#### **ORIGINAL PAPER**



# Effects of Mindfulness-Based Interventions on Self-compassion in Health Care Professionals: a Meta-analysis

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#### **Abstract**

**Objectives** Health care professionals have elevated rates of burnout and compassion fatigue which are correlated with poorer quality of life and patient care, and inversely correlated with self-compassion. Primary studies have evaluated the extent to which mindfulness-based interventions increase self-compassion with contradictory findings. A meta-analytic review of the literature was conducted to quantitatively synthesize the effects of mindfulness-based interventions on self-compassion among health care professionals.

**Methods** Twenty-eight treatment outcome studies were identified eligible for inclusion. Five cumulative effect sizes were calculated using random-effects models to evaluate differences of changes in self-compassion for treatment and control groups. Within and between group comparisons were evaluated. Sub-group and moderator analyses were conducted to explore potential moderating variables.

**Results** Twenty-seven articles (k = 29, N = 1020) were utilized in the pre-post-treatment meta-analysis. Fifteen samples (52%) included health care professionals and fourteen (48%) professional health care students. Results showed a moderate effect size between pre-post-treatment comparisons (g = .61, 95% CI = .47 to .76) for self-compassion and a strong effect size for pre-treatment to follow-up (g = .76, 95% CI = .41 to 1.12). The effect size comparing post-treatment versus post-control was moderate. One exploratory moderator analysis was significant, with stronger effects for interventions with a retreat component. **Conclusions** Findings suggest mindfulness-based interventions improve self-compassion in health care professionals. Additionally, a variety of mindfulness-based programs may be useful for employees and trainees. Future studies with rigorous methodology evaluating effects on self-compassion and patient care from mindfulness-based interventions are warranted to extend findings and explore moderators.

**Keywords** Mindfulness · Self-compassion · Health care professionals · Treatment outcome

Health care professionals, individuals who provide expert caregiving services to others, have emotionally demanding careers and high rates of emotional exhaustion, compassion fatigue, and burnout. They experience exposure to human suffering and death, and interact with challenging patients, families, and co-workers on a regular basis. Simultaneously,

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Department of Psychology, Bowling Green State University, 822 E Merry Ave, Bowling Green, OH 43403, USA health care professionals must perform their job tasks efficiently, accurately, and ethically (Dyrbye et al. 2005; Lee et al. 2009; Murphy et al. 2009; Regehr et al. 2014; Stewart et al. 1997). Many health care professionals also struggle with fluctuating and long work hours, insurance difficulties, changing workplace roles, and low staffing (Regehr et al. 2014; Rutledge et al. 2009; Stucky et al. 2009; Wallace et al. 2009). Given these factors, it comes as no surprise that this population is particularly susceptible to stress (Aiken et al. 2002).

Stress in health care professionals is a serious concern because it can adversely affect their mental health, quality of life, and job performance (Galantino et al. 2005; McVicar 2003; Spickard Jr et al. 2002). Stress is also correlated with reduced ability to establish strong relationships with patients (Pastore et al. 1995), which is a critical component to positive



therapeutic outcomes. Thus, health care professionals experiencing high levels of stress may deliver sub-optimal patient care and are more likely to make medical errors (Leiter et al. 1998; Shanafelt et al. 2002; Vahey et al. 2004; Williams et al. 2007).

Furthermore, prolonged stress is a precursor to burnout. Burnout is a syndrome of emotional exhaustion, excessive stress, loss of meaning in work, feelings of ineffectiveness, and a tendency to view people as objects rather than people (Maslach et al. 1996). A 2014 survey of 6880 US physicians found a 54.4% prevalence rate of burnout (Shanafelt et al. 2015). Moreover, Shanafelt et al. (2012) found burnout to be more common among physicians relative to the general US working population. Similar to stress outcomes, these high rates of burnout in health care professionals are important because they are correlated with poorer mental health, quality of life, and quality of patient care (Brazeua et al. 2010; Poghosyan et al. 2010; Rosen et al. 2006; Shanafelt and Dyrbye 2012; Verdon et al. 2008).

In addition to stress and burnout, health care professionals can also experience compassion fatigue. Compassion is when one experiences feelings of concern for another who is suffering, coupled with the desire to help alleviate their suffering (Klimecki and Singer 2012). The other-oriented essence of compassion promotes prosocial behavior. Compassion fatigue is characterized by a reduced capacity for compassion (Klimecki and Singer 2012). Arising from repetitive and intense exposure to persons experiencing trauma and suffering, compassion fatigue is thought of as a form of secondary traumatic stress. Similar to Post-Traumatic Stress Disorder, compassion fatigue is characterized by intrusive thoughts of patient suffering and trauma, chronic physiological activation, and avoidance of interactions where suffering is involved (Cocker and Joss 2016; Gallagher 2013). This reduction of compassion is of concern for health care professionals because it may adversely affect their ability to be sensitive, nonjudgmental, and respectful to patients (Gilbert 2005; Wiklund Gustin and Wagner 2013).

The construct of compassion is related, yet distinct from empathy (Klimecki and Singer 2012). Empathy refers to sharing the same feeling as the suffering person and usually occurs prior to feelings of compassion. However, empathetic distress is self-oriented and involves being overwhelmed by experiencing the emotion of the sufferer, often leading to withdrawal behavior (Klimecki and Singer 2012). For health care professionals, it is important to encourage a compassionate response and reduce compassion fatigue to allow them to continue engaging in prosocial helping behaviors to aid their patients and clients. With this skill intact, they can acknowledge the suffering of their clients and care for them, without being overwhelmed by the painful emotions themselves.

Given the negative consequences of stress, burnout, and compassion fatigue among health care professionals, it is critical to identify ways to prevent or reduce their occurrence and severity. One promising area of research is focused on self-compassion. The construct of self-compassion originates in Buddhist philosophy and was defined by Neff (2003) as three interconnected components: mindfulness, self-kindness, and common humanity. Self-compassion involves the employment of these three components during times of pain, failure, and difficulty. Mindfulness skills are particularly important because they promote an enhanced present-moment awareness and a willingness to experience emotions with openness, curiosity, and acceptance. Self-kindness refers to letting go of judgment and criticism, and employing kindness toward the self. Finally, common humanity is the concept that other human beings experience difficulties in life which can help prevent self-pity, isolation, and feelings of shame. The Self-Compassion Scale (SCS) was developed by Neff in 2003 and assesses the positive and negative aspects of these three main self-compassion components: mindfulness versus overidentification, self-kindness versus self-judgment, and common humanity versus isolation. This measure has demonstrated adequate construct and convergent validity (Neff 2003), and the SCS-Short Form has shown good test-retest reliability and internal consistency ( $\alpha = .87$ ; Raes et al. 2011).

Self-compassion is distinct from the construct of mindfulness. Both constructs involve turning one's awareness toward their inner experiences with an accepting stance (Neff and Dahm 2015). However, the general mindfulness construct focuses on paying attention to any experience, not exclusively painful ones. Additionally, self-compassion includes elements of self-kindness and common humanity, which may or may not occur through mindfulness alone. Furthermore, mindfulness practice focuses on the internal experience while self-compassion emphasizes the "experiencer" of the suffering (Neff and Dahm 2015).

Since the development of the SCS, research on the construct of self-compassion and correlates has expanded. Within the general population, findings indicate that individuals who are more self-compassionate tend to report less burnout, anxiety, depression, shame, and fear of failure, and greater life satisfaction, social connectedness, emotional intelligence, and happiness (Barnard and Curry 2011; Mills et al. 2007; Neff et al. 2005; Neff et al. 2007a; Williams et al. 2008). Self-compassion has also been positively correlated with positive affect and negatively correlated with negative affect, emotional exhaustion, and shame (Barnard and Curry 2011; Leary et al. 2007; Neff et al. 2007b; Neff and Vonk 2009). Additionally, self-compassion has been found to be negatively correlated with rumination (Neff 2003).

Given the breadth of this literature, MacBeth and Gumley (2012) conducted a meta-analysis on the association between self-compassion and psychopathology. The study examined 20 samples reporting data from 4007 participants. Most of the participants were students, with health care professionals



accounting for about 5% of the total. They found that self-compassion was inversely correlated with stress (r=-.54, p<.0001), depression (r=-.52, p<.0001), and anxiety symptoms (r=-.51, p<.0001). Finally, higher levels of self-compassion have been associated with higher levels of empathic concern, altruism, perspective-taking, and forgiveness of others, all desirable traits in health care professionals (Neff and Pommier 2013). Overall, there is considerable evidence that higher levels of self-compassion are associated with positive aspects of well-being and inversely correlated with negative constructs.

Self-compassion is an important skill for health care professionals because it allows them to maintain their emotional sensitivity to patients. Self-compassion provides the health care professional in-the-moment self-care to alleviate personal empathetic distress and, therefore, proceed with compassionate care (Neff and Germer 2018). Preliminary research supports the relationship of higher self-compassion and overall positive variables of well-being in health care professionals (Beaumont et al. 2016a). Some promising early research in this area includes the recent cross-sectional survey of 213 health care professionals (Kemper et al. 2015). They reported that self-compassion was inversely correlated with sleep difficulties (r = -.27, p < .01), and positively correlated with resilience (r = .54, p < .01). Similarly, Richardson et al. (2016) reported the results of a cross-sectional survey of 307 medical students and residents indicating that self-compassion significantly and inversely predicted burnout ( $\beta = -.375$ ; p < .05). Finally, Duarte et al. (2016) conducted a cross-sectional study of 280 nurses and found that self-compassion with mindful awareness was associated with lower levels of burnout and compassion fatigue. Based on these findings, selfcompassion may play an important role in reducing burnout symptoms and enhancing well-being among health care professionals. The emerging research on self-compassion highlights the importance of fostering this ability in health care professionals who are particularly vulnerable to burnout, stress, and compassion fatigue.

One way to promote self-compassion in health care professionals may be through training in mindfulness meditation. Mindfulness has been defined as "the awareness that emerges through paying attention on purpose, in the present moment, and nonjudgmentally, to things as they are" (Kabat-Zinn 1994). Mindfulness-based therapies have gone by many different names and forms. The more common contemporary therapies are mindfulness-based stress reduction (MBSR), mindfulness-based cognitive therapy (MBCT), dialectical-behavioral therapy (DBT), and acceptance and commitment therapy (ACT)—although meditation is not a required component of ACT.

Mindfulness-based interventions share core components that include fostering awareness, increasing present-moment experiences, cultivating response flexibility, and improving affect tolerance. These skills are thought to help individuals become more aware of automatic thinking and acting, interrupt rumination about past experiences and worries about future events, promote "considered action," and learn to allow emotional experiences to rise and fall without behaviorally responding, while simultaneously engaging in more adaptive behaviors. Mindfulness-based therapies can be implemented for individuals, small groups, or at an organizational level.

General mindfulness-based interventions (e.g., MBSR, MBCT) sometimes implicitly communicate components of self-compassion. Other more focused mindfulness interventions such as compassion-focused therapy and mindful selfcompassion explicitly teach self-compassion skills with an emphasis on conveying the importance of being kind to others and oneself during times of difficulty. Interestingly, even though general mindfulness-based interventions do not explicitly teach self-compassion, researchers have argued and demonstrated that self-compassion may be a key mediator to the positive outcomes observed from these types of interventions (Neff and Dahm 2015). For example, Birnie et al. (2010) found that after completing an MBSR intervention, community-sample participants showed an increase in selfcompassion and decreased personal distress, while no significant change was observed in empathic concern.

Researchers have begun to evaluate the effects of mindfulness training on self-compassion among health care professionals. For example, Shapiro et al. (2005) conducted a randomized controlled trial that evaluated an 8-week MBSR program for twenty-eight health care professionals who were randomly assigned to the treatment or a wait-list control group. Shapiro et al. (2005) assessed levels of self-compassion, perceived stress, psychological distress, burnout, and satisfaction with life before and after the intervention. Results showed significant treatment (n = 28) versus control (n = 18) differences on measures of perceived stress (F(2, 24) = 4.4,p = .04, d = 0.65) and self-compassion (F(2, 24) = 9.85, p = .004, d = 0.97). Additionally, a separate regression analysis showed that changes in self-compassion significantly predicted positive changes in perceived stress.

Conducting mindfulness-based interventions for time-limited health care professionals is logistically challenging. Obtaining "buy-in" from the organization and employees/students may take considerable conscious effort to implement an intervention effectively (Byron et al. 2015). Even when health care professionals are invested in a mindfulness-based intervention, work conflicts can often arise and affect attendance and home practice adherence (Luberto et al. 2017). Some studies also revealed nonsignificant changes in self-compassion among health care professionals participating in similar interventions (e.g., Brooker et al. 2013; Mahon et al. 2017; Romcevich et al. 2018). The current inconsistencies in the intervention methods (e.g., treatment type, length of



intervention, home practice) and discrepancies in corresponding results throughout the literature make it difficult to draw conclusions about the efficacy of these programs. There is, however, a large enough body of research to conduct a meta-analytic investigation.

A few meta-analyses exist in the current literature that relates to the current investigation. For example, Khoury et al. (2013) conducted a meta-analysis on mindfulness-based therapies and observed moderate to large changes in anxiety, depression, mindfulness, and stress outcomes. Kirby et al. (2017) investigated compassion-based interventions and identified significant changes for compassion, self-compassion, mindfulness, depression, and anxiety. For both studies, the population included was restricted to adults, providing little insight on important demographic variables, such as professions and lifestyles. Burton et al. (2017) conducted a metaanalysis on mindfulness-based intervention focused exclusively on health care professionals. Stress significantly improved; however, the authors noted that the focus on only one outcome was limiting. Burton et al. (2017) emphasized the need for future studies to investigate "dosage" and "active ingredients" in mindfulness-based interventions to help condense these programs to meet the time-limited needs of this population.

The specific aims of this project were to (1) provide a systematic methodological review of the literature on treatment outcome studies evaluating the extent to which mindfulness-based interventions produce change in self-compassion in health care professionals, (2) calculate the effect sizes associated with mindfulness-based interventions targeting self-compassion among health care professionals, and (3) explore potential moderators of mindfulness-based intervention effect sizes.

#### Methods

#### **Article Identification and Selection**

An unstructured review of the literature up until June 2018 was conducted. Informed by the preliminary literature review, the following databases were determined to be relevant to this meta-analysis: PsycINFO, Academic Search Complete, MEDLINE, Psychological and Behavioral Sciences Collection, CINAHL Plus, and Humanities International Complete. The preliminary literature review indicated that the following search terms would provide a sufficiently wide nomological network of variables: mindfulness, mindfulness-based cognitive therapy, mindfulness-based stress reduction, acceptance and commitment therapy, dialectic behavior therapy, stress management, yoga, meditation, mindfulness-based compassion training, compassion, and self-compassion. All the possible combinations of the intervention and self-

compassion variables were used in the literature search with no search constraints.

For the current study, mindfulness-based interventions were defined as any intervention that was based on either (a) a previously established mindfulness-based therapy (e.g., MBSR) or (b) included explicit mindfulness skills training. As explained previously, mindfulness-based interventions have common core components that theoretically translate across intervention types. Health care professional terms were not used in the article search because of the many different terms used to describe this population. By not including these terms, all studies in this area were screened for health care professional participants. We also included students in training who are seeing patients. The purpose of this decision was to include all health care professionals who are providing patient care. Even though the skill levels and experiences vary, burnout and compassion fatigue, which may affect patient care, has been documented across all included participant groups. Overall, the rationale for the wide scope of intervention type and health care professional participants allowed us to gather the most information we could about the possible relationships among interventions and outcomes.

An additional search for articles was conducted by descendency and ascendency through the authors and references, respectively. Authors of articles with insufficient data were contacted via email to request missing information. Finally, a search for articles was conducted of the references section in related review and meta-analysis articles (Chiappetta et al. 2018; Irving et al. 2009; Kirby et al. 2017; Khoury et al. 2013; West et al. 2016).

All potential articles identified in the literature search were evaluated for relevance to this meta-analysis using the following inclusion criteria: (a) a mindfulness-based intervention was provided; (b) the sample included health care professionals (e.g., medical students, medical and psychological trainees, physicians, nurses, psychologists, midwives); (c) an experimental pre-post-design was used to evaluate outcomes; (d) self-compassion was used as an outcome variable; (e) the article was published in an academic journal or dissertation (if full text was available); and (f) the article was written in the English language. Randomized and non-randomized trials were included. Inclusion criteria also required that the study use an explicit quantitative measure of self-compassion, such as the Self-Compassion Scale (SCS; 26 items; Neff 2003) or the SCS-Short Form (SCS-SF; 12 items; Raes et al. 2011).

Articles were evaluated by two independent raters through multiple steps in succession in the following order: title, abstract, and full-text reviews. Studies that explicitly contained rule-out criteria (e.g., book reviews, literature reviews, cross-sectional studies, correlational studies) were excluded from further review. All disagreements were discussed and coding rules were updated as necessary. The search results produced 683 articles, all of which were screened for inclusion by two



independent raters. Cohen's kappa was calculated to determine interrater reliability. The interrater reliability for title selection was k = .80, p < .0001 and 206 articles were retained for further review. The interrater reliability for abstract selection was k = .85, p < .01 and 74 articles retained for full-text review. At this stage, seven additional articles were identified in a review of references. Both raters reviewed 20 of these 81 articles in order to establish an agreement rate of 95%. The remaining studies were reviewed independently by each coder for full-text review. See Fig. 1 for a flowchart of article selection. Two raters independently coded nineteen data points (i.e., means, standard deviations, effect sizes) from five articles and demonstrated a strong reliability (k = 1.00). The remaining data points were coded by one researcher.

### **Effect Size Calculation and Analyses**

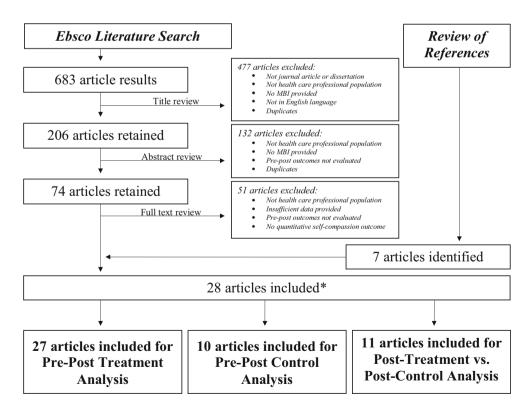
Becker's d was designed to be analogous to the standard treatment-control effect size (Becker 1988). Becker argued that the treatment might affect not only the mean of a sample but also the standard deviation. Thus, the formula used  $SD_{pretreatment}$  in the denominator to standardize the difference between pre-treatment and post-treatment means as follows: Becker's  $d = M_{post} - M_{pre}/SD_{pre}$ . Morris and DeShon (2002) established that the denominator of Becker's d should be adjusted by the pre-treatment to post-treatment correlation in

order to render a more accurate estimate of the population parameter. Smith and Beretvas (2009) demonstrated that using a pre-treatment-post-treatment correlation—adjusted pooled standard deviation in the denominator was superior to Becker's formula which used only the pre-treatment standard deviation (Smith and Beretvas 2009).

In this investigation, we thus used the effect size formula developed by Morris and DeShon (2002) and recommended by Smith and Beretvas (2009) which is  $d_{\rm rm} = M_{\rm post} - M_{\rm pro}/SD_{\rm difference}$ , where  $SD_{\rm difference} = \sqrt{Sd^2}_{\rm pre} + SD^2_{\rm post} - 2r_{\rm prepost}SD_{\rm pre}SD_{\rm post}$ . The  $d_{\rm rm}$  was then converted to Hedges  $g_{\rm rm}$  by multiplying it by the standard bias adjustment (Hedges 1982): Hedges  $g_{\rm rm} = d_{\rm rm} \times (1-3/(4(n-1)-1))$ . The sampling variance of  $d_{\rm rm}$  was then calculated as Variance  $d_{\rm rm} = (1/n)(n-1/n-3) \left[1 + (nd^2_{\rm rm}] - d^2_{\rm rm}/[c(n-1)]^2\right]$ , where c(n-1) = 1 - (3/(4(n-1)-1)). Given that the Variance  $d_{\rm rm}$  includes the bias adjustment (i.e., c(n-1)), the same value was used as the variance for Hedges g (Smith and Beretvas 2009).

To calculate the  $d_{\rm rm}$  using the formulas above, it was necessary to know the pre-treatment to post-treatment correlation. This statistics is infrequently reported in research articles (Morris and DeShon 2002). Among the studies included in this meta-analysis, none reported the pre-treatment to post-treatment correlation. We were able to impute the pre-treatment to post-treatment correlation using alternative data. Our review of all included articles indicated that 9 provided

**Fig. 1** Flowchart of article selection



<sup>\*</sup>Articles were utilized in multiple analyses; MBI, mindfulness-based intervention; Ebsco literature search included all databases described in methods section



pre-post-t test results that could be used to calculate the pre-treatment to post-treatment correlation using Morris and DeShon's (2002) two formulas:  $SD_{difference} = n(M_{post} - M_{pre})^2/t^2_{pre-post}$  and  $r = SD^2_{pre} + SD^2_{post} - SD_{difference}/2(SD_{pre} \times SD_{post})$ . The 9 imputed correlations were then weighted by sample size and a final weighted average pre-treatment to post-treatment correlation of r = .62 was calculated. This correlation was then used on all subsequent calculations of effect sizes, effect size variance, and sampling error.

The final cumulative effect sizes, confidence intervals, homogeneity, bias, and fail-safe values were calculated using MetaWin (Rosenberg et al. 2000). A random-effects model was utilized because different health professional samples were used and multiple effects were expected in addition to sampling error. Jamovi version 1.0.5 was used to conduct moderator analyses and to generate forest plots and funnel plots (The Jamovi Project 2019). In all, the following cumulative effect sizes were calculated: pre-post-treatment for treatment groups, pre-post for control groups, pre-treatment to follow-up for treatment groups, pre-post-comparison between treatment and control groups, and treatment versus control group comparisons at post-treatment (note: treatment versus control group comparisons at follow-up were not conducted because there were too few effect sizes). Confidence intervals were generated for all cumulative effect sizes. Heterogeneity of the cumulative effect sizes were evaluated using Ototal (Huedo-Medina et al. 2006). Since Qtotal can be underpowered with small samples,  $I^2$  was also calculated (Higgins and Thompson 2002). The formula for  $I^2$  was [Qtotal - (k-1)/ $Qtotal \times 100$ . In cases where the Qtotal < k-1,  $I^2$  was truncated to zero, as recommended by Higgins and Thompson (2002).

Forest plots were generated to illustrate the results of the included samples. Three methods were used to assess publication bias. First, funnel plots were used to display the relationship between sample size and effect size. If there is no publication bias, the diagram is expected to maintain symmetry (Egger et al. 1997). Second, Kendall's Tau was calculated to evaluate the relationship between effect size and sample size. Finally, Rosenthal's fail-safe number was calculated, which provides a number that indicates how many nonsignificant hypothesis tests would be needed to raise the overall *p* value to greater than .05. The robustness of Rosenthal's fail-safe *N* was evaluated by comparing it against the 5N+10 threshold (Rosenberg 2005). When the fail-safe *N* exceeded that threshold, the cumulative effect size is classified as robust.

The current manuscript met the Preferred Reporting Items for Systematic Review and Meta-Analyses (PRISMA) guidelines (Moher et al. 2009). See Supplementary Materials for complete PRISMA checklist. Risk of bias was assessed across studies using the previously stated tests of publication bias and heterogeneity. All included studies were evaluated using the

revised Jadad criteria to assess risk of bias within studies (Jadad et al. 1996; Piet and Hougaard 2011). Two raters independently assigned one point for each fulfilled criterion. Each study was assigned a total score which ranged from 0 to 4 (see Table 1). Disagreements between raters were resolved by discussion. Additional sub-analyses were conducted to examine possible effects of risk of bias and study quality.

### **Results**

#### **Methodological Characteristics of Studies**

See Tables 2, 3, and 4 for detailed information on samples included for each meta-analysis. Fifteen samples (52%) included health care professionals, while the remaining fourteen (48%) included professional health care students, with nurses as the most common primary population (k = 7, 24%). Seventeen samples (59%) included medical professionals and eleven (38%) included professionals in a psychological or social field. There were eleven (38%) samples that investigated manualized treatment protocols, while the remaining 18 (62%) utilized an adapted, modified, or abbreviated intervention. Most of the samples (k = 13, 45%) implemented either MBSR (k = 5, 17%) or an adapted version of MBSR (k = 8, 28%). Other types of mindfulness-based interventions were MBCT, ACT, compassion-focused therapy, yoga, and other mindfulness- and compassion-based specialized trainings and programs (see Table 2).

A majority of the samples provided in-person, group-based interventions (k = 25, 86%), with two online-based interventions (7%), and two interventions that utilized both in-person and online format (7%). Five samples (17%) indicated that the intervention was part of an academic class, course, or curriculum. Out of the in-person, group-based interventions, 22 (76%) samples described the qualifications of the intervention leader or facilitator. Of those samples that reported information on the intervention leader, eight (36%) explicitly stated that sessions were led by a licensed mental health professional, ten (45%) by an experienced or trained individual, three (14%) were listed as faculty, and one (5%) was described as a registered yoga teacher. Twelve samples were conducted in the USA (41%), five in Australia (17%), four in the UK (14%), four in Canada (14%), and the following had one sample each: India (3%), Portugal (3%), Brazil (3%), and the Netherlands (3%). A strong majority of samples (k = 26, 90%) were peer-reviewed journal articles, and the remaining were dissertations (k = 3, 10%).

The average number of sessions was 7.6, ranging from 1 to 40 sessions. The average duration of a session was 1.82 hour, ranging from 1 to 2.5 h. The average span of an intervention was 6.61 weeks, ranging from 1 to 12 weeks. The average total intervention time (not including home practice) was



 Table 1
 Methodological quality of included studies

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ומסוב ו זאוכוווס	Methodological quanty of included studies	inty of included	siddles								
Study	Did study include a comparison group? <sup>a</sup>	Was the trial randomized?	Was the trial Was the randomized? randomization procedure described and was it appropriate?	Was the treatment allocation concealed?	Were groups similar at baseline on prognostic indicators?	Were blind outcome assessments conducted?	Was the number of withdrawals/ dropouts in each group mentioned?	In addition to stating the number of withdrawals/dropouts, were reasons given for each group?	Was an analysis conducted on the intention-to-treat sample?	Was a power calculation described?	Jadad score (revised, maximum score = 4)
Bazarko et al. (2013)	No	No	No	No	No	No	Yes	No	No	No	0
Beaumont et al. (2016b) a	No	No	No	No	No	No	No	No	No	No	0
Beaumont et al. (2016b) b	No	No	No	No	No	No	No	No	No	No	0
Beaumont (2016b) c	No	No	No	No	No	No	No	No	No	No	0
Beck et al. (2017)	Yes	Yes	No	No	Yes	No	No	No	No	No	
Bond et al. (2013)	No	No	No	No	No	No	Yes	No	No	No	0
Brooker et al.	No	No	No	No	No	No	Yes	No	No	No	0
Crowder and	Yes	Yes	No	No	Yes	No	Yes	No	No	No	1
Sears (2017) Danilewitz et al. (2016)	Yes	Yes	No	No	Yes	No	Yes	No	No	No	
Duarte and Pinto-Gouve-	Yes	No	No	No	Yes	Yes	Yes	Yes	No	No	2
ia (2010) Erogul et al. (2014)	Yes	Yes	Yes	No	Yes	Yes	Yes	Yes	No	Yes	4
Finlay-Jones et al. (2017)	No	No	No	No	No	No	Yes	No	No	Yes	0
Irving (2011)	No	No	No	No	No	No	Yes	Yes	No	No	-
Lopes et al. (2018)	No	No	No	No	No	No	Yes	No	No	Yes	0
Marx et al. (2014)	No	No	No	No	No	No	Yes	Yes	No	No	_
Mathad et al. (2017)	Yes	Yes	No	No	Yes	No	Yes	Yes	No	Yes	2
Pakenham (2015)	No	No	No	No	No	No	Yes	No	No	No	0
Penque (2009)	No	No	No	No	No	No	Yes	Yes	No	Yes	1
Raab et al. (2015)	No	No	No	No	No	No	No	No	No	No	0
	No	No	No	No	No	No	No	No	No	No	0



Table 1 (continued)

Study	Did study include a comparison group?a	Was the trial Was the randomized? randomid procedul described it appropriate the control of the	zation e d and was riate?	Was the treatment allocation concealed?	Were groups similar at baseline on prognostic indicators?	Were blind outcome assessments conducted?	Was the number of withdrawals/ dropouts in each group mentioned?	In addition to stating the number of withdrawals/ dropouts, were reasons given for each group?	Was an analysis conducted on the intention-to- treat sample?	Was a power calculation described?	Jadad score (revised, maximum score = 4)
Rao and Kemper (2017) Rimes and Wingrove	oN S	No	No 1	oN o	°Z	No	Yes	Yes	°Z	°Z	_
(2011) Rodrigues et al.	No	No	No ]	No	N <sub>o</sub>	No	Yes	Yes	No	No	1
Romcevich et al. (2018)	No	No	No ]	No	No	No	Yes	No	No	No	0
Runyan et al. (2016)	No	No	No	No	No	No	Yes	No	No	Yes	0
Scarlet et al. (2017)	No	No	No	No	No	No	Yes	Yes	No	No	
Shapiro et al. (2007)	Yes	No	No	No	Yes	No	Yes	No	No	No	0
Shapiro et al. (2005)	Yes	Yes	No	No	No	No	Yes	Yes	No	No	2
Slatyer et al. (2017)	Yes	No	No	No	Yes	No	Yes	Yes	No	Yes	1
3rown am	Yes	No	No	No	Yes	No	Yes	Yes	Yes	No	1
(2012) Verweij et al. (2017)	Yes	Yes	Yes	No	Yes	No	Yes	Yes	Yes	Yes	33

Columns in italics constitute the revised Jadad Scale



<sup>&</sup>lt;sup>a</sup> Not part of Jadad scale, supplemented to provide context for scores

 Table 2
 Summary of included studies for pre-post-treatment comparisons for treatment groups

Study	Total sample, $N$	Participants	Intervention	Intervention format	Intervention leader	# of sessions	Session length (in hours)	# of weeks
<sup>a</sup> Bazarko et al. (2013)	36	Nurses	Telephonic MBSR	In-person and telephonic groups	Experienced mindfulness instructor	9	1.5	∞
Beaumont et al. (2016b) a	11	Nurses and midwives	<sup>b</sup> Compassion-focused therapy	In-person group	Experienced trainer and practitioner in CFT	3	NA	_
Beaumont et al. (2016b) b	10	Counselors and therapists	<sup>b</sup> Compassion-focused therapy	In-person group	Experienced trainer and practitioner in CFT	8	NA	_
Beaumont (2016b)	7	Health care providers	<sup>b</sup> Compassion-focused	In-person group	Experienced trainer and practitioner in	3	NA	_
Beck et al. (2017)	20	Communication sciences and disorders and speech-language pathology students	Mindfulness & yoga practice	In-person group	Registered yoga instructor	∞	NA	∞
Bond et al. (2013)	24	Medical students	Embodied health: yoga and mindfulness elective	In-person group	University professor	11	1.5	11
Brooker et al. (2013)	34	Disability support staff	Occupational mindfulness (adapted MBSR)	In-person group	Mental health professionals with formal MBCT training	∞	2	∞
<sup>a</sup> Crowder and Sears (2017)	7	Social workers	<sup>b</sup> MBSR	In-person group	Faculty of social work	∞	2.5	∞
Danilewitz et al. (2016)	13	Medical students	<sup>b</sup> Mindfulness meditation program (adapted MBSR)	In-person group	Medical student with MBSR training and psychologist	∞	1.5	∞
Duarte and Pinto-Gouveia	29	Oncology nurses	<sup>b</sup> Adapted MBSR	In-person group	Clinical psychologist trained in MBSR	9	7	9
<sup>a</sup> Erogul et al. (2014)	28	1st year medical students	Abridged MBSR	In-person group	NA	∞	1.5	∞
<sup>a</sup> Finlay-Jones et al. (2017)	20	Psychology trainees	<sup>b</sup> Self-compassion online program	Online modules	NA	9	NA	9
Irving (2011)	51	Social work graduate students	Adapted MBSR (mindfulness-based medical practice)	In-person groups	PhD psychologist and palliative care physician	∞	2.5	∞
Lopes et al. (2018)	49	Nursing workers	<sup>b</sup> MBPM (adapted MBSR)	In-person group	NA	8	1	∞
<sup>a</sup> Marx et al. (2014)	37	Mental health service staff	MBCT	In-person groups	Experienced mindfulness teacher and trainee co-teacher	∞	7	∞
Mathad et al. (2017)	40	Nursing students	Yoga	In-person groups	NA	40		∞
Pakenham (2015)	32	Postgraduate clinical psychology trainees	ACT	In-person class	NA	12	2	12
Penque (2009) Raab et al. (2015) Rao and Kemper (2017)	61 22 153	Registered nurses Mental health professionals Health professionals	Adapted MBSR <sup>b</sup> MBSR <sup>b</sup> Online meditation training	In-person groups In-person groups Online modules	MBSR trained, licensed therapist Certified MBSR instructor NA	∞ ∞ €	2 2.5 1	% % N



Rimes and	20	Clinical psychology doctoral trainees	MBSR	In-person training	Post-doctoral clinical psychologist	8	2.5	∞
Wingrove (2011)		}		course				
Rodrigues et al. (2018)	33	Nurses	<sup>b</sup> Skills in working in pediatric pain program	In-person groups	Clinical psychologist	-	1.5	1
<sup>a</sup> Romcevich et al. (2018)	7	Residents	Mind-body skills training	In-person groups and online modules	Resident with mindfulness experience	4	1.5	4
Runyan et al. (2016)	6	2nd year residents	<sup>b</sup> Family medicine resident wellness curriculum	In-person class	Behavioral science faculty member	4	2	4
<sup>a</sup> Scarlet et al. (2017)	62	Health care workers	<sup>b</sup> Compassion cultivation training	In-Person Groups	NA	∞	2	∞
Shapiro et al. (2007)	22	Therapist in training	$^{ m b}$ MBSR	In-person groups (part of course)	PhD-level instructor	∞	2	10
<sup>a</sup> Slatyer et al. (2017)	09	Nurses	<sup>b</sup> Mindful self-care and resiliency (MSCR)	In-person group (educational workshop)	Clinical psychologist experienced in MSCR	$\kappa$	1.75	4
<sup>a</sup> Stafford-Brown and Pakenham (2012)	28	Clinical psychology trainees	Adapted ACT	In-person groups	Registered psychologist with ACT training	4	2.5	4
Verweij et al. (2017)	80	Residents	MBSR	In-person groups	Trainer who met good practice guidelines for teaching mindfulness-based courses	∞	2.5	∞

Study	Full-day retreat	Full-day retreat Total session hours	Home practice activities	Self-compassion measure	Location of study	Hedges g	Location of Hedges Pre- to follow-up study g Hedges g
<sup>a</sup> Bazarko et al. (2013)	2 days	25	Formal and informal: CD, DVD, workbook, readings, SCS meditation	SCS	SO	1.40	1.64
Beaumont et al. (2016b) a	None	NA	None	SCS-SF	UK	0.29	1
Beaumont et al. (2016b) b	None	NA	None	SCS-SF	UK	0.65	-
Beaumont (2016b) c	None	NA	None	SCS-SF	UK	0.94	1
Beck et al. (2017)	None	NA	None	SCS	NS	0.50	1
Bond et al. (2013)	None	16.5	Readings and informal practice	SCS	NS	0.53	I
Brooker et al. (2013)	None	16	Formal & informal: mindfulness activities, stretching, audio recordings	SCS	Australia	0.03	ı
<sup>a</sup> Crowder and Sears (2017)	1 day	28	Experiential tasks	SCS	Canada	1.21	.74
Danilewitz et al. (2016)	None	12	Recorded meditation	SCS	Canada	.39	ı
Duarte and Pinto-Gouveia (2016)	None	12	Formal and informal: recorded meditation, mindfulness activities	SCS	Portugal	0.32	I
<sup>a</sup> Erogul et al. (2014)	1 day (5 h)	17	Meditation, audiorecordings, breathing, body scan, & SCS gentle voga	SCS	NS	1.11	1.30
<sup>a</sup> Finlay-Jones et al. (2017)	None	NA	NA NA	SCS	Australia	0.83	1.08
Irving (2011)	1 day	28	Formal and informal: CDs, daily mindfulness, yoga,	SCS	Canada	0.57	1
Lopes et al. (2018)	None	∞	& oreaumg Recorded meditation	SCS	Brazil	0.95	I



Table 2 (continued)

<sup>a</sup> Marx et al. (2014)	None	16	NA	SCS	SN	0.59	.61
Mathad et al. (2017)	None	40	NA	SCS-SF	India	0.39	I
Pakenham (2015)	None	24	Self-care practice	SCS	Australia	0.15	I
Penque (2009)	1 day (7 h)	15	Readings, recordings, body scan, and yoga	SCS	SO	1.89	I
Raab et al. (2015)	1 day of silence	28	Yes (not described)	SCS	Canada	0.54	I
Rao and Kemper (2017)	None	3	Guided and informal practice	SCS-SF	SO	0.52	I
Rimes and Wingrove (2011)	None	20	Yes (not described)	SCS	UK	0.53	I
Rodrigues et al. (2018)	None	1.5	NA	SCS-SF	SO	0.52	I
<sup>a</sup> Romcevich et al. (2018)	None	14 (includes 8 h of online modules)	Mindfulness plan and maintenance sessions	SCS-SF	SN	0.35	50
Runyan et al. (2016)	None	8	Self-reflection, journaling, readings mindful breathing SCS	SCS	SN	0.61	ı
<sup>a</sup> Scarlet et al. (2017)	None	16	Formal and Informal encouraged	SCS-SF	Australia	98.0	.91
Shapiro et al. (2007)	None	16	Mindfulness practices	SCS	SO	0.81	I
<sup>a</sup> Slatyer et al. (2017)	1 day workshop 11.5 (6 h)	11.5	Formal and informal, body scan, breathing space, and SCS-SF mindful eating	SCS-SF	Australia	0.29	.38
<sup>a</sup> Stafford-Brown and Pakenham (2012)	None	10	Yes (not described)	SCS	Australia	0.40	.45
Verweij et al. (2017)	1 day of silence 24 (6 h)	24	Body scan, yoga, walking meditation, & informal practice	SCS-SF	Netherlands	0.45	ı

<sup>a</sup> Included in pre-treatment to follow-up comparisons for treatment groups; <sup>b</sup> Explicit self-compassion skills training; *MBSR*, mindfulness-based stress reduction; *ACT*, acceptance and commitment therapy; *ABPM*, mindfulness-based pain and illness management; *MBCT*, mindfulness-based cognitive therapy; *SCS*, Self-Compassion Scale; *SCS-SF*, Self-Compassion Scale-Short Form; *Hedges g*, effect size; for articles with full-day retreats that did not indicate the # of hours, 8 h was used to calculate the total # of hours of intervention; *NA*, not available



 Table 3
 Summary of included studies for pre-post-control comparisons for control groups

Study	N	Participants	Control format	# of weeks	Location of study	Article type	Hedges g
Beck et al. (2017)	17	Communication sciences and disorders and speech-language pathology students	Inactive	8	U.S.	Journal	21
Crowder and Sears (2017)	6	Social workers	Wait-list	8	Canada	Dissertation	.54
Danilewitz et al. (2016)	9	Medical students	Wait-list	8	Canada	Journal	.24
Duarte and Pinto-Gouveia (2016)	19	Oncology nurses	Wait-list	6	Portugal	Journal	14
Erogul et al. (2014)	29	1 <sup>st</sup> year medical students	Inactive	8	U.S.	Journal	.15
Mathad et al. (2017)	40	Nursing students	Wait-list	8	India	Journal	10
Shapiro et al. (2007)	32	Therapist in training	Inactive	10	U.S.	Journal	05
Slatyer et al. (2017)	19	Nurses	Wait-list	4	Australia	Journal	.00
Stafford-Brown and Pakenham (2012)	28	Clinical psychology trainees	Wait-list	4	Australia	Journal	.32
Verweij et al. (2017)	67	Residents	Wait-list	8	Netherlands	Journal	.17

SCS, Self-Compassion Scale; Hedges g, effect size

17.06 h, ranging from 1.5 to 40 h. Twenty (69%) samples had interventions that included 10 or more hours of intervention. Twenty-two (76%) samples indicated that their intervention

included home practice; however, only sixteen (73%) studies described the home practice activities. Eight (28%) samples included at least a 1-day retreat as part of the intervention;

 Table 4
 Summary of included studies for post-treatment versus post-control comparisons

Study	N	Tx n	Cx n	Participants	Intervention	Self- compassion measure	Randomized?	Location of study	Hedges g
Beck et al. (2017)	37	20	17	Communication sciences and disorders and speech-language pathology students	Mindfulness practice	SCS	Non-randomized	US	0.40
Crowder and Sears (2017)	14	7	7	Social workers	MBSR	SCS	Randomized	Canada	0.98
Danilewitz et al. (2016)	22	13	9	Medical students	Mindfulness meditation program (adapted MBSR)	SCS	Randomized	Canada	0.35
Duarte and Pinto-Gouveia (2016)	48	29	19	Oncology nurses	Abbreviated MBSR	SCS	Non-randomized	Portugal	0.03
Erogul et al. (2014)	57	28	29	Medical students	MBSR	SCS	Randomized	US	0.88
Mathad et al. (2017)	80	40	40	Yoga	In-person groups	SCS-SF	Randomized	India	0.03
Shapiro et al. (2007)	54	22	32	Trainee therapists	MBSR	SCS	Non-randomized	US	0.42
Shapiro et al. (2005)	38	18	20	Health care professionals	MBSR	SCS	Randomized	US	1.20
Slatyer et al. (2017)	76	60	16	Nurses	Mindful self-care and resiliency (MSCR)	SCS-SF	Non-randomized	Australia	0.20
Stafford-Brown and Pakenham (2012)	56	28	28	Psychology interns	ACT stress management	SCS	Non-randomized	Australia	0.61
Verweij et al. (2017)	138	71	67	Residents	MBSR	SCS-SF	Randomized	Netherlands	0.38

N, total sample size; n, sample size; Tx, treatment group; Cx, control group; MBSR, mindfulness-based stress reduction; ACT, acceptance and commitment therapy; SCS, Self-Compassion Scale; SCS-SF, Self-Compassion Scale-Short Form; Hedges g, effect size



however, most did not indicate if the "day of silence" was together with the group or completed individually.

The methodological quality of the included studies using the revised Jadad criteria was reported in Table 1. Across all studies, scores ranged from 0 to 4 points (M = .80, SD = 1.00). When limited to only those studies with a comparison group, scores ranged from 0 to 4 points (M = 1.64, SD = 1.12).

# **Pre-Post-treatment Comparisons Among Treatment Groups**

After all inclusion and exclusion criteria were confirmed, the final number of articles included in the within group pre-posttreatment meta-analysis was 27 (k = 29, N = 1020). Table 2 provides a summary of the included studies. The overall effect size was moderate (Hedges g = .61, 95% CI = .47 to .76; see Table 5). The confidence interval did not contain zero; therefore, the null hypothesis of no treatment effect was rejected. The Ototal was nonsignificant (*Qtotal* (28) = 30.03, p = .36). The  $l^2$  of 6.67% also indicated that there was minimal effect size variation. Figure 2 displays forest plots of effect sizes and confidence intervals. Visual inspection of the funnel plot indicated symmetry suggesting little evidence of publication bias (see Fig. 3). Kendall's Tau was not significant (Tau = .22, p = .10) which further indicates an absence of publication bias. Rosenthal's fail-safe test suggested that there would have to be at least 790 unpublished, nonsignificant comparisons to raise the overall p value to greater than .05, which is considered robust.

### Pre-treatment to Follow-up Comparisons Among Treatment Groups

The final number of samples included in this within group meta-analysis was 9 (k = 9, N = 285). Articles in Table 2 marked with an "a" were included in the pre-treatment to follow-up comparison. The average follow-up time period was about 15 weeks, ranging from 4 to 24 weeks after the

conclusion of the intervention. All samples included only one follow-up time point.

The cumulative effect size was large (Hedges g = .76, 95% CI = .41 to 1.12; see Table 5). The confidence interval did not contain zero; therefore, the null hypothesis of no treatment effect was rejected. The *Qtotal* was nonsignificant (*Qtotal* (8) = 11.16, p = .19). The  $I^2$  of 28.32% indicated that there was a minimal effect size variation. Figure 4 displays forest plots of individual effect sizes and confidence intervals. Visual inspection of the funnel plot indicated symmetry which suggests publication bias was not present (see Fig. 5). Kendall's Tau was not significant (Tau = .00, p = 1.00). Results from Rosenthal's fail-safe test suggested that there would have to be at least 70 unpublished, nonsignificant comparisons to raise the overall p value to greater than .05, which is considered robust.

#### **Pre-Post-control Comparisons for Control Groups**

The final number of samples included in the within group prepost-control meta-analysis was  $10 \ (k=10, N=266)$ . Table 3 provides a summary of included samples. Five (50%) of the samples were randomized. None of the control groups included an active control treatment. Seven (70%) samples included a wait-list control group, while the remaining three (30%) were inactive (or classes as usual).

A cumulative very small and nonsignificant effect size was observed (Hedges g=.04, 95% CI=-.11 to .20). The confidence interval contained zero; therefore, the null hypothesis cannot be rejected. The *Qtotal* was nonsignificant (*Qtotal* (10) = 5.59, p=.78). The  $I^2$  of 0 indicated there was minimal effect size variation. Figure 6 displays forest plots of included samples. Visual inspection of the funnel plot indicated demonstrated symmetry which suggests an absence of publication bias (see Fig. 7). Kendall's Tau was not significant (Tau=.16, p=.51). Rosenthal's fail-safe did not need to be calculated because the overall effect size was already nonsignificant.

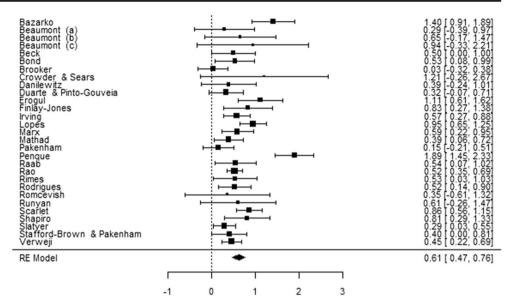
Table 5 Random-effects model meta-analyses summary of results

	N	k	Hedges g	95% CI [lower, upper]	QTotal	<i>I</i> <sup>2</sup> (%)	Rosenthal's fail-safe	Robust fail-safe cutoff
Pre-post-treatment groups	1020	29	.61	[.47, .76]	30.03	6.76	790	155
Pre-post-treatment groups Randomized trials sub-analysis	311	5	.58	[.29, .87]	7.23	44.67	60	35
Pre-post-control groups	266	10	.04	[11, .20]	5.59	0.00	-	-
Pre- to follow-up treatment groups	261	9	.76	[.41, 1.12]	11.16	28.32	70	55
Post-treatment versus post-control	620	11	.48	[.27, .69]	10.84	0.00	145	65
Post-treatment versus post-control Randomized trials sub-analysis	349	6	.58	[.19, .97]	11.44	56.29	42	40

Robust fail-safe cutoff formula: 5(k) + 10; k, number of effect sizes



**Fig. 2** Forest plot of Hedges *g* for pre-post-treatment samples



# **Pre-control to Follow-up Comparisons Among Control Groups**

Only two samples included a follow-up for control groups; therefore, the pre-control to follow-up meta-analysis was not conducted.

# Comparisons of Pre-Post-effect Sizes for the Treatment and Control Groups

The 95% confidence interval for the cumulative prepost-effect size for the treatment groups was .47 to .76. The 95% confidence interval for the cumulative pre-post-effect size for the control groups was -.11 to .20. There was no overlap between the two confidence intervals, indicating that the cumulative effect size for the treatment groups was reliably larger than the cumulative effect size for the control groups.

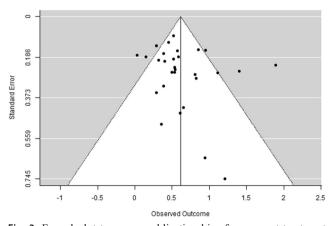


Fig. 3 Funnel plot to assess publication bias for pre-post-treatment samples

Another way to evaluate the differences between the pre-treatment to post-treatment effect sizes for the treatment and control groups is to calculate the average difference between the two effect sizes (Becker 1988). Ten studies yielded pre-post-effect sizes for both the treatment groups and control groups. The average difference (i.e.,  $\Sigma$ (Hedges  $d_{\rm rm}$  for treatment groups – Hedges  $d_{\rm rm}$ for control groups)/k) was .50, SD = 30. This indicated that the pre-treatment to post-treatment intervention effect size remained moderate after controlling for time effects (Morris and DeShon 2002). These differences were also examined using a paired t test which indicated that the average pre-post-effect size for the treatment group (Hedges g = .64) was significantly larger than the average pre-post-effect size for the control group (Hedges g = .09;  $t_{\text{paired}}$  (9) = 5.33, p = .001).

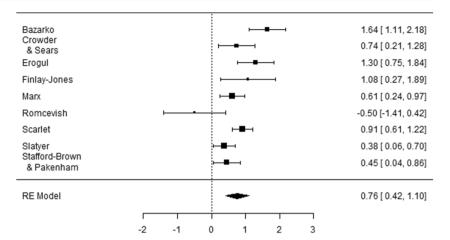
# Comparisons Between Post-treatment and Post-control Groups

After all inclusion and exclusion criteria were verified, the final number of samples included in the between groups post-treatment versus post-control meta-analysis was 11 (k = 11, N = 620). See Table 4 for a summary of the included samples.

There was a moderate cumulative effect size comparing treatment group versus control group self-compassion scores at post-intervention (Hedges g = .48, 95% CI = .27 to .69; see Table 5). The confidence interval did not contain zero; therefore, the null hypothesis of no treatment effect was rejected. The *Qtotal* was nonsignificant (*Qtotal* (11) = 10.84, p = .46). The  $I^2$  of 0% indicated that there was a very little effect size variation. Figure 8 displays a forest plot of individual effect sizes and confidence



**Fig. 4** Forest plot of Hedges *g* for pre-treatment to follow-up samples



intervals. Visual inspection of the funnel plot indicated symmetry suggesting that there is little evidence of publication bias (see Fig. 9). Kendall's Tau was not significant (Tau = .03, p = .89). Results from Rosenthal's fail-safe suggested that there would have to be at least 145 unpublished, nonsignificant comparisons to raise the overall p value to greater than .05, which is considered robust.

### **Moderator Analyses**

All primary analyses indicated that there was no significant heterogeneity in effect sizes (Table 5). However, for exploratory purposes, several possible moderators were conducted for pre-post-treatment overall effect size for the treatment groups. These moderators included participant characteristics (e.g., students vs. professionals; medical field vs. psychological field), intervention characteristics (e.g., manualized interventions vs. modified interventions; 12 or less total interventions hours vs. more 12 total intervention hours), and risk of bias. One moderator was identified as significant, indicating that studies with a retreat had a significantly larger effect size than those without a retreat. No other moderator analyses

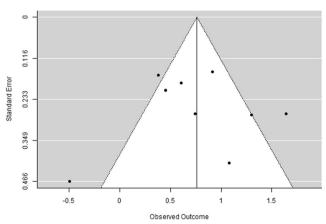


Fig. 5 Funnel plot to assess publication bias for pre-treatment to followup samples



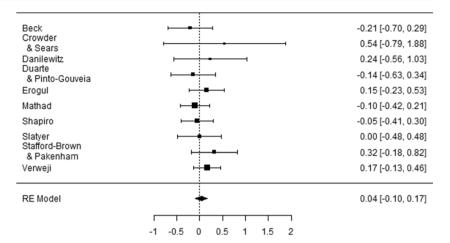
indicated that other methodological variations were associated with different levels of intervention effects (see Table 6). A moderator analysis was also conducted to investigate the effects of randomization for the post-treatment versus post-control comparison and was nonsignificant.

#### **Discussion**

This meta-analysis evaluated the effects of mindfulness-based interventions on self-compassion among health care professionals. An intermediate and reliable effect size was observed for pre-post-differences among treatment groups as well as treatment group versus control group comparisons at posttreatment. A large effect size was observed for pre-treatment to follow-up differences among treatment groups. The prepost effect size for control groups was small, nonsignificant, and significantly smaller than the pre-post-difference for treatment groups. There was no evidence of publication bias among effect sizes; however, there is possible risk of bias within most of the studies. Studies with retreats included as part of intervention demonstrated a significantly larger effect size than those without retreats. No other significant moderators were identified. Finally, Rosenthal's fail-safe values indicated that a large number of nonsignificant comparisons from unpublished studies would need to be stored in researchers' file drawers to disconfirm the meta-analytic findings that mindfulness-based interventions exerted moderate to large and statistically significant improvements in self-compassion.

These findings parallel with meta-analytic results from Kirby et al. (2017) on self-compassion interventions for the general population (d = .70, k = 13, 95% CI = .53–.87, p < .001). Moreover, the current findings suggest that mindfulness-based interventions help improve self-compassion specifically in health care students and professionals. Importantly, the consistency among effect sizes and lack of moderator effects indicates that a range of intervention formats, leader types, number of sessions, total number of

**Fig. 6** Forest plot of Hedges *g* for pre-post-control samples



hours of intervention, and home practice activities yield positive effects on self-compassion. In turn, this information suggests that medical settings and training facilities can have flexibility in the format and implementation of these kinds of programs and interventions for their employees and trainees. The current findings also extend upon Burton et al.'s (2017) interpretation from their meta-analysis that multiple forms of mindfulness-based interventions, not just MBSR, can reduce stress and benefit health care professionals.

The current meta-analytic findings combined with the prior meta-analytic findings suggest that organizations can implement mindfulness-based interventions for health care professionals with an expectation that improvements in self-compassion, stress, depression, and anxiety will be observed (Burton et al. 2017; Dharmawardene et al. 2016). Further, it is reasonable to speculate that improvements in these outcomes may engender better-quality patient care. Finally, some studies suggest that the most significant changes in mindful care occurs when mindfulness training is implemented at an organizational level, generating a compassionate and supportive environment (Barratt 2017; Leonard 2016).

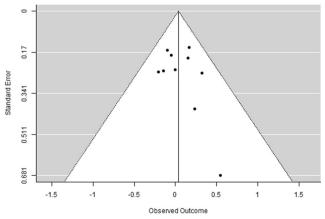


Fig. 7 Funnel plot to assess publication bias for pre-post-control samples

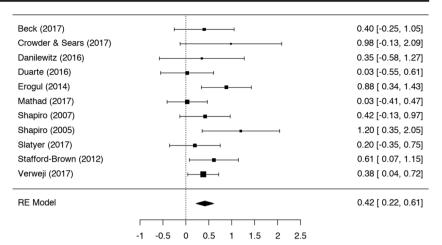
There was a larger effect size for pre-treatment to follow-up comparisons relative to the pre-post-treatment comparisons. This may suggest that the effects of a mindfulness-based intervention are not only maintained but continue to strengthen over time. This is an important finding that should encourage researchers to collect more long-term outcome data when possible. Of note, most of the samples included in the pre-treatment to follow-up analysis had distributed information and resources to participants to continue their practice, and only one offered an optional "booster" session. If institutions are to invest in a mindfulness-based intervention for their employees or students, it would likely be worthwhile for them to include home resources (e.g., audio recordings, home practice plans, handouts, workbooks) to maintain or further develop the benefits gained from the initial intervention.

In regard to the types of interventions utilized, a majority of the studies implemented a version of MBSR. MBSR was the first manualized psychological treatment that incorporated mindfulness. Even though there have been several popular, empirically supported, manualized treatments that have a strong mindfulness component (e.g., MBCT, ACT, DBT), it appears that training and organizational settings have a preference for MBSR. It is possible that institutions prefer a "stress reduction" program, compared with programs that may seem to be designed to address "problems" or psychopathology. Therefore, MBSR may be more acceptable to health care students and professionals. Additionally, MBSR was designed for a medical setting, which may make it more appropriate for these populations.

Health care professionals have significant time constraints, making it difficult for them to commit to weekly mindfulness-based interventions. The question of the appropriate "dosage" of mindfulness interventions is not a new one (Bartlett et al. 2019; Carmody and Baer 2009). There is some support in the literature to suggest that higher dose mindfulness training produces stronger effects



**Fig. 8** Forest plot of Hedges *g* for post-treatment vs. post-control samples



in well-being compared with lower dose training (e.g., single-day training) in a general working population (Chin et al. 2019). The current meta-analysis found that studies that included a retreat component (e.g., day of silence) demonstrated a significantly larger effect size than those without. It may be that the scheduling of a retreat component allowed the participants to experience a higher dose of the mindfulness intervention. At the same time, a retreat element would allow participants to be more disengaged from the everyday stressors and work-related time constraints. Further exploration of the relative advantages of a retreat element for health care professionals is warranted.

It is noteworthy that the included articles date back only to 2005. This is likely due to the relatively recent emergence of the valid and reliable measurement self-compassion (Neff 2003). The timing of the construct overlaps with the recent movement to employ mindfulness-based interventions for health care professionals. Over the past decade, there has been a growth in the number of studies implementing and evaluating the effects of

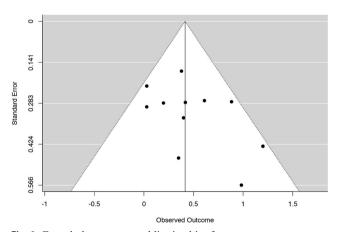


Fig. 9 Funnel plot to assess publication bias for post-treatment vs. post-control samples



interventions for this population (Regehr et al. 2014; West et al. 2016). Based on the growth in the number of these implemented programs and empirical evidence of their effectiveness, it is likely that researchers in this field will continue to publish similar articles. With this assumption, there are a few recommendations for future research based on the findings and limitations of the current study.

#### **Limitations and Future Research**

First, many studies did not meet inclusion criteria for the current meta-analysis due to insufficient data presented. Future studies should report, at the least, means, standard deviations, sample sizes, and pre-post correlations among dependent variables. Second, as more studies are published on this topic, possible outcome moderators may be identified, as the moderator analyses were likely underpowered (Hedges and Pigott 2001). Third, there may be other unmeasured moderators that can help distinguish aspects of interventions that may be associated with effectiveness and thereby better inform clinical practice (e.g., intervention type, duration, content, home practice). Specifically, it would be important for future researchers to explore differences in outcomes after inventions with explicit versus implicit focus on teaching self-compassion. A fourth limitation was the analysis of only self-report measures. It would be important to have a better understanding of whether change in self-compassion corresponds with change in objective indices of self-care and patient care behaviors. A fifth limitation was the broad scope of our inclusion criteria for types of health care professionals and types of interventions. While our purpose was to evaluate the extent to which mindfulness interventions influenced self-compassion in professional caregivers who are at risk of compassion fatigue and burnout, regardless of their job title and years of work, we may have missed important nuances between interventions and specific

 Table 6
 Moderator analyses for pre-post-treatment groups

	k	Hedges $g$	95% CI [lower, upper]	Z	p
Students	14	.50	[.39, .61]	1.00	.32
Professionals	15	.70	[.49, .96]		
Manualized protocol	11	.56	[.39, .72]	0.41	.68
Modified protocol	18	.64	[.43, .85]		
Medical field HCP	15 <sup>a</sup>	.68	[.43, .92]	0.65	.51
Psychological/social field HCP	10	.53	[.38, .67]		
≤ 12 h total intervention time	6	.54	[.33, .75]	0.52	.60
> 12 h total intervention time	18	.65	[.43, .87]		
Home practice	21	.64	[.45, .83]	0.73	.64
No home practice	8	.53	[.36, .69]		
Explicit self-compassion (at minimum, loving-kindness meditation)	16	.44	[.44, .72]	0.33	.75
Implicit self-compassion	13	.65	[.37, .93]		
Retreat	8	.89	[.48, 1.30]	2.04	.04*
No retreat	21	.52	[.39, .64]		
In-person	25	.59	[.43, .74]	0.58	.56
Part or full online <sup>b</sup>	4	.80	[.35, 1.26]		
≤6 sessions	12	.55	[.36, .75]	0.03	.98
>6 sessions	17	.65	[.44, .86]		
<sup>c</sup> Randomized	6	.58	[.19, .97]	0.87	.38
<sup>c</sup> Non-randomized	5	.34	[.09, .60]		
High Jadad quality rating (2–4)	4	.52	[.25, .80]	0.43	.67
Low Jadad quality rating (0–1)	25	.63	[.46, .78]		

<sup>\*</sup>Significant difference; k = number of effect sizes

types of health care providers. Sixth, results from the study quality ratings indicated that many studies did not use a randomized control trial design and did not sufficiently report their methodological procedures. It is recommended that primary researchers refer to standardized reporting guidelines, such as Jadad et al. (1996), to increase methodological quality. Finally, there is a lack of active control group comparisons which limits the external validity of our findings. Moreover, it makes it difficult to determine whether mindfulness-based techniques per se are promoting improvements in self-compassion over and above nonspecific intervention and group variables such as social support, empathy, and problem-solving. The proposed recommendations parallel previous researchers' conclusions, more studies are warranted to better understand the effects of specific intervention components and "dosage," as well as the need for more rigorous study designs (Burton et al. 2017; Dharmawardene et al. 2016).

This meta-analysis indicated that mindfulness-based interventions can promote improvements in self-compassion among a variety of health care professionals. Previous research has shown that health care professionals experience high rates of stress, burnout, and compassion fatigue, which have been correlated with poorer patient satisfaction and more medical errors. Future studies with rigorous methodological designs evaluating the impact of increased self-compassion via mindfulness-based interventions on objective indictors of self-care, quality of patient care, and job performance would be beneficial. Studies evaluating the differences between key variables, such as interventions with implicit

versus explicit compassion skills and levels of training (students vs. employees), are warranted. Additionally, self-compassion as a protective factor for burnout and compassion fatigue in this population should be investigated.

**Author's Contributions** RW designed and executed the study, completed data analyses, and wrote the paper. CB collaborated with the methodological design and editing of the final manuscript. WO collaborated with the methodological design and assisted with data analyses and writing of the study. All authors approved the final version of the manuscript for submission.

#### **Compliance with Ethical Standards**

**Conflict of Interest** The authors declare that they have no conflict of interest.

**Ethical Approval** This article does not contain any studies with human participants or animals performed by any of the authors.

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Articles with an asterisk denotes that it was used in the metaanalysis.

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<sup>&</sup>lt;sup>a</sup> Only 25 effect sizes were utilized because 4 had mixed populations

<sup>&</sup>lt;sup>b</sup> Two studies were all online, while two were a mix of in-person groups with an online supplement

<sup>&</sup>lt;sup>c</sup> Moderator analysis conducted for post-treatment versus post-control comparison

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