

Erratum to: Utility maximization in incomplete markets with random endowment

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Received: 8 February 2017 / Accepted: 27 March 2017 / Published online: 7 June 2017
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Abstract K. Larsen, M. Soner and G. Žitković kindly pointed out to us an error in our paper (Cvitanić et al. in *Finance Stoch.* 5:259–272, 2001) which appeared in 2001 in this journal. They also provide an explicit counterexample in Larsen et al. (<https://arxiv.org/abs/1702.02087>, 2017).

In Theorem 3.1 of Cvitanić et al. (*Finance Stoch.* 5:259–272, 2001), it was incorrectly claimed (among several other correct assertions) that the value function $u(x)$ is continuously differentiable. The erroneous argument for this assertion is contained in Remark 4.2 of Cvitanić et al. (*Finance Stoch.* 5:259–272, 2001), where it was claimed that the dual value function $v(y)$ is strictly concave. As the functions u and v are mutually conjugate, the continuous differentiability of u is equivalent to the strict convexity of v . By the same token, in Remark 4.3 of Cvitanić et al. (*Finance Stoch.* 5:259–272, 2001), the assertion on the uniqueness of the element \hat{y} in the supergradient of $u(x)$ is also incorrect.

Similarly, the assertion in Theorem 3.1(ii) that \hat{y} and x are related via $\hat{y} = u'(x)$ is incorrect. It should be replaced by the relation $x = -v'(\hat{y})$ or, equivalently, by requiring that \hat{y} is in the supergradient of $u(x)$.

The online version of the original article can be found under doi:[10.1007/PL00013534](https://doi.org/10.1007/PL00013534).

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To the best of our knowledge, all the other statements in Cvitanic et al. (Finance Stoch. 5:259–272, 2001) are correct.

As we believe that the counterexample in Larsen et al. (<https://arxiv.org/abs/1702.02087>, 2017) is beautiful and instructive in its own right, we take the opportunity to present it in some detail.

Keywords Utility maximization · Incomplete markets · Random endowment · Duality

Mathematics Subject Classification (2010) 91B16 · 91G10 · 91G20

JEL Classification G11 · G12 · C61

Erratum to: Finance Stoch (2001) 5:259–272
DOI [10.1007/PL00013534](https://doi.org/10.1007/PL00013534)

1 Discussion

We sketch the counterexample in [4] in a slightly modified and self-contained way. In the sequel, we suppose that the reader is familiar with the paper [1] as well as with [3, Example 5.1 bis]. We briefly recall the notation of that example which is the basis of the counterexample in [4].

The stock price process $S = (S_0, S_1)$ is defined by $S_0 = 1$ and by letting S_1 assume the value $x_0 = 2$ with probability $p_0 = 1 - \alpha$ and, for $n \geq 1$, the value $x_n = \frac{1}{n}$ with probability $p_n = \alpha 2^{-n}$, for $0 < \alpha < 1$ sufficiently small.

For logarithmic utility $U(x) = \ln x$, we obtain, for given endowment $x > 0$, that it is optimal to invest the entire endowment into the stock S so that we end up at time $t = 1$ with the random wealth $\hat{X}(x) = xS_1$. For the value function $u(x)$, we thus obtain (see [3, Example 5.1 bis] for details) the expected utility of $U(\hat{X}(x))$, i.e.,

$$u(x) = \mathbb{E}[U(\hat{X}(x))] = \ln x + \sum_{n=0}^{\infty} p_n \ln x_n. \quad (1.1)$$

Using the notation from [1], the point of this example is that the dual optimizer $\hat{Q}(y)$ is not an element of L^1 , but only of its bidual $(L^\infty)^*$. In other words, $\hat{Q}(y)$ defines a finitely additive probability measure on (Ω, \mathcal{F}) which fails to be countably (or σ -) additive. We write $\hat{Q}(y) = \hat{Q}^r(y) + \hat{Q}^s(y)$ for the decomposition of $\hat{Q}(y)$ into its regular and singular parts.

To verify that $\|\hat{Q}^r(y)\| < \|\hat{Q}(y)\|$, fix $x = y = 1$. Writing \hat{X} for $\hat{X}(1)$, \hat{Q} for $\hat{Q}(1)$ and q_n for $\hat{Q}[S_1 = x_n]$, we have the relation [3, Theorem 2.2]

$$U'(\hat{X}) = \frac{d\hat{Q}^r}{d\mathbb{P}}, \quad (1.2)$$

so that

$$(x_n)^{-1} = \frac{q_n}{p_n}, \quad n \geq 0,$$

which yields

$$\sum_{n=0}^{\infty} q_n = \frac{1 - \alpha}{2} + \alpha \sum_{n=1}^{\infty} np_n.$$

This term is smaller than 1 (recall that $0 < \alpha < 1$ is small), which readily shows that the regular part \hat{Q}^r has a smaller mass than \hat{Q} .

So far, we have just recalled [3, Example 5.1 bis]. For the next step, we follow [2] and distinguish between the odd and even numbers $n \in \mathbb{N}_0$ to define

$$A = \{S_1 = x_n \text{ for odd } n \geq 1\}, \quad B = \{S_1 = x_n \text{ for even } n \geq 0\}.$$

Now comes the beautiful idea from [4]. Define the process $\tilde{S} = (\tilde{S}_0, \tilde{S}_1)$ by setting $\tilde{S}_0 = S_0 = 1$ and

$$\tilde{S}_1 = S_1 + \frac{1}{2}\mathbb{1}_A. \tag{1.3}$$

We also consider the random variable e_T which we define as

$$e_T = -\frac{1}{2}\mathbb{1}_A. \tag{1.4}$$

For the utility maximization problem, subject to the additional random endowment e_T , we define the value function $\tilde{u}(x)$ as in [1]. This boils down to the formula

$$\tilde{u}(x) = \sup_{\lambda \in \mathbb{R}} \mathbb{E}[\ln(x + \lambda(\tilde{S}_1 - \tilde{S}_0) + e_T)]. \tag{1.5}$$

For example, for $x = 1$, we again find that the optimizer $\hat{\lambda}$ in (1.5) equals 1, i.e., it again is optimal to invest the entire initial endowment $x = 1$ into the stock \tilde{S} ; in this case, the terms $\mathbb{1}_A$ in (1.3) and (1.4) cancel out perfectly. In fact, for all $x \geq 1$, we find that the optimal $\hat{\lambda}$ in (1.5) equals $\hat{\lambda}(x) = x$, just as in (1.1). Indeed, for $x \geq 1$, the crucial constraint is that we cannot invest more than the amount $\hat{\lambda} = x$ into the stock due to the definition of \tilde{S} and e_T on the set B , which corresponds to the even numbers $n \in \mathbb{N}_0$.

On the other hand, for $x < 1$, the picture changes; now the binding constraint is given by the odd numbers $n \in \mathbb{N}_0$, i.e., by the behaviour of \tilde{S} and e_T on the set A . For $\frac{1}{2} < x \leq 1$, we find that the optimizer $\hat{\lambda}$ in (1.5) equals $\hat{\lambda}(x) = 1 - 2(1 - x) = 2x - 1$. The remaining amount $x - (2x - 1) = 1 - x$ of the initial wealth x is kept in the bond. Note that for $x \leq \frac{1}{2}$, there is no admissible solution λ in (1.5), i.e., $\tilde{u}(x) = -\infty$ in (1.5). We thus obtain for the optimal terminal wealth $\hat{X}(x) = x + \hat{\lambda}(\tilde{S}_1 - \tilde{S}_0) + e_T$, for $\frac{1}{2} < x \leq 1$,

$$\hat{X}(x) = \begin{cases} (2x - 1)x_n + (1 - x) & \text{for } n \text{ even,} \\ (2x - 1)(x_n + \frac{1}{2}) + (\frac{1}{2} - x) & \text{for } n \text{ odd} \end{cases} \tag{1.6}$$

and, for $1 \leq x < \infty$,

$$\hat{X}(x) = \begin{cases} xx_n & \text{for } n \text{ even,} \\ x(x_n + \frac{1}{2}) - \frac{1}{2} & \text{for } n \text{ odd.} \end{cases} \tag{1.7}$$

Note that we always have that $\hat{X}(x)$ is an a.s. strictly positive random variable whose essential infimum is zero. The latter property is obtained by considering the sets $\{S_n = x_n\}$ with n tending to infinity, where we have to consider the odd n in the case $\frac{1}{2} < x \leq 1$, and the even n in the case $x \geq 1$.

Clearly, the definitions (1.6) and (1.7) coincide for $x = 1$ in which case we obtain $\hat{X}(1) = S_1$. In particular, the value function

$$\tilde{u}(x) = \mathbb{E}[\ln \hat{X}(x)]$$

is continuous at $x = 1$, as must be the case. We shall see that \tilde{u} has a kink at $x = 1$. Indeed, we may calculate the derivative of $\tilde{u}(x)$, for $x \in (\frac{1}{2}, 1)$ as well as for $x \in (1, \infty)$, by using the formula

$$\frac{d}{dx} \tilde{u}(x) = \mathbb{E} \left[\frac{d}{dx} \ln \hat{X}(x) \right], \quad x \in (1/2, 1) \cup (1, \infty).$$

Hence the difference $\Delta \tilde{u}'(1) = \lim_{x \searrow 1} (\frac{d}{dx} \tilde{u}(x)) - \lim_{x \nearrow 1} (\frac{d}{dx} \tilde{u}(x))$ of the right and left derivatives of \tilde{u} at $x = 1$ can be explicitly computed as

$$\Delta \tilde{u}'(1) = \sum_{n=0}^{\infty} p_n \left(\frac{1}{x_n} - \frac{2}{x_n} \right) = - \left(\frac{p_0}{2} + \sum_{n=1}^{\infty} n p_n \right),$$

which clearly shows that the function $\tilde{u}(x)$ fails to be differentiable at $x = 1$.

Summing up, following [4], we have constructed an example where the value function $\tilde{u}(\cdot)$ fails to be differentiable.

We still want to have a closer look at the dual problem associated to the above example. In particular, we want to spot precisely where the erroneous argument in [1] has occurred. Define

$$y_1 = \lim_{x \searrow 1} \left(\frac{d}{dx} \tilde{u}(x) \right) \quad \text{and} \quad y_2 = \lim_{x \nearrow 1} \left(\frac{d}{dx} \tilde{u}(x) \right).$$

As the dual value function \tilde{v} (see [1] for the definition)

$$\tilde{v}(y) = \min_{Q \in \mathcal{D}} \left\{ \mathbb{E} \left[V \left(y \frac{dQ^r}{d\mathbb{P}} \right) \right] + y \langle Q, e_T \rangle \right\} \tag{1.8}$$

is conjugate to \tilde{u} (see [1]), we know from the fact that $\tilde{u}(x)$ has a kink at $x = 1$ that $\tilde{v}(y)$ is an affine function with slope -1 on the interval $[y_1, y_2]$, in view of the basic relation

$$\tilde{v}(y) = \sup_x (\tilde{u}(x) - xy).$$

What are the dual optimizers \hat{Q}_y for $y \in [y_1, y_2]$, given by [1, Theorem 3.1 and Lemma 4.1]? We know from [1] that the regular parts \hat{Q}_y^r are unique and given by the formula

$$U'(\hat{X}(x)) = y \frac{d\hat{Q}_y^r}{d\mathbb{P}},$$

as in (1.2) above. The number x is associated to y via the relation $-\tilde{v}'(y) = x$, which yields $x = 1$ for $y \in [y_1, y_2]$. This implies the amazing fact that *the regular parts $y\hat{Q}_y^r$ of the dual optimizers $y\hat{Q}_y$ are identical, for all $y \in [y_1, y_2]$* . Note that the total mass of the elements $y\hat{Q}_y \in (L^\infty)^*$ equals $\|y\hat{Q}_y\| = y$. If we pass as usual to the normalized finitely additive probability measures \hat{Q}_y , their regular parts \hat{Q}_y^r scale by the factor y^{-1} .

As regards the singular part \hat{Q}_y^s of \hat{Q}_y , it is clear that \hat{Q}_y^s is supported by each of the sets

$$C_N = \bigcup_{n=N}^\infty \{S_n = x_n\}.$$

Indeed, for each $\epsilon > 0$, the singular measure \hat{Q}_y^s is supported by the set $\{\hat{X}(x) < \epsilon\}$, where $-\tilde{v}'(y) = x$. This follows from the analysis in [3, Example 5.1 bis]. But now the additional aspect of the odd and even n arises: How much of this singular mass sits on $C_N \cap A$, and how much on $C_N \cap B$?

It follows from (1.6), (1.7) and the subsequent discussion that for $\frac{1}{2} < x < 1$ and $\tilde{u}'(x) = y$, the singular measure \hat{Q}_y^s is supported by A , while for $1 < x < \infty$ and $\tilde{u}'(x) = y$, the singular measure \hat{Q}_y^s is supported by B . One may also pass to the limits $x \nearrow 1$ and $x \searrow 1$ to show that $\hat{Q}_{y_1}^s$ is supported by A , while $\hat{Q}_{y_2}^s$ is supported by B . It turns out that for general $y \in [y_1, y_2]$ of the form $y = \mu y_1 + (1 - \mu)y_2$, we have the affine relations

$$y\hat{Q}_y^s[A] = \mu y_1 \hat{Q}_{y_1}^s[A] \tag{1.9}$$

and

$$y\hat{Q}_y^s[B] = (1 - \mu)y_2 \hat{Q}_{y_2}^s[B]. \tag{1.10}$$

Indeed, because $\mathbb{E}_Q[\tilde{S}_1 - \tilde{S}_0] = 0$ for any equivalent martingale measure Q for \tilde{S} , we also obtain $\langle \hat{Q}_y, \tilde{S}_1 - \tilde{S}_0 \rangle = 0$ by weak-star continuity. As $y\hat{Q}_y^r$ does not depend on y , we obtain that $\langle y\hat{Q}_y^s, \tilde{S}_1 - \tilde{S}_0 \rangle$ does not depend on y either, for $y \in [y_1, y_2]$. On the set $C_N \cap A$ (resp. $C_N \cap B$), the random variable $\tilde{S}_1 - \tilde{S}_0$ equals $-\frac{1}{2}$ (resp. -1), up to an error of at most $\frac{1}{N}$ which disappears in the limit $N \mapsto \infty$. This implies that $-\frac{1}{2}y\hat{Q}_y^s[A] - y\hat{Q}_y^s[B]$ is constant when y varies in $[y_1, y_2]$ and readily yields the affine relations (1.9) and (1.10).

Finally, let us have a closer look where the mistake in Remark 4.2 of [1] occurred. In that argument, we fixed numbers $0 < y_1 < y_2$ (which may or may not coincide with the y_1, y_2 considered above) and considered the value function $\tilde{v}(y)$ as in (1.8). For $y_1 \neq y_2$, we have that $y_1\hat{Q}_{y_1}$ is different from $y_2\hat{Q}_{y_2}$, as shown in [1]. Up to

this point, the reasoning was correct. We then tacitly (and incorrectly) assumed that this implies that their regular parts $y_1 \hat{Q}_{y_1}^r$ and $y_2 \hat{Q}_{y_2}^r$ must be different too! This would imply the strict inequality claimed in Remark 4.2 of [1]. But as we just have seen, it may happen that these two measures coincide. In addition, the singular parts $y \hat{Q}_y^s$ satisfy the affine relations (1.9) and (1.10), which also prevent the inequality in Remark 4.2 of [1] becoming strict.

Acknowledgements We thank K. Larsen, M. Soner and G. Žitković for providing this illuminating counterexample which we expect to have applications and allow additional insight also beyond the present context.

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