REVIEW ARTICLE OPEN



The efficacy of transcranial magnetic stimulation (TMS) for negative symptoms in schizophrenia: a systematic review and meta-analysis

Rasmus Lorentzen (1)^{1,2}, Tuan D. Nguyen^{1,2}, Alexander McGirr (1)^{3,4,5}, Fredrik Hieronymus (1)^{1,2,6} and Søren D. Østergaard (1)^{1,2,8}

Several trials have shown preliminary evidence for the efficacy of transcranial magnetic stimulation (TMS) as a treatment for negative symptoms in schizophrenia. Here, we synthesize this literature in a systematic review and quantitative meta-analysis of double-blind randomized controlled trials of TMS in patients with schizophrenia. Specifically, MEDLINE, EMBASE, Web of Science, and PsycINFO were searched for sham-controlled, randomized trials of TMS among patients with schizophrenia. The effect of TMS vs. sham on negative symptoms in each study was quantified by the standardized mean difference (SMD, Cohen's *d*) with 95% confidence intervals (95%CI) and pooled across studies using an inverse variance random effects model. We identified 57 studies with a total of 2633 participants that were included in the meta-analysis. The pooled analysis showed statistically significant superiority of TMS (SMD = 0.41, 95%CI: 0.26; 0.56, *p*-value < 0.001), corresponding to a number needed to treat of 5. Furthermore, stratified analyses suggested that TMS targeting the left dorsolateral prefrontal cortex and using a stimulation frequency >1 Hz was most efficacious. There was, however, substantial heterogeneity and high risk of bias among the included studies. In conclusion, TMS appears to be an efficacious treatment option for patients with schizophrenia suffering from negative symptoms, but the optimal TMS parameters are yet to be established.

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INTRODUCTION

Pharmacological treatment is the cornerstone of care in schizophrenia and other psychotic disorders¹. Though positive symptoms (e.g., delusions and hallucinations) respond relatively well to pharmacological treatment, negative symptoms often do not respond to the same degree^{2–4}. The negative symptoms of schizophrenia are those representing absence/lessening of normal functions and include affective flattening, alogia, apathy and social withdrawal⁵. Patients with predominantly negative symptoms are more resistant to treatment than patients with primarily positive symptoms, and negative symptoms are strongly associated with low daily functioning and poor long-term prognosis^{6–8}. Therefore, identification and development of efficacious treatments of negative symptoms is a priority^{3,9}.

Transcranial magnetic stimulation (TMS) is a non-invasive neuromodulation technique in which a localized electrical field is elicited in underlying brain parenchyma through electromagnetic induction, generally limited to superficial cortical regions ^{10,11}. Whether TMS increases or decreases the activity of the targeted neurons depends on the frequency of the magnetic pulses with 1 Hz and below (low frequency) being inhibitory and >1 Hz being excitatory (high frequency) ¹². Repetitive TMS (rTMS) is the most widely used modality with a single session lasting 20–40 min, typically delivering between 1200-3000 magnetic pulses ¹³. Other types include deep TMS in which the magnetic field reaches deeper subcortical regions of the brain as well as theta burst stimulation (TBS) in which the frequency of stimulation is 50 Hz administered five times per second to mimic endogenous theta waves either continuously or intermittently ¹⁴.

TMS is approved for the treatment of major depression (https://www.accessdata.fda.gov/cdrh_docs/pdf6/K061053.pdf) and since the negative symptoms of schizophrenia and the symptoms of major depression both represent deficits of normal functions, TMS has also been explored as a potential treatment of negative symptoms among patients with schizophrenia^{15–18}. Since the most recent reviews of the literature on TMS for treatment of negative symptoms^{15–18}, several trials have been conducted—some using novel stimulation parameters as well as neuronavigation to improve targeting^{19–27}. Due to these developments in the field, an updated synthesis would be of relevance. Therefore, we conducted a systematic review and quantitative meta-analysis of randomized controlled trials reporting on the efficacy of rTMS in the treatment of negative symptoms among patients with schizophrenia.

METHODS

Protocol and registration

The study protocol was registered at the International Prospective Register of Systematic Reviews (PROSPERO, ID: CRD42021238828) (https://www.crd.york.ac.uk/PROSPERO) and carried out in accordance with the Preferred Reporting Items for Systematic Reviews and Meta-Analyses (PRISMA) guidelines²⁸.

Information sources and screening

MEDLINE (PubMed), PsycINFO, Web of Science and EMBASE were searched for relevant studies. Earlier reviews on the subject,

¹Department of Affective Disorders, Aarhus University Hospital – Psychiatry, Aarhus, Denmark. ²Department of Clinical Medicine, Aarhus University, Aarhus, Denmark. ³Hotchkiss Brain Institute, University of Calgary, Canada. ⁵Department of Psychiatry, Cumming School of Medicine, University of Calgary, Canada. ⁵Mathison Centre for Mental Health Research and Education, University of Calgary, Calgary, Canada. ⁶Department of Pharmacology, Sahlgrenska Academy, University of Gothenburg, Gothenburg, Sweden. ^{Se}email: soeoes@rm.dk

clinicaltrials.gov, as well as citations of included studies were reviewed in order to find further eligible studies. The search was carried out on May 1st 2021 using the following search string in MEDLINE: ("schizophreni*" OR "schizoaffective disorder" OR "schizophreniform disorder" OR "schizophrenia" [MeSH Terms] OR "negative symptom*" OR "CHR" OR "Clinical High Risk" OR "Ultra High Risk" OR "UHR" OR "Psychotic Disorders" [MeSH Terms] OR "Psychotic Disorder*") AND ("transcranial magnetic stimulation" OR "TMS" OR "rTMS" OR "theta burst" OR "iTBS" OR "cTBS" OR Transcranial Magnetic Stimulation*[MeSH Terms]). The analogue search strings used for the other databases are available in the Supplementary Material. Titles and abstracts of studies identified via the search strategy described above were screened independently by two authors (RL and TDN) assisted by Covidence²⁹. Full text versions of the studies deemed relevant after initial screening were subsequently assessed for eligibility.

Eligibility criteria

In line with earlier reviews in the field, the following inclusion criteria were employed^{15,17,18}. Notably, no language restrictions were employed:

- Randomized, sham-controlled trials of transcranial magnetic stimulation (e.g., rTMS or theta burst stimulation)
- Participants with a primary diagnosis of schizophrenia, schizoaffective disorder or another psychotic disorder (e.g., acute/transient/brief psychotic disorder or persistent delusional disorder), according to the DSM-IV, DSM-5, or ICD-10.
- Adult participants (≥18 years)
- Outcome measured using an established psychometric scale for negative symptoms in schizophrenia (e.g., the negative subscale of the Positive and Negative Syndromes Scale (PANSS-N)³⁰ or the Scale for Assessment of Negative Symptoms (SANS)³¹).

The following exclusion criterion was employed:

 Co-initiation of other treatments, e.g., pharmacological treatment, as the results of such studies could be affected by an interaction effect between TMS and the co-initiated treatment.

Data extraction

The following items were extracted from each included study: Author name, publication year, country, study type (cross-over or parallel), analysis-type (per protocol or intention-to-treat (ITT)), number of participants, drop-out rates, mean age of participants, sex distribution of participants, diagnostic distribution, whether samples were selected for predominantly negative symptoms, frequency and intensity of TMS including the total number of stimuli and number of treatments, TMS target, nature of the sham intervention, outcome measure (rating scale), post treatment scores, follow-up scores and post treatment depression scores (if available). If these data were not reported, the authors were contacted by e-mail with a request to provide the data. If authors did not reply, data from graphs (if available) were extracted using the GetData Graph Digitizer (http://getdata-graph-digitizer.com/). Previous meta-analyses were screened for post-treatment outcome data required to compute effect sizes. Studies where data was not available upon request, via graphs or through previous meta-analyses, were excluded from the analyses.

Evaluation of risk of bias

The included studies were evaluated according to five domains of bias (articles in non-English languages were not evaluated) using the Cochrane Risk of Bias Tool 2.0:(https://methods.cochrane.org/risk-bias-2) (A) Randomization process (allocation sequence generation and concealment), (B) Deviations from intended interventions (bias arising from non-protocol interventions), (C) Missing outcome data (dropouts), (D) Measurement of the

outcome (using a validated tool), and (E) Selection of the reported result (alignment with protocol and method section). In accordance with the instructions for the Cochrane Risk of Bias Tool 2.0, the highest risk score assigned in one of these domains defined the overall risk of bias score for each study (https://methods.cochrane.org/risk-bias-2). Furthermore, potential publication bias was explored via a funnel plot and Egger's regression test.

Statistical analysis

The effect of TMS vs. sham on negative symptoms in each included study was quantified by the standardized mean difference (SMD, Cohen's *d*) with 95% confidence intervals (95% CI) based on endpoint scores or change scores (with endpoint scores being preferred). If multiple outcome measures were used, PANSS-N was preferred to improve methodological homogeneity because it was used in 89% (51/57) of the included studies. If a study did not provide standard deviations (SD) or data that could be used to calculate SD (e.g., standard error), the mean standard deviation across all studies of the same outcome measure was used. For cross-over studies, data was extracted after the first treatment phase (before cross-over) to exclude possible carry-over effects of treatment and thus regarded as a parallel design study.

SMDs were pooled using the inverse variance random effects model in Review Manager 5.3^{32} . This model takes into account both in-study and between-study variability. For the primary analysis, number needed to treat (NNT) was estimated using the method proposed by Kraemer and Kupfer³³. Heterogeneity was assessed using the l^2 -test with l^2 -values $\geq 50\%$ suggesting considerable heterogeneity. For multi-arm studies, data from different active TMS treatment arms were pooled in the calculation of overall efficacy as to not duplicate data from the sham group, using the formulas provided in table 6.5a in the Cochrane Handbook (https://training.cochrane.org/handbook/current/chapter-06#section-6-5-2-10).

Following the main analysis, the following secondary/subgroup analyses (yielding effect sizes) were carried out: (i) focusing on long term effect using data from at least four weeks after the last treatment (the last follow up in each study was used), (ii) focusing on patients with predominantly negative symptoms, (iii) focusing on the effect size of TMS for depressive symptoms (all depression measures allowed with a preference for the Calgary Depression Scale in Schizophrenia³⁴), (iv) after stratifying by target site (v) after stratifying by type of TMS, (vi) after stratifying by stimulation frequency, (vii) after stratifying by stimulation intensity, and (viii) after stratifying by age. Finally, the following sensitivity analyses (yielding effect sizes) were conducted: (I) after excluding studies with data extracted from graphs, (II) after excluding studies with high risk of bias, (III) after excluding studies reporting changefrom-baseline scores, (IV) after excluding studies with data extracted from other reviews, and (V) after stepwise exclusion of the 10 most outlying studies compared to the overall efficacy estimate.

RESULTS

Study selection

The search yielded 3287 articles of which 1573 were duplicates, resulting in 1714 studies that underwent title- and abstract screening (Fig. 1). Following this screening, 1565 were excluded. This left 149 articles to be assessed in full text, of which 80 studies did not meet the eligibility criteria. The most common reason for exclusion in this final step was "not randomized sham-controlled trials" (31 studies), which included non-blinded studies, studies with no sham-treated group and studies, which were not clinical trials. The second most common reason for exclusion was "did not include eligible outcome measure" (14 studies), which included studies that did not measure negative symptoms, but had other

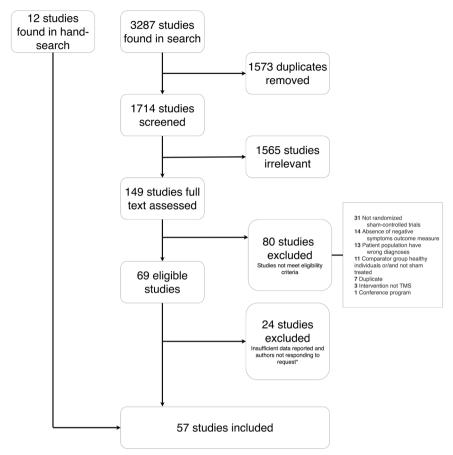


Fig. 1 PRISMA flowchart illustrating the literature screening. *Authors were contacted by e-mail. If data was not provided and data could not be taken from graphs, the study was excluded.

primary endpoints such as biomarkers/neuroimaging or cognitive function. Of the 69 eligible studies, 35 reported insufficient data and thus the authors were contacted by e-mail requesting additional data. The means and standard deviations of PANSS-N was the most common missing piece of information (i.e., from studies where only the total PANSS scores were reported). From these 35 studies, three author groups provided data^{21,26,35}, and data were extracted from graphs in an additional eight studies^{27,36–42}. Hence, 24 articles were excluded due to nonavailable data⁴³⁻⁶⁶. In total, the search yielded 45 includable studies, with 51 comparisons as a result of studies including multiple interventions 19–27,35–42,67–94. No additional studies were found in citations or in the database of clinicaltrials.gov. Summary data was available from 15 studies reviewed by Wang and colleagues¹⁸ from non-English reports^{90–92,95–106}, of which 3 were also identified by the database search, leaving a total of 57 studies and 63 comparisons 19-27,35-42,66-106.

Study characteristics

Table 1 shows the characteristics of the 57 included studies.

The 57 studies included 2633 participants, of whom 1481 received active treatment and 1152 sham treatment. In the 55 studies (n = 2525) that reported the specific diagnoses of the participants, 98.9% had schizophrenia and 1.1% had schizoaffective disorder. The two remaining studies reported that the participants had either schizophrenia or schizoaffective disorder, without providing the distribution between the two. The studies were conducted in 15 different countries, of which China was the most common (n = 25). Almost all included studies reported the outcome using PANSS-N or SANS with only one study using

the Brief Psychiatric Rating Scale – Negative/Disorganized factor (BPRS-N/D)⁶⁹.

Several different active TMS modalities were used in the included trials, with some testing more than one active modality against sham treatment: rTMS (48 studies, 10 used \leq 1 Hz, 39 used >1 Hz, 42 used unilateral treatment, and eight bilateral or midline treatment), theta burst stimulation (TBS) (9 studies, 5 used intermittent TBS (iTBS), 1 used continuous TBS (cTBS), and 3 used unspecified TBS), and deep-TMS (2 studies). The mean total number of TMS pulses per trial was 25,684 varying from 1200 to 80,000 with an average of 1455 pulses per treatment. The majority of the studies (n=39) had the left dorsolateral prefrontal cortex (L-DLPFC) as the primary stimulation target.

Risk of bias of individual studies

Eight studies were regarded as having low risk of bias, 10 studies with "some concerns", and 23 studies with high risk of bias (Supplementary Table 1). The most common reason for "some concerns" was insufficient reporting whether the randomization sequence was concealed adequately (domain A). Improper analysis (e.g. "per protocol" analysis, domain B) and missing outcome data (domain C) were the most common reasons (n=12 and n=21, respectively, with n=11 having both) for a study being regarded as having high risk of bias.

Results of individual studies

Standardized mean differences for the included studies are shown in Fig. 2. Eighteen studies showed a statistically significant superior effect of TMS compared to sham treatment^{21,22,38,78,79,81,82,88–90,95,96,100–103,105,106} and one study found

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	Rosa	2007								25%	%0	PANSS-N	rTMS	Sham coil	-	%06	10	0096	L-TPC

Author	Year Country	Z e	N _c Mear	nage Pi	Mean age PNS Analysis	sis Study type	e % males	Drop- out rate	Outcome score TMS type	TMS type	Sham type	Ŧ	W %	Treatment sessions	Total number of stimuli	Brain target
Rosenberg	2012 Israel	5 5	39.2	Z	ЬР	Par	78%	44%	SANS	deep-TMS	Sham coil	-	110%	10	0009	L-TPC
Saba	2006 France	8	30.6	Z	ЬР	Par	81%	11%	PANSS-N	rTMS	Sham coil	-	%08	14	4200	L-TPC
Schneider	2008 USA	33 15	5 41.1	>	ЬР	Par	33%	%9	SANS	rTMS	Sham coil	1-10	110%	20	Var	L-DLPFC
Singh	2020 India	15 15	5 31.0	>	E	Par	21%	13%	PANSS-N, SANS	rTMS	Sham coil	20	100%	20	40,000	L-DLPFC
Tikka	2017 India	8 7	26.5	Z	ЬР	Par	NA	25%	PANSS-N	cTBS	Sham coil	20	%08	10	0006	R-IPL
Wang	2015 China	41 42	2 42.3	AN	A NA	ΑN	NA	NA	PANSS-N	rTMS	NA	10	110%	20	48,000	L-DLPFC
Wang	2020 China	25 25	5 NA	z	Ν	Par	NA	ΑΝ	PANSS-N, SANS	iTBS	ΑN	Ν	ΝΑ	14	ΑN	L-DLPFC
Wen-Xiang	2012 China	76 31	1 42.7	Z	E	Par	71%	12%	PANSS-N	rTMS	Sham coil	_	%08	20	16,000	L-DLPFC
Wobrock	2015 Germany	76 81	1 35.3	>	E	Par	75%	10%	PANSS-N	rTMS	45°	10	110%	15	15,000	L-DLPFC
Xiu	2020 China	67 30	0 52.4	>	E	Par	100%	19%	PANSS-N	rTMS	Sham coil	10-20	110%	40	Var	L-DLPFC
χn	2006 China	33 34	4 38.0	AN	A NA	Ϋ́	NA	NA	PANSS-N	rTMS	NA	NA	%08	10	Var	L-DLPFC, B-PC
χn	2015 China	60 30	0 45.3	Ν	A NA	ΥN	NA	NA	PANSS-N	rTMS	NA	5-10	%08	10	25,000	L-DLPFC
Zhang	2010 China	15 15	5 38	NA	A NA	ΥN	NA	ΑΝ	PANSS-N	TBS	ΑN	NA	%08	20	48,000	L-DLPFC
Zhang	2015 China	35 34	4 39.9	NA	A NA	ΑN	NA	NA	PANSS-N	rTMS	NA	10	Ä	20	16,000	L-DLPFC
Zhao	2014 China	71 22	2 47.9	>	ЬР	Par	%95	3%	PANSS-N, SANS	rTMS, iTBS	180°	10-50	Var	20	Var	L-DLPFC
Zheng	2012 China	56 17	7 56.3	NA	A NA	Par	NA	ΑΝ	PANSS-N	rTMS, TBS	NA	10-20	%08	5	2000-6000	L-DLPFC
Zhuo	2019 China	33 27	7 30.0	>	ЬР	Par	37%	14%	PANSS-N, SANS	rTMS	180°	20	%06	20	40,000	L-DLPFC

Parallel, PANSS-N Positive and Negative Syndrome Scale – Negative subscale, SANS Scale for the Assessment of Negative Symptoms, BPRS-N/D Brief Psychiatric Rating Scale – Negative/Disorganization factor. TMS Transcranial Magnetic Stimulation, Hz Herz (frequency of stimulation), % MT Percent of Motor Threshold, rTMS repetitive Transcranial Magnetic Stimulation, TBS theta burst stimulation, B-DLPFC Bilateral dorsolateral prefrontal cortex, L-DLPFC Left dorsolateral prefrontal cortex, R-DLPFC Right dorsolateral prefrontal cortex, MC Medullar cerebellum, *L-TPC* Left tempero-parietal cortex, *L-57*5 Left superior temporal sulcus, *B-PFC* Bilateral prefrontal cortex, *L-PFC* Left prefrontal cortex, *R-PFC* Right prefrontal cortex, *B-PFC* Bilateral parietal cortex, *L-PTC* Left prefrontal cortex, *R-MTC* Left medial temporal gyrus, *R-PP* Right inferior parietal lobule. ° in the sham type column indicates if the coil was rotated as sham method and by how many degrees.

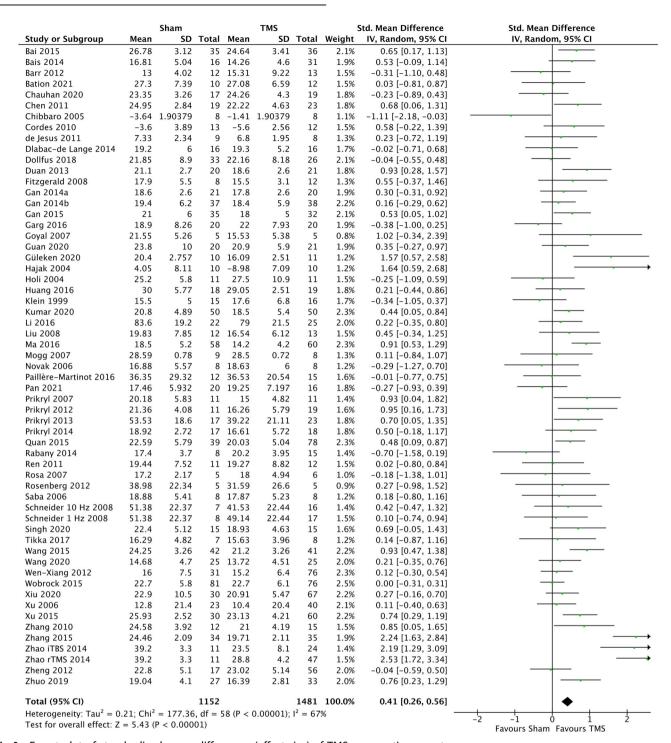


Fig. 2 Forest plot of standardized mean differences (effect size) of TMS on negative symptoms.

a statistically significant superior effect of sham treatment³⁶. The remaining studies did not show a statistically significant difference between the treatment groups. There was considerable heterogeneity between the included studies ($l^2 = 67\%$).

Synthesis of results

As evident from the forest plot in Fig. 2, the overall SMD was 0.41 (95%CI: 0.26; 0.56, p < 0.001) in favour of TMS, corresponding to a NNT of 5. The results of the subgroup analyses are available in Table 2A and the results of sensitivity analyses I–IV in Table 2B. Using follow-up data from at least four weeks after end of

treatment (see Supplementary Table 2 for further details) yielded an SMD of 0.27 (95%Cl: 0.05; 0.49). The effect in participants with predominantly negative symptoms was substantial (SMD = 0.50, 95%Cl: 0.25; 0.74), while no effect on depressive symptoms was observed (SMD = 0.02, 95%Cl: -0.17; 0.20) (see Supplementary Table 3). Stimulation of the L-DLPFC had a statistically significant (p=0.0002) larger effect than other sites (SMD = 0.55, 95%Cl: 0.38; 0.72 vs. SMD = 0.04, 95%Cl: -0.18; 0.25), however, there was considerable methodological heterogeneity in the "other sites" category. The SMD of different types of TMS did not statistically significantly differ (p=0.28) (SMD = 0.49, 95%Cl: 0.03; 0.95 for TBS, SMD = 0.43, 95%Cl: 0.27; 0.59 for rTMS and SMD = -0.32,

A. Overall effect size of TMS on negative symptoms and results from the subgroup analyses.	e symptoms and results f	rom the subgro	oup analyses.			
		Z	SMD (95% confidence interval)	ф	12	p difference between groups
Analyses						
Overall standardized mean difference (SMD)	2633	0.41	(0.26, 0.56)	<0.00001	%29	
Follow-up (≥4 weeks after end of treatment	nent	695	0.27 (0.05, 0.49)	0.02	42%	
Participants with PNS		1037	0.50 (0.25, 0.74)	0.0001	%89	
Depressive symptoms		743	0.02 (-0.17, 0.20)	0.86	30%	
Stratified analyses						
Site	L-DLPFC	2101	0.55 (0.38, 0.72)	<0.00001	%02	0.0002
	Other	532	0.04 (-0.18, 0.25)	0.74	27%	
Type	rTMS	2259	0.43 (0.27, 0.59)	<0.00001	%89	0.28
	TBS	268	0.49 (0.03, 0.95)	0.04	%69	
	Deep-TMS	33	-0.32 (-1.24, 0.61)	0.50	35%	
Stimulation Frequency (rTMS only)	1 Hz	297	0.05 (-0.21, 0.30)	0.72	2%	0.003
	> 1 Hz	1972	0.51 (0.33, 0.59)	<0.00001	%02	
Stimulation intensity	<100% of MT	1088	0.33 (0.11, 0.55)	0.003	63%	0.89
	≥100% of MT	1368	0.35 (0.20, 0.50)	<0.00001	40%	
Age ^a	≤median	1026	0.34 (0.17, 0.51)	0.0001	41%	0.43
	>median	1439	0.46 (0.22, 0.70)	0.0002	78%	
B. Results from the sensitivity analyses.						
Analyses	N	ĮS	SMD (95% confidence interval)	d		12
Excluding data from graphs	2451	Ö	0.43 (0.27, 0.59)	<0.00001		%69
Excluding high risk of Bias studies ^b	751	Ö	0.43 (0.14, 0.73)	0.004		71%
Excluding change-from- baseline studies	2592	Ö	0.43 (0.28, 0.57)	<0.00001		%29

Bold values indicate statistical significance.
L-DLPFC left dorsolateral prefrontal cortex, rTMS repetitive transcranial magnetic stimulation, MT motor threshold, TBS theta burst stimulation, deep-TMS deep transcranial magnetic stimulation, PNS predominant negative symptoms.

%19

0.0002

0.32 (0.15, 0.49)

1725

Excluding data taken from earlier reviews

*Median = 35.5 years. bNon-English language articles also excluded as they were not evaluated. Some studies had several active treatment groups and are included in several analyses.

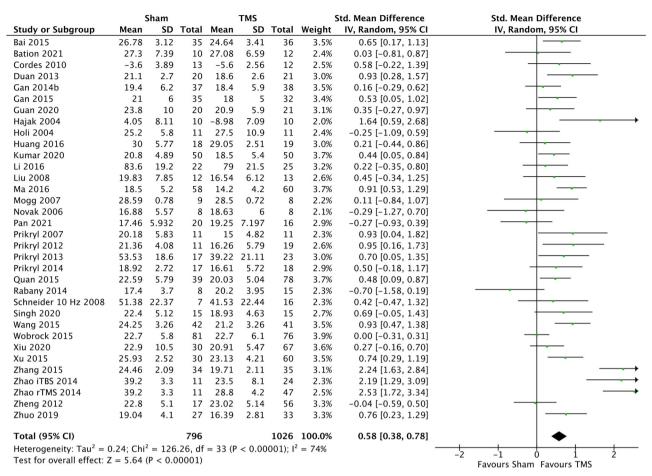


Fig. 3 Forest plot of standardized mean differences (effect sizes) of TMS targeting negative symptoms via stimulation of the left dorsolateral prefrontal cortex (L-DLPFC) at > 1 Hz.

95%CI: -1.24; 0.61 for deep-TMS). For forest plots of these subgroup analyses, see Supplementary Figs. 1–5. The sensitivity analyses following exclusion of (I) studies with data taken from graphs (SMD = 0.43, 95%CI: 0.27; 0.59), (II) studies with a high risk of bias (SMD = 0.42, 95%CI: 0.14; 0.73), (III) studies using change-from-baseline scores (SMD = 0.43, 95%CI: 0.28; 0.57), or (IV) studies with data extracted from other reviews (SMD = 0.32, 95%CI: 0.15; 0.49), did not impact the effect estimate substantially. A separate forest plot of the standardized mean differences (effect sizes) of the TMS studies targeting the left dorsolateral prefrontal cortex (L-DLPFC) at >1 Hz is shown in Fig. 3.

Risk of bias across studies

Based on the funnel plot examining study precision versus effect size (Fig. 4), we saw no qualitative evidence of asymmetry and this was confirmed with Egger's regression (p=0.9498). There were, however, outlying studies on both sides of the confidence interval. Notably in this regard, the sensitivity analysis (no. V) involving stepwise exclusion of the 10 most outlying studies yielded no substantial change in the efficacy estimate (see Supplementary Table 4).

DISCUSSION

Based on meta-analysis of 57 studies with a total of 2633 participants with schizophrenia (the vast majority) or schizoaffective disorder, we found a superior effect of active TMS on negative symptoms compared to sham treatment. The SMD was 0.41 (95% Cl: 0.26; 0.56) in favour of active TMS, translating to a NNT of 5.

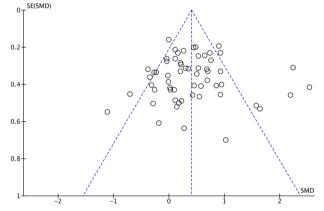


Fig. 4 Funnel plot examining study precision versus effect size. SE Standard error, SMD Standardized mean difference.

This result aligns with those from prior meta-analyses on the subject, as Aleman et al.¹⁵ found an SMD of 0.64 (95%Cl: 0.32; 0.96), He et al.¹⁶ an SMD of 0.41 (95%Cl: -0.35; 1.16), Wang et al.¹⁸ an SMD of 0.40 (95%Cl: 0.18; 0.62), and Osoegawa et al.¹⁷ an SMD of 0.19 (95%Cl: 0.07; 0.32). The superiority of active TMS on negative symptoms remained statistically significant in sensitivity analyses following (a) exclusion of data extracted from graphs, (b) exclusion of studies deemed to be at high risk of bias, and (c) exclusion of studies reporting change-from-baseline scores, respectively. Subgroup analyses suggested that using >1 Hz

stimulation (SMD = 0.51 vs. SMD = 0.05, p = 0.003) and targeting the L-DLPFC (SMD = 0.55 vs. SMD = 0.04, p = 0.0002) may be more effective. However, there was considerable heterogeneity across the included studies and these results should therefore be considered tentative. In contrast to the meta-analysis by Aleman et al.¹⁵, we found no support for the suggestion that TMS should have a particularly beneficial effect upon negative symptoms among younger patients (SMD = 0.34 vs. SMD = 0.46, p = 0.43). The meta-analysis by Aleman et al. was, however, based on data from a much smaller number of studies/participating patients compared to the present work. Indeed, while there are prior metaanalysis on the effect of TMS on negative symptoms, our updated version covers substantially more studies and participants (138%) more studies and 219% more participants than Aleman et al.¹⁵, 714% more studies and 539% more participants than He et al. 16 97% more studies and 83% more participants than Wang et al. 18, and 138% more studies and 142% more participants than Osoegawa et al.¹⁷) and should therefore be more representative of the state-of-the-art.

Several different brain areas were targeted by TMS in the studies included in this synthesis, in which subgroup analyses suggested that stimulation of the L-DLPFC may be particularly beneficial (SMD = 0.55). These results align with earlier studies that have found an inverse correlation between frontal lobe size and glucose metabolism, and negative symptom severity 107,108. Together with small sample size, this variability in target could in part explain why only eighteen studies found statistically significant effects of TMS since 17 of these targeted L-DLPFC. Hence, increased activity in the L-DLPFC due to magnetic stimulation could be the mechanism of action underlying the effect on negative symptoms of TMS as proposed in several of the largest included studies^{25,87,88}. Moreover, there is an increasing body of data suggesting that the DLPFC has a privileged relationship with other structures implicated in negative symptoms, including the midline cerebellum¹⁰⁹. For these reasons, circuitries involving the DLPFC will likely receive substantial attention in future efforts to relieve negative symptoms of schizophrenia and related psychotic disorders. Conversely, TMS of "other targets" than the L-DLPFC yielded no positive effect compared to sham treatment (SMD = 0.06). While there is considerable methodological heterogeneity among these studies, the quantitative synthesis converged on a null effect with low statistical heterogeneity. Furthermore, several of the studies targeting other sites than the L-DLPFC had other primary endpoints such as the severity of auditory hallucinations (the temporo-parietal cortex as target) with negative symptom severity as a secondary outcome^{36,40,67,71,77,84}, which likely contributes to the lack of effect compared to sham treatment.

There are limitations to this study, which should be acknowledged by the readers. First, as there are phenomenological overlaps between negative and depressive symptoms and since depression responds well to TMS^{110,111}, the relief of depressive symptoms during treatment could potentially confound the estimation of the effect on negative symptoms. However, the results from our analysis of data from studies measuring depressive symptoms in the context of schizophrenia (no effect of TMS) do not support this explanation. Second, 56% of the evaluated studies were regarded as having "high risk of bias", which is a substantially larger proportion compared to the 13% reported in the review by Wang et al.¹⁸. This difference, however, is predominantly a consequence of classification as we used the Cochrane Risk of Bias Tool 2.0, while Wang et al. used the Cochrane Risk of Bias Tool 1.0. The most common reason for studies being considered as "high risk" in the context of the present review was missing data. We employed a relatively conservative 10% cut-off for missing data, but there is no agreed upon threshold (https://training.cochrane.org/ handbook/current/chapter-08#section-8-5) and the proportion of studies classified as "high risk of bias" can thus vary considerably between reviews. Also, publication bias seems unlikely as Egger's regression test showed non-significant results (p = 0.9498), but cannot be ruled out. The two most outlying studies^{88,103} did not differ substantially in their treatment parameters as both targeted the L-DLPFC with >1 Hz over 20 days of treatment, however, the stimulation intensity (percent of MT) was not described for Zhang 2015. Third, we used a broad search strategy, but relevant studies may have been missed nevertheless. However, assuming that such potential misses occur at random, it should not have affected the reported efficacy estimates. Fourth, the inclusion of data drawn from reviews and graphs is suboptimal. However, the analyses excluding these data yielded results equivalent to those from the primary analysis. Fifth, there was significant heterogeneity in outcome across the included studies, with l^2 assessments at 50% or above in all but seven cases (66% in the primary analysis; Fig. 2). While this is likely partly due to the considerable heterogeneity of the TMS treatment provided across the included studies, other sources of heterogeneity, such as differences in sham conditions, patient populations, outcome measures, or random chance, are also likely contributors. Relatedly, in the review by He et al.¹⁶, a univariate metaregression of stimulation frequency, total simulation, motor threshold, stimulation site, study design, and type of coil was conducted. None of these factors were shown to be the main source of heterogeneity. Seventh, 24 studies could not be included due to unavailable data. These studies were, however, generally smaller, which reduces the impact of this limitation.

In conclusion, this systematic review and meta-analysis of data from sham-controlled studies suggests that TMS is efficacious in the treatment of negative symptoms of schizophrenia. Although it appears that targeting the L-DLPFC and using a stimulation frequency >1 Hz are the most efficacious settings, the optimal treatment parameters are yet to be established.

DATA AVAILABILITY

All data are available in the manuscript, figures, tables, and supplementary files.

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AUTHOR CONTRIBUTIONS

The study was designed by all authors. The screening and selection of articles was carried out by R.L. and T.D.N. The data extraction was performed by R.L. and F.H. The assessment of bias was carried out by R.L. The statistical analyses were conducted by R.L. and F.H. The results were interpreted by all authors. The manuscript was drafted by R.L. and S.D.Ø. T.D.N., A.M. and F.H. revised the manuscript for important intellectual content. The final manuscript was approved by all authors prior to submission.

COMPETING INTERESTS

A.M. has served as an advisor for Allergan Canada. S.D.Ø. received the 2020 Lundbeck Foundation Young Investigator Prize. Furthermore, S.D.Ø. owns units of mutual funds with stock tickers DKIGI and WEKAFKI, as well as units of exchange-traded funds with stock tickers TRET, QDVE, QDVH, EUNL and SADM.

ADDITIONAL INFORMATION

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Correspondence and requests for materials should be addressed to Søren D. Østergaard.

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