# Multivariate analysis of infant death in England and Wales in 2005–06, with focus on socio-economic status and deprivation

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Current health inequality targets include the goal of reducing the differential in infant mortality between social groups. This article reports on a multivariate analysis of risk factors for infant mortality, with specific focus on deprivation and socio-economic status. Data on all singleton live births in England and Wales in 2005-06 were used and deprivation quintile (Carstairs index) was assigned to each birth using postcode at birth registration. Deprivation had a strong independent effect on infant mortality, risk of death tending to increase with increasing levels of deprivation. The strength of this relationship depended, however, on whether the babies were low birthweight, preterm or small-for-gestational-age. Trends of increasing mortality risk with increasing deprivation were strongest in the postneonatal period. Uniquely, this article reports the number and proportion of all infant deaths which would potentially be avoided if all levels of deprivation were reduced to that of the least deprived group. It estimates that one quarter of all infant deaths would potentially be avoided if deprivation levels were reduced in this way.

#### Introduction

There are many established risk factors for infant mortality; prematurity,¹ low birthweight² and multiplicity³ being the most significant in terms of strength of association and consistency. Risk factors are known to vary according to age at death. For example, the effect of low birthweight and prematurity is stronger in the neonatal period than the postneonatal period.⁴

Socio-economic status is strongly associated with deaths under one year, with a clear trend observed for increased mortality among births occurring to more socially disadvantaged mothers. In 2007, infants of fathers in the routine occupations class had an infant mortality rate of 5.8 per 1,000 live births compared with a rate of 2.8 per 1,000 live births among infants born to fathers in the large employers and higher managerial occupations class. Social inequalities in health are a key public health focus. Public Service Agreement (PSA) targets set in 2001 and updated in 2004, to be achieved by 2010, include the goal of reducing the gap between the infant mortality rate for the routine and manual class and the rate for the population as a whole by at least 10 per cent.

Social class derived from father's occupation is a frequently used indicator of socio-economic status and the indicator used in the setting of the PSA target. However, it is not the only indicator of socio-economic status available. Deprivation indices, such as Carstairs scores are area-based measures of economic and social deprivation and have been used as another indicator of socio-economic status. Deprivation indices are subject to the usual limitations of ecological summary measures, but may provide a measure of socio-economic status where other indicators are unreliable or unavailable. An advantage of area-based measures is that they are not reliant on the

availability of individual data. This is a particular issue with the National Statistics Socio-economic Classification (NS-SEC), as sole registered births (births occurring outside marriage, registered by the mother alone) are not coded to NS-SEC because father's occupation is not available. In addition, only a 10 per cent sample of live birth records (which have occurred inside marriage or jointly registered by both parents) are coded to NS-SEC.

Previous studies have used deprivation indices to investigate the role of deprivation in infant mortality, observing similar trends to occupational social class. <sup>7,8,9</sup> The literature on the relationship between social class and deprivation indices is mixed, with some authors suggesting that deprivation indices can be used as a proxy for social class, and other authors expressing caution and suggesting that area-based deprivation has an independent effect on mortality.

National infant mortality rates have previously been investigated with respect to important risk factors such as multiplicity, sex, mother's age, marital status/type of registration and parity inside marriage and the results presented separately for each risk factor in turn, <sup>10</sup> but little work has been conducted using multivariate analysis. The recent exception to this was an analysis of risk factors for neonatal mortality among singleton infants weighing 2,500-5,499 grams. 11 This analysis found that even after excluding low birthweight babies, birthweight was still the strongest risk factor for neonatal mortality. Sex and older maternal age both remained independently associated with neonatal mortality in this sample after adjustment for other factors.

Recent methodological improvements have provided useful opportunities for the investigation of infant mortality in England and Wales. The introduction of the NHS Numbers for Babies (NN4B) programme in 2005 (when NHS numbers began to be allocated at birth) has enabled the linkage of key variables (such as gestation and ethnicity of the infant) to birth registration records. Information on these variables is not collected at live birth registration, and so previously these data have been unavailable for analysis (and were not included in the multivariate analysis quoted earlier11).

In this article we report our investigation into risk factors for infant mortality in England and Wales in 2005-06, focusing on the contribution of socio-economic status and deprivation.

#### Methods

### Source data

Data on all live births that occurred in England and Wales in 2005 and 2006 were extracted from birth registration records linked to the corresponding NN4B record. A detailed description of this linkage process as applied to births occurring in the first quarter of 2005 is reported in an earlier Health Statistics Quarterly article. 12

For births which had resulted in infant death, further details of timing and cause of death were obtained by linkage to death registration records. Details on this linkage have been published previously. 13 The data extract was taken in mid-2008 to ensure the inclusion of all late birth and death registrations. Multiple births were excluded from the data extract prior to the main analysis, since risk factors for mortality are likely to differ substantially for these births.

#### Variables available for analysis

Socio-economic status: NS-SEC and deprivation index The role of social inequalities in infant mortality was a key focus of this investigation. NS-SEC replaced previously-used classifications of Social Class and Socio-economic Groups in 2001, and represents

the current system of allocating socio-economic status to reported occupational position. A detailed description of NS-SEC is published on the Office for National Statistics (ONS) website. 14 The principal version of NS-SEC is an eight or nine group analytic version, collapsible to five or three group versions (Box One). The analysis reported here uses the three class version of NS-SEC, as this includes the routine and manual group which was used in the setting of the PSA target on infant mortality. Parental occupation is recorded at birth and death registration, but only NS-SEC status coded at birth registration was used for this analysis. However, since NS-SEC status was missing for over 90 per cent of the data (being only available for a 10 per cent sample of married/jointly registered births and not present at all for sole registrations) most analyses in this paper did not concentrate on this measure of socio-economic status.

# **Box** one

#### **National Statistics Socio-economic Classification**

Principal version	5 group version	3 group version			
1.1 Large employers and higher managerial occupations					
1.2 Higher professional occupations	1 Managerial and professional occupations	1 Managerial and professional occupations			
2 Lower managerial and professional occupations					
3 Intermediate occupations	2 Intermediate occupations	2 Intermediate			
4 Small employers and own account workers	3 Small employers and own account workers	occupations			
5 Lower supervisory and technical occupations	4 Lower supervisory and technical occupations				
6 Semi-routine occupations	5 Semi-routine and routine occupations	3 Routine and manual occupations			
7 Routine occupations	- Toutine occupations				
8 Never worked and long-term unemployed	Never worked and long-term unemployed	Never worked and long-term unemployed			

The Carstairs deprivation index was used to assign a deprivation score to all births, according to the postcode given at birth registration. This deprivation index was chosen as it is frequently used for other ONS health-related analyses and correlates well with other deprivation indices. A full description of how Carstairs scores are calculated is in an earlier article. 15 Briefly, Carstairs scores are based on the un-weighted combination of four variables from the 2001 Census (unemployment, overcrowding, car ownership and low social class) (Box Two), and are assigned to the postcode. The Carstairs index is thus an ecological (rather than individual) based measure, since postcodes only identify the address to within an average of 15 properties per postcode (and can be up to 100 properties). The Carstairs scores for the total (all ages) population can be used to rank electoral wards from the least to the most deprived and divided into percentile groups. For this analysis electoral wards were divided into quintiles (equal fifths) of deprivation using 2001 experimental ward total population estimates. Since these quintiles are based on the total (all age) population, numbers of live births within each Carstairs deprivation group in this analysis will not be equal.

# **Box** two

# 2001 Census variables used in the calculation of the Carstairs deprivation index

Unemployment: unemployed males 16 and over as a proportion of all economically active males aged 16 and over

Overcrowding: persons in households with one or more persons per room as a proportion of all residents in households

Car ownership: residents in households with no car as a proportion of all residents in households

Low Social Class: residents in households with an economically active head of household in Social Class IV or V approximated from NS-SEC

#### Other social and biological risk factors

The following information was taken from birth registration records: maternal age at birth, maternal country of birth, birthweight of infant, year of birth of infant and sex of infant. Parity of the mother was available from birth registration records but was excluded as it only refers to parity within marriage and is therefore not a reliable indicator of 'true' parity. The ethnicity of the infant (as defined by the mother using pre-specified categories) was taken from NN4B data, along with gestational age at birth. A composite variable incorporating maternal country of birth (UK versus non-UK) and infant's ethnicity (collapsed into four categories: Asian, Black, White, Other) was also created. Sex-standardised birthweight for gestational age was calculated using within-cohort gestation- and sex-specific centiles. Infants with birthweights below the 5th centile were classified as 'small-for-gestational-age'.

#### Classification of infant deaths

Timing of death was classified as early neonatal, late neonatal, or postneonatal, according to established definitions (Box Three). Causes of death were grouped using ONS cause groups, 16 based on the established Wigglesworth classification system. Deaths were classified according to causes occurring before the onset of labour (congenital anomalies, antepartum infections, or immaturity related conditions), in or shortly after labour (asphyxia, anoxia or intrapartum trauma), postnatally (other specific conditions, sudden infant deaths), and other (other conditions not mentioned above).

# **Box** three

# Definitions used in this paper

Early neonatal deaths: deaths at ages under 7 days Late neonatal deaths: deaths at ages 7 days and over but under 28 days Neonatal deaths: deaths at ages under 28 days

Postneonatal deaths: deaths at ages 28 days and over but under one year Infant deaths: deaths under one year

Early neonatal mortality rate: early neonatal deaths per 1,000 live births Late neonatal mortality rate: late neonatal deaths per 1,000 live births surviving the early neonatal period

Neonatal mortality rate: neonatal deaths per 1,000 live births

Postneonatal mortality rate: postneonatal deaths per 1,000 live births surviving the neonatal period

Infant mortality rate: infant deaths per 1,000 live births

#### Statistical analysis

All analyses in this paper were performed using Stata statistical software (version 10, Stata Corp, College Station, TX, USA). All P-values quoted are two-sided and values less than 0.05 have been taken to indicate statistical significance.

Mortality rates were calculated using the number of infants still alive and at risk at the beginning of each time period (Box Three).

The multivariate analysis followed a strategy which included consideration of both plausible confounding (Box Four) and potential causal pathways. Carstairs deprivation indices and (among the 10 per cent sample where this was coded) NS-SEC were considered to be the main factors of interest in this analysis, with maternal age, marital/registration status, maternal country of birth, and sex and ethnicity of the baby being considered as potential confounding variables.

# **Box** four

# **Confounding and Interaction**

Confounding is the situation where an association between an exposure and an outcome is entirely or partially due to another exposure (called the confounder). A variable will only confound an association if it satisfies three conditions:

- It must be associated with the exposure of interest
- It must be a risk factor for the outcome of interest
- It must not be on the causal pathway (be an intervening or mediating variable) between the exposure of interest and the outcome of interest

An example is the finding that coffee drinking is associated with risk of coronary heart disease (CHD). In fact coffee drinking is not a risk factor for CHD, but the observation is driven by the fact that people who drink coffee are more likely to smoke than people who do not drink coffee. When the data are analysed separately for smokers and non-smokers (stratified by smoking status) we find that coffee drinking is not, in itself, a risk factor. This is an example of complete confounding, but most examples are of partial confounding.

There are several ways of making adjustments for confounding effects, and all involve the stratification of data according to different levels of the potential confounding factor. It is important to note that in confounding, the two factors (the exposure of interest and the potential confounder) are associated with each other (for example, coffee drinkers are more likely to smoke than non-coffee drinkers) but do not act together —or rely on each other — to produce an effect. They act independently.

Interaction (or effect modification) is an effect of two exposures (or risk factors) on an outcome, where the effects are not independent. They act together to produce an effect on the outcome which is different than the effect of each factor separately. In this situation it is not possible to 'adjust' for one of the factors, and the results must be presented separately for different levels of the effect modifier.

Analyses were conducted separately for neonatal, postneonatal and total infant mortality. The association between mortality and deprivation, controlling for confounding, was explored using logistic regression analysis, effects on risk being estimated by odds ratios (OR) with 95 per cent confidence intervals (CI), and statistical significance being tested using likelihood ratio tests. <sup>17</sup> The odds ratios are sometimes referred to as generic 'relative risks' in the text.

The association between deprivation and socio-economic status and infant mortality in relation to the more 'proximal' factors (along the causal pathway) of gestation and birthweight (and sex-standardised weight for gestation) was explored further by examining interaction terms in the models (Box Four). To maximise statistical power, interaction terms with Carstairs index or NS-SEC were calculated using a dichotomised version of the potential modifying variable (for example, interaction of deprivation with birthweight: 10 categories, 5 deprivation quintiles, and two birthweight categories (<2,500 grams, 2,500 grams and over). Since there is substantial overlap between preterm delivery, low birthweight and small-for-gestational-age, these three factors could only be considered in separate analyses. Stratified measures of effect were reported when statistically significant interactions were found.

Although NS-SEC is the variable most relevant to the PSA target, the small proportion of births coded to NS-SEC, together with the exclusion of sole registrations from the coding led to conducting the majority of the analyses using deprivation as the main factor of interest. An analysis of the interrelationship between NS-SEC and Carstairs deprivation among the 10 per cent sample of births where the parents were married or jointly registered the birth was included in order to test the validity of this decision.

Population excess (or attributable) deaths and population excess (or attributable) deaths per cent were calculated using adjusted odds ratios (Box Five). These measure the number and proportion of deaths in the population that would potentially be avoided if the risk in the 'exposed'

# **Box** five

# **Calculation of Population Excess Deaths** (Population Attributable Risk) and Population Excess Deaths per cent (Population Attributable Risk Fraction)

Population excess per cent, also called the population attributable risk fraction, refers to the proportion of deaths in a population that can be attributed to an exposure, and is the proportion of deaths that could theoretically be saved or avoided in the population if the exposure was removed. There are several ways of calculating this, but in this paper the following were used:

(i) Where there are two strata only (exposed and unexposed):

$$p (RR-1)/(RR-1) + 1$$

where p=proportion of population exposed, and RR is the adjusted relative risk between exposure and outcome. Odds Ratios are used to estimate the relative risk

(ii) Where there were more than two strata:

$$p_i (RR_i - 1) / RR_i$$

where i refers to each stratum, and p<sub>i</sub> = proportion of 'exposed' (in groups above the baseline category) deaths in stratum i

Population excess deaths refer to the number of deaths that can be attributed to a factor, and that theoretically could be 'saved' or 'avoided' if the factor was removed from the population. In this paper these were calculated as the total number of deaths within the population (or stratum) multiplied by the population (or stratum) excess per cent.

group was the same as that in the baseline or 'unexposed' group. It should be emphasised that this is a theoretical construct and is used in this work for illustrative purposes only.

## Results

Overall there were 1,315,352 live births that occurred in 2005 and 2006 and were registered in England and Wales, 3 per cent (39,154) of which were multiple births. All but 1,387 of these births (0.1 per cent) were successfully linked to the corresponding NN4B birth record. Among these births, 6,491 were registered as infant deaths (3,444 early neonatal, 1,099 late neonatal and 1,948 postneonatal deaths). Mortality rates were higher among multiple births, with an infant mortality rate of 23.5 per 1,000 live births compared with 4.4 per 1,000 live births for singleton births. Early neonatal mortality was particularly high among multiple births compared to singletons (14.4 compared with 2.3 per 1,000 live births). Multiple births were excluded from all further analyses, as their risk profile is likely to differ substantially from that of singleton babies.

## **Descriptive analyses**

A description of infant mortality rates by year of death and the available social and biological risk factors is presented in Table 1. Among 1,276,198 singleton live births registered in 2005-06 there were a total of 5,571 infant deaths, giving an overall infant mortality rate of 4.4 per 1,000 live births. There was no difference in infant mortality between 2005 and 2006.

Death rates were more than twice as high in the first 28 days of life as in the subsequent 11 months. There were marked trends of increasing mortality with increasing deprivation (all babies) and increasing mortality with declining social class as measured by NS-SEC (among the subset for which this was coded).

Babies born to married parents had the lowest infant mortality, with the highest risks observed among babies registered by parents living at different addresses and those where the birth was registered by the mother alone (sole registrations).

Infants born to mothers who were born outside the UK had higher risks than those born to UK-born mothers. Non-White babies had higher mortality rates than White babies, particularly those classified as Asian or Black. Interestingly, within each ethnic group mortality was lower among those born to non-UK-born mothers, the only exception being White babies in the neonatal period, where the rates were the same.

There was a u-shaped relationship between mortality and maternal age, more marked in the neonatal period, with the lowest risk among babies whose mothers were aged 30-34 at birth.

With regard to the characteristics of the babies themselves, males had around 20 per cent higher mortality than females. Highest mortality rates were observed among preterm infants with rates declining dramatically with increasing gestational age: among preterm infants the rates ranged from 372.9 to 6.5 deaths per 1,000 in the neonatal period, and from 81.1 to 4.4 deaths per 1,000 in the postneonatal period, the rates then reducing to under 1 per 1,000 in both periods for babies born at 37 weeks and over.

A similar pattern was observed for low birthweight babies compared with those weighing 2,500 grams or more, the rates reducing from 170.6 and 7.6 per 1,000 live births (neonatal period) and 38.9 and 5.3 per 1,000 (postneonatal period) among extremely and very low birthweight babies respectively, to under 1 per 1,000 in both periods for babies weighing 2,500 grams or more.

Similarly, babies whose weight was below the 5 per cent centile of sex-specific weight-for-gestation had infant mortality rates that were almost 3.5 times those of babies at or above the 5th centile. The neonatal mortality rate for small-for-gestational-age babies was 3.2 times, and the postneonatal mortality rate 4 times, that for non-small-for-gestational-age babies.

Table 1

# Singleton neonatal, postneonatal and infant deaths (numbers and rates) by social and biological risk factors, 2005–06

			Nu	ımbers			D-4	
		Live birthe		Deaths		1	Rates	
		Live births	Neonatal	Postneonatal	Infant	Neonatal <sup>1</sup>	Postneonatal <sup>2</sup>	Infant1
Total		1,276,198	3,797	1,774	5,571	3.0	1.4	4.4
Year	2005	626,917	1,875	867	2,742	3.0	1.4	4.4
	2006	649,281	1,922	907	2,829	3.0	1.4	4.4
Carstairs deprivation index	1 (least deprived)	199,540	399	176	575	2.0	0.9	2.9
carstans acprivation mack	2	213,613	488	222	710	2.3	1.0	3.3
	3	239,341	619	279	898	2.6	1.2	3.8
	4	276,132	891	416	1,307	3.2	1.5	4.7
	5 (most deprived)	347,182	1,387	676	2,063	4.0	2.0	5.9
	Missing	390	13	5	18	33.3	13.3	46.2
NS-SEC	Higher and professional	44,651	100	33	133	2.2	0.7	3.0
	Intermediate	23,103	54	30	84	2.3	1.3	3.6
	Routine and manual	44,136	152	56	208	3.4	1.3	4.7
	Other (unemployed/students/ns)	7,234	35	14	49	4.8	1.9	6.8
	Missing	1,157,074	3,456	1,641	5,097	3.0	1.4	4.4
Registration type/Marital status	Married	722,871	1,895	840	2,735	2.6	1.2	3.8
	Sole registration <sup>3</sup>	88,242	309	243	552	3.5	2.8	6.3
	Joint registration/same address <sup>4</sup>	351,429	1,147	461	1,608	3.3	1.3	4.6
	Joint registration/different address <sup>4</sup>	113,656	446	230	676	3.9	2.0	5.9
Maternal age (years)	<20	89,108	353	200	553	4.0	2.3	6.2
	20–24	245,328	795	469	1,264	3.2	1.9	5.2
	25–29	328,102	1,014	411	1,425	3.1	1.3	4.3
	30–34	364,249	914	374	1,288	2.5	1.0	3.5
	35–39	205,633	559	263	822	2.7	1.3	4.0
	40 and over	43,777	162	57	219	3.7	1.3	5.0
Maternal country of birth	UK	1,002,722	2,827	1,368	4,195	2.8	1.4	4.2
	non-UK	273,439	969	406	1,375	3.5	1.5	5.0
Sex of baby	Male	653,362	2,127	1,002	3,129	3.3	1.5	4.8
	Female	622,836	1,670	772	2,442	2.7	1.2	3.9
Ethnicity of baby	White	889,241	2,265	1,125	3,390	2.5	1.3	3.8
, ,	Asian	96,908	433	228	661	4.5	2.4	6.8
	Black	53,999	273	98	371	5.1	1.8	6.9
	Other	102,133	348	160	508	3.4	1.6	5.0
	Missing	133,917	478	326	804	3.6	2.4	6.0
Baby's ethnicity and maternal country of birth	White, UK born	812,702	2,072	1,056	3,128	2.5	1.3	3.8
	White, non-UK born	76,517	193	69	262	2.5	0.9	3.4
	Asian, UK born	32,203	161	84	245	5.0	2.6	7.6
	Asian, non-UK born	64,700	272	144	416	4.2	2.2	6.4
	Black, UK born	11,992	67	26	93	5.6	2.2	7.8
	Black, non-UK born	42,004	206	72	278	4.9	1.7	6.6
	Other, UK born	42,093	165	74	239	3.9	1.8	5.7
	Other, non-UK born Missing (not stated/not linked)	60,037 133,950	182 479	86 163	268 642	3.0 3.6	1.4 1.2	4.5 4.8
Gestation (weeks)	21–27	4,934	1,840	251	2,091	372.9	81.1	423.8
	28–31	8,866	335	132	467	37.8	15.5	52.7
	32–36 37–41	64,330 1,131,844	418 1,027	280 1,045	698 2,072	6.5 0.9	4.4 0.9	10.9 1.8
	42 and over	56,260	52	46	98	0.9	0.9	1.7
	Missing	9,964	125	20	145	12.5	2.0	14.6
Birthweight (grams)	<1,500	12,132	2,070	391	2,461	170.6	38.9	202.9
Zirani (granis)	1,500–2,499	65,050	497	341	838	7.6	5.3	12.9
	2,500–4,499	1,169,415	1,034	1,017	2,051	0.9	0.9	1.8
	4,500 and over	21,709	19	10	29	0.9	0.5	1.3
	Missing	7,892	177	15	192	22.4	1.9	24.3
Small for gestational age	SGA (<5%)	62,136	478	295	773	7.7	4.8	12.4
	Not SGA (≥5%)	1,203,756	2,860	1,459	4,316	2.4	1.2	3.6
	Missing	10,306	462	20	482	44.8	2.0	46.8

Neonatal and infant mortality rates are given per 1,000 live births.
 Postneonatal mortality rates are given per 1,000 live births surviving the neonatal period.
 Birth takes place outside marriage, only mother registers the birth.
 Birth takes place outside marriage, both mother and father register the birth.

Mortality rates appeared high among those with missing information on gestation and birthweight. The reason for this is not known, but one possible explanation may be due to prioritising emergency/intensive care of high risk babies over recording data, particularly on the NN4B.

#### Cause of death

Immaturity-related conditions and congenital anomalies (42 per cent and 30 per cent of all infant deaths respectively) caused the largest numbers and rates of infant death (Table 2). In the neonatal period immaturity-related death predominated at 1.6 per 1,000 live births, with congenital anomalies accounting for a further 0.8 per 1,000 live births; these two causes accounted for 80 per cent of all neonatal deaths. In the postneonatal period congenital anomalies were the most common cause of death, 0.5 per 1,000 (36 per cent of all postneonatal death). Three other causes - immaturity, sudden infant death and other conditions – accounted for a further 41 per cent of deaths, each with a rate 0.2 per 1,000 (Table 2).

Table 2

Singleton neonatal, postneonatal and infant deaths by cause of death and deprivation quintile, 2005-06

**England and Wales** 

		Numbers			Rates	
Cause group		Deaths			nates	
	Neonatal	Postneonatal	Infant	Neonatal <sup>1</sup>	Postneonatal <sup>2</sup>	Infant1
Congenital anomalies	1,008	646	1,654	0.8	0.5	1.3
Antepartum infections	121	19	140	0.1	_	0.1
Immaturity related conditions	2,045	267	2,312	1.6	0.2	1.8
Asphyxia, anoxia or trauma (intrapartum)	369	11	380	0.3	_	0.3
External conditions	18	88	106	_	0.1	0.1
Infections	55	170	225	_	0.1	0.2
Other specific conditions	21	39	60	_	_	_
Sudden infant deaths	49	252	301	_	0.2	0.2
Other conditions	111	282	393	0.1	0.2	0.3

- Neonatal and infant mortality rates are given per 1,000 live births.
- Postneonatal mortality rates are given per 1,000 live births surviving the neonatal period.

Mortality from immaturity related conditions increased with increasing deprivation in both neonatal and postneonatal death periods, though this was more marked in the neonatal period (Figure 1). Conversely, trends for deaths from causes related to congenital anomalies were less marked with increasing deprivation in the neonatal period, but stronger in the postneonatal period. Death from antepartum infection and other conditions also showed strong trends with increasing deprivation in the neonatal period, but there were no other consistent trends in the first 28 days of life. In the postneonatal period mortality from sudden infant death increased markedly with increasing levels of deprivation, as did deaths from external factors, infections, and other causes.

## **National Statistics Socio-economic Classification and Carstairs deprivation index**

Infant mortality rates by Carstairs classification and NS-SEC are presented in Table 3 for the 10 per cent sample of live births leading to infant deaths where NS-SEC was coded.

Overall, the mortality rates for NS-SEC and Carstairs deprivation index mapped well to each other (Table 3). The limitations of the ecological nature of the Carstairs measure are demonstrated in the body of the table, where it can be seen that within each (individually-coded) NS-SEC grouping there were births coded to all five deprivation categories, even in the highest (managerial and professional) group. The one exception to this was among babies of fathers who were unemployed or students (or no stated occupation), where there were no deaths and only 297 live births in the least deprived Carstairs category.

#### Multivariate analysis

Association between deprivation and co-variates (potential confounding variables and more proximal risk factors for infant mortality)

Maternal age, marital status/registration type, maternal country of birth and baby's ethnicity were all strongly related to Carstairs deprivation index (all P<0.001). As levels of deprivation increased, mothers tended to be younger, less likely to be married or living with the father, and more likely to be non-UK-born. Babies who were Asian, Black or of other non-White ethnicity tended to have higher levels of deprivation.

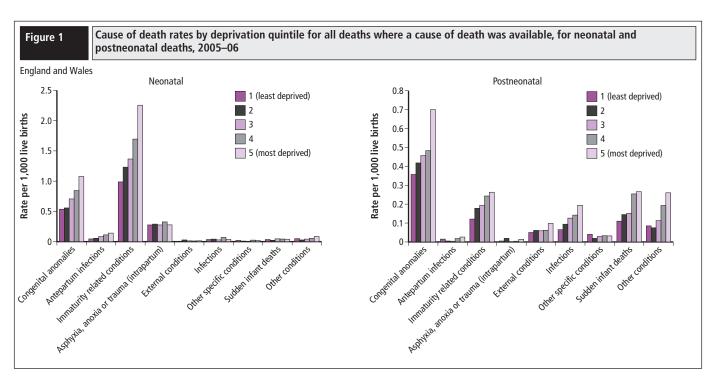


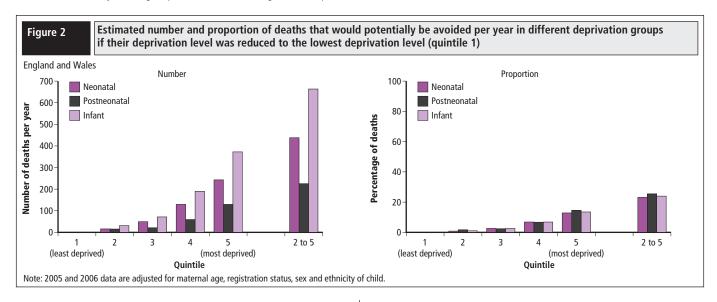
Table 3

#### Neonatal, postneonatal and infant deaths rates by Carstairs quintile and NS-SEC, 2005-06

**England and Wales** 

						NS-	-SEC							All	
Carstairs quintiles	Manage	rial and pro	fessional	Intermediate			Routine and manual				Other				
·	Neonatal <sup>1</sup>	Post- neonatal <sup>2</sup>	Infant <sup>1</sup>	Neonatal <sup>1</sup>	Post- neonatal <sup>2</sup>	Infant¹	Neonatal <sup>1</sup>	Post- neonatal <sup>2</sup>	Infant <sup>1</sup>	Neonatal <sup>1</sup>	Post- neonatal <sup>2</sup>	Infant¹	Neonatal <sup>1</sup>	Post- neonatal <sup>2</sup>	Infant <sup>1</sup>
1 (least deprived)	1.6	0.6	2.2	1.8	0.8	2.5	2.0	1.3	3.3	_	_	-	1.7	0.8	2.4
2	2.0	0.1	2.1	1.8	0.4	2.2	2.7	1.7	4.3	6.0	4.0	10.0	2.2	0.7	3.0
3	2.4	0.9	3.3	2.2	1.1	3.3	2.8	1.5	4.2	4.8	2.4	7.3	2.6	1.2	3.8
4	2.3	1.0	3.4	2.6	1.9	4.5	3.8	1.1	4.9	3.5	2.9	6.4	3.1	1.4	4.5
5 (most deprived)	3.3	1.3	4.6	3.2	2.0	5.2	4.3	1.1	5.4	5.7	1.3	6.9	4.0	1.3	5.4
All	2.2	0.7	3.0	2.3	1.3	3.6	3.4	1.3	4.7	4.8	1.9	6.8	2.9	1.1	4.0

- Neonatal and infant mortality rates are given per 1,000 live births.
- Postneonatal mortality rates are given per 1,000 live births surviving the neonatal period.



As might be expected, there was no association between sex of the baby and deprivation level (P>0.05). Preterm delivery, low birthweight and small-for-gestational-age (<5th centile, adjusted for gender) all showed marked increasing trends in prevalence with increasing levels of deprivation (Appendix Table A1).

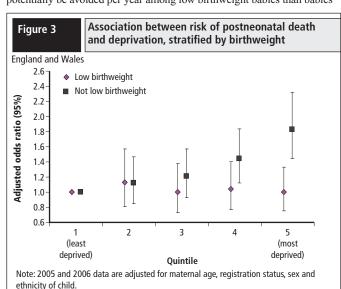
Association between deprivation/socio-economic status and infant mortality

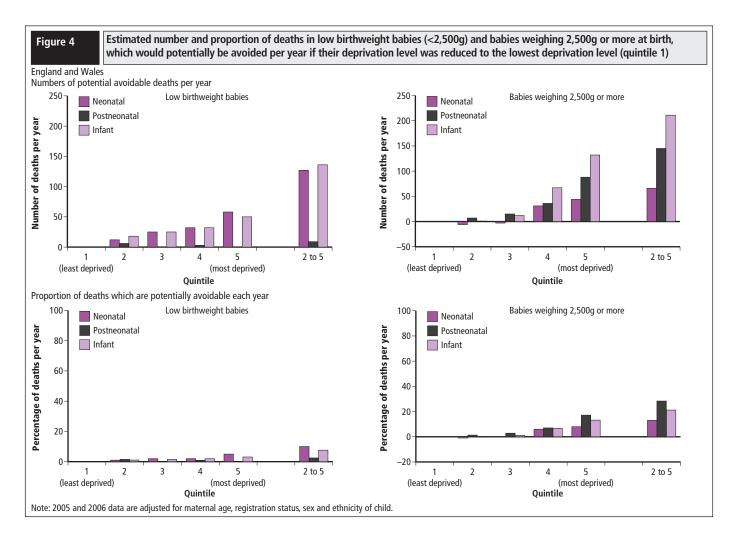
#### Carstairs deprivation index

There was a highly statistically significant independent association between deprivation and risk of infant mortality, similar in both neonatal and postneonatal periods, with a marked trend of increasing risk of infant mortality with increasing levels of deprivation (Tables 4 to 6). There was evidence of some confounding of this relationship by maternal age, registration status and sex and ethnicity of the baby, though the association remained after adjustment for these factors. For example, after adjustment, risk of overall infant mortality in the most deprived group was estimated to be around 1.6 times that in the least deprived group (OR (95 per cent CI): 1.57 (1.41, 1.75)), having reduced from a crude OR of around two (Table 6). The number and proportion of excess deaths, which potentially would be avoided if deprivation levels reduced to the lowest level, rose with increasing levels of deprivation. Twenty-four per cent of all infant deaths, equivalent to around 666 deaths per year, would potentially be avoided if all mothers had the lowest deprivation level (Table 6 and Figure 2).

There was strong evidence of an interaction between deprivation and (separately) birthweight, gestation and small-for-gestational-age (Table 7). Among low birthweight babies, there was only a very weak, if any,

association between deprivation and infant mortality risk. Among babies weighing 2,500 grams or more, however, there was a marked trend of increasing risk of infant mortality with increasing deprivation, strongest in the postneonatal period; ORs (95 per cent CI) rising steadily to 1.35 (1.09, 1.67) and 1.83 (1.44, 2.32) for the most deprived compared to least deprived groups in the neonatal and postneonatal periods respectively (Table 7 and Figure 3). With regard to actual impact on infant mortality in the general population, however, in the neonatal period more deaths could potentially be avoided per year among low birthweight babies than babies





weighing 2,500 grams or more if the level of deprivation were reduced within each category to that of the least deprived group (127 deaths versus 66 deaths per year). In the postneonatal period, more deaths could potentially be avoided in this way among non-low birthweight babies (145 versus 9 deaths per year) (Table 7 and Figure 4).

With regard to the interaction between deprivation and gestation, in the neonatal period there was a steadily increasing trend in mortality risk with increasing deprivation among preterm babies (Table 7). For term babies, although odds ratios were raised in the two most deprived categories, there was no evidence of a difference in mortality risk among the first three deprivation groups. Conversely, in the postneonatal period, there was no consistent trend of mortality risk with deprivation among preterm babies (though the odds ratios were raised for all categories relative to the least deprived group), but there was a marked trend of increasing mortality risk with increasing levels of deprivation among babies who were not preterm. Figure 5 displays the numbers and proportions of deaths attributable to deprivation in preterm and term babies. If deprivation were reduced to the lowest level, 227 neonatal deaths could be avoided in the preterm group (representing 18 per cent of all deaths in this group), and 131 postneonatal deaths could be avoided in the term group (24 per cent of all deaths in this group).

Examining the interaction between deprivation and small-forgestational-age, the effect of deprivation was much stronger in non-smallfor-gestational-age than small-for-gestational-age babies in both neonatal and postneonatal periods (Table 7). Among small-for-gestational-age babies there was a shallow trend of increasing risk with increasing level of deprivation in the neonatal period, but in the postneonatal period there was no evidence of an effect of deprivation on risk of death. Among

babies whose sex-specific weight for gestation was at the 5th centile or above, there were marked trends of increasing odds ratios with increasing level of deprivation in both time periods (Table 7). With regard to potentially avoidable deaths, if deprivation levels reduced to the lowest level, 303 neonatal deaths and 220 postneonatal deaths in non-small-forgestational-age babies could be saved. These numbers represent 21 per cent and 30 per cent respectively of all deaths in these groups (Table 7).

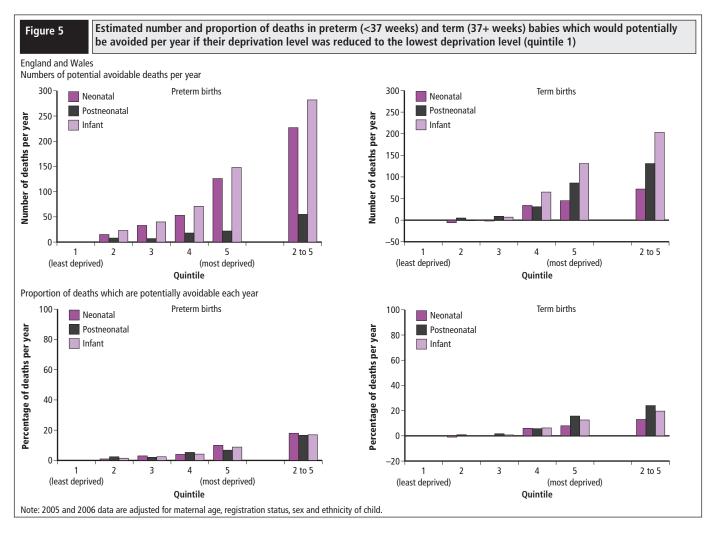
#### Socio-economic status

Among the 10 per cent sample of births where the mother and father were both named on the birth certificate (inside marriage, or outside marriage registered by both parents), and hence where NS-SEC could be coded, there was a highly statistically significant independent association between NS-SEC and risk of infant mortality, similar in both neonatal and postneonatal periods, with a marked trend of increasing risk of infant mortality with decreasing socio-economic class (Tables 4 to 6). As with deprivation, there was evidence of some confounding of this relationship by maternal age, sex of baby, registration status and ethnicity of the baby, but the association remained after adjustment for these factors, though many of the odds ratios were no longer statistically significant. There was strong evidence of an interaction between NS-SEC and (separately) birthweight, gestation and smallness-for-gestational-age, with similar patterns of odds ratios to those observed for deprivation status. However due to small numbers, the results are not presented here.

Association between infant mortality and potential confounding factors/proximal risk factors (birthweight and gestation)

# Confounding factors

As expected, all potential confounding factors under examination



were strongly independently related to risk of infant mortality in both neonatal and postneonatal periods (Tables 4 to 6). Relative risks of mortality tended to be highest for babies where the mother was young (aged under 25) or older (aged 40 and over) (compared with mothers aged 25-29), where the parents were unmarried (particularly where the birth was registered by the mother alone or registered by both parents living at different addresses, compared with those where the parents were married), and where mothers were born outside the UK (compared with those born in the UK). Large increases in mortality odds ratios for babies described by their mother as being of non-White ethnic origin were also observed (relative to those described as White), particularly among those described as Asian or Black, and odds ratios for male babies were consistently high relative to females. None of these effects was substantially reduced in either the neonatal or the postneonatal group, after adjustment for maternal age, baby's sex, registration status, deprivation and ethnicity of the baby (as appropriate). Maternal country of birth could not be included in any models including ethnicity of the baby, due to the high correlation between these factors. However, with the exception of baby's sex, where estimates remained almost unchanged, when birthweight or gestation were additionally included in the models (Model 4 and 5) all odds ratios tended to reduce sharply, though the overall effects of registration status, maternal age and baby's ethnicity remained statistically significant independent predictors of mortality (all P<0.03). Baby's sex was a highly significant independent risk factor for infant mortality, apparently unaffected by any other factor available in this analysis (P<0.005 for neonatal, postneonatal and infant mortality).

#### Registration status:

In the neonatal period, for sole registrations, and joint registrations by parents living at different addresses, there was no apparent raised risk after additional adjustment for birthweight or gestation plus maternal age, deprivation, and sex and ethnicity of the baby. Babies whose parents lived at the same address, but were not married, had an apparent 14 per cent and 20 per cent raised risk compared with those born inside marriage after additionally controlling for these factors (ORs (95 per cent CI): 1.14 (1.04, 1.25) and 1.20 (1.10, 1.31) adjusted for birthweight and gestation respectively) (Table 4). This rather unexpected result may simply be a chance finding due to the large number of statistical estimations involved in the analyses. Conversely, in the postneonatal period there was no apparent increase in risk after additional adjustment for birthweight or gestation among joint registrations with the same address relative to births inside marriage. However, sole registrations had 67 per cent and 82 per cent increased risks even after further adjustment for these more proximal risk factors (ORs (95 per cent CI): 1.67 (1.40, 1.98) and 1.82 (1.52, 2.15) adjusted additionally for birthweight and gestation respectively). Calculations of excess deaths showed that after full adjustment around 5 per cent of all postneonatal deaths were independently associated with sole registration (Table 5). Babies whose parents did not live at the same address also had raised risks of postneonatal mortality (ORs (95 per cent CI): 1.32 (1.11, 1.57) and 1.40 (1.17, 1.66) after additional adjustment for birthweight and gestation respectively) (Table 5).

#### Maternal age:

In the neonatal period, odds ratios for very young (<20 years) and older (40 and over) mothers were raised relative to mothers aged 25-29 after additional adjustment for birthweight, but not statistically significantly so (except for a borderline result for mothers aged 40 and over) (Table 4). In the postneonatal period, odds ratios were significantly raised for both mothers under the age of 25 (ORs (95 per cent CI) 1.38 (1.14, 1.67) and

1.36 (1.18, 1.57) (maternal age <20), and 1.41 (1.16, 1.71) and 1.39 (1.21, 1.61) (maternal age 20-24)), and for those aged 35 and over (ORs (95 per cent CI) 1.23 (1.04, 1.45) and 1.22 (1.04, 1.44) (maternal age 35–39), after adjustment for birthweight or gestation respectively) (Table 5). Over 12 per cent of postneonatal deaths could be independently attributed to maternal age above or below age 25-29 (Table 5).

#### Baby's sex:

Inclusion of other factors in the model made very little difference to the effect of sex. In both neonatal and postneonatal periods relative risk of mortality was consistently and independently raised among boys, risks being around 15–30 per cent higher than those of girls (ORs (95 per cent CI): 1.33 (1.25, 1.43) and 1.12 (1.04, 1.20) in the neonatal period, and 1.28 (1.16, 1.43) and 1.16 (1.06, 1.30) in the postneonatal period, after adjustment for all other risk factors, and birthweight and gestation respectively) (Tables 4 and 5). In the fully adjusted model including birthweight, male sex accounted for 13 per cent of all infant deaths (Table 6).

#### Ethnicity of the baby:

Babies of non-White ethnic origin (as described by the mother) had raised odds ratios in both neonatal and postneonatal periods, though this increase was restricted to babies described as Asian in the latter period. After adjustment for all other potential confounders and birthweight, ORs (95 per cent CIs) in the neonatal period were up to 40 per cent higher compared with babies described as White (1.13 (1.00, 1.27), 1.42 (1.24, 1.64) and 1.11 (0.98, 1.25) in Asian, Black and Other non-White babies respectively). Adjusted additionally for gestation rather than birthweight, the estimates were even higher, at 1.63 (1.44, 1.85), 1.60 (1.40, 1.83) and 1.24 (1.10, 1.40) (Table 4). In the postneonatal period, the corresponding odds ratios for babies described as Asian were 1.49 (1.26, 1.77) and 1.86 (1.57, 2.20) respectively. Around 5 per cent of all infant deaths in the population could be attributed to non-white ethnicity, independent of birthweight (Table 6).

Proximal risk factors (further along the causal pathway than deprivation) gestation and birthweight

Almost 60 per cent of all infant deaths could be attributed to either low birthweight or preterm delivery (Table 6; 'Excess deaths %' column; Model 3).

#### Birthweight:

Birthweight was a major risk factor for infant mortality - having a greater effect than any other factor except for gestation (there is considerable overlap between the two factors) (Tables 4 to 6). Maternal age- and sex-adjusted odds ratios (relative to babies weighing 2,500 to 4,499 grams) rose to around 230 for neonatal mortality and over 45 for postneonatal mortality in extremely low birthweight infants (Tables 4 and 5 respectively). There was, however, a highly statistically significant interaction between the effects of birthweight (low birthweight) and Carstairs deprivation on infant mortality, stronger than that between gestation and deprivation, and present in all time periods (all P≤0.01) (Table 8). The effect of low birthweight on infant mortality decreased markedly with increasing levels of deprivation, adjusted (for maternal age, registration status and baby's sex and ethnicity) odds ratios decreasing by a third between the highest (least deprived) and lowest (most deprived) groups from 28.73 (24.00, 34.39) to 19.46 (17.65, 21.45). The effect modification was most marked in the postneonatal period, where adjusted relative risks reduced by almost half from the least to most deprived groups, but was also substantial in the neonatal period, with a 15 to 30 per cent decline between the effect of low birthweight in the top and bottom deprivation categories. Excess neonatal deaths attributed to low birthweight per year varied between 460 for the most deprived and 121 for the least deprived, and equivalent figures for postneonatal deaths were 121 for the most deprived and 32 for the least deprived (Table 8).

#### Gestational age:

The effect of gestation on risk of infant mortality was higher than that of any other factor by an order of magnitude (Tables 4 to 6), with maternal age- and sex-adjusted odds ratios (relative to term babies) of up to around 650 for neonatal mortality and over 90 for postneonatal mortality in extremely preterm babies (Tables 4 and 5 respectively).

Again, there was a statistically significant interaction between the effects of gestation (preterm delivery) and Carstairs deprivation index on infant mortality (P=0.01) (Table 8). This was driven by a strong interaction between these factors in the postneonatal period (P= 0.03) where there was a 30 per cent decline in adjusted (for maternal age, registration status and baby's sex and ethnicity) odds ratios between the least deprived (10.13 (7.28, 14.09)) and most deprived (7.40 (6.27, 8.74)) groups. There was a large number of postneonatal deaths attributed to preterm delivery in the most deprived group (104) compared to that in the least deprived group (27), although the proportions of these excess deaths were similar at around 30 per cent for each group. There was no evidence of an interaction between gestation and deprivation in the neonatal period (P=0.14) (Table 8).

#### Small-for-gestational-age:

Live-born babies whose weight was below the 5th centile of the distribution expected for their sex and gestation had more than three times the risk of dying in the first year of life as those whose weight was above this level (Tables 4 to 6). Though the odds ratios in all age periods were high, their magnitude in relation to those of absolute birthweight and gestation indicate that the latter are the more critical factors in terms of predicting risk of infant death. As with gestation and birthweight, there was an interaction between the effects of small-for-gestational-age and Carstairs deprivation on infant mortality, most marked in the postneonatal period where the effect of being small-for-gestational-age in the least deprived group was almost double that in the most deprived group (OR (95 per cent CI): 5.40 (3.49, 8.36) versus 2.88 (2.35, 3.53)). The excess number of postneonatal deaths attributed to small-for-gestationalage ranged between 11 for the least deprived group and 42 for the most deprived group (Table 8).

#### Discussion

Inequalities in infant mortality in England and Wales in 2005 and 2006 by deprivation, and socio-economic and biological factors was investigated in this analysis. Multivariate analysis was undertaken to obtain information on the independent contributions of individual factors to the risk of death. The approach was to use both the strength of association between various factors and risk of death, measured as adjusted odds ratios, and the theoretical impact of these factors on the number and proportion of 'excess' deaths in the population that could potentially be saved, or avoided, if the factor was not present.

The emphasis of this work was on socio-economic inequalities in infant mortality. The national health inequalities PSA target was set in 2001 and was updated in 2004. The target is to 'Reduce health inequalities by 10 per cent by 2010 as measured by infant mortality and life expectancy at birth'.6 More specifically this target is underpinned by an objective to reduce by at least 10 per cent the gap in mortality between the routine and manual group and the population as a whole by 2010. A second objective is, also by 2010, to reduce by at least 10 per cent the gap in life expectancy at birth between the fifth of areas with the worst health and deprivation indicators (the Spearhead Group) and the population as a whole.

This analysis was limited by the number of births coded to NS-SEC (only 10 per cent are coded), and hence the main analyses were on the area level of deprivation, using the Carstairs index of deprivation. Despite its limitations, this indicator performed well as a marker for socio-economic status and, since it was available for all births and deaths, it enabled multivariate analyses to be conducted with good statistical power.

The findings are consistent with those of other studies based on England and Wales data that have found variation in infant mortality by birthweight, gestation, mother's age, registration status, sex and ethnicity of the child. 4,10,11 The very strong links between risk of neonatal death and both low birthweight and preterm delivery were confirmed. After controlling for maternal age and sex of child, over two-thirds of all neonatal deaths could be attributed to these factors in 2005 and 2006. The multivariate approach also confirmed independent effects of registration status, ethnicity of child, and male sex, after allowing for birthweight or gestational age and other confounding factors.

Low birthweight and preterm delivery are the main proximal risk factors for death under one year, especially under one month, and both factors are highly associated with socio-economic status and deprivation. The finding that approximately one quarter of all infant deaths could be avoided if all mothers had the lowest deprivation level is at least partly mediated through low birthweight and/or prematurity. When the impact of deprivation on risk of death over and above that of low birthweight and preterm delivery was examined, perhaps not surprisingly, evidence of interaction was found. Stratified analyses, for example within term babies only, revealed residual associations between deprivation and risk of death with a consistent pattern of stronger associations for babies who were not low birthweight, and term, than for babies who were low birthweight or preterm. These associations were much stronger for postneonatal than for neonatal death, an observation reported in a previous study by Leon. 18 An estimated 28 per cent of postneonatal deaths in babies weighing 2,500 grams or more, equivalent to 145 postneonatal deaths per year in this group, could be explained by deprivation levels being above the lowest (least deprived) level. Similarly, 24 per cent of postneonatal deaths in term babies, equivalent to 131 postneonatal deaths among term babies per year, were explained by deprivation levels above the least deprived category. These deaths are potentially avoidable if the deprivation gap is eliminated for all mothers.

There was no information from birth, death or NN4B records on more specific, and possibly more proximal, environmental, social, economic or behavioural exposures related to deprivation, or on access to and uptake of health or social care, which may explain these findings. Smoking, obesity, poor nutrition and inadequate antenatal care are likely to be more prevalent in groups of mothers with higher levels of deprivation, and may contribute to the effects reported here. Further work to investigate the explanatory causal pathways between deprivation and infant mortality, perhaps for specific causes of death, is required. This will inform interventions to reduce the inequalities in infant death described in this paper and thus contribute towards reducing the infant mortality gap identified in the PSA target.6

# **Key** findings

- Among all singleton live births in England and Wales, 2005–06, as levels of deprivation increased, mothers tended to be younger, less likely to be married or living with the father, and more likely to be non-UK-born
- Babies who were Asian, Black or of other non-white ethnicity tended to have higher levels of deprivation
- · Preterm birth, low birthweight and small-for-gestational-age increased in prevalence with increasing levels of deprivation
- · Deprivation, births outside marriage, non-white ethnicity of infant, maternal age under the age of 20 and male sex were all independently associated with an increased risk of infant mortality
- Lower birthweight, earlier gestation and small-for-gestational-age were strongly associated with infant mortality, with the effect of birthweight and gestation most marked in the neonatal period
- The effect of deprivation on death in the first year of life was partly confounded by maternal age, registration status/marital type and ethnicity of baby, but a trend of increasing risk of death with increasing deprivation persisted after adjustment. One quarter of all deaths under one year would potentially be avoided if all births were to women with the lowest level of deprivation
- Trends of increasing mortality risk with increasing deprivation tended to be strongest in the postneonatal period
- Statistically significant interactions were found between deprivation and low birthweight, preterm and small-forgestational-age. There was a stronger association between deprivation and infant mortality among babies who were not low birthweight, and among non-small-for-gestational-age babies. With regard to gestation, the strongest effect of deprivation was observed among preterm births in the neonatal period, and among term births in the postneonatal period
- Around 28 per cent of postneonatal deaths (45 per year) among non-low birthweight babies, and 24 per cent of postneonatal deaths (131 per year) among babies born at term, would potentially be avoided if all levels of deprivation were reduced to that of the least deprived group
- The National Statistics Socio-economic Classification showed similar associations with infant mortality to deprivation, but small numbers prohibited the calculation of fully adjusted models using this factor

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#### Table 4

#### Unadjusted and adjusted odds ratios (OR) for the association between social and biological risk factors and singleton neonatal mortality, 2005-06

England and Wales									
	Unadjusted	Model 1	Neonatal mo Model 2	rtality Model 3	Model 4	Model 5	-		
		Adjusted for age and sex <sup>1</sup>	Model 1 + adjusted for deprivation and registration status <sup>1,2</sup>	Model 2 + adjusted for ethnicity of baby <sup>1</sup>	Model 3 + adjusted for interaction between birthweight and deprivation	Model 3 + adjusted for interaction between gestation and deprivation	Exces	s deaths	s %³
Donrivation and co	OR (95% CI) ocio-economic status	OR (95% CI)	OR (95% CI)	OR (95% CI)	OR (95% CI)	OR (95% CI)		-	
Carstairs deprivation									
1 (least deprived) 2 3 4 5 (most deprived)	1.00 1.14 (1.00, 1.30)* 1.29 (1.14, 1.47)*** 1.62 (1.43, 1.82)*** 2.00 (1.79, 2.24)	1.00 1.13 (0.99, 1.30) 1.27 (1.12, 1.44)*** 1.57 (1.39, 1.77)*** 1.94 (1.73, 2.17)***	1.00 1.12 (0.98, 1.28) 1.25 (1.10, 1.42)** 1.53 (1.36, 1.72)*** 1.90 (1.69, 2.13)***	1.00 1.07 (0.92, 1.23) 1.19 (1.04, 1.37)* 1.41 (1.24, 1.61)*** 1.54 (1.36, 1.75)***	Pooled odds ratios not calculated due to interaction. See Table 8	Pooled odds ratios not calculated due to interaction. See Table 8	Model 3 0.84 2.61 6.85 12.85 23.16		
NS-SEC Higher and professional Intermediate Routine and manual Other (unemployed/	1.04 (0.75, 1.45) 1.54 (1.20, 1.98)*	1.00 1.05 (0.75, 1.47) 1.58 (1.22, 2.06)*	1.00 1.03 (0.74, 1.44) 1.53 (1.17, 2.00)*	1.00 1.00 (0.70, 1.43) 1.33 (1.00, 1.77)	Pooled odds ratios not calculated due to interaction with	Pooled odds ratios not calculated due to interaction with	Model 3 - 11.06		
students/ns)	2.17 (1.47, 3.18)***	2.19 (1.47, 3.27)***	2.18 (1.46, 3.26)***	1.67 (1.08, 2.58)*	birthweight	gestation	4.12 15.18		
Potential confound Registration type/Ma Married		1.00	1.00	1.00	1.00	1.00	Model 31	Model 4	Model 5
Sole registration <sup>4</sup>	1.34 (1.18, 1.51)***	1.24 (1.10, 1.41)**	1.14 (1.00. 1.29)*	1.26 (1.10, 1.45)*	0.86 (0.75, 1.00)*	0.93 (0.81, 1.07)			-0.62
Joint registration/same address <sup>5</sup>	1.25 (1.16, 1.34)***	1.20 (1.11, 1.29)***	1.18 (1.09, 1.28)***	1.36 (1.24, 1.48)***	1.14 (1.04, 1.25)*	1.20 (1.10, 1.31)***	7.99	3.71	5.03
Joint registration/ different address <sup>5</sup>	1.50 (1.35, 1.66)***	1.39 (1.24, 1.55)***	1.28 (1.14, 1.43)***	1.43 (1.27, 1.62)***	1.05 (0.93, 1.19)	1.13 (1.00, 1.28)	3.53 13.21	0.56 2.93	1.35 5.76
Maternal age (years) <20 20–24 25–29	1.28 (1.14, 1.45)*** 1.05 (0.96, 1.15) 1.00	1.28 (1.14, 1.44)*** 1.05 (0.95, 1.15) 1.00	1.10 (0.97, 1.25) 0.95 (0.86, 1.05) 1.00	1.16 (1.01, 1.32)* 0.96 (0.89, 1.06) 1.00	1.10 (0.95, 1.26) 0.94 (0.85, 1.04) 1.00	1.13 (0.99, 1.30) 0.98 (0.88, 1.08) 1.00	Model 3   1.28 -0.87	0.85	Model 5 1.07 –0.43 –
30–34 35–39 40 and over	0.81 (0.74, 0.89)*** 0.88 (0.79, 0.97) 1.20 (1.01, 1.41)	0.81 (0.74, 0.89)*** 0.88 (0.79, 0.97)* 1.20 (1.01, 1.41)*	0.89 (0.81, 0.97)* 0.97 (0.88, 1.08) 1.31 (1.11, 1.54)**	0.92 (0.84, 1.02) 1.03 (0.92, 1.15) 1.41 (1.18, 1.68)***	0.94 (0.85, 1.03) 0.98 (0.87, 1.10) 1.21 (1.00, 1.45)*	0.90 (0.81, 0.99)* 0.95 (0.85, 1.06) 1.17 (0.98, 1.41)	0.43 1.24	-1.54 -0.30 0.74 -1.59	-2.67 -0.77 0.42 -2.38
Maternal country of UK Non-UK	birth 1.00 1.26 (1.17, 1.35)***	1.00 1.29 (1.19, 1.38)***	1.00 1.21 (1.12, 1.31)***	Unable to be calculated due to high correlation of country of birth & ethnicity	Unable to be calculated due to high correlation of country of birth & ethnicity	Unable to be calculated due to high correlation of country of birth & ethnicity	Model 2 4.43		
Sex of baby Female Male	1.00 1.22 (1.14, 1.30)***	1.00 1.22 (1.14, 1.30)***	1.00 1.22 (1.14, 1.30)***	1.00 1.22 (1.14, 1.30)***	1.00 1.33 (1.25,1.43)***	1.00 1.12 (1.04,1.20)***	Model 3 10.08		Model 5 5.69
Ethnicity of baby White	1.00	1.00	1.00		1.00	1.00	Model 2	Model 4	Model 5
Asian (Bangladeshi/ Pakistani/ Indian)	1.76 (1.58, 1.95)***	1.77 (1.60, 1.97)***	1.83 (1.62, 2.06)***	Equivalent to Model 2	1.13 (1.00, 1.27)	1.63 (1.44, 1.85)***	5.92	1.50	5.04
Black African/ Black Caribbean Other	1.99 (1.75, 2.26)*** 1.34 (1.20, 1.50)***	1.99 (1.76, 2.26)*** 1.34 (1.19, 1.50)***	1.74 (1.52, 1.99)*** 1.28 (1.14, 1.44)***	1	1.42 (1.24, 1.64)*** 1.11 (0.98, 1.25)	1.60 (1.40, 1.83)*** 1.24 (1.10, 1.40)***	3.50 2.29 11.71	2.43 1.04 4.97	3.08 2.03 10.16
Factors on causal p Birthweight (grams) <1,500 1,500–2,499 2,500–4,499 4,500 and over	232.46 (215, 251.11)*** 8.70 (7.81, 9.68)* 1.00 0.99 (0.63, 1.56)	231.70 (214.46, 250.33)*** 8.71 (7.82, 9.70)*** 1.00 0.96 (0.61, 1.52)	Interaction between birthweight and deprivation. See Table 7	Interaction between birthweight and deprivation. See Table 7	Interaction between birthweight and deprivation. See Table 7	Interaction between birthweight and deprivation. See Table 7	Model 1 56.94 12.15 - -0.02		
Gestation (weeks) 21–27 28–31 32–36 37–41 42 and over	654.82 (602.00, 712. 27)*** 43.24 (38.15, 49.00)*** 7.20 (6.42, 8.07)*** 1.00 1.02 (0.77, 1.35)	649.46 (596.98, 706.56)*** 42.74 (37.71, 48.33)*** 7.14 (6.47, 7.99)*** 1.00 1.01 (0.77, 1.34)	Interaction between gestation and deprivation. See Table 7	Interaction between gestation and deprivation. See Table 7	Interaction between gestation and deprivation. See Table 7	Interaction between gestation and deprivation. See Table 7	69.06 Model 1 50.03 8.91 9.79 - 0.01 68.74		
Small for gestational SGA (<5%) Not SGA (≥5%)	age 3.26 (2.96, 3.59)*** 1.00	3.19 (2.89, 3.51)*** 1.00	Interaction between SGA and deprivation. See Table 7	Interaction between SGA and deprivation. See Table 7	Interaction between SGA and deprivation. See Table 7	Interaction between SGA and deprivation. See Table 7	Model 1 9.83		

<sup>\*\*</sup> p <0.005 \*\*\* p <0.001

Adjusted for covariate only where it is not the main explanatory variable in model.
 Carstairs deprivation not adjusted for where NS-SEC is main explanatory variable.

Proportion of all deaths in population which potentially would be avoided if risk was same as that in baseline group. Birth takes place outside marriage, only mother registers the birth.

Birth takes place outside marriage, both mother and father register the birth.

## Table 5

#### Unadjusted and adjusted odds ratios (OR) for the association between social and biological risk factors and singleton postneonatal mortality, 2005-06

England and Wales				. 10.					
-	Unadjusted	Model 1	Postneonatal m Model 2	Model 3	Model 4	Model 5	-		
	Î	Adjusted for age and sex <sup>1</sup>	Model 1 + adjusted for deprivation and registration status <sup>1,2</sup>	Model 2 + adjusted for ethnicity of baby <sup>1</sup>	Model 3 + adjusted for interaction between birthweight and deprivation	Model 3 + adjusted for interaction between gestation and deprivation	Exce	ss death	IS % <sup>3</sup>
Danimetian and as di	OR (95% CI)	OR (95% CI)	OR (95% CI)	OR (95% CI)	OR (95% CI)	OR (95% CI)	L		
Deprivation and socion Carstairs deprivation in									
1 (least deprived) 2 3 4 5 (most deprived)	1.00 1.18 (0.97, 1.44) 1.32 (1.09, 1.60)** 1.71 (1.43, 2.04)*** 2.21 (1.88, 2.61)***	1.00 1.15 (0.94, 1.40) 1.25 (1.03, 1.50)* 1.55 (1.30, 1.86)*** 1.97 (1.66, 2.34)***	1.00 1.13 (0.93, 1.38) 1.22 (1.00, 1.47)* 1.49 (1.24, 1.78)*** 1.86 (1.57, 2.20)***	1.00 1.16 (0.94, 1.43) 1.18 (0.97, 1.45) 1.40 (1.16, 1.70)** 1.62 (1.35, 1.96)***	Pooled odds ratios not calculated due to interaction. See Table 8	Pooled odds ratios not calculated due to interaction. See Table 8	Model 3 1.73 2.41 6.72 14.62 25.48		
NS-SEC Higher and professional Intermediate Routine and manual Other (unemployed/ students/ns)	1.00 1.76 (1.07, 2.88)* 1.72 (1.12, 2.64)* 2.63 (1.41, 4.91)*	1.00 1.59 (0.96, 2.62) 1.41 (0.90, 2.21) 1.96 (1.02, 3.78)*	1.00 1.61 (0.98, 2.66) 1.44 (0.91, 2.27) 2.00 (1.04, 3.87)*	1.00 1.59 (0.94, 2.67) 1.39 (0.86, 2.24) 1.82 (0.91, 3.65)	Pooled odds ratios not calculated due to interaction with birthweight	Pooled odds ratios not calculated due to interaction with gestation	Model 3 8.37 11.81 4.74 24.93		
Potential confounding Registration type/Marit									
Married Sole registration <sup>4</sup>	1.00 2.38 (2.06, 2.74)***	1.00 1.96 (1.68, 2.29)***	1.00 1.82 (1.56, 2.12)***	1.00 2.07 (1.75, 2.46)***	1.00 1.67 (1.40, 1.98)***	1.00 1.82 (1.52, 2.15)***	Model 3 7.05	Model 4 5.47	Model 6.15
Joint registration/ same address <sup>5</sup>	1.13 (1.01, 1.27)*	1.00 (0.89, 1.13)	1.00 (0.89, 1.13)	1.20 (1.05, 1.37)*	1.09 (0.95, 1.25)	1.14 (1.00, 1.30)	4.34	2.15	3.20
Joint registration/ different address <sup>5</sup>	1.74 (1.51, 2.02)***	1.42 (1.21, 1.67)***	1.34 (1.14, 1.57)***	1.59 (1.34, 1.89)***	1.32 (1.11, 1.57)*	1.40 (1.17, 1.66)***	4.81 16.20	3.14 10.77	3.70 13.05
Maternal age (years)							Model 3		
<20 20–24	1.79 (1.52, 2.13)*** 1.53 (1.34, 1.74)***	1.79 (1.52, 2.13)*** 1.53 (1.34, 1.74)***	1.38 (1.16, 1.66)*** 1.35 (1.18, 1.54)***	1.40 (1.16, 1.70)*** 1.36 (1.18, 1.57)***	1.38 (1.14, 1.67)** 1.36 (1.18, 1.57)***	1.41 (1.16, 1.71)*** 1.39 (1.21, 1.61)***	3.22 7.00	3.10 7.00	3.28 7.42
25–29 30–34 35–39 40 and over	1.00 0.82 (0.71, 0.94)* 1.02 (0.87, 1.19) 1.04 (0.79, 1.37)	1.00 0.82 (0.71, 0.94)* 1.02 (0.87, 1.19) 1.04 (0.79, 1.37)	1.00 0.89 (0.77, 1.03) 1.13 (0.97, 1.32) 1.06 (0.80, 1.41)	1.00 0.96 (0.83, 1.11) 1.27 (1.08, 1.50)** 1.10 (0.81, 1.50)	1.00 0.97 (0.83, 1.12) 1.23 (1.04, 1.45)* 1.03 (0.76, 1.40)	1.00 0.95 (0.82, 1.11) 1.22 (1.04, 1.44)* 1.01 (0.74, 1.37)	-0.88 3.15 0.29	-0.65 2.77 0.09	-1.11 2.67 0.03
							12.78	12.32	12.29
Maternal country of bir UK Non-UK	th 1.00 1.09 (0.97, 1.22)	1.00 1.15(1.03, 1.28)*	1.00 1.04 (0.92, 1.18)	Unable to be calculated due to high correlation between country of birth & ethnicity	Unable to be calculated due to high correlation between country of birth & ethnicity	Unable to be calculated due to high correlation between country of birth & ethnicity	Model 2 0.88		
Sex of baby Female Male	1.00 1.23 (1.12, 1.37) ***	1.00 1.23 (1.12, 1.37) ***	1.00 1.23 (1.12, 1.37) ***	1.00 1.22 (1.10,1.35)***	1.00 1.28 (1.16, 1.43)***	1.00 1.16 (1.06,1.30)*	Model 3 10.17	Model 4 12.43	Model ! 7.91
Ethnicity of baby White	1.00	1.00	1.00		1.00	1.00	Model 2	Model 4	Model
Asian (Bangladeshi/ Pakistani/ Indian)	1.86 (1.62, 2.15)***	1.89 (1.63, 2.18)***	1.95 (1.65, 2.30)***	Equivalent to Model 2	1.49 (1.26, 1.77)***	1.86 (1.57, 2.20)***	6.89	4.65	6.54
Black African/ Black Caribbean Other	1.44 (1.17, 1.77)** 1.24 (1.05, 1.46)*	1.47 (1.19, 1.81)*** 1.24 (1.05, 1.47)*	1.17 (0.95, 1.45) 1.13 (0.95, 1.34)		1.09 (0.88, 1.36) 1.06 (0.90, 1.26)	1.14 (0.92, 1.41) 1.12 (0.94, 1.32)	0.88 1.14 8.92	0.50 0.56 5.72	2.81 1.22 10.58
Factors on causal pat Birthweight (grams)	hway						Model 1		
<1,500 1,500–2,499 2,500–4,499	46.41 (41.23, 52.24)*** 6.10 (5.39, 6.89)*** 1.00	45.55 (40.46, 51.29)*** 5.96 (5.27, 6.75)*** 1.00	Interaction between birthweight and	Interaction between birthweight and	Interaction between birthweight and	Interaction between birthweight and	21.74 16.13		
4,500 and over	0.53 (0.28, 0.99)*	0.53 (0.28, 0.99)*	ueprivation. See Table /	deprivation. See Table 7	ueprivation. See Table /	ueprivation. See Table /	-0.50 37.37		
Gestation (weeks) 21–27 28–31 32–36 37–41	95.45 (82.76, 110.78)*** 17.00 (14.16, 20.39)*** 4.76 (4.17, 5.43)*** 1.00	93.40 (80.95, 107.76)*** 16.45 (13.70, 19.74)*** 4.66 (4.08, 5.32)*** 1.00	Interaction between gestation and deprivation. See Table 7	Interaction between gestation and deprivation. See Table 7	Interaction between gestation and deprivation. See Table 7	Interaction between gestation and deprivation. See Table 7	Model 1 14.16 7.07 12.54		
42 and over	0.88 (0.66, 1.19)	0.88 (0.65, 1.18)					-0.36 33.41		
Small for gestational ag SGA (<5%) Not SGA (≥5%)	3.29 (2.96, 3.66)*** 1.00	3.13 (2.83, 3.51)*** 1.00	Interaction between SGA and deprivation. See Table 7	Interaction between SGA and deprivation. See Table 7	Interaction between SGA and deprivation. See Table 7	Interaction between SGA and deprivation. See Table 7	Model 1 11.45		

<sup>\*</sup> p <0.05 \*\* p <0.005 \*\*\* p <0.001

1 Adjusted for covariate only where it is not the main explanatory variable in model.

2 Carstairs deprivation not adjusted for where NS-SEC is main explanatory variable.

3 Proportion of all deaths in population which potentially would be avoided if risk was same as that in baseline group.

4 Birth takes place outside marriage, only mother registers the birth.

5 Birth takes place outside marriage, both mother and father register the birth.

## Table 6

Unadjusted and adjusted odds ratios (OR) for the association between social and biological risk factors and singleton infant mortality, 2005-06

Eligialiu aliu wales			Infant mort	ality			T
	Unadjusted	Model 1	Model 2	Model 3	Model 4	Model 5	-
	,	Adjusted for age and sex <sup>1</sup>	Model 1 + adjusted for deprivation and registration status <sup>1,2</sup>	Model 2 + adjusted for ethnicity of baby <sup>1</sup>	Model 3 + adjusted for interaction between birthweight and deprivation	Model 3 + adjusted for interaction between gestation and deprivation	Excess deaths % <sup>3</sup>
	OR (95% CI)	OR (95% CI)	OR (95% CI)	OR (95% CI)	OR (95% CI)	OR (95% CI)	1
Deprivation and soc Carstairs deprivation i 1 (least deprived)		1.00	1.00	1.00			Model 3
2 3 4 5 (most deprived)	1.15 (1.03, 1.29)* 1.30 (1.17, 1.45)*** 1.65 (1.49, 1.81)*** 2.07 (1.88, 2.27)***	1.14 (1.02, 1.27)* 1.26 (1.14, 1.40)*** 1.57 (1.42, 1.73)*** 1.95 (1.77, 2.14)***	1.13 (1.01, 1.26)* 1.24 (1.11, 1.38)*** 1.52 (1.37, 1.68)***	1.09 (0.97, 1.23) 1.19 (1.06, 1.33)** 1.41 (1.27, 1.57)*** 1.57 (1.41, 1.75)***	Pooled odds ratios not calculated due to interaction. See Table 8	Pooled odds ratios not calculated due to interaction. See Table 8	1.06 2.58 6.84 13.49 23.97
NS-SEC Higher and professional Intermediate Routine and manual Other (unemployed/ students/ns)	1.00 1.22 (0.93, 1.61) 1.58 (1.27, 1.97)*** 2.28 (1.64, 3.17)***	1.00 1.20 (0.91, 1.58) 1.53 (1.22, 1.92)*** 2.13 (1.51, 2.99)***	1.00 1.18 (0.90, 1.56) 1.50 (1.19, 1.89)* 2.13 (1.51, 3.00)***	1.00 1.16 (0.86, 1.55) 1.34 (1.05, 1.72)* 1.71 (1.18. 2.47)**	Pooled odds ratios not calculated due to interaction with birthweight	Pooled odds ratios not calculated due to interaction with gestation	Model 3 2.44 11.13 4.29
Potential confoundi							17.87
Registration type/Mari Married Sole registration <sup>4</sup>	tal status 1.00 1.66 (1.51, 1.82)***	1.00 1.48 (1.35, 1.64)***	1.00 1.36 (1.23, 1.5)***	1.00 1.52 (1.37, 1.69)***	1.00 1.11 (1.00, 1.24)	1.00 1.20 (1.07, 1.33)**	Model 3 Model 4 Model 5 3.40 0.98 1.65
Joint registration/ same address <sup>5</sup>	1.21 (1.28, 1.29)***	1.13 (1.06, 1.21)***	1.12 (1.05, 1.20)***	1.31 (1.21, 1.40)***	1.13 (1.04, 1.21)**	1.18 (1.10, 1.27)***	6.83 3.32 4.40
Joint registration/ different address <sup>5</sup>	1.57 (1.45, 1.71)***	1.40 (1.28, 1.53)***	1.30 (1.18, 1.42)***	1.48 (1.23, 1.64)***	1.14 (1.03, 1.26)*	1.21 (1.09, 1.34)***	3.93 1.49 2.11 14.16 5.79 8.16
Maternal age (years) <20 20–24	1.43 (1.30, 1.58)*** 1.19 (1.10, 1.28)***	1.43 (1.30, 1.58)*** 1.19 (1.10, 1.28)***	1.07 (0.99, 1.15)	1.23 (1.10, 1.38)*** 1.08 (1.00, 1.17)	1.19 (1.06, 1.33)** 1.07 (0.98, 1.17)	1.22 (1.09, 1.37)** 1.11 (1.02, 1.20)*	Model 3 Model 4 Model 5 1.86 1.58 1.79 1.68 1.48 2.25
25–29 30–34 35–39 40 and over	1.00 0.81 (0.75, 0.88)*** 0.92 (0.84, 1.00) 1.15 (1.00, 1.33)	1.00 0.81 (0.75, 0.88)*** 0.92 (0.84, 1.00) 1.15 (1.00, 1.33)	1.00 0.89 (0.82, 0.96)** 1.02 (0.94, 1.11) 1.23 (1.07, 1.43)**	1.00 0.93 (0.86, 1.01) 1.10 (1.00, 1.21)* 1.32 (1.14, 1.54)***	1.00 0.95 (0.87, 1.03) 1.16 (0.99, 1.36) 1.11 (0.94, 1.31)	1.00 0.91 (0.84, 0.99)* 1.13 (0.97, 1.33) 1.12 (0.94, 1.33)	
Maternal country of bi UK Non-UK	rth 1.00 1.20 (1.13, 1.28)***	1.00 1.24 (1.68, 1.32)***	1.00 1.16 (1.08. 1.23)***	Unable to calculate due to high correlation between country of birth & ethnicity	Unable to calculate due to high correlation between country of birth & ethnicity	Unable to calculate due to high correlation between country of birth & ethnicity	Model 2 3.40
Sex of baby Female Male	1.00 1.22 (1.16, 1.28)***	1.00 1.22 (1.16, 1.28)***	1.00 1.22 (1.16, 1.30)***	1.00 1.22 (1.15,1.28)***	1.00 1.32 (1.25,1.41)***	1.00 1.14 (1.08,1.20)***	Model 3 Model 4 Model 5 10.11 13.48 5.97
Ethnicity of baby White	1.00	1.00	1.00		1.00	1.00	Model 2 Model 4 Model 5
Asian (Bangladeshi/ Pakistani/ Indian)	1.79 (1.65, 1.95)***	1.81 (1.67, 1,97)***	1.87 (1.69, 2.05)***	Equivalent to Model 2	1.24 (1.12, 1.37)***	1.71 (1.55, 1.89)***	6.24 2.60 5.57
Black African/ Black Caribbean Other	1.81 (1.62, 2.01)*** 1.31 (1.19, 1.43)***	1.82 (1.64, 2.03)*** 1.31 (1.18, 1.43)***	1.54 (1.38, 1.73)*** 1.23 (1.12, 1.35)***	·	1.31 (1.17, 1.48)*** 1.10 (0.99, 1.21)	1.44 (1.29, 1.62)*** 1.44 (1.09, 1.32)***	2.64 1.78 2.30 1.93 0.94 3.15 10.80 5.31 11.01
Factors on causal pa Birthweight (grams) <1,500	144.84 (136.14, 154.09)***	1/3 70 /125 05 152 00\***					Model 1 45.43
1,500–2,499 2,500–4,499	7.43 (6.85, 8.05)*** 1.00	7.38 (6.80, 8.00)*** 1.00	Interaction between birthweight and deprivation. See Table 7	Interaction between birthweight and deprivation. See Table 7	Interaction between birthweight and deprivation. See Table 7	Interaction between birthweight and deprivation. See Table 7	13.47
4,500 and over	0.76 (0.53, 1.10)	0.75 (0.52, 1.08)					-0.18 58.72
Gestation (weeks) 21–27 28–31 32–36 37–41	401.03 (373.53, 430.55)*** 30.32 (27.36, 33.60)*** 5.98 (5.49, 6.52)*** 1.00	396.52 (369.27, 425.79)*** 29.74 (26.83, 32.96)*** 5.90 (5.41, 6.43)*** 1.00	Interaction between gestation and deprivation. See Table 7	Interaction between gestation and deprivation. See Table 7	Interaction between gestation and deprivation. See Table 7	Interaction between gestation and deprivation. See Table 7	Model 1 38.44 8.32 10.68
42 and over	0.95 (0.78, 1.16)	0.95 (0.77, 1.16)					-0.10 57.35
Small for gestational a SGA (<5%) Not SGA (≥5%)	3.95 (3.49, 4.48)*** 1.00	3.79 (3.34, 4.29)*** 1.00	Interaction between SGA and deprivation. See Table 7	Interaction between SGA and deprivation. See Table 7	Interaction between SGA and deprivation. See Table 7	Interaction between SGA and deprivation. See Table 7	Model 1 11.18 –

<sup>\*\*</sup> p < 0.005 \*\*\* p <0.001

Adjusted for covariate only where it is not the main explanatory variable in model. Carstairs deprivation not adjusted for where NS-SEC is main explanatory variable.

Proportion of all deaths in population which potentially would be avoided if risk was same as that in baseline group.

Birth takes place outside marriage, only mother registers the birth.

Birth takes place outside marriage, both mother and father register the birth.

Table 7

Adjusted odds ratios for the effect of deprivation on neonatal, postneonatal and infant mortality, stratified by birthweight, gestation and small for gestational age status, 2005-06.

Population excess deaths (number and proportion) attributed to deprivation also presented.

			Total excess					
	Quintile 1         Quintile 2         Quintile 3         Quintile 4         Quintile 5         dea							
	OR (95% CI) <sup>1</sup>	OR (95% CI) 1	OR (95% CI) 1	OR (95% CI) 1	OR (95% CI) 1	& %4		
leonatal mortality	'						,	
Birthweight	1.00	1.00 (0.00, 1.30)	1 12 (0 00 1 25)	1 12 /0 05 1 21\	1 14 (0 07 1 24)			
ow birthweight (<2,500 grams) Jumber excess deaths per year <sup>3</sup>	1.00	1.08 (0.90, 1.29) 12	1.13 (0.96, 1.35) 25	1.12 (0.95, 1.31) 32	1.14 (0.97, 1.34) 58	127	0.01	
% excess deaths per year⁴	_	0.96	1.96	2.48	4.56	9.96	0.0.	
Not low birthweight (≥2,500 grams)	1.00	0.92 (0.72, 1.19)	0.97 (0.76, 1.23)	1.31 (1.05, 1.64)*	1.35 (1.09, 1.67)*			
Not low birthweight (=2,500 grains)  Number excess deaths per year <sup>3</sup>	-	-6	-3	31	44	66		
% excess deaths per year⁴	-	-1.22	-0.48	5.84	8.47	12.61		
Gestation								
reterm (<37 weeks)	1.00	1.10 (0.92, 1.32)	1.18 (1.00, 1.40)	1.22 (1.04, 1.44)*	1.35 (1.15, 1.58)***	227	0.44	
Number excess deaths per year³ % excess deaths per year⁴	_	15 1.15	33 2.59	53 4.13	126 9.74	227 17.60	0.14	
. ,								
erm (≥37 weeks) Jumber excess deaths per year³	1.00	0.93 (0.72, 1.20) -6	0.98 (0.77, 1.25) -2	1.34 (1.07, 1.67)*** 34	1.33 (1.07, 1.65)* 45	72		
6 excess deaths per year	_	-1.03	-0.31	6.34	8.39	13.38		
. ,								
imall for gestational age imall for gestational age (<5%)	1.00	0.97 (0.67, 1.40)	1.16 (0.83, 1.62)	1.13 (0.82, 1.56)	1.28 (0.94, 1.73)			
lumber excess deaths per year <sup>3</sup>	_	` <b>–</b> 1	5	6	23	33	0.46	
6 excess deaths per year4	_	-0.32	2.14	2.56	9.56	13.94		
Not small for gestational age (≥5%)	1.00	1.09 (0.92, 1.28)	1.17 (1.00, 1.38)*	1.44 (1.24, 1.67)***	1.43 (1.32, 1.78)***			
lumber excess deaths per year <sup>3</sup>		16 1.12	34	105 7.34	148	303		
% excess deaths per year⁴	_	1.12	2.41	7.34	10.42	21.28		
Postneonatal mortality								
irthweight ow birthweight (<2,500 grams)	1.00	1.13 (0.81, 1.57)	1.00 (0.73, 1.38)	1.04 (0.77, 1.40)	1.00 (0.75, 1.33)			
lumber excess deaths per year <sup>3</sup>	_	6	` - '	3		9	< 0.00	
6 excess deaths per year4	_	1.51	-	0.93	-	2.45		
lot low birthweight (≥2,500 grams)	1.00	1.12 (0.85, 1.47)	1.21 (0.93, 1.57)	1.44 (1.12, 1.84)*	1.83 (1.44, 2.32)***			
Jumber excess deaths per year³ 6 excess deaths per year⁴		7 1.32	15 2.84	36 7.01	88 17.17	145 28.34		
o excess deaths per year	_	1.32	2.04	7.01	17.17	20.34		
Gestation	1.00	1 22 (0 99 1 75)	1 15 (0 92 1 61)	1 27 (0 02 1 7/1)	1 22 (0 01 1 67)			
reterm (<37 weeks) Iumber excess deaths per year <sup>3</sup>	1.00	1.22 (0.88, 1.75) 8	1.15 (0.82, 1.61) 7	1.27 (0.93, 1.74) 18	1.23 (0.91, 1.67) 22	55	0.03	
% excess deaths per year4	_	2.43	2.09	5.31	6.80	16.64		
erm (≥37 weeks)	1.00	1.08 (0.83, 1.41)	1.12 (0.87, 1.45)	1.34 (1.06, 1.71)*	1.68 (1.34, 2.12)***			
lumber excess deaths per year <sup>3</sup>	-	5	9	31	86	131		
% excess deaths per year⁴	_	0.90	1.68	5.73	15.72	24.04		
mall for gestational age								
mall for gestational age (<5%) Number excess deaths per year <sup>3</sup>	1.00	1.00 (0.66, 1.52)	0.73 (0.48, 1.10)	0.95 (0.65, 1.37)	0.99 (0.69, 1.40) -1	-10	0.02	
6 excess deaths per year	_	-	–8 –5.14	−2 −1.18	-0.44	-6.76	0.02	
	1.00	1 10 (0 02 1 50)	1 20 /1 02 1 (4)*	1 40 /1 10 1 00\***	1 70 /1 42 2 40\***			
Not small for gestational age (≥5%) Number excess deaths per year³	1.00	1.18 (0.92, 1.50) 14	1.39 (1.03, 1.64)* 33	1.49 (1.19, 1.86)*** 57	1.76 (1.42, 2.18)*** 116	220		
6 excess deaths per year4	-	1.98	4.55	7.80	15.89	30.23		
nfant mortality								
Birthweight	4.00	4.00 (0.03.4.30)	4.40 (0.04.4.20)	4.00 (0.04.4.35)	4.00 (0.03.4.35)			
ow birthweight (<2,500 grams) Iumber excess deaths per year³	1.00	1.09 (0.93, 1.28) 18	1.10 (0.94, 1.28) 25	1.09 (0.94, 1.25) 32	1.09 (0.93, 1.25) 50	125	< 0.00	
6 excess deaths per year4	_	1.07	1.51	1.93	3.08	7.60	νο.οι	
lot low birthweight (≥2,500 grams)	1.00	1.01 (0.84, 1.22)	1.08 (0.91, 1.29)	1.39 (1.18, 1.63)***	1.60 (1.37, 1.88)***			
Not low birthweight (22,500 grams)  Number excess deaths per year <sup>3</sup>	-	1.01 (0.04, 1.22)	1.00 (0.51, 1.25)	69	137	220		
6 excess deaths per year4	-	0.13	1.18	6.68	13.21	21.20		
iestation								
reterm (<37 weeks)	1.00	1.13 (0.96, 1.33)	1.17 (1.01, 1.36)*	1.22 (1.06, 1.42)*	1. 31 (1.14, 1.51)***			
umber excess deaths per year <sup>3</sup> b excess deaths per year <sup>4</sup>	_	24 1.48	40 2.44	68 4.20	143 8.83	275 16.95	0.00	
						10.33		
erm (≥37 weeks) umber excess deaths per year³	1.00	1.00 (0.83, 1.20)	1.05 (0.88, 1.26)	1.36 (1.15, 1.60)***	1.53 (1.30, 1.78)***	212		
excess deaths per year4	_	_	8 0.74	68 6.29	136 12.59	212 19.62		
. ,								
mall for gestational age mall for gestational age (<5%)	1.00	0.98 (0.75, 1.30)	0.97 (0.75, 1.26)	1.05 (0.82, 1.34)	1.15 (0.92, 1.45)			
lumber excess deaths per year <sup>3</sup>	-	-0.1	-2	4	22	23	0.06	
6 excess deaths per year4	-	-0.22	-0.46	1.06	5.68	6.07		
lot small for gestational age (≥5%)	1.00	1.11 (0.97, 1.28)	1.22 (1.07, 1.39)*	1.45 (1.28, 1.65)***	1.60 (1.42, 1.81)***			
lumber excess deaths per year <sup>3</sup>	_	-28	64	160	285	480		

<sup>\*\*\*</sup> p <0.001 \*\* p <0.005

Adjusted for maternal age, registration status, sex of child and ethnicity of child.

P value from interaction taken from likelihood ratio test.

Number of deaths in sub-population (eg low birthweight babies) which potentially would be avoided each year if deprivation level was 1. Adjusted for maternal age, registration status, sex of child and ethnicity of child.

Proportion of deaths in sub-population (eg low birthweight babies) which potentially would be avoided each year if deprivation level was 1. Adjusted for maternal age, registration status, sex of child and ethnicity of child.

Table 8

Adjusted odds ratios for the effect of low birthweight, preterm and small for gestational age on neonatal, postneonatal and infant mortality, stratified by deprivation, 2005-06.

Population excess deaths (number and proportion) attributed to low birthweight, preterm delivery and small gestation for age also presented.

	Birthwe	eight				Gesta	tion				Small for ge	stational	age	
Not low birthweight (≥2,500 grams)	Low birthw	eight (<2,	500grams	s)	Term (≥37 weeks)	Preteri	n (<37 w	eeks)			Small for g	gestationa	l age (<5	%)
OR 95% CI¹)	OR (95% CI)1	p value²	Number of excess deaths per year <sup>3</sup>	% excess deaths per year <sup>4</sup>	OR (95% CI) <sup>1</sup>	OR (95% CI) <sup>1</sup>	p value²	Number excess deaths per year <sup>3</sup>	deaths	UK	OR (95% CI) <sup>1</sup>	p value²	Number excess deaths per year <sup>3</sup>	% excess deaths per year <sup>4</sup>
lity														
1.00	39.40 (31.52, 49.24)		121	62.71	1.00	35.00 (27.93, 43.84)		124	64.09	1.00	3.25 (2.23, 4.75)		12	7.37
1.00	45.96 (37.13, 56.89)		161	67.68	1.00	41.67 (33.60, 51.67)		160	67.25	1.00	2.96 (2.13, 4.11)		17	7.59
1.00	46.39 (38.28, 56.22)	< 0.01	213	71.22	1.00	42.25 (34.88, 51.18)	0.14	214	71.04	1.00	3.28 (2.52, 4.27)	0.7	26	9.40
1.00	33.44 (28.64, 39.04)		288	67.52	1.00	32.05 (27.50, 37.34		286	68.72	1.00	2.68 (2.16, 3.33)		33	8.40
1.00	33.16 (29.06, 37.84)		460	71.24	1.00	35.34 (31.11, 40.16)		471	70.66	1.00	2.73 (2.33, 3.19)		66	11.03
ortality														
1.00	14.49 (10.48, 20.02)		32	35.97	1.00	10.13 (7.28, 14.09)		27	30.73	1.00	5.40 (3.49, 8.36)		11	12.50
1.00	14.65 (11.08, 19.38)		45	40.32	1.00	11.68 (8.82, 15.47)		41	36.66	1.00	4.13 (2.80, 6.10)		13	11.27
1.00	12.04 (9.34, 15.50)	< 0.001	50	36.28	1.00	10.39 (8.05, 13.41)	0.03	48	34.58	1.00	3.31 (2.30, 4.76)	0.11	14	10.33
1.00	10.48 (8.53, 12.87)		80	38.85	1.00	9.60 (7.81, 11.81)		74	35.96	1.00	3.15 (2.39, 4.17)		23	10.96
1.00	7.91 (6.72, 9.31)		121	36.46	1.00	7.40 (6.27, 8.74)		104	31.31	1.00	2.88 (2.35, 3.53)		42	12.60
,														
1.00	28.73 (24.00, 34.39)		153	54.35	1.00	23.59 (19.72, 28.23)		151	53.65	1.00	3.96 (2.98, 5.27)		24	9.12
1.00	30.87 (26.18, 36.39)		206	58.97	1.00	26.65 (22.62, 31.41)		201	57.50	1.00	3.38 (2.62, 4.34)		29	8.88
1.00	29.19 (15.20, 33.80)	< 0.001	264	60.12	1.00	26.18 (22.61, 30.30)	< 0.01	262	58.20	1.00	3.30 (2.66, 4.08)	0.16	40	9.73
1.00	22.48 (19.93, 25.35)		368	58.12	1.00	21.32 (18.93, 24.01)		360	56.69	1.00	2.85 (2.40, 3.38)		56	9.30
1.00	19.46 (17.65, 21.45)		582	59.45	1.00	20.25 (18.41, 22.27)		575	57.60	1.00	2.79 (2.46, 3.16)		108	11.60
	birthweight (≥2,500 grams)  OR 95% CI')  lity  1.00	Not low birthweight (≥2,500 grams)  OR 95% Cl¹)  OR (95% Cl)1  lity  1.00 39.40 (31.52, 49.24) 1.00 45.96 (37.13, 56.89) 1.00 46.39 (38.28, 56.22) 1.00 33.44 (28.64, 39.04) 1.00 33.16 (29.06, 37.84)  ortality  1.00 14.49 (10.48, 20.02) 1.00 14.65 (11.08, 19.38) 1.00 12.04 (9.34, 15.50) 1.00 10.48 (8.53, 12.87) 1.00 7.91 (6.72, 9.31)  7.91 (6.72, 9.31) 7.91 (15.20, 33.80) 1.00 29.19 (15.20, 33.80) 1.00 22.48 (19.93, 25.35)	Not low birthweight (≥2,500 grams)  OR 95% Cl¹)  OR (95% Cl)1  p value²  lity  1.00 39.40 (31.52, 49.24) 1.00 45.96 (37.13, 56.89) 1.00 46.39 (38.28, 56.22) <0.01 1.00 33.44 (28.64, 39.04) 1.00 33.16 (29.06, 37.84)  ortality  1.00 14.49 (10.48, 20.02) 1.00 14.65 (11.08, 19.38) 1.00 12.04 (9.34, 15.50) <0.001 1.00 10.48 (8.53, 12.87) 1.00 7.91 (6.72, 9.31)  7.1.00 28.73 (24.00, 34.39) 1.00 30.87 (26.18, 36.39) 1.00 29.19 (15.20, 33.80) <0.001 1.00 22.48 (19.93, 25.35)	Not low birthweight	Not low birthweight (≥2,500 grams)   Low birthweight (≥2,500 grams)   Low birthweight (<2,500 grams)   P value   Number of excess deaths per year   P value   P val	Not low birthweight (≥2,500 grams)   Low birthweight (≥2,500 grams)   Low birthweight (≥2,500 grams)   DR (95% CI)1   p value²   Number of excess deaths, per year³   P value²   OR (95% CI)1   p value²   Number of excess deaths, per year³   P value²   OR (95% CI)1   OR (95% C	Not low birthweight   Low birthweight (<2,500 grams)	Not low birthweight   Low birthweight (<2,500 grams)	Not low birthweight   Low birthweight (<2,500grams)	Not low birthweight (<2,500grams)	Not low birthweight   Low birthweight (<2,500 grams)	Not low birthweight (\$\geq 2,500 grams)	Not low birthweight (22,500 grams)   Low birthweight (22,500 grams)   Low birthweight (22,500 grams)   Low birthweight (22,500 grams)   Preterm (-37 weeks)   Preterm (-37 weeks)   Rot small for gestational age (25%)   Preterm (-37 weeks)   Preterm	Not low birthweight (22,500 grams)

Adjusted for maternal age, registration status, sex of child and ethnicity of child. P value from interaction taken from likelihood ratio test.

Number of deaths in sub-population (for example, Carstairs deprivation level 5) which potentially could be avoided each year if all babies had risk of that in the baseline group, adjusted for maternal age, registration status, sex of child and ethnicity of child.

Proportion of deaths in sub-population (for example, Carstairs deprivation level 5) which potentially would be avoided each year if all babies had risk of that in the baseline group, adjusted for maternal age, registration status, sex of child and ethnicity of child.

# Appendix

Table A1

# Covariates by Carstairs distribution, 2005–06

England and Wales												Ι		
		Quintile	1 /loast			Carstairs	quintiles			Ouintilo	ile 5 (most Missing		Al	ı
		depri	ved)	Quin	tile 2	Quin		Quin		depr	ived)	IVIISSIIIG	A	
		Number	% non- missing	Number	% non- missing	Number	% non- missing	Number	% non- missing	Number	% non- missing	Number	Number	% non- missing
Year	2005	98,318	49.3	105,261	49.3	117,779	49.2	135,268	49.0	170,096	49.0	195	626,917	49.1
	2006	101,222	50.7	108,352	50.7	121,562	50.8	140,864	51.0	177,086	51.0	195	649,281	50.9
Registration type/Marital status	Married	144,080	72.2	136,222	63.8	133,228	55.7	132,276	47.9	176,787	50.9	278	722,871	56.6
	Sole registration	5,029	2.5	8,637	4.0	13,914	5.8	23,338	8.5	37,301	10.7	23	88,242	6.9
	Joint registration/same address	43,420							32.4					
	Joint registration/different address	7,011	3.5	11,672	5.5	18,795	7.9	30,914	11.2	45,242	13.0	22	113,656	8.9
Maternal age (years)	<20	5,289	2.7	9,513	4.5	15,599	6.5	25,030	9.1	33,664	9.7	' 13	89,108	7.0
	20–24	17,870	9.0	28,116	13.2	43,220	18.1	63,765	23.1	92,294	26.6	63	245,328	19.2
	25–29	40,866	20.5	52,185	24.4	62,980	26.3	75,030	27.2	96,928	27.9	113	328,102	25.7
	30–34	75,286				71,337			25.0					
	35–39	49,631	24.9						13.0					
	40 and over	10,597	5.3	8,637	4.0	7,729	3.2	7,570	2.7	9,218	2.7	26	43,777	3.4
Sex of baby	Male	102,114	51.2	109,651	51.3	123,009	51.4	140,892	51.0	177,493	51.1	203	653,362	51.2
	Female	97,426	48.8	103,962	48.7	116,332	48.6	135,240	49.0	169,689	48.9	187	622,836	48.8
Ethnicity of baby	White	162,820	92.3	168,370	89.3	183,744	86.2	198,483	79.3	175,629	56.0	195	889,241	77.8
	Asian	3,584				9,600		16,482	6.6				,	
	Black	1,282				4,538		11,758	4.7					
	Other	8,626	4.9	11,621	6.2	15,204	7.1	23,486	9.4	43,164	13.8	32	102,133	8.9
	Missing	23,228	13.2	25,034	13.3	26,255	12.3	25,923	10.4	33,338	10.6	139	133,917	11.7
Baby's ethnicity and maternal														
county of birth	White, UK born	149,746	84.9	155,085	82.2	170,002	79.8	183,026	73.2	154,693	49.3	150	812,702	71.1
	White, non-UK born	13,071	7.4	13,280	7.0	13,740	6.4	15,453	6.2	20,928	6.7	45	76,517	6.7
	Asian, UK born	1,469	0.8	2,266	1.2	3,298	1.5	5,522	2.2	19,645	6.3	3	32,203	2.8
	Asian, non-UK born	2,115	1.2	3,919	2.1	6,302	3.0	10,960	4.4	41,396	13.2	8	64,700	5.7
	Black, UK born	324				1,040			1.1	7,212			,	
	Black, non-UK born	958				3,498			3.6					
	Other, UK born	4,064				6,608			4.0	16,110				
	Other, non-UK born	4,561	2.6						5.4					
	Missing (not stated/not linked)	23,232		25,039		26,257		25,927		33,355		140	133,950	
Birthweight (grams)	<1,500	1,297	0.7	1,641	0.8	2,067		2,814	1.0	4,282	1.2		12,132	1.0
	1,500–2,499	7,156				10,924	4.6	15,163	5.5	23,146	6.7	33	65,050	5.1
	2,500–4,499	186,320		198,140		220,978		252,110		311,550			1,169,415	
	4,500 and over	4,200							1.6					
	Missing	567		965		1,008	i	1,521		3,826		5	7,892	
Low birthweight	Yes	8,453	4.2	10,269	4.8	12,991	5.5	17,977	6.5	27,428	8.0	64	77,182	6.1
	No	190,520	95.8	202,379	95.2	225,342	94.5	256,634	93.5	315,928	92.0	321	1,191,124	93.9
	Missing	567		965		1,008		1,521		3,826		5	7,892	
Gestation (weeks)	21–27	491	0.2	658	0.3	816	0.3	1,161	0.4	1,791	0.5	17	4,934	0.4
, ,,	28–31	1,072							0.7	2,985				
	32–36	8,576							5.4	19,673			64,330	5.1
	37–41	179,690	90.5	191,555	90.1	213,682	89.7	243,796	88.9	302,815	88.5	306	1,131,844	89.4
	42 and over	8,726	4.4	9,407	4.4	10,513	4.4	12,601	4.6	15,008	4.4	! 5	56,260	4.4
	Missing (inconsistent gest-bwt)	19		25		31		46		65		1	187	
	Missing (gest not stated)	804		889		885		1,507		4,339		2	8,426	
	Missing (no link to nn4b)	162		163		209		304		506		7		
	Missing	985		1,077		1,125		1,857		4,910		10	9,964	
Preterm	Yes	10,139	5.1	11,574	5.4	14,021	5.9	17,878	6.5	24,449	7.1	69	78,130	6.2
	No	188,416		200,962		224,195		256,397		317,823			1,188,104	
	Missing	985		1,077		1,125		1,857		4,910		10	9,964	
Small for gestational age	SGA (<5%)	6,076	3.1	7,873	2 7	10,026	42	14,133	5 2	24,017	7.0	11	62,136	4.9
5an for gestational age	Not SGA (≥5%)	192,446		204,623		228,135		260,070		318,115			1,203,756	
	Missing	1,018		1,117		1,180		1,929	3 1.0	5,050		12		