

A New Proof of Kesten's Theorem on the Growth of the Sum of Independent and Identically Distributed Random Variables

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0. In this paper, a more direct and probabilistically intuitive proof of Kesten's theorem on the growth of the sum of independent and identically distributed random variables ([2]) is presented. The technique used is a modified version of "centering" employed by Wolfowitz ([4]) and Kesten ([3]).

1. Theorem (Kesten [2], Corollary 3). *Let $\{X_i\}_{i=1}^{\infty}$ be a sequence of independent and identically distributed random variables with $EX_i^+ = EX_i^- = +\infty$. Then either*

$$(i) \quad \lim_{n \rightarrow \infty} \frac{S_n}{n} = +\infty \quad \text{w.p.l, or}$$

$$(ii) \quad \lim_{n \rightarrow \infty} \frac{S_n}{n} = -\infty \quad \text{w.p.l, or}$$

$$(iii) \quad \limsup_{n \rightarrow \infty} \frac{S_n}{n} = +\infty \quad \text{and} \quad \liminf_{n \rightarrow \infty} \frac{S_n}{n} = -\infty \quad \text{w.p.l,}$$

where $S_n = \sum_{i=1}^n X_i$, $n = 1, 2, \dots$

Proof. Let $D = \liminf_{n \rightarrow \infty} S_n/n$ where $-\infty \leq D \leq +\infty$. By the Kolmogorov Zero-One Law, D is constant w.p.l. To prove the theorem, it suffices to show that $D = +\infty$ or $D = -\infty$. Suppose $-\infty < D < +\infty$. Without loss of generality we may assume $D > 0$; for otherwise we may replace X_i by $Y_i = X_i - D + \frac{1}{2}$ and $H_n = \sum_{i=1}^n Y_i$ by noticing that $\liminf_{n \rightarrow \infty} H_n/n = \frac{1}{2} > 0$ and $\liminf_{n \rightarrow \infty} H_n/n$ is finite iff $\liminf_{n \rightarrow \infty} S_n/n$ is finite.

Consider the two-sided sequence $\{X_i\}_{i=-\infty}^{+\infty}$ which is the extension of $\{X_i\}_{i=1}^{\infty}$ into "the past" ([1] Ch. 6). Define $\tilde{S}_k = \sum_{i=-k+1}^0 X_i$ and $\tilde{S} = 0 = S_0$. It follows that

* Supported under NSF grant MPS-75-07228

$\liminf_{n \rightarrow \infty} \tilde{S}_n/n = D$ w.p.l. Define stopping times $\gamma_0 \equiv 0$,

$$\begin{aligned} \gamma_1 &= \min \{n \geq 1: \tilde{S}_n > 0\} \\ &= +\infty \text{ if no such } n \text{ exists.} \end{aligned} \tag{1}$$

Since $D > 0$, $\tilde{S}_n \rightarrow +\infty$ as $n \rightarrow \infty$ so $\gamma_1 < \infty$ w.p.l. In the same manner, define

$$\begin{aligned} \gamma_k &= \min \{n > \gamma_{k-1}: \tilde{S}_n - \tilde{S}_{\gamma_{k-1}} > 0\} \\ &= +\infty \text{ if no such } n \text{ exists for } k=2, 3, \dots \end{aligned} \tag{2}$$

Now $\{\tilde{S}_{\gamma_k} - \tilde{S}_{\gamma_{k-1}}\}_{k=1}^\infty$ are i.i.d. random variables and are positive, so the Birkhoff Ergodic Theorem implies that

$$\begin{aligned} \lim_{n \rightarrow \infty} \tilde{S}_{\gamma_n}/n &= \lim_{n \rightarrow \infty} \frac{1}{n} \sum_{k=1}^n (\tilde{S}_{\gamma_k} - \tilde{S}_{\gamma_{k-1}}) \\ &= E(\tilde{S}_{\gamma_1} - \tilde{S}_{\gamma_0}) \text{ w.p.l.} \end{aligned}$$

To conclude the proof of the theorem, we need the following facts:

$$\liminf_{n \rightarrow \infty} X_{-\gamma_n-1}/n = -\infty, \tag{3}$$

$$0 < E(\gamma_1 - \gamma_0) < +\infty, \text{ and} \tag{4}$$

$$0 < E(\tilde{S}_{\gamma_1} - \tilde{S}_{\gamma_0}) < \infty. \tag{5}$$

For (4) implies that $\lim_{n \rightarrow \infty} \tilde{S}_{\gamma_n}/\gamma_n$ exists and then (3) and (5) imply that

$\lim_{n \rightarrow \infty} \tilde{S}_{\gamma_n+1}/(\gamma_n+1) = -\infty$ which would contradict the fact that D is finite.

Now $X_{-\gamma_k-1} = \tilde{S}_{\gamma_{k+1}} - \tilde{S}_{\gamma_k}$ so that $\{X_{-\gamma_k-1}\}_{k=1}^\infty$ is an i.i.d sequence of random variables with the distribution of $X_{-\gamma_1-1}$ the same as the distribution of X_1 ([1]). Hence $EX_{-\gamma_k-1}^+ = EX_{-\gamma_k-1}^- = +\infty$ and so (3) holds.

To prove (4) and (5) we adopt the notation of Kesten ([3]). In particular

$$\begin{aligned} v_0 &= \min \{j \geq 0: \inf_{n > j} S_n - S_j > 0\} \text{ and} \\ v_{k+1} &= \min \{j > v_k: \inf_{n > j} S_n - S_j > 0\}. \end{aligned} \tag{6}$$

Since $D > 0$, $\lim_{n \rightarrow \infty} S_n = +\infty$ w.p.l so that $v_0 < \infty$ and $v_{k+1} < \infty$. Furthermore, as was shown in ([3]), $q \equiv P(v_0 = 0) > 0$. This allows us to define a new probability measure Q on the Borel subsets of \mathbb{R} by

$$Q(A) = \frac{1}{q} P(A, v_0 = 0). \tag{7}$$

Let

$$v^* = \min \{m \geq 1: \sum_{-m+1 \leq i \leq v_0} X_i > 0\}. \tag{8}$$

Then since $\{X_0, X_{-1}, \dots\}$ and $\{X_1, X_2, \dots\}$ generate independent σ -fields

$$\begin{aligned}
 P(\gamma_1 = k) &= \frac{P(\gamma_1 = k, v_0 = 0)}{P(v_0 = 0)} \\
 &= \frac{1}{q} P\left(\sum_{-j+1 \leq i \leq 0} X_i \leq 0 \text{ for } 1 \leq j < k, v_0 = 0, \sum_{j=k} X_i > 0, v_0 = 0\right) \\
 &= \frac{1}{q} P(v^* = k, v_0 = 0) \\
 &= \frac{1}{q} P(v_0 = 0, v_1 = k) = Q(v_1 - v_0 = k). \tag{9}
 \end{aligned}$$

Thus

$$0 < \int \gamma_1 dP = \int (v_1 - v_0) dQ = \frac{1}{q} < \infty. \tag{10}$$

where the last equality is a simple result which may be found in [3].

To prove (5) we use ince again the independence of the σ -fields generated by X_0, X_{-1}, \dots and X_1, X_2, \dots respectively to obtain

$$\begin{aligned}
 P(\tilde{S}_{\gamma_1} - \tilde{S}_{\gamma_0} \in A) &= \frac{1}{q} P(\tilde{S}_{\gamma_1} - \tilde{S}_{\gamma_0} \in A, v_0 = 0) \\
 &= \frac{1}{q} \sum_{l=1}^{\infty} P(v^* = l, v_0 = 0, \tilde{S}_l \in A), \tag{11}
 \end{aligned}$$

where A is a Borel set.

Now increasing the indices by l on the set

$$\{v^* = l, v_0 = 0, \tilde{S}_l \in A\}$$

yields the set

$$\{v_1 = l, v_0 = 0, S_l \in A\}.$$

From this and (10) it follows that

$$P(\tilde{S}_{\gamma_1} - \tilde{S}_{\gamma_0} \in A) = Q(S_{v_1} - S_{v_0} \in A) \tag{12}$$

Therefore

$$0 \leq \int (\tilde{S}_{\gamma_1} - \tilde{S}_{\gamma_0}) dP = \int (S_{v_1} - S_{v_0}) dQ \leq +\infty. \tag{13}$$

Now by [3], $\lim_{n \rightarrow \infty} S_{v_n}/k \equiv \beta$ exists w.p.1, $\beta > 0$, and $E(\beta) = \int (S_{v_1} - S_{v_0}) dQ$. Suppose $E(\beta) = +\infty$. Then for any integer $K > 0$ $Q(\beta/\alpha > K) > 0$, where $\alpha \equiv \lim_{n \rightarrow \infty} v_n/k$, since if $\beta/\alpha < K$ w.p.1 Q , then

$$EQ(\beta) < KE_Q(\alpha) = K \int (v_1 - v_0) dQ < \infty. \tag{14}$$

Since we assumed $D < \infty$, we may choose $K_0 > D$. Then since $P(\liminf_{n \rightarrow \infty} S_n/n > K_0) = 0$ and $Q \ll P$, it follows that

$$\begin{aligned} 0 &= Q(\liminf_{n \rightarrow \infty} S_n/n > K_0) \\ &\geq Q(\liminf_{n \rightarrow \infty} S_n/n > K_0, \beta/\alpha > K_0) \\ &= Q(\beta/\alpha > K_0) \geq 0. \end{aligned} \tag{15}$$

Hence $E(\beta) < \infty$ and so (5) is proved. As stated earlier, this, in conjunction with (3) and (4), proves the theorem.

References

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Received May 8, 1976