

RESEARCH

Open Access

Complete convergence for negatively orthant dependent random variables

Dehua Qiu^{1*}, Qunying Wu² and Pingyan Chen³

*Correspondence:
qjudhua@sina.com

¹School of Mathematics and Statistics, Guangdong University of Finance and Economics, Guangzhou, 510320, P.R. China
Full list of author information is available at the end of the article

Abstract

In this paper, necessary and sufficient conditions of the complete convergence are obtained for the maximum partial sums of negatively orthant dependent (NOD) random variables. The results extend and improve those in Kuczmaszewska (Acta Math. Hung. 128(1-2):116-130, 2010) for negatively associated (NA) random variables.

MSC: 60F15; 60G50

Keywords: NOD; complete convergence

1 Introduction

The concept of complete convergence for a sequence of random variables was introduced by Hsu and Robbins [1] as follows. A sequence $\{U_n, n \geq 1\}$ of random variables *converges completely* to the constant θ if

$$\sum_{n=1}^{\infty} P(|U_n - \theta| > \varepsilon) < \infty \quad \text{for all } \varepsilon > 0.$$

Moreover, they proved that the sequence of arithmetic means of independent identically distribution (i.i.d.) random variables converges completely to the expected value if the variance of the summands is finite. This result has been generalized and extended in several directions by many authors. One can refer to [2–16], and so forth. Kuczmaszewska [8] proved the following result.

Theorem A *Let $\{X_n, n \geq 1\}$ be a sequence of negatively associated (NA) random variables and X be a random variables possibly defined on a different space satisfying the condition*

$$\frac{1}{n} \sum_{i=1}^n P(|X_i| > x) = DP(|X| > x)$$

for all $x > 0$, all $n \geq 1$ and some positive constant D . Let $\alpha p > 1$ and $\alpha > 1/2$. Moreover, additionally assume that $EX_n = 0$ for all $n \geq 1$ if $p \geq 1$. Then the following statements are equivalent:

- (i) $E|X|^p < \infty$,
- (ii) $\sum_{n=1}^{\infty} n^{\alpha p - 2} P(\max_{1 \leq j \leq n} |\sum_{i=1}^j X_i| \geq \varepsilon n^{\alpha}) < \infty, \forall \varepsilon > 0$.

The aim of this paper is to extend and improve Theorem A to negatively orthant dependent (NOD) random variables. The tool in the proof of Theorem A is the Rosenthal

maximal inequality for NA sequence (cf. [17]), but no one established the kind of maximal inequality for NOD sequence. So the truncated method is different and the proofs of our main results are more complicated and difficult.

The concept of negatively associated (NA) and negatively orthant dependent (NOD) was introduced by Joag-Dev and Proschan [18] in the following way.

Definition 1.1 A finite family of random variables $\{X_i, 1 \leq i \leq n\}$ is said to be negatively associated (NA) if for every pair of disjoint nonempty subset A_1, A_2 of $\{1, 2, \dots, n\}$,

$$|\text{Cov}(f_1(X_i, i \in A_1), f_2(X_j, j \in A_2))| \leq 0,$$

where f_1 and f_2 are coordinatewise nondecreasing such that the covariance exists. An infinite sequence of $\{X_n, n \geq 1\}$ is NA if every finite subfamily is NA.

Definition 1.2 A finite family of random variables $\{X_i, 1 \leq i \leq n\}$ is said to be

- (a) negatively upper orthant dependent (NUOD) if

$$P(X_i > x_i, i = 1, 2, \dots, n) \leq \prod_{i=1}^n P(X_i > x_i)$$

for $\forall x_1, x_2, \dots, x_n \in R$,

- (b) negatively lower orthant dependent (NLOD) if

$$P(X_i \leq x_i, i = 1, 2, \dots, n) \leq \prod_{i=1}^n P(X_i \leq x_i)$$

for $\forall x_1, x_2, \dots, x_n \in R$,

- (c) negatively orthant dependent (NOD) if they are both NUOD and NLOD.

A sequence of random variables $\{X_n, n \geq 1\}$ is said to be NOD if for each n, X_1, X_2, \dots, X_n are NOD.

Obviously, every sequence of independent random variables is NOD. Joag-Dev and Proschan [18] pointed out that NA implies NOD, neither being NUOD nor being NLOD implies being NA. They gave an example that possesses NOD, but does not possess NA, which shows that NOD is strictly wider than NA. For more details of NOD random variables, one can refer to [3, 6, 11, 14, 19–21], and so forth.

In order to prove our main results, we need the following lemmas.

Lemma 1.1 (Bozorgnia et al. [19]) *Let X_1, X_2, \dots, X_n be NOD random variables.*

- (i) *If f_1, f_2, \dots, f_n are Borel functions all of which are monotone increasing (or all monotone decreasing), then $f_1(X_1), f_2(X_2), \dots, f_n(X_n)$ are NOD random variables.*
- (ii) *$E \prod_{i=1}^n X_i^+ \leq \prod_{i=1}^n EX_i^+, \forall n \geq 2.$*

Lemma 1.2 (Asadian et al. [22]) *For any $q \geq 2$, there is a positive constant $C(q)$ depending only on q such that if $\{X_n, n \geq 1\}$ is a sequence of NOD random variables with $EX_n = 0$ for*

every $n \geq 1$, then for all $n \geq 1$,

$$E \left| \sum_{i=1}^n X_i \right|^q \leq C(q) \left\{ \sum_{i=1}^n E|X_i|^q + \left(\sum_{i=1}^n EX_i^2 \right)^{q/2} \right\}.$$

Lemma 1.3 For any $q \geq 2$, there is a positive constant $C(q)$ depending only on q such that if $\{X_n, n \geq 1\}$ is a sequence of NOD random variables with $EX_n = 0$ for every $n \geq 1$, then for all $n \geq 1$,

$$E \max_{1 \leq j \leq n} \left| \sum_{i=1}^j X_i \right|^q \leq C(q) (\log(4n))^q \left\{ \sum_{i=1}^n E|X_i|^q + \left(\sum_{i=1}^n EX_i^2 \right)^{q/2} \right\}.$$

Proof By Lemma 1.2, the proof is similar to that of Theorem 2.3.1 in Stout [23], so it is omitted here. \square

Lemma 1.4 (Kuczmaszewska [8]) Let β, γ be positive constants. Suppose that $\{X_n, n \geq 1\}$ is a sequence of random variables and X is a random variable. There exists constant $D > 0$ such that

$$\sum_{i=1}^n P(|X_i| > x) \leq DnP(|X| > x), \quad \forall x > 0, \forall n \geq 1; \tag{1.1}$$

- (i) if $E|X|^\beta < \infty$, then $\frac{1}{n} \sum_{j=1}^n E|X_j|^\beta \leq CE|X|^\beta$;
- (ii) $\frac{1}{n} \sum_{j=1}^n E|X_j|^\beta I(|X_j| \leq \gamma) \leq C\{E|X|^\beta I(|X| \leq \gamma) + \gamma^\beta P(|X| > \gamma)\}$;
- (iii) $\frac{1}{n} \sum_{j=1}^n E|X_j|^\beta I(|X_j| > \gamma) \leq CE|X|^\beta I(|X| > \gamma)$.

Recall that a function $h(x)$ is said to be slowly varying at infinity if it is real valued, positive, and measurable on $[0, \infty)$, and if for each $\lambda > 0$

$$\lim_{x \rightarrow \infty} \frac{h(\lambda x)}{h(x)} = 1.$$

We refer to Seneta [24] for other equivalent definitions and for a detailed and comprehensive study of properties of slowly varying functions.

We frequently use the following properties of slowly varying functions (cf. Seneta [24]).

Lemma 1.5 If $h(x)$ is a function slowly varying at infinity, then for any $s > 0$

$$C_1 n^{-s} h(n) \leq \sum_{i=n}^{\infty} i^{-1-s} h(i) \leq C_2 n^{-s} h(n)$$

and

$$C_3 n^s h(n) \leq \sum_{i=1}^n i^{-1+s} h(i) \leq C_4 n^s h(n),$$

where $C_1, C_2, C_3, C_4 > 0$ depend only on s .

Throughout this paper, C will represent positive constants of which the value may change from one place to another.

2 Main results and proofs

Theorem 2.1 *Let $\alpha > 1/2$, $p > 0$, $\alpha p > 1$ and $h(x)$ be a slowly varying function at infinity. Let $\{X_n, n \geq 1\}$ be a sequence of NOD random variables and X be a random variables possibly defined on a different space satisfying the condition (1.1). Moreover, additionally assume that for $\alpha \leq 1$, $EX_n = 0$ for all $n \geq 1$. If*

$$E|X|^p h(|X|^{1/\alpha}) < \infty, \tag{2.1}$$

then the following statements hold:

$$(i) \sum_{n=1}^{\infty} n^{\alpha p - 2} h(n) P\left(\max_{1 \leq j \leq n} |S_j| \geq \varepsilon n^\alpha\right) < \infty, \quad \forall \varepsilon > 0; \tag{2.2}$$

$$(ii) \sum_{n=2}^{\infty} n^{\alpha p - 2} h(n) P\left(\max_{1 \leq k \leq n} |S_n^{(k)}| \geq \varepsilon n^\alpha\right) < \infty, \quad \forall \varepsilon > 0; \tag{2.3}$$

$$(iii) \sum_{n=1}^{\infty} n^{\alpha p - 2} h(n) P\left(\max_{1 \leq j \leq n} |X_j| \geq \varepsilon n^\alpha\right) < \infty, \quad \forall \varepsilon > 0; \tag{2.4}$$

$$(iv) \sum_{n=1}^{\infty} n^{\alpha p - 2} h(n) P\left(\sup_{j \geq n} j^{-\alpha} |S_j| \geq \varepsilon\right) < \infty, \quad \forall \varepsilon > 0; \tag{2.5}$$

$$(v) \sum_{n=1}^{\infty} n^{\alpha p - 2} h(n) P\left(\sup_{j \geq n} j^{-\alpha} |X_j| \geq \varepsilon\right) < \infty, \quad \forall \varepsilon > 0. \tag{2.6}$$

Here $S_n = \sum_{i=1}^n X_i$, $S_n^{(k)} = S_n - X_k$, $k = 1, 2, \dots, n$.

Proof First, we prove (2.2). Choose q such that $1/\alpha p < q < 1$. Let $X_i^{(n,1)} = -n^{\alpha q} I(X_i < -n^{\alpha q}) + X_i I(|X_i| \leq n^{\alpha q}) + n^{\alpha q} I(X_i > n^{\alpha q})$, $X_i^{(n,2)} = (X_i - n^{\alpha q}) I(X_i > n^{\alpha q})$, $X_i^{(n,3)} = -(X_i + n^{\alpha q}) I(X_i < -n^{\alpha q})$, $\forall n \geq 1, 1 \leq i \leq n$. Note that

$$X_i = X_i^{(n,1)} + X_i^{(n,2)} - X_i^{(n,3)}$$

and

$$\begin{aligned} & \sum_{n=1}^{\infty} n^{\alpha p - 2} h(n) P\left(\max_{1 \leq j \leq n} |S_j| > \varepsilon n^\alpha\right) \\ & \leq \sum_{n=1}^{\infty} n^{\alpha p - 2} h(n) P\left(\max_{1 \leq j \leq n} \left| \sum_{i=1}^j X_i^{(n,1)} \right| > \varepsilon n^\alpha / 3\right) \\ & \quad + \sum_{n=1}^{\infty} n^{\alpha p - 2} h(n) P\left(\sum_{i=1}^n X_i^{(n,2)} > \varepsilon n^\alpha / 3\right) + \sum_{n=1}^{\infty} n^{\alpha p - 2} h(n) P\left(\sum_{i=1}^n X_i^{(n,3)} > \varepsilon n^\alpha / 3\right) \\ & \stackrel{\text{def}}{=} I_1 + I_2 + I_3. \end{aligned} \tag{2.7}$$

In order to prove (2.2), it suffices to show that $I_l < \infty$ for $l = 1, 2, 3$. Obviously, for $0 < \eta < p$, the condition (2.1) implies $E|X|^{p-\eta} < \infty$. Therefore, we choose $0 < \eta < p$, $\alpha(p - \eta) > \alpha(p -$

$\eta)q > 1$ and $p - \eta - 1 > 0$ if $p > 1$. In order to prove $I_1 < \infty$, we first prove that

$$\lim_{n \rightarrow \infty} n^{-\alpha} \max_{1 \leq j \leq n} \left| \sum_{i=1}^j EX_i^{(n,1)} \right| = 0. \tag{2.8}$$

This holds when $\alpha \leq 1$. Since $\alpha p > 1, p > 1$. By $EX_i = 0, i \geq 1$, and Lemma 1.4, we have

$$\begin{aligned} n^{-\alpha} \max_{1 \leq j \leq n} \left| \sum_{i=1}^j EX_i^{(n,1)} \right| &\leq n^{-\alpha} \max_{1 \leq j \leq n} \sum_{i=1}^j \{E|X_i|I(|X_i| > n^{\alpha q}) + n^{\alpha q}P(|X_i| > n^{\alpha q})\} \\ &\leq 2n^{-\alpha} \sum_{i=1}^n E|X_i|I(|X_i| > n^{\alpha q}) \leq Cn^{1-\alpha} E|X|I(|X| > n^{\alpha q}) \\ &\leq Cn^{-(\alpha(p-\eta)q-1)-\alpha(1-q)} E|X|^{p-\eta} \\ &\rightarrow 0, \quad n \rightarrow \infty. \end{aligned}$$

When $\alpha > 1, p > 1$,

$$\begin{aligned} n^{-\alpha} \max_{1 \leq j \leq n} \left| \sum_{i=1}^j EX_i^{(n,1)} \right| &\leq n^{-\alpha} \max_{1 \leq j \leq n} \sum_{i=1}^j \{E|X_i|I(|X_i| \leq n^{\alpha q}) + n^{\alpha q}P(|X_i| > n^{\alpha q})\} \\ &\leq n^{-\alpha} \sum_{i=1}^n E|X_i| \leq Cn^{1-\alpha} E|X| \\ &\rightarrow 0, \quad n \rightarrow \infty. \end{aligned}$$

When $\alpha > 1, p \leq 1$,

$$\begin{aligned} n^{-\alpha} \max_{1 \leq j \leq n} \left| \sum_{i=1}^j EX_i^{(n,1)} \right| &\leq n^{-\alpha} \max_{1 \leq j \leq n} \sum_{i=1}^j \{E|X_i|I(|X_i| \leq n^{\alpha q}) + n^{\alpha q}P(|X_i| > n^{\alpha q})\} \\ &\leq n^{-\alpha} \sum_{i=1}^n \{E|X_i|I(|X_i| \leq n^{\alpha q}) + n^{\alpha q}P(|X_i| > n^{\alpha q})\} \\ &\leq n^{-\alpha} \sum_{i=1}^n (n^{\alpha(1-p+\eta)q} E|X_i|^{p-\eta}) \\ &\leq Cn^{-(\alpha(p-\eta)q-1)-\alpha(1-q)} E|X|^{p-\eta} \\ &\rightarrow 0, \quad n \rightarrow \infty. \end{aligned}$$

Therefore, (2.8) holds. So, in order to prove $I_1 < \infty$, it is enough to prove that

$$I_1^* := \sum_{n=1}^{\infty} n^{\alpha p-2} h(n) P \left(\max_{1 \leq j \leq n} \left| \sum_{i=1}^j (X_i^{(n,1)} - EX_i^{(n,1)}) \right| > \varepsilon n^\alpha / 6 \right) < \infty. \tag{2.9}$$

By Lemma 1.1 for $\forall n \geq 1, \{X_i^{(n,1)} - EX_i^{(n,1)}, 1 \leq i \leq n\}$ is a sequence of NOD random variables. When $0 < p \leq 2$, by $\alpha(p - \eta) > 1$ and $0 < q < 1$, we have $\alpha - \frac{1}{2} - \alpha(1 - \frac{p-\eta}{2})q > \alpha - \frac{1}{2} - \alpha(1 - \frac{p-\eta}{2}) > 0$. Taking ν such that $\nu > \max\{2, p, (\alpha p - 1)/(\alpha - 1/2), (\alpha p - 1)/(\alpha -$

$\frac{1}{2} - \alpha(1 - \frac{p-\eta}{2})q, \frac{p-(p-\eta)q}{1-q}$, we get by the Markov inequality, the C_r inequality, the Hölder inequality, and Lemma 1.3,

$$I_1^* \leq C \sum_{n=1}^{\infty} n^{\alpha p - \alpha v - 2} h(n) (\log(4n))^v \sum_{i=1}^n E |X_i^{(n,1)}|^v \\
 + C \sum_{n=1}^{\infty} n^{\alpha p - \alpha v - 2} h(n) (\log(4n))^v \left(\sum_{i=1}^n E |X_i^{(n,1)}|^2 \right)^{v/2} \stackrel{\text{def}}{=} I_{11}^* + I_{12}^*.$$

By the C_r inequality, Lemma 1.4, and Lemma 1.5, we have

$$I_{11}^* \leq C \sum_{n=1}^{\infty} n^{\alpha p - \alpha v - 2} h(n) (\log(4n))^v \sum_{i=1}^n E \{ |X_i|^v I(|X_i| \leq n^{\alpha q}) + n^{\alpha q v} P(|X_i| > n^{\alpha q}) \} \\
 \leq C \sum_{n=1}^{\infty} n^{\alpha p - \alpha v - 1} h(n) (\log(4n))^v E \{ |X|^v I(|X| \leq n^{\alpha q}) + n^{\alpha q v} P(|X| > n^{\alpha q}) \} \\
 \leq C \sum_{n=1}^{\infty} n^{\alpha \{ -(1-q)v + p - q(p-\eta) \} - 1} h(n) (\log(4n))^v E |X|^{p-\eta} < \infty.$$

By the C_r inequality and Lemma 1.4,

$$I_{12}^* \leq C \sum_{n=1}^{\infty} n^{\alpha p - \alpha v - 2} h(n) (\log(4n))^v \left\{ \sum_{i=1}^n (E |X_i|^2 I(|X_i| \leq n^{\alpha q}) + n^{2\alpha q} P(|X_i| > n^{\alpha q})) \right\}^{v/2} \\
 \leq C \sum_{n=1}^{\infty} n^{\alpha p - 2 - (\alpha - 1/2)v} h(n) (\log(4n))^v \{ E |X|^2 I(|X| \leq n^{\alpha q}) + n^{2\alpha q} P(|X| > n^{\alpha q}) \}^{v/2}.$$

When $p > 2$,

$$I_{12}^* \leq C \sum_{n=1}^{\infty} n^{\alpha p - 2 - (\alpha - 1/2)v} h(n) (\log(4n))^v (E X^2)^{v/2} < \infty.$$

When $0 < p \leq 2$,

$$I_{12}^* \leq C \sum_{n=1}^{\infty} n^{\alpha p - 2 - (\alpha - 1/2)v} h(n) (\log(4n))^v (E |X|^{p-\eta})^{v/2} n^{\alpha q \{ 2 - (p-\eta) \} v/2} \\
 \leq C \sum_{n=1}^{\infty} n^{\alpha p - 2 - \{ \alpha - \frac{1}{2} - \alpha(1 - \frac{p-\eta}{2})q \} v} h(n) (\log(4n))^v < \infty.$$

Therefore, (2.9) holds for I_2 . Define $Y_i^{(n,2)} = (X_i - n^{\alpha q}) I(n^{\alpha q} < X_i \leq n^\alpha + n^{\alpha q}) + n^\alpha I(X_i > n^\alpha + n^{\alpha q})$, $1 \leq i \leq n$, $n \geq 1$, since $X_i^{(n,2)} = Y_i^{(n,2)} + (X_i - n^{\alpha q} - n^\alpha) I(X_i > n^\alpha + n^{\alpha q})$, we have

$$I_2 \leq \sum_{n=1}^{\infty} n^{\alpha p - 2} h(n) P \left(\sum_{i=1}^n Y_i^{(n,2)} > \varepsilon n^\alpha / 6 \right) \\
 + \sum_{n=1}^{\infty} n^{\alpha p - 2} h(n) P \left(\sum_{i=1}^n (X_i - n^{\alpha q} - n^\alpha) I(X_i > n^\alpha + n^{\alpha q}) > \varepsilon n^\alpha / 6 \right) \\
 \stackrel{\text{def}}{=} I_{21} + I_{22}. \tag{2.10}$$

By Lemma 1.5, (2.1), and a standard computation, we have

$$\begin{aligned}
 I_{22} &\leq \sum_{n=1}^{\infty} n^{\alpha p-2} h(n) \sum_{i=1}^n P(X_i > n^\alpha + n^{\alpha q}) \leq \sum_{n=1}^{\infty} n^{\alpha p-2} h(n) \sum_{i=1}^n P(|X_i| > n^\alpha) \\
 &\leq C \sum_{n=1}^{\infty} n^{\alpha p-1} h(n) P(|X| > n^\alpha) \leq C + CE|X|^p h(|X|^{1/\alpha}) < \infty.
 \end{aligned}
 \tag{2.11}$$

Now we prove $I_{21} < \infty$. By (2.1) and Lemma 1.4, we have

$$\begin{aligned}
 0 &\leq n^{-\alpha} \sum_{i=1}^n EY_i^{(n,2)} \\
 &\leq \begin{cases} n^{-\alpha} \sum_{i=1}^n EX_i I(X_i > n^{\alpha q}), & \text{if } p > 1, \\ n^{-\alpha} \sum_{i=1}^n \{E|X_i| I(|X_i| \leq 2n^\alpha) + n^\alpha P(|X_i| > 2n^{\alpha q})\}, & \text{if } 0 < p \leq 1 \end{cases} \\
 &\leq \begin{cases} Cn^{-\{\alpha(p-\eta)q-1\}-\alpha(1-q)} E|X|^{p-\eta}, & \text{if } p > 1, \\ Cn^{1-\alpha(p-\eta)q} E|X|^{p-\eta}, & \text{if } 0 < p \leq 1 \end{cases} \rightarrow 0, \quad n \rightarrow \infty.
 \end{aligned}$$

Therefore, in order to prove $I_{21} < \infty$, it is enough to prove that

$$I_{21}^* \leq \sum_{n=1}^{\infty} n^{\alpha p-2} h(n) P\left(\sum_{i=1}^n (Y_i^{(n,2)} - EY_i^{(n,2)}) > \varepsilon n^\alpha / 12\right) < \infty.
 \tag{2.12}$$

Taking ν such that $\nu > \max\{2, \frac{\alpha p-1}{\alpha-1/2}, \frac{2(\alpha p-1)}{\alpha(p-\eta)-1}\}$, we get by Lemma 1.1, the Markov inequality, the C_r inequality, the Hölder inequality, and Lemma 1.2,

$$\begin{aligned}
 I_{21}^* &\leq C \sum_{n=1}^{\infty} n^{\alpha p-\alpha\nu-2} h(n) E \left| \sum_{i=1}^n (Y_i^{(n,2)} - EY_i^{(n,2)}) \right|^\nu \\
 &\leq C \sum_{n=1}^{\infty} n^{\alpha p-\alpha\nu-2} h(n) \sum_{i=1}^n E|Y_i^{(n,2)}|^\nu + C \sum_{n=1}^{\infty} n^{\alpha p-\alpha\nu-2} h(n) \left(\sum_{i=1}^n E(Y_i^{(n,2)})^2 \right)^{\nu/2} \\
 &\stackrel{\text{def}}{=} I_{211}^* + I_{212}^*.
 \end{aligned}$$

By the C_r inequality, Lemma 1.4, Lemma 1.5, (2.1), and a standard computation, we have

$$\begin{aligned}
 I_{211}^* &= C \sum_{n=1}^{\infty} n^{\alpha p-\alpha\nu-2} h(n) \sum_{i=1}^n E|Y_i^{(n,2)}|^\nu \\
 &\leq C \sum_{n=1}^{\infty} n^{\alpha p-\alpha\nu-2} h(n) \sum_{i=1}^n \{EX_i^\nu I(n^{\alpha q} < X_i \leq n^{\alpha q} + n^\alpha) + n^{\alpha\nu} P(X_i > n^{\alpha q} + n^\alpha)\} \\
 &\leq C \sum_{n=1}^{\infty} n^{\alpha p-\alpha\nu-2} h(n) \sum_{i=1}^n \{E|X_i|^\nu I(|X_i| \leq 2n^\alpha) + n^{\alpha\nu} P(|X_i| > n^\alpha)\} \\
 &\leq C \sum_{n=1}^{\infty} n^{\alpha p-\alpha\nu-1} h(n) \{E|X|^\nu I(|X| \leq 2n^\alpha) + n^{\alpha\nu} P(|X| > n^\alpha)\} \\
 &\leq C + CE|X|^p h(|X|^{1/\alpha}) < \infty
 \end{aligned}$$

and

$$\begin{aligned}
 I_{212}^* &\leq C \sum_{n=1}^{\infty} n^{\alpha p - \alpha v - 2} h(n) \left\{ \sum_{i=1}^n (EX_i^2 I(n^{\alpha q} < X_i \leq n^{\alpha q} + n^\alpha) + n^{2\alpha} P(X_i > n^{\alpha q} + n^\alpha)) \right\}^{v/2} \\
 &\leq C \sum_{n=1}^{\infty} n^{\alpha p - \alpha v + v/2 - 2} h(n) \{EX^2 I(|X| \leq 2n^\alpha) + n^{2\alpha} P(|X| > n^\alpha)\}^{v/2} \\
 &\leq \begin{cases} C \sum_{n=1}^{\infty} n^{\alpha p - (\alpha - 1/2)v - 2} h(n) (EX^2)^{v/2}, & \text{if } p > 2, \\ C \sum_{n=1}^{\infty} n^{\alpha p - 2 - \{\alpha(p - \eta) - 1\}v/2} h(n) (E|X|^{p-\eta})^{v/2}, & \text{if } p \leq 2 \end{cases} \\
 &\leq \begin{cases} C \sum_{n=1}^{\infty} n^{\alpha p - (\alpha - 1/2)v - 2} h(n), & \text{if } p > 2, \\ C \sum_{n=1}^{\infty} n^{\alpha p - 2 - \{\alpha(p - \eta) - 1\}v/2} h(n), & \text{if } p \leq 2 \end{cases} \\
 &< \infty.
 \end{aligned}$$

Therefore, (2.12) holds. By (2.10)-(2.12) we get $I_2 < \infty$. In a similar way of $I_2 < \infty$ we can obtain $I_3 < \infty$. Thus, (2.2) holds.

(2.2) \Rightarrow (2.3). Note that $|S_n^{(k)}| = |S_n - X_k| \leq |S_n| + |X_k| = |S_n| + |S_k - S_{k-1}| \leq |S_n| + |S_k| + |S_{k-1}| \leq 3 \max_{1 \leq j \leq n} |S_j|$, we have $(\max_{1 \leq k \leq n} |S_n^{(k)}| \geq \varepsilon n^\alpha) \subseteq (\max_{1 \leq j \leq n} |S_j| \geq \varepsilon n^\alpha / 3)$, hence, from (2.2), (2.3) holds.

(2.3) \Rightarrow (2.4). Since $\frac{1}{2}|S_n| \leq \frac{n-1}{n}|S_n| = |\frac{1}{n} \sum_{k=1}^n S_n^{(k)}| \leq \max_{1 \leq k \leq n} |S_n^{(k)}|$, $\forall n \geq 2$, and $|X_k| = |S_n - S_n^{(k)}| \leq |S_n| + |S_n^{(k)}|$, we have $(\max_{1 \leq k \leq n} |X_k| \geq \varepsilon n^\alpha) \subseteq (|S_n| \geq \varepsilon n^\alpha / 2) \cup (\max_{1 \leq k \leq n} |S_n^{(k)}| \geq \varepsilon n^\alpha / 2) \subseteq (\max_{1 \leq k \leq n} |S_n^{(k)}| \geq \varepsilon n^\alpha / 4)$, $\forall n \geq 1$, hence, from (2.3), (2.4) holds.

(2.2) \Rightarrow (2.5). By Lemma 1.5 and (2.3), we have

$$\begin{aligned}
 &\sum_{n=1}^{\infty} n^{\alpha p - 2} h(n) P\left(\sup_{j \geq n} j^{-\alpha} |S_j| \geq \varepsilon\right) \\
 &= \sum_{i=1}^{\infty} \sum_{2^{i-1} \leq n < 2^i} n^{\alpha p - 2} h(n) P\left(\sup_{j \geq n} j^{-\alpha} |S_j| \geq \varepsilon\right) \\
 &\leq C \sum_{i=1}^{\infty} 2^{i(\alpha p - 1)} h(2^i) P\left(\sup_{j \geq 2^{i-1}} j^{-\alpha} |S_j| \geq \varepsilon\right) \\
 &\leq C \sum_{i=1}^{\infty} 2^{i(\alpha p - 1)} h(2^i) \sum_{k=i}^{\infty} P\left(\max_{2^{k-1} \leq j < 2^k} |S_j| \geq \varepsilon 2^{\alpha(k-1)}\right) \\
 &\leq C \sum_{k=1}^{\infty} P\left(\max_{2^{k-1} \leq j < 2^k} |S_j| \geq \varepsilon 2^{\alpha(k-1)}\right) \sum_{i=1}^k 2^{i(\alpha p - 1)} h(2^i) \\
 &\leq C \sum_{k=1}^{\infty} 2^{k(\alpha p - 1)} h(2^k) P\left(\max_{1 \leq j < 2^k} |S_j| \geq \varepsilon 2^{\alpha(k-1)}\right) < \infty.
 \end{aligned}$$

(2.5) \Rightarrow (2.6). The proof of (2.5) \Rightarrow (2.6) is similar to that of (2.2) \Rightarrow (2.4), so it is omitted. \square

Theorem 2.2 *Let $\alpha > 1/2$, $p > 0$, $\alpha p > 1$ and $h(x)$ be a slowly varying function at infinity. Let $\{X_n, n \geq 1\}$ be a sequence of NOD random variables and X be a random variables possibly*

defined on a different space. Moreover, additionally assume that for $\alpha \leq 1$, $EX_n = 0$ for all $n \geq 1$. If there exist constant $D_1 > 0$ and $D_2 > 0$ such that

$$\frac{D_1}{n} \sum_{i=n}^{2n-1} P(|X_i| > x) \leq P(|X| > x) \leq \frac{D_2}{n} \sum_{i=n}^{2n-1} P(|X_i| > x), \quad \forall x > 0, n \geq 1,$$

then (2.1)-(2.6) are equivalent.

Proof From the proof of Theorem 2.1, in order to prove Theorem 2.2, it is enough to show that (2.4) \Rightarrow (2.6) and (2.6) \Rightarrow (2.1). The proof of (2.4) \Rightarrow (2.6) is similar to that of (2.2) \Rightarrow (2.5). Now, we prove (2.6) \Rightarrow (2.1). Firstly we prove that

$$\lim_{n \rightarrow \infty} P\left(\sup_{j \geq n} j^{-\alpha} |X_j| \geq \varepsilon\right) = 0, \quad \forall \varepsilon > 0. \tag{2.13}$$

Otherwise, there are $\varepsilon_0 > 0$, $\delta > 0$, and a sequence of positive integers $\{n_k, k \geq 1\}$, $n_k \uparrow \infty$ such that $P(\sup_{j \geq n_k} j^{-\alpha} |X_j| \geq \varepsilon_0) \geq \delta$, $\forall k \geq 1$. Without loss of generality, we can assume that $n_{k+1} \geq 2n_k$, $\forall k \geq 1$. Therefore, we have

$$P\left(\sup_{j \geq 2n_k} j^{-\alpha} |X_j| \geq \varepsilon_0\right) \geq \delta, \quad \forall k \geq 1.$$

By $\alpha p > 1$ we have

$$\begin{aligned} & \sum_{n=1}^{\infty} n^{\alpha p - 2} h(n) P\left(\sup_{j \geq n} j^{-\alpha} |X_j| \geq \varepsilon_0\right) \\ & \geq \sum_{k=1}^{\infty} \sum_{n=n_k+1}^{2n_k} n^{\alpha p - 2} h(n) P\left(\sup_{j \geq n} j^{-\alpha} |X_j| \geq \varepsilon_0\right) \\ & \geq C \sum_{k=1}^{\infty} n_k^{\alpha p - 1} h(n_k) P\left(\sup_{j \geq 2n_k} j^{-\alpha} |X_j| \geq \varepsilon_0\right) = \infty, \end{aligned}$$

which is in contradiction with (2.6), thus, (2.13) holds. By Lemma 1.1, we get

$$\begin{aligned} P\left(\sup_{j \geq n} j^{-\alpha} |X_j| \geq \varepsilon\right) & \geq P\left(\max_{n \leq j < 2n} j^{-\alpha} |X_j| \geq \varepsilon\right) \\ & \geq P\left(\max_{n \leq j < 2n} |X_j| \geq (2n)^\alpha \varepsilon\right) \\ & \geq 1 - P\left(\max_{n \leq j < 2n} X_j < (2n)^\alpha \varepsilon\right) = 1 - E\left(\prod_{j=n}^{2n-1} I(X_j < (2n)^\alpha \varepsilon)\right) \\ & \geq 1 - \prod_{j=n}^{2n-1} P(X_j < (2n)^\alpha \varepsilon) = 1 - \prod_{j=n}^{2n-1} (1 - P(X_j \geq (2n)^\alpha \varepsilon)) \\ & \geq 1 - \exp\left(-\sum_{j=n}^{2n-1} P(X_j \geq (2n)^\alpha \varepsilon)\right). \end{aligned}$$

By (2.13), we have $\lim_{n \rightarrow \infty} \sum_{j=n}^{2n-1} P(X_j \geq (2n)^\alpha \varepsilon) = 0, \forall \varepsilon > 0$. Therefore, when n is large enough, we have

$$\begin{aligned} P\left(\max_{n \leq j < 2n} j^{-\alpha} |X_j| \geq \varepsilon\right) &\geq 1 - \left\{1 - \sum_{j=n}^{2n-1} P(X_j \geq (2n)^\alpha \varepsilon) + \frac{1}{2} \left(\sum_{j=n}^{2n-1} P(X_j \geq (2n)^\alpha \varepsilon)\right)^2\right\} \\ &\geq C \sum_{j=n}^{2n-1} P(X_j \geq (2n)^\alpha \varepsilon), \quad \forall \varepsilon > 0. \end{aligned}$$

In a similar way, when n is large enough,

$$P\left(\max_{n \leq j < 2n} j^{-\alpha} |X_j| \geq \varepsilon\right) \geq C \sum_{j=n}^{2n-1} P(-X_j \geq (2n)^\alpha \varepsilon), \quad \forall \varepsilon > 0.$$

Thus, when n is large enough, we have

$$P\left(\max_{n \leq j < 2n} j^{-\alpha} |X_j| \geq \varepsilon\right) \geq C \sum_{j=n}^{2n-1} P(|X_j| \geq (2n)^\alpha \varepsilon) \geq Cn P(|X| \geq (2n)^\alpha \varepsilon), \quad \forall \varepsilon > 0. \quad (2.14)$$

Taking $\varepsilon = 2^{-\alpha}$, by (2.6), (2.14), Lemma 1.5, and a standard computation, we have

$$\begin{aligned} \infty &> \sum_{n=1}^{\infty} n^{\alpha p-2} h(n) P\left(\sup_{j \geq n} j^{-\alpha} |X_j| \geq 2^{-\alpha}\right) \geq \sum_{n=1}^{\infty} n^{\alpha p-2} h(n) P\left(\max_{n \leq j < 2n} j^{-\alpha} |X_j| \geq 2^{-\alpha}\right) \\ &\geq C \sum_{n=1}^{\infty} n^{\alpha p-1} h(n) P(|X| \geq n^\alpha) \\ &\geq CE |X|^p h(|X|^{1/\alpha}). \end{aligned}$$

Thus, (2.1) holds. □

In the following, let $\{\tau_n, n \geq 1\}$ be a sequence of non-negative, integer valued random variables and τ a positive random variable. All random variables are defined on the same probability space.

Theorem 2.3 *Let $\alpha > 1/2, p > 0, \alpha p > 1$ and $h(x) > 0$ be a slowly varying function as $x \rightarrow +\infty$. Let $\{X_n, n \geq 1\}$ be a sequence of NOD random variables and X be a random variables possibly defined on a different space satisfying the condition (1.1) and (2.1). Moreover, additionally assume that for $\alpha \leq 1, EX_n = 0$ for all $n \geq 1$. If there exists $\lambda > 0$ such that $\sum_{n=1}^{\infty} n^{\alpha p-2} h(n) P(\frac{\tau_n}{n} < \lambda) < \infty$, then*

$$\sum_{n=1}^{\infty} n^{\alpha p-2} h(n) P(|S_{\tau_n}| \geq \varepsilon \tau_n^\alpha) < \infty, \quad \forall \varepsilon > 0. \quad (2.15)$$

Proof Note that

$$\left\{|S_{\tau_n}| \geq \varepsilon \tau_n^\alpha\right\} \subseteq \left\{\tau_n/n < \lambda\right\} \cup \left\{|S_{\tau_n}| \geq \varepsilon \tau_n^\alpha, \tau_n \geq \lambda n\right\} \subseteq \left\{\tau_n/n < \lambda\right\} \cup \left\{\sup_{j \geq \lambda n} j^{-\alpha} |S_j| \geq \varepsilon\right\}.$$

Thus, by (2.5) of Theorem 2.1, we have (2.15). □

Theorem 2.4 *Let $\alpha > 1/2$, $p > 0$, $\alpha p > 1$ and $h(x)$ be a slowly varying function at infinity. Let $\{X_n, n \geq 1\}$ be a sequence of NOD random variables and X be a random variables possibly defined on a different space satisfying the condition (1.1) and (2.1). Moreover, additionally assume that for $\alpha \leq 1$, $EX_n = 0$ for all $n \geq 1$. If there exists $\theta > 0$ such that $\sum_{n=1}^{\infty} n^{\alpha p - 2} h(n) P(|\frac{\tau_n}{n} - \tau| > \theta) < \infty$ with $P(\tau \leq B) = 1$ for some $B > 0$, then*

$$\sum_{n=1}^{\infty} n^{\alpha p - 2} h(n) P(|S_{\tau_n}| \geq \varepsilon n^\alpha) < \infty, \quad \forall \varepsilon > 0. \tag{2.16}$$

Proof Note that

$$\begin{aligned} (|S_{\tau_n}| \geq \varepsilon n^\alpha) &\subseteq \left(\left| \frac{\tau_n}{n} - \tau \right| > \theta \right) \cup \left(|S_{\tau_n}| \geq \varepsilon n^\alpha, \left| \frac{\tau_n}{n} - \tau \right| \leq \theta \right) \\ &\subseteq \left(\left| \frac{\tau_n}{n} - \tau \right| > \theta \right) \cup (|S_{\tau_n}| \geq \varepsilon n^\alpha, \tau_n \leq (\tau + \theta)n) \\ &\subseteq \left(\left| \frac{\tau_n}{n} - \tau \right| > \theta \right) \cup (|S_{\tau_n}| \geq \varepsilon n^\alpha, \tau_n \leq (B + \theta)n) \\ &\subseteq \left(\left| \frac{\tau_n}{n} - \tau \right| > \theta \right) \cup \left(\max_{1 \leq j \leq (B + \theta)n} |S_j| \geq \varepsilon n^\alpha \right). \end{aligned}$$

Thus, by (2.2) of Theorem 2.1, we have (2.16). □

Competing interests

The authors declare that they have no competing interests.

Authors' contributions

All authors contributed equally to the writing of this paper. All authors read and approved the final manuscript.

Author details

¹School of Mathematics and Statistics, Guangdong University of Finance and Economics, Guangzhou, 510320, P.R. China. ²College of Science, Guilin University of Technology, Guilin, 541004, P.R. China. ³Department of Mathematics, Jinan University, Guangzhou, 510630, P.R. China.

Acknowledgements

The authors would like to thank the referees and the editors for the helpful comments and suggestions. This work was supported by the National Natural Science Foundation of China (Grant No. 11271161).

Received: 14 November 2013 Accepted: 26 March 2014 Published: 09 Apr 2014

References

1. Hsu, P, Robbins, H: Complete convergence and the law of large numbers. *Proc. Natl. Acad. Sci. USA* **33**, 25-31 (1947)
2. Baum, IE, Katz, M: Convergence rates in the law of large numbers. *Trans. Am. Math. Soc.* **120**, 108-123 (1965)
3. Baek, J, Park, ST: Convergence of weighted sums for arrays of negatively dependent random variables and its applications. *J. Stat. Plan. Inference* **140**, 2461-2469 (2010)
4. Bai, ZD, Su, C: The complete convergence for partial sums of i.i.d. random variables. *Sci. China Ser. A* **5**, 399-412 (1985)
5. Chen, P, Hu, TC, Liu, X, Volodin, A: On complete convergence for arrays of rowwise negatively associated random variables. *Theory Probab. Appl.* **52**, 323-328 (2007)
6. Gan, S, Chen, P: Strong convergence rate of weighted sums for negatively dependent sequences. *Acta. Math. Sci. Ser. A* **28**, 283-290 (2008) (in Chinese)
7. Gut, A: Complete convergence for arrays. *Period. Math. Hung.* **25**, 51-75 (1992)
8. Kuczmaszewska, A: On complete convergence in Marcinkiewica-Zygmund type SLLN for negatively associated random variables. *Acta Math. Hung.* **128**(1-2), 116-130 (2010)
9. Liang, HY, Wang, L: Convergence rates in the law of large numbers for B-valued random elements. *Acta Math. Sci. Ser. B* **21**, 229-236 (2001)
10. Peligrad, M, Gut, A: Almost-sure results for a class of dependent random variables. *J. Theor. Probab.* **12**, 87-104 (1999)
11. Qiu, DH, Chang, KC, Antonini, RG, Volodin, A: On the strong rates of convergence for arrays of rowwise negatively dependent random variables. *Stoch. Anal. Appl.* **29**, 375-385 (2011)
12. Sung, SH: Complete convergence for weighted sums of random variables. *Stat. Probab. Lett.* **77**, 303-311 (2007)

13. Sung, SH: A note on the complete convergence for arrays of rowwise independent random elements. *Stat. Probab. Lett.* **78**, 1283-1289 (2008)
14. Taylor, RL, Patterson, R, Bozorgnia, A: A strong law of large numbers for arrays of rowwise negatively dependent random variables. *Stoch. Anal. Appl.* **20**, 643-656 (2002)
15. Wang, XM: Complete convergence for sums of NA sequence. *Acta Math. Appl. Sin.* **22**, 407-412 (1999)
16. Zhang, LX, Wang, JF: A note on complete convergence of pairwise NQD random sequences. *Appl. Math. J. Chin. Univ. Ser. A* **19**, 203-208 (2004)
17. Shao, QM: A comparison theorem on moment inequalities between negatively associated and independent random variables. *J. Theor. Probab.* **13**, 343-356 (2000)
18. Joag-Dev, K, Proschan, F: Negative association of random variables with applications. *Ann. Stat.* **11**, 286-295 (1983)
19. Bozorgnia, A, Patterson, RF, Taylor, RL: Limit theorems for dependent random variables. In: *Proc. of the First World Congress of Nonlinear Analysts '92*, vol. II, pp. 1639-1650. de Gruyter, Berlin (1996)
20. Ko, MH, Han, KH, Kim, TS: Strong laws of large numbers for weighted sums of negatively dependent random variables. *J. Korean Math. Soc.* **43**, 1325-1338 (2006)
21. Ko, MH, Kim, TS: Almost sure convergence for weighted sums of negatively dependent random variables. *J. Korean Math. Soc.* **42**, 949-957 (2005)
22. Asadian, N, Fakoor, V, Bozorgnia, A: Rosenthal's type inequalities for negatively orthant dependent random variables. *J. Iran. Stat. Soc.* **5**(1-2), 69-75 (2006)
23. Stout, WF: *Almost Sure Convergence*. Academic Press, New York (1974)
24. Seneta, E: *Regularly Varying Function*. Lecture Notes in Math., vol. 508. Springer, Berlin (1976)

10.1186/1029-242X-2014-145

Cite this article as: Qiu et al.: Complete convergence for negatively orthant dependent random variables. *Journal of Inequalities and Applications* 2014, **2014**:145

Submit your manuscript to a SpringerOpen[®] journal and benefit from:

- ▶ Convenient online submission
- ▶ Rigorous peer review
- ▶ Immediate publication on acceptance
- ▶ Open access: articles freely available online
- ▶ High visibility within the field
- ▶ Retaining the copyright to your article

Submit your next manuscript at ▶ springeropen.com
