



# Ischemic Stroke and Cerebral Microbleeds: A Two-Sample Bidirectional Mendelian Randomization Study

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## ABSTRACT

**Introduction:** Recent observational studies have reported the association between ischemic stroke (IS) and cerebral microbleeds (CMBs). Whether this reflects a causal association remains to be established. Herein, we adopted a two-sample bidirectional Mendelian randomization (MR) analysis to comprehensively evaluate the causal association of IS and CMBs.

**Methods:** The summary-level genome-wide association studies (GWASs) data of IS were obtained from the GIGASTROKE consortium (62,100 European ancestry cases and 1,234,808 European ancestry controls). All IS cases could be further divided into large-vessel atherosclerosis stroke (LVS,  $n = 6399$ ), cardio-embolic stroke (CES,  $n = 10,804$ ) and small-vessel occlusion stroke (SVS,  $n = 6811$ ). Meanwhile,

we used publicly available summary statistics from published GWASs of CMBs (3556 of the 25,862 European participants across 2 large initiatives). A bidirectional MR analysis was conducted using inverse-variance weighting (IVW) as the major outcome, whereas MR-Egger and weighted median (WM) were used to complement the IVW estimates as they can provide more robust estimates in a broader set of scenarios but are less efficient (wider CIs). A Bonferroni-corrected threshold of  $p < 0.0125$  was considered significant, and  $p$  values between 0.0125 and 0.05 were considered suggestive of evidence for a potential association.

**Results:** We detected that higher risk of IS [IVW odds ratio (OR) 1.47, 95% confidence interval (CI) 1.04–2.07,  $p = 0.03$ ] and SVS (IVW OR 1.62, 95% CI 1.07–2.47,  $p = 0.02$ ) were significantly associated with CMBs. Reverse MR analyses found no significant evidence for a causal effect of CMBs on IS and its subtypes.

**Conclusions:** Our study provides potential evidence that IS and SVS are causally linked to increased risk of CMBs. Further research is needed to determine the mechanisms of association between IS and CMBs.

**Keywords:** Ischemic stroke; Cerebral microbleeds; Mendelian randomization study

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### Key Summary Points

#### *Why carry out this study?*

Recent observational studies have reported the association between ischemic stroke (IS) and cerebral microbleeds (CMBs), but the causal relationship between IS and CMBs has not been established

In this study, we examined the bidirectional causal relationship between IS and CMBs using a genetically informed method

#### *What was learned from the study?*

Our study provides potential evidence that IS and SVS are causally linked to the increased risk of CMBs

Reverse MR analyses found no significant evidence for a causal effect of CMBs on IS

## INTRODUCTION

Ischemic stroke (IS) occurs primarily when cerebral blood flow is interrupted, causing severe neural damage [1–3]. It is ranked second among the leading causes of death worldwide, causing an estimated 5.9 million deaths and 102 million disability-adjusted life years lost [2]. With the ever-changing development of medical imaging technology, especially the magnetic resonance (MR) imaging system, an increasing number of patients with IS are found to be accompanied by different degrees of cerebral microbleeds (CMBs) [4].

CMBs are typically detected on susceptibility-weighted imaging (SWI) as small (< 10 mm), hypointense (black), ovoid or rounded regions which develop because of red blood cell leakage from arteries and capillaries [5, 6]. The potential increased risk of intracranial hemorrhage associated with CMBs is well known [7]. Evidence, however, suggests CMBs may not only provide information about bleeding propensity but also

about future ischemic events. According to Cordonnier et al., approximately one-third of IS patients have one or more CMBs [8]. Additionally, several observational studies have demonstrated that CMBs increase the chances of subsequent intracranial hemorrhage and recurrent IS in those who have suffered from IS [9–12]. Despite this, observational studies have some limitations, such as reverse causation and unmeasured confounding factors that are difficult to eliminate, which may limit the ability of causal inference. In consequence, the causal association between CMBs with IS has not been examined systematically.

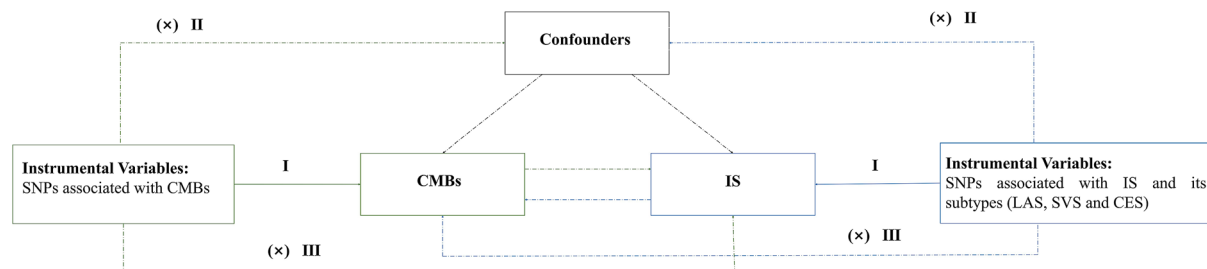
Mendelian randomization (MR), using genetic variants as instrumental variables (IVs) to assess the causal effects of risk factors related to diseases, can overcome confounding biases inherent in observational studies. Previous studies have found the causality relationship between IS and other diseases by using MR analysis [13–15]; however, the causal association between IS and CMBs has not been demonstrated yet.

Herein, we adopted a two-sample bidirectional MR analysis to comprehensively evaluate the causal association of CMBs and IS.

## METHODS

### Study Design

In this study, a two-sample bidirectional MR study was conducted to explore the causal relationship between IS and CMBs. Figure 1 shows the flowchart of the overall study design. MR must be premised on the following three basic criteria: (1) relevance: the genetic variant, usually single nucleotide polymorphism (SNP), should be closely associated with the exposure; (2) independence: genetic variants should not be associated with any potential confounders; (3) exclusion-restriction: genetic variants should not be associated with the outcome except via the way of exposure [16].



**Fig. 1** Conceptual framework for the Mendelian randomization analysis of the causal association of cerebral micobleeds and ischemic stroke. The design follows the following three basic criteria: (1) relevance: genetic variants, usually single nucleotide polymorphisms, should

be close to the exposure; (2) independence: genetic variants should not be associated with any potential confounders; (3) exclusion-restriction: genetic variants should not be associated with the outcome except via the way of exposure

## Data Source Description

The summary-level genome-wide association study (GWAS) data of IS were obtained from the GIGASTROKE consortium (62,100 European ancestry cases and 1,234,808 European ancestry controls) [17]. Based on the Trial of ORG 10172 in Acute Stroke Treatment (TOAST) classification, all IS cases could be further classified as large-vessel atherosclerosis stroke (LVS,  $n = 6399$ ), small-vessel occlusion stroke (SVS,  $n = 6811$ ) and cardio-embolic stroke (CES,  $n = 10,804$ ) [18, 19]. Meanwhile, we used publicly available summary statistics from published GWASs of CMBs (3556 out of 25,862 participants had CMBs), which performed genome-wide association studies in 11 population-based cohort studies and 3 case-control or case-only stroke cohorts [20] (<https://cd.hugeamp.org/datasets.html>). This study used publicly available de-identified data from participant studies that were approved by an ethical standards committee with respect to human experimentation. No separate ethical approval was required in this study.

## Selection of Genetic Instruments

All selected single-nucleotide polymorphisms (SNPs) should be clearly associated with exposures [ $p < 5 \times 10^{-7}$ , linkage disequilibrium (LD)  $r^2 < 0.01$ , and  $F$ -statistics  $> 10$ ] [21]. The SNPs ( $P < 5 \times 10^{-6}$ ) in the PhenoScanner database, a comprehensive information platform on the

relationship between genotypes and phenotypes, were eliminated with potential confounders and outcomes, such as, drinking, diabetes, smoking, hypertension, treatment with aspirin and treatment with warfarin.

## Mendelian Randomization Analyses

A bidirectional MR analysis was conducted using inverse-variance weighting (IVW) as the preferred method, whereas MR-Egger and weighted median (WM) were used to complement the IVW estimates as they could provide more robust estimates in a broader set of scenarios but are less efficient (wider CIs). Furthermore, we adopted MR-Egger intercept and MR-PRESSO as secondary analyses to detect heterogeneity and horizontal pleiotropy [22]. The MR-Egger method relaxes the IVW assumption that the average pleiotropic effect is zero, allowing all genetic variants to have a pleiotropic effect (not via the exposure) [23]. In addition, we performed a leave-one-out analysis to further investigate the impact of outlying and/or pleiotropic genetic variants. If estimates of these approaches in our study were inconsistent, a tighten instrument  $p$  value threshold was set, and then the MR analysis was performed again [23]. Furthermore, we calculated the  $r^2$  ( $r^2 = \text{beta}^2 / (\text{se}^2 \times (n - 2) + \text{beta}^2)$ ) of each SNP and summed them up to calculate the overall  $r^2$  [ $r^2 \times (N - 2) / (1 - r^2)$ ] and  $F$  statistics using the sample size of exposure GWASs [24].

## Statistical analyses

Statistical analyses and data visualization were performed using the R software, version 4.2.0 (<http://www.r-project.org>). All two-sample MR analyses were performed using the MR-Base R package (“TwoSampleMR”) “TwoSampleMR” (<https://github.com/MRCIEU/TwoSampleMR>) [25]. The mRnd was used to calculate the statistical power for Mendelian randomization (<https://cnsgenomics.shinyapps.io/mRnd/>) [26].

## Strength of Evidence

Bonferroni correction was applied to correct the threshold of statistical significance for multiple comparisons, and a  $p$ -value below 0.0125 (where  $p = 0.05/4$ ) was considered strong evidence of significant association;  $p$  values between 0.0125 and 0.05 were considered suggestive of evidence for a potential association [27].

## RESULTS

### The Causal Effects of IS and its subtypes on CMBs

#### *Causal Effect of IS on CMBs*

An overall analysis found 25 SNPs associated with IS both significantly and independently (Table 1). The IVW method showed that IS was causally associated with an increase in risk of CMBs significantly (OR 1.47, 95% CI: 1.04–2.07,  $p = 0.029$ ) (Fig. 2).

No heterogeneity was observed with a Cochran Q derived in MR study between IS with CMBs ( $p = 0.657$ ). Furthermore, Egger intercepts did not detect any pleiotropy (intercept =  $-0.02$ ; SE = 0.04.  $p = 0.66$ ) (Supplementary Fig. 1A). No single SNP was strongly violating the overall effect of IS on CMBs in the leave-one-out sensitivity analysis (Supplementary Fig. 2A). The funnel plot was symmetrical, indicating no heterogeneity (Supplementary Fig. 1B). Similarly, the MR-PRESSO global test showed the absence of pleiotropy

( $p = 0.10$ ). The power and F-statistics of MR analysis are displayed in Table S1.

#### *Causal Effect of LVS on CMBs*

Using the seven LVS-related SNPs (Table 1), no evidence suggested a potential causal effect of LVS on the risk of CMBs (IVW OR 1.15, 95% CI 0.83–1.57,  $p = 0.40$ ). However, opposing results were observed using the MR-Egger approach (OR 0.50, 95% CI 0.19–1.31,  $p = 0.22$ ). Since the MR estimates of MR-Egger and IVW were inconsistent, we tightened the instrument  $p$  value threshold to  $1.2 \times 10^{-7}$ , and two SNPs (rs142395500 and rs476762) were used as instrument tools. The MR estimates still indicated that LVS had no causality on CMBs (OR 1.04, 95% CI: 0.27–3.93,  $p = 0.96$ ) (Fig. 2). However, heterogeneity was observed with a Cochran Q-derived  $p$  value  $< 0.05$  in MR study between LVS with CMBs ( $p = 0.002$ ). As we used the random effects IVW as main result, heterogeneity was acceptable [22]. MR-Egger and MR-PRESSO global tests could not be performed because only two SNPs were included in the analysis.

#### *Causal Effect of SVS on CMBs*

Five SNPs were taken as IVs for SVS (Table 1). We found that SVS increased the risk for CMBs significantly (IVW OR 1.39, 95% CI: 1.05–1.85,  $p = 0.022$ ), while opposing results were observed using the MR-Egger approach (OR 0.86, 95% CI: 0.35–2.14,  $p = 0.77$ ). Then, we tightened the instrument  $p$  value threshold to  $3 \times 10^{-7}$  and two SNPs were identified, which were significantly and independently associated with SVS (rs12445022 and rs7766042). The result of IVW indicated that SVS had causality on CMBs (IVW OR 1.62, 95% CI: 1.07–2.47,  $p = 0.02$ ) (Fig. 2). No obvious heterogeneity was observed (the Cochran Q-test derived  $p$  value was 0.35). However, MR-Egger and MR-PRESSO global tests could not be performed.

#### *Causal Effect of CES on CMBs*

Using the six CES-related SNPs (Table 1), we found that CES had no obvious causal effect on CMBs (OR 0.98, 95% CI: 0.64–1.50,  $p = 0.93$ ) (Fig. 2), while opposing results were observed

**Table 1** Characteristics of included SNPs associated with exposures

Exposures	rsID	Chr	Gene	EA	$\beta$	SE	<i>p</i> value	<i>F</i>	
IS	rs12445022	16	RP11-482M8.1	A	0.0500	0.0077	8.39E−11	42.16	
	rs10886430	10	GRK5	A	− 0.0778	0.0114	8.32E−12	46.57	
	rs12879705	14	KTN1-AS1	C	0.0661	0.0129	2.99E−07	26.25	
	rs13148045	4	LINC02172	T	− 0.0502	0.0096	1.70E−07	27.34	
	rs1487504	9	RP11-132E11.2	A	0.0622	0.0114	4.87E−08	29.76	
	rs150986675	11	RNU7-159P	T	− 0.0618	0.0098	2.86E−10	39.76	
	rs17148926	5	CTC-441N14.4	A	0.0606	0.0096	2.75E−10	39.84	
	rs2500281	1	PRDM16	A	0.0602	0.0112	7.66E−08	28.89	
	rs2501966	6	CENPQ	A	0.0380	0.0071	8.69E−08	28.64	
	rs2526381	17	TSPOAP1	T	− 0.0543	0.0095	1.09E−08	32.67	
	rs2736613	1	PMF1-BGLAP	T	0.0487	0.0073	2.54E−11	44.50	
	rs34003787	16	ZFHX3	T	− 0.0717	0.0141	3.67E−07	25.85	
	rs36229526	6	TAP1	T	0.0741	0.0136	5.08E−08	29.68	
	rs3756011	4	F11	A	0.0439	0.0071	6.29E−10	38.23	
	rs3787382	20	BMP7	T	0.0498	0.0099	4.90E−07	25.30	
	rs4759076	12	COPZ1	T	0.0371	0.0071	1.74E−07	27.30	
	rs56010410	4	FGA	T	− 0.0651	0.0078	7.05E−17	69.65	
	rs6496123	15	LINC00924	A	− 0.0483	0.0085	1.33E−08	32.28	
	rs7174762	15	PDE8A	A	0.0386	0.0073	1.24E−07	27.95	
	rs72631113	10	ARMS2	T	− 0.0452	0.0072	3.43E−10	39.41	
	rs7304841	12	PDE3A	A	0.0401	0.0073	3.95E−08	30.17	
	rs77851364	17	NLRP1	T	− 0.0802	0.0155	2.29E−07	26.77	
	rs7820415	8	ZYXP1	T	− 0.0644	0.0126	3.20E−07	26.12	
	rs79318212	6	RP11-157J24.2	A	− 0.0864	0.0115	5.78E−14	56.44	
	rs842365	13	LRCH1	A	0.0446	0.0081	3.67E−08	30.31	
	LVS	rs476762	11	MMP3	A	0.2010	0.0353	1.24E−08	32.42
		rs10760966	9	LINC01492	A	− 0.1462	0.0278	1.45E−07	27.65
		rs11670056*	19	ELL	T	0.2246	0.0435	2.43E−07	26.65
		rs142395500*	18	TRAPPC8	A	0.3966	0.0749	1.19E−07	28.03
		rs180789*	3	OXTR	A	− 0.1311	0.026	4.60E−07	25.42
rs67401230*		5	AC116606.1	T	0.1676	0.0331	4.12E−07	25.63	
rs78957137*		8	EIF3H	A	0.2808	0.0541	2.10E−07	26.94	
SVS	rs12445022	16	RP11-482M8.1	A	0.1301	0.0216	7.27E−08	36.27	

**Table 1** continued

Exposures	rsID	Chr	Gene	EA	$\beta$	SE	<i>p</i> value	<i>F</i>
	rs111338112 <sup>*</sup>	22	PES1	T	0.3479	0.0683	3.51E−07	25.94
	rs16890461 <sup>*</sup>	8	SFRP1	T	− 0.1674	0.0406	3.95E−07	17.00
	rs72932716 <sup>*</sup>	2	ICA1L	T	0.1595	0.0302	4.58E−07	27.89
	rs7766042	6	RP11-157J24.2	T	− 0.2129	0.0351	1.37E−07	36.79
CES	rs1015037	10	SH3PXD2A	T	0.1437	0.0249	7.88E−09	33.31
	rs113580960	5	PDZD2	T	− 0.2829	0.0544	1.99E−07	27.04
	rs117015542 <sup>*</sup>	8	LINC00588	T	− 0.2185	0.0432	4.24E−07	25.58
	rs2948098	3	CRIP1P2	A	− 0.0945	0.0182	2.08E−07	26.96
	rs616154	9	ABO	T	0.0949	0.0160	3.01E−09	35.17
	rs6536024	4	LRAT	T	− 0.0903	0.0161	2.04E−08	31.45
CMBs	rs6950978	7	ABCB1	A	− 0.1698	0.0331	2.90E−07	26.31
	rs769449	19	APOE	A	0.2779	0.0496	2.11E−08	31.39

*CMBs* cerebral microbleeds; *IVW* inverse-variance weighted; *WM* weighted median; *IS* ischemic stroke; *LVS* large-vessel atherosclerosis stroke; *SVS* small-vessel occlusion stroke; *CES* cardio-embolic stroke; *Chr* chromosome; *EA* effect allele; *SE* standard error

<sup>a</sup>Since the MR estimates of MR-Egger and IVW were inconsistent, we tightened the instrument *p*-value threshold, and those SNPs were eliminated

using the MR-Egger approach (OR 1.11, 95% CI: 0.27–4.52, *p* = 0.89). Then, we tightened the instrument *p* value threshold to  $3 \times 10^{-7}$ , and five SNPs were identified. The result of IVW still indicated that CES had no causality on CMBs (OR 1.07, 95% CI: 0.71–1.60, *p* = 0.75). The *p* value for MR-Egger intercept is > 0.05. No outliers were identified in the leave-one-out plot (Supplementary Fig. 2B). No pleiotropy and heterogeneity were observed (Supplementary Fig. 1C and D). No pleiotropy was found in the MR-PRESSO global test (*p* = 0.179).

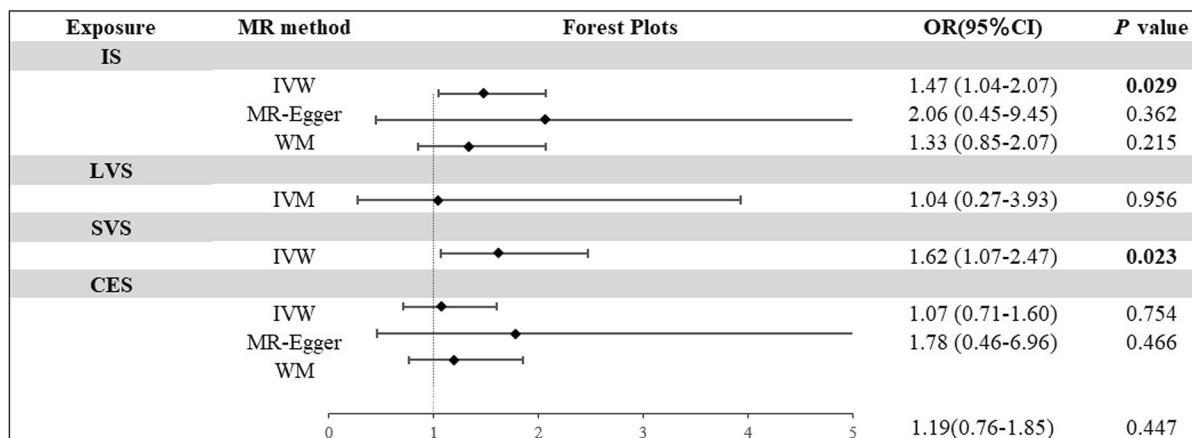
### Causal Effects of CMBs on IS and Its Subtypes

We further performed bidirectional MR analysis to estimate the causal effects of CMBs on IS and its subtypes. At the global level, two SNPs were

taken as IVs for CMBs (Table 1). We found that CMBs had no causal relationship with IS (IVW OR 1.01, 95% CI 0.94–1.08, *p* = 0.818), LVS (IVW OR 1.10, 95%CI 0.92–1.33, *p* = 0.30), SVS (IVW OR 1.07, 95% CI 0.88–1.31, *p* = 0.49) and CES (IVW OR 0.93, 95% CI 0.79–1.11, *p* = 0.43). No obvious heterogeneity was observed. MR-Egger and MR-PRESSO global tests could not be performed. Details are presented in Table S2.

## DISCUSSION

To our knowledge, this is the first two-sample bidirectional MR study to comprehensively evaluate the causal relation between IS and CMBs. In the present MR study, we found that IS and SVS causally increased the risk of CMBs. However, no significant causal effect of CES and LVS on CMBs was observed, and no evidence to



IVW, Inverse-variance weighted; WM, weighted median; CMBs, cerebral microbleeds; IS, ischemic stroke; LVS, large-vessel atherosclerosis stroke; SVS, small-vessel occlusion stroke; CES cardio-embolic stroke.

**Fig. 2** Mendelian randomization estimates from ischemic stroke and its subtypes on genetically predicted cerebral micobleeds. *CMBs* cerebral microbleeds; *IVW* inverse-variance weighted; *WM* weighted median; *IS* ischemic

stroke; *LVS* large-vessel atherosclerosis stroke; *SVS* small-vessel occlusion stroke; *CES* cardio-embolic stroke

support the causal links of CMBs on IS and its subtypes.

Previous studies reported that the incidence of CMBs in the early onset of acute IS patients was about 12–39% [28], and 12.7% of IS patients developed new CMBs within 1 week after the onset [29]. In addition, it was found that IS patients with CMBs have a more than threefold increased risk of stroke recurrence compared with those without CMBs [30, 31], including hemorrhagic transformation or recurrent IS [32]. As of yet, the relationship between IS and CMBs has remained unclear. In all previous observational studies, it was difficult to avoid violations from confounding risk factors, while this present study, by applying MR methods, enabled us to confidently prove causality without bias because of a better study design. The primary MR analyses in our study were performed by IVW method, which provides the most precise estimates. If the IVW method result is significant ( $p < 0.05$ ), even if the results of other methods are not, and no pleiotropy and heterogeneity were identified, it can be regarded as a positive result, provided that the beta values of the other methods are in the same direction. Therefore, through this MR study, we have sufficient reason to show that CMBs occur

more frequently in patients with any IS and SVS. However, when exploring the causal effects of SVS on CMBs, because of the presence of only two significant and independent SNPs (rs12445022 and rs7766042) associated with SVS, sensitivity studies such as MR-Egger and MR-PREESO cannot be conducted. Furthermore, to explore whether CMBs have causal effects on IS and its subtypes, we also performed reverse MR analysis and found that CMBs are not causally connected with IS. A meta-analysis of patients with IS and transient ischemic attacks (TIA) found that CMBs increased the risk of recurrent IS [33]. A European study also discovered that CMBs led to a ninefold increased risk of recurrent IS [11]. Accordingly, CMBs may not contribute to IS, but they may increase the risk of stroke recurrence.

Although we found that IS increased the occurrence of CMBs, the pathophysiologic mechanisms and histopathologic basis of the relationship between CMBs and IS remain poorly understood and highly debated [10]. It was hypothesized that small vessels could be damaged by CMBs, leading to thrombosis in situ and reduced arterial circulation distal to CMBs, which not only reflect vessel fragility and endothelial instability, but also increase the risk

of IS and hemorrhagic stroke [10, 11]. Additionally, intracranial microvessels are particularly sensitive to excessive pressure and hemodynamic pulsatility. For patients with IS, the intracranial microvascular resistance decreases, and the excessive pressure in the carotid artery will be directly transmitted to the intracranial microcirculation, increasing cerebral blood perfusion, thereby playing a compensatory role [34]. However, high-pulse blood flow may cause microvascular damage to the brain, which may lead to damage to smooth muscle cells and vascular endothelial cells, thereby promoting the development of CMBs [35]. A study has also shown new CMBs are most likely to appear in the early stage of acute IS patients, and early control of blood pressure will count for much to prevent the occurrence and development of CMBs [29].

Several strengths of this study include the use of summarized statistics derived from very large genetic association studies. Additionally, several conservative MR methods were used to assess the consistency of the results. Due to the fact that genetic variation is allocated at the time of conception, MR analysis can also prevent reverse causation. There are, however, several limitations to our study. First, although there is no cohort overlap between IS ( $n = 1,296,908$ ) and CMBs ( $n = 55,280$ ), both population-based studies were obtained from mostly European research and might therefore have some sample overlap, resulting in inflation of test results. However, to avoid biases in causal estimates introduced by sample overlap, we chose the maximal sample sizes in GWAS while minimizing sample overlap between exposures and outcomes [36]. Second, the enrolled patients were all European; hence, there is no evidence that IS and CMBs are causally linked in other populations. Another factor to consider is that potential violations of instrumental variable assumptions could bias the MR analysis. Causing effect estimates to be biased may be the result of directional pleiotropy, which is difficult to completely eliminate. It has been observed, however, that pleiotropic effects are not evident in MR-Egger regression analysis or sensitivity analyses with other instruments, and

other robust models show mostly similar results.

## CONCLUSIONS

This is the first two-sample bidirectional MR study to comprehensively evaluate the potential causal relation between IS and its subtypes with CMBs. Accordingly, our study provides evidence that IS and SVS are causally linked to an increased risk of CMBs. Further research is needed to determine the mechanisms of association between IS and CMBs.

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**Author Contributions.** Renjie Liu proposed the idea and wrote the draft of the manuscript. Haoyuan Yin, Jiahui Feng and Yuhao Zhao contributed to the data analysis and manuscript revision. Jianmin Piao, Zhongxi Yang and Xin Shi prepared the figures and tables. Xuan Chen supervised the whole research and is responsible for the integrity of the study. All authors read and approved the final manuscript.

**Disclosures.** Renjie Liu, Xin shi, Jiahui Feng, Jianmin Piao, Zhongxi Yang, Yuhao Zhao, Haoyuan Yin and Xuan Chen have nothing to disclose.

**Compliance with Ethics Guidelines.** This study used publicly available de-identified data from participant studies that were approved by an ethical standards committee with respect to human experimentation. No separate ethical approval was required in this study.

**Data Availability.** All data generated or analyzed during this study are included in this



article and its additional materials. Summary statistics generated by the GIGASTROKE consortium across ancestries and stroke subtypes are available in the GWAS Catalog (GCST90104540–GCST90104543). Summary statistics from published GWASs of CMBs are available in the Cerebrovascular Disease KP genetic association datasets (<https://cd.hugeamp.org/datasets.html>).

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