# Change in mammography screening attendance after removing the out-of-pocket fee: a population-based study in Sweden (2014-2018) 

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#### Abstract

Purpose To assess the change in mammography screening attendance in Sweden-overall and in sociodemographic groups at risk of low attendance-after removal of the out-of-pocket fee in 2016. Methods Individual-level data on all screening invitations and attendance between 2014 and 2018 were linked to sociodemographic data from Statistics Sweden. Odds ratios and $95 \%$ confidence intervals (CIs) for attendance by time period and sociodemographic factor were computed using mixed logistic regression to account for repeated measures within women. The study sample included 1.4 million women, aged 40-75, who had a mammography screening appointment in 2014-2015 and/or 2017-2018 in 14 of Sweden's 21 health care regions. Results Overall screening attendance was $83.8 \%$ in 2014-2015 and $84.1 \%$ in 2017-2018 (+ 0.3 percentage points, $95 \%$ CI $0.2-0.4)$. The greatest increase in attendance was observed in non-Nordic women with the lowest income, where attendance rose from 62.9 to $65.8 \% ~(+2.9$ points, $95 \%$ CI $2.3-3.6$ ), and among women with four or more risk factors for low attendance, where attendance rose from 59.2 to $62.0 \%$ ( +2.8 points, $95 \%$ CI $2.2-3.4$ ). Conclusion Screening attendance did not undergo any important increase after implementing free screening, although attendance among some sociodemographic groups increased by almost three percentage points after the policy change.


Keywords Mammography • Breast cancer screening • Women's health • Socioeconomic aspects of health

## Abbreviations

OR Odds ratio
CI Confidence interval
SEK Swedish kronor (currency)

## Introduction

Most European countries have national population-based programs that offer mammography screening to women in varying age ranges between 40 and 74 [1]. Since the public health impact of population-based screening depends on

[^0]high attendance in order to reduce breast cancer mortality, monitoring, as well as understanding and considering the factors influencing attendance, is important. In the European guidelines for quality assurance in breast cancer screening and diagnosis, attendance is listed as one of the key performance indicators where $>70 \%$ is stated as the acceptable level and $>75 \%$ as the desirable level [2].

Sweden has offered a nationwide outreach mammography screening program since 1997 [3, 4]. Overall attendance is typically about $80 \%$ [1], but is lower in groups who may be socioeconomically vulnerable, such as women who were born abroad [5-8], have a low income [6-8], are unmarried or living without a partner [5, 7-10], are not gainfully employed [5-7, 10], or who have a lower education [5].

As part of the Swedish government's efforts to improve health care equality and women's health, mammography screening became free of charge on July 12016 [11], before which time most regions charged a small out-of-pocket fee ( $\leq 200$ SEK $\approx \$ 23$ USD). According to research conducted in the USA, reducing or removing the out-of-pocket fee can increase screening attendance [12-14]. However, a Swedish
study found no correlation between fees $(0-170$ SEK $\approx$ \$0-20 USD) and attendance (66-91\%) in different regions in Sweden in 1995-1996 [4]. In a more recent study conducted in Stockholm County (Sweden), attendance rose from 68 to $70 \%$ after the screening fee was removed in 2012 [15].

The objective of this study was to assess the change in mammography screening attendance in Sweden-overall and in sociodemographic groups at risk of low attendanceafter removal of the out-of-pocket screening fee in 2016. We examined the change in attendance from 2014-2015 to 2017-2018, stratified by region and sociodemographic factors.

## Materials and methods

This longitudinal population-based register study was conducted in Sweden, where women between the ages of 40 and 74 are invited to mammography screening every 18-24 months depending on age and regional capacity. All invitations are sent by post and offer a pre-booked appointment date and time, which does not need to be confirmed and can be rescheduled or canceled. Since each health care region individually conducts and administers their screening, there are differences in intervals between screening appointments, the layout and content of the invitation letter, hours of operation, ways of canceling or rescheduling appointments, reminders, etc. All but two (Stockholm and Östergötland) of the 21 health care regions charged a small out-of-pocket fee between 80 and 200 SEK ( $\approx \$ 9-23$ USD) before the implementation of free screening in 2016.

A study period between 2014 and 2018 was chosen to study the change in screening attendance during the two-year period before and after removal of the out-of-pocket fee in 2016. Individual screening-related data were extracted for all women invited to the screening program in 15 of 21 health care regions in Sweden. These regions used the same company (Sectra AB) for their radiological information system (RIS) to administer and track invitations, attendance, and results throughout the entire study period, which enabled high quality and consistency of data between regions. Of the six regions initially excluded, four (Jönköping, Kronoberg, Norrbotten, and Uppsala) used different radiological information systems for all or part of the study period, and two (Sörmland and Östergötland) did not grant us permission to extract data. Two of three programs operating in the Stockholm region granted permission. However, these were excluded from the final study sample, since they had already removed the out-of-pocket fee before the study period. This study was approved by the local ethics committee at Lund University (Nos. 2018/576 and 2018/965). Active informed consent as a requirement for data collection was waived.

The extracted data included the screening appointment date, age at the screening appointment, and attendance outcome (attended, canceled, missed, and unavailable), for each regional mammography program separately, and were combined into one dataset. The unique personal identity number assigned to every resident in Sweden was used to merge screening data with information on individual-level sociodemographic characteristics obtained from population registers at Statistics Sweden (the Longitudinal Integration Database for Health Insurance and Labour Market Studies (LISA) [16], the Total Population Register [17], and the Geodatabase [18]). To secure anonymity, Statistics Sweden replaced this number with an arbitrary code before releasing the data to the research group. The most recent sociodemographic information was used for each screening appointment. Sameyear sociodemographic data were linked to each screening appointment in 2014, 2015, and 2017; in 2018, same-year data were available only for home ownership and type of municipality, and data on income, education, and cohabitation from 2017 were used.

Initially, the dataset included a total of 4,582,477 appointments among $1,780,164$ women in 15 regions (including Stockholm). The flow chart in Fig. 1 describes the different steps of exclusion, resulting in a final selection of $2,381,142$ appointments among 1,350,654 women in 14 regions (Fig. 2). These 14 regions encompass about $59 \%$ of the women eligible for mammography screening in Sweden and about $81 \%$ of the women affected by the fee removal. The most recent screening appointment for each woman was selected, aged 40-75, within each time period (2014-2015 and 2017-2018), excluding appointments during the transition year of 2016 . Women 75 years of age were included to allow for overflow from the age limit of 74 years due to administrative reasons, e.g., rescheduling. Appointments were excluded when personal identity numbers lacked a match, had duplicates, or were suspected to have been recycled according to data from Statistics Sweden. Furthermore, appointments with examination or cancelation codes that were not related to mammography screening were excluded. Duplicate appointments within the same program (both identical and non-identical) at different locations and within the same year were excluded as well.

## Outcome variable

The outcome variable in this study was mammography screening attendance (yes/no), irrespective of whether it was the original or a rescheduled appointment date, according to the most recent screening appointment for each woman during the periods 2014-2015 and 2017-2018. The rationale for studying two-year time periods was to allow for longer screening cycles, which is common in several regions.

Fig. 1 Selection of the final study sample


## Sociodemographic and program-related variables

Sociodemographic variables and categorizations are
presented in Table 1 and include age group, cohabitation (in which only couples who have children together were categorized as cohabiting), level of education, income (individual


## Population size

14,194-40,971 54,916-59,074 62,746-65,107 72,971-96,597 286,238-482,383

Region
1 Blekinge
2 Dalarna
3 Gotland
4 Gävleborg
5 Halland
6 Jämtland
7 Jönköping
8 Kalmar
9 Kronoberg
10 Norrbotten
11 Skåne 12 Stockholm 13 Södermanland 14 Uppsala 15 Värmland 16 Västerbotten 17 Västernorrland 18 Västmanlands 19 Västra Götaland 20 Örebro 21 Östergötland

Fig. 2 Map of the Swedish health care regions showing the population of women aged 40-75 in 2018. Excluded regions are hatched
share of equivalized disposable household income in SEK), main source of income, home ownership, country of birth, and type of municipality (based on categorization by the Swedish Association of Local Authorities and Regions) [19]. Program-related variables included region, out-of-pocket fee in 2015, and year of screening appointment. Variables were categorized based on the way in which they were provided by Statistics Sweden, and by logically and conceptually combining categories without losing important differences in the attendance outcome, in order to minimize the number of categories. Missing values, which were excluded from the analyses, ranged from none (region, year, and age) to $1.2 \%$ for education and $2.3 \%$ for home ownership in 2014-2015 and were below $0.5 \%$ for all other variables (Table 1). All variables were analyzed as categorical variables.

## Statistical analysis

Standardized differences were calculated to examine change in the sociodemographic distribution of the study sample between time periods, with a value greater than $10 \%$ considered potentially meaningful. This measure describes the difference between the groups divided by the pooled standard deviation. The percentage of screening appointments attended was calculated for each time period, by region and sociodemographic factor, and reported with $95 \%$ confidence intervals (CIs). Change in attendance before and after
removing the out-of-pocket fee was reported in percentage points with $95 \%$ CIs. We identified six large sociodemographic groups ( $n>50,000$ ) with attendance below $80 \%$ in 2014-2015 and further examined change in attendance before and after the fee removal in each of these groups and combinations thereof. Since the same women can be included in both time periods, mixed logistic regression was used to account for the correlation of observations within individuals. Results are presented as odds ratios (ORs) and 95\% CIs for mammography attendance in 2017-2018 vs. 2014-2015. We calculated unadjusted estimates, as well as estimates adjusted for the potential confounding effect of various sociodemographic factors. Statistical software used for the analyses were SPSS, version 25 , and R , version 4.0.

## Results

The study sample among the 14 included health care regions contained a total of $2,381,142$ appointments among $1,350,654$ women, with $1,191,609$ appointments in 2014-2015 and 1,189,533 appointments in 2017-2018. A total of $1,032,810$ women had an appointment in each time period.

Descriptive characteristics of the study sample in each two-year time period are presented in Table 1. In 2014-2015, the mean age at the time of the screening appointment was 56 years; $61 \%$ of women were living with a partner; $16 \%$ had a low level of education; $12 \%$ got their main income from social assistance and benefits and $85 \%$ from employment or retirement income; $74 \%$ owned their home; $84 \%$ were born in Sweden; $25 \%$ lived in a large city or surrounding commuting areas; and $37 \%$ lived in smaller cities and rural areas. Regions with the largest share of the sample were Västra Götaland (28.7\%) and Skåne (22.3\%). Most proportions were stable between time periods; all of the standardized differences were below $10 \%$.

The overall attendance for all regions combined was $83.8 \%$ in 2014-2015 and $84.1 \%$ in 2017-2018 ( +0.3 percentage points, $95 \%$ CI $0.2-0.4$ ). Attendance per region and by sociodemographic factor are presented in Table 2. There was no change in attendance between 2014-2015 and 2017-2018 in the group of regions with the lowest fees (80-120 SEK) in 2015. A larger increase occurred in the regions that charged 150 SEK than among regions charging 200 SEK. In 2014-2015, Gotland had the lowest attendance ( $81.0 \%$ ), followed by Örebro ( $82.0 \%$ ) and the highest attendance was achieved in Jämtland/Härjedalen (86.1\%), followed by Värmland (85.9\%) and Västmanland (85.5\%). A statistically significant increase in attendance from 2014-2015 to 2017-2018 occurred in seven regions, with the highest increase in Västernorrland (+3.1 points), Västerbotten (+ 1.4 points), and Halland ( +1.3 points). In four regions, there

Table 1 Sociodemographic and other characteristics of the study sample at the time of the most recent mammography screening appointment in Sweden (2014-2015 and 2017-2018)

| Characteristic | $\begin{aligned} & 2014-2015 \\ & N=1,191,609 \end{aligned}$ |  | $\begin{aligned} & 2017-2018 \\ & N=1,189,533 \end{aligned}$ |  | Standardized difference ${ }^{\text {a }}$ \% |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  | $N$ | \% | $N$ | \% |  |
| Mean age (SD) | 56.12 | (10.06) | 56.37 | (10.10) | 7.87 |
| Age group (years) |  |  |  |  |  |
| 40-44 | 192,887 | 16.19 | 185,197 | 15.57 | 1.69 |
| 45-49 | 186,686 | 15.67 | 176,442 | 14.83 | 2.32 |
| 50-54 | 172,539 | 14.48 | 188,375 | 15.84 | 3.78 |
| 55-59 | 164,171 | 13.78 | 162,726 | 13.68 | 0.28 |
| 60-64 | 163,671 | 13.74 | 159,841 | 13.44 | 0.87 |
| 65-69 | 173,291 | 14.54 | 159,718 | 13.43 | 3.22 |
| 70-75 | 138,364 | 11.61 | 157,234 | 13.22 | 4.87 |
| Cohabitation (living with partner) |  |  |  |  |  |
| Yes | 721,892 | 60.58 | 721,839 | 60.68 | 0.21 |
| No | 464,111 | 38.95 | 464,613 | 39.06 | 0.23 |
| Missing | 5,606 | 0.47 | 3,081 | 0.26 | 3.51 |
| Level of education |  |  |  |  |  |
| Low (elementary school, $\leq 9$ years) | 193,073 | 16.20 | 171,287 | 14.40 | 5.01 |
| Intermediate (secondary school) | 550,549 | 46.20 | 542,545 | 45.61 | 1.19 |
| High (post-secondary) | 433,252 | 36.36 | 462,339 | 38.87 | 5.18 |
| Missing | 14,735 | 1.24 | 13,362 | 1.12 | 1.05 |
| Income category ${ }^{\text {b }}$ |  |  |  |  |  |
| Lowest (decile 1) | 118,754 | 9.97 | 118,677 | 9.98 | 0.04 |
| Low-medium (decile 2-4) | 355,513 | 29.83 | 355,832 | 29.91 | 0.17 |
| Medium-high (decile 5-10) | 711,736 | 59.73 | 711,944 | 59.85 | 0.25 |
| Missing | 5,606 | 0.47 | 3,080 | 0.26 | 3.51 |
| Main source of income |  |  |  |  |  |
| Employment | 693,208 | 58.17 | 716,433 | 60.23 | 4.18 |
| Retirement pension | 315,836 | 26.51 | 299,031 | 25.14 | 0.32 |
| Student finance | 4,574 | 0.38 | 4293 | 0.36 | 0.38 |
| Care of a sick child or relative | 5,777 | 0.48 | 6,090 | 0.51 | 0.39 |
| Social assistance and benefits |  |  |  |  |  |
| Sickness benefit | 25,046 | 2.10 | 25,792 | 2.17 | 0.46 |
| Sickness compensation | 74,548 | 6.26 | 67,550 | 5.68 | 1.19 |
| Unemployment insurance/benefit | 7382 | 0.62 | 6298 | 0.53 | 2.44 |
| Labour market program | 15,125 | 1.27 | 14,261 | 1.20 | 0.64 |
| Financial assistance | 20,339 | 1.71 | 23,107 | 1.94 | 1.76 |
| No income | 24,168 | 2.03 | 23,598 | 1.98 | 0.32 |
| Missing | 5,606 | 0.47 | 3,080 | 0.26 | 3.51 |
| Home ownership |  |  |  |  |  |
| Yes (house or apartment) | 883,642 | 74.16 | 882,378 | 74.18 | 1.63 |
| No | 280,078 | 23.50 | 287,870 | 24.20 | 0.05 |
| Missing | 27,889 | 2.34 | 19,285 | 1.62 | 5.16 |
| Country of birth |  |  |  |  |  |
| Sweden | 1,001,904 | 84.08 | 979,249 | 82.32 | 4.70 |
| Nordic country (except Sweden) | 73,001 | 6.13 | 78,051 | 6.56 | 1.79 |
| Europe (except Sweden and Nordic countries) | 72,832 | 6.11 | 92,728 | 7.80 | 6.62 |
| Other | 43,813 | 3.68 | 39,427 | 3.31 | 1.97 |
| Missing | 59 | 0.00 | 78 | 0.01 | 0.21 |
| Region |  |  |  |  |  |
| Blekinge | 32,735 | 2.75 | 33,110 | 2.78 | 0.22 |

Table 1 (continued)

| Characteristic | $\begin{aligned} & 2014-2015 \\ & N=1,191,609 \end{aligned}$ |  | $\begin{aligned} & 2017-2018 \\ & N=1,189,533 \end{aligned}$ |  | Standardized difference ${ }^{\text {a }}$ \% |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  | $N$ | \% | $N$ | \% |  |
| Dalarna | 63,497 | 5.33 | 60,682 | 5.10 | 1.02 |
| Gotland | 13,577 | 1.14 | 13,832 | 1.16 | 0.22 |
| Gävleborg | 62,516 | 5.25 | 60,775 | 5.11 | 0.62 |
| Halland | 64,814 | 5.44 | 65,874 | 5.54 | 0.43 |
| Jämtland/Härjedalen | 25,501 | 2.14 | 26,699 | 2.24 | 0.71 |
| Kalmar | 51,321 | 4.31 | 50,630 | 4.26 | 0.25 |
| Skåne | 265,192 | 22.25 | 271,723 | 22.84 | 1.41 |
| Värmland | 60,219 | 5.05 | 59,993 | 5.04 | 0.05 |
| Västerbotten | 50,257 | 4.22 | 49,076 | 4.13 | 0.46 |
| Västernorrland | 50,328 | 4.22 | 50,988 | 4.29 | 0.31 |
| Västmanland | 53,995 | 4.53 | 47,125 | 3.96 | 2.83 |
| Västra Götaland | 342,100 | 28.71 | 348,371 | 29.29 | 1.27 |
| Örebro | 55,557 | 4.66 | 50,655 | 4.26 | 1.96 |
| Type of municipality |  |  |  |  |  |
| Large cities (> 200,000) ${ }^{\text {c }}$ | 294,396 | 24.71 | 305,857 | 25.71 | 2.32 |
| Mid-sized cities (50,000-200,000) ${ }^{\text {d }}$ | 455,161 | 38.20 | 451,125 | 37.92 | 0.56 |
| Smaller cities, towns, and rural areas | 436,446 | 36.63 | 429,897 | 36.14 | 1.01 |
| Missing | 5,606 | 0.47 | 2,654 | 0.22 | 4.21 |
| Out-of-pocket fee (2015) in SEK ${ }^{\text {e }}$ |  |  |  |  |  |
| 80 | 55,557 | 4.66 | 50,655 | 4.26 | 1.96 |
| 100 | 393,421 | 33.02 | 399,001 | 33.54 | 1.12 |
| 120 | 297,927 | 25.00 | 304,833 | 25.63 | 1.44 |
| 150 | 204,140 | 17.13 | 204,243 | 17.17 | 0.10 |
| 200 | 240,564 | 20.19 | 230,801 | 19.40 | 1.97 |
| Year of scheduled appointment |  |  |  |  |  |
| 2014 | 542,148 | 45.50 |  |  |  |
| 2015 | 649,461 | 54.50 |  |  |  |
| 2017 |  |  | 542,602 | 45.61 |  |
| 2018 |  |  | 646,931 | 54.39 |  |

${ }^{\text {a }}$ The difference between the groups divided by the pooled standard deviation; a value greater than $10 \%$ is interpreted as a meaningful difference
${ }^{\mathrm{b}}$ Income categories for 2014-15: Lowest: $\leq 92,700$ SEK; low-medium: 92,800-150,400 SEK; mediumhigh: $\geq 150,500$ SEK. Income categories for 2017-18: Lowest: $\leq 98,300$ SEK; low-medium: 98,400162,700 SEK; medium-high: $\geq 162,800$ SEK
${ }^{\text {c }}$ Includes commuting zone
${ }^{\mathrm{d}}$ Includes neighboring municipalities
${ }^{\mathrm{e}} 80$ (Örebro), 100 (Kalmar and Västra Götaland), 120 (Blekinge and Skåne), 150 (Dalarna, Halland, Jämtland, and Västernorrland), 200 (Gotland, Gävleborg, Värmland, Västerbotten, and Västmanland)
was a statistically significant decrease in attendance, with the largest decrease in Örebro ( -1.3 points) and Blekinge ( -1.1 points). Attendance according to age groups varied from $82.2 \%$ among women in their forties to $86.0 \%$ among women in their sixties in 2014-2015. The largest increase over time was found among women 70-75 years of age (+ 1.0 points). In 2014-2015, attendance was least prevalent among women who lived without a partner ( $76.9 \%$ ), who had a low level of education level ( $78.0 \%$ ), who were in the lowest income decile ( $71.8 \%$ ), whose main source of income
was not employment or retirement pension (e.g., $68.8 \%$ among women with social assistance or benefits and $56.0 \%$ among those with no income), who did not own their home ( $73.5 \%$ ), and who were not born in Sweden and especially among those born outside of the Nordic countries ( $70.8 \%$ ).

When assessing how the time period associated with attendance among all women in the study sample, the odds of attending were not statistically higher in 2017-2018 than in 2014-2015, neither in an unadjusted analysis (OR 1.01, $95 \%$ CI $0.99-1.02$ ) nor when adjusting for region, age,

Table 2 Mammography screening attendance with 95\% confidence intervals (CI) and change in percentage points (PP) from 2014-2015 to 2017-2018 by sociodemographic factor ( $\mathrm{n}=$ number of attenders)

| Variable | $\begin{aligned} & 2014-2015 \\ & (N=1,191,609) \end{aligned}$ |  |  |  | $\begin{aligned} & 2017-2018 \\ & (N=1,189,533) \end{aligned}$ |  |  |  | Change (95\% CI) |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | $n$ | \% | Lower CI | Upper CI | $n$ | \% | Lower CI | Upper CI | PP | Lower CI | Upper CI |
| Total | 998,533 | 83.80 | 83.73 | 83.86 | 999,913 | 84.06 | 83.99 | 84.13 | 0.26 | 0.17 | 0.36 |
| Age group |  |  |  |  |  |  |  |  |  |  |  |
| 40-44 | 158,851 | 82.35 | 82.18 | 82.52 | 153,410 | 82.84 | 82.66 | 83.01 | 0.48 | 0.24 | 0.72 |
| 45-49 | 153,046 | 81.98 | 81.81 | 82.15 | 145,817 | 82.64 | 82.47 | 82.82 | 0.66 | 0.41 | 0.91 |
| 50-54 | 142,228 | 82.43 | 82.25 | 82.61 | 155,948 | 82.79 | 82.62 | 82.96 | 0.35 | 0.11 | 0.60 |
| 55-59 | 138,116 | 84.13 | 83.95 | 84.31 | 136,605 | 83.95 | 83.77 | 84.13 | -0.18 | -0.43 | 0.07 |
| 60-64 | 139,594 | 85.29 | 85.12 | 85.46 | 136,189 | 85.20 | 85.03 | 85.38 | -0.09 | -0.33 | 0.16 |
| 65-69 | 150,219 | 86.69 | 86.53 | 86.85 | 138,083 | 86.45 | 86.29 | 86.62 | -0.23 | - 0.46 | 0.00 |
| 70-75 | 116,479 | 84.18 | 83.99 | 84.38 | 133,861 | 85.13 | 84.96 | 85.31 | 0.95 | 0.69 | 1.21 |
| Cohabitation |  |  |  |  |  |  |  |  |  |  |  |
| Yes | 639,756 | 88.62 | 88.55 | 88.70 | 639,126 | 88.54 | 88.47 | 88.61 | -0.08 | -0.18 | 0.02 |
| No | 357,050 | 76.93 | 76.81 | 77.05 | 359,585 | 77.39 | 77.27 | 77.51 | 0.46 | 0.29 | 0.63 |
| Level of education |  |  |  |  |  |  |  |  |  |  |  |
| Low | 150,608 | 78.01 | 77.82 | 78.19 | 132,025 | 77.08 | 76.88 | 77.28 | -0.93 | - 1.20 | - 0.66 |
| Intermediate | 465,077 | 84.48 | 84.38 | 84.57 | 458,064 | 84.43 | 84.33 | 84.53 | -0.05 | -0.18 | 0.09 |
| High | 376,740 | 86.96 | 86.86 | 87.06 | 403,215 | 87.21 | 87.12 | 87.31 | 0.26 | 0.12 | 0.39 |
| Income |  |  |  |  |  |  |  |  |  |  |  |
| Lowest decile | 85,270 | 71.80 | 71.55 | 72.06 | 85,670 | 72.19 | 71.93 | 72.44 | 0.38 | 0.02 | 0.74 |
| Decile 2-4 | 289,236 | 81.36 | 81.23 | 81.49 | 290,381 | 81.61 | 81.48 | 81.73 | 0.25 | 0.07 | 0.43 |
| Decile 5-10 | 622,300 | 87.43 | 87.36 | 87.51 | 622,661 | 87.46 | 87.38 | 87.54 | 0.03 | -0.08 | 0.13 |
| Main source of income |  |  |  |  |  |  |  |  |  |  |  |
| Employment/retirement | 877,908 | 87.00 | 86.94 | 87.07 | 882,134 | 86.87 | 86.80 | 86.94 | -0.13 | -0.23 | - 0.04 |
| Student finance | 3,463 | 75.71 | 74.47 | 76.95 | 3,290 | 76.64 | 75.37 | 77.90 | 0.93 | -0.85 | 2.70 |
| Care of sick child/relative | 3,952 | 68.41 | 67.21 | 69.61 | 4,415 | 72.50 | 71.37 | 73.62 | 4.09 | 2.45 | 5.73 |
| Social assistance/benefits | 97,949 | 68.77 | 68.52 | 69.01 | 95,447 | 69.67 | 69.42 | 69.91 | 0.90 | 0.56 | 1.24 |
| No income | 13,534 | 56.00 | 55.37 | 56.63 | 13,426 | 56.89 | 56.26 | 57.53 | 0.89 | 0.01 | 1.78 |
| Home ownership |  |  |  |  |  |  |  |  |  |  |  |
| Yes | 774,600 | 87.66 | 87.59 | 87.73 | 773,668 | 87.68 | 87.61 | 87.75 | 0.02 | -0.08 | 0.12 |
| No | 205,775 | 73.47 | 73.31 | 73.63 | 213,390 | 74.13 | 73.97 | 74.29 | 0.66 | 0.43 | 0.89 |
| Country of birth |  |  |  |  |  |  |  |  |  |  |  |
| Sweden | 861,164 | 85.95 | 85.88 | 86.02 | 845,066 | 86.30 | 86.23 | 86.37 | 0.34 | 0.25 | 0.44 |
| Nordic country | 34,017 | 77.64 | 77.25 | 78.03 | 30,975 | 78.56 | 78.16 | 78.97 | 0.92 | 0.36 | 1.48 |
| Europe | 51,696 | 70.82 | 70.49 | 71.15 | 55,945 | 71.68 | 71.36 | 71.99 | 0.86 | 0.41 | 1.32 |
| Other | 51,625 | 70.88 | 70.55 | 71.21 | 67,883 | 73.21 | 72.92 | 73.49 | 2.32 | 1.89 | 2.76 |
| Region |  |  |  |  |  |  |  |  |  |  |  |
| Blekinge | 28,741 | 87.80 | 87.44 | 88.15 | 28,715 | 86.73 | 86.36 | 87.09 | - 1.07 | -1.58 | -0.56 |
| Dalarna | 53,141 | 83.69 | 83.40 | 83.98 | 51,165 | 84.32 | 84.03 | 84.61 | 0.63 | 0.22 | 1.03 |
| Gotland | 11,002 | 81.03 | 80.37 | 81.69 | 11,368 | 82.19 | 81.55 | 82.82 | 1.15 | 0.23 | 2.07 |
| Gävleborg | 52,158 | 83.43 | 83.14 | 83.72 | 50,837 | 83.65 | 83.35 | 83.94 | 0.22 | -0.20 | 0.63 |
| Halland | 55,339 | 85.38 | 85.11 | 85.65 | 57,119 | 86.71 | 86.45 | 86.97 | 1.33 | 0.95 | 1.70 |
| Jämtland/Härjedalen | 21,969 | 86.15 | 85.73 | 86.57 | 22,963 | 86.01 | 85.59 | 86.42 | -0.14 | -0.74 | 0.45 |
| Kalmar | 43,380 | 84.53 | 84.21 | 84.84 | 42,708 | 84.35 | 84.04 | 84.67 | -0.17 | -0.62 | 0.27 |
| Skåne | 218,281 | 82.31 | 82.17 | 82.46 | 222,339 | 81.83 | 81.68 | 81.97 | -0.48 | -0.69 | -0.28 |
| Värmland | 51,733 | 85.91 | 85.63 | 86.19 | 51,355 | 85.60 | 85.32 | 85.88 | -0.31 | $-0.70$ | 0.09 |
| Västerbotten | 42,857 | 85.28 | 84.97 | 85.59 | 42,554 | 86.71 | 86.41 | 87.01 | 1.43 | 1.00 | 1.87 |
| Västernorrland | 42,277 | 84.00 | 83.68 | 84.32 | 44,425 | 87.13 | 86.84 | 87.42 | 3.13 | 2.69 | 3.56 |
| Västmanland | 46,155 | 85.48 | 85.18 | 85.78 | 40,509 | 85.96 | 85.65 | 86.27 | 0.48 | 0.05 | 0.91 |

Table 2 (continued)

| Variable | $\begin{aligned} & 2014-2015 \\ & (N=1,191,609) \end{aligned}$ |  |  |  | $\begin{aligned} & 2017-2018 \\ & (N=1,189,533) \end{aligned}$ |  |  |  | Change (95\% CI) |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | $n$ | \% | Lower CI | Upper CI | $n$ | \% | Lower CI | Upper CI | PP | Lower CI | Upper CI |
| Västra Götaland | 285,926 | 83.58 | 83.46 | 83.70 | 292,982 | 84.10 | 83.98 | 84.22 | 0.52 | 0.35 | 0.69 |
| Örebro | 45,574 | 82.03 | 81.71 | 82.35 | 40,874 | 80.69 | 80.35 | 81.03 | -1.34 | -1.81 | -0.87 |
| Type of municipality |  |  |  |  |  |  |  |  |  |  |  |
| Large cities | 240,174 | 81.58 | 81.44 | 81.72 | 250,738 | 81.98 | 81.84 | 82.12 | 0.40 | 0.20 | 0.59 |
| Mid-sized cities | 384,838 | 84.55 | 84.44 | 84.65 | 381,547 | 84.58 | 84.47 | 84.68 | 0.03 | -0.12 | 0.18 |
| Smaller cities/rural areas | 371,794 | 85.19 | 85.08 | 85.29 | 366,731 | 85.31 | 85.20 | 85.41 | 0.12 | -0.03 | 0.27 |
| Out-of-pocket fee (2015) |  |  |  |  |  |  |  |  |  |  |  |
| 80 | 45,574 | 82.03 | 81.71 | 82.35 | 40,874 | 80.69 | 80.35 | 81.03 | - 1.34 | -1.81 | -0.87 |
| 100 | 329,306 | 83.70 | 83.59 | 83.82 | 335,690 | 84.13 | 84.02 | 84.25 | 0.43 | 0.27 | 0.59 |
| 120 | 247,022 | 82.91 | 82.78 | 83.05 | 251,054 | 82.36 | 82.22 | 82.49 | -0.56 | -0.75 | -0.36 |
| 150 | 172,726 | 84.61 | 84.46 | 84.77 | 175,672 | 86.01 | 85.86 | 86.16 | 1.40 | 1.18 | 1.62 |
| 200 | 203,905 | 84.76 | 84.62 | 84.90 | 196,623 | 85.19 | 85.05 | 85.34 | 0.43 | 0.23 | 0.63 |

cohabitation, education, income, main source of income, home ownership, and country of birth (OR 1.00, $95 \%$ CI 0.99-1.01).

Table 3 presents attendance before and after the fee removal in sociodemographic sub-groups at risk of low attendance. Although the change between time periods was statistically significant in all six of the main sub-groups, it was small and only exceeded one percentage point for those born in non-Nordic countries-where attendance increased by 1.7 points from $2014-2015$ to 2017-2018. This effect remained statistically significant in the multivariable analysis. Among women with a low level of education, attendance decreased by 0.9 points and this effect persisted in the multivariable analysis. Attendance increased from 59.2 to $62.0 \%$ ( +2.8 points, $95 \%$ CI $2.2-3.4$ ) among women with a combination of any four or more risk factors for low attendance, but did not change substantially among those with three or fewer risk factors. When risk factors were combined, the groups decreased considerably in size and women with four or more risk factors constituted only $3.2 \%$ of all women invited during 2014-2015. When two specific risk factors were combined, the lowest attendance in 2014-2015 was found among women who had the lowest income and were living without a partner (53.0\%) and among women whose main source of income was social assistance and benefits and either lived without a partner ( $60.9 \%$ ) or did not own their home ( $61.2 \%$ ). The biggest increase in attendance was found among non-Nordic women with the lowest income (+ 2.9 points, $95 \%$ CI 2.3-3.6). Other larger sub-groups where attendance increased by more than two points were women born outside of the Nordic countries who were either living alone ( +2.1 points, $95 \%$ CI 1.6-2.6), did not own their home ( +2.0 points, $95 \%$ CI 1.5-2.5), or whose main source of income was social assistance and benefits (+2.1 points,
$95 \%$ CI 1.5-2.7) and women whose main source of income was social assistance and benefits and who did not own their home ( +2.3 points, $95 \%$ CI 1.8-2.8). In all combined risk groups, the association between time period and attendance remained statistically significant in the multivariable analysis.

## Discussion

In this longitudinal population-based register study of mammography screening attendance in Sweden, we found that overall attendance remained relatively stable at $84 \%$ after removal of the out-of-pocket fee in 2016, but rose by 2.9 percentage points among non-Nordic women with the lowest income (from 62.9 to $65.8 \%$ ) and by 2.8 points among women with four or more risk factors for low attendance (from 59.2 to $62.0 \%$ ).

To our knowledge, no other European studies have reported on the effect of removing the fee for mammography screening on screening attendance, except for a study in Stockholm showing an increase of two percentage points after removing the out-of-pocket fee in 2012 [15]. However, a study in Finland found that the likelihood of screening attendance decreased after a fee was introduced, independently of socioeconomic status [20]. A Swedish randomized study examining the effect of offering free cervical cancer screening exams to women in socioeconomically disadvantaged areas failed to establish a statistically significant effect on attendance [21].

Although some US research has demonstrated a positive effect of reducing or removing the out-of-pocket fee on screening attendance [12-14], more recent studies investigating the effect of cost-share removal among Medicare

Table 3 Effect of time period on mammography screening attendance in selected sub-groups with low attendance. Mixed logistic regression analysis of attendance in 2017-2018 vs. 2014-2015

| Sub-group | 2014-2015 |  | 2017-2018 |  | Change (95\% CI) |  |  | UnadjustedOR (95\% CI) | Adjusted <br> OR (95\% CI) |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | $n$ | \% | $n$ | \% | PP | Lower CI | Upper CI |  |  |
| Risk factor for low attendance |  |  |  |  |  |  |  |  |  |
| Living without partner (alone) | 357,050 | 76.93 | 359,585 | 77.39 | 0.46 | 0.29 | 0.63 | 1.02 (1.01-1.03) | $1.01(0.99-1.02)^{\text {a }}$ |
| Low level of education | 150,608 | 78.01 | 132,025 | 77.08 | -0.93 | $-1.20$ | - 0.66 | 0.94 (0.92-0.96) | 0.95 (0.93-0.97) ${ }^{\text {b }}$ |
| Lowest income | 85,270 | 71.80 | 85,670 | 72.19 | 0.38 | 0.02 | 0.74 | 1.02 (0.99-1.04) | $1.02(1.00-1.05)^{\text {c }}$ |
| Social assistance/benefits | 97,949 | 68.77 | 95,447 | 69.67 | 0.90 | 0.56 | 1.24 | 1.05 (1.03-1.07) | $1.04(1.02-1.06)^{\text {d }}$ |
| Not owning home | 205,775 | 73.47 | 213,390 | 74.13 | 0.66 | 0.43 | 0.89 | 1.04 (1.02-1.05) | $1.01(1.00-1.03)^{\text {e }}$ |
| Non-Nordic | 103,321 | 70.85 | 123,828 | 72.51 | 1.66 | 1.34 | 1.97 | 1.08 (1.06-1.10) | $1.05(1.03-1.07)^{\mathrm{f}}$ |
| Index of the above 6 risk factors ${ }^{\text {g }}$ |  |  |  |  |  |  |  |  |  |
| 0 risk factor | 411,288 | 91.63 | 413,250 | 92.03 | 0.41 | 0.29 | 0.52 | 1.12 (1.08-1.15) | 1.11 (1.07-1.15) |
| 1 | 313,432 | 86.11 | 311,851 | 86.11 | 0.00 | -0.16 | 0.16 | 0.98 (0.95-1.01) | 0.97 (0.93-1.00) |
| 2 | 176,436 | 78.15 | 173,189 | 77.95 | -0.20 | - 0.44 | 0.05 | 0.98 (0.97-1.00) | 0.98 (0.96-1.00) |
| 3 | 65,709 | 66.04 | 66,390 | 66.83 | 0.79 | 0.37 | 1.20 | 1.03 (1.01-1.06) | 1.03 (1.01-1.06) |
| 4 | 22,675 | 58.61 | 24,727 | 61.32 | 2.71 | 2.03 | 3.39 | 1.12 (1.08-1.16) | 1.13 (1.10-1.18) |
| 5 | 8,116 | 60.89 | 9,398 | 63.71 | 2.82 | 1.69 | 3.96 | 1.19 (1.10-1.28) | 1.23 (1.14-1.33) |
| 6 | 877 | 59.14 | 1,108 | 62.32 | 3.18 | -0.19 | 6.55 | 1.20 (0.97-1.49) | 1.22 (0.98-1.52) |
| 4 or more | 31,668 | 59.19 | 35,233 | 61.97 | 2.78 | 2.20 | 3.36 | 1.12 (1.09-1.16) | 1.14 (1.10-1.17) |
| 5 or 6 | 8,993 | 60.71 | 10,506 | 63.56 | 2.85 | 1.77 | 3.92 | 1.13 (1.05-1.21) | 1.21 (1.13-1.30) |
| Selected combinations ${ }^{\text {h }}$ |  |  |  |  |  |  |  |  |  |
| Lowest income + living alone | 15,854 | 53.02 | 16,989 | 54.81 | 1.79 | 1.00 | 2.58 | 1.07 (1.03-1.11) | $1.07(1.02-1.12)^{\text {i }}$ |
| Lowest income + social assist | 18,462 | 64.83 | 20,734 | 67.33 | 2.50 | 1.73 | 3.26 | 1.17 (1.12-1.22) | $1.15(1.09-1.21)^{j}$ |
| Lowest income + not owning | 24,956 | 64.28 | 27,945 | 66.26 | 1.98 | 1.32 | 2.63 | 1.09 (1.06-1.13) | $1.10(1.05-1.14)^{\mathrm{k}}$ |
| Non-Nordic + lowest income | 26,128 | 62.87 | 31,038 | 65.79 | 2.92 | 2.29 | 3.55 | 1.14 (1.10-1.17) | $1.17(1.12-1.21)^{1}$ |
| Non-Nordic + living alone | 39,079 | 64.13 | 47,236 | 66.21 | 2.09 | 1.57 | 2.60 | 1.09 (1.06-1.12) | $1.06(1.02-1.09)^{\mathrm{m}}$ |
| Non-Nordic + social assist | 26,803 | 66.85 | 29,459 | 68.96 | 2.11 | 1.47 | 2.75 | 1.10 (1.06-1.14) | $1.13(1.08-1.17)^{\mathrm{n}}$ |
| Non-Nordic + not owning | 44,735 | 66.56 | 54,933 | 68.58 | 2.02 | 1.54 | 2.50 | 1.10 (1.07-1.13) | $1.07(1.04-1.10)^{\circ}$ |
| Social assist + living alone | 48,165 | 60.86 | 48,068 | 62.60 | 1.74 | 1.26 | 2.22 | 1.08 (1.05-1.11) | $1.07(1.04-1.10)^{\mathrm{p}}$ |
| Social assist + not owning | 41,403 | 61.19 | 42,936 | 63.51 | 2.32 | 1.80 | 2.83 | 1.11 (1.08-1.14) | $1.08(1.05-1.11)^{9}$ |

ORs Odds Ratios, CIs $95 \%$ confidence intervals, $n$ number of attenders, $P P$ percentage points
${ }^{\text {a }}$ Estimates were adjusted for region, age group, education, income, main source of income, home ownership, and country of birth
${ }^{\mathrm{b}}$ Estimates were adjusted for region, age group, cohabitation, income, main source of income, home ownership, and country of birth
${ }^{c}$ Estimates were adjusted for region, age group, cohabitation, education, main source of income, home ownership, and country of birth
${ }^{\mathrm{d}}$ Estimates were adjusted for region, age group, cohabitation, education, income, home ownership, and country of birth
${ }^{\mathrm{e}}$ Estimates were adjusted for region, age group, cohabitation, education, income, main source of income, and country of birth
${ }^{\mathrm{f}}$ Estimates were adjusted for region, age group, cohabitation, education, income, main source of income, and home ownership
${ }^{\text {g }}$ Estimates were adjusted for region and age group
${ }^{\mathrm{h}}$ A combination of two risk factors was selected for inclusion where the change exceeded one percentage point and the change in the combined group exceeded the change observed for either individual risk factor. Estimates were adjusted for region, age group, and risk variables not selected for respective subgroup
${ }^{i}$ Estimates were adjusted for region, age group, education, main source of income, home ownership, and country of birth
${ }^{j}$ Estimates were adjusted for region, age group, cohabitation, education, home ownership, and country of birth
${ }^{\mathrm{k}}$ Estimates were adjusted for region, age group, cohabitation, education, main source of income, and country of birth
${ }^{1}$ Estimates were adjusted for region, age group, cohabitation, education, main source of income, and home ownership
${ }^{m}$ Estimates were adjusted for region, age group, education, income, main source of income, and home ownership
${ }^{\mathrm{n}}$ Estimates were adjusted for region, age group, cohabitation, education, income, and home ownership
${ }^{\circ}$ Estimates were adjusted for region, age group, cohabitation, education, income, and main source of income
${ }^{\mathrm{p}}$ Estimates were adjusted for region, age group, education, income, home ownership, and country of birth
${ }^{q}$ Estimates were adjusted for region, age group, cohabitation, education, income, and country of birth
beneficiaries show conflicting results. Some studies found no improvement in uptake in general [22-24], nor in low socioeconomic groups [23, 24], whereas others found an increase in general attendance [25, 26], as well as among unmarried women [25], but not in areas with lower education [26]. Yet another study detected a smaller decrease in attendance in the intervention group compared to a control group, but found no difference by neighborhood socioeconomic status [27].

The lack of change in attendance in the present study is somewhat surprising, considering that the odds of viewing the screening exam as too expensive were threefold among non-attenders compared to attenders according to a previous Swedish cross-sectional study [28]. However, such an opinion, no matter how strong, is just one of many factors and may not ultimately be the deciding one. Other rationales and life circumstances play into the complex decision-making process that some qualitative research studies have depicted [29, 30]. The actual out-of-pocket fee was fairly modest, and there could be other inconveniences and costs that influence the decision to attend and remain obstacles to attendance, especially among socioeconomically vulnerable groups. Such obstacles might include transportation to the screening site, making the effort to reschedule, and taking time off from work [31]. This may explain why only removing the out-of-pocket cost does not appear sufficient to impact uptake in these groups.

## Strengths and limitations

Removal of the screening fee was implemented simultaneously in all health care regions included in this study. However, the absence of control groups or a staggered design, and the lack of means to adequately control for other changes and interventions that may have been introduced in different regions during the study period, limit our ability to evaluate a causal effect of introducing free mammography screening on screening attendance. Another requirement for establishing a causal relationship is evidence of a dose-response relationship, which our results do not demonstrate. Although no change in average attendance was noticeable among regions charging 80-120 SEK, a larger average increase occurred among regions that charged 150 SEK than among those charging 200. An absence of a correlation between the fee and the attendance is corroborated by a Swedish study examining mammography screening attendance in 1995-1996 [4].

This study assesses screening attendance up to and including 2018, i.e., two years after eliminating the screening fee, which may not have been a sufficient amount of time for change to occur. There was no standardized process of informing women about the fee removal across regionsalthough a survey among the regions indicated that most of them at least included information about the screening
exam being free of charge in the letter of invitation-and the level of awareness of this change in the study population is unknown. Such awareness is a necessary condition for establishing a causal effect on attendance.

The screening attendance, based on the actual appointment date within each time period 2014-2015 and 2017-2018, that is presented in this study, might be somewhat higher than screening attendance based on the primary appointment date specified in the invitation. In about $45 \%$ of the cases, these dates were the same (based on data for 2017 and 2018), but a majority of the primary appointment dates were rescheduled. For 2017 and 2018, we had relatively complete data for the primary appointment date, but not for 2014 and 2015. For comparative reasons, we had to use the same definition of attendance for both time periods. Among those who were invited to be screened during 2017-2018, $82.8 \%$ attended within 90 days of the primary appointment, which is 1.3 percentage points lower than the attendance rate that we present. However, we believe that this difference would have been similar in size for the time period 2014-2015.

Despite these limitations, our population study has many strengths, including its large size, covering the majority of Sweden's regional mammography programs. While many studies may struggle with low response rates and selection bias and rely on self-reported data for both exposure and outcome measures, we obtained this information from high-quality register data, originating from several different official sources with high coverage [16], thus minimizing measurement error and misclassification.

## Conclusion

Screening attendance did not undergo any important increase after implementing free screening, although attendance among some sociodemographic groups increased by almost three percentage points after the policy change. These findings may apply to similar population-based screening programs in countries with universal health care, high screening uptake, and small out-of-pocket fees.

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Data availability The data can be made available upon reasonable request.

## Declarations

Conflict of interest The authors declare that they have no conflict of interest.

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