An Approximation Theorem for Sums of Certain Randomly Selected Indicators

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1. Introduction

The Poisson approximation for sums of independent Bernoulli random variables has been of considerable interest in the literature [see Prohorov (1953), Hodges and LeCam (1960), LeCam (1960), Kerstan (1964) and Vervaat (1969)]. The problem was generalised by Chen (1975) to include certain classes of dependent Bernoulli random variables. In this paper a similar approximation theorem is proved for the distribution of $\sum_{i=1}^{n} X_{i\pi(i)}$, where X_{ij} , i, j = 1, 2, ..., n, are independent Bernoulli random variables and $(\pi(1), \pi(2), ..., \pi(n))$ a random permutation of (1, 2, ..., n) independent of the X_{ij} 's. The nature of the dependence among $X_{1\pi(1)}, ..., X_{n\pi(n)}$ differs from that considered in [2].

A number of corollaries are derived from the main theorem. These include results of LeCam (1960), but with larger absolute constants in the bounds. The largeness of the absolute constants can be attributed to the greater generality of the present problem. One of the corollaries is a Poisson counterpart of a theorem of Wald and Wolfowitz (1944), where the latter is actually a limit theorem. [See also Noether (1949), Hoeffding (1951) and Robinson (1972).] Another corollary is an approximation theorem for the hypergeometric distribution.

Throughout this paper, all summations will be from 1 to n unless otherwise stated.

2. The Main Theorem

We first state the theorem as follows:

Theorem 2.1. Let X_{ij} , i, j = 1, 2, ..., n be independent Bernoulli random variables with $P(X_{ij}=1)=1-P(X_{ij}=0)=p_{ij}$ and let $(\pi(1), \pi(2), ..., \pi(n))$ be a random perpemutation of (1, 2, ..., n) independent of the X_{ij} 's. Then for $n \ge 5$ and every realvalued function h defined on the non-negative integers such that $|h| \le 1$, we have

$$|Eh(\sum_{i} X_{i\pi(i)}) - \mathcal{P}_{\lambda}h| \leq 15.75 \min(\lambda^{-\frac{1}{2}}, 1) \{\sum_{i} \bar{p}_{i+}^{2} + \sum_{j} \bar{p}_{+j}^{2}\}$$
(2.1)

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(2.7)

and

$$|Eh(\sum_{i} X_{i\pi(i)}) - \mathscr{P}_{\lambda}h| \leq 45.25 \,\lambda^{-1} \left\{ \sum_{i} \bar{p}_{i+}^{2} + \sum_{j} \bar{p}_{+j}^{2} \right\}$$
(2.2)

where

$$\mathcal{P}_{\lambda}h = e^{-\lambda} \sum_{k=0}^{\infty} h(k)\lambda^{k}/k!,$$

$$\bar{p}_{i+} = \sum_{j} p_{ij}/n, \qquad \bar{p}_{+j} = \sum_{i} p_{ij}/n,$$

$$\lambda = \sum_{i} \bar{p}_{i+} = \sum_{j} \bar{p}_{+j} = \sum_{i} \sum_{j} p_{ij}/n.$$

The proof of the theorem is based on the derivation of an identity similar to that in [2] and a few lemmas. (An interesting application of a special case of these identities is in [1].)

Let *I*, *J*, *K*, *L*, *M* be random variables, each uniformly distributed on $\{1, 2, ..., n\}$ and let $\pi = (\pi(1), \pi(2), ..., \pi(n)), \tilde{\pi} = (\tilde{\pi}(1), \tilde{\pi}(2), ..., \tilde{\pi}(n))$ and $\tilde{\tilde{\pi}} = (\tilde{\tilde{\pi}}(1), \tilde{\tilde{\pi}}(2), ..., \tilde{\tilde{\pi}}(n))$ be random permutations of (1, 2, ..., n) such that

- { $I, J, K, L, M, \pi, \tilde{\pi}, \tilde{\pi}$ } is independent of { X_{ij} : i, j = 1, 2, ..., n}, (2.3)
- (I, K) and (L, M) are uniformly distributed on $\{(i, k): i \neq k; i, k = 1, 2, ..., n\}$, (2.4)
- J, (I, K), (L, M) and $\tilde{\pi}$ are mutually independent, (2.5)

 $J_{\tau}(I, K)$ and $\tilde{\pi}$ are mutually independent, (2.6)

I and π are independent,

$$\tilde{\pi}(\alpha) = \begin{cases} \tilde{\tilde{\pi}}(\alpha), & \alpha \neq I, K, \tilde{\tilde{\pi}}^{-1}(L), \tilde{\tilde{\pi}}^{-1}(M) \\ L, & \alpha = I \\ M, & \alpha = K \\ \tilde{\tilde{\pi}}(I), & \alpha = \tilde{\pi}^{-1}(L) \\ \tilde{\tilde{\pi}}(K), & \alpha = \tilde{\tilde{\pi}}^{-1}(M) \end{cases}$$

$$(2.8)$$

and

$$\pi(\alpha) = \begin{cases} \tilde{\pi}(\alpha), & \alpha \neq I, \, \tilde{\pi}^{-1}(J) \\ J, & \alpha = I \\ \tilde{\pi}(I), & \alpha = \tilde{\pi}^{-1}(J), \end{cases}$$
(2.9)

where $\tilde{\pi}(\tilde{\pi}^{-1}(\alpha)) = \alpha$ and $\tilde{\tilde{\pi}}(\tilde{\tilde{\pi}}^{-1}(\alpha)) = \alpha$.

The consistency of the conditions (2.3)-(2.9) can easily be verified.

Now let (Ω, \mathcal{B}, P) be a common probability space on which all the above random vectors are defined and let

 \mathscr{F} = the σ -algebra generated by π and the X_{ij} 's,

$$\begin{split} W &= \sum_{i} X_{i\pi(i)}, \qquad \tilde{W} = \sum_{i} X_{i\tilde{\pi}(i)}, \qquad \tilde{\tilde{W}} = \sum_{i} X_{i\tilde{\pi}(i)}, \\ W^* &= \sum_{i \neq I} X_{i\pi(i)}, \qquad \tilde{W}^* = \sum_{i \neq I} X_{i\tilde{\pi}(i)}, \qquad \tilde{W}^{**} = \sum_{i \neq I, K} X_{i\tilde{\pi}(i)}. \end{split}$$

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Also define the operator Δ by $\Delta f(w) = f(w+1) - f(w)$. Then, using the basic properties of conditional expectations, the fact that each X_{ij} takes on 0 and 1, the conditional independence of $X_{I\pi(I)}$ and W^* given (I, π) and the independence of I and W, we obtain, for every real-valued function f defined on $\{0, 1, 2, ...\}$.

$$E[Wf(W)] = nE \{ [E^{\mathscr{P}} X_{I\pi(I)}] f(W) \} = nE [X_{I\pi(I)} f(W)] = nE [p_{I\pi(I)} f(W^{*} + 1)]$$

$$= nE [(p_{I\pi(I)} - \bar{p}_{I_{+}}) f(W^{*} + 1)] + nE \{ \bar{p}_{I_{+}} [f(W^{*} + 1) - f(W + 1)] \}$$

$$+ \lambda E f(W + 1)$$

$$= nE [(p_{IJ} - \bar{p}_{I_{+}}) f(W^{*} + 1)] - nE [\bar{p}_{I_{+}} p_{IJ} \Delta f(W^{*} + 1)]$$

$$+ \lambda E f(W + 1),$$
(2.10)

where $E^{\mathscr{F}}$ denotes conditional expectation given the σ -algebra \mathscr{F} . Using again the fact that each X_{ij} takes on 0 and 1, the conditional independence of p_{IJ} and \tilde{W}^* given I and the conditional independence of

$$(X_{\tilde{\pi}^{-1}(J), \tilde{\pi}(I)}, X_{\tilde{\pi}^{-1}(J), J})$$
 and $\sum_{\alpha \neq I, \tilde{\pi}^{-1}(J)} X_{\alpha \tilde{\pi}(\alpha)}$

given $(I, J, \tilde{\pi})$, we obtain

$$nE[(p_{IJ} - \bar{p}_{I+})f(W^{*} + 1)]$$

$$= nE\{(p_{IJ} - \bar{p}_{I+})[f(W^{*} + 1) - f(\tilde{W}^{*} + 1)]\}$$

$$= nE\{(p_{IJ} - \bar{p}_{I+})(X_{\tilde{\pi}^{-1}(J), \tilde{\pi}(I)} - X_{\tilde{\pi}^{-1}(J), J})\Delta f(\sum_{\alpha \neq I, \tilde{\pi}^{-1}(J)} X_{\alpha \tilde{\pi}(\alpha)} + 1)\}$$

$$= nE\{(p_{IJ} - \bar{p}_{I+})(p_{\tilde{\pi}^{-1}(J), \tilde{\pi}(I)} - p_{\tilde{\pi}^{-1}(J), J})\Delta f(\sum_{\alpha \neq I, \tilde{\pi}^{-1}(J)} X_{\alpha \tilde{\pi}(\alpha)} + 1)\}$$

$$= n(n-1)E\{(p_{IM} - \bar{p}_{I+})(p_{KL} - p_{KM})\chi(J = M)\Delta f(\tilde{W}^{**} + 1)\}$$
(2.11)

where $\chi(A)$ is the indicator function of the set A. Now, as in [2], we choose f such that

$$wf(w) - \lambda f(w+1) = h(w) - \mathcal{P}_{\lambda}h \tag{2.12}$$

where h is another real-valued and bounded function defined on $\{0, 1, 2, ...\}$, and let $S_{\lambda}h(w)$ denote the solution of the difference equation (2.12) for $w \ge 1$ (the solution is unique except at w = 0). Then (2.10) and (2.11) yield

$$Eh(W) = \mathscr{P}_{\lambda}h + n(n-1)E[(p_{IM} - \bar{p}_{I+})(p_{KL} - p_{KM})\chi(J = M)\Delta S_{\lambda}h(\tilde{W}^{**} + 1) - nE[\bar{p}_{I+}p_{IJ}\Delta S_{\lambda}h(W^{*} + 1)].$$
(2.13)

In order to bound the error terms on the right hand side of (2.13), we need a few lemmas.

Lemma 2.1 [2]. For $|h| \le 1$ and $w \ge 1$, $|\Delta S_{\lambda}h(w)| \le 6 \min (\lambda^{-\frac{1}{2}}, 1)$. Lemma 2.2. [2] For $|h| \le 1$ and $w \ge 1$, $|\Delta S_{\lambda}h(w)| \le \lambda^{-1} \{2+4|w-\lambda| \min (\lambda^{-\frac{1}{2}}, 1)\}$. Lemma 2.3. For $Z = \tilde{W}$ or $\tilde{\tilde{W}}$, $E(Z-\lambda)^2 \le \lambda$. Proof. Direct computations yield

$$E(Z-\lambda)^{2} = \lambda + \lambda^{2}/(n-1) - n \sum_{i} \overline{p}_{i+}^{2}/(n-1) - n \sum_{j} \overline{p}_{+j}^{2}/(n-1) + \sum_{i} \sum_{j} p_{ij}^{2}/n(n-1).$$

This together with the inequalities $\lambda^2 \leq n \sum_i \overline{p}_{i+}^2$ and $\sum_i p_{ij}^2 \leq (\sum_i p_{ij})^2$ proves the lemma.

Lemma 2.4.

$$n(n-1) E |(p_{IM} - \bar{p}_{I+})(p_{KL} - p_{KM}) \chi(J = M)| \leq (3n-2) \lambda^2 / n(n-1) + \sum_{i} \bar{p}_{+i}^2.$$

Proof. By direct computations.

We now prove Theorem 2.1. In the following, we shall take h in (2.13) to be such that $|h| \leq 1$. We first bound the second error term on the right hand side of (2.13). By Lemma 2.1, we obtain

$$|nE[\bar{p}_{I+}p_{IJ} \Delta S_{\lambda}h(W^*+1)]| \leq 6\min(\lambda^{-\frac{1}{2}},1)\sum_{i}\bar{p}_{i+}^2.$$
(2.14)

Also, by Lemma 2.2, the inequality $|W^*+1-\tilde{W}| \leq 2$ and the independence of I, J and \tilde{W} , we obtain

$$|nE[\bar{p}_{I+}p_{IJ}\Delta S_{\lambda}h(W^{*}+1)]| \leq \lambda^{-1} [nE\bar{p}_{I+}p_{IJ}] [10+4\min(\lambda^{-\frac{1}{2}},1)E|\tilde{W}-\lambda|]$$

which by Jensen's inequality and Lemma 2.3

$$\leq 14\lambda^{-1}\sum_{i}\bar{p}_{i+}^{2}.$$
(2.15)

Next we consider the first error term on the right hand side of (2.13). By Lemmas 2.1 and 2.4, we obtain

$$|n(n-1)E[(p_{IM}-\bar{p}_{I+})(p_{KL}-p_{KM})\chi(J=M)\Delta S_{\lambda}h(\tilde{W}^{**}+1)]| \leq 6\min(\lambda^{-\frac{1}{2}},1)\{(3n-2)\lambda^{2}/n(n-1)+\sum_{j}\bar{p}_{+j}^{2}\}.$$
(2.16)

Also, by Lemma 2.2, the inequality $|\tilde{W}^{**} + 1 - \tilde{\tilde{W}}| \leq 3$, the independence of J, (I, K), (L, M) and $\tilde{\tilde{W}}$, and Lemma 2.4, we obtain as in (2.15)

$$|n(n-1)E[(p_{IM}-\bar{p}_{I+})(p_{KL}-p_{KM})\chi(J=M)\Delta S_{\lambda}h(\tilde{W}^{**}+1)]| \leq 18\,\lambda^{-1}\{(3n-2)\lambda^{2}/n(n-1)+\sum_{i}\bar{p}_{+i}^{2}\}.$$
(2.17)

Finally, noting that $n \ge 5$, that $\lambda^2 \le n \sum_i \overline{p}_{i+}^2$ and that $\lambda^2 \le n \sum_j \overline{p}_{+j}^2$, we obtain (2.1) from (2.13), (2.14) and (2.16), and obtain (2.2) from (2.13), (2.15) and (2.17). This completes the proof of Theorem 2.1.

3. Corollaries

Except for Corollary 3.2, all the corollaries in this section have already been mentioned in the Introduction. Corollaries 3.3 and 3.4 are actually corollaries to

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Corollary 3.2. Unless otherwise stated, all notations are the same as in the preceding sections.

Corollary 3.1 (LeCam). Let $X_1, X_2, ..., X_n$ be independent Bernoulli random variables with $P(X_i=1)=1-P(X_i=0)=p_i$. Then for $n \ge 5$ and $|h| \le 1$, we have

$$\left|Eh\left(\sum_{i}X_{i}\right)-\mathscr{P}_{\lambda}h\right| \leq 31.5 \min\left(\lambda^{-\frac{1}{2}},1\right)\sum_{i}p_{i}^{2}$$

and

$$\left| Eh\left(\sum_{i} X_{i}\right) - \mathscr{P}_{\lambda}h \right| \leq 90.5 \,\lambda^{-1} \sum_{i} p_{i}^{2}$$

where

$$\lambda = \sum_{i} p_{i}$$

Proof. Let $p_{ij} = p_j$ for all *i* and *j* and observe that $\lambda^2 \leq n \sum_i p_i^2$.

Corollary 3.2. Let a_{ij} , i, j = 1, 2, ..., n, be an array of 0's and 1's and let $(\pi(1), \pi(2), ..., \pi(n))$ be a random permutation of (1, 2, ..., n). Then for $n \ge 5$ and $|h| \le 1$, we have

$$\left|Eh\left(\sum_{i}a_{i\pi(i)}\right)-\mathscr{P}_{\lambda}h\right| \leq 15.75\min\left(\lambda^{-\frac{1}{2}},1\right)\left\{\sum_{i}p_{i}^{2}+\sum_{j}q_{j}^{2}\right\}$$

and

$$\left|Eh\left(\sum_{i}a_{i\pi(i)}\right)-\mathscr{P}_{\lambda}h\right| \leq 45.25\,\lambda^{-1}\left\{\sum_{i}p_{i}^{2}+\sum_{j}q_{j}^{2}\right\}$$

where

$$p_i = \sum_j a_{ij}/n, \qquad q_j = \sum_i a_{ij}/n,$$

$$\lambda = \sum_i p_i = \sum_j q_j = \sum_i \sum_j a_{ij}/n.$$

Proof. Let $p_{ij} = a_{ij}$ for every *i* and *j*.

Corollary 3.3. Let $a_1, \ldots, a_n, b_1, \ldots, b_n$ be 0's and 1's and let $(\pi(1), \ldots, \pi(n))$ be a random permutation of $(1, \ldots, n)$. Then for $n \ge 5$ and $|h| \le 1$, we have

$$\left|Eh\left(\sum_{i}a_{i}b_{\pi(i)}\right)-\mathscr{P}_{\lambda}h\right| \leq 45.25(\bar{a}+\bar{b})$$

where

$$\overline{a} = \sum_{i} a_{i}/n, \quad \overline{b} = \sum_{i} b_{i}/n, \quad \lambda = n \overline{a} \overline{b}.$$

Proof. This follows from Corollary 3.2 with $a_{ij} = a_i b_j$.

Corollary 3.4. Let

$$h(r; n, a, b) = \begin{cases} \binom{a}{r} \binom{n-a}{b-r} / \binom{n}{b} & \text{if } \max(a+b-n, 0) \leq r \leq a \\ 0 & \text{otherwise.} \end{cases}$$

Then

$$\sum_{r=0}^{\infty} |h(r; n, a, b) - e^{-\lambda} \lambda^r / r!| \leq 45.25 (a+b) / n$$

where

 $\lambda = ab/n$.

Proof. This follows from Corollary 3.2 with $a_{ij}=1$ if $1 \le i \le a$, $1 \le j \le b$ and =0 otherwise, and with h(r)=1 or -1 according as $h(r; n, a, b) \ge$ or $\langle e^{-\lambda} \lambda^r / r!$.

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