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Association between preoperative serum sodium and postoperative 30-day mortality in adult patients with tumor craniotomy

Yufei Liu^{1,2,3†}, Haofei Hu^{3,4†}, Zongyang Li^{1,3†}, Yuandi Yang^{1,3}, Fanfan Chen^{1,3}, Weiping Li^{1,3}, Liwei Zhang² and Guodong Huang^{1,3*}

Abstract

Background Limited data exist regarding preoperative serum sodium (Na) and 30-day mortality in adult patients with tumor craniotomy. Therefore, this study investigates their relationship.

Methods A secondary retrospective analysis was performed using data from the ACS NSQIP database (2012–2015). The principal exposure was preoperative Na. The outcome measure was 30-day postoperative mortality. Binary logistic regression modeling was conducted to explore the link between them, and a generalized additive model and smooth curve fitting were applied to evaluate the potential association and its explicit curve shape. We also conducted sensitivity analyses and subgroup analyses.

Results A total of 17,844 patients (47.59% male) were included in our analysis. The mean preoperative Na was 138.63 ± 3.23 mmol/L. The 30-day mortality was 2.54% (455/17,844). After adjusting for covariates, we found that preoperative Na was negative associated with 30-day mortality. (OR=0.967, 95% CI:0.941, 0.994). For patients with Na ≤ 140, each increase Na was related to a 7.1% decreased 30-day mortality (OR=0.929, 95% CI:0.898, 0.961); for cases with Na > 140, each increased Na unit was related to a 8.8% increase 30-day mortality (OR=1.088, 95% CI:1.019, 1.162). The sensitivity analysis and subgroup analysis indicated that the results were robust.

Conclusions This study shows a positive and nonlinear association between preoperative Na and postoperative 30-day mortality in adult patients with tumor craniotomy. Appropriate preoperative Na management and maintenance of serum Na near the inflection point (140) may reduce 30-day mortality.

Keywords Sodium, Brain tumor, Craniotomy, Nonlinear, Mortality

[†]Yufei Liu, Haofei Hu and Zongyang Li contributed equally.

*Correspondence:

Guodong Huang

huangguodong@email.szu.edu.cn

¹Department of Neurosurgery, Shenzhen Key Laboratory of Neurosurgery, Shenzhen Second People's Hospital, the First Affiliated Hospital of Shenzhen University, No. 3002 Sungang west Road, Futian District, Shenzhen, Guangdong Province 518035, China





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Background

Craniotomy is a basic surgical procedure for intracranial tumor resection. Unfortunately, craniotomies for intracranial tumors bring significant risks of all kinds of complications, as well as mortality [1-3]. The 30-day postoperative mortality, also known as postoperative 30-day mortality, provides an efficient and objective evaluation of the patient's access to surgery and determining the safety of the operation and anesthesia [4, 5]. Postoperative 30-day mortality ranged from 0.95 to 8.62% based on a study of English patients with a brain tumor who underwent a craniotomy in 2008–2010 [6]. A study from a US National Surgical Quality Improvement Program (NSQIP) database reported a 30-day mortality rate of 2.6% after craniotomy for primary brain malignancies [3]. In addition, Hankinson TC et al. maintained that the 30-day postoperative mortality of diagnostic neurosurgery for a primary childhood intracranial tumor was between 1.16% and 1.72%, and the mortality of a U.S. pediatric population was in line with contemporary data from Europe [5].

According to the serum sodium (Na) levels, patients can be divided clinically into hypernatremia (Na≥146 mmol/L), normal serum sodium level (Na \geq 135 mmol/L, <146 mmol/L) and hyponatremia (Na<135 mmol/L) groups [7–9]. Abnormal serum Na levels have been shown to be associated with unfavorable outcomes and even increased mortality in patients admitted to the ICU after surgery [10], patients after cardiac surgery [9], patients with open abdominal surgery [11], patients with liver transplantation [12], and patients with various cancers, including pediatric brain tumors [13], head and neck squamous cell carcinoma [14], esophageal cancer [15], lung carcinoma [16], gastric malignancy [17], ovarian cancer [18], and upper tract urothelial carcinoma [19]. Plasma Na concentration disturbances expose cells to hypertonic or hypotonic stress. Although all cells of the body are affected, the clinical manifestations of hypernatremia and hyponatremia are predominantly neurologic, and rapid changes in Na concentrations in any orientation can cause permanent, severe, and occasionally lethal brain damage [20]. Hyponatremia and hypernatremia are common in some patients after neurosurgery, especially patients with sellar area lesions and thalamic lesions [21–23]. A correct diagnosis and positive management of sodium-related cerebral salt-wasting syndrome and syndrome of inappropriate secretion of antidiuretic hormone are important in the practice of dealing with neurosurgical patients [21, 24]. Accumulating studies have focused on the link between preoperative serum Na levels and postoperative prognosis [11, 14, 17, 18, 25–28].

To date, few studies have investigated the nonlinear relationship and conducted subgroup analyses between

preoperative serum Na levels and 30-day postoperative mortality in cancer patients. The relationship between preoperative serum Na levels and postoperative 30-day mortality has yet to be investigated in tumor craniotomy patients older than 18 years of age. Thus, this study aimed to examine their relationship from data from the American College of Surgeons NSQIP (ACS NSQIP) database. These studies may provide clinicians with the understanding of the role of preoperative sodium in assessing short-term postoperative outcomes of tumor craniotomy patients.

Participants and methods

Study design

We utilized data from the ACS NSQIP database (2012–2015) for this cross-sectional study. In our study, the independent and dependent variables were serum Na before the operation and postoperative 30-day mortality, respectively.

Data source

Jingwen Zhang et al., the authors of "Sepsis and septic shock after craniotomy: Predicting a significant patient safety and quality outcome measure [29], originally uploaded [CSV] data named [S1 Data] from the ACS NSQIP database (2012–2015). The original study was published as an open access article under the Creative Commons Attribution License, which permits unrestricted distribution, use, and reproduction. Therefore, our study could use the uploaded database for secondary analysis of nonviolation of the rights of authors.

Participants

The original research included 18,642 adult participants with brain tumors. A total of 17,844 participants were included in our analysis after excluding patients with missing values for preoperative serum Na (N=798) (Fig. 1). Because our research was a secondary analysis of historical data and the original personal information was not identified, consent forms were not needed.

Variables

Initially in the uploaded data, serum Na before surgery was recorded as a continuous variable [29]. Data collected for Na was obtained by the Current Procedural Terminology code.

Thirty-day postoperative mortality

Originally, the authors defined 30-day postoperative mortality as mortality within the first 30 days following a surgical procedure [29]. The NSQIP tracked mortality for the first 30 postoperative days.



Fig. 1 Flowchart of the study participants

Covariates

Our choice of covariates was based on previous literature reports and our clinical experience. Weight (kg) divided by height (m) squared (kg/m^2) was used to calculate body mass index (BMI). Standard conditions were followed when collecting and treating the original data. The continuous variables in this study were as follows: indicators of preoperative blood tests (blood urea nitrogen (BUN), white blood cell (WBC) count, creatine (Cr), hematocrit (HCT), platelet (PLT) count) and BMI; and categorical variables were as follows: sex (male or female), age ranges (year-old) (18-40, 41-60, 60-80, or >80), race (White, Asian, African American or unknown race), a history of hypertension (No, Yes), diabetes (No, Yes (noninsulindependent) or Yes (insulin-dependent)), smoking status (No, Yes), severe chronic obstructive pulmonary disease (COPD) (No, Yes), congestive heart failure (CHF) (No, Yes), renal failure (No, Yes), dialysis (No, Yes), steroid use for a chronic condition (No, Yes), disseminated cancer (No, Yes) and preoperative systemic sepsis (No, systemic inflammatory response syndrome (SIRS) or septic shock). The authors provided more details in the original study [29].

Statistical analysis

The number and proportion of patients with missing values for BMI, BUN, Cr, WBC, HCT and PLT were 709 (3.97%), 798 (4.47%), 139 (0.78%), 198 (1.11%), 113 (0.63%), and 193 (1.08%), respectively. The median or mean value replaced the missing values for statistical analysis.

According to the serum Na levels, patients were divided clinically into three groups: hyponatremia

(Na < 135 mmol/L), normal serum sodium level (Na \ge 135 mmol/L, <146 mmol/L) and hypernatremia (Na \geq 146) groups. We reported the mean±standard deviation (SD) (normally distributed variables) or median (interquartile range) (nonnormally distributed variables) for continuous variables; we extracted percentages and frequencies for categorical outcomes. χ^2 tests (for categorical variables), Kruskal-Wallis H tests or one-way ANOVAs were conducted to test for statistically significant differences among the three serum Na groups. To explore the relationship between serum Na level and postoperative 30-day mortality, distinctive univariate and multivariate binary logistic regression models were established complying with the Strengthening the Reporting of Observational Studies in Epidemiology (STROBE) statement guidelines. The models included a nonadjusted model (not adjusted for any covariates), a minimally adjusted model (only adjusted for sex and age range), and a fully adjusted model (adjusted for those variables that were significant in univariate analysis including serum WBC, BUN, Cr, PLT and HCT, sex, age range, renal failure, diabetes, severe COPD, CHF, hypertension, dialysis, disseminated cancer, preoperative systemic sepsis, steroid use) for the covariates presented in Table 1. We recorded 95% confidence intervals and effect sizes and adjusted the effect sizes when covariates were added to the model and if the OR changed by 10% or more [30].

We conducted a generalized additive model (GAM) and smooth curve fitting to address the specific link between serum Na level and postoperative 30-day mortality. If nonlinearity was detected, the inflection point was calculated using a recursive algorithm, and a piecewise binary logistic regression model with one piece on each side of the inflection point was subsequently constructed. We employed the log-likelihood ratio test to determine the most suitable model for interpreting the link between preoperative serum Na level and postoperative 30-day mortality.

We conducted subgroup analyses using a stratified binary logistic regression model for the various subgroups, including sex, a history of severe COPD, hypertension, metastatic cancer and steroid use. First, we adjusted each stratification for the factors (sex, renal failure, serum Na, BUN, Cr, WBC count, PLT count, age range, steroid use, diabetes, disseminated cancer, CHF, severe COPD, hypertension, dialysis, preoperative systemic sepsis) in addition to the stratification factor itself. Second, we tested for interactions by using the likelihood ratio test for models with and without interaction terms [31, 32].

We performed sensitivity analyses to verify the robustness of our findings. We converted serum Na level into a categorical variable based on the definitions of hyponatremia and hypernatremia. We also investigated the potential for unknown confounders of the link between serum Na level and 30-day postoperative mortality by calculating E-values [33]. All results were reported in compliance with the STROBE statement guidelines. We followed the methods of Liu et al. [34].

Statistical analyses were performed using the statistical software EmpowerStats (http://www.empowerstats.com, X&Y Solutions, Inc., Boston, MA) and the R package (http://www.R-project.org, The R Foundation). For all the tests, a two-sided p value <0.05 was considered statistically significant.

Results

Study population

Demographic and clinical characteristics of the study participants are provided in Table 1. A total of 17,844 patients (47.59% male) were included in our analysis. The age (years old) distribution proportions were 15.89% (18-40), 41.54% (41-60), 39.23% (61-80) and 3.33% (>81). The mean preoperative Na level was 138.63 ± 3.23 (mmol/L). The postoperative 30-day mortality of the included patients was 2.54% (454/17,844). Compared with those of the participants with hypernatremia, the hyponatremia group was significantly positively associated with BMI, sex, preoperative serum BUN, WBC, Cr, HCT and PLT count, race, age range, diabetes, disseminated cancer, smoking status, severe COPD, CHF, hypertension, dialysis, steroid use and preoperative systemic sepsis (all P values<0.05), except renal failure (P values>0.05). These results may be due to the sample size.

The results of the univariate analyses using a binary logistic regression model

The univariate analysis was performed and indicated that the following significant variables in patients were positively associated with postoperative 30-day mortality: female (OR=0.652, 95% CI: 0.540 to 0.788), 41-60 years old (OR=2.730, 95% CI: 1.706 to 4.370), 61-80 vears old (OR=4.933, 95% CI: (3.119 to 7.803), >81 years old (OR=14.629, 95% CI: 8.708 to 24.576), had diabetes (Noninsulin-dependent) (OR=1.538, 95% CI: 1.129 to 2.096), diabetes (Insulin-dependent) (OR=2.587, 95% CI: 1.881 to 3.557), severe COPD (OR=2.496, 95% CI: 1.830 to 3.405), CHF (OR=7.157, 95% CI: 3.494 to 14.661), hypertension (OR=2.246, 95% CI: 1.859 to 2.713), renal failure (OR=10.509, 95% CI: 2.922 to 37.800), dialysis (OR=6.348, 95% CI: 2.989 to 13.483), disseminated cancer (OR=2.835, 95% CI: 2.347 to 3.425), steroid use (OR=2.278, 95% CI: 1.850 to 2.806), preoperative SIRS (OR=2.557, 95% CI: 1.824 to 3.585), preoperative septic shock (OR=9.019, 95% CI: 1.943 to 41.874), hyponatremia group (OR=2.372, 95% CI: 1.868 to 3.012), hypernatremia group (OR=2.500, 95% CI: 1.349 to 4.631), BUN (OR=1.043, 95% CI: 1.036 to 1.050), WBC (OR=1.072, 95% CI: 1.055 to 1.089), PLT (OR=0.998, 95% CI: 0.997 to 0.999), HCT (OR=0.924, 95% CI: 0.909 to 0.940), BUN (OR=1.043, 95% CI: 1.035 to 1.050), and Cr (OR=1.241, 95% CI: 1.114 to 1.383). By comparison, participants who were Asian (OR=0.757, 95% CI: 0.401 to 1.429), African American (OR=0.878, 95% CI: 0.590 to 1.306), of unknown race (OR=1.152, 95% CI: 0.916 to 1.450) and those who smoked (OR=1.130, 95% CI: 0.900 to 1.419) or had an abnormal BMI (OR=0.987, 95% CI: 0.972 to 1.002) were not associated with postoperative 30-day mortality (Table 2).

The results of the multivariate analyses using the binary logistic regression model

To explore the link between preoperative serum Na level and postoperative 30-day mortality, we established three models by binary logistic regression models. We discovered a significant link between preoperative serum Na and postoperative 30-day mortality in the three models (Table 3). In the nonadjusted model, we found that a increase of 1 mmol/L in preoperative Na was related to an 8.9% decrease in postoperative 30-day mortality (OR=0.911, 95% CI:0.888 to 0.936), which was statistically significant. In the minimally adjusted model, when we only adjusted for sex and age ranges, each elevation in the preoperative Na unit may result in a 7.3% decrease in the postoperative 30-day mortality (OR=0.927, 95% CI:0.902 to 0.952). The findings on the link between preoperative serum Na level and postoperative 30-day mortality obtained from the model were statistically significant. In the fully adjusted model, a increase of 1

Table 1 The baseline characteristics of the participants

Na (mmol/L)	Hypernatremia group (Na ≥ 146)	Normal serum Na group (Na≥135, <146)	Hyponatremia group (Na < 135)	P value
N	204	15,925	1715	
BMI (Mean±SD)	29.105±6.701	28.813±6.644	27.892±6.371	< 0.001
BUN (Median (Q1-Q3))	17.000 (14.000-22.500)	16.000 (12.000-20.000)	18.000 (14.000-25.000)	< 0.001
WBC (Median (Q1-Q3))	7.700 (5.700-10.310)	8.400 (6.400-11.300)	10.900 (7.800–14.300)	< 0.001
Cr (Median (Q1-Q3))	0.921 ± 0.383	0.859±0.419	0.890 ± 0.677	0.004
HCT(Mean±SD)	37.969 ± 5.967	40.389±4.718	39.653±5.320	< 0.001
PLT (Median (Q1-Q3))	217.500 (166.000-268.250)	237.000 (195.000-284.000)	244.000 (194.000-297.000)	< 0.001
Sex, N (%)				< 0.001
Male	95 (46.569%)	7464 (46.870%)	933 (54.402%)	
Female	109 (53.431%)	8461 (53.130%)	782 (45.598%)	
Age ranges (years), N (%)				< 0.001
18–40	26 (12.745%)	2660 (16.703%)	150 (8.746%)	
41–60	79 (38.725%)	6719 (42.192%)	614 (35.802%)	
61–80	90 (44.118%)	6043 (37.947%)	868 (50.612%)	
>81	9 (4.412%)	503 (3.159%)	83 (4.840%)	
Race, N (%)				< 0.001
White	143 (70.098%)	11,396 (71.560%)	1162 (67.755%)	
Asian	10 (4.902%)	467 (2.932%)	46 (2.682%)	
African American	25 (12.255%)	1089 (6.838%)	107 (6.239%)	
Unknown race	26 (12.745%)	2973 (18.669%)	400 (23.324%)	
Hypertension, N (%)				< 0.001
No	109 (53.431%)	10,009 (62.851%)	815 (47.522%)	
Yes	95 (46.569%)	5916 (37.149%)	900 (52.478%)	
Diabetes, N (%)				< 0.001
No	171 (83.824%)	14,165 (88.948%)	1382 (80.583%)	
Yes (Noninsulin)	21 (10.294%)	1129 (7.089%)	193 (11.254%)	
Yes (Insulin)	12 (5.882%)	631 (3.962%)	140 (8.163%)	
Smoking status, N (%)	X ,		, , , , , , , , , , , , , , , , , , ,	< 0.001
No	167 (81.863%)	12,893 (80.961%)	1316 (76.735%)	
Yes	37 (18.137%)	3032 (19.039%)	399 (23.265%)	
Severe COPD, N (%)				< 0.001
No	193 (94.608%)	15,266 (95.862%)	1569 (91.487%)	
Yes	11 (5.392%)	659 (4.138%)	146 (8.513%)	
Congestive heart failure, N (%)				< 0.001
No	201 (98.529%)	15.885 (99.749%)	1700 (99.125%)	
Yes	3 (1.471%)	40 (0.251%)	15 (0.875%)	
Renal failure, N (%)				0.086
Νο	203 (99.510%)	15.914 (99.931%)	1713 (99.883%)	
Yes	1 (0.490%)	11 (0.069%)	2 (0.117%)	
Dialvsis, N (%)		(,		< 0.001
No	203 (99.510%)	15.887 (99.761%)	1697 (98.950%)	
Yes	1 (0.490%)	38 (0.239%)	18 (1.050%)	
Steroid use for a chronic condition. N (%)				< 0.001
No	175 (85 784%)	13 676 (85 878%)	1284 (74 869%)	(0.00)
Yes	29 (14,216%)	2249 (14.122%)	431 (25.131%)	
Disseminated cancer N (%)				< 0.001
Νο	174 (85,294%)	12.641 (79.378%)	1104 (64.373%)	
Yes	30 (14,706%)	3284 (20,622%)	611 (35.627%)	
Preoperative systemic sensis, N (%)				< 0.001
No	187 (91.667%)	15.375 (96.546%)	1613 (94.052%)	
SIRS	16 (7 843%)	541 (3 397%)	101 (5 889%)	
Sentic Shock	1 (0.490%)	9 (0.057%)	1 (0.058%)	

|--|

Na (mmol/L)	Hypernatremia group (Na≥146)	Normal serum Na group (Na≥135, <146)	Hyponatremia group (Na < 135)	P value
30-day mortality events, N (%)				< 0.001
No	193 (94.608%)	15,570 (97.771%)	1627 (94.869%)	
Yes	11 (5.392%)	355 (2.229%)	88 (5.131%)	

mmol/L of preoperative Na was accompanied by a 3.3% decrease in 30-day postoperative mortality (OR=0.967, 95% CI:0.941 to 0.994). The distribution of CIs suggested that the relationship between the Na level and the postoperative 30-day mortality obtained by the model was reliable (Table 3).

The nonlinearity addressed by the generalized additive model

We found that the relationship between preoperative serum Na and postoperative 30-day mortality was nonlinear through the GAM and smooth curve fitting (as shown in Fig. 2). Hence, we fit the data to a piecewise binary logistic regression model that allowed two different slopes and a standard binary logistic regression model based on the sensitivity analysis. The best fit model was chosen through the log-likelihood ratio test (as shown in Table 4). The P value for the log-likelihood ratio test was less than 0.05 in our analysis. Therefore, a piecewise model was used to fit the association between preoperative serum Na and postoperative 30-day mortality. With a recursive algorithm, we obtained an inflection point of 140 for the first time and computed the effect sizes (OR) and confidence intervals to the right and left of the inflection point with the piecewise binary logistic regression model. The effect size (OR) was 0.929, and the 95% CI was from 0.898 to 0.961 on the left side of the inflection point. The OR was 1.088, and the 95% CI was from 1.019 to 1.162 on the right side of the inflection point.

Subgroup analyses

We conducted subgroup analyses in view of other factors that might influence the link between preoperative serum Na level and postoperative 30-day mortality. We employed sex, severe COPD, hypertension, steroid use and disseminated cancer as stratification variables to test the trend of effect sizes. The results revealed that all listed subgroups did not affect the link between serum Na level and postoperative 30-day mortality (P for interaction >0.05) (as shown in Table 5).

Sensitivity analyses

We conducted sensitivity analyses to test the robustness of our results. We first converted the preoperative Na value from a continuous variable to a categorical variable (dividing the patients into three groups based on hyponatremia and hypernatremia) and then changed the previous Na variable in the model to the categoricaltransformed Na. The trend of the effect sizes in the different groups was equidistant after serum Na was converted to a categorical variable. Second, we performed a GAM to insert the continuity covariate into the equation as a curve. In addition, to assess the sensitivity to unmeasured confounders, we also calculated an E-value and obtained 1.21, which was greater than the relative risk of unmeasured confounders influencing the link between serum Na level and postoperative 30-day mortality, indicating that unmeasured or unknown confounders had little effect on the link between them.

Discussion

Numerous studies have focused on the relationship between postoperative sodium disturbance and the outcome of brain tumor surgery. However, few studies have concentrated on the link between preoperative serum Na levels and the outcomes of brain tumors. Muhlestein WE et al. suggested that preoperative sodium abnormality was one of the strongest variables that influenced inpatient length of stay after brain tumor surgery [35]. In addition, Sorba EL et al. maintained that lower preoperative serum Na was an independent risk factor for SIADH after pituitary surgery [36]. These studies showed that preoperative abnormal blood Na levels may be a potential indicator for patient prognosis.

The authors used the ACS NSQIP database (2012-2015) to investigate whether there was a link between preoperative serum Na level and postoperative 30-day mortality in the present study. To the best of our knowledge, the present data analysis is the first to observe a nonlinear link between them in tumor craniotomy patients. The authors used a two-piecewise linear regression model to elucidate a nonlinear link between the serum Na level and 30-day postoperative mortality and first obtained the inflection point (140) of Na after adjusting for confounders. When the preoperative serum Na was below 140 mmol/L, a increase of 1 mmol/L of Na was related to a 7.1% decrease in postoperative 30-day mortality (OR=0.929, 95% CI:0.898, 0.961). However, when preoperative serum Na was >140 mmol/L, each increase in Na unit was related to a 8.8% increase in postoperative 30-day mortality (OR=1.088, 95% CI:1.019, 1.162). Hyponatremia was associated with significant morbidity and mortality in a broad variety of malignant tumors [37], and hypernatremia at admission could be a

Table 2 The results of the univariate analysis

	Statistics	OR	95% CI	P value
Sex, N (%)				
Male	8492 (47.590%)	Ref.		
Female	9352 (52.410%)	0.652	(0.540, 0.788)	< 0.00001
Race, N (%)				
White	12,701 (71.178%)	Ref.		
Asian	523 (2.931%)	0.757	(0.401, 1.429)	0.38976
African American	1221 (6.843%)	0.878	(0.590, 1.306)	0.51996
Unknown race	3399 (19.048%)	1.152	(0.916, 1.450)	0.22611
Age range, N (%)				
18–40	2836 (15.893%)	Ref.		
41–60	7412 (41.538%)	2.730	(1.706, 4.370)	0.00003
61–80	7001 (39.234%)	4.933	(3.119, 7.803)	< 0.00001
> 81	595 (3.334%)	14.629	(8.708, 24.576)	< 0.00001
Diabetes, N (%)				
No	15,718 (88.086%)	Ref.		
Yes (Noninsulin-dependent)	1343 (7.526%)	1.538	(1.129, 2.096)	0.00632
Yes (Insulin-dependent)	783 (4.388%)	2.587	(1.881, 3.557)	< 0.00001
Smoking status, N (%)				
No	14,376 (80.565%)	Ref.		
Yes	3468 (19.435%)	1.130	(0.900, 1.419)	0.29262
Severe COPD, N (%)				
No	17,028 (95.427%)	Ref.	(4.000.0.405)	
Yes	816 (4.573%)	2.496	(1.830, 3.405)	< 0.00001
CHF, N (%)	17 706 (00 6750)			
No	17,786 (99.675%)	Ket.	(2.404.14((1))	.0.00001
Yes	58 (0.325%)	/.15/	(3.494, 14.661)	< 0.00001
Appertension, N (%)	10,022 (61,2700()	Dof		
NO	(011 (28 720%)	Rei.		< 0.00001
res Repel failure N/0()	0911 (38.730%)	2.240	(1.859, 2.713)	< 0.00001
No	17 820 (00 02204)	Pof		
Vos	14 (0.078%)	10 500	(2022 37 800)	0.0003.2
Dialysis	14 (0.07070)	10.505	(2.922, 37.000)	0.00032
No	17 787 (99 681%)	Ref		
Yes	57 (0.319%)	6.348	(2.989, 13,483)	< 0.00001
Disseminated cancer, N (%)			()	
Νο	13.919 (78.004%)	Ref.		
Yes	3925 (21.996%)	2.835	(2.347, 3.425)	< 0.00001
Steroid use, N (%)				
No	15,135 (84.818%)	Ref.		
Yes	2709 (15.182%)	2.278	(1.850, 2.806)	< 0.00001
Preoperative systemic sepsis N (%)				
No	17,175 (96.251%)	Ref.		
SIRS	658 (3.688%)	2.557	(1.824, 3.585)	< 0.00001
Septic Shock	11 (0.062%)	9.019	(1.943, 41.874)	0.00499
Na				
Normal serum Na group (Na≥135, <146)	15,925 (89.246%)	Ref.		
Hyponatremia group (Na < 135)	1715 (9.611%)	2.372	(1.868, 3.012)	< 0.00001
Hypernatremia group (Na≥146)	204 (1.143%)	2.500	(1.349, 4.631)	0.00359
WBC	9.524 ± 4.460	1.072	(1.055, 1.089)	< 0.00001
нст	40.291 ± 4.806	0.924	(0.909, 0.940)	< 0.00001
PLT	244.222 ± 76.614	0.998	(0.997, 0.999)	0.00144
BUN	17.348 ± 8.155	1.043	(1.035, 1.050)	< 0.00001
Cr	0.863 ± 0.450	1.241	(1.114, 1.383)	0.00009
BMI	28.728 ± 6.624	0.987	(0.972, 1.002)	0.08094

Table 3 The results of the multivariate analysis

Exposure	Model 1 (OR, 95% Cl,P)	Model 2 (OR, 95% Cl,P)	Model 3 (OR, 95% Cl,P)
Na	0.911 (0.888, 0.936) < 0.00001	0.927 (0.902, 0.952) < 0.00001	0.967 (0.941, 0.994) 0.01588
Na groups			
Hypernatremia group (Na≥146)	Ref	Ref	Ref
Normal serum Na group (Na≥135, <146)	0.400 (0.216, 0.741) 0.00359	0.431 (0.231, 0.803) 0.00805	0.621 (0.315, 1.224) 0.16846
Hyponatremia group (Na < 135)	0.949 (0.498, 1.808) 0.87345	0.864 (0.451, 1.658) 0.66079	0.831 (0.409, 1.687) 0.60853

Model 1 (nonadjusted model): not adjusted for any covariate. Model 2 (minimally adjusted model): adjusted for sex and age range. Model 3 (fully adjusted model): adjusted for sex, age range, diabetes, serum WBC, HCT, PLT, BUN and Cr, severe COPD, CHF, hypertension, renal failure, dialysis, steroid use for chronic condition, disseminated cancer, and preoperative systemic sepsis. OR, odds ratio; CI: confidence, Ref: reference



Fig. 2 The nonlinear relationship between preoperative serum Na and 30-day postoperative mortality

meaningful indicator of poor outcomes in patients with terminal cancer [38]. These factors may explain the nonlinear association between serum Na level and postoperative 30-day mortality in our study. This is the first study to provide an inflection point as a reference for the management of preoperative serum Na in tumor craniotomy patients in the U.S. population. Hence, these results have prominent clinical value.

Our study has the following strengths. It is the large sample size that allows such analysis. Most covariates

Table 4	The results	of the	standard	linear	and	piecewise	linear
rearessic	n models						

	Postoperative 30-day mor- tality (OR, 95% Cl, P value)
Fitting model by standard linear regression	0.967 (0.941,
	0.994) 0.0159
Fitting model by two-piecewise linear regression	
Infection point of Na	140
≤ 140	0.929 (0.898,
	0.961) < 0.0001
> 140	1.088 (1.019,
	1.162) 0.0118
P value for the log likelihood ratio test	< 0.001

The model adjusted for sex, age range, diabetes, serum WBC, HCT, PLT, BUN and Cr, severe COPD, CHF, hypertension, renal failure, dialysis, steroid use, disseminated cancer, and preoperative systemic sepsis. OR, odds ratio; CI: confidence, Ref: reference

Table 5	Results of the	interaction anal	ysis and	subgroup	analysis

Characteristic	No of participants	OR (95% CI)	P for in- teraction
Sex			0.5130
Male	8492	0.960 (0.927, 0.994)	
Female	9352	0.977 (0.937, 1.020)	
Severe COPD			0.9466
No	17,028	0.965 (0.937, 0.994)	
Yes	816	0.962 (0.884, 1.047)	
Hypertension			0.4143
No	10,933	0.954 (0.913, 0.996)	
Yes	6911	0.976 (0.943, 1.011)	
Disseminated cancer			0.4315
No	13,919	0.974 (0.940, 1.010)	
Yes	3925	0.953 (0.915, 0.993)	
Steroid use for a chronic condition			0.7373
No	15,135	0.969 (0.938, 1.002)	
Yes	2709	0.960 (0.914, 1.007)	

have complete information, with few missing. To the best of our knowledge, this is the first study to observe the association between preoperative serum Na levels and 30-day postoperative mortality in adult patients with tumor craniotomy. We found a nonlinear relationship between them, so our study has greater clinical significance, which previous studies have not investigated. We employed an interaction analysis and subgroup analysis. We used sensitivity analyses to test the robustness of our results.

Our research has the following shortcomings: this was a retrospective study on a national database; because our research was a secondary analysis of previously published data, we could not exclude some residual and/or unmeasured confounding factors that could interfere with the estimated link (e.g., pharmacological treatments, dietary habits, socioeconomic factors, pathological classification, perioperative fluid therapy, characteristics of benign and malignant tumors, and different types of brain tumors). The database used in this study lacked preoperative blood potassium and chlorine data. However, we computed the E-value to quantify the potential influence of unmeasured confounders, and we could not explore the link between preoperative serum Na level and long-term prognosis. Despite the above limitations, this research used collected data from a heterogeneous and large group of brain tumor patients. Therefore, the association and results postulated remain highly plausible. Besides, the database did not provide information about the clear cause of death. Future studies should collect those mentioned possible confounders and data that were not included in this study.

Conclusions

This study proves a positive and nonlinear association between preoperative serum Na level and postoperative 30-day mortality in adult patients with tumor craniotomy and a threshold effect between them. These findings are expected to provide clinicians (especially neurosurgeons) with a suitable reference for preoperative serum Na. The use of the sodium inflection point of 140 mmol/L and proper preoperative management could reduce 30-day postoperative mortality after surgery in this population. Consequently, abnormal preoperative serum Na signals may assist clinicians in identifying high-risk groups for 30-day postoperative mortality in this particular population, which will facilitate clinicians planning and initiating appropriate management strategies.

Abbreviations

Na	sodium
HCT	hematocrit
SD	standard deviation
CI	confidence interval
OR	odds ratio
ACS NSQIP	American College of Surgeons National Surgical Quality
	Improvement Program
BMI	body mass index
COPD	chronic obstructive pulmonary disease
GAM	generalized additive model
CHF	congestive heart failure
BUN	blood urea nitrogen
Cr	creatine
WBC	white blood cell count
PLT	platelet count

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Author contributions

All authors contributed to the study conception and design. Material preparation, data collection and analysis were performed by Guodong Huang, Weiping Li and Liwei Zhang. Yufei Liu, Haofei Hu and Zongyang Li wrote the main manuscript text and all authors commented on previous versions of the manuscript. Fanfan Chen and Yuandi Yang were responsible for data verification. All authors reviewed the manuscript.

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Data Availability

Anyone can download the raw data from https://journals.plos.org/plosone/article?id=https://doi.org/10.1371/journal.pone.0235273.

Declarations

Ethics approval and consent to participate

Patient data were anonymous, and we analyzed previously collected data. Therefore, informed consent was not necessary. Our research was exempted from the Clinical Research Ethics Committee of Shenzhen Second People's Hospital due to the nature of the database (No. 20220407005).

Consent for publication

Not Applicable (NA)

Declarations of interest

The authors have no relevant financial or nonfinancial interests to disclose.

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References

- Lonjaret L, Guyonnet M, Berard E, Vironneau M, Peres F, Sacrista S, et al. Fourcade and T. Geeraerts: postoperative complications after craniotomy for brain tumor surgery. Anaesth Crit Care Pain Med. 2017;36(4):213–8. https:// doi.org/10.1016/j.accpm.2016.06.012.
- Noordzij PG, Poldermans D, Schouten O, Bax JJ, Schreiner FA, Boersma E. Postoperative mortality in the Netherlands: a population-based analysis of surgery-specific risk in adults. Anesthesiology. 2010;112(5):1105–15. https:// doi.org/10.1097/ALN.0b013e3181d5f95c.
- Senders JT, Muskens IS, Cote DJ, Goldhaber NH, Dawood HY, Gormley WB, et al. Thirty-day outcomes after craniotomy for primary malignant brain tumors: a National Surgical Quality Improvement Program Analysis. Neurosurgery. 2018;83(6):1249–59. https://doi.org/10.1093/neuros/nyy001.
- Watters DA, Hollands MJ, Gruen RL, Maoate K, Perndt H, McDougall RJ, et al. Perioperative mortality rate (POMR): a global indicator of access to safe surgery and anaesthesia. World J Surg. 2015;39(4):856–64. https://doi. org/10.1007/s00268-014-2638-4.
- Hankinson TC, Dudley RW, Torok MR, Patibandla MR, Dorris K, Poonia S, et al. Short-term mortality following surgical procedures for the diagnosis of pediatric brain tumors: outcome analysis in 5533 children from SEER, 2004–2011. J Neurosurg Pediatr. 2016;17(3):289–97. https://doi.org/10.3171/2015.7.P EDS15224.
- Williams M, Treasure P, Greenberg D, Brodbelt A, Collins P. Surgeon volume and 30 day mortality for brain tumours in England. Br J Cancer. 2016;115(11):1379–82. https://doi.org/10.1038/bjc.2016.317.
- Dmitrieva NI, Liu D, Wu CO, Boehm M. Middle age serum sodium levels in the upper part of normal range and risk of heart failure. Eur Heart J. 2022. https:// doi.org/10.1093/eurheartj/ehac138.

- Alukal JJ, John S, Thuluvath PJ. Hyponatremia in cirrhosis: an update. Am J Gastroenterol. 2020;115(11):1775–85. https://doi.org/10.14309/ ajg.000000000000786.
- Khan FW, Fatima B, Lahr BD, Greason KL, Schaff HV, Dearani JA, et al. Hyponatremia: an overlooked risk factor Associated with adverse outcomes after cardiac surgery. Ann Thorac Surg. 2021;112(1):91–8. https://doi.org/10.1016/j. athoracsur.2020.08.030.
- Sakr Y, Rother S, Ferreira AM, Ewald C, Dunisch P, Riedemmann N, et al. Fluctuations in serum sodium level are associated with an increased risk of death in surgical ICU patients. Crit Care Med. 2013;41(1):133–42. https://doi. org/10.1097/CCM.0b013e318265f576.
- Ebrahim M, Larsen PB, Hannani D, Liest S, Jorgensen LN, Jorgensen HL. Preoperative risk factors including serum levels of potassium, sodium, and creatinine for early mortality after open abdominal surgery: a retrospective cohort study. BMC Surg. 2021;21(1):62. https://doi.org/10.1186/s12893-021-01070-0.
- Leise MD, Yun BC, Larson JJ, Benson JT, Yang JD, Therneau TM, et al. Effect of the pretransplant serum sodium concentration on outcomes following liver transplantation. Liver Transpl. 2014;20(6):687–97. https://doi.org/10.1002/ lt.23860.
- Williams CN, Riva-Cambrin J, Presson AP, Bratton SL. Hyponatremia and poor cognitive outcome following pediatric brain tumor surgery. J Neurosurg Pediatr. 2015;15(5):480–7. https://doi.org/10.3171/2014.10.PEDS14368.
- Feinstein AJ, Davis J, Gonzalez L, Blackwell KE, Abemayor E, Mendelsohn AH. Hyponatremia and perioperative complications in patients with head and neck squamous cell carcinoma. Head Neck. 2016;38(Suppl 1):E1370–4. https://doi.org/10.1002/hed.24229.
- Wang Q, Peng L, Han Y, Li T, Dai W, Wang Y, et al. Preoperative serum Sodium Level as a prognostic and predictive biomarker for adjuvant therapy in Esophageal Cancer. Front Oncol. 2020;10:555714. https://doi.org/10.3389/ fonc.2020.555714.
- Li W, Chen X, Wang L, Wang Y, Huang C, Wang G, et al. The prognostic effects of hyponatremia and hyperchloremia on postoperative NSCLC patients. Curr Probl Cancer. 2019;43(5):402–10. https://doi.org/10.1016/j. currproblcancer.2018.12.006.
- Xu J, Chen X, Wang X, Zhu C, Hu Y, Yang X, et al. Preoperative hyponatremia and Hypocalcemia Predict Poor Prognosis in Elderly Gastric Cancer Patients. Cancer Manag Res. 2019;11:8765–80. https://doi.org/10.2147/CMAR.S211603.
- Martin JY, Goff BA, Urban RR. Preoperative hyponatremia in women with ovarian cancer: an additional cause for concern? Gynecol Oncol. 2016;142(3):471–6. https://doi.org/10.1016/j.ygyno.2016.06.018.
- Yeh HC, Li CC, Wen SC, Singla N, Woldu SL, Robyak H, et al. Validation of Hyponatremia as a prognostic predictor in Multiregional Upper Tract Urothelial Carcinoma. J Clin Med. 2020;9(4). https://doi.org/10.3390/jcm9041218.
- Sterns RH. Disorders of plasma sodium–causes, consequences, and correction. N Engl J Med. 2015;372(1):55–65. https://doi.org/10.1056/ NEJMra1404489.
- Castle-Kirszbaum M, Kyi M, Wright C, Goldschlager T, Danks RA, Parkin WG. Hyponatraemia and hypernatraemia: Disorders of Water Balance in Neurosurgery. Neurosurg Rev. 2021;44(5):2433–58. https://doi.org/10.1007/ s10143-020-01450-9.
- Hasegawa H, Shin M, Makita N, Shinya Y, Kondo K, Saito N. Delayed postoperative hyponatremia following endoscopic transsphenoidal surgery for non-adenomatous parasellar tumors. Cancers (Basel). 2020;12(12). https://doi. org/10.3390/cancers12123849.
- Voglis S, van Niftrik C, Staartjes VE, Brandi G, Tschopp O, Regli L, et al. Feasibility of machine learning based predictive modelling of postoperative hyponatremia after pituitary surgery. Pituitary. 2020;23(5):543–51. https://doi. org/10.1007/s11102-020-01056-w.
- Rahman M, Friedman WA. Hyponatremia in neurosurgical patients: clinical guidelines development. Neurosurgery. 2009;65(5). https://doi. org/10.1227/01.NEU.0000358954.62182.B3. 925 – 35; discussion 935-6.

- Leung AA, McAlister FA, Rogers SJ, Pazo V, Wright A, Bates DW. Preoperative hyponatremia and perioperative complications. Arch Intern Med. 2012;172(19):1474–81. https://doi.org/10.1001/archinternmed.2012.3992.
- Pennington Z, Bomberger TT, Lubelski D, Benzel EC, Steinmetz MP, Mroz TE. Preoperative hyponatremia and perioperative complications in cervical spinal Fusion. World Neurosurg. 2020;141:e864–72. https://doi.org/10.1016/j. wneu.2020.06.068.
- 27. Crestanello JA, Phillips G, Firstenberg MS, Sai-Sudhakar C, Sirak J, Higgins R, et al. Preoperative hyponatremia predicts outcomes after cardiac surgery. J Surg Res. 2013;181(1):60–6. https://doi.org/10.1016/j.jss.2012.06.004.
- Ramburuth M, Moodley Y, Gopalan PD. Preoperative serum sodium measurements and postoperative inpatient mortality: a case-control analysis of data from the South African Surgical Outcomes Study. S Afr Med J. 2018;108(10):847–51. https://doi.org/10.7196/SAMJ.2018.v108i10.13178.
- Zhang J, Li YI, Pieters TA, Towner J, Li KZ, Al-Dhahir MA, et al. Sepsis and septic shock after craniotomy: Predicting a significant patient safety and quality outcome measure. PLoS ONE. 2020;15(9). https://doi.org/10.1371/journal. pone.0235273. e0235273.
- Vandenbroucke JP, von Elm E, Altman DG, Gotzsche PC, Mulrow CD, Pocock SJ. Schlesselman and M. Egger: strengthening the reporting of Observational Studies in Epidemiology (STROBE): explanation and elaboration. PLoS Med. 2007;4(10). https://doi.org/10.1371/journal.pmed.0040297. e297.
- Mullee A, Romaguera D, Pearson-Stuttard J, Viallon V, Stepien M, Freisling H, et al. Association between soft drink consumption and mortality in 10 european countries. JAMA Intern Med. 2019;179(11):1479–90. https://doi. org/10.1001/jamainternmed.2019.2478.
- Keidel D, Anto JM, Basagana X, Bono R, Burte E, Carsin AE, et al. The role of Socioeconomic Status in the association of lung function and Air Pollution-A Pooled Analysis of three adult ESCAPE cohorts. Int J Environ Res Public Health. 2019;16(11). https://doi.org/10.3390/ijerph16111901.
- Haneuse S, VanderWeele TJ, Arterburn D. Using the E-Value to assess the potential effect of unmeasured confounding in Observational Studies. JAMA. 2019;321(6):602–3. https://doi.org/10.1001/jama.2018.21554.
- 34. Liu Y, Hu H, Li Z, Yang J, Zhang X, Chen L, et al. Association between preoperative platelet and 30-day postoperative mortality of adult patients undergoing craniotomy for brain tumors: data from the American College of Surgeons National Surgical Quality Improvement Program database. BMC Neurol. 2022;22(1):465. https://doi.org/10.1186/s12883-022-03005-5.
- Muhlestein WE, Akagi DS, Davies JM, Chambless LB. Predicting Inpatient length of stay after brain tumor surgery: developing machine learning ensembles to improve predictive performance. Neurosurgery. 2019;85(3):384–93. https://doi.org/10.1093/neuros/nyy343.
- Sorba EL, Staartjes VE, Voglis S, Tosic L, Brandi G, Tschopp O, et al. Diabetes insipidus and syndrome of inappropriate antidiuresis (SIADH) after pituitary surgery: incidence and risk factors. Neurosurg Rev. 2021;44(3):1503–11. https://doi.org/10.1007/s10143-020-01340-0.
- Kitchlu A, Rosner MH. Hyponatremia in patients with cancer. Curr Opin Nephrol Hypertens. 2019;28(5):433–40. https://doi.org/10.1097/ MNH.00000000000525.
- Seo MS, Hwang IC, Jung J, Lee H, Choi JH, Shim JY. Hypernatremia at admission predicts poor survival in patients with terminal cancer: a retrospective cohort study. BMC Palliat Care. 2020;19(1):94. https://doi.org/10.1186/ s12904-020-00607-z.

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