

Extreme Value Theory for Continuous Parameter Stationary Processes*

M.R. Leadbetter¹ and Holger Rootzén²

¹ University of North Carolina, Dept. of Statistics, Chapel Hill, N.C. 27514, USA

² University of Copenhagen, Institute of Mathematical Statistics, Universitetsparken 5,
DK-2100 Copenhagen Q, Denmark

Summary. In this paper the central distributional results of classical extreme value theory are obtained, under appropriate dependence restrictions, for maxima of continuous parameter stochastic processes. In particular we prove the basic result (here called Gnedenko's Theorem) concerning the existence of just three types of non-degenerate limiting distributions in such cases, and give necessary and sufficient conditions for each to apply. The development relies, in part, on the corresponding known theory for stationary sequences.

The general theory given does not require finiteness of the number of upcrossings of any level x . However when the number per unit time is a.s. finite and has a finite mean $\mu(x)$, it is found that the classical criteria for domains of attraction apply when $\mu(x)$ is used in lieu of the tail of the marginal distribution function. The theory is specialized to this case and applied to give the general known results for stationary normal processes for which $\mu(x)$ may or may not be finite).

A general Poisson convergence theorem is given for high level upcrossings, together with its implications for the asymptotic distributions of r^{th} largest local maxima.

1. Introduction

In this paper we shall be concerned primarily with asymptotic distributional properties of the maximum

$$M(T) = \sup \{ \xi(t) : 0 \leq t \leq T \}$$

of a continuous parameter stationary process $\{ \xi(t) : t \geq 0 \}$. (We write also $M(I)$ to denote the supremum in an interval or set I .) A great deal is known about such properties in the important special case when the process is normal (cf.

* This work was supported by the Office of Naval Research under Contract N00014-75-C-0809, and in part by the Danish natural Science research Council

[2, 16]). Our purpose here is to delineate the types of limiting behavior which are possible when the process is not necessarily normal, obtaining, in particular, versions of the central results of classical extreme value theory which apply in this context.

The classical theory is concerned with properties of the maximum $M_n = \max\{\xi_1, \xi_2, \dots, \xi_n\}$ of n i.i.d. random variables as n becomes large. Central to the theory is the result which asserts that if M_n has a non-degenerate limiting distribution (under linear normalizations), i.e. if $P\{a_n(M_n - b_n) \leq x\} \rightarrow G(x)$ for sequences $\{a_n > 0\}$, $\{b_n\}$, then G must be one of only three general types:

$$\begin{array}{lll} \text{Type I} & G(x) = \exp(-e^{-x}) & -\infty < x < \infty \\ \text{Type II} & G(x) = \exp(-x^{-\alpha}) & x > 0 \quad \alpha > 0 \\ \text{Type III} & G(x) = \exp(-(-x)^\alpha) & x < 0 \end{array}$$

(linear transformations of the variable x being permitted). This result, which arose from work of Fréchet [5] and Fisher and Tippett [4], was later given a complete form by Gnedenko [6] and is here referred to as “Gnedenko’s Theorem.”

Gnedenko also obtained necessary and sufficient conditions for the domains of attraction for each of the three limiting types. These and other versions obtained subsequently (cf. [7]) concern the rate of decay of the tail $1 - F(x)$ of the distribution F of each ξ_n as x increases.

A further result – trivially proved in the classical case – is that for any sequence $\{u_n\}$, $\tau > 0$, $P\{M_n \leq u_n\} \rightarrow e^{-\tau}$ if and only if $1 - F(u_n) \sim \tau/n$. This is sometimes useful in calculation of the constants a_n , b_n in Gnedenko’s Theorem (when $u_n = x/a_n + b_n$).

In more recent years there has been considerable interest in extending these and other results of the classical theory to apply to stationary sequences which exhibit a “decay of dependence” which is not too slow. In particular the early work of Watson [17] concerning convergence of $P\{M_n \leq u_n\}$ applied under m -dependence, Loynes [14] proved Gnedenko’s Theorem under strong mixing assumptions, and Berman [1] obtained detailed results for normal sequences under a mild condition involving correlation decay. More recently we have obtained a theory (cf. [9]) involving weak “distributional mixing” conditions, which unifies these results and provides a rather satisfying extension of the classical distributional theory to include stationary sequences.

It is not too surprising that such an extension is possible for stationary sequences, at least under suitable dependence restrictions. What may seem surprising at first sight is that a corresponding theory is possible for continuous parameter stationary processes. However this becomes intuitively clear by recognizing that the maximum up to time n , say, is just the maximum of n random variables – the “submaxima” in the fixed intervals $(i-1, i)$, $1 \leq i \leq n$. Our procedure will be, in fact, to use the existing theory for stationary sequences by means of (a slightly modified version of) this precise approach. The sequence results which will be needed are stated in Sect. 2.

In Sect. 3 we will obtain Gnedenko’s Theorem for continuous parameter stationary processes, showing under appropriate conditions that if

$$P\{a_T(M(T)-b_T)\leq x\} \rightarrow G(x) \quad \text{as } T \rightarrow \infty$$

for some constants $a_T > 0$, b_T , then G must be one of the extreme value forms.

In Sect. 4 we obtain a related result – again extending a classical theorem – to give necessary and sufficient conditions for the convergence of $P\{M(T)\leq u_T\}$ for sequences not necessarily of the form $u_T = x/a_T + b_T$ implicit in Gnedenko’s Theorem.

As a corollary of this result we obtain necessary and sufficient criteria for the domains of attraction occurring in Gnedenko’s Theorem. In the classical i.i.d. sequence case, the criteria for domains of attraction involve the rate of decay of the marginal distribution $1-F(x)$ as x increases. For the present case the very same criteria apply, provided $1-F(x)$ is replaced by another function $\psi(x)$. For processes whose mean number $\mu(x)$ of upcrossings of any level x is finite, the function $\psi(x)$ is precisely $\mu(x)$, a readily calculated quantity.

The general theory will not require that the mean number of upcrossings of a level per unit time be finite, and in fact will include the class of stationary Gaussian processes with covariances of the form $r(\tau) = 1 - C|\tau|^\alpha + o(|\tau|^\alpha)$ as $\tau \rightarrow 0$ for $0 < \alpha < 2$. In Sect. 5 we consider such processes, as well as (possibly non-Gaussian) cases for which the mean number of upcrossings per unit time is finite. Finally in Sect. 6 we note the general Poisson limit for the point processes of upcrossings of increasingly high levels and its implications regarding limit theorems for the distribution of the r^{th} largest local maximum of $\zeta(t)$ in $0 \leq t \leq T$.

2. Two Results for Stationary Sequences

As noted, our development of extremal theory for stationary processes will rely in part on the existing sequence theory. Specifically we shall require the following definitions and results (which may be found e.g. in [10]).

Let $\{\xi_n\}$ be a stationary sequence and write $F_{i_1 \dots i_n}(x_1 \dots x_n)$ for the joint distribution function of $\xi_{i_1} \dots \xi_{i_n}$. For brevity write also $F_{i_1 \dots i_n}(u)$ to denote $F_{i_1 \dots i_n}(u, u \dots u) = P\{\xi_{i_1} \leq u \dots \xi_{i_n} \leq u\}$. If $\{u_n\}$ is a sequence of real constants, we say that *the sequence $\{\xi_n\}$ satisfies the (dependence) condition $D(u_n)$* if for each n , $1 \leq i_1 < i_2 \dots < i_p < j_1 \dots < j_{p'} \leq n$, $j_1 - i_p \geq l$,

$$|F_{i_1 \dots i_p j_1 \dots j_{p'}}(u_n) - F_{i_1 \dots i_p}(u_n) F_{j_1 \dots j_{p'}}(u_n)| \leq \alpha_{n,l} \tag{2.1}$$

where

$$\alpha_{n,l,n} \rightarrow 0 \quad \text{for some sequence } l_n = o(n), \text{ as } n \rightarrow \infty. \tag{2.2}$$

Note that $\alpha_{n,l}$ can (and will) be taken to be decreasing in l for each n by simply replacing it by the smallest value it can take to make (2.1) hold (i.e. the maximum value of the left-hand side of (2.1) over all allowable sets of integers $i_1 \dots i_p, j_1 \dots j_{p'}$). Note also that (2.2) may then be shown equivalent to the condition (cf. [12] for proof)

$$\alpha_{n,[n\lambda]} \rightarrow 0 \quad \text{as } n \rightarrow \infty \text{ for each } \lambda > 0 \tag{2.3}$$

The condition $D(u_n)$ indicates a degree of “approximate independence” of members of the sequence separated by increasing distances. However this condition, which we refer to as “distributional mixing,” is clearly potentially far less restrictive than, for example, “strong mixing”. In the case of normal sequences, it is in fact satisfied when the covariance sequence $\{r_n\}$ tends to zero even just fast enough so that $r_n \log n \rightarrow 0$.

The following result is basic to the sequence theory and will be required in later sections.

Lemma 2.1. *Let $\{\xi_n\}$ be a stationary sequence satisfying $D\{u_n\}$ for a given sequence $\{u_n\}$ of constants and write $M_n = \max(\xi_1, \xi_2 \dots \xi_n)$. Then for any integer $k \geq 1$ (writing $[\]$ to denote integer part),*

$$P\{M_n \leq u_n\} - P^k\{M_{[n/k]} \leq u_n\} \rightarrow 0 \quad \text{as } n \rightarrow \infty.$$

This lemma indicates a degree of independence between the submaxima when the first n integers are divided into k groups. We shall also need the sequence form of Gnedenko’s Theorem, which is given (e.g. in [10]) as follows:

Theorem 2.2. *Let $\{\xi_n\}$ be a stationary sequence such that $M_n = \max(\xi_1, \xi_2 \dots \xi_n)$ satisfies $P\{a_n(M_n - b_n) \leq x\} \rightarrow G(x)$ as $n \rightarrow \infty$ for some non-degenerate d.f. G and constants $\{a_n > 0\}$, $\{b_n\}$. Suppose that $D(u_n)$ holds for all u_n of the form $x/a_n + b_n$, $-\infty < x < \infty$. Then G is one of the three extreme value distributional types.*

The other classical result quoted – concerning convergence of $P\{M_n \leq u_n\}$ for arbitrary sequences $\{u_n\}$ – is also important and holds under appropriate conditions for stationary sequences $\{\xi_n\}$. This will not be discussed here since the corresponding continuous parameter result will be independently derived.

3. Gnedenko’s Theorem for Stationary Processes

As indicated above, it will be convenient to relate the maximum $M(T)$ of the continuous parameter stationary process $\xi(t)$ to the maximum of n terms of a sequence of “submaxima.” Specifically if $h > 0$ we write

$$\zeta_i = \sup\{\xi(t) : (i-1)h \leq t \leq ih\} \quad (3.1)$$

so that for $n=1, 2, 3, \dots$,

$$M(nh) = \max(\zeta_1, \zeta_2, \dots, \zeta_n). \quad (3.2)$$

The following preliminary form of Gnedenko’s Theorem (involving conditions on the ζ -sequence) is immediate.

Theorem 3.1. *Suppose that for some families of constants $\{a_T > 0\}$, $\{b_T\}$ we have*

$$P\{a_T(M(T) - b_T) \leq x\} \rightarrow G(x) \quad \text{as } T \rightarrow \infty \quad (3.3)$$

for some non-degenerate G , and that the $\{\zeta_i\}$ sequence defined by (3.1) satisfies $D(u_n)$ whenever $u_n = x/a_{nh} + b_{nh}$ for some fixed $h > 0$ and all real x . Then G is one of the three extreme value types.

Proof. Since (3.3) holds in particular as $T \rightarrow \infty$ through values nh and the ζ_n -sequence is clearly stationary, the result follows by replacing ξ_n by ζ_n in Theorem 2.2 and using (3.2). \square

Corollary 3.2. *The result holds in particular if the $D(u_n)$ conditions are replaced by the assumption that $\{\xi(t)\}$ is strongly mixing. For then the sequence $\{\zeta_n\}$ is strongly mixing and hence satisfies $D(u_n)$. \square*

We now introduce the continuous analog of the condition $D(u_n)$, stated in terms of the finite dimensional distribution functions F_{t_1, \dots, t_n} of $\xi(t)$ (again writing $F_{t_1, \dots, t_n}(u)$ for $F_{t_1, \dots, t_n}(u \dots u)$.)

The condition $D_c(u_T)$ will be said to hold for the process $\xi(t)$ and the family of constants $\{u_T: T > 0\}$, with respect to a family $\{q_T \rightarrow 0\}$, if for any points $s_1 < s_2 \dots < s_p < t_1 \dots < t_p$, belonging to $(kq_T: 0 \leq kq_T \leq T)$ and satisfying $t_1 - s_p \geq \tau$, we have

$$|F_{s_1 \dots s_p t_1 \dots t_p}(u_T) - F_{s_1 \dots s_p}(u_T) F_{t_1 \dots t_p}(u_T)| \leq \alpha_{T, \tau} \quad (3.4)$$

where $\alpha_{T, \gamma_T} \rightarrow 0$ for some family $\gamma_T = o(T)$ or, equivalently, where

$$\alpha_{T, \lambda T} \rightarrow 0 \quad \text{as } T \rightarrow \infty \quad (3.5)$$

for each $\lambda > 0$. By replacing $\alpha_{T, \tau}$ by the maximum of the left-hand side of (3.4) over all permitted choices of s_i and t_j , we may (and do) take $\alpha_{T, \tau}$ decreasing in τ for fixed T .

The $D(u_n)$ condition for $\{\zeta_n\}$ required in Theorem 3.1 will now be related to $D_c(u_T)$ by approximating crossings and extremes of the continuous parameter process, by corresponding quantities for a sampled version. To achieve the approximation we require two conditions involving the maximum of $\xi(t)$ in fixed and in very small time intervals. These conditions are given here in a form which applies very generally – readily verifiable sufficient conditions for important cases are given in Sect. 5.

Specifically we suppose that there is a function $\psi(u)$ such that, for some $h_0 > 0$, $0 < h \leq h_0$,

$$\limsup_{u \rightarrow \infty} \frac{P\{M(h) > u\}}{h\psi(u)} \leq 1, \quad (3.6)$$

and that for each $a > 0$, there is a family of constants $q = q_a(u) \rightarrow 0$ as $u \rightarrow \infty$ such that for any fixed $h > 0$,

$$\limsup_{u \rightarrow \infty} P\{M(h) > u, \xi(jq) \leq u, 0 \leq jq \leq h\} / \psi(u) \rightarrow 0 \quad \text{as } a \rightarrow 0. \quad (3.7)$$

A condition which is sometimes more readily verified, and which, together with (3.6), implies (3.7) (see Lemma 3.3), is

$$\limsup_{u \rightarrow \infty} \frac{P\{\xi(0) < u, \xi(q) < u, M(q) > u\}}{q\psi(u)} \rightarrow 0 \quad \text{as } a \rightarrow 0. \quad (3.8)$$

Note that Eq. (3.6) specifies an asymptotic upper bound for the tail distribution of the maximum in a fixed interval, whereas (3.8) limits the probability

that the maximum in a short interval exceeds u , but the process itself is less than u at both endpoints. The following result now enables us to approximate the maximum in an interval of length h by the maximum at discrete points in that interval.

Lemma 3.3. (i) *If (3.6) holds, then $P\{M(q) > u\} = o(\psi(u))$ as $u \rightarrow \infty$ for any $q = q(u) \rightarrow 0$. Also $P\{\xi(0) > u\} = o(\psi(u))$.*

(ii) *(3.6) and (3.8) imply (3.7). If (3.6) and (3.7) hold, and I is an interval of length h , then there are constants λ_a such that*

$$0 \leq \limsup_{u \rightarrow \infty} [P\{\xi(jq) \leq u, jq \in I\} - P\{M(I) \leq u\}] / \psi(u) \leq \lambda_a \rightarrow 0 \quad \text{as } a \rightarrow 0, \quad (3.9)$$

where $q = q_a(u)$ is as in (3.7), the convergence being uniform in all intervals of this fixed length h .

Proof. If (3.6) holds and $q \rightarrow 0$ as $u \rightarrow \infty$, then for any fixed $h > 0$, q is eventually smaller than h and $P\{M(q) > u\} \leq P\{M(h) > u\}$, so that

$$\limsup_{u \rightarrow \infty} P\{M(q) > u\} / \psi(u) \leq \limsup_{u \rightarrow \infty} P\{M(h) > u\} / \psi(u) \leq h$$

by (3.6), from which it follows that $P\{M(q) > u\} / \psi(u) \rightarrow 0$, as stated. The remaining statement of (i) also follows since $P\{\xi(0) > u\} \leq P\{M(q) > u\}$.

To prove (ii), note that there are at most $[h/q]$ intervals $[(j-1)q, jq]$ in $[0, h]$, with perhaps a smaller interval remaining so that

$$\begin{aligned} & P\{M(h) > u, \xi(jq) \leq u, 0 \leq jq \leq h\} \\ & \leq \frac{h}{q} P\{\xi(0) < u, \xi(q) < u, M(q) > u\} + P\{M(q) > u\}, \end{aligned}$$

so that (3.7) easily follows from (3.6) and (3.8) (using (i)).

Finally, let (3.6) and (3.7) hold. It is readily shown, using stationarity and the fact that the number of points jq in I and in $[0, h]$ differ by at most 2, that the (non-negative) difference in probabilities in (3.9) does not exceed

$$\lambda_{a,u} = P\{M(h) > u, \xi(jq) \leq h, 0 \leq jq \leq h\} + 2P\{\xi(0) > u\},$$

from which (3.9) follows by (3.7), (3.6), and (i) on writing

$$\lambda_a = \limsup_u \lambda_{a,u} / \psi(u). \quad \square$$

It is now relatively straightforward to relate $D(u_n)$ for the sequence $\{\xi_n\}$ to the condition $D_c(u_T)$ for the process $\xi(t)$, as the following lemma shows. For later use, a slightly more general result will be proved than needed here. In this, for $h > 0$, $\{T_n\}$ will denote any sequence of time points such that $T_n \in [nh, (n+1)h]$ and $v_n = u_{T_n}$.

Lemma 3.4. *Suppose that (3.6) holds with some function $\psi(u)$ and let $\{q_a(u)\}$ be a family of constants for each $a > 0$ with $q_a(u) > 0$, $q_a(u) \rightarrow 0$ as $u \rightarrow \infty$, and such that (3.7) holds. If $D_c(u_T)$ is satisfied with respect to the family $q_T = q_a(u_T)$ for each $a > 0$, and $T\psi(u_T)$ is bounded, then the sequence $\{\xi_n\}$ defined by (3.1) satisfies $D(v_n)$, where $v_n = u_{T_n}$ is as above.*

Proof. For a given n , let $i_1 < i_2 \dots < i_p < j_1 \dots < j_{p'} < n$, $j_1 - i_p \geq l$. Write $I_r = [(i_r - 1)h, i_r h]$, $J_s = [(j_s - 1)h, j_s h]$. For brevity let q denote one of the families $\{q_a(\cdot)\}$ and

$$\begin{aligned} A_q &= \bigcap_{r=1}^p \{\xi(jq) \leq v_n, jq \in I_r\}, & A &= \bigcap_{r=1}^p \{\zeta_{i_r} \leq v_n\} \\ B_q &= \bigcap_{s=1}^{p'} \{\xi(jq) \leq v_n, jq \in J_s\}, & B &= \bigcap_{s=1}^{p'} \{\zeta_{j_s} \leq v_n\}. \end{aligned}$$

It follows in an obvious way from Lemma 3.3 that

$$\begin{aligned} 0 \leq \limsup_{n \rightarrow \infty} \{P(A_q \cap B_q) - P(A \cap B)\} &\leq \limsup_{n \rightarrow \infty} (p + p') \psi(v_n) \lambda_a \\ &\leq \limsup n \psi(v_n) \lambda_a \leq K \lambda_a \end{aligned}$$

for some constant K (since $nh \sim T_n$, and $T_n \psi(v_n)$ is bounded), and where $\lambda_a \rightarrow 0$ as $a \rightarrow 0$. Similarly

$$\limsup |P(A_q) - P(A)| \leq K \lambda_a, \quad \limsup |P(B_q) - P(B)| \leq K \lambda_a.$$

Now

$$\begin{aligned} |P(A \cap B) - P(A)P(B)| &\leq |P(A \cap B) - P(A_q \cap B_q)| + |P(A_q \cap B_q) - P(A_q)P(B_q)| \\ &\quad + P(A_q)|P(B_q) - P(B)| + P(B)|P(A_q) - P(A)| \\ &= R_{n,a} + |P(A_q \cap B_q) - P(A_q)P(B_q)| \end{aligned} \quad (3.10)$$

where $\limsup_{n \rightarrow \infty} R_{n,a} \leq 3K \lambda_a$.

Since the largest jq in any I_r is at most $i_p h$, and the smallest in any J_s is at least $(j_1 - 1)h$, their difference is at least $(l - 1)h$. Also the largest jq in $J_{p'}$ does not exceed $j_{p'} h \leq nh \leq T_n$ so that from (3.4) and (3.10)

$$|P(A \cap B) - P(A)P(B)| \leq R_{n,a} + \alpha_{T_n, (l-1)h}^{(a)} \quad (3.11)$$

(in which the dependence of $\alpha_{T,l}$ on a is explicitly indicated). Write now $\alpha_{n,l}^* = \inf_{a > 0} \{R_{n,a} + \alpha_{T_n, (l-1)h}^{(a)}\}$. Since the left-hand side of (3.11) does not depend on a we have

$$|P(A \cap B) - P(A)P(B)| \leq \alpha_{n,l}^*,$$

which is precisely the desired conclusion of the lemma, provided we can show that $\lim_{n \rightarrow \infty} \alpha_{n, \lambda n}^* = 0$ for any $\lambda > 0$ (cf. (2.3)). But for any $a > 0$

$$\alpha_{n, [\lambda n]}^* \leq R_{n,a} + \alpha_{T_n, ([\lambda n] - 1)h}^{(a)} \leq R_{n,a} + \alpha_{T_n, \frac{1}{2} \lambda T_n}^{(a)}$$

when n is sufficiently large (since $\alpha_{T,l}^{(a)}$ decreases in l), and hence by (3.5)

$$\limsup_{n \rightarrow \infty} \alpha_{n, [\lambda n]}^* \leq 3K \lambda_a,$$

and since a is arbitrary and $\lambda_a \rightarrow 0$ as $a \rightarrow 0$, it follows that $\alpha_{n, [\lambda n]}^* \rightarrow 0$ as desired. \square

The general continuous version of Gnedenko's Theorem is now readily restated in terms of conditions on $\xi(t)$ itself.

Theorem 3.5. *With the above notation for the stationary process $\xi(t)$ satisfying (3.6) for some function ψ , suppose that, for some families of constants $\{a_T > 0\}$, $\{b_T\}$,*

$$P\{a_T(M(T) - b_T) \leq x\} \rightarrow G(x)$$

for a non-degenerate G . Suppose that $T\psi(u_T)$ is bounded and $D_c(u_T)$ holds for $u_T = x/a_T + b_T$ for each real x , with respect to families of constants $\{q_a(u)\}$ satisfying (3.7). Then G is one of the three extreme value distributional types.

Proof. This follows at once from Theorem 3.1 and Lemma 3.4 by choosing $T_n = nh$. \square

As noted the conditions of this theorem are of a general kind, and more specific sufficient conditions will be given in the applications in Sect. 5.

4. Convergence of $P\{M(T) \leq u_T\}$

Gnedenko's Theorem involved consideration of $P\{a_T\{M(T) - b_T\} \leq x\}$, which may be rewritten as $P\{M(T) \leq u_T\}$ with $u_T = a_T^{-1}x + b_T$. We turn now to the question of convergence of $P\{M(T) \leq u_T\}$ as $T \rightarrow \infty$ for families u_T which are not necessarily linear functions of a parameter x . (This is analogous to the convergence of $P\{M_n \leq u_n\}$ for sequences, of course.) These results are of interest in their own right, but also since they make it possible to simply modify the classical criteria for domains of attraction to the three limiting distributions, to apply in this continuous parameter context.

Our main purpose is to demonstrate the equivalence of the relations $P\{M(h) > u_T\} \sim \tau/T$ and $P\{M(T) \leq u_T\} \rightarrow e^{-\tau}$ under appropriate conditions. The following condition will be referred to as $D'_c(u_T)$ and is analogous to a condition $D'(u_n)$ defined e.g. in [10] for sequences.

$D'_c(u_T)$: The condition $D'_c(u_T)$ will be said to hold for the process $\{\xi(t)\}$ and the family of constants $\{u_T; T > 0\}$, with respect to the constants $q_T \rightarrow 0$, if $\limsup_{T \rightarrow \infty} (T/q) \sum_{h < jq < \varepsilon T} P\{\xi(0) > u_T, \xi(jq) > u_T\} \rightarrow 0$ as $\varepsilon \rightarrow 0$ for some $h > 0$.

One further condition which will play a central role is the following stronger version of (3.6):

$$P\{M(h) > u\} \sim h\psi(u) \quad \text{as } u \rightarrow \infty \text{ for } 0 < h \leq h_0 \text{ and some } h_0 > 0. \quad (4.1)$$

The following lemma will be useful in obtaining the desired equivalence.

Lemma 4.1. Suppose that (4.1) holds for some function ψ , and let $\{u_T\}$ be a family of levels such that $D'_c(u_T)$ holds with respect to families $\{q_a(u)\}$ satisfying (3.7), for each $a > 0$, with h in $D'_c(u_T)$ not exceeding $h_0/2$ in (4.1). Then $T\psi(u_T)$ is bounded, and writing $n' = \lfloor n/k \rfloor$, for n and k integers,

$$0 \leq \limsup_{n \rightarrow \infty} [n' P\{M(h) > v_n\} - P\{M(n'h) > v_n\}] = o(k^{-1}), \quad \text{as } k \rightarrow \infty, \quad (4.2)$$

with $v_n = u_{T_n}$, for any sequence $\{T_n\}$ with $T_n \in [nh, (n+1)h)$.

Proof. We shall use the extra assumption

$$\liminf_{T \rightarrow \infty} T\psi(u_T) > 0, \quad (4.3)$$

in proving $T\psi(u_T)$ bounded and (4.2). It is then easily checked (e.g. by replacing $T\psi(u_T)$ by $\max(1, T\psi(u_T))$ in the proof) that the result holds also without the extra assumption.

Now, write $I_j = [(j-1)h, jh]$, $j=1, 2, \dots$ and $M_q(I) = \max\{\xi(jq); jq \in I\}$, for any interval I . We shall first show that (assuming (4.3) holds)

$$0 \leq \limsup_{n \rightarrow \infty} \frac{1}{T_n \psi(v_n)} [n' P\{M(h) > v_n\} - P\{M(n'h) > v_n\}] = o(k^{-1}) \quad (4.4)$$

as $k \rightarrow \infty$. The expression in (4.4) is clearly non-negative, and by stationarity and the fact that $M \geq M_q$, does not exceed

$$\begin{aligned} & \limsup_{n \rightarrow \infty} \frac{1}{T_n \psi(v_n)} \sum_{j=1}^{n'} [P\{M(I_j) > v_n\} - P\{M_q(I_j) > v_n\}] \\ & + \limsup_{n \rightarrow \infty} \frac{1}{T_n \psi(v_n)} \left[\sum_{j=1}^{n'} P\{M_q(I_j) > v_n\} - P\{M_q(n'h) > v_n\} \right]. \end{aligned}$$

By Lemma 3.3(ii), the first of the upper limits does not exceed $\lambda_a \limsup_{n \rightarrow \infty} n'/T_n = \lambda_a/(hk)$, where $\lambda_a \rightarrow 0$ as $a \rightarrow 0$. The expression in the second upper limit may be written as

$$\begin{aligned} & \frac{1}{T_n \psi(v_n)} \left[\sum_{j=1}^{n'} P\{M_q(I_j) > v_n\} - \sum_{j=1}^{n'} P\left\{M_q(I_j) > v_n, M_q\left(\bigcup_{l=j+1}^{n'} I_l\right) \leq v_n\right\} \right] \\ & \leq \frac{1}{T_n \psi(v_n)} \sum_{j=1}^{n'} P\{M_q(I_j) > v_n, M_q(I_{j+1}) > v_n\} \\ & + \frac{1}{T_n \psi(v_n)} \sum_{j=1}^{n'} P\left\{M_q(I_j) > v_n, M_q\left(\bigcup_{l=j+2}^{n'} I_l\right) > v_n\right\}. \end{aligned} \quad (4.5)$$

Now the application of (4.1) to $M(I_j)$, $M(I_{j+1})$, and $M(I_j \cup I_{j+1})$ leads simply to the relation $P\{M(I_j) > v_n, M(I_{j+1}) > v_n\} = o(\psi(v_n))$. Since $M \geq M_q$, the first term of (4.5) is $(n'/T_n) o(1)$. The second term is clearly dominated by

$$\frac{n'h}{q T_n \psi(v_n)} \sum_{h \leq jq \leq n'h} P\{\xi(0) > v_n, \xi(jq) > v_n\}.$$

By $D'_c(u_T)$ and (4.3), it is seen that the upper limit (over n) of this last term is $o(k^{-1})$ for each $a > 0$, and (4.4) follows by gathering these facts.

Further, by (4.4) and (4.1)

$$\begin{aligned} \liminf_{n \rightarrow \infty} \frac{1}{T_n \psi(v_n)} & \geq \liminf_{n \rightarrow \infty} \frac{1}{T_n \psi(v_n)} P\{M(n'h) > v_n\} \\ & \geq \liminf_{n \rightarrow \infty} \frac{1}{T_n \psi(v_n)} n' P\{M(h) > v_n\} \\ & \quad - \limsup_{n \rightarrow \infty} \frac{1}{T_n \psi(v_n)} [n' P\{M(h) > v_n\} - P\{M(n'h) > v_n\}] \\ & = \frac{1}{k} - o\left(\frac{1}{k}\right), \end{aligned}$$

and hence $\liminf_{n \rightarrow \infty} (T_n \psi(v_n))^{-1} > 0$. Thus $T_n \psi(u_{T_n})$ is bounded for any sequence $\{T_n\}$ satisfying $nh \leq T_n \leq (n+1)h$, which readily implies that $T\psi(u_T)$ is bounded. Finally, (4.2) then follows at once from (4.4). \square

Corollary 4.2. *Under the conditions of the lemma, if*

$$\lambda_{n,k} = |n'h\psi(v_n) - P\{M(n'h) > v_n\}|,$$

then

$$\limsup_{n \rightarrow \infty} \lambda_{n,k} = o(k^{-1}) \quad \text{as } k \rightarrow \infty.$$

Proof. Noting that $n'\psi(v_n)$ is bounded, this follows at once from the lemma by (4.1). \square

Our main result now follows readily.

Theorem 4.3. *Suppose that (4.1) holds for some function ψ , and let $\{u_T\}$ be a family of constants such that for each $a > 0$, $D_c(u_T)$ and $D'_c(u_T)$ hold with respect to the family $\{q_a(u)\}$ of constants satisfying (3.7), with h in $D'_c(u_T)$ not exceeding $h_0/2$ in (4.1). Then*

$$T\psi(u_T) \rightarrow \tau > 0 \tag{4.6}$$

if and only if

$$P\{M(T) \leq u_T\} \rightarrow e^{-\tau}. \tag{4.7}$$

Proof. If (4.1), (3.7), and $D'_c(u_T)$ hold as stated, then $T\psi(u_T)$ is bounded according to Lemma 4.1 and by Lemma 3.4 the sequence of “submaxima” $\{\zeta'_n\}$ defined by (3.1) satisfies $D(v_n)$, with $v_n = u_{T_n}$, for any sequence $\{T_n\}$ with $T_n \in [nh, (n+1)h)$. Hence from Lemma 2.1, writing $n' = [n/k]$,

$$P\{M(nh) \leq v_n\} - P^k\{M(n'h) \leq v_n\} \rightarrow 0 \quad \text{as } n \rightarrow \infty. \tag{4.8}$$

Clearly it is enough to prove that

$$T_n \psi(v_n) \rightarrow \tau > 0 \tag{4.9}$$

if and only if

$$P\{M(T_n) \leq v_n\} \rightarrow e^{-\tau}, \tag{4.10}$$

for any sequence $\{T_n\}$ with $T_n \in [nh, (n+1)h)$. Further, $T\psi(u_T)$ bounded implies that $\psi(u_T) \rightarrow 0$ as $T \rightarrow \infty$ so that

$$\begin{aligned} 0 &\leq P\{M(nh) \leq v_n\} - P\{M(T_n) \leq v_n\} \\ &\leq P\{M(h) > v_n\} \sim h\psi(v_n) \rightarrow 0, \end{aligned}$$

and thus (4.10) holds if and only if

$$P\{M(nh) \leq v_n\} \rightarrow e^{-\tau}. \tag{4.11}$$

Hence it is sufficient to prove that (4.9) and (4.11) are equivalent under the hypothesis of the theorem.

Suppose now that (4.9) holds so that in particular

$$n'h\psi(v_n) \rightarrow \tau/k \quad \text{as } n \rightarrow \infty. \quad (4.12)$$

With the notation of Corollary 4.2 we have

$$1 - n'h\psi(v_n) - \lambda_{n,k} \leq P\{M(n'h) \leq v_n\} \leq 1 - n'h\psi(v_n) + \lambda_{n,k} \quad (4.13)$$

so that, letting $n \rightarrow \infty$,

$$\begin{aligned} 1 - \tau/k - o(k^{-1}) &\leq \liminf_{n \rightarrow \infty} P\{M(n'h) \leq v_n\} \\ &\leq \limsup_{n \rightarrow \infty} P\{M(n'h) \leq v_n\} \\ &\leq 1 - \tau/k + o(k^{-1}). \end{aligned}$$

By taking k -th powers throughout and using (4.8) we obtain

$$\begin{aligned} (1 - \tau/k - o(k^{-1}))^k &\leq \liminf_{n \rightarrow \infty} P\{M(nh) \leq v_n\} \\ &\leq \limsup_{n \rightarrow \infty} P\{M(nh) \leq v_n\} \\ &\leq (1 - \tau/k + o(k^{-1}))^k, \end{aligned}$$

and letting k tend to infinity proves (4.11).

Hence (4.9) implies (4.11) under the stated conditions. We shall now show that conversely (4.11) implies (4.9). The first part of the above proof still applies so that (4.8) and the conclusion of Corollary 4.2, and hence (4.13), hold. A rearrangement of (4.13) gives

$$\begin{aligned} 1 - P\{M(n'h) \leq v_n\} - \lambda_{n,k} &\leq n'h\psi(v_n) \\ &\leq 1 - P\{M(n'h) \leq v_n\} + \lambda_{n,k}. \end{aligned}$$

But it follows from (4.8) and (4.11) that $P\{M(n'h) \leq v_n\} \rightarrow e^{-\tau/k}$ and hence, using Corollary 4.2, that

$$\begin{aligned} 1 - e^{-\tau/k} - o(k^{-1}) &\leq \liminf_{n \rightarrow \infty} n'h\psi(v_n) \\ &\leq \limsup_{n \rightarrow \infty} n'h\psi(v_n) \\ &\leq 1 - e^{-\tau/k} + o(k^{-1}). \end{aligned}$$

Multiplying through by k and letting $k \rightarrow \infty$ shows that $T_n\psi(v_n) \sim nh\psi(v_n) \rightarrow \tau$, and concludes the proof that (4.11) implies (4.9). \square

Theorem 4.3 may be related to the corresponding results for i.i.d. sequences in the following way.

Theorem 4.4. *Let $\{u_T\}$ be a family of constants such that the conditions of Theorem 4.3 hold, let $0 < \rho < 1$, and let h be chosen as in (4.1). Suppose that $\psi(u_T) \sim \psi(u_{h\lceil T/h \rceil})$ as $T \rightarrow \infty$ (which will be the case if, e.g., $u_T = u_{nh}$ for $nh \leq T < (n+1)h$). Then*

$$P\{M(T) \leq u_T\} \rightarrow \rho \quad \text{as } T \rightarrow \infty \quad (4.14)$$

if and only if there is a sequence $\{\zeta_n\}$ of i.i.d. random variables with common d.f. F satisfying $1 - F(u) \sim h\psi(u)$ as $u \rightarrow \infty$ and such that $\hat{M}_n = \max(\zeta_1, \zeta_2, \dots, \zeta_n)$ satisfies

$$P\{\hat{M}_n \leq u_{nh}\} \rightarrow \rho. \quad (4.15)$$

Proof. If there is an i.i.d. sequence $\{\zeta_n\}$ with common d.f. F such that (4.15) holds then (as noted in the introduction) we have $1 - F(u_{nh}) \sim \tau/n$, where $\rho = e^{-\tau}$. Since $1 - F(u) \sim h\psi(u)$ we have $\psi(u_{nh}) \sim \tau/nh$, from which it follows easily (using the stated assumption about ψ) that $\psi(u_T) \sim \tau/T$. Hence Theorem 4.3 gives $P\{M(T) \leq u_T\} \rightarrow e^{-\tau}$ so that (4.14) holds.

Conversely if (4.14) holds it follows from Theorem 4.3 that $T\psi(u_T) \rightarrow \tau$ and hence $nh\psi(u_{nh}) \rightarrow \tau$. Let $\{\zeta_n\}$ be i.i.d. random variables with the same d.f. F , say, as $M(h)$, so that by (4.1)

$$1 - F(u_{nh}) \sim h\psi(u_{nh}) \sim \tau/n,$$

from which it follows that $\hat{M}_n = \max(\zeta_1, \zeta_2, \dots, \zeta_n)$ satisfies $P\{\hat{M}_n \leq u_n\} \rightarrow e^{-\tau} = \rho$, as required. \square

These results show how the function ψ may be used in the classical criteria for domains of attraction to determine the asymptotic distribution of $M(T)$. We write $\mathcal{D}(G)$ for the domain of attraction to the (extreme value) d.f. G , i.e. the set of all d.f.'s F such that $F^n(x/a_n + b_n) \rightarrow G(x)$ for some sequences $\{a_n > 0\}$, $\{b_n\}$.

Theorem 4.5. *Suppose that the conditions of Theorem 4.4 hold for all families $u_T = x/a_T + b_T$, $-\infty < x < \infty$, when $\{a_T > 0\}$, $\{b_T\}$ are given constants and*

$$P\{a_T(M(T) - b_T) \leq x\} \rightarrow G(x). \quad (4.16)$$

Then

$$\psi(u) \sim 1 - F(u) \quad \text{as } u \rightarrow \infty \quad \text{for some } F \in \mathcal{D}(G). \quad (4.17)$$

Conversely if (4.1) holds and $\psi(u)$ satisfies (4.17) there are families of constants $\{a_T > 0\}$, $\{b_T\}$ such that (4.16) holds, provided that the conditions of Theorem 4.4 are satisfied for each $u_T = x/a_T + b_T$, $-\infty < x < \infty$.

Proof. If (4.16) holds together with the conditions stated, Theorem 4.4 shows that

$$P\{a_{nh}(\hat{M}_n - b_{nh}) \leq x\} \rightarrow G(x)$$

where \hat{M}_n is the maximum of n i.i.d. random variables with a common d.f. F_0 , say, and where $h\psi(u) \sim 1 - F_0(u)$ as $u \rightarrow \infty$, and $F_0 \in \mathcal{D}(G)$. We may choose a d.f. F such that $1 - F(u) = \frac{1}{h}(1 - F_0(u))$ when u is large and the classical domain of attraction criteria show that $F \in \mathcal{D}(G)$. But $\psi(u) \sim 1 - F(u)$ as desired, showing (4.17).

Conversely if (4.17) holds and $h > 0$ we may choose $F_0 \in \mathcal{D}(G)$ such that $h\psi(u) \sim 1 - F_0(u)$ and hence define an i.i.d. sequence $\{\zeta_n\}$ with common d.f. F_0 , $\hat{M}_n = \max(\zeta_1, \zeta_2, \dots, \zeta_n)$, such that

$$P\{a'_n(\hat{M}_n - b'_n) \leq x\} \rightarrow G(x)$$

for some constants $a'_n > 0$, b'_n . Define $a_T = a'_n$, $b_T = b'_n$ for $nh \leq T < (n+1)h$, $n = 0, 1, 2, \dots$. Then (4.15) holds with $\rho = G(x)$. If the conditions of Theorem 4.4 hold for each $u_T = x/a_T + b_T$ then (4.14) holds, which yields (4.16). \square

5. Particular Classes of Processes

In this section we first show how the conditions required for the previous theory may be simplified when the mean number $\mu(u)$ of upcrossings of each level u by $\xi(t)$ per unit time is finite, and then briefly indicate applications to stationary normal processes (whether or not $\mu(u) < \infty$). Throughout $N_u(I)$ ($N_u(t)$) will denote the number of upcrossings of the level u in the interval I (or in $(0, t)$ respectively).

First we write for $q > 0$

$$I_q(u) = P\{\xi(0) < u < \xi(q)\}/q. \tag{5.1}$$

Clearly $I_q(u) \leq P\{N_u(q) \geq 1\}/q \leq \mathcal{E}N_u(q)/q = \mu$. Further, it is readily shown (by a standard dissection of the unit interval into subintervals of length q) that

$$\mu(u) = \lim_{q \rightarrow 0} I_q(u), \tag{5.2}$$

which, for now, we assume finite for each u . It is apparent from (5.2) that $\mu(u)$ may, at least in principle, be readily calculated from the bivariate distributions of the process. It may also happen (as for many normal processes) that $I_q(u) \sim \mu(u)$ as $u \rightarrow \infty$ when q depends on u , $q = q(u) \rightarrow 0$. For greater flexibility we shall use the following variant of such a property. Specifically we shall assume, when needed, that for each $a > 0$ there is a family $\{q_a(u) \rightarrow 0$ as $u \rightarrow \infty\}$ such that (with $q_a = q_a(u)$, $\mu = \mu(u)$)

$$\liminf_{u \rightarrow \infty} I_{q_a}(u)/\mu \geq v_a \tag{5.3}$$

where $v_a \rightarrow 1$ as $a \rightarrow 0$. As indicated below, for many normal processes we may take $q_a(u) = a/u$ and more generally as $aP\{\xi(0) > u\}/\mu(u)$.

We shall assume as needed that

$$P\{\xi(0) > u\} = o\mu(u) \quad \text{as } u \rightarrow \infty, \tag{5.4}$$

which holds under general conditions. For example, 5.4 is readily verified if for some $q = q(u) \rightarrow 0$ as $u \rightarrow \infty$,

$$\limsup_{u \rightarrow \infty} \frac{P\{\xi(0) > u, \xi(q) > u\}}{P\{\xi(0) > u\}} < 1 \tag{5.5}$$

since (5.5) implies that $\liminf_{u \rightarrow \infty} qI_q(u)/P\{\xi(0) > u\} > 0$, from which it follows that $P\{\xi(0) > u\}/I_q(u) \rightarrow 0$, and hence (5.4) holds since $I_q(u) \leq \mu(u)$.

We may now recast the conditions (3.6) and (3.8) in terms of the function $\mu(u)$.

Lemma 5.1. (i) Suppose $\mu(u) < \infty$ for each u and that (5.4) (or the sufficient condition (5.5)) holds. Then (3.6) holds with $\psi(u) = \mu(u)$.

(ii) If (5.3) holds (for some family $\{q_a(u)\}$) then (3.8) holds with $\psi(u) = \mu(u)$.

Proof. Since clearly

$$P\{M(h) > u\} \leq P\{N_u(h) \geq 1\} + P\{\xi(0) > u\} \leq \mu h + P\{\xi(0) > u\},$$

(3.6) follows at once from (5.4), which proves (i).

Now if (5.3) holds, then with $q = q_a(u)$, $\mu = \mu(u)$,

$$\begin{aligned} P\{\xi(0) < u, \xi(q) < u, M(q) > u\} &= P\{\xi(0) < u, M(q) > u\} - P\{\xi(0) < u < \xi(q)\} \\ &\leq P\{N_u(q) \geq 1\} - qI_q(u) \\ &\leq \mu q - \mu q v_a(1 + o(1)) \end{aligned}$$

so that

$$\limsup_{u \rightarrow \infty} P\{\xi(0) < u, \xi(q) < u, M(q) > u\} / (q\mu) \leq 1 - v_a,$$

which tends to zero as $a \rightarrow 0$, giving (3.8). \square

In view of this lemma, Gnedenko's Theorem now applies to processes of this kind using the more readily verifiable conditions (5.3) and (5.4), as follows.

Theorem 5.2. Theorem 3.5 holds for a stationary process $\xi(t)$ with $\psi(u) = \mu(u) < \infty$ for each u if the conditions (3.6) and (3.7) are replaced by (5.4) and (5.3). \square

Note that while (5.3) and (5.4) are especially convenient to give (3.6) and (3.8), the verification of (4.1) still requires obtaining

$$\liminf_{u \rightarrow \infty} P\{M(h) > u\} / h\psi(u) \geq 1 \quad \text{for } 0 < h \leq h_0.$$

For stationary normal processes with $\mu = \mathcal{E}N_u(1) < \infty$, there are a number of relatively simple derivations available. However, we turn here to a brief consideration of more general normal cases, where μ can be infinite.

Specifically, assume now that $\xi(t)$ is a (zero mean) stationary normal process with covariance function

$$r(\tau) = 1 - C|\tau|^\alpha + o|\tau|^\alpha \quad \text{as } \tau \rightarrow 0 \tag{5.6}$$

for some α , $0 < \alpha \leq 2$. (The case $\alpha = 2$ gives $\mu < \infty$.) There is a considerable literature dealing with extremal properties of such processes, and of slightly more general cases (which could be included here) in which the term $|\tau|^\alpha$ is multiplied by a slowly varying function as $\tau \rightarrow 0$ (cf. [2, 16]). Of course a number of the arguments (which in some cases are rather intricate) used in these papers are required to verify our general conditions here. We will not attempt to reproduce these arguments but rather to simply indicate the basic considerations used and where they may be found. However it will be convenient to summarize these results as a theorem even though formal proofs are not given.

Theorem 5.4. *Let $\xi(t)$ be a zero mean stationary normal process with covariance function $r(t)$ satisfying (5.6). Then*

(i) (3.6), and in fact (4.1), hold with $\psi(u) = C^{1/\alpha} H_\alpha u^{2/\alpha} \phi(u)/u$, in which ϕ is the standard normal density, C is as in (5.6), and H_α is a constant depending only on α .

(ii) (3.7) holds with $q_a(u) = au^{-2/\alpha}$.

(iii) $D_c(u_T)$ holds with respect to a family $\{q\}$ if $T\psi(u_T)$ is bounded and

$$\frac{T}{q} \sum_{\lambda T \leq kq \leq T} |r(kq)| e^{-u^2/(1+|r(kq)|)} \rightarrow 0 \quad \text{as } T \rightarrow \infty \quad (5.7)$$

for each $\lambda > 0$. This holds, in particular, if $T\psi(u_T)$ is bounded (with ψ defined as in (i)) and $r(t) \log t \rightarrow 0$ as $t \rightarrow \infty$.

(iv) If $r(t) \log t \rightarrow 0$ and $T\psi(u_T) \rightarrow \tau > 0$, then $D'_c(u_T)$ holds and $P\{M(T) \leq u_T\} \rightarrow e^{-\tau}$.

(v) If $r(t) \log t \rightarrow 0$, $M(T)$ has the limiting distribution given by

$$P\{a_T(M(T) - b_T) \leq x\} \rightarrow e^{-e^{-x}}$$

where

$$\begin{aligned} a_T &= (2 \log T)^{\frac{1}{\alpha}}, \\ b_T &= (2 \log T)^{\frac{1}{\alpha}} + (2 \log T)^{-\frac{1}{\alpha}} \left\{ \left(\frac{1}{\alpha} - \frac{1}{2} \right) \log \log T \right. \\ &\quad \left. + \log(2^{\alpha-1} - 1) \pi^{-\frac{1}{\alpha}} C^{1/\alpha} H_\alpha \right\}. \end{aligned}$$

Indications and Sources of Proof

(i) A derivation of (4.1) (from which (3.6) follows) appears in several developments of the normal theory (e.g. Theorem 2.1 of [16]). In the case $\alpha=2$, (3.6) is incidentally simply obtained from ‘‘Rice’s formula’’ $\mu = [(-r''(0))^{\frac{1}{2}}/2\pi] e^{-u^2/2}$.

(ii) This may be shown, for example, along the lines of Lemma 2.4 of [16], although a more direct derivation is obtainable from the normal theory given in [13].

(iii) The proof of this involves a standard calculation using ‘‘Slepian’s Lemma’’ (cf. Lemma 3.5 of [15]), from which it follows that for two sets of standard normal random variables $\xi_1 \dots \xi_n, \eta_1 \dots \eta_n$ with covariance matrices $[\lambda_{ij}], [v_{ij}]$, $|\lambda_{ij}| \geq |v_{ij}|$

$$\left\{ P \left\{ \bigcap_{j=1}^n (\xi_j \leq u) \right\} - P \left\{ \bigcap_{j=1}^n (\eta_j \leq u) \right\} \right\} \leq K \sum_{i < j} |\lambda_{ij} - v_{ij}| (1 - \lambda_{ij}^2)^{-\frac{1}{2}} e^{-u^2/(1+|\lambda_{ij}|)}.$$

In this application (using the notation of (3.4)), the ξ_i are identified with the r.v.’s $\xi(s_1) \dots \xi(s_p), \xi(t_1) \dots \xi(t_{p'})$ and the η_i with $p+p'$ standard normal r.v.’s having the same correlations except that $\text{cov}(\xi(s_i), \xi(t_j))$ is replaced by zero for $1 \leq i \leq p, 1 \leq j \leq p'$.

The fact that boundedness of $T\psi(u_T)$ together with $r(t)\log t \rightarrow 0$ implies (5.7) follows by standard calculations (cf. [1] or Lemma 3.1 of [13]).

(iv) If $r(t)\log t \rightarrow 0$ and $T\psi(u_T) \rightarrow \tau > 0$ then $D'_c(u_T)$ may be simply obtained by employing Slepian's Lemma with $n=2$. It then follows from Theorem 4.3 that $P\{M(T) \leq u_T\} \rightarrow e^{-\tau}$.

(v) This follows at once from the (relatively) straightforward verification of the fact that $T\psi(u_T) \rightarrow \tau = e^{-x}$ when $u_T = x/a_T + b_T$, using the above results. \square

6. Poisson and Related Properties

In this section we shall just briefly indicate the Poisson properties associated with high level upcrossings. We confine the discussion to the case where the number $N_u(I)$ of upcrossings in a bounded interval I has a finite mean, writing again $\mu = \mu(u) = \mathcal{E}N_u(I)$. Cases where this is not so are similarly dealt with in terms of so-called ε -upcrossings (cf. [15]).

Our objective is to show, under D_c and D'_c conditions, that the point process of upcrossings of a high level takes on a Poisson character – as is well-known in the case when the stationary process $\zeta(t)$ is normal. Since the upcrossings of increasingly high levels will tend to become rare, a normalization is required. To that end we consider a time period T and a level u_T , both increasing in such a way that $T\mu \rightarrow \tau$, ($\mu = \mu(u_T)$), and define a normalized point process of upcrossings by

$$N_T^*(I) = N_{u_T}(TI), \quad (N_T^*(t) = N_{u_T}(tT))$$

for each interval (or more general Borel set) I , so that, in particular,

$$\mathcal{E}N_T^*(1) = \mathcal{E}N_{u_T}(T) = \mu T \rightarrow \tau. \quad (6.1)$$

This shows that the “intensity” (i.e. mean number of events per unit time) of the (normalized) upcrossing point process converges to τ . Our task is to show that the upcrossing point process actually converges (weakly) to a Poisson process with mean τ .

The derivation of this result is based on the following two extensions of Theorem 4.3, which are proved by similar arguments to those used in obtaining Theorem 4.3.

Theorem 6.1. *Under the conditions of Theorem 4.3 with $\psi(u) = \mu(u)$, if $\theta < 1$ and $\mu T \rightarrow \tau$, then*

$$P\{M(\theta T) \leq u_T\} \rightarrow e^{-\theta\tau} \quad \text{as } T \rightarrow \infty. \quad \square \quad (6.2)$$

Theorem 6.2. *If $I_1, I_2 \dots I_k$ are disjoint subintervals of $[0, 1]$ and $I_j^* = TI_j = \{t: t/T \in I_j\}$, then under the conditions of Theorem 4.3 with $\psi(u) = \mu(u)$ if $\mu T \rightarrow \tau$,*

$$P\left\{\bigcap_{j=1}^k (M(I_j^*) \leq u_T)\right\} - \prod_{j=1}^k P\{M(I_j^*) \leq u_T\} \rightarrow 0, \quad (6.3)$$

so that by Theorem 6.1

$$P\left\{\bigcap_{j=1}^k (M(I_j^*) \leq u_T)\right\} \rightarrow e^{-\tau \sum \theta_j}, \quad (6.4)$$

where θ_j is the length of I_j , $1 \leq j \leq k$. \square

It is now a relatively straightforward matter to show that the point processes N_T^* converge (in the full sense of weak convergence) to a Poisson process N with intensity τ .

Theorem 6.3. *Under the conditions of Theorem 4.3, with $\psi = \mu$, if $T\mu \rightarrow \tau$ where $\mu = \mu(u_T)$, then the family N_T^* of (normalized) point processes of upcrossings of u_T on the unit interval converges in distribution to a Poisson process N with intensity τ on the unit interval as $T \rightarrow \infty$.*

Proof. By Theorem 4.7 of [8] it is sufficient to prove that

(i) $\mathcal{E}N_T^*\{(a, b]\} \rightarrow \mathcal{E}N\{(a, b]\} = \tau(b-a)$ as $T \rightarrow \infty$ for all a, b , $0 \leq a \leq b \leq 1$.

(ii) $P\{N_T^*(B) = 0\} \rightarrow P\{N(B) = 0\}$ as $T \rightarrow \infty$ for all sets B of the form $\bigcup_1^n B_i$ where n is any integer and B_i are disjoint intervals $(a_i, b_i] \subset (0, 1]$.

Now (i) follows trivially since

$$\mathcal{E}N_T^*\{(a, b]\} = \mu T(b-a) \rightarrow \tau(b-a).$$

To obtain (ii) we note that

$$\begin{aligned} 0 &\leq P\{N_T^*(B) = 0\} - P\{M(TB) \leq u_T\} \\ &= P\{N_u(TB) = 0, M(TB) > u_T\} \\ &\leq \sum_{i=1}^n P\{\xi(Ta_i) > u_T\} \end{aligned}$$

since if the maximum in $TB = \bigcup_{i=1}^n (Ta_i, Tb_i]$ exceeds u_T , but there are no upcrossings of u_T in these intervals, then ξ must exceed u at the initial point of at least one such interval. But the last expression is just $nP\{\xi(0) > u_T\} \rightarrow 0$ as $T \rightarrow \infty$. Hence

$$P\{N_T^*(B) = 0\} - P\{M(TB) \leq u_T\} \rightarrow 0.$$

But $P\{M(TB) \leq u_T\} = P\left\{\bigcap_{i=1}^n (M(TB_i) \leq u_T)\right\} \rightarrow e^{-\tau \sum (b_i - a_i)}$ by Theorem 6.2 so that

(ii) follows since $P\{N(B) = 0\} = e^{-\tau \sum (b_i - a_i)}$. \square

Corollary 6.4. *If B_i are disjoint (Borel) subsets of the unit interval and if the boundary of each B_i has zero Lebesgue measure then*

$$P\{N_T^*(B_i) = r_i, 1 \leq i \leq n\} \rightarrow \prod_{i=1}^n e^{-\tau m(B_i)} \frac{[\tau m(B_i)]^{r_i}}{r_i!}$$

where $m(B_i)$ denotes the Lebesgue measure of B_i .

Proof. This is an immediate consequence of the full weak convergence proved (cf. Lemma 4.4 of [8]). \square

The above results concern convergence of the point processes of upcrossings of u_T in the unit interval to a Poisson process in the unit interval. A slight modification (requiring D_c and D'_c to hold for all families $u_{\theta T}$ in place of u_T for all $\theta > 0$) enables a corresponding result to be shown for the upcrossings on the whole positive real line, but we do not pursue this here. Instead we show how Theorem 6.3 yields the asymptotic distribution of the r^{th} largest local maximum in $(0, T)$.

Suppose, then, that $\xi(t)$ has a continuous derivative a.s. and define $N'_u(T)$ to be the number of local maxima in the interval $(0, T)$ for which the process value exceeds u , i.e. the number of downcrossing points t of zero by ξ' in $(0, T)$ such that $\xi(t) > u$. Clearly $N'_u(T) \geq N_u(T) - 1$ since at least one local maximum occurs between two upcrossings. It is also reasonable to expect that if the sample function behavior is not too irregular there will tend to be just one local maximum between most successive upcrossings of u when u is large, so that $N'_u(T)$ and $N_u(T)$ will tend to be approximately equal. The following result makes this precise.

Theorem 6.5. *With the above notation let $\{u_T\}$ be constants such that $T\mu(=T\mu(u_T)) \rightarrow \tau > 0$. Suppose that $\mathcal{E}N'_u(1)$ is finite for each u and that $\mathcal{E}N'_u(1) \sim \mu(u)$ as $u \rightarrow \infty$. Then, writing $u_T = u$, $\mathcal{E}|N'_u(T) - N_u(T)| \rightarrow 0$.*

If also the conditions of Theorem 6.3 hold (so that $P\{N'_u(T) = r\} \rightarrow e^{-\tau} \tau^r / r!$) it follows that $P\{N'_u(T) = r\} \rightarrow e^{-\tau} \tau^r / r!$.

Proof. As noted above, $N'_u(T) \geq N_u(T) - 1$, and it is clear, moreover, that if $N'_u(T) = N_u(T) - 1$, then $\xi(T) > u$. Hence

$$\begin{aligned} \mathcal{E}|N'_u(T) - N_u(T)| &= \mathcal{E}\{N'_u(T) - N_u(T)\} + 2P\{N'_u(T) = N_u(T) - 1\} \\ &\leq T\mathcal{E}N'_u(1) - \mu T + 2P\{\xi(T) > u\}, \end{aligned}$$

which tends to zero as $T \rightarrow \infty$ since $P\{\xi(T) > u_T\} = P\{\xi(0) > u_T\} \rightarrow 0$ and $T\mathcal{E}N'_{u_T}(1) - \mu T = \mu T[(1 + o(1)) - 1] \rightarrow 0$, so that the first part of the theorem follows. The second part now follows immediately since the integer-valued r.v. $N'_u(T) - N_u(T)$ tends to zero in probability, giving $P\{N'_u(T) \neq N_u(T)\} \rightarrow 0$ and hence $P\{N'_u(T) = r\} - P\{N_u(T) = r\} \rightarrow 0$ for each r . \square

Now write $M^{(r)}(T)$ for the r^{th} largest local maximum in the interval $(0, T)$. Since the events $\{M^{(r)}(T) \leq u\}$, $\{N'_u(T) < r\}$ are identical we obtain the following corollary:

Corollary 6.6 *Under the conditions of the theorem*

$$P\{M^{(r)}(T) \leq u_T\} \rightarrow e^{-\tau} \sum_{s=0}^{r-1} \tau^s / s!. \quad \square$$

As a further corollary we obtain the limiting distribution of $M^{(r)}(T)$ in terms of that for $M(T)$.

Corollary 6.7. *Suppose that $P\{a_T(M(T) - b_T) \leq x\} \rightarrow G(x)$ and that the conditions of Theorem 4.3 hold with $u_T = x/a_T + b_T$ for each real x (and $\psi = \mu$). Suppose also that $\mathcal{E}N'_u(1) \sim \mathcal{E}N_u(1)$ as $u \rightarrow \infty$. Then*

$$P\{a_T(M^{(r)}(T) - b_T) \leq x\} \rightarrow G(x) \sum_{s=0}^{r-1} [-\log G(x)]^s / s!,$$

where $G(x) > 0$ (and zero if $G(x) = 0$).

Proof. This follows from Corollary 6.6 by writing $G(x) = e^{-\tau}$ since Theorem 4.3 implies that $T\mu \rightarrow \tau$. \square

Note that for a stationary normal process with finite second and fourth spectral moments λ_2, λ_4 it may be shown (Sect. 11.6 of [3]) that

$$\mathcal{E}N'_u(1) = \mu \Phi(u\lambda_2/\Delta^{\frac{1}{2}}) + (2\pi)^{-1} (\lambda_4/\lambda_2)^{\frac{1}{2}} [1 - \Phi\{u(\lambda_4/\Delta)^{\frac{1}{2}}\}]$$

where $\Delta = \lambda_4 - \lambda_2^2$ and Φ is the standard normal d.f., so that clearly $\mathcal{E}N'_u(1) \sim \mu$ as $u \rightarrow \infty$.

The relation (6.5) gives the asymptotic distribution of the r^{th} largest local maximum $M^{(r)}(T)$ as a corollary of the Poisson result, Theorem 6.3. This Poisson result may itself be generalized to apply to joint convergence of upcrossings of several levels to a point process in the plane composed of successive “thinnings” of a Poisson process. From a result of this kind it is possible to obtain the joint asymptotic distribution of any number of the $M^{(r)}(T)$, and also of their time locations.

Acknowledgement. We are very grateful to Georg Lindgren for uncountably many helpful conversations regarding this and related topics.

References

1. Berman, S.M.: Maxima and high level excursions of stationary Gaussian processes. Trans. Amer. Math. Soc. **60**, 65-85 (1961)
2. Berman, S.M.: Limit theorems for the maximum term in stationary sequences. Ann. Math. Statist. **35**, 502-516 (1964)
3. Cramér, H., Leadbetter, M.R.: Stationary and Related Stochastic Processes. New York: John Wiley 1967
4. Fisher, R.A., Tippett, L.H.C.: Limiting forms of the frequency distribution of the largest or smallest member of a sample. Proc. Cambridge Philos. Soc. **24**, 180-190 (1928)
5. Fréchet, M.: Sur la loi de probabilité de l'écart maximum. Ann. de la Soc. Polonaise de Math (Cracow), p. 93 (1927)
6. Gnedenko, B.V.: Sur la distribution limite du terme maximum d'une serie aleatoire. Ann. Math. **44**, 423-453 (1943)
7. Haan, L. de: On regular variation and its application to the weak convergence of sample extremes. Amsterdam Math. Centre Tract **32** (1970)
8. Kallenberg, O.: Random measures. Berlin: Akademie-Verlag 1975; and London and New York: Academic Press 1976
9. Leadbetter, M.R. (1974). On extreme values in stationary sequences. Z. Wahrscheinlichkeitstheorie verw. Geb. **28**, 289-303 (1974)

10. Leadbetter, M.R.: Extreme value theory under weak mixing conditions. *M.A.A. Studies in Mathematics 18 and Studies in Probability Theory*, ed. M. Rosenblatt, 46–110 (1978)
11. Leadbetter, M.R.: On extremes of stationary processes. *Institute of Statistics Mimeo Series #1194*, University of North Carolina at Chapel Hill (1978)
12. Leadbetter, M.R., Lindgren, G., Rootzen, H.: Extremal and related properties of stationary processes I: Extremes of stationary sequences. *Institute of Statistics Mimeo Series #1227*, University of North Carolina at Chapel Hill (1979)
13. Leadbetter, M.R., Lindgren, G., Rootzen, H.: Extremal and related properties of stationary processes II: Extremes of continuous parameter stationary processes. *Institute of Statistics Mimeo Series #1307*, University of North Carolina at Chapel Hill.
14. Loynes, R.M.: Extreme values in uniformly mixing stationary stochastic processes. *Ann. Math. Statist.* **36**, 993–999 (1965)
15. Pickands, J.: Upcrossing probabilities for stationary Gaussian processes. *Trans. Amer. Math. Soc.* **145**, 51–73 (1969)
16. Qualls, C., Watanabe, H.: Asymptotic properties of Gaussian processes. *Ann. Math. Statist.* **43**, 580–596 (1972)
17. Watson, G.S.: Extreme values in samples from m -dependent stationary stochastic processes. *Ann. Math. Statist.* **25**, 798–800 (1954)

Received May 7, 1980; in revised form July 30, 1981