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Complete convergence and complete moment convergence for randomly weighted sums of martingale difference sequence

Huanhuan Ma^{1*} ond Yan Sun²

*Correspondence: huanhuanma16@126.com ¹College of Mathematics and Information Science, Henan Normal University, Xinxiang, China Full list of author information is available at the end of the article

Abstract

In this paper, we extend some known results about complete convergence and establish the complete convergence and complete moment convergence for randomly weighted sums of martingale difference sequence. Our results can generalize some conclusions related to Hsu–Robbins–Erdös strong laws and Baum–Katz type theorems for martingales.

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1 Introduction

The complete convergence plays a key role in the development of probability theory, especially in establishing the rate of convergence. Hsu and Robbins [1] introduced the concept of complete convergence as follows. A sequence $\{X_n, n \ge 1\}$ is said to converge completely to *C* if

$$\sum_{n=1}^{\infty} P(|X_n - C| \ge \epsilon) < \infty \quad \text{for all } \epsilon > 0,$$
(1.1)

where *C* is a constant. By the Borel–Cantelli lemma, it follows that $X_n \to C$ almost surely as $n \to \infty$. If $\{X_n, n \ge 1\}$ is independent and identically distributed (i.i.d.) random variables, the converse is true.

Suppose that $\{X_n, n \ge 1\}$ is a random variable sequence defined on the fixed probability space (Ω, \mathcal{F}, P) . Denote $S_n = \sum_{i=1}^n X_i$, $S_0 = 0$, $\log x = \log(\max\{e, x\})$, $x^+ = xI(x \ge 0)$, and $\mathcal{F}_0 = \{\Omega, \emptyset\}$. Let $\{\mathcal{F}_n, n \ge 1\}$ be an increasing sequence of σ fields with $\mathcal{F}_n \subset \mathcal{F}$ for each $n \ge 1$. If X_n is \mathcal{F}_n measurable for each $n \ge 1$, then σ fields $\{\mathcal{F}_n, n \ge 1\}$ are thought to be adapted to the random variable sequence $\{X_n, n \ge 1\}$ and $\{X_n, \mathcal{F}_n, n \ge 1\}$ is thought to be an adapted stochastic sequence. The following theorem is a generalization of some known results.

Theorem 1.1 (Hsu–Robbins–Erdös strong law [1, 2]) Let $\{X_n, n \ge 1\}$ be a sequence of independent and identically distributed random variables. Assume that $EX_n = 0$ and set





$$S_n = \sum_{i=1}^n X_i$$
, $n \ge 1$. Then $EX_n^2 < \infty$ is equivalent to the condition that

$$\sum_{n=1}^{\infty} P(|S_n| \ge \epsilon n) < \infty \quad \text{for all } \epsilon > 0.$$
(1.2)

In probability theory, Hsu–Robbins–Erdös strong law as a basic theorem has been extended in several directions by some authors. The following theorem is given by Baum and Katz [3] to establish a rate of convergence.

Theorem 1.2 (Baum and Katz strong law) Let $\alpha > 1/2$, $\alpha p > 1$, and let $\{X_n, n \ge 1\}$ be a sequence of independent and identically distributed random variables. Assume that $EX_n = 0$ if $\alpha \le 1$, and set $S_n = \sum_{i=1}^n X_i$, $n \ge 1$. Then $E|X_n|^p < \infty$ is equivalent to the condition that

$$\sum_{n=1}^{\infty} n^{\alpha p-2} P(|S_n| \ge \epsilon n^{\alpha}) < \infty \quad \text{for all } \epsilon > 0$$
(1.3)

and also equivalent to the condition that

$$\sum_{n=1}^{\infty} n^{\alpha p-2} P\left(\max_{1 \le k \le n} |S_k| \ge \epsilon n^{\alpha}\right) < \infty \quad \text{for all } \epsilon > 0.$$

$$(1.4)$$

Motivated by the above results for i.i.d. random variables, many authors have studied them for dependent cases. The case for weighted sums of extended negatively dependent (END) random variable sequence was investigated by Shen et al. [4]. Miao et al. [5] improved some known results and studied the Baum–Katz type convergence rate in the Marcinkiewicz–Zygmund strong law for martingales. Chen et al. [6] also gave some extended results for the sequence of martingale difference.

The aims of the present paper are to extend the results on complete convergence for the sequence of martingale difference. The following definitions will be used frequently in this paper.

Definition 1.1 If $\{X_n, \mathcal{F}_n, n \ge 1\}$ is an adapted stochastic sequence with

$$E(X_n|\mathcal{F}_{n-1})=0 \quad \text{a.s.}$$

and $E|X_n| < \infty$ for each $n \ge 1$, then the sequence $\{X_n, \mathcal{F}_n, n \ge 1\}$ is called a martingale difference sequence.

Definition 1.2 A real-valued function l(x), positive and measurable on $(0, \infty)$, is said to be slowly varying if

$$\lim_{x\to\infty}\frac{l(\lambda x)}{l(x)}=1$$

for each $\lambda > 0$.

Definition 1.3 A sequence $\{X_n, n \ge 1\}$ of random variables is said to be stochastically dominated by a random variable *X* if there exists a positive constant *C*, such that

$$P(|X_n| > x) \le CP(|X| > x)$$

for all $x \ge 0$ and $n \ge 1$.

Now let us recall some known results for complete convergence of martingales.

Theorem 1.3 ([7,8]) Let $\{X_n, \mathcal{F}_n, n \ge 1\}$ be an L^p -bounded martingale difference sequence. If $0 < 1/\alpha < 2 < p$ or 1 , then

$$\sum_{n=1}^{\infty} n^{\alpha p-2} P(|S_n| \ge \epsilon n^{\alpha}) < \infty \quad for \ all \ \epsilon > 0.$$
(1.5)

If $p = \alpha = 1$, the martingale difference sequence satisfies

$$\sup_{n\geq 1} E|X_n|\log|X_n|<\infty,$$

then (1.5) holds.

Wang and Hu [9] further studied the Baum–Katz type theorem for the maximal partial sum of martingale difference sequence.

Theorem 1.4 [9] Let $\{X_n, \mathcal{F}_n, n \ge 1\}$ be a sequence of martingale difference, which is stochastically dominated by a random variable X. Let l(x) > 0 be a slowly varying function as $x \to \infty$. Let $\alpha > 1/2$, $p \ge 1$ and $\alpha p \ge 1$. When $p \ge 2$, we further assume that

$$E\left[\sup_{i\geq 1} E\left(X_i^2 | \mathcal{F}_{i-1}\right)\right]^{q/2} < \infty$$

for some $q > \frac{2(\alpha p-1)}{2\alpha-1}$. *If*

$$E|X|^{p}l(|X|^{1/\alpha}) < \infty, \tag{1.6}$$

then for any $\epsilon > 0$ *,*

$$\sum_{n=1}^{\infty} n^{\alpha p-2} l(n) P\left(\max_{1 \le j \le n} |S_j| \ge \epsilon n^{\alpha}\right) < \infty.$$
(1.7)

Yang et al. [10] generalized the results of Stoica [7, 8] and Wang et al. [11] for the nonweighted sums of martingale difference sequence to the case of randomly weighted sums.

Theorem 1.5 [10] Let $\{X_n, \mathcal{F}_n, n \ge 1\}$ be a martingale difference sequence stochastically dominated by a nonnegative random variable X with $EX^p < \infty$. Assume that $\{A_n, n \ge 1\}$

is a random sequence, and it is independent of $\{X_n, n \ge 1\}$. Denote $\mathcal{G}_0 = \{\emptyset, \Omega\}$ and $\mathcal{G}_n = \sigma(X_1, \ldots, X_n), n \ge 1$. Let $\alpha > 1/2, 1 , and <math>1 \le \alpha < 2$. If

$$\sum_{i=1}^{n} EA_i^2 = O(n), \tag{1.8}$$

then

$$\sum_{n=1}^{\infty} n^{\alpha p-2-\alpha} E\left(\max_{1\le k\le n} \left|\sum_{i=1}^{n} A_i X_i\right| - \epsilon n^{\alpha}\right)^+ < \infty \quad for \ any \ \epsilon > 0.$$

$$(1.9)$$

If $\alpha > 1/2$, $p \ge 2$, and for some $q > \frac{2(\alpha p - 1)}{2\alpha - 1}$, we assume that

$$E\left[\sup_{i\geq 1} E\left(X_i^2|\mathcal{G}_{i-1}\right)\right]^{q/2} < \infty.$$

Let

$$\sum_{i=1}^{n} E|A_i|^q = O(n), \tag{1.10}$$

then (1.9) holds.

If $\alpha > 0$ and p = 1, the martingale difference sequence is stochastically dominated by a nonnegative random variable X with $E[X \log(1 + X)] < \infty$, and (1.8) holds, then

$$\sum_{n=1}^{\infty} n^{-2} E\left(\max_{1 \le k \le n} \left| \sum_{i=1}^{n} A_i X_i \right| - \epsilon n^{\alpha} \right)^+ < \infty \quad \text{for any } \epsilon > 0.$$

$$(1.11)$$

We shall study the complete convergence and complete moment convergence for randomly weighted sums of martingale difference sequence. The paper is organized as follows. The next section is devoted to the descriptions of our main results, and their proofs will be given in Sect. 3. Throughout the paper, we use the constant C to denote a universal real number that is not necessarily the same in each appearance.

2 Main results

Theorem 2.1 Let $\{X_n, \mathcal{F}_n, n \ge 1\}$ be a sequence of martingale difference, which is stochastically dominated by a random variable X. Let l(x) > 0 be a slowly varying function as $x \to \infty$. Suppose that $\{b_n, n \ge 1\}$ and $\{c_n, n \ge 1\}$ are sequences of positive constants such that, for p > 1, $\alpha > 0$, $\alpha p \ge 1$, and some $q \ge \max\{2, p\}$,

$$\sum_{n=1}^{m} \frac{nb_n l(n)}{c_n} = O(c_m^{p-1} l(c_m^{1/\alpha})), \qquad \sum_{n=m}^{\infty} \frac{nb_n l(n)}{c_n^q} = O(c_m^{p-q} l(c_m^{1/\alpha})), \tag{2.1}$$

and

$$\sum_{n=1}^{\infty} \frac{n^{\frac{q}{2}} b_n l(n)}{c_n^q} < \infty, \qquad E \Big[\sup_{i \ge 1} E \big(X_i^2 | \mathcal{G}_{i-1} \big) \Big]^{\frac{q}{2}} < \infty,$$
(2.2)

where $c_n \to \infty$ as $n \to \infty$. Assume that $\{A_n, n \ge 1\}$ is a random sequence independent of $\{X_n, n \ge 1\}$ such that

$$\sum_{i=1}^{n} E|A_i|^q = O(n).$$
(2.3)

If

$$E|X|^{p}l(|X|^{1/\alpha}) < \infty, \tag{2.4}$$

then for any $\epsilon > 0$,

$$\sum_{n=1}^{\infty} b_n l(n) P\left(\max_{1 \le j \le n} \left| \sum_{i=1}^{j} A_i X_i \right| \ge \epsilon c_n \right) < \infty.$$
(2.5)

Corollary 2.1 Under the conditions of Theorem 2.1, we take $b_n = n^{\alpha p-2}$, $c_n = n^{\alpha}$ for $\alpha > 1/2$, p > 1, and $\alpha p \ge 1$. If

$$\begin{cases} q > \max\{p, \frac{2(\alpha p - 1)}{2\alpha - 1}\}, & p \ge 2; \\ q = 2, & 1$$

then for any $\epsilon > 0$ *,*

$$\sum_{n=1}^{\infty} n^{\alpha p-2} l(n) P\left(\max_{1 \le j \le n} \left| \sum_{i=1}^{j} A_i X_i \right| \ge \epsilon n^{\alpha} \right) < \infty.$$
(2.6)

Remark 2.1 Obviously, (2.6) can be checked by Theorem 2.1 and Lemma 3.4. Under the conditions of Corollary 2.1, if we take $A_i \equiv 1$, $i \ge 1$, then we have (1.7), i.e., the conclusion of Wang and Hu [9] holds for p > 1. On the other hand, if we take $l(x) \equiv 1$, then we can get Remark 2.1 in Yang et al. [10]. So our results can imply these known results.

Example 2.1 Under the conditions of Theorem 2.1, we take $b_n = n^{r-2}$, $l(n) = \log n$, and $c_n = n^{r/p}$ for p > 1 and r > p. If

$$\begin{cases} q > \max\{p, \frac{2p(1-r)}{p-2r}\}, & p \ge 2; \\ q = 2, & 1$$

then for any $\epsilon > 0$,

$$\sum_{n=1}^{\infty} n^{r-2} \log n P\left(\max_{1 \le j \le n} \left| \sum_{i=1}^{j} A_i X_i \right| \ge \epsilon n^{r/p} \right) < \infty.$$

Example 2.2 Under the conditions of Theorem 2.1, we take $b_n = \frac{\log n}{n}$, $l(n) = \log n$, and $c_n = (n \log n)^{\frac{1}{p}}$ for 1 . If

$$\begin{cases} q > 6, \quad p = 2; \\ q = 2, \quad 1$$

then for any $\epsilon > 0$,

$$\sum_{n=1}^{\infty} \frac{(\log n)^2}{n} P\left(\max_{1 \le j \le n} \left| \sum_{i=1}^{j} A_i X_i \right| \ge \epsilon (n \log n)^{\frac{1}{p}} \right) < \infty.$$

Theorem 2.2 Let $\{X_n, \mathcal{F}_n, n \ge 1\}$ be a sequence of martingale difference, which is stochastically dominated by a random variable X. Let l(x) > 0 be a slowly varying function as $x \to \infty$. Suppose that $\{b_n, n \ge 1\}$ and $\{c_n, n \ge 1\}$ are sequences of positive constants such that, for p > 1, $\alpha > 0$, $\alpha p \ge 1$, and some $q \ge \max\{2, p\}$,

$$\sum_{n=1}^{m} nb_n l(n) = O(c_m^{p-1}l(c_m^{1/\alpha})), \qquad \sum_{n=m}^{\infty} \frac{nb_n l(n)}{c_n^{q-1}} = O(c_m^{p-q}l(c_m^{1/\alpha})),$$
(2.7)

and

$$\sum_{n=1}^{\infty} \frac{n^{\frac{q}{2}} b_n l(n)}{c_n^{q-1}} < \infty, \qquad E \Big[\sup_{i \ge 1} E \big(X_i^2 | \mathcal{G}_{i-1} \big) \Big]^{\frac{q}{2}} < \infty,$$
(2.8)

where $c_n \to \infty$ as $n \to \infty$. Assume that $\{A_n, n \ge 1\}$ is a random sequence independent of $\{X_n, n \ge 1\}$ such that

$$\sum_{i=1}^{n} E|A_i|^q = O(n).$$
(2.9)

If

$$E|X|^{p}l(|X|^{1/\alpha}) < \infty, \tag{2.10}$$

then for any $\epsilon > 0$ *,*

$$\sum_{n=1}^{\infty} b_n l(n) E\left(\max_{1 \le j \le n} \left| \sum_{i=1}^{j} A_i X_i \right| - \epsilon c_n \right)^+ < \infty.$$
(2.11)

Theorem 2.3 Let $\{X_n, \mathcal{F}_n, n \ge 1\}$ be a sequence of martingale difference, which is stochastically dominated by a random variable X. Let l(x) > 0 be a slowly varying function as $x \to \infty$. Suppose that $\{b_n, n \ge 1\}$ and $\{c_n, n \ge 1\}$ are sequences of positive constants such that, for p = 1, $\alpha > 0$, $\alpha p \ge 1$, and some $q \ge 2$,

$$\sum_{n=1}^{m} nb_n l(n) = O((\log c_m) l(c_m^{1/\alpha})), \qquad \sum_{n=m}^{\infty} \frac{nb_n l(n)}{c_n} = O(c_m^{1-q} l(c_m^{1/\alpha})), \tag{2.12}$$

where $c_n \to \infty$ as $n \to \infty$. Assume that $\{A_n, n \ge 1\}$ is a random sequence independent of $\{X_n, n \ge 1\}$ and satisfying (2.9). If

$$E\left[|X|\left(\log|X|\right)l\left(|X|^{1/\alpha}\right)\right] < \infty, \tag{2.13}$$

then we have formula (2.11).

Corollary 2.2 Under the conditions of Theorem 2.2 for p > 1, we take $b_n = n^{\alpha p - 2 - \alpha}$, $c_n = n^{\alpha}$ for $\alpha > 1/2$ and $\alpha p \ge 1$. If

$$\begin{cases} q > \max\{p, \frac{2(\alpha p - 1)}{2\alpha - 1}\}, & p \ge 2; \\ q = 2, & 1$$

then for any $\epsilon > 0$,

$$\sum_{n=1}^{\infty} n^{\alpha p-2-\alpha} l(n) E\left(\max_{1 \le j \le n} \left| \sum_{i=1}^{j} A_i X_i \right| - \epsilon n^{\alpha} \right)^+ < \infty.$$
(2.14)

Under the conditions of Theorem 2.3 for p = 1, if we take $b_n = n^{-2}$, $c_n = n^{\alpha}$ for $\alpha > 0$. Then, for any $\epsilon > 0$, we have

$$\sum_{n=1}^{\infty} n^{-2} l(n) E\left(\max_{1 \le j \le n} \left| \sum_{i=1}^{j} A_i X_i \right| - \epsilon n^{\alpha} \right)^+ < \infty.$$
(2.15)

Remark 2.2 Obviously, (2.14) and (2.15) can be checked by Theorem 2.2, Theorem 2.3, and Lemma 3.4. Under the conditions of Corollary 2.2, if we take $l(x) \equiv 1$, then we have (1.9) and (1.11), i.e., the results of Yang et al. [10] can be generalized by our conclusions. On the other hand, if we take $A_i \equiv 1$, $i \ge 1$, then we can get Theorem 3.3 and Theorem 3.4 in Wang and Hu [9]. Hence, our conclusions can extend these known results.

Example 2.3 Under the conditions of Theorem 2.2, we take $b_n = n^{r-2-r/p}$, $l(n) = \log n$, and $c_n = n^{r/p}$ for p > 1 and r > p. If

$$\begin{cases} q > \max\{p, \frac{2p(1-r)}{p-2r}\}, & p \ge 2; \\ q = 2, & 1$$

then for any $\epsilon > 0$,

$$\sum_{n=1}^{\infty} n^{r-2-r/p} \log nE\left(\max_{1\leq j\leq n}\left|\sum_{i=1}^{j}A_{i}X_{i}\right|-\epsilon n^{r/p}\right)^{+}<\infty.$$

Example 2.4 Under the conditions of Theorem 2.2, we take $b_n = \frac{(\log n)^{1-1/p}}{n^{1+1/p}}$, $l(n) = (\log n)^{1-1/p}$ and $c_n = (n \log n)^{\frac{1}{p}}$ for 1 . If

$$\begin{cases} q > 5, \quad p = 2; \\ q = 2, \quad 1$$

then for any $\epsilon > 0$,

$$\sum_{n=1}^{\infty} \frac{(\log n)^{2-2/p}}{n^{1+1/p}} E\left(\max_{1 \le j \le n} \left| \sum_{i=1}^{j} A_i X_i \right| - \epsilon (n \log n)^{\frac{1}{p}} \right)^+ < \infty.$$

Remark 2.3 If the conditions of Theorem 2.2 or Theorem 2.3 hold, then for any $\epsilon > 0$, we can get

$$\sum_{n=1}^{\infty} b_n c_n l(n) P\left(\max_{1 \le j \le n} \left| \sum_{i=1}^{j} A_i X_i \right| \ge \epsilon c_n \right) < \infty.$$

In fact, it can be checked that for any $\epsilon > 0$,

$$\sum_{n=1}^{\infty} b_n l(n) E\left(\max_{1 \le j \le n} \left| \sum_{i=1}^{j} A_i X_i \right| - \epsilon c_n \right)^+$$

$$= \sum_{n=1}^{\infty} b_n l(n) \int_0^{\infty} P\left(\max_{1 \le j \le n} \left| \sum_{i=1}^{j} A_i X_i \right| - \epsilon c_n > t \right) dt$$

$$\geq \sum_{n=1}^{\infty} b_n l(n) \int_0^{\epsilon c_n} P\left(\max_{1 \le j \le n} \left| \sum_{i=1}^{j} A_i X_i \right| - \epsilon c_n > t \right) dt$$

$$\geq \epsilon \sum_{n=1}^{\infty} b_n c_n l(n) P\left(\max_{1 \le j \le n} \left| \sum_{i=1}^{j} A_i X_i \right| \ge 2\epsilon c_n \right).$$

Remark 2.4 If $A_n = a_n$, $n \ge 1$ is non-random (the case of constant weighted), then we can get the results of Theorems 2.1–2.3 for the non-random weighted sums of martingale difference sequence.

3 Proofs for the main results

Throughout this section, we use the constant C to denote a generic real number that is not necessarily the same in each appearance.

3.1 Several lemmas

To prove the main results of the paper, we need to recall the following lemmas.

Lemma 3.1 ([12]) If $\{X_i, \mathcal{F}_i, 1 \le i \le n\}$ is a sequence of martingale difference and q > 0, then there exists a constant *C* depending only on *p* such that

$$E\left(\max_{1\leq k\leq n}\left|\sum_{i=1}^{k} X_{i}\right|^{q}\right)\leq C\left\{E\left(\sum_{i=1}^{n} E\left(X_{i}^{2}|\mathcal{F}_{i-1}\right)\right)^{q/2}+E\left(\max_{1\leq i\leq n}|X_{i}|^{q}\right)\right\}.$$

Lemma 3.2 ([13–15]) *Let* { X_n , $n \ge 1$ } *be a sequence of random variables, which is stochastically dominated by a random variable X. Then, for any* a > 0 *and* b > 0*, the following two statements hold:*

$$E\left[|X_n|^a I\left(|X_n| \le b\right)\right] \le CE\left[|X|^a I\left(|X| \le b\right)\right] + b^a P\left(|X| > b\right)$$

and

$$E\left[|X_n|^a I(|X_n| > b)\right] \le CE\left[|X|^a I(|X| > b)\right].$$

Lemma 3.3 [16] Let $\{Y_n, n \ge 1\}$ and $\{Z_n, n \ge 1\}$ be sequences of random variables. Then, for any q > 1, $\epsilon > 0$, and a > 0,

$$E\left(\max_{1\leq k\leq n}\left|\sum_{i=1}^{k} (Y_i+Z_i)\right|-\epsilon a\right)^+ \leq \left(\frac{1}{\epsilon^q} + \frac{1}{q-1}\right)\frac{1}{a^{q-1}}E\left(\max_{1\leq k\leq n}\left|\sum_{i=1}^{k} Y_i\right|^q\right) + E\left(\max_{1\leq k\leq n}\left|\sum_{i=1}^{k} Z_i\right|\right).$$

Lemma 3.4 ([17]) If l(x) > 0 is a slowly varying function as $x \to \infty$, then

- (1) $\lim_{x\to\infty} \frac{l(x)}{l(x)} = 1 \text{ for each } t > 0; \\ \lim_{x\to\infty} \frac{l(x+u)}{l(x)} = 1 \text{ for each } u > 0; \\ (2) \quad \lim_{k\to\infty} \sup_{2^k \le x < 2^{k+1}} \frac{l(x)}{l(2^k)} = 1; \\ (3) \quad \lim_{x\to\infty} x^{\delta} l(x) = \infty, \\ \lim_{x\to\infty} x^{-\delta} l(x) = 0 \text{ for each } \delta > 0; \\ (4) \quad C_1 2^{kr} l(\epsilon 2^k) \le \sum_{j=1}^k 2^{jr} l(\epsilon 2^j) \le C_2 2^{kr} l(\epsilon 2^k) \text{ for every } r > 0, \\ \epsilon > 0, \\ positive integer k, \end{cases}$ and some $C_1 > 0$, $C_2 > 0$;
- (5) $C_1 2^{kr} l(\epsilon 2^k) \leq \sum_{j=k}^{\infty} 2^{jr} l(\epsilon 2^j) \leq C_2 2^{kr} l(\epsilon 2^k)$ for every $r < 0, \epsilon > 0$, positive integer k, and some $C_1 > 0$, $C_2 > 0$.

3.2 Proof of Theorem 2.1

For fixed $n \ge 1$, denote

$$Y_{ni} = A_i X_i I(|X_i| \le c_n) - E[A_i X_i I(|X_i| \le c_n) | \mathcal{G}_{i-1}], \quad i = 1, 2, \dots$$

Since

$$A_i X_i = A_i X_i I(|X_i| > c_n) + Y_{ni} + E[A_i X_i I(|X_i| \le c_n) |\mathcal{G}_{i-1}],$$

we have

$$\sum_{n=1}^{\infty} b_n l(n) P\left(\max_{1 \le j \le n} \left| \sum_{i=1}^{j} A_i X_i \right| \ge \epsilon c_n \right)$$

$$\leq \sum_{n=1}^{\infty} b_n l(n) P\left(\max_{1 \le j \le n} \left| \sum_{i=1}^{j} A_i X_i I(|X_i| > c_n) \right| \ge \epsilon c_n / 3 \right)$$

$$+ \sum_{n=1}^{\infty} b_n l(n) P\left(\max_{1 \le j \le n} \left| \sum_{i=1}^{j} E[A_i X_i I(|X_i| \le c_n) |\mathcal{G}_{i-1}] \right| \ge \epsilon c_n / 3 \right)$$

$$+ \sum_{n=1}^{\infty} b_n l(n) P\left(\max_{1 \le j \le n} \left| \sum_{i=1}^{j} Y_{ni} \right| \ge \epsilon c_n / 3 \right)$$

$$=: H + I + J. \tag{3.1}$$

To prove (2.5), it is enough to show $H < \infty$, $I < \infty$, and $J < \infty$. Obviously, it follows from Hölder's inequality, Lyapunov's inequality, and (2.3) that

$$\sum_{i=1}^{n} E|A_i| \le \left(\sum_{i=1}^{n} E|A_i|^q\right)^{\frac{1}{q}} \left(\sum_{i=1}^{n} 1\right)^{1-\frac{1}{q}} = O(n).$$
(3.2)

By the fact that $\{A_n, n \ge 1\}$ is independent of $\{X_n, n \ge 1\}$, it is easy to check by Markov's inequality, Lemma 3.2, (3.2), (2.1), and (2.4) that

$$H \leq C \sum_{n=1}^{\infty} \frac{b_n l(n)}{c_n} E\left(\max_{1 \leq j \leq n} \left| \sum_{i=1}^{j} A_i X_i I(|X_i| > c_n) \right| \right)$$

$$\leq C \sum_{n=1}^{\infty} \frac{b_n l(n)}{c_n} \sum_{i=1}^{n} E|A_i| E[|X_i| I(|X_i| > c_n)]$$

$$\leq C \sum_{n=1}^{\infty} \frac{n b_n l(n)}{c_n} E[|X| I(|X| > c_n)]$$

$$= C \sum_{n=1}^{\infty} \frac{n b_n l(n)}{c_n} \sum_{m=n}^{\infty} E[|X| I(c_m < |X| \leq c_{m+1})]$$

$$= C \sum_{m=1}^{\infty} E[|X| I(c_m < |X| \leq c_{m+1})] \sum_{n=1}^{m} \frac{n b_n l(n)}{c_n}$$

$$\leq C \sum_{m=1}^{\infty} E[|X| I(c_m < |X| \leq c_{m+1})] c_m^{p-1} l(c_m^{1/\alpha})$$

$$\leq C E|X|^p l(|X|^{1/\alpha}) < \infty.$$
(3.3)

For *I*, since $\{X_n, \mathcal{F}_n, n \ge 1\}$ is a sequence of martingale difference, we can see that $\{X_n, \mathcal{G}_n, n \ge 1\}$ is also a sequence of martingale difference. Combining with the fact that $\{A_n, n \ge 1\}$ is independent of $\{X_n, n \ge 1\}$, we have

$$E(A_n X_n | \mathcal{G}_{n-1}) = E[E(A_n X_n | \mathcal{G}_n) | \mathcal{G}_{n-1}]$$
$$= E[X_n E(A_n | \mathcal{G}_n) | \mathcal{G}_{n-1}]$$
$$= EA_n E(X_n | \mathcal{G}_{n-1})$$
$$= 0 \quad \text{a.s., } n \ge 1.$$

Consequently, by Markov's inequality and the proof of (3.3), we have

$$I \leq C \sum_{n=1}^{\infty} \frac{b_n l(n)}{c_n} E\left(\max_{1\leq j\leq n} \left| \sum_{i=1}^{j} E[A_i X_i I(|X_i| \leq c_n) |\mathcal{G}_{i-1}] \right| \right)$$

$$\leq C \sum_{n=1}^{\infty} \frac{b_n l(n)}{c_n} E\left(\max_{1\leq j\leq n} \left| \sum_{i=1}^{j} E[A_i X_i I(|X_i| > c_n) |\mathcal{G}_{i-1}] \right| \right)$$

$$\leq C \sum_{n=1}^{\infty} \frac{b_n l(n)}{c_n} \sum_{i=1}^{n} E|A_i| E[|X_i| I(|X_i| > c_n)]$$

$$\leq C E|X|^p l(|X|^{1/\alpha}) < \infty.$$
(3.4)

Next, we shall show that $J < \infty$. Let $X_{ni} = X_i I(|X_i| \le c_n)$ and $\hat{Y}_{ni} = a_i X_{ni} - E(a_i X_{ni} | \mathcal{G}_{i-1})$. It can be found that for fixed real numbers a_1, \ldots, a_n , $\{\hat{Y}_{ni}, \mathcal{G}_i, 1 \le i \le n\}$ is also a sequence of martingale difference. Note that $\{A_1, \ldots, A_n\}$ is independent of $\{X_{n1}, \ldots, X_{nn}\}$. So, by

Markov's inequality and Lemma 3.1, we have

$$\begin{split} & J \leq C \sum_{n=1}^{\infty} \frac{b_n l(n)}{c_n^q} E\left(\max_{1 \leq j \leq n} \left| \sum_{i=1}^j Y_{ni} \right| \right)^q \\ &= C \sum_{n=1}^{\infty} \frac{b_n l(n)}{c_n^q} E\left\{ E\left(\max_{1 \leq j \leq n} \left| \sum_{i=1}^j \left[a_i X_{ni} - E(a_i X_{ni} | \mathcal{G}_{i-1}) \right] \right|^q \right) \middle| A_1 = a_1, \dots, A_n = a_n \right\} \\ &\leq C \sum_{n=1}^{\infty} \frac{b_n l(n)}{c_n^q} E\left\{ E\left(\sum_{i=1}^n E(\hat{Y}_{ni}^2 | \mathcal{G}_{i-1})\right)^{q/2} + \sum_{i=1}^n E|\hat{Y}_{ni}|^q \middle| A_1 = a_1, \dots, A_n = a_n \right\} \\ &= C \sum_{n=1}^{\infty} \frac{b_n l(n)}{c_n^q} \sum_{i=1}^n E|Y_{ni}|^q + C \sum_{n=1}^{\infty} \frac{b_n l(n)}{c_n^q} E\left(\sum_{i=1}^n E(Y_{ni}^2 | \mathcal{G}_{i-1})\right)^{q/2} \\ &=: J_1 + J_2. \end{split}$$
(3.5)

For J_1 , we have by C_r -inequality, Lemma 3.2 with $b = c_n$, and (2.3) that

$$J_{1} \leq C \sum_{n=1}^{\infty} \frac{b_{n}l(n)}{c_{n}^{q}} \sum_{i=1}^{n} E|A_{i}|^{q} E\left[|X_{i}|^{q} I\left(|X_{i}| \leq c_{n}\right)\right]$$

$$\leq C \sum_{n=1}^{\infty} \frac{b_{n}l(n)}{c_{n}^{q}} \sum_{i=1}^{n} E|A_{i}|^{q} E\left[|X|^{q} I\left(|X| \leq c_{n}\right)\right] + C \sum_{n=1}^{\infty} \frac{b_{n}l(n)}{c_{n}^{q}} \sum_{i=1}^{n} c_{n}^{q} P\left(|X| > c_{n}\right)$$

$$= C \sum_{n=1}^{\infty} \frac{nb_{n}l(n)}{c_{n}^{q}} E\left[|X|^{q} I\left(|X| \leq c_{n}\right)\right] + C \sum_{n=1}^{\infty} nb_{n}l(n) P\left(|X| > c_{n}\right)$$

$$\leq C \sum_{n=1}^{\infty} \frac{nb_{n}l(n)}{c_{n}^{q}} E\left[|X|^{q} I\left(|X| \leq c_{n}\right)\right] + C \sum_{n=1}^{\infty} \frac{nb_{n}l(n)}{c_{n}} E\left[|X|I\left(|X| > c_{n}\right)\right]$$

$$=: CJ_{11} + CJ_{12}.$$
(3.6)

For J_{11} , we have by (2.1) and (2.4) that

$$J_{11} = \sum_{n=1}^{\infty} \frac{nb_n l(n)}{c_n^n} \sum_{m=1}^n E[|X|^q I(c_{m-1} < |X| \le c_m)]$$

$$= \sum_{m=1}^{\infty} E[|X|^q I(c_{m-1} < |X| \le c_m)] \sum_{n=m}^{\infty} \frac{nb_n l(n)}{c_n^q}$$

$$\le C \sum_{m=1}^{\infty} E[|X|^q I(c_{m-1} < |X| \le c_m)] c_m^{p-q} l(c_m^{1/\alpha})$$

$$\le C E|X|^p l(|X|^{1/\alpha}) < \infty.$$
(3.7)

By the proof of (3.3), it follows

$$J_{12} = \sum_{n=1}^{\infty} \frac{n b_n l(n)}{c_n} E[|X| I(|X| > c_n)]$$

$$\leq C E|X|^p l(|X|^{1/\alpha}) < \infty.$$
(3.8)

Furthermore, by Hölder's inequality and (2.3), for any 1 , we have

$$\sum_{i=1}^{n} E|A_i|^p \le \left(\sum_{i=1}^{n} E|A_i|^q\right)^{\frac{p}{q}} \left(\sum_{i=1}^{n} 1\right)^{1-\frac{p}{q}} = O(n).$$
(3.9)

Obviously, for $1 \le i \le n$, it has

$$E(Y_{ni}^{2}|\mathcal{G}_{i-1}) = E[A_{i}^{2}X_{i}^{2}I(|X_{i}| \leq c_{n})|\mathcal{G}_{i-1}] - [E(A_{i}X_{i}I(|X_{i}| \leq c_{n})|\mathcal{G}_{i-1})]^{2} \leq E[A_{i}^{2}X_{i}^{2}I(|X_{i}| \leq c_{n})|\mathcal{G}_{i-1}] \leq EA_{i}^{2}E(X_{i}^{2}|\mathcal{G}_{i-1}), \quad \text{a.s.}$$
(3.10)

Combining (3.9) and (2.2), we obtain that

$$J_{2} \leq \sum_{n=1}^{\infty} \frac{b_{n}l(n)}{c_{n}^{q}} E\left(\sum_{i=1}^{n} EA_{i}^{2}E(X_{i}^{2}|\mathcal{G}_{i-1})\right)^{q/2}$$

$$\leq \sum_{n=1}^{\infty} \frac{b_{n}l(n)}{c_{n}^{q}} \left(\sum_{i=1}^{n} EA_{i}^{2}\right)^{q/2} E\left(\sup_{i\geq 1} E(X_{i}^{2}|\mathcal{G}_{i-1})\right)^{q/2}$$

$$\leq C\sum_{n=1}^{\infty} \frac{n^{q/2}b_{n}l(n)}{c_{n}^{q}} < \infty.$$
(3.11)

By (3.1) and (3.3)-(3.11), we can get (2.5). This completes the proof of Theorem 2.1.

3.3 Proof of Theorem 2.2

As the proof of Theorem 2.1,

$$A_{i}X_{i} = A_{i}X_{i}I(|X_{i}| > c_{n}) + Y_{ni} + E[A_{i}X_{i}I(|X_{i}| \le c_{n})|\mathcal{G}_{i-1}],$$

where $Y_{ni} = A_i X_i I(|X_i| \le c_n) - E[A_i X_i I(|X_i| \le c_n) | \mathcal{G}_{i-1}], i = 1, 2, ...$ By Lemma 3.3 with $a = c_n$, we have

$$\begin{split} &\sum_{n=1}^{\infty} b_n l(n) E\left(\max_{1 \le j \le n} \left| \sum_{i=1}^{j} A_i X_i \right| - \epsilon c_n \right)^+ \\ &\leq C \sum_{n=1}^{\infty} \frac{b_n l(n)}{c_n^{q-1}} E\left(\max_{1 \le j \le n} \left| \sum_{i=1}^{j} Y_{ni} \right|^q \right) \\ &+ \sum_{n=1}^{\infty} b_n l(n) E\left(\max_{1 \le j \le n} \left| \sum_{i=1}^{j} \left[A_i X_i I(|X_i| > c_n) + E(A_i X_i I(|X_i| \le c_n) | \mathcal{G}_{i-1}) \right] \right| \right) \\ &\leq \sum_{n=1}^{\infty} b_n l(n) E\left(\max_{1 \le j \le n} \left| \sum_{i=1}^{j} A_i X_i I(|X_i| > c_n) \right| \right) \\ &+ \sum_{n=1}^{\infty} b_n l(n) E\left(\max_{1 \le j \le n} \left| \sum_{i=1}^{j} E[A_i X_i I(|X_i| \le c_n) | \mathcal{G}_{i-1}] \right| \right) \end{split}$$

$$+ C \sum_{n=1}^{\infty} \frac{b_n l(n)}{c_n^{q-1}} E\left(\max_{1 \le j \le n} \left| \sum_{i=1}^j Y_{ni} \right|^q \right)$$

=: $H_1 + H_2 + H_3.$ (3.12)

To prove (2.11), it is enough to show $H_1 < \infty$, $H_2 < \infty$, and $H_3 < \infty$.

Since $\{A_n, n \ge 1\}$ is independent of $\{X_n, n \ge 1\}$, we have by Lemma 3.2, (3.2), (2.7), and (2.10) that

$$H_{1} \leq \sum_{n=1}^{\infty} b_{n} l(n) \sum_{i=1}^{n} E|A_{i}|E[|X_{i}|I(|X_{i}| > c_{n})]$$

$$\leq C \sum_{n=1}^{\infty} n b_{n} l(n) E[|X|I(|X| > c_{n})]$$

$$= C \sum_{n=1}^{\infty} n b_{n} l(n) \sum_{m=n}^{\infty} E[|X|I(c_{m} < |X| \le c_{m+1})]$$

$$= C \sum_{m=1}^{\infty} E[|X|I(c_{m} < |X| \le c_{m+1})] \sum_{n=1}^{m} n b_{n} l(n)$$

$$\leq C \sum_{m=1}^{\infty} E[|X|I(c_{m} < |X| \le c_{m+1})] c_{m}^{p-1} l(c_{m}^{1/\alpha})$$

$$\leq C E|X|^{p} l(|X|^{1/\alpha}) < \infty.$$
(3.13)

For H_2 , by a similar proof of (3.4), we have $E(A_nX_n|\mathcal{G}_{n-1}) = 0$ a.s., $n \ge 1$. Combining with (3.13), we get

$$H_{2} = \sum_{n=1}^{\infty} b_{n} l(n) E\left(\max_{1 \le j \le n} \left| \sum_{i=1}^{j} E\left[A_{i} X_{i} I\left(|X_{i}| \le c_{n}\right)|\mathcal{G}_{i-1}\right] \right| \right)$$

$$= \sum_{n=1}^{\infty} b_{n} l(n) E\left(\max_{1 \le j \le n} \left| \sum_{i=1}^{j} E\left[A_{i} X_{i} I\left(|X_{i}| > c_{n}\right)|\mathcal{G}_{i-1}\right] \right| \right)$$

$$\leq \sum_{n=1}^{\infty} b_{n} l(n) \sum_{i=1}^{n} E|A_{i}| E\left[|X_{i}| I\left(|X_{i}| > c_{n}\right)\right]$$

$$\leq E|X|^{p} l\left(|X|^{1/\alpha}\right) < \infty.$$
(3.14)

Next, from a similar proof of Theorem 2.1 (see (3.5)), we turn to prove $H_3 < \infty$.

$$H_{3} = C \sum_{n=1}^{\infty} \frac{b_{n}l(n)}{c_{n}^{q-1}} E\left(\max_{1 \le j \le n} \left|\sum_{i=1}^{j} Y_{ni}\right|\right)^{q}$$

$$\leq C \sum_{n=1}^{\infty} \frac{b_{n}l(n)}{c_{n}^{q-1}} \sum_{i=1}^{n} E|Y_{ni}|^{q} + C \sum_{n=1}^{\infty} \frac{b_{n}l(n)}{c_{n}^{q-1}} E\left(\sum_{i=1}^{n} E(Y_{ni}^{2}|\mathcal{G}_{i-1})\right)^{q/2}$$

$$=: H_{31} + H_{32}.$$
(3.15)

For H_{31} , by C_r -inequality, Lemma 3.2, and (2.9), we have

$$H_{31} \leq C \sum_{n=1}^{\infty} \frac{b_n l(n)}{c_n^{q-1}} \sum_{i=1}^n E|A_i|^q E[|X_i|^q I(|X_i| \leq c_n)]$$

$$\leq C \sum_{n=1}^{\infty} \frac{b_n l(n)}{c_n^{q-1}} \sum_{i=1}^n E|A_i|^q E[|X|^q I(|X| \leq c_n)]$$

$$+ C \sum_{n=1}^{\infty} n b_n l(n) c_n P(|X| > c_n)$$

$$\leq C \sum_{n=1}^{\infty} \frac{n b_n l(n)}{c_n^{q-1}} E[|X|^q I(|X| \leq c_n)]$$

$$+ C \sum_{n=1}^{\infty} n b_n l(n) E[|X| I(|X| > c_n)]$$

$$=: C \widehat{H_{31}} + C \widehat{H_{32}}.$$
(3.16)

From the condition $q > \max\{2, p\}$, (2.7), and (2.10), we get

$$\widehat{H_{31}} = \sum_{n=1}^{\infty} \frac{nb_n l(n)}{c_n^{q-1}} \sum_{m=1}^n E[|X|^q I(c_{m-1} < |X| \le c_m)]$$

$$= \sum_{m=1}^{\infty} E[|X|^q I(c_{m-1} < |X| \le c_m)] \sum_{n=m}^{\infty} \frac{nb_n l(n)}{c_n^{q-1}}$$

$$\le C \sum_{m=1}^{\infty} E[|X|^q I(c_{m-1} < |X| \le c_m)] c_m^{p-q} l(c_m^{1/\alpha})$$

$$\le C E|X|^p l(|X|^{1/\alpha}) < \infty.$$
(3.17)

By the proof of (3.13), it follows

$$\widehat{H_{32}} = \sum_{n=1}^{\infty} nb_n l(n) E\Big[|X|I\big(|X| > c_n\big)\Big] \le CE|X|^p l\big(|X|^{1/\alpha}\big) < \infty.$$
(3.18)

For H_{32} , from a similar proof of Theorem 2.1 (see (3.9)–(3.11)), combining (3.10), (3.9), and (2.8), we get

$$H_{32} \leq \sum_{n=1}^{\infty} \frac{b_n l(n)}{c_n^{q-1}} E\left(\sum_{i=1}^n EA_i^2 E(X_i^2 | \mathcal{G}_{i-1})\right)^{q/2}$$

$$\leq \sum_{n=1}^{\infty} \frac{b_n l(n)}{c_n^{q-1}} \left(\sum_{i=1}^n EA_i^2\right)^{q/2} E\left(\sup_{i\geq 1} E(X_i^2 | \mathcal{G}_{i-1})\right)^{q/2}$$

$$\leq C \sum_{n=1}^{\infty} \frac{n^{q/2} b_n l(n)}{c_n^{q-1}} < \infty.$$
(3.19)

Therefore, we can get (2.11) by (3.12)-(3.19). This completes the proof of Theorem 2.2.

3.4 Proof of Theorem 2.3

By a similar proof of Theorem 2.2, we take q = 2. It is enough to prove $H_1 < \infty$, $H_2 < \infty$, and $H_3 < \infty$. Combining (3.13) with conditions (2.12), (2.13), we have

$$H_{1} = \sum_{n=1}^{\infty} b_{n} l(n) E\left(\max_{1 \le j \le n} \left| \sum_{i=1}^{j} A_{i} X_{i} I(|X_{i}| > c_{n}) \right| \right)$$

$$\leq \sum_{n=1}^{\infty} b_{n} l(n) \sum_{i=1}^{n} E|A_{i}| E[|X_{i}|I(|X_{i}| > c_{n})]$$

$$\leq C \sum_{m=1}^{\infty} E[|X|I(c_{m} < |X| \le c_{m+1})] \sum_{n=1}^{m} n b_{n} l(n)$$

$$\leq C \sum_{m=1}^{\infty} E[|X|I(c_{m} < |X| \le c_{m+1})] (\log c_{m}) l(c_{m}^{1/\alpha})$$

$$\leq C E[|X|(\log |X|) l(|X|^{1/\alpha})] < \infty.$$
(3.20)

By the proof of (3.14) and (3.20), we get

$$H_{2} = \sum_{n=1}^{\infty} b_{n} l(n) E\left(\max_{1 \le j \le n} \left| \sum_{i=1}^{j} E[A_{i} X_{i} I(|X_{i}| \le c_{n}) |\mathcal{G}_{i-1}] \right| \right)$$

$$\leq \sum_{n=1}^{\infty} b_{n} l(n) \sum_{i=1}^{n} E|A_{i}| E[|X_{i}| I(|X_{i}| > c_{n})]$$

$$\leq E[|X| (\log |X|) l(|X|^{1/\alpha})] < \infty.$$
(3.21)

For H_3 , from a similar proof of Theorem 2.1 (see (3.5) for q = 2), we have

$$\begin{aligned} H_{3} &= C \sum_{n=1}^{\infty} \frac{b_{n}l(n)}{c_{n}} E\left(\max_{1 \le j \le n} \left|\sum_{i=1}^{j} Y_{ni}\right|^{2}\right) \\ &= C \sum_{n=1}^{\infty} \frac{b_{n}l(n)}{c_{n}} E\left\{E\left(\max_{1 \le j \le n} \left|\sum_{i=1}^{j} \left[a_{i}X_{ni} - E(a_{i}X_{ni}|\mathcal{G}_{i-1})\right]\right|^{2}\right) \middle| A_{1} = a_{1}, \dots, A_{n} = a_{n}\right\} \\ &\leq C \sum_{n=1}^{\infty} \frac{b_{n}l(n)}{c_{n}} E\left\{\left(\sum_{i=1}^{n} E|\hat{Y}_{ni}|^{2}\right) \middle| A_{1} = a_{1}, \dots, A_{n} = a_{n}\right\} \\ &= C \sum_{n=1}^{\infty} \frac{b_{n}l(n)}{c_{n}} \sum_{i=1}^{n} E|Y_{ni}|^{2}. \end{aligned}$$

Then, according to (2.12) and (3.20), we can get

$$H_{3} \leq C \sum_{n=1}^{\infty} \frac{b_{n}l(n)}{c_{n}} \sum_{i=1}^{n} E|A_{i}|^{2} E[|X_{i}|^{2} I(|X_{i}| \leq c_{n})]$$

$$\leq C \sum_{n=1}^{\infty} \frac{nb_{n}l(n)}{c_{n}} \sum_{i=1}^{n} E[|X|^{2} I(|X| \leq c_{n})] + C \sum_{n=1}^{\infty} nb_{n}l(n)c_{n}P(|X| > c_{n})$$

$$\leq C \sum_{n=1}^{\infty} \frac{nb_n l(n)}{c_n} \sum_{m=1}^{n} E[|X|^q I(c_{m-1} < |X| \le c_m)] + C \sum_{n=1}^{\infty} nb_n l(n) E[|X| I(|X| > c_n)] = \sum_{m=1}^{\infty} E[|X|^q I(c_{m-1} < |X| \le c_m)] \sum_{n=m}^{\infty} \frac{nb_n l(n)}{c_n} + CE[|X| (\log |X|) l(|X|^{1/\alpha})] \leq C \sum_{m=1}^{\infty} E[|X|^q I(c_{m-1} < |X| \le c_m)] c_m^{1-q} l(c_m^{1/\alpha}) + C < \infty.$$
(3.22)

Hence, the desired result follows from (3.20)-(3.22). This completes the proof of Theorem 2.3.

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Authors' contributions

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Author details

¹ College of Mathematics and Information Science, Henan Normal University, Xinxiang, China. ²Science College, Beijing Information Science and Technology University, Beijing, China.

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