



Sensitivity to volatile anesthetics in patients with dementia: a case-control analysis

Sensibilité aux anesthésiques volatils chez des patients atteints de démence: une analyse cas-témoin

Silvia Perez-Protto, MD · Mariya Geube, MD · Daniel Ontaneda, MD · Jarrod E. Dalton, PhD · Andrea Kurz, MD · Daniel I. Sessler, MD

Received: 22 May 2013 / Accepted: 1 April 2014 / Published online: 25 April 2014
© Canadian Anesthesiologists' Society 2014

Abstract

Background Patients with dementia are thought to be more sensitive to anesthesia, although volatile anesthetic requirement has not specifically been evaluated in this population. We tested the hypothesis that patients with dementia having non-cardiac surgery have a lower ratio of bispectral index (BIS) to minimal alveolar concentration (MAC) during the five minutes immediately preceding

incision, thus exhibiting deeper hypnotic levels at a given MAC fraction.

Methods We obtained records from our database registry on patients who had volatile general anesthesia during their most recent operation. We excluded patients premedicated with midazolam. Patients with dementia were identified and their diagnosis was confirmed by chart review. Each patient with dementia was successfully matched with a maximum of five patients without dementia using a multivariate nearest-neighbor distance-matching algorithm restricted to the following criteria: American Society of Anesthesiologists physical status, age, five-minute pre-incision time-weighted average (TWA) estimated effect-site concentration of propofol and fentanyl, and use of remifentanyl. Our primary outcome was the TWA BIS-to-MAC ratio during the five minutes immediately preceding incision.

This article is accompanied by an editorial. Please see Can J Anesth 2014; 61: this issue.

Author contributions Silvia Perez-Protto, Mariya Geube, Daniel Ontaneda, Jarrod E. Dalton, Andrea Kurz, and Daniel I. Sessler were involved in the research study design. Silvia Perez-Protto, Mariya Geube, and Daniel Ontaneda were involved in the data collection. Jarrod E. Dalton was involved in the data analysis. Silvia Perez-Protto, Mariya Geube, Daniel Ontaneda, Jarrod E. Dalton, Andrea Kurz, and Daniel I. Sessler were involved in writing the manuscript.

Results We analyzed 31 patients with dementia matched with 151 patients without dementia. Median [quartiles] TWA BIS-to-MAC ratios for the matched patients were 85 [73, 100] for the patients with dementia and 78 [73, 84] for the patients without dementia. The percent difference in mean BIS-to-MAC ratios between patients with dementia and patients without dementia was 9% (95% confidence interval: -9% to 29%) ($P = 0.35$, Wald test).

S. Perez-Protto, MD · M. Geube, MD
Anesthesiology Institute, Cleveland Clinic Foundation,
Cleveland, OH, USA

Present Address:
M. Geube, MD
Department of Anesthesia, Critical Care and Pain Medicine,
Massachusetts General Hospital, Boston, MA, USA

D. Ontaneda, MD
Neurological Institute, Cleveland Clinic Foundation, Cleveland,
OH, USA

J. E. Dalton, PhD
Department of Quantitative Health Sciences, Cleveland Clinic
Foundation, Cleveland, OH, USA

J. E. Dalton, PhD · A. Kurz, MD · D. I. Sessler, MD (✉)
Department of Outcomes Research, Cleveland Clinic
Foundation, 9500 Euclid Avenue, Cleveland, OH 44195, USA
e-mail: ds@or.org

Conclusions Our results do not support the hypothesis that patients with dementia are more sensitive to volatile anesthetics than patients without dementia.

Résumé

Contexte On estime que les patients atteints de démence sont plus sensibles à l'anesthésie, alors que le besoin en anesthésiques volatils n'a pas été spécifiquement évalué dans cette population. Nous avons testé l'hypothèse selon laquelle les patients atteints de démence subissant une

chirurgie non cardiaque auraient un plus bas rapport d'indice bispectral (BIS) à la concentration alvéolaire minimum (MAC) au cours des cinq minutes précédant immédiatement l'incision, présentant ainsi des niveaux hypnotiques plus profonds pour une fraction MAC donnée.

Méthodes Nous avons obtenu des enregistrements de notre registre de données sur des patients ayant eu une anesthésie générale par des agents volatils au cours de leur plus récente intervention chirurgicale. Nous avons exclu les patients prémédiqués par midazolam. Les patients atteints de démence ont été identifiés et leur diagnostic a été confirmé par l'analyse des dossiers. Chaque patient atteint de démence a été apparié avec succès avec un maximum de cinq patients ne souffrant pas de démence, à l'aide d'un algorithme multifactoriel d'appariements des distances au voisin le plus proche, restreint aux critères suivants: statut physique selon l'American Society of Anesthesiologists, âge, effet estimé/concentration au site moyenne pondérée en fonction du temps (TWA) cinq minutes avant incision pour le propofol et le fentanyl, et utilisation de rémifentanyl. Notre critère d'évaluation principal était le rapport TWA BIS sur MAC au cours des cinq minutes précédant immédiatement l'incision.

Résultats Nous avons analysé 31 patients atteints de démence appariés à 151 patients indemnes de démence. Les ratios TWA BIS sur MAC médians [quartiles] pour les patients appariés ont été 85 [73, 100] pour les patients atteints de démence et 78 [73, 84] pour les patients sans démence. La différence en pourcentage des ratios BIS sur MAC moyens entre les patients atteints de démence et les patients sans démence a été de 9% (intervalle de confiance à 95%: -9% à 29%) ($P = 0.35$, test de Wald).

Conclusions Nos résultats ne confortent pas l'hypothèse selon laquelle les patients atteints de démence sont plus sensibles aux anesthésiques volatils que les patients indemnes de démence.

Life expectancy in the United States has progressively improved; consequently, the proportion of geriatric patients needing surgery is increasing.¹ Anesthesiologists are thus faced with an increasing number of patients with dementia because age is a well-established risk factor for neurodegenerative diseases.^{2,3} Dementia has been preliminarily associated with postoperative complications, and providing anesthesia to this population imposes an important challenge to anesthesiologists.⁴

Patients with dementia are thought to be more sensitive to anesthesia. Although anesthetic requirement has not specifically been evaluated in this population, there are some published opinions suggesting that the dose of

volatile anesthetics be reduced in patients with dementia.^{5,6} Sensitivity to volatile anesthetics is important because both excessive anesthesia⁷ and anesthetic sensitivity⁸ are associated with postoperative mortality.⁹

Accordingly, our goal was to compare volatile anesthetic sensitivity in patients with and without dementia. Specifically, we tested the hypothesis that patients with dementia having non-cardiac surgery have a lower ratio of bispectral index (BIS) to minimal alveolar concentration (MAC) and exhibit deeper hypnotic levels at a given MAC fraction during the five minutes immediately preceding incision.

Methods

We conducted a case-control study with approval from the Cleveland Clinic Foundation Institutional Review Board (IRB) (July 31, 2011). The IRB waived the requirement for informed consent.

Data sources

We obtained records on 77,238 patients who had elective non-cardiac surgery at the Cleveland Clinic Main Campus during January 6, 2005 to September 30, 2010. Prospectively collected data were obtained from the Perioperative Health Documentation System (PHDS) registry. The PHDS contains the entire electronic anesthesia record along with International Classification of Disease (ICD-9) codes, mortality, and a host of other electronic data about every surgical patient at the Cleveland Clinic's Main Campus. We included adults with American Society of Anesthesiologists (ASA) physical status I-III who had volatile general anesthesia during their most recent operations.

We excluded patients premedicated with midazolam because the elderly are intrinsically sensitive to benzodiazepine's sedative effect and can display paradoxical responses.¹⁰ We also excluded patients in whom a full five minutes of anesthetic data immediately preceding incision were unavailable, including minute-to-minute BIS XP™ (A-2000, version 3.21, single disposable 4-Electrode Sensor, Covidien, Dublin, Ireland).

Initially, we used the ICD-9 Clinical Modification (ICD-9-CM) diagnosis codes to identify and match patients with dementia with control patients without dementia. The ICD codes have been validated to identify a diagnosis of dementia with a specificity of 84-85%, a sensitivity of 69%, and a positive predictive value of 81% in a single-center study.¹¹ In an additional recent systematic review,

an overall specificity of 85% was found for ICD diagnoses of dementia/Alzheimer disease.¹²

We used patients' electronic medical records as well as the Knowledge Project Center for Brain Health database at the Cleveland Clinic Foundation. The Knowledge Project Center has developed an information technology program which is integrated with the electronic medical record to combine data entered by both patient and physician.¹³ The diagnosis was based on previous formal testing by neurology staff and classified as Alzheimer's disease, vascular dementia, dementia with Lewy bodies, and senile dementia.¹⁴ Patients whose dementia status could not be confirmed by manual chart review were removed from analysis along with their five matched controls. We similarly deleted any matched control patient from the analysis when manual chart review revealed a clinical diagnosis of dementia.

Our primary outcome was the time-weighted average (TWA) BIS-to-MAC ratio during the five minutes immediately preceding incision. Time-weighted averages are similar to ordinary averages apart from accounting for potential gaps between measurements (specifically, by interpolating observed data pairs and integrating the interpolated profiles). The two are equivalent if measurements are made at equally-spaced time intervals, as was virtually always the case in our electronic records.

Bispectral index was selected as our indicator of anesthetic requirement because it is the best-validated monitor of hypnosis during general anesthesia.¹⁵ We evaluated the five minutes just before skin incision because this period of time is usually stable and avoids the confounding effect of surgical stimulation on the electroencephalogram.¹⁶ Scatterplots and histograms were used to visualize distributions of TWA BIS, TWA MAC, and TWA BIS-to-MAC ratio.

Secondary outcomes were TWA BIS, proportion of the five-minute pre-incision window with BIS < 45, and proportion of the five-minute pre-incision window with hypotension (defined as mean arterial pressure < 60 mmHg). Minimal alveolar concentration-equivalent doses were calculated from end-tidal volatile anesthetic partial pressures using a 1 MAC-equivalent concentration of desflurane (6.6%), sevoflurane (1.8%), and isoflurane (1.17%).¹⁷ Concentrations of plasma fentanyl as well as plasma propofol were estimated using Schneider's pharmacokinetic model.¹⁸

We matched each patient with dementia with a maximum of five patients without dementia using a multivariate nearest-neighbor distance-matching algorithm. Data were randomly sorted before matching to remove any biases attributable to the ordering of observations (e.g., due to date of surgery). Successful matches were restricted to those with common ASA

physical status scores, those with common remifentanyl use (yes or no) during the five minutes preceding incision, and those within one standard deviation of each other according to the following factors (removing control patients whose values lied outside the observed range for the patients with dementia): age, five-minute pre-incision TWA estimated effect-site concentration of fentanyl, and five-minute pre-incision TWA estimated effect-site concentration of propofol. The standard deviations among the patients available for matching were 16.2 yr, 0.3 and 0.3 ng·mL⁻¹, respectively.

Linear mixed-effect regression modelling was used to compare the matched patients on the pre-incision TWA BIS-to-MAC ratio.¹⁹ This model accounts for the grouped nature of the matched sample by assuming that outcomes among patients within a matched group are correlated (and outcomes among patients in differing matched groups are independent). We used the compound symmetric correlation structure in our mixed-effect model. A logarithmic transformation for the outcome was used prior to modelling in order to model percent differences in medians between groups (technically, percent differences in geometric means under an assumed log-normal distribution for the outcome) and to normalize model errors. We adjusted for any of the above patient- and anesthesia-related factors which exhibited imbalance between patients with and without dementia after matching (Table 1).

We similarly modelled the secondary outcome, mean pre-incision BIS. The percentage of pre-incision minutes with BIS < 45, MAP < 60 mmHg, and systolic blood pressure < 90 mmHg were modelled using a generalized linear mixed model with a quasibinomial link function.²⁰ The offset term in these models was the number of observations recorded within the five-minute pre-incision window.

Wald tests were used to assess statistical significance of all model parameters for the primary and secondary hypotheses. The type I error rate was set at 0.05 for both the primary and secondary outcomes, and we used the Bonferroni correction to account for three simultaneous tests for the secondary outcomes (specifically, we used nominally-adjusted confidence levels for the confidence intervals and nominally-adjusted significance criteria for tests).²¹ R statistical software version 2.14.1 (The R Foundation for Statistical Computing, Vienna, Austria) was used for the statistical analysis.

Power considerations

Our study was retrospective in nature. As such, sample size was fixed (but unknown). In anticipation of analyzing data on 30 patients with dementia and 90 matched controls, we

Table 1 Summary of patient- and anesthesia-related characteristics among patients with and without dementia meeting study inclusion/exclusion criteria before and after matching

Factor	Patients Meeting Inclusion/Exclusion Criteria			Matched Patients*		
	Patients without Dementia (n = 3,778)	Patients with Dementia (n = 32)	ASD	Patients without Dementia (n = 151)	Patients with Dementia (n = 31)	ASD
Age (yr), mean (SD)	64 (17)	80 (7)	10.19	79 (7)	80 (7)	0.10
ASA Physical Status (%)						
I	2	0	00.49	0	0	0.01
II	33	15		17	16	
III	65	85		83	84	
Five-minute pre-incision TWA propofol concentration (mg), median [quartiles]	0.21 [0.11, 0.37]	0.13 [0.08, 0.24]	00.40	0.16 [0.07, 0.28]	0.13 [0.08, 0.21]	0.23
Five-minute pre-incision TWA fentanyl concentration (µg), median [quartiles]	0.5 [0.3, 0.7]	0.4 [0.3, 0.6]	00.26	0.4 [0.4, 0.6]	0.4 [0.3, 0.6]	0.03
Remifentanyl used during five-minute pre-incision window (%)	4	3	00.07	3	3	0.00

Absolute standardized difference (ASD) scores, defined as the difference in means, mean rankings, or proportions divided by a pooled estimate of standard deviation, are presented. Absolute standardized difference values greater than 0.10 were considered as indicative of imbalance between groups; we adjusted for any of the displayed variables exhibiting an ASD > 0.10 after matching in all comparisons of primary and secondary outcomes

*One patient initially coded as having dementia was actually not demented. This patient and the corresponding five control patients were removed from the matched sample. Similarly, we removed four matched patients without an ICD-9-CM code for dementia who proved to have a diagnosis of dementia

ASA = American Society of Anesthesiologists; SD = standard deviation; TWA = time-weighted average

conducted a power analysis prior to data collection in order to identify the minimal effect size (ratio of median TWA BIS-to-MAC ratios between patients with and without dementia) for which we had 90% power to detect. This power analysis was performed using PASS software version 11 (NCSS, LLC, Kaysville, UT, USA).

For this analysis, we assumed a coefficient of variation of 1.0 for the primary outcome (based on the observed variability in BIS-to-MAC ratio in the PHDS registry), a lognormal distribution for the outcome, and a type I error rate of 5%. Under this setup, we anticipated approximately 90% power to detect a ratio of median TWA BIS-to-MAC ratios between patients with and without dementia of ≥ 1.78 (i.e., 78% larger BIS-to-MAC ratio in patients with dementia).

Results

Among the 77,238 records extracted from the PHDS database, 20,821 (27.0%) received general anesthesia with sevoflurane or isoflurane. Of these, 16,878 (81.1%) receiving midazolam were removed, and 133 (0.8%) of the remaining 3,943 patients had unavailable anesthetic data within the five minutes immediately preceding incision. Thus, data on 3,810 patients (32 with an ICD-9-CM code for dementia and 3,778 without) were available

for matching. All 32 patients with an ICD-9-CM code for dementia were successfully matched to five patients without such a code.

Upon reviewing the charts and the Knowledge Project database, we found that one patient initially coded as having dementia was actually not demented; consequently, this patient and the corresponding five control patients were removed from the matched sample. Similarly, we removed four matched patients without an ICD-9-CM code for dementia who proved to have a diagnosis of dementia. To protect against hidden bias, these patients were not included in the dementia group as they did not meet the pre-specified criteria for study inclusion. Our final matched sample thus included 31 patients with dementia and 151 patients without dementia. Among the 31 matched patients with dementia, 26 (84%) had Alzheimer's disease, 2 (6.5%) had dementia with Lewy bodies, 2 (6.5%) had vascular dementia, and 1 (3.2%) had senile dementia.

Median [quartiles] times from induction to incision were 34 [26, 47] min for patients without dementia and 35 [29, 46] min for patients with dementia. All patients received a non-depolarizing muscle relaxant for tracheal intubation and their lungs were mechanically ventilated.

Median [quartiles] TWA BIS absolute values for the matched patients during the five minutes immediately preceding incision were 46.0 [41.6, 53.1] for the patients without dementia and 46.6 [38.6, 55.2] for the patients with

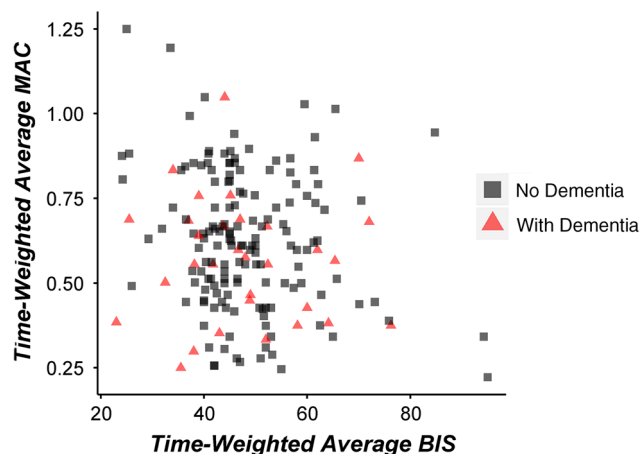


Fig. 1 Scatterplot of time-weighted average minimum alveolar concentration (MAC) vs time-weighted average bispectral index (BIS) during the five minutes immediately preceding incision. Patients with dementia are represented as triangles and matched control patients are represented as squares

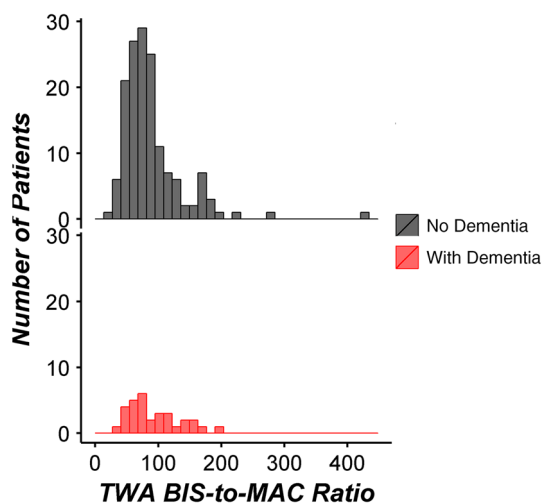


Fig. 2 Histogram of the ratio of time-weighted average bispectral index (BIS) to time-weighted average minimum alveolar concentration (MAC), separately for patients with dementia and matched control patients

dementia. The TWA MAC absolute values at the same time frame were 0.62 [0.49, 0.79] for the patients without dementia and 0.57 [0.41, 0.68] for the patients with dementia (Fig. 1). The TWA BIS-to-MAC ratios were 75 [61, 99] for the patients without dementia and 78 [62, 116] for the patients with dementia (Fig. 2).

The percent difference [95% confidence interval (CI)] in median BIS-to-MAC ratios between patients with dementia and patients without dementia was 9% (95% CI: -9 to 29), which was not statistically significant ($P = 0.38$, Wald test). This univariable result did not control for age and five-minute pre-incision propofol concentration, which

continued to show slight imbalance after matching (Table 1). Adjusting for these factors, the estimates were largely unchanged (difference, 9%; 95% CI: -9 to 29; $P = 0.35$).

Secondary outcomes did not differ significantly between patients with dementia and matched patients without dementia. Table 2 summarizes adjusted means for each group and the comparison between groups for all outcomes.

Discussion

By virtue of our multivariate distance matching procedure, our patients with dementia and patients without dementia were well matched on potential confounding factors and anesthetic management. Our results showed a small and non-significant increase in BIS-to-MAC ratio in patients with dementia. We therefore conclude that patients with and without dementia are similarly sensitive to volatile anesthetics. This result is surprising given the widespread view that dementia increases sensitivity to anesthesia.^{5,22,23} On the other hand, the result is consistent with recent work showing comparable intravenous anesthesia requirements in patients with cognitive impairment.²⁴ In contrast, there are two studies showing that a murine model displaying histopathologic and clinical features of Alzheimer's disease has an increased resistance to inhalational anesthetics when compared with their wild-type age-matched controls.^{25,26}

The classical definition of inhaled anesthetic potency is MAC (the alveolar concentration needed to prevent movement in 50% of subjects in response to skin incision).^{27,28} Minimum alveolar concentration immobility is primarily a spinal reflex and would not be expected to change much with dementia.²⁹⁻³² When anesthetic agents are delivered preferentially to the brain excluding spinal cord perfusion, it has been found that MAC immobility increases, emphasizing the critical role of the spinal cord as a target of the inhalational anesthetic action and primary site of motor response to noxious stimuli under general anesthesia.³³ One of the advantages of MAC is its small inter-individual variance. We did not use age-adjusted MAC (MAC years) because the patients were about the same age in both groups. The same linear adjustment would apply to both groups, which would not alter the relative difference or statistical significance.

Bispectral index is a near-instantaneous measure of cortical function.³⁴ Although BIS is the best-established measure of hypnotic depth,^{15,35,36} it is surely imperfect,⁹ mainly due to its intrinsic variability. A recent publication proposed a composite variability index comprising BIS and electromyography variability to predict the occurrence of

Table 2 Results of primary and secondary comparisons between matched patients with and without dementia (denoted by 'D' and 'ND', respectively)

	Adjusted Mean (95% CI)		Comparison (D vs ND)	Estimate (95% CI)	P value
	Patients without Dementia (ND) (n = 151)	Patients with Dementia (D) (n = 31)			
Primary Outcome:					
TWA BIS-to-MAC Ratio	78 (73 to 84)	85 (73 to 100)	% Difference in Means (95% CI)	9% (-9 to 29)	0.35
Secondary Outcomes:					
TWA BIS	48 (46 to 51)	48 (43 to 53)	Difference in Means (95% CI)	-0.28 (-5.72 to 5.17)	0.90
Proportion of five-minute pre-incision window with BIS < 45	0.41 (0.33 to 0.50)	0.41 (0.24 to 0.60)	Odds Ratio (95% CI)	0.97 (0.41 to 2.32)	0.94
Proportion of five-minute pre-incision window with hypotension (*)	0.15 (0.10 to 0.21)	0.20 (0.09, 0.37)	Odds Ratio (95% CI)	1.42 (0.52 to 3.88)	0.41

All results adjusted for age and five-minute pre-incision time-weighted average propofol concentration

(*) Hypotension was defined as SBP < 90 mmHg or MAP < 60 mmHg

BIS = bispectral index; CI = confidence interval; MAC = minimal alveolar concentration; MAP = mean arterial pressure; SBP = systolic blood pressure; TWA = time-weighted average

somatic events during general anesthesia. This index was found to predict the occurrence of somatic events and to be more reliable with respect to the variability of the mean heart rate and blood pressure.³⁷ Patients with dementia in the awake state show a significantly lower baseline BIS value on average than age-matched controls.^{24,38} Whether it is accompanied by a comparable reduction in BIS during inhalational anesthesia and surgery remains unanswered. Anesthesiologists should, as always, titrate volatile anesthesia based on individual responses; but our data suggest that patients with dementia will not have a lesser requirement than other patients.

We excluded patients who received midazolam for induction of anesthesia because the elderly are intrinsically sensitive to benzodiazepine's sedative effect and can display a paradoxical effect.¹⁰ Furthermore, benzodiazepines reduce MAC, which would have compromised our primary outcome.^{39,40} Eliminating patients given midazolam markedly reduced the number of patients with dementia available for analysis; however, it presumably improved the reliability of our estimates of the true physiologic effect of dementia on anesthetic requirement. Anesthesia was induced with propofol in virtually all cases. The estimated propofol plasma concentrations at the time anesthetic sensitivity was assessed were similar in the patients with and without dementia and low, i.e., $\approx 0.15 \mu\text{g}\cdot\text{mL}^{-1}$, which is less than a tenth of the effective dose of 50% for loss of consciousness (CP₅₀).^{41,42} Estimated fentanyl plasma concentrations were also similar and low. Fentanyl reduces MAC by 59% at a plasma concentration of $3 \text{ ng}\cdot\text{mL}^{-1}$, which is ten times the calculated concentration in our population

($\approx 0.4 \text{ ng}\cdot\text{mL}^{-1}$).⁴³ Based on the estimated plasma concentrations, neither drug is likely to have had any substantial effect on MAC.

Our analysis was restricted to the five-minute period just before incision to eliminate the likely influence of the surgical stimulation on BIS. For example, studies reported nociceptive stimuli to have a wide variety of effects on brain electrical activity, either increasing⁴⁴ or decreasing alpha frequency.⁴⁵ Specifically, the effect of skin incision on the BIS during general anesthesia is an overall tendency for it to become more continuous with significant loss of episodic activity. The changes in BIS after surgical incision depend on the anesthetic technique, with less correlation when opioid analgesics are given.⁴⁶ But even moderately deep anesthesia does not consistently obtund BIS response.¹⁶

The major limitation of our study is that we only had power to identify a relatively large difference of 78% or more in the BIS-to-MAC ratio. Even so, there was not even a trend towards reduced anesthetic sensitivity in the patients with dementia. Low power resulted from the small number of patients with dementia that we were able to identify after applying strict exclusion criteria. Another limitation of this study was using the ICD-9 code for dementia and subsequent chart review as a way to identify the most severe cases of dementia. There may be difficulty in completely uncoupling the relationship between increasing age and cognitive dysfunction in older patients due to the strong correlation that exists between them. Nevertheless, by taking only patients with a formal diagnosis of dementia in the medical record, we were able to identify the patients who had more severe symptoms of dementia. The possibility of undiagnosed

dementia in the controls could not be completely excluded, but these cases were probably mild as they had not yet been clinically diagnosed. Our sample was optimized to collect only the most severe cases of dementia and controls where the likelihood of dementia was low and, if present, mild in severity.

The relatively low proportion of patients with dementia from our database also deserves comment. There are several likely explanations for this finding. Probably the most important is selecting patients without dementia to be eligible for surgery. A second possibility is the presence of undiagnosed patients with mild dementia in the control group. Although this is possible, all elderly patients scheduled for surgery at our institution have a rigorous preoperative evaluation by both internal medicine and anesthesiology. If patients were erroneously identified as not having dementia based on the initial PHDS documentation, it is likely that these patients had mild dementia at most. A subsequent chart review of the detailed medical record failed to disclose a diagnosis of dementia in the control group. A consequence is that prevalence of dementia in our database should not be estimated from our sample.

In summary, our results do not support the hypothesis that patients with dementia are more sensitive to volatile anesthetics than patients without dementia during maintenance of anesthesia. Volatile anesthetics do not appear to be appreciably different based on the limits of the 95% CI for the ratio of means.

Competing interests None declared.

References

1. National Center for Health Statistics. Discharges with at least one procedure in nonfederal short-stay hospitals, by sex, age, and selected procedures: United States, selected years 1990 through 2008-2009. Health, United States, 2011 - With special feature on socioeconomic status and health. Hyattsville, MD: National Center for Health Statistics; 2012. Available from URL: www.cdc.gov/nchs/data/abus/2010/103.pdf (accessed February 2014).
2. Qiu C, De Ronchi D, Fratiglioni L. The epidemiology of the dementias: an update. *Curr Opin Psychiatry* 2007; 20: 380-5.
3. Matthews FE, Arthur A, Barnes LE, et al. A two-decade comparison of prevalence of dementia in individuals aged 65 years and older from three geographical areas of England: results of the Cognitive Function and Ageing Study I and II. *Lancet* 2013; 382: 1405-12.
4. Hu CJ, Liao CC, Chang CC, Wu CH, Chen TL. Postoperative adverse outcomes in surgical patients with dementia: a retrospective cohort study. *World J Surg* 2012; 36: 2051-8.
5. Funder KS, Steinmetz J, Rasmussen LS. Anaesthesia for the patient with dementia undergoing outpatient surgery. *Curr Opin Anaesthesiol* 2009; 22: 712-7.
6. Di Nino G, Adversi M, Dekel BG, Fodale V, Rosa G, Melotti RM. Peri-operative risk management in patients with Alzheimer's disease. *J Alzheimers Dis* 2010; 22(Suppl 3): 121-7.
7. Monk TG, Saini V, Weldon BC, Sigl JC. Anesthetic management and one-year mortality after noncardiac surgery. *Anesth Analg* 2005; 100: 4-10.
8. Sessler DI, Sigl JC, Kelley SD, et al. Hospital stay and mortality are increased in patients having a "triple low" of low blood pressure, low bispectral index, and low minimum alveolar concentration of volatile anesthesia. *Anesthesiology* 2012; 116: 1195-203.
9. Leslie K, Myles PS, Forbes A, Chan MT. The effect of bispectral index monitoring on long-term survival in the B-aware trial. *Anesth Analg* 2010; 110: 816-22.
10. Stoelting RK, Hillier SC. *Pharmacology & Physiology in Anesthetic Practice*. 4th ed. Philadelphia: Lippincott Williams & Wilkins; 2006 .
11. Bharmal MF, Weiner M, Sands LP, Xu H, Craig BA, Thomas J 3rd. Impact of patient selection criteria on prevalence estimates and prevalence of diagnosed dementia in a Medicaid population. *Alzheimer Dis Assoc Disord* 2007; 21: 92-100.
12. St Germaine-Smith C, Metcalfe A, Pringsheim T, et al. Recommendations for optimal ICD codes to study neurologic conditions: a systematic review. *Neurology* 2012; 79: 1049-55.
13. Katzan I, Speck M, Dopler C, et al. The Knowledge Program: an innovative, comprehensive electronic data capture system and warehouse. *AMIA Annu Symp Proc* 2011; 2011: 683-92.
14. Bolla LR, Filley CM, Palmer RM. Dementia DDx. Office diagnosis of the four major types of dementia. *Geriatrics* 2000; 55: 34-7, 41-2, 45-6.
15. Glass PS, Bloom M, Kears L, Rosow C, Sebel P, Manberg P. Bispectral analysis measures sedation and memory effects of propofol, midazolam, isoflurane, and alfentanil in healthy volunteers. *Anesthesiology* 1997; 86: 836-47.
16. Sleight JW, Leslie K, Voss L. The effect of skin incision on the electroencephalogram during general anesthesia maintained with propofol or desflurane. *J Clin Monit Comput* 2010; 24: 307-18.
17. Nickalls RW, Mapleson WW. Age-related iso-MAC charts for isoflurane, sevoflurane and desflurane in man. *Br J Anaesth* 2003; 91: 170-4.
18. Schnider TW, Minto CF, Gambus PL, et al. The influence of method of administration and covariates on the pharmacokinetics of propofol in adult volunteers. *Anesthesiology* 1998; 88: 1170-82.
19. Pinheiro JC, Bates DM. *Mixed-Effects Models in S and S-PLUS*. New York: Springer; 2000 .
20. Faraway JJ. *Extending the Linear Model with R: Generalized Linear, Mixed Effects and Nonparametric Regression Models*. Boca Raton: Chapman & Hall/CRC; 2006 .
21. Bonferroni CE. Il calcolo delle assicurazioni su gruppi di teste. *Studi in Onore del Professore Salvatore Ortu Carboni* 1935: 13-60.
22. Funder KS, Steinmetz J, Rasmussen LS. Anesthesia for the patient with dementia. *J Alzheimers Dis* 2010; 22(Suppl 3): 129-34.
23. Burton DA, Nicholson G, Hall GM. Anaesthesia in elderly patients with neurodegenerative disorders: special considerations. *Drugs Aging* 2004; 21: 229-42.
24. Erdogan MA, Demirbilek S, Erdil F, et al. The effects of cognitive impairment on anaesthetic requirement in the elderly. *Eur J Anaesthesiol* 2012; 29: 326-31.
25. Bianchi SL, Caltagaroni BM, Laferla FM, Eckenhoff RG, Kelz MB. Inhaled anesthetic potency in aged Alzheimer mice. *Anesth Analg* 2010; 110: 427-30.
26. Eckel B, Richtsfeld M, Starker L, Blobner M. Transgenic Alzheimer mice have a larger minimum alveolar anesthetic concentration of isoflurane than their nontransgenic littermates. *Anesth Analg* 2010; 110: 438-41.
27. Merkel G, Eger EI 2nd. A comparative study of halothane and halopropane anesthesia including method for determining equipotency. *Anesthesiology* 1963; 24: 346-57.

28. *Eger EI 2nd*. A brief history of the origin of minimum alveolar concentration (MAC). *Anesthesiology* 2002; 96: 238-9.
29. *Rampil IJ*. Anesthetic potency is not altered after hypothermic spinal cord transection in rats. *Anesthesiology* 1994; 80: 606-10.
30. *Rampil IJ, Mason P, Singh H*. Anesthetic potency (MAC) is independent of forebrain structures in the rat. *Anesthesiology* 1993; 78: 707-12.
31. *Antognini JF, Carstens E, Sudo M, Sudo S*. Isoflurane depresses electroencephalographic and medial thalamic responses to noxious stimulation via an indirect spinal action. *Anesth Analg* 2000; 91: 1282-8.
32. *Antognini JF, Wang XW, Carstens E*. Isoflurane action in the spinal cord blunts electroencephalographic and thalamic-reticular formation responses to noxious stimulation in goats. *Anesthesiology* 2000; 92: 559-66.
33. *Antognini JF, Schwartz K*. Exaggerated anesthetic requirements in the preferentially anesthetized brain. *Anesthesiology* 1993; 79: 1244-9.
34. *Katoh T, Suzuki A, Ikeda K*. Electroencephalographic derivatives as a tool for predicting the depth of sedation and anesthesia induced by sevoflurane. *Anesthesiology* 1998; 88: 642-50.
35. *Morimoto Y, Hagihira S, Koizumi Y, Ishida K, Matsumoto M, Sakabe T*. The relationship between bispectral index and electroencephalographic parameters during isoflurane anesthesia. *Anesth Analg* 2004; 98: 1336-40.
36. *Schraag S, Bothner U, Gajraj R, Kenny GN, Georgieff M*. The performance of electroencephalogram bispectral index and auditory evoked potential index to predict loss of consciousness during propofol infusion. *Anesth Analg* 1999; 89: 1311-5.
37. *Mathews DM, Clark L, Johansen J, Matute E, Seshagiri CV*. Increases in electroencephalogram and electromyogram variability are associated with an increased incidence of intraoperative somatic response. *Anesth Analg* 2012; 114: 759-70.
38. *Renna M, Handy J, Shah A*. Low baseline bispectral index of the electroencephalogram in patients with dementia. *Anesth Analg* 2003; 96: 1380-5.
39. *Quasha AL, Eger EI 2nd, Tinker JH*. Determination and applications of MAC. *Anesthesiology* 1980; 53: 315-34.
40. *Gyulai FE, Mintun MA, Firestone LL*. Dose-dependent enhancement of in vivo GABA(A)-benzodiazepine receptor binding by isoflurane. *Anesthesiology* 2001; 95: 585-93.
41. *Smith C, McEwan AI, Jhaveri R, et al*. The interaction of fentanyl on the Cp50 of propofol for loss of consciousness and skin incision. *Anesthesiology* 1994; 81: 820-8.
42. *Kazama T, Ikeda K, Morita K*. Reduction by fentanyl of the Cp50 values of propofol and hemodynamic responses to various noxious stimuli. *Anesthesiology* 1997; 87: 213-27.
43. *Katoh T, Ikeda K*. The effects of fentanyl on sevoflurane requirements for loss of consciousness and skin incision. *Anesthesiology* 1998; 88: 18-24.
44. *Kox WJ, von Heymann C, Heinze J, Prichep LS, John ER, Rundshagen I*. Electroencephalographic mapping during routine clinical practice: cortical arousal during tracheal intubation? *Anesth Analg* 2006; 102: 825-31.
45. *Kochs E, Bischoff P, Pichlmeier U, Schulte am Esch J*. Surgical stimulation induces changes in brain electrical activity during isoflurane/nitrous oxide anesthesia. A topographic electroencephalographic analysis. *Anesthesiology* 1994; 80: 1026-34.
46. *Sebel PS, Lang E, Rampil IJ, et al*. A multicenter study of bispectral electroencephalogram analysis for monitoring anesthetic effect. *Anesth Analg* 1997; 84: 891-9.