



The QT_{c-Bazett} Interval in Former Very Preterm Infants in Adolescence and Young Adulthood is Not Different from Term-Born Controls

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Accepted: 21 June 2023 / Published online: 28 July 2023
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Abstract

Introduction Although relevant for precision pharmacovigilance, there are conflicting data on whether former preterm birth is associated with QT_{c-Bazett} prolongation in later life.

Methods To explore QT_{c-Bazett} interval differences between former preterm and/or extremely low birth weight (ELBW) cases and term-born controls in adolescence and young adulthood, we analyzed pooled individual data after a structured search on published cohorts. To test the absence of a QT_{c-Bazett} difference, a non-inferiority approach was applied (one-sided, upper limit of the 95% confidence interval [CI] mean QT_{c-Bazett} difference, 5 and 10 ms). We also investigated the impact of characteristics, either perinatal or at assessment, on QT_{c-Bazett} in the full dataset (cases and controls). Data were reported as median and range.

Results The pooled dataset contained 164 former preterm and/or ELBW (cases) and 140 controls born full-term from three studies. The median QT_{c-Bazett} intervals were 409 (335–490) and 410 (318–480) ms in cases and controls. The mean QT_{c-Bazett} difference was 1 ms, with an upper 95% CI of 6 ms ($p > 0.05$ and $p < 0.01$ for 5 and 10 ms, respectively). In the full dataset, females had a significantly longer QT_{c-Bazett} than males (415 vs. 401 ms; $p < 0.0001$).

Conclusions QT_{c-Bazett} intervals are not significantly different between former preterm and/or ELBW cases and term-born controls, and we rejected a potential prolongation > 10 ms in cases. When prescribing QTc-prolonging drugs, pharmacovigilance practices in this subpopulation should be similar to the general public (NCT05243537).

1 Introduction

Preterm birth is the primary cause of neonatal mortality [1]. Cardiovascular issues affect preterm infants during neonatal life and afterwards. Elevated blood pressure, changes in heart structure and function, or impaired vascular growth are consequences of preterm birth [2–4]. Former preterm subjects still have a higher overall mortality risk in infancy, childhood, and even early adulthood, including sudden death [5]. Despite these findings, the association between preterm birth and cardiac conduction or repolarization abnormalities in later life has been investigated less often, and associations with established cardiovascular risk factors are poorly explored [4, 5]. Such abnormalities may provide a mechanistic explanation for the higher mortality and may facilitate precision screening and prevention, including pharmacovigilance [4–6].

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Key Points

There are conflicting data on whether former preterm birth is associated with a prolonged QT_{c-Bazett} interval in later life. We therefore analyzed a pooled dataset of published case-control studies.

The pooled dataset contained 164 very preterm or extremely low birth weight (ELBW, < 1 kg) cases and 140 term-born controls from three studies. The mean QT_{c-Bazett} intervals were not significantly different between former preterm subjects and/or ELBW and term-born controls (409 vs. 410 ms). A potential QT_{c-Bazett} prolongation > 10 ms in cases was rejected.

Consequently, when prescribing QTc-prolonging drugs, pharmacovigilance practices in this subpopulation should be similar to those applied in the general public.

Bassareo et al. (Italy) reported that a heart rate-corrected QT time according to the Bazett formula ($QT_{c-Bazett}$) in former extremely low birth weight (ELBW) infants in young adulthood (mean age 23 years) was at the upper limit of the normal range and correlated with gestational age. Relevant for precision pharmacovigilance, Bassareo et al. observed a significant difference in mean $QT_{c-Bazett}$ time (417 vs. 369.9 ms) between former ELBW cases and controls [7]. On the contrary, neither Gervais et al. (Canada) nor Salaets et al. (Belgium) confirmed these differences in former preterm subjects in young adulthood and late childhood to adolescence, respectively [8, 9]. More clarity on the presence or absence of a difference in $QT_{c-Bazett}$ is relevant for precision pharmacovigilance, as certain QTc-prolonging drugs, such as antipsychotics or attention deficit and hyperactivity disorder (ADHD) drugs, are more commonly prescribed in former premature infants [10, 11].

Therefore, the primary aim of this study was to explore potential differences in $QT_{c-Bazett}$ intervals between former preterm and/or ELBW cases and term-born healthy controls by pooling individual data as published. On a second level, we explored the impact of covariates on the $QT_{c-Bazett}$ intervals in cases and controls.

2 Methods

2.1 Ethics, Study Registration, and Data Handling

The Ethics Committee Research of University Hospitals Leuven approved the study protocol (7 January 2022; S66020) and the study was registered at ClinicalTrials.gov (NCT05243537). When not available, individual data extracted from figures or tables in the source document using a web-based, valid extraction program (WebPlotDigitizer) was our second option, be that this commonly results in a more restricted dataset [12].

2.2 Search Strategy

Two authors (JV, MVP) conducted a search for relevant articles using PubMed Advanced to retrieve case-control cohorts in November 2021. Following internal discussion, it was agreed to use ‘(long) QT’ and ‘preterm’ as search terms. Based on the aims of this study, papers had to report on QTc data in such cohorts after neonatal stay, and had to report on more than one specific covariate. Retained papers were checked for potential additional relevant references or citations (electronic supplementary material [ESM] Table S1).

2.3 Quality Assessment

Quality assessment was performed in the retained articles, applying the Scottish Intercollegiate Guidelines Network (SIGN) methodology checklist for case-control studies [13]. The SIGN questionnaire contains 15 questions rating aspects on internal validity (selection of subjects, assessment, confounding, statistical analysis) and overall assessment. When the majority of criteria are met and results are unlikely to be changed by further research, this is classified as *high quality* (++); when most criteria are met, with some flaws in the study with an associated risk of bias, and conclusions may change in the light of further studies, this is classified as *acceptable* (+); and when most criteria were not met or significant flaws related to key aspects of the study design, so that conclusions are likely to change in the light of further studies, this is classified as *low quality* [13]. This effort was made to provide transparency on the quality (risk of bias) of the studies retained. Two authors (JV, MVP) individually completed the SIGN questionnaire. In the event of discrepancy for a given question, a third author (KA) also assessed the paper to find consensus.

2.4 Primary Outcome

$QT_{c-Bazett}$ obtained at rest was chosen as the primary outcome measure. The Bazett formula ($QT_{c-Bazett} = QT/\sqrt{RR}$) was used for heart rate correction.

2.5 Statistical Analysis

Descriptive statistics are presented as median (range) for continuous variables, or frequency (%) for categorical variables. For the comparison of $QT_{c-Bazett}$ intervals between preterm and/or ELBW cases and term-born controls, a non-inferiority approach was followed. This approach aims to test the null hypothesis that the mean $QT_{c-Bazett}$ is prolonged in cases, compared with controls, by a clinically relevant margin. For the primary analysis, a non-inferiority margin of 5 ms was considered, based on the US Food and Drug Administration (FDA) and European Medicines Agency (EMA) paired study guidelines as the most stringent criteria for positive control effects [14, 15]. In a post hoc analysis (after publication of the protocol on the ClinicalTrials.gov website), we also considered 10 ms as another non-inferiority margin because a 10 ms margin better reflects the clinically relevant margin, as the International Conference on Harmonisation (ICH) E14 guidelines

define a negative ‘thorough QT/QTc study’ as the one in which the upper bound of the 95% one-sided confidence interval (CI) for the largest time-matched mean effect of the drug on the QTc interval excludes 10 ms [15].

The analysis was performed by estimating the upper limit of a one-sided 95% CI around the difference in means (cases to controls) using a pooled variance *t*-test, and non-inferiority was demonstrated if the upper limit falls below the non-inferiority margin.

Comparative statistical analyses were used (Spearman’s rank, Mann–Whitney U, or *t*-test) to explore associations of QT_{c-Bazett} measurements to characteristics, either perinatal or at assessment.

Analyses were performed using SAS[®] software (version 9.4, SAS System for Windows; SAS Institute, Inc., Cary, NC, USA) or MedCalc[®] (version 20.111; MedCalc Software Ltd, Ostend, Belgium). A *p*-value < 0.05 was considered to be statistically significant.

3 Results

3.1 Search Results

Based on the search strategy described, a limited number of papers were retrieved (25 for ‘long QT and preterm’), of which we retained three articles [7–9]. Twenty-two papers were not retained because they were reviews (*n* = 2), case reports or clinical observations on congenital long QT (either maternal, fetal, or neonatal; *n* = 13), drug-related observations during neonatal care (cisapride, 4; erythromycin, 1; serotonin reuptake inhibitors, 1), or related to car seating in former preterm subjects (*n* = 1) (ESM Table S1). Further efforts (‘QT and preterm’), references, or citation tracking (snowball and citation search) did not provide additional documents.

3.2 Quality Assessment

Using the SIGN approach, two papers (Salaets et al. [9] and Gervais et al. [8]) were classified as ‘high quality’. The study by Bassareo et al. did not meet these criteria due to uncertainties on the representation of the population (marked difference in sex distribution, with 20/24 females in both cases and controls) and on the blinding procedures, and was hence classified as having ‘low quality’ [7]. The final consensus assessment is provided in ESM Table S2.

3.3 Data Acquisition, Cohort and Study Characteristics

Reaching out to the corresponding authors resulted in data sharing and access to two of three cohorts (Salaets et al., 93 cases, 87 controls; Gervais et al., 47 cases, 53 controls) for the individual QT_{c-Bazett} and related characteristics [7–9]. Despite repeated attempts, we failed to receive any response from the Bassareo et al. study team [7]; however, based on the figures provided in the paper by Bassareo et al., we were able to extract raw individual data on QT_{c-Bazett} intervals and gestational age in the 24 cases. Based on the information in the public domain, we were unable to link this to either their birth weight or sex, and neither could we extract raw individual data in the 24 controls [12].

Gervais et al. characterized cases as born preterm (< 30 weeks gestational age). Controls were term-born with a birth weight >2500 g, and were matched for age (at assessment) and sex. Controls were either friends or siblings of the preterm cases or found through advertising. The mean age at assessment was 23.9 years [8]. Similarly, Salaets et al. defined cases as preterm-born children (gestational age of 23–33 weeks) born with a birth weight of ≤ 1000 g (ELBW). Controls were either term-born friends of the cases or were recruited from an elementary school nearby the research center (Hamont Achel, Belgium). The median age at assessment was 11 (8–14) years [9]. Bassareo et al. recruited former ELBW cases and controls as healthy, term-born subjects, matched for sex, age and body mass index (BMI). Furthermore, all subjects were contacted in alphabetical order from the records of the Neonatal Intensive Care Unit of the University of Cagliari. The mean age at assessment was 23.2 years [7].

In the Gervais cohort, continuous 12-lead ECG (GE Case Stress System V6.5 and 6.73, GE Medical Systems Information Technologies GmbH, Freiburg, Germany) was recorded at a speed of 25 mm/s. The ECG of participants at rest was manually analyzed. Tracings were scanned and measurements were performed using the magnifier and ruler function of Adobe Photoshop (version 19.1.6; Adobe Systems, San Jose, CA, USA) by two trained operators blinded to the exposure status, under the supervision of a staff cardiologist. Intraobserver correlation, assessed on a subset of 24 QTc measurements performed twice on separate days, was very good, with an intraclass correlation coefficient of 0.94. Three measurements were taken in lead DII or V5 on three QRS complexes, and averaged [8].

In the Salaets cohort, a 12-lead ECG was collected at rest after 30 min in the supine position by a trained assistant. As a quality control measure, research assistants received periodical training on skin preparation, electrode placement, and positioning of the participants. The Cardiax[®] device (RDSM Medical Devices, Hasselt, Belgium) was used for ECG acquisition and automatic determination of standard ECG parameters on the digital traces, thereby excluding observer-related bias. PR and QT intervals, QRS duration, and QRS were automatically measured to the nearest 1 ms (they were not calculated manually). Based on visual inspection of printouts, the quality of the ECGs used for analysis was assessed by one assessor, blinded for group allocation [9].

Although not explicitly mentioned, Bassareo et al. did not specify in which state the ECGs were taken, but we indirectly understood—based on an additional letter—that the Bassareo et al. data were also collected at rest [7, 16]. For additional details on the individual cohorts, including the equipment used and automatic versus manual measurements, we refer to the initial papers and related letters as published source documents [7–9, 16, 17].

After pooling, the final cohort was based on 164 ELBW and/or preterm cases and 140 healthy, term-born controls. Table 1 describes the characteristics (either perinatal or at assessment [ECG related or biochemical findings]) of the pooled study population. The heart rate at rest in the pooled dataset was 78 (48–161) and 72 (49–129) in the cases and controls, respectively ($p > 0.05$). Differences in clinical characteristics (shorter, lower weight, BMI) and blood pressure between cases and controls confirmed the previously reported differences in clinical characteristics [8, 9].

While there was no information on QT-prolonging medications in the Bassareo et al. cohort [7], asthma medicines (9 vs. 2 cases vs. controls), or psychiatry/ADHD (8 vs. 3) or any QT_c-prolonging medicines (11 vs. 2) were more commonly used in former preterm-born adults in the Gervais cohort [8]. In the Salaets cohort, eight medicines in four ELBW cases and four controls were reported to be associated with QT_c prolongation (inhaled formoterol [$n = 2$] or oral methylphenidate [$n = 6$]) [9]. For both cohorts, the impact of these drugs on the QT_c interval was not significant, as reported in the individual papers [8, 9].

3.4 Comparison of the QT_{c-Bazett} Interval Between Cases and Controls

In the analysis of the pooled individual QT_{c-Bazett} observations, there was no statistical difference between cases and controls upon superiority testing (cases to controls: 409 vs. 410 ms; $p > 0.05$) (Table 1, Fig. 1). For QT_{c-Bazett}, we observed a mean difference between cases and controls of

1 ms (95% upper CI limit 6 ms). This means that the upper limit does not fall below the initially set non-inferiority margin of 5 ms ($p > 0.05$). However, the upper limit does fall below the clinically relevant non-inferiority margin of 10 ms ($p < 0.01$, post hoc analysis).

3.5 Covariate Analysis of QT_{c-Bazett} in the Pooled Dataset of Cases and Controls

In the covariate analysis, we did not observe any significant association between QT_{c-Bazett} and birth weight or gestational age (perinatal characteristics). The same holds true for age, height, weight, BMI, systolic and diastolic blood pressure (clinical characteristics at assessment), or sodium, potassium or calcium (biochemical findings). The mean QT_{c-Bazett} was significantly higher in female subjects compared with male subjects (415 vs. 401 ms; $p < 0.0001$). There was also a significant correlation (Spearman's rank 0.151; $p < 0.05$) between QT_{c-Bazett} and phosphate level.

4 Discussion

In an effort to pool individual data of published cohorts, we observed no significant differences in the QT_{c-Bazett} time interval between former ELBW and/or preterm-born cases and term-born controls in late childhood, adolescence, and young adulthood. Applying a one-sided non-inferiority approach, a difference of > 10 ms (but not > 5 ms) in former ELBW and/or preterm subjects was hereby excluded. The absence of a significant correlation between QT_{c-Bazett} and birth weight or gestational age in the covariate analysis provides further support for the absence of a relevant QT_{c-Bazett} prolongation in former ELBW and/or preterm subjects. At assessment, sex (female to male, 415 to 401 ms) was a strong covariate, while phosphate levels were a weak (Spearman 0.151, $p < 0.05$) covariate of the QT_{c-Bazett} time interval. Both can be expected, as it is generally known that both sex and electrolytes, including calcium/phosphate balance, have been reported as significant covariates of the QT_{c-Bazett} time interval, as also reflected in the FDA guidance and EMA guidelines [14, 15, 18]. Although evidence is inconsistent, a higher BMI in adults was associated with a longer QT_{c-Bazett} in some studies [19, 20]. In the current pooled dataset, this was not the case.

Our main motivation was driven by targeted or precision pharmacovigilance, as certain QT-prolonging drugs, such as antipsychotics or drugs used to treat ADHD, are more commonly prescribed in former preterm subjects [10, 11]. As these drugs are associated with prolongation of the QT_{c-Bazett} interval, the absence of any *a priori* prolongation in former ELBW and/or preterm subjects matters. In essence, these

Table 1 Characteristics (perinatal, at assessment, ECG-related, and biochemical findings) of the pooled study population in either cases or controls [7–9]

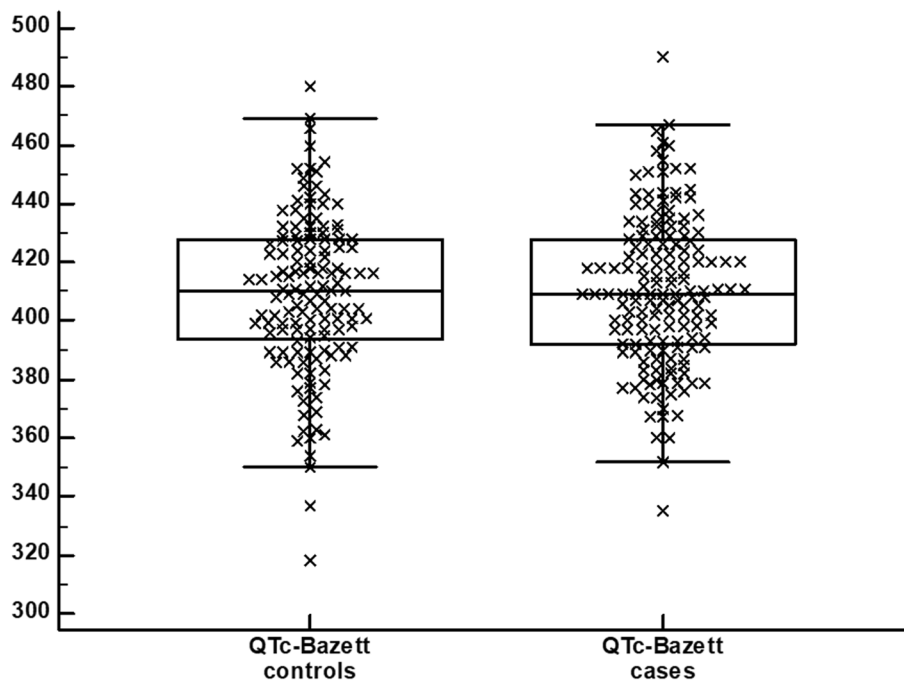
	Preterms	No.	Term controls	No.	<i>p</i> -value ^a
<i>Perinatal</i>					
Birth weight, g	840 (430–1460)	164	3400 (2300–5000)	133	< 0.001
Gestational age, weeks	27 (24–33)	163	39 (37–42)	132	< 0.001
<i>At assessment</i>					
Age, years	12 (9–33)	140	12 (9–30)	140	> 0.05
Height, cm	151 (124–189)	140	157 (129–194)	140	< 0.01
Weight, kg	41.9 (21.4–90)	140	47.1 (25.4–149)	140	< 0.001
BMI, kg/m ²	18.1 (12.8–32.6)	140	19.5 (13.7–48.6)	140	< 0.01
Males/females	65/75	140	63/77	140	< 0.05
BPS, mmHg	114 (92–153)	140	108 (88–163)	140	< 0.01
BDP, mmHg	68 (47–103)	140	65 (50–105)	139	< 0.001
<i>ECG related</i>					
Heart rate, /min	78 (48–161)	135	72 (49–129)	138	> 0.05
RR, ms	761 (372–1255)	135	830.5 (464–1236)	138	> 0.05
PR, ms	127 (80–378)	137	136 (94–232)	137	< 0.01
QT, ms	362 (237–438)	134	368 (272–442)	137	> 0.05
QT _{c Bazett} , ms	409 (335–490)	158	410 (318–480)	137	> 0.05
QRS time, ms	84 (65–104)	138	84 (59–112)	137	> 0.05
<i>Biochemical findings</i>					
Sodium, mmol/L	140 (134–146)	102	140 (134–146)	122	> 0.05
Potassium, mmol/L	4.2 (3.6–5.9)	101	4.2 (3.6–5.36)	122	> 0.05
Calcium, mmol/L	2.455 (2.2–2.82)	90	2.41 (2.14–2.67)	105	> 0.05
Phosphate, mmol/L	1.37 (0.84–1.96)	89	1.42 (0.74–1.84)	105	> 0.05

Data are expressed as median (range) or frequencies

BMI body mass index, BPS systolic blood pressure, BPD diastolic blood pressure, ECG electrocardiogram, RR interval between two R waves, PR interval between P-wave and R-wave, QT interval between Q-wave and T-wave, QT_{c Bazett} corrected (for heart rate) interval between Q-wave and T-wave

^aBolded values indicate statistical significance

Fig. 1 Individual QT_{c-Bazett} time (dot plot, ms) intervals as observed in cases or term-born controls



findings support the statement that pharmacovigilance procedures in this subpopulation should be similar to the general public, and that additional precautions or precision pharmacovigilance on this topic are obsolete. Obviously, there are limitations to our study and reporting.

First, we only had partial access (cases only, and individual $QT_{c-Bazett}$ values only linked gestational age, not birth weight) to the first case-control study (24 cases and 24 controls) [7], and assessed the quality of this case-control study to be poorer compared with the two other cohorts [8, 9]. Despite these limitations, we decided to retain these data in cases in the pooled analysis as the Bassareo et al. cohort was the only study that reported on a prolonged $QT_{c-Bazett}$ time interval in former ELBW cases [7]. Our approach also limited the use of other QT_c formulae, despite the FDA recommendations to explore different formulae, because the best correction factor remains controversial, and the problems with Bazett QT_c correction in pediatric screening for prolonged QT_c [14, 21].

Second, in our initial study protocol (ClinicalTrials.gov NCT05243537), we defined non-inferiority using an upper limit of 5 ms, based on our initial understanding of the FDA guidance and the EMA paired guidelines on the paired study design [14, 15]. However, this is a very stringent criterion that we did not meet. A post hoc power analysis based on our pooled data distribution determined that we would need a very large sample size of 1160 subjects to document the absence of a difference 5 ms in a non-inferiority study (with a power of 80% and a significance level of 5%). Regulatory guidance indeed considers a 5 ms increase in QT/QT_c to be of clinical relevance in drug trials and sets this as the threshold for a ‘thorough QT/QT_c study’, but does not recommend a non-inferiority approach using 5 ms as the maximal upper limit of the 95% CI. In order to exclude a relevant increase in QT/QT_c as reliable and feasible, regulatory authorities define a negative ‘thorough QT/QT_c study’ as one in which the upper bound of the 95% one-sided CI for the largest time-matched mean effect of the drug on the QT_c interval excludes 10 ms. This 10 ms criterion is very reasonable as a clinical target, as also recently suggested [3]. Thus, our post hoc analysis, using a cut-off of 10 ms—to show non-inferiority of the former preterm-born subjects to controls with regard to $QT_{c-Bazett}$ —is in line with regulatory guidance and excludes this 10 ms cut-off.

Finally, most participants included in our study design had a Caucasian background and all were recruited in Italy, Canada, and Belgium. Although research on inter-ethnic differences in baseline QT_c intervals is limited, the available data suggest that there is no relevant difference in QT_c time intervals related to ethnicity, while the FDA guidance and EMA guideline also state that it is not expected that clinical evaluation of QT/QT_c interval prolongation and

proarrhythmic potential for non-antiarrhythmic drugs would be affected by ethnic factors [14, 15].

5 Conclusions

There was no significant difference in the $QT_{c-Bazett}$ interval between preterm and/or ELBW cases and term-born controls at late childhood/adolescence and young adulthood, and a potential prolongation in $QT_{c-Bazett}$ interval of > 10 ms was rejected. When prescribing QT_c -prolonging drugs, pharmacovigilance practices in this subpopulation should be similar to the general public.

Supplementary Information The online version contains supplementary material available at <https://doi.org/10.1007/s40264-023-01335-y>.

Declarations

Funding TS is supported by the Frans Van de Werf Fund for Clinical Cardiovascular Research for his research fellowship (2022 recipient), and V.Z.W. ‘De Kleine Hartjes’ and F.W.O. Vlaanderen. The PREMATCH study was supported by the Agency for Innovation by Science and Technology in Flanders (IWT) through the SAFE-PEDRUG project [IWT/SBO 130033]. KU Leuven Internal Funds (STG-18-00379) supported the Research Unit Hypertension and Cardiovascular Epidemiology, Department of Cardiovascular Sciences, Leuven. The research activities of AS are supported by the Clinical Research and Education Council of the University Hospitals, Leuven.

Conflict of interest Jill Vanthienen, Marine Vassilev Petrov, Thuy Mai Luu, Anik Cloutier, Anke Raaijmakers, Jan A. Staessen, Zhenyu Zhang, Thomas Salaets, Annouschka Laenen, Anne Smits, Anne-Monique Nuyt, Adrien Flahault, and Karel Allegaert have no conflicts of interest to declare that are directly relevant to the contents of this study.

Ethics approval The Ethics Committee Research (EC Research) of University Hospitals Leuven (UZ Leuven) approved the study protocol for the pooled analysis (7 January 2022; S66020). The Ethics Committee of CHU Sainte-Justine and the Université de Montréal also approved the study protocol for the pooled analysis (26 January 2022; reference number 2022-3773).

Consent to participate All data pooled in this study have been previously published. These initial studies and sources were conducted following Ethics Committee approval at the individual institutes and following informed consent and, if applicable, assent.

Consent to publish Not applicable.

Availability of data materials The corresponding author can be contacted to share the raw data based on reasonable request and a study protocol. The individual study groups remain the sole owner and controller of their datasets.

Code availability Not applicable.

Authors' contributions Conceptualization: AMN, AF, TML, AR, TS, KA. Funding acquisition: KA, JAS, AR. Data curation: KA, TML. Data acquisition and verification: AMN, AF, TML, AR, JAS, KA, AC, TS. Structured search and quality assessment: JV, MVP, KA. Statistics: AL, KA. Project administration: KA, TML. Writing original draft:

JV, MVP, KA. Writing – reviewing and editing: AMN, AF, TML, AR, JAS, ZZ, TS, AL, KA, JV, MVP, AC, AS. All authors have read and approved the final version of the paper.

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