#### **RESEARCH ARTICLE**



# Can CO<sub>2</sub> emissions and energy consumption determine the economic performance of South Korea? A time series analysis

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#### **Abstract**

Following the United Nations Sustainable Development Goals (UN-SDGs), which place emphasis on relevant concerns that encompass access to energy (SDG-7) and sustainable development (SDG-8), this research intends to re-examine the relationship between urbanization, CO<sub>2</sub> emissions, gross capital formation, energy use, and economic growth in South Korea, which has not yet been assessed using recent econometric techniques, based on data covering the period between 1965 and 2019. The present study utilized the autoregressive distributed lag (ARDL), dynamic ordinary least square (DOLS), and fully modified ordinary least squares (FMOLS) methods, while the gradual shift and wavelet coherence techniques are utilized to determine the direction of the causality. The ARDL bounds test reveals a long-run linkage between the variables of interest. Empirical evidence shows that CO<sub>2</sub> emissions trigger economic growth. Thus, based on increasing environmental awareness across the globe, it is necessary to change the energy mix in South Korea to renewables to enable the use of sustainable energy sources and establish an environmentally sustainable ecosystem. Moreover, the energy-induced growth hypothesis is validated. This result is supported by the causality analysis, which shows a one-way causality running from energy consumption to GDP in South Korea. This suggests that South Korea cannot embark on conservative energy policies, as such actions will damage economic progress. Additionally, a unidirectional causality is seen from CO<sub>2</sub> emissions and energy consumption to economic growth. These findings have far-reaching consequences for GDP growth and macroeconomic indicators in South Korea.

Keywords CO<sub>2</sub> emissions · Economic growth · Urban population · Energy consumption · Gross capital formation · South Korea

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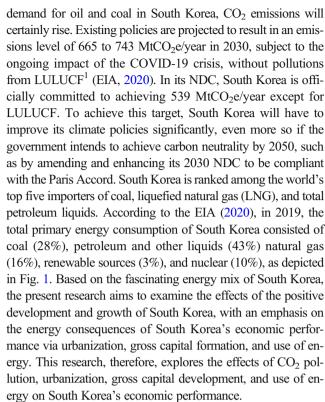


### Introduction

The threat of global warming has raised the level of awareness throughout the world regarding the need to minimize the critical situation confronting all societies (Hasanov et al. 2021). Human activities on the planet's surface are the sole contributor to global warming, resulting in environmental destruction due to global warming (Kirikkaleli & Adebayo 2020). With the advent of global warming, countries have been charged with finding individual and collaborative ways of thinking and working towards mitigating global warming. Climate change is a global problem that has strengthened the international and domestic consciousness to identify ways of mitigating the growing trend (Olanrewaju et al. 2021; Adebayo 2021a). Emissions from diverse energy sources, particularly fossil fuels and other non-renewable sources of energy, are dispersed as pollutants into the air. These are responsible for adversely affecting both the climate and the welfare of the people. Not only are the pollutants released into the environment but they also have connections to bodies of water and wetlands, which can damage or poison marine life (Kirikkaleli & Adebayo2021). The pollution of both water bodies and air has a detrimental impact on society by impacting the populace's living conditions, health, and nutrition (Zhang et al. 2021; Udemba et al. 2021).

According to various measures, growth in the economy has been identified to have a disastrous impact as a result of pollution. Various economic practices, both directed at and based on economic growth, contribute to pollutants' emissions (Ayobamiji & Kalmaz 2020; Udemba 2020). Such practices from multiple sectors (petroleum sector, manufacturing, oil extraction, agriculture) of the economy that cause GDP growth also trigger pollution (Umar et al. 2020). This pollution from various sources and economic sectors negatively impacts human well-being through various types of diseases such as cancer and heart disease (Adedoyin et al. 2020; Oluwajana et al. 2021; He et al. 2021). Manufacturing practices such as the use of heavy duty equipment with the potential to burn huge amounts of fossil fuels, production processes, the delivery of products from the point of origin to the last or final customer with vehicles releasing CO2 emissions via exhaust pipes, waste dumping in bodies of water, and electricity generation using other fossil fuels and coal all lead to environmental degradation.

Since 2016, South Korea's total energy use has remained stable after a period of constant growth (2.7%/year between 2000 and 2016). South Korea's economic performance's driving forces are oil and coal, which account for 37% and 26% of energy needs, respectively. This shows the reliance of the South Korean economy on oil and coal as an energy source. In 2019, South Korea consumed 2.5 million barrels per day of petroleum and other liquids, ranking it the 8th largest consumer globally (EIA, 2020). This implies that, as a consequence of the rising



It is important to recognize that, as a foremost economy in terms of growth, it is necessary to explore the economy and render appropriate economic sustainability suggestions based on analytical results. Based on the results, a detailed analysis of the South Korean economy's sustainability will allow us to develop sustainable policies to answer questions including (a) Can South Korea diversify its policies regarding its energy mix by embracing renewable energy to boost its green economy? (b) Can South Korea introduce innovative pollution mitigating measures without weakening sustainable economic development? It is also essential to remember that, as the South Korean economy grows, the position of the nation in the international rankings of CO2 is very vulnerable and the levels are rising at the same rate as those of other nations considered to be the major global emitters such as India, China, US, Japan, and Russia. In South Korea and Japan, which have the same economic characteristics and demographic composition, there is a disparity between the trend of output growth and the level of pollution. In both countries, this pattern has contributed to the emissions of CO<sub>2</sub>. That being said, considering the rapid population growth rate, significant efforts have been made to mitigate the detrimental consequences of global warming without impacting GDP growth. This is the inspiration for the researchers to investigate the variables illustrated in this report, as we strive to utilize the implications of this research to make policy recommendations for government administrators and stakeholders.



<sup>&</sup>lt;sup>1</sup> Land use, land use change, and forestry

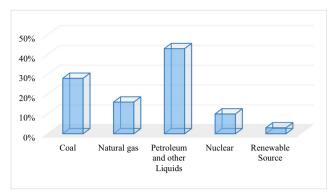


Fig. 1 South Korea total primary energy consumption by fuel 2019

Given this progress, despite the optimistic and substantial growth of the South Korean economy, there has been limited emphasis on examining this trend's significance. Nevertheless, this research intends to examine the economic performance in South Korea amidst CO<sub>2</sub> emissions. The current research is distinct from the existing studies because it accounts for other economic growth determinants, including energy usage, CO<sub>2</sub> emissions, gross capital formation, and urbanization. In addition to the autoregressive distributed lag (ARDL), fully modified ordinary least squares (FMOLS), dynamic ordinary least squares (DOLS), and Gradual shift tests, the study employs the wavelet coherence test. The advantage of the wavelet coherence test is that it can capture both correlation/causality between series at different frequencies and time periods. This report expands/complements the discussion on the South Korean economy on the growth-energy and pollution nexus and expands on India's research by Udemba et al. (2021). The research is inspired by the Sustainable Development Goals (SDGs-7, 8, 12, and 13) and discusses specific energy use concerns (SDG-7) with a particular emphasis on green and sustainable use of energy (SDGs 7 and 12) to meet the 2020 Agenda. This is to avoid problems associated with economic growth (SDGs-8) and climate change (SDGs-13). The present research is considered particularly timely and deserving of inquiry, especially in the current era in which responsible energy use and environmental protection are increasingly being targeted.

The concluding part of this report is planned in the following way: a short synopsis of the previous studies and theoretical framework is presented in the "Literature review" section. The "Data and methodology" section covers the data and the methodologies, whereas the "Findings and discussion" section presents the analytical results. The "Conclusion and policy directions" section presents the conclusion and policy directions.

### Literature review

This part is divided into two, namely, (i) empirical reviewwhich discussed the previous studies conducted regarding the relationship between economic growth, gross capital formation, CO<sub>2</sub> emissions, urbanization, and energy consumption, and (ii) the theoretical framework-which discuss the environmental Kuznets curve theory.

### **Empirical review**

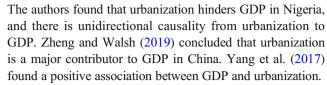
An overview of the relevant literature on this topic will be discussed by reviewing the connections observed between the dependent variable (GDP) and its regressors (CO<sub>2</sub> emissions, urbanization, gross capital formation, and energy use). There is no consensus in the literature on the relation between GDP and these regressors as a result of the mixed outcomes, which has led to an increase in interest in this subject matter. The study of Teng et al. (2020) found that GDP increased CO<sub>2</sub> emissions for ten different OECD economies during the period between 1985 and 2018. However, Ayobamiji and Kalmaz (2020) employed the wavelet technique to capture the timefrequency dependency between CO<sub>2</sub> and real output, which is consistent with the results of Teng et al. (2020). The study of Ahmed et al. (2020) revealed that GDP exerts a positive impact on CO<sub>2</sub> emissions in G7 economies. Aye and Edoja (2017) found a negative link between GDP and CO<sub>2</sub> emission in 31 developing countries. Salahuddin et al. (2018) showed no association between CO<sub>2</sub> and real output. In Kuwait, Wasti and Zaidi (2020) concluded that energy consumption and CO<sub>2</sub> emission accelerate GDP. Chontanawat (2020) and Gorus and Aydin (2019) suggested that there was no causal association between CO2 and real GDP in ASEAN economies for the period from 1971 to 2015. Kirikkaleli and Adebayo (2020), Wang et al. (2019), Kirikkaleli et al. (2020), Aydoğan and Vardar (2020), and Jafari et al. (2015) revealed a one-way casual interaction from GDP to CO<sub>2</sub> emissions. However, while Gao and Zhang's (2021) study showed that there is a unidirectional causal link from CO<sub>2</sub> emission to GDP, a bidirectional causal link between CO2 emission and GDP was revealed by Wu et al. (2018). The study of Ahmed et al. (2019) revealed a positive relationship between GDP and CO2 emissions. Bouznit and Pablo-Romero (2016) examined the interconnection between CO<sub>2</sub> emissions and GDP in Algeria from 1970 to 2010 utilizing the ARDL approach. The results showed that there was a positive association between CO<sub>2</sub> emission and GDP. The study of Odugbesan and Adebayo (2020) in South Africa also revealed a positive interaction between CO<sub>2</sub> emissions and GDP. In Japan, the study of Ahmed et al. (2021) and Adebayo (2021b) established a positive association between GDP growth and environmental degradation. Al-Mulali (2011) established a positive connection between CO<sub>2</sub> emissions and GDP in the MENA region. Furthermore, they found evidence of a two-way causal link between CO<sub>2</sub> emissions and GDP. Awosusi et al. (2020) employed panel data from 1980 to 2018 for the MINT economies. The results showed



that there is no significant link between CO<sub>2</sub> emission and GDP. However, economic growth Granger causes CO<sub>2</sub> emissions. Adebayo (2020) also employed the ARDL and wavelet coherence methods to examine the long-run and causal relationship between CO<sub>2</sub> emissions and GDP in Mexico. The results showed a positive link between these variables. In terms of causality, they identified a two-way interaction between CO<sub>2</sub> emissions and GDP. Zhang et al.'s (2021) study revealed a different causal interconnection between CO<sub>2</sub> emissions and GDP in Malaysia, namely a one-way link from GDP to CO<sub>2</sub> emissions.

The study of Khobai and Le Roux (2017) established a bidirectional link between GDP and energy usage. Muhammad (2019) examined the link between energy usage and GDP in the MENA economies from 2001 to 2017, suggesting a negative linkage between energy use and GDP. Shahbaz et al. (2018) established a positive interaction between energy consumption and GDP in the top ten energyconsuming economies utilizing the quantile-on-quantile (QQ) approach for the period 1960Q1 to 2015Q4. This is consistent with the study conducted by Magazzino (2018) in Italy. Mutascu (2016) explored the causal association between economic growth and energy use in the G7 nations. The author revealed a bidirectional link between economic growth and energy use in the USA, Canada, and Japan, but no causality was evident in the UK and Italy. Yang and Zhao (2014) examined the association between energy consumption and economic growth from 1970 to 2008 in India, utilizing the Granger causality test and DAG. The author revealed that there was a unidirectional linkage from energy consumption to GDP. Faisal et al. (2016) utilized the TY causality to examine the link between GDP and Russia's energy consumption. The authors revealed no causal link between these two variables. Ha and Ngoc (2020) also employed Toda-Yamamoto causality on data covering the period from 1971 to 2017 in Vietnam. The authors revealed a two-way causal link between GDP and energy use. Baz et al. (2019) confirmed a positive shock moving from energy consumption to GDP in Pakistan from 1971 to 2014. Rahman et al. (2020) revealed that energy consumption positively affected China's GDP, covering the period from 1981-2016.

The study of Nathaniel and Bekun (2020) examined the association between urbanization and GDP in Nigeria covering the period from 1971 to 2014 by employing the Bayer and Hanck cointegration test, ARDL, FMOLS, DOLS, CCR, and VECM Granger causality. They found that urbanization negatively inhibits GDP, and there is a bidirectional link between urbanization and GDP. Nguyen and Nguyen (2018) found that urbanization positively affects GDP in the ASEAN countries. Ali et al. (2020a) examined the association between urbanization and GDP using the Maki cointegration test, FMOLS, DOLS, CCR, and VECM Granger causality covering the period from 1971 to 2014.



Numerous studies have been conducted in terms of the linkage between gross capital formation and GDP growth, although their findings are mixed. For instance, Topcu et al. (2020) explored the interaction between gross capital accumulation and GDP by using the panel vector (PVAR) for the period between 1980 and 2018 for 124 economies. The author concluded that the impact of gross capital formation differs based on the country's income level. Etokakpan et al. (2020) examined the association between gross capital accumulation and GDP in Malaysia covering the period 1980-2014, employing the Bayer and Hanck cointegration tests, ARDL, and Granger causality. The authors concluded that an increase in gross capital formation would increase GDP. Kong et al. (2020) employed recent panel techniques to examine the relationship between gross capital formation and GDP for 39 African economies. The authors established a positive link between gross capital formation and GDP. Furthermore, a bidirectional causal link was also evident between these two variables. Boamah et al. (2018) also found similar results for 18 Asian nations by employing panel data covering the period from 1990 to 2017. Table 1 presents a synopsis of related studies.

#### Theoretical framework

This study's theoretical work is based on the EKC which was built on the Kuznets curve of Kuznets (1955), which was centered on income inequality. This theory explains the increasing trend of inequality and income per capita. There is a turning point along the curve that shows where the farmers' per capita income, who exit the farming practices to take up white collar jobs in urban regions is increasing, which closes the large gap between rich and poor. Environmental economists such as (Panayotou 1997; Grossman and Krueger 1991) improved this theory by examining the association between economic growth and environmental quality. The effect of GDP growth on the quality of any economy's environment arises in 3 phases—scale effect, structural effect, and composite effects. In the first phase, environmental degradation is experienced but reaching a point (turning point), the environmental quality begins to improve due to development in innovations and increasing environmental consciousness. This first phase is termed the scale effects. This phase is related to developing nations because non-renewable energy sources promote their economic and production activities. The structural and composite effects are regarding as the turning point. This is associated with developed countries, where most of their economic activities are service and technology driven.



Table 1 Synopsis of studies

| Investigator (s)                              | Timeframe       | Nation (s)                        | Technique(s)  | Findings  |
|---|-----------------|-----------------------------------|---|---|
| CO <sub>2</sub> emissions and economic growth | th<br>1005-2010 |                                   | Total Orient  | (1) ado (0)   |
| 1 eng et al. (2020)                           | 1985–2018       | 10 UECD economies                 | FMG-AKDL  | $CO_2 \rightarrow GDP(+)$                           |
| Rjoub et al. (2021)                           | 1960–2018       | Turkey                            | FMOLS, DOLS   | $CO_2 \rightarrow GDP (+)$                          |
| Zhang et al. (2021)                           | 1960–2018       | Malaysia                          | Maki cointegration, wavelet, and gradual shift  | $GDP \to CO_2 (+)$                                  |
| -   |                 |                                   |   | Î   |
| Odugbesan and Adebayo (2020a)                 | 19/1–2016       | South Africa                      | AKDL & wavelet coherence $\frac{1}{2}$  | $CO_2 \rightarrow GDP(+)$                           |
| Al-Mulalı (2011)                              | 1980–2009       | MENA                              | Panel Granger causality   | 1   |
| Odugbesan and Adebayo (2020b)                 | 1980–2016       | Nigeria                           | ARDL, NARDL   | $\uparrow$  |
| Ayobamiji and Kalmaz (2020)                   | 1971–2015       | Nigeria                           | ARDL, FMOLS, DOLS, wavelet coherence  | $CO_2 \rightarrow GDP (+)$                          |
| Wasti and Zaidi (2020)                        | 1990–2014       | Kuwait                            | ARDL  | $CO_2 \rightarrow GDP (+)$                          |
| Aye and Edoja (2017)                          | 1971–2013       | 31 emerging nations               | Panel techniques  | $CO_2 \rightarrow GDP(-)$                           |
| Kalmaz and Kirikkaleli (2019)                 | 1960-2016       | Turkey                            | ARDL, FMOLS, DOLS, wavelet coherence  | 1   |
| Kirikkaleli et al. (2020)                     | 1950–2016.      | China                             | Maki cointegration, wavelet, and gradual shift  | $CO_2 \rightarrow GDP$                              |
| Bouznit & Pablo-Romero (2016)                 | 1970-2010       | Algeria                           | ARDL  | $CO_2 \rightarrow GDP (+)$                          |
| Aydoğan and Vardar (2020)                     | 1990–2014       | E-7                               | Panel VECM  | $CO_2 \rightarrow GDP$                              |
| Wu et al. (2018).                             | 1995–2017       | World                             | PRISMA  | $CO_2 \leftrightarrow GDP$                          |
| Khan et al. (2020)                            | 1987–2017       | China                             | GMM   | $CO_2 \rightarrow GDP (+)$                          |
| Adebayo (2020)                                | 1971–2016       | Mexico                            | ARDL & wavelet coherence  | $CO_2 \leftrightarrow GDP(+)$                       |
| Jafari et al. (2015)                          | 1980–2007       | Bahrain                           | TY causality  | $CO_{2} \rightarrow GDP$                            |
| Li et al. (2021)                              | 1995–2017       | 20 provinces in China             | Panel CSARDL, AMG   | $CO_2 \rightarrow GDP(+)$                           |
| Salahuddin et al. (2018)                      | 1980-2017       | South Africa                      | ARDI,   | $CO_2 \neq GDP$                                     |
| Autoensi et al. (2020)                        | 1980 2018       | MINIT aconomies                   | ABDI and Danal Grancar conceptive   | CO. + GDP   |
| Awosusi et al. (2020)                         | 1960-2016       | MINI COMORNES                     | ANCH and I and Causanty   | CO2 + CD1   |
| Akinsola and Adebayo (2021)                   | 1971–2016       | Thailand                          | ARDL & wavelet coherence, Granger and Toda-Yamamoto causality                           | $CO_2 \rightarrow GDP(+)$<br>$CO_2 \rightarrow GDP$ |
| Ali et al. (2020b)                            | 1990–2017       | Top 10 emitter countries          | Panel CSARDL  | $CO_2 \rightarrow GDP (+)$                          |
| Gao and Zhang (2021)                          | 1980-2010       | 13 Asian developing countries     | FMOLS and panel Granger causality tests   | $CO_{2} \rightarrow GDP$                            |
| Energy consumption and economic growth        | growth          | 0                                 |   | 1   |
| Shahbaz et al. (2018)                         | 1960Q1-2015Q4   | Top 10 energy-consuming countries | Quantile-on-quantile (QQ) approach  | $EC \rightarrow GDP (+)$                            |
| Khobai and Le Roux (2017)                     | 1971–2013       |                                   | Johansen cointegration and VECM Granger causality tests                                 | $EC \leftrightarrow GDP$                            |
| Ha and Ngoc (2020)                            | 1971–2017       | Vietnam                           | ARDL and Toda-Yamamoto causality  | $EC \leftrightarrow GDP$                            |
| Rahman et al. (2020)                          | 1981–2016       | China                             | Hatemi-J and FMOLS  | $EC \rightarrow GDP (+)$                            |
| Baz et al. (2019)                             | 1971–2014       | Pakistan                          | NARDL   | $EC \rightarrow GDP(+)$                             |
| Faisal et al. (2016)                          | 1990–2011       | Russia                            | Toda-Yamamoto causality   | No causal link                                      |
| Yang and Zhao (2014)                          | 1970-2008       | India                             | Granger causality and DAG   | EC → GDP  |
| Mutascu (2016)                                | 1970-2012       | G7 economies                      | Granger causality tests   | EC ⇔ GDP  |
| Muhammad (2019)                               | 2001–2017       | MENA                              | MENA  | EC → GDP (-)  |
| Urbanization and economic growth              |                 |                                   |   |   |
| Nathaniel and Bekun (2020)                    | 1971–2014       | Nigeria                           | Bayer and Hanck cointegration tests, ARDL, FMOLS, DOLS, CCR, and VECM Granger causality | $URB \leftrightarrow GDP (-)$                       |
| Nguyen and Nguyen (2018)                      | 1971–2014       | ASEAN                             | D-GMM and PMG   | $URB \rightarrow GDP (+)$                           |
| Ali et al. (2020a)                            | 1971–2014       | Nigeria                           | Maki cointegration, FMOLS, DOLS, CCR, and VECM Granger causality                        | $URB \rightarrow GDP(-)$                            |
| Adebayo (2021b)                               | 1971–2015       | Japan                             | Wavelet coherence, ARDL, FMOLS, DOLS  | $URB \rightarrow GDP (+)$                           |
| Yang et al. (2017)                            | 2000-2010       | China                             | Pooled ordinary least squares (POLS), fixed effects (FE), and random effect (RE         | $URB \rightarrow GDP (+)$                           |
| Zheng and Walsh (2019)                        | 2001–2012       | 29 provinces in China             | FE and sys-GMM estimated methods  | $URB \rightarrow GDP (+)$                           |
| Gross capital formation and economic growth   | nic growth      |                                   |   |   |
| Topcu et al. (2020)                           | 1980–2018       | 124 countries                     | PVAR  | $GCF \rightarrow GDP (+)$                           |
| Etokakpan et al. (2020)                       | 1980–2014       | Malaysia                          | Bayer and Hanck cointegration tests, ARDL and Granger causality                         | $GCF \rightarrow GDP (+)$                           |
|   |                 |                                   |   |   |



Urban ↔ GDP (+)  $Urban \rightarrow GDP(+)$  $GCF \rightarrow GDP(+)$  $GCF \rightarrow GDP(+)$ gradual Shift ARDL, wavelet coherence, AMG and CCEMG Fechnique(s) POLS 39 African countries 18 Asian nations Nation (s) Malaysia **Fimeframe** 1997-2017 990-2017 1971-2016 
 Fable 1 (continued)
 Boamah et al. (2018) Zhang et al. (2021) Kong et al. (2020) nvestigator (s)

Note: EC, energy consumption; GDP, economic growth;  $CO_2$ , carbon emission;  $\rightarrow$  (+), positive relationship;  $\rightarrow$  ( $\rightarrow$ ), negative relationship;  $\rightarrow$ , one-way causality;  $\leftrightarrow$ , bidirectional causality; urban, urbanization; GCF, gross capital formation

# **Data and methodology**

### **Data**

The present research explores the impact of CO<sub>2</sub> emissions (CO<sub>2</sub>) on economic growth (GDP) and also the role of gross capital formation (GCF), energy use (EC), and urbanization (URB) in South Korea using data spanning between 1965 and 2018. In the case of South Korea, the current research was conceived with the perspective of examining the connections between GDP growth, CO<sub>2</sub> pollution, urbanization, and energy use. The empirical modeling is based on the ARDL technique. This analysis is based on the study by Udemba (2020) and Nathaniel et al. (2020) by adjusting for further catalysts of growth that have been overlooked in the literature, including growth theory caused by the urban population. As shown in the model of Solow growth regarding capital and labor contribution. For South Korean cases that have the same economic characteristics, urban populations are included in our sample scenario. The parameters utilized are transmuted into a logarithm. This was conducted to ensure data is normally distributed (Rjoub et al. 2021; Kirikkaleli et al. 2020). Table 2 illustrates the data source, measurement, and unit of measurement. Also, the flow of analysis is depicted in Fig. 2. The study economic function and econometric model are depicted in Eqs. (1) and (2):

$$GDP_t = f(CO_{2t}, URB_t, EC_t GCF_t)$$
 (1)

$$GDP_t = \vartheta_0 + \vartheta_1 CO_{2t} + \vartheta_2 URB_t + \vartheta_3 EC_t + \vartheta_3 GCF_t + \varepsilon_t$$
 (2)

In Eq. (1), GDP, CO<sub>2</sub>, GCF, EC, and URB represent economic growth, CO<sub>2</sub> emissions, gross capital formation, energy consumption, and urbanization.

### Methodology

Analysis of correlation is used to verify the association of two time series data. The correlation can be defined as follows:

$$Corr (X, Y) = \frac{Cov (X, Y)}{\sqrt{Var(X)Var(Y)}}$$
(3)

where the covariance between the two time series (X,Y) is denoted as Cov(X,Y) while Var(X) and Var(Y) denote the value of the two time series of X and Y, respectively.

To achieve this purpose, ADF, PP, and Zivot and Andrews (ZA) unit root tests were employed to establish the order of integration. However, the ZA can capture both the stationarity property and structural break. This study employed the ARDL approach because this technique accommodates a limited number of observations (Ayobamiji and Kalmaz 2020; Kirikkaleli et al. 2018; Adebayo and Kirikkaleli 2021). Its



**Table 2** Variable units and sources

| Variable | Description               | Units  | Sources    |
|----------|---------------------------|--|------------|
| GDP      | Economic growth           | GDP per capita constant \$US, 2010                             | WDI (2021) |
| $CO_2$   | Environmental degradation | Metric Tonnes Per Capita                                       | WDI (2021) |
| GCF      | Gross capital formation   | % of GDP   | WDI (2021) |
| URB      | Urbanization              | Urban population   | WDI (2021) |
| EC       | Energy use                | Primary energy consumption is measured in terawatt-hours (TWh) | BP (2021)  |

model is suitable for a model with different lags and mixed order of integration. It is also beneficial because of the attributes of revealing coefficients in the short and long run simultaneously and solves the problem of autocorrelation. This makes the credibility of the formulated policy of this study to be effective. ARDL modeling is defined in the equation below:

$$\Delta \text{InGDP}_{t} = \alpha_{0} + \sum_{i=1}^{t} \alpha_{1} \Delta \text{InGDP}_{t-i} + \sum_{i=1}^{t} \alpha_{2} \Delta \text{InCO}_{2t-i}$$

$$+ \sum_{i=1}^{t} \alpha_{3} \Delta \text{InEN}_{t-i} + \sum_{i=1}^{t} \alpha_{4} \Delta \text{InGCF}_{t-i}$$

$$+ \sum_{i=1}^{t} \alpha_{5} \Delta \text{InURB}_{t-i} + \beta_{1} \text{InGDP}_{t-1}$$

$$+ \beta_{2} \text{InCO}_{2t-1} + \beta_{3} \text{InEN}_{t-1} + \beta_{4} \text{InGCF}_{t-1}$$

$$+ \beta_{5} \text{InURB}_{t-1} + \rho \text{ECT}_{t-i} + \varepsilon_{t}$$

$$(4)$$

where  $\alpha_{i=5}$  and  $\beta_{i=5}$  are the long- and short-run parameters respectively,  $\rho$  denotes parameter for  $ECT_{t-i}$ ,  $\varepsilon_t$  and  $\Delta$  denote the error term and first difference respectively, and  $ECT_{t-i}$  represents the error correction term, which is the adjusted speed to long-run balance from short-run shock. The ARDL hypotheses are written below. The null hypothesis reiterates that there is no cointegration presence in the model, while the alternate hypothesis affirms a contradictory view, which is the presence of cointegration.

$$H_0: \alpha_0 = \alpha_1 = \alpha_2 = \alpha_3 = \alpha_4 = \alpha_5$$
 (5)

$$H_1: \alpha_0 \neq \alpha_1 \neq \alpha_2 \neq \alpha_3 \neq \alpha_4 \neq \alpha_5 \tag{6}$$

The testing procedure in the ARDL model is by comparing the *F* or *T* statistics calculated with the critical bound (lower and upper bound). Normality test, heteroscedasticity test, Ramsey RESET, and serial correlation are the diagnostic tests undertaken to examine for BLUE. Model stability was checked employing the cumulative sum of recursive residuals (CUSUM) and cumulative sum of squares of recursive residuals (CUSUM of squares). Furthermore, the long-run coefficients of the ARDL model were verified using the FMOLS and DOLS tests.

Wavelets coherence is employed to detect the time-frequency dependence of energy consumption, CO<sub>2</sub> emissions, urbanization, and gross capital formation on economic growth. Time-frequency dependence puts into account the changes over time and how the relationship varies from one frequency to another becomes essential and strategic in the formulation of policies (Beton Kalmaz and Adebayo 2020; Mutascu 2018; Alola and Kirikkaleli 2019; Umar et al. 2020; Alola and Kirikkaleli 2020). The Morlet wavelet function was employed since it brings balance between phase and amplitude. Morlet wavelet function is defined as follows:

$$\supseteq(\backslash) = \pi^{-\frac{1}{4}} e^{-i \supseteq \backslash} e^{-\frac{1}{2} \backslash^2} \tag{7}$$

Note: non-dimensional frequency was used by  $\exists$ ; i denotes  $\sqrt{-1}$  p(\)). Using the time and space, with \ = 0,1,2,3.....N-1, the time series continuous wavelet transformation (CWT) is defined as:

$$\supseteq_{k,f}(\backslash) = \frac{1}{\sqrt{h}} \supseteq \left(\frac{n-k}{f}\right), \quad k, f \in \mathbb{R}, f \neq 0$$
 (8)

where *k* and *f* symbolize time and frequency, respectively. CWT helps the cross-wavelet analysis to interrelate between two variables (Kirikkaleli 2019). The CWT equation is written below as follows:

$$\mathcal{W}_{p}(k,f) = \int_{-\infty}^{\infty} p(n) \frac{1}{\sqrt{f}} \, \Box \left( \frac{\overline{n-k}}{f} \right) dn \tag{9}$$

The local variance was revealed using the wavelet power spectrum (WPS). The equation defining the WPS is as follows:

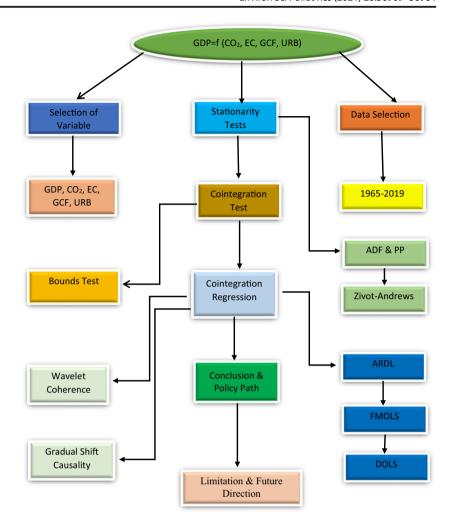
$$WPS_p(k,f) = \left| W_p(k,f) \right|^2 \tag{10}$$

To examine the co-movement between two time series, the wavelet coherence approach (WTC) was used, which is defined in the equation below:

$$R^{2}(k,f) = \frac{\left| S(f^{-1}W_{pj}(k,f)) \right|^{2}}{S(f^{-1}|W_{p}(k,f)|^{2})S(f^{-1}|W_{j}(k,f)|^{2})}$$
(11)



Fig. 2 Analysis flowchart



where the smoothing operator to both time and scale with  $0 \le R^2(k,f) \le 1$  is denoted as *S*. WTC can also detect the phase difference  $\phi_{pq}$  of the two time series; it is defined in this form:

$$\phi_{pq}(k,f) = \tan^{-1}\left(\frac{L\{S(f^{-1}W_{pj}(k,f))\}}{O\{S(f^{-1}W_{pj}(k,f))\}}\right)$$
(12)

where L denotes an imaginary operator while O stands for a real part operator.

Apart from the wavelet coherence approach, the gradual shift causality test developed by Nazlioglu et al. (2016) was utilized to establish the direction of causation between two variables. Nazlioglu et al. (2016) employed the Toda and Yamamoto (1995) and Fourier approximation, which captures the structural changes during the period of coverage (Kirikkaleli and Gokmenoglu 2020; Gokmenoglu et al. 2019). This technique helps to overcome the inaccuracies and inconsistencies associated with the VAR model. Using the modified VAR model stated in the equation below:

$$y_t = \sigma(t) + \beta_1 y_{t-1} + \dots + \beta_{p+d\max} y_{t-(p+d\max)} + \varepsilon_t$$
 (13)



where  $y_t$  symbolizes variable used;  $\sigma$  symbolizes intercept;  $\beta$  symbolizes coefficient matrices;  $\varepsilon$  symbolizes the error term; and t symbolizes time function. The Fourier approximation with cumulative frequencies is defined as:

$$\sigma(t) = \sigma_0 + \sum_{k=1}^{n} \gamma_{1k} \sin\left(\frac{2\pi kt}{T}\right) + \sum_{k=1}^{n} \gamma_{2k} \cos\left(\frac{2\pi kt}{T}\right)$$
 (14)

where  $\gamma_{2k}$  and  $\gamma_{1k}$  measures the displacement and frequency amplitude respectively and the number of frequencies is denoted as n. Fourier Toda-Yamamoto causality with cumulative frequencies (CF) is defined as follows in:

$$y_{t} = \sigma_{0} + \sum_{k=1}^{n} \gamma_{1k} \sin\left(\frac{2\pi kt}{T}\right) + \sum_{k=1}^{n} \gamma_{2k} \cos\left(\frac{2\pi kt}{T}\right) + \beta_{1} y_{t-1} + \dots + \beta_{p+d\max} y_{t-(p+d\max)} + \varepsilon_{t}$$

$$(15)$$

where approximation frequency is symbolized as k. For the Fourier Toda-Yamamoto causality with single

frequencies, single-frequency components are defined in Eq. (11) as follows:

$$\sigma(t) = \sigma_0 + \gamma_1 \sin\left(\frac{2\pi kt}{T}\right) + \gamma_2 \cos\left(\frac{2\pi kt}{T}\right)$$
 (16)

The Fourier Toda-Yamamoto causality with single frequencies (SF) is defined as follows:

$$y_{t} = \sigma_{0} + \gamma_{1} \sin\left(\frac{2\pi kt}{T}\right) + \gamma_{2} \cos\left(\frac{2\pi kt}{T}\right) + \beta_{1} y_{t-1}$$

$$+ \dots + \beta_{p+d} y_{t-(p+d)} + \varepsilon_{t}$$

$$(17)$$

## Findings and discussion

The present paper aims to examine the connection between CO<sub>2</sub> emissions (CO<sub>2</sub>) and economic growth (GDP) as well as the role of urbanization (URB), gross capital formation (GCF), and energy usage (EC) in South Korea between 1965 and 2018. Figure 3 (correlation box) depicts the correlation among the parameters. In Fig. 3, the blue color illustrates a positive correlation, while the red color illustrates a negative correlation between variables. The correlation outcomes show that all the variables have a strong correlation with each other with the exemption of GCF with a weak correlation with other indicators. Consequently, the elementary summary statistical characteristics that report the measure of central tendencies and dispersion outlined in Table 3 show that urbanization reveals the highest average followed by economic growth, energy use, and the least gross capital formation. All series shows negative skewness with light tail, and the kurtosis revealed that data are normally distributed with the exemption of GCF. Furthermore, the study utilized the aforementioned methods to test the stationarity features of the data. The outcomes show that the parameters are integrated in mixed order, i.e., I(I) and I(0), as depicted in Tables 4 and 5.

After the stationarity characteristics of the series are affirmed, we proceed to estimate the ARDL framework, which is reported in Table 6. The outcomes are presented in such a manner that the existence of a long-run association between the parameters is verified if the estimated value of the *F*-test is greater than the values of both limits (lower and upper bound). In this scenario, we fail to accept the null hypothesis of no cointegration. Nonetheless, if the *F*-statistics is less than the lower bound critical value, the alternative explanation of the presence of cointegration is dismissed, although, if the *F*-stats breakdown between the two thresholds critical values, the finding is called inconclusive.

We proceed to estimate the long-run and short-run association between variables of concern, and findings are depicted in Table 7. Appropriate lag selection is essential when applying the ARDL. Thus, we utilized the AIC criteria proposed by Akaike (1987). As stated by Udemba et al. (2021) and Zhang e al. (2021), the AIC is preferred for selection lag due to its superior characteristics. The model goodness of fit is depicted by the  $R^2$  (0.99) and Adj  $R^2$  (0.98), respectively. The outcomes of the  $R^2$  and Adj  $R^2$  illustrate that 99% and 98% variation in GDP can be explained by urbanization, gross capital formation, CO<sub>2</sub> emissions, and energy consumption, and the remaining percentage can be attributed to error. The speed of adjustment is seen to facilitate long-term convergence between the parameters with a significant and negative error correction model coefficient (ECM). The outcome of the ECT is 0.22, which illustrates evidence of cointegration among the parameters, and this signifies the capability of the model to witness 22% speed of adjustment to verify the alignment to equilibrium in the long run on GDP due to the effect of the regressors (URB, EC, CO<sub>2</sub>, and GCF). The outcomes of the ARDL illustrate the different linkage between GDP and the regressors (URB, EC, CO<sub>2</sub>, and GCF). For instance, there is proof of a negative linkage between GDP and CO<sub>2</sub> emissions in the short run. This illustrates that an increase in CO<sub>2</sub>

Table 3 Descriptive statistics

|              | GDP       | $CO_2$    | EC        | GCF        | URB        |
|--------------|-----------|-----------|-----------|------------|------------|
| Mean         | 3.899157  | 0.722277  | 1.905183  | 1.494485   | 7.435024   |
| Median       | 3.990954  | 0.812789  | 2.046541  | 1.508928   | 7.520742   |
| Maximum      | 4.457504  | 1.141450  | 2.389607  | 1.615256   | 7.624351   |
| Minimum      | 3.062850  | -0.059921 | 0.968732  | 1.169181   | 6.967840   |
| Std. Dev.    | 0.433817  | 0.337479  | 0.434017  | 0.075982   | 0.196133   |
| Skewness     | -0.388375 | -0.696901 | -0.558883 | - 1.695635 | - 0.921605 |
| Kurtosis     | 1.819464  | 2.317733  | 1.986679  | 7.925874   | 2.583351   |
| Jarque-Bera  | 4.576475  | 5.518732  | 5.216344  | 81.96133   | 8.183577   |
| Probability  | 0.101445  | 0.063332  | 0.073669  | 0.000000   | 0.016709   |
| Observations | 55        | 55        | 55        | 55         | 55         |

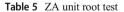


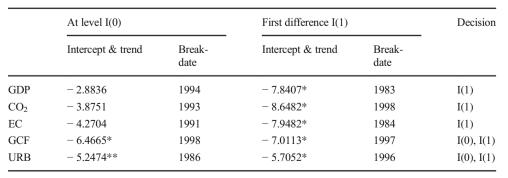
Table 4 Traditional unit root tests

|            | At level I(0)<br>Intercept & trend | First difference I(1)<br>Intercept & trend | Decision   |
|------------|------------------------------------|--|------------|
| ADF unit   | root test                          |  |            |
| GDP        | 0.3979                             | - 0.9324*                                  | I(1)       |
| $CO_2$     | - 0.9710                           | - 8.0537*                                  | I(1)       |
| EC         | - 0.5678                           | - 7.0742*                                  | I(1)       |
| GCF        | - 4.7370*                          | - 6.3073*                                  | I(0), I(1) |
| Urban      | - 4.9187*                          | - 2.1810                                   | I(0)       |
| PP unit ro | oot test                           |  |            |
| GDP        | 0.6869                             | - 6.9458*                                  | I(1)       |
| $CO_2$     | - 0.9463                           | - 8.0591*                                  | I(1)       |
| EC         | - 0.5908                           | - 7.0742                                   | I(1)       |
| GCF        | - 4.7149*                          | - 15.3632                                  | I(0), I(1) |
| URB        | - 4.0221**                         | - 2.6810                                   | I(0)       |

Note: 1% and 5% level of significance is illustrated by \* and \*\* respectively

emissions will mitigate GDP, which infers that an increase in CO<sub>2</sub> by 1% in the short run is accompanied by a 0.28% decrease in GDP when other indicators are held constant. This result concurs with the findings of Lee (2013) for G-20 nations and Udemba et al. (2021) for India. Furthermore, urbanization and energy consumption exert a positive impact on economic growth in the short run, which illustrates that 1% in urbanization and energy consumption will increase economic growth by 1.92% and 0.23%, respectively. This is a strong indicator that the intensity of energy and the populace are critical indicators in the growth of the South Korean economy. The industrial sector uses a momentous quantity of energy as one of the main contributors and generators of the South Korean economy, which in turn affects income positively. This outcome aligns with the results of Zhang et al. (2021) for Malaysia and Kirikkaleli and Demet (2020) for Turkey. Moreover, there is no evidence of interconnection between gross capital formation and GDP in the short run, which aligns with the study of Olanrewaju et al. (2021) for Thailand and Zhang et al. (2021) for Malaysia.





Note: 1% and 5% level of significance is illustrated by \* and \*\* respectively



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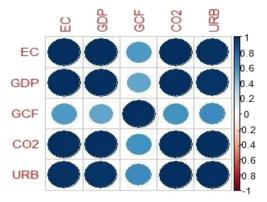


Fig. 3 Correlation between GDP, URB, EC, GCF, and CO<sub>2</sub>

In the long run, there is evidence of a long-run association between GDP and CO<sub>2</sub> emissions, which illustrates that a 0.15% increase in GDP is linked with a 1% upsurge in CO<sub>2</sub> emissions. This is not unexpected given that the South Korean economy is primarily an investment-oriented and manufacturing economy that relies heavily on the utilization of energy; nonetheless, a positive side of this is the potential to minimize CO<sub>2</sub> by shifting the energy mix to include more renewable options such as wind and solar energy (renewables). Furthermore, urbanization influences economic growth positively in South Korea, which is in harmony with a positive linkage with GDP growth and urbanization. We see that a 1.92% increase in GDP growth is due to a 1% increase in the urban population. We make the claim based on the empirical revelation that the growing population in South Korea is productive to her economic trajectory. Nevertheless, there is a necessity for cautiousness on the part of policymakers to match urban infrastructure and amenities in the rural area. This is to avoid the rush to urban cities, given that most government officials develop urban areas more than rural areas. Otherwise, the urban infrastructure might be overwhelmed and might impede economic growth in the long run. We noticed that energy use improves economic growth as we noticed that a 0.23% increase in economic growth is due to an increase in energy use by a magnitude of 1%. This outcome gives credence to the energy-induce growth hypothesis, which complies with the study of Nathaniel et al. (2020) for Nigeria

and also in Pakistan by Shahbaz et al. (2012). This outcome implies that the South Korean economy is energy driven and cannot embark on conservative energy strategies as such action will compromise economic growth.

Furthermore, various post-estimation tests are conducted. The outcomes of the normality, serial correlation, heteroscedasticity, and Ramsey tests show that the model is well specified, and there is no serial correlation. Moreover, the outcomes of the CUSUM and CUSUMSQ depict in Figure 4a, and 4b, respectively, which exemplifies that the model is stable.

To confirm the ARDL long-run estimations' outcomes, the current study utilized the FMOLS and DOLS, which are portrayed in Table 8. The outcomes show that CO<sub>2</sub> emissions. Energy consumption and urbanization enhance economic growth in the long run, while there is no significant interconnection between gross capital formation and economic growth.

The current research further utilizes the wavelet coherence (WTC) test to catch the causality and correlation between economic growth and the regressors (CO2, EC, URB, and GCF). This method is crafted from physics to obtain formerly undetected information. Therefore, the research investigates the connection in the short, medium, and long run between GDP and its regressors. The cone of influence (COI) is the white cone where discussion is carried out in the WTC. The thick black contour illustrates a level of significance based on simulations of Monte Carlo. Figure 5a–5d 0–4, 4–8, and 8–16 illustrate short, medium, and long term, respectively. Moreover, the vertical and horizontal axis in the figure depicts frequency and time, respectively. The blue and yellow colors depict low and high dependence between the series. In-phase and out-of-phase connections are depicted by rightward and leftward arrows, respectively. Moreover, the rightward-down (leftward-up) illustrates that the first variable lead (cause) the second parameter while the rightward-up (leftward-down) depicts that second parameter lead (cause) the first parameter. Figure 5a illustrates the WTC between GCF and GDP between 1965 and 2019. At different frequencies between 1970 and 2016, the arrows are rightward-up, showing the positive interconnection between the series with GCF leading. Figure 5 b depicts the WTC between EN and GDP between 1965 and 2018 in South Korea. The majority of the arrows are

 Table 7
 ARDL long-run and short-run results

| Variables       | Long-run re | esults              |       | Short-run results |                     |       |
|-----------------|-------------|---------------------|-------|-------------------|---------------------|-------|
|                 | Coefficient | <i>t</i> -statistic | Prob  | Coefficient       | <i>t</i> -statistic | Prob  |
| CO <sub>2</sub> | 0.1552**    | 2.0736              | 0.045 | - 0.2804*         | - 2.999             | 0.006 |
| EC              | 0.2370**    | 2.1121              | 0.041 | 0.4142*           | 3.993               | 0.000 |
| GCF             | 0.0486      | 1.1719              | 0.248 | 0.0695            | 1.691               | 0.103 |
| URB             | 1.9218**    | 2.1331              | 0.039 | 4.0658*           | 2.789               | 0.010 |
| ECT(-1)         | -           | -                   | -     | - 0.2273*         | -4.030              | 0.000 |
| $R^2$           | 0.99        |                     |       |                   |                     |       |
| $Adj R^2$       | 0.98        |                     |       |                   |                     |       |

Note: 1% and 5% level of significance is illustrated by \* and \*\* respectively

rightward-up, which illustrates a strong connection (dependency) at different frequencies with energy use leading. Figure 5c illustrates the WTC between URB and GDP in South Korea between 1965 and 2019. At different scales, from 1970 to 2016, the bulk of arrows are facing rightward-up, which show positive association (dependency) at different frequencies. Furthermore, the rightward-up arrows show that URB lead (cause) GDP. Figure 5 d illustrates the WTC between URB and GDP in South Korea between 1965 and 2019. From 1995 to 2003, the bulk of arrows are facing rightward-up at high and medium frequencies, which shows a positive association (dependency) between GDP and GCF. Furthermore, the rightward-up arrows show that GCF lead (cause) GDP. The outcomes from the wavelet coherence test comply with the results of DOLS, FMOLS, and ARDL.

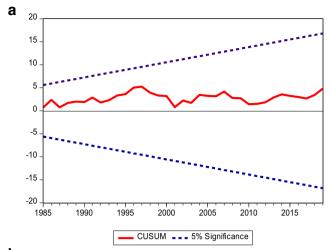
To capture the causal association among the indicators, the gradual shift test is applied, and the results of this technique are presented in Table 9. The outcomes of the gradual shift causality revealed: (i) a one-way causality from energy consumption to GDP at a significance level of 1%. This suggests that South Korea's economic growth is not immune from development caused by fossil fuels. This is because South Korea's economy is heavily reliant on production and manufacturing operations. This outcome corresponds to the findings of Ramakrishna and Rena (2013) and Udemba et al. (2021); (ii) a unidirectional causality from CO<sub>2</sub> emissions to

Table 6 Bound test

| Model                                 | F-<br>statistics | Cointegration | $\chi^2$ ARCH | $\chi^2$ RESET | $\chi^2$ normality | $\chi^2$ LM |
|---------------------------------------|------------------|---------------|---------------|----------------|--------------------|-------------|
| GDP=f(CO <sub>2</sub> , EC, GCF, URB) | 5.61*            | Yes           | 0.57 (0.56)   | 0.72 (0.47)    | 0.91 (0.63)        | 0.58 (0.44) |
| , ,                                   | 10%              |               | 5%            |                | 1%                 |             |
|                                       | LB               | UB            | LB            | UB             | LB                 | UB          |
|                                       | 2.26             | 3.35          | 2.62          | 3.79           | 3.41               | 4.68        |

<sup>\*</sup>Represents a 1% level of significance and LB and UB denote lower bound and upper bound critical value





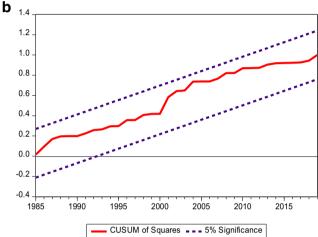


Fig. 4 a CUSUM. b CUSUM of squares

GDP growth at a 1% level of significance, which implies that  $CO_2$  emissions are a strong predictor of GDP growth in South Korea. In this regard, South Korean policymakers should develop policies in line with the nation's energy portfolio diversification. This finding is consistent with the results of Zhang et al. (2021) for Malaysia, Adebayo (2020) for Indonesia, and

Table 8 FMOLS and DOLS outcomes

| Variable                 | FMOLS       |                     |       | DOLS        |                     |       |
|--------------------------|-------------|---------------------|-------|-------------|---------------------|-------|
|                          | Coefficient | <i>t</i> -statistic | Prob. | Coefficient | <i>t</i> -statistic | Prob. |
| $\overline{\text{CO}_2}$ | 0.1443**    | 2.5057              | 0.017 | 0.2370**    | 2.4821              | 0.017 |
| EC                       | 0.2437*     | 2.8833              | 0.006 | 0.2332**    | 2.3277              | 0.025 |
| GCF                      | 0.0574      | 1.3338              | 0.185 | 0.0486      | 1.3770              | 0.177 |
| URB                      | 2.0115**    | 2.5782              | 0.014 | 1.9218**    | 2.5067              | 0.016 |
| $R^2$                    | 0.98        |                     |       | 0.98        |                     |       |
| $Adj R^2$                | 0.97        |                     |       | 0.97        |                     |       |

Note: \* and \*\* represents 1% and 5% level of significance respectively



 Table 9
 Gradual shift causality test

| Causality path         | Wald stat | No of Fourier | P-value | Decision         |
|------------------------|-----------|---------------|---------|------------------|
| GDP → EC               | 10.940    | 3             | 0.1412  | Do not reject Ho |
| $EC \rightarrow GDP$   | 18.828*   | 3             | 0.0087  | Reject Ho        |
| GDP → URB              | 21.986*   | 1             | 0.0025  | Reject Ho        |
| URB <b>→</b> GDP       | 5.868     | 1             | 0.5551  | Do not reject Ho |
| $GDP \rightarrow CO_2$ | 10.721    | 2             | 0.1512  | Do not reject Ho |
| $CO_2 \rightarrow GDP$ | 23.974*   | 2             | 0.0011  | Reject Ho        |
| $GDP \rightarrow GCF$  | 10.721    | 2             | 0.1512  | Do not reject Ho |
| $GCF \rightarrow GDP$  | 5.1177    | 2             | 0.6455  | Do not reject Ho |

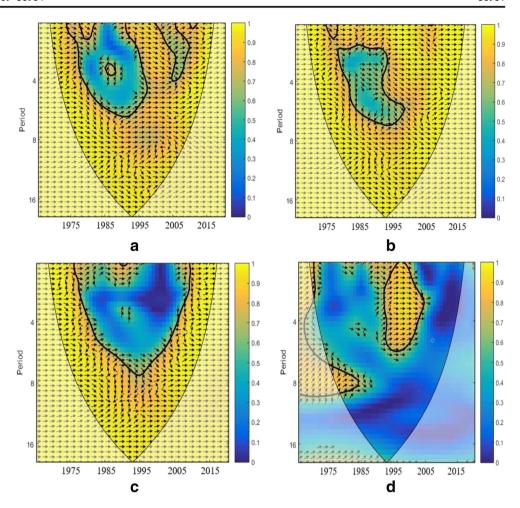
Note: 1%, 5%, and 10% level of significance is illustrated by \*, \*\*, and \*\*\* respectively

Awosusi et al. (2020) for the MINT economies; (iii) a one-way causality from GDP to urbanization at the significance level of 1%, which suggests that GDP can predict URB. This finding has consequences for South Korea's economic growth in terms of rapid urbanization, where investment will be attached mainly to the vulnerable population with the promise of improvement in welfare. This outcome concurs with the findings of Awosusi et al. (2020), Udemba et al. (2021), and Zhang et al. (2021).

# **Conclusion and policy directions**

The current study adds to the previously existing literature by assessing the linkage between economic growth, CO<sub>2</sub> emissions, energy usage, urbanization, and gross capital formation in South Korea using yearly data stretching between 1965 and 2019. To accomplish the stated objectives, the ARDL bounds test, the gradual shift causality test, and the novel wavelet coherence test are utilized. The outcomes show a mix (significant and insignificant) of associations between economic growth and the regressors. The outcome of the bounds test reveals that all the indicators have a long-run interconnection. Furthermore, the outcomes of the ARDL long-run and shortrun estimations show that energy usage, urbanization, and CO<sub>2</sub> emissions enhance the economic performance of South Korea, while gross capital formation exerts an insignificant impact on the economic performance of South Korea. Furthermore, we applied the novel wavelet test to capture the correlation and causal association between economic growth and the regressors. The wavelet analysis findings revealed a positive connection between economic growth and the regressors with the exemption of gross capital formation, which has a weak interconnection with economic performance. Furthermore, the wavelet coherence test outcomes provide further support for the ARDL, FMOSL, and DOLS tests. The gradual shift causality test outcomes provide

Fig. 5 a WTC between GDP and CO<sub>2</sub> emissions. b WTC between GDP and energy consumption. c WTC between GDP and urbanization. d WTC between GDP and gross capital formation



intuition and credibility to the linkage among economic growth and urbanization, energy usage, gross capital formation, and CO<sub>2</sub> emissions.

This research's outcomes have aided us in embracing the promotion of energy intensity diversification of South Korea. This could be achieved by implementing a more ambitious green energy initiative that will maintain the nation's economic momentum. The design and execution of successful policies to regulate South Korean energy and manufacturing sector practices will improve its sustainable growth. This will continue to regulate the CO<sub>2</sub> pollution levels in the nation if the government sets emission restrictions on firms and factories who are emitting CO<sub>2</sub> emissions. The threat of punitive action or high taxes on infringers of this policy will deter environmental pollution. Also, energy usage should be embraced by incorporating sustainable (renewable) energy sources, including hydropower, oceanic, and wind energy sources. Furthermore, South Korea should be careful when formulating policies that will stimulate economic growth at the expense of environmental degradation. Implementing the aforementioned policies will help to maintain sustainable economic development and South Korea's proven environmental performance. This study's outcome could also have a positive impact on neighboring nations who are willing to take the steps suggested in this paper to strengthen their sustainable growth. In conclusion, this study has examined the nexus between energy use, urbanization, energy consumption, gross capital formation, and economic growth in South Korea using recent time series data. Further studies can be conducted for other emerging nations while considering asymmetric in the econometrics modeling or the use of micro-disaggregated data. Furthermore, other studies can account for other drivers of growth that have not been explored in this study.

Acronyms ADF, Augmented Dickey-Fuller; ARDL, Autoregressive distributed lag; BRICS, Brazil, Russia, India, China, and South Africa; CCO2, Consumption-based carbon emissions; CO2, Carbon dioxide; COP21, UN Climate Change Conference in Paris; DOLS, Dynamic ordinary least square; ECM, Error correction model; ECT, Error correction term; EKC, Environmental Kuznets curve; EC, Energy consumption; FMOLS, Fully modified ordinary least square; GCF, Gross capital formation; GMM, General method of moment; GDP, Economic growth; GHGs, Greenhouse gas emissions; GMM, Generalized method of moments; IEA, International Energy Agency; OECD, The Organization for Economic Co-operation and Development; PMG-ARDL, Pooled mean group autoregressive distributed lag; PP, Phillips-Perron; TY, Toda and Yamamoto; URB, Urbanization; VAR, Vector autoregression; ZA, Zivot & Andrews



**Symbols**  $\beta$ , Coefficient of the regressors;  $\rho$ , Speed of adjustment;  $e_b$  Error term;  $ECT_{t-1}$ , Error correction term;  $H_0$ , Null hypothesis;  $H_a$ . Alternative hypothesis

Author contribution Madhy Nyota Mwamba and Tomiwa Sunday Adebayo designed the experiment and collect the dataset. The introduction and literature review sections are written by Tomiwa Sunday Adebayo, Gbenga Daniel Akinsola, and Abraham Ayobamiji Awosusi. Dervis Kirikkaleli and Abraham Ayobamiji Awosusi constructed the methodology section and empirical outcomes in the study. Tomiwa Sunday Adebayo and Gbenga Daniel Akinsola contributed to the interpretation of the outcomes. All the authors read and approved the final manuscript.

Availability of data Data is readily available at https://data.worldbank.org/country/SouthKorea.

### **Declarations**

Ethical approval This study follows all ethical practices during writing.

Consent to participate Not applicable.

Consent to publish Not applicable.

**Competing interests** The authors declare no competing interests.

**Transparency** The authors confirm that the manuscript is an honest, accurate, and transparent account of the study was reported; that no vital features of the study have been omitted; and that any discrepancies from the study as planned have been explained.

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