

Convergence Rates of the Strong Law for Stationary Mixing Sequences

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Summary. In this note we estimate the rate of convergence in Marcinkiewicz-Zygmund strong law, for partial sums S_n of strong stationary mixing sequences of random variables. The results improve the corresponding ones obtained by Tze Leung Lai (1977) and Christian Hipp (1979).

1. Introduction

Let $\{X_n\}_n$ be a sequence of random variables and let $S_n = \sum_{k=1}^n X_k$. Baum and

Katz (1965) estimated the rate of convergence of the strong law for partial sums S_n of i.i.d., showing that if $\alpha > 1/2$, $p\alpha > 1$ and assuming $EX_1 = 0$ if $\alpha \leq 1$ then

$$E|X_1|^p < \infty \tag{1.1}$$

is equivalent to

$$\sum_n n^{p\alpha-2} P(\max_{j \leq n} |S_j| > \varepsilon n^\alpha) < \infty. \tag{1.2}$$

Motivated by applications to sequential analysis of time series, Lai (1977) extended this theorem from i.i.d. case to other dependent cases namely for some classes of ϕ and strong mixing sequences of random variables satisfying the following additional assumption: There exists $\beta > 1$ and a positive integer m such that as $x \rightarrow \infty$

$$(C) \quad \sup_{i \geq m} P(|X_1| > x, |X_i| > x) = O(P^\beta(|X_1| > x)).$$

The purpose of this note is to prove that the equivalence of (1.1) and (1.2) holds for ϕ and ρ -mixing sequences without the additional assumption (C) and under an improved mixing rate (logarithmic).

We shall denote the L_p norm by $\|\cdot\|_p$, the greatest integer contained in x by $[x]$, the indicator of A by I_A and we shall use the Vinogradov symbol \ll instead of O . Denote $\mathfrak{F}_n^m = \sigma(X_i; n \leq i < m)$.

We shall use the following mixing coefficients:

$$\rho(n) = \sup_m \sup_{\{X \in L_2(\mathfrak{F}_m^n), Y \in L_2(\mathfrak{F}_{n+m}^n)\}} \frac{|E(X - EX)(Y - EY)|}{\|X - EX\|_2 \cdot \|Y - EY\|_2}$$

and

$$\phi(n) = \sup_m \sup_{\{A \in \mathfrak{F}_m^n, B \in \mathfrak{F}_{n+m}^n, P(A) \neq 0\}} |P(B|A) - P(B)|.$$

The sequence $(X_n; n \geq 1)$ is said to be ρ -mixing or ϕ -mixing, according to $\rho(n) \rightarrow 0$ or $\phi(n) \rightarrow 0$. Bradley (1983) has shown that the ρ -mixing sequences are equivalent with λ -mixing where

$$\lambda(n) = \sup_m \sup_{\{A \in \mathfrak{F}_m^n, B \in \mathfrak{F}_{n+m}^n, P(A) \neq 0, P(B) \neq 0\}} \frac{|P(A \cap B) - P(A)P(B)|}{[P(A)P(B)]^{1/2}}.$$

By Lemma (1.17) of [4], we have

$$\rho(n) \leq 2\phi^{1/2}(n) \tag{1.3}$$

and by [2]

$$\lambda(n) \leq \rho(n). \tag{1.4}$$

2. The Results

We shall establish the following result.

Theorem 1. Let $\{X_i\}_i$ be a strong stationary ρ -mixing sequence of random variables, $\alpha p > 1$, $\alpha > 1/2$ and assume that $EX_1 = 0$ for $\alpha \leq 1$,

if

$$\sum_i \rho^{2/k}(2^i) < \infty \quad \text{where } k = \begin{cases} 0 & \text{for } 0 < p < 1 \\ 2 & \text{for } 1 \leq p < 2 \\ [(\alpha p - 1)/(\alpha - 1/2)] + 1 & \text{for } p \geq 2 \end{cases} \tag{2.1}$$

then (1.1) \Rightarrow (1.2)

if

$$\sum_i \rho^m(i) < \infty \quad \text{where } m = \begin{cases} 2(p\alpha - 1)/(2 - p\alpha) & 1 < p\alpha < 4/3 \\ 1 & p\alpha \geq 4/3 \end{cases} \tag{2.2}$$

then (1.2) \Rightarrow (1.1).

This theorem, (1.3) and Lemma (5), (ii) of [5] imply:

Theorem 2. Let $\{X_i\}_i$ be a strict stationary ϕ -mixing sequence of random variables, $p\alpha > 1$, $\alpha > 1/2$.

Assume that $EX_1 = 0$ for $\alpha \leq 1$, and $\sum \phi^{1/k}(2^i) < \infty$,

$$\text{where } k = \begin{cases} 0 & \text{for } 0 < p < 1 \\ 2 & \text{for } 1 \leq p < 2 \\ [(\alpha p - 1)/(\alpha - 1/2)] + 1 & \text{for } p \geq 2 \end{cases}$$

then (1.1) \Leftrightarrow (1.2).

In order to prove Theorems 1 and 2 we need the following:

Lemma 1. *Suppose $\{X_i\}_i$ is a stationary ρ -mixing sequence of random variables and for some integer $k \geq 2$, $E|X_1|^k < \infty$ and $\sum \rho^{2^i/k}(2^i) < \infty$. Then, there exists a positive constant K_k , depending only on $\{\rho(n)\}_n$ and on k such that for every $n \geq 1$,*

$$\|S_n\|_k \leq K_k(n^{1/k} \|X_1\|_k + n^{1/(k-1)} \|X_1\|_{k-1} + \dots + n^{1/2} \|X_1\|_2 + n|EX_1|). \quad (2.3)$$

Proof. We shall prove this Lemma by induction on k . For $k=2$ cf. [7] Lemma (3.4) there exists a constant K_2 depending only on $\{\rho(n)\}_n$ such that

$$\|S_n\|_2 \leq K_2(n^{1/2} \|X_1\|_2 + n|EX_1|).$$

We assume (2.3) holds for any integer $l, l < k$. We shall show first that we can find two positive constants C_1 and C_2 depending only on k , such that for every $n \geq 1$ and $m \geq 1$,

$$\|S_{2n}\|_k \leq 2^{1/k}(1 + C_1 \rho^{2/k}(m))^{1/k} \|S_n\|_k + C_2 \|S_n\|_{k-1} + 2m \|X_1\|_k. \quad (2.4)$$

Denote by $\bar{S}_n = \sum_{j=n+m+1}^{2n+m} X_j$. From the equation

$$S_{2n} = S_n + \sum_{j=n+1}^{n+m} X_j + \bar{S}_n - \sum_{j=2n+1}^{2n+m} X_j$$

we find that

$$\|S_{2n}\|_k \leq \|S_n + \bar{S}_n\|_k + 2m \|X_1\|_k. \quad (2.5)$$

Obviously there exists a positive constant C , depending on k such that, by stationarity,

$$E|S_n + \bar{S}_n|^k \leq 2E|S_n|^k + C \sum_{i=1}^{k-1} E|S_n|^i |\bar{S}_n|^{k-i}.$$

Using Hölder inequality and then the definition of ρ -mixing we obtain for $i \leq k/2$

$$\begin{aligned} E|S_n|^i |\bar{S}_n|^{k-i} &\leq (E|S_n|^{k/2} |\bar{S}_n|^{k/2})^{2i/k} (E|\bar{S}_n|^k)^{1-2i/k} \\ &\leq [(E|S_n|^{k/2})^2 + \rho(m)E|S_n|^k]^{2i/k} (E|S_n|^k)^{1-2i/k} \\ &\leq \rho^{2i/k}(m)E|S_n|^k + \|S_n\|_k^{k-2i} \|S_n\|_{k-1}^{2i}. \end{aligned}$$

Therefore:

$$\begin{aligned} E|S_n + \bar{S}_n|^k &\leq 2(1 + kC \rho^{2/k}(m)) \|S_n\|_k^k + 2C \sum_{i=1}^{[k/2]} \|S_n\|_{k-1}^{2i} \|S_n\|_k^{k-2i} \\ &\leq (2^{1/k}(1 + kC \rho^{2/k}(m)))^{1/k} \|S_n\|_k + 2C \|S_n\|_{k-1}^k \end{aligned} \quad (2.6)$$

and (2.4) follows from (2.5) and (2.6). Taking now into account the induction assumption, from (2.4) we deduce

$$\begin{aligned} \|S_{2n}\|_k &\leq 2^{1/k}(1 + C_1 \rho^{2/k}(m))^{1/k} \|S_n\|_k \\ &\quad + C_2 K_{k-1} \left[\sum_{i=1}^{k-2} n^{1/(k-i)} \|X_1\|_{k-i} + n|EX_1| \right] + 2m \|X_1\|_k. \end{aligned}$$

Writing this inequality for $n = 2^{r-1}$ and $m = \lceil n^{1/(k+1)} \rceil$, and denoting $[1 + C_1 \rho^{2/k} (\lceil 2^{i/(k+1)} \rceil)]^{1/k} = a_i$ we obtain

$$\begin{aligned} \|S_{2^r}\|_k &\leq \left(\prod_{i=0}^{r-1} a_i \right) (2^{r/k} \|X_1\|_k \\ &\quad + \sum_{j=0}^{r-1} 2^{j/k} \left[C_2 K_{k-1} \left(\sum_{i=1}^{k-2} 2^{(r-j-1)/(k-i)} \|X_1\|_{k-i} \right. \right. \\ &\quad \left. \left. + 2^{r-j-1} |EX_1| \right) + 2 \times 2^{(r-j-1)/(k+1)} \|X_1\|_k \right]. \end{aligned}$$

Therefore there exists a positive constant C depending only on k and $\{\rho(n)\}_n$ such that:

$$\|S_{2^r}\|_k \leq C \left(\prod_{i=0}^{r-1} a_i \right) \left(2^{r/k} \|X_1\|_k + \sum_{i=1}^{k-2} 2^{r/(k-i)} \|X_1\|_{k-i} + 2^r |EX_1| \right).$$

Since $\sum_i \rho^{2/k} (2^i) < \infty$ we have $\prod_{i=1}^{\infty} a_i < \infty$.

Writing n in binary form we obtain from the preceding inequality that for every n , the relation (2.3) holds.

We also need the following variant of Theorem 5 in [6]:

Lemma 2. *Let $r \geq 1$ be a given real. Suppose that*

$$E|S_m|^r \leq m \lambda^r(m) \quad \text{for all } m \leq n$$

where $\lambda(n)$ is a nondecreasing sequence of positive numbers. Then

$$E(\max_{i \leq n} |S_i|^r) \ll n \left(\sum_{j=0}^{\lceil \log_2 n \rceil} \lambda(\lceil n/2^{j+1} \rceil) \right)^r.$$

Proof of Theorem 1.

I. We prove first that (1.1) implies (1.2). Let us denote by $b_k = P\{k \leq |X_1| < k+1\}$. We note that

$$E|X_1|^p < \infty \Leftrightarrow \sum_k k^p b_k < \infty.$$

1. We consider first the case $p \geq 1$. Without loss of generality we may assume that the random variables are centered at expectations for all $\alpha > 1/2$, because for $\alpha > 1$, and n large enough we have

$$\begin{aligned} P(\max_{i \leq n} |S_i| > \varepsilon n^\alpha) &\leq P \left(\max_{i \leq n} \left| \sum_{j=1}^i (X_j - EX_j) \right| > \varepsilon n^\alpha - nE|X_1| \right) \\ &\leq P \left(\max_{i \leq n} \left| \sum_{j=1}^i (X_j - EX_j) \right| > (\varepsilon/2)n^\alpha \right). \end{aligned}$$

Let k be an integer as in (2.1). Obviously $k > p$. For some $\beta > (1+k)/(k-p)$ let us define:

$$X_i^n(1) = X_i I_{\{|X_i| > n^\alpha\}} - EX_i I_{\{|X_i| > n^\alpha\}}$$

$$X_i^n(2) = X_i I_{\{|X_i| \leq n^\alpha / (\log_2 n)^\beta\}} - EX_i I_{\{|X_i| \leq n^\alpha / (\log_2 n)^\beta\}}$$

and

$$X_i^n(3) = X_i I_{\{n^\alpha / (\log_2 n)^\beta < |X_i| \leq n^\alpha\}} - EX_i I_{\{n^\alpha / (\log_2 n)^\beta < |X_i| \leq n^\alpha\}}.$$

Let us put $S_m^n(j) = \sum_{i=1}^m X_i^n(j)$ for $j = 1, 2, 3$.
 We note that

$$\sum_n n^{p\alpha-2} P(\max_{i \leq n} |S_i^n| > \varepsilon n^\alpha) \leq \sum_{j=1}^3 \sum_n n^{p\alpha-2} P(\max_{i \leq n} |S_i^n(j)| > (\varepsilon/3)n^\alpha).$$

i) We prove first that for every $\varepsilon > 0$, $\sum_n n^{p\alpha-2} P(\max_{i \leq n} |S_i^n(1)| > \varepsilon n^\alpha) = I < \infty$.

Indeed we have successively

$$\begin{aligned} I &\leq \varepsilon^{-1} \sum_n n^{p\alpha-\alpha-1} E|X_1^n(1)| \leq \varepsilon^{-1} \sum_n n^{p\alpha-\alpha-1} \sum_{k \geq n^{\alpha-1}} (k+1)b_k \\ &\leq \varepsilon^{-1} \sum_k (k+1)b_k \sum_{n \leq (k+1)^{1/\alpha}} n^{p\alpha-\alpha-1} \ll \sum_k (k+1)^p b_k < \infty. \end{aligned}$$

ii) We prove now that $\sum_n n^{p\alpha-2} P(\max_{i \leq n} |S_i^n(2)| > \varepsilon n^\alpha) = II < \infty$, for every $\varepsilon > 0$.

Taking into account that the random variables $X_i^n(2)$ are centered, by Lemma 1 we have for every $m \leq n$,

$$E|S_m^n(2)|^k \leq \left(K_k \sum_{i=0}^{k-2} m^{1/(k-i)} \|X_1^n(2)\|_{k-i} \right)^k.$$

Using Lemma 2 with $\lambda(m) = K_k \sum_{i=0}^{k-2} m^{i/k(k-i)} \|X_1^n(2)\|_{k-i}$ we obtain that:

$$\begin{aligned} E(\max_{i \leq n} |S_i^n(2)|^k) &\leq n \left\{ \sum_{j=0}^{[\log_2 n]} \left(K_k \sum_{i=0}^{k-2} [n/2^{j+1}]^{i/k(k-i)} \|X_1^n(2)\|_{k-i} \right)^k \right\} \\ &\ll n(\log_2 n)^k \|X_1^n(2)\|_k^k + \sum_{i=1}^{k-2} n^{k/(k-i)} \|X_1^n(2)\|_{k-i}^k. \end{aligned}$$

Therefore

$$\begin{aligned} II &\ll \sum_n n^{p\alpha-2-\alpha k} E(\max_{i \leq n} |S_i^n(2)|^k) \\ &\ll \sum_n n^{\alpha(p-k)-2} \left\{ n(\log_2 n)^k \|X_1^n(2)\|_k^k + \sum_{i=1}^{k-[p]-1} n^{k/(k-i)} \|X_1^n(2)\|_{k-i}^k \right. \\ &\quad \left. + \sum_{i=k-[p]}^{k-2} n^{k/(k-i)} \|X_1^n(2)\|_p^k \right\} = A + B + C. \end{aligned}$$

By the definition of $X_i^n(2)$ we have

$$A \ll \sum_n n^{\alpha(p-k)-1} (\log_2 n)^k [n^\alpha / (\log_2 n)^\beta]^{k-p} \ll \sum_n n^{-1} (\log_2 n)^{k-\beta(k-p)}$$

which converges for the chosen value of β .

The series obtained for $1 \leq i \leq k-2$ appear only for $p \geq 2$.

For $1 \leq i \leq k - [p] - 1$ we have

$$\begin{aligned}
 B &\ll \sum_n \sum_{i=1}^{k-[p]-1} n^{\alpha(p-k)-2+k/(k-i)+\alpha k(k-i-p)/(k-i)} \\
 &\ll \sum_n \sum_{i=1}^{k-[p]-1} n^{\alpha p-2+k(1-p\alpha)/(k-i)} \ll \sum_n n^{-1-(p\alpha-1)(k/(k-1)-1)},
 \end{aligned}$$

which converge because $p\alpha > 1$.

For the series obtained for $k - [p] \leq i \leq k-2$ we have

$$C \ll \sum_n n^{\alpha(p-k)-2+k/2}$$

which converge by the definition of k .

iii) We prove that $\sum_n n^{p\alpha-2} P(\max_{i \leq n} |S_i^n(3)| > \varepsilon n^\alpha) = III < \infty$.

By Lemma 1 we have

$$\begin{aligned}
 III &\ll \sum_n n^{p\alpha-2} P\left(\sum_{i=1}^n |X_i^n(3)| > \varepsilon n^\alpha\right) \ll \sum_n n^{p\alpha-2-k\alpha} \sum_{i=0}^{k-1} n^{k/(k-i)} \\
 &\quad \times (E|X_1^n(3)|^{k-i})^{k/(k-i)}.
 \end{aligned}$$

For $i=0$ we have successively

$$\begin{aligned}
 \sum_n n^{p\alpha-1-k\alpha} E|X_1^n(3)|^k &\ll \sum_n n^{p\alpha-1-k\alpha} \sum_{j \leq n^\alpha} j^k b_j \\
 &\ll \sum_j j^k b_j \sum_{n \geq j^{1/\alpha}} n^{p\alpha-1-k\alpha} \ll \sum_j j^p b_j < \infty.
 \end{aligned}$$

The proof of the fact that the series obtained for $1 \leq i \leq k-2$ converge is similar with the proof of the convergence of the series A and B which appear at the point ii) of this proof. For $i=k-1$ we have

$$\begin{aligned}
 \sum_n n^{p\alpha-2-k\alpha} n^k (E|X_1^n(3)|)^k &\ll \sum_n n^{p\alpha-2-k\alpha+k} ((\log_2 n)^\beta / n^\alpha)^{k(p-1)} \\
 &\ll \sum_n n^{-1-(k-1)(\alpha p-1)} (\log_2 n)^{\beta k(p-1)} < \infty.
 \end{aligned}$$

2. We consider now the case $p < 1$. We have

$$S_m = \sum_{i=1}^m X_i I_{\{|X_i| \leq n^\alpha\}} + \sum_{i=1}^m X_i I_{\{|X_i| > n^\alpha\}} = S_m^n + \bar{S}_m^n.$$

We have

$$\sum_n n^{p\alpha-2} P(\max_{i \leq n} |S_i^n| > \varepsilon n^\alpha) \leq \varepsilon^{-1} \sum_n n^{p\alpha-\alpha-1} \sum_{k \leq n^\alpha} (k+1) b_k \ll \sum_k (k+1)^p b_k < \infty.$$

We also have

$$\sum_n n^{p\alpha-2} P(\max_{i \leq n} |\bar{S}_i^n| > \varepsilon n^\alpha) \leq \varepsilon^{-p/2} \sum_n n^{p\alpha/2-1} \sum_{k \geq n^{\alpha-1}} (k+1)^{p/2} b_k \ll \sum_k (k+1)^p b_k < \infty.$$

We note that for $0 < p < 1$, (1.1) \Rightarrow (1.2) was proved without mixing assumptions.

II. We prove now that (1.2) implies (1.1). This proof is inspired from Lemma (5) of [5]. First we show that

$$nP(|X_1| > \varepsilon n^\alpha) \rightarrow 0. \tag{2.7}$$

By (1.2) we have

$$\sum_n n^{p\alpha-2} P(\max_{j \leq n} |X_j| > \varepsilon n^\alpha) < \infty$$

for every $\varepsilon > 0$.

Then, as $n \rightarrow \infty$, we have

$$n^{p\alpha-1} P(\max_{j \leq n} |X_j| > \varepsilon n^\alpha) = O\left(\sum_{k=n}^{2n} k^{p\alpha-2} P(\max_{j \leq k} |X_j| > \varepsilon(k/2)^\alpha)\right) \rightarrow 0. \tag{2.8}$$

If $p\alpha \geq 2$, (2.8) implies (2.7). If $1 < p\alpha < 2$ we put $q = n^{p\alpha-1}$. Then

$$\begin{aligned} P(\max_{j \leq n} |X_j| > \varepsilon n^\alpha) &\geq P\left(\bigcup_{i=1}^{[n/q]} |X_{i_q}| > \varepsilon n^\alpha\right) \\ &= \sum_{i=1}^{[n/q]} P(\max_{j \leq i} |X_{j_q}| \leq \varepsilon n^\alpha, |X_{i_q}| > \varepsilon n^\alpha) \\ &\geq \sum_{i=1}^{[n/q]} \{P(\max_{j \leq i} |X_{j_q}| \leq \varepsilon n^\alpha) P(|X_1| > \varepsilon n^\alpha) - \rho(q) P^{1/2}(\max_{j \leq i} |X_{j_q}| > \varepsilon n^\alpha) \\ &\quad \times P^{1/2}(|X_1| > \varepsilon n^\alpha)\}. \end{aligned}$$

The last relation follows from (1.4), taking into account that

$$P(A \cap B) - P(A)P(B) = P(A \cap CB) - P(A)P(CB).$$

Obviously $P(\max_{j \leq i} |X_{j_q}| > \varepsilon n^\alpha) \leq iP(|X_1| > \varepsilon n^\alpha)$. Therefore

$$P(\max_{j \leq n} |X_j| > \varepsilon n^\alpha) \geq [n/q] P(|X_1| > \varepsilon n^\alpha) \{P(\max_{1 \leq j \leq n} |X_j| \leq \varepsilon n^\alpha) - [n/q]^{1/2} \rho(q)\}.$$

Because by (2.8) $P(\max_{1 \leq j \leq n} |X_j| \leq \varepsilon n^\alpha) \rightarrow 1$ and by (2.2) $[n/q]^{1/2} \rho(q) \rightarrow 0$ from this last inequality we deduce (2.7).

By condition (2.2), we also deduce that we can choose an integer r such that

$$\sum_i \rho(ri) < 1. \tag{2.9}$$

By (1.4) we have

$$\begin{aligned} P(\max_{j \leq n} |X_j| > \varepsilon n^\alpha) &\geq [n/r] P(|X_1| > \varepsilon n^\alpha) - \sum_{1 \leq j < i \leq [n/r]} P(|X_{ri}| > \varepsilon n^\alpha, |X_{rj}| > \varepsilon n^\alpha) \\ &\geq [n/r] P(|X_1| > \varepsilon n^\alpha) - [n/r]^2 P^2(|X_1| > \varepsilon n^\alpha) - [n/r] P(|X_1| > \varepsilon n^\alpha) \times \sum_{i=1}^{[n/r]} \rho(ri) \end{aligned}$$

whence by (2.7) and (2.9) we obtain the existence of a constant K such that for n sufficiently large

$$[n/r] P(|X_1| > \varepsilon n^\alpha) \leq K P(\max_{j \leq n} |X_j| > \varepsilon n^\alpha).$$

Therefore by (1.2) $\sum_n n^{p\alpha-1} P(|X_1| > \varepsilon n^\alpha) < \infty$, which implies (1.1).

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