent to participate

This study was approved by the Mbale Research Ethics committee (Reference: 2013–050). Verbal consent was sought from parents or guardians for researchers to measure the child's weight, observe and weigh the child's food intake over 1 day, and for two separate 24 h dietary recalls to be undertaken the following day. The child's assent and ongoing verbal parental/guardian consent was also sought prior to each activity.

Author details

¹Faculty of Medicine, Nutrition and Dietetic Research Group, Division of Diabetes, Endocrinology and Metabolism, Department of Investigative Medicine, Imperial College London, Hammersmith Campus, London W12 0NN, UK. ²Busitema University Faculty of Health Sciences (BUFHS), Mbale Clinical Research Centre, Mbale, Uganda. ³Mbale Regional Referral Hospital Clinical Research Unit (MCRU), Mbale, Uganda. ⁴Department of Paediatrics, Soroti Regional Referral Hospital, Soroti, Uganda. ⁵Kilifi Clinical Trials Facility, KEMRI-Wellcome Trust Research Programme, PO Box 230, Kilifi, Kenya. ⁶Wellcome Trust Centre for Clinical Tropical Medicine, and Department of Paediatrics, Faculty of Medicine, Imperial College, London W2 1PG, UK.

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