ORIGINAL ARTICLE



# Effects of non-invasive ventilation in patients with acute respiratory failure excluding post-extubation respiratory failure, cardiogenic pulmonary edema and exacerbation of COPD: a systematic review and meta-analysis

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Received: 9 April 2017 / Accepted: 10 July 2017 / Published online: 24 July 2017 © Japanese Society of Anesthesiologists 2017

#### Abstract

*Background* This meta-analysis compared the effects of non-invasive ventilation (NIV) with invasive mechanical ventilation (InMV) and standard oxygen ( $O_2$ ) therapy on mortality and rate of tracheal intubation in patients presenting acute respiratory failure (ARF).

*Methods* We searched the MEDLINE, EMBASE and Cochrane Central Register of clinical trials databases between 1949 and May 2015 to identify randomized trials of NIV for ARF. We excluded the ARF caused by extubation, cardiogenic pulmonary edema, and COPD.

*Results* The meta-analysis included 21 studies and 1691 patients, of whom 846 were assigned to NIV and 845 to control (InMV or standard  $O_2$  therapy). One hundred ninety-one patients (22.6%) in the NIV group and 261 patients (30.9%) in the control group died before discharge from hospital. The pooled odds ratio (OR) for

**Electronic supplementary material** The online version of this article (doi:10.1007/s00540-017-2389-0) contains supplementary material, which is available to authorized users.

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short-term mortality (in-hospital mortality) was 0.56 (95% CI 0.40–0.78). When comparing NIV with standard  $O_2$  therapy, the short-term mortality was 155 (27.4%) versus 204 (36.0%), respectively. For this comparison, the pooled OR of short-term mortality was 0.56 (95% CI 0.36–0.85). When comparing NIV with InMV, the short-term mortality was 36 (12.9%) versus 57 (20.5%) patients, respectively. For this comparison, the pooled OR of short-term mortality was 0.56 (95% CI 0.34–0.90). Tracheal intubation was performed in 106 patients (22.7%) in the NIV and in 183 patients (39.4%) in the standard  $O_2$  group, representing a pooled OR of 0.37 (95% CI 0.25–0.55). There were publication biases and the quality of the evidence was graded as low.

Conclusion Compared with standard  $O_2$  therapy or InMV, NIV lowered both the short-term mortality and the rate of tracheal intubation in patients presenting with ARF.

**Keywords** Non-invasive ventilation · Invasive ventilation · Acute respiratory failure · Tracheal intubation · Survival analysis

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#### Abbreviations

ALI	Acute lung injury
ARDS	Acute respiratory distress syndrome
ARF	Acute respiratory failure
CI	Confidence interval
COPD	Chronic obstructive pulmonary disease
CPAP	Continuous positive airway pressure
CPE	Cardiogenic pulmonary oedema
GRADE	Grading of Recommendations Assessment,
	Development and Evaluation
NIV	Non-invasive ventilation
OR	Odds ratio
PEEP	Positive end expiratory pressure
PERF	Post-extubation respiratory failure

# Introduction

Acute respiratory failure (ARF) is a widely prevalent medical emergency, which must be treated in a timely manner to prevent developments of life-threatening hypoxia or/ and hypercapnia [1]. Non-invasive ventilation (NIV) is one of the treatments of hypoxic respiratory failure, which effectively improves the gas exchange in selected patients although it is still unknown if it affects the prognosis [2].

The first report of NIV, used in 10 patients presenting with ARF due to intrinsic diseases, was published in 1989 [3]. Since then, NIV has been widely used because it did not require tracheal intubations and/or specialized medical personnel, which seemed to save time and lower costs. The mask was generally well tolerated and there were no obvious complications, such as vomiting or aspiration. In addition, the physiologic improvements observed in that study were similar to those achieved with intubation and mechanical ventilation. After the benefits conferred by NIV observed in that study, several randomized controlled trials (RCTs) confirmed that NIV prevented the need for tracheal intubation and increased the blood oxygen (O2) concentration in hypoxic patients [4–6]. However, the effects of NIV on mortality rate have remained unknown. Some have claimed that it increases the survival rate in the acute care setting, while others have stated that it does not increase survival since failure of NIV management is associated with a significantly higher mortality [7–12]. The controversial results of the previous studies regarding the potential benefits of NIV may be explained by the variable degree of hypoxia. Systematic reviews to assess the efficacy of NIV for chronic obstructive pulmonary disease (COPD) or cardiogenic pulmonary edema (CPE) demonstrated its efficacy and safety [13, 14]. However, it has remained unknown whether NIV is effective among patients with acute respiratory failure, excluding ARF due to these etiologies.

This study examined whether NIV increases the survival rate of patients presenting with ARF excluding COPD and CPE.

The efficacy of NIV is mainly attributed to increasing lung volumes and decreasing work of breathing of the patients [15]. It is, therefore, currently considered first-line treatment of disorders such as post-extubation respiratory failure (PERF), CPE, and exacerbation of COPD, in which several prospective studies have confirmed its efficacy [16–18]. NIV lowered the risk of intubation by 65%, and the length of hospitalization by 1.9 days compared with InMV [16]. However, these distinct disorders, particularly CPE, are triggered by cardiac diseases rather than by respiratory failure. PERF and COPD are not limited to lung tissue. Therefore, after excluding these 3 etiologies of ARF, we want to discuss the majority of patients who presented with "pure" ARF: this is an important difference compared to other previous systematic reviews [9, 10].

NIV is contraindicated in patients presenting with respiratory arrest or upper airway obstructions when: the mask does not fit or the secretions cannot be managed properly; compliance with the mask is poor; or there is instability in hemodynamic status. Several complications associated with failure of treatment have also been reported, such as skin lesions or major air leaks [7, 19]. The failure of NIV may influence the intubation rate and mortality, and is more prevalent in ARF complicated by hypoxia [20]. The NIV failure rate has ranged from 10 to 40% among various studies [21], suggesting that its outcome is influenced by personal experiences or techniques. Its effectiveness and the benefit it has conferred on survival has, indeed, been variable among medical centers and experimental protocols designed to standardize its use and mitigate the variability of judgment and decision-making among caregivers [22]. Consequently, a meta-analysis of randomized trials seemed the best means of resolving differences attributable to the temporal variability in the collection of data by different institutions and countries. The aim of this meta-analysis was to examine the effects of NIV on mortality and tracheal intubation rate in ARF not due to PERF, CPE or exacerbation of COPD, with a view to help caregivers choose an optimal first-line of treatment in specific clinical settings.

# Methods

#### Data sources and search strategies

We searched the MEDLINE, EMBASE and the Cochrane Central Register databases of clinical trials, published between January 1949 and 6 May 2015, January 1949 and 2 June 2015, and January 1949 and 1 June 2015, respectively.

The full search strategies for each database are described in the Online Supplement 1.

# Study selection

The titles and abstracts of references retrieved from the databases, and literature searches were independently conducted by the four investigators (YK, JK, AK, RS), prior to the full article reviews. Divergences of opinion were resolved by consensus. We used the following criteria to identify studies to be included: (1) randomized trial design; (2) inclusion of patients presenting with ARF and hypoxemia, defined by each study; (3) comparisons of NIV with mechanical ventilation or standard O<sub>2</sub> therapy for treatment of ARF. The exclusion criteria for studies were: (1) < 18 years of age; (2) patients with CPE as a single etiology; (3) patients with exacerbation of COPD as a single etiology; (4) ARF following extubation; (5) studies performed in an ambulatory setting; (6) studies published in a language other than English.

#### **Data extraction**

The four investigators independently extracted the data from each eligible study. We contacted the corresponding authors of eligible articles via e-mail to request missing data. The data extracted included: author, year of publication, study design, number of patients, interventions (NIV, invasive mechanical ventilation and standard  $O_2$  therapy), outcome measures and study results, including (a) shortterm mortality, defined as death in the intensive care unit or in the hospital, and (b) tracheal intubation. Divergences of opinion were resolved by consensus.

### Study endpoints

The primary outcome was short-term mortality as defined earlier in each of the following comparisons: (1) NIV versus InMV or standard  $O_2$  therapy, (2) NIV versus standard  $O_2$  therapy defined as any oxygen concentration delivered by mask, (3) NIV versus InMV. The secondary outcome was tracheal intubation rate in the non-invasive versus standard  $O_2$  therapy comparison.

# Assessment of methodological quality: risk of bias assessment and GRADE approach

We adapted the Cochrane risk of bias tool to assess the quality of the studies included in the meta-analysis [23, 24]. Each study was assessed for: (1) random sequence generation (selection bias); (2) allocation concealment (selection bias); (3) blinding of participants and staff (performance bias); (4) blinding of related outcomes

assessment (detection bias); (5) incomplete outcome data (attrition bias); (6) selective reporting (reporting bias); (7) other biases. We classified the studies as low, intermediate or high risk of bias in each domain. In addition, we graded the quality of evidence of each finding based on the criteria established by the Grading of Recommendations Assessment, Development and Evaluation (GRADE) working group [25]. The quality of the study methodology was classified by the four independent investigators as high, intermediate, low or very low, based on the study design, risk of biases, indirectness, inconsistency, imprecision and publication bias. The publication biases were assessed visually by inspecting funnel plots as well as analytical appraisals based on Eggar's linear regression test [26]. A two-sided pvalue  $\leq 0.10$  was regarded as significant in Eggar's linear regression test.

#### Statistical analysis

We pooled the eligible patients for each outcome and calculated the odds ratios (OR) and corresponding 95% confidence intervals (CI), using the Der Simonian–Laird random-effects model with weights calculated by the inverse variance method. We verified the heterogeneity of the studies, using the estimated Cochrane chi-square test and the  $l^2$ statistic with  $l^2 > 50\%$ .

We applied unadjusted p values for the significance assessment in this study. It was set at the two-tailed 0.05 level for hypothesis testing and at the 0.10 level for testing of heterogeneity.

The meta-analyses were performed using the Review Manager, Cochrane systematic review software, version 5.3.5 for Windows (Copenhagen: The Nordic Cochrane Centre; The Cochrane Collaboration, 2014, http://tech. cochrane.org/revman). The publication biases were analyzed with Stata version 13<sup>®</sup> (Stata Corp LP, 2013).

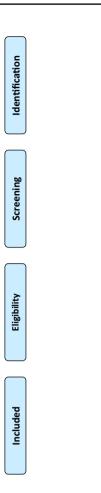
# Results

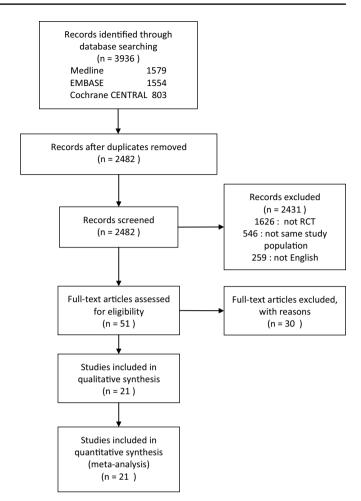
We identified 2482 studies from the electronic databases after elimination of duplicates. We excluded 1626 studies because their design was not randomized, 546 studies because the patients did not fit our selection criteria, and 259 studies because they were not published in English. We retained 51 studies for review of the full-length reports, and included 21 studies [5, 27–46] in the final analysis (Fig. 1).

### **Study characteristics**

The 21 trials [5, 27–46] included 1691 patients, of whom 846 were randomly assigned to NIV and 845 to controls. In 17 trials [5, 28–33, 35–37, 39, 41–46] 566 patients

Fig. 1 Flow of study between preliminary searches of databases and selection of the 21 articles entered in the metaanalysis





were assigned to NIV and 567 were assigned to standard O<sub>2</sub> therapy. In four trials [27, 34, 38, 40] 280 patients were assigned to NIV and 278 to invasive ventilation. In all trials, patients were assigned each intervention as an initial therapy for respiratory failure, and patients who received NIV or InMV before assignment were excluded from each study. Among the 17 trials, which compared NIV with standard O<sub>2</sub> therapy, two [31, 35] reported no death. Of these 17 trials, 16 [5, 28-33, 35-37, 39, 42-46] reported tracheal intubation as a secondary outcome and two [31, 35] of these 16 trials reported no instance of intubation. The non-invasive ventilators used were BiPAP Vision® (Respironics Inc. Koninklijke Philips N.V., Amsterdam, The Netherlands) in 6 studies (28.6%), Puritan Bennett<sup>TM</sup> 7200<sup>®</sup> (Covidian, Minneapolis, MN) in 5 (23.8%), and Dräger Evita<sup>®</sup> Infinity<sup>®</sup> V500 (Drägerwerk AG & Co. KGaA, Lübeck, Germany) in 4 studies (19.0%). The NIV interfaces were full-face masks in 11 studies (52.4%), nasal masks in 5 (23.8%), face masks in 4 (19.1%), helmets in 3 (14.3%), and an oro-nasal mask in 1 study (4.8%). The NIV mode was BiPAP in 15 (71.4%) and CPAP in 6 studies (28.6%). The individual characteristics of the trials included in this meta-analysis are detailed in Table 1.

# Outcomes

In the NIV group, 191 patients (22.6%) died in the hospital, before or after leaving the intensive care unit, while 261 patients (30.9%) died in the hospital during standard  $O_2$  therapy or InMV. The pooled OR of short-term mortality (Fig. 2) was 0.56 (95% CI 0.40–0.78). When comparing NIV with standard  $O_2$  therapy, the short-term mortality was 155 patients (27.4%) versus 204 (36.0%) patients, respectively. For this comparison (Fig. 3), the pooled OR of short-term mortality was 0.56 (95% CI 0.36–0.85). When comparing NIV with InMV, the short-term mortality was 36 patients (12.9%) versus 57 (20.5%) patients, respectively. For this comparison (Fig. 4), the pooled OR of short-term mortality was 0.56 (95% CI 0.34–0.90). Tracheal intubation (Fig. 5) was performed in 106 patients (22.7%) in the

# Table 1 Detail of included studies

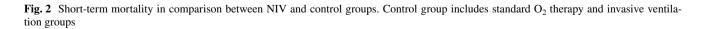
Interventions	9							
First author, year	Country	Number of study partici- pants	Experi- mental group	Control group	NIV ventilator	NIV interface	NIV mode	Outcomes
Antonelli, 2000	Italy	40	NIV	Standard oxygen therapy	Bennett 7200, Puritan or Servo 900 C, Siemens	Full face mask	BiPAP	Mortality and tracheal intu- bation
Antonelli, 1998	Italy	64	NIV	Invasive ventila- tion	Bennett 7200, Puritan or Servo 900 C, Siemens	Full face mask	BiPAP	Mortality
Brambilla, 2014	Italy	81	NIV	Standard oxygen therapy	High-flow generator (90–140 l/min; VitalSigns Inc.)	Helmet	СРАР	Mortality and tracheal intu- bation
Confalonieri, 1999	Italy	56	NIV	Standard oxygen therapy	Cesar Thaema, Bennett 7200 Puritan, Vential Saime, or Servo 900 C Siemens	Full face mask	BiPAP	Mortality and tracheal intu- bation
Cosentini, 2010	Italy	47	NIV	Standard oxygen therapy	High-flow generator (90–140 l/min; VitalSigns Inc.)	Helmet	СРАР	Mortality and tracheal intu- bation
Delclaux, 2000	France	123	NIV	Standard oxygen therapy	High-flow generator (90–140 l/min; VitalSigns Inc.)	Full face mask	СРАР	Mortality and tracheal intu- bation
Ferrer, 2003	Spain	105	NIV	Standard oxygen therapy	BiPAP Vision; Respironics Inc.	Full face mask or nasal mask	BiPAP	Mortality and tracheal intu- bation
Gunduz, 2005	Turkey	43	NIV	Invasive ventila- tion	Evita 4, Drager	Face mask	CPAP	Mortality
Gupta, 2010	India	53	NIV	Standard oxygen therapy	Servo-i, Maquet	Oronasal mask	BiPAP	Mortality and tracheal intu- bation
Hernandez, 2010	Spain	50	NIV	Standard oxygen therapy	BiPAP Vision; Respironics Inc.	Full face mask or face mask	BiPAP	Mortality and tracheal intu- bation
Hilbert, 2001	France	52	NIV	Standard oxygen therapy	Evita, Drager	Full face mask	BiPAP	Mortality and tracheal intu- bation
Honrubia, 2005	Spain	64	NIV	Invasive ventila- tion	Evita, Drager	Face mask	BiPAP	Mortality
Kramer, 1995	America	31	NIV	Standard oxygen therapy	BiPAP Vision; Respironics Inc.	Nasal mask	BiPAP	Mortality and tracheal intu- bation
Martin, 2000	America	61	NIV	Standard oxygen therapy	BiPAP Vision; Respironics Inc.	Nasal mask	BiPAP	Mortality and tracheal intu- bation
Matic, 2007	Croatia	387	NIV	Invasive ventila- tion	Evita, Drager, or Bennett 7200, Puritain	Nasal mask and face mask	BiPAP	Mortality

# Table 1 (continued)

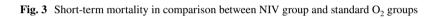
Interventions								
First author, year	Country	Number of study partici- pants	Experi- mental group	Control group	NIV ventilator	NIV interface	NIV mode	Outcomes
Nava, 2013	Italy	200	NIV	Standard oxygen therapy	PV 102, Breas Medical	Full face mask	CPAP	Mortality
Squadrone, 2010	Canada	40	NIV	Standard oxygen therapy	High-flow generator (Whisperflow, Caradyne)	Helmet	CPAP	Mortality and tracheal intubation
Wermke, 2012	Germany	86	NIV	Standard oxygen therapy	VS-Integra, Saime or Respicare SC, Drager	Full face mask	BiPAP	Mortality and tracheal intubation
Wood, 1998	America	27	NIV	Standard oxygen therapy	BiPAP Vision; Respironics Inc.	Nasal mask	BiPAP	Mortality and tracheal intubation
Wysocki, 1995	France	41	NIV	Standard oxygen therapy	Bennett 7200, Puritain	Full face mask	BiPAP	Mortality and tracheal intubation
Zhan, 2012	China	40	NIV	Standard oxygen therapy	BiPAP Vision; Respironics Inc.	Full face mask	BiPAP	Mortality and tracheal intubation

NIV non-invasive ventilation

	Experim	nental	Contr	ol		Odds Ratio	Odds Ratio
Study or Subgroup	Events	Total	Events	Total	Weight	IV, Random, 95% CI	IV, Random, 95% CI
Antonelli 1998	9	32	15	32	6.4%	0.44 [0.16, 1.25]	
Antonelli 2000	7	20	11	20	4.9%	0.44 [0.12, 1.57]	
Brambilla 2014	2	40	7	41	3.3%	0.26 [0.05, 1.32]	
Confalonieri 1999	7	28	6	28	5.0%	1.22 [0.35, 4.24]	
Cosentini 2010	0	20	0	27		Not estimable	
Delclaux 2000	19	61	18	62	9.0%	1.11 [0.51, 2.39]	
Ferrer 2003	9	51	21	54	7.6%	0.34 [0.14, 0.83]	
Gunduz 2005	2	22	7	21	3.1%	0.20 [0.04, 1.11]	
Gupta 2010	0	28	0	25		Not estimable	
Hernandez 2010	1	25	1	25	1.3%	1.00 [0.06, 16.93]	
Hilbert 2001	12	26	18	26	5.7%	0.38 [0.12, 1.19]	
Honrubia 2005	10	31	14	33	6.6%	0.65 [0.23, 1.80]	
Kramer 1995	1	16	2	15	1.6%	0.43 [0.04, 5.35]	
Martin 2000	5	32	10	29	5.2%	0.35 [0.10, 1.20]	
Matic 2007	15	195	21	192	9.9%	0.68 [0.34, 1.36]	
Nava 2013	61	99	66	101	11.4%	0.85 [0.48, 1.51]	— <u> </u>
Squadrone 2010	3	20	15	20	3.5%	0.06 [0.01, 0.29]	
Wermke 2012	16	42	14	44	7.7%	1.32 [0.54, 3.21]	
Wood 1998	4	16	0	11	1.1%	8.28 [0.40, 171.29]	
Wysocki 1995	7	21	10	20	4.9%	0.50 [0.14, 1.77]	
Zhan 2012	1	21	5	19	1.9%	0.14 [0.01, 1.33]	· · · · · · · · · · · · · · · · · · ·
Total (95% CI)		846		845	100.0%	0.56 [0.40, 0.78]	•
Total events	191		261				
Heterogeneity: Tau <sup>2</sup> =	= 0.16; Ch	$i^2 = 27.$	26, df =	18 (P =	0.07); l <sup>2</sup>	= 34%	0.01 0.1 1 10 10
Test for overall effect	,		,	,	.,		0.01 0.1 1 10 10 Favours [experimental] Favours [control]

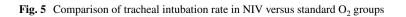


	ental	Contr	01		Odds Ratio	Odds Ratio
Events	Total	Events	Total	Weight	IV, Random, 95% CI	IV, Random, 95% CI
7	20	11	20	6.9%	0.44 [0.12, 1.57]	
2	40	7	41	5.0%	0.26 [0.05, 1.32]	
7	28	6	28	7.1%	1.22 [0.35, 4.24]	
0	20	0	27		Not estimable	
19	61	18	62	11.2%	1.11 [0.51, 2.39]	
9	51	21	54	9.9%	0.34 [0.14, 0.83]	<b>.</b>
0	28	0	25		Not estimable	
1	25	1	25	2.1%	1.00 [0.06, 16.93]	
12	26	18	26	7.9%	0.38 [0.12, 1.19]	
1	16	2	15	2.5%	0.43 [0.04, 5.35]	
5	32	10	29	7.2%	0.35 [0.10, 1.20]	
61	99	66	101	13.3%	0.85 [0.48, 1.51]	
3	20	15	20	5.2%	0.06 [0.01, 0.29]	
16	42	14	44	10.0%	1.32 [0.54, 3.21]	
4	16	0	11	1.8%	8.28 [0.40, 171.29]	
7	21	10	20	7.0%	0.50 [0.14, 1.77]	
1	21	5	19	3.0%	0.14 [0.01, 1.33]	· · · · · · · · · · · · · · · · · · ·
	566		567	100.0%	0.56 [0.36, 0.85]	•
155		204				
0.28; Ch	$i^2 = 25.$	11, df =	14 (P =	• 0.03); l <sup>2</sup>	= 44%	0.01 0.1 1 10 100
Z = 2.68	(P = 0.0)	007)				0.01 0.1 1 10 100 Favours [experimental] Favours [control]
	7 2 7 0 19 9 0 1 12 1 5 61 3 16 4 7 1 155 0.28; Ch	$\begin{array}{cccccc} 7 & 20 \\ 2 & 40 \\ 7 & 28 \\ 0 & 20 \\ 19 & 61 \\ 9 & 51 \\ 0 & 28 \\ 1 & 25 \\ 12 & 26 \\ 1 & 16 \\ 5 & 32 \\ 61 & 99 \\ 3 & 20 \\ 16 & 42 \\ 4 & 16 \\ 7 & 21 \\ 1 & 21 \\ \hline 566 \\ 155 \\ 0.28; \mathrm{Chi}^2 = 25. \end{array}$	$\begin{array}{cccccccccccccccccccccccccccccccccccc$	$\begin{array}{c ccccccccccccccccccccccccccccccccccc$	$\begin{array}{c ccccccccccccccccccccccccccccccccccc$	7       20       11       20       6.9%       0.44       [0.12, 1.57]         2       40       7       41       5.0%       0.26       [0.05, 1.32]         7       28       6       28       7.1%       1.22       [0.35, 4.24]         0       20       0       27       Not estimable         19       61       18       62       11.2%       1.11       [0.51, 2.39]         9       51       21       54       9.9%       0.34       [0.14, 0.83]         0       28       0       25       Not estimable         1       25       1       25       2.1%       1.00       [0.06, 16.93]         12       26       18       26       7.9%       0.38       [0.12, 1.19]         1       16       2       15       2.5%       0.43       [0.04, 5.35]         5       32       10       29       7.2%       0.35       [0.10, 1.20]         61       99       66       101       13.3%       0.85       [0.48, 1.51]         3       20       15       20       5.2%       0.06       [0.01, 0.29]         16       42       1



	Experim	ental	Conti	rol		Odds Ratio	Odds Ratio
Study or Subgroup	Events	Total	Events	Total	Weight	IV, Random, 95% CI	IV, Random, 95% CI
Antonelli 1998	9	32	15	32	21.6%	0.44 [0.16, 1.25]	
Gunduz 2005	2	22	7	21	7.9%	0.20 [0.04, 1.11]	
Honrubia 2005	10	31	14	33	22.3%	0.65 [0.23, 1.80]	
Matic 2007	15	195	21	192	48.2%	0.68 [0.34, 1.36]	
Total (95% CI)		280		278	100.0%	0.56 [0.34, 0.90]	◆
Total events	36		57				
Heterogeneity: Tau <sup>2</sup> =	= 0.00; Ch	$i^2 = 1.9$	5, df = 3	(P = 0)	.58); I <sup>2</sup> =	0%	0.01 0.1 1 10 100
Test for overall effect	: Z = 2.39	(P = 0.0	02)				0.01 0.1 1 10 100 Favours [experimental] Favours [control]

	Experim	imental Control			Odds Ratio	Odds Ratio	
Study or Subgroup	Events	Total	Events	Total	Weight	IV, Random, 95% CI	IV, Random, 95% CI
Antonelli 2000	4	20	14	20	5.8%	0.11 [0.03, 0.46]	
Brambilla 2014	2	40	1	40	2.4%	2.05 [0.18, 23.59]	
Confalonieri 1999	6	28	14	28	8.0%	0.27 [0.08, 0.88]	
Cosentini 2010	0	20	0	27		Not estimable	
Delclaux 2000	21	61	24	62	13.8%	0.83 [0.40, 1.73]	
Ferrer 2003	13	51	28	54	12.3%	0.32 [0.14, 0.73]	<b>_</b>
Gupta 2010	2	28	0	25	1.6%	4.81 [0.22, 105.18]	
Hernandez 2010	3	25	12	25	5.9%	0.15 [0.04, 0.62]	
Hilbert 2001	12	26	20	26	7.8%	0.26 [0.08, 0.85]	
Kramer 1995	5	16	11	15	5.2%	0.17 [0.03, 0.78]	
Martin 2000	9	32	17	29	9.1%	0.28 [0.09, 0.80]	
Squadrone 2010	2	20	8	20	4.5%	0.17 [0.03, 0.92]	
Wermke 2012	6	42	11	44	8.7%	0.50 [0.17, 1.50]	
Wood 1998	7	16	5	11	5.3%	0.93 [0.20, 4.37]	
Wysocki 1995	13	21	14	20	6.9%	0.70 [0.19, 2.56]	
Zhan 2012	1	21	4	19	2.7%	0.19 [0.02, 1.85]	
Total (95% CI)		467		465	100.0%	0.37 [0.25, 0.55]	•
Total events	106		183				
Heterogeneity: Tau <sup>2</sup> =	= 0.16; Ch	$i^2 = 19.$	33, df =	14 (P =	= 0.15); I <sup>2</sup>	= 28%	
Test for overall effect	,		,	`	- , ,		0.01 0.1 1 10 100 100
							Favours [experimental] Favours [control]



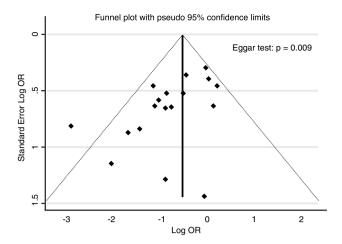


Fig. 6 Funnel plot of short-term mortality in the comparison between the NIV and control groups. Control group includes standard  $O_2$  therapy and invasive ventilation groups

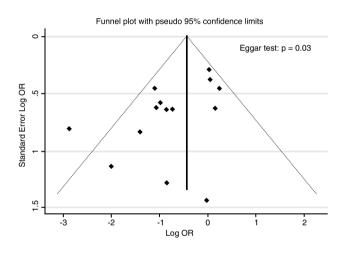


Fig. 7 Funnel plot for short-term mortality in the comparison between NIV and standard  $O_2$  therapy groups

NIV and in 183 patients (39.4%) in the standard  $O_2$  group, representing a pooled OR of 0.37 (95% CI 0.25–0.55).

### Heterogeneity

A statistically significant heterogeneity in short-term mortality was observed between the NIV and the control (standard O<sub>2</sub> therapy and InMV) groups ( $l^2 = 34.0\%$ ,  $\chi^2 = 27.3$ , p = 0.07) and between the NIV and the standard O<sub>2</sub> groups ( $l^2 = 42.0\%$ ,  $\chi^2 = 24.3$ , p = 0.04). A statistical heterogeneity was observed in neither short-term mortality between the NIV and the InMV groups ( $l^2 = 0.0\%$ ;  $\chi^2 = 1.95$ ; p = 0.58) nor in the tracheal intubation rate between the NIV and the standard O<sub>2</sub> groups ( $l^2 = 28.0\%$ ,  $\chi^2 = 19.3$ , p = 0.15).

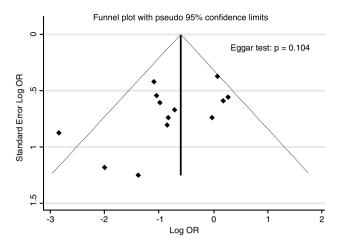


Fig. 8 Funnel plot for the tracheal intubation rate in comparison between NIV and standard  $O_2$  therapy groups

#### Publication biases, risk of bias and quality of evidence

We tested for the presence of publication biases for each outcome, except the short-term mortality between the NIV and the invasive ventilation group, because of a small sample size. A visual inspection of the funnel plots and the Eggar linear regression test suggested the existence of publication biases in short-term mortality (Figs. 6, 7) between (a) the NIV and the control (standard O<sub>2</sub> therapy and InMV) groups (p < 0.01) and (b) the NIV and standard O<sub>2</sub> therapy group (p < 0.01). The funnel plot for the tracheal intubation rate between the NIV and standard O<sub>2</sub> groups was symmetric (Fig. 8), excluding the existence of a small publication bias (p = 0.10).

In the nature of the intervention, blinding was categorized as high risk for all the trials and selective outcome reporting was assessed as uncertain risk for nearly all the trials due to the unavailability of study protocols (Online Supplements 2, 3 and 4). The quality of evidence was rated as low for the effect of NIV on short-term mortality compared with the control group, including standard  $O_2$  therapy or InMV. The grade was lowered by 2 points, due to the major inconsistency and publication bias, for which the Cochrane chi-square test showed a significant heterogeneity. The quality of evidence was rated as low for the effect of NIV on short-term mortality, compared with standard O2 therapy. The grade was lowered by 2 points due to a very serious risk of bias; the domain of blinding in risk of bias was rated as high in nearly all studies because the decision of intubation depended on each clinician. The quality of evidence was rated as low for the effect of NIV on shortterm mortality compared with invasive ventilation. We lowered the grade by 2 points because of a major imprecision, since the size of the criterion of optimal information was

# Table 2 Summary of findings table

Type of comparison	Outcome	Illustrative com CI)	nparative risks (95%	Relative effect (95% CI)	Number of participants	Quality of evidence (GRADE)	
		Assumed risk	Corresponding risk		(studies)		
		Control group	NIV				
NIV versus standard oxygen therapy and InMV <sup>a</sup>	Short-term mortality	318 per 1000	207 per 1000 (157 to 267)	OR 0.56 (0.40, 0.78)	1691 (21 studies)	Low <sup>a,b</sup>	
NIV versus standard oxygen therapy	Short-term mortality	290 per 1000	186 per 1000 (128 to 258)	OR 0.56 (0.36, 0.85)	1133 (17 studies)	Low <sup>c,d</sup>	
	Tracheal intubation	467 per 1000	245 per 1000 (180 to 325)	OR 0.37 (0.25, 0.55)	932 (16 studies)	Low <sup>e</sup>	
NIV versus InMV <sup>a</sup>	Short-term mortality	379 per 1000	255 per 1000 (172 to 355)	OR 0.56 (0.34, 0.90)	558 (4 studies)	Low <sup>f,g</sup>	

The basis for the assumed risk is the median control group risk across studies. The corresponding risk (and its 95% CI) is based on the assumed risk in the comparison group and the relative effect of the intervention (and its 95% CI)

GRADE Working Group grades of evidence

High quality: further research is very unlikely to change our confidence in the estimate of effect

Moderate quality: further research is likely to have an important impact on our confidence in the estimate of effect and may change the estimate Low quality: further research is very likely to have an important impact on our confidence in the estimate of effect and is likely to change the estimate

Very low quality: we are very uncertain about the estimate

CI confidence interval, OR odds ratio, NIV non-invasive ventilation

<sup>a</sup> In this comparison, quality of evidence for the effect of NIV on tracheal intubation was not rated because patients in the InMV group were all intubated

<sup>b</sup> Serious inconsistency: Cochrane chi-squared test indicated significant heterogeneity ( $\chi^2 = 27.3$ , p = 0.07). Downgraded by 1

<sup>c</sup> Publication bias was detected: visual inspection and Eggar test suggested publication bias. Downgraded by 1

<sup>d</sup> Serious inconsistency: Cochrane chi-squared test indicated significant heterogeneity ( $\chi^2 = 24.3, p = 0.04$ ). Downgraded by 1

<sup>e</sup> Publication bias was detected: visual inspection and Eggar test suggested publication bias. Downgraded by 1

 $^{\rm f}$  Very serious risk of bias: the domain of blinding in risk of bias was rated as a high risk of bias in almost all studies because the decision to intubate depended on each clinician. Downgraded by 2

<sup>g</sup> Serious imprecision: criterion of optimal information size was not met. Downgraded by 1

<sup>h</sup> Publication bias was detected: only four small sample size studies were included. Downgraded by 1

insufficient, and because of a publication bias (Table 2). The detail of the evidence profile is shown in Online Supplement 5.

# Discussion

This meta-analysis of 21 randomized trials, which includes 1691 patients with ARF, suggests that NIV lowers the short-term mortality compared with standard  $O_2$  therapy and InMV. Subgroup analyses were performed comparing NIV with the standard  $O_2$  therapy and InMV control groups separately. The short-term mortality was significantly higher in both control groups than in the NIV group. In addition, NIV lowered the tracheal intubation rate. We excluded patients presenting with PERF, CPE or exacerbation of COPD to discuss only ARF caused by lung disease such as ARDS.

Previous meta-analyses [9–14, 16, 39] of patients presenting with ARF have been published, which have reported favorable survival rates attributable mainly to the inclusion of PERF, CPE and exacerbation of COPD [22, 47, 48]. On the other hand, this is the first study which evaluates the efficacy of NIV among patients with ARF excluding CPE or COPD. Because of the paucity of randomized trials limited to acute lung injury ALI/ARDS, we are aware of only two relevant meta-analyses of NIV for this indication. In 2010, Agarwal et al. examined the role of NIV in the management of ALI and ARDS [49]. They searched the Pubmed and EMBASE databases for relevant studies published between 1995 and 2009, and included studies that reported rates of tracheal intubation, or death or both in patients with ALI/ARDS treated with NIV. They found 13 studies, including a total of 540 patients who met their inclusion criteria. Their analysis revealed a nearly 50% failure of NIV, prompting the authors to conclude that NIV should be used cautiously in this population. They also found a significant statistical heterogeneity for both intubations ( $I^2 = 76\%$ , 95% CI 55–85, Cochran Q statistic 50, p = 0.001) and deaths ( $I^2 = 79\%$ , 95% CI 61–86, Cochran Q statistic 56, p = 0.001). The quality of that analysis was lowered by the inclusion of several observational studies and by the presentation of insufficient overall evidence [49]. In 2014, Luo et al. pointed out in their meta-analysis that the role of NIV in the management of ALI/ARDS was controversial [50]. They included studies that reported the tracheal intubation rate and/or mortality in patients treated for ALI/ARDS with NIV. They found 6 RCTs, including a total of 227 patients who met their inclusion criteria. In their meta-analysis, they did not find an improvement of inhospital mortality, although the rate of tracheal intubation was significantly lowered by NIV. The heterogeneity for tracheal intubation and in-hospital mortality was ( $l^2 = 43\%$ ,  $\chi^2 = 8.82, p = 0.12$ ) and  $(I^2 = 61\%, \chi^2 = 5.12, p = 0.08)$ , respectively [50]. These two studies were both limited by the small sample sizes and by the inclusion of high heterogeneity only, which may explain the absence of effects on the survival rate. To obviate these limitations, we included 21 studies dedicated to the treatments, including nearly 1700 patients with ARF, and excluded patients suffering from PERF, CPE or COPD. This large sample size, combined with the lower heterogeneity rates observed in our study (compared with the studies of Agarwal et al. and Luo et al.) increased the reliability of the outcome estimate, in which we found significantly lower rates of both short-term mortality and tracheal intubation conferred by NIV.

Another contribution of our study was in the comparisons between NIV with InMV. Several studies have been limited to comparisons of NIV with standard  $O_2$  therapy, although it is important to compare non-invasive versus InMV in clinical settings such as emergency departments or intensive care units [49–51]. It has been suggested that the failure of NIV in patients with ARF is independently correlated with poor outcomes compared with patients intubated without prior NIV [20]. In our study, short-term mortality was significantly lower in the NIV than in the InMV group.

From a physiological perspective, the benefits conferred by NIV are not only from continuous positive end-expiratory pressure, since both NIV and InMV can create physiological conditions of positive end-expiratory pressure. NIV interferes with neither the native upper airways nor the glottis function. It can alleviate the respiratory efforts and improve gas exchange while preserving the ability to swallow, cough, and speak. Furthermore, NIV may obviate the adverse effects of InMV, including deep sedation, administration of muscle relaxants, delirium, ventilator-induced lung injury, and ventilator-associated infections [52, 53].

#### Limitations of our study

All randomized trials included in our meta-analysis were not double-blinded because the NIV is a visible intervention and cannot be concealed. Therefore, we assigned a high risk of bias to all the trials. Furthermore, the publication bias and selective outcome reporting were assessed as unclear risks of bias because we could not obtain the actual study protocols for most of the studies. A second limitation of this meta-analysis was the inclusion of various kinds of diseases causing ARF, as well as obvious differences in the patients' baseline characteristics among the studies. Third, treatment and prognosis of ARF are changing year by year and it is no longer the same situation as in the old days. Finally, our analysis was limited to short-term mortality because most studies did not report long-term survival data. Whether a decrease in short-term mortality increases the long-term survival remains to be clarified by further investigations.

# Conclusion

In this meta-analysis of 21 randomized trials and 1691 patients with ARF, NIV lowered both the short-term mortality and the rate of tracheal intubation compared with standard  $O_2$  or InMV. Despite unavoidable risks of biases, this study suggests that NIV is worth considering as a treatment option for patients with ARF.

Acknowledgements We thank all members of the Clinical Practice Guideline for ARDS 2016, from the Japanese Society of Respiratory Care Medicine, the Japanese Society of Intensive Care Medicine, and the Japanese Respiratory Society, as well as the librarians Yumi Yamashita, Yoshiko Nakagawa and Takaaki Suzuki, from Kyoto Prefectural University of Medicine library (YY and YN) and Nara Medical University library (TS), for their assistance in the search for relevant publications. Rodolphe Ruffy, M.D., http://www.cardioscript. com, reviewed the manuscript for style and language and was compensated for his services.

**Author contributions** YK, JK, AK, and RS designed the study. YK, JK, AK, and RS identified the studies entered in the meta-analysis, and extracted and analysed the data. YK and JK drafted the manuscript. AK, RS, EN and SH critically reviewed the manuscript. All authors have read and approved the final version of the manuscript.

#### Compliance with ethical standards

Ethics approval and consent to participate Not applicable.

Consent for publication Not applicable.

Availability of data and material The data and material used for this meta-analysis are contained in references [15-35] of our list of references.

**Conflict of interest** The authors have no competing interest to declare.

**Funding** This study was supported in part the Grant-in-Aid for Scientific Research (B) 24390404, 2012–2016, awarded to Satoru Hashimoto, M.D., Ph.D., by the Japanese Ministry of Education, Science, Sports and Culture.

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