



Capital Market Implications of Corporate Disclosure: German Evidence

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Abstract

This paper investigates the relationship between annual report disclosure, market liquidity, and capital cost for firms registered on the Deutsche Börse. Disclosure is comprehensively measured using the innovative Artificial Intelligence Measurement of Disclosure (AIMD). Results show that annual report disclosure enhances market liquidity by changing investors' expectations and inducing portfolio adjustments. Trading frictions are negatively associated with disclosure. The study provides evidence for a capital-cost-reduction effect of disclosure based on the analysis of investors' return requirements and market values. Altogether, no evidence is found that the information processing at the German capital market is structurally different from other markets.

Keywords: annual report disclosure, capital market, market liquidity, cost of capital

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1 Introduction

The impact of corporate disclosure on capital markets has been discussed for decades—and the interest has not decreased over time. [Lev and Ohlson \(1982: 249\)](#) stated in their review of research in the field: “A decade and a half of the most concerted and ambitious research effort in accounting history is evaluated here”. Even today neither capital market implications nor other consequences of corporate disclosure have been fully explored or understood.

Corporate disclosure is intended to provide information on the operating activities, and the financial and business situation of a firm to an interested audience. Firms' disclosure is therefore not limited to financial information, but also addresses a great variety of non-financial information.

It is undisputed that corporate disclosure increases the amount of information available to addressees. This information effect is only a means to an end and firms will hardly provide information out of philanthropic motives. In fact, it is purely economic reasoning that drives disclosure of corporate information. In a related vein [Dye \(2001\)](#) noted that capital market impacts of disclosure are in the spotlight while implications for merchandise markets are of peripheral interest. Prior research mainly

discussed the relation between disclosure and an increase in wealth of shareholders, but also evaluated other market consequences of disclosure.

Following [Healy and Palepu \(2001\)](#) two types of disclosure consequences exist: (1) the liquidity of markets and individual securities and (2) the cost of capital to firms. Most prior studies investigated the implications of annual report disclosure on capital markets outside Germany or relied on specific, narrow proxies for disclosure with questionable validity (e.g., [Leuz and Verrecchia 2000](#), [Leuz 2003](#)). I therefore examine these relations in the German capital market in 2006 using the innovative Artificial Intelligence Measurement of Disclosure (AIMD) that directly derives disclosure measures from annual reports.

Capital-market-related research mainly addresses voluntary disclosure as from a mandatory disclosure that does not vary between firms (of a certain industry, size and exchange segment) no firm-specific endogenous market reactions can be expected. A research design to address the mandatory disclosure issue must be either normative or requires sufficient variance in mandatory disclosure levels that can often be found in international samples only. In line with, e.g., [Chen, Chen, and Cheng \(2008\)](#), [Langberg](#)

and Sivaramakrishnan (2010), and Zechman (2010) I therefore disregard mandatory disclosure implications but focus on firm-specific disclosure consequences.

The remaining paper is organized as follows. In section two the relations between disclosure and liquidity aspects are discussed. Section three looks at disclosure from a cost-of-capital perspective. The research design for the empirical evaluation of the four hypotheses developed in sections two and three is described in section four. Section five reports the results of the empirical investigation. Finally, section six summarizes the findings.

2 Liquidity Hypotheses Development

Different aspects of liquidity exist. On the one hand share turnover measures the overall trading activity. On the other hand the bid-ask spread also addresses liquidity issues as the friction that burdens traders.

2.1 Trading activity

Trading activity (i.e. share turnover) is an obvious measure for market liquidity as liquidity increases with more liquid funds used to buy securities.

The market price balances different individual expectations and reflects the consensus expectations of all market participants while changing consensus expectations imply changes in market prices (Elton, Gruber, and Gultekin 1981, Karpoff 1987, Ryan and Taffler 2004, and Ryan 2005). Consensus expectations only change, however, if the individual expectations of some investors change and they adjust their portfolios by buying or selling securities. Changes in individual expectations therefore induce share turnover.

Individual expectations are co-determined by the level of information that is available to the respective market participant. Roll (1988) identified three different factors to influence capital market participants' expectations: (1) common macroeconomic knowledge, (2) industry-specific information, and (3) firm-specific information. All of them could potentially be addressed by corporate disclosure. If disclosure provides information that is new to some investors their individual expectations change. As a result, portfolios need to be adjusted and share turnover occurs. This effect not only exists if the information is new to some investors, but also if certain market participants interpret available in-

formation differently against the background of the new information (Karpoff 1986, Kim and Verrecchia 1991).

Prior empirical research revealed a strong and sustainable impact on trading activity for various disclosure instruments. The influence of earnings announcements on share turnover is very well documented. In the 1960s empirical studies started to report an increasing share turnover in connection with earnings announcements (e.g., Beaver 1968). Multiple papers followed to examine this relationship based on larger samples, more advanced empirical methods, and in different capital markets (e.g., Morse 1981, Bamber 1986, Sivakumar and Waymire 1994). They consistently showed a positive effect of disclosure on share turnover. On the other hand, laboratory experiments fail to reproduce a turnover increase following an earnings announcement in simulated markets (Gillette, Stevens, Watts, and Williams 1999).

Share turnover is also influenced by other types of information that could be provided by corporate disclosure, including macroeconomic data (McGuinness 1999), social responsibility disclosure (Patten 1990), and ad-hoc information disclosure (Ryan and Taffler 2004, Muntermann and Guettler 2007). Based on the idea that the information content of annual reports decreases with increasing temporal distance to the accounting year end, the positive relation between timeliness of disclosure and share turnover demonstrated by Leventis and Weetman (2004) can also be interpreted as evidence for a positive association of disclosure and trading activity. Furthermore, Bailey, Li, Mao, and Zhong (2003) and Francis, Nanda, and Wang (2006) reported an increase in share turnover following increased non-discrimination disclosure requirements of U.S. regulation F-D.

A different stream of literature discusses a negative relation between corporate disclosure and share turnover under certain conditions. Share turnover increases with increasing uncertainty of investors and a greater variance in individual expectations (Barron 1995, Bamber, Barron, and Stober 1997, Bamber, Barron, and Stober 1999, Linsmeier, Thornton, Venkatachalam, and Welker 2002). The higher the variance of individual expectations the higher the information advantage of informed investors that must be transformed into real economic benefits by market transactions—causing share turnover. With higher information asymmetry therefore

more insider transactions occur. Corporate disclosure is expected to reduce information asymmetry and hence reduce insider transactions and trading activities. Eventually, the negative effect of corporate disclosure is based on the assumption that a considerable proportion of all stock market transactions are insider transactions and their decrease cannot be compensated by additional transactions of previously uninformed investors.

Altogether, the idea that corporate disclosure develops to an extent where information asymmetry vanishes and previously informed investors leave the market, whereas the decrease in informed transactions cannot be compensated by an increase in share turnover from other (now better informed) investors seems to be implausible. In line with the first stream of literature an increasing share turnover with an increasing information provision by corporate disclosure is therefore hypothesized.

H1: Corporate disclosure and share turnover are positively associated.

2.2 Bid-Ask Spread

Information asymmetry reduces, i.e. information levels of market participants are more even, if (1) information that has been private is made public or (2) disclosure provides information that is entirely new to the market (and the relative proportion of public information increases to the disadvantage of private information).

In this respect two different views on how information asymmetry reduction affects the bid-ask spread are discussed in the literature. On the one hand, a lower information asymmetry reduces the risk of passive traders to lose against better informed active traders. As passive traders compete the actual spread reduces accordingly. Bloomfield and O'Hara (1999) reproduced this competitive effect in a laboratory setting. On the other hand, a lower spread makes the price-offering service of passive traders less attractive as the risk and the resulting compensation provided by active traders are lower (Diamond and Verrecchia 1991). Consequently, passive traders exit the market and market liquidity decreases. With lower information asymmetry uninformed traders may also become more aware of being uninformed and leave the market. Altogether, liquidity decreases and the spread increases (Wu

and Zhang 2002). Diamond and Verrecchia (1991) argued that the effect of a spread reduction due to reduced risk outweighs the contrary impact of a spread increase due to trader exits. A negative relation between corporate disclosure and bid-ask spread is therefore expected.

In contrast to main stream literature some papers predict a positive association of disclosure and information asymmetry. McNichols and Trueman (1994) argued that market participants respond to disclosure with a varying motivation to search for additional information, i.e. to reduce information asymmetry on their own. Verrecchia (1982), Diamond (1985), and Kim and Verrecchia (1994) assumed that market participants differ in their information-processing capabilities and therefore information asymmetry increases if more information is available.

Various empirical evidence supports a negative association of disclosure and spread. Early papers relied on a contingency approach where certain firm characteristics are expected to accompany higher information asymmetry. For instance, a high proportion of insiders holding stocks, a high proportion of institutional investors and a large influence of principal shareholders—that are all considered to accompany high information asymmetry—imply a high spread (e.g., Chiang and Venkatesh 1988, Heflin and Shaw 2000). Later studies directly addressed the relation between corporate disclosure, information asymmetry, and spread. For instance, Welker (1995), Healy, Hutton, and Palepu (1999), Leuz and Verrecchia (2000), Brockman and Chung (2001), Heflin, Shaw, and Wild (2005), and Attig, Fong, Gadhoun, and Lang (2006) provided evidence that firms' voluntary disclosure policy is negatively associated with spreads. The spread is also reducing with the disclosure of particular information, including the disclosure of explored and unexplored oil reserves (Raman and Tripathy 1993, Boone 1998, Boone, Luther, and Raman 1998), segmental reporting information (Greenstein and Sami 1994), and management forecasts of quarterly earnings (Coller and Yohn 1997). However, the disclosure effect on the spread seems to depend on the type of information (Guo, Lev, and Zhou 2004).

In summary, previous theoretical as well as empirical work suggests the following hypothesis.

H2: Corporate disclosure and the bid-ask spread are negatively associated.

3 Equity Cost Hypothesis Development

The expected reduction of equity cost is probably the most important motive for firms to disclose information (Choi 1973, Cooke 1993, Verrecchia 2001, Lambert, Leuz, and Verrecchia 2007). The basic idea of the cost reduction effect is twofold (Core 2001): First, comprehensive disclosure reduces investors' information uncertainty as discussed in the previous section. Second, if information asymmetry decreases, investors require a lower risk premium (Bailey, Karolyi, and Salva 2006), i.e. cost of capital reduces.

Two main streams of literature exist that examine the relation of cost of capital and corporate disclosure in detail. On the one hand, information asymmetry is considered to drive capital cost as investors require compensation for higher transaction costs that result from higher bid-ask spreads (Amihud and Mendelson 1986, Botosan, Plumlee, and Xie 2004). In line with the negative relation between disclosure and information asymmetry cost of capital reduces with an increasing disclosure level. On the other hand, corporate disclosure influences the non-diversifiable prognosis risk for unknown returns (Klein and Bawa 1976, Coles, Loewenstein, and Suay 1995, Clarkson, Guedes, and Thompson 1996). Here, investors are assumed to decide based on uncertain returns while the precise distribution is unknown and distribution parameters are estimated from past experience and other available information. Uncertain returns are modeled using estimators for the distribution function parameters that depend on an investor's level of information. As the prognosis risk is non-diversifiable it affects cost of capital but is not considered in popular capital market models (Botosan 2006).

A number of approaches are discussed in the literature to quantify the cost of capital concept empirically.

3.1 Investors' Return Requirements

Cost of equity is opportunity cost of alternative investment opportunities (e.g., Shapiro 1978, Kerins, Smith, and Smith 2004, Indjejikian 2007) and equals the rate of return investors require to hold stocks of a firm. Easley and O'Hara (2004) regarded an imperfect capital market where differently informed investors with rational expectations hold risky securities. The overall amount of information

is separated into public and private, i.e. information asymmetry is proportional to the percentage of private information. Easley and O'Hara (2004) concluded that in equilibrium firms with lower information asymmetry are ceteris-paribus confronted with lower return expectations of their investors. Corporate disclosure is introduced into the model and increases the number of investors that have access to private information or reduces the amount of private information for the benefit of public information. Both effects reduce the required rate of return and decrease firms' capital cost accordingly. In their empirical study Easley, Hvidkjaer, and O'Hara (2002) reported an increase in returns with higher information asymmetry. Therefore, the following hypothesis is formulated.

H3: Corporate disclosure and cost of capital are negatively associated.

3.2 Other direct measures for cost of capital

The Capital Asset Pricing Model (CAPM) (Sharpe 1964, Lintner 1965, Mossin 1966) is inappropriate because of the following theoretical concerns. None of the mechanisms discussed in section two, neither the transaction-cost-based nor the information-based one, can be modeled within the CAPM. While the first is incompatible with the imputed frictionless capital market, the latter is incompatible with the assumption of well-known distribution functions of all relevant market parameters. Accordingly, the CAPM does not leave room for effects of information asymmetry and corporate disclosure because of the assumed high information efficiency and identical expectations of all market participants (Easley, Hvidkjaer, and O'Hara 2002). A descriptive analysis (available from the BuR – Business Research website) reveals no association of disclosure and the beta for the German sample used in sections four and five. Similarly there is no room for the transaction-cost-based as well as the prognosis-risk-related impact of corporate disclosure on capital cost within the Arbitrage Pricing Theory (APT) (Lehmann and Modest 1988, Fama and French 1996) framework (Easley and O'Hara 2004). I cannot apply expected income discounting methods (e.g., Gebhardt, Lee, and Swaminathan 2001, Botosan and Plumlee 2002) as no IBES prognoses data could be obtained for the sample. Cost of capital can also not be measured by the direct welfare loss of a

capital-raising firm, i.e. the underpricing when issuing securities (Rock 1986, Beatty and Ritter 1986) as a considerable amount of the underpricing is not driven by information asymmetry, but results from non-rational decisions and principal-agent problems (Ritter and Welch 2002). Furthermore, most sample firms do not have IPOs within the time period under review.

Aside from the above capital cost concepts other approaches are also infrequently used to evaluate the relation between corporate disclosure and cost of capital empirically—and provide mixed evidence. Some reveal an unconditioned capital cost reduction because of disclosure (Nikolaev and Van Lent 2005, Cheng, Collins, and Huang 2006, Eaton, Nofsinger, and Weaver 2007, Black, Carnes, Jandik, and Henderson 2007), some document a relation that depends on the kind of information (Richardson and Welker 2001), and some a relation that depends on how the information is disclosed (Botosan and Plumlee 2002).

3.3 Market Value

Maximizing market value is a means to an end of reducing cost of equity. As market value is the future expected cash flows from the firm to its investors discounted with the equity cost rate, with a lower equity cost rate the market value increases. Within short periods of time the number of shares issued can be considered to be constant. Market value and share price therefore contain similar information with respect to disclosure effects.

Many papers regarded “value relevance”, i.e. a positive influence of corporate disclosure on a firm’s market value as the ultimate aim of disclosure (e.g., Hughes 2000, Espahbodi, Espahbodi, Rezaee, and Tehranian 2002, Young and Guenther 2003, Şabac, Scott, and Wier 2005, Bowen, Davis, and Matsu-moto 2005).

To analyze this effect accounting information and market capitalization can be linked through abnormal returns. For an indefinite period of time a firm’s market value is the sum of all discounted future cash payments to investors (Peasnell 1982, Feltham and Ohlson 1995, Bernard 1995). The Ohlson and Feltham framework allows a firm’s market value to be approximated using earnings prognoses and equity book values if the clean-surplus condition is fulfilled. It requires that all changes in the equity book value that do not result from transactions with

investors are recognized in the income statement. Hence, $equity_t = equity_{t-1} + earnings_t + dividends_t$ (Ohlson 1995, Feltham and Ohlson 1995) and market capitalization of equity can be restated as

$$(1) \text{ market value} = equity_t + \sum_{\tau=1}^{\infty} E(earnings_{t+\tau} - r \cdot equity_{t+\tau-1}) / (1+r)^\tau$$

where r is the risk-free interest rate (Lundholm 1995). In addition to equity book values the formula only contains expected values of future abnormal returns above a risk-free equity return. Bernard (1995) concluded from his analysis of the U.S. capital market that valid estimates can be obtained using a short forecasting horizon of only $T = 4$ periods. Any conceivable mandatory disclosure includes the equity book value for the reporting period. Further mandatory or voluntary disclosure should therefore provide information that is useful to estimate abnormal returns. While dividend prognoses are also influenced by investors’ decision on whether to distribute or retain profits, abnormal earnings can be forecast using information about operating activities and the financial and business situation of a firm only.

Still, the above arguments and economic intuition do not allow to predict the direction of the relation, i.e. whether there is a proportional or inversely proportional relationship. Corporate disclosure can either contain “good” or “bad” news, i.e. increase or reduce market expectations of abnormal returns. Considering the “good news bias” that is well documented in the literature (e.g., Clarkson, Kao, and Richardson 1994, Karamanou and Vafeas 2005) I hypothesize a positive relation between disclosure and a firm’s market value.

Similarly, differences between accounting standards’ and markets’ perceptions can be explained by assets and liabilities. It is well known that changes in accounting standards induce price reactions at the capital market (Dyckman and Smith 1979, Collins, Rozeff, and Dhaliwal 1981, Healy and Palepu 2001). Disclosure is an instrument to bridge the gap between market perceptions and accounting standards by providing further information about the value of assets and liabilities (Robinson and Lucas 2007). On the one hand, differences between market values and accounting values result as the balance sheet does not recognize certain assets and liabilities at market price. Market and book values of equity differ as a consequence. On the other hand,

what is perceived as an asset differs between an accounting standard and the market. Although certain intangibles are not recognized in the balance sheet they are valued by the market and co-determine a firm's market value. Accordingly, market value can also be explained by varying recognition of assets/liabilities following accounting standards and market perceptions.

Often book values do not exceed market values of assets while the reverse is true for liabilities. Furthermore, the value of intangible assets considerably exceeds that of intangible liabilities. The book value of total equity is therefore lower than market capitalization. As intangible assets are much more important for firms than intangible liabilities I will restrict myself to the former. A difference between market capitalization and equity book values may result because (1) valuation differences between accounting standards and the market are small, (2) the firm has few, low-value intangible assets, and (3) the firm has difficulties in explaining the inherent value of intangible assets to the capital market (Beattie and Thomson 2005). The first effect only relates to mandatory disclosure that is not addressed here as discussed in the introduction. The second effect is not linked to disclosure, but the third can easily be influenced by a firm's disclosure policy. Given the bias of disclosure towards "good news" market value should increase following disclosure.

Empirical papers researching the topic often rely on multiple regression models with market capitalization as the dependent, certain disclosure measures as independent, and controls for items that are not (or not at market values) recognized in the balance sheet, e.g., equity book values, total assets and liabilities. They overwhelmingly support a positive relation between disclosure and market capitalization (e.g., Landsman 1986, Barth 1991, Healy, Hutton, and Palepu 1999, Ritter and Welch 2002). Contradictory, Mangena and Tauringana (2007) failed to reproduce a relation between disclosure level and market capitalization.

Summarizing, based on theory and empirical evidence I hypothesize the following.

H4: Corporate disclosure and market capitalization are positively associated.

4 Research Design

This section documents the empirical design to evaluate the four hypotheses from the preceding sections.

4.1 Regression Models

First I establish the regression models. In addition to corporate disclosure a number of control variables used in previous studies are included to ascertain the explanatory power of disclosure for the respective dependent variable.

4.1.1. Trading Activity

Prior empirical studies suggested that a number of parameters besides corporate disclosure affect trading activity. In line with Leuz and Verrecchia (2000) I use the free float that measures shareholder concentration as a control for trading activity. Empirical evidence of whether firm size is associated with share turnover is mixed. Some studies find a negative association of firm size and share turnover (e.g., Tkac 1999); some a positive (e.g., Bessembinder, Chan, and Seguin 1996, Leuz and Verrecchia 2000). The positive relation between price volatility and share turnover is well documented in the empirical literature (see the review of Bessembinder and Seguin 1993). Membership to stock exchange segments may also influence share turnover (Tkac 1999) as inclusion into major stock exchange indices is found to increase trading activity (Shleifer 1986, Harris and Gurel 1986, Vijh 1994, Beneish and Whaley 1996, Lynch and Mendenhall 1997). Other studies documented increasing trading activities for cross-listed firms (e.g., Bancel and Mittoo 2001 and the overview at Karolyi 1998).

Therefore, I evaluate H1 using the following multiple regression equation.

$$(2) \text{ SHARE_TURNOVER} = a_0 + a_1 \cdot \text{DISCLOSURE} + a_2 \cdot \text{FREEFLOAT} + a_3 \cdot \text{SIZE} + a_4 \cdot \text{VOLATILITY} + a_5 \cdot \text{EXCHANGE_SEGMENT} + a_6 \cdot \text{CROSSLISTING} + \varepsilon.$$

4.1.2. Bid-Ask Spread

Prior research developed various approaches to measuring spreads. Following Stoll (2000) I here use the proportional quoted half-spread defined as (ask price – bid price) / (2 · closing price) as all

required information is easily available. I measure this spread as the annual average based on daily ask, bid, and closing prices. The proportional quoted half-spread does not allow to separate the transaction-cost-based from information-based spread components (Glosten and Harris 1988). These two components cannot be separated here because information about each single trade within a certain period of time is not available.

However, I argue that the approximation error is acceptable as the transaction-cost spread component is low relative to the information-based on the Deutsche Börse. Between the mid-1980s and 1995 the proportion of the information-based spread on the NYSE increased from 43 percent to between 47 and 54 percent (Stoll 1989, Menyah and Paudyal 2000). Transaction costs further decrease with increasing electronic trading (Weber 2006) and the spread is more sensitive to incoming new information in electronic trading systems compared to floor trading (Aitken, Frino, Hill, and Jarnecic 2004). In 2006 about 92 percent of all trades on the Deutsche Börse were processed using the electronic Xetra system. Therefore, the information-based spread is supposed to be considerably larger than the transaction-cost-based spread and the approximation error is acceptable. Furthermore, for the regression analysis the total level of the transaction-cost spread is irrelevant as long as it is identical for all securities. This is most likely true for order processing cost, but not for inventory cost that varies with the risk of an individual security. However, inventory cost contributes to the total spread only marginally (Stoll 1989).

I consider a number of control variables when evaluating the relation between corporate disclosure and the spread. Empirical evidence suggests that the spread increases with large, especially institutional investors as they tend to have insider knowledge (e.g., Chiang and Venkatesh 1988, Leuz 2003). The analyst following has a depressant effect on the spread (e.g., Brennan and Subrahmanyam 1995, Leuz and Verrecchia 2000), but no data are available for the German capital market, unfortunately. Stock exchange turnover as a measure of trading activity, the closing price, and market capitalization also inhibit the spread (e.g., Lev 1988, Leuz and Verrecchia 2000, Stoll 2000, Brockman and Chung 2003, Leuz 2003, Heflin, Shaw, and Wild 2005,

Chang, Chen, Liao, and Mishra 2006). I do not consider the closing price and market capitalization as additional controls because the proportional quoted half-spread already accounts for the closing price and market value information beyond firm size. Prior research found higher spreads with higher volatility of prices (e.g., Leuz and Verrecchia 2000, Stoll 2000, Leuz 2003, Heflin, Shaw, and Wild 2005) and lower spreads for larger firms (e.g., Leuz and Verrecchia 2000, Leuz 2003, Heflin, Shaw, and Wild 2005, Ali, Chen, and Radhakrishnan 2007).

I use the following multiple regression equation to evaluate H2.

$$(3) \text{ SPREAD} = a_0 + a_1 \cdot \text{DISCLOSURE} + a_2 \cdot \text{FREEFLOAT} + a_3 \cdot \text{SIZE} + a_4 \cdot \text{SHARE_TURNOVER} + a_5 \cdot \text{VOLATILITY} + \varepsilon.$$

4.1.3. Investors' Return Requirements

Return expectations, i.e. ex-ante returns of securities are not easy to measure from an empirical perspective. Following Francis, LaFond, Olsson, and Schipper (2004) I therefore use ex-post returns measured as annual average of ex-post daily stock return deviations from the industry average. This deviation also controls for industry effects. In line with Easley, Hvidkjaer, and O'Hara (2002) I use firm size as a control variable in addition to industry and the following multiple regression equation results for H3.

$$(4) \text{ ABNORMAL_RETURN} = a_0 + a_1 \cdot \text{DISCLOSURE} + a_2 \cdot \text{SIZE} + \varepsilon.$$

4.1.4. Market Capitalization

Accounting information is not completely value irrelevant. Therefore, I control for book value of equity in line with Hughes (2000) and Bowen, Davis, and Rajgopal (2002). Following Bowen, Davis, and Rajgopal (2002) I also control for other firm size effects using an additional size measure. In line with Abdolmohammadi (2005) I also use abnormal returns beyond the industry average as a control variable, although no significant influence was documented in prior research. Altogether the following multiple regression equation results for H4.

$$(5) \text{ MARKET_VALUE} = a_0 + a_1 \cdot \text{DISCLOSURE} + a_2 \cdot \text{COMMON_EQUITY} + a_3 \cdot \text{SIZE} + a_4 \cdot \text{ABNORMAL_RETURN} + \varepsilon.$$

4.2 Empirical Measures

This section documents the measures used for testing the models set up in section 4.1.

4.2.1. Proxies for Corporate Disclosure

Previous studies used a wide variety of methods to measure corporate disclosure, including the popular rating provided by the Association for Investment and Management (AIMR) (e.g., [Lang and Lundholm 1993, 1996](#), [Welker 1995](#), [Sengupta 1998](#), [Healy, Hutton, and Palepu 1999](#), [Bushee and Noe 2000](#), [Botosan and Plumlee 2002](#), [Lundholm and Myers 2002](#)), scores from the Center for International Financial Analysis & Research (CIFAR) (e.g., [Carlin and Mayer 2003](#), [Hope 2003](#), [Bushee 2004](#), [Bushman, Piotroski, and Smith 2004](#), [Francis, Khurana, and Pereira 2005](#)), Standard & Poor's Scoring (e.g., [Khanna, Palepu, and Srinivasan 2004](#), [Durnev and Kim 2005](#), [Ali, Chen, and Radhakrishnan 2007](#), [Doidge, Karolyi, and Stulz 2007](#)), and scores derived from manual, proprietary text analysis. No raw data are available for the German sample.

Therefore, I derive disclosure measures from firms' annual reports using the Artificial Intelligence Measurement of Disclosure (AIMD) ([Grüning 2011](#)). It measures disclosure in ten distinctive information dimensions (financial information, information about customers, information about the value chain, information about employees, information about R&D, information about corporate strategy, information about corporate governance, information about the stock market, and information about social and environmental responsibility). Appendix A provides further details about AIMD.

For the empirical analysis I aggregate the AIMD measures for the ten disclosure dimensions (DISCLOSURE_01 ... DISCLOSURE_10) using factor analysis. Based on a KMO of 0.928 and one eigenvalue above one (8.123) a single factor (DISCLOSURE) is extracted.

4.2.2. Other measures

The remaining variables are taken from either Datastream Worldscope or the Hoppenstedt Firmen-datenbank, depending on information availability and sample coverage.

Following [Francis, LaFond, Olsson, and Schipper \(2004\)](#) I proxy abnormal returns beyond the industry average (ABNORMAL_RETURN) by the 2006 annual average of daily abnormal stock returns against the industry average using four industry groups described in Table 1. Daily stock returns are

obtained from Datastream Worldscope. Equity book values (COMMON_EQUITY) are measured by the 2006 book value of total equity obtained from Hoppenstedt. In line with [Archambault and Archambault \(2003\)](#) CROSSLISTING is the number of countries a firm is listed outside Germany based on the Datastream list of stock exchanges. EXCHANGE_SEGMENT distinguishes between firms listed in the Prime (coding 1) and General Standard (coding 2) based on Deutsche Börse information. The percentage of shares not held by strategic investors (FREEFLOAT) measures shareholder concentration in line with [Leuz and Verrecchia \(2000\)](#). Data are obtained from the Deutsche Börse Weighting Files. MARKET_VALUE is the market value of the firm obtained from Datastream. Trading activity is measured as the turnover of shares (SHARE_TURNOVER) in 2006 obtained from Datastream. Following [Hagerman and Zmijewski \(1979\)](#) and [Ecker, Francis, Kim, Olsson, and Schipper \(2006\)](#), SIZE is measured as the natural logarithm of total sales of the firm in 2006 obtained from Hoppenstedt. In line with [Stoll \(2000\)](#) information asymmetry is measured by the annual average of daily proportionally quoted half-spreads 2006 (SPREAD). Ask, bid, and closing prices are obtained from Datastream. VOLATILITY is the standard deviation of daily prices 2006 divided by the mean price 2006. Daily prices are taken from Datastream.

4.3 Sample and Descriptive Results

The hypotheses are evaluated using a sample that initially consists of all 600 German firms listed in the Prime and General Standard of the Deutsche Börse in 2006. I exclude foreign firms listed on the Deutsche Börse. Table 1 reports selected characteristics of the initial sample.

Table 1: Sample characteristics

Panel A: Industry of sample firms

industry	number of firms in sample
production (NACE groups A, B, C, D, E, F)	264
commerce (NACE group G)	33
financial (NACE group J)	44
services (NACE groups H, I, K, L, M, N, O, P, Q)	259
	600

Panel B: Size of sample firms

total sales	number of firms in sample*
less than 10 ^{7.5} €	159
less than 10 ^{8.0} €	109
less than 10 ^{8.5} €	119
less than 10 ^{9.0} €	68
less than 10 ^{9.5} €	51
less than 10 ¹⁰ €	22
less than 10 ^{10.5} €	17
more than 10 ^{10.5} €	11
	556

* 44 missing values

The effective sample size decreases considerably because of missing values. Particularly, I could only obtain English-language annual reports for 390 firms. The effective sample size decreases further depending on availability of other measures in the Datastream and Hoppenstedt databases. I apply a list-wise deletion approach in order to retain a large proportion of the sample variance. Appendix B provides selected descriptive statistics for the sample. I address the delay in the availability of annual reports by regressing disclosure measures for the annual report of 2005 that became available sometime in 2006 with controls and a dependent variable for 2006.¹

Table 2: Multivariate regression of share turnover on disclosure

Panel A: Robust regression

	Predicted Sign	(Standardized) Coefficient	t-Statistics	p	VIF
DISCLOSURE	+	0.4546***	(5.12)	0.000	2.33
FREEFLOAT	+	0.0580*	(1.82)	0.070	1.13
SIZE	?	0.0564	(1.45)	0.147	2.36
VOLATILITY	+	0.0510	(1.52)	0.130	1.14
EXCHANGE_SEGMENT	-	0.0106	(0.45)	0.652	1.15
CROSSLISTING	+	0.4414***	(8.75)	0.000	1.19
Condition Number		10.71			
N		348			
R ²		0.6084			

Panel B: Robust rank regression

	Predicted Sign	(Standardized) Coefficient	t-Statistics	p	VIF
DISCLOSURE	+	0.2253***	(3.81)	0.000	2.03
FREEFLOAT	+	0.3473***	(7.26)	0.000	1.12
SIZE	?	0.0589	(0.99)	0.324	2.06
VOLATILITY	+	0.1295***	(3.01)	0.003	1.14
EXCHANGE_SEGMENT	-	-0.2562***	(-5.23)	0.000	1.17
CROSSLISTING	+	0.1687***	(5.95)	0.000	1.14
Condition Number		9.79			
N		348			
R ²		0.4456			

The table reports standardized coefficients, t-statistics, significance values (p), and variance inflation factors (VIF) from a linear and a ranked OLS regression with heteroscedasticity-corrected errors. The dependent variable is the share turnover (SHARE_TURNOVER) 2006 obtained from Datastream. DISCLOSURE is the factor (KMO 0.928; one eigenvalue of 8.123 above one) extracted from the ten AIMD disclosure measures derived from annual reports of 2005. FREEFLOAT is the percentage of shares not held by strategic investors derived from the Deutsche Börse Weighting files 2006. SIZE is the natural logarithm of total sales 2006 obtained from the Hoppenstedt Firmendatenbank. VOLATILITY is the standard deviation of daily prices 2006 divided by the mean price 2006 with prices being obtained from Datastream. EXCHANGE_SEGMENT distinguishes between firms listed in the Prime (coding 1) and General Standard (coding 2). CROSSLISTING is the number of countries a firm is listed outside Germany based on the Datastream list of stock exchanges. All metric variables are winsorized at the 0.02 level at both extremes to adjust for potential outliers. * p<0.1, ** p<0.05, *** p<0.01.

¹ Raw data are available from the BuR – Business Research website.

5 Empirical Results

This section reports the results of the hypotheses testing and the robustness analysis.

5.1 Regression Results

Table 2 reports the results for a multivariate regression of share turnover on disclosure with heteroscedasticity robust errors in panel A. To address potential non-linear relations between the dependent and independent variables, panel B also lists the results for a robust rank regression. The models do not provide evidence for multicollinearity.

The signs of the controls are in line with expectations. A comparison of panels A and B reveals indications of a potential non-linearity between freefloat and trading activity that could result as the freefloat is constrained between 0 and 100 percent while the

trading activity measure has no natural upper limit. Similarly nonlinearities might exist for the volatility and the exchange segment measure. All models congruently report a significant positive association between disclosure and trading activity and support H1.

Table 3 reports the results for a heteroscedasticity robust multivariate regression of spreads on disclosure in panel A. Again, panel B lists the results of robust rank regression to address potential non-linear relations. There is no indication of severe multicollinearity.

All models support the hypothesized negative association between disclosure and spread. With the exception of price volatility the controls are significant and have the expected sign. Overall, the models support the inhibiting effect stated in H2.

Table 3: Multivariate regression of bid-ask spread on disclosure

Panel A: Robust regression

	Predicted Sign	(Standardized) Coefficient	t-Statistics	p	VIF
DISCLOSURE	-	- 0.1589***	(-3.08)	0.002	2.70
FREEFLOAT	-	- 0.2219***	(-4.76)	0.000	1.07
SIZE	-	- 0.5549***	(-10.04)	0.000	2.28
SHARE_TURNOVER	-	0.1181***	(3.71)	0.000	1.84
VOLATILITY	+	0.0463	(0.84)	0.399	1.15
Condition Number		5.89			
N		354			
R ²		0.4820			

Panel B: Robust rank regression

	Predicted Sign	(Standardized) Coefficient	t-Statistics	p	VIF
DISCLOSURE	-	-0.1589***	(-3.08)	0.002	2.70
FREEFLOAT	-	-0.2219***	(-4.76)	0.000	1.07
SIZE	-	-0.5549***	(-10.04)	0.000	2.28
SHARE_TURNOVER	-	0.1181***	(3.71)	0.000	1.84
VOLATILITY	+	0.0463	(0.84)	0.399	1.15
Condition Number		7.99			
N		354			
R ²		0.6164			

The table reports standardized coefficients, t-statistics, significance values (p), and variance inflation factors (VIF) from a linear and a ranked OLS regression with heteroscedasticity-corrected errors. The dependent variable (SPREAD) is the annual average of daily proportionally quoted half-spreads 2006 obtained from Datastream. DISCLOSURE is the factor (KMO 0.928; one eigenvalue of 8.123 above one) extracted from the ten AIMD disclosure measures derived from annual reports of 2005. FREEFLOAT is the percentage of shares not held by strategic investors derived from the Deutsche Börse Weighting files 2006. SIZE is the natural logarithm of total sales 2006 obtained from the Hoppenstedt Firmendatenbank. SHARE_TURNOVER is the 2006 turnover in shares obtained from Datastream. VOLATILITY is the standard deviation of daily prices 2006 divided by the mean price 2006 with prices being obtained from Datastream. All metric variables are winsorized at the 0.02 level at both extremes to adjust for potential outliers. * p<0.1, ** p<0.05, *** p<0.01.

Table 4: Multivariate regression of abnormal returns on disclosure

Panel A: Robust regression

	Predicted Sign	(Standardized) Coefficient	t-Statistics	p	VIF
DISCLOSURE	–	–0.2504***	(–3.61)	0.000	2.08
SIZE	+	0.5772***	(9.43)	0.000	2.08
Condition Number		1.01			
N		361			
R ²		0.1875			

Panel B: Robust rank regression

	Predicted Sign	(Standardized) Coefficient	t-Statistics	p	VIF
DISCLOSURE	–	–0.1538**	(–2.47)	0.014	1.77
SIZE	+	0.5215***	(9.16)	0.000	1.77
Condition Number		4.97			
N		361			
R ²		0.1898			

The table reports standardized coefficients, t-statistics, significance values (p), and variance inflation factors (VIF) from a linear and a ranked OLS regression with heteroscedasticity-corrected errors. The dependent variable (ABNORMAL_RETURN) is the 2006 annual average of daily abnormal stock returns against the industry average using four industry groups based on the NACE industry classification obtained from Datastream. DISCLOSURE is the factor (KMO 0.928; one eigenvalue of 8.123 above one) extracted from the ten AIMD disclosure measures derived from annual reports of 2005. SIZE is the natural logarithm of total sales 2006 obtained from the Hoppenstedt Firmendatenbank. All metric variables are winsorized at the 0.02 level at both extremes to adjust for potential outliers. *p<0.1, **p<0.05, ***p<0.01.

Table 4 reports the results of a robust multivariate regression of abnormal returns on disclosure in panel A and the results of a robust rank regression in panel B. There are no indications for multicollinearity.

The models congruently support hypothesis H₃ that disclosure and abnormal returns are negatively associated.

Table 5 reports the results of a robust multivariate regression of market value on disclosure in panel A.

Panel B again lists the results of a robust rank regression to address potential nonlinear relations.

The condition number for the model in panel A is at the lower end of the continuum between moderate and strong multicollinearity according to [Belsley, Kuh, and Welsch \(1980: 105\)](#) and therefore suggests a moderate multicollinearity problem. The alternative variance inflation factors do not indicate a severe multicollinearity. The model in panel B does not show any indication of severe multicollinearity.

Table 5: Multivariate regression of market value on disclosure

Panel A: Robust regression

	Predicted Sign	(Standardized) Coefficient	t-Statistics	p	VIF
DISCLOSURE	+	0.1219***	(3.60)	0.000	2.31
COMMON_EQUITY	+	0.7241***	(15.38)	0.000	4.94
SIZE	?	0.1342**	(2.55)	0.011	5.08
ABNORMAL_RETURN	?	–0.0428*	(–1.84)	0.067	1.27
Condition Number		44.65			
N		345			
R ²		0.8539			

Table 5 continued: Multivariate regression of market value on disclosure

Panel B: Robust rank regression

	Predicted Sign	(Standardized) Coefficient	t-Statistics	p	VIF
DISCLOSURE	+	0.1424***	(4.21)	0.000	1.88
COMMON_EQUITY	+	0.7139***	(14.09)	0.000	3.79
SIZE	?	0.0990*	(1.80)	0.073	4.12
ABNORMAL_RETURN	?	-0.0128	(-0.43)	0.671	1.28
Condition Number		11.31			
N		345			
R ²		0.8029			

The table reports standardized coefficients, t-statistics, significance values (p), and variance inflation factors (VIF) from a linear and a ranked OLS regression with heteroscedasticity-corrected errors. The dependent variable is the natural logarithm of market value (MARKET_VALUE) obtained from Datastream. DISCLOSURE is the factor (KMO 0.928; one eigenvalue of 8.123 above one) extracted from the ten AIMD disclosure measures derived from annual reports of 2005. COMMON_EQUITY is the natural logarithm of the book value of total equity obtained from Hoppenstedt Firmendatenbank. SIZE is the natural logarithm of total sales 2006 obtained from the Hoppenstedt Firmendatenbank. ABNORMAL_RETURN is the 2006 annual average of daily abnormal stock returns against the industry average using four industry groups based on the NACE industry classification obtained from Datastream. All metric variables are winsorized at the 0.02 level at both extremes to adjust for potential outliers. * p<0.1, ** p<0.05, *** p<0.01.

It is not surprising to find a positive relation between market value and common equity. In addition, both models provide evidence for a positive association of disclosure and market value. In line with [Abdalmohammadi \(2005\)](#), no significant association between abnormal returns and market value could be found. Altogether, the results support H4 that disclosure is value relevant.

5.2 Robustness Analysis

Various robustness tests have been conducted to find out if sampling issues or the particular model specifications affect the findings.

I test the regression models from section 5.1 for a sample selection bias using the two-step methodology of [Heckman \(1979\)](#). Here, the population is equivalent to all 600 listed firms. The samples consists of the 345 to 361 firms that could be considered for the respective regression models. I use the exchange segment and the natural logarithm of total sales as determinants in the selection equation to address the fact that data might be more likely available for larger, Prime Standard firms. Lambda values for the share turnover and market value regressions are insignificant indicating no sample selection problem. Lambda values are significant for the spread and the abnormal return regressions. For these models I apply a Heckman correction. The Heckman-corrected results (see Table A-3 in Appendix C) are structurally identical to the results shown in Tables 3 and 4 and support the significant

negative association of disclosure and the respective dependent variables. Altogether, the Heckman analysis provides no evidence that the findings are affected by sample selection bias.

To examine whether the models are robust against varying specifications I use various alternative dependent and independent proxies. The factor extracted from the natural logarithm of the ten AIMD measures, the sum of the ten AIMDs and the sum of the natural logarithm of the ten AIMDs are used as alternative disclosure measures. I use total sales, the number of employees and total assets, and the natural logarithm of both as alternative size proxies. A binary variable if firms are crosslisted in the U.S. is used as an alternative proxy of crosslisting. Finally, I test if weekly and monthly averages of spreads and volatility as an alternative to daily averages affect the results. I found structurally identical results. Tables A-4 to A-7 in Appendix C provide the regression results of a selection of these alternative model specifications.

6 Conclusion

Corporate disclosure is discussed in the literature to have considerable capital market implications. In line with [Healy and Palepu \(2001\)](#) I examine market liquidity and capital cost implications of annual report disclosure for a sample of German firms listed on the Deutsche Börse. Different from other papers that rely on specific, narrow proxies for disclosure I derive disclosure measures from annual



reports using AIMD, a comprehensive, innovative computerized approach using artificial intelligence. The findings provide evidence that annual report disclosure positively affects market liquidity in the German capital market. Information disclosure changes market participants' individual expectations which in turn require portfolio adjustments leading to trading activities. The friction that hinders trading activities proxied by the bid-ask spread is also vulnerable to disclosure influences and the information and transaction-cost-based spread reduces with more information being disclosed. I also found that market liquidity benefits from the provision of additional information via disclosure in the German capital market in line with U.S. evidence (e.g., [Welker 1995](#), [Healy, Hutton, and Palepu 1999](#), [Heflin, Shaw, and Wild 2005](#)).

Popular approaches to measure cost of capital, including the CAPM and the APT, do not provide insights into a potential equity-cost-reducing effect of disclosure as they do not allow to model information asymmetry reductions. Consequently, I found no association of the systematic risk (beta) and the level of annual report disclosure. This finding is contradictory to the influential [Botosan \(1997\)](#) analysis but in line with theory and in line with the empirical studies of [Auer \(1998\)](#), [De Alencar \(2005\)](#), and [Linsley and Shrives \(2006\)](#). I found that investors' return requirements proxied by ex-post abnormal returns are associated with a higher level of annual report disclosure for the German sample. In line with [Easley, Hvidkjaer, and O'Hara \(2002\)](#) this supports the view that capital cost reduction is a disclosure consequence within the [Easley and O'Hara \(2004\)](#) framework for the German capital market.

The findings also support a capital-cost-reduction effect using market capitalization as an indirect measure of cost of capital. In line with e.g., [Healy, Hutton, and Palepu \(1999\)](#), [Ritter and Welch \(2002\)](#), and [Linsley and Shrives \(2006\)](#) I found that market value is positively associated with annual report disclosure.

Using a sample of listed German firms the study supports the view that corporate disclosure increases market liquidity and reduces cost of capital. It therefore does not reveal evidence that the information processing at the German capital market might be structurally different from that on other capital markets.

Appendices

Appendix A: Artificial Intelligence Measurement of Disclosure (AIMD)

The Artificial Intelligence Measure of Disclosure (AIMD) measures how informative a corporate text (e.g., annual report) is across ten information dimensions:

1. sales-market and customer-related information,
2. employee-related information,
3. corporate environment information,
4. financial information,
5. corporate governance information,
6. research-and-development-related information,
7. social and environmental responsibility information,
8. capital-market-related information,
9. corporate strategy information, and
10. information relating to the value chain.

AIMD differs from other disclosure measures with various respects. As described above it covers a broad understanding of disclosure, including non-financial and voluntary disclosure. Furthermore, AIMD extends word-count systems (e.g., [Hussainey, Schleicher, and Walker 2003](#)) by analysing more complex semantical units ("phrases"), and partly considers the grammatical structure of sentences. It does not require human interaction and judgement to derive disclosure measures from annual reports and therefore goes beyond current applications that only assist human beings in a manual coding process (e.g., [Beattie, McInnes, and Fearnley 2004](#)). Accordingly, AIMD is cost-effective and can be used with large samples.

AIMD derives individual disclosure measures across the ten distinctive information dimensions listed above. As for most of these information dimensions it is not possible to discriminate between voluntary and mandatory disclosure ([Cooke and Zeff 2007](#)) AIMD does not distinguish between mandatory and voluntary disclosure. This does not weaken the explanatory power of the results as a fixed, mandatory disclosure component will not affect the estimated coefficients except in the form of a constant.

AIMD implements an information-retrieval vector space model ([Salton, Wong, and Yang 1975](#)) using artificial-intelligence (AI) techniques and consists of the two usual AI phases. First, in the training phase, a coding scheme is established based on a sample of annual reports. Second, in the application phase,



AIMD applies the coding scheme to a large number of new texts.

Subsequently, some technical details are discussed. [Grüning \(2011\)](#) provided further documentation. AIMD measures disclosure based on the occurrence of certain N-grams, i.e. ordered sequences of N words. For example, “order backlog” is a bigram and “order backlog development” is a trigram.

In the training phase, a coding scheme is created on the basis of a sample of representative annual reports, attempting to classify and quantify the diversity of corporate disclosures. Empirical evidence suggests that the coding scheme quickly converges and a training sample of 24 exemplary annual reports is comprehensive and complete, i.e. covers all relevant aspects of corporate disclosure ([Grüning 2011](#)).

To reduce complexity and to improve applicability the different orthographical (e.g., “labor” vs. “labour”) and grammatical forms (plural, conjugation, and word order within an N-gram) in the raw coding scheme are standardized to a morphological canonical form.

Accordingly, the final coding scheme is independent of (1) stopwords without meaning (e.g., “and”, “or”), (2) the morphology of elements of N-grams and (3) permutations of the elements of N-grams. For example, the standardized bigram “employee number” also represents the tetragram “number of the employees”. AIMD uses the stopword list of the Information Retrieval Group of the University of Glasgow for the first step. For the grammatical and orthographical normalization in the second step it performs a dictionary-based stemming process using the Automatically Generated Inflection Database (AGID) that contains 112,503 roots of 281,904 inflected forms and the Variant Conversion Info (Var-Con) that lists spelling variants of 16,019 words.

Following the morphological normalization, permutations of the elements of an N-gram are standardized. Different grammatical structures of the English language result in varying word orders within a sentence. When measuring the level of corporate disclosure only semantic differences are of importance. For instance, there is no relevant semantic difference between the N-grams “flexible arrangement of the work time” and “an arrangement of flexible working time”. The final coding scheme comprises 13,314 standardized N-grams, covering

about 10^{24} different semantic units relevant for corporate disclosure.

In the application phase AIMD applies the coding scheme from the training phase to a larger number of annual reports. Following [Weber \(1990\)](#) AIMD establishes a disclosure measure for each information dimension using category counts. The entire application phase is automated using proprietary software written in C#.

As it does not contain any random component the application phase of AIMD is perfectly reliable, being entirely objective. [Grüning \(2011\)](#) demonstrated that AIMD has criterion validity and construct validity based on large U.S. and German samples. Construct validity, i.e. the convergence of the results of different measurement techniques, is tested against several common disclosure proxies. The analysis reveals strong and highly significant correlations against the AIMR rating, Standard & Poor’s Transparency and Disclosure Scoring, the ratings of the business magazines *Manager Magazin* and *Capital/Focus Money* as well as against firms’ disclosure self evaluations. AIMD also has criterion validity, i.e. can replicate theoretically predicted relations empirically, in the information asymmetry reduction proxied by the spread and the probability of information based trading (PIN) for a sample of about 13,700 and 11,600 U.S. firm-year observations, respectively.

AIMD disclosure measures can be derived from different forms of disclosure, including annual/quarterly reports, records from analysts’ meetings, press news, conference calls, and corporate web pages. AIMD raw data is available from the author on request.

Appendix B: Descriptive Statistics

Table A-1 provides descriptive statistics for the dependent and independent variables. The table lists the size of the available subsample, the mean, quantiles including the median, and the standard deviation. The disclosure level varies considerably between industries (not reported) with highest mean disclosure for financial and lowest for service firms. It also varies with the exchange segment with higher mean disclosure for Prime Standard firms.

Table A-2 reports the correlation matrix between the variables. The correlations do not provide indications of multicollinearity problems.



Table A-1: Descriptive Statistics

Variable	Sub-sample	Mean	0.25 Quantile	Median	0.75 Quantile	Standard Deviation
DISCLOSURE_01	390	4,904	3,305	4,364	5,944	2,439
DISCLOSURE_02	390	1,445	833	1,235	1,794	866
DISCLOSURE_03	390	2,175	1,370	1,891	2,661	1,195
DISCLOSURE_04	390	1,093	684	953	1,378	602
DISCLOSURE_05	390	439	214	349	578	328
DISCLOSURE_06	390	1,408	859	1,174	1,719	828
DISCLOSURE_07	390	2,114	1,407	1,944	2,548	1,040
DISCLOSURE_08	390	2,025	1,264	1,818	2,479	1,102
DISCLOSURE_09	390	673	367	554	838	432
DISCLOSURE_10	390	399	199	286	491	335
DISCLOSURE	390	0.000	-0.691	-0.227	0.422	0.994
ABNORMAL_RETURN	599	-0.000	0.003	0.011	0.026	0.122
COMMON_EQUITY [mio €]	525	1,022	17	52	234	4,438
CROSSLISTING	569	0.111	0.000	0.000	0.000	0.707
EXCHANGE_SEGMENT	600	1.463	1.000	1.000	2.000	0.499
FREEFLOAT	597	0.521	0.278	0.517	0.770	0.302
MARKET_VALUE [mio €]	597	1,745	21	69	435	6,644
SHARE_TURNOVER	590	65,041	1,351	5,119	19,045	319,306
SIZE (total sales [mio €])	556	2,407	27	115	484	11,061
SIZE (employees)	596	8,981	140	578	2,772	38,543
SIZE (total assets [mio €])	569	7,499	40	131	548	65,853
SPREAD	597	0.019	0.006	0.010	0.018	0.032
VOLATILITY	599	0.144	0.077	0.113	0.171	0.115

Table A-2: Correlation Matrix

	(01)	(02)	(03)	(04)	(05)	(06)	(07)	(08)	(09)	(10)	(11)	(12)	(13)
(01) DISCLOSURE		0.24	0.67	0.28	-0.24	0.20	0.69	0.41	0.67	0.66	0.67	-0.62	-0.21
(02) ABNORMAL_RETURN	0.22		0.35	0.12	0.14	0.02	0.34	-0.03	0.47	0.46	0.41	-0.29	-0.17
(03) COMMON_EQUITY	0.59	0.11		0.33	-0.14	0.14	0.88	0.34	0.85	0.78	0.92	-0.77	-0.34
(04) CROSSLISTING	0.36	0.13	0.57		-0.10	0.18	0.33	0.33	0.33	0.34	0.33	-0.32	-0.14
(05) EXCHANGE_SEGMENT	-0.22	0.13	-0.07	-0.08		-0.25	-0.14	-0.39	-0.04	-0.01	-0.05	0.18	0.05
(06) FREEFLOAT	0.20	0.02	0.13	0.16	-0.29		0.09	0.48	0.14	0.10	0.15	-0.31	0.02
(07) MARKET_VALUE	0.62	0.11	0.89	0.63	-0.07	0.12		0.39	0.80	0.71	0.87	-0.83	-0.29
(08) SHARE_TURNOVER	0.52	-0.00	0.76	0.31	-0.09	0.13	0.71		0.30	0.24	0.34	-0.46	0.03
(09) SIZE (total sales)	0.58	0.16	0.86	0.61	-0.06	0.11	0.78	0.56		0.92	0.91	-0.72	-0.34
(10) SIZE (employees)	0.59	0.17	0.76	0.58	-0.08	0.12	0.74	0.58	0.91		0.82	-0.65	-0.36
(11) SIZE (total assets)	0.58	0.13	0.93	0.60	-0.06	0.11	0.83	0.66	0.95	0.83		-0.75	-0.30
(12) SPREAD	-0.53	-0.23	-0.27	-0.21	0.21	-0.31	-0.31	-0.22	-0.29	-0.29	-0.27		0.23
(13) VOLATILITY	-0.23	-0.25	-0.13	-0.09	0.01	0.06	-0.16	-0.09	-0.13	-0.13	-0.12	0.25	

The table lists in the lower left triangle Pearson correlations and in the upper-right triangle Spearman correlations.

Appendix C: Results of the Robustness Analysis

This appendix provides results of the robustness analysis described in section 5.2. Table A-3 reports the Heckman-corrected results for the spread and abnormal return regressions that might potentially be affected by a sample selection error because of the significant lambda values. The Heckman-corrected results are structurally identical to the results shown in Tables 3 and 4 and therefore provide no evidence for a sample selection bias.

Table A-4 reports results for selected alternative specifications of the share turnover regression from Table 2. The first column repeats the results of Panel A in Table 2 to allow a convenient comparison. I also examined non-logarithmic size measures and combinations of alternative measures. The results of these additional 135 regression models and the 143 rank regression models are not reported here but can be replicated using the publically available dataset. The results suggest that the share turnover increase with disclosure is robust.

Table A-3: Heckman-corrected multivariate regressions

Panel A: Multivariate regression of bid-ask spread on disclosure

	(Standardized) Coefficient	t-Statistics
DISCLOSURE	-0.0372*	(-1.91)
FREEFLOAT	-0.0649***	(-4.83)
SIZE	-0.1701***	(-9.55)
SHARE_TURNOVER	0.0273**	(2.06)
VOLATILITY	0.0157	(1.13)
EXCHANGE_SEGMENT	-2.4778***	(-14.63)
SIZE	0.6441***	(3.71)
LAMBDA	0.002***	(3.39)
N	555	

Panel B: Multivariate regression of abnormal return on disclosure

	(Standardized) Coefficient	t-Statistics
DISCLOSURE	-0.2012***	(-2.98)
SIZE	0.5840***	(8.70)
EXCHANGE_SEGMENT	-2.7695***	(-13.68)
SIZE	0.8247***	(4.29)
LAMBDA	0.005***	(3.66)
N	555	

The table reports standardized coefficients (except of LAMBDA where a non-standardized coefficient is reported) and t-statistics from linear OLS regressions using a Heckman selection model with two-step consistent estimates. Panel A reports results for the Heckman-corrected regression from Panel A of Table 3; Panel B reports results for the Heckman-corrected regression from Panel A of Table 4. The dependent variable in Panel A (SPREAD) is the annual average of daily proportionally quoted half-spreads 2006 obtained from Datastream. The dependent variable in Panel B (ABNORMAL_RETURN) is the 2006 annual average of daily abnormal stock returns against the industry average using four industry groups based on the NACE industry classification obtained from Datastream. DISCLOSURE is the factor (KMO 0.928; one eigenvalue of 8.123 above one) extracted from the ten AIMD disclosure measures derived from annual reports of 2005. FREEFLOAT is the percentage of shares not held by strategic investors derived from the Deutsche Börse Weighting files 2006. SIZE is the natural logarithm of total sales 2006 obtained from the Hoppenstedt Firmendatenbank. SHARE_TURNOVER is the 2006 turnover in shares obtained from Datastream. VOLATILITY is the standard deviation of daily prices 2006 divided by the mean price 2006 with prices being obtained from Datastream. EXCHANGE_SEGMENT distinguishes between firms listed in the Prime (coding 1) and General Standard (coding 2). All metric variables are winsorized at the 0.02 level at both extremes to adjust for potential outliers. * $p < 0.1$, ** $p < 0.05$, *** $p < 0.01$.



Table A-4: Alternative multivariate regressions of share turnover on disclosure

	Original Model	Alternative 1	Alternative 2	Alternative 3	Alternative 4	Alternative 5	Alternative 6	Alternative 7	Alternative 8
DISCLOSURE_1	0.4546***				0.4150***	0.3317***	0.4343***	0.4543***	0.4543***
DISCLOSURE_2		0.2101***							
DISCLOSURE_3			0.4457***						
DISCLOSURE_4				0.2053***					
FREEFLOAT	0.0580*	0.0774**	0.0610*	0.0760**	0.0681*	0.0738*	0.0548*	0.0599*	0.0585*
SIZE_1	0.0564	0.2213***	0.0688*	0.2281***			0.1053***	0.0561	0.0558
SIZE_2					0.0687				
SIZE_3						0.1686***			
VOLATILITY_1	0.0510	0.0480	0.0492	0.0474	0.0471	0.0619**	0.0619*		
VOLATILITY_2								0.0487	
VOLATILITY_3									0.0499
EXCHANGE_SEGMENT	0.0106	-0.0166	0.0135	-0.0169	-0.0173	-0.0337	-0.0022	0.0108	0.0111
CROSSLISTING_1	0.4414***	0.4827***	0.4435***	0.4796***	0.4671***	0.4695***		0.4420***	0.4415***
CROSSLISTING_2							0.4290***		
N	348	348	348	348	376	372	355	346	346
R ²	0.6084	0.5406	0.6094	0.5402	0.6190	0.6451	0.6150	0.6080	0.6081

The table reports standardized coefficients from linear OLS regressions with heteroscedasticity-corrected errors. The results for the original model are also reported in Panel A of Table 2. The dependent variable is the share turnover (SHARE_TURNOVER) 2006 obtained from Datastream. DISCLOSURE_1 is the factor (KMO 0.928; one eigenvalue of 8.123 above one) extracted from the ten AIMD disclosure measures derived from annual reports of 2005. The alternative measure DISCLOSURE_2 is the factor (KMO 0.935; one eigenvalue of 8.216 above one) extracted from the natural logarithm of the ten AIMD disclosure measures derived from annual reports of 2005. The alternative measure DISCLOSURE_3 is the unweighted sum of the ten AIMD disclosure measures derived from annual reports of 2005. The alternative measure DISCLOSURE_4 is the unweighted sum of the natural logarithm of the ten AIMD disclosure measures derived from annual reports of 2005. FREEFLOAT is the percentage of shares not held by strategic investors derived from the Deutsche Börse Weighting files 2006. SIZE_1 is the natural logarithm of total sales 2006 obtained from the Hoppenstedt Firmendatenbank. The alternative measure SIZE_2 is the natural logarithm of the number of group employees 2006 obtained from the Hoppenstedt Firmendatenbank. The alternative measure SIZE_3 is the natural logarithm of total assets 2006 obtained from the Hoppenstedt Firmendatenbank. VOLATILITY_1 is the standard deviation of daily prices 2006 divided by the mean price 2006 with prices being obtained from Datastream. The alternative VOLATILITY_2 is the standard deviation of monthly prices 2006 divided by the mean price 2006 with prices being obtained from Datastream. The alternative VOLATILITY_3 is the standard deviation of weekly prices 2006 divided by the mean price 2006 with prices being obtained from Datastream. EXCHANGE_SEGMENT distinguishes between firms listed in the Prime (coding 1) and General Standard (coding 2). CROSSLISTING_1 is the number of countries a firm is listed outside Germany based on the Datastream list of stock exchanges. The alternative CROSSLISTING_2 equals 1 if the firm is crosslisted in the U.S. and 0 otherwise. All metric variables are winsorized at the 0.02 level at both extremes to adjust for potential outliers. * p<0.1, ** p<0.05, *** p<0.01.



Table A-5: Alternative multivariate regressions of bid-ask spread on disclosure

	Original Model	Alternative 1	Alternative 2	Alternative 3	Alternative 4	Alternative 5	Alternative 6	Alternative 7	Alternative 8	Alternative 9	
DISCLOSURE_1	-0.1589***					-0.2971***	-0.2630**	-0.1557***	-0.1542***	-0.1578***	-0.1624***
DISCLOSURE_2		-0.1490***									
DISCLOSURE_3			-0.1420***								
DISCLOSURE_4				-0.1831***							
FREEFLOAT	-0.2219***	-0.2218***	-0.2234***	-0.2186***	-0.2206***	-0.2463***	-0.2227***	-0.2255***	-0.2464***	-0.2459***	
SIZE_1	-0.5549***	-0.5594***	-0.5661***	-0.5327***			-0.5632***	-0.5578***	-0.5028***	-0.5072***	
SIZE_2					-0.3749***						
SIZE_3						-0.4353***					
SHARE_TURNOVER	0.1181***	0.0866***	0.1128***	0.0943***	0.0990***	0.1500***	0.1210***	0.1188***	0.1170***	0.1057***	
VOLATILITY_1	0.0463	0.0490	0.0472	0.0490	0.0686	0.0738			0.0579	0.0446	
VOLATILITY_2							0.0323				
VOLATILITY_3								0.0496			
N	354	354	354	354	382	376	352	352	353	353	
R ²	0.4820	0.4836	0.4805	0.4896	0.4178	0.4241	0.4821	0.4833	0.4356	0.4468	

The table reports standardized coefficients from linear OLS regressions with heteroscedasticity-corrected errors. The results for the original model are also reported in Panel A of Table 3. The dependent variable of the original model and alternatives 1 to 7 is the annual average of daily proportionally quoted half-spreads 2006 obtained from Datastream. The dependent variable of alternative 8 is the annual average of monthly proportionally quoted half-spreads 2006 obtained from Datastream. The dependent variable of alternative 9 is the annual average of weekly proportionally quoted half-spreads 2006 obtained from Datastream. DISCLOSURE_1 is the factor (KMO 0.928; one eigenvalue of 8.123 above one) extracted from the ten AIMD disclosure measures derived from annual reports of 2005. The alternative measure DISCLOSURE_2 is the factor (KMO 0.935; one eigenvalue of 8.216 above one) extracted from the natural logarithm of the ten AIMD disclosure measures derived from annual reports of 2005. The alternative measure DISCLOSURE_3 is the unweighted sum of the ten AIMD disclosure measures derived from annual reports of 2005. The alternative measure DISCLOSURE_4 is the unweighted sum of the natural logarithm of the ten AIMD disclosure measures derived from annual reports of 2005. FREEFLOAT is the percentage of shares not held by strategic investors derived from the Deutsche Börse Weighting files 2006. SIZE_1 is the natural logarithm of total sales 2006 obtained from the Hoppenstedt Firmendatenbank. The alternative measure SIZE_2 is the natural logarithm of the number of group employees 2006 obtained from the Hoppenstedt Firmendatenbank. The alternative measure SIZE_3 is the natural logarithm of total assets 2006 obtained from the Hoppenstedt Firmendatenbank. SHARE_TURNOVER is the 2006 turnover in shares obtained from Datastream. VOLATILITY_1 is the standard deviation of daily prices 2006 divided by the mean price 2006 with prices being obtained from Datastream. The alternative VOLATILITY_2 is the standard deviation of monthly prices 2006 divided by the mean price 2006 with prices being obtained from Datastream. The alternative VOLATILITY_3 is the standard deviation of weekly prices 2006 divided by the mean price 2006 with prices being obtained from Datastream. All metric variables are winsorized at the 0.02 level at both extremes to adjust for potential outliers.
 * p<0.1, ** p<0.05, *** p<0.01.

Table A-5 reports results for selected alternative specifications of the spread regression from Table 3. Again, the first column repeats the results of Panel A in Table 3 to allow a convenient comparison. Non-logarithmic size measures and combinations of alternative measures have also been examined. In line with the share turnover regression the results of these additional 368 regression models and the 377 alternative rank regression models are not reported here. The publically available dataset allows to replicate these regressions. The results provide evidence that the spread-reducing effect of disclosure is robust against particular specifications of the model.

Table A-6 reports results for selected alternative specifications of the abnormal return regression from Table 4. Similar to the previous analyses the first column repeats the results of Panel A in Table 4 to allow a convenient comparison. Results for additional 18 regression models that consider non-logarithmic size measures and combinations of

alternative measures and 24 alternative rank regression models are not reported here. They can be easily replicated using the publically available dataset. Altogether, the results suggest that the negative association of disclosure and abnormal returns is robust against particular model specifications.

Table A-7 reports results for selected alternative specifications of the market value regression from Table 5. The first column again repeats the results of Panel A in Table 5 to allow a convenient comparison. I also considered non-logarithmic size measures and combinations of alternative measures. The results of these additional 18 regression models and the 24 alternative rank regression models are not reported as they can be easily replicated using the publically available dataset. The robustness analysis provides evidence that the positive association of market value and disclosure is robust against the particular model specification.

Table A-6: Alternative multivariate regressions of abnormal returns on disclosure

	Original Model	Alternative 1	Alternative 2	Alternative 3	Alternative 4	Alternative 5
DISCLOSURE_1	- 0.2504***				- 0.3157***	-0.0543
DISCLOSURE_2		- 0.1642***				
DISCLOSURE_3			- 0.2571***			
DISCLOSURE_4				-0.1398***		
SIZE_1	0.5772***	0.5088***	0.5780***	0.4905***		
SIZE_2					0.5388***	
SIZE_3						0.2071**
N	361	361	361	361	389	381
R ²	0.1875	0.1717	0.1906	0.1681	0.1527	0.0293

The table reports standardized coefficients from linear OLS regressions with heteroscedasticity-corrected errors. The results for the original model are also reported in Panel A of Table 4. The dependent variable (ABNORMAL_RETURN) is the 2006 annual average of daily abnormal stock returns against the industry average using four industry groups based on the NACE industry classification obtained from Datastream. DISCLOSURE_1 is the factor (KMO 0.928; one eigenvalue of 8.123 above one) extracted from the ten AIMD disclosure measures derived from annual reports of 2005. The alternative measure DISCLOSURE_2 is the factor (KMO 0.935; one eigenvalue of 8.216 above one) extracted from the natural logarithm of the ten AIMD disclosure measures derived from annual reports of 2005. The alternative measure DISCLOSURE_3 is the unweighted sum of the ten AIMD disclosure measures derived from annual reports of 2005. The alternative measure DISCLOSURE_4 is the unweighted sum of the natural logarithm of the ten AIMD disclosure measures derived from annual reports of 2005. SIZE_1 is the natural logarithm of total sales 2006 obtained from the Hoppenstedt Firmendatenbank. The alternative measure SIZE_2 is the natural logarithm of the number of group employees 2006 obtained from the Hoppenstedt Firmendatenbank. The alternative measure SIZE_3 is the natural logarithm of total assets 2006 obtained from the Hoppenstedt Firmendatenbank. All metric variables are winsorized at the 0.02 level at both extremes to adjust for potential outliers. * p<0.1, ** p<0.05, *** p<0.01.

Table A-7: Alternative multivariate regressions of market value on disclosure

	Original Model	Alternative 1	Alternative 2	Alternative 3	Alternative 4	Alternative 5
DISCLOSURE_1	0.1219***				0.1265***	0.1171***
DISCLOSURE_2		0.1062***				
DISCLOSURE_3			0.1105***			
DISCLOSURE_4				0.1136***		
COMMON_EQUITY	0.1342**	0.1365***	0.1386***	0.1402***		
SIZE_1					0.0219	
SIZE_2						0.2094***
SIZE_3	0.7241***	0.7408***	0.7302***	0.7330***	0.8198***	0.6455***
ABNORMAL_RETURN	-0.0428*	-0.0467**	-0.0426*	-0.0478**	-0.0244	-0.0191
N	345	345	345	345	371	372
R ²	0.8539	0.8532	0.8530	0.8541	0.8686	0.8730

The table reports standardized coefficients from linear OLS regression with heteroscedasticity-corrected errors. The results for the original model are also reported in Panel A of Table 5. The dependent variable is the natural logarithm of market value (MARKET_VALUE) obtained from Datastream. DISCLOSURE_1 is the factor (KMO 0.928; one eigenvalue of 8.123 above one) extracted from the ten AIMD disclosure measures derived from annual reports of 2005. The alternative measure DISCLOSURE_2 is the factor (KMO 0.935; one eigenvalue of 8.216 above one) extracted from the natural logarithm of the ten AIMD disclosure measures derived from annual reports of 2005. The alternative measure DISCLOSURE_3 is the unweighted sum of the ten AIMD disclosure measures derived from annual reports of 2005. The alternative measure DISCLOSURE_4 is the unweighted sum of the natural logarithm of the ten AIMD disclosure measures derived from annual reports of 2005. COMMON_EQUITY is the natural logarithm of the book value of total equity obtained from Hoppenstedt Firmendatenbank. SIZE_1 is the natural logarithm of total sales 2006 obtained from the Hoppenstedt Firmendatenbank. The alternative measure SIZE_2 is the natural logarithm of the number of group employees 2006 obtained from the Hoppenstedt Firmendatenbank. The alternative measure SIZE_3 is the natural logarithm of total assets 2006 obtained from the Hoppenstedt Firmendatenbank. ABNORMAL_RETURN is the 2006 annual average of daily abnormal stock returns against the industry average using four industry groups based on the NACE industry classification obtained from Datastream. All metric variables are winsorized at the 0.02 level at both extremes to adjust for potential outliers. * $p < 0.1$, ** $p < 0.05$, *** $p < 0.01$.

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