


Long-Term Effectiveness of the TIME Intervention to Improve Diabetes Outcomes in Low-Income Settings: a 2-Year Follow-Up



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BACKGROUND: We previously found that a 6-month multidimensional diabetes program, TIME (Telehealth-Supported, Integrated Community Health Workers, Medication-Access) resulted in improved clinical outcomes.

OBJECTIVE: To follow TIME participant clinical outcomes for 24 months

PARTICIPANTS: Low-income Latino(a)s with type 2 diabetes

DESIGN AND INTERVENTION: We collected post-intervention clinical data for five cohorts ($n = 101$, mean $n = 20$ /cohort) who participated in TIME programs from 2018 to 2020 in Houston, Texas.

MAIN MEASURES: We gathered HbA1c (primary outcome), weight, body mass index (BMI), and blood pressure data at baseline, 6 months (intervention end), and semiannually thereafter until 24 months after baseline to assess sustainability. We also evaluated participant loss to follow-up until 24 months.

KEY RESULTS: Participants decreased HbA1c levels during the intervention ($p < 0.0001$) and maintained these improvements at each timepoint from baseline to 24 months (p range: < 0.0001 to 0.015). Participants reduced blood pressure levels during TIME and maintained these changes at each timepoint from baseline until 18 months (systolic p range: < 0.0001 to 0.0005 , diastolic p range: < 0.0001 to 0.008) but not at 24 months (systolic: $p = 0.065$; diastolic: $p = 0.85$). There were no significant weight changes during TIME or post-intervention: weight (p range = 0.07 to 0.77), BMI (p range = 0.11 to 0.71). Attrition rates (loss to follow-up during the post-intervention period) were 5.9% (6 months), 24.8% (12 months), 35.6% (18 months), and 41.8% (24 months).

CONCLUSIONS: It is possible for vulnerable populations to maintain long-term glycemic and blood pressure improvements using a multiple dimensional intervention. Attrition rates rose over time but show promise given the

majority of post-intervention timepoints occurred during the COVID-19 pandemic when low-income populations were most susceptible to suboptimal healthcare access. Future studies are needed to evaluate longitudinal outcomes of diabetes interventions conducted by local clinics rather than research teams.

KEY WORDS: diabetes; telehealth or telemedicine; community health worker (CHW); group visits or shared medical appointments; access to care; disparities; low-income.

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INTRODUCTION

There are 34.2 million people with diabetes in the USA, constituting 13.0% of the adult population and including an estimated 7.3 million individuals who are unaware of their disease.¹ Individuals with diabetes incur more than twice the medical expenditures compared to those without, including \$90 billion for lost productivity. Diabetes care costs \$327 billion annually, consuming 25% of US healthcare dollars.²

Individuals with diabetes from low socioeconomic status face numerous barriers to care. These include medication access, transportation, language, healthy food availability, health literacy, and insurance.^{3,4} Latino(a)s are twice as likely to develop diabetes than non-Hispanic whites and suffer from more complications, hyperglycemia, and poor quality of life.^{1,5} Numerous targeted diabetes intervention studies have demonstrated improved outcomes among low-income minorities.^{6–8} However, a key gap is follow-up data to evaluate long-term maintenance of the outcomes achieved during active interventions.

We previously evaluated the effects of combining four individually validated strategies (telehealth, community health workers (CHWs), medication access, group visit education) in

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the *TIME* programs conducted at community clinics serving low-income Latino(a)s.^{9–11} *TIME* included a 6-month intervention followed by 6 months of CHW follow-up to bridge care. It was the first study to our knowledge to integrate CHWs into group visit leadership teams as well as utilize telehealth to provide ongoing CHW support. During the intervention, participants significantly improved diabetes outcomes compared to the control arm including lower HbA1c (−1.43% vs. −0.45%, $p = 0.002$), improved blood pressure levels (systolic: −6.89 mmHg vs. 0.003 mmHg, $p = 0.023$; diastolic: −3.36 mmHg vs. 0.2 mmHg, $p = 0.046$), and greater American Diabetes Association guideline adherence ($p < 0.001$) from baseline to 6 months, respectively, but did not result in significant body mass index (BMI) or weight reductions. The program also showed the value of CHWs in identifying barriers to care that may have otherwise been unrecognized.^{9–11} The clinical outcomes achieved during *TIME* are important, but it is not known if these clinical outcomes would be sustainable over time or worsen as would be expected without an intervention.^{12–15}

In this study, we evaluated post-intervention data of *TIME* participants. Specifically, we recorded semiannual HbA1c (primary outcome), blood pressure, weight, and BMI data from baseline to 24 months to assess sustainability. We hypothesized that the significant clinical improvements achieved during the *TIME* programs, in particular HbA1c and blood pressure reduction, would be sustained during the post-intervention period.

METHODS

Study Design and Setting

This was an observational follow-up study of five cohorts who participated in the *TIME* program from January 2018 to February 2020 in Houston, Texas. The five cohorts included the two original intervention arms ($n = 22$, each)^{9–11} and three cohorts ($n = 15$, $n = 22$, $n = 20$) in subsequent programs. The methodology of the *TIME* program has been described previously^{9–11}; sites were nonprofit, 501c(3) community clinics that served only uninsured and low-income individuals earning $\leq 150\%$ of the federal poverty level. Individuals who were insured or earned greater than the allotted income were not eligible for clinic services. Study inclusion criteria were adults ≥ 18 years, Latino(a)s, type 2 diabetes (HbA1c $\geq 6.5\%$), and Spanish-speaking. Individuals were excluded if they did not attend at least one group visit ($n = 0$); were pregnant ($n = 1$); had severe cognitive impairment ($n = 0$); required frequent provider encounters, e.g., for labile glucose levels ($n = 1$); or had a treatment or condition that could alter reliable HbA1c measurements ($n = 2$).¹⁶ Recruitment processes included contacting potential participants identified by a clinic database or provider referrals, inviting interested individuals to an orientation for written consent, and baseline data collection, i.e.,

HbA1c. This study was approved by the Baylor College of Medicine Institutional Review Board.

Intervention: Months 1–6. The *TIME* intervention is described elsewhere.^{9–11} It consisted of 3-h, CHW-led monthly group visits for 6 months. Group visits included large-group diabetes education, small-group breakout sessions that addressed barriers to care, healthy meal education with food examples, and a 1:1 medical encounter with a provider. Telehealth (video conferencing, mobile health (mHealth) phone/text) was utilized to enhance CHW-participant and CHW-provider communication.^{17,18} Providers met with CHWs weekly via video conferencing for training and support; CHWs contacted participants via mHealth weekly to reiterate educational topics from group visits, obtain glucose values, ensure medication adherence, and answer questions.^{9–11}

Post-Intervention: Months 7–24. Starting in month seven, participants returned to usual care in the clinic with their primary care providers in lieu of receiving care from the study physicians. During months 7–12, the CHWs continued to contact participants via mHealth to bridge care to the clinic. During months 13–24, participants continued to receive usual care in the clinic but did not receive CHW communication.

Measures

Clinical Outcomes. *TIME* baseline occurred at month one, and *TIME* ended at month six. Post-intervention timepoints were at 12, 18, and 24 months after baseline. We assessed clinical sustainability using electronic medical records (EMRs) (EPIC, Athena Health) and collecting quarterly HbA1c (primary outcome), weight, BMI, and systolic and diastolic blood pressure values from baseline to 24 months. Data at each timepoint was the mean ± 3 months, e.g., the 12-month value was the average of 9, 12, and 15 months. We chose this method to more accurately capture trends due to the high levels of between-visit clinical variability in low-income areas.^{12,19}

Medication Access. Medication access and adherence to treatment are often limited by cost, education, and access to care in low-income settings, resulting in glucose variability.¹⁹ Therefore, we recorded whether participants received ≥ 1 high-cost medication ($> \$10/\text{month}$) during and after *TIME*. We also documented if individuals received insulin during the intervention and post-intervention.²⁰

CHW-Participant mHealth Outcomes. We evaluated the frequency of successful CHW-participant mHealth contact, defined as communicating with participants via phone or text message.

Attrition Outcomes. We also assessed participant loss to follow-up (attrition), defined as not receiving treatment by a provider for diabetes, i.e., primary care appointment at each timepoint. We excluded visits that were not diabetes-related or not with a primary care provider (e.g., mammogram or nutritionist appointments) in order to capture the ability to collect long-term diabetes outcomes in this population.

Statistical Analysis

We examined change over time in each clinical outcome: HbA1c, weight, BMI, and blood pressure (systolic and diastolic). For each outcome, we constructed two growth curve models. We constructed a linear multilevel growth curve model with the outcome as the dependent variable, and the baseline variable of the outcome and time as independent variables. Time was coded as such: baseline = 0, 6 months = 1, 12 months = 2, 18 months = 3, and 24 months = 4. Time was included as both a fixed and a random effect, and an unstructured covariance matrix was specified. For each outcome, we conducted a quadratic multilevel growth curve model. The only difference was the addition of a time \times time interaction as a predictor. For HbA1c, we examined whether change in values over time varied as a function of whether it was controlled at 6 months ($< 7\%$ for < 65 years, $< 7.5\%$ for > 65 years). We repeated the linear model adding the time \times control interaction term and repeated the quadratic model, adding the time \times time \times control interaction term.

We then used repeated-measures *t* tests to examine improvement in each of the five outcomes during the active period (i.e., baseline to 6 months). We repeated these tests to examine maintenance of effects from baseline to each follow-up period. For each outcome, three additional repeated-measures *t* tests were conducted to examine maintenance from baseline to 12, 18, and 24 months. If HbA1c levels were not available but glucose levels were recorded (1% of timepoints), we converted these values to HbA1c using the equation $(46.7 + (\text{mean glucose})/26.7)^{21,22}$. To account for potential variations between glucose and HbA1c data, we recorded the mean glucose values if more than one was available and examined whether any proxy glucose values were outliers (i.e., exceeded two standard deviations beyond individuals' mean HbA1c data). Missing HbA1c data ($n = 101$ except 24 months ($n = 79$)) were baseline (0%), 6 months (3.96%), 12 months (25.74%), 18 months (36.64%), and 24 months (45.57%). A total of 21.33% (103/483) of all data points were missing.

We conducted a series of independent-samples *t* tests to determine if mean age or baseline clinical scores differed between those missing data and those with complete data. All models were repeated using intent-to-treat (ITT) methods and employed the multiple imputation procedures PROC MI and MI ANALYZE in SAS Version 9.4 (SAS Institute, Cary, NC) to handle missing data. Analyses were reported using

paired *t* test with multiple imputations if missing data were present. For completeness, analyses were also conducted using paired *t* test without multiple imputations. Binomial data (CHW contact, participant attrition, medication data) were recorded as yes/no and proportions were evaluated.

RESULTS

Study data reflects the combined five cohorts' outcomes. Table 1 illustrates baseline characteristics of individuals ($n = 101$, mean 20.2/cohort) who participated in TIME programs. Mean participant age was 54.8 years, more females (67.3%) participated, and nearly half (45.5%) did not have a legal status. Most individuals worked in domestic (44.6%) or manual labor (22.8%) employments and received oral hypoglycemics alone (60.4%) or oral(s) with injectable(s) (36.6%). Individuals averaged 12.3 years since diabetes diagnosis. Mean baseline clinical levels were blood pressure 132.7/76.4 mmHg, HbA1c 8.56%, BMI 33.4 kg/m², total cholesterol 180.4 mg/dL, triglycerides 180.9 mg/dL, HDL-C 48.9 mg/dL, and LDL-C 99.3 mg/dL. Those missing any data did not differ from those with complete data on age or any baseline clinical scores (all $p > 0.05$).

Medication Access and Attrition Outcomes

Clinic and medication eligibility concerns were the most common participant issue during the post-intervention period. The

Table 1 Baseline Characteristics of Low-Income Latino(a)s with Type 2 Diabetes Who Participated in TIME Programs ($n = 101$)

Variable	Mean \pm SD or n(%)
Age (years)	54.8 \pm 8.4
Sex (<i>n</i>)	
Male	33 (32.7)
Female	68 (67.3)
Legal status (<i>n</i>)	
Documented	56 (55.5)
Employment (<i>n</i>)	
Domestic	45 (44.6)
Manual labor	23 (22.8)
Food service	12 (11.9)
Unemployed	14 (13.9)
Other/unknown	7 (6.8)
Diabetes therapy (<i>n</i>)	
Lifestyle only	3 (3.0)
Oral hypoglycemic(s) only	61 (60.4)
Oral hypoglycemic(s) with injectable(s)	37 (36.6)
Injectable(s) only	0 (0.0)
Time since diabetes diagnosis (years)	12.3 \pm 8.7
HbA1c (%)	8.56 \pm 2.1
Weight (lb)	183.12 (46.43)
Body mass index (kg/m ²)	33.4 \pm 21.1
Blood pressure (mmHg)	
Systolic	132.7 \pm 15.1
Diastolic	76.4 \pm 9.8
Cholesterol (mg/dL)	
Total	180.4 \pm 43.7
Triglycerides	180.9 \pm 89.5
HDL-C	48.9 \pm 14.5
LDL-C	99.3 \pm 37.4

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proportion of individuals receiving high-cost medications and insulin decreased during TIME but increased in the post-intervention period when they returned to clinic. A total of 59.4% received high-cost medications at baseline. This proportion decreased to 53.5% during the intervention but increased to 62.8% during the post-intervention period. Similarly, 33.7%, 26.7%, and 31.4% received insulin at baseline, 6 months, and post-intervention, respectively (Fig. 1). The differences in high-cost medications and insulin use at baseline, 6 months, and during the post-intervention period were not significant: high-cost medications (baseline to 6 months: $p = 0.397$, 6 to 24 months: $p = 0.113$), insulin (baseline to 6 months: $p = 0.307$, 6 to 24 months: $p = 0.513$).

CHW-Participant mHealth Outcomes

Participant group visit attendance averaged 4.2/6 (72.0%) sessions, and most (85.2%) were present for three or more classes. CHWs successfully contacted participants via mHealth an average of 15.2/20 (76.0%) of attempts during the intervention (months 1–6) and an average of 6.1/12 (50.1%) of attempts in months 7–12. Common concerns participants relayed to CHWs included medication refills, eligibility, and other access to care issues.

Clinical Outcomes

All proxy glucose values ($n = 5$) were within two standard deviations of individuals' mean HbA1c data and were, therefore, retained for analyses. We conducted linear and quadratic growth curve models for each outcome at baseline and at 6, 12, 18, and 24 months (Table 2). Controlling for respective baseline scores, growth curve models using multiple imputations revealed a significant decrease with some loss of gains over time for HbA1c (linear: $p = 0.07$, quadratic: < 0.0001), systolic blood pressure (linear: $p = 0.001$, quadratic: 0.0007), and diastolic blood pressure (linear: $p = 0.12$, quadratic: 0.003) but not weight (linear: $p = 0.42$, quadratic: 0.30) or BMI (linear: $p = 0.43$, quadratic: $p = 0.22$). For every 6-month

period, systolic blood pressure decreased, on average, by 1.74 points. Quadratic multilevel growth curve models using multiple imputations revealed a significant initial improvement that lessened over time for HbA1c ($p = 0.001$) and diastolic ($p = 0.02$) and systolic blood ($p = 0.03$) pressure but not for weight ($p = 0.12$) and BMI ($p = 0.12$). Furthermore, HbA1c change over time did not vary as a function of whether it was controlled or uncontrolled at 6 months (linear time \times control effect: $b(\text{SE}) = -0.02 (0.08)$, $t(272) = -0.24$, $p = 0.81$, quadratic time \times time \times control effect: $b(\text{se}) = 0.02 (0.05)$, $t(399) = 0.42$, $p = 0.67$).

Table 3 and Figure 2 show HbA1c, weight, BMI, and blood pressure values at baseline to 6, 12, 18, and 24 months. HbA1c levels decreased significantly during the intervention (baseline to 6 months, $p < 0.0001$) and were sustained from baseline to each timepoint through 24 months (range: $p < 0.0001$ to $p = 0.015$). Blood pressure levels decreased significantly during the intervention (systolic: $p < 0.0001$, diastolic: $p = 0.004$), and these improvements were sustained at each timepoint through 18 months but not at 24 months (12 to 18-month systolic range: $p < 0.0001$ to $p = 0.0005$, 24 months $p = 0.065$; 12 to 18-month diastolic range: $p < 0.0001$ to $p = 0.008$; 24 months: $p = 0.85$). There were no significant weight or BMI changes during intervention or post-intervention (weight range: $p = 0.07$ – 0.77 , BMI range: $p = 0.11$ – 0.71). Comparisons of analyses with and without multiple imputations revealed that significance did not change with the exception of baseline to 18-month diastolic blood pressure (multiple imputations with: $p = 0.008$ and without: $p = 0.07$).

Figure 2 illustrates these clinical trends in addition to attrition rates. Attrition levels were $< 50\%$ at all timepoints: 5.9% (6 months), 24.7% (12 months), 35.6% (18 months), and 41.8% (24 months). Attrition was due to loss of eligibility to attend the clinic (22.8% obtained insurance, 15.8% had insufficient paperwork), relocation or deceased (7.0%), or unknown (54.4%). However, 78.5% of the scheduled timepoints occurred during the state-widestay-at-home orders or reduce clinic access due to COVID-19.

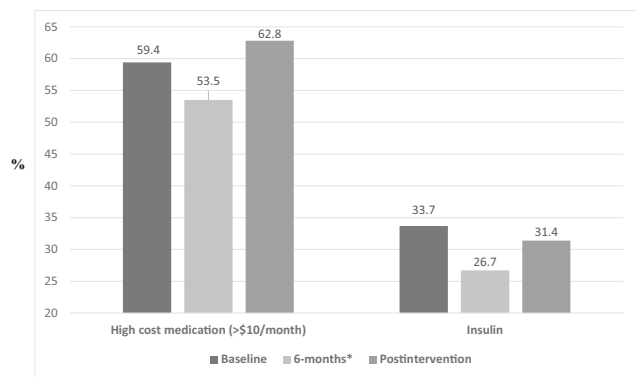


Figure 1 Proportion of low-income Latino(a)s with type 2 diabetes who participated in diabetes programs ($n = 101$) who received high-cost medications and insulin during and post-intervention (*intervention end).

DISCUSSION

This post-intervention analysis evaluated follow-up data from low-income Latino(a)s who participated in the 6-month TIME program. We found HbA1c and blood pressure improvements achieved during the TIME intervention were statistically significantly sustained through 18 and 24 months, respectively. This encourages a model in which a focused, TIME-like multidimensional intervention by an expert group may be sufficient to provide longer-term benefits and could be implemented in routine clinical settings.

Studies have shown that the natural progression of HbA1c and blood pressure is to increase over time.^{12–15} Investigators have also demonstrated that if participants remain in an intervention long term, clinical markers can remain stable.²³ For

Table 2 Results of Linear and Quadratic Growth Curve Models for Each Clinical Outcome ($n = 101$)

Variable	Linear model				Quadratic model			
	<i>b</i> (SE)	<i>t</i> value	df	<i>p</i> value	<i>b</i> (SE)	<i>t</i> value	df	<i>p</i> value
HbA1c (%)								
Effect of time	-0.07 (0.04)	-1.84	49.12	0.07	-0.49 (0.12)	-4.19	75.85	< 0.0001
Quadratic effect of time	N/A	N/A	N/A	N/A	0.10 (0.03)	3.45	61.73	0.001
Weight (lb)								
Effect of time	-0.29 (0.35)	-0.82	52.62	0.42	1.05 (0.81)	1.29	65.48	0.30
Quadratic effect of time	N/A	N/A	N/A	N/A	-0.34 (0.21)	-1.58	45.47	0.12
Body mass index (kg/m ²)								
Effect of time	-0.05 (0.06)	-0.80	56.06	0.43	0.19 (0.15)	1.24	62.58	0.22
Quadratic effect of time	N/A	N/A	N/A	N/A	-0.06 (0.04)	-1.56	48.23	0.12
Blood pressure (systolic) (mmHg)								
Effect of time	-1.74 (0.51)	-3.42	40.59	0.001	-4.98 (1.39)	-3.58	64.36	0.0007
Quadratic effect of time	N/A	N/A	N/A	N/A	0.81 (0.36)	2.29	52.68	0.03
Blood pressure (diastolic) (mmHg)								
Effect of time	-0.46 (0.29)	-1.59	44.37	0.12	-2.56 (0.83)	-3.08	65.06	0.003
Quadratic effect of time	N/A	N/A	N/A	N/A	0.53 (0.21)	2.45	52.28	0.02

Analyses were intent to treat and used multiple imputation methods; models included the respective outcome variable at baseline as a covariate

example, a 5-year study of 120 individuals randomized to group (intervention) or individual (control) diabetes care revealed that HbA1c levels remained stable for intervention but not control participants (intervention: 0.1%; control: 1.7%).²³ However, long-term interventions may not be feasible in low-resource areas. Other investigators demonstrated that the prolonged COVID-19 lockdowns alone played a major role in worsening glucose control.²⁴ Our study showed

that participants may not need to remain in an active intervention trial for a long period of time, even during COVID-19, to maintain significant clinical improvements in HbA1c and blood pressure.

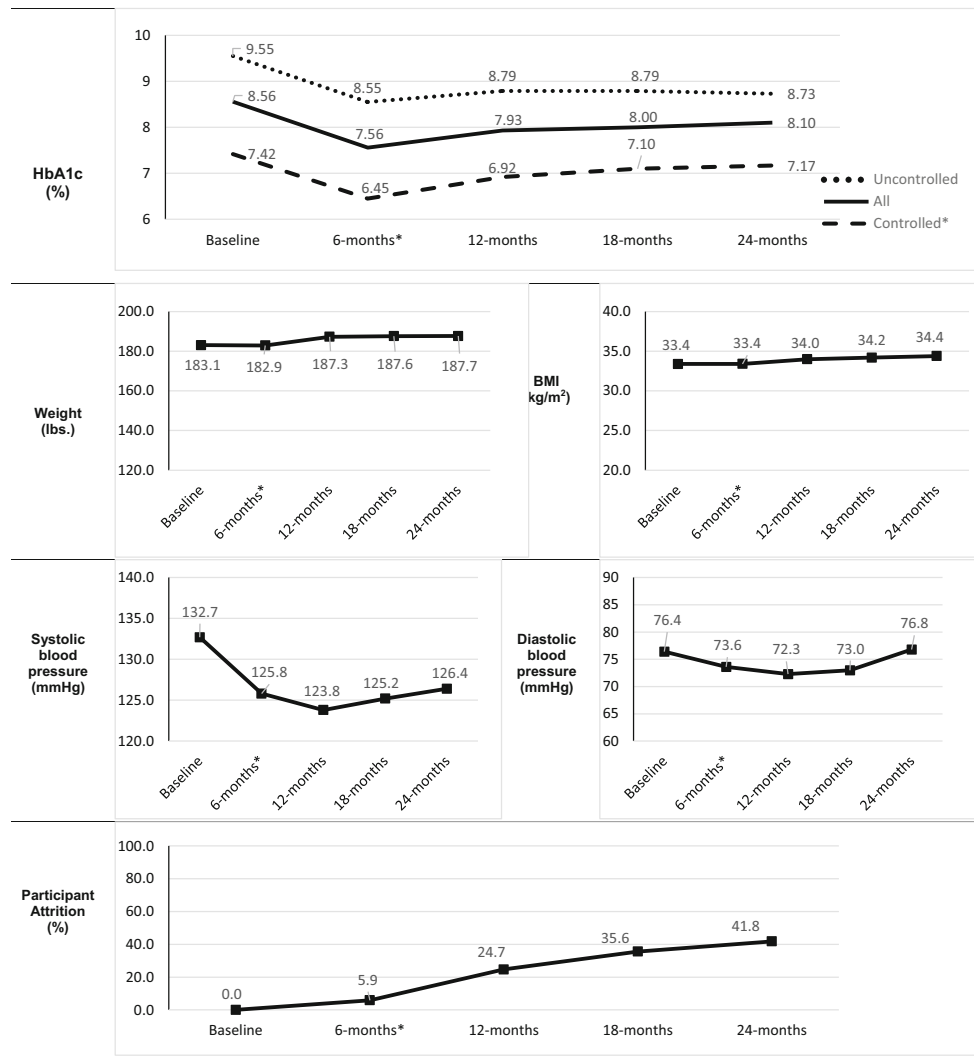
Post-intervention information is critical to determine efficacy and sustainability in addition to maximizing resource utilization. However, there is a paucity of post-intervention data involving CHWs in diabetes initiatives. A study of Mexican

Table 3 Clinical Outcomes of Low-Income Latino(a)s with Type 2 Diabetes Who Participated in TIME ($n = 101$)

Variable	Mean (SD)	Results of paired <i>t</i> test with multiple imputation representing change from baseline to each follow-up assessment ^a			
		<i>b</i> (SE)	<i>t</i> value	df	<i>p</i> value
HbA1c (%)					
Baseline ($n = 101$)	8.56 (2.09)				
6 months (intervention end) ($n = 97$)	7.56 (1.65)	1.00 (0.14)	6.91	54363.0	< 0.0001
12 months ($n = 75$)	7.93 (1.76)	0.64 (0.15)	4.17	660.2	< 0.0001
18 months ($n = 64$)	8.00 (1.77)	0.43 (0.17)	2.54	133.2	0.012
24 months ($n = 43$)	8.10 (1.49)	0.46 (0.18)	2.50	84.1	0.015
Weight (lb)					
Baseline ($n = 101$)	183.12 (46.43)				
6 months (intervention end) ($n = 90$)	182.93 (46.80)	0.19 (0.67)	0.29	100.0	0.77
12 months ($n = 69$)	187.38 (45.48)	-1.06 (0.98)	-1.08	264.0	0.28
18 months ($n = 61$)	187.60 (39.55)	1.15 (1.54)	-0.74	240.6	0.46
24 months ($n = 35$)	187.71 (54.23)	3.76 (1.93)	1.83	68.6	0.07
Body mass index (kg/m ²)					
Baseline ($n = 101$)	33.4 (7.13)				
6 months (intervention end) ($n = 90$)	33.4 (7.13)	0.04 (1.25)	0.37	100.0	0.71
12 months ($n = 67$)	34.0 (6.71)	-0.22 (0.20)	-1.09	158.86	0.28
18 months ($n = 61$)	34.2 (5.79)	-0.23 (0.28)	-0.82	255.78	0.41
Baseline to 24 months ($n = 35$)	34.4 (7.06)	0.54 (0.34)	1.61	79.56	0.11
Blood pressure (systolic) (mmHg)					
Baseline ($n = 101$)	132.73 (15.13)				
6 months (intervention end) ($n = 87$)	125.88 (14.20)	6.85 (1.59)	4.31	100.0	< 0.0001
12 months ($n = 70$)	123.79 (12.38)	8.43 (1.54)	5.46	1434.5	< 0.0001
18 months ($n = 55$)	125.19 (15.20)	6.76 (1.92)	3.53	180.7	0.0005
24 months ($n = 26$)	126.42 (16.21)	6.13 (3.27)	1.88	69.8	0.065
Blood pressure (diastolic) (mmHg)					
Baseline ($n = 101$)	76.41 (9.83)				
6 months (intervention end) ($n = 87$)	73.61 (8.26)	2.79 (0.94)	2.97	100.0	0.004
12 months ($n = 70$)	72.30 (8.19)	3.93 (1.02)	3.87	1435.9	< 0.0001
18 months ($n = 55$)	73.00 (7.60)	3.19 (1.19)	2.67	302.7	0.008
24 months ($n = 26$)	76.83 (8.62)	-0.34 (1.77)	-0.19	69.9	0.85

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^aFor completeness, analyses were also conducted using paired *t* test without multiple imputations. These analyses were compared, and significance did not change with the exception of baseline to 18-month diastolic blood pressure: multiple imputations with: $p = 0.008$, without: $p = 0.07$



*Controlled HbA1c: Defined as HbA1c <7% (or <7.5% if ≥65 years) at baseline²⁰ **intervention end

Figure 2 Clinical and attrition trends for individuals (n = 101) who participated in diabetes programs from baseline to 24 months. *Controlled HbA1c: defined as HbA1c < 7% (or < 7.5% if ≥ 65 years) at baseline.²⁰ ** Intervention end.

Americans participating in 36 CHW-led home visits delivered for 2 years resulted in HbA1c reductions of 0.69%, but post-intervention data are not included.²⁵ A 12-month CHW intervention with follow-up data revealed that HbA1c improvements were sustainable (p = 0.002), but post-intervention data duration was only 6 months.²⁶ Long-term data trends provided by our study were of particular importance since the greatest clinical regression and increases in attrition levels occurred immediately after participants completed the intervention.

There are several possibilities for the clinical regression and increased attrition rates during months 7–12. The majority (78.5%) of data timepoints occurred during the COVID-19 pandemic. This affected clinic eligibility, medication availability, and other aspects of access to care including obtaining routine laboratory values, e.g., HbA1c. Additionally, the clinic ensured that all participants maintained their clinic eligibility in order to complete TIME during the active intervention period, including those who may have lost access otherwise, e.g., incomplete paperwork or obtaining insurance. This likely

contributed to the initial increased attrition after the intervention. Further, some clinic providers changed participant medications to high-cost agents or insulin. Recent Medicaid expansions make it easier to obtain high-cost medications, but many in our population would not meet the requirements, e.g., inability to provide citizenship or residency status.²⁷ In addition, barriers to insulin in underserved populations including lower education attainment, food insecurity, and access to care significantly increase the likelihood of medication non-adherence, hypoglycemic episodes, and poorly controlled diabetes.^{28–30} On the other hand, blood pressure medications are typically generic, reducing access barriers. This may provide rationale for more stable blood pressure levels post-intervention.

Though not entirely clear, there are few possibilities why BMI and weight trends increased during the post-intervention period (months 7–12). The majority of study timepoints occurred during COVID-19, which negatively impacted weight. Investigators found that nearly half of individuals gained

weight, reporting poor dietary habits and high stress levels; average BMI and weight levels increased 0.54 kg/m² and 0.62 kg, respectively, during the pandemic.^{24,31} In our population, most employments were active, e.g., manual labor, which decreased or was omitted during COVID-19, further risking inactivity. Additionally, TIME focuses on prescribing hypoglycemics that are available to low-income populations. Two of the three available (pioglitazone, sulfonylureas) are associated with weight gain.²⁰

It is likely that the effects of CHW involvement during the post-intervention period were likely not immediate. For example, the timeframe for participants to receive post-intervention appointments, and realize and communicate barriers to CHWs, and for CHWs to provide assistance often lasted several months. It is probable, however, that CHW involvement during months 7–12 enhanced the sustainability of clinical outcomes. Future studies are needed to evaluate the efficacy of longitudinal CHW-participant mHealth communication in improving engagement in healthcare systems.

This study had several strengths. One strength of this study is that we obtained long-term data of clinical trends in an underserved population. These trends included the transition to “real-world” care and other relevant parameters of diabetes management beyond HbA1c levels, i.e., blood pressure, weight, BMI, attrition rates, and medication access. To our knowledge, the TIME program is the first to involve CHWs as group visit leaders as well as key members in the research team, providing valuable input in study recruitment and implementation processes.

The study also had limitations. Though attrition was expected in the post-intervention period, it enhanced difficulty of collecting data, limited power, and potentially resulted in selection bias.³² In addition, COVID-19 imposed restrictions, likely decreasing access to care and increasing missing data including HbA1c levels, resulting in using glucose values for some points. However, more than half of the participants remained active in the clinic for diabetes care at month 24 during the pandemic. Additionally, the data are retrospective and observational without a control arm. Future, prospective randomized studies managed by routine clinic personnel are warranted.

CONCLUSIONS

This study revealed that a multidimensional diabetes program may be a valuable modality for vulnerable populations to sustain long-term clinical outcomes. Future studies are warranted to assess long-term outcomes of diabetes interventions conducted by local clinics rather than research teams.

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Declarations:

Ethics Approval: The Institutional Review Board of Baylor College of Medicine approved the study.

Conflict of Interest: The authors declare that they do not have a conflict of interest.

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