Yuji Hamana · Harry Kesten

# A large-deviation result for the range of random walk and for the Wiener sausage

Received: 20 April 2000 / Revised version: 1 September 2000 / Published online: 26 April 2001 – © Springer-Verlag 2001

**Abstract.** Let  $\{S_n\}$  be a random walk on  $\mathbb{Z}^d$  and let  $R_n$  be the number of different points among  $\mathbf{0}, S_1, \ldots, S_{n-1}$ . We prove here that if  $d \ge 2$ , then  $\psi(x) := \lim_{n\to\infty} (-1/n) \log P\{R_n \ge nx\}$  exists for  $x \ge 0$  and establish some convexity and monotonicity properties of  $\psi(x)$ . The one-dimensional case will be treated in a separate paper.

We also prove a similar result for the Wiener sausage (with drift). Let B(t) be a *d*-dimensional Brownian motion with constant drift, and for a bounded set  $A \subset \mathbb{R}^d$  let  $\Lambda_t = \Lambda_t(A)$  be the *d*-dimensional Lebesgue measure of the 'sausage'  $\bigcup_{0 \le s \le t} (B(s) + A)$ . Then  $\phi(x) := \lim_{t \to \infty} (-1/t) \log P\{\Lambda_t \ge tx\}$  exists for  $x \ge 0$  and has similar properties as  $\psi$ .

### 1. Introduction

Let  $X, X_1, X_2, ...$  be i.i.d.  $\mathbb{Z}^d$ -valued random variables such that  $P\{X = \mathbf{0}\} < 1$ . Let  $S_0 = \mathbf{0}, S_k = \sum_{i=1}^k X_i$  and let |A| denote the cardinality of the set A. The range (at time n) of the random walk  $S = \{S_k\}$  is

$$R_n = |\{\mathbf{0}, S_1, \dots, S_{n-1}\}|$$
  
= number of different points among  $\mathbf{0}, S_1, \dots, S_{n-1}$ . (1.1)

(Note that in this definition we take the last point to be  $S_{n-1}$  rather than  $S_n$ ; this gives a somewhat more convenient subadditivity relation for the range.) It was first shown by Spitzer in [S3], pp. 38-40 that

$$\frac{R_n}{n} \to \pi := P\{S_n \neq \mathbf{0} \text{ for all } n \ge 1\} \text{ a.s.}$$
(1.2)

(cf. [S2], [De] for later references and improvements). Moreover, since

$$R_{n+m} \le R_n + |\{S_n, S_{n+1}, \dots, S_{n+m-1}\}| = R_n + |\{\mathbf{0}, S_{n+1} - S_n, S_{n+2} - S_n, \dots, S_{n+m-1} - S_n\}|, \quad (1.3)$$

Mathematics Subject Claasification (2000): Primary 60K35; Secondary 82B43

Key words or phrases: Large deviations - Range of random walk - Wiener sausage

Y. Hamana: Faculty of Mathematics, Kyushu University 36, Fukuoka 812-8581, Japan. e-mail: hamana@math.kyushu-u.ac.jp

H. Kesten: Department of Mathematics, Malott Hall, Cornell University, Ithaca, NY 14853, USA. e-mail: kesten@math.cornell.edu

one easily sees that

$$P\{R_{n+m} \le (n+m)x\} \ge P\{R_n \le nx\}P\{R_m \le mx\}$$
(1.4)

for  $x \ge 0, n, m \ge 1$ . It follows from this subadditivity relation (see [PS], problem I.98) that

$$\zeta(x) := \lim_{n \to \infty} \frac{-1}{n} \log P\{R_n \le nx\} \text{ exists.}$$
(1.5)

Of course, it follows from (1.2) that  $\zeta(x) = 0$  for  $x > \pi$ . It came as a bit of a surprise when Donsker and Varadhan (see [DV]) proved that for random walks in the domain of normal attraction of a symmetric stable law of index  $\alpha$ , and  $\lambda \ge 0$ ,

$$\lim_{n\to\infty} n^{-d/(d+\alpha)} \log E e^{-\lambda R_n}$$
 exists and is finite.

This shows that the main contributions to  $E \exp[-\lambda R_n]$  do not come from values for  $R_n$  of order n. Perhaps  $P\{R_n \le nx\}$  does not decrease exponentially in n for any x > 0. This can indeed be proven by easy lower bounds on  $P\{|S_i| \le n^\beta \text{ for } 1 \le i \le n\}$  for  $\beta \le 1/d$ . Thus, for many random walks  $\zeta(x) = 0$  for all x > 0. It turns out that a different normalization for log  $P\{R_n \le nx\}$  should be used. Indeed, van den Berg, Bolthausen, den Hollander ([BBH]) recently evaluated the limit

$$I(x) := -\lim_{t \to \infty} t^{-(d-2)/d} \log P\{\Lambda_t \le tx\},$$
(1.6)

where  $\Lambda_t$  denotes the volume of the Wiener sausage, which will be defined more precisely in (1.18) below. They gave a variational characterization of I(x) and found some peculiar dimension dependence for the associated variational problem. One can expect that a similar situation prevails for

$$\widetilde{\zeta}(x) := -\lim_{n \to \infty} n^{-(d-2)/d} \log P\{R_n \le nx\},\tag{1.7}$$

if the random walk  $\{S_k\}$  has mean zero and bounded variance, and  $d \ge 3$ .

In this paper we consider large deviations for  $R_n$  in the *upwards* direction, that is, we study

$$\psi(x) := \lim_{n \to \infty} \frac{-1}{n} \log P\{R_n \ge nx\}.$$
(1.8)

Since there is no obvious analogue of (1.4) when  $P\{R_n \le nx\}$  is replaced by  $P\{R_n \ge nx\}$ , it is not clear that the limit in (1.8) exists for  $x \ge \pi$ . Our first theorem shows that this is indeed the case for essentially all random walks. Throughout we assume that *X* has a genuinely *d*-dimensional distribution and that the corresponding random walk is aperiodic, that is, we assume that

the group generated by the support of X is all of 
$$\mathbb{Z}^d$$
 (1.9)

(see Section 2 for some discussion of this assumption).

**Theorem 1.** Let  $S_n$ ,  $R_n$  and  $\pi$  be as above and assume that (1.9) holds. Then

$$\psi(x) = \lim_{n \to \infty} \frac{-1}{n} \log P\{R_n \ge nx\} \text{ exists}$$
(1.10)

for all x (but  $\psi(x)$  may equal  $+\infty$ ).  $\psi(\cdot)$  has the following properties:

$$\psi(x) = 0 \text{ for } x \le \pi, \tag{1.11}$$

$$0 < \psi(x) < \infty \text{ for } \pi < x \le 1,$$
 (1.12)

$$\psi(x) = \infty \text{ for } x > 1, \tag{1.13}$$

$$x \mapsto \psi(x)$$
 is continuous on [0, 1], and in case  $d \ge 2$ ,  
 $x \mapsto \psi(x)$  is also convex on [0, 1], (1.14)

and

$$x \mapsto \psi(x)$$
 is strictly increasing on  $[\pi, 1]$ . (1.15)

*Remark 1.* Our proof also shows convexity of  $\psi$  if d = 1 and |X| does *not* have an exponentially bounded tail.  $\psi$  is also convex when d = 1 and  $P\{X > 0\}P\{X < 0\} = 0$ , but we were unable to prove convexity for all one-dimensional cases. The proof of Theorem 1 for d = 1 will be given in a separate paper, though.

The following is a straightforward consequence of Theorem 1. It gives a partial large deviation principle for the range of random walks. The proof of this corollary is given at the end of Section 2.

**Corollary 1.** Let  $\mu_n$  be the probability distribution of the random variable  $R_n/n$ . In the set-up of Theorem 1, we have that

$$\limsup_{n \to \infty} \frac{1}{n} \log \mu_n(F) \le -\inf_{x \in F} \psi(x) \tag{1.16}$$

*for each closed subset*  $F \subset [\pi, \infty)$  *and that* 

$$\liminf_{n \to \infty} \frac{1}{n} \log \mu_n(G) \ge -\inf_{x \in G} \psi(x) \tag{1.17}$$

*for each open subset*  $G \subset [\pi, \infty)$ *.* 

One can consider analogous problems for the cardinality of the set  $\{A, S_1 + A, ..., S_n + A\}$  when *A* is a fixed finite subset of  $\mathbb{Z}^d$  (compare [S3], Problem 27.14). We shall not do this, but instead shall go directly to continuous space and time and consider the Wiener sausage. Let  $\{B(t)\}$  be a *d*-dimensional Brownian motion with  $B(0) = \mathbf{0}$ . Contrary to the usual convention we allow the Brownian motion to have a constant drift  $\mu \neq \mathbf{0}$ . Further, let  $A \subset \mathbb{R}^d$  be a bounded, Lebesgue measurable, nonempty set. The set  $\bigcup_{0 \le s \le t} (B(s) + A)$  is called the *Wiener sausage* (associated with *A*). Its volume is

$$\Lambda_t = \Lambda_t(A) := \lambda_d \left( \bigcup_{0 \le s \le t} (B(s) + A) \right), \tag{1.18}$$

where  $\lambda_d(\cdot)$  denotes the *d*-dimensional Lebesgue measure. The analogue of (1.2) is now that

$$\theta := \lim_{t \to \infty} \frac{\Lambda_t(A)}{t} \text{ exists a.s.}$$
(1.19)

(cf. [S2], [IM, Problem 7.8.4]). When  $\mu = 0$ 

$$\lim_{t \to \infty} \frac{\Lambda_t(A)}{t} = \begin{cases} \operatorname{cap}(A) & \text{if } d \ge 3\\ 0 & \text{if } d = 1 \text{ or } 2 \end{cases}$$
(1.20)

a.s., where

$$\operatorname{cap}(A) = \begin{cases} d \text{-dimensional Newtonian capacity of } A & \text{if } d \ge 3\\ \text{logarithmic capacity of } A & \text{if } d = 2. \end{cases}$$
(1.21)

It can also be shown that  $\theta$  in (1.19) is strictly positive and finite for either  $d \ge 3$ , cap (A) > 0, or d = 2, cap (A) > 0,  $\mu \neq 0$ , or d = 1,  $A \neq \emptyset$ ,  $\mu \neq 0$ . (We are grateful to an anonymous referee for this last remark.)

 $\Lambda_t$  in the result of [BBH] ((1.6) above) stands for  $\Lambda_t(A)$  with A a ball of radius a, and for this choice of A, (1.6) supplements (1.20). We have, however, no nice expression for  $\theta$  for general drift. Some related results for the case  $\mu \neq 0$  are in [EL].

Here we again consider deviations from (1.19) in the upwards direction. Some control of the probabilities  $P{\Lambda_t \ge tx}$  is provided by [BB] and [BT]. These papers show the existence of

$$\rho(\lambda) := \lim_{t \to \infty} \frac{1}{t} \log E e^{\lambda \Lambda_t} \quad \text{for} \quad \lambda \ge 0, \tag{1.22}$$

again under the assumptions that A is a ball and  $\mu = 0$ . They further derive some estimates for  $\rho(\lambda)$  as  $\lambda \downarrow 0$  or  $\lambda \uparrow \infty$ . We shall indicate here a proof of the following analogue of Theorem 1:

**Theorem 2.** If A is a bounded, Lebesgue measurable, nonempty set in  $\mathbb{R}^d$ , then for any constant (possibly zero) drift  $\mu$ 

$$\phi(x) := \lim_{t \to \infty} \frac{-1}{t} \log P\{\Lambda_t(A) \ge tx\} \text{ exists}$$
(1.23)

for all  $x \in \mathbb{R}$ .  $\phi(\cdot)$  satisfies

$$\phi(x) = 0 \quad for \quad x \le \theta, \tag{1.24}$$

where  $\theta$  is defined by (1.19). If d = 1, then

$$\phi(x) = \begin{cases} \frac{x^2}{2} & \text{if } \mu = 0, \ x \ge 0\\ 0 & \text{if } 0 \le x \le |\mu| \\ \frac{(x - |\mu|)^2}{2} & \text{if } 0 < |\mu| \le x. \end{cases}$$
(1.25)

If  $d \ge 2$  and cap(A) > 0, then  $\phi$  has the following properties:

$$0 < \phi(x) < \infty \text{ for } x > \theta, \tag{1.26}$$

$$x \mapsto \phi(x)$$
 is convex and continuous on  $[0, \infty)$ , (1.27)

and

$$x \mapsto \phi(x)$$
 is strictly increasing on  $[\theta, \infty)$ . (1.28)

*Remark 2.* The case  $d \ge 2$  with cap(A)=0 is not interesting, because in that case

$$P{B(t) \in A \text{ for some } t > 0} = 0.$$

For a Brownian motion without drift this was proven by Kakutani in [Ka] for d = 2, and the higher dimensional case seems to have been treated first in [Do1]; see also Section 2.IX.5 in [Do2]. This statement remains valid even if the drift is nonzero, because on each finite time interval Brownian motion with drift and without drift are absolutely continuous with respect to each other (see [KS, Section 3.5, especially 3.5C]). But then

$$E\Lambda_t(A) = \int_{\mathbb{R}^d} P\left\{ y \in \bigcup_{s \le t} (B(s) + A) \right\} dy$$
  
=  $\int_{\mathbb{R}^d} P\{B(s) \in y - A \text{ for some } s \le t\} dy = 0.$ 

Thus if cap(A) = 0, then  $P\{\Lambda_t(A) = 0\} = 1$  for each fixed *t*, and then by the monotonicity of  $t \mapsto \Lambda_t$ ,  $P\{\Lambda_t(A) = 0 \text{ for all } t\} = 1$ .

The following corollary is the analogue of Corollary 1 for the Wiener sausage. Its proof is entirely analogous to that of Corollary 1 and will be left to the reader.

**Corollary 2.** Let A be a bounded, Lebesgue measureable, nonempty set in  $\mathbb{R}$  and let  $v_t$  be the distribution of  $\Lambda_t(A)/t$ . If  $d \ge 2$  assume that cap(A) > 0. The constant  $\theta$  is the same as in (1.19). Then we have that

$$\limsup_{t \to \infty} \frac{1}{t} \log v_t(F) \le -\inf_{x \in F} \phi(x)$$

for each closed set  $F \subset [\theta, \infty)$  and that

$$\liminf_{t\to\infty}\frac{1}{t}\log\nu_t(G)\geq-\inf_{x\in G}\phi(x)$$

for each open set  $G \subset [\theta, \infty)$ .

The principal **open question** now is to find a manageable expression or characterization for  $\psi(x)$  and  $\phi(x)$ .

# *1.1. Outline of the proof of* (1.10) *for* $d \ge 2$

As we already stated in Remark 1, we only prove Theorem 1 for  $d \ge 2$  in this paper. The proof is based on an approximate subaddivity relation. We build a path of length n + m with  $R_{n+m} \ge y + z - E(n, m)$  for some error term E(n, m) from two paths,  $\mathscr{P}_1$  and  $\mathscr{P}_2$ , say.  $\mathscr{P}_1(\mathscr{P}_2)$  has length n(m) and range greater than or equal to y(z, respectively). The error term comes from the fact that some points are counted in the range of both  $\mathscr{P}_1$  and  $\mathscr{P}_2$ . In order to make this overlap small we do not put the initial point of  $\mathscr{P}_2$  at the endpoint of  $\mathscr{P}_1$ , but at some nearby point. We show that we can place the initial point of  $\mathscr{P}_2$  at a distance at most of order  $(nm)^{1/(d+1)}$  from the endpoint of  $\mathscr{P}_1$  so as to get an overlap of order  $(nm)^{1/(d+1)}$ . The two paths are then connected at not too large a cost in probability. The result is the relation

$$P\{R_{n+m} \ge y + z - (2d+2)(nm)^{1/(d+1)}\}$$
  
$$\ge \frac{1}{2}\zeta^{d(nm)^{1/(d+1)} + d}P\{R_n \ge y\}P\{R_m \ge z\}.$$
 (1.29)

for some  $\zeta > 0$ . (The idea of putting the initial point of  $\mathscr{P}_2$  at a point which is near, but not necessarily at, the endpoint of  $\mathscr{P}_1$ , was also used in [Ke] for estimating numbers of self-avoiding walks. It is of course not surprising that tools for selfavoiding walks are useful here, since the event  $\{R_n \ge n\}$  is just the event that the initial piece  $S_0, \ldots, S_{n-1}$  of the random walk is selfavoiding.)

When  $d \ge 2$ , then  $(nm)^{1/(d+1)}$  is small with respect to  $(n \lor m)$ . From this one deduces by more or less standard subadditivity arguments that  $\lim_{n\to\infty}(-1/n)\log P$  $\{R_n \ge nx\}$  exists at all continuity points  $x \in (0, 1)$  of  $\psi(x) := \liminf_{n\to\infty}(-1/n)\log P$  $\{R_n \ge nx\}$ . It is then easy to obtain from (1.29) that the restriction of  $\lim_{n\to\infty}(-1/n)\log P\{R_n \ge nx\}$  to the continuity points of  $\psi$  in (0, 1) is convex (see proof of Lemma 3 for the precise meaning of this statement). This is enough to conclude that  $\psi$  is in fact continuous on (0, 1), just as in the usual proof of continuity of a convex function. Hence  $\lim_{n\to\infty}(-1/n)\log P\{R_n \ge nx\}$  exists for all  $x \in (0, 1)$ . The existence of this limit for  $x \notin (0, 1)$  is easily shown directly.

We remark that it is also possible to prove Theorem 1 by the methods of Lemmas 3 and 6 of [HK], which are used there for the one-dimensional case. For the higher dimensional case the proof of this paper is more direct.

# 2. A subadditivity argument; proof of Theorem 1 for $d \ge 2$

Throughout this section  $X, X_1, X_2, ...$  are i.i.d.  $\mathbb{Z}^d$ -valued random variables and  $S_n, R_n$  are as in Section 1.

Before we begin any proofs we point out that assumption (1.9) does not entail any loss of generality. Indeed we can always change coordinates so that (1.9) holds, by Proposition 7.1 in [S3]. More specifically, if  $\mathscr{G}$  denotes the subgroup of  $\mathbb{Z}^d$  generated by the support of X, then there exists a  $d' \leq d$  and d' linearly independent vectors  $v_1, v_2, \ldots, v_{d'} \in \mathscr{G}$  such that the group generated by  $v_1, v_2, \ldots, v_{d'}$  is equal to  $\mathscr{G}$ , and is isomorphic to  $\mathbb{Z}^{d'}$ . In particular, there exist unique random variables

$$X' = (X'(1), \dots, X'(d')), X'_j = (X'_j(1), \dots, X'_j(d')) \in \mathbb{Z}^{d'}$$
 such that

$$X = \sum_{\ell=1}^{d} X'(\ell) v_{\ell}, \quad X_{j} = \sum_{\ell=1}^{d'} X'_{j}(\ell) v_{\ell},$$

and the group generated by the support of X' equals  $\mathbb{Z}^{d'}$  Accordingly, we can view  $\{S_n\}$  as an aperiodic random walk on  $\mathbb{Z}^{d'}$ . From now on we therefore drop the primes from our notation and assume that the problem has been set up from the beginning so that (1.9) holds.

We begin with a fundamental subadditivity relation.

**Lemma 1.** If (1.9) holds, then there exists a constant  $\zeta \in (0, 1)$  such that for all integers  $n, m \ge 0$  and  $y, z \in [0, \infty)$ , it holds that

$$P\{R_{n+m} \ge y + z - (2d+2)(nm)^{1/(d+1)}\}$$
  
$$\ge \frac{1}{2} \zeta^{d(nm)^{1/(d+1)} + d} P\{R_n \ge y\} P\{R_m \ge z\}.$$
(2.1)

*Proof.* We introduce a number of quantities. The relevance of these quantities will become clear in a little while. We let  $\widehat{X}_1, \widehat{X}_2, \ldots$  be an independent copy of  $X_1, X_2, \ldots$  In analogy with our previous notation we define  $\widehat{S} = \{\widehat{S}_n\}_{n\geq 0}$  by  $\widehat{S}_0 = \mathbf{0}, \widehat{S}_n = \sum_{i=1}^n \widehat{X}_i$ . We define

$$R[a, b] = |\{S_a, \ldots, S_b\}|$$
 and  $\widehat{R}[a, b] = |\{\widehat{S}_a, \ldots, \widehat{S}_b\}|.$ 

Note that  $R_n = R[0, n-1]$  in this notation. Next we define for  $w \in \mathbb{Z}^d$ 

$$N_{n,m}(w) = N_{n,m}(w, S, \widehat{S}) = |\{u \in \mathbb{Z}^d : u \in \{S_0, S_1, \dots, S_{n-1}\} and u \in S_n + w + \{\widehat{S}_0, \widehat{S}_1, \dots, \widehat{S}_{m-1}\}\}|.$$

Thus  $N_{n,m}(w)$  counts the number of points which are visited during the time interval [0, n-1] by the walk  $\hat{S}$  and also visited during [0, m-1] by the walk  $\hat{S}$  shifted by  $S_n + w$ .

For any fixed integers  $p \ge 0$ ,  $n \ge 0$ , consider the random walk defined by

$$T_k = T_k^{(p,n)} = \begin{cases} S_k & \text{for } k \le n+p \\ S_{n+p} + \widehat{S}_{k-n-p} & \text{for } k > n+p. \end{cases}$$

Of course,  $\{T_k\}_{k\geq 0}$  has the same distribution as  $\{S_k\}_{k\geq 0}$ , and hence also

$$P\{R_{n+p+m} \ge \ell\} = P\{R[0, n+p+m-1] \ge \ell\}$$
  
=  $P\{|\{T_0, \dots, T_{n+p+m-1}\}| \ge \ell\}.$  (2.2)

We claim that on the event

$$\{S_{n+p} - S_n = w\},\tag{2.3}$$

it holds that

$$|\{T_0, \dots, T_{n+p+m-1}\}| \ge R_n + \widehat{R}_m - N_{n,m}(w).$$
(2.4)

This is immediate from the fact that  $R_n = |\{T_0, \ldots, T_{n-1}\}|$ , and  $\widehat{R}_m = |\{T_{n+p}, \ldots, T_{n+p+m-1}\}|$ , and on the event (2.3),

$$N_{n,m}(w) = N_{n,m}(T_{n+p} - T_n)$$
  
=  $|\{T_0, \dots, T_{n-1}\} \cap \{T_{n+p}, \dots, T_{n+p+m-1}\}|.$ 

At this stage we remind the reader of (1.9). This allows us to pick *d* linearly independent vectors  $v_1, \ldots, v_d \in \mathbb{Z}^d$  for which  $P\{X = v_i\} > 0$ . We can then choose  $0 < \zeta < 1$  such that  $P\{X = v_i\} \ge \zeta$  for  $i = 1, \ldots, d$ . We set

$$\Xi_q = \left\{ \sum_{i=1}^d k_i v_i : 0 \le k_i \le q \right\} \subset \mathbb{Z}^d.$$

For any  $w = \sum_{i=1}^{d} k_i v_i \in \Xi_q$ , we then have for  $p = p(w) = \sum_{i=1}^{d} k_i \le dq$  that

$$P\{S_{n+p} - S_n = w\} = P\{S_p = w\} \ge \zeta^p \ge \zeta^{dq}.$$
(2.5)

Moreover,

$$|\Xi_q| = (\text{number of vectors } w \in \Xi_q) = (q+1)^d$$
.

We take

$$q = q(n, m) = \lceil (nm)^{1/(d+1)} \rceil,$$
 (2.6)

where  $\lceil a \rceil$  denotes the smallest integer  $\geq a$ . As a result of (2.2) and (2.4) we have for each  $w \in \Xi_q$ ,

$$P\{R_{n+dq+m} \ge y + z - 2(nm)^{1/(d+1)}\}$$
  

$$\ge P\{R_{n+p(w)+m} \ge y + z - 2(nm)^{1/(d+1)}\}$$
  

$$\ge P\{R_n \ge y, \widehat{R}_m \ge z, S_{n+p(w)} - S_n = w, N_{n,m}(w) \le 2(nm)^{1/(d+1)}\}. (2.7)$$

The event (2.3) depends only on the  $X_i$  with  $n < i \le n + p$ , and is therefore independent of the events  $\{R_n \ge y\}$ ,  $\{\widehat{R}_m \ge z\}$  and of the random variable  $N_{n,m}(w)$ (for fixed w). Consequently,

$$P\{R_{n+dq+m} \ge y + z - 2(nm)^{1/(d+1)}\}$$
  

$$\ge P\{S_{n+p} - S_n = w\}P\{R_n \ge y, \ \widehat{R}_m \ge z, N_{n,m}(w) \le 2(nm)^{1/(d+1)}\}$$
  

$$\ge \zeta^{dq}P\{R_n \ge y, \ \widehat{R}_m \ge z, \ N_{n,m}(w) \le 2(nm)^{1/(d+1)}\}.$$
(2.8)

Since this inequality holds for all  $w \in \Xi_q$ , we can take its average over  $\Xi_q$  to obtain

$$P\{R_{n+dq+m} \ge y + z - 2(nm)^{1/(d+1)}\}$$
  

$$\ge \frac{\zeta^{dq}}{|\Xi_q|} \sum_{w \in \Xi_q} P\{N_{n,m}(w) \le 2(nm)^{1/(d+1)}, R_n \ge y, \widehat{R}_m \ge z\}$$
  

$$\ge \frac{\zeta^{dq}}{|\Xi_q|} E\Big\{ \Big| \{w \in \Xi_q : N_{n,m}(w) \le 2(nm)^{1/(d+1)}\} \Big| I[R_n \ge y] I[\widehat{R}_m \ge z] \Big\}.$$
(2.9)

We shall soon show that always

$$\left| \{ w \in \Xi_q : N_{n,m}(w) \le 2(nm)^{1/(d+1)} \} \right| \ge \frac{1}{2}(q+1)^d.$$
 (2.10)

Before we do this we show that this will complete the the proof of the lemma. Indeed,  $\{R_n \ge y\}$  and  $\{\widehat{R}_m \ge z\}$  are independent because they depend only on the  $X_i$  and  $\widehat{X}_j$ , respectively. Their respective probabilities equal  $P\{R_n \ge y\}$  and  $P\{R_m \ge z\}$ . Thus, if (2.10) is true, then (2.9) yields

$$P\{R_{n+dq+m} \ge y + z - 2(nm)^{1/(d+1)}\}$$
  
$$\ge \frac{\zeta^{dq}}{(q+1)^d} \frac{1}{2} (q+1)^d P\{R_n \ge y\} P\{R_m \ge z\}.$$
(2.11)

Finally, by removing the last dq steps from  $S_0 = 0, S_1, ..., S_{n+dq+m}$ , we see that  $\{R_{n+dq+m} \ge \ell\}$  implies  $\{R_{n+m} \ge \ell - dq\}$ . Thus, (2.11) shows that

$$P\{R_{n+m} \ge y + z - 2(nm)^{1/(d+1)} - d\lceil (nm)^{1/(d+1)}\rceil\}$$
  
$$\ge \frac{1}{2}\zeta^{dq}P\{R_n \ge y\}P\{R_m \ge z\},$$
(2.12)

which will indeed prove the lemma.

We conclude with the promised proof of (2.10). We have

$$\begin{split} \sum_{w \in \Xi_q} N_{n,m}(w) &\leq \sum_{w \in \mathbb{Z}^d} N_{n,m}(w) \\ &= \sum_{u \in \mathbb{Z}^d} \sum_{w \in \mathbb{Z}^d} I[u \in \{S_0, S_1, \dots, S_{n-1}\}] \\ &\times I[u \in S_n + w + \{\widehat{S}_0, \widehat{S}_1, \dots, \widehat{S}_{m-1}\}] \\ &= \sum_{u \in \mathbb{Z}^d} I[u \in \{S_0, S_1, \dots, S_{n-1}\}] \\ &\times |\{u - S_n - \widehat{S}_0, u - S_n - \widehat{S}_1, \dots, u - S_n - \widehat{S}_{m-1}\}| \\ &= \sum_{u \in \mathbb{Z}^d} I[u \in \{S_0, S_1, \dots, S_{n-1}\}] \widehat{R}_m \\ &= R_n \widehat{R}_m \leq nm. \end{split}$$

It follows that

$$\begin{aligned} \left| \{ w \in \Xi_q : N_{n,m}(w) > 2(nm)^{1/(d+1)} \} \right| \\ &\leq \frac{nm}{2(nm)^{1/(d+1)}} = \frac{1}{2}(nm)^{d/(d+1)} \leq \frac{1}{2} |\Xi_q|, \end{aligned}$$

and hence

$$|\{w \in \Xi_q : N_{n,m}(w) \le 2(nm)^{1/(d+1)}\}| \ge \frac{1}{2}|\Xi_q|,$$

as desired in (2.10).

*Remark 3.* As stated, Lemma 1 is not useful when d = 1; the error factor  $\zeta^{(nm)^{1/(d+1)}}$  becomes too dominant in this case. However, when d = 1 and  $\limsup_{k\to\infty} [P\{|X| \ge k\}]^{1/k} = 1$  we can still derive an estimate similar to (2.1), by taking a different set  $\Xi_q$  in the proof of Lemma 1. This replacement for (2.1) will not be enough by itself to deduce (1.10), but it will give us an upper bound on  $\limsup_{n\to\infty}(-1/n)\log P\{R_n \ge nx\}$ .

For  $x \in \mathbb{R}$  we now define

$$\psi(x) = \liminf_{n \to \infty} \frac{-1}{n} \log P\{R_n \ge nx\}.$$
(2.13)

Observe that  $\psi(x)$  is nondecreasing in x. Moreover, it is bounded on [0, 1] because

$$P\{R_n \ge n\} = P\{S_0, \dots, S_{n-1} \text{ are distinct}\} \ge [P\{X(i) > 0\}]^n, \qquad (2.14)$$

where X(i) denotes the *i*-th component of *X*. Hence, for all  $1 \le i \le d$ ,

$$\psi(1) \le -\log P\{X(i) > 0\}. \tag{2.15}$$

If  $P{X(i) > 0} = 0$  for all  $1 \le i \le d$ , then we can replace X(i) > 0 by X(i) < 0 in this estimate. This always gives a finite upper bound for  $\psi(x)$ .

We have to prove for (1.10) that the liminf in (2.13) can be replaced by lim. We first show that this is permissible for any  $x \in [0, 1)$  at which  $\psi$  is continuous from the right.

**Proposition 2.** If (1.9) holds and  $d \ge 2$  and if  $\psi$  is right continuous at a given  $x \in [0, 1)$ , then

$$\psi(x) = \lim_{n \to \infty} \frac{-1}{n} \log P\{R_n \ge nx\}.$$
(2.16)

*Proof.* The proof here will be based only on (2.11) and the fact that  $P\{R_n \ge z\}$  is increasing in *n* and decreasing in *z*. It would be somewhat simpler to use (2.1), but a proof which only uses (2.11) has the advantage that it can also be used in the next section for the Wiener sausage.

For simplicity we write  $\eta = (d-1)/(d+1)$  and  $\xi = 2/(d+1)$ . To start with, we define for any integer  $N \ge 1$ ,

$$\sigma(0) = \sigma_N(0) = N,$$
  

$$\sigma(k+1) = \sigma_N(k+1) = 2\sigma(k) + d\left[\left[\sigma(k)\right]^{\xi}\right], \ k \ge 0.$$

It is immediate from these definitions that

$$\frac{\sigma(i-1)}{\sigma(i)} \le \frac{1}{2}, \quad \sigma(i) \ge 2^i N, \tag{2.17}$$

and for some constants  $c_1$ ,  $N_0 < \infty$  and  $N \ge N_0$ 

$$1 \leq \frac{\sigma(k)}{2^k N} = \prod_{i=1}^k \frac{\sigma(i)}{2\sigma(i-1)}$$
$$= \prod_{i=1}^k \left[ 1 + \frac{d\left[ [\sigma(i-1)]^{\xi} \right]}{2\sigma(i-1)} \right] \leq \exp\left[ \sum_{i=1}^\infty d[\sigma(i-1)]^{-\eta} \right]$$
$$\leq \exp\left[ \sum_{j=0}^\infty d(2^j N)^{-\eta} \right] \leq 1 + \frac{c_1}{N^{\eta}} \leq 2.$$
(2.18)

Also, for  $c_2 = 2^{\xi}/(1 - 2^{-\eta})$  and  $N \ge N_0$ ,

$$\sum_{i=0}^{k-1} 2^{k-i} [\sigma(i)]^{\xi} \leq \sum_{i=0}^{k-1} 2^{k-i} 2^{\xi(i+1)} N^{\xi}$$
$$\leq c_2 2^k N^{\xi} \leq c_2 N^{-\eta} \sigma(k).$$
(2.19)

Now let  $x \in [0, 1)$  be such that  $\psi$  is right continuous at x and let  $\varepsilon > 0$ . Take  $\delta \in (0, 1)$  such that

$$\psi(x+4\delta) \le \psi(x) + \varepsilon. \tag{2.20}$$

Then take  $c_3 = \zeta^d/2 < 1$  and fix  $\ell \ge 2$  such that

$$(1 - 2^{-\ell+2})(x + 2\delta) \ge x + \delta.$$
 (2.21)

Finally, fix  $N \ge N_0$  so that

$$P\{R_N \ge N(x+4\delta)\} \ge \exp[-N(\psi(x+4\delta)+\varepsilon)]$$
  
$$\ge \exp[-N(\psi(x)+2\varepsilon)], \qquad (2.22)$$

$$1 + \frac{c_1}{N^{\eta}} \le \frac{x + 4\delta}{x + 3\delta},\tag{2.23}$$

$$N^{-\eta} < \min\left\{\frac{\delta}{c_2}, \frac{-2\varepsilon}{c_2 d \log \zeta}, \frac{1}{2d}\right\},\tag{2.24}$$

$$2\ell(3d+2)(N^{\xi}+1) < \delta N, \tag{2.25}$$

and

$$\frac{2}{N}|\log c_3| < \varepsilon. \tag{2.26}$$

We shall first consider  $P\{R_n \ge nx\}$  for  $n \in \{\sigma_N(k)\}_{k \ge 0}$ . (2.11) with  $m = n = \sigma(k-1)$  and

$$y = z = 2^{k-1}N(x+4\delta) - \sum_{i=0}^{k-2} 2^{k-1-i} [\sigma(i)]^{\xi}$$

now gives for  $k \ge 1$ 

$$P\left\{R_{\sigma(k)} \ge 2^{k}N(x+4\delta) - \sum_{i=0}^{k-1} 2^{k-i}[\sigma(i)]^{\xi}\right\}$$
$$\ge c_{3}\zeta^{d[\sigma(k-1)]^{\xi}} \left[P\left\{R_{\sigma(k-1)} \ge 2^{k-1}N(x+4\delta) - \sum_{i=0}^{k-2} 2^{k-1-i}[\sigma(i)]^{\xi}\right\}\right]^{2}. (2.27)$$

We also have

$$2^{k}N(x+4\delta) - \sum_{i=0}^{k-1} 2^{k-i} [\sigma(i)]^{\xi}$$
  

$$\geq \sigma(k)(x+3\delta) - \sum_{i=0}^{k-1} 2^{k-i} [\sigma(i)]^{\xi} \quad (by (2.18) \text{ and } (2.23))$$
  

$$\geq \sigma(k)(x+2\delta) \quad (by (2.19) \text{ and } (2.24)).$$

Combining this with (2.27) we obtain

$$P\{R_{\sigma(k)} \ge \sigma(k)(x+2\delta)\}$$

$$\ge P\left\{R_{\sigma(k)} \ge 2^{k}N(x+4\delta) - \sum_{i=0}^{k-1} 2^{k-i}[\sigma(i)]^{\xi}\right\}$$

$$\ge c_{3}\zeta^{d[\sigma(k-1)]^{\xi}} \left[P\left\{R_{\sigma(k-1)} \ge 2^{k-1}N(x+4\delta) - \sum_{i=0}^{k-2} 2^{k-1-i}[\sigma(i)]^{\xi}\right\}\right]^{2}$$

$$\ge \cdots \ge$$

$$\ge [c_{3}]^{2^{k+1}} \exp\left[d\sum_{j=1}^{k} 2^{j-1}[\sigma(k-j)]^{\xi}\log\zeta\right] \left[P\{R_{\sigma(0)} \ge N(x+4\delta)\}\right]^{2^{k}}$$

$$\ge [c_{3}]^{2^{k+1}} \exp\left[c_{2}d2^{k-1}N^{\xi}\log\zeta\right] \exp\left[-2^{k}N(\psi(x)+2\varepsilon)\right]$$
(by (2.19) and (2.22))

$$= [c_3]^{2^{k+1}} \exp\left[-2^k N\left(\psi(x) + 2\varepsilon - \frac{1}{2}c_2 dN^{-\eta}\log\zeta\right)\right]$$
  

$$\geq [c_3]^{2^{k+1}} \exp\left[-2^k N(\psi(x) + 3\varepsilon)\right].$$
(2.28)

The last inequality here was obtained from (2.24). With the help of (2.26), the estimate (2.28) implies in particular that

$$\limsup_{k \to \infty} \frac{-1}{\sigma(k)} \log P\{R_{\sigma(k)} \ge \sigma(k)(x+2\delta)\}$$
$$\le \psi(x) + 3\varepsilon - \frac{2}{N} \log c_3 \le \psi(x) + 4\varepsilon,$$

but we shall need the explicit bound of (2.28) for the next step.

In order to deal with general *n*, we "expand *n* (approximately) into a linear combination of the  $\sigma_N(k)$ ." More precisely, recall that we fixed  $\ell$  in (2.21). Now let  $n \ge \sigma_N(2\ell)$ . Take

$$\widehat{n} = n - 2d\ell \lceil n^{\xi} \rceil,$$

and choose  $k_i$ ,  $\alpha_i$  recursively in the following manner ( $\sigma = \sigma_N$  again here):  $k_1$  is determined by

$$\sigma(k_1) \le \widehat{n} < \sigma(k_1 + 1),$$

and

$$\alpha_1 = \begin{cases} 1 & \text{if } \sigma(k_1) \le \widehat{n} < 2\sigma(k_1) \\ 2 & \text{if } 2\sigma(k_1) \le \widehat{n} < \sigma(k_1+1); \end{cases}$$

then  $k_r$  and  $\alpha_r$  for  $r \ge 2$  are determined by

$$\sigma(k_r) \le \widehat{n} - \sum_{i=1}^{r-1} \alpha_i \sigma(k_i) < \sigma(k_r + 1)$$
(2.29)

and

$$\alpha_r = \begin{cases} 1 & \text{if } \sigma(k_r) \le \widehat{n} - \sum_{i=1}^{r-1} \alpha_i \sigma(k_i) < 2\sigma(k_r) \\ 2 & \text{if } 2\sigma(k_r) \le \widehat{n} - \sum_{i=1}^{r-1} \alpha_i \sigma(k_i) < \sigma(k_r+1). \end{cases}$$

We only choose these  $k_i, \alpha_i$  as long as  $k_i \ge 0$ . Note that  $n \ge \sigma(2\ell)$  implies  $k_1 \ge 2\ell - 1$ . Also, by virtue of (2.24),

$$\frac{\sigma(k+1)}{\sigma(k)} \le 2 + 2d[\sigma(k)]^{-\eta} \le 2 + 2dN^{-\eta} < 3,$$

so that we obtain from our choice of  $k_r$ ,  $\alpha_r$  that

$$0 \leq \widehat{n} - \sum_{i=1}^{r} \alpha_i \sigma(k_i) < \sigma(k_r).$$

As a consequence  $k_1 > k_2 > \cdots$ . Let *p* be such that  $k_p > k_1 - \ell \ge k_{p+1}$ . By (2.29), the monotonicity of  $\sigma(i)$  in *i*, and (2.18) we then have that

$$0 \le \widehat{n} - \sum_{i=1}^{p} \alpha_i \sigma(k_i) < \sigma(k_{p+1}+1) \le \sigma(k_1 - \ell + 1)$$
  
$$< 2^{k_1 - \ell + 2} N \le 2^{-\ell + 2} \sigma(k_1) \le 2^{-\ell + 2} n.$$
(2.30)

Of course we also have  $p \leq \ell$ .

We set  $\beta := \sum_{i=1}^{p} \alpha_i$  and let  $n_1 < n_2 < \cdots < n_{\beta}$  be the numbers of the form  $\sum_{i=1}^{j} \alpha_i \sigma(k_i)$  or  $\sum_{i=1}^{j} \alpha_i \sigma(k_i) - \sigma(k_j)$ ; the latter form is included only if  $\alpha_j = 2$ . If  $n_{\gamma}$  has the first form for some  $1 \le j \le p-1$ , then  $n_{\gamma+1} = n_{\gamma} + \sigma(k_{j+1})$ . If  $n_{\gamma}$  has the second form for some  $1 \le j \le p-1$ , then  $n_{\gamma+1} = n_{\gamma} + \sigma(k_j)$ . We now apply (2.11), with *n* replaced by  $n_{\gamma} + d\gamma \lceil n^{\xi} \rceil$  and  $m = n_{\gamma+1} - n_{\gamma}$ . We further take  $y = n_{\gamma}(x + 2\delta) - 2\gamma n^{\xi}$  and  $z = (n_{\gamma+1} - n_{\gamma})(x + 2\delta) = m(x + 2\delta)$ . Note that

$$\beta \le 2p \le 2\ell, \tag{2.31}$$

and for  $\gamma \leq \beta - 1$ ,

$$(n_{\gamma} + d\gamma \lceil n^{\xi} \rceil)(n_{\gamma+1} - n_{\gamma}) \le (\widehat{n} + d\beta \lceil n^{\xi} \rceil)\widehat{n} \le n^{2}.$$

Using the monotonicity properties of  $P\{R_r \ge z\}$  we then find

$$\begin{split} &P\{R_{n_{\gamma+1}+d(\gamma+1)\lceil n^{\xi}\rceil} \geq n_{\gamma+1}(x+2\delta) - 2(\gamma+1)n^{\xi}\}\\ &\geq \frac{1}{2}\zeta^{dn^{\xi}+d}P\{R_{n_{\gamma}+d\gamma\lceil n^{\xi}\rceil} \geq n_{\gamma}(x+2\delta) - 2\gamma n^{\xi}\}\\ &\times P\{R_{n_{\gamma+1}-n_{\gamma}} \geq (n_{\gamma+1}-n_{\gamma})(x+2\delta)\}. \end{split}$$

We use this inequality successively for  $\gamma = \beta - 1, \beta - 2, ..., 1$  (with  $n_0 = 0$ ) to obtain finally

$$P\{R_{n_{\beta}+d\beta\lceil n^{\xi}\rceil} \ge n_{\beta}(x+2\delta) - 2\beta n^{\xi}\}$$

$$\ge [c_{3}]^{\beta} \zeta^{\beta dn^{\xi}} \prod_{\gamma=0}^{\beta-1} P\{R_{n_{\gamma+1}-n_{\gamma}} \ge (n_{\gamma+1}-n_{\gamma})(x+2\delta)\}$$

$$\ge [c_{3}]^{2\ell} \zeta^{2\ell dn^{\xi}} \prod_{j=1}^{p} \left[ P\{R_{\sigma(k_{j})} \ge \sigma(k_{j})(x+2\delta)\} \right]^{\alpha_{j}}$$
(2.32)
(by (2.31) and the definition of  $n_{\gamma}$ )

(by (2.31) and the definition of  $n_{\gamma}$ )

$$\geq [c_3]^{2\ell + \sum_{j=1}^p \alpha_j 2^{k_j + 1}} \zeta^{2\ell dn^{\xi}} \exp\left[-\sum_{j=1}^p \alpha_j 2^{k_j} N(\psi(x) + 3\varepsilon)\right]$$

(by (2.28)).

Now observe that by (2.30) and (2.31)

$$n_{\beta}(x+2\delta) - 2\beta n^{\xi} = \sum_{i=1}^{p} \alpha_{i}\sigma(k_{i})(x+2\delta) - 2\beta n^{\xi}$$
  

$$\geq (\widehat{n} - 2^{-\ell+2}n)(x+2\delta) - 2\beta n^{\xi}$$
  

$$\geq (1 - 2^{-\ell+2})n(x+2\delta) - 2d\ell(x+2\delta)\lceil n^{\xi}\rceil - 4\ell n^{\xi}$$
  

$$\geq n(x+\delta) - 2\ell(3d+2)\lceil n^{\xi}\rceil \text{ (by (2.21)).}$$

Then, by (2.25), we have that  $n^{\beta}(x+2\delta) - 2\beta n^{\xi} \ge nx$  for all  $n \ge N$ . Note also that

$$n_{\beta} + d\beta \lceil n^{\xi} \rceil \leq \widehat{n} + d\beta \lceil n^{\xi} \rceil \leq n,$$

and that

$$\sum_{j=1}^{p} \alpha_j 2^{k_j} N \leq \sum_{j=1}^{p} \alpha_j \sigma(k_j) \leq \widehat{n} \leq n.$$

Thus,  $P\{R_n \ge nx\}$  is at least as large as the left hand side of (2.32), and we have

$$-\frac{1}{n}\log P\{R_n \ge nx\}$$
  
$$\leq -\left(\frac{2\ell}{n} + \frac{2}{N}\right)\log c_3 - \frac{2\ell d}{n^{\eta}}\log \zeta + \psi(x) + 3\varepsilon$$
  
$$\leq -\frac{2\ell}{n}\log c_3 - \frac{2\ell d}{n^{\eta}}\log \zeta + \psi(x) + 4\varepsilon.$$

The proposition follows by taking the limits  $n \to \infty$ , and then  $\varepsilon \downarrow 0$ .

**Lemma 3.** If (1.9) holds and  $d \ge 2$ , then  $\psi$  is convex and continuous on (0, 1).

*Proof.* We first show that  $\psi$  is convex at its continuity points in [0, 1), that is, for  $x, y, z \in [0, 1)$ , all of which are continuity points of  $\psi$  and such that  $x = \alpha y + (1 - \alpha)z$  for some  $0 < \alpha < 1$  it holds that

$$\psi(x) = \psi(\alpha y + (1 - \alpha)z) \le \alpha \psi(y) + (1 - \alpha)\psi(z).$$
(2.33)

To see this we use Lemma 1 and the obvious monotonicity in w of  $P\{R_n \ge w\}$ . Recall that  $\xi = 2/(d+1)$  and take  $m = m_n = n - d\lceil n^{\xi} \rceil$  for large n. Since  $\lfloor \alpha m \rfloor \cdot \lfloor (1-\alpha)m \rfloor \le n^2$ , and  $\lfloor \alpha m \rfloor (y+\varepsilon) + \lfloor (1-\alpha)m \rfloor z - 2\lceil n^{\xi} \rceil \ge nx$  for large n, we obtain from (2.11) for  $\varepsilon \in (0, 1)$  and n large that

$$\frac{1}{2} \zeta^{dn^{\xi}+d} P\{R_{\lfloor \alpha m \rfloor} \ge \lfloor \alpha m \rfloor (y+\varepsilon)\} P\{R_{\lfloor (1-\alpha)m \rfloor} \ge \lfloor (1-\alpha)m \rfloor z\}$$
  
$$\leq P\{R_{\lfloor \alpha m \rfloor + \lfloor (1-\alpha)m \rfloor + d\lceil n^{\xi}\rceil} \ge \lfloor \alpha m \rfloor (y+\varepsilon) + \lfloor (1-\alpha)m \rfloor z - 2\lceil n^{\xi}\rceil\}$$
  
$$\leq P\{R_n \ge nx\},$$

where  $\lfloor a \rfloor$  denotes the largest integer  $\leq a$ .

Now restrict  $\varepsilon > 0$  to such  $\varepsilon$  for which  $y + \varepsilon$  is a continuity point of  $\psi$ . As we already observed,  $\psi$  is nondecreasing so that this holds for all but at most countably many  $\varepsilon$ . If we now take the logarithm and divide by *n* and let  $n \to \infty$ , then we obtain from Proposition 2 that

$$\psi(x) \le \alpha \psi(y + \varepsilon) + (1 - \alpha) \psi(z).$$

The required (2.33) now follows by letting  $\varepsilon \downarrow 0$  such that  $y + \varepsilon$  runs only through continuity points of  $\psi$ .

Now assume, to derive a contradiction, that  $\psi$  is not continuous at some  $x \in (0, 1)$ . Since  $\psi$  is nondecreasing, this means that  $\psi$  has an upwards jump at x. There therefore exist  $x_n \downarrow x$  and  $y_n \uparrow x$  such that  $x_n$ ,  $y_n$  are continuity points of  $\psi$  and such that

$$\lim_{n \to \infty} \frac{\psi(x_n) - \psi(y_n)}{x_n - y_n} = \infty.$$

However, if  $z \in (x, 1)$  is also a continuity point of  $\psi$ , then (2.33) implies

$$\frac{\psi(z) - \psi(x_n)}{z - x_n} \ge \frac{\psi(x_n) - \psi(y_n)}{x_n - y_n}$$

as soon as  $0 < y_n < x_n < z$ . Thus  $(\psi(z) - \psi(x_n))/(z - x_n) \rightarrow \infty$ , which contradicts the boundedness of  $\psi$  on [0,1] (see (2.14), (2.15)). This proves the continuity of  $\psi$  on (0, 1).

**Proposition 4.** If (1.9) holds and  $d \ge 2$ , then  $\psi$  has the properties (1.10)–(1.15), so that Theorem 1 holds when  $d \ge 2$ .

*Proof.* We first note that  $\psi$  is continuous at 0. Indeed, we clearly have  $P\{R_n \ge nx\} = 1$  and  $\psi(x) = 0$  for  $x \le 0$ , while for  $\delta \in (0, 1)$ ,

 $P\{R_n \ge \delta n\} \ge P\{S_0, \dots, S_{\lfloor \delta n \rfloor} \text{ are distinct}\} \ge [P\{X(i) > 0\}]^{\delta n},$ 

exactly as in (2.14). It follows that  $\psi(\delta) \leq -\delta \log P\{X(i) > 0\}$ , as in (2.15). Hence also  $\lim_{\delta \downarrow 0} \psi(\delta) = 0$ .

This result, together with Lemma 3 and Proposition 2 shows that (1.10) holds for all x < 1. Moreover,  $\lim_{n\to\infty} (-1/n) P\{R_n \ge n\}$  exists by subadditivity, since  $\{R_n \ge n\} = \{S_0, S_1, \dots, S_{n-1}, \text{ are distinct}\}$ , so that

$$P\{R_{n+m} \ge n+m\} \le P\{R_n \ge n\}P\{R_m \ge m\}.$$

In addition  $P{R_n > n} = 0$ , so that (1.10) holds for all x and also (1.13) holds.

Next we show that  $\psi(x)$  is left continuous at x = 1. Assume that this is not the case. Since we already know that  $\psi(1)$  is finite, there must then exist some  $\beta > 0$  such that

$$\psi(1) = \psi(1-) + 2\beta = \lim_{x \uparrow 1} \psi(x) + 2\beta.$$
(2.34)

By (1.10) we also know that there exists a constant  $c_4 \ge 1$  such that

$$P\{R_n \ge n\} \le c_4 \exp[-n(\psi(1) - \beta)], \ n \ge 0.$$
(2.35)

Now if  $R_n \ge nx$ , then there are at most  $\lfloor (1 - x)n \rfloor$  indices  $0 \le k \le n - 1$  for which  $S_k = S_i$  for some  $0 \le i < k$  (compare the proof of Theorem 4.1 in [S3]). Let these indices be  $0 \le k_1 < k_2 < \cdots < k_\rho \le n - 1$ , (with  $\rho \le (1 - x)n$ ). The number of ways in which these indices can be chosen is for  $3/4 \le x \le 1$  at most

$$\sum_{\rho=0}^{\lfloor (1-x)n \rfloor} \binom{n}{\rho} \le c_5 x^{-xn} (1-x)^{-(1-x)n},$$

for some constant  $c_5$ , independent of x, n. All times  $j \notin \{k_1, \ldots, k_\rho\}$  must be "first visit times", that is, times j with  $S_j \neq S_i$  for  $0 \le i < j$ , so that  $R[k_i, k_{i+1} - 1] = k_{i+1} - k_i$ ,  $0 \le i \le \rho$ , where we take  $k_0 = 0$  and  $k_{\rho+1} = n$ . We conclude from this that

$$P\{R_n \ge nx\} \le \sum_{\rho} \sum_{k_1, \dots, k_{\rho}} \prod_{i=0}^{\rho} P\{R[k_i, k_{i+1} - 1] \ge (k_{i+1} - k_i)\}$$
  
$$\le \sum_{\rho} \sum_{k_1, \dots, k_{\rho}} [c_4]^{\rho+1} \exp\left[-\sum_{i=0}^{\rho} (k_{i+1} - k_i)(\psi(1) - \beta)\right]$$
  
$$= \sum_{\rho} \sum_{k_1, \dots, k_{\rho}} [c_4]^{\rho+1} \exp[-n(\psi(1) - \beta)]$$
  
$$\le [c_4]^{(1-x)n+1} c_5 x^{-xn} (1-x)^{-(1-x)n} \exp[-n(\psi(1) - \beta)].$$

It follows that

$$\psi(x) \ge -(1-x)\log c_4 + x\log x + (1-x)\log(1-x) + \psi(1) - \beta.$$

By taking the limit  $x \uparrow 1$ , we see that

$$\lim_{x\uparrow 1}\psi(x)\geq\psi(1)-\beta.$$

This contradicts (2.34) unless  $\beta = 0$ , and the left continuity of  $\psi$  at x = 1 follows.

Now that we have continuity of  $\psi$  on [0, 1] we obtain the convexity of  $\psi$  on [0, 1] from Lemma 3. We also have continuity at  $\pi$ , so that also (1.11) holds.

Next, we prove (1.12). By Markov's inequality,

$$P\{R_n \ge n(\pi + \varepsilon)\} \le e^{-n(\pi + \varepsilon)\lambda} E e^{\lambda R_n}$$

for any  $\lambda \geq 0$ . Consequently,

$$\lim_{n\to\infty}\frac{-1}{n}\log P\{R_n\geq n(\pi+\varepsilon)\}\geq (\pi+\varepsilon)\lambda+\lim_{n\to\infty}\frac{-1}{n}\log Ee^{\lambda R_n}.$$

The last limit exists by subadditivity, because  $Ee^{\lambda R_{n+m}} \leq Ee^{\lambda R_n} Ee^{\lambda R_m}$  for  $\lambda \geq 0$ . Moreover, it is proven in [H] that the right derivative at  $\lambda = 0$  of this limit equals  $-\pi$ . (This reference only considers simple random walk, but the proof applies to any random walk on  $\mathbb{Z}^d$ ). Thus for sufficiently small  $\lambda > 0$ ,

$$\lim_{n \to \infty} \frac{-1}{n} \log E e^{\lambda R_n} \ge \lim_{n \to \infty} \frac{-1}{n} \log e^{0 \cdot R_n} - \left(\pi + \frac{\varepsilon}{2}\right) \lambda = -\left(\pi + \frac{\varepsilon}{2}\right) \lambda$$

and

$$\lim_{n\to\infty}\frac{-1}{n}\log P\{R_n\geq n(\pi+\varepsilon)\}\geq\frac{\varepsilon}{2}\lambda.$$

This establishes (1.12).

Finally, (1.12) also implies (1.15). Indeed, convexity of  $\psi$  implies that for  $h > 0, x \ge \pi$  one has  $\psi(x + h) - \psi(x) \ge \psi(\pi + h) - \psi(\pi)$ .

The proof of Theorem 1 for d = 1 will be given in a separate paper. We conclude this section with the proof of Corollary 1.

*Proof of Corollary 1.* The upper bound (1.16) for  $\mu_n(F)$  can easily be obtained from a standard large deviation estimate with the help of (1.15) (compare the proof of Theorem 2.2.3 in [DZ]). We therefore concentrate on proving the lower bound (1.17) for  $\mu_n(G)$ .

Let *G* be an open set contained in the interval  $[\pi, \infty)$ . If  $\inf G \ge 1$ , then the infimum of  $\psi$  on *G* is infinity by (1.13) since x > 1 for all  $x \in G$ . Therefore the lower bound is trivial in this case. In the case that  $\inf G < 1$ , we can find for each  $x \in G \cap [\pi, 1)$  some constant  $\delta = \delta(x) > 0$  such that  $[x, x + \delta)$  is contained in  $(\pi, 1)$ . Noting that

$$\mu_n(G) \ge P\{R_n \ge x\} - P\{R_n \ge (x+\delta)n\},\$$

and using (1.15), we easily obtain for each  $x \in G \cap [\pi, 1)$  that

$$\liminf_{n \to \infty} \frac{1}{n} \log \mu_n(G) \ge -\psi(x).$$

This implies that

$$\liminf_{n\to\infty}\frac{1}{n}\log\mu_n(G)\geq-\inf_{x\in G\cap[\pi,1)}\psi(x).$$

In view of (1.13), the right hand side here is unchanged even if we take the infimum over all of *G* instead of  $G \cap [\pi, 1)$ .

#### 3. The Wiener sausage

We indicate here how to prove Theorem 2. We shall first deal with the one-dimensional case of Theorem 2. Because of the continuity of Brownian motion one can more or less write down the distribution of  $\Lambda_t$  and find  $\phi(x)$  explicitly when d = 1. For  $d \ge 2$  there are only some technical differences with the proof of Theorem 1 and we shall skip some details.

**Lemma 5.** If d = 1 and  $A \subset \mathbb{R}$  is any bounded, Lebesgue measurable, nonempty set, then for  $x \ge 0$ 

$$\phi(x) = \lim_{t \to \infty} \frac{-1}{t} \log P\{\Lambda_t(A) \ge tx\} = \begin{cases} \frac{x^2}{2} & \text{if } \mu = 0\\ 0 & \text{if } 0 \le x \le \mu \\ \frac{(x-\mu)^2}{2} & \text{if } 0 < \mu \le x. \end{cases}$$
(3.1)

*Proof.* Since Brownian motion is continuous, it is easy to see that for any  $A \subset [-K, K]$ 

$$\max_{s \le t} B(s) - \min_{s \le t} B(s) \le \Lambda_t(A) \le 2K + \max_{s \le t} B(s) - \min_{s \le t} B(s).$$

For the purposes of (3.1) we may therefore replace  $\Lambda_t(A)$  by  $\widetilde{\Lambda}_t := \max_{s \le t} B(s) - \min_{s \le t} B(s)$ . The density of this quantity when  $\mu = 0$  is explicitly given in [F1], and this immediately gives the result for  $\mu = 0$ . When  $\mu > 0$  it is still possible to derive an infinite series for the density of  $\widetilde{\Lambda}_t$  as in [F1], but it is easier to use the following crude estimates.

$$P\{\widetilde{\Lambda}_t \ge tx\} \ge P\{B(t) \ge tx\} = \frac{1}{\sqrt{2\pi t}} \int_{tx}^{\infty} e^{-(y-t\mu)^2/(2t)} dy$$

When  $x \le \mu$ , the integral in the right hand side is at least 1/2, and when  $x > \mu$ , then this integral behaves asymptotically like

$$\frac{1}{(x-\mu)\sqrt{2\pi t}} \exp\left[-(x-\mu)^2 t/2\right]$$
(3.2)

(see [F2], Lemma VII.1.2). To prove (3.1) we only need an upper bound for  $P\{\widetilde{\Lambda}_t \ge tx\}$  when  $x > \mu$ . For this we can use

$$P\{\widetilde{\Lambda}_t \ge tx\} \le \sum_{k \le t} P\{\max_{k \le s \le k+1} |B(s) - B(k)| \ge t^{3/4}\} + \sum_{k,\ell \le t} P\{|B(k) - B(\ell)| \ge tx - 2t^{3/4}\}.$$

The first sum in the right hand side here is by the familiar reflection principle ([IM], Section 1.7)  $O(t \exp[-(t^{3/4} - \mu)^2/2])$ . The summand in the second sum equals

$$\frac{2}{\sqrt{2\pi}} \int_{|k-\ell|^{-1/2} [tx-2t^{3/4}-|k-\ell|\mu]}^{\infty} e^{-y^2/2} dy = O\left(\exp\left[-\frac{(t(x-\mu)-2t^{3/4})^2}{2t}\right]\right),$$
(3.3)

(by virtue of [F2], lemma VII.1.2 again). (3.1) now follows easily.

For the remainder of this section we take  $d \ge 2$  and  $A \subset \mathbb{R}^d$  a fixed bounded, Lebesgue measurable set with  $\operatorname{cap}(A) > 0$ , where  $\operatorname{cap}(A)$  is as in (1.21). We usually abbreviate  $\Lambda_t(A)$  to  $\Lambda_t$ . The constants  $c_i$  which appear in the proof do not necessarily have the same value as in the previous sections. Without loss of

generality we assume that *A* is contained in  $\mathscr{S} := \{x \in \mathbb{R}^d : |x| \le 1\}$ , the unit ball centered at the origin. Unless otherwise indicated the initial point *B*(0) equals **0**. We shall write *B*(*t*; *i*) for the *i*-th coordinate of *B*(*t*). Finally, we define

$$V[a,b] = \bigcup_{a \le s \le b} (B(s) + A)$$

In this notation,

$$\Lambda_t = \lambda_d(V[0, t]).$$

Before we can give the analogue of Lemma 1 we need some crude a priori bounds on the distribution of  $\Lambda_t$ . The following lemma is from [ABP].

**Lemma 6.** Let  $d \ge 2$  and let A be a Lebesgue measurable set contained in  $\mathcal{S}$ , and with cap(A) > 0. Then

$$P\{\Lambda_{1/2}(A) \ge x\} > 0 \quad for \ all \quad x \ge 0.$$
 (3.4)

*Proof.* This is proven in [ABP] for a Brownian motion without drift. For a Brownian motion with drift this then follows from the fact that on any finite time interval Brownian motions with and without drift are absolutely continuous with respect to each other.

**Lemma 7.** For all  $x \ge 0$  there exists an  $L_1 = L_1(x, A, d) < \infty$  such that

$$P\{\Lambda_t \ge tx\} \ge e^{-L_1 t}, \quad t \ge 1.$$
(3.5)

*Proof.* Clearly, for  $t \ge 1$ 

$$P\{\Lambda_t \ge tx\} \ge P\{\Lambda_{\lfloor t \rfloor} \ge 2\lfloor t \rfloor x \ge tx\}.$$

Thus, at the cost of replacing x by 2x we may restrict ourselves to t an integer. Now, by Lemma 6, for any  $x \ge 0$ ,  $P\{\Lambda_{1/2} \ge x\} > 0$ . There then exists a constant  $c_1$  such that even

$$P\left\{\Lambda_{1/2} \ge x, \sup_{0 \le s \le 1/2} |B(s; 1)| \le c_1\right\} > 0.$$
(3.6)

Now let  $\mathscr{E}_k = \mathscr{E}_{k,1} \cap \mathscr{E}_{k,2} \cap \mathscr{E}_{k,3}$ , where

$$\mathscr{E}_{k,1} := \{ B(k + \frac{1}{4}; 1) - B(k; 1) > c_1 + 1 \};$$

$$\mathscr{E}_{k,2} := \left\{ \lambda_d \left( V[k + \frac{1}{4}, k + \frac{3}{4}] \right) \ge x \\ \text{and} \quad \left| B(k+s; 1) - B(k + \frac{1}{4}; 1) \right| \le c_1 \quad \text{for} \quad \frac{1}{4} \le s \le \frac{3}{4} \right\};$$

$$\mathscr{E}_{k,3} := \Big\{ B(k+1;1) - B(k+\frac{3}{4};1) > 2c_1 + 2 \Big\}.$$

Then the events  $\mathscr{E}_k$ ,  $k = 0, 1 \dots$  are independent and all have the same probability. This probability equals

$$P\left\{B\left(\frac{1}{4};1\right) > c_1 + 1\right\} P\left\{B\left(\frac{1}{4};1\right) > 2c_1 + 2\right\} \\ \times P\{\Lambda_{1/2} \ge x, \sup_{0 \le s \le 1/2} |B(s;1)| \le c_1\}.$$

Let us denote this probability by  $\alpha$ . Then  $\alpha > 0$  by virtue of (3.6), and

$$P\left\{\bigcap_{0\leq k\leq t-1}\mathscr{E}_k\right\}\geq \alpha^t.$$
(3.7)

Finally, note that if  $\mathscr{E}_k$  and  $\mathscr{E}_\ell$  occur for some  $k < \ell$ , then

$$V[k + \frac{1}{4}, k + \frac{3}{4}] \cap V[\ell + \frac{1}{4}, \ell + \frac{3}{4}] = \emptyset,$$
(3.8)

because the points in V[k+1/4, k+3/4] lie within distance 1 of  $\{B(s) : k+1/4 \le s \le k+3/4\}$ , and therefore have their first coordinate in

$$\begin{bmatrix} B(k+\frac{1}{4};1) - c_1 - 1, B(k+\frac{1}{4};1) + c_1 + 1 \end{bmatrix} \subset (B(k;1), B(k+1;1)).$$

In addition,

$$B(k + 1; 1) > B(k + \frac{3}{4}) + 2c_1 + 2$$
  

$$\geq B(k + \frac{1}{4}; 1) + c_1 + 2 > B(k; 1) + 2c_1 + 3,$$

so that the intervals (B(k; 1), B(k + 1; 1)) for different k are disjoint. Thus (3.8) indeed holds, and on the event  $\bigcap_{0 \le k \le t-1} E_k$  one has

$$\Lambda_t \ge \sum_{k=0}^{t-1} \lambda_d \left( V \left[ k + \frac{1}{4}, k + \frac{3}{4} \right] \right) \ge tx.$$

The lemma now follows from (3.7), by taking  $L_1 = -\log \alpha$ .

**Lemma 8.** For all  $0 \le L < \infty$ , there exist a constant  $x_1 = x_1(L, A, d) < \infty$  such that

$$P\{\Lambda_t \ge tx_1\} \le e^{-Lt}, \quad t \ge 1.$$
(3.9)

*Proof.* When  $\mu = 0$  this is essentially contained in Theorem 1 of [BT]. When  $\mu \neq 0$  we cannot use the scaling property of Brownian motion, but we can still get our estimate (even for  $\mu = 0$ ) by following part of the proof in [BT]. We do not attempt to get sharp estimates such as given in [BB]. Probably this is still possible even if  $\mu \neq 0$ , but crude estimates are good enough for our purposes.

Since  $\Lambda_t$  is increasing in A, it is sufficient to restrict ourselves to  $A = \mathscr{S}$ . We define

$$\theta_0 = 0, \ \theta_{n+1} = \inf\{s > \theta_n : \|B(s) - B(\theta_n)\| > 1\},\\ \tau_n = \theta_n - \theta_{n-1}, \quad n \ge 1,$$

and

$$\nu(t) = \max\{n : \theta_n \le t\}.$$

These are the definitions of Section 4 of [BT] with y = 1. As in (4.5) of [BT], there then exist constants  $c_2$  and h such that  $\Lambda_t \le c_2 + h\nu(t)$ . Since the  $\tau_n$ ,  $n \ge 1$ , are i.i.d. we therefore obtain for  $t \ge 1$ ,  $x \ge 2c_2$ ,

$$P\{\Lambda_t \ge tx\} \le P\{hv(t) \ge \frac{1}{2}tx\}$$
$$= P\left\{\sum_{i=1}^{\lceil tx/(2h)\rceil} \tau_i \le t\right\}$$
$$\le e^t \left[E\exp(-\tau_1)\right]^{tx/(2h)}.$$

Clearly  $E \exp(-\tau_1) < 1$ , so that (3.9) is satisfied for

$$x_1 \ge \frac{2h(L+1)}{-\log E \exp(-\tau_1)} + 2c_2.$$

We can now prove the following analogue of Lemma 1:

**Lemma 9.** For any fixed  $x_2 \ge 0$ , there exist constants  $c_3, c_4 \in (0, \infty)$  such that for  $s, t \ge 1, y \le sx_2, z \le tx_2$ , it holds that

$$P\{\Lambda_{s+t+(st)^{1/(d+1)}} \ge y+z-c_3(st)^{1/(d+1)}\}$$
  

$$\ge c_4 \exp[-d(st)^{1/(d+1)}]P\{\Lambda_s \ge y\}P\{\Lambda_t \ge z\}.$$
(3.10)

*Proof.* To prove this we bring in an additional Brownian motion  $\widehat{B} = \{\widehat{B}(t)\}_{t \ge 0}$  which is independent of  $B = \{B(t)\}_{t \ge 0}$ , but has the same distribution as B.  $\widehat{\Lambda}_t$  is defined by (1.18) with B replaced by  $\widehat{B}$ . We further define for  $w \in \mathbb{R}^d$ 

$$M_{s,t}(w) = M_{s,t}(w, B, \widehat{B})$$
  
=  $\lambda_d \Big( \bigcup_{r \le s} (B(r) + A) \cap \bigcup_{r \le t} (\widehat{B}(r) + B(s) + w + A) \Big),$ 

and for  $p, r \ge 0$  we define

$$T(r) = T^{(p,s)}(r) = \begin{cases} B(r) & \text{if } r \le s+p \\ B(s+p) + \widehat{B}(r-s-p) & \text{if } r > s+p. \end{cases}$$

Finally we define

$$\Xi_q = [-q, q]^d$$

and take

$$q = q(s, t) = (st)^{1/(d+1)}$$

Analogously to the proof of Lemma 1 we now have for any  $c_3$ ,  $\beta \ge 0$ ,

$$P\{\Lambda_{s+t+q} \ge y+z-c_{3}q\}$$

$$\ge \int_{w\in\Xi_{q}} \frac{1}{q} \int_{0\le p\le q} P\{B(s+p)-B(s)\in dw\}dp$$

$$\times P\{\Lambda_{s}\ge y, \widehat{\Lambda}_{t}\ge z, M_{s,t}(w, B, \widehat{B})\le c_{3}q\}$$

$$\ge \int_{w\in\Xi_{q}} \frac{1}{q} \int_{0\le p\le q} P\{B(s+p)-B(s)\in dw\}dp$$

$$\times P\{y\le \Lambda_{s}\le s\beta, z\le \widehat{\Lambda}_{t}\le t\beta, M_{s,t}(w, B, \widehat{B})\le c_{3}q\}. (3.11)$$

Of course

$$P\{B(s+p) - B(s) \in dw\} = \frac{1}{(2\pi p)^{d/2}} \exp[-\|w\|^2/(2p)]dw,$$

and for some constant  $c_5 > 0$  and all  $w \in \Xi_q$ ,

$$\frac{1}{q} \int_{0 \le p \le q} \frac{1}{(2\pi p)^{d/2}} \exp[-\|w\|^2/(2p)] dp$$
  
$$\ge \frac{1}{q(2\pi q)^{d/2}} \int_{q/2}^q \exp[-dq^2/q] dp \ge c_5 q^{-d/2} \exp[-dq].$$

Thus, the right hand side of (3.11) is at least

$$c_{5}q^{-d/2} \exp[-dq] E \left\{ I[y \le \Lambda_{s} \le s\beta] I[z \le \widehat{\Lambda}_{t} \le t\beta] \times \lambda_{d} \left( \{ w \in \Xi_{q} : M_{s,t}(w, B, \widehat{B}) \le c_{3}q \} \right) \right\}.$$
(3.12)

Moreover, for fixed realizations of *B* and  $\widehat{B}$  with  $\Lambda_s \leq s\beta$ ,  $\widehat{\Lambda}_t \leq t\beta$  we have

$$\begin{split} &\int_{\Xi_q} M_{s,t}(w,B,\widehat{B})dw \\ &\leq \int_{\mathbb{R}^d} dw \int_{\mathbb{R}^d} db I \Big[ b \in \bigcup_{r \leq s} (B(r) + A) \Big] I \Big[ b \in \bigcup_{r \leq t} (\widehat{B}(r) + B(s) + w + A) \Big] \\ &= \int_{\mathbb{R}^d} db I \Big[ b \in \bigcup_{r \leq s} (B(r) + A) \Big] \widehat{\Lambda}_t \\ &= \Lambda_s \widehat{\Lambda}_t \leq \beta^2 st. \end{split}$$

Thus, if  $\Lambda_s \leq s\beta$ ,  $\widehat{\Lambda}_t \leq t\beta$ , and if we choose  $c_3 \geq \beta^2$  (for a  $\beta$  to be chosen later), then

$$\lambda_d \left( \{ w \in \Xi_q : M_{s,t}(w, B, \widehat{B}) \le c_3 q \} \right) \ge \lambda_d(\Xi_q) - \frac{\beta^2 st}{c_3 q}$$
$$= 2^d q^d - \frac{\beta^2}{c_3} q^d \ge q^d.$$

In this situation the expression in (3.12) is at least

$$c_6 q^{-d/2} \exp[-dq] q^d P\{y \le \Lambda_s \le s\beta\} P\{z \le \widehat{\Lambda}_t \le t\beta\}$$
  
$$\ge c_6 \exp[-dq] P\{y \le \Lambda_s \le s\beta\} P\{z \le \widehat{\Lambda}_t \le t\beta\}.$$

Finally,

$$P\{y \le \Lambda_s \le s\beta\} = P\{\Lambda_s \ge y\} - P\{\Lambda_s > s\beta\}.$$

But if  $y \leq sx_2$ , then

$$P\{\Lambda_s \ge y\} \ge P\{\Lambda_s \ge sx_2\} \ge e^{-L_1s},$$

for  $L_1 = L_1(x_2, A, d)$  as in Lemma 7. By Lemma 8 we can therefore first choose  $\beta = c_7(x_2, A, d)$  (and then  $c_3 \ge \beta^2$ ) such that for  $s \ge 1$ ,

$$P\{\Lambda_s > sc_7\} \le e^{-(L_1+1)s} \le \frac{1}{2}P\{\Lambda_s \ge y\},$$

and

$$P\{y \le \Lambda_s \le sc_7\} \ge \frac{1}{2}P\{\Lambda_s \ge y\}.$$

Similarly

$$P\{z \le \widehat{\Lambda}_t \le tc_7\} \ge \frac{1}{2}P\{\Lambda_t \ge z\}.$$

Combining all these estimates we find that for  $c_4 = c_6/4$ ,

$$P\{\Lambda_{s+t+q} \ge y+z-c_3q\} \ge c_4 \exp[-dq]P\{\Lambda_s \ge y\}P\{\Lambda_t \ge z\}. \qquad \Box$$

We now define

$$\phi(x) = \liminf_{n \to \infty} \frac{-1}{n} \log P\{\Lambda_n(A) \ge nx\}.$$

This  $\phi(x)$  is finite for all *x*, by virtue of Lemma 7. We can repeat the proof of Proposition 2 with (3.10) taking the role of (2.11) to obtain that

$$\lim_{n \to \infty} \frac{-1}{n} \log P\{\Lambda_n(A) \ge nx\} = \phi(x)$$

at each  $x \ge 0$  at which  $\phi$  is right continuous. This further implies that if  $x \pm \varepsilon \ge 0$  are continuity points of  $\phi$ , then

$$\limsup_{t \to \infty} \frac{-1}{t} \log P\{\Lambda_t(A) \ge tx\} \le \limsup_{t \to \infty} \frac{-1}{t} \log P\{\Lambda_{\lfloor t \rfloor}(A) \ge \lfloor t \rfloor \frac{t}{\lfloor t \rfloor}x\}$$
$$\le \lim_{n \to \infty} \frac{-1}{n} \log P\{\Lambda_n(A) \ge n(x+\varepsilon)\}$$
$$= \phi(x+\varepsilon),$$

as well as

$$\liminf_{t \to \infty} \frac{-1}{t} \log P\{\Lambda_t(A) \ge tx\} \ge \liminf_{t \to \infty} \frac{-1}{t} \log P\{\Lambda_{\lceil t \rceil}(A) \ge \lceil t \rceil \frac{t}{\lceil t \rceil}x\}$$
$$\ge \lim_{n \to \infty} \frac{-1}{n} \log P\{\Lambda_n(A) \ge n(x-\varepsilon)\}$$
$$= \phi(x-\varepsilon).$$

Thus, at any continuity point x > 0 of  $\phi$ ,

$$\lim_{t \to \infty} \frac{-1}{t} \log P\{\Lambda_t(A) \ge tx\} = \phi(x).$$
(3.13)

Further, one can show exactly as in Lemma 3 that

$$\phi(x) \le \alpha \phi(y) + (1 - \alpha)\phi(z),$$

when x, y, z are continuity points of  $\phi$  such that  $x, y, z > 0, 0 < \alpha < 1$ , and  $x = \alpha y + (1 - \alpha)z$ . From this convexity property and the finiteness of  $\phi$  we then conclude that  $\phi$  is convex, continuous and finite for all x > 0, and that (3.13) holds for all x > 0 (compare Lemma 3). Moreover, we trivially have

$$\phi(x) = \lim_{t \to \infty} \frac{-1}{t} \log P\{\Lambda_t(A) \ge tx\}$$
  
= 
$$\lim_{t \to \infty} \frac{-1}{t} \log P\{\Lambda_t(A) \ge 0\} = 0 \quad \text{for} \quad x \le 0.$$

Also, analogously to the proof of Proposition 4, for  $0 < \delta < 1$ ,

$$P\{\Lambda_t(A) \ge t\delta\} \ge P\{\Lambda_{t\delta}(A) \ge t\delta\} \ge e^{-L_1 t\delta}$$

for  $L_1 = L_1(1, A, d)$ , by Lemma 7. Thus  $\phi(\delta) \le L_1\delta$ , and  $\phi$  is also continuous at 0. Thus (3.13) holds for all  $x \in \mathbb{R}$ . This proves (1.23) and (1.27). Relation (1.24) follows trivially from (1.19). Finally, (1.26) and (1.28) can be proven in exactly the same way as (1.12) and (1.15).

*Acknowledgements.* We are indebted to O. Angel, I. Benjamini and Y. Peres for producing a proof of Lemma 6 in response to a question of ours. We also are grateful to Frank den Hollander for a useful correspondence in connection with this paper, and for several references. The research for this paper was carried out during a year long visit by Y. Hamana in 1999–2000 to Cornell University in the Program for Overseas Researchers of the Japan Ministry of Education. Y. H. thanks Cornell University for its hospitality. The research of H. Kesten was supported by the NSF through a grant to Cornell University.

### References

- [ABP] Angel, O., Benjamini, I., Peres, Y.: A large Wiener sausage from crumbs, Electron. Comm. Probab. 5, 67–71 (2000)
- [BB] van den Berg, M., Bolthausen, E.: Asymptotics of the generating function for the volume of the Wiener sausage, Probab. Theory Relat. Fields, 99, 389–397 (1994)
- [BBH] van den Berg, M., Bolthausen, E., den Hollander, F.: Moderate deviations for the volume of the Wiener sausage, Report No 9933, University of Nijmegen (1999)

- [BT] van den Berg, M., Tóth, B.: Exponential estimates for the Wiener sausage, Probab. Theory Relat. Fields, 88, 249–259 (1991)
- [De] Dekking, F.M.: On transience and recurrence of generalized random walks, Z. Wahrsch. verw. Gebiete, 61, 459–465 (1982)
- [DZ] Dembo, A., Zeitouni, O.: Large Deviations Techniques and Applications, second ed., Springer-Verlag (1998)
- [DV] Donsker, M.D., Varadhan, S.R.S.: On the number of distinct sites visited by a random walk, Comm. Pure Appl. Math., **32**, 721–747 (1979)
- [Do1] Doob, J.L.: Semimartingales and subharmonic functions, Trans. Amer. Math. Soc., 77, 86–121 (1954)
- [Do2] Doob, J.L.: Classical Potential Theory and its Probabilistic Counterpart, Springer-Verlag (1984)
- [EL] Eisele, T., Lang, R.: Asymptotics for the Wiener sausage with drift, Probab. Theory Relat. Fields, **74**, 125–140 (1987)
- [F1] Feller, W.: The asymptotic distribution of the range of sums of independent random variables, Ann. Math. Statist., **22**, 427–432 (1951)
- [F2] Feller, W.: Introduction to Probability Theory and its Applications, vol I, third ed., John Wiley & Sons (1968)
- [H] Hamana, Y.: Asymptotics of the moment generating function for the range of random walks, J. Theoret. Probab., (to appear)
- [IM] Itô, K., McKean, H.P., Jr.: Diffusion Processes and their Sample Paths, Springer-Verlag, (1965)
- [HK] Hamana, Y., Kesten, H.: Large deviations for the range of an integer valued random walk, to appear in Ann. Inst. M. Poincaré (2001)
- [Ka] Kakutani, S.: Two-dimensional Brownian motion and harmonic functions, Proc. Acad. Japan, 20, 706–714 (1944)
- [Ke] Kesten, H.: On the number of self-avoiding walks II, J. Math. Phys., **5**, 1128–1137 (1964)
- [KS] Karatzas, I., Shreve, S.E.: Brownian Motion and Stochastic Calculus, second ed., Springer-Verlag (1991)
- [PS] Pólya, G., Szegö, G.: Aufgaben und Lehrsätze aus der Analysis, second ed., Springer-Verlag (1954)
- [S1] Spitzer, F.: Electrostatic capacity, heat flow and Brownian motion, Z. Wahrsch. verw. Gebiete, 3, 110–121 (1964)
- [S2] Spitzer, F.: Discussion of "Subadditive ergodic theory" by J. F. C. Kingman, Ann. Probab., 1, 904–905 (1973)
- [S3] Spitzer, F.: Principles of Random Walk, second ed., Springer-Verlag (1976)