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Agent-based model with multi-level herding for complex financial systems

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In complex financial systems, the sector structure and volatility clustering are respectively important features of the spatial and temporal correlations. However, the microscopic generation mechanism of the sector structure is not yet understood. Especially, how to produce these two features in one model remains challenging. We introduce a novel interaction mechanism, i.e., the multi-level herding, in constructing an agent-based model to investigate the sector structure combined with volatility clustering. According to the previous market performance, agents trade in groups, and their herding behavior comprises the herding at stock, sector and market levels. Further, we propose methods to determine the key model parameters from historical market data, rather than from statistical fitting of the results. From the simulation, we obtain the sector structure and volatility clustering, as well as the eigenvalue distribution of the cross-correlation matrix, for the New York and Hong Kong stock exchanges. These properties are in agreement with the empirical ones. Our results quantitatively reveal that the multi-level herding is the microscopic generation mechanism of the sector structure, and provide new insight into the spatio-temporal interactions in financial systems at the microscopic level.

inancial markets are complex systems with many-body interactions. In recent years, large amounts of historical financial data have sparked the interest of scientists in various fields, including physicists, to quantitatively investigate the properties of the markets. With physical concepts and methods, plenty of results have been obtained¹⁻¹⁹.

From the view of physicists, the dynamic behavior and community structure of complex financial systems can be characterized by temporal and spatial correlation functions. In stock markets, it is well-known that the volatilities are long-range correlated in time, which is the so-called volatility clustering^{2,3,6,7,20}. As to higher-order time correlations, it is discovered that the correlation between past returns and future volatilities is negative for almost all the stock markets in the world^{4,11,21-24}, while currently the correlation is found to be positive only for the Chinese stock market^{11,25}. In other words, the positive and negative returns influence the volatilities asymmetrically, which is known as the volatility asymmetry. During the financial crisis, the volatility asymmetry experiences local minima for developed economies, while reaches local maxima for transition economies²⁶. The spacial structure of the stock markets is explored by investigating the cross-correlation of stocks^{12,18,27-35}. With the random matrix theory (RMT), communities can be identified, which are usually associated with business sec $tors^{31-33}$. The cross-correlation matrix C of price returns is analyzed to investigate the interactions between stocks^{12,18,29-35}. The largest eigenvalue of C deviates significantly from the theoretical distribution of the Wishart matrix, which is the cross-correlation matrix of non-correlated time series. This eigenvalue represents the market mode, i.e., the price co-movement of the entire market, and the components of the corresponding eigenvector is relatively uniform for all stocks. For developed markets, each eigenvector of other large eigenvalues is dominated by the stocks in a certain business sector^{18,30,31}. These large eigenvalues stand for the sector modes.

The spatio-temporal correlations are theoretically crucial in the understanding of the price dynamics, and practically useful for the optimization of the investment portfolio. The sector structure and volatility clustering are, respectively, important features of the spatial and temporal correlations, which we focus on in this paper. In recent years, various models have been proposed to study volatility clustering with certain success^{6,15,36–39}, but the models for the sector structure are phenomenological and usually without interactions of investors^{12,40}. On the other hand, although many activities have been devoted to the sector structure, its microscopic generation mechanism is not yet understood. Both the sector structure and volatility clustering are important characteristics of stock markets, and it remains challenging how to produce these two properties in one model.

As a powerful simulation technique, agent-based modeling is widely applied in various fields^{37,38,41-45}. Recently, an agent-based model is proposed for simulating the cumulative distribution of returns and volatility clustering in

stock markets¹⁵. The concept in constructing the model is to determine the key parameters from empirical data instead of setting them artificially. In this paper, we construct an agent-based model with a novel interaction mechanism, i.e., the multi-level herding, to investigate the sector structure combined with volatility clustering. Further, we propose methods to determine the key model parameters from historical market data rather than from statistical fitting of the results.

Results

In stock markets, the temporal evolution of stock prices and interactions between stocks are complicated. The price dynamics of a market naturally contains that of each individual stock. Recent research has reported that the price dynamics of a market can be decomposed into different modes of motion, such as the market mode and sector mode^{18,30,31}. The market mode is driven by interactions that are common for all stocks in the market, and the sector mode is dominated by interactions of stocks within a sector. Therefore, the price dynamics of a market is multi-level. In financial markets, herding is one of the collective behaviors^{46–50}. Investors cluster into groups when making decisions, and these groups can be large. Since investors' herding behavior is essential to the price dynamics, it may be multi-level as well. In our model, we suppose that agents' herding is composed of three different levels, that is, herding at stock, sector and market levels.

Multi-level herding. Our model is constructed based on the agents' daily trading, i.e., buying, selling and holding stocks. In the model, there are *N* agents, *n* stocks and n_{sec} sectors. Each sector contains *n*/ n_{sec} stocks. Every agent holds only one stock, which is randomly chosen from the *n* stocks. In a real market, an investor may hold different stocks. For simplicity, we suppose a reasonable investor would trade his stocks separately, according to the performance of each stock, even if his operation is based on an investment portfolio. Thus, the scenario for one investor holding, e.g., three stocks is basically the same as that for three investors with each holding one stock.

The stock price of the *i*-th stock on day *t* is denoted by $Y_i(t)$, and the logarithmic price return is $R_i(t) = \ln[Y_i(t)/Y_i(t-1)]$. Since investors' trading decisions in a real market is based on the previous stock performance of different time scales, the investment horizon is introduced in our model for better description of agents' behavior. It has been reported in Ref. 15 that the relative portion ξ_l of investors with a *l* days investment horizon follows a power-law decay, $\xi_l \propto l^{-1.12}$. The maximum investment horizon is denoted by *L*. With the condition of

 $\sum_{l=1}^{L} \xi_l = 1$, we normalize ξ_l to be $\xi_l = l^{-1.12} / \sum_{l=1}^{L} l^{-1.12}$. Agents' trading decisions are made according to the past price returns. On day t + 1, for an agent holding stock i with a investment horizon of l days, $\sum_{m=0}^{l-1} R_i(t-m)$ represents the basis for estimating the previous stock performance. We introduce a weighted average return $R_i'(t)$ to describe the basis of all agents holding stock i. Since ξ_l is the weight of $\sum_{m=0}^{l-1} R_i(t-m)$, $R_i'(t)$ is defined as

$$R_{i}'(t) = k \cdot \sum_{l=1}^{L} \left[\xi_{l} \sum_{m=0}^{l-1} R_{i}(t-m) \right].$$
(1)

We set the coefficient $k=1/\left(\sum_{l=1}^{L}\sum_{m=l}^{L}\xi_{m}\right)$ to ensure that the fluctuation scale of $R_{i}(t)$ is consistent with the one of $R_{i}(t)$, i.e., $|R_{i}(t)|_{\max} = |R_{i}(t)|_{\max}$ (see Supplementary Information S1). If L = 1, $R_{i}(t)$ is just identical to $R_{i}(t)$.

In the model, agents' herding behavior comprises the herding at stock, sector and market levels. For convenience, we denote an agent holding stock i by "agent in stock i". If stock i belongs to sector j, this agent can be denoted by "agent in sector j" as well. A group formed by agents in stock i or by agents in sector j" as well. A group formed by "group in stock i" or "group in sector j". The schematic diagram of the multi-level herding is displayed in Fig. 1(a). The agents in a same stock first cluster into groups. This herding behavior at stock level is similar to the herding in other models which simulate only one stock. In a real market, the stocks in a same sector share the characteristics of the sector. Thus in our model, the groups in each sector further form larger groups, which is the herding at sector level. At last, the groups formed at sector level cluster into even larger ones, since all sectors share common features of the whole market. This is the herding at market level.

(i) Herding at stock level. The agents in each individual stock first cluster into groups, which are called I-groups. We introduce a herding degree D^{i} to quantify the herding behavior at this level. On day *t*, the herding degree for the *i*-th stock is

$$D_i^I(t) = \bar{n}_i(t)/N_i, \qquad (2)$$

where $\bar{n}_i(t)$ denotes the average number of agents in each I-group, and N_i is the number of agents in the *i*-th stock. Agents' herding behavior is based on their estimation of the previous stock performance. Since agents' basis for estimation on day *t* is $|R_i'(t-1)|$, we set $\bar{n}_i(t) = |R_i'(t-1)|$. Thus,



Figure 1 | The schematic diagram of (a) the multi-level herding; (b) the procedure of simulation. (a) "I-herding", "S-herding" and "M-herding" denotes the herding at stock level, sector level and market level, respectively.

$$D_{i}^{I}(t) = \left| R_{i}^{'}(t-1) \right| / N_{i}.$$
(3)

In the *i*-th stock, the number of I-groups is $N_i/\bar{n}_i(t) = 1/D_i^I(t)$, and the agents randomly join in one of the I-groups. After the herding at stock level for all the *n* stocks, the number of I-groups in the *j*-th sector and in the whole market are, respectively, denoted by $N_j^I(t)$ and $N_M^I(t)$,

$$\begin{cases} N_j^I(t) = \sum_{i \in j} \left[1/D_i^I(t) \right] \\ N_M^I(t) = \sum_i \left[1/D_i^I(t) \right] \end{cases}$$
(4)

Here $i \in j$ represents the stock i in sector j.

(ii) Herding at sector level. The stocks in a same sector share the characteristics of the sector. At this level, agents' herding behavior is driven by the price co-movement of the sector, i.e., the prices of stocks in a sector tend to rise and fall at the same time. Thus the I-groups in a same sector would further form larger groups, which are called S-groups. Here we introduce H_M and H_j to characterize the price co-movement degrees for stocks in the whole market and in sector *j*, respectively. For the *j*-th sector, the average number of I-groups in each S-group is set to be $n \cdot (H_j - H_M)$, which represents the pure price co-movement of the sector. Therefore the herding degree is

$$D_j^{\rm S}(t) = n \cdot \left(H_j - H_M\right) / N_j^{\rm I}(t).$$
⁽⁵⁾

In sector *j*, the number of S-groups is $1/D_j^S(t)$, and each I-group joins in one of the S-groups.

(iii) Herding at market level. Agents' herding behavior at this level is driven by the price co-movement of the entire market. The Sgroups in different sectors share common features of the whole market, and thus cluster into larger groups. These groups are called M-groups. In the model, both the herding degrees at sector and market levels are computed based on the I-groups. The co-movement degree H_i represents the percentage of connected I-groups, i.e., I-groups which co-move with each other. $\bar{H} = \left(\sum_{j} H_{j} / n_{\text{sec}}\right)$ stands for the average percentage of connected I-groups. Since the group formation at this level is driven by the price co-movement of the whole market, we suppose that the number of connected I-groups should be the same for different sectors and equal to $\overline{H} \cdot N_M^I(t)$. We denote $\overline{H} \cdot N_M^I(t) / H_j$ by $N_j^M(t)$, and call it the effective number of I-groups for the *j*-th sector. $N_j^M(t)$ satisfies $H_j \cdot N_j^M(t) = \overline{H} \cdot N_M^I(t)$. On the other hand, $n \cdot H_M$ represents the price co-movement for stocks in the whole market. Thus for the S-groups in sector *j*, the herding degree at market level is

$$D_j^M(t) = n \cdot H_M / N_j^M(t), \tag{6}$$

and the number of M-groups is $1/D_j^M(t)$. The total number of Mgroups in the market is the maximum of $1/D_j^M(t)$ for different *j*. With all M-groups numbered, an S-group in sector *j* joins in one of the first $1/D_i^M(t)$ M-groups.

In the formation of S-groups, the I-groups in a same stock tend not to join in a same S-group, otherwise these I-groups would have gathered together during the herding at stock level. Similarly, in the formation of M-groups, the S-groups in a same sector tend not to join in a same M-group. In other words, an I-group prefers to join in an S-group with no other I-groups from the same stock, and an Sgroup prefers to join in an M-group with no other S-groups from the same sector. After the herding for the three levels, all agents cluster into Mgroups. Since intraday trading is not persistent in empirical trading data⁵¹, we suppose that each day only one trading decision is made by every agent. The agents in a same M-group make a same trading decision with a same probability. Considering each agent operates one share, we denote the decision of the α -th agent on day *t* by

$$\phi_{\alpha}(t) = \begin{cases} 1 & \text{buy} \\ -1 & \text{sell} \\ 0 & \text{hold} \end{cases}$$
(7)

and the probabilities of buying, selling and holding decisions of Mgroups are denoted by P_{buy} , P_{sell} and P_{hold} , respectively. The same as the previous models^{15,17}, we suppose that the buying and selling probabilities are equal, i.e., $P_{buy} = P_{sell} = P$, thus $P_{hold} = 1 - 2P$. The return of the *i*-th stock is defined as the difference of the demand and supply of this stock, i.e., the difference between the number of agents buying and selling the stock,

$$R_i(t) = \sum_{\alpha \in i} \phi_{\alpha}(t).$$
(8)

Here $\alpha \in i$ represents the agent α in stock *i*.

Determination of parameters H_M and H_j . The New York Stock Exchange (NYSE) and Hong Kong Stock Exchange (HKSE) are the two representative stock markets considered in this paper. The NYSE is one of the world's most mature markets in the West, and the HKSE is an important market in Asia. We collect the daily closing price data of 150 stocks in the NYSE and HKSE, respectively (Methods). For the comparison of different time series of returns, the normalized return $r_i(t)$ is introduced,

$$r_i(t) = [R_i(t) - \langle R_i(t) \rangle] / \sigma, \qquad (9)$$

where $\langle \cdots \rangle$ represents the average over time *t*, and $\sigma = \sqrt{\langle R_i^2(t) \rangle - \langle R_i(t) \rangle^2}$ is the standard deviation of $R_i(t)$.

The parameters H_M and H_j are introduced to characterize the price co-movement degrees for stocks in the whole market and in sector j, respectively. Actually, the price co-movement of stocks can be characterized by the similarities in the signs and amplitudes of the returns for different stocks. We denote the number of stocks in a sector by n_s , thus $n_s = n/n_{sec}$. In the calculation of H_M , we simply set $n_s = n$. On each day t, according to the sign of $r_i(t)$, these stocks are grouped into two market trends, i.e., the rising and the falling. The amplitudes of the rising and falling trends on day t are defined as $v^+(t)$ and $v^-(t)$, respectively,

$$\begin{cases} v^{+}(t) = \sum_{i,r_{i}(t) > 0} r_{i}^{2}(t) / n_{s} \\ v^{-}(t) = \sum_{i,r_{i}(t) < 0} r_{i}^{2}(t) / n_{s} \end{cases}$$
(10)

The two trends are usually not in balance, and we suppose that these n_s stocks are dominated by either of the two trends, according to the magnitudes of $v^+(t)$ and $v^-(t)$. The amplitude $v^d(t)$ of the dominating trend and the amplitude $v^n(t)$ of the non-dominating one are

$$\begin{cases} v^{d}(t) = \max[v^{+}(t), v^{-}(t)] \\ v^{n}(t) = \min[v^{+}(t), v^{-}(t)] \end{cases}.$$
(11)

We call the stocks grouped into the dominating trend the "dominating stocks". The price co-movement is a common property for all the n_s stocks, hence the total amplitude is $v^d(t) - v^n(t)$. Besides, we take into consideration the similarity in the signs of the returns for these stocks. This similarity is defined as the percentage of the dominating stocks. The number of the dominating stocks is denoted by $n^d(t)$, and the percentage is $\zeta(t) = n^d(t)/n_s$. We take the average over time t for $\zeta(t)$ and $v^d(t) - v^n(t)$, denoted respectively by $\langle \zeta(t) \rangle$ and $\langle v^d(t) - v^n(t) \rangle$. Then, the co-movement degree H_M and H_j are

Table 1 | The values of parameters H_M and H_j for the NYSE and HKSE. H_M and H_j are introduced to characterize the price co-movement degrees for stocks in the whole market and in sector j, respectively. We determine these parameters from the historical market data for each stock exchange

	H _M	H ₁	H ₂	H ₃	H₄	H5
NYSE	0.363	0.491	0.414	0.438	0.431	0.546
HKSE	0.306	0.426	0.406	0.364	0.361	0.340

$$\begin{cases} H_M = \langle \zeta(t) \rangle \cdot \langle v^d(t) - v^n(t) \rangle |_{\text{market}} \\ H_j = \langle \zeta(t) \rangle \cdot \langle v^d(t) - v^n(t) \rangle |_{j-\text{th sector}}. \end{cases}$$
(12)

Here $|_{\text{market}}$ and $|_{j-\text{th sector}}$ represent the calculations for the stocks in the whole market and in the *j*-th sector, respectively. These parameters for the NYSE and HKSE are shown in Table 1.

Simulation of the model. The number of stocks is n = 150 and the number of sectors is $n_{sec} = 5$, which are the same as those of the empirical data we collect for the NYSE and HKSE (Methods). We set the number of agents N to be 600, 000, and the maximum investment horizon L to be 1000 trading days (Methods). In our model, the investment horizons of 94 percent of agents are shorter than 500 days (Methods), which is similar to the previous research¹⁵. Estimated from the historical market data and investment report, the buying or selling probability is P = 0.363 for the NYSE and P = 0.317 for the HKSE (Methods). With H_M and H_j determined for the NYSE and HKSE respectively, our model produces the time series $R_i(t)$ of each stock. The schematic diagram of the simulation procedure is displayed in Fig. 1(b).

Initially, the returns of the first *L* time steps are set to be 0 for all the *n* stocks. On day *t*, we calculate agents' basis $R_i(t)$ according to equation (1), and then $D_i^I(t)$ according to equation (3) for each stock *i*. The agents in the *i*-th stock randomly join in one of the $1/D_i^I(t)$ I-groups. Next, the I-groups in the *j*-th sector join in one of the $1/D_j^S(t)$ S-groups, and then each S-group in the *j*-th sector joins in one of the $1/D_j^M(t)$ M-groups. After the herding for these three levels, the agents in a same M-group make a same trading decision (buy, sell or hold) with the same probability (P_{buy} , P_{sell} or P_{hold}). Thus the returns for each stock on day *t* are calculated with equation (8). The groups disband after their decisions are made. Repeating this procedure, we obtain the time series of returns for all the stocks in the market. For equilibration, the first 10,000 data points of returns are abandoned for each stock, and the length of the time series from simulation is the same as that of the empirical data.

Simulation results. From the calculation for the simulated returns, we obtain the sector structure and volatility clustering. For each stock *i*, the volatility clustering is characterized by the auto-correlation function of volatilities^{2,6}, which is defined as

$$A_{i}(t) = \left[\langle |r_{i}(t')||r_{i}(t'+t)|\rangle - \langle |r_{i}(t')|\rangle^{2} \right] / A_{i}^{0}.$$
(13)

Here $A_i^0 = \langle |r_i(t')|^2 \rangle = \langle |r_i(t')| \rangle^2$, and $\langle \cdots \rangle$ represents the average over time *t'*. Thus, the auto-correlation function of volatilities averaged over all stocks is $A(t) = \sum_{i} A_i(t)/n$. As shown in Fig. 2, for both the NYSE and HKSE, A(t) for the simulations is in agreement with that for the empirical data.

To characterize the spacial structure, we first compute the equaltime cross-correlation matrix $C^{12,29,52}$, of which each element is

$$C_{ij} = \langle r_i(t)r_j(t) \rangle. \tag{14}$$

Here $\langle \cdots \rangle$ represents the average over time *t*, and C_{ij} measures the correlation between the returns of the *i*-th and *j*-th stocks. From the

definition, *C* is a real symmetric matrix with $C_{ii} = 1$, and the values of other elements C_{ii} are in the interval [-1, 1]. The first, second and third largest eigenvalues of C are denoted by λ_0 , λ_1 and λ_2 , respectively. Now we focus on the components $u_i(\lambda)$ of the eigenvector for the three largest eigenvalues. The empirical result of the NYSE is displayed in Fig. 3(a). For λ_0 , the components of the corresponding eigenvector are relatively uniform. The eigenvectors of λ_1 and λ_2 are dominated by sector (5) and sector (1) respectively, with the components significantly larger than those in other sectors. These features are reproduced in our simulation, and the results are shown in Fig. 3(b). The empirical result of the HKSE is displayed in Fig. 4(a). The eigenvectors of λ_1 and λ_2 are respectively dominated by sector (1) and sector (2), and these features are simulated by our model, shown in Fig. 4(b). For the HKSE, the scenario is somewhat complicated⁵³, since a company in the HKSE usually runs various business. As a result, the components of the eigenvector of λ_0 are not so uniform as those in the NYSE. The dominating sectors for λ_1 and λ_2 are less prominent, especially for λ_2 .

Also, the distribution of the eigenvalues is calculated from the simulated returns. As displayed in Fig. 5, for the NYSE and HKSE, the bulk of the distribution of eigenvalues and the three largest eigenvalues from the simulation are in agreement with those from the empirical data.

Discussion

In financial markets, the sector structure and volatility clustering are respectively important features of the spatial and temporal correlations. However, the microscopic generation mechanism of the sector structure is not yet understood. Especially, how to produce these two features in one model remains challenging.

To investigate the sector structure combined with volatility clustering, we construct an agent-based model with a novel interaction mechanism, that is, the multi-level herding. The model is based on the individual and collective behaviors of investors in real markets. According to the previous market performance, agents trade in



Figure 2 | The average auto-correlation functions of volatilities for the NYSE and HKSE, and for the corresponding simulations. For clarity, the curves for the HKSE have been shifted down by a factor of 20.



Figure 3 | The absolute values of the eigenvector components $u_i(\lambda)$ corresponding to the three largest eigenvalues for the cross-correlation matrix *C* calculated from (a) the empirical data in the NYSE; (b) the simulated returns for the NYSE. Stocks are arranged according to business sectors separated by dashed lines. (1): Basic Materials; (2): Consumer Goods; (3): Industrial Goods; (4): Services; (5): Utility.

groups, and their herding behavior comprises the herding at stock, sector and market levels. The key parameters, H_M and H_j , are introduced to characterize the price co-movement degrees for stocks in the whole market and in sector *j*, respectively. We propose methods to determine these parameters from historical market data rather than from statistical fitting of the results. Other parameters *L* and *P* are also estimated from the empirical findings.

With parameters determined for the NYSE and HKSE respectively, our model produces the corresponding time series of returns. From these time series, we obtain the sector structure and volatility clustering, as well as the eigenvalue distribution of the cross-correlation matrix *C*. These properties are in agreement with the empirical ones. Our results quantitatively reveal that the multi-level herding is the microscopic generation mechanism of the sector structure, and provide new insight into the spatio-temporal interactions in financial systems at the microscopic level. The mechanism of the multi-level herding, including the concept of characterizing the price co-movement with parameters H_M and H_j , can also be applied to other complex systems with similar community structures.

Methods

Data. Our data are obtained from "Yahoo! Finance" (http://finance.yahoo.com). We collect the daily data of closing prices of 150 large-cap stocks from 5 business sectors in the NYSE, i.e., the Basic Materials, Consumer Goods, Industrial Goods, Services and Utility, with 30 stocks from each sector. The data are from Jan., 1990 to Dec., 2006 with 4286 data points for every stock. For comparison, we also collect the daily closing price data of 150 stocks in the HKSE to form 5 sectors, with 30 stocks in each sector, and most of these stocks are large-cap stocks. The data are from Jan, 2003 to Sep., 2011 with 2146 data points for each stock. The sector structure of the HKSE is somewhat complicated53, since a company in the HKSE usually runs various business. According to the dominating stocks of eigenvectors of the cross-correlation matrix C, the sectors in the HKSE are not so strict as those in the NYSE, and may be composed of two business sectors. Specifically, the second sector comprises 14 stocks from the Conglomerates and 16 stocks from the Industrial Goods. The third sector consists of 12 stocks from the Basic Materials and 18 stocks from the Technology. The stocks in other three sectors are, respectively, from the Real Estate Development, Services and Consumer Goods.

Parameter *N* and *L*. To simulate the properties of a real stock market, the number of agents *N* should not be too small, since we suppose that every agent operates only one share of a stock. We set N = 600, 000, i.e., 4000 agents holding a same stock on average. The simulation results are not sensitive to *N*. For example, we obtain almost the same results for N = 450, 000 or N = 750, 000. *N* has no influence on the sector



Figure 4 | The absolute values of the eigenvector components $u_i(\lambda)$ corresponding to the three largest eigenvalues for the cross-correlation matrix *C* calculated from (a) the empirical data in the HKSE; (b) the simulated returns for the HKSE. Stocks are arranged according to business sectors separated by dashed lines. Sector (2) and (3) are composed of two business sectors, respectively (Methods). (1): Real Estate Development; (2): Conglomerates - Industrial Goods; (3): Basic Materials - Technology; (4): Services; (5): Consumer Goods.



Figure 5 | The probability distribution of the eigenvalues of the crosscorrelation matrix *C* for the NYSE and HKSE, and for the corresponding simulations. The probability distribution of the three largest eigenvalues is shown in the inset. For the NYSE, the three largest eigenvalues are $(\lambda_0, \lambda_1, \lambda_2) = (26.01, 7.45, 5.13)$, in comparison with (24.62, 7.93, 3.82) for the simulation. For the HKSE, the three largest eigenvalues are $(\lambda_0, \lambda_1, \lambda_2) = (21.70, 3.06, 1.89)$, in comparison with (21.43, 4.57, 2.97) for the simulation.

structure and the eigenvalue distribution of the matrix *C*, and slightly affects the amplitude of the auto-correlation function A(t) of volatilities averaged over stocks. The investment horizons of investors range from one day to more than one year⁵⁴. Considering there could be some investors with long investment horizon in the stock market, we set L = 1000 trading days in our model. Thus according to the relative portion of investors, the investment horizons of 94 percent of agents are shorter than 500 days, suggesting that most agents do not estimate previous market performance over a too long time period. *L* affects the temporal and spatial properties little. Similar with *N*, the simulation results are almost the same for, e.g., L = 800 or L = 1200.

Determination of parameter P. We first determine the daily buying, selling and holding probabilities of a single investor in a real market, which is denoted by p_{buy} , p_{sell} and p_{hold} respectively. p_{buy} and p_{sell} are supposed to be equal, i.e., $p_{buy} = p_{sell} = p$. The time series of returns of the 150 stocks in the NYSE are from Jan., 1990 to Dec., 2006. According to "The 2010 Institutional Investment Report"55, the average percentage of institutional holdings of shares in the top 1000 U.S. corporations during these years is 60.3 percent (see Supplementary Information S2). There are two kinds of investors in stock markets, that is individual and institutional investors. The percentage of holding shares for the individual investors is thus 39.7 percent. According to Ref. 15, the yearly average ratio between the number of shares an investor trades and the number of shares he holds is 1.64. This ratio corresponds to the yearly average trading times of an investor. Since institutional investors contribute little of the trading times¹⁵, we ignore their trades. So the yearly trading times for an individual investor is 1.64/0.397 = 4.13. Since there are 250 trading days in every year, the daily trading probability is $4.13/250 = p_{buy}(t) + p_{sell}(t) = 2p$. Therefore, p is 0.00826 for the NYSE. For the HKSE, the corresponding empirical data are not available to us, and we assume that p is 0.00826 as well.

The agents in a same M-group are connected. We suppose that if one agent in the group decides to buy or sell the stock, the whole group would make the same decision. In the model, the average number of agents in an M-group is $n \cdot H_M$. Therefore, the buying or selling probability of an M-group is $P = 1 - (1-p)^{n'H_M}$. Thus, *P* is 0.363 for the NYSE and 0.317 for the HKSE.

- Mantegna, R. N. & Stanley, H. E. Scaling behaviour in the dynamics of an economic index. *Nature* 376, 46–49 (1995).
- Gopikrishnan, P., Plerou, V., Amaral, L. A. N., Meyer, M. & Stanley, H. E. Scaling of the distribution of fluctuations of financial market indices. *Phys. Rev. E* 60, 5305 (1999).
- 3. Liu, Y. *et al.* Statistical properties of the volatility of price fluctuations. *Phys. Rev. E* **60**, 1390 (1999).
- Bouchaud, J. P., Matacz, A. & Potters, M. Leverage effect in financial markets: the retarded volatility model. *Phys. Rev. Lett.* 87, 228701 (2001).
- 5. Tino, P., Schittenkopf, C. & Dorffner, G. Financial volatility trading using recurrent neural networks. *IEEE Trans. Neural Netw.* **12**, 865–874 (2001).
- Krawiecki, A., Ho lyst, J. A. & Helbing, D. Volatility clustering and scaling for financial time series due to attractor bubbling. *Phys. Rev. Lett.* 89, 158701 (2002).

- Gabaix, X., Gopikrishnan, P., Plerou, V. & Stanley, H. E. A theory of power-law distributions in financial market fluctuations. *Nature* 423, 267–270 (2003).
- 8. Sornette, D. Critical market crashes. Phys. Rep. 378, 1-98 (2003).
- Bonanno, G., Caldarelli, G., Lillo, F. & Mantegna, R. N. Topology of correlationbased minimal spanning trees in real and model markets. *Phys. Rev. E* 68, 046130 (2003).
- Johnson, N. F., Jefferies, P. & Hui, P. M. Financial Market Complexity. (Oxford Univ. Press, Oxford, 2003).
- 11. Qiu, T., Zheng, B., Ren, F. & Trimper, S. Return-volatility correlation in financial dynamics. *Phys. Rev. E* 73, 065103 (2006).
- Shen, J. & Zheng, B. Cross-correlation in financial dynamics. EPL 86, 48005 (2009).
- 13. Zhao, L. et al. Herd behavior in a complex adaptive system. Proc. Nati. Acad. Sci. USA 108, 15058–15063 (2011).
- Preis, T., Schneider, J. J. & Stanley, H. E. Switching processes in financial markets. Proc. Nati. Acad. Sci. USA 108, 7674–7678 (2011).
- Feng, L., Li, B., Podobnik, B., Preis, T. & Stanley, H. E. Linking agent-based models and stochastic models of financial markets. *Proc. Nati. Acad. Sci. USA* 109, 8388–8393 (2012).
- 16. Preis, T., Moat, H. S. & Stanley, H. E. Quantifying trading behavior in financial markets using google trends. *Sci. Rep.* **3**, 1684 (2013).
- Chen, J. J., Zheng, B. & Tan, L. Agent-based model with asymmetric trading and herding for complex financial systems. *PloS one* 8, e79531 (2013).
- Jiang, X. F., Chen, T. T. & Zheng, B. Structure of local interactions in complex financial dynamics. *Sci. Rep.* 4, 5321 (2014).
- Meng, H. *et al.* Systemic risk and spatiotemporal dynamics of the us housing market. *Sci. Rep.* 4, 3655 (2014).
- Ding, Z., Granger, C. W. & Engle, R. F. A long memory property of stock market returns and a new model. J. Empir. Financ. 1, 83–106 (1993).
- Black, F. Studies of stock price volatility changes. Proceedings of the 1976 Meetings of the American Statistical Association, Business and Economical Statistics Section, 177–181 (1976).
- Glosten, L. R., Jagannathan, R. & Runkle, D. E. On the relation between the expected value and the volatility of the nominal excess return on stocks. *J. Financ.* 48, 1779–1801 (1993).
- 23. Engle, R. F. & Ng, V. K. Measuring and testing the impact of news on volatility. *J. Financ.* **48**, 1749–1778 (1993).
- Zakoian, J. M. Threshold heteroskedastic models. J. Econ. Dyn. Control 18, 931–955 (1994).
- Shen, J. & Zheng, B. On return-volatility correlation in financial dynamics. *EPL* 88, 28003 (2009).
- Tenenbaum, J. et al. Comparison between response dynamics in transition economies and developed economies. *Phys. Rev. E* 82, 046104 (2010).
- Erb, C. B., Harvey, C. R. & Viskanta, T. E. Forecasting international equity correlations. *Financ. Anal. J.* 50, 32–45 (1994).
- Solnik, B., Boucrelle, C. & Le Fur, Y. International market correlation and volatility. *Financ. Anal. J.* 52, 17–34 (1996).
- Laloux, L., Cizeau, P., Bouchaud, J. P. & Potters, M. Noise dressing of financial correlation matrices. *Phys. Rev. Lett.* 83, 1467 (1999).
- Gopikrishnan, P., Rosenow, B., Plerou, V. & Stanley, H. E. Quantifying and interpreting collective behavior in financial markets. *Phys. Rev. E* 64, 035106 (2001).
- 31. Plerou, V. *et al.* Random matrix approach to cross correlations in financial data. *Phys. Rev. E* **65**, 066126 (2002).
- Utsugi, A., Ino, K. & Oshikawa, M. Random matrix theory analysis of cross correlations in financial markets. *Phys. Rev. E* 70, 026110 (2004).
- Pan, R. K. & Sinha, S. Collective behavior of stock price movements in an emerging market. *Phys. Rev. E* 76, 046116 (2007).
- Podobnik, B., Wang, D., Horvatic, D., Grosse, I. & Stanley, H. E. Time-lag crosscorrelations in collective phenomena. *EPL* 90, 68001 (2010).
- Jiang, X. F. & Zheng, B. Anti-correlation and subsector structure in financial systems. *EPL* 97, 48006 (2012).
- 36. Lux, T. & Marchesi, M. Volatility clustering in financial markets: A microsimulation of interacting agents. *International Journal of Theoretical and Applied Finance* 3, 675–702 (2000).
- Giardina, I., Bouchaud, J. P. & Mézard, M. Microscopic models for long ranged volatility correlations. *Physica A* 299, 28–39 (2001).
- Challet, D., Marsili, M. & Zhang, Y. C. Stylized facts of financial markets and market crashes in minority games. *Physica A* 294, 514–524 (2001).
- Ren, F., Zheng, B., Qiu, T. & Trimper, S. Score-dependent payoffs and minority games. *Physica A* 371, 649–657 (2006).
- Ma, W. J., Hu, C. K. & Amritkar, R. E. Stochastic dynamical model for stock-stock correlations. *Phys. Rev. E* 70, 026101 (2004).
- Bonabeau, E. Agent-based modeling: methods and techniques for simulating human systems. Proc. Nati. Acad. Sci. USA 99, 7280–7287 (2002).
- Evans, T. P. & Kelley, H. Multi-scale analysis of a household level agent-based model of landcover change. J. Environ. Manage. 72, 57–72 (2004).
- Ren, F., Zheng, B., Qiu, T. & Trimper, S. Minority games with score-dependent and agent-dependent payoffs. *Phys. Rev. E* 74, 041111 (2006).
- Farmer, J. D. & Foley, D. The economy needs agent-based modelling. *Nature* 460, 685–686 (2009).

- Schwarz, N. & Ernst, A. Agent-based modeling of the diffusion of environmental innovationsan empirical approach. *Technol. Forecast. Soc.* 76, 497–511 (2009).
- Eguiluz, V. M. & Zimmermann, M. G. Transmission of information and herd behavior: an application to financial markets. *Phys. Rev. Lett.* 85, 5659–5662 (2000).
- 47. Cont, R. & Bouchaud, J. P. Herd behavior and aggregate fluctuations in financial markets. *Macroecon. Dyn.* 4, 170–196 (2000).
- Hwang, S. & Salmon, M. Market stress and herding. J. Empir. Financ. 11, 585–616 (2004).
- Zheng, B., Qiu, T. & Ren, F. Two-phase phenomena, minority games, and herding models. *Phys. Rev. E* 69, 046115–1 (2004).
- 50. Kenett, D. Y. *et al.* Index cohesive force analysis reveals that the US market became prone to systemic collapses since 2002. *PLoS one* **6**, e19378 (2011).
- 51. Eisler, Z. & Kertesz, J. Liquidity and the multiscaling properties of the volume traded on the stock market. *EPL* 77, 28001 (2007).
- Plerou, V., Gopikrishnan, P., Rosenow, B., Amaral, L. A. N. & Stanley, H. E. Universal and nonuniversal properties of cross correlations in financial time series. *Phys. Rev. Lett.* 83, 1471 (1999).
- 53. Ouyang, F. Y., Zheng, B. & Jiang, X. F. Spatial and temporal structures of four financial markets in greater china. *Physica A* **402**, 236–244 (2014).
- 54. Menkhoff, L. The use of technical analysis by fund managers: international evidence. J. Bank. Financ. 34, 2573–2586 (2010).
- Tonello, M. & Rabimov, S. R. The 2010 institutional investment report: trends in asset allocation and portfolio composition. *The Conference Board Research Report* No. R-1468-10-RR (2010).

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Author contributions

J.J.C., L.T. and B.Z. conceived the study; J.J.C. and L.T. designed and performed the research; J.J.C. and L.T. performed the statistical analysis of the data; J.J.C., L.T. and B.Z. discussed the results and contributed to the text of the manuscript.

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