On Large Deviations

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Received April 1, 1966

A well-known lemma by K. ESSEEN ([1], § 39, Bd. 1) and its sharpenings obtained by A. C. BERRY [2] and V. M. ZOLOTAREV (cf. (25)) allow to estimate the distance sup |F(x) - G(x)| of two distribution functions F and G in terms of the closeness of their characteristic functions. Similarly, using the Laplace transform

$$\int_{-\infty}^{\infty} e^{zx} d(F(x) - G(x)), \quad |z| \leq \Delta$$

one might try to estimate the ratios $\frac{1F(-x)}{1-G(x)}$ and $\frac{F(-x)}{G(-x)}$ in the interval $0 \le x \le \delta \Delta$ and in this way simplify the derivation of limit theorems including the calculation of large deviations. We consider here the case where G is the normal law $\Phi = N(0, 1)$.

Lemma. Let ξ be a random variable with distribution function F, mean $m = M \xi$, dispersion $\sigma^2 = D\xi$ and finite moments of any order $M |\xi|^k < \infty, k = 1, 2, ...$ Let \mathfrak{S}_k be the k-th cumulant of ξ and

$$\Delta = \sigma \inf\left(\frac{k! H \sigma^2}{|\tilde{\mathfrak{S}}_k|}\right)^{\frac{1}{k-2}} \tag{1}$$

where H > 0. Then in the interval

$$1\!\leq\!x\!\leq\!ar{\delta}\!arDelta$$
 , $ar{\delta}$

the following relations hold:

$$\frac{1 - F(m + x\sigma)}{1 - \Phi(x)} = e^{\frac{x^3}{4}\lambda} \left(\frac{x}{4}\right) \left(1 + f_1(\bar{\delta}, H)\frac{x}{4}\right),\tag{2}$$

$$\frac{F(m-x\sigma)}{\Phi(-x)} = e^{-\frac{x^3}{\Delta}\lambda} \left(-\frac{x}{\Delta}\right) \left(1 + f_2(\bar{\delta}, H) \frac{x}{\Delta}\right).$$
(2')

Here,

$$\left|f_{i}(\bar{\delta}, H)\right| < \frac{8H\{1+7, 2(1+2\delta+\min\{\frac{1}{3}(1-\bar{\delta})^{3}H^{-1}, \frac{1}{2}H^{-\frac{1}{4}}\})^{4}\}}{(1-\delta)^{4}(1-\varrho)^{\frac{3}{2}}}.$$
(3)

i=1,2, the number δ and δ_H with $0<\delta<\delta_H$ are determined by the equations

$$\bar{\delta} = \frac{\delta(1+\delta)}{2}, \quad \varrho = \frac{6 H \delta}{(1-\delta)^3}, \quad \bar{\delta}_H = \frac{\delta_H (1+\delta_H)}{2}, \quad (4)$$

 δ_H is a real root of the equation $\varrho = 1$ and $\lambda(t) = \sum_{k=0}^{\infty} \lambda_k t^k$ is the power series of Cramèr which converges for $|t| < \delta_H$, where

$$\left|\lambda_{k}\right| \leq \frac{\delta_{H}}{(k+3)\overline{\delta}_{H}^{k+2}}, k = 0, 1, \dots.$$

$$(5)$$

We remark that

$$\bar{\delta}_{H} > \frac{1}{1 + 14,55 \max{\{H, H^{\frac{1}{2}}\}}}$$
(6)

A more precise estimate for $h(\overline{\delta}, H)$ is given in (38) and (39). The coefficients λ_k are determined by the first k + 3 cumulants (cf. (42), (43)). We always set $\frac{1}{0} = \infty$, $\frac{1}{\infty} = 0$. Moreover, θ always denotes some quantity not exceeding 1 in absolute value.

Instead of the formular (1) in terms of semiinvariants, which is equivalent to

$$\left|\mathfrak{S}_{k}\right| \leq \frac{k! \, H \, \sigma^{k}}{\varDelta^{k-2}}, \quad k = 3, 4, \dots$$

$$(1')$$

it is sometimes more convenient to use immediately the Laplace transform

$$\psi(z) = \int_{-\infty}^{+\infty} e^{zx} dF(x+m)$$

in order to define H and Δ .

Remark. If there exist $H < \infty$ and $\Delta < \infty$ satisfying

$$\left|\ln \varphi(z)\right|_{|z|=\frac{A}{\sigma}} \leq H \Delta^2, \qquad (1'')$$

then the assumptions of the lemma are satisfied and (2) and (2') hold with these H and Δ .

Here we take the principle value of logarithm. Before entering the proof of the lemma we obtain with its help some theorems on probabilities of large deviations.

Let $\{\xi(t), 0 \leq t < \infty\}$ be a measurable stochastic process. We say that $\xi(t)$ belongs to the class $T^{(k)}$ if

$$M \left| \xi(t) \right|^k \leq C_k < \infty \, .$$

Consider the random variable

$$\zeta_T = \int_0^T \xi(t) \, dt$$

and let

$$m_T = M \zeta_T, \quad \sigma_T^2 = \mathfrak{D} \zeta_T$$

and $s_{\xi}^{(k)}(t_1, \ldots, t_k)$ be the correlation function of k-th order of the process $\xi(t)$, which is simply the semiinvariant of the random vector $(\xi(t_1), \ldots, \xi(t_k))$.

Theorem 1. If $\xi(t) \in T^{(\infty)}$ and

$$\sigma_T^{-k} \left| \int_0^T \cdots \int_0^T s_{\xi}^{(k)}(t_1, \dots, t_k) dt_1 \cdots dt_k \right| \le \frac{k! H_1}{\varDelta_T^{k-3}}$$
(7)

for all $k \ge 3$ then for $F(x) = P\{\zeta_T < x\}$ the relations (2) and (2') hold with $m = m_T$, $\sigma = \sigma_T$, $H = H_1$, $\Delta = \Delta_T$ and

$$\mathfrak{S}_{k} = \int_{0}^{T} \cdots \int_{0}^{T} s_{\xi}^{(k)}(t_{1}, \dots, t_{k}) dt_{1} \dots dt_{k}.$$
(8)

In fact, \mathfrak{S}_k as defined by the equality (8) is the cumulant of order k of the random variable ζ_T (cf. [9], (1.14)), and by comparing (7) with (1') we obtain the proof of the theorem.

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Let

$$\xi_1, \dots, \xi_n, \quad n \ge 1 \tag{9}$$

be independent random variables with $M\xi_j = 0, j = 1, 2, ..., n$,

$$S_n = \sum_{j=1}^n \xi_j, \quad \sigma_j^2 = \mathfrak{D}\xi_j, \quad B_n^2 = \sum_{j=1}^n \sigma_j^2$$

and Γ_{kn} the cumulant of S_n of order k.

Theorem 2. If the condition of S. N. BERNSTEIN:

$$\left| M\xi_{j}^{k} \right| \leq k! H_{2} K^{k-2} \mathfrak{D}\xi_{j}, \qquad j = 1, \dots, n$$

$$(10)$$

is satisfied for all $k \geq 3$, where H_2 and K are some positive numbers, then for

$$F(x) = P\{S_n < x\}$$

the relations (2) and (2') hold with m = 0, $\sigma = B_n$, $H = \frac{3}{2}$,

$$\varDelta = \frac{B_n}{\max \left\{ K(1+2H_2), \sqrt{2} \max \sigma_j \right\}}_{\substack{1 \le j \le n}}$$

and $\mathfrak{S}_k = \Gamma_{kn}$.

We set

$$z_0 = (\max\{K(1+2H_2), \sqrt{2}\max_{1 \le j \le n} \sigma_j\})^{-1}.$$

Then, taking into account (10), we obtain in the circle $|z| \leq z_0$

$$\varphi_{\xi_j}(z) = M e^{z \,\xi_j} = 1 + \sum_{k=2}^{\infty} \frac{z^k}{k!} M \xi_j^k = 1 + \theta |z|^2 \sigma_j^2.$$

Moreover, $\ln(1+w) = w + \theta |w|^2$ for $|w| \leq \frac{1}{2}$, and since $|z^2 \sigma_j^2| \leq \frac{1}{2}$, we have

$$\ln \varphi_{S_n}(z) = \sum_{\substack{j=1 \ j \in I}}^n \ln \varphi_{\xi_j}(z) = \theta |z|^2 B_n^2 + \theta |z|^4 B_n^2 \max_{\substack{1 \le j \le n}} \sigma_j^2 = \frac{3}{2} \theta |z|^2 B_n^2$$

 \mathbf{or}

$$\left| \ln \varphi_{S_n}(z) \right|_{|z|=z_0} \leq \frac{3}{2} z_0^2 B_n^2.$$

The inequality thus obtained shows that (1'') holds for $\sigma = B_n$, $H = \frac{3}{2}$, $\Delta = z_0 B_n$. This proves the theorem.

Limit theorems of Cramèr's type with calculation of large deviations for sums of independent random variables (9) had been proved by various authors (cf. H. CRAMÈR [4], V. V. PETROV [5], [6], W. RICHTER [7]) under the condition of CRAMÈR and PETROV: there exist positive numbers A, L and l such that

$$l \leq |Me^{z\xi_j}| \leq L$$
 for $|z| \leq A$, $j = 1, ..., n$.

In addition these authors assumed that $B_n^2 \ge c^2 n$ with some c > 0. We remark that in this case we can choose

$$H = \frac{\delta}{A^2 c^2}, \quad \Delta = A B_n \ge A c \sqrt{n},$$

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where $\mathscr{E} = 2\pi + \max\{|\ln l|, |\ln L|\}$ since

$$\left|\ln \varphi S_n(z)\right|_{|z|=A} \leq \sum_{j=1}^n \{\left|\ln \left|\varphi_{\xi_j}(z)\right|\right|_{|z|=A} + 2\pi\} \leq n \,\mathscr{E}.$$

Consider random variables X_1, \ldots, X_n , forming a Markov chain with n moments of time, transition probabilities $P_j(w, A)$ from the state w at time j into the measurable set of states A at time j + 1, initial probability distribution $P_0(A)$ and coefficients of ergodicity

$$\alpha_n = 1 - \max_{\substack{0 \leq j < n \\ \omega, \widetilde{\omega}, A}} \sup_{\omega, \widetilde{\omega}, A} (P_j(\omega, A) - P_j(\widetilde{\omega}, A)) |.$$

 \mathbf{Set}

$$S_n = \sum_{j=1}^n X_j, \quad B_n^2 = \mathfrak{D} S_n$$

Theorem 3. If with probability 1

$$|X_j| \leq C_n < \infty, \quad j = 1, 2, ..., n, \quad \alpha_n > 0$$

then there exists an absolute constant $H_3 > 0$, $H_4 > 0$ such that for $F(x) = P\{S_n < x\}$ the relations (2) and (2') hold with

$$m = M S_n, \quad \sigma = B_n, \quad H = H_3, \quad \Delta = \frac{\alpha_n B_n}{C_n H_4}$$

The *proof* of this theorem follows from the fact that the k-th cumulant \mathfrak{S}_k of the sum S_n can be estimated in the following way:

$$\left|\mathfrak{S}_{k}\right| \leq \frac{k! H_{3}C_{n}^{k-2}B_{n}^{2}H_{4}^{k-2}}{\alpha_{n}^{k-2}}, \quad k = 3, 4, \dots$$
 (11)

Let $\zeta(s, t)$ be a random function which is an additive function of the interval (s, t), that is,

$$\zeta(s, u) + \zeta(u, t) = \zeta(s, t)$$

with probability 1 for all $0 \leq s < u < t \leq T$

$$m(s,t) = M\zeta(s,t), \quad \sigma^2(s,t) = \mathfrak{D}\zeta(s,t),$$

e.g.

$$\zeta(s,t) = \int_{s}^{t} \xi(u) \, du \quad \text{or} \quad \zeta(s,t) = \sum_{s \le k \le t} \xi(k)$$

where $\xi(t)$ is some stochastic process. Let \mathscr{F}^t denote the σ -algebra generated by the events $\{\zeta(u, v) < x\}, s \leq u < v < t$.

Theorem 4. Assume that $\zeta(s, t)$ satisfies the following strong mixing condition of M. ROSENBLATT

$$\sup_{\substack{0 \le t \le T - \tau}} \sup_{\substack{A \in \mathfrak{T}_{c^{t}} \\ B \in \mathfrak{T}_{t+\tau}^{T}}} \left| P(A B) - P(A) P(B) \right| \le e^{-\alpha T t}$$

where $\alpha_T > 0$. In addition, assume that there are

$$\mathbf{l} \leq T_0 \leq \left[\frac{1}{\alpha_T}\right] + 1 \quad and \quad C_{T_0,T} < \infty$$

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such that

$$\frac{\left|\left|\zeta(s,s+T_0)\right|\right|}{T_0} \leq C_{T_0,T}$$

with probability 1 for all $0 \leq s \leq T - T_0$. Then there exists an absolute constant $H_5 > 0$, $H_6 > 0$ such that for $F(x) = P\{\zeta(0, T) < x\}$ the relations (2) and (2') hold with m = m(0, T), $\sigma = \sigma(0, T)$, $H = H_5$ and

$$\varDelta = \frac{\alpha_T^3 \sigma(0, T)}{C_{T_0, T} H_6 T_0}$$

The theorem results from the following inequality for cumulants of the random variables $\zeta(0, T)$:

$$\left|\mathfrak{S}_{k}\right| \leq \frac{k! H_{5} C_{T_{0}, T}^{k-2} \sigma^{2}(0, T) H_{6}^{k-2}}{\alpha_{T}^{3^{(k-2)}}} , \quad k = 3, 4, \dots$$
(12)

The proof of the inequalities (11) and (12) is involved and will not be given here.

Proof of the lemma. Without restricting the generality we set m = 0, $\sigma = 1$, $\Delta > 0$. On account of (1') the series

$$K(z) = \sum_{k=2}^{\infty} \frac{\mathfrak{S}_k}{k!} z^k = \frac{z^2}{2} \left(1 + \frac{2\theta |z| H}{\Delta \left(1 - \frac{|z|}{\Delta} \right)} \right)$$
(13)

converges for $|z| < \Delta$, and K(z) and $\tilde{\varphi}(z) = \exp\{K(z)\}$ are analytic in the circle $|z| < \Delta$. Since the moments $\mu_s = M \xi^s$ exist and can be expressed by the cumulants in closed form, we have $\mu_s = \frac{d^s \hat{\varphi}(z)}{dz^s} \Big|_{z=0}$ and, by (13) and Cauchy's inequality, we find that for

$$|\mu_s| \leq \frac{s!}{\delta_1^s \Delta^s} \exp\left\{\frac{\delta_1^s \Delta^2}{2} \left(1 + \frac{2H\delta_1}{(1-\delta_1)}\right)\right\}, \ s = 1, 2, \dots$$

Therefore the Laplace transform $\varphi(z) = M e^{z\xi}$ exists and is analytic in the circle $|z| < \Delta$ and there $\ln \varphi(z) = K(z)$.

Let us start from the transformation of ESSCHER [10] and CRAMER [4]. For arbitrary $0 \leq h < \Delta$ we have

$$1 - F(x) = \int_{x}^{\infty} dF(x) = \varphi(h) \int_{x}^{\infty} e^{-hy} dF_h(y), \qquad (14)$$

where the distribution function F_h of the random variable $\xi(h)$ is determined by the relation

$$dF_{h}(y) = \frac{e^{hy} dF(y)}{\varphi(h)}.$$
(15)

The main mass of the approximating distribution $1 - \Phi(x)$ is concentrated in the neighborhood of the point x. Therefore we should choose h in such a way that $\frac{1}{\varphi(h)} \exp\{hy\}$ takes its maximum at the point y = x, that is, h should be defined by the equation

$$x = \frac{d}{dh} \ln \varphi(h) = m(h).$$
(16)

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From (15) we easily find that $m(h) = M\xi(h)$ and $\sigma^2(h) = \frac{dm(h)}{dh} = \mathfrak{D}\xi(h)$. If \overline{F}_h is the distribution function of the normalised random variable

$$\overline{\xi}(h) = \frac{\xi(h) - m(h)}{\sigma(h)},$$

we derive from (14) and (16):

$$1 - F(x) = \varphi(h) e^{-hm(h)} \int_{0}^{\infty} e^{-h\sigma(h)y} d\bar{F}_{h}(y).$$
 (17)

In the following let $0 \leq h = \delta \Delta$ where $0 < \delta < \delta_H$, and δ_H is determined by the condition

$$\frac{6\,H\,\delta_H}{(1-\delta_H)^3} = 1\,. \tag{18}$$

For the characteristic function $f_h(t) = M \exp\{it\overline{\xi}(h)\}$ of the distribution function \overline{F}_h we have

$$f_h(t) \equiv e^{rac{-itm(h)}{\sigma(h)}} rac{\varphi\left(ht rac{it}{\sigma(h)}
ight)}{\varphi(h)}.$$

Expanding $\ln \varphi(z)$ in its Taylor series in the neighborhood of the point h given by $\left| h + \frac{|t|}{\sigma(h)} \right| \leq \delta_2 \Delta, \ \delta < \delta_2 < 1$ that is, by

$$|t| \leq \sigma(h) \left(\delta_2 - \delta\right) \Delta, \qquad (19)$$

we find

$$\ln f_{h}(t) = -it \frac{m(h)}{\sigma(h)} - \ln \varphi(h) + \ln \varphi\left(h + \frac{it}{\sigma(h)}\right) = = -\frac{t^{2}}{2} + \frac{(it)^{3}}{6\sigma^{3}(h)} (\ln \varphi(z))^{\prime\prime\prime} z = h + \frac{i\theta t}{\sigma(n)} = = -\frac{t^{2}}{2} + \theta \frac{|t|^{3}H}{\sigma^{3}(h)(1 - \delta_{2})^{4}\varDelta},$$
(20)

so that

$$(\ln \varphi(z))^{\prime\prime\prime} = \sum_{k=3}^{\infty} \frac{\mathfrak{S}_k z^{k-3}}{(k-3)!} = \theta \frac{H}{\Delta} \sum_{k=3}^{\infty} (k-2) (k-1) \cdot k \left(\frac{|z|}{\Delta}\right)^{k-3} = \frac{6 \theta H}{\Delta \left(1 - \frac{|z|}{\Delta}\right)^4}, \quad \text{if} \quad |z| < \Delta.$$

In the same way

$$\sigma^{2}(z) = (\ln \varphi(z))'' = 1 + \sum_{k=3}^{\infty} \frac{\mathfrak{S}_{k} z^{k-2}}{(k-2)!} = 1 + \frac{\theta H|z|}{\Delta} \sum_{k=3}^{\infty} (k-1) k \left(\frac{|z|}{\Delta}\right)^{k-3} = 1 + \frac{\theta H|z|}{\Delta \left(1 - \frac{|z|}{\Delta}\right)^{3}}$$

and

$$\sigma^2(h) = 1 + \theta \varrho \,. \tag{21}$$

We set

$$T_0 = \frac{1}{H} \sigma^3(h) (1 - \delta_2)^4 \varDelta, \quad T = \varepsilon T_0,$$
 (22)

where $0 < \varepsilon < \frac{1}{2}$ will be chosen later. Let $\delta_2(\delta < \delta_2 < 1)$ satisfy the relation

$$\frac{\delta_2 - \delta}{(1 - \delta_2)^4} = \frac{\varepsilon \sigma^2(h)}{H}.$$
(23)

Then, on account of (19), (22), (23) and the inequality $|e^{\alpha} - 1| \leq |\alpha| e^{|\alpha|}$, for $|t| \leq T$ we obtain

$$\left|f_{h}(t) - e^{-t^{2}/2}\right| \leq e^{-t^{2}/2} (e^{|t^{2}/2| + \ln f_{h}(t)|} - 1) \leq \frac{|t|^{3}}{T_{0}} e^{-(1/2 - \varepsilon)t^{2}}.$$
 (24)

Next we exploit an inequality which follows from a lemma of V. M. ZOLOTAREV ([13], lemma 3): let F be some distribution function, G a function of bounded variation with the properties

$$q = \sup |G'(x)| < \infty$$
, $G(-\infty) = 1 - G(\infty) = 0$

and f and g the Fourier-Stieltjes transforms of these functions. Then for all T > 0 and all $\lambda > A$ the inequality

$$\sup |F(x) - G(x)| \le 2q \frac{\lambda(s(\lambda) + Q(T))}{T(4s(\lambda) - \lambda)}$$
(25)

holds, where

$$s(\lambda) = \lambda \int_0^\lambda \frac{1 - \cos \lambda}{\pi \lambda^2} d\lambda.$$

A is a positive solution of the inequality $4s(\lambda) = \lambda$,

$$Q(T) = \frac{T}{2\pi q} \int_0^T \left(1 - \frac{t}{T}\right) \left|f(t) - g(t)\right| \frac{dt}{t}.$$

We apply the inequality (25) for an estimate of $\sup |\bar{F}_{\hbar}(y) - \Phi(y)|$. In our case $q = \frac{1}{\sqrt{2\pi}}$, $T = \varepsilon T_0$,

$$Q(T) \leq \frac{\varepsilon T_0}{\sqrt{2\pi}} \int_0^{\varepsilon T_0} f_h(t) - e^{-t^2/2} \left| \frac{dt}{t} < \frac{\varepsilon}{\sqrt{2\pi}} \int_0^{\infty} t^2 e^{-(1/2-\varepsilon)t^2} dt = \frac{\varepsilon}{2(1-2\varepsilon)^{\frac{1}{2}}},$$

hence for arbitrary $0 < \varepsilon < \frac{1}{2}$

$$\sup \left| \vec{F}_{\hbar}(y) - \Phi(y) \right| \leq \frac{2d(\varepsilon)}{\sqrt{2\pi} T_0}, \qquad (26)$$

where

$$d(\varepsilon) = \min_{\lambda > A} \frac{\frac{s(\lambda)}{\varepsilon} + \frac{1}{2(1-2\varepsilon)^{\frac{3}{2}}}}{4\frac{s(\lambda)}{\lambda} - 1}.$$
 (27)

A calculation shows that the minimum is attained for

$$\varepsilon = 0,282$$
, $\lambda = 3,600$

and equals 10,79.

Now we can approach the evaluation of the righ t-hand side of equation (17). We have

$$1 - F(x) = \varphi(h) e^{-hm(h)} I(h),$$

$$I(h) = I_1(h) + I_2(h),$$

$$I_1(h) = \int_0^{\infty} e^{-h\sigma(h)y} d(\bar{F}_h(y) - \Phi(y)),$$

$$I_2(h) = \int_0^{\infty} e^{-h\sigma(h)y} d\Phi(y).$$
(28)

By definition, h > 0 and $\sigma(h) > 0$, hence the estimate (26) allows to state that

$$|I_1(h)| \le |F_h(0) - \Phi(0)| + \sup |F_h(y) - \Phi(y)| \le \frac{4d}{\sqrt{2\pi} T_0}.$$
 (29)

Moreover

$$x = m(h) = h + \sum_{k=3}^{\infty} \frac{\mathfrak{S}_{k}}{(k-1)!} h^{k-1} = h + \frac{\theta H h^{3}}{\varDelta} \sum_{k=3}^{\infty} k \left(\frac{h}{\varDelta}\right)^{k-3} = h + \frac{3\theta H h^{2}}{\varDelta \left(1 - \frac{h}{\varDelta}\right)^{2}} = h \left(1 + \frac{\theta \varrho}{2} \left(1 - \delta\right)\right),$$

$$(30)$$

where, as before,

$$\varrho = \frac{6 H \delta}{(1-\delta)^3}.$$

Similarly,

$$h \sigma^{2}(h) - x = \sum_{k=3}^{\infty} \mathfrak{S}_{k} \left(\frac{1}{(k-2)!} - \frac{1}{(k-1)!} \right) h^{k-1} = \frac{\theta H h^{2}}{\Delta} \sum_{k=3}^{\infty} h (k-2) \left(\frac{h}{\Delta} \right)^{k-2} = \frac{\theta h \varrho}{2}.$$
(31)

The relations (30) and (31) show that

$$h\sigma(h) = x + \frac{\theta h\varrho}{2}.$$
(32)

In fact, in the case $\sigma(h) \leq 1$ they imply

$$x - rac{h \varrho}{2} \leq h \sigma^2(h) \leq h \sigma(h) \leq h \leq x + rac{h \varrho (1 - \delta)}{2}$$

and in the case $\sigma(h) \geq 1$ we have

$$x - \frac{h\varrho(1-\delta)}{2} \leq h \leq h\sigma(h) \leq h\sigma^2(h) \leq x + \frac{h\varrho}{2}$$
,

which gives (32). Therefore

$$I_2(h) = \frac{1}{\sqrt{2\pi}} \int_0^\infty e^{-xy - y^2/2} \left(1 + \theta \sum_{k=1}^\infty \frac{h^k \varrho^k y^k}{2^k k!} \right) dy = e^{x^2/2} (1 - \Phi(x)) + R$$

where by the equality

$$\int_{0}^{\infty} e^{-xy} y^k dy = \frac{k!}{x^{k+1}}$$

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and the estimate

$$rac{harrho}{2x}\!\leq\!rac{arrho}{2-arrho(1-\delta)}\!<\!rac{arrho}{1+\delta}\!<\!1$$
 ,

derived from (30), we obtain

$$|R| \leq \sum_{k=1}^{\infty} \frac{1}{\sqrt{2\pi x}} \left(\frac{h\varrho}{2x}\right)^k \leq \frac{1}{\sqrt{2\pi x}} \frac{\varrho}{2(1-\varrho)+\varrho\delta}.$$

If $x \ge 1$, then

$$\max\left\{\frac{1}{\sqrt{2\pi x}}\left(1-\frac{1}{x^{2}}\right),\frac{3}{4}\frac{1}{\sqrt{2\pi x}}\right\} \leq (1-\Phi(x))e^{x^{2}/2} \leq \frac{1}{\sqrt{2\pi x}},$$
(33)

thus we finally find

$$I_{2}(h) = e^{x^{2}/2} \left(1 - \Phi(x)\right) \left(1 + \frac{4}{3} \frac{\theta \varrho}{2(1 - \varrho) + \varrho \delta}\right).$$
(34)

The relations (28), (29), (22), (23), (33) and (34) lead to

$$\frac{1-F(x)}{1-\Phi(x)} = \varphi(h) e^{x^2/2 - hx} \left((1+f(\delta, H)\frac{x}{\Delta}), \right)$$
(35)

where

$$|f(\delta, H)| \leq \frac{16dH}{3(1-\varrho)^{\frac{3}{2}}(1-\delta_2)^4} + \frac{4\Delta}{3x} \frac{\varrho}{2(1-\varrho)+\varrho\delta}$$

Remembering that

$$\frac{\Delta}{x} = \frac{\hbar}{\delta x} \leq \frac{1}{\delta \left(1 - \frac{\varrho}{2} \left(1 - \delta\right)\right)}, \quad \varrho = \frac{6 H \delta}{(1 - \delta)^3},$$

we can also write

$$|f(\delta, H)| \leq \frac{16dH}{3(1-\varrho)^{\frac{3}{2}}(1-\delta_2)^4} + \frac{4H}{(1-\delta)^3 \left(1-\frac{\varrho}{2}+\frac{\varrho\delta}{2}\right) \left(1-\varrho+\frac{\varrho\delta}{2}\right)}.$$
 (36)

We had set $\delta = \frac{h}{\Delta}$. However, the relations (35), (36) remain true if we select h and δ arbitrarily, but so as to satisfy the conditions $0 < h \leq \delta \Delta$, $\delta < \delta_H$. Since then $\sigma^2(h) = \frac{d}{dh} m(h) > 0$ by the equation (16) x = m(h) there corresponds to each value x a value h. In order to fulfil the inequality $0 < h \leq \delta \Delta$ it is necessary to consider values x which satisfy the condition

$$0 < x \leq \overline{\delta} \Delta$$
, $\overline{\delta} < \overline{\delta}_H$, (37)

where

$$\overline{\delta} = rac{\delta}{2} \left(2 - \varrho + \varrho \, \delta
ight)$$
 ,

because $x = h\left(1 + \frac{\theta \varrho}{2}(1 - \delta)\right)$, by (30).

Thus

$$f_{1}(\bar{\delta}, H) = f(\delta, H) = \frac{16 \,\theta \,d(\varepsilon) H}{3 \,(1-\varrho)^{\frac{3}{2}} (1-\delta_{2})^{4}} + \frac{4 \,\theta \,H}{(1-\delta)^{3} \left(1-\frac{\varrho}{2}+\frac{\varrho \,\delta}{2}\right) \left(1-\varrho+\frac{\varrho \,\delta}{2}\right)},$$
(38)

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where δ and $\overline{\delta}$ are related by the equality $\overline{\delta} = \frac{\delta}{2}(2 - \rho + \delta \rho)$, δ_2 with $\delta < \delta_2 < 1$ is the positive solution of the equation

$$\frac{\delta_2-\delta}{(1-\delta_2)^4} = \frac{\varepsilon \sigma^2(h)}{H}, \ \sigma^2(h) \ge 1-\varrho, \quad 0 < \varepsilon < \frac{1}{2}$$

and $d(\varepsilon)$ fulfills the inequality (27).

In passing we make the following remark. If, in order to estimate Q(T), we use instead of (24) the inequality

$$|f_h(t) - e^{-t^2/2}| \le e^{-t^2/2} (e^{|t|^3/T_0} - 1) \le \frac{|t|^3}{T_0} e^{-t^2/2} + \frac{|t|^6}{T_0^2} e^{-(1/2-\varepsilon)t^2}$$

then we obtain

$$d(\varepsilon) \leq \frac{\varepsilon^{-1}s(\lambda) + \frac{1}{2} + (2\sqrt{2\pi T_0}(\frac{1}{2} - \varepsilon)^3)^{-1}}{4\lambda^{-1}s(\lambda) - 1},$$

for arbitrary $0 < \varepsilon < \frac{1}{2}$ and $\lambda > A$.

If we chose λ so as to minimize the expression

$$\frac{2 \, s \left(\lambda\right) + \frac{1}{2}}{4 \, \lambda^{-1} s \left(\lambda\right) - 1},$$

then $\lambda = 3,2467$, thus

$$d(\varepsilon) \leq \frac{2,23}{\varepsilon} + 0.87 + \frac{0.35 H}{(1-\varrho)^{\frac{3}{2}} (1-\delta_2)^4 (\frac{1}{2}-\varepsilon)^3 \varDelta}.$$

This estimate is valuable for large Δ and relatively small x. Note that

$$f_1(\bar{\delta}, H) | \leq 4 H \left\{ 1 + \frac{5,33 x^2}{x^2 - 1} \left(1 + \min \left\{ \frac{1}{2 H}, \frac{1}{\sqrt[4]{2 H}} \right\} \right)^4 + o(1) \right\},$$

if $\Delta \to \infty$ and $x = o(\Delta)$.

Continuing the proof, it is not difficult to verify that

$$\frac{1}{(1-\delta_2)^4} \leq \frac{(1+6\varepsilon\delta + \min\{\varepsilon^{\frac{1}{4}}(1-\delta)^{\frac{3}{4}}H^{-\frac{1}{4}}, \varepsilon(1-\delta)^{\frac{3}{4}}H^{-1}\})^4}{(1-\delta)^4}$$

Therefrom we obtain a more crude, but in turm clearer estimate:

$$|f_1(\bar{\delta}, H)| \leq \frac{8H(1 + \frac{2}{3}d(\varepsilon)(\Delta + 6\varepsilon\delta + \min\{\varepsilon^{\frac{1}{4}}(1 - \delta)^3 H^{-1}, \varepsilon^{\frac{3}{4}} H^{-\frac{1}{4}}\})^4)}{(1 - \delta)^4 (1 - \varrho)^{\frac{3}{2}}}.$$
 (39)

Obviously, $\overline{\delta} > \frac{\delta(1+\delta)}{2}$, and only for $\delta = \delta_H$, that is for $\varrho = 1$, the equality $\overline{\delta}_H = \frac{\delta_H(1+\delta_H)}{2}$ is attained. In order to simplify the relation between δ and $\overline{\delta}$ in the formulation of the lemma (cf. (4)), we had set $\overline{\delta} = \frac{\delta(1+\delta)}{2}$. Therefore in the evaluation of the right-hand side of the inequality (39) we have to replace $1 - \delta$ by $1 - \overline{\delta}$. Having done so, and putting $\varepsilon = 0.282$, $d(\varepsilon) = 10.79$, we obtain (3).

It remains to consider $L(x) = \frac{x^2}{2} - hx + \ln \varphi(h)$ in (35). It follows from what we have said earlier that h can be expanded in a power series in x which converges for $|x| < \overline{\delta}_H \Delta$:

$$h = h(x) = \sum_{k=1}^{\infty} a_k x^k$$
 (40)

The coefficients a_k are determined by the first k + 1 cumulants. Moreover, from Cauchy's inequality for the coefficients of a power series we find

$$\left|a_{k}\right| \leq \frac{\delta_{H}}{\delta_{H}^{k} d^{k-1}}, k = 1, 2, \dots,$$

$$(41)$$

because $|h(z)|_{|z|=\overline{\delta}H\varDelta} \leq \delta_H \varDelta$ by (30). It is easy to verify that

$$a_{1} = 1, \quad a_{2} = -\frac{\mathfrak{S}_{3}}{2}, \quad a_{3} = -\frac{\mathfrak{S}_{4} - 3 \,\mathfrak{S}_{3}^{2}}{6}, \\ a_{4} = -\frac{\mathfrak{S}_{5} - 10 \,\mathfrak{S}_{4} \mathfrak{S}_{3} + 15 \,\mathfrak{S}_{3}^{2}}{24}, \dots$$

$$(42)$$

Using the expansion (40) we obtain

$$L(x) = \frac{x^2}{2} - hx + \sum_{k=2}^{\infty} \frac{\mathfrak{S}_k}{k!} h^k = \frac{x^2}{2} - \sum_{k=2}^{\infty} \frac{k-1}{k!} \mathfrak{S}_k h^k = \sum_{k=3}^{\infty} C_k x^k,$$

where

$$C_{k} = -\sum_{\nu=2}^{k} \frac{\nu-1}{\nu!} \mathfrak{S}_{\nu} \sum_{k_{1}+\dots+k_{\nu}=k} a_{k_{1}}\dots a_{k_{\nu}} = -\frac{a_{k-1}}{k}$$

We set

$$L(x) = \frac{x^3}{\varDelta} \lambda\left(\frac{x}{\varDelta}\right).$$

The inequality (41) shows that the coefficients of Cramèr's series

$$\lambda(t) = \sum_{k=0}^{\infty} \lambda_k t^k$$

are subject to the estimate (5):

$$\lambda_k = -\frac{a_{k+2}}{k+3} \Delta^{k+1} = \theta \frac{\delta_H}{(k+3)\,\overline{\delta}_{H}^{k+2}}.$$
(43)

Hence

$$\varphi(h)e^{x^2/2 - hx} = \exp L(x) = \exp \left\{\frac{x^3}{\Delta}\lambda\left(\frac{x}{\Delta}\right)\right\}$$

which together with (51) concludes the proof of the relation (2) as well as the proof of the lemma, because the proof of (2') runs analogously. The estimate (6) for $\overline{\delta}_H$ from above is verified by (4) and a simple calculation. The validity of the remark is obvious, because it folloews from the inequality (1'') that the function $\ln \varphi(z)$ is analytic in the circle $|z| < \frac{\Delta}{\sigma}$, and therefore, by Cauchy's inequalities, (1') is satisfied.

We remark that limit theorems on deviations of type Δ for sums of independent random variables and additive functions had been considered in [8].

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