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Speciation and subsequent secondary contact in two edaphic endemic primroses driven by Pleistocene climatic oscillation

Masaya Yamamoto ¹ · Daiki Takahashi² · Kiyoshi Horita³ · Hiroaki Setoguchi²

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Abstract

Climatic perturbation during the Pleistocene era has played a major role in plant evolutionary history by altering species distribution range. However, the relative roles of climatic and geographic factors in the distribution dynamics remain poorly understood; in particular, the edaphic endemics. In this paper, we examine the evolutionary history of two ultramafic primroses, *Primula hidakana* and *Primula takedana*. These species are ecologically and morphologically distinct with disjunct distributions on Hokkaido Island, Japan. *Primula hidakana* is found on various rocks in southern Hokkaido and *P. takedana* in serpentine areas in northern Hokkaido. We performed population genetics analyses on nuclear and chloroplast data sets and tested alternative phylogenetic models of divergence using approximate Bayesian computation (ABC) analyses. Nuclear microsatellite loci clearly distinguished the two sister taxa. In contrast, chloroplast sequence variations were shared between *P. takedana* and *P. hidakana* to *P. takedana*. Paleodistribution modeling also supported the divergence model, and predicted their latitudinal range shifts leading to past secondary contact. Our findings highlight the importance of the distribution dynamics during the Pleistocene climatic oscillations in the evolution of serpentine plants, and demonstrate that tight species cohesion between serpentine and nonserpentine sister taxa has been maintained despite past interspecific gene flow across soil boundaries.

Introduction

Pleistocene climatic oscillations have played a key role in altering species distribution ranges and in providing opportunities for speciation through isolation (Comes and Kadereit 1998; Hewitt 2000, 2004). In other circumstances,

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Masaya Yamamoto mayamamo@hyogo-u.ac.jp

- ¹ Hyogo University of Teacher Education, 942-1 Shimokume, Katocity, Hyogo 673-1494, Japan
- ² Graduate School of Human and Environmental Studies, Kyoto University, Yoshida Nihonmatsu, Sakyo-ku, Kyoto 606-8501, Japan
- ³ Medical Plant Garden & Northern Ecological Garden, Faculty of Pharmaceutical Science, Health Sciences University of Hokkaido, 1757 Kanazawa, Tobetsu-cho, Ishikari-gun, Hokkaido 061-0293, Japan

repeated range shifts during climatic perturbation have also created opportunities for secondary contact between closely related species that were previously distributed in an allopatric manner. Accordingly, such distributional dynamics possibly result in interspecific hybridization (Rieseberg et al. 2003; Liu et al. 2014), reinforcement of reproductive barriers (Dobzhansky 1940; Kay and Schemske 2008), and transfer of adaptive alleles (Arnold et al. 2016). Interspecific gene exchange (i.e., introgression) in particular leads to homogenization during secondary contact of divergent gene pools, which may result in "reverse speciation" (Taylor et al. 2006). Therefore, gaining a greater insight into the latent roles of interspecific hybridization under Pleistocene climatic oscillations in a species' evolutionary history is of general interest to both molecular ecologists and evolutionary biologists, as this can improve our understanding of how species are generated and maintained in nature.

Severe climatic perturbation may eliminate edaphic endemics (plants restricted to soils with unsuitable chemical or physical characteristics for many plants) from their original location, because their narrow edaphic niches limit successful migration (Van de Ven et al. 2007). On the other

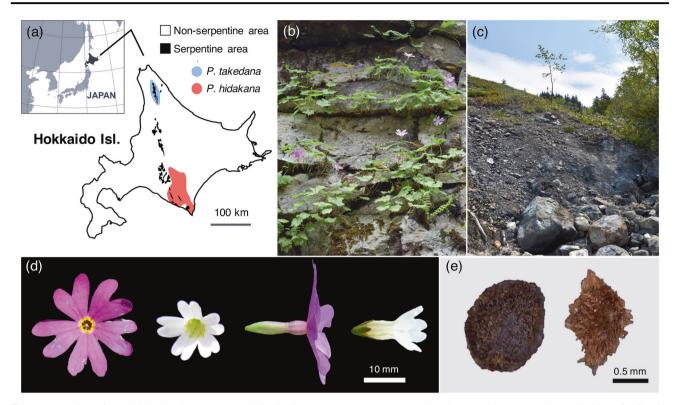


Fig. 1 a Location of Hokkaido Island and presumed distribution ranges of *Primula takedana* and *P. hidakana*. b Typical habitat of *P. hidakana*, Utabetsu-gawa stream, Erimo-cho, Horoizumi-gun, and Hokkaido. c Typical habitat of *P. takedana*, Nupuromapporo-gawa

stream, Horonobe-cho, Teshio-gun, and Hokkaido. **d** Corolla morphologies of *P. hidakana* (pink) and *P. takedana* (white). **e** Seed morphologies of *P. hidakana* (left) and *P. takedana* (right)

hand, because the niches of edaphic endemics are defined by soil type rather than climate, they may be able to endure climate change in their geographical ranges. Rather, several lines of evidence suggest that climatic cooling during glacial periods provides edaphic endemics with newly available but edaphically unsuitable habitats. Serpentine endemic plants, Arenaria norvegica and Minuartia laricifolia subsp. ophiolitica, both expanded out of serpentine outcrops when new nonserpentine habitats became available as the forest retreated during glacial periods (Rune 1953; Moore et al. 2013). These perspectives trigger another idea that Pleistocene climatic oscillations provided opportunities for secondary contact between edaphic endemics and their sister species, which were distributed allopatrically during the interglacial periods. Climate has long been thought to have played a key role in the evolutionary origin and persistence of serpentine plants and other edaphic endemics (Anacker and Harrison 2012; Anacker 2014; Rajakaruna 2018), yet the importance of past climatic oscillations in the evolution of edaphic endemics and their potential population dynamics, including interactions with their edaphically nonspecialized relatives, remains poorly understood.

Here, we examined the evolutionary history of recently diverged allopatric sister species, *Primula hidakana* (Miyabe and Kudo) Nakai and *Primula takedana* Tatew. (Primulaceae). We used phylogeographic and population genetics approaches to verify whether the Pleistocene climatic fluctuations drove interspecific gene flow between closely related taxa occurring in disjunct ultramafic habitats. Both species are diploid (2n = 24) self-incompatible perennial herbs that occur on the peridotite-rich metamorphic and serpentine belts of Hokkaido Island, Japan (Fig. 1a). Previous molecular phylogenies using both chloroplast and nuclear DNA showed that these two species have a robust sister relationship (Mast et al. 2001; Richards 2003; Yamamoto et al. 2017a), and inferred that P. takedana diverged recently (mid- to late Pleistocene) from P. hidakana (Yamamoto et al. 2017a). P. hidakana is found across the Hidaka ranges in southern Hokkaido (Fig. 1a), and occurs on wet, shaded rocky cliffs along a stream and a ridgeline in the mountains (Fig. 1b). Although the rock type is multifaceted, including peridotite, diabase, gneiss, and amphibolite (Watanabe 1971), P. hidakana does not occur in serpentine soils. Conversely, P. takedana is endemic to the denuded slopes of serpentine rocks partly weathered into clay along a stream in northern Hokkaido (Fig. 1a, c), and has a narrow and patchy distribution (Saito 1970). Their distribution ranges are separated by ~200 km in a north-south direction (Fig. 1a), and no hybrids have been reported in their distribution ranges or in the interval zone.

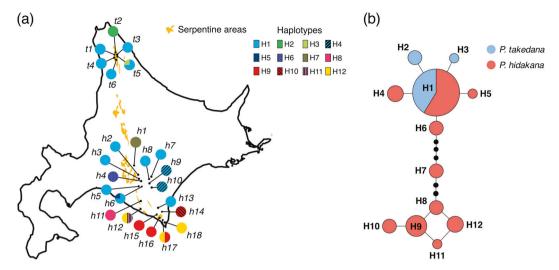


Fig. 2 Geographical distribution (**a**) and a median joining network (**b**) of the 12 chloroplast DNA haplotypes detected in *P. hidakana* and *P. takedana*. Location of population sampling sites (t_1-t_6, h_1-h_18) and

chloroplast haplotypes (H1-H12) are shown in italics and bold type, respectively

Although both species have been classified as ultramafic *Primula* (Kitamura 1950), there are differences between the species in relation to their morphological traits and physiological responses: nickel accumulation in *P. takedana* plants is much higher than in *P. hidakana* (Horie et al. 2000); and *P. takedana* has the typical morphological traits associated with adaptation to serpentine soils, including smaller flower size (Fig. 1d) and well-developed creeping rhizomes (Richards 2003); and the seed surface of *P. takedana* is covered with minute membrane organizations (Fig. 1e). These multifaceted observations indicate that *P. takedana* is the only species adapted to serpentine soils of the two ultramafic *Primula*, yet common garden studies testing the serpentine tolerance of *P. hidakana* are currently lacking.

In this study, we used nuclear microsatellite markers as well as multiple chloroplast DNA sequences to reconstruct the evolutionary history of the two ultramafic *Primula* in Hokkaido, Japan, utilizing an approximate Bayesian computational (ABC) approach (see Csilléry et al. 2012) and ecological niche modeling (ENM). The aims of this study were to elucidate the origin of serpentine endemic *P. takedana*, test alternative phylogenetic hypotheses for the divergence process, reconstruct historical range shifts of the currently disjunctive species, and better understand the role of past climatic changes in the speciation process of the edaphic endemic primroses.

Materials and methods

Population sampling, chloroplast DNA sequencing, and nuclear microsatellite genotyping

A total of 384 individuals from 24 populations (18 of *P. hidakana* and 6 of *P. takedana*) were collected throughout

the entire range of the two species on Hokkaido Island, Japan (Fig. 2a). Fresh leaves from 16 individuals from each population were collected in silica gel, and genomic DNA was extracted from silica gel–dried leaf material following the cetyltrimethyl ammonium bromide method (Doyle 1990).

Polymorphisms in several plastid sequences using previous phylogeny (*trnL-F*, *trnH-psbA*, *trnD-T*, and *rps16*; Yamamoto et al. 2017a) were tested in four individuals of *P. hidakana* and *P. takedana*. This trial indicated that the two noncoding regions *rps16* and *trnD-T* were the most informative markers. Each plastid region was amplified in four individuals per population using the primer pairs *rpsF* (5'-GTGGTAGAAAGCAACGTGCGACTT-3')/*rpsR2* (5'-TCGGGATCGAACATCAATTGCAAC-3'), and *trnD* (5'-AGGGCGGTACTCTAACCAA-3')/*trnT* (5'-CGATGA CTTACGCCTTACCA-3') (Oxelman et al. 1997; Honjo et al. 2004). All sequencing experiments followed the methods described by Yamamoto et al. (2017a), and all new sequences have been deposited in the GenBank/DDBJ/ EMBL database (LC330946–LC330969).

The nuclear genetic diversity of all 384 individuals was assessed using the following 12 microsatellite loci: Pok25, Pok32, Pok39, Pre02, Pre28, Pre51, Pre52, Pto06, Pto12, Pto21, Pto29, and Pto38 (Yamamoto et al. 2018a, 2018b). For all primer pairs, one of three different M13 universal sequences (5'-CACGACGTTGTAAAACGAC-3', 5'-TGT GGAATTGTGA GCGG-3', or 5'-CTATAGGGCACGC GTGGT-3') and a pigtail sequence (5'-GTTTCTT-3') were added to the 5' end of each forward primer and the 3' end of each reverse primer, respectively. The PCR reaction mix contained 5 μ L Multiplex PCR Master Mix (Qiagen, Hilden, Germany), sterile water, 0.01 μ M forward primer, 0.2 μ M reverse primer, 0.1 μ M fluorescence-labeled M13 primer, and 10–50 ng of template DNA in a total volume of 10 μ L. The PCR amplification sequence was an initial 15 min of denaturation at 94 °C and 35 cycles consisting of 30 s of denaturation at 94 °C, 3 min of annealing at 60 °C, and 1 min of extension at 68 °C, with a final 20 min extension at 68 °C. Amplified products were mixed with highly deionized (Hi-Di) formamide and GeneScan 500 ROX size standard (Applied Biosystems, Foster City, CA, USA) and electrophoresed using an ABI 3130 Genetic Analyzer (Applied Biosystems). The fragment sizes were analyzed using GeneMapper (Applied Biosystems) and calibrated with AUTOBIN v0.9 (Salin 2013).

Data analyses of chloroplast haplotypes

Chloroplast regions rps16 and trnD-T were assembled and aligned with the Muscle program implemented in MEGA7 software (Kumar et al. 2016) using default settings with minor manual adjustment. The two plastid sequences were combined into a single sequence using SequenceMatrix v1.7.8 (Vaidya et al. 2011). Each individual was assigned to the different haplotypes using DnaSP v5 (Librado and Rozas 2009). DnaSP was also used to estimate molecular diversity indexes, including the number of haplotypes, haplotype diversity (h), and nucleotide diversity (π) . Visualization of haplotype diversity distribution was performed via a median-joining haplotype network reconstructed using PopART v1.7 (Leigh and Bryant 2015). An analysis of molecular variance (AMOVA) was performed to partition variation within and among species and populations using Arlequin v3.5.1 (Excoffier and Lischer 2010). The significance of Φ -statistics was determined with 1000 permutations.

Data analyses, nuclear microsatellites

For all 12 nuclear microsatellite loci, the absence of linkage disequilibrium (LD), deviation from Hardy-Weinberg equilibrium (HWE), and the null allele frequency between loci in each population were detected using the program GENEPOP v4.2 (Raymond and Rousset 1995). LD and HWE tests were verified using a Markov chain method with 1000 dememorization steps, 100 batches, and 1000 iterations per batch. Null allele frequencies were estimated by maximum-likelihood the estimator based on the expectation-maximization algorithm (Dempster et al. 1977) with default settings. In addition, the effect of null alleles on population genetic differentiation (F_{ST}) was evaluated by comparing Weir's F_{ST} (Weir 1996) before and after correction for null alleles using FreeNA software (Chapuis and Estoup 2007).

At the species level, for each locus and across loci, the following genetic diversity measurements were calculated using FSTAT v2.9.3 (Goudet 1995): *A*, number of alleles; $A_{\rm R}$, allelic richness; $H_{\rm O}$, observed heterozygosity within populations; $H_{\rm S}$, mean expected heterozygosity within populations; $H_{\rm T}$, total expected heterozygosity across populations; and $F_{\rm IS}$, fixation index. We also calculated the species-specific alleles and lineage-specific alleles that correspond to Bayesian clustering analysis (K = 2) in *P*. *hidakana* (see Fig. 5c).

At the population level, three genetic diversity statistics were calculated with FSTAT: $A_{\rm R}$, allelic richness; $H_{\rm O}$, observed heterozygosity; and $H_{\rm E}$, gene diversity (Nei 1987). In addition, the number of locally common alleles (i.e., frequency of more than 5% occurring in less than 25% of populations; $A_{\rm C}$) was calculated using GenAlEx v6.5 (Peakall and Smouse 2006). To graphically represent spatial patterns in population genetic diversity, smoothed interpolated maps were produced using an inverse distance weighting (IDW) interpolation method in Quantum GIS v2.18 (OSGeo, Beaverton, OR, USA) for the genetic diversity measurements of allelic richness, gene diversity, and locally common alleles. IDW in Quantum GIS was set at 2.0.

To evaluate genetic differentiation within species, we first calculated global F_{ST} (Weir and Cockrham 1984) and $R_{\rm ST}$ (Rousset 1996) using the SPAGeDi v1.5 (Hardy and Vekemans 2002). Index F_{ST} is based on only allele frequency variation, whereas R_{ST} is an index of genetic differentiation accounting for allele frequency and variance in allele size for microsatellite loci under a stepwise mutation model (Slatkin 1995). Following Hardy et al. (2003), SPAGeDi can be used to determine the relative importance of genetic drift versus mutation in genetic differentiation between populations by testing whether the observed R_{ST} value was significantly greater than the permuted $R_{\rm ST}$ values $(pR_{ST}; 10,000 \text{ permutations})$. If the observed R_{ST} is significantly greater than pR_{ST} , a stepwise mutation model is the most likely explanation for genetic differentiation, suggesting a phylogeographic signal. Moreover, insignificant differences between R_{ST} and pR_{ST} suggest that allele size is not important, and that F_{ST} should be used instead of $R_{\rm ST}$. Because $R_{\rm ST}$ was not significantly greater than $pR_{\rm ST}$ in this study (see Results), AMOVA was carried out based on $F_{\rm ST}$ using Arlequin.

Population genetic structures were examined using a Bayesian clustering analysis in STRUCTURE v2.3.2 (Pritchard et al. 2000). STRUCTURE analysis was conducted for all samples across the two species, and then the analysis was run separately within each species. Under an admixture model with correlated allele frequency, 20 independent simulations were run for each K (K = 1-20 for all samples, K = 1-15 for *P. hidakana*, and K = 1-10 for *P. takedana*) with 5×10^5 Markov chain Monte Carlo (MCMC) steps and a burn-in period of 10^5 iterations. For

each dataset, the most likely value of *K* was determined by the ΔK method (Evanno et al. 2005) using STRUCTURE HARVESTER v0.6.94 (Earl and vonHoldt 2012). CLUM-PAK (Kopelman et al. 2015) was used to average the outputs from multiple STRUCTURE runs and produce the graphical results.

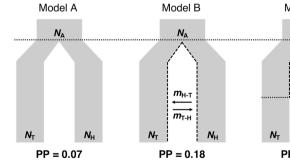
To further detect the genetic grouping of microsatellite data, principal component analyses (PCA) were performed on a matrix of covariance values calculated from population allele frequency using GENODIVE (Meirmans and Van Tienderen 2004). We also used the Neighbor-Net algorithm (Bryant and Moulton 2004) to construct phylogenetic relationships between individuals based on microsatellite data. Construction of a Neighbor-Net phenogram based on the Bruvo genetic distance (Bruvo et al. 2004) was performed with SPLITSTREE4 v4.14.4 (Huson and Bryant 2005). In addition, spatial principal component analysis (sPCA) was performed to investigate genetic clines across the species distribution range using the R package adegenet v2.0.0 (Jombart 2008; Jombart and Ahmed 2011). sPCA was conducted in between- and within-species data sets using a distance-based connection network that assumes that all samples have at least one connection.

Evaluating speciation models by approximate Bayesian computation

We compared the following four hypothesized models of the speciation of *P. hidakana* and *P. takedana* (represented in Fig. 3) through analyses of the nuclear microsatellite data set using the ABC approach, as implemented in the ABC-TOOLBOX software package (Wegmann et al. 2010): Model A, which assumed no gene flow throughout the divergence process (isolation model); Model B, which assumed constant gene flow since divergence (isolationwith-migration model); Model C, which assumed gene flow only in the initial phase of the divergence process (sympatric speciation model); and Model D, which assumed secondary contact after speciation and gene flow subsequent to an initial allopatric phase of divergence (isolation with secondary contact model). The parameters for prior population size (the size of each population remained constant), timing, and migration were assumed to have a log-uniform distribution (prior intervals for all model parameters are detailed in Supplementary Table 1).

We examined nine statistics summarizing population genetic parameter information in the ABC analyses. These statistics were calculated using ARLSUMSTAT v3.5.2 (Excoffier and Lischer 2010) and are detailed in Supplementary Table 2. In the coalescent simulation step, one million simulations were generated for each model using the standard algorithm of ABCTOOLBOX and the program FASTSIMCOL2 v2.5.2.21 (Excoffier and Foll 2011). When simulating, the average mutation rate across loci was estimated to be 1.0×10^{-6} – 1.0×10^{-3} substitutions/site/year (Thuillet et al. 2002; Vigouroux et al. 2002; Resende-Moreira et al. 2017).

To evaluate which scenario best fit the observed summary statistics, the posterior probabilities (the marginal density of the model of interest divided by the sum of the marginal densities of all models) were estimated using the rejection method, and analyses were conducted with a range of tolerance values ($\delta = 0.01-0.001$). For the best-fitting model, the sufficiency of the summary statistics and the robustness of our results were evaluated by posterior predictive checks (Csilléry et al. 2012). Population parameters in the best model were inferred with the neural network procedure (Blum and François 2010). Although accurate generation times (between seed germination and first flowering) of *P. hidakana* and *P. takedana* have not been recorded, the generation time of *P. reinii* var. *rhodotricha*, which is placed in the same section, has been estimated to



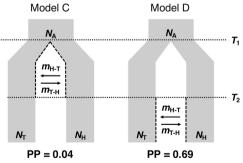


Fig. 3 Schematic representation of the four evolutionary models of speciation compared using approximate Bayesian computation (ABC): Model A, divergence without gene flow; Model B, isolation with migration; Model C, sympatric speciation; Model D, isolation with secondary contact model. Estimated posterior probabilities (PP) for each model are shown below each model (results based on $\delta = 0.001$).

Abbreviations are as follows: $N_{\rm T}$ and $N_{\rm H}$, current effective population size of *P. takedana* and *P. hidakana*, respectively; $N_{\rm A}$, effective population size of their ancestor; T_1 , divergence time; T_2 , the time at which gene flow began; $m_{\rm H-T}$ and $m_{\rm T-H}$, asymmetrical migration rate from *P. hidakana* to *P. takedana* and from *P. takedana* to *P. hidakana*, respectively

be 2–3 years (Yamamoto et al. 2017b). Taking this into account, we assumed a generation time of three years for *P*. *hidakana* and *P*. *takedana* to estimate divergence time between the two species. All analyses described above were performed using the package *abc* (Csilléry et al. 2012) in R v3.4.1 (R Development Core Team 2017).

Bayesian population demographic analyses

In addition to evaluating speciation models using ABC analyses, we used the VarEff method implemented in the R package vareff (Nikolic and Chevalet 2014) to examine the past historically effective population sizes of P. takedana and P. hidakana. This method estimates prior changes in effective population size from microsatellite data (motif distance frequencies between alleles) by resolving coalescence theory and using an approximate likelihoods MCMC approach (Nikolic and Chevalet 2014). We used the prior parameter settings for each species (detailed in Supplementary Table 3) based on the results of our ABC analyses and a variety of preliminary runs. We ran each analysis under a two-phased mutation model with a proportion of 0.20 for multistep mutations (Peery et al. 2012), for 10^4 MCMC steps (NumberBatch = 10,000, LengthBatch = 10), sampling every 10 steps (SpaceBatch = 10), with an acceptance ratio of 0.25 (AccRate = 0.25). To assess time since changes in years, we used the estimated mutation rate 6.61×10^{-5} as estimated by our ABC results.

Ecological niche modeling

For ENM, we combined climate and nonclimate data coupled with species' locality data at a spatial resolution of 30" $(\sim 1 \text{ km}^2)$. Coordinates from 33 localities for *P. hidakana* and 24 for P. takedana were obtained from our field collections and recent distribution research on *P. takedana* (Ito et al. 2016). The 19 bioclimatic variables for the current era (ca. 1950-2000) and the last interglacial (LIG; ca. 140-120 ka) were downloaded from the WorldClim database (http://www.worldclim.org/), whereas those for the last glacial maximum (LGM; ca. 21-18 ka) were obtained from two models; the Community Climate System Model (CCSM3: Collins et al. 2006) and the Model for Interdisciplinary Research on Climate (MIROC 3.2: Hasumi and Emori 2004). To avoid issues with multicollinearity (Dormann et al. 2013), six climatic variables (annual mean temperature, mean diurnal range as the mean of the monthly difference between the maximum temperature and the minimum temperature, maximum temperature of the warmest month, mean temperature of the warmest quarter, precipitation of the wettest quarter, and precipitation of the coldest quarter) were selected as uncorrelated climatic data. For nonclimate datasets, geology data were downloaded from a seamless digital geological map of Japan (1:200,000; Geological Survey of Japan 2015), whereas slope data were extracted from the National Land Numerical Information, Japan (http://nlftp.mlit.go.jp/).

ENMs were conducted using a maximum-entropy model implemented in MAXENT v3.3 (Phillips and Dudík 2008), which models presence-only species distribution data. All analyses were performed using the logistic output and a convergence threshold of 10^{-5} with 500 iterations, and then projected for the paleoclimate layers using both clamping and extrapolation methods (Merow et al. 2013). A total of 75% of the species' records were included for training, and 25% for testing the model. Jackknife testing was conducted to assess the importance of each variable in the selected model. Model performances were evaluated using area under the curve (AUC) and true skill statistics (TSS; Allouche et al. 2006) based on 20 replicate runs. The 10th percentile training presence logistic threshold from MAXENT was used as the threshold for TSS because it produces accurate results when true absence data is lacking (Escalante et al. 2013).

We first ran the model for each species in MAXENT using all selected variables (i.e., "full model") and then assessed the transferability of six climatic variables in the paleoclimate layers, because projecting species niche models into different time periods with possible nonanalog climate is not only liable to error, but also leads to unconvincing outcomes (e.g., Heikkinen et al. 2006; Fitzpatrick and Hargrove 2009). Based on the preliminary analyses, we removed the climatic variables that showed a broad extrapolated area and little contribution to the full model. Finally, we selected a model with two climatic variables (mean temperature of warmest quarter and precipitation of wettest quarter) and two nonclimate variables (geology and slope). The AUC and TSS values of the selected model were very similar to those of the full models. The selected model contained only very few or no values outside the training data range, which shows that no areas were strongly affected by variables beyond their training range.

The potential distribution maps of each species in three different time periods (current, LGM, and LIG) were illustrated with Quantum GIS. In addition, the overlap between the ranges of the two species at different time periods was computed using ENMTOOLS v1.3 (Warren et al. 2010) with probabilities greater than 0.307 (the 10th percentile training presence logistic threshold).

Results

Chloroplast haplotypes

The concatenated cpDNA sequences (rps16 and trnD-T) were aligned with a total length of 1355 bp, including 14

Table 1 Genetic diversity and
genetic differentiation of
Primula takedana and *P.*
hidakana were obtained by
chloroplast and nuclear DNA

cpDNA		Nuclear microsatellite								
h	$\pi (\times 10^{-2})$	A	A_{P}	$A_{\rm R}$	H _O	$H_{\rm S}$	H_{T}	$F_{\rm IS}$	$F_{\rm ST}$	R _{ST}
										0.021
	$\frac{1}{h}$ 0.359	$\frac{1}{h} \qquad \pi (\times 10^{-2})$	$\frac{1}{h} \qquad \pi (\times 10^{-2}) \qquad \overline{A}$ 0.359 0.028 86	$\frac{1}{h} \qquad \pi \ (\times 10^{-2}) \qquad \overline{A} \qquad A_{\rm P} \\ 0.359 \qquad 0.028 \qquad 86 \qquad 15$	$\frac{1}{h} \qquad \pi (\times 10^{-2}) \qquad \overline{A} \qquad A_{\rm P} \qquad A_{\rm R}$ 0.359 0.028 86 15 3.42	$\begin{array}{c ccccccccccccccccccccccccccccccccccc$	$\begin{array}{c ccccccccccccccccccccccccccccccccccc$	$\begin{array}{c ccccccccccccccccccccccccccccccccccc$	$\frac{1}{h} \qquad \pi (\times 10^{-2}) \qquad \overline{A} \qquad A_{\rm P} \qquad A_{\rm R} \qquad H_{\rm O} \qquad H_{\rm S} \qquad H_{\rm T} \qquad F_{\rm IS}$ $0.359 \qquad 0.028 \qquad 86 \qquad 15 \qquad 3.42 \qquad 0.362 \qquad 0.412 \qquad 0.441 \qquad 0.121$	$\begin{array}{c ccccccccccccccccccccccccccccccccccc$

h haplotype diversity, π nucleotide diversity, *A* number of alleles, A_p number of private alleles, A_R allelic richness, H_0 observed heterozygosity, H_S mean expected heterozygosity within populations, H_T total expected heterozygosity, F_{IS} fixation index, F_{ST} among-population differentiation based on differences in allele frequencies, R_{ST} among-population based on differences in allele size

Table 2 Results of the analysisof molecular variance(AMOVA) for chloroplast DNAand nuclear microsatellite data

	Source of variation	d.f.	SS	VC	TVP	Φ -statistics
cpDNA						
Total	Between species	1	2.344	0.028	7.0	$\Phi_{\rm CT} = 0.070$
	Among populations within species	22	29.50	0.323	80.9	$\Phi_{\rm SC} = 0.869^{*}$
	Within populations	72	3.50	0.049	12.2	$\Phi_{\rm ST} = 0.878^{*}$
P. takedana	Among populations	5	3.375	0.158	79.2	$\Phi_{\rm ST} = 0.792^{*}$
	Within populations	18	0.75	0.042	20.8	
P. hidakana	Among populations	17	24.069	0.341	87.0	$\Phi_{\rm ST} = 0.870^{*}$
	Within populations	54	2.75	0.051	13.0	
Nuclear micros	satellite					
Total	Between species	1	397.96	2.038	20.5	$\Phi_{\rm CT} = 0.205^{*}$
	Among populations within species	22	35.76	2.055	20.6	$\Phi_{\rm SC} = 0.259^{*}$
	Within populations	360	5.86	5.686	58.9	$\Phi_{\rm ST} = 0.411^{*}$
P. takedana	Among populations	5	47.797	0.222	8.2	$\Phi_{\rm ST} = 0.083^{*}$
	Within populations	186	457.531	2.460	91.7	
P. hidakana	Among populations	17	525.875	0.869	21.7	$\Phi_{\rm ST} = 0.217*$
	Within populations	558	1748.219	3.133	78.3	

d.f. degrees of freedom, *SS* sum of squares, *VC* variance components, *TVP* total variance percentages *P < 0.05

polymorphic sites and 19 insertion/deletion polymorphisms (indels). The chloroplast haplotype diversity (*h*) and nucleotide diversity (π) for each species are summarized in Table 1. The widespread *P. hidakana* had much higher levels of genetic diversity (h = 0.813; $\pi = 3.59 \times 10^{-3}$) than the geographically more restricted *P. takedana* (h = 0.359; $\pi = 2.80 \times 10^{-4}$).

A total of 12 chloroplast haplotypes was identified among the 96 accessions including *P. hidakana* and *P. takedana*. Of these, nine haplotypes were specific to *P. hidakana* (H4–H12) and two to *P. takedana* (H2 and H3), whereas H1 was the most widespread and common haplotype for both species (Fig. 2). The haplotype network showed that two haplotypes specific to *P. takedana* (H2 and H3) and three haplotypes found in the northern part of the geographical distribution of *P. hidakana* (H4–H6) were linked to the main haplotype H1 by single steps in a "star-like" network (Fig. 2b). Conversely, haplotypes H8–H12 found in *P. hidakana* were six steps away from the main haplotype, H1, and were all confined to the southern part of their distribution range (Fig. 2b). The hierarchal AMOVAs revealed that only 7.0% of the genetic variation was partitioned between species with low and statistically nonsignificant genetic differentiation $(\Phi_{\rm CT} = 0.070; P > 0.05)$ (Table 2). On the other hand, high and significant among-population differentiation was observed for both species ($\Phi_{\rm ST} = 0.870$ for *P. hidakana* and $\Phi_{\rm ST} = 0.792$ for *P. takedana*).

Nuclear microsatellites

Each locus used in this study has sufficient polymorphism to conduct population genetic analyses (Supplementary Table 4). Furthermore, for all nuclear microsatellite markers, no evidence was found for a systematic departure from HWE at any particular locus in all populations, and no pairs of loci consistently showed LD in all populations. The estimated frequency of null alleles across populations varied from 0.02 to 0.31, with an average frequency of 0.06 in all 12 loci (Supplementary Table 5). However, no clear differences between uncorrected $F_{\rm ST}$ values and $F_{\rm ST}$ values corrected for all null alleles were found (Supplementary Table 6).

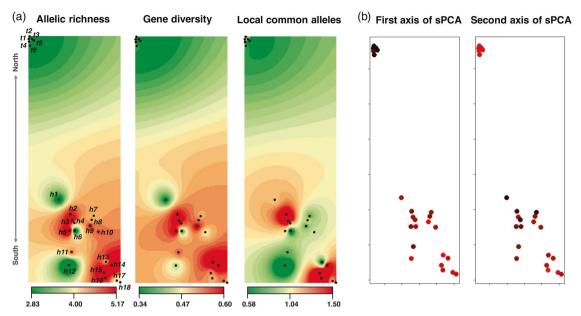


Fig. 4 Geographical distribution of the genetic diversity among *P. takedana* (t1-t6) and *P. hidakana* (h1-h18) populations (**a**). The axis scores of the spatial principal component analysis (sPCA) (**b**); the

locations of the samples are represented by circles that are colored according to a black to red gradient corresponding to the eigenvalues of the first (left) and second (right) axis of the sPCA, respectively

Genetic diversity measurements are summarized for all loci at the species (Table 1) and population levels (Supplementary Table 7). At the species level, all genetic diversity measurements $(A_{\rm R}, H_{\rm O}, H_{\rm S}, \text{ and } H_{\rm T})$ were higher in P. hidakana (e.g., $H_{\rm S} = 0.525$) than in P. takedana ($H_{\rm S} =$ 0.412). Furthermore, nonuniform spatial distributions of genetic diversity were observed for P. hidakana, in contrast to P. takedana, which consistently showed low genetic diversity among populations. The geographically northern (h2-h5) and southern (h13-h16) populations indicated higher genetic diversity than the central area of the distribution range of P. hidakana (Fig. 4a). The h3 population harbored the highest genetic diversity ($A_{\rm R} = 4.92$, $H_{\rm S} =$ 0.603, and $A_{\rm C} = 1.50$), closely followed by the h14 population $(A_{\rm R} = 5.17, H_{\rm S} = 0.591, \text{ and } A_{\rm C} = 1.17)$ (Supplementary Table 7).

The number of species-specific alleles was larger in *P. hidakana* than that in *P. takedana* (Table 1), while in the three subdivisions investigated [*P. takedana* populations, and northern (h1–2, h4–7, and h11) and southern (h3, h8-18) *P. hidakana* populations], the southern *P. hidakana* population exhibited a greater number of private alleles per locus (= 3.33) compared with *P. takedana* (= 1.25) and northern *P. hidakana* populations (= 1.50).

Based on the $F_{ST}-R_{ST}$ comparison, no significant difference was found in *P. hidakana* populations, whereas the R_{ST} value was significantly lower than the F_{ST} value in *P. takedana* populations (Table 1), suggesting that genetic drift is more important in this case. The AMOVAs using F_{ST} showed that 20.5% of the total nuclear microsatellite genetic variation was partitioned between species with moderate and significant genetic differentiation ($\Phi_{\rm CT} = 0.205$, P < 0.05). In contrast to the pronounced levels of amongpopulation genetic differentiation detected in *P. hidakana* ($F_{\rm ST} = 0.213$, $\Phi_{\rm ST} = 0.217$), weak to moderate population differentiation was observed for *P. takedana* ($F_{\rm ST} = 0.079$, $\Phi_{\rm ST} = 0.083$) (see Tables 1 and 2).

In the STRUCTURE analysis using all samples, the ΔK statistic peaked at K = 2, which indicates optimal division of the nuclear microsatellite data into two species (Supplementary Fig. 1a). However, some admixed individuals were also found in the P. hidakana populations (particularly in h3-5 and h14). These results were consistent throughout with other structuring results (see Fig. 5a). Further analysis indicated divergence of the northern populations of P. *hidakana* from the other populations at K = 3. For K = 4. some of the individuals in the cluster (blue or purple) in K =3 were reassigned to a new cluster (green), which was distributed and scattered across P. hidakana populations. Further analyses within each species are shown in Fig. 5b, c. The ΔK statistics peaked at K = 3 and K = 2 for *P*. takedana and P. hidakana, respectively (Supplementary Fig. 1). Within P. takedana, although a weak genetic structure was detected, genetic admixture within three clusters was observed in all populations (Fig. 5b). Within P. hidakana, the two assigned clusters largely corresponded with the results of K = 3 using all samples across species (Fig. 5c).

The results of PCA and Neighbor-Net tree also supported the existence of two distinct genetic units corresponding to each species (Supplementary Fig. 2). sPCA distinguished the two species in the score of the first principle component, whereas the second axis differentiated the northern

.15

MA

11,18

101

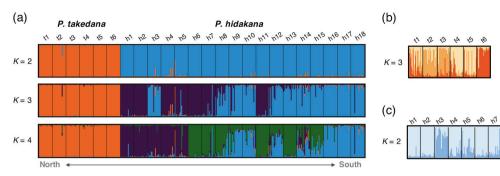


Fig. 5 Relationship between *P. takedana* and *P. hidakana* based on 12 nuclear microsatellite loci. **a** Bayesian clustering of all *P. takedana* and *P. hidakana* individuals (K = 2-4). **b** Bayesian clustering of *P.*

takedana individuals (K = 3). c Bayesian clustering of *P. hidakana* individuals (K = 2)

n8 n9 . 10

12.13

 Table 3 Estimates of the posterior distribution of demographic parameters revealed by Approximate Bayesian Computation (ABC) for the best model (model D)

Parameter	N _A	$N_{ m H}$	N _T	<i>m</i> _{H-T}	m _{T-H}	M _{H-T}	$M_{ m T-H}$	T_1 (years)	T_2 (years)
Mode					2.17×10^{-5}				
HPD 95 low	3.12×10^{5}	1.30×10^4	1.11×10^{3}	1.76×10^{-7}	1.27×10^{-7}	7.54×10^{-3}	5.62×10^{-4}	8.02×10^{4}	3.15×10^4
HPD 95 high	1.74×10^{7}	6.21×10^{5}	9.09×10^{4}	4.89×10^{-4}	1.13×10^{-4}	2.10×10^1	5.04×10^{-1}	4.62×10^{5}	3.06×10^{5}

 $N_{\rm A}$ effective population size of ancestral population, $N_{\rm H}$ current effective population size of *P. hidakana*, $N_{\rm T}$ current effective population size of *P. takedana*, $m_{\rm H-T}$ asymmetric migration rate from *P. hidakana* to *P. takedana*, $m_{\rm T-H}$ asymmetric migration rate form *P. hidakana* to *P. takedana*, $m_{\rm T-H}$ asymmetric migration rate form *P. hidakana*, $M_{\rm H-T}$ effective numbers of migrants per generation from *P. hidakana* to *P. takedana*, computed as $m_{\rm H-T} \times N_{\rm H}$, $M_{\rm T-H}$ effective numbers of migrants per generation from *P. hidakana*, computed as $m_{\rm T-H} \times N_{\rm T}$, T_1 divergence time between *P. hidakana* and *P. takedana*, T_2 the timing at which migration starts

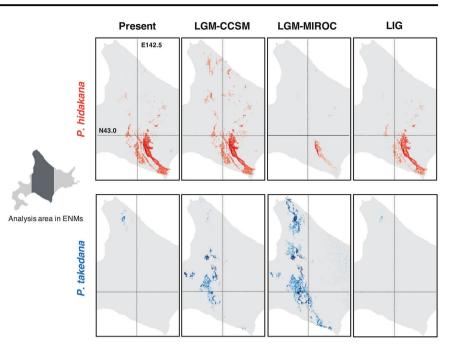
populations of *P. hidakana* from the others (Fig. 4b). At the species levels, although a latitudinal genetic cline was clearly detected within *P. hidakana* (Supplementary Fig. 6), several populations (e.g., h3 and h9) were discriminated from other northern populations. Conversely, such a spatial structure was weaker in the case of *P. takedana*.

Evaluating speciation models using approximate Bayesian computation

Based on ABC analyses, the isolation with secondary contact model (model D) was the most probable of the four proposed models for the speciation process of P. hidakana and P. takedana; the other three models received weaker support at all tolerance levels (Fig. 3, see also Supplementary Fig. 3). Posterior predictive checks confirmed that our observed summary statistics fell within distributions produced by parameter estimates from 10,000 randomly chosen accepted simulations in model D (Supplementary Fig. 4). Marginal posterior distributions of model parameters in model D are shown in Supplementary Fig. 5 and summarized in Table 3. Posterior parameter estimates suggested that P. hidakana and P. takedana diverged around 112 ka (95% highest posterior density interval, HPDI: 462–80 ka) and migration started at 45 ka (HPDI: 306-30 ka), assuming an average generation time of 3 years. The effective number of migrants per generation was unidirectional and quite high from *P. hidakana* to *P. take-dana* ($M_{\rm H-T} = 4.04$, HPDI: 7.54×10^{-3} to 2.10×10^{1}), while the converse was almost zero. Although the modern effective population size of each species was much lower than the ancestral one ($N_{\rm A} = 1.39 \times 10^{6}$; HPDI: 3.12×10^{5} to 1.74×10^{7}), *P. hidakana* had a larger effective population size ($N_{\rm H} = 1.30 \times 10^{4}$; HPDI: 4.39×10^{4} to 6.21×10^{5}) than *P. takedana* ($N_{\rm T} = 4.44 \times 10^{3}$; HPDI: 1.11×10^{3} to 9.09×10^{4}).

Bayesian population demographic analyses

The demographic pattern outlined by VarEff indicates that the effective population sizes of each primrose were equivalent directly after isolation (Ne = 73,088 for *P. hidakana*, Ne = 51,493 for *P. takedana*). However, we also found that the effective population size of *P. takedana* has declined drastically over the past 25,000 years, whereas the curve of *P. hidakana* represents a gradual decline beginning 60,000 years ago. These results suggest that the timing and trend of the population decline was not identical between the two primroses; a severe population reduction in *P. takedana* began at or after the LGM, while *P. hidakana* experienced a gradual population decline during the LG period. **Fig. 6** Potential distribution ranges of *P. takedana* and *P. hidakana* at the present, last glacial maximum (LGM) and last interglacial (LIG) periods inferred by MAXENT. Darker colors represent areas with more suitable predicted conditions. CCSM, Community Climate System Model; MIROC, Model for Interdisciplinary Research on Climate



Ecological niche modeling

The selected models showed high predictive performance in both *P. hidakana* (AUC = 0.91 and TSS = 0.76) and *P. takedana* (AUC = 0.99 and TSS = 0.87), suggesting that the niche models for the two species capture their current actual distributions well (Fig. 6). The selected model's internal jackknife test of variable importance showed that "geology" was the most important predictor for habitat distribution of *P. takedana*, whereas in addition to "geology", "precipitation of wettest quarter" also contributed to the habitat prediction for *P. hidakana* (Supplementary Fig. 8).

According to the models used for both species, their suitable habitats during the LIG period were consistently smaller than in the present climate, whereas the LGM projections differed depending on whether CCSM or MIROC climatic reconstructions were used (Fig. 6). The CCSM models predicted that suitable habitats for both species expanded between the LIG and LGM; the main distribution of P. takedana expanded and shifted towards the south, whereas some local expansions were also identified in the case of P. hidakana. In contrast, the MIROC models inferred that each species experienced contrasting distribution changes during the LGM period; the potential distribution range for P. hidakana largely contracted, whereas the one of P. takedana remained widespread throughout the serpentine belts of the Island. As in the present and LIG period, the computed range overlap between the two species during the LGM was zero in the MIROC model. However, their overlap based on the CCSM reconstructions was 0.21, which indicates that the ranges of the currently allopatric sister species overlapped slightly during the LGM period.

Discussion

Past interspecific gene flow between serpentine endemics and their progenitors throughout the Pleistocene climatic oscillations was not elucidated in previous studies; however, our population genetic analyses and ecological inferences were consistent in supporting the hypothesis that historical gene flow occurred between the currently disjunct progenitor-derivative species pair P. hidakana and P. takedana. ABC analyses supported a secondary contact model and inferred their population dynamics and recent evolutionary history. ENMs strongly supported the estimated speciation mode from our molecular data, and illustrated that their latitudinal range shifts in response to Pleistocene climatic changes resulted in secondary contact. Overall, our results imply that the late Pleistocene climatic oscillations contributed significantly to the alteration of the distribution ranges of serpentine endemics and created opportunities for secondary contact with their progenitors, leading to interspecific gene flow across soil boundaries.

Genetic diversity and structure

P. takedana was characterized by low genetic diversity and population differentiation in this study. Conversely, the nonuniform spatial distribution of genetic diversity and a phylogeographic signal were detected in *P. hidakana*. Genetic diversity was high in the northern and southern

edges of their distribution range (Fig. 4a). These areas may mirror the persistence of *P. hidakana* during the Pleistocene climatic fluctuations and are candidate locations for their refugia. Thereby, the clear genetic structure along north–south direction (Fig. 5c) may be shaped by the refugial isolation during glacial cycles. This scenario can be reinforced by deep haplotype divergence between the northern and southern populations as observed in cpDNA (Fig. 2).

Meanwhile, Bayesian clustering using the nuclear dataset also detected heterogeneous populations within the northern area: although h3 and H8-10 populations were included in the same cpDNA haplotype group together with other northern populations (Fig. 2), these populations were clustered with geographically southern populations in nuclear microsatellites (Fig. 5c). Similar trends were also found in sPCA (Supplementary Fig. 6a). This complex phylogeographic pattern may be explained by the secondary contact between northern and southern lineages after colonization events. Considering the large number of private alleles found in the southern region, high allelic richness in the northern edge area (e.g., h3) might reflect recent gene flow from the southern to northern populations. Altogether, the current genetic structure of and spatial pattern of genetic diversity in P. hidakana was likely shaped by both historical isolation and recent gene flow between the northern and southern populations.

Origin of the serpentine endemic P. takedana

The serpentine-endemic *P. takedana* appears to have evolved recently from *P. hidakana* by vicariant separation. According to our results based on the chloroplast data set, there was extensive sharing of chloroplast haplotypes between species (Fig. 2) and no significant divergence (Table 2). This reflects the fact that *P. takedana* is likely to have originated recently. In fact, the ABC analyses suggested that the divergence occurred around 112 ka (Table 3), corresponding to the LIG period (Ono and Igarashi 1991). Although the time of divergence, as estimated by the ABC analyses, remain controversial due to the wide 95% HPDI ($T_1 = 463-80$ ka; Table 3), the process of divergence had started before the major transition into the Würm glaciation at 75 ka (Ruddiman 1977).

Our ABC analysis supported a scenario with an initial phase of divergence without interspecific gene flow, implying that the initial splitting between species occurred in allopatry. In addition, despite the fact that *P. takedana* is rare and has lower genetic diversity than *P. hidakana*, our demographic reconstructions indicated that the two species have kept similar effective population size in the past (Supplementary Fig. 7). These results suggest that the initial phase of the divergence proceeded via a vicariant separation

event rather than a peripheral isolate. Each primrose prefers cool and moist habitats (as discussed below), and it is likely that climatic amelioration during the mid to late Pleistocene period triggered speciation via vicariance range disjunction of their common ancestor that was once more widespread along the metamorphic belts of Hokkaido Island. The spatial distributions of cpDNA haplotype H1 and the star-like networks (see Fig. 2) plausibly reflect the range expansion of their ancestor, which occurred before the speciation events. Such vicariant separation scenarios have been hypothesized for the evolution of other serpentine plants (Kropf et al. 2008; Michl et al. 2010; Douglas et al. 2011; Moore et al. 2013). Our results, together with those of previous studies, highlight the relative role of Pleistocene climatic oscillations in the evolution of serpentine plants.

Secondary contact during the Pleistocene climatic oscillations

According to our ABC analyses, a model of isolation followed by secondary contact provides the best fit to our data. This result could be corroborated by chloroplast DNA analysis, which indicates that a major haplotype is shared across species (Fig. 2). Perhaps the indistinct species boundary implies the past introgression due to secondary contacts. On the contrary, a phylogeographic signature caused by secondary contact was unclear, particularly in the nuclear dataset. For example, genetic admixture at secondary contact zones often results in a genetic cline across species (Barton and Hewitt 1985). Although a few highly admixed individuals were found in several populations from the results of Bayesian clustering (e.g. h3 and h4), no clear genetic cline across species was detected in sPCA (Fig. 4b). In addition, secondary contacts between distinct lineages often lead to high local genetic diversity (e.g., Comps et al. 2001); however, past secondary contact between two Primula sister species could occur without leaving a trace on the spatial distribution of their genetic diversity, possibly due to a weak unidirectional gene flow. In our study, therefore, it would be premature to infer a speciation mode based on only the results of molecular data, which may lead to misinterpretation.

Nevertheless, our ecological inferences would support the secondary contact scenario. From the potential distributions estimated by the ENMs (Fig. 6), *P. takedana* obviously underwent a range expansion during the LGM, followed by a more recent contraction. Because the predicted distributions of both species were separated by a large region of unsuitable habitat when the climate was warm (i.e., at present and during the LIG period), this prediction makes it reasonable to postulate that the past secondary contact between the two species was caused by the expansion of their distributions and subsequent range overlap during a glacial period. Contrary to these inferences, the ENMs results for *P. hidakana* provide two contrasting range dynamics scenarios during the climatic cycles. The CCSM model indicated a wider LGM distribution of *P. hidakana* in comparison with the present and the LIG distributions, suggesting that the secondary contact was caused by the overlap during cool periods. In fact, our quantitative analysis predicts this range overlap between taxa during the LGM period. Conversely, according to the MIROC model, the distribution of *P. hidakana* was limited to their central geographical distribution area during the LGM period, and then spread from this area during the following postglacial warming period. This implies that the secondary contact occurred before or after the coolest part of the glacial period.

It is hard to distinguish the two contradictory scenarios suggested by the ENMs, because no palynological evidence has been found in either species. Nevertheless, our temporal framework provides further insights into their distributional dynamics. According to the time estimations provided by the ABC analyses, migration started prior to the LGM $(T_2 = 306 - 30 \text{ ka in } 95\% \text{ HPDI}; \text{ Table } 3)$, which contradicts the hypothesis that the secondary contact was caused by climatic amelioration after the coolest period, as suggested by the MIROC model. Moreover, the results of the Bayesian demographic inference also suggest interpretations of their distributional dynamics. If the species' distribution range has changed over time, then evidence for such a change may be predicted based on the corresponding effective population sizes (Excoffier et al. 2009; Slatkin and Excoffier 2012). According to the results of our VarEff analyses (Supplementary Fig. 7), until at least the beginning of the LGM, the effective population sizes of P. takedana and P. hidakana were remarkably larger than at present. This implies that both species occupied wider distributional ranges during the LG period, as suggested by the CCSM model. Regardless of whether the CCSM or MIROC results are more accurate, however, our demographic and ecological inference are consistent with a migration during the cold period in the mid- to late Pleistocene, which provided opportunities for secondary contact between the two now disjunct species. In this case, the mode value of the time estimation suggests that secondary contact between species started during the LG period and the currently disjunctive and scattered distributions were shaped by the climatic amelioration between the LGM and the Holocene.

The serpentine endemic *P. takedana* is a clonal pioneer species (Saito 1970) and probably has high dispersal ability, as discussed below. Therefore, climatically harsh environments with relatively competition-free habitats during the glacial period may have facilitated the colonization of *P. takedana* into neighboring serpentine patches via edaphically unsuitable habitats. Considering the topographic

characteristics of the Island, it is likely that, in addition to the Pleistocene climatic cycles, the latitudinal distribution of the serpentine belts and the absence of mountain chains dividing the lowlands would also have played an important role in the historical range shifts of *P. takedana*. Thus, our study suggests the relative roles of geological and climatic factors in the evolutionary origin and persistence of serpentine plants on Hokkaido Island, which harbors a wide variety of serpentine plants with high endemism (Brooks 1987).

Asymmetric gene flow possibly due to ecological differences

The estimated asymmetrical gene flow from P. hidakana to P. takedana during the past secondary contact phase is another key feature of this system. Such asymmetrical interspecific gene flow across soil boundaries was also reported in other edaphic endemic pairs, and was explained by pollinator behavior, dispersal ability or transfer of adaptive alleles (e.g. Douglas et al. 2011; Kay et al. 2018; Arnold et al. 2016). Although a thorough investigation of the explanations for asymmetrical migration is beyond the scope of this study, we conceived several potential explanations. First, the differences in corolla morphology between the two primroses may have contributed to the observed unidirectional gene flow. P. takedana flowers with funnel-shaped corollas (Fig. 1d) tend to expose their stigmas to various small insects, such as syrphid flies and shorttongued bees (including relatively ineffective pollinators of P. hidakana). On the other hand, P. hidakana flowers with narrow tubes (Fig. 1d) are mostly inaccessible to these short-tongued insects (Yamamoto et al. 2018c), which suggests that P. hidakana cannot receive P. takedana pollen. Hence, our findings suggest the possibility of the occurrence of asymmetrical prezygotic reproductive isolation due to corolla shape differentiation between recently diverged sister taxa. Such a unidirectional hybridization has frequently been reported between the co-occurring Primula species (e.g., Ma et al. 2014; Xie et al. 2017), possibly due to pollinator fauna differentiation.

Second, the differences in dispersal abilities among species may play an important role in promoting unidirectional gene flow. Serpentine plants tend to exhibit high genetic differentiation and low gene flow among populations (e.g. Wolf et al. 2000; Mengoni et al. 2003; Salmerón-Sánchez et al. 2014). In contrast, *P. takedana* is characterized by low population differentiation (Tables 1 and 2) despite their rarity and habitat isolation. As the natural habitats of *P. takedana* are strongly linked with riparian environments (Saito 1970; Ito et al. 2016), occasional seed-mediated gene flow by water (i.e., hydrochory) may be one of the most plausible explanations for the low genetic

differentiation among populations. Interestingly, unlike closely related species, *P. takedana* exhibits seed buoyancy and the seed surface is densely covered with minute membrane organizations that may contribute to increased surface tension of the seeds (Fig. 1e) (Yamamoto, personal observation). It is thought that the relatively high dispersal ability of *P. takedana* may promote the spread of *P. hidakana* alleles into the entire population range of *P. takedana* from the past contact zones. Similar explanations have been discussed for the sympatric serpentine endemic *Monardella* species (Kay et al. 2018); introgression was predicted from *M. stebinsii* to *M. folletti*, which is characterized by lower F_{ST} than *M. stebinsii*.

Finally, this unidirectional migration may be explained by an adaptive process. Increase in metalliferous tolerant genotypes are usually believed to be rare in populations growing on nonmetalliferous soils (Kruckeberg 1951; Wu et al. 1975). Hence, alleles adapted for living in serpentine soils would be strongly selected against in nonserpentine soils. Conversely, migrant alleles from *P. hidakana* may have facilitated the adaptation of *P. takedana* to harsh environments and climatic perturbations, as has been found in *Arabidopsis* species (Arnold et al. 2016). Although we cannot provide further evidence to test these hypotheses, future analyses of large-scale genomic data sets, such as by restriction site-associated DNA sequencing, may shed light on the introgression and selection of alleles of sister species during the past secondary contact phase.

Conclusions

The predicted historical interspecific gene flow across soil boundaries provides a more complex scenario than a simple allopatric speciation model, which has been discussed in several edaphic endemics, and highlights the relative importance of Pleistocene climatic oscillations for the evolution of serpentine plants and their distributional dynamics. Although interspecific gene exchange has the potential to eliminate species boundaries (Taylor et al. 2006), divergence in reproductive traits and habitat preferences between P. hidakana and P. takedana might be key factors for the maintenance and reinforcement of their species integrity, as has been discussed in other plant groups (Matsuda et al. 2017; Kay et al. 2018). However, latent effects against the speciation process and local adaptation of ultramafic Primula arising from interspecific hybridization across soil boundaries remain unclear. Thus, studies of reproductive isolation between species, reciprocal transplantation, tests for local adaptation, and selection against hybrids are needed to better understand the roles of recent hybridization and speciation in the two ultramafic plants.

Data archiving

Sequence data have been submitted to GenBank/DDBJ/ EMBL: accession numbers LC330946–LC330969. Microsatellite genotyping data have been submitted to Dryad: https://doi.org/10.5061/dryad.7r1p25r.

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Compliance with ethical standards

Conflict of interest The authors declare that they have no conflict of interest.

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