

An Invariance Principle for the Law of the Iterated Logarithm*

By

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Summary

Let S_n be the sum of the first n of a sequence of independent identically distributed r. v. s. having mean 0 and variance 1. One version of the law of the iterated logarithm asserts that with probability one the set of limit points of the sequence

$$\left((2n \log \log n)^{-1/2} S_n \right)_{n \geq 3}$$

coincides with $\langle -1, 1 \rangle = \{x : x \text{ real and } |x| \leq 1\}$ (see HARTMAN-WINTNER [6]). Now consider the continuous function η_n on $\langle 0, 1 \rangle$ obtained by linearly interpolating $(2n \log \log n)^{-1/2} S_i$ at i/n . Then we prove (theorem 3) that with probability one the set of limit points of the sequence $(\eta_n)_{n \geq 3}$ with respect to the uniform topology coincides with the set of absolutely continuous functions x on $\langle 0, 1 \rangle$ such that

$$x(0) = 0$$

and

$$\int \dot{x}^2 dt \leq 1.$$

As applications we obtain, e. g.,

$$Pr \left\{ \limsup_{n \rightarrow \infty} n^{-1-(a/2)} (2 \log \log n)^{-(a/2)} \sum_{i=1}^n |S_i|^a = \frac{2(a+2)^{(a/2)-1}}{\left(\int_0^1 \frac{dt}{\sqrt{1-t^a}} \right)^a a^{a/2}} \right\} = 1$$

for any $a \geq 1$, and

$$Pr \left\{ \limsup_{n \rightarrow \infty} v_n = 1 - \exp \left\{ -4 \left(\frac{1}{c^2} - 1 \right) \right\} \right\} = 1,$$

where v_n is the frequency of the events

$$S_i > c(2i \log \log i)^{1/2}$$

among the first n integers i ($0 \leq c \leq 1$).

To prove theorem 3, we first derive an analogous result for the (k -dimensional) Brownian Motion using well-known ideas of KOLMOGOROV [7] and of ERDÖS and KAC [4]. Also CHUNG's profound paper [1] is to be mentioned here. Then we prove

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an invariance principle for the law of the iterated logarithm by a powerful device of SKOROKHOD's [11], designed by him to yield improvements of the ordinary invariance principle. The paper assumes no knowledge of the classical law of the iterated logarithm.

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1. Brownian Motion

Let ζ be the Brownian Motion in R^k . Define

$$\zeta_n(t) = (2n \log \log n)^{-1/2} \zeta(nt)$$

for $t \in \langle 0, 1 \rangle$ and $n \geq 3$. Let C be the Banach space of continuous maps from $\langle 0, 1 \rangle$ to R^k endowed with the supremum norm $\| \cdot \|$, using the euclidean norm in R^k . ζ_n is then a r. v. with values in C . Let K be the set of absolutely continuous $x \in C$ such that

$$x(0) = 0$$

and

$$\int_0^1 (\dot{x}(t))^2 dt \leq 1$$

(where \dot{x} denotes the derivative of x determined almost everywhere with respect to Lebesgue measure and the square is meant as inner product). K is a norm-compact subset of C : In fact for $a \leq b$

$$\begin{aligned} (1) \quad |x(b) - x(a)| &= \left| \int_a^b \dot{x} dt \right| \leq \\ &\leq \left(\int_a^b dt \int_a^b \dot{x}^2 dt \right)^{1/2} \leq \\ &\leq \sqrt{b-a}, \end{aligned}$$

so that K is relatively norm-compact. That K is closed follows immediately from a result of F. RIESZ (see [10], p. 68, lemma).

If one considers an $x \in C$ as the motion of a mass point with mass 2 from time 0 to time 1, then K consists of those motions for which the mean kinetic energy is ≤ 1 . This interpretation has a nice feature: The obscure factor $2^{-1/2}$ in the definition of ζ_n which has been kept for historical reasons cancels with the factor 2 of the mass of our masspoint, if one connects them by theorem 1. A similar remark applies to theorem 3.

Theorem 1. *With probability one the sequence $(\zeta_n)_{n \geq 3}$ is relatively norm-compact and the set of its (norm-)limit points coincides with K .*

Proof. Let $\varepsilon > 0$, K_ε be the set of all points in C which have a distance $< \varepsilon$ from K . Then if m is any positive integer and $r > 1$ real

$$\begin{aligned} \Pr \{ \zeta_n \notin K_\varepsilon \} &\leq \Pr \left\{ 2m \sum_{i=1}^{2m} \left(\zeta_n \left(\frac{i}{2m} \right) - \zeta_n \left(\frac{i-1}{2m} \right) \right)^2 > r^2 \right\} + \\ + \Pr \left\{ 2m \sum_{i=1}^{2m} \left(\zeta_n \left(\frac{i}{2m} \right) - \zeta_n \left(\frac{i-1}{2m} \right) \right)^2 \leq r^2 \quad \text{and} \quad \zeta_n \notin K_\varepsilon \right\} &= I + II \quad (\text{say}). \end{aligned}$$

Now

$$\begin{aligned} \text{I} &= \Pr \left\{ \chi_{2mk}^2 > 2r^2 \log \log n \right\} = \frac{1}{\Gamma(mk)} \int_{r^2 \log \log n}^{\infty} t^{mk-1} e^{-t} dt \\ &\sim \frac{(r^2 \log \log n)^{mk-1} e^{-r^2 \log \log n}}{\Gamma(mk)} \end{aligned}$$

for $n \rightarrow \infty$ (recall that k is the dimension of the Brownian Motion ζ). Moreover, let η_n be the r. v. with values in C obtained by linearly interpolating the points $\zeta_n \left(\frac{i}{2m} \right)$ at $\frac{i}{2m}$ ($i = 1, \dots, 2m$). Then

$$2m \sum_{i=1}^{2m} \left(\zeta_n \left(\frac{i}{2m} \right) - \zeta_n \left(\frac{i-1}{2m} \right) \right)^2 \leq r^2$$

means just $\frac{1}{r} \eta_n \in K$ and we get

$$\text{II} = \Pr \left\{ \frac{1}{r} \eta_n \in K, \zeta_n \notin K_\varepsilon \right\} \leq \Pr \left\{ \frac{1}{r} \eta_n \in K, \left\| \frac{1}{r} \eta_n - \zeta_n \right\| \geq \varepsilon \right\}.$$

Define the r. v. T by

$$T = \begin{cases} \min \left\{ t : t \in \langle 0, 1 \rangle, \left| \frac{1}{r} \eta_n(t) - \zeta_n(t) \right| \geq \varepsilon \right\} & \text{if this set is nonempty} \\ 2 & \text{otherwise,} \end{cases}$$

and let F be its d. f., so that

$$\begin{aligned} \text{II} &\leq \int_0^1 \Pr \left\{ \frac{1}{r} \eta_n \in K \mid T = t \right\} dF(t) \\ &= \int_0^1 \Pr \left\{ \frac{1}{r} \eta_n \in K, \left| \frac{1}{r} \eta_n(t) - \zeta_n(t) \right| = \varepsilon \mid T = t \right\} dF(t). \end{aligned}$$

If $i(t)$ is the smallest i with $\frac{i}{2m} \geq t$, the statement

$$\frac{1}{r} \eta_n \in K$$

implies

$$\left| \eta_n \left(\frac{i(t)}{2m} \right) - \eta_n(t) \right| \leq r \int_t^{i(t)/2m} \left| \frac{1}{r} \dot{\eta}_n \right| ds \leq r \int_t^{i(t)/2m} \left(\frac{1}{r} \dot{\eta}_n \right)^2 ds \frac{1}{\sqrt{2m}} \leq \frac{r}{\sqrt{2m}}.$$

The two statements

$$\frac{1}{r} \eta_n \in K$$

and

$$\left| \frac{1}{r} \eta_n(t) - \zeta_n(t) \right| = \varepsilon$$

together therefore imply

$$\begin{aligned} \left| \zeta_n \left(\frac{i(t)}{2m} \right) - \zeta_n(t) \right| &\geq \left| \eta_n(t) - \zeta_n(t) \right| - \left| \eta_n \left(\frac{i(t)}{2m} \right) - \eta_n(t) \right| \\ \left(\text{remember that } \eta_n \left(\frac{i(t)}{2m} \right) &= \zeta_n \left(\frac{i(t)}{2m} \right) \right) \\ &\geq r\varepsilon - (r-1) - \frac{r}{\sqrt{2m}} \geq \frac{\varepsilon}{2} \end{aligned}$$

if r is chosen close enough to 1 and m is sufficiently large. Then

$$\begin{aligned} \text{II} &\leq \int_0^1 \Pr \left\{ \left| \zeta_n \left(\frac{i(t)}{2m} \right) - \zeta_n(t) \right| \geq \frac{\varepsilon}{2} \mid T = t \right\} dF(t) \leq \Pr \left\{ \left| \zeta_n \left(\frac{1}{2m} \right) \right| \geq \frac{\varepsilon}{2} \right\} \int_0^1 dF(t) \\ &\leq \Pr \left\{ \left| \zeta \left(\frac{n}{2m} \right) \right| \geq \frac{\varepsilon}{2} \sqrt{2n \log \log n} \right\} \sim \Gamma \left(\frac{k}{2} \right)^{-1} ((\varepsilon^2 m \log \log n)/2)^{k/2-1} \times \\ &\quad \times \exp \{ (-\varepsilon^2 m \log \log n)/2 \} \end{aligned}$$

(compare the estimation of I). By choosing m and r appropriately and using the above estimates for I and II, it is easily seen that for some $r > 1$ and sufficiently large n

$$\Pr \{ \zeta_n \notin K_\varepsilon \} \leq e^{-r^2 \log \log n}.$$

If $n_j = [c^j] + 1$, where $c > 1$, then

$$\sum_j \Pr \{ \zeta_{n_j} \notin K_\varepsilon \} \leq (\log c)^{-r^2} \sum_j j^{-r^2} < \infty$$

so that eventually $\zeta_{n_j} \in K_\varepsilon$ with probability 1. For c sufficiently close to 1 this implies that eventually $\zeta_n \in K_{2\varepsilon}$ with probability one.

This shows that almost surely at most the points of K are limit points of $(\zeta_n)_{n \geq 3}$ and also that almost surely this sequence is relatively compact (for $\{\zeta_n : n \geq 3\}$ is totally bounded). To prove the theorem it is therefore sufficient (because of the compactness of K) to show the following: given $x \in K$ and $\varepsilon > 0$, the probability that ζ_n is infinitely often in the open ε -sphere $\{x\}_\varepsilon$ around x equals one. Let $m \geq 1$ be an integer, $0 < \delta < 1$ and x^κ, ζ_n^κ be the κ -th coordinate of x and ζ_n respectively ($1 \leq \kappa \leq k$). We denote the event

$$\left\{ \left| \zeta_n^\kappa \left(\frac{i}{m} \right) - \zeta_n^\kappa \left(\frac{i-1}{m} \right) - \left(x^\kappa \left(\frac{i}{m} \right) - x^\kappa \left(\frac{i-1}{m} \right) \right) \right| < \delta \right. \\ \left. \text{for all } i \text{ with } 2 \leq i \leq m \text{ and all } \kappa \right\}$$

by A_n . Then

$$\begin{aligned} \Pr(A_n) &\geq \prod_{i=2}^m \prod_{\kappa=1}^k \frac{1}{\sqrt{2\pi}} \frac{\int_{\sqrt{2m \log \log n} \left(\left| x^\kappa \left(\frac{i}{m} \right) - x^\kappa \left(\frac{i-1}{m} \right) \right| + \delta \right)}^{\infty} e^{-(s^2/2)} ds}{\sqrt{2m \log \log n} \left| x^\kappa \left(\frac{i}{m} \right) - x^\kappa \left(\frac{i-1}{m} \right) \right|} \\ &\geq \text{const.} \prod_{i=2}^m \prod_{\kappa=1}^k \frac{\exp \left\{ -m \left(x^\kappa \left(\frac{i}{m} \right) - x^\kappa \left(\frac{i-1}{m} \right) \right)^2 \log \log n \right\}}{\sqrt{m \log \log n}} \end{aligned}$$

for n sufficiently large (having used $\int_a^b \frac{e^{-(s^2/2)}}{\sqrt{2\pi}} ds \geq \frac{1}{b\sqrt{2\pi}} e^{-(a^2/2)} (1 - e^{-1/2(b^2 - a^2)})$

for $0 \leq a < b$). So by summing up the exponents and using Schwarz's inequality we get

$$\Pr(A_n) \geq \frac{\text{const.}}{\log n \sqrt{m \log \log n}}$$

for large n ("large" depending on m and δ). We now put $n_j = m^j$. Then the A_{n_j} 's

are mutually independent and

$$\sum_{j=1}^{\infty} \Pr(A_{n_j}) = \infty$$

because $\sum_j \frac{1}{j\sqrt{\log j}}$ diverges. Hence by Borel-Cantelli's lemma infinitely many events A_n happen almost surely. By what we previously have proved ζ_n is eventually close to K , and therefore almost surely we have eventually

$$(2) \quad |\zeta_n(t) - \zeta_n(s)| \leq \sqrt{|t-s|} + \delta$$

for any $s, t \in \langle 0, 1 \rangle$. Now if $y \in C$ the two statements

$$|y(t) - y(s)| \leq \sqrt{|t-s|} + \delta$$

for all $s, t \in \langle 0, 1 \rangle$ and

$$\left| y^{\varkappa} \left(\frac{i}{m} \right) - y^{\varkappa} \left(\frac{i-1}{m} \right) - \left(x^{\varkappa} \left(\frac{i}{m} \right) - x^{\varkappa} \left(\frac{i-1}{m} \right) \right) \right| < \delta$$

for all i and \varkappa with $2 \leq i \leq m$ and $1 \leq \varkappa \leq k$ together imply

$$\|y - x\| < \varepsilon,$$

provided m is sufficiently large and δ is sufficiently small ("small" depending also on the choice of m).

Looking at the definition of A_n , at the fact that A_n happens infinitely often a. s. and at (2) we conclude that

$$\Pr\{\|\zeta_n - x\| < \varepsilon \text{ infinitely often}\} = 1.$$

This proves the theorem.

The discreteness of n is inessential for the previous considerations. So if $u > e$ (base for the natural logarithm) is real and we put

$$\zeta_u(t) = (2u \log \log u)^{-1/2} \zeta(ut)$$

for $t \in \langle 0, 1 \rangle$, we have the following

Corollary 1. *With probability one the net $(\zeta_u)_{u>e}$ is relatively norm-compact and the set of its norm-limit points as u tends to ∞ coincides with K .*

2. The invariance principle

Let Y_1, Y_2, \dots be a sequence of independent and identically distributed r. v. s. Assume

$$E Y_1 = 0$$

and

$$E(Y_1^2) = 1$$

(no further moments are needed). Put

$$S_n = \sum_{i=1}^n Y_i, \quad S_0 = 0$$

and

$$\eta(t) = ([t] + 1 - t)S_{[t]} + (t - [t])S_{[t]+1},$$

i. e., the function η is obtained by linearly interpolating S_n at n .

Theorem 2. *The 1-dimensional Brownian Motion ζ and the above η can be redefined on a common probability space without changing their respective laws (the Wiener measure in the case of ζ), in such a way that*

$$\Pr \left\{ \lim_{t \rightarrow \infty} (2t \log \log t)^{-1/2} \sup_{\tau \leq t} |\zeta(\tau) - \eta(\tau)| = 0 \right\} = 1.$$

Proof. This follows easily from an important result of SKOROKHOD [11], p. 180:

“If a sequence Y_1, Y_2, \dots of independent identically distributed real r. v.’s satisfying

$$E Y_1 = 0$$

and

$$E(Y_1^2) < \infty$$

is defined together with a 1-dimensional Brownian Motion ζ on a probability space such that

$$Y_1, Y_2, \dots$$

and

$$\zeta$$

are mutually independent, then there is a sequence

$$\tau_1, \tau_2, \dots$$

of independent identically distributed nonnegative r. v.’s defined on the same space such that

$$E \tau_1 = E(Y_1^2)$$

and such that the process

$$\zeta(\tau_1), \zeta(\tau_1 + \tau_2) - \zeta(\tau_1), \dots, \zeta\left(\sum_{i \leq n} \tau_i\right) - \zeta\left(\sum_{i \leq n-1} \tau_i\right), \dots$$

and the process

$$Y_1, Y_2, \dots, Y_n, \dots$$

have the same distribution.”

To apply this result, let us assume that our original sequence Y_1, Y_2, \dots and a Brownian Motion ζ are already defined on the same probability space and are mutually independent (this of course can always be done). Put $\tau_0 = 0$ and

$$\tilde{Y}_j = \zeta\left(\sum_{i=0}^j \tau_i\right) - \zeta\left(\sum_{i=0}^{j-1} \tau_i\right)$$

for $j \geq 1$, also

$$\tilde{S}_n = \sum_{i=1}^n \tilde{Y}_i = \zeta\left(\sum_{i=1}^n \tau_i\right), \quad \tilde{S}_0 = 0$$

and

$$\tilde{\eta}(t) = ([t] + 1 - t) \tilde{S}_{[t]} + (t - [t]) \tilde{S}_{[t]+1}$$

for any $t \geq 0$. Then $\tilde{\eta}$ has the same distribution as η . Moreover

$$|\tilde{\eta}(t) - \zeta(t)| \leq \max \left\{ \left| \zeta\left(\sum_{1 \leq i \leq t} \tau_i\right) - \zeta(t) \right|, \left| \zeta\left(\sum_{1 \leq i \leq t+1} \tau_i\right) - \zeta(t) \right| \right\}.$$

Let $\varepsilon > 0$. By KOLMOGOROFF'S law of large numbers we have for sufficiently large T

$$\Pr\left\{\left|\sum_{1 \leq i \leq t} \tau_i - t\right| > t\varepsilon \text{ for some } t > T\right\} < \frac{\varepsilon}{2}$$

and

$$\Pr\left\{\left|\sum_{1 \leq i \leq t+1} \tau_i - t\right| > t\varepsilon \text{ for some } t > T\right\} < \frac{\varepsilon}{2},$$

so that with probability $> 1 - \varepsilon$ for all $t > T$

$$\begin{aligned} & \max\left\{\left|\zeta\left(\sum_{1 \leq i \leq t} \tau_i\right) - \zeta(t)\right|, \left|\zeta\left(\sum_{1 \leq i \leq t+1} \tau_i\right) - \zeta(t)\right|\right\} \\ & \leq \sup\left\{\left|\zeta(s) - \zeta(t)\right| : s \in \langle t(1 - \varepsilon), t(1 + \varepsilon) \rangle\right\} \\ & \leq 2 \sup\left\{\left|\zeta(s) - \zeta(t(1 + \varepsilon))\right| : s \in \langle t(1 - \varepsilon), t(1 + \varepsilon) \rangle\right\}. \end{aligned}$$

Therefore with probability $> 1 - \varepsilon$ we have for all $t > T$

$$\begin{aligned} & (2t \log \log t)^{-1/2} \left| \tilde{\eta}(t) - \zeta(t) \right| \\ & \leq 2 \sup\left\{(2t \log \log t)^{-1/2} \left|\zeta(s) - \zeta(t(1 + \varepsilon))\right| : s \in \langle t(1 - \varepsilon), t(1 + \varepsilon) \rangle\right\} \\ & \leq 2 \sup\left\{\left((1 + \varepsilon) \frac{\log \log (t(1 + \varepsilon))}{\log \log t}\right)^{1/2} (2u \log \log u)^{-1/2} \left|\zeta(s'u) - \zeta(u)\right| : \right. \\ & \qquad \qquad \qquad \left. \frac{1 - \varepsilon}{1 + \varepsilon} \leq s' \leq 1\right\} \end{aligned}$$

where we put $u = t(1 + \varepsilon)$. For large t this may be continued

$$\leq 4 \sup\left\{\left|\zeta_{t(1 + \varepsilon)}(s') - \zeta_{t(1 + \varepsilon)}(1)\right| : 1 - 2\varepsilon \leq s' \leq 1\right\}.$$

Applying corollary 1, we see that if we restrict attention to a suitable event of somewhat smaller but still large probability (say $> 1 - 2\varepsilon$), then for $t > T'$ (nonrandom) we have $\zeta_{t(1 + \varepsilon)} \in K_\varepsilon$, so that

$$4 \sup\left\{\left|\zeta_{t(1 + \varepsilon)}(s') - \zeta_{t(1 + \varepsilon)}(1)\right| : 1 - 2\varepsilon \leq s' \leq 1\right\} \leq 4(\sqrt{2\varepsilon} + 2\varepsilon)$$

(recall (1)). So given $\varepsilon > 0$ there is a T' such that

$$\Pr\left\{(2t \log \log t)^{-1/2} \left|\tilde{\eta}(t) - \zeta(t)\right| \leq 4(\sqrt{2\varepsilon} + 2\varepsilon) \text{ for all } t > T'\right\} \geq 1 - 2\varepsilon.$$

This implies

$$\Pr\left\{\lim_{t \rightarrow \infty} (2t \log \log t)^{-1/2} \left|\tilde{\eta}(t) - \zeta(t)\right| = 0\right\} = 1$$

and this implies the theorem.

Put

$$\eta_n(t) = (2n \log \log n)^{-1/2} \eta(nt)$$

for $t \in \langle 0, 1 \rangle$ and $n \geq 3$, i. e., η_n is the r. v. with values in C (where $k = 1$) which is obtained by interpolating linearly

$$(2n \log \log n)^{-1/2} S_i$$

at $\frac{i}{n}$. Then we have

Theorem 3. *With probability one the sequence $(\eta_n)_{n \geq 3}$ is relatively norm-compact and the set of its norm-limit points coincides with K .*

Proof. Replacing t by n in theorem 2 we get

$$\Pr \left\{ \lim_{n \rightarrow \infty} \|\zeta_n - \eta_n\| = 0 \right\} = 1.$$

Theorem 3 then follows from theorem 1.

3. Some applications and comments

Let again

$$Y_1, Y_2, \dots$$

be a sequence of independent identically distributed real r. v.'s with

$$E Y_1 = 0$$

and

$$E(Y_1^2) = 1.$$

(i) The ordinary law of the iterated logarithm for the sequence Y_1, Y_2, \dots follows from theorem 3. In fact (1) with $\alpha = 0$ and $b = 1$ yields

$$\sup_{x \in K} x(1) = 1,$$

where the supremum is attained for and only for $x = t$. In view of theorem 3 this means

$$\Pr \left\{ \limsup_{n \rightarrow \infty} (2n \log \log n)^{-1/2} S_n = 1 \right\} = 1.$$

Moreover for small $\varepsilon > 0$ and sufficiently large n ("large" depending on ε and on chance)

$$(2n \log \log n)^{-1/2} S_n > 1 - \varepsilon$$

happens only if the sequence

$$S_1, S_2, \dots, S_n$$

has an approximately linear shape (here we make use of the compactness of K).

(ii) Let f be any Riemann integrable real function on $\langle 0, 1 \rangle$,

$$F(t) = \int_t^1 f(s) ds$$

for $t \in \langle 0, 1 \rangle$. Then

$$(3) \quad \Pr \left\{ \limsup_{n \rightarrow \infty} (2n^3 \log \log n)^{-1/2} \sum_{i=1}^n f\left(\frac{i}{n}\right) S_i = \left(\int_0^1 F(t)^2 dt \right)^{1/2} \right\} = 1.$$

In particular (putting $f(t) = t^\alpha$)

$$\Pr \left\{ \limsup_{n \rightarrow \infty} (2n^{2\alpha+3} \log \log n)^{-1/2} \sum_{i=1}^n i^\alpha S_i = \frac{1}{\sqrt{\left(\alpha + \frac{3}{2}\right)(\alpha + 2)}} \right\} = 1$$

for any $\alpha > -1$.

To prove (3) we make use of the following

Corollary of Theorem 3. *If φ is a continuous map from C (= Banach space of continuous real functions on $\langle 0, 1 \rangle$) to some Hausdorff space H (in all following applications H will be the set of real numbers), then with probability 1 the sequence $(\varphi(\eta_n))_{n \geq 3}$ is relatively compact and the set of its limit points coincides with $\varphi(K)$.*

Proof. In general if a relatively compact sequence $(y_n)_{n \geq 1}$ of points in C has some (compact) K as the set of its limit points, then $\varphi(K)$ is the set of limit points of $(\varphi(y_n))_{n \geq 1}$. The corollary follows if we assume the basic probability space to be complete (so as to be sure that the event of coincidence between the set of limit points of $(\varphi(\eta_n))_{n \geq 3}$ and $\varphi(K)$ is measurable).

We apply this corollary to the function φ defined by

$$\varphi(x) = \int_0^1 x(t) f(t) dt, \quad (x \in C)$$

to get

$$\Pr \left\{ \limsup_{n \rightarrow \infty} \int_0^1 \eta_n(t) f(t) dt = \sup \varphi(K) \right\} = 1.$$

But

$$\begin{aligned} \sup \varphi(K) &= \sup_{x \in K} \int_0^1 x(t) f(t) dt \\ &= \sup_{x \in K} \int_0^1 F(t) \dot{x}(t) dt \\ &= \sup \left\{ \int_0^1 F(t) y(t) dt : \int_0^1 y^2(t) dt \leq 1 \right\} \\ &= \left(\int_0^1 (F(t))^2 dt \right)^{1/2} \end{aligned}$$

(evaluation of the supremum of a linear functional on the unit sphere of a Hilbert space).

An elementary consideration yields

$$\Pr \left\{ \limsup_{n \rightarrow \infty} \frac{1}{n} \sum_{i=1}^n f\left(\frac{i}{n}\right) (2 n \log \log n)^{-1/2} S_i = \limsup_{n \rightarrow \infty} \int_0^1 f(t) \eta_n(t) dt \right\} = 1,$$

so that (3) is proved.

(iii) Let $a \geq 1$ be real. Then

$$(4) \quad \Pr \left\{ \limsup_{n \rightarrow \infty} n^{-1-(a/2)} (2 \log \log n)^{-a/2} \sum_{i=1}^n |S_i|^a = \frac{2(a+2)^{(a/2)-1}}{\left(\int_0^1 \frac{dt}{\sqrt{1-t^a}} \right)^a a^{a/2}} \right\} = 1,$$

in particular

$$\Pr \left\{ \limsup_{n \rightarrow \infty} n^{-3/2} (2 \log \log n)^{-1/2} \sum_{i=1}^n |S_i| = \frac{1}{\sqrt{3}} \right\} = 1$$

and

$$\Pr \left\{ \limsup_{n \rightarrow \infty} n^{-2} (2 \log \log n)^{-1} \sum_{i=1}^n S_i^2 = \frac{4}{\pi^2} \right\} = 1.$$

One reduces the proof of (4) in an entirely analogous way as above to that of the following assertion:

$$(5) \quad \sup_{x \in K} \int_0^1 |x(t)|^a dt = \frac{2(a+2)^{(a/2)-1}}{\left(\int_0^1 \frac{dt}{\sqrt{1-t^a}} \right)^{a/2}}.$$

Now if \mathcal{A}^2 is the Hilbert space of all absolutely continuous $x \in C$ such that

$$x(0) = 0$$

and

$$\int_0^1 \dot{x}^2 dt < \infty$$

endowed with the inner product

$$((x, y)) = \int_0^1 \dot{x} \dot{y} dt,$$

we have

$$\sup_{x \in K} \int_0^1 |x|^a dt = \sup \left\{ \int_0^1 |x|^a dt : x \in \mathcal{A}^2 \text{ and } ((x, x)) = 1 \right\}.$$

The right-hand side can be evaluated by classical methods of the calculus of variations. We know that the supremum is obtained by some x (because K is a norm compact subset of C and any $x \in K$ which maximizes $\int |x|^a dt$ satisfies $((x, x)) = 1$). Without loss of generality we may assume $x \geq 0$. A necessary condition for x is the existence of a Lagrange multiplier β such that for all $y \in \mathcal{A}^2$

$$\int_0^1 a x^{a-1} dt = \beta 2 \int_0^1 \dot{x} \dot{y} dt$$

(the left-hand side is the derivative of the functional $\int |x|^a dt$ at x applied to y , i. e.,

$$\left(\frac{\partial}{\partial \varepsilon} \int_0^1 |x + \varepsilon y|^a dt \right)_{\varepsilon=0}$$

(see [2]), the right-hand side is β times the derivative of the functional $((x, x))$ at x applied to y). Partial integration of the left-hand side yields

$$\int_0^1 \int_t^1 a(x(s))^{a-1} ds \dot{y}(t) dt = 2 \beta \int_0^1 \dot{x} \dot{y} dt,$$

therefore

$$(6) \quad \int_t^1 a x^{a-1} ds = 2 \beta \dot{x}(t)$$

which shows that \dot{x} has a continuous derivative (of course $\beta \neq 0$) and also that

$$\dot{x}(1) = 0.$$

Differentiating, multiplying with $\dot{x}(t)$ and integrating again yields

$$x^a + \beta \dot{x}^2 = x(1)^a + \beta \dot{x}(1)^2 = x(1)^a.$$

From the significance of x and from our assumption $x \geq 0$ it follows that x is nondecreasing, $x(1) > 0$, so from the above equation

$$\begin{aligned} \dot{x}(0) &> 0, \\ \beta = \frac{x(1)^a}{x(0)^2} &> 0. \end{aligned}$$

Using (6) we see also that

$$\dot{x}(t) > 0$$

for all t . The last equation becomes

$$(7) \quad x^a + \frac{x(1)^a}{x(0)^2} \dot{x}^2 = x(1)^a.$$

Separation of variables and integration yields

$$(8) \quad t = \int_0^{x(t)} \frac{du}{\dot{x}(0) \sqrt{1 - \frac{u^a}{x(1)^a}}}$$

so that

$$1 = \frac{x(1)}{x(0)} \int_0^1 \frac{dv}{\dot{x}(0) \sqrt{1 - v^a}} = \frac{x(1)}{x(0)} \gamma \quad (\text{say}).$$

Now, on the one hand, using (7) and $\int \dot{x}^2 dt = 1$

$$\int_0^1 x^a dt = x(1)^a \left(1 - \frac{1}{\dot{x}(0)^2} \right),$$

on the other hand, using (8)

$$\int_0^1 x^a dt = \int_0^{x(1)} \frac{u^a du}{\dot{x}(0) \sqrt{1 - (u^a/x(1)^a)}} = \frac{2x(1)^{a+1}}{(a+2)\dot{x}(0)} \gamma.$$

Eliminating $x(1)$ and $\dot{x}(0)$ from the last 3 equations we get the desired conclusion (5).

Remarks: If a is an integer, $|S_i|^a$ in (4) can of course be replaced by S_i^a . Also: For sufficiently large n ("large" being random) whenever

$$n^{-1-(a/2)} (2 \log \log n)^{-a/2} \sum_{i=1}^n |S_i|^a$$

is close to its limes superior

$$\frac{2(a+2)^{(a/2)-1}}{\gamma^a a^{a/2}}$$

then η_n is close to either ψ or $-\psi$, where

$$t = \frac{1}{\gamma} \int_0^{\gamma^{(a/a+2)^{1/2}\psi(t)}} \frac{dv}{\sqrt{1-v^a}}$$

(e. g., if $a = 2$, we have $\psi(t) = (\sqrt{8/\pi}) \sin(\pi/2)t$).

One might inquire about

$$\limsup_{i \rightarrow \infty} n_i^{-2} (2 \log \log n_i)^{-1} \sum_{j=1}^{n_i} S_j^2,$$

where n_i runs through all n such that

$$S_n S_{n-1} \leq 0.$$

It is easy to see that the above quantity equals almost surely

$$\sup \left\{ \int_0^1 x^2 dt : x \in K \text{ and } x(1) = 0 \right\}.$$

This supremum equals $1/\pi^2$ and is obtained by the two functions $\pm \sqrt{2}/\pi \sin \pi t$, as can be derived from (8) using a symmetry argument, or directly from WIRTINGER's lemma.

$$(iv) \quad \Pr \left\{ \limsup (2n \log \log n)^{-1/2} \frac{\sum_{i=1}^n S_i^2}{\sum_{i=1}^n |S_i|} = 2p \right\} = 1$$

where p is the largest solution of

$$\sqrt{1-p} \sin \frac{\sqrt{1-p}}{p} + \cos \frac{\sqrt{1-p}}{p} = 0$$

for $0 < p < 1$.

Again one has only to prove

$$\sup \left\{ \frac{\int x^2 dt}{\int |x| dt} : x \in A^2 \text{ and } ((x, x)) = 1 \right\} = 2p.$$

The supremum is attained, say at x . We may assume $x \geq 0$, in fact $x(t) > 0$ for all $t > 0$. Taking derivatives we get with a Lagrange multiplier β

$$\left(\int x dt \int 2xy dt - \int x^2 dt \int y dt \right) (\int x dt)^{-2} = \beta \int 2x \dot{y} dt$$

for all $y \in A^2$ as a necessary condition on x . Substituting for y functions of the form

$$\dot{y}_s(t) = \begin{cases} \frac{1}{s} & \text{if } t \in \langle t_0, s \rangle \\ 0 & \text{otherwise} \end{cases}$$

we see that \dot{x} has a continuous derivative. Putting

$$\alpha = \left(\int x dt \right)^2 \beta$$

we obtain by a partial integration from the above equation

$$\int x dt \int 2xy dt - \int x^2 dt \int y dt = 2\alpha \dot{x}(1) y(1) - 2\alpha \int \ddot{x} y dt$$

and therefore

$$(9) \quad 2x \int x dt - \int x^2 dt + 2\alpha \ddot{x} = 0$$

and

$$\dot{x}(1) = 0.$$

Multiplying (9) by \dot{x} and integrating we get

$$x^2 \int x dt - x \int x^2 dt + \alpha(x^2 - \dot{x}(0)^2) = 0.$$

Integrating from 0 to 1

$$\dot{x}(0) = 1.$$

Any solution of (9) has the form

$$z(t) = p + q \cos \lambda t + r \sin \lambda t.$$

The conditions $z(0) = 0$, $\dot{z}(0) = 1$, $\dot{z}(1) = 0$ and $\int z^2 dt = 1$ specify p, q, r, λ :

$$z(t) = p - p \cos \frac{\sqrt{1-p}}{p} t + \frac{p}{\sqrt{1-p}} \sin \frac{\sqrt{1-p}}{p} t$$

where $-\infty < p < 1$, $p \neq 0$ and

$$(10) \quad \sqrt{1-p} \sin \frac{\sqrt{1-p}}{p} + \cos \frac{\sqrt{1-p}}{p} = 0.$$

For any such p the corresponding z actually solves (9) and satisfies the side conditions. Moreover

$$\int z^2 dt = \frac{2p^2}{1-p},$$

$$\int z dt = \frac{p}{1-p},$$

so that

$$\frac{\int z^2 dt}{\int z dt} = 2p.$$

One easily checks that the z corresponding to the largest value of $p < 1$ satisfying (10) is positive in $(0, 1)$, so that it must coincide with x .

(v) In view of the ordinary law of the iterated logarithm it seems natural to ask about the relative frequency of events

$$S_n > (1 - \varepsilon) (2n \log \log n)^{1/2}.$$

Let $0 \leq c \leq 1$ and

$$c_i = \begin{cases} 1 & \text{if } S_i > c(2i \log \log i)^{1/2} \\ 0 & \text{otherwise.} \end{cases}$$

Then

$$(11) \quad \Pr \left\{ \limsup_{n \rightarrow \infty} \frac{1}{n} \sum_{i=3}^n c_i = 1 - \exp \left\{ -4 \left(\frac{1}{c^2} - 1 \right) \right\} \right\} = 1.$$

For $c = \frac{1}{2}$ as an example we get the somewhat surprising result that with probability one for infinitely many n the percentage of times $i \leq n$ when

$$S_i > \frac{1}{2} (2i \log \log i)^{1/2}$$

exceeds 99.999, but only for finitely many n exceeds 99.9999.

It suffices to prove (11) for $0 < c < 1$. One easily shows, using the fact that for any $\alpha \in (0, 1)$

$$\log \log \alpha n \sim \log \log n,$$

that with probability one

$$\limsup_{n \rightarrow \infty} \frac{1}{n} \sum_{i=[\alpha n]}^n c_i \leq \sup_{x \in K} m \{t : \alpha \leq t \leq 1 \text{ and } x(t) \geq c \sqrt{t}\}$$

and

$$\limsup_{n \rightarrow \infty} \frac{1}{n} \sum_{i=[\alpha n]}^n c_i \geq \sup_{x \in K} m \{t : \alpha \leq t \leq 1 \text{ and } x(t) \geq c' \sqrt{t}\},$$

where $c' < c < c''$ arbitrary and m denotes Lebesgue measure (note that

$$m \{t : \alpha \leq t \leq 1 \text{ and } x(t) \geq c \sqrt{t}\}$$

is not a continuous function of x , so that one has to go back to theorem 3 instead of using its corollary). For continuity reasons it is therefore sufficient to prove

$$(12) \quad \sup_{x \in K} m \{t : x(t) \geq c \sqrt{t}\} = 1 - \exp \left\{ -4 \left(\frac{1}{c^2} - 1 \right) \right\}.$$

Now $m \{t : x(t) \geq c \sqrt{t}\}$ is upper semicontinuous in x , so that the supremum is attained, say by $x_0 \in K$. It is easy to see that all functions y_0 for which the supremum is attained have to satisfy

$$(13) \quad \int \dot{y}_0^2 dt = 1$$

(one uses the fact that for $c > 0$ the set $\{t : y_0(t) < c \sqrt{t}\}$ is not empty). Generally if $0 \leq t_0 < t_1 \leq 1$ and $x \in K$ (only $x \in A^2$ is needed), then

$$(14) \quad \int_{t_0}^{t_1} \dot{x}^2 dt \geq \frac{(x(t_1) - x(t_0))^2}{t_1 - t_0},$$

where equality holds if x is linear. This follows from Jensen's inequality (see [3]) and is known in point mechanics. The linear connection of $x(t_0)$ with $x(t_1)$ will be called the (t_0, t_1) -secant of x .

The set $\{t : x_0(t) > c \sqrt{t}\}$ is empty. Otherwise there would exist t_0, t_1 such that $0 \leq t_0 < t_1 \leq 1$, x_0 is not linear in $\langle t_0, t_1 \rangle$ and the (t_0, t_1) -secant of x_0 would still be greater than $c \sqrt{t}$ in $\langle t_0, t_1 \rangle$. Replacing x_0 by its (t_0, t_1) -secant in $\langle t_0, t_1 \rangle$ we would get a y_0 contradicting (13).

The point 0 is not an accumulation point of $\{t : x_0(t) = c \sqrt{t}\}$. Otherwise there would be $1 \geq t_1 > t_2 > \dots$ such that for all i

$$\sqrt{t_{i+1}} < \frac{1}{2} \sqrt{t_i}$$

and

$$x(t_i) = c \sqrt{t_i}.$$

But then using (14)

$$1 \geq \int_0^{t_1} \dot{x}_0^2 dt \geq \sum_{i=1}^{\infty} \int_{t_{i+1}}^{t_i} \dot{x}_0^2 dt \geq \sum_{i=1}^{\infty} \frac{c^2 (\sqrt{t_i} - \sqrt{t_{i+1}})^2}{t_i - t_{i+1}} > \sum_{i=1}^{\infty} \frac{c^2}{4} = \infty.$$

Therefore the open set

$$\{t : x_0(t) < c \sqrt{t}\},$$

which as any open set is the disjoint union of open intervals, contains among its components an interval of the form $(0, s_0)$. By (13) and (14) x_0 is linear in $\langle 0, s_0 \rangle$ and

$$\int_0^{s_0} \dot{x}_0^2 dt = c^2$$

(note that this value does not depend on s_0). If $\{t: x_0(t) < c\sqrt{t}\} \neq (0, s_0)$, there is another nonempty open interval (s_1, s_2) such that $x_0(t) < c\sqrt{t}$ on (s_1, s_2) and $x_0(s_1) = c\sqrt{s_1}, x_0(s_2) = c\sqrt{s_2}$. But then consider the function y_0 defined by

$$y_0(t) = \begin{cases} \frac{tc}{\sqrt{s_0 + s_2 - s_1}} & \text{if } 0 \leq t \leq s_0 + s_2 - s_1 \\ c\sqrt{s_0 + s_2 - s_1} + x_0(t - s_2 + s_1) - x_0(s_0) & \text{if } s_0 + s_2 - s_1 \leq t \leq s_2 \\ c\sqrt{s_0 + s_2 - s_1} + x_0(s_1) - x_0(s_0) + x_0(t) - x_0(s_2) & \text{if } s_2 \leq t \leq 1. \end{cases}$$

We have

$$y_0 \in A^2,$$

$$m\{t: y_0(t) \geq c\sqrt{t}\} \geq m\{t: x_0(t) \geq c\sqrt{t}\}$$

$$\int_0^1 \dot{y}_0^2 dt = c^2 + \int_{s_0}^{s_1} \dot{x}_0^2 dt + \int_{s_2}^1 \dot{x}_0^2 dt = 1 - \int_{s_1}^{s_2} \dot{x}_0^2 dt < 1,$$

contradicting (13). So x_0 is linear in $\langle 0, s_0 \rangle$ and coincides with $c\sqrt{t}$ in $\langle s_0, 1 \rangle$. $\int \dot{x}_0^2 dt = 1$ determines s_0 as $\exp\{-4((1/c^2) - 1)\}$, which proves (11).

(vi) Though $K \subseteq A^2$ and $\eta_n \in A^2$ for all $n \geq 3$, the sequence $(\eta_n)_{n \geq 3}$ with probability one has no limit points in A^2 with respect to the Hilbert space norm. In fact

$$\Pr\left\{\lim_{n \rightarrow \infty} \int (\dot{\eta}_n)^2 dt = \infty\right\} = 1.$$

For any strictly increasing sequence of integers $a_n \geq 0$ such that $a_0 = 0$ and

$$\lim_{n \rightarrow \infty} \frac{a_{n+1}}{a_n} = 1$$

let $\tilde{\eta}_n \in C$ be obtained by linearly interpolating

$$(2 a_n \log \log a_n)^{-1/2} S_{a_i}$$

at a_i/a_n for $0 \leq i \leq n$. It would be interesting to know for which sequences $(a_n)_{n \geq 0}$ with probability one the set of limit points of $(\tilde{\eta}_n)_{n \geq 3}$ with respect to the Hilbert space norm in A^2 coincides with the unit sphere K in A^2 . In this connection see also LAMPERTI [8].

It would also be very interesting to find the strong form (in the sense of FELLER [5]) of our law of the iterated logarithm, at least for Brownian Motion.

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