



# Interactions between folate metabolism-related nutrients and polymorphisms on colorectal cancer risk: a case-control study in the Basque country

Sara Corchero-Palacios<sup>1</sup> · Iker Alegria-Lertxundi<sup>1</sup> · Marian M. de Pancorbo<sup>2,3</sup> · Marta Arroyo-Izaga<sup>1,3</sup>

Received: 14 August 2023 / Accepted: 1 March 2024  
© The Author(s) 2024

## Abstract

Folate-mediated one-carbon metabolism (FOCM) plays an important role in colorectal carcinogenesis. Previous studies have assessed the role of folate-mediated one-carbon metabolism (FOCM)-related gene-diet interaction in the aetiology of colorectal cancer (CRC), however, the results remained inconclusive. Thus, this study aimed to investigate dietary factors and genetic variants related to FOCM, as well as potential nutrient-gene and nutrient-lifestyle interactions, on CRC risk. This observational study included 229 patients diagnosed with CRC and 229 age- and sex-matched subjects as controls from a population-based bowel cancer screening program. Conditional logistic regression was used to calculate odds ratios (ORs) and 95% confidence intervals (95%CI) for CRC risk. A Bonferroni-corrected threshold of  $\alpha=0.005$  was considered significant, and  $P$  values less than 0.05 were considered to be suggestive of an association. After Bonferroni correction, a high dietary intake of betaine was associated with a decreased risk of CRC in the adjusted model (OR, 95% CI: 0.21, 0.10–0.40,  $P<0.001$ ). Two SNPs, rs1476413 and rs17824591, exhibited significant gene-diet interactions with total choline and vitamin B<sub>12</sub> intakes, respectively, in adjusted models (total choline, tertile 3 vs. 1, OR, 95% CI: 0.25, 0.11–0.66,  $P_{\text{interaction}} = 0.012$ ; vitamin B<sub>12</sub>, tertile 2 vs. tertile 1, OR, 95% CI: 2.48, 1.04–5.00,  $P_{\text{interaction}} = 0.003$ ). These findings suggest that betaine intake and interactions between some dietary factors and variants in *MTHFR* and *MTHFD1* genes have an influence on CRC risk in the population studied. If these results are confirmed, specific nutritional intervention strategies could be designed.

**Keywords** Colorectal cancer · Risk factor · Nutrient · Genetic polymorphism · Gene-nutrient interaction · Case-control study

✉ Marta Arroyo-Izaga  
marta.arroyo@ehu.es

Sara Corchero-Palacios  
saracorchero@gmail.com

Iker Alegria-Lertxundi  
lertxundi83@gmail.com

Marian M. de Pancorbo  
marian.mdepancorbo@ehu.es

<sup>1</sup> Department of Pharmacy and Food Sciences, Faculty of Pharmacy, University of the Basque Country UPV/EHU, Vitoria-Gasteiz (Araba/Álava) 01006, Spain

<sup>2</sup> Department of Z. and Cellular Biology A., Faculty of Pharmacy, University of the Basque Country UPV/EHU, Vitoria-Gasteiz (Araba/Álava) 01006, Spain

<sup>3</sup> BIOMICs Research Group, Microfluidics & BIOMICs Cluster, Lascaray Research Center, University of the Basque Country UPV/EHU, Bioaraba, BA04.03, 01006 Vitoria-Gasteiz (Araba/Álava), Spain

## Abbreviations

1CM	One-carbon metabolism
BCSP	Population-based bowel cancer screening program
CI	Confidence interval
CRC	Colorectal cancer
DI	Deprivation index
DNMT	DNA methyltransferases
FFQ	Food frequency questionnaire
FOCM	Folate-mediated one-carbon metabolism
Met	Methionine
MTHFD1	Methylene tetrahydrofolate dehydrogenase 1
MTHFR	Methylene tetrahydrofolate reductase
OR	Odds ratio
PE	Physical exercise
PRM	Predictive risk modelling

SD	Standard deviation
SNP	Single nucleotide polymorphism
T	Tertile

### Novelty and Impact – What’s new?

Previous studies have assessed the role of folate-mediated one-carbon metabolism (FOCM)-related gene-diet interaction in the aetiology of colorectal cancer (CRC), however, the results remained inconclusive. Here, the authors investigate this type of interaction. The findings highlight the importance of interactions between total choline and vitamin B<sub>12</sub> intakes, and variants in *MTHFR* and *MTHFD1* genes on CRC risk. If these results are confirmed, they may provide valuable risk stratification guidance for diet recommendations.

## Introduction

Colorectal cancer (CRC) is the third most frequent cancer and the second highest mortality in cancer patients worldwide [1]. In Spain, CRC is currently the most frequently diagnosed tumour, with 41,661 new cases (25,415 in men and 16,246 in women) detected in 2022. This incidence is comparable to that found in high-risk zones of Occidental Europe, North America, Australia, and Japan [3].

Although screening for early detection of CRC is effective in decreasing trends in mortality rates, understanding the factors involved in daily life are CRC-diagnosed also important for a proactive approach to prevent this type of cancer [4]. The primary prevention for CRC is mostly associated with diet, lifestyle factors, and metabolic diseases. Regarding dietary factors, one-carbon metabolism (ICM)-related nutrients (such as folate, other B vitamins, methionine (Met), choline, and betaine) have been considered anticarcinogenic and chemotherapeutic agents in the 1 C metabolic network [5], whereas alcohol antagonizes ICM, and its high consumption has been related to higher CRC risk [6]. In addition, the observed inverse association between folate status and CRC risk was further modified by genetic polymorphisms of the enzymes involved in folate metabolism, most notably methylene tetrahydrofolate reductase (*MTHFR*) [5].

However, not only the influence of polymorphisms but also the influences of ICM-related nutrients on genetic polymorphisms in relation to interaction CRC risk remain largely unexplored. Most studies on this type of gene-diet interaction have focused on folate, B vitamin, and methionine intake [7]. To date, there are no studies in which the intake of choline and/or betaine has been evaluated. To better elucidate the role of genetic factors and environmental

conditions, especially diet, on CRC risk, this study had a triple aim: (i) to investigate dietary factors (dietary methyl donors and dietary components that potentially modulate the bioavailability of methyl groups) and genetic variants in methyl metabolizing enzymes; (ii) to determine the potential nutrient-gene interactions; and (iii) to analyse the potential nutrient-lifestyle interactions, that is, interactions between the consumption of the dietary factors mentioned on the first objective and other lifestyle factors.

These aims refer to the CRC risk, in a sample of cases and controls, matched on age and sex, from the population-based bowel cancer screening program (BCSP) of the Osakidetza/Basque Health Service. In particular, the dietary factors investigated were: intakes of folate, vitamins B<sub>2</sub>, B<sub>6</sub>, and B<sub>12</sub>, Met, choline, betaine, and ethanol; the genetic variants: DNA methyltransferases (*DNMT3B* and *DNMT1*), *MTHFR*, methylene tetrahydrofolate dehydrogenase 1 (*MTHFD1*), and *Met synthase reductase*; and the other lifestyle factors: physical exercise (PE), smoking, and alcohol consumption.

## Methods

### Study participants

Overall, this epidemiologic study is an observational analytic case-control study designed to address possible gene-diet interactions in relation to CRC. Participants in this study were recruited from among patients attending any of the three hospitals of the Osakidetza/Basque Health Service (Basurto, Galdakao, and Donostia) members of the Basque Country’s BCSP. To be eligible for this BCSP, the patients had to be aged between 50 and 69, asymptomatic for colorectal symptoms, and registered with the Osakidetza/Basque Health Service.

These inclusion criteria were applied to both case and control groups; that is, controls fulfilled the same eligibility criteria defined for the cases, except for the disease (outcome). The age- and sex-matched controls were patients with positive results (abnormal) for an immunochemical faecal occult blood test and negative colonoscopy results (normal). Recruitment and data collection through questionnaires were conducted between 2014 and 2016. The start date of the study was 2014 because the BCSP in the Basque Country reached the whole target population (approximately 586,700 people) at the beginning of this year.

The characteristics of the sampling and the cases (pathological staging, location of cancer, tumour grade, and treatments) have been described before [8]. Briefly, 72% were diagnosed with early-stage (I/II) CRC and 76% had a distal location of cancer. The total sample consisted of 308 cases who were diagnosed with CRC and 308 age- and

sex-matched controls. However, in the present study, data from 229 CRC patients and 229 controls were analysed, since this is the sample from which biological samples and associated data were obtained. All participants had data on the main dietary factors (folate, vitamin B<sub>2</sub>, B<sub>6</sub>, B<sub>12</sub>, Met, choline, and betaine) and genetic variants that were included in the present study.

This study was conducted according to the guidelines laid down in the Declaration of Helsinki, and all procedures involving patients were approved by the Clinical Research Ethics Committee of the Basque Country (protocol code PI2011006, data of approval 03/23/2012; and PI2014042, data of approval 05/28/2014). Written informed consent was obtained from all the study participants.

### Dietary assessment

Diets were assessed using a short food frequency questionnaire (FFQ) that was a modified version of the Rodríguez et al. questionnaire [9]. This adaptation was validated with multiple 24-h recalls in the Basque general population [10] and CRC-diagnosed patients in a pilot of the present study [11]. It consisted of 67 items and requires the subjects to recall the number of times each food item was consumed either per week or per month. Moreover, the respondents could also record the consumption of other foods that were not included on the food list, as well as the use of dietetic products and nutritional supplements (generic and brand-name, dose, and frequency).

Once the completed FFQ was received, it was reviewed by a dietitian. Consumption frequencies were standardised to “per day” and multiplied by standard serving sizes (grams) [12]. For items that included several foods, each food’s contribution was estimated with weighting coefficients that were obtained from the usual consumption data [13]. All food items that were consumed were entered into DIAL 2.12 (2011 ALCE INGENIERIA), a type of dietary assessment software, to estimate energy intake (kcal/day), dietary fibre, and ICM-related vitamins (B<sub>2</sub>, B<sub>6</sub>, folate, and B<sub>12</sub>). The intakes of other ICM-related compounds, in particular, Met, total choline, and betaine were estimated using the United States Department of Agriculture (USDA) Food Data central database [14] and from the publication of Zeisel et al. [15].

### Assessment of covariates

Potential confounders of CRC risk were selected based on published evidence from European Prospective Investigation into Cancer and Nutrition (EPIC), the International Agency for Research on Cancer (IARC), and the World Cancer Research Fund (WCRF), which included: overweight

or obesity based on self-reported body mass index (BMI), age, sex, PE (expressed as daily minutes of cycling/sports), smoking (never vs. past/currently, and intensity of smoking measured by the number of cigarettes smoked per day), and alcohol consumption (reported as grams of ethanol per day). In addition, the use of drugs related to decreasing CRC risk (antiplatelet, including non-steroidal anti-inflammatory drugs, and anticoagulants) [16] was recorded. These questions were taken from the Spanish Health Questionnaire [17]. BMI, estimated from self-reported height and weight, was classified according to the WHO criteria for those under 65 years of age [18] and according to the criteria proposed by Silva Rodrigues et al. for those 65 and older [19]. The characteristics of the sample is shown in the Supplementary Material (Table S1).

On the other hand, the FFQ used to assess dietary intake included specific questions about the frequency of intake of the following five major types of alcoholic beverages: beer, wine, cider, aperitif with alcohol, and liquor. These consumption frequencies were standardised to “per day” and multiplied by standard serving sizes (ml) [20]. The alcohol consumption data were expressed as grams of ethanol/day that were estimated with the software DIAL 2.12 (2011 ALCE INGENIERIA) and standard drink units [21]. We used the standard drink unit defined for Spain (one standard drink unit is the equivalent of 10 g of ethanol). With this information, the participants were categorised into those who did and did not meet the recommendations [20]. Those participants who did not meet the recommendations were categorised as “high-risk consumption.”

The differences in general characteristics (age, BMI, PE, smoking, and alcohol consumption, among others) between cases and controls were previously described [22]. Briefly, significant differences between cases and controls were found for smoking and weight status, with a higher percentage of cases with past or current smoking status and with overweight/obesity compared to controls ( $P < 0.01$ ). However, no significant differences were found in alcohol consumption between cases and controls ( $P > 0.05$ ).

Additionally, in both cases and controls, socioeconomic level, and health status (specifically health resource consumption) data were assessed with two indices that were obtained from the clinical databases developed by the Health Department of the Basque Government, namely the socioeconomic deprivation index (DI) and predictive risk modelling (PRM), respectively. The first one was estimated using the MEDEA project criteria, as has been described elsewhere, [23] and was divided into quintiles, with the first being the least disadvantaged and the fifth being the most disadvantaged. The PRM is an index that is based on Adjusted Clinical Groups [24] and Clinical Risk Groups [25]. This index combines information about diagnoses,

prescriptions, previous costs, and the use of specific procedures. It can predict the use of health resources, and it was stratified into four levels: the first included participants with a risk of high health resource consumption, and the fourth included those with low health resource consumption. The differences in these two indices (DI and PRM) between cases and controls were previously described [22].

### Biological samples and genotyping

In this study, healthy tissues or saliva samples of cases and controls were collected and genotyped. Samples were provided by the Basque Biobank for Research-OEHUN ([www.biobancovasco.org](http://www.biobancovasco.org)) and were processed following standard operating procedures with appropriate approval of the Ethical and Scientific Committees. DNA was extracted using AllPrep DNA / RNA kit (Qiagen) for paraffin-embedded tissue samples and AutoGenFlex Tissue DNA Extraction kit (Autogen) for mouthwash saliva samples and then was quantified with NanoDrop™ Spectrophotometer (ThermoFisher). Double-stranded DNA was quantified by fluorometry using the Quant-iT™ PicoGreen1 dsDNA Assay Kit (Invitrogen, CA) on a DTX 880 Multimode Detector (Beckman Coulter) to normalize DNA concentration.

After an updated summary of the published genetic variants in methyl metabolizing enzymes related to CRC risk, a total of ten single nucleotide polymorphisms (SNPs) were identified, in particular: *DNMT3B* (rs2424913, rs406193), [26] *DNMT1* (rs2228612), [27] *MTHFR* (rs1476413, rs1801131, rs1801133), [28, 29] *MTHFD1* (rs8003379, rs17824591) and *Met synthase reductase* (rs1801394, rs10380) [29]. These SNPs were organised in the context of the gene(s) at or near the locus and chromosome locus. The allelic discrimination was assessed using the MassARRAY1 System (Agena Bioscience) on CeGen-PRB2-ISCI (Nodo USC) following the procedure provided by the manufacturer. Quality control samples were included in the genotyping assays.

### Statistical analysis

The sample size was estimated to be 286 in each group (cases and controls) to detect an odds ratio (OR) of 2.0 with 80% power at a two-sided level of significance of 5%, under an exposure prevalence of 10%, using the Epidat 3.0 program [30].

Data were analysed using IBM SPSS Statistics for Windows, version 22.0 (IBM Corp., Armonk, NY, USA) and STATA 16.0 (StataCorp LP, Texas, USA). Categorical variables are shown as a percentage, and continuous variables are as the means and standard deviation. Normality was checked using the Kolmogorov–Smirnov–Lilliefors test

and since all continuous variables followed a non-normal distribution, the Wilcoxon rank-sum test was used for two related means comparison. The Chi-square test was used to evaluate differences between categorical variables. When expected frequencies were lesser than five, Fisher's exact test was used. Tests for associations and deviation from Hardy-Weinberg equilibrium were performed separately in cases and controls.

Conditional logistic regression was used to calculate ORs and 95% confidence intervals (95%CI) for CRC risk according to (i) tertiles (T) of dietary compound intake, (ii) dominant and recessive models of SNPs, (iii) nutrient-gene interactions, and (iii) nutrient-lifestyle interactions. The logistic regression models were applied to the total sample and the subgroup of cases with distal location ( $n = 178$ ) and their respective controls (Tables S2, S7, and S8). When the sample size was small ( $\leq 10$  per group), the conditional exact logistic regression model was applied. The logistic regression model could not be applied to the subgroup of cases with proximal location due to the (too small) sample size. The intake of dietary compounds was categorised into Ts based on the distribution in the control group, taking into account sex differences when they were significant. Specifically, different cutoff points were applied to estimate Ts in men and women when significant sex differences were identified, that is, in the case of folate and Met intake. The lowest T was used as the reference group. The most frequent genotype (homozygous) was considered the reference group to calculate ORs in a dominant model, and the most frequent genotype (homozygous) and the heterozygous genotype containing the risk allele were considered the reference group in the recessive model. The nutrient-gene interaction analyses were carried out using a dominant model for genotypes.

The analyses of logistic regression were done for the unadjusted (model I) and adjusted models (models II and III). Models II and III were adjusted for known risk factors for CRC: [31, 32] age, sex, BMI, PE level, smoking status, the intensity of smoking (in current and past smokers), socio-economic level (DI) and health status (PRM), energy intake, dietary fibre, ethanol intake, and antiplatelet and anticoagulants use. The reference categories were those that, according to the literature, have a lower CRC risk. For the BMI variable, normal weight was considered as the reference category, and underweight was included as a separate category. We included participants with missing data for the covariates as a separate category.

Quantitative covariates such as intensity of smoking (cigarettes/day) were dichotomised by mean or median, according to the normality test. We used the cut-off of Romaguera et al. [33] to create two PE levels expressed in min/day of cycling/sports: sedentary-light ( $< 15$  min/day)

and moderate-vigorous ( $\geq 15$  min/day). Age was dichotomised using the same age ranges that were used in the sample selection process (50–59 years old vs. 60–69 years old). Qualitative ones, such as DI and PRM were dichotomised considering the distribution of frequencies to obtain similar sample sizes for each category (DI, quintile 1–3 vs. quintile 4–5; PRM, level 3–4 vs. level 1–2). Energy, dietary fibre, and ethanol intake were included as quantitative variables in the adjusted models.

In model II, dietary compound intake or SNPs were included separately, whereas model III was only used in the analysis of dietary compound intake and, in this case, all the compounds were included at the same time. In addition, to study the possible association between the intake of betaine and total choline, and the CRC risk, model II (adjusted model) of the regression analysis was also applied, including folate as covariate (these data are shown in text form in the [Results](#) section). The significance level was corrected using a Bonferroni correction by dividing the standard  $P$  value (two-tailed) ( $\alpha = 0.05$ ) by the total number of SNPs analysed ( $n = 10$ ), assuming alpha was equal to 0.005 ( $\alpha = 0.05/10$ ).

## Results

### Colorectal cancer risk according to nutrient intake

The intakes of folate, vitamins B<sub>2</sub> and B<sub>6</sub>, Met, choline, and betaine were significantly higher in controls than those in cases ( $P < 0.001$ ), whereas the consumption of vitamin B<sub>12</sub> and alcohol was higher in cases than controls ( $P < 0.001$ ) (Table 1). The ORs for CRC risk by the intake of nutrients are presented in Table 2. The adjusted ORs for CRC risk decreased with higher intakes of choline and betaine ( $P < 0.005$ ). Although in the case of choline, the association was only significant in the unadjusted model. These

**Table 1** Daily intake of nutrients and alcohol in cases and control studied

Nutrient and ethanol intake, mean (SD)	Cases (n=229)	Controls (n=229)	$P^a$
Folate, $\mu\text{g/day}$	266.5 (81.6)	270.9 (77.8)	<b>&lt; 0.001</b>
Vitamin B <sub>2</sub> , mg/day	1.5 (0.5)	1.6 (0.5)	<b>&lt; 0.001</b>
Vitamin B <sub>6</sub> , mg/day	1.8 (0.5)	1.9 (0.6)	<b>&lt; 0.001</b>
B <sub>12</sub> , $\mu\text{g/day}$	5.0 (1.8)	4.9 (1.6)	<b>&lt; 0.001</b>
Met, mg/day	1783.0 (655.9)	1884.0 (756.9)	<b>&lt; 0.001</b>
Total choline, mg/day	136.6 (87.3)	165.2 (86.9)	<b>&lt; 0.001</b>
Betaine, mg/day	111.5 (54.8)	149.6 (61.5)	<b>&lt; 0.001</b>
Ethanol, g/day	8.3 (8.1)	7.4 (8.9)	<b>&lt; 0.001</b>

Abbreviations: Met, methionine; SD, standard deviation

<sup>a</sup>Wilcoxon test. A value of  $P < 0.005$  was considered significant after the Bonferroni correction (assuming alpha was equal to 0.005,  $\alpha = 0.05/10$ ). Significant results are highlighted in bold

results were confirmed in the subgroup of cases with distal tumour location (Table S2). In the total sample, after further adjustment for folate, moderate intake of betaine remained associated with a reduced risk of CRC (Model II including folate as a covariate,  $\text{OR}_{\text{T2 vs. T1}} = 0.36$ , 95% CI: 0.20–0.65,  $P = 0.001$ ;  $\text{OR}_{\text{T3 vs. T1}} = 0.20$ , 95% CI: 0.09–0.40,  $P < 0.01$ ).

### Colorectal cancer risk according to polymorphism genotypes

The distribution of genotypes at SNPs selected in the CRC group and in the control group that deviated from the Hardy-Weinberg equilibrium is shown in the Supplementary Material (Table S3). The SNP that was not following the Hardy-Weinberg equilibrium in controls was rs1801394 ( $P < 0.05$ ), however, this SNP showed no significance when the conservative Bonferroni method is used. None of the genotype frequencies for the SNPs analysed reached statistically significant differences between cases and controls, after the Bonferroni correction.

Supplementary Table S4 presents the associations between genotype variants and CRC risk. No significant association was found between the genotype for any SNP analysed and CRC risk, except a decreased risk of CRC among those with rs2424913-TT variant ( $\text{OR}_{\text{TT vs. CC}} = 0.56$ , 95% CI: 0.33–0.96,  $P < 0.05$ ), even though this association was not significant after the Bonferroni correction.

### Colorectal cancer risk according to nutrient-gene interactions

Associations between SNP genotypes and CRC risk, stratified by dietary factors in unadjusted and adjusted models are shown in Supplementary Tables S5 and S6, respectively. A summary of all observed associations between folate metabolism-related nutrients and SNPs on CRC risk is provided in Table 3. In the unadjusted model, the rs1476413-CC genotype was associated with a decreased risk of CRC among individuals with high total choline intake ( $\text{OR}_{\text{T3 vs. T1}} = 0.29$ , 95% CI: 0.15–0.55) ( $P_{\text{interaction}} = 0.002$ ). This result was also confirmed in the adjusted model, although in this case,  $P$ -value did not remain significant after applying the Bonferroni correction ( $P_{\text{interaction}} = 0.012$ ). Moreover, in the adjusted model, the rs17824591-GG genotype was associated with an increased risk of CRC among individuals with moderate vitamin B<sub>12</sub> intake ( $\text{OR}_{\text{T2 vs. T1}} = 2.48$ , 95% CI: 1.04–5.00) ( $P_{\text{interaction}} = 0.003$ ). Although, in the unadjusted model, these results were not confirmed in the subgroup of cases with distal tumour location (Table S7); in the adjusted model, the rs1476413-CC genotype was associated with a decreased risk of CRC among subjects with high betaine

**Table 2** Colorectal cancer risk according to nutrient intake

Nutrient intake <sup>a</sup>	Model I <sup>b</sup>			Model II <sup>c</sup>		Model III <sup>d</sup>	
	Cases/Controls, n	OR (95% CI)	P <sup>e</sup>	OR (95% CI)	P <sup>e</sup>	OR (95% CI)	P <sup>e</sup>
<b>Folate</b>							
T1	81/83	1.00	-	1.00	-	1.00	-
T2	80/71	1.15 (0.74–1.79)	0.539	1.21 (0.64–2.33)	0.563	1.11 (0.49–2.54)	0.792
T3	68/75	0.94 (0.61–1.46)	0.791	1.33 (0.54–3.22)	0.640	1.82 (0.60–5.57)	0.299
<b>Vitamin B<sub>2</sub></b>							
T1	85/83	1.00	-	1.00	-	1.00	-
T2	91/71	1.24 (0.81–1.89)	0.318	0.93 (0.51–1.70)	0.798	1.19 (0.51–2.72)	0.704
T3	53/75	0.64 (0.39–1.04)	0.072	0.45 (0.20–0.98)	0.048	0.56 (0.18–1.77)	0.324
<b>Vitamin B<sub>6</sub></b>							
T1	68/67	1.00	-	1.00	-	1.00	-
T2	99/85	1.11 (0.72–1.70)	0.639	0.76 (0.39–1.60)	0.477	0.61 (0.25–1.60)	0.310
T3	62/77	0.77 (0.46–1.26)	0.297	0.53 (0.24–1.23)	0.141	0.60 (0.19–2.21)	0.454
<b>Vitamin B<sub>12</sub></b>							
T1	70/73	1.00	-	1.00	-	1.00	-
T2	74/80	0.95 (0.60–1.48)	0.809	0.95 (0.52–1.83)	0.895	0.85 (0.40–1.88)	0.686
T3	85/76	1.21 (0.76–1.91)	0.429	1.15 (0.51–2.72)	0.699	1.73 (0.63–4.79)	0.287
<b>Met</b>							
T1	72/76	1.00	-	1.00	-	1.00	-
T2	89/78	1.23 (0.77–1.95)	0.383	1.13 (0.62–2.11)	0.690	0.96 (0.47–2.00)	0.922
T3	68/75	0.96 (0.60–1.53)	0.864	0.51 (0.24–0.99)	0.049	0.42 (0.22–0.92)	0.026
<b>Choline</b>							
T1	105/75	1.00	-	1.00	-	1.00	-
T2	75/77	0.72 (0.46–1.11)	0.134	0.60 (0.30–0.95)	0.030	0.64 (0.32–1.25)	0.183
T3	49/77	0.44 (0.27–0.72)	<b>0.001</b>	0.53 (0.29–1.06)	0.060	0.84 (0.38–1.84)	0.656
<b>Betaine</b>							
T1	150/77	1.00	-	1.00	-	1.00	-
T2	41/75	0.27 (0.17–0.46)	<b>&lt;0.001</b>	0.30 (0.18–0.64)	<b>0.001</b>	0.31 (0.14–0.72)	<b>0.003</b>
T3	38/77	0.28 (0.17–0.46)	<b>&lt;0.001</b>	0.21 (0.10–0.45)	<b>&lt;0.001</b>	0.21 (0.10–0.40)	<b>&lt;0.001</b>

Abbreviations: CI, confidence interval; Met, methionine; OR, odds ratio; T, tertile

<sup>a</sup>Nutrient intake was categorised into tertiles based on the distribution in the control group, taking into account sex differences when they were significant. Specifically, different cutoff points were applied to estimate tertiles in men and women when significant sex differences were identified. Tertiles of nutrient intake: folate (µg/day), for males, T1 ≤ 220.0, T2 220.1–289.0, T3 > 289.0, and females, T1 ≤ 245.0, T2 245.1–300.0, T3 > 300.0; vitamin B<sub>2</sub> (mg/day), T1 ≤ 1.3, T2 1.4–1.7, T3 > 1.7; vitamin B<sub>6</sub> (mg/day), T1 ≤ 1.5, T2 1.6–2.0, T3 > 2.0; vitamin B<sub>12</sub> (µg/day), T1 ≤ 3.9, T2 4.0–5.3, T3 > 5.3; Met (mg/day), for males, T1 ≤ 1324.0, T2 1324.1–1985.0, T3 > 1985.0, and females, T1 ≤ 1564.0, T2 1564.1–2623.0, T3 > 2623.0; choline (mg/day), T1 ≤ 114.0, T2 114.1–190.0, T3 > 190.0; betaine (mg/day), T1 ≤ 119.0, T2 119.1–165.0, T3 > 165.0

<sup>b</sup>Model I, analysis was performed using crude conditional logistic regression

<sup>c</sup>Model II, analyses were performed using conditional logistic regression analysis adjusted for the following variables (reference categories are underlined): sex (women; men) age (50–59 y old, 60–69 y old), BMI (normal weight, overweight/obesity), physical exercise (< 15 min/day of cycling/sports, ≥ 15 min/day), smoking status (never, past/currently: smoker: ≤ 15 cigarettes/day, > 15 cigarettes/day), Deprivation Index (quintile 1–3, quintile 4–5), Predictive Risk Modelling (level 1–2, level 3–4), energy intake (kcal/day), dietary fibre (g/day), ethanol intake (g/day), antiplatelet (including non-steroidal anti-inflammatory drugs) and anticoagulants use (dichotomised variable, yes vs. no), including nutrients separately; participants with missing data for the confounding variables were included as a separate category for these variables

<sup>d</sup>Model III, model II including all the nutrients analysed

<sup>e</sup>A value of  $P < 0.005$  was considered significant after the Bonferroni correction (assuming alpha was equal to 0.005,  $\alpha = 0.05/10$ ). Significant results are highlighted in bold

intake (OR<sub>T3 vs. T1</sub> = 0.22, 95% CI: 0.09–0.49) ( $P_{\text{interaction}} = 0.004$ ) (Table S8).

Additionally, even if the remaining combination of SNPs and nutrient intakes did not show any significant interaction for CRC risk, the following variants were associated with a decreased risk of CRC among individuals with moderate-high betaine and/or total choline intake, in both the unadjusted and adjusted models (Tables S5 and S6): *DNMT3B* (rs2424913, rs406193), *DNMT1* (rs2228612),

*MTHFR* (rs1801131, rs1801133), *MTHFD1* (rs8003379, rs17824591), and *Met synthase reductase* (rs1801394, rs10380). These results were confirmed in the subgroup of cases with distal tumour location (Table S7 and S8).

**Table 3** Summary of observed associations between nutrients and SNPs on colorectal cancer risk

Model	Genes	SNP ID (rs), genotypes	Nutrients (tertile)	CRC risk ( $P_{\text{interaction}}^a$ )
Unadjusted <sup>b</sup>	<i>MTHFR</i> (Chr 1)	rs1476413-CC	Choline (T3)	↓ ( <b>0.002</b> )
		rs1801131-TT	Choline (T3)	↓ (0.019)
Adjusted <sup>c</sup>	<i>MTHFR</i> (Chr 1)	rs1476413-CC	Choline (T3)	↓ (0.012)
		<i>MTHFD1</i> (Chr 14)	Vitamin B <sub>12</sub> (T2)	↑ ( <b>0.003</b> )

Abbreviations: C, cytosine; Chr, chromosome; CRC, colorectal cancer; G, guanine; *MTHFD*, methylene tetrahydrofolate dehydrogenase; *MTHFR*, methylene tetrahydrofolate reductase; rs, reference single nucleotide polymorphism; SNP, single nucleotide polymorphism; T, thymine; T1, first tertile; T2, second tertile; T3, third tertile  
Tertiles of nutrient intake: vitamin B<sub>12</sub> (μg/day), T1 ≤ 3.9, T2 4.0–5.3, T3 > 5.3; total choline (mg/day), T1 ≤ 114.0, T2 114.1–190.0, T3 > 190.0; betaine (mg/day), T1 ≤ 119.0, T2 119.1–165.0, T3 > 165.0  
↑, increased risk; ↓, decreased risk

<sup>a</sup>A value of  $P < 0.005$  was considered significant after the Bonferroni correction. Significant results are highlighted in bold

<sup>b</sup>Analysis was performed using crude conditional logistic regression

<sup>c</sup>Analyses were performed using conditional logistic regression analysis adjusted for the following variables (reference categories are underlined): sex (women; men) age (50–59 y old, 60–69 y old), BMI (normal weight, overweight/obesity), physical exercise (< 15 min/day of cycling/sports, ≥ 15 min/day), smoking status (never, past/currently: smoker: ≤ 15 cigarettes/day, > 15 cigarettes/day), Deprivation Index (quintile 1–3, quintile 4–5), Predictive Risk Modelling (level 1–2, level 3–4), energy intake (kcal/day), dietary fibre (g/day), ethanol intake (g/day), antiplatelet (including non-steroidal anti-inflammatory drugs) and anticoagulants use (dichotomised variable, yes vs. no), including SNPs separately; participants with missing data for the confounding variables were included as a separate category for these variables

### Colorectal cancer risk according to nutrient-lifestyle interactions

On the other hand, the combined effects of nutrient intake and lifestyle factors (PE, smoking, and alcohol consumption) on CRC risk were also examined. Individuals who reported both a low and a moderate-high level of PE (OR<sub>T3 vs. T1</sub> = 0.27, 95% CI: 0.13–0.66; OR<sub>T3 vs. T1</sub> = 0.12, 95% CI: 0.07–0.44, respectively) and a low or no alcohol consumption (OR<sub>T3 vs. T1</sub> = 0.34, 95% CI: 0.23–0.61) and had high betaine intake showed a low CRC risk, even if no significant interactions were found (Table 4).

### Discussion

The present study aimed to determine the association between dietary factors and genetic variants related to the FOCM and CRC risk, as well as possible nutrient-gene

interactions. Moreover, the combined effects of nutrient intake and lifestyle factors (PE, smoking, and alcohol consumption) on CRC risk were also examined. Our results suggest that betaine intake and interactions between some dietary factors and variants in *MTHFR* and *MTHFD1* genes have an influence on CRC risk in the population studied. The results were confirmed in the subgroup with distal tumour location. On the other hand, no significant interactions were observed between nutrient intake and lifestyle factors on CRC risk.

As we mentioned in the [introduction](#) section, to date, few epidemiologic studies have examined the association between betaine and cancer risk, and those who have studied this possible relationship have obtained inconsistent results. Several researchers found an inverse association between betaine intake and breast cancer risk [34]. However, other studies found no evidence that higher intakes of this nutrient reduced the risk of breast cancer [35]. Some studies have reported that a higher intake of betaine was associated with a reduced risk of lung cancer [36], whereas no association was found for epithelial ovarian cancer [37]. It has been suggested that the underlying mechanisms by which a high intake of betaine would reduce the risk of some cancers would be similar to those of folate. Betaine can donate the methyl group to homocysteine as does folate, although the donation of the methyl group by betaine is limited to the liver and the kidney [34]. A high intake of betaine could help prevent the adverse effect resulting from hypomethylation of DNA or restore DNA repair mechanism, and therefore, would lead to reduced cancer risk [36].

Inconsistent results were also observed on the relationship between betaine intake and CRC risk. The Health Professionals Follow-up Study conducted in the United States [38] and an investigation carried out in a Chinese population [39] have examined this possible association and in both studies, no association was found. This result has been attributed, in part, to the fact that the levels of this nutrient would not be critical in folate-nourished populations, because folate and choline metabolic pathways are highly interrelated, and betaine is derived from choline and increase in response to a higher choline intake [40].

Our study, like another case-control study, where plasma betaine was analysed, [41] confirmed the inverse association between betaine intake and CRC risk, even among subjects with an average total folate intake below population recommendations and below the intakes recorded in other studies, such as the Health Professionals Follow-up Study mentioned above, [38] with a total folate intake from diet and supplements of 479–858 μg/day in the total sample. In the present study, the average folate intake was 270.9 μg/day among controls and therefore the intake level of folate was not very high.

**Table 4** Associations between lifestyle factors and colorectal cancer risk, stratified by dietary factors (adjusted model)<sup>a</sup>

Lifestyle factors, stratified by dietary factors	Nutrient intake									<i>P</i> <sub>interaction</sub> <sup>b</sup>
	T1			T2			T3			
	Cases/ Con- trols, n	OR (95%CI) <sup>a</sup>	<i>P</i> <sup>2</sup>	Cases/ Con- trols, n	OR (95%CI) <sup>a</sup>	<i>P</i> <sup>2</sup>	Cases/ Con- trols, n	OR (95%CI) <sup>a</sup>	<i>P</i> <sup>2</sup>	
Folate										
Physical exercise										
≥15 min/day	32/20	1.00	0.760	26/15	1.45 (0.44–4.78)	0.522	21/17	1.63 (0.42–6.49)	0.498	0.938
<15 min/day	49/63	0.89 (0.34–2.22)		54/56	0.95 (0.33–2.28)	0.916	47/58	1.07 (0.34–3.41)	0.892	
Smoking status										
Never	23/35	1.00	0.464	22/25	1.33 (0.42–3.99)	0.539	24/29	1.56 (0.46–5.16)	0.483	0.233
Ever	58/48	1.42 (0.31–3.48)		58/46	1.51 (0.60–3.81)	0.391	44/46	1.81 (0.62–5.54)	0.322	
Alcohol consumption <sup>c</sup>										
Abstemious/low risk	70/73	1.00	0.987	68/63	1.34 (0.66–2.66)	0.450	61/68	1.35 (0.53–3.35)	0.543	0.891
High risk	11/10	1.01 (0.31–3.30)		12/8	0.80 (0.24–2.65)	0.650	7/7	1.32 (0.35–6.01)	0.660	
Vitamin B <sub>2</sub>										
Physical exercise										
≥15 min/day	34/21	1.00	0.389	31/17	0.83 (0.29–2.52)	0.719	14/14	0.55 (0.16–1.89)	0.298	0.836
<15 min/day	51/62	0.72 (0.33–1.61)		60/54	0.69 (0.30–1.60)	0.350	39/61	0.25 (0.10–0.90)	0.031	
Smoking status										
Never	28/33	1.00	0.337	22/29	0.34 (0.14–0.98)	0.049	19/27	0.36 (0.11–1.01)	0.051	0.400
Ever	57/50	0.70 (0.32–1.55)		69/42	1.03 (0.44–2.41)	0.975	34/48	0.35 (0.13–1.00)	0.050	
Alcohol consumption <sup>c</sup>										
Abstemious/low risk	75/74	1.00	0.438	76/64	0.81 (0.30–1.47)	0.512	48/66	0.47 (0.20–1.12)	0.073	0.350
High risk	10/9	0.61 (0.18–2.23)		15/7	1.93 (0.53–8.32)	0.422	5/9	0.20 (0.09–0.79)	0.023	
Vitamin B <sub>6</sub>										
Physical exercise										
≥15 min/day	28/16	1.00	0.359	34/22	0.65 (0.26–1.91)	0.432	17/14	0.48 (0.07–2.03)	0.291	0.885
<15 min/day	40/51	0.65 (0.23–1.62)		65/63	0.55 (0.20–1.46)	0.233	45/63	0.37 (0.14–1.09)	0.070	
Smoking status										
Never	16/28	1.00	0.964	33/29	0.75 (0.25–2.14)	0.613	20/32	0.38 (0.12–1.21)	0.105	0.299
Ever	52/39	1.02 (0.41–2.64)		66/56	0.83 (0.32–2.25)	0.712	42/45	0.68 (0.22–2.10)	0.479	
Alcohol consumption <sup>c</sup>										
Abstemious/low risk	62/60	1.00	0.401	82/76	0.70 (0.37–1.40)	0.301	55/7	0.50 (0.20–1.27)	0.187	0.389
High risk	6/7	0.57 (0.15–2.30)		17/9	1.10 (0.32–3.60)	0.885	7/9	0.33 (0.12–1.30)	0.103	
Vitamin B <sub>12</sub>										
Physical exercise										
≥15 min/day	27/16	1.00	0.060	25/19	0.41 (0.12–1.33)	0.135	27/17	0.70 (0.20–2.09)	0.419	0.720
<15 min/day	43/57	0.41 (0.12–1.09)		49/61	0.51 (0.19–1.33)	0.166	58/59	0.61 (0.18–2.05)	0.409	
Smoking status										
Never	21/26	1.00	0.890	23/35	0.73 (0.29–2.02)	0.539	25/28	0.99 (0.31–3.20)	0.991	0.829
Ever	49/47	1.09 (0.43–2.77)		51/45	1.11 (0.43–2.76)	0.869	60/48	1.22 (0.41–3.65)	0.712	
Alcohol consumption <sup>c</sup>										
Abstemious/low risk	63/65	1.00	0.282	61/74	0.88 (0.46–1.59)	0.577	75/65	1.10 (0.47–2.67)	0.802	0.353
High risk	7/8	0.52 (0.13–1.81)		13/6	1.70 (0.44–6.67)	0.459	10/11	1.02 (0.22–2.71)	0.912	
Met										
Physical exercise										



**Table 4** (continued)

Lifestyle factors, stratified by dietary factors	Nutrient intake									$P_{\text{interaction}}^b$
	T1			T2			T3			
	Cases/ Con- trols, n	OR (95%CI) <sup>a</sup>	$P^2$	Cases/ Con- trols, n	OR (95%CI) <sup>a</sup>	$P^2$	Cases/ Con- trols, n	OR (95%CI) <sup>a</sup>	$P^2$	
≥15 min/day	20/18	1.00	0.701	36/20	1.34 (0.48–3.71)	0.572	23/14	0.90 (0.20–3.80)	0.878	0.740
<15 min/day	52/58	1.16 (0.43–3.09)		53/58	1.21 (0.47–3.31)	0.680	45/61	0.52 (0.18–1.33)	0.177	
<b>Smoking status</b>										
Never	21/31	1.00	0.879	31/31	1.20 (0.44–3.18)	0.755	17/27	0.41 (0.16–1.10)	0.079	0.477
Ever	51/45	1.08 (0.44–2.60)		58/47	1.25 (0.46–3.15)	0.643	51/48	0.65 (0.27–1.73)	0.369	
<b>Alcohol consumption<sup>c</sup></b>										
Abstemious/low risk	63/68	1.00	0.925	75/71	1.12 (0.60–2.23)	0.707	61/65	0.49 (0.30–1.09)	0.080	0.798
High risk	9/8	0.95 (0.32–3.45)		14/7	1.10 (0.29–3.37)	0.822	7/10	0.42 (0.10–1.73)	0.221	
<b>Choline</b>										
<b>Physical exercise</b>										
≥15 min/day	37/13	1.00	0.364	29/20	0.59 (0.16–1.80)	0.330	13/19	0.25 (0.10–1.13)	0.070	0.710
<15 min/day	68/62	0.60 (0.20–1.81)		46/57	0.32 (0.13–0.88)	0.029	36/58	0.40 (0.12–1.08)	0.069	
<b>Smoking status</b>										
Never	28/31	1.00	0.678	23/30	0.49 (0.17–1.54)	0.232	18/28	0.60 (0.22–1.52)	0.251	0.998
Ever	77/44	1.16 (0.53–2.69)		52/47	0.74 (0.39–1.72)	0.450	31/49	0.66 (0.21–1.56)	0.269	
<b>Alcohol consumption<sup>c</sup></b>										
Abstemious/low risk	92/72	1.00	0.361	65/64	0.66 (0.36–1.16)	0.163	42/68	0.61 (0.26–1.05)	0.077	0.290
High risk	13/3	2.31 (0.53– 11.97)		10/13	0.27 (0.11–0.82)	0.021	7/9	0.79 (0.21–2.89)	0.730	
<b>Betaine</b>										
<b>Physical exercise<sup>c</sup></b>										
≥15 min/day	57/15	1.00	0.508	13/18	0.48 (0.15–1.69)	0.242	9/19	0.12 (0.07–0.44)	<b>0.002</b>	0.218
<15 min/day	93/62	0.76 (0.33–1.78)		28/57	0.22 (0.09–0.55)	<b>0.002</b>	29/58	0.27 (0.13–0.66)	<b>0.004</b>	
<b>Smoking status<sup>c</sup></b>										
Never	45/30	1.00	0.330	10/32	0.27 (0.09–0.83)	0.022	14/27	0.50 (0.18–1.38)	0.184	0.351
Ever	105/47	1.51 (0.66–3.48)		31/43	0.56 (0.22–1.46)	0.235	24/50	0.26 (0.10–0.66)	0.005	
<b>Alcohol consumption<sup>c</sup></b>										
Abstemious/low risk	131/72	1.00	0.558	37/65	0.41 (0.20–0.78)	0.009	6/13	0.34 (0.23–0.61)	<b>0.001</b>	0.298
High risk	19/5	1.62 (0.37–6.61)		4/10	0.36 (0.08–0.93)	0.044	32/64	0.31 (0.10–0.98)	0.048	

Abbreviations: CI, confidence interval; Met, methionine; OR, odds ratio; T1, first tertile; T2, second tertile; T3, third tertile

Tertiles of nutrient intake: folate (µg/day), for males, T1 ≤ 220.0, T2 220.1–289.0, T3 > 289.1, and females, T1 ≤ 245.0, T2 245.1–300.0, T3 > 300.0; vitamin B<sub>2</sub> (mg/day), T1 ≤ 1.3, T2 1.4–1.7, T3 > 1.70; vitamin B<sub>6</sub> (mg/day), T1 ≤ 1.5, T2 1.6–2.0, T3 > 2.0; vitamin B<sub>12</sub> (µg/day), T1 ≤ 3.9, T2 4.0–5.3, T3 > 5.3; Met (mg/day), for males, T1 ≤ 1324.0, T2 1324.1–1985.0, T3 > 1985.0, and females, T1 ≤ 1564.0, T2 1564.1–2623.0, T3 > 2623.0; choline (mg/day), T1 ≤ 114.0, T2 114.1–190.0, T3 > 190.0; betaine (mg/day), T1 ≤ 119.0, T2 119.1–165.0, T3 > 165.0

<sup>a</sup>Analyses were performed using conditional logistic regression analysis adjusted for the following variables (reference categories are underlined): sex (women; men) age (50–59 y old, 60–69 y old), BMI (normal weight, overweight/obesity), physical exercise (< 15 min/day of cycling/sports, ≥ 15 min/day), smoking status (never, past/currently: smoker: ≤ 15 cigarettes/day, > 15 cigarettes/day), Deprivation Index (quintile 1–3, quintile 4–5), Predictive Risk Modelling (level 1–2, level 3–4), energy intake (kcal/day), dietary fibre (g/day), ethanol intake (g/day), antiplatelet (including non-steroidal anti-inflammatory drugs) and anticoagulants use (dichotomised variable, yes vs. no), including lifestyle factors separately; participants with missing data for the confounding variables were included as a separate category for these variables. In the analyses of the variables physical exercise, smoking status, and alcohol consumption these same variables were excluded as an adjustment variable

<sup>b</sup>A value of  $P < 0.005$  was considered significant after the Bonferroni correction. Significant results are highlighted in bold

<sup>c</sup>Conditional exact logistic regression

Nevertheless, the few published on the association between betaine intake and CRC risk are confusing. Neither the Health Professionals Follow-up Study [38] nor the one carried out by Lu et al. [39] found significant associations for betaine intakes. These discrepancies could be due to differences in the characteristics of participants, the intake and status of other nutrients involved in FOCM, and the design type. In summary, in the Health Professionals Follow-up Study, [38] the participants were US male health professionals aged 40 to 75 years, and in Lu et al.'s study, [39], Chinese males and females aged 30 to 75 years. The folate intake in the total sample of Lee et al.' study [38] was of 479–858  $\mu\text{g}/\text{day}$ , whereas, in the control sample of Lu et al.'s study [39] was 240.3  $\mu\text{g}/\text{d}$ . Lee et al.' study [38] is a prospective cohort study, whereas, Lu et al.'s study [39] is a case-control study.

Regarding the possible associations between genetic variants related to the FOCM and CRC risk, in the current study, a genotype in *DNMT3B* (rs2424913) was related to CRC risk, even though this relationship was not significant after the Bonferroni correction. Other authors did not find a close correlation between this genetic variant and the development of CRC among the Chinese population [42]. As regards potential nutrient-gene interactions on CRC risk, rs1476413, and rs17824591 exhibited significant nutrient-gene interactions with total choline and vitamin B<sub>12</sub> intakes, respectively. These findings suggest that these variants in *MTHFR* and *MTHFD1* genes interact with dietary factors to modify the risk for CRC. It should be noted that the mechanism by which choline intake/status affects DNA integrity is not entirely clear but may be through effects on mitochondrial membrane integrity and oxidative stress [43]. Regarding vitamin B<sub>12</sub>, the underlying mechanism by which a high intake of this vitamin would reduce CRC risk could be related to tumour methylation, as other authors have pointed out in colon cancer [44].

In another case-control study, however, no interactions were found between the *MTHFR* rs1476413 SNP and dietary factors (including folate and Met) on CRC risk, although other *MTHFR* and *MTHFD1* SNPs exhibited gene-diet interactions with Met intake [28]. The lack of data on the possible interaction between the aforementioned SNPs, and their interaction with betaine, makes it difficult to compare our results to other studies.

Most studies assessing *MTHFR* and CRC risk have focused on the rs1801133 SNP. The variant allele in this SNP causes an increase in thermolability of the *MTHFR* enzyme [45] which is associated with decreased plasma folate and increased plasma homocysteine [46]. The potential influence of *MTHFR* activity on DNA methylation and the availability of uridylates and thymidylates for DNA synthesis and repair makes *MTHFR* an attractive candidate for

a cancer-predisposing gene. Even if other studies found an association between this SNP and the CRC risk and interaction with Met, [42] in the present study no significant association was found between these factors. Furthermore, few studies have assessed polymorphisms in *MTHFD1* in relation to risk for CRC [7, 42]. Nevertheless, those SNPs (rs2295638, rs2236225) in which they found interaction with nutrients, specifically with Met [7, 42, 47] were not any of those analysed in the present study.

In any case, even if in the present study no more nutrient-gene interactions were detected, individuals with moderate-high betaine and/or total choline intakes showed a decreased risk, for all genotypes analysed, both in the total sample and in the subgroup with distal tumour location. Finally, concerning interactions between nutrient intakes and lifestyle factors, although no interactions were found, individuals who reported low or no alcohol consumption, and had moderate-high levels of PE and high betaine intake showed the lowest risk of CRC. Even though no data have been found in the literature on these interactions, our findings agree with previous studies about the interaction effects of folate status and lifestyle factors on CRC risk [48].

The main strength of this study compared to other previous published [49] is that colonoscopy was used as a diagnosis criterion to identify both cases and controls to avoid false positives and negatives. To our knowledge, to date, only one other study of the association between diet and CRC risk has been published, in which it was confirmed that controls were free of the disease through colonoscopy [50]. Another strength is the fact that information is provided based on a standardised protocol including not only dietary factors but also other possible determinants of CRC such as health determinants and weight status among others. However, some limitations should be mentioned. First, recall bias is also of concern in case-control studies. Second, the small sample size makes it difficult to detect possible associations and nutrient-gene and nutrient-lifestyle interactions and disease risk, since some genotypes and categories according to lifestyle factors showed low frequencies in our population.

Another disadvantage of the small sample size is that they can produce false-positive results; to avoid it, the Bonferroni correction was used. Third, self-reported data could be subject to measurement errors and the problem of food omissions due to memory failure and under-reporting of unhealthy habits among disease subjects. However, previous validation studies indicate that the self-reported dietary information is reported with sufficient accuracy for use in epidemiology analysis [51]. Fourth, although the FFQ used to collect information on dietary intake in the present study has been validated among people who lived in the same region, this validation did not include specific nutrients such

as betaine or total choline, but it included the main food sources of these methyl-donors. So, the possible measurement errors are most likely non-differential and thus do not explain the inverse associations observed in our study.

Fifth, even though data on lifestyles (including diet) were recorded retrospectively—that is, the questions referred to behaviours before participating in the BCSP—it should be also noted that dietary changes are usually modest after participating in the BCSP due to a lack of information and personalised advice [52]. In addition, adults generally maintain relatively stable eating habits for a long time [53]. Therefore, the results of this study are unlikely to be greatly affected by potential changes in eating. Sixth, although we have adjusted for several confounding factors, some residual confounding may result from the misclassification of those variables and confounding by unmeasured variables. Finally, to avoid selection bias of controls, we obtained controls from the same BCSP and in the same period as cases. It should also be worth noting that we have compared our results with those of the Health Professionals Follow-up Study, [38] that only includes male participants.

In conclusion, the present study suggests that high betaine intake is associated with decreased risk of CRC among the population studied. Moreover, our results support the existing hypothesis of genetic-nutrient interactions in colorectal carcinogenesis. Total choline and vitamin B<sub>12</sub> intake, and the SNPs rs1476413 and rs17824591 may be related, respectively, to CRC risk in this population. High total choline intake together with the *MTHFR* rs1476413-CC genotype reduces CRC risk, whereas moderate vitamin B<sub>12</sub> intake together with the *MTHFD1* rs17824591-GG genotype increases CRC risk.

Further studies are necessary to confirm these associations and understand in depth their role in colorectal carcinogenesis, including participants under 50 years old. Understanding the interaction between nutrition and genetic variation can be useful to distinguish between individuals who will and who will not benefit from diet intervention strategies.

**Supplementary Information** The online version contains supplementary material available at <https://doi.org/10.1007/s00394-024-03371-5>.

**Acknowledgements** The author would like to thank the patients who enrolled in this study for their participation, and the DNA Bank unit of the General Genomics and Proteomics Service at the UPV/EHU and the Basque Biobank for Research-OEHUN for its collaboration. The genotyping service was carried out at CEGEN-PRB2-ISCI; it is supported by grant PT13/0001, ISCI-SGEFI / FEDER.

**Author Contributions** Marta Arroyo-Izaga: conceptualisation, data curation, formal analysis, software, resources, project administration, writing-original draft, validation, visualisation, writing review & editing, funding acquisition, and supervision. Iker Alegria-Lertxundi:

conceptualisation, data curation, formal analysis, software, writing-original draft, and writing review & editing. Sara Corchero-Palacios: formal analysis, and writing-original draft. Marian M. de Pancorbo: conceptualisation, validation, visualisation, writing review & editing, funding acquisition, and supervision. The work reported in the paper has been performed by the authors unless clearly specified in the text.

**Funding** This research was funded by the Basque Government: Department of Health and Consumer Affairs (201111153), the Saiotek program (S-PE12UN058), and Education Department (BIOMICs Research Group, MICROFLUIDICs & BIOMICs Cluster of the University of the Basque Country UPV/EHU, No. IT1633-22). I.A.-L. was funded by a pre-doctoral grant from the Basque Government (PRE\_2014\_1\_161, PRE\_2015\_2\_0084, EP\_2016\_1\_0098, EP\_2016\_1\_0098, and PRE\_2017\_2\_0006). Open Access funding provided by the University of the Basque Country UPV/EHU. The funders had no role in study design, data collection, data analysis, data interpretation, or writing of the report. Open Access funding provided thanks to the CRUE-CSIC agreement with Springer Nature.

**Data availability** The data that support the findings of this study are available from the corresponding author upon reasonable request.

## Declarations

**Conflict of interest** The authors declare no conflict of interest.

**Ethics Statement** The study was approved by the Clinical Research Ethics Committee of the Basque Country (protocol code PI2011006, data of approval 03/23/2012; and PI2014042, data of approval 05/28/2014). All participants gave written informed consent before enrolment in the study, which was conducted in accordance with the principles of the Declaration of Helsinki.

**Open Access** This article is licensed under a Creative Commons Attribution 4.0 International License, which permits use, sharing, adaptation, distribution and reproduction in any medium or format, as long as you give appropriate credit to the original author(s) and the source, provide a link to the Creative Commons licence, and indicate if changes were made. The images or other third party material in this article are included in the article's Creative Commons licence, unless indicated otherwise in a credit line to the material. If material is not included in the article's Creative Commons licence and your intended use is not permitted by statutory regulation or exceeds the permitted use, you will need to obtain permission directly from the copyright holder. To view a copy of this licence, visit <http://creativecommons.org/licenses/by/4.0/>.

## References

1. Morgan E, Arnold M, Gini A, Lorenzoni V, Cabasag CJ, Laversanne M, Vignat J, Ferlay J, Murphy N, Bray F (2023) Global burden of colorectal cancer in 2020 and 2040: incidence and mortality estimates from GLOBOCAN. *Gut* 72(2):338–344. <https://doi.org/10.1136/gutjnl-2022-327736>
2. Cancer Observatory of the Spanish Association against Cancer (2022) Accessed December 28, 2023. <https://observatorio.contraelcancer.es/>
3. World Health Organization. Cancer Today. Global Cancer Observatory [GCO] (2020) Accessed December 28, 2023. <http://gco.iarc.fr/today>

4. Keum N, Giovannucci E (2019) Global burden of colorectal cancer: emerging trends, risk factors and prevention strategies. *Nat Rev Gastroenterol Hepatol* 16(12):713–732. <https://doi.org/10.1038/s41575-019-0189-8>
5. Choi SW, Mason JB (2002) Folate status: effects on pathways of colorectal carcinogenesis. *J Nutr* 132(8 Suppl). <https://doi.org/10.1093/jn/132.8.2413s>. :2413S-2418S
6. Fedirko V, Tramacere I, Bagnardi V, Rota M, Scotti L, Islami F, Negri E, Straif K, Romieu I, La Vecchia C, Boffetta P, Jenab M (2011) Alcohol drinking and colorectal cancer risk: an overall and dose-response meta-analysis of published studies. *Ann Oncol* 22(9):1958–1972. <https://doi.org/10.1093/annonc/mdq653>
7. Liu AY, Scherer D, Poole E, Potter JD, Curtin K, Makar K, Slatery ML, Caan BJ, Ulrich CM (2013) Gene-diet-interactions in folate-mediated one-carbon metabolism modify colon cancer risk. *Mol Nutr Food Res* 57(4):721–734. <https://doi.org/10.1002/mnfr.201200180>
8. Alegria-Lertxundi I, Aguirre C, Bujanda L, Fernández FJ, Polo F, Ordovás JM, Etxezarraga MC, Zabalza I, Larzabal M, Portillo I, M de Pancorbo M, Palencia-Madrid L, Garcia-Etxebarria K, Rocandio AM, Arroyo-Izaga M (2020) Gene-diet interactions in colorectal cancer: survey design, instruments, participants and descriptive data of a case-control study in the Basque Country. *Nutrients* 12(8):2362. <https://doi.org/10.3390/nu12082362>
9. Rodríguez IT, Ballart JF, Pastor GC, Jordà EB, Val VA (2008) Validación De Un Cuestionario De frecuencia de consumo alimentario corto: reproducibilidad y validez [Validation of a short questionnaire on frequency of dietary intake: reproducibility and validity]. *Nutr Hosp* 23(3):242–252 PMID: 18560701
10. Telleria-Aramburu N, Alegria-Lertxundi I, Arroyo-Izaga M (2021) Adaptation, validation and reproducibility of a short FFQ to assess food group intake in the population resident in the basque country (Spain). *Public Health Nutr* 24(3):436–448. <https://doi.org/10.1017/s1368980020001822>
11. Alegria-Lertxundi I, Alvarez M, Rocandio AM, de Pancorbo MM, Arroyo-Izaga M (2016) Nutritional adequacy and diet quality in colorectal cancer patients postsurgery: a pilot study. *Nutr Cancer* 68(4):577–588. <https://doi.org/10.1080/01635581.2016.1158299>
12. Carbajal A, Sánchez-Muniz FJ, Secretariat of Publications and Audiovisual Media (2003) *Guía de prácticas en Nutrición y Dietética [Guide to Practices in Nutrition and Dietetics]*. University of León, ; :1–3. Accessed December 10, 2020. <https://www.ucm.es/data/cont/docs/458-2019-01-04-Guia-Practicas-2019-web.pdf>
13. Department of Agriculture, Fisheries and Food, Basque Government. Estudio cuantitativo del consumo de alimentos en la CAPV [Quantitative study of food consumption in the autonomous community of the Basque Country]. Basque Government: Central Publication Service (2008) Accessed December 9, 2020. [https://www.euskadi.eus/contenidos/informacion/coleccion\\_elika/es\\_dapa/adjuntos/Guia\\_Elika\\_8.pdf](https://www.euskadi.eus/contenidos/informacion/coleccion_elika/es_dapa/adjuntos/Guia_Elika_8.pdf)
14. U.S. Department of Agriculture, Agricultural Research Service. USDA's FoodData Central (2017) Accessed August 9, 2018. <https://fdc.nal.usda.gov/>
15. Zeisel SH, Mar MH, Howe JC, Holden JM (2003) Concentrations of choline-containing compounds and betaine in common foods. *J Nutr* 133(5):1302–1307. <https://doi.org/10.1093/jn/133.5.1302>
16. Iles-Shih L, Collins JF, Holub JL, Lieberman DA (2010) Prevalence of significant neoplasia in FOBt-positive patients on warfarin compared with those not on warfarin. *Am J Gastroenterol* 105(9):2030–2035. <https://doi.org/10.1038/ajg.2010.264>
17. Ministry of Health, Social Services and Equality, Government of Spain / National Statistics Institute. Encuesta Nacional de Salud. España 2011/12 [National Health Survey, Spain 2011/12]. Accessed December 20 (2019) <http://www.msssi.gob.es/estadEstudios/estadisticas/sisInfSanSNS/nivelSalud.htm>
18. World Health Organization. Obesity: preventing and managing the global epidemic: report of a WHO consultation, 2000. Accessed December 5 (2019) <https://apps.who.int/iris/handle/10665/42330>
19. Silva Rodrigues RA, Martinez Espinosa M, Duarte Melo C, Rodrigues Perracini M, Rezende Fett WC, Fett CA (2014) New values anthropometry for classification of nutritional status in the elderly. *J Nutr Health Aging* 18(7):655–661. <https://doi.org/10.1007/s12603-014-0451-2>
20. Serra-Majem L, Aranceta J, SENC Working Group on Nutritional Objectives for the Spanish Population (2001) Spanish Society of Community Nutrition. Nutritional objectives for the Spanish population. Consensus from the Spanish Society of Community Nutrition. *Public Health Nutr* 4(6A):1409–1413. <https://doi.org/10.1079/phn2001229>
21. Ministry of Health and Consumption, Government of Spain. Informes de la Comisión Clínica 2017: Alcohol [Clinical Commission Reports: Alcohol]. Accessed December 10 (2019) <http://www.pnsd.msc.es/Categoria2/publica/pdf/InformeAlcohol.pdf>
22. Alegria-Lertxundi I, Aguirre C, Bujanda L, Fernández FJ, Polo F, Ordovás JM, Etxezarraga MC, Zabalza I, Larzabal M, Portillo I, de Pancorbo MM, Palencia-Madrid L, Rocandio AM, Arroyo-Izaga M (2019) Single nucleotide polymorphisms associated with susceptibility for development of colorectal cancer: case-control study in a basque population. *PLoS ONE* 14(12):e0225779. <https://doi.org/10.1371/journal.pone.0225779>
23. Orueta JF, Nuño-Solinis R, Mateos M, Vergara I, Grandes G, Esnaola S (2013) Predictive risk modelling in the Spanish population: a cross-sectional study. *BMC Health Serv Res* 13:269. <https://doi.org/10.1186/1472-6963-13-269>
24. School of Public Health, Johns Hopkins University. The Johns Hopkins University ACG Case-Mix System. Accessed August 10 (2012) [http://www.acg.jhsph.org/index.php?option=com\\_content&view=article&id=46&Itemid=61](http://www.acg.jhsph.org/index.php?option=com_content&view=article&id=46&Itemid=61)
25. Verisk Analytics, Inc. DxCG® Introduces Disease Management Calculator. Accessed August 10 (2012) <https://www.verisk.com/archived/dxcg-introduces-disease-management-calculator/>
26. de Vogel S, Wouters KA, Gottschalk RW, van Schooten FJ, de Goeij AF, de Bruïne AP, Goldbohm RA, van den Brandt PA, van Engeland M, Weijnenberg MP (2011) Dietary methyl donors, methyl metabolizing enzymes, and epigenetic regulators: diet-gene interactions and promoter CpG island hypermethylation in colorectal cancer. *Cancer Causes Control* 22(1):1–12. <https://doi.org/10.1007/s10552-010-9659-6>
27. Li H, Liu JW, Sun LP, Yuan Y (2017) A Meta-analysis of the association between *DNMT1* polymorphisms and cancer risk. *Biomed Res Int* 2017:3971259. <https://doi.org/10.1155/2017/3971259>
28. Ashmore JH, Lesko SM, Muscat JE, Gallagher CJ, Berg AS, Miller PE, Hartman TJ, Lazarus P (2013) Association of dietary and supplemental folate intake and polymorphisms in three FOCM pathway genes with colorectal cancer in a population-based case-control study. *Genes Chromosomes Cancer* 52(10):945–953. <https://doi.org/10.1002/gcc.22089>
29. Pardini B, Kumar R, Naccarati A, Prasad RB, Forsti A, Polakova V, Vodickova L, Novotny J, Hemminki K, Vodicka P (2011) MTHFR and MTRR genotype and haplotype analysis and colorectal cancer susceptibility in a case-control study from the Czech Republic. *Mutat Res* 721(1):74–80. <https://doi.org/10.1016/j.mrgentox.2010.12.008>
30. Service of Epidemiology (2004) Dirección Xeral de Saúde Pública, Consellería de Sanidade, Xunta de Galicia, Spain; Unit of Health Analysis and Information Systems, Pan-American Health Organization (PAHO), Washington, DC. *Epidat 3.0 program*. Santiago de Compostela, A Coruña, Spain

31. Torres Stone RA, Waring ME, Cutrona SL, Kiefe CI, Allison J, Doubeni CA (2017) The association of dietary quality with colorectal cancer among normal weight, overweight and obese men and women: a prospective longitudinal study in the USA. *BMJ Open* 7(6):e015619. <https://doi.org/10.1136/bmjopen-2016-015619>
32. Doubeni CA, Laiyemo AO, Major JM, Schootman M, Lian M, Park Y, Graubard BI, Hollenbeck AR, Sinha R (2012) Socioeconomic status and the risk of colorectal cancer: an analysis of more than a half million adults in the National Institutes of Health-AARP Diet and Health Study. *Cancer* 118(14):3636–3644. <https://doi.org/10.1002/ncr.26677>
33. Romaguera D, Vergnaud AC, Peeters PH, van Gils CH, Chan DS, Ferrari P, Romieu I, Jenab M, Slimani N, Clavel-Chapelon F, Fagherazzi G, Perquier F, Kaaks R, Teucher B, Boeing H, von Ruesten A, Tjønneland A, Olsen A, Dahm CC, Overvad K, Quirós JR, Gonzalez CA, Sánchez MJ, Navarro C, Barricarte A, Dorransoro M, Khaw KT, Wareham NJ, Crowe FL, Key TJ, Trichopoulos A, Lagiou P, Bamia C, Masala G, Vineis P, Tumino R, Sieri S, Panico S, May AM, Bueno-de-Mesquita HB, Büchner FL, Wirfält E, Manjer J, Johansson I, Hallmans G, Skeie G, Benjamínsen Borch K, Parr CL, Riboli E, Norat T (2012) Is concordance with World Cancer Research Fund/American Institute for Cancer Research guidelines for cancer prevention related to subsequent risk of cancer? Results from the EPIC study. *Am J Clin Nutr* 96(1):150–163. <https://doi.org/10.3945/ajcn.111.031674>
34. Zhang CX, Pan MX, Li B, Wang L, Mo XF, Chen YM, Lin FY, Ho SC (2013) Choline and betaine intake is inversely associated with breast cancer risk: a two-stage case-control study in China. *Cancer Sci* 104(2):250–258. <https://doi.org/10.1111/cas.12064>
35. Cho E, Holmes MD, Hankinson SE, Willett WC (2010) Choline and betaine intake and risk of breast cancer among postmenopausal women. *Br J Cancer* 102(3):489–494. <https://doi.org/10.1038/sj.bjc.6605510>
36. Ying J, Rahbar MH, Hallman DM, Hernandez LM, Spitz MR, Forman MR, Gorlova OY (2013) Associations between dietary intake of choline and betaine and lung cancer risk. *PLoS ONE* 8(2):e54561. <https://doi.org/10.1371/journal.pone.0054561>
37. Kotsopoulos J, Hankinson SE, Tworoger SS (2010) Dietary betaine and choline intake are not associated with risk of epithelial ovarian cancer. *Eur J Clin Nutr* 64(1):111–114. <https://doi.org/10.1038/ejcn.2009.109>
38. Lee JE, Giovannucci E, Fuchs CS, Willett WC, Zeisel SH, Cho E (2010) Choline and betaine intake and the risk of colorectal cancer in men. *Cancer Epidemiol Biomarkers Prev* 19(3):884–887. <https://doi.org/10.1158/1055-9965.EPI-09-1295>
39. Lu MS, Fang YJ, Pan ZZ, Zhong X, Zheng MC, Chen YM, Zhang CX (2015) Choline and betaine intake and colorectal cancer risk in Chinese population: a case-control study. *PLoS ONE* 10(3):e0118661. <https://doi.org/10.1371/journal.pone.0118661>
40. Yan J, Jiang X, West AA, Perry CA, Malysheva OV, Devapatla S, Pressman E, Vermeylen F, Stabler SP, Allen RH, Caudill MA (2012) Maternal choline intake modulates maternal and fetal biomarkers of choline metabolism in humans. *Am J Clin Nutr* 95(5):1060–1071. <https://doi.org/10.3945/ajcn.111.022772>
41. Myte R, Gylling B, Schneede J, Ueland PM, Häggström J, Hultdin J, Hallmans G, Johansson I, Palmqvist R, Van Gulpen B (2016) Components of one-carbon Metabolism Other than Folate and Colorectal Cancer Risk. *Epidemiology* 27(6):787–796. <https://doi.org/10.1097/EDE.0000000000000529>
42. Bao Q, He BS, Chen LP, Gu L, Nie ZL, Wang SK (2012) Correlation between polymorphism in the promoter of DNA methyltransferase-3B and the risk of colorectal cancer. *Zhonghua Yu Fang Yi Xue Za Zhi* 46(1):53–57 PMID: 22490141
43. da Costa KA, Niculescu MD, Craciunescu CN, Fischer LM, Zeisel SH (2006) Choline deficiency increases lymphocyte apoptosis and DNA damage in humans. *Am J Clin Nutr* 84:88–94. <https://doi.org/10.1093/ajcn/84.1.88>
44. Mokarram P, Naghibalhossaini F, Firoozi MS, Hosseini SV, Izadpanah A, Salahi H, Malek-Hosseini SA, Talei A, Mojallal M (2008) Methylenetetrahydrofolate reductase C677T genotype affects promoter methylation of tumor-specific genes in sporadic colorectal cancer through an interaction with folate/vitamin B12 status. *World J Gastroenterol* 14(23):3662–3671. <https://doi.org/10.3748/wjg.14.3662>
45. Frosst P, Blom HJ, Milos R, Goyette P, Sheppard CA, Matthews RG, Boers GJ, den Heijer M, Kluijtmans LA, van den Heuvel LP, Rozen R (1995) A candidate genetic risk factor for vascular disease: a common mutation in methylenetetrahydrofolate reductase. *Nat Genet* 10(1):111–113. <https://doi.org/10.1038/ng0595-111>
46. Ma J, Stampfer MJ, Hennekens CH, Frosst P, Selhub J, Horsford J, Malinow MR, Willett WC, Rozen R (1996) Methylenetetrahydrofolate reductase polymorphism, plasma folate, homocysteine, and risk of myocardial infarction in US physicians. *Circulation* 94(10):2410–2416. <https://doi.org/10.1161/01.cir.94.10.2410>
47. Razzak AA, Oxentenko AS, Vierkant RA, Tillmans LS, Wang AH, Weisenberger DJ, Laird PW, Lynch CF, Anderson KE, French AJ, Haile RW, Harnack LJ, Potter JD, Slager SL, Smyrk TC, Thibodeau SN, Cerhan JR, Limburg PJ (2012) Associations between intake of folate and related micronutrients with molecularly defined colorectal cancer risks in the Iowa women's Health Study. *Nutr Cancer* 64(7):899–910. <https://doi.org/10.1080/01635581.2012.714833>
48. Panprathip P, Petmitr S, Tungtrongchitr R, Kaewkungwal J, Kwanbunjan K (2019) Low folate status, and MTHFR 677C>T and MTR 2756A>G polymorphisms associated with colorectal cancer risk in Thais: a case-control study. *Nutr Res* 72:80–91. <https://doi.org/10.1016/j.nutres.2019.10.008>
49. Castelló A, Amiano P, Fernández de Larrea N, Martín V, Alonso MH, Castaño-Vinyals G, Pérez-Gómez B, Olmedo-Requena R, Guevara M, Fernandez-Tardon G, Dierssen-Sotos T, Llorens-Ivorra C, Huerta JM, Capelo R, Fernández-Villa T, Díez-Villanueva A, Urtiaga C, Castilla J, Jiménez-Moleón JJ, Moreno V, Dávila-Batista V, Kogevinas M, Aragonés N, Pollán M (2019) MCC-Spain researchers. Low adherence to the western and high adherence to the Mediterranean dietary patterns could prevent colorectal cancer. *Eur J Nutr* 58(4):1495–1505. <https://doi.org/10.1007/s00394-018-1674-5>
50. Feldblyum IV, Alyeva MK, Markovich NI (2016) The association between diet and the probability of colorectal cancer among the population of Perm Krai: epidemiological study. *Vopr Pitan* 85(5):60–67 PMID: 29381303
51. Vereecken CA, Covents M, Sichert-Hellert W, Alvira JM, Le Donne C, De Henauw S, De Vriendt T, Phillipp MK, Béghin L, Manios Y, Hallström L, Poortvliet E, Matthys C, Plada M, Nagy E, Moreno LA, HELENA Study Group (2008) Development and evaluation of a self-administered computerized 24-h dietary recall method for adolescents in Europe. *Int J Obes (Lond)* 32(Suppl 5):S26–34. <https://doi.org/10.1038/ijo.2008.180>
52. Anderson AS, Steele R, Coyle J (2013) Lifestyle issues for colorectal cancer survivors—perceived needs, beliefs and opportunities. *Support Care Cancer* 21(1):35–42. <https://doi.org/10.1007/s00520-012-1487-7>
53. Berstad P, Løberg M, Larsen IK, Kalager M, Holme Ø, Botteri E, Bretthauer M, Hoff G (2015) Long-term lifestyle changes after colorectal cancer screening: randomised controlled trial. *Gut* 64(8):1268–1276. <https://doi.org/10.1136/gutjnl-2014-307376>