

Upper Bounds for k -th Maximal Spacings

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Summary. If K_n denotes the k -th maximal spacing generated by an i.i.d. sequence of random variables uniformly distributed on $(0, 1)$, we show that for any $p \geq 3$,

$$P(nK_n \geq \text{Log } n + (2\text{Log}_2 n + \text{Log}_3 n + \dots + (1 + \varepsilon)\text{Log}_p n)/k \text{ i.o.}) = 1 \quad \text{or} \quad 0$$

according as $\varepsilon \leq 0$ or $\varepsilon > 0$. We also obtain strong limiting bounds for the record times and inter-record times of K_n .

1. Introduction

Let U_1, U_2, \dots be a sequence of independent uniformly distributed random variables on $[0, 1]$. Let $U_0^{(n)} = 0 < U_1^{(n)} < \dots < U_n^{(n)} < U_{n+1}^{(n)} = 1$ be the order statistic corresponding to $0, 1, U_1, \dots, U_n$. The corresponding spacings are defined by

$$S_i^{(n)} = U_i^{(n)} - U_{i-1}^{(n)}, \quad 1 \leq i \leq n+1.$$

Let $K_{n,1} > K_{n,2} > \dots > K_{n,n+1}$ be the order statistic of the spacings. For any $k \leq n$, $K_{n,k}$ will be called the k -th largest spacing of order n , and $M_n = K_{n,1}$ the maximal spacing of order n .

The upper class for $M_n, n \uparrow \infty$ is yet known and given by

$$P(nM_n \geq \text{Log } n + 2\text{Log}_2 n + \text{Log}_3 n + \dots + (1 + \varepsilon)\text{Log}_p n \text{ i.o.}) = 0 \quad \text{or} \quad 1, \quad (1)$$

according as $\varepsilon > 0$ (Devroye [4]) or $\varepsilon \leq 0$ (Deheuvels [2]), for any $p \geq 3$, where Log_j is the j -th times iterated logarithm.

For $k \geq 2$, however, the best results available up to now are:

$$P\left(nK_{n,k} \geq \text{Log } n + \frac{1}{k}(2\text{Log}_2 n + \text{Log}_3 n + \dots + (1 + \varepsilon)\text{Log}_p n) \text{ i.o.}\right) = 0 \quad (2)$$

for any $\varepsilon > 0$ and $p \geq 3$, and

$$P\left(nK_{n,k} \geq \text{Log } n + \frac{2-\varepsilon}{k} \text{Log}_2 n \text{ i.o.}\right) = 1, \tag{3}$$

for any $\varepsilon > 0$ (Devroye [4]).

The aim of the following is to make this result more precise, by proving:

Theorem 1. *For any fixed $k \geq 1$, if $K_n = K_{n,k}$ is the k -th largest spacing of order n , then*

$$P\left(nK_n \geq \text{Log } n + \frac{1}{k}(2 \text{Log}_2 n + \text{Log}_3 n + \dots + (1+\varepsilon) \text{Log}_p n)\right) = 0 \text{ or } 1,$$

according as $\varepsilon > 0$ or $\varepsilon \leq 0$.

In the proof, we shall make use of a sequence of random stopping times defined on U_1, U_2, \dots in the following way. For any $n \geq k$, note by $I_n = I_{n,k}$ the union of the spacing intervals corresponding to $K_{n,1}, K_{n,2}, \dots, K_{n,k}$. The length of I_n will be denoted by $L_n = K_{n,1} + \dots + K_{n,k}$. Put now

$$\begin{aligned} N_1 &= k \\ N_n &= \text{Inf}\{m > N_{n-1}; L_m < L_{N_{n-1}}\}, \quad n = 2, 3, \dots \end{aligned} \tag{4}$$

The definition of N_1, N_2, \dots corresponds to the fact that $K_{n,1}, \dots, K_{n,k}$ remain unchanged when n varies between N_{k-1} and N_k . When n takes the value N_k , one of them decreases, U_n falling then in $I_{n,k}$.

Our main results about N_1, N_2, \dots are the following:

Theorem 2. *For any fixed $k \geq 1$, if N_1, N_2, \dots is defined by (4), then*

$$N_n = \exp\left(\sqrt{2n/k} + \frac{\alpha_n \sqrt{\text{Log}_2 n}}{\sqrt{k}}\right), \tag{5}$$

where

$$\text{LimSup}_{n \rightarrow \infty} \alpha_n = 1 \text{ a.s.}, \quad \text{LimInf}_{n \rightarrow \infty} \alpha_n = -1 \text{ a.s.},$$

and, for any $j \geq 4$, almost surely

$$\begin{aligned} \text{LimSup}_{n \rightarrow \infty} \left\{ k \left(\frac{N_{n+1} - N_n}{N_n} \right) \text{Log } N_n - 2 \text{Log}_2 N_n - \text{Log}_3 N_n - \dots - \text{Log}_{j-1} N_n \right\} / \text{Log}_j N_n \\ = 1 \text{ a.s.} \end{aligned} \tag{6}$$

Remarks. 1°) Theorem 2 makes precise Theorem 1 of [2], where it was proved in the case $k=1$ that

$$N_n = \exp(\sqrt{2n} + O(\text{Log } n)) \text{ a.s.}$$

2°) The study of the sequence $\{N_n, n \geq 1\}$ has some interest in itself. In fact, it is for the maximal spacings the equivalent of the record times for maxima of i.i.d. sequences (see [3]). It behaves though in a quite different way. We shall, in the following paragraph, describe the main properties of this sequence which we have named as the *record time sequence of the k -th maximal spacing*.

2. The Record Time Sequence of the Maximal Spacings

We shall, in the following, use the notations of §1 and assume that $k \geq 1$ is fixed.

Lemma 1. *The sequence $N_1 = k < N_2 < \dots < N_n < \dots$ is an increasing sequence of stopping times on $\mathfrak{U}_1 \subset \mathfrak{U}_2 \subset \dots \subset \mathfrak{U}_n = \sigma(U_1, \dots, U_n) \subset \dots$. Furthermore, if $\mathfrak{U}(N_n)$ denotes the σ -field generated by the events measurable on U_1, \dots, U_{N_n} , then*

$$P(N_{n+1} - N_n \geq r | \mathfrak{U}(N_n)) = (1 - L_{N_n})^r, \quad r = 1, 2, \dots \tag{7}$$

Proof. The first part of the Lemma is straightforward. For the second, note that if L_{N_n} is given, the probability that $N_{n+1} - N_n$ is greater than r is the probability that at least r independent U_i fall outside I_{N_n} . The result follows.

Lemma 2. *For any $\varepsilon > 0$, there exists almost surely an n_ε such that for any $n \geq n_\varepsilon$,*

$$-k(1 + \varepsilon) \text{Log}_3 n \leq nL_n - k \text{Log} n \leq (2 + \varepsilon) \left(\sum_{i=1}^k \frac{1}{i} \right) \text{Log}_2 n, \tag{8}$$

and

$$-k(1 + \varepsilon) \text{Log}_3 N_n \leq N_n L_{N_n} - k \text{Log} N_n \leq (2 + \varepsilon) \left(\sum_{i=1}^k \frac{1}{i} \right) \text{Log}_2 N_n. \tag{9}$$

Proof. Note first that by (4), $N_n \geq n$ and the sequence N_n increases to infinity. Hence, it suffices to prove (8), which in turn follows from (2) and (see Devroye [5])

$$\text{Lim Inf}_{n \rightarrow \infty} (nK_n - \text{Log} n + \text{Log}_3 n) = -\text{Log} 2 \text{ a.s.} \tag{10}$$

Lemma 3. *On the same probability space where U_1, U_2, \dots is defined (eventually extended), there exists a sequence ζ_1, ζ_2, \dots , independent of U_1, U_2, \dots , of independent uniformly distributed on $[0, 1]$ random variables such that if for $n = 1, 2, \dots$,*

$$\omega_n = (-\text{Log}(1 - L_{N_n}))(N_{n+1} - N_n - 1) - \text{Log}(1 - \zeta_n(1 - L_{N_n})), \tag{11}$$

then the sequence $\{\omega_n, n \geq 1\}$ is an i.i.d. sequence of exponentially $E(1)$ distributed random variables such that, for any $n \geq 1$, ω_n is independent of $\mathfrak{U}(N_n)$ and satisfies

$$N_{n+1} - N_n = \frac{\omega_n}{-\text{Log}(1 - L_{N_n})} + 1 - \frac{\text{Log}(1 - \zeta_n L_{N_n})}{\text{Log}(1 - L_{N_n})} = \left[\frac{\omega_n}{-\text{Log}(1 - L_{N_n})} \right] + 1, \tag{12}$$

where $[u]$ denotes the integer part of u .

Proof. It follows as a direct corollary of [3], Theorem 1. Note that Lemma 3 is essentially the same as [1], Lemma 2.

In the following, we shall use the notations

$$\lambda_n = L_{N_n}, \quad n = 1, 2, \dots$$

Lemma 4. *Let, for $0 < z, \zeta < 1$, $\phi(z) = \frac{1}{-\text{Log}(1 - z)} - \frac{1}{z} + \frac{1}{2}$ and*

$$\psi(\zeta, z) = -\frac{\text{Log}(1-\zeta z)}{\text{Log}(1-z)} + \zeta, \quad \text{then,}$$

$$-\frac{z}{2} < \phi(z) < 0 \quad \text{and} \quad 0 < \psi(\zeta, z) < \zeta z.$$

Proof. See [3], Lemma 3 and Lemma 4.

Lemma 5. For any $r > 0$, $\text{Lim}_{n \rightarrow \infty} N_n/n^r = +\infty$.

Proof. The result is clearly true for any $r: 0 < r < 1$, since by (4) for any $n \geq 1$, $N_n \geq n$. Let us assume that it is true for some $r > 0$. If we use the fact that (see (9)) as $n \uparrow \infty$, $\lambda_n \sim k(\text{Log } N_n)/N_n$ a.s., we obtain that $\lambda_n = o(\text{Log } n/n^r)$ a.s. Next, by (12) and Lemma 4,

$$N_{n+1} - N_n \geq \frac{\omega_n}{-\text{Log}(1-\lambda_n)} \geq \omega_n \left\{ \frac{1}{\lambda_n} - \frac{1}{2} - \frac{\lambda_n}{2} \right\}.$$

By summing up and by Kronecker's lemma, it follows that

$$N_n = k + \sum_{i=1}^{n-1} (N_{i+1} - N_i) = \frac{n^{r+1}}{(\text{Log } n) o(1)} \quad \text{a.s.}$$

Hence, the result is still true for $r + \frac{1}{2}$ and also for any $r > 0$.

Lemma 6. Almost surely as $n \uparrow \infty$,

$$N_{n+1} - N_n \sim \frac{\omega_n N_n}{k \text{Log } N_n}.$$

Proof. It suffices to prove by (12) and (9) that $\text{Lim}_{n \rightarrow \infty} \frac{\omega_n N_n}{\text{Log } N_n} = +\infty$ a.s.

This follows from Lemma 5 and:

Lemma 7. For any $p \geq 1$,

$$P(\omega_n \geq \text{Log } n + \dots + (1 + \varepsilon) \text{Log}_p n \text{ i.o.}) = P(\omega_n \leq (n(\text{Log } n) \dots (\text{Log}_p n)^{1+\varepsilon})^{-1} \text{ i.o.}) = 0 \text{ or } 1$$

according as $\varepsilon > 0$ or $\varepsilon \leq 0$.

Proof. By Borel-Cantelli.

Lemma 8. Almost surely as $n \uparrow \infty$,

$$\text{Log } N_n \sim \sqrt{2n/k}.$$

Proof. By Lemma 6 and Lemma 7,

$$\text{Log } N_{n+1} = \text{Log } N_n + \frac{\omega_n}{k \text{Log } N_n} (1 + o(1)) \text{ a.s.}$$

By taking squares,

$$\text{Log}^2 N_{n+1} - \text{Log}^2 N_n = \frac{2\omega_n}{k} (1 + o(1)) \text{ a.s.}$$

It follows by summing up and using the law of large numbers that

$$\text{Log}^2 N_n \sim 2n/k \text{ a.s. as } n \uparrow \infty.$$

Lemma 9. *Let for $n \geq 1$*

$$N_{n+1} = N_n \left\{ 1 + \frac{\omega_n}{k \text{Log} N_n} (1 + \delta_n) \right\}.$$

Then

$$\text{LimSup}_{n \rightarrow \infty} \delta_n \frac{\text{Log} N_n}{\text{Log}_2 N_n} \leq \frac{4}{k} \sum_{i=1}^k \frac{1}{i} \text{ a.s.}$$

Proof. By (9), if $C = (2 + \varepsilon) \sum_{i=1}^k \frac{1}{i}$, as $n \uparrow \infty$, we get

$$\left| \frac{1}{N_n \lambda_n} - \frac{1}{k \text{Log} N_n} \right| \leq \frac{C \text{Log}_2 N_n}{k N_n \lambda_n \text{Log} N_n} \sim \frac{C \text{Log}_2 N_n}{k^2 \text{Log}^2 N_n}.$$

Next, from (12), using the notations of Lemma 4,

$$N_{n+1} - N_n = \frac{\omega_n}{\lambda_n} - \frac{1}{2} \omega_n + 1 - \zeta_n + \omega_n \phi(\lambda_n) + \psi(\zeta_n, \lambda_n).$$

Clearly, by Lemma 4 and Lemma 7, since by (9), $\lambda_n \sim \frac{k \text{Log} N_n}{N_n}$ a.s.,

$$N_{n+1} - N_n = N_n \left\{ \frac{\omega_n}{k \text{Log} N_n} + \omega_n \left(\frac{1}{N_n \lambda_n} - \frac{1}{k \text{Log} N_n} \right) + O \left(\frac{\text{Log}_2 N_n}{N_n} \right) \right\} \text{ a.s.}$$

Finally, we use the fact that, as $n \uparrow \infty$, $\omega_n \leq (1 + \varepsilon) \text{Log} n \sim 2 \text{Log}_2 N_n$ a.s.

Lemma 10. *Let for $n \geq 1$*

$$\text{Log} N_{n+1} = \text{Log} N_n + \frac{\omega_n(1 + \delta_n)}{k \text{Log} N_n} - \frac{\varepsilon_n \omega_n^2}{2k^2 \text{Log}^2 N_n}, \tag{13}$$

then

$$\text{Lim}_{n \rightarrow \infty} \varepsilon_n = 1 \text{ a.s.}$$

Proof. It follows from Lemma 9, by taking logarithms.

We now take squares in (13). It gives

$$\begin{aligned} & \text{Log}^2 N_{n+1} - \text{Log}^2 N_n \\ &= \frac{2\omega_n(1 + \delta_n)}{k} - \frac{\varepsilon_n \omega_n^2}{k^2 \text{Log} N_n} - \frac{\varepsilon_n \omega_n^3(1 + \delta_n)}{k^3 \text{Log}^3 N_n} + \frac{\omega_n^2(1 + \delta_n)^2}{k^2 \text{Log}^2 N_n} + \frac{\varepsilon_n^2 \omega_n^4}{4k^4 \text{Log}^4 N_n}. \end{aligned}$$

Next, since $\text{Log} N_n \sim \sqrt{2n/k}$ by Lemma 8, we get easily the following evaluations:

$$\begin{aligned} \sum_{i=1}^n \frac{\varepsilon_i \omega_i^2}{k^2 \text{Log } N_i} &\sim \sum_{i=1}^n \frac{\omega_i^2}{k^2 \sqrt{2i/k}} \sim (2/k)^{3/2} \sqrt{n} \text{ a.s.}, \\ \sum_{i=1}^n \frac{\omega_i^2}{k^2 \text{Log}^2 N_i} &\sim \sum_{i=1}^n \frac{\omega_i^2}{2ki} \sim k^{-1} \text{Log } n \text{ a.s.}, \\ \sum_{i=1}^{\infty} \frac{\varepsilon_n \omega_n^3}{k^3 \text{Log}^3 N_n} < \infty \quad \text{and} \quad \sum_{i=1}^{\infty} \frac{\varepsilon_n \omega_n^4}{4k^4 \text{Log}^4 N_n} < \infty \text{ a.s.} \end{aligned}$$

From there the following result can be proved.

Theorem 3. *If $\{N_n, n \geq 1\}$ is defined by (4), then, on the same probability space (eventually extended), it is possible to define an i.i.d. sequence $\{\omega_n, n \geq 1\}$ of exponentially $E(1)$ distributed random variables such that*

$$\text{Lim}_{n \rightarrow \infty} n^{-1/2} \left\{ \text{Log}^2 N_n - \frac{2}{k} \sum_{i=1}^n \omega_i \right\} = R \in]-\infty, +\infty[\text{ a.s.} \tag{14}$$

Corollary 1. *If $N(0, 1)$ denotes the standard normal distribution, then, as $n \rightarrow \infty$,*

$$\frac{k}{2\sqrt{n}} \left\{ \text{Log}^2 N_n - \frac{2n}{k} \right\} - \frac{Rk}{2} \xrightarrow{w} N(0, 1)$$

and

$$\sqrt{2k} \{ \text{Log } N_n - \sqrt{2n/k} \} - \frac{Rk}{2} \xrightarrow{w} N(0, 1).$$

Proof. It follows directly from Theorem 3 and the central limit theorem.

Proof of Theorem 2. Note first that (5) is a direct consequence of Theorem 3 and of the law of the iterated logarithm applied to $\sum_{i=1}^n \omega_i$. Next, we prove the following Lemma.

Lemma 11. *Almost surely as $n \uparrow \infty$,*

$$\begin{aligned} \text{Log}_2 N_n &= \frac{1}{2} \text{Log } n + \frac{1}{2} \text{Log}(2/k) + O(n^{-1/2} \text{Log } n), \\ \text{Log}_3 N_n &= \text{Log}_2 n - \text{Log } 2 + O(1/\text{Log } n), \end{aligned}$$

and, for any $p \geq 4$,

$$\text{Log}_p N_n = \text{Log}_{p-1} n + O\left(1 \left/ \prod_{j=2}^{p-2} \text{Log}_j n \right.\right).$$

Proof. It suffices to take the p -th iterated logarithm of both sides of (5).

It is possible by Lemma 11 and Lemma 7 to prove that, for any $p \geq 3$, $P(\omega_n \geq 2 \text{Log}_2 N_n + \text{Log}_3 N_n + \dots + (1 + \varepsilon) \text{Log}_p N_n \text{ i.o.}) = 0$ when $\varepsilon > 0$ and 1 when $\varepsilon < 0$. Finally, by Lemma 9, almost surely as $n \uparrow \infty$,

$$k \left\{ \frac{N_{n+1} - N_n}{N_n} \right\} \text{Log } N_n = \omega_n \left(1 + O\left(\frac{\text{Log } n}{\sqrt{n}}\right) \right). \tag{15}$$

This suffices to complete the proof of (6).

Let us now consider (15), and evaluate:

$$\text{Log } k + \text{Log}(N_{n+1} - N_n) = \text{Log } N_n - \text{Log}_2 N_n + \text{Log } \omega_n + O\left(\frac{\text{Log } n}{\sqrt{n}}\right) \text{ a.s.}$$

An easy deduction from Theorem 2 and Lemma 7 gives the following result.

Theorem 4. *If $\{N_n, n \geq 1\}$ is defined by (4), then,*

$$\limsup_{n \rightarrow \infty} (\text{Log}_2 n)^{-1} \{ \text{Log}(N_{n+1} - N_n) - \sqrt{2n/k} + \frac{1}{2} \text{Log } n \} = 1 \text{ a.s.,}$$

and

$$\liminf_{n \rightarrow \infty} (\text{Log}_2 n)^{-1} \{ \text{Log}(N_{n+1} - N_n) - \sqrt{2n/k} + \frac{3}{2} \text{Log } n \} = -1 \text{ a.s.}$$

Further expansions of the upper and of the lower class of $N_{n+1} - N_n$ may be deduced easily from Lemma 9, Lemma 7, and Theorem 2. It may be remarked here that these classes differ from the classes of N_n .

3. Upper Bounds for the k -th Maximal Spacing

We shall now give the proof of Theorem 1. Even though this proof follows closely the case of $k=1$, treated in [2], the extension from this case to an arbitrary k is not trivial and needs to be detailed.

We first define the sequence $\{n_l, l \geq 1\}$ by $n_l = [\exp(\sqrt{2l/k})]$, $l=1, 2, \dots$, and put $T_l = \omega_{n_l}$. Next, we consider the random sequence defined by

$$\begin{aligned} l(1) &= \text{Min} \{ l \geq 1; L_{n_l} > L_{n_{l+1}} \}, \\ l(r) &= \text{Min} \{ l > l(r-1); L_{n_l} > L_{n_{l+1}} \}, \quad r=2, 3, \dots \end{aligned} \tag{16}$$

Lemma 12. *If $\{l(r), r \geq 1\}$ is defined by (16), then $\{n_{l(r)+1}, r \geq 1\}$ is an increasing sequence of stopping times on $\{\mathfrak{U}_n, n \geq 1\}$ and if for $r=1, 2, \dots$, $\theta_r = \omega_{l(r)+1}$, then $\{\theta_r, r \geq 1\}$ is an i.i.d. sequence of exponentially $E(1)$ distributed random variables.*

Proof. See [2], Lemma 4.

Lemma 13. *With the hypothesis of Lemma 12,*

$$\lim_{r \rightarrow \infty} l(r)/r = \frac{e^{1/k}}{e^{1/k} - 1} \text{ a.s.}$$

Proof. See [2], Lemma 5.

Following exactly the proof in [2], it can be seen that Theorem 1 will be proved for an arbitrary $k \geq 1$ if the following Lemma is true.

Lemma 14. *For any $j \geq 4$ and $c > 0$,*

$$\lim_{N \rightarrow \infty} \frac{1}{N} \sum_{l=1}^N I(n_l K_{n_l} - \text{Log } n_l \geq -c \text{Log } n_l) = 1 \text{ a.s.,}$$

where $I(A)$ denotes the indicator function of the event A .

To see that it is indeed the case, it can be verified that then, as in [2] Lemma 8 and Lemma 10, there exists a.s. an infinite set of indices n such that:

$$nK_n \geq \text{Log } n - c \text{Log}_j n, \tag{17}$$

and hence, since $L_n \geq kK_n$, that

$$nL_n \geq k \text{Log } n - ck \text{Log}_j n,$$

and such that there exists an $m > n$ with $L_m = L_n$ and $K_m = K_n$, satisfying:

$$\left(\frac{m-n}{n}\right) \text{Log } n \geq \frac{1}{k} (2 \text{Log}_2 n + \text{Log}_3 n + \dots + \text{Log}_j n), \tag{18}$$

by an inequality analogous to (8).

It follows from (17) and (18) that

$$mK_m > \text{Log } m + \frac{1}{k} (2 \text{Log}_2 m + \text{Log}_3 m + \dots + \text{Log}_{j-1} m) + \left(\frac{1}{k} - c\right) \text{Log}_j m + o(1),$$

proving Theorem 1.

An important point of this proof is to note that any index n where K_n decreases is also an index where L_n decreases. Hence, if $L_n = L_m$, we must have $K_n = K_m$.

We shall not repeat here the steps detailed in [2], which are, after the preceding remarks, identical in the case $k=1$ and $k \geq 1$. It remains only to prove Lemma 14.

To do so, we shall use the following evaluation given by Devroye [4], Lemma 3.2:

Lemma 15. *If $a_n \rightarrow 0$ and $a_n \text{Log } n \rightarrow \infty$ as $n \rightarrow \infty$, then,*

$$P(nK_n / \text{Log } n - 1 < -a_n) \sim n^{(k-1)a_n} \exp(-n^{a_n}) / (k-1)!. \tag{19}$$

If we put $a_n = (c \text{Log}_j n) / \text{Log } n$, we obtain that

$$\begin{aligned} P(nK_n < \text{Log } n - c \text{Log}_j n) &\sim (1/(k-1)!) \exp(-\exp(a_n \text{Log } n) + (k-1)a_n \text{Log } n) \\ &= (1/(k-1)!) \exp(-(\text{Log}_{j-1} n)^c + c(k-1) \text{Log}_j n). \end{aligned}$$

Let $\eta_l = I(n_l K_{n_l} - \text{Log } n_l < -c \text{Log}_j n_l)$. By the preceding evaluation,

$$E(\eta_l) \sim (1/(k-1)!) \exp(-(\text{Log}_{j-2} \sqrt{2l/k})^c + c(k-1) \text{Log}_{j-1} \sqrt{2l/k}) \rightarrow 0 \text{ as } l \rightarrow \infty.$$

The Lemma 14 will be proved if we prove that $\text{Lim}_{n \rightarrow \infty} \frac{1}{N} \sum_{i=1}^n (\eta_i - E(\eta_i)) = 0$, or equivalently (see [12]), if $\xi_n = \frac{1}{N} \sum_{i=1}^N \eta_i$, if there exists an $a > 1$ such that

$$\sum_{n=1}^{\infty} D^2(\xi_{[a^n]}) < \infty, \text{ where } D^2(\xi_N) = N^{-2} \sum_{i=1}^N \sum_{l=1}^N (E(\eta_i \eta_l) - E(\eta_i) E(\eta_l)).$$

Let us now put $c > 1, j \geq 4, b_i = \lceil i/(\text{Log } i)^{1+\varepsilon} \rceil$, where $\varepsilon > 0$,

$$A_N = N^{-2} \sum_{i=4}^N \sum_{l=i-b_i}^{i+b_i} \text{Cov}(\eta_i, \eta_l), \text{ and } B_N = N^{-2} \sum_{i=4}^N \sum_{l=i+b_i+1}^N \text{Cov}(\eta_i, \eta_l).$$

Clearly $E(\eta_l) = o((\text{Log}_{j-3} l)^{-2}), l \rightarrow \infty$, and, as in [2], (26), it follows that $A_N = O((\text{Log } N)^{-1-\varepsilon} (\text{Log}_{j-3} N)^{-2})$ as $N \rightarrow \infty$. Consequently if $a > 0$, then $\sum_{n \geq 1} A_{\lfloor a^n \rfloor} < \infty$.

Following [2], it suffices to prove that $\sum_{n \geq 1} B_{\lfloor a^n \rfloor} < \infty$. To do so, we have to evaluate $E(\eta_i \eta_l)$ when $l > i + b_i$. By a similar proof as in [2], we obtain the following upper bound for B_N :

$$B_N \leq N^{-2} \sum_{i=4}^N \sum_{l=i+b_i+1}^N E(\eta_i) \{P(n_i K_{n_i-n_i} < \text{Log } n_i - c \text{Log}_j n_l) - E(\eta_l)\}. \tag{20}$$

Next, it follows likewise from [2] that

$$\begin{aligned} n_i(\text{Log } n_l)/n_l &\leq n_i(\text{Log } n_{i+b_i})/n_{i+b_i} \\ &= c'_i = \exp(-b_i(1+o(1))/\sqrt{2ki}) \rightarrow 0 \text{ as } i \rightarrow \infty. \end{aligned}$$

Hence, we get

$$P(n_i K_{n_i-n_i} < \text{Log } n_l - c \text{Log}_j n_l) = P((n_l - n_i) K_{n_l-n_i} < \text{Log}(n_l - n_i) - c \text{Log}_j n_l + O(c'_i)).$$

Finally we get from (20) $B_N \leq N^{-2} \sum_{i=4}^N \sum_{l=i+b_i+1}^N E(\eta_i) \gamma_i$, where

$$\gamma_i = \{P((n_l - n_i) K_{n_l-n_i} < \text{Log}(n_l - n_i) - c \text{Log}_j n_l + O(c'_i)) - P(n_l K_{n_l} < \text{Log } n_l - c \text{Log}_j n_l)\}.$$

Following the bounds given in Devroye [4] (3.3) and Lemma 3.1, if L'_n is the k -th largest of n independent identically distributed random variables with exponential density and whose sum is T_n , then

$$\begin{aligned} P(L'_n < (1-a-b) \text{Log } n) - P(T_n < n(1-b)) &\leq P(nK_n/\text{Log } n < 1-a) \\ &\leq P(L'_n < (1-a+b) \text{Log } n) + P(T_n \geq n(1+b)), \end{aligned} \tag{21}$$

and for n large enough,

$$P(|T_n - n| \geq bn) \leq 2 \exp(-nb^2/4).$$

Let us take now in (21) $a = a_n$ and $b = n^{-1/4}$. It follows that

$$P(|T_n - n| \geq bn) \leq 2 \exp(-\sqrt{n}/4). \text{ Next, } P(L'_n < x) = \sum_{j=0}^{k-1} \binom{n}{j} (1 - e^{-x})^{n-j} e^{-jx}.$$

Put for $j = 0, \dots, k, C_n^j(x) = \binom{n}{j} (1 - e^{-x})^{n-j} e^{-jx}$. We get, for any $\varepsilon > 0$ as $n \rightarrow \infty$,

$$C_n^j((1-a \pm b) \text{Log } n) = \frac{1}{j!} \left(1 + O\left(\frac{1}{n}\right)\right) \exp(-n^{a \mp b} + O(n^{\varepsilon-1}) + j(a \mp b) \text{Log } n).$$

If now $a = a_n \sim c(\text{Log}_j n)/\text{Log } n$, it follows that for any $\theta > 0$,

$$C_n^j((1-a \pm b) \text{Log } n) = n^{ja} \exp(-n^a)(1 + o(n^{\theta-1/4})).$$

Hence,

$$P(nK_n/\text{Log } n < 1-a) = \sum_{j=0}^{k-1} n^{ja} \exp(-n^a)(1 + O(n^{\theta-1/4})).$$

From there it follows as in [2] (27) that

$$B_N \leq N^{-2} \sum_{i=4}^N \sum_{l=i+b_i+1}^N E(\eta_i) E(\eta_l) o(n_i^{\theta-1/4}) = O(N^{-1}).$$

Hence $\sum_{n=1}^{\infty} B_{[a^n]} < \infty$ and the proof of Theorem 1 is complete.

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