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Comparison of the prognostic values of three calculation methods for echocardiographic relative wall thickness in acute decompensated heart failure

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Abstract

Purpose: Left ventricular (LV) wall thickness can be measured at the posterior wall (PW) and the intraventricular septum (IVS) in a parasternal long axis view by transthoracic echocardiography. Thus, there are three methods to calculate relative wall thickness as follows: $RWT_{PW} = 2 \times PWth / LVDd$; $RWT_{IVS + PW} = (IVSth + PWth) / LVDd$; and $RWT_{IVS} = 2 \times IVSth / LVDd$ (IVSth = interventricular septum thickness; LVDd = LV internal dimension at end-diastole; PWth = posterior wall thickness). The aim was to compare the prognostic values of these RWTs in patients with acute decompensated heart failure (ADHF).

Method: This was a single-center, retrospective, observational study at a Japanese community hospital. A total of 389 hospitalized ADHF patients were divided into two groups based on the three median RWT values. The primary outcome was all-cause death. Survival analysis was performed, and Cox proportional hazard models unadjusted and adjusted by Get With The Guideline score were used.

Results: High- RWT_{PW} had poor survival (log-rank, $P = 0.009$) and was a significant risk (unadjusted HR (95%CI), 1.72 (1.14–2.61), $P = 0.01$; adjusted HR, 1.95 (1.28–2.98), $P = 0.02$). High- $RWT_{IVS + PW}$ was not associated with poor survival on survival analysis or the unadjusted Cox model. Only the adjusted Cox model showed that High- $RWT_{IVS + PW}$ was associated with a significant risk of the primary outcome (unadjusted HR (95%CI), 1.45 (0.96–2.17), $P = 0.07$; adjusted HR, 1.53 (1.01–2.32), $P = 0.045$). High- RWT_{IVS} did not have significant prognostic value.

Conclusions: When calculating RWT, RWT_{PW} should be recommended for evaluating the mortality risk in ADHF.

Keywords: Concentric left ventricular structure, Relative wall thickness, Acute decompensated heart failure, Transthoracic echocardiography, Prognosis

Introduction

A concentric left ventricular (LV) structure is the result of remodeling that occurs with LV wall thickening relative to the LV cavity to compensate for pressure overload [1, 2]. A concentric LV structure is a risk factor for cardiovascular events in hypertensive patients [3, 4]. Furthermore, we previously reported that a concentric LV

structure evaluated by transthoracic echocardiography (TTE) was associated with poor survival in patients with acute decompensated heart failure (ADHF) [5].

Relative wall thickness (RWT) is an index of LV concentricity. RWT is the ratio of LV wall thickness to the LV internal dimension at end diastole (LVDd) [6]. LV wall thickness, which can be measured in a parasternal long-axis view by TTE, is represented by the Interventricular septum wall thickness (IVSth) and the posterior wall thickness (PWth) [6]. Therefore, there are three methods to calculate the RWT: $RWT_{PW} = 2 \times PWth / LVDd$; $RWT_{IVS + PW} = (IVSth + PWth) / LVDd$; and

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$RWT_{IVS} = 2 \times IVS_{th}/LVDD$. The American Society of Echocardiography (ASE) recommends RWT_{PW} for calculating RWT [6]. However, some studies found that RWT_{IVS+PW} had clinical significance [7, 8]. The difference in clinical significance among the three methods of measuring RWT is unclear.

To compare the clinical significance of RWT_{PW} , RWT_{IVS+PW} , and RWT_{IVS} , the prognostic values of the RWTs were examined and compared in patients with ADHF.

Materials and methods

Participants

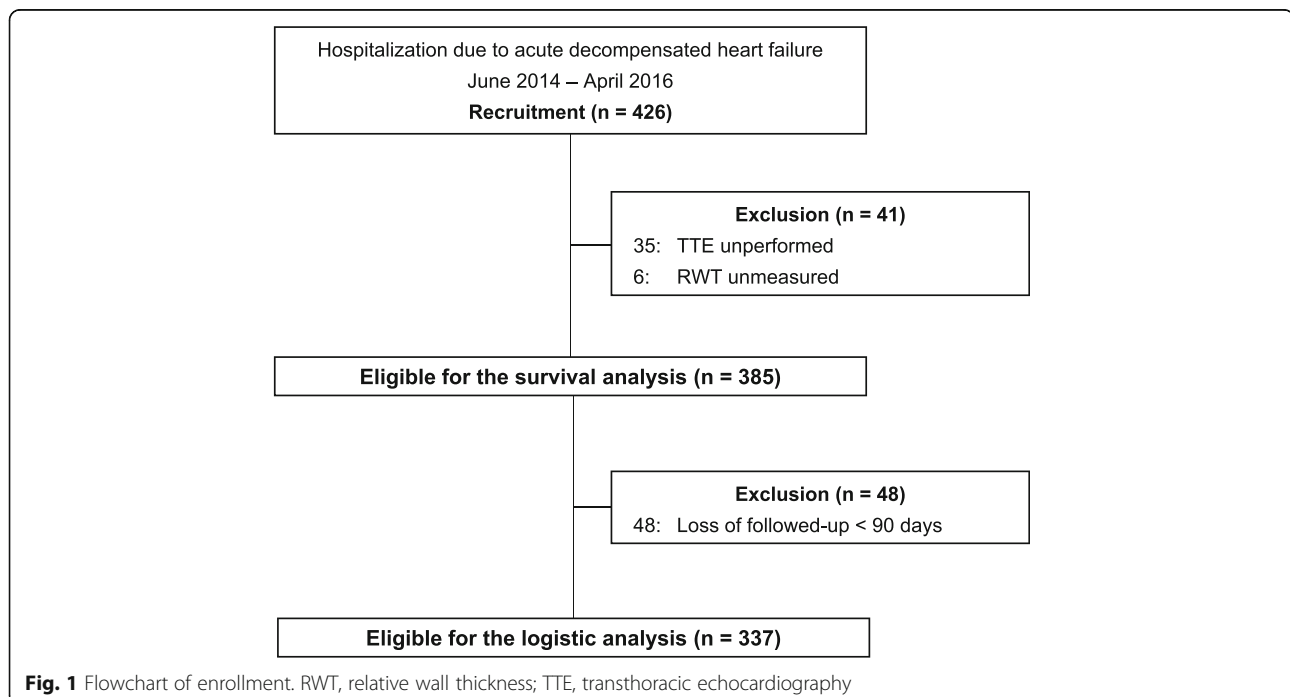
This was a single-center, retrospective, observational study conducted at a Japanese community hospital. In total, 426 consecutive patients admitted due to ADHF through the clinic or emergency room were recruited between June 2014 and April 2016 and followed-up from June 2014 to September 2016. A total of 41 patients were excluded for any of the following reasons: no TTE on admission ($n = 35$); and RWT not measured ($n = 6$). Finally, 385 patients were eligible for the analysis (Fig. 1). We previously documented the enrolled patients in detail [5].

The present study followed the tenets of the Declaration of Helsinki and the Ethical Guidelines for Medical and Health Research Involving Human Subjects proposed by the Ministry of Health and Welfare in Japan. The institutional ethics committee at Tomishiro Central Hospital approved the present study and waived informed consent because of the observational nature of the study.

Transthoracic echocardiography

Comprehensive TTE (Vivid 7 ultrasound system, GE Vingmed Ultrasound, Horten, Norway) was performed during hospital admission by four medical technicians who had at least 5 years of experience performing TTE. Their measurements followed established and standardized methods recommended by the ASE and the European Society of Cardiology. At least two attending cardiologists certified by the Japanese Circulation Society and an experienced sonographer reviewed the echocardiography reports immediately after comprehensive TTE. LV geometry, including PW_{th} , IVS_{th} , and $LVDD$, was measured in M-mode in a parasternal long-axis view [6]. All measurements were performed from the leading edge to the leading edge [6]. RWTs were calculated by the three measurement methods and defined as follows: $RWT_{PW} = 2 \times PW_{th}/LVDD$; $RWT_{IVS+PW} = (IVS_{th} + PW_{th})/LVDD$; and $RWT_{IVS} = 2 \times IVS_{th}/LVDD$. The patients were divided into two groups based on the median RWT_{PW} (low- and high- RWT_{PW}), median RWT_{IVS+PW} (low- and high- RWT_{IVS+PW}), or median RWT_{IVS} (low- and high- RWT_{IVS}).

Left ventricular ejection fraction (LVEF) was assessed using the biplane Simpson's method [6]. Heart failure with preserved ejection fraction (HFpEF) was defined as an ejection fraction $\geq 50\%$ [9]. LV mass was computed by the Cube formula [6]. LV end-diastolic volume (LVEDVI) was estimated by the Teichholz equation [10]. Peak transmitral early diastolic wave (E wave) velocity, atrial contraction wave (A wave) velocity, and deceleration time (DCT) were measured by the pulse wave Doppler signals of the mitral



inflow in the apical four-chamber view [11]. Valvular diseases were evaluated using a semiquantitative 4-grade scale (none, mild, moderate, and severe) [12].

Data collection

Cardiologists followed the patients at Tomishiro Central Hospital Clinic every 1–3 months after hospital discharge. Medical clerks confirmed the patients' condition if the patients canceled the appointment.

Patients' medical charts were reviewed to collect their demographic characteristics and clinical data, including medications, laboratory tests, and hemodynamic data on hospital admission. The primary outcome was all-cause death. Death was confirmed by the medical chart, telephone call with a patient's family, or obituary in local newspapers.

Statistical analysis

Continuous variables with normal and skewed distributions are presented as means (SD) and medians [25th, 75th percentiles], respectively. Categorical variables are presented as numbers with a percentage.

In two-group comparisons, Student's *t*-test and the Mann-Whitney U test were used to compare normally distributed and non-normally distributed continuous variables, respectively. Fisher's exact test was used for categorical variables.

Survival analysis

During follow-up (235 [92, 425] days), 95/385 (25%) patients died. Survival analysis for all-cause death was performed. Kaplan-Meier curves were stratified by RWT_{PW} , RWT_{IVS+PW} , and RWT_{IVS} . The log-rank test was used to compare survival curves. High- RWT_{PW} , high- RWT_{IVS+PW} , and high- RWT_{IVS} were examined by univariate Cox proportional hazard models and a Cox proportional hazard model adjusted by the Get With The Guideline score (GWTG) [13, 14], an established risk score for mortality in patients with acute heart failure, to obtain hazard ratios (HRs) and 95% confidence intervals (95% CIs).

Logistic regression model for 90-day mortality

A total of 48 patients who were lost to follow-up were excluded to evaluate the risk of 90-day mortality. Logistic regression models were used to obtain the odds ratios (ORs) of 90-day mortality and 95% CIs. High- RWT_{PW} , high- RWT_{IVS+PW} , and high- RWT_{IVS} were examined in univariate logistic regression models and a logistic regression model adjusted by GWTG.

Receiver operating curves for 90-day mortality

Receiver operating curves for 90-day mortality were drawn using RWT_{PW} , RWT_{IVS+PW} , and RWT_{IVS} to

obtain c-statistics, and the best RWT cut-off values were determined by the maximum Youden index [15].

Sensitivity analysis of the survival analysis by stratified RWTs by the best cut-off

To confirm the consistency of the survival analysis, the participants were divided based on the best RWT cut-off value derived from the Youden index.

Survival analysis was performed to compare low and high-RWTs. High- RWT_{PW} , high- RWT_{IVS+PW} , and high- RWT_{IVS} were also examined with univariate and adjusted proportional Cox hazard models.

Relationships between RWTs and clinical characteristics

Spearman's correlation coefficient (ρ) was used to identify significant associations between RWTs and clinical characteristics: age, the natural logarithm of brain natriuretic peptide (logBNP), LVEF, LVEDV, and systolic blood pressure (SBP).

Reliability of measurement of PWth and IVSth

The reliabilities of the TTE measurements of PWth, IVSth, and LVDD were examined in 25 patients whose TTE image quality was good, and all of the patients underwent TTE performed by the same one of four medical technicians. The medical technician and two other examiners re-measured PWth and IVSth in the TTE image stored in the local server on hospital admission, using an off-line image analysis system (Nahri Aqua, Mehergen Group, Fukuoka, Japan). Comparing every two examiners' measurements, Bland-Altman plots were used to assess the agreement between the measurement by the same examiner and different examiners [16]. The inter-class coefficient (ICC) was computed to assess agreement [17].

The reliabilities of RWT_{PW} , RWT_{IVS+PW} , and RWT_{IVS} were also examined. RWTs were computed using PWth, IVSth, and LVDD measured by three examiners. Bland-Altman plots were drawn, and the ICC and *P* values were calculated.

Software

The statistical software used was R 3.4.3 (R Foundation for Statistical Computing, Vienna, Australia). All reported *P* values are two-tailed, and a *P* value < 0.05 was considered significant.

Results

Participants

The participants' median age was 81 years, and there were 181/385 (47%) men in the overall population. Comparing low- and high- RWT_{PW} , high- RWT_{PW} had more elderly patients and more females, whereas in comparisons between low- and high- RWT_{IVS+PW}

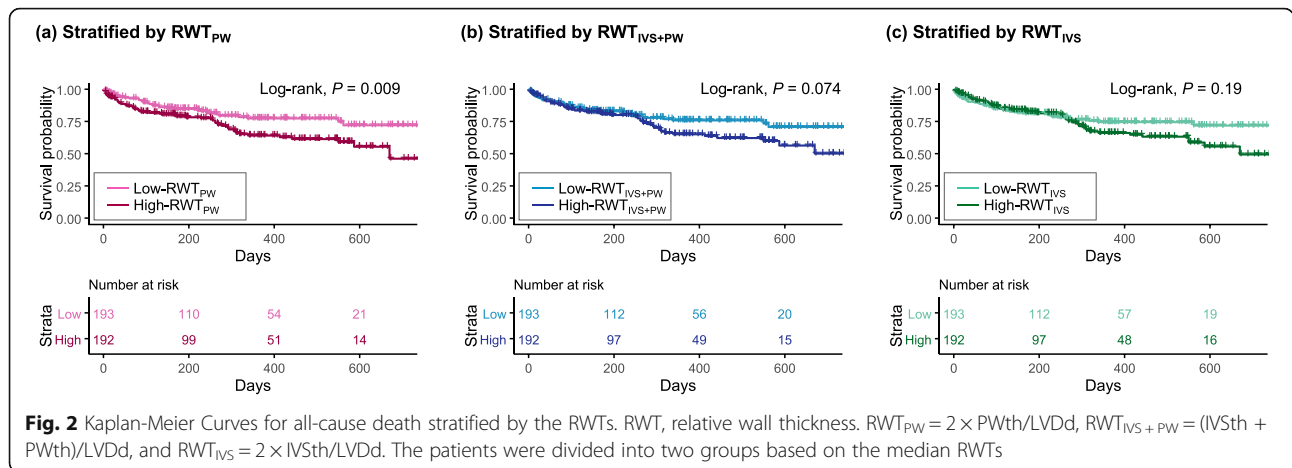
Table 1 Demographic data and echocardiographic parameters

	Overall		RWTP _{PW} ^b		P value	RWTP _{IVS+PW} ^b		P value	RWTP _{IVS} ^b		P value
	Low	High	Low	High		Low	High		Low	High	
Age, y	81 [70, 88]	83 [73, 89]	80 [68, 87]	83 [73, 89]	0.021	80 [69, 87]	83 [73, 89]	80 [69, 87]	83 [73, 88]	0.067	0.082
Male, n (%)	181/385 (47)	77/192 (40)	104/193 (54)	77/192 (40)	0.008	99/193 (51)	82/192 (43)	98/193 (51)	83/192 (43)	0.1	0.15
Height, cm	154 ± 10	153 ± 9.9	156 ± 9.7	153 ± 9.9	0.002	155 ± 10	153 ± 10	155 ± 10	153 ± 10	0.048	0.15
Body weight, kg	60 ± 16	60 ± 17	60 ± 15	60 ± 17	0.95	61 ± 15	59 ± 17	60 ± 16	59 ± 16	0.45	0.59
Body mass index, kg/m ²	22.8 ± 4.6	23.0 ± 4.8	22.6 ± 4.5	23.0 ± 4.8	0.42	22.8 ± 4.6	22.8 ± 4.7	22.9 ± 4.7	22.8 ± 4.6	0.99	0.87
Body surface area, m ²	1.51 ± 0.22	1.49 ± 0.23	1.53 ± 0.22	1.49 ± 0.23	0.11	1.53 ± 0.22	1.50 ± 0.23	1.52 ± 0.22	1.50 ± 0.22	0.2	0.3
Get With The Guideline score	38 ± 7	38 ± 8	38 ± 6	38 ± 8	0.99	38 ± 7	38 ± 8	38 ± 7	38 ± 8	0.55	0.86
Hospital stay, days	13 [8, 20]	12 [8, 19]	13 [8, 21]	12 [8, 19]	0.55	13 [8, 21]	12 [8, 19]	13 [8, 20]	12 [8, 19]	0.35	0.76
Past medical history, n (%)											
Hypertension	187/385 (49)	95/192 (50)	92/193 (48)	95/192 (50)	0.76	92/193 (48)	95/192 (50)	93/193 (48)	94/192 (49)	0.76	0.92
Diabetes mellitus	132/385 (34)	64/192 (33)	68/193 (35)	64/192 (33)	0.75	75/193 (39)	57/192 (30)	80/193 (42)	52/192 (27)	0.068	0.004
Chronic obstructive pulmonary disease	18/385 (4.7)	11/192 (5.7)	7/193 (3.6)	11/192 (5.7)	0.35	7/193 (3.6)	11/192 (5.7)	8/193 (4.1)	10/192 (5.2)	0.35	0.64
Old myocardial infarction	62/385 (16)	25/192 (13)	37/193 (19)	25/192 (13)	0.13	35/193 (18)	27/192 (14)	34/193 (18)	28/192 (15)	0.33	0.49
Echocardiographic parameters											
RWT _{PW} ^a	0.36 ± 0.12	0.45 ± 0.12	0.28 ± 0.05	0.45 ± 0.12	< 0.001	0.28 ± 0.05	0.44 ± 0.12	0.30 ± 0.07	0.43 ± 0.13	< 0.001	< 0.001
RWT _{IVS+PW} ^a	0.37 ± 0.13	0.45 ± 0.12	0.29 ± 0.06	0.45 ± 0.12	< 0.001	0.28 ± 0.05	0.46 ± 0.12	0.29 ± 0.06	0.46 ± 0.12	< 0.001	< 0.001
RWT _{IVS} ^a	0.38 ± 0.14	0.46 ± 0.15	0.30 ± 0.09	0.46 ± 0.15	< 0.001	0.29 ± 0.06	0.48 ± 0.14	0.28 ± 0.06	0.48 ± 0.13	< 0.001	< 0.001
IVSth, mm	9.4 ± 2.4	10.4 ± 2.4	8.5 ± 2.0	10.4 ± 2.4	< 0.001	8.1 ± 1.8	10.7 ± 2.2	7.9 ± 1.5	11.0 ± 2.1	< 0.001	< 0.001
PWth, mm	9.0 ± 2.1	10.3 ± 2.0	7.8 ± 1.3	10.3 ± 2.0	< 0.001	8.1 ± 1.4	10 ± 2.2	8.3 ± 1.6	9.8 ± 2.3	< 0.001	< 0.001
LVDD, mm	52 ± 9.7	46.8 ± 7.8	57.1 ± 8.8	46.8 ± 7.8	< 0.001	57.7 ± 8.2	46.2 ± 7.6	57 ± 8.8	46.8 ± 7.8	< 0.001	< 0.001
LVEF, (%)	47 ± 17	51 ± 16	41 ± 16	51 ± 16	< 0.001	39 ± 16	52 ± 16	40 ± 16	51 ± 16	< 0.001	< 0.001
HFpEF (LVEF ≥ 50%), n (%)	157/383 (41)	101/190 (53)	56/193 (29)	101/190 (53)	< 0.001	47/193 (24)	110/190 (58)	50/193 (26)	107/190 (56)	< 0.001	< 0.001
LVM, g	168 [131, 211]	164 [132, 208]	173 [135, 211]	164 [132, 208]	0.51	174 [138, 211]	164 [123, 208]	170 [135, 207]	166 [130, 212]	0.18	0.9
LVEDV, mL	130 [92, 167]	102 [79, 126]	160 [130, 194]	102 [79, 126]	< 0.001	160 [130, 194]	97 [74, 124]	160 [124, 194]	102 [79, 130]	< 0.001	< 0.001
LVM/LVEDV	1.43 ± 0.56	1.75 ± 0.61	1.10 ± 0.23	1.75 ± 0.61	< 0.001	1.08 ± 0.20	1.77 ± 0.59	1.09 ± 0.21	1.76 ± 0.6	< 0.001	< 0.001
E wave, cm/sec	97 ± 29	97 ± 28	97 ± 30	97 ± 28	0.94	95 ± 28	99 ± 29	98 ± 30	96 ± 28	0.37	0.58
A wave, cm/sec	76 ± 32	82 ± 34	70 ± 29	82 ± 34	0.011	73 ± 29	79 ± 34	74 ± 30	77 ± 33	0.13	0.57
E/A	1.21 [0.84, 1.83]	1.15 [0.82, 1.68]	1.29 [0.89, 2.04]	1.15 [0.82, 1.68]	0.18	1.21 [0.84, 1.84]	1.21 [0.86, 1.81]	1.22 [0.84, 1.83]	1.17 [0.85, 1.83]	0.81	0.6
Deceleration time, ms	150 [123, 195]	150 [128, 192]	150 [121, 196]	150 [128, 192]	0.49	149 [119, 185]	150 [129, 200]	147 [118, 181]	152 [129, 201]	0.043	0.013
Aortic valve stenosis, n (%)	29/385 (7.5)	22/192 (12)	7/193 (3.6)	22/192 (12)	0.004	8/193 (4.1)	21/192 (11)	8/193 (4.1)	21/192 (11)	0.012	0.012
Aortic valve regurgitation, n (%)	24/385 (6.2)	10/192 (5.2)	14/193 (7.3)	10/192 (5.2)	0.53	14/193 (7.3)	10/192 (5.2)	11/193 (5.7)	13/192 (6.8)	0.53	0.68

Table 1 Demographic data and echocardiographic parameters (Continued)

	Overall		RWT _{PW} ^b		P value	RWT _{IVS+PW} ^b		P value	RWT _{IVS} ^b		P value			
	n	n (%)	Low	High		Low	High		Low	High				
Mitral valve regurgitation, n (%)	59/385	(15)	39/193	(20)	0.01	40/193	(21)	19/192	(9.9)	41/193	(21)	18/192	(9.4)	0.002
Laboratory data														
Blood urea nitrogen, mg/dL	24	[17, 35]	24	[17, 34]	0.77	24	[17, 35]	24	[17, 35]	24	[17, 36]	23	[17, 35]	0.76
Creatinine, mg/dL	1.14	[0.81, 1.52]	1.15	[0.83, 1.54]	0.31	1.18	[0.83, 1.56]	1.09	[0.79, 1.50]	1.17	[0.83, 1.55]	1.07	[0.79, 1.50]	0.19
Hemoglobin, g/dL	12.0 ± 2.4		12.0 ± 2.4	11.9 ± 2.4	0.84	11.9 ± 2.4		12.0 ± 2.4		11.9 ± 2.5		12 ± 2.3		0.68
Brain natriuretic peptide, pg/mL	666	[427, 1266]	737	[449, 1376]	0.056	765	[472, 1376]	636	[401, 1092]	683	[437, 1349]	645	[413, 1190]	0.28
Medication, n (%)														
ACE-I and/or ARB	124/285	(32)	71/193	(37)	0.069	67/193	(35)	57/192	(30)	66/193	(34)	58/192	(30)	0.47
Beta blocker	153/385	(40)	78/193	(40)	0.069	75/193	(39)	78/192	(41)	74/193	(38)	79/192	(41)	0.65
Hemodynamic data														
Systolic blood pressure, mmHg	132 ± 26		128 ± 24	135 ± 29	0.006	131 ± 24		133 ± 29		130 ± 24		134 ± 29		0.13
Diastolic blood pressure, mmHg	78 ± 21		75 ± 19	73 ± 17	0.009	74 ± 18		77 ± 21		74 ± 18		77 ± 21		0.1
Heart rate, bpm	84 ± 21		83 ± 21	84 ± 21	0.81	83 ± 19		84 ± 22		83 ± 19		83 ± 22		0.78

A wave, late mitral valve inflow velocity; ACE-I angiotensin converting enzyme inhibitor; ARB angiotensin receptor blocker; E wave, early mitral valve inflow velocity; IVSth interventricular septum thickness; LVEDV left ventricular end diastolic volume; LVDD left ventricular internal dimension at end-diastole; LVEF left ventricular ejection fraction; LVM left ventricular mass; PWth posterior wall thickness; RWT relative wall thickness
^aRWT was the ratio of left ventricular wall thickness to LVDD. Left ventricular wall thickness was measured at interventricular septum as IVSth and posterior wall as PWth. Three measurement methods to compute RWT were as follows: RWT_{PW} = 2 × PWth/LVDD, RWT_{IVS+PW} = (PWth + IVSth)/LVDD, and RWT_{IVS} = 2 × IVSth/LVDD
^bThe patients were divided into two groups based on the median of RWT_{PW}, RWT_{IVS+PW}, and RWT_{IVS}



and between low- and high- RWT_{IVS} , there were no significant differences in baseline characteristics (Table 1).

Transthoracic echocardiography

The mean RWT_{PW} , RWT_{IVS+PW} , and RWT_{IVS} values in the overall population were 0.36 ± 0.12 , 0.37 ± 0.13 , and 0.38 ± 0.14 , respectively.

On comparing the three RWTs (low- vs. high- RWT_{PW} , RWT_{IVS+PW} , RWT_{IVS}), high-RWTs had thicker IVSth and PWth, smaller LVDD, greater LVEF, smaller LV end-diastolic volume, high LVM/LVEDV, and less severe mitral regurgitation than low-RWTs (Table 1).

Survival analysis

During follow-up (235 [92, 425] days), 95/385 (25%) patients died in the overall population.

Comparing low- and high- RWT_{PW} , there was a significant difference in the incidence of all-cause death (low 36/193 (19%) vs. high- RWT_{PW} 59/192 (31%), $P = 0.007$). Kaplan-Meier curves showed that high- RWT_{PW} had worse survival than low- RWT_{PW} (P for log-rank test = 0.009; Fig. 2a).

Comparing low- and high- RWT_{IVS+PW} , there was no significant difference in all-cause death (low 40/193 (21%) vs. high- RWT_{PW} 55/192 (29%), $P = 0.077$) or survival (P for log-rank test = 0.074; Fig. 2b).

In a comparison between low- and high- RWT_{IVS} , there was no significant difference in all-cause death

(low 42/193 (22%) vs. high- RWT_{IVS} 53/192 (28%), P incidence = 0.2) or survival (P for log-rank test = 0.19; Fig. 2c).

Cox proportional hazard models for all-cause death

In the unadjusted and adjusted Cox proportional hazard models, high- RWT_{PW} was a significant risk factor for all-cause death (unadjusted Cox model, HR (95% CI), 1.72 (1.41–2.61), $P = 0.01$; adjusted Cox model, 1.95 (1.28–2.98), $P = 0.02$; Table 2).

High- RWT_{IVS+PW} was not a significant risk factor for all-cause death in the unadjusted Cox proportional model (unadjusted Cox model, HR, 1.45 (0.96–2.17), $P = 0.075$), but it was in the adjusted Cox proportional hazard model (adjusted Cox model, 1.53 (1.01–2.32), $P = 0.045$; Table 2).

High- RWT_{IVS} was not a significant factor in either the unadjusted or the adjusted Cox proportional hazard model (Table 2).

Logistic regression models for 90-day mortality

The OR of high- to low- RWT_{PW} was significant (univariate, OR, 2.19, 95%CI, 1.15–2.19, $P = 0.017$; adjusted, OR, 2.26, 95%CI, 1.16–4.4, $P = 0.017$) on univariate analysis and the adjusted logistic regression model (Table 3). In contrast, the OR of neither high to low- RWT_{IVS+PW} nor RWT_{IVS} was significant on univariate analysis or the adjusted logistic regression models.

Table 2 Cox proportional hazard model for evaluate the risk of RWTs for all-cause mortality

Calculate method and factor	Unadjusted					Adjusted by GWTG				
	Event/cases	HR	95% CI		P value	Event/cases ^a	HR	95% CI		P value
High- to low- RWT_{PW}	95/385	1.72	1.14 – 2.61	–	0.01	93/380	1.95	1.28 – 2.98	–	0.002
High- to low- RWT_{IVS+PW}	95/385	1.45	0.96 – 2.17	–	0.075	93/380	1.53	1.01 – 2.32	–	0.045
High- to low- RWT_{IVS}	95/385	1.31	0.87 – 1.96	–	0.19	93/380	1.36	0.9 – 2.06	–	0.14

CI confidence interval; GWTG Get With The Guideline score; HR hazard ratio; RWT relative wall thickness

^a5 cases were removed because of GWTG missing

Table 3 Logistic models for evaluating the risk of 90 days mortality

Calculate method and factor	Unadjusted				Adjusted by GWTG			
	Event/cases	OR	95% CI	P value	Event/cases	OR	95% CI	P value
High- to low-RWT _{PW}	48/337	2.19	1.15 – 2.19	0.017	48/337	2.26	1.16 – 4.4	0.017
High- to low-RWT _{IVS + PW}	48/337	1.26	0.68 – 1.26	0.46	48/337	1.19	0.63 – 2.25	0.6
High- to low-RWT _{IVS}	48/337	0.86	0.47 – 0.86	0.64	48/337	0.8	0.42 – 1.52	0.5

CI confidence interval; GWTG Get With The Guideline score, OR odds ratio; RWT relative wall thickness

Receiver operating curves for 90-day mortality

A total of 48 (13%) patients died within 90 days from hospital admission. Figure 3 shows the receiver operating characteristic (ROC) curves for 90-day mortality using the RWTs. The c-statistic of the ROC curve using RWT_{PW} was 62.6%, and the best cut-off value was 0.35. The c-statistic of the ROC curve using RWT_{IVS + PW} was 59.7%, and the best cut-off value was 0.55. The c-statistic of the ROC curve using RWT_{IVS} was 43.1%, and the best cut-off value was 0.36.

Sensitivity analysis of the survival analysis by stratified RWTs by the best cut-off

Additional file 1: Table S1 shows the demographic data and echocardiographic data with stratification by the best RWT cut-off. High-RWT_{PW} had worse survival than low-RWT_{PW} (P for log-rank test = 0.03; Additional file 2: Figure S1a). High-RWT_{IVS + PW} also had a worse prognosis than low-RWT_{IVS + PW} (P for log-rank test < 0.001; Additional file 2: Figure S1b). In contrast, there was no significant difference in survival between low- and high-RWT_{IVS} (P for log-rank test = 0.077; Additional file 2: Figure S1c).

In the unadjusted and adjusted Cox proportional hazard models, high-RWT_{PW} and high-RWT_{IVS + PW} were associated with mortality (high-RWT_{PW}, unadjusted Cox model, HR (95% CI), 1.55 (1.04–2.33), P = 0.033; adjusted Cox model, 1.72 (1.14–2.59), P = 0.01; high-RWT_{IVS + PW}, unadjusted Cox model, HR (95% CI), 3.88 (2.34–6.43), P < 0.001; adjusted Cox model, 3.42 (2.04–5.72), P < 0.001; Additional file 3: Table S2). High-RWT_{IVS} was

not a significant risk factor in the unadjusted and adjusted Cox proportional hazard models.

Relationship between RWTs and clinical characteristics

There were significant positive correlations between the three RWTs and age and LVEF, and negative correlations between the RWTs and LogBNP and LVEDV (Table 4). RWT_{IVS + PW} and RWT_{IVS} did not have significant correlations with systolic blood pressure, but RWT_{PW} did (ρ = 0.15, P = 0.004).

Reliability of TTE measurement of PWth, IVSth, and LVDD

Intra-observer agreement of TTE measurement of PWth was significant (ICC = 0.73, P < 0.001; Fig. 4a). Inter-observer agreements of TTE measurement of PWth were also significant (observer 1 vs. 2, ICC = 0.76, P < 0.001; observer 1 vs. 3, ICC = 0.6, P < 0.001; observer 2 vs. 3, ICC = 0.7, P < 0.001; Fig. 3a). There were no systematic biases in the intra- and inter-observer agreements in PWth measurement (Fig. 4a).

Intra-observer agreement of TTE measurement of IVSth was significant (ICC = 0.88, P < 0.001; Fig. 4b). Inter-observer agreements of TTE measurement of IVSth were also significant (observer 1 vs. 2, ICC = 0.81, P < 0.001; observer 1 vs. 3, ICC = 0.77, P < 0.001; observer 2 vs. 3, ICC = 0.73, P < 0.001; Fig. 4b). There were no systematic biases in the intra- and inter-observer agreements in IVSth measurement (Fig. 4b).

Intra-observer agreement of TTE measurement of LVDD was significant (ICC = 0.94, P < 0.001; Fig. 4c). Inter-observer agreements of TTE measurement of LVDD were also significant (observer 1 vs. 2, ICC = 0.71,

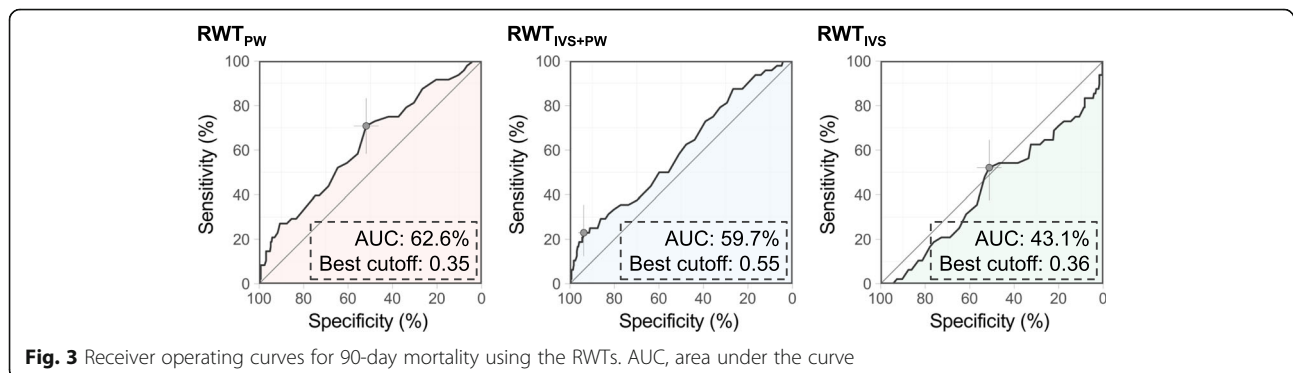


Fig. 3 Receiver operating curves for 90-day mortality using the RWTs. AUC, area under the curve

Table 4 Relationship between RWTs and clinical characteristics

	RWT _{PW}		RWT _{IVS+PW}		RWT _{IVS}	
	ρ	<i>P</i> value	ρ	<i>P</i> value	ρ	<i>P</i> value
Age, y	0.15	0.003	0.17	0.003	0.17	0.001
LogBNP, log (pg/mL)	-0.2	< 0.001	-0.15	0.003	-0.11	0.039
LVEF, %	0.42	< 0.001	0.47	< 0.001	0.43	< 0.001
LVEDV, mL	-0.67	< 0.001	-0.74	< 0.001	-0.69	< 0.001
Systolic blood pressure, mmHg	0.15	0.004	0.094	0.065	0.063	0.22

LogBNP logarithmed brain natriuretic peptide; LVEDV left ventricular end-diastolic volume; LVEF left ventricular ejection fraction; ρ , Spearman's correlation coefficient

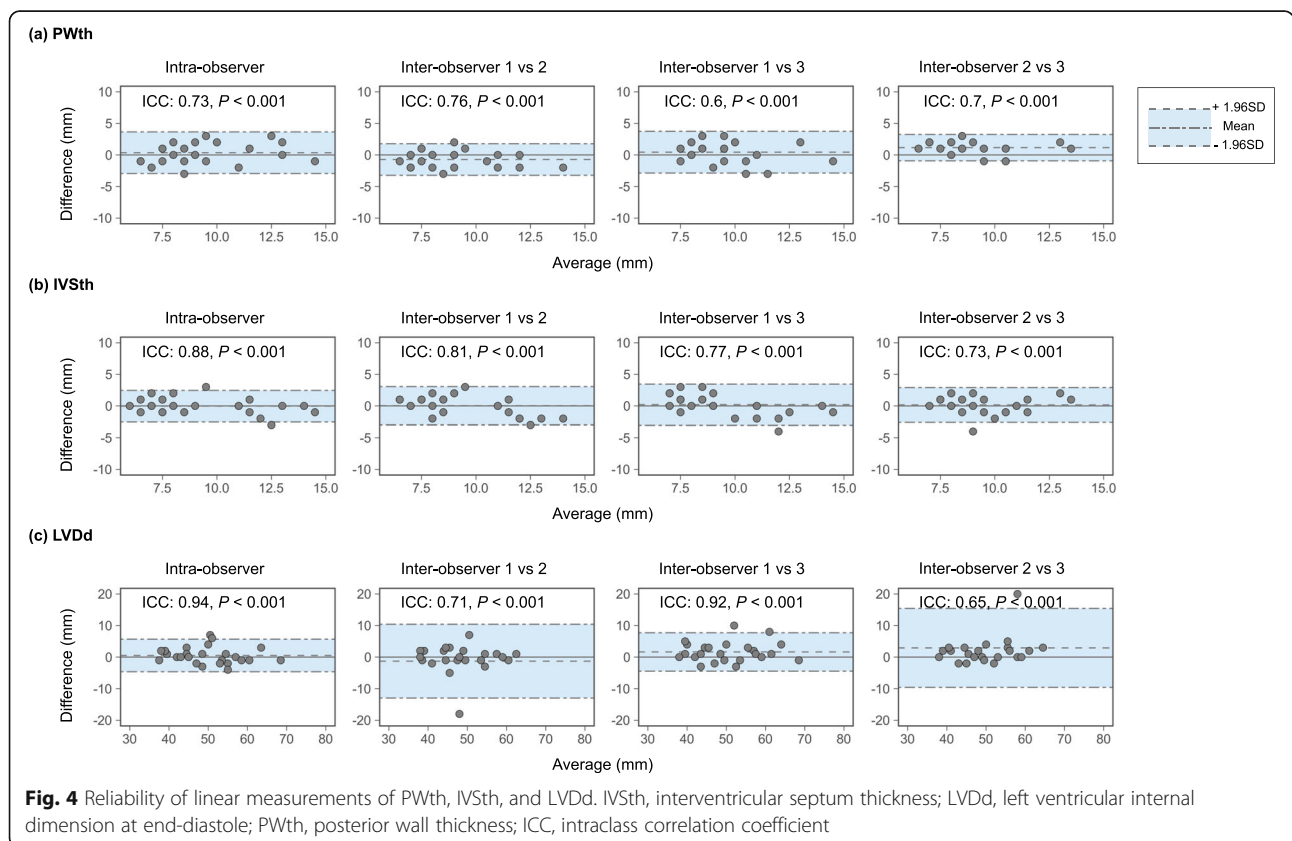
$P < 0.001$; observer 1 vs. 3, ICC = 0.92, $P < 0.001$; observer 2 vs. 3, ICC = 0.65, $P < 0.001$; Fig. 4c). There were no systematic biases in the intra- and inter-observer agreements in LVDD measurement (Fig. 4c).

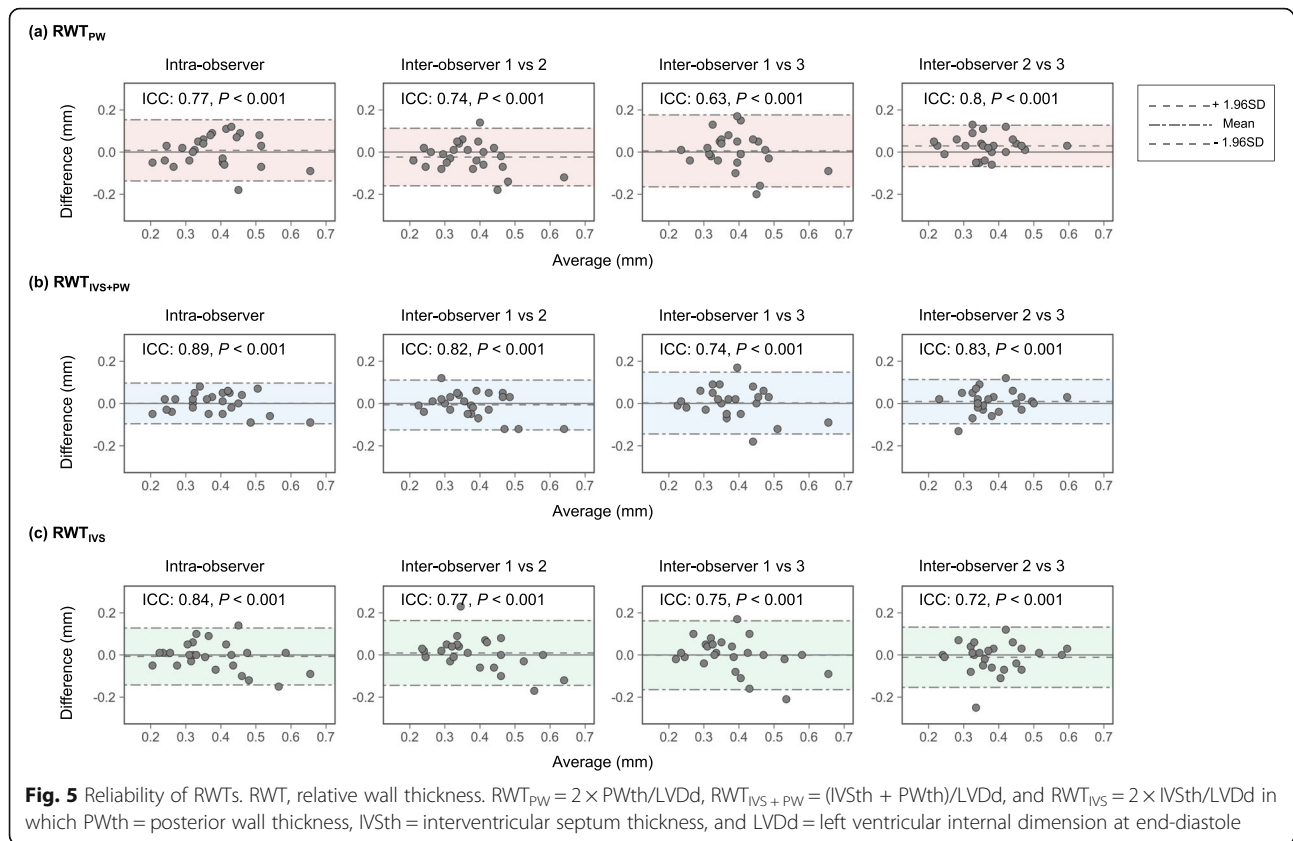
Reliability of RWTs obtained from TTE measurement

Intra-observer agreement of RWT_{PW} was significant (ICC = 0.77, $P < 0.001$; Fig. 5a). Inter-observer agreements of RWT_{PW} were significant (observer 1 vs. 2, ICC = 0.74, $P < 0.001$; observer 1 vs. 3, ICC = 0.63, $P < 0.001$; observer 2 vs. 3, ICC = 0.8, $P < 0.001$). There were no systematic biases in the intra- and inter-observer agreements in RWT_{PW}.

Intra-observer agreement of RWT_{IVS+PW} was significant (ICC = 0.89, $P < 0.001$; Fig. 5b). Inter-observer agreements of RWT_{PW} were also significant (observer 1 vs. 2, ICC = 0.82, $P < 0.001$; observer 1 vs. 3, ICC = 0.74, $P < 0.001$; observer 2 vs. 3, ICC = 0.83, $P < 0.001$). There were no systematic biases in the intra- and inter-observer agreements in RWT_{IVS+PW}.

Intra-observer agreement of RWT_{IVS} was significant (ICC = 0.84, $P < 0.001$; Fig. 5c). Inter-observer agreements of RWT_{IVS} were also significant (observer 1 vs. 2, ICC = 0.77, $P < 0.001$; observer 1 vs. 3, ICC = 0.75, $P < 0.001$; observer 2 vs. 3, ICC = 0.72, $P < 0.001$). There were no systematic biases in the intra- and inter-observer agreements in RWT_{IVS}.





Discussion

To the best of our knowledge, this is the first study to show the difference in the clinical significance of the three RWTs. The present study demonstrated that, compared to RWT_{IVS+PW} and RWT_{IVS} , RWT_{PW} is the best to stratify the risk for all-cause death in ADHF patients. This may be consistently supported by three findings. First, high- RWT_{PW} had a significantly worse prognosis than low- RWT_{PW} . In contrast, on survival analysis, there was no significant difference between high- and low- RWT_{IVS+PW} or RWT_{IVS} . Second, in the logistic regression model for 90-day mortality, only high- RWT_{PW} was significant among the three RWTs (Table 3). Third, ROC for 90-day all-cause death using RWT_{PW} had the highest c-statistic among the three ROCs.

Explanations of the differences in the prognostic values among the three RWTs

High- RWT_{PW} was associated with a poor prognosis on survival analysis and Cox proportional hazard models (Fig. 2a; Table 2). High- RWT_{IVS+PW} was not associated with poor survival on survival analysis (Fig. 1b), whereas high- RWT_{IVS+PW} was a significant risk only in the Cox proportional hazard model adjusted by GWTG, not in the unadjusted model (Table 2). High- RWT_{IVS} did not show worse survival than low- RWT_{IVS} (Fig. 1c; Table 2).

The equations of RWT_{PW} and RWT_{IVS+PW} contain PWth. PWth or the ratio of PWth to LVDd, therefore, may represent the LV remodeling related to a worse prognosis better than IVStH or IVStH to LVDd in patients with ADHF. Patients with high- RWT_{PW} had higher systolic blood pressure than those with low- RWT_{PW} (Table 1), while there was no such difference either between low- and high- RWT_{IVS+PW} or between low- and high- RWT_{IVS} . RWT_{PW} had a positive correlation with systolic blood pressure (Table 4), while either RWT_{IVS+PW} or RWT_{IVS} did not. This may suggest that thickening of PWth, rather than IVStH, is likely to counterbalance pressure overload and may lead to LV diastolic dysfunction leading to a poor prognosis. A higher A wave in high RWT_{PW} patients than in low RWT_{PW} patients may support this assumption (Table 1).

In terms of methodological validity, there were no differences in inter- and intra-observer agreements for each RWT. Given that fairly good reproducibility was observed in all measurements, differences in prognostic values among the three RWTs may not result from technical aspects of TTE.

Paradoxically, high- RWT_{PW} patients had lower BNP than low- RWT_{PW} patients (Table 1). High- RWT_{PW} included 101 (53%) patients with HFpEF. Generally, BNP increases modestly in HFpEF [18]. Furthermore,

the prognostic value of BNP has not been confirmed in patients with HFpEF [19]. High RWT_{PW} might be of clinically utility, especially, in patients with HFpEF.

Limitations

The present study had several limitations. The present study did not have pressure data such as LV end-diastolic pressure or pulmonary artery wedge pressure. LV wall thickness was not evaluated by other modalities, such as magnetic resonance imaging or computed tomography. Patients having valvular diseases with various etiologies were included, which might affect the prognostic value of RWTs.

In conclusion, high- RWT_{PW} had a higher systolic pressure and A wave than low- RWT_{PW} . This finding was not observed in the comparison between low- and high- RWT_{IVS+PW} or between low- and high- RWT_{IVS} . PW_{th} may represent pressure overload better than IVS_{th} . When calculating RWT, RWT_{PW} should be recommended for evaluating the mortality risk in ADHF.

Supplementary information

Supplementary information accompanies this paper at <https://doi.org/10.1186/s12947-019-0179-6>.

Additional file 1: Table S1. Demographic data and echocardiographic parameters.

Additional file 2: Figure S1. Kaplan-Meier Curves for all-cause mortality stratified by the stratified RWTs by the best cut-off.

Additional file 3: Table S2. Cox proportional hazard model for evaluate the risk of high RWTs for all-cause mortality.

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Authors' contributions

SY originated the concept of this study, collected data, and wrote the original manuscript. MA and TA collected the data and discussed the study results. OA took responsibility for conducting the study. MS and SU supervised the study concept and discussed the study results in depth. All authors approved the final manuscript submission.

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Availability of data and materials

Not available. We are not allowed the any study data to share by the ethical committee.

Ethics approval and consent to participate

The institutional ethics committee at Tomishiro Central Hospital approved the present study and waived informed consent because of the observational nature of the study.

Consent for publication

Not applicable.

Competing interests

The authors declare that they have no competing interests.

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References

- Grossman W, Paulus WJ. Myocardial stress and hypertrophy: a complex interface between biophysics and cardiac remodeling. *J Clin Invest*. 2013; 123(9):3701–3.
- Gjesdal O, Bluemke DA, Lima JA. Cardiac remodeling at the population level—risk factors, screening, and outcomes. *Nat Rev Cardiol*. 2011;8(12):673–85.
- Li L, Shigematsu Y, Hamada M, Hiwada K. Relative wall thickness is an independent predictor of left ventricular systolic and diastolic dysfunctions in essential hypertension. *Hypertens Res*. 2001;24(5):493–9.
- Pierdomenico SD, Lapenna D, Bucci A, Manente BM, Cucurullo F, Mezzetti A. Prognostic value of left ventricular concentric remodeling in uncomplicated mild hypertension. *Am J Hypertens*. 2004;17(11 Pt 1):1035–9.
- Yamaguchi S, Abe M, Arasaki O, Shimabukuro M, Ueda S. The prognostic impact of a concentric left ventricular structure evaluated by transthoracic echocardiography in patients with acute decompensated heart failure: a retrospective study. *Int J Cardiol*. 2018;287:73–80.
- Lang RM, Badano LP, Mor-Avi V, Afilalo J, Armstrong A, Ernande L, Flachskampf FA, Foster E, Goldstein SA, Kuznetsova T, et al. Recommendations for cardiac chamber quantification by echocardiography in adults: an update from the American Society of Echocardiography and the European Association of Cardiovascular Imaging. *J Am Soc Echocardiogr*. 2015;28(1):1–39.e14.
- Wang S, Song K, Guo X, Xue H, Wang N, Chen J, Zou Y, Sun K, Wang H, He J, et al. The association of metabolic syndrome with left ventricular mass and geometry in community-based hypertensive patients among Han Chinese. *J Res Med Sci*. 2015;20(10):963–8.
- Chahal NS, Lim TK, Jain P, Chambers JC, Kooner JS, Senior R. New insights into the relationship of left ventricular geometry and left ventricular mass with cardiac function: a population study of hypertensive subjects. *Eur Heart J*. 2010;31(5):588–94.
- Huis AE, de Man FS, van Rossum AC, Handoko ML. How to diagnose heart failure with preserved ejection fraction: the value of invasive stress testing. *Netherlands Heart J*. 2016;24(4):244–51.
- Folland ED, Parisi AF, Moynihan PF, Jones DR, Feldman CL, Tow DE. Assessment of left ventricular ejection fraction and volumes by real-time, two-dimensional echocardiography. A comparison of cineangiographic and radionuclide techniques. *Circulation*. 1979;60(4):760–6.
- Nagueh SF, Smiseth OA, Appleton CP, Byrd BF 3rd, Dokainish H, Edvardsen T, Flachskampf FA, Gillebert TC, Klein AL, Lancellotti P, et al. Recommendations for the evaluation of left ventricular diastolic function by echocardiography: an update from the American Society of Echocardiography and the European Association of Cardiovascular Imaging. *J Am Soc Echocardiogr*. 2016;29(4):277–314.
- Zoghbi WA, Adams D, Bonow RO, Enriquez-Sarano M, Foster E, Grayburn PA, Hahn RT, Han Y, Hung J, Lang RM, et al. Recommendations for noninvasive evaluation of native Valvular regurgitation: a report from the American Society of Echocardiography developed in collaboration with the Society for Cardiovascular Magnetic Resonance. *J Am Soc Echocardiogr*. 2017;30(4):303–71.
- Peterson PN, Rumsfeld JS, Liang L, Albert NM, Hernandez AF, Peterson ED, Fonarow GC, Masoudi FA. American Heart Association get with the guidelines-heart failure P: a validated risk score for in-hospital mortality in patients with heart failure from the American Heart Association get with the guidelines program. *Circ Cardiovasc Qual Outcomes*. 2010;3(1):25–32.

14. Shiraishi Y, Kohsaka S, Abe T, Mizuno A, Goda A, Izumi Y, Yagawa M, Akita K, Sawano M, Inohara T, et al. Validation of the get with the guideline-heart failure risk score in Japanese patients and the potential improvement of its discrimination ability by the inclusion of B-type natriuretic peptide level. *Am Heart J*. 2016;171(1):33–9.
15. Kamarudin AN, Cox T, Kolamunnage-Dona R. Time-dependent ROC curve analysis in medical research: current methods and applications. *BMC Med Res Methodol*. 2017;17(1):53.
16. Giavarina D. Understanding bland Altman analysis. *Biochem Med (Zagreb)*. 2015;25(2):141–51.
17. Koo TK, Li MY. A guideline of selecting and reporting Intraclass correlation coefficients for reliability research. *J Chiropr Med*. 2016;15(2):155–63.
18. van Veldhuisen DJ, Linssen GC, Jaarsma T, van Gilst WH, Hoes AW, Tijssen JG, Paulus WJ, Voors AA, Hillege HL. B-type natriuretic peptide and prognosis in heart failure patients with preserved and reduced ejection fraction. *J Am Coll Cardiol*. 2013;61(14):1498–506.
19. Kitada S, Kikuchi S, Tsujino T, Masuyama T, Ohte N. The prognostic value of brain natriuretic peptide in patients with heart failure and left ventricular ejection fraction higher than 60%: a sub-analysis of the J-MELODIC study. *ESC Heart Fail*. 2018;5(1):36–45.

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