

# Reciprocal Processes

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## Introduction

The concept of a reciprocal process was first formulated by Bernstein in 1932 [1]. In 1961 Slepian exploited the reciprocal property of a particular Gaussian process to obtain explicitly a first passage time density for the process [13]. The real-valued reciprocal processes which are stationary and Gaussian are classified in [8]. The first two sections of this paper are devoted to a systematic study of reciprocal processes whose time parameter is a finite closed interval. In the second section, we define the notion of a reciprocal transition probability function. The main result is that given any reciprocal transition probability function there is a probability space supporting a reciprocal process whose transitions are governed by the given transition function. In the third section we give a method of constructing reciprocal processes from Markov processes. Given a Markov process  $\{Y_t, a \leq t \leq b\}$  with state space  $(S, \Sigma)$  whose transition function has with respect to some measure  $\lambda$  on  $\Sigma$  an everywhere positive transition density  $q(s, x, t, y)$ ,  $a \leq s < t \leq b$ ,  $x, y$  in  $S$ , we obtain a reciprocal process  $\{X_t, a \leq t \leq b\}$  by first tying down  $\{Y_t, a \leq t \leq b\}$  at  $Y_a = x$  and  $Y_b = y$  and then giving  $(x, y)$  an arbitrary probability distribution on  $\Sigma \times \Sigma$ . This method is a generalization of one due to Schrödinger ([11, 12]) and discussed by Bernstein [1] (see also Miller's appendix on p. 202–223 of [10]). Since any Markov process is a reciprocal process, a question arises as to whether all of the processes which are constructed by this method are not only reciprocal but Markovian. We prove the following result: An endpoint distribution  $\mu$  gives rise to a Markov process  $\{X_t, a \leq t \leq b\}$  if and only if there is a product measure  $\pi$  on  $\Sigma \times \Sigma$  for which  $d\mu/d\pi = q$ , where  $q(x, y) = q(a, x; b, y)$ . For example, it is easy to see that if we reproduce the original process by taking for  $\mu$  the original joint distribution of  $Y_a$  and  $Y_b$ , it is of this form (as indeed it must be if the result is at all valid). Two questions arise. First, are there any other probability distributions on  $\Sigma \times \Sigma$  which are of this form? (If not, the original process  $\{Y_t, a \leq t \leq b\}$  is the only one of the derived processes  $\{X_t, a \leq t \leq b\}$  which is Markov.) We show that under quite general conditions, the answer is yes: In fact, given any probability measures  $\mu_1$  and  $\mu_2$  on  $\Sigma$  there is a measure  $\mu$  having  $\mu_1$  and  $\mu_2$  for marginals for which  $d\mu/d\pi = q$  for some product measure  $\pi$  on  $\Sigma \times \Sigma$ . Thus our construction yields a Markov process  $\{X_t, a \leq t \leq b\}$  with prescribed distributions for  $X_a$  and  $X_b$ . If  $\mu_1$  and  $\mu_2$  are absolutely continuous with respect to  $\lambda$ , finding such a  $\mu$  amounts to solving a pair of nonlinear functional equations first derived by Schrödinger ([11] and [12]) in a completely different way based on considerations partly physical and partly probabilistic, which seem to have no

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connection with the Markovian or non-markovian nature of the process so constructed. The problem of the existence and uniqueness of solutions to Schrödinger's functional equations was first treated systematically by Fortet [7]. Beurling [2] has formulated and analyzed a more general problem which includes ours as a special case. He obtains not only existence but uniqueness of the solution in case  $S$  is locally compact,  $q$  is bounded and continuous, and  $\iint \log q(x, y) \mu_1(dx) \mu_2(dy)$  is finite. We are able to remove this last condition. The uniqueness part of the result answers a second question which arises, namely, do perhaps all probability measures  $\mu$  on  $\Sigma \times \Sigma$  satisfy  $d\mu/d\pi = q$  for some product measure  $\pi$  on  $\Sigma \times \Sigma$ ? If this were so, our construction would not yield any reciprocal processes which are not Markov. (We remark that Bernstein [1] seemed unaware that Schrödinger's construction, with endpoint measures obtained via his functional equations, yields only Markov processes.) However, if we are given probability measures  $\mu_1$  and  $\mu_2$  on  $\Sigma$ , exactly one of the processes  $\{X_t, a \leq t \leq b\}$  with the distributions of  $X_a$  and  $X_b$  given by  $\mu_1$  and  $\mu_2$  respectively is Markov, all the rest being reciprocal but not Markov. (There are as many processes constructed with the distributions of  $X_a$  and  $X_b$  so prescribed as there are probability measures  $\mu$  in  $\Sigma \times \Sigma$  with marginals  $\mu_1$  and  $\mu_2$ .)

The reciprocal processes constructed from Markov processes by the method of the third section have transition functions which are absolutely continuous with respect to the reference measure  $\lambda$ . In the last section we examine the question of whether the converse holds: that is, given a reciprocal process whose transition function is absolutely continuous with respect to  $\lambda$ , is there a Markov process from which it can be constructed by our method? Our answer is in the partial affirmative.

There are a number of equations in this paper in which strict equality is indicated, but which actually hold almost everywhere with respect to some measure. The necessity for such a qualification will in each case be clear from the context.

## § 1

We begin by defining our basic notion.  $(S, \Sigma)$  is an arbitrary measurable space.

*Definition.* Let  $\{X_t, a \leq t \leq b\}$  be an  $(S, \Sigma)$ -valued stochastic process on the finite closed interval  $[a, b]$  with underlying probability space  $(\Omega, \mathcal{A}, P)$ . We say that  $\{X_t, a \leq t \leq b\}$  is a *reciprocal process* if, for each  $a \leq s < t \leq b$ ,

$$P(AB|X_s, X_t) = P(A|X_s, X_t)P(B|X_s, X_t)$$

whenever  $A$  belongs to the  $\sigma$ -field generated by the random variables  $\{X_r: a \leq r < s$  or  $t < r \leq b\}$  and  $B$  to the  $\sigma$ -field generated by  $\{X_r: s < r < t\}$ .

The following two lemmas are proved in [8].

**Lemma 1.1.** *The process  $X_t, a \leq t \leq b$  is reciprocal if and only if*

$$E\{f(X_n)|X_{s_1}, \dots, X_{s_n}, X_t, X_v\} = E\{f(X_n)|X_t, X_v\} \quad (1.1)$$

for each  $a \leq t < u < v \leq b$ ,  $\{s_1, \dots, s_n\} \subset [a, b] - (t, v)$ , and bounded Borel-measurable  $f$ .

**Lemma 1.2.** *If  $\{X_t, a \leq t \leq b\}$  is a Markov process, then it is a reciprocal process.*

The following lemma is referred to in the next section.

**Lemma 1.3.** Suppose  $\{X_t, a \leq t \leq b\}$  is a reciprocal process, that  $a \leq s < t < u < v \leq b$ , and that  $f$  and  $g$  are bounded Borel functions. Then

$$E\{f(X_t) E\{g(X_u)|X_t, X_v\} | X_s, X_v\} = E\{g(X_u) E\{f(X_t)|X_s, X_u\} | X_s, X_v\}. \quad (1.2)$$

*Proof.* Using the reciprocal property, we have

$$\begin{aligned} & E\{f(X_t) E\{g(X_u)|X_t, X_v\} | X_s, X_v\} \\ &= E\{f(X_t) E\{g(X_u)|X_s, X_t, X_v\} | X_s, X_v\} \\ &= E\{E\{f(X_t) g(X_u)|X_s, X_t, X_v\} | X_s, X_v\} \\ &= E\{f(X_t) g(X_u) | X_s, X_v\} \\ &= E\{E\{f(X_t) g(X_u)|X_s, X_u, X_v\} | X_s, X_v\} \\ &= E\{g(X_u) E\{f(X_t)|X_s, X_u, X_v\} | X_s, X_v\} \\ &= E\{g(X_u) E\{f(X_t)|X_s, X_u\} | X_s, X_v\}. \end{aligned}$$

**Lemma 1.4.** If  $\{X_t, a \leq t \leq b\}$  is a reciprocal process, and either  $X_a$  or  $X_b$  is a.s. constant, then it is a Markov process.

*Proof.* First, suppose  $X_b$  is constant a.s. Then, if  $a \leq t_1, < \dots < t_n < u \leq b$ , and if  $f$  is bounded measurable,

$$\begin{aligned} E\{f(X_u) | X_{t_1}, \dots, X_{t_n}\} &= E\{f(X_u) | X_{t_1}, \dots, X_{t_n}, X_b\} \\ &= E\{f(X_u) | X_{t_n}, X_b\} \\ &= E\{f(X_u) | X_{t_n}\}. \end{aligned}$$

Thus  $\{X_t, a \leq t \leq b\}$  is Markov. Since the Markov and reciprocal properties are both preserved under reversal of the time direction, the conclusion also holds if  $X_a$  is constant a.s.

## § 2

We begin by defining axiomatically a class of reciprocal transition probability functions which are to reciprocal processes what transition probability functions are to Markov processes (for the latter, see [9], Section 38.2). First let  $I = [a, b]$  be a closed interval of real numbers. Let  $(S, \Sigma)$  be a measurable space. We use  $\mathcal{D}$  to denote the set of all ordered sextuples  $(s, x, t, E, u, y)$  for which  $x$  and  $y$  are in  $S$ ,  $a \leq s < t < u \leq b$ , and  $E \in \Sigma$ . A real valued function  $P$  on  $\mathcal{D}$  is called a *reciprocal transition probability function* if the following three conditions are satisfied:

A 1. For each  $x$  and  $y$  in  $S$  and  $a \leq s < t < u \leq b$ , the map

$$E \rightarrow P(s, x, t, E, u, y), \quad E \in \Sigma$$

defines a probability measure on  $\Sigma$ .

A 2. For each  $E \in \Sigma$  and  $a \leq s < t < u < v \leq b$ , the map

$$(x, y) \rightarrow P(s, x, t, E, u, y)$$

is  $\Sigma \times \Sigma$ -measurable.

A3. For each  $a \leq s < t < u < v \leq b$ ,  $C \in \Sigma$ ,  $D \in \Sigma$ ,  $x \in S$ , and  $y \in S$ ,

$$\begin{aligned} & \int_D P(s, x, u, d\xi, v, y) P(s, x, t, C, u, \xi) \\ &= \int_C P(s, x, t, d\eta, v, y) P(t, \eta, u, D, v, y). \end{aligned}$$

Intuitively,  $P(s, x, t, E, u, y)$  is the probability that a particle located at  $x$  at time  $s$  and at  $y$  at time  $u$  is in the set  $E$  at time  $t$ . To help keep this in mind, we write  $P(s, x; t, E; u, y)$  for  $P(s, x, t, E, u, y)$ . The following are two consequences of A1–A3. The first is obtained by setting  $C = A$  and  $D = S$  in A3 and applying A1, the second by setting  $C = S$  and  $D = A$  in A3.

For each  $a \leq s < t < u < v \leq b$ ,  $A \in \Sigma$ ,  $x \in S$ , and  $y \in S$ ,

$$\text{A4.} \quad \int P(s, x; t, d\eta; v, y) P(t, \eta; u, A; v, y) = P(s, x; u, A; v, y)$$

and

$$\text{A5.} \quad \int P(s, x; u, d\xi; v, y) P(s, x; t, A; u, \xi) = P(s, x; t, A; v, y).$$

Let  $\Omega$  be the set of all  $S$ -valued functions on  $[a, b]$ . For each  $t \in [a, b]$ , we denote by  $X_t$  the function on  $\Omega$  for which  $X_t(\omega) = \omega(t)$ ,  $\omega \in \Omega$ . The smallest  $\sigma$ -field  $\mathcal{G}$  on  $\Omega$  relative to which  $X_t$  is  $\mathcal{G} - \Sigma$  measurable for each  $t \in [a, b]$  is denoted by  $\mathcal{I}$ .

**Theorem 2.1.** *Assume that  $S$  is a  $\sigma$ -compact Hausdorff space, with  $\Sigma$  the  $\sigma$ -field generated by the open sets. Let  $P(s, x; t, E; u, y)$  be a reciprocal transition probability function as defined above, and let  $\mu$  be a probability measure on  $\Sigma \times \Sigma$ . Then there is a probability measure  $P_\mu$  on  $\mathcal{I}$  such that, relative to the probability space  $(\Omega, \mathcal{I}, P_\mu)$ ,  $\{X_t, a \leq t \leq b\}$  is a reciprocal process for which*

$$(i) \quad P_\mu \{X_a \in A, X_b \in B\} = \mu(A \times B), \quad A \in \Sigma, B \in \Sigma,$$

and

$$(ii) \text{ for all } a \leq s < t < u \leq b \text{ and } A \in \Sigma,$$

$$P_\mu(X_t \in A | X_s, X_u) = P(s, X_s; t, A; u, X_u).$$

There is only one such measure, and its finite-dimensional distributions are given as follows. Suppose  $a < t_1 < \dots < t_n < b$ ,  $A \in \Sigma$ ,  $B \in \Sigma$ , and  $E_i \in \Sigma$ ,  $i = 1, \dots, n$ . Let

$$A = \{X_a \in A, X_{t_1} \in E_1, \dots, X_{t_n} \in E_n, X_b \in B\}. \quad (2.1)$$

Then  $P_\mu(A)$  is equal to

$$\begin{aligned} & \int_{A \times B} d\mu(x, y) \int_{E_1} P(a, x; t_1, dz_1; b, y) \dots \\ & \int_{E_{n-1}} P(t_{n-2}, z_{n-2}; t_{n-1}, dz_{n-1}; b, y) P(t_{n-1}, z_{n-1}; t, E_n; b, y). \end{aligned} \quad (2.2)$$

*Proof.* We begin by showing that if  $\{X_t, a \leq t \leq b\}$  is a reciprocal process on  $(\Omega, \mathcal{I}, P)$  relative to which (i) and (ii) hold (with “ $P_\mu$ ” replaced by “ $P$ ”), then, if  $A$  is given by (2.1),  $P(A)$  is given by (2.2). (This will, of course, establish the uniqueness

asserted by the theorem.) To this end let  $a \leq t_1 < \dots < t_n < b$ , and  $E_i \in \Sigma$ ,  $i = 1, \dots, n$ . I claim

$$\begin{aligned} & P\{X_{t_2} \in E_2, \dots, X_{t_n} \in E_n | X_{t_1}, X_b\} \\ &= \int_{E_2} P(t_1, X_{t_1}; t_2, dz_2; b, X_b) \int_{E_3} P(t_2, x_2; t_3, dx_3; b, X_b) \dots \\ & \int_{E_{n-1}} P(t_{n-2}, x_{n-2}; t_{n-1}, dx_{n-1}; b, X_b) P(t_{n-1}, x_{n-1}; t_n, E_n; b, X_b). \end{aligned} \quad (2.3)$$

We prove (2.3) by induction on  $n$ . For  $n = 1$ , it reduces to (ii), which we are assuming. Note that (ii) also implies that

$$E\{f(X_{t_1}, X_b) | X_a, X_b\} = \int P(a, X_a; t_1, dz; b, X_b) f(x, X_b)$$

for any bounded  $\Sigma \times \Sigma$ -measurable  $f$  on  $S \times S$ , and that there is such an  $f$  for which the right hand side of (2.3) is  $f(X_{t_1}, X_b)$ . Assuming that (2.3) holds as it stands, we have, if  $a \leq t_0 < t_1 < \dots < t_n < b$

$$\begin{aligned} & P\{X_{t_1} \in E_1, \dots, X_{t_n} \in E_n | X_{t_0}, X_b\} \\ &= E\{I_{E_1}(X_{t_1}) P\{X_{t_2} \in E_2, \dots, X_{t_n} \in E_n | X_{t_0}, X_{t_1}, X_b\} | X_{t_0}, X_b\} \\ &= E\{I_{E_1}(X_{t_1}) P\{X_{t_2} \in E_2, \dots, X_{t_n} \in E_n | X_{t_1}, X_b\} | X_{t_0}, X_b\} \\ &= E\{I_{E_1}(X_{t_1}) f(X_{t_1}, X_b) | X_{t_0}, X_b\} \\ &= \int_{E_1} P(t_0, X_{t_0}; t_1, dz_1; b, X_b) f(X_{t_1}, X_b) \\ &= \int_{E_1} P(t_0, x_{t_0}; t_1, z_1; b, X_b) \int_{E_2} P(t_1, X_{t_1}; t_2, dz_2; b, X_b) \dots \\ & \int_{E_{n-1}} P(t_{n-2}, X_{n-2}; t_{n-1}, dx_{n-1}; b, X_b) P(t_{n-1}, X_{n-1}; t, E_n, b, X_b). \end{aligned} \quad (2.4)$$

This shows that (2.3) holds for all  $n$ . Using (2.4) for  $t_0 = a$ , the fact that

$$P(A) = E\{I_A(X_a) I_B(X_b) P\{X_{t_1} \in E_1, \dots, X_{t_n} \in E_n | X_a, X_b\}\},$$

and (i), we conclude that  $P(A)$  is equal to expression (2.2).

Next, we construct  $P_\mu$ . Rather than using a consistency argument to extend the set function defined by (2.2) to  $\mathcal{A}$ , we proceed indirectly. Fix  $y \in S$ . For each  $a \leq s < t < b$ ,  $z \in S$ , and  $E \in \Sigma$ , set

$$Q_y(s, z; t, E) = P(s, z, t, E; b, y).$$

Then, if  $a \leq s < t < u < b$ , we have, using (A 4),

$$\begin{aligned} \int Q_y(s, z; t, d\eta) Q_y(t, \eta; u, E) &= \int P(s, z; t, d\eta; b, y) P(t, \eta; u, E; b, y) \\ &= P(s, z; u, E; b, y) \\ &= Q_y(s, z; u, E). \end{aligned}$$

It follows that  $Q_y(s, z; t, E)$  is a (Markov) transition probability function. Let  $\Omega_0$  be the set of all functions from  $[a, b)$  into  $S$ , and  $\mathcal{F}_0$  the smallest  $\sigma$ -field on  $\Omega_0$  rendering measurable all the coordinate functions  $X_t$ ,  $a \leq t < b$ . Because of our assumptions on  $(S, \Sigma)$  it follows ([4], p. 16) that given any probability measure  $\gamma$

on  $\Sigma$  there is a measure  $\tilde{Q}_{y,\gamma}$  on  $\mathcal{S}_0$  such that, relative to  $(\Omega_0, \mathcal{S}_0, \tilde{Q}_{y,\gamma})$ ,  $\{X_t, a \leq t < b\}$  is a Markov process with  $\gamma$  as initial measure and  $Q_y(s, z; t, E)$  as transition probability function ([9], p. 569). Now,  $(S \times S, \Sigma \times \Sigma, \mu)$  is a probability space. Let  $X$  and  $Y$  be the random variables defined thereon by  $X(x, y) = x$  and  $Y(x, y) = y$  for all  $(x, y) \in S \times S$ . Let  $\nu$  be the conditional distribution of  $X$  given  $Y$  ([9], p. 359). Then  $\nu$  is defined on  $S \times \Sigma$ ,  $\nu(y, \cdot)$  is a probability measure on  $\Sigma$  for each  $y \in S$  and  $\nu(\cdot, E)$  is  $\Sigma$ -measurable for each  $E \in \Sigma$ . Using  $\nu(y, \cdot)$  as the initial measure we define  $\tilde{Q}_y$  on  $\mathcal{S}_0$  as above. Checking first the case where  $A$  is a cylinder with finite-dimensional base, we see that  $\tilde{Q}_y(A)$  is a  $\Sigma$ -measurable function of  $y$  for each  $A \in \mathcal{S}_0$ . Let  $\eta$  be the distribution of  $Y$ ; that is,  $\eta(F) = \mu(S \times F)$  for each  $F \in \Sigma$ . We define  $P_\mu$  on  $\mathcal{S}_0 \times \Sigma$  by

$$P_\mu(A \times F) = \int_F \eta(dy) \tilde{Q}_y(A) \quad A \in \mathcal{S}_0, F \in \Sigma. \tag{2.5}$$

It is observed on p. 359 of [9] that this indeed defines a measure on  $\mathcal{S}_0 \times \Sigma$ . The measure  $P_\mu$  is not yet defined on  $\mathcal{S}$  as promised. But the correspondence  $\omega \leftrightarrow (\omega_0, \omega(b))$  between  $\Omega$  and  $\Omega_0 \times S$ , where  $\omega_0$  is the restriction to  $[a, b)$  of  $\omega \in \Omega$ , is one-to-one and  $\mathcal{S} - \mathcal{S}_0 \times \Sigma$  bimeasurable, permitting us to identify the measurable spaces  $(\Omega, \mathcal{S})$  and  $(\Omega_0 \times S, \mathcal{S}_0 \times \Sigma)$ . Accordingly, (2.5) does define a probability measure on  $\mathcal{S}$ .

Next, we verify that if  $A$  is as in (2.1), then  $P_\mu(A)$  is given by (2.2). First, suppose that  $f$  is a bounded  $\Sigma \times \Sigma$ -measurable function on  $S \times S$ . Then the definitions of  $\gamma$  and  $\nu$  easily yield  $\int \gamma(dy) \int \nu(y, dx) f(x, y) = \int f d\mu$ ; consequently,

$$\int_B \gamma(dy) \int_A \nu(y, dx) f(x, y) = \int_{A \times B} f d\mu \tag{2.6}$$

for any  $A \in \Sigma, B \in \Sigma$ . Now let

$$f(x, y) = \int_{E_1} Q_y(a, x; t_1, dz_1) \dots \int_{E_{n-1}} Q_y(t_{n-1}, dx_{n-1}; t_n, A_n)$$

and observe ([9], p. 569) that if

$$A = \{\omega \in \Omega_0 : X_a(\omega) \in A, X_{t_1}(\omega) \in E_1, \dots, X_{t_n}(\omega) \in E_n\},$$

then

$$\tilde{Q}_y(A) = \int_A \nu(y, dx) f(x, y). \tag{2.7}$$

If  $A$  is given by (2.1), we identify  $A$  with  $A \times B$ , so combining (2.5), (2.6) and (2.7) we see that  $P_\mu(A)$  is indeed given by (2.2). It is evident from (2.2) that (i) holds.

We next show that (ii) holds. Suppose  $a < t < u < v < b$ . It is easy to see from the form (2.2) of the finite dimensional distributions that

$$\int h(X_t, X_v) dP_\mu = \int d\mu(x, y) P(a, x; t, dw; b, y) \int P(t, w; v, dz; b, y) h(w, z) \tag{2.8}$$

for all bounded  $\Sigma \times \Sigma$ -measurable functions  $h$  on  $S \times S$ . Let  $B \in \Sigma, C \in \Sigma, D \in \Sigma$ . Let

$$h(w, z) = P(t, w; u, C; v, z) I_B(w) I_D(z),$$

and apply (2.8) to obtain

$$\int_{\{X_t \in B, X_v \in D\}} P(t, X_t; u, C; v; X_v) dP_\mu = \int d\mu(x, y) \int_B P(a, x; t, dw; b, y) \int_D P(t, w; v, dz; b, y) P(t, w; u, C; v, z). \quad (2.9)$$

By (A 3), however,

$$\int_D P(t, w; v, dz; b, y) P(t, w; u, C; v, z) = \int_C P(t, w; u, d\eta; b, y) P(u, \eta; v, D; b, y).$$

Substituting the right hand side of this last expression into the right hand side of (2.9), and referring to (2.2), we see that

$$\int_{\{X_t \in B, X_v \in D\}} P(t, X_t; u, C; v; X_v) dP_\mu = P_\mu \{X_t \in B, X_u \in C, X_v \in D\}.$$

Since this holds for all  $B, D$  in  $\Sigma$  it follows that

$$P_\mu \{X_u \in C | X_t, X_v\} = P_\mu(t, X_t; u, C; v, X_v).$$

A similar argument shows that this last also holds if  $t=a$  or  $v=b$ . Thus (ii) is proved.

We complete the proof of the theorem by establishing the reciprocal property of  $\{X_t, a \leq t \leq b\}$  relative to  $(\Omega, \mathcal{F}, P_\mu)$ . Suppose that  $a < t_n < \dots < t_1 < t < u < v < v_1 < \dots < v_m < b$ , and that  $C \in \Sigma$ . We will show that

$$P_\mu \{X_u \in C | X_a, X_{t_n}, \dots, X_{t_1}, X_t, X_v, X_{v_1}, \dots, X_{v_m}, X_b\} = P(t, X_t; u, C; v, X_v).$$

To do this, we must show that

$$\int_{AA} P(t, X_t; u, C; v, X_v) dP_\mu = P_\mu(AA \{X_u \in C\}), \quad (2.10)$$

whenever

$$A = \{X_a \in A, X_{t_n} \in D_n, \dots, X_{t_1} \in D_1, X_t \in D\}$$

and

$$\Delta = \{X_v \in E, X_{v_1} \in E_1, \dots, X_{v_m} \in E_m, X_b \in B\}$$

with  $A, D_n, \dots, D_1, D, E, E_1, \dots, E_m$  all in  $\Sigma$ .

To this end, let

$$K(y, z) = \int_{E_2} P(v, y; v_1, dy_1; b, z) \dots \int_{E_{m-1}} P(v_{m-2}, y_{m-2}; v_{m-1}, dy_{m-1}; b, z) P(v_{m-1}, y_{m-1}; v_m, E_m; b, z).$$

It follows from (2.2) that if  $f$  is any bounded  $\Sigma \times \Sigma$ -measurable function on  $S \times S$ , then

$$\int_{AA} f(X_t, X_v) dP_\mu = \int_{A \times B} d\mu(w, z) \int_{D_n} P(a, w; t_n, dx_n; b, z) \dots \int_{D_n} P(t_n, x_n; t_{n-1}, dx_{n-1}; b, z) \int_D P(t_1, x_1; t, dx; b, z) F(x, z), \quad (2.11)$$

where

$$F(x, z) = \int_D P(t, x; v, dy; b, z) f(x, y) K(y, z).$$

In particular, if  $f(x, y) = P(t, x; u, C; v, y)$ , the left hand side of (2.10) is equal to the right hand side of (2.11) with

$$F(x, z) = \int_D P(t, x; v, dy; b, z) P(t, x; u, C; v, y) K(x, z). \quad (2.12)$$

Ba (A3),

$$\int_D P(t, x; v, dy; b, z) P(t, x; u, C; v, y) = \int_C P(t, x; u, d\eta; b, z) P(u, \eta; v, D; b, z),$$

and it easily follows that

$$\begin{aligned} & \int_D P(t, x; v, dy; b, z) P(t, x; u, C; v, y) K(y, z) \\ &= \int_C P(t, x; u, d\eta; b, z) \int_D P(u, \eta; v, dy; b, z) K(y, z). \end{aligned} \quad (2.13)$$

Substituting the right hand side of (2.13) for  $F(x, z)$  into the right hand side of (2.11), and referring to (2.2) again, we obtain (2.10). Thus  $\{X_t, a \leq t \leq b\}$  (as a process on  $(\Omega, \mathcal{F}, P_a)$ ) is reciprocal, and the proof of the theorem is complete.

If  $\{X_t, a \leq t \leq b\}$  is a reciprocal process, and if we define  $P(s, x; t, E; u, y)$  to be a conditional distribution satisfying (ii) of the theorem with appropriate almost everywhere qualifications, A3 must hold (with similar qualifications), as is seen by setting  $f = I_C$  and  $g = I_D$  in Lemma 1.3. This shows that A3 is not too strong a condition to impose on reciprocal transition functions.

### § 3

Suppose  $\{Y_t, a \leq t \leq b\}$  is a Markov process with Markov transition probability function  $Q(s, x, t, E), a \leq s < t \leq b, x \in S, E \in \Sigma$ . We assume that  $Q$  is given by a positive density relative to some  $\sigma$ -finite measure  $\lambda$  on  $\Sigma$ ; that is, there is a strictly positive function  $q(s, x; t, y)$  defined for  $a \leq s < t \leq b$  and  $(x, y) \in S \times S$ ,  $\Sigma$ -measurable in  $(x, y)$  for each  $s$  and  $t$ , and for which

$$Q(s, x, t, E) = \int_E q(s, x; t, y) \lambda(dy) \quad a \leq t \leq b, x \in S, E \in \Sigma. \quad (3.1)$$

We define

$$p(s, x; t, y; u, z) = \frac{q(s, x; t, y) q(t, y; u, z)}{q(s, x; u, z)}, \quad a \leq s < t < u \leq b, (x, y, z) \in S \times S \times S, \quad (3.2)$$

and

$$\begin{aligned} P(s, x; t, E; u, y) &= \int_E p(s, x; t, z; u, y) \lambda(dz), \\ &a \leq s < t < u \leq b, (x, y) \in S \times S, E \in \Sigma. \end{aligned} \quad (3.3)$$

It is easy to verify that  $P(s, x; t, E; u, y)$  is a reciprocal transition probability function; we say that it is *derived from*  $q(s, x; t, y)$ . We observe that  $P(s, Y_s; t, E; u, Y_u)$  is a version of  $P(Y_t \in E | Y_s, Y_u)$ .

Let  $\mu$  be an arbitrary probability measure on  $\Sigma \times \Sigma$ . By virtue of theorem 2.1, if  $S$  is a  $\sigma$ -compact Hausdorff space with  $\Sigma$  its topological Borel sets there is a



unique measure  $P_\mu$  on the measurable space  $(\Omega, \mathcal{F})$  of paths such that the coordinate functions  $\{X_t, a \leq t \leq b\}$  constitute a reciprocal process for which

$$(i) \quad P_\mu(X_a \in A, X_b \in B) = \mu(A \times B) \quad A \in \Sigma, B \in \Sigma,$$

and

$$(ii) \quad P_\mu(X_t \in A | X_s, X_u) = P(s, X_s; t, A; u, X_u), \quad A \in \Sigma, a \leq s < t \leq u \leq b.$$

We call  $\mu$  the (joint) endpoint distribution of  $\{X_t, a \leq t \leq b\}$ . The measures  $\mu_a$  and  $\mu_b$  defined by  $\mu_a(E) = \mu(E \times S)$  and  $\mu_b(E) = \mu(S \times E)$  are called the marginal endpoint distributions. We denote the joint distribution of  $X_s$  and  $X_t$  by  $\mu_{s,t}$  for  $a \leq s < t \leq b$ . Thus  $\mu_{a,b} = \mu$ . The distribution of  $X_s$  is denoted by  $\mu_s, a \leq s \leq b$ . If either  $\mu_a$  or  $\mu_b$  concentrates all its mass on a single point of  $S$ ,  $\{X_t, a \leq t \leq b\}$  is not only reciprocal but Markovian by virtue of Lemma 1.4. In the following theorem we characterize for  $S$  metric all endpoint distributions  $\mu$  for which  $\{X_t, a \leq t \leq b\}$  is a Markov process.

**Theorem 3.1.** *Let  $Q(s, x; t, E), a \leq s < t \leq b, x \in S, E \in \Sigma$  be a Markov transition probability function. Assume that  $S$  is a  $\sigma$ -compact metric space and that  $\Sigma$  is the  $\sigma$ -field of topological Borel sets  $C$ . (Then  $\Sigma$  is generated by a countable class of sets.) Suppose there is a  $\sigma$ -finite measure  $\lambda$  on  $\Sigma$  and a function  $q(s, x; t, y), a \leq s < t \leq b, (x, y) \in S \times S$  which is strictly positive,  $\Sigma \times \Sigma$ -measurable in  $(x, y)$ , and for which (3.1) holds. Let  $P(s, x; t, E; u, y), a \leq s < t \leq b, (x, y) \in S \times S, E \in \Sigma$ , be the reciprocal probability function derived from  $q(s, x; t, y)$ , let  $\mu$  be a probability measure on  $\Sigma \times \Sigma$ , and let  $X_t, \{a \leq t \leq b\}$  be the corresponding reciprocal process with endpoint distribution  $\mu$ . The following are equivalent:*

- (a)  $\{X_t, a \leq t \leq b\}$  is a Markov process.
- (b) There are measures  $\nu_a$  and  $\nu_b$  on  $\Sigma$  such that

$$\mu(G) = \int_G q(a, x; b, y) d(\nu_a \times \nu_b)(x, y), \quad G \in \Sigma \times \Sigma.$$

*Proof.* (b)  $\Rightarrow$  (a). Suppose (b) holds. Let  $a < t_1 < \dots < t_n < b$ , and  $E_i \in \Sigma, i = 1, \dots, n$ . For each  $(z_1, \dots, z_n) \in S^n$  let

$$\alpha(z_1, \dots, z_n) = q(t_1, z_1; t_2, z_2) \cdots q(t_{n-1}, z_{n-1}; t_n, z_n).$$

Let  $f$  be any non-negative  $\Sigma$ -measurable function on  $S$ . Referring to (2.2), (3.2), and (3.3), we see, after some cancellations, that

$$\begin{aligned} & \int_{\{X_{t_1} \in E_1, \dots, X_{t_n} \in E_n\}} f(X_{t_n}) dP \\ &= \int_{S \times E_1 \times \dots \times E_n \times S} q(a, x; t_1, z_1) \alpha(z_1, \dots, z_n) q(t_n, z_n; b, y) f(z_n) d\gamma(x, z_1, \dots, z_n, y), \end{aligned}$$

where  $\gamma$  is the product measure  $\nu_a \times \lambda^n \times \nu_b$ ,  $\lambda^n$  being the  $n$ -fold product of  $\lambda$  with itself. This last expression can be written as

$$\begin{aligned} & \int_{S \times E_1 \times \dots \times E_n} q(a, x; t_1, z_1) \alpha(z_1, \dots, z_n) f(z_n) \\ & \cdot \left[ \int_S q(t_n, z_n; b, y) \nu_b(dy) \right] d\bar{\gamma}(x, z_1, \dots, z_n), \end{aligned}$$

with  $\bar{\gamma} = v_a \times \lambda^n$ . Suppose  $f$  is defined by

$$f(w) = \frac{\int_{F \times S} q(t, z; b, y) d(\lambda \times v_b)(z, y)}{\int_S q(t_n, w; b, y) v_b(dy)}$$

where  $F \in \Sigma$ . Substituting in the previous expression, we have

$$\begin{aligned} \int_{\{X_{t_1} \in E_1, \dots, X_{t_n} \in E_n\}} f(X_{t_n}) dP_\mu &= \int_{S \times E_1 \times \dots \times E_n} q(a, x; t_1, z_1) \alpha(z_1, \dots, z_n) \\ &\cdot \left[ \int_{F \times S} q(t, z; b, y) d(\lambda \times v_b)(z, y) \right] d\bar{\gamma}(x, z_1, \dots, z_n) \\ &= \int_{S \times E_1 \times \dots \times E_n \times F \times S} q(a, x; t_1, z_1) \alpha(z_1, \dots, z_n) \\ &\cdot q(t, z; b, y) d\rho(x, z_1, \dots, z_n, z, y), \end{aligned}$$

where  $\rho = v_a \times \lambda^{n+1} \times v_b$ . Again using (2.2), (3.2), and (3.3), we see that this last expression is equal to  $P_\mu(X_{t_1} \in E_1, \dots, X_{t_n} \in E_n, X_t \in F)$ . Since all this is independent of the choice of  $E_1, \dots, E_n$ , what we have shown is that

$$P(X_t \in F | X_{t_1}, \dots, X_{t_n}) = f(X_{t_n}),$$

whence  $P(X_t \in F | X_{t_1}, \dots, X_{t_n}) = P(X_t \in F | X_{t_n})$ . Similar calculations lead to the same conclusion if  $t_1 = a$  or  $t_n = b$  or both. Thus (a) holds.

(a)  $\Rightarrow$  (b). Suppose (a) holds. Then there is a Markov transition probability function  $\bar{Q}(s, x; t, E)$ ,  $a \leq s < t \leq b$ ,  $E \in \Sigma$ , for the Markov processes  $\{X_t, a \leq t \leq b\}$ , and we may assume that  $\bar{Q}$  satisfies the Chapman-Kolmogorov equations in the following sense: for each  $E \in \Sigma$  and  $a \leq s < t < u \leq b$ ,

$$\bar{Q}(s, x; u, E) = \int \bar{Q}(s, x; t, dy) \bar{Q}(t, y; b, E) \quad (3.4)$$

for  $\mu_s$ -almost all  $x \in S$ . Then

$$\mu(E \times F) = \int_E \mu_a(dx) \bar{Q}(a, x; b, F) \quad (3.5)$$

for  $E, F$  in  $\Sigma$ .

(i) For each  $a \leq s < t < b$ ,  $\bar{Q}(s, x; t, \cdot)$  is equivalent to  $\lambda$  for  $\mu_s$ -almost all  $x \in S$ .

To prove (i) it suffices to verify that

$$\bar{Q}(s, X_s; t, F) = \int \bar{Q}(s, X_s; b, dy) \int_F p(s, X_s; t, z; b, y) \lambda(dz) \quad (3.6)$$

for each  $F \in \Sigma$  except for a  $P_\mu$ -null set of  $\omega \in \Omega$ . For each  $\omega \in \Omega$ , however, both sides of (3.6) are, as functions of  $F$ , probability measures on  $\Sigma$ . Since  $\Sigma$  is generated by a countable subfield, it suffices to show that, for each  $F \in \Sigma$ , (3.6) holds  $P_\mu$ -almost surely. It follows from the definition of  $P_\mu$  that

$$P_\mu(X_s \in E, X_t \in F) = \int_{E \times S} d\mu_{s,b}(x, y) \int_F p(s, x; t, z; b, y) \lambda(dz). \quad (3.7)$$

Also

$$\mu_{s,b}(A \times B) = \int_A \mu_s(dx) \bar{Q}(s, x; b, B), \quad (3.8)$$

which implies that for any non-negative  $\Sigma \times \Sigma$ -measurable  $f$ ,

$$\int_{E \times S} d\mu_{s,b} f(x, y) = \int_E \mu_s(dx) \int_S \bar{Q}(s, x; b, dy) f(x, y).$$

If

$$f(x, y) = \int_F p(s, x; t, z; b, y) \lambda(dz),$$

we have, using (3.7),

$$P_\mu(X_s \in E, X_t \in F) = \int_E \mu_s(dx) \int \bar{Q}(s, x; b, dy) \int_F p(s, x; t, z; b, y) \lambda(dz).$$

Since this last equation holds for each  $E \in \Sigma$ , it follows that the right hand side of (3.6) is a version of  $P_\mu(X_t \in F | X_s)$ . But so is  $\bar{Q}(s, X_s; t, F)$ , whence (3.6) follows  $P_\mu$ -almost surely. This completes the proof of (i).

(ii) For each  $a \leq s < b$ ,  $\bar{Q}(s, x; b, \cdot)$  is absolutely continuous with respect to  $\mu_b$  for  $\mu_s$ -almost all  $x \in S$ .

To prove (ii), pick  $t \in (s, b)$  and observe that since  $\mu_b(E) = \int \mu_s(dx) \bar{Q}(s, x; b, E)$  we have

$$\mu_b(E) = \int \mu_s(dx) \int \bar{Q}(s, x; t, dy) \bar{Q}(t, y; b, E) \tag{3.9}$$

by virtue of (3.4). Let  $S' = \{x: \bar{Q}(s, x; t, \cdot) \sim \lambda\}$ . The complement of  $S'$  is  $\mu_s$ -null by (i). Let  $S''$  be the set of  $x \in S$  such that (3.4) holds for all  $E \in \Sigma$ . Since  $\Sigma$  is generated by a countable field, the complement of  $S''$  is  $\mu_s$ -null. Let  $S_0 = S' \cap S''$ . Suppose  $x_0 \in S_0$ , and  $\mu_b(E) = 0$ . Then  $\bar{Q}(s, x; b, E) = 0$  for  $\mu_s$ -almost all  $x \in S$ , hence for some  $x_1 \in S_0$ . Since (3.4) holds for  $x = x_1$ ,  $\{y: \bar{Q}(t, y; b, E) > 0\}$  has  $\bar{Q}(s, x_1; t, \cdot)$ -measure zero, hence  $\lambda$ -measure zero, hence  $\bar{Q}(s, x_0; t, \cdot)$ -measure zero. Since (3.4) holds for  $x = x_0$ ,  $\bar{Q}(s, x_0; b, E) = 0$ . Since  $E$  is arbitrary,  $\bar{Q}(s, x_0; b, \cdot) \ll \lambda$ . Since the complement of  $S_0 = S' \cap S''$  is  $\mu_s$ -null, this proves (ii).

Fix  $t \in (a, b)$ . Let

$$\gamma(x, z) = \int \bar{Q}(s, x; b, dy) p(s, x; t, z; b, y) \quad (s, z) \in S \times S.$$

By virtue of the proof of (i) (see (3.6)),  $\gamma(x, \cdot)$  is, for  $\mu_s$ -almost all  $x$ , a density of  $\bar{Q}(a, x; t, \cdot)$  with respect to  $\lambda$ . Let  $\nu$  be the probability measure on  $\Sigma \times \Sigma$  determined by

$$\nu(E \times F) = \int_E \mu_t(dz) \bar{Q}(t, z; b, F), \quad E, F \in \Sigma.$$

We may infer from (ii) that  $\nu$  is absolutely continuous with respect to the product measure  $\mu_t \times \mu_b$  on  $\Sigma \times \Sigma$ . By an argument of Doob ([5], Chapt. VII, § 8) the Radon-Nikodym derivative  $d\nu/d(\mu_t \times \mu_b)$  has a  $\Sigma \times \Sigma$ -measurable version  $\delta(z, y)$ , so that

$$\nu(E \times F) = \int_E \mu_t(dz) \int_F \mu_b(dy) \delta(z, y).$$

Comparing the last two expressions for  $\nu(E \times F)$ , we see that for  $\mu_t$ -almost all  $z$ ,  $\delta(z, \cdot)$  is a density of  $\bar{Q}(t, z; b, \cdot)$  with respect to  $\mu_b$ . Because  $\bar{Q}(s, x; t, E)$  is a transi-

tion function for the Markov process  $\{X_t, a \leq t \leq b\}$ , we have

$$P_\mu(X_a \in A, X_t \in B, X_b \in C) = \int_A \mu_a(dx) \int_B \lambda(dz) \lambda(x, z) \int_C \delta(z, y) \mu_b(dy). \quad (3.10)$$

for each  $A, B, C$  in  $\Sigma$ . Using Doob's argument again, we find a  $\Sigma \times \Sigma$ -measurable function  $\rho(x, y)$  such that  $\rho(x, \cdot)$  is the Radon-Nikodym derivative of  $Q(a, x; b, \cdot)$  with respect to  $\mu_b$  for  $\mu_a$ -almost all  $x \in S$ . By (3.5) and the definition of  $\{X_t, a \leq t \leq b\}$  as the reciprocal process with endpoint distribution  $\mu$  and reciprocal transition function given by (3.2), we have

$$P_\mu(X_a \in A, X_t \in B, X_b \in C) = \int_A \mu_a(dx) \int_C \mu_b(dy) \rho(x, y) \int_B p(a, x; t, z; b, y) \lambda(dz), \quad (3.11)$$

for all  $A, B, C$  in  $\Sigma$ . But (3.10) and (3.11) imply that for  $\mu_a \times \lambda \times \mu_b$ -almost all  $(x, z, y)$ ,

$$\rho(x, y) p(a, x; t, z; b, y) = \gamma(x, z) \delta(z, y).$$

By Fubini's theorem there is a  $z_0$  such that for  $\mu_a \times \mu_b$ -almost all  $(x, y)$ ,

$$\rho(x, y) p(a, x; t, z_0; b, y) = \gamma(x, z_0) \delta(z_0, y).$$

Referring to (3.2), we see that for these  $(x, y)$ ,

$$\rho(x, y) = f(x) q(a, x; b, y) g(y), \quad (3.12)$$

where  $f$  and  $g$  are defined by

$$f(x) = \frac{\gamma(x, z_0)}{q(a, x; t, z_0)}, \quad g(y) = \frac{\delta(z_0, y)}{q(t, z_0; b, y)}.$$

From (3.12) and (3.5) we get

$$\begin{aligned} \mu(A \times B) &= \int_A \mu_a(dx) f(x) \int_B q(a, x; b, y) g(y) \mu_b(dy) \\ &= \int_A \nu_a(dx) \int_B q(a, x; b, y) \nu_b(dy), \end{aligned}$$

where  $\nu_a = f \cdot d\mu_a$ ,  $\nu_b = g \cdot d\mu_b$ . This shows that (a)  $\Rightarrow$  (b), which completes the proof of the theorem.

*Remark.* Condition (b) on  $\mu$  is simply that there exists a product measure  $\pi$  on  $\Sigma \times \Sigma$  such that  $d\mu/d\pi = q$ , where  $q(x, y) = q(a, x; b, y)$ .

Consider the following problem. Suppose that  $(S, \Sigma, \lambda)$  is a  $\sigma$ -finite measure space, and that  $q(x, y)$  is an everywhere positive,  $\Sigma \times \Sigma$ -measurable function on  $S \times S$  for which  $\int q(x, y) \lambda(dy) = 1$  for each  $x \in S$ . Suppose  $\mu_1$  and  $\mu_2$  are probability measures on  $\Sigma$ . Is there a probability measure  $\mu$  on  $\Sigma \times \Sigma$  which has  $\mu_1$  and  $\mu_2$  for marginals and which satisfies condition (b) of the theorem? That is, can we find measures  $\nu_1$  and  $\nu_2$  on  $\Sigma$  such that, if  $\mu$  is defined on  $\Sigma \times \Sigma$  by

$$\mu(E \times F) = \int_E \nu_1(dx) \int_F q(x, y) \nu_2(dy), \quad E \in \Sigma, F \in \Sigma \quad (3.13)$$

then

$$\begin{aligned} \mu(E \times S) &= \mu_1(E) & E \in \Sigma \\ \mu(S \times F) &= \mu_2(F) & F \in \Sigma \end{aligned} \tag{3.14}$$

both hold? Also, is  $\mu$  uniquely determined by (3.13) and (3.14)? Since  $\mu$  determines and is uniquely determined by the pair of measures  $\nu_1$  and  $\nu_2$ , the problem we are posing is that of the existence and uniqueness of solutions  $\nu_1$  and  $\nu_2$  for the functional equations

$$\mu_1(E) = \int_E \nu_1(dx) q(x, y) \nu_2(dy) \quad E \in \Sigma, \tag{3.15}$$

and

$$\mu_2(F) = \int_F \nu_2(dy) q(x, y) \nu_1(dx) \quad F \in \Sigma. \tag{3.16}$$

Suppose that  $\mu_i \ll \lambda$ ,  $i=1, 2$ . Let  $f = d\mu_1/d\lambda$ ,  $g = d\mu_2/d\lambda$ . Since  $d\mu/d(\nu_1 \times \nu_2) = q$ ,  $d(\nu_1 \times \nu_2)/d\mu = 1/q$ , from which it easily follows that  $\nu_i \ll \mu_i$ ,  $i=1, 2$ . Thus  $\nu_i \ll \lambda$ ,  $i=1, 2$ . Let  $\psi = d\nu_1/d\lambda$  and  $\phi = d\nu_2/d\lambda$ . Then (3.15) and (3.16) are equivalent to

$$f(x) = \psi(x) \int q(x, y) \phi(y) \lambda(dy) \quad x \in S, \tag{3.17}$$

$$g(y) = \phi(y) \int q(x, y) \psi(x) \lambda(dx) \quad y \in S. \tag{3.18}$$

These equations, with  $(S, \Sigma, \lambda)$  being the real line with Lebesgue measure, and with

$$q(x, y) = \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{(x-y)^2}{2\sigma^2}}, \tag{3.19}$$

were derived by Schrödinger ([11, 12]), who conjectured the existence and uniqueness (up to multiplicative constants) of the functions  $\phi$  and  $\psi$  except perhaps when  $f$  or  $g$  are especially “tückisch”. In [1], S. Bernstein stated without proof that the pair of functional equations, with  $q$  given by (3.19), has a solution provided  $f$  and  $g$  are continuous. In [7], Fortet used the method of successive approximations to prove the existence and uniqueness of non-negative solutions (3.17) and (3.18) for a wide class of continuous functions  $q(x, y)$  including (3.19), but with  $(S, \Sigma, \lambda)$  still the real line. In [2], Beurling formulated a problem which (when his  $n=2$ ) is that of the existence and uniqueness of solutions  $\nu_1$  and  $\nu_2$  to (3.15) and (3.16), except that  $S$  is a locally compact Hausdorff space,  $q$  is required to be continuous and the requirement that  $\int q(x, y) \lambda(dy) = 1$  for all  $x$  is dropped, there being no underlying measure  $\lambda$ . It turns out that if  $0 < a \leq q < b < \infty$ , then (3.15) and (3.16) have uniquely determined solutions  $\nu_1$  and  $\nu_2$  (if  $q$  is a Markovian density relative to  $\lambda$ , this requires that  $\lambda$  be finite, and so excludes the case for which  $\lambda$  is Lebesgue measure on the real line). Relaxing the assumption that  $q$  be bounded away from 0, he proves existence and uniqueness of positive but not necessarily finite measures  $\nu_1$  and  $\nu_2$  for which (3.15) and (3.16) hold if  $q > 0$  and if in addition

$$\iint \log q(x, y) \mu_1(dx) \mu_2(dy)$$

is finite. Beurling shows that this last condition can be replaced by a weaker but more complicated one. His uniqueness proof, however, is valid without his condition and we can extend his proof to yield existence as well.

**Theorem 3.2.** *Suppose  $S$  is a  $\sigma$ -compact metric space, that  $\mu_1$  and  $\mu_2$  are probability measures on its  $\sigma$ -field  $\Sigma$  of Borel sets, and that  $q$  is an everywhere continuous, strictly positive function on  $S \times S$ . Then there is a unique pair  $\mu, \pi$  of measures on  $\Sigma \times \Sigma$  for which*

(a)  $\mu$  is a probability measure and  $\pi$  is a  $\sigma$ -finite product measure.

(b) 
$$\mu(E \times S) = \mu_1(E), \quad \mu(S \times E) = \mu_2(E), \quad E \in \Sigma,$$

(c) 
$$\frac{d\mu}{d\pi} = q.$$

*Proof.* To say that  $S$  is  $\sigma$ -compact means that there is an increasing sequence  $A_1, A_2, \dots$  of compact subsets of  $S$  for which  $S = \bigcup_{n=1}^{\infty} A_n$ . Let  $B_n = A_n \times A_n$ , and let  $\Sigma_n = \Sigma \cap A_n = \{E \cap A_n : E \in \Sigma\}$ . Then  $\Sigma_n \times \Sigma_n$  is the class of Borel subsets of  $B_n$ . On  $B_n$ ,  $q$  is bounded above and away from zero below. By theorem I of [2] there exists a finite product measure  $\pi^n$  on  $\Sigma_n \times \Sigma_n$  and a measure  $\mu^{(n)}$  on  $\Sigma_n \times \Sigma_n$  such that

(i) 
$$\left. \begin{aligned} \mu^{(n)}(E \times A_n) &= \mu_1(E) \\ \mu^{(n)}(A_n \times E) &= \mu_2(E) \end{aligned} \right\} E \in \Sigma_n$$

(ii) 
$$\frac{d\mu^{(n)}}{d\pi^n} = q \quad \text{on } B_n.$$

We extend  $\mu^{(n)}$  and  $\pi^n$  to all of  $\Sigma$  by setting them equal to 0 on sets  $E \in \Sigma$  disjoint from  $B_n$ .  $\pi^n$  remains a product measure and (ii) holds throughout  $S \times S$   $\pi^n$ -almost surely. Let  $\mu_i^{(n)}$  be the marginals of  $\mu^{(n)}$  as so extended. There is a sequence  $\{\mu_k\}$  such that the restriction of  $\mu^{(m_k)}$  to  $B_m$  converges weakly for each  $m=1, 2, \dots$ . It is easy to see that this implies the existence of a measure  $\mu$  on  $\Sigma \times \Sigma$  whose restriction to  $\Sigma_m \times \Sigma_m$  is for each  $m$  the weak limit relative to  $C(B_m)$  of the sequence formed by the restrictions of  $\mu^{(m_k)}$  to  $\Sigma_m$ . Then  $\iint g d\mu^{(m_k)} \rightarrow \iint g d\mu$  for any continuous  $g$  on  $S \times S$  with support contained in one of the compact sets  $B_m$ . I claim that this convergence holds provided only that  $g$  is bounded and continuous on  $S \times S$ . Since  $\mu^{(n)}(S) = \mu^{(n)}(B_n) = \mu^{(n)}(A_n \times A_n) = \mu_1(A_n) \leq 1$  by (i), this certainly holds if we establish that  $\mu$  is a probability measure. It is clear from (ii), however, that the marginals  $\mu_i^{(n)}$  of  $\mu^{(n)}$  converge weakly to  $\mu_i$ ,  $i=1, 2$ . Since these are probability measures, the sequence  $\{\mu^{(n)}\}$  of probability measures is tight and the limit  $\mu$  of  $\mu^{(m_k)}$  is a probability measure ([3], p. 30) with  $\mu_1, \mu_2$  as marginals, which establishes (b) and half of (a). Now fix  $m$  and assume  $f \in C(S \times S)$  has support in  $B_m$ . The restriction of  $f/q$  to  $B_m$  belongs to  $C(B_m)$ , so by (ii)

$$\int f d\mu^{m_k} = \int (f/q) d\mu^{(m_k)} \rightarrow \int (f/q) d\mu \tag{3.20}$$

as  $n \rightarrow \infty$ . This shows that the restriction of  $\pi^{m_k}$  to  $B_m$  converges weakly to a limit  $\pi_m$  as  $k \rightarrow \infty$ . Again, it is easy to see that there is a measure  $\pi$  on  $\Sigma \times \Sigma$  whose restriction to  $B_m$  is  $\pi_m$ ,  $m=1, 2, \dots$ . It follows from (3.20) that  $d\pi/d\mu = 1/q$ , whence  $d\mu/d\pi = q$ . Since  $\pi = \pi_m$  is finite on  $B_m$ ,  $\pi$  is  $\sigma$ -finite.  $\pi^n$  is a product measure for each  $n$ , and an easy argument shows that each  $\pi_m$ , hence  $\pi$ , must therefore be a

product measure. This shows the existence of measures  $\pi$  and  $\mu$  as described in the theorem.

To establish that  $\pi$  and  $\mu$  are unique, assume that  $\pi'$  is a product measure and  $\mu'$  a probability measure for which (a) and (b) hold. Then

$$\mu_1(E) = \int_{E \times S} q d\pi = \int_{E \times S} q d\pi' \tag{3.21}$$

and

$$\mu_2(E) = \int_{S \times E} q d\pi = \int_{S \times E} q d\pi'$$

for each  $E \in \Sigma$ . Suppose  $\pi = v_1 \times v_2$ ,  $\pi' = v'_1 \times v'_2$ . Let  $h_1(x) = \int q(x, y) v_2(dy)$ ,  $x \in S$ ,  $h_2(y) = \int q(x, y) v_1(dx)$ ,  $y \in S$ , and let  $h(x, y) = h_1(x) h_2(y)$ ,  $(x, y) \in S \times S$ . Let  $k_1, k_2$ , and  $k$  be similarly defined but with  $v'_i$  replacing  $v_i$ ,  $i = 1, 2$ . Let  $g_1$  and  $g_2$  be bounded  $\Sigma$ -measurable functions on  $S$ , and let  $g(x, y) = g_1(x) g_2(y)$ ,  $(x, y) \in S \times S$ . By virtue of (3.21),  $\int g_i d\mu_i = \int g_i h_i d\pi_i$ ,  $i = 1, 2$ . Multiplying corresponding sides of these two equations, we have

$$\int g d(\mu_1 \times \mu_2) = \int g h d\pi.$$

Since  $h$  is strictly positive, we can rewrite this as

$$\int g h^{-1} d(\mu_1 \times \mu_2) = \int g d\pi. \tag{3.22}$$

(Of course  $h^{-1}$  denotes the reciprocal, not the inverse, of  $h$ .) Similarly

$$\int g k^{-1} d(\mu_1 \times \mu_2) = \int g d\pi'. \tag{3.23}$$

The definition of  $\Sigma \times \Sigma$  as the  $\sigma$ -field generated by the field of finite disjoint unions of rectangles  $E \times F$  with  $E, F \in \Sigma$  ensures that (5) and (6) hold for all non-negative  $\Sigma \times \Sigma$ -measurable functions  $g$ . Let  $\sigma_1$  and  $\sigma_2$  be bounded  $\Sigma$ -measurable functions on  $S$ , and let  $\sigma(x, y) = \sigma_1(x) + \sigma_2(y)$ ,  $(x, y) \in S \times S$ . Then

$$\begin{aligned} \int \sigma d(\mu_1 \times \mu_2) &= \int \sigma_1 d\mu_1 + \int \sigma_2 d\mu_2 \\ &= \iint \sigma_1(x) q(x, y) v_1(dx) v_2(dy) + \iint \sigma_2(y) q(x, y) v_1(dx) v_2(dy) \\ &= \int \sigma q d(v_1 \times v_2) = \int \sigma q h^{-1} d(\mu_1 \times \mu_2) \end{aligned}$$

by virtue of (3.21) and (3.22). Using  $v'_i$  instead of  $v_i$ ,  $i = 1, 2$ , we obtain similarly  $\int \sigma d(\mu_1 \times \mu_2) = \int \sigma q k^{-1} d(\mu_1 \times \mu_2)$ . We conclude that

$$\int \sigma q h^{-1} d(\mu_1 \times \mu_2) = \int \sigma q k^{-1} d(\mu_1 \times \mu_2). \tag{3.24}$$

Since  $\sigma$  is bounded and since  $q d\pi$  is a probability measure, the common value of the two sides of (3.24) is finite by virtue of (3.22) and (3.23). Thus (3.24) yields

$$\int \sigma q (h^{-1} - k^{-1}) d(\mu_1 \times \mu_2) = 0.$$

In particular, this last equation holds if

$$\sigma(x, y) = \frac{h_1^{-1}(x)}{h_1^{-1}(x) + k_1^{-1}(x)} - \frac{h_2^{-1}(y)}{h_2^{-1}(y) + k_2^{-1}(y)}, \quad (x, y) \in S \times S \tag{3.25}$$

from which we deduce

$$\int q(h^{-1} - k^{-1})^2 / r d(\mu_1 \times \mu_2) = 0, \quad (3.26)$$

where  $r(x, y)$  is the product of the denominators of the two fractions on the right hand side of (3.25). Since  $q > 0$ ,  $h = k$  on the support of  $\mu_1 \times \mu_2$ . It now follows from (3.22) and (3.23) that  $\pi = \pi'$ , and it follows from  $d\mu/d\pi = d\mu'/d\pi'$  that  $\mu = \mu'$ . This completes the proof of the theorem. (The very elegant proof of uniqueness is due to Beurling; we have changed his notation to conform with ours, and re-arranged his proof to exhibit its independence from his condition (8.1) on p. 198 of [2].)

#### § 4

If a reciprocal transition function  $P(s, x; t, \cdot; u, z)$  is absolutely continuous relative to a  $\sigma$ -finite measure  $\lambda$  on  $\Sigma$ , then there is a function  $p(s, x; t, y; u, z)$  for which

$$P(s, x; t, E; u, z) = \int_E p(s, x; t, y; u, z) \lambda(dy), \quad a \leq s < t < u \leq b \quad x, z \in S, E \in \Sigma. \quad (4.1)$$

If  $P(s, x; t, \cdot, u, z)$  is derived from a Markov transition density  $q(s, x; t, y)$  we have in fact

$$p(s, x; t, y; u, z) = \frac{q(s, x; t, y) q(t, y; u, z)}{q(s, x; u, z)}, \quad a \leq s < t < u \leq b, \quad x, y, z \in S. \quad (4.2)$$

Is any reciprocal transition density derived from a Markov transition density? More precisely, given that a function  $p(s, x; t, y; u, z)$  satisfies (4.1), does there exist a Markov transition density  $q(s, x; t, y)$  such that (4.2) holds? In this section we give partial answers to this question. First, to motivate our definition of reciprocal transition density, we list those properties of  $p(s, x; t, y; u, z)$  which follow by virtue of (4.1) and properties (A 1), (A 2) and (A 3) of  $P(s, x; t, \cdot; u, z)$ . As usual,  $a \leq s < t < u \leq b$ ,  $x, y, z$  are in  $S$ , and  $E \in \Sigma$ .

(a 1)  $y \rightarrow p(s, x; t, y; u, v)$  is  $\lambda$ -measurable, with

$$p(s, x; t, y; u, v) \geq 0 \quad \lambda\text{-almost all } y$$

and

$$\int p(s, x; t, y; u, v) \lambda(dy) = 1.$$

(a 2)  $(x, y) \rightarrow \int_E p(s, x; t, y; u, z) \lambda(dy)$  is  $\Sigma \times \Sigma$  measurable.

(a 3) For each  $a \leq s < t < u < v \leq b$ , and each  $x, w$  in  $S$ ,

$$p(s, x; u, z; v, w) p(s, x; t, y; u, z) = p(s, x; t, y; v, w) p(t, y; u, z; v, w)$$

for  $\lambda \times \lambda$ -almost all  $(y, z) \in S \times S$ .

This last property is an almost immediate consequence of (A 3), which in turn is analogous to the Chapman-Komogorov equation satisfied by Markov transition functions. However, its consequence (a 3) for densities is not an integral equation as in the Markov case but a pointwise, nonintegrated equality which right away yields our first result. We require of our definition sharper versions of (a 1)–(a 3).



As in Section 2,  $[a, b]$  is a non-degenerate closed interval, and  $(S, \Sigma)$  is a measurable space. We use  $\mathcal{E}$  to denote the set of all ordered sextuples  $(s, x, t, z, u, y)$  for which  $x, y,$  and  $z$  are in  $S$  and  $a \leq s < t < u \leq b$ . Let  $\lambda$  be a  $\sigma$ -finite measure on  $\Sigma$ . A function  $p$  on  $\mathcal{E}$  to the (positive) non-negative reals is called a (strictly positive) reciprocal transition probability  $\lambda$ -density if the following conditions are satisfied.

(b1) For each  $a \leq s < t < u \leq b$ , the map  $(s, y, z) \rightarrow p(s, x, t, y, u, z)$  is  $\Sigma \times \Sigma \times \Sigma$ -measurable, and

$$\int p(s, x, t, y, u, z) \lambda(dz) = 1 \quad x, y \text{ in } S.$$

(b2) For each  $a \leq s < t < u < b \leq b$  and  $x, y, z, w$  in  $S$ ,

$$p(s, x, u, z, v, w) p(s, x, t, y, u, z) = p(s, x, t, y, v, w) p(t, y, u, z, v, w).$$

If (b1) and (b2) are satisfied then the function  $P$  on  $\mathcal{D}$  (see Section 2) defined by

$$P(s, x, t, E, u, z) = \int_E p(s, x, t, z, u, y) \lambda(dz) \tag{4.3}$$

is a reciprocal transition probability function. We write  $p(s, x; t, y; u, z)$  for  $p(s, x, t, y, u, z)$ . We pose but otherwise ignore the question of whether a density  $p(s, x; t, y; u, z)$  satisfying (a1), (a2), and (a3) has a version satisfying (b1) and (b2).

**Theorem 4.1.** *Let  $p(s, x; t, y; u, v)$  be a strictly positive reciprocal transition  $\lambda$ -density on  $[a, b]$ . Then for each  $b' \in (a, b)$  there is a Markov transition  $\lambda$ -density for which*

$$p(s, x; t, y; u, v) = \frac{q(s, x; t, y) q(t, y; u, z)}{q(s, x; u, z)}, \quad a \leq s < t < u \leq b', \quad x, y, z \text{ in } S.$$

*Proof.* In property (b2) set  $v = b'$ , fix  $w \in S$  and let  $q(s, x; t, y) = p(s, x; t, y; w, b')$ .

There are processes defined on  $(-\infty, \infty)$  which are reciprocal on an interval  $[a, b]$  but on no strictly larger super-interval (for example, the process discussed by Slepian in [12]). Thus we wish to replace  $b' < b$  by  $b$  itself. One would think it possible to concoct some simple limiting argument and let  $b' \rightarrow b$ . We are able to obtain the result only under some restrictions on  $p(s, x; t, y; u, z)$  and by rather involved reasoning. We first give an example to show that not all discrete-parameter reciprocal processes are derived from Markov transition functions. Given  $X_1, \dots, X_n$  reciprocal, then  $X_1, \dots, X_{n-1}$  is derived from a Markov transition function, but there may be an "endpoint effect" ensuring that  $X_1, \dots, X_n$  is not so derived. Any process  $X_1, X_2, X_3$  is reciprocal for the same trivial sort of reason that any process  $X_1, X_2$  is Markovian. Take  $S = \{0, 1\}$ , and let  $p(x|y|z) = P(X_2 = y | X_1 = x, X_2 = z)$ , where  $x, y,$  and  $z$  range over  $\{0, 1\}$ . For our example of a reciprocal process not derived from a Markov process we choose  $p(x|y|z)$  so that there is no system of Markov transition functions  $q(i, x; j, y), 1 \leq i < j \leq 3, x, y$  in  $S$  for which

$$p(x|y|z) = \frac{q(1, x; 2, y) q(2, y; 3, z)}{q(1, x; 3, z)} \quad x, y, z = 0, 1. \tag{4.4}$$

We determine  $p(x|y|z)$  by the condition that

$$p(x|0|z) = \begin{cases} \frac{1}{3} & x = z \\ \frac{2}{3} & x \neq z \end{cases} \quad x, z = 0, 1. \tag{4.5}$$

Define  $F(x)$  to be the quotient of  $p(x|0|0)/p(x|0|1)$  by  $p(x|1|0)/p(x|1|1)$ . Suppose (4.4) holds. Then  $p(x|y|z) = f(x, y) g(y, z) h(x, z)$  for some functions  $f, g, h$ , and

$$F(x) = \frac{f(x, 0) g(0, 0) h(x, 0)}{f(x, 0) g(0, 1) h(x, 1)} \cdot \frac{f(x, 1) g(1, 1) h(x, 1)}{f(x, 1) g(1, 0) h(x, 0)}$$

which is independent of  $x$ . We see from (4.5), however, that  $F(0) = 1/4$ ,  $F(1) = 1/3$ . Contradiction: (4.4) cannot hold, and  $p(x|y|z)$  is not derived from a Markov transition function.

We need the following lemma, which is of interest in its own right as a partial converse to the results of Section 3.

**Lemma 4.2.** *Let  $(S, d)$  be a  $\sigma$ -compact metric space with  $\Sigma$  the  $\sigma$ -field generated by the open sets of  $S$ , and let  $\lambda$  be a  $\sigma$ -finite measure on  $\Sigma$ . Let  $p(s, x; t, y; u, z)$  be a reciprocal transition  $\lambda$ -density on  $[a, b]$ . Let  $\mu$  be a probability measure on  $\Sigma \times \Sigma$  both of whose marginals are absolutely continuous with respect to  $\lambda$  and have strictly positive densities. Let  $\{X_t, a \leq t \leq b\}$  be the reciprocal process with transition function given by (4.3) and endpoint distribution  $\mu$ . If  $\{X_t, a \leq t \leq b\}$  is Markov, then  $\mu \ll \lambda \times \lambda$ ,  $d\mu/d(\lambda \times \lambda)$  has a strictly positive version, and there is a Markov transition density  $q(s, x; t, y)$  such that (4.2) holds for each  $a \leq s < t < u \leq b$  and  $\lambda \times \lambda \times \lambda$ -almost all  $(x, y, z)$ .*

*Proof.* Assume the hypotheses of the theorem. There are everywhere positive measurable functions  $f$  and  $g$  on  $S$  for which

$$\mu(E \times S) = \int_E f d\lambda, \quad \mu(S \times E) = \int_E g d\lambda, \quad E \in \Sigma.$$

Let  $x \rightarrow \mu(\cdot, x)$  and  $x \rightarrow \mu(x, \cdot)$  be conditional distributions of  $\mu$  given the sub- $\sigma$ -fields  $\{E \times S: E \in \Sigma\}$  and  $\{S \times E: E \in \Sigma\}$  respectively. For each  $x, y$  in  $S$ , let

$$\begin{aligned} r(s, x, t, y) &= \int p(a, x'; s, x; b, z) p(s, x; t, y; b, z) d\mu(x', z), & a < s < t < b, \\ r(a, x, t, y) &= f(x) \int p(a, x; t, y; b, z) \mu(x, dz), & a < t < b, \\ r(t, x, b, y) &= g(y) \int p(a, x'; t, x; b, y) \mu(dx', y), & a < t < b. \end{aligned} \tag{4.6}$$

For each choice of  $(s, t)$  with  $a \leq s < t \leq b$  other than  $(s, t) = (a, b)$ ,  $r(s, x; b, y)$  is the value at  $(x, y)$  of the joint density of  $(X_s, X_t)$  with respect to  $\lambda \times \lambda$ . This can be checked using (2.2). For example,

$$\begin{aligned} \int_{E \times F} r(a, x, t, y) d(\lambda \times \lambda)(x, y) &= \iiint_{E \times F} p(a, x; t, y; b, z) \lambda(dy) \mu(x, dz) f(x) \lambda(dx) \\ &= \int_{E \times S} P(a, x; t, F; b, z) d\mu(x, z) \\ &= P_\mu(X_a \in E, X_t \in F), \end{aligned}$$

and the others are similar. Let

$$\rho(t, x) = \int r(a, x', t, x) \lambda(dx') \quad a < t < b;$$

then  $\rho(t, x)$  is the value at  $x$  of the conditional density of  $X_t$ . For each  $x, y$  in  $S$  define

$$q(a, x; t, y) = \frac{r(a, x, t, y)}{f(x)} \quad a < t < b,$$

$$q(t, x; b, y) = \frac{r(t, x, b, y)}{\rho(t, x)} \quad a < t < b,$$

$$q(s, x; t, y) = \frac{r(s, x, t, y)}{\rho(s, x)} \quad a < s < t < b.$$

Then, for each  $(s, t)$  with  $a \leq s < t \leq b$ , except  $(s, t) = (a, b)$ ,  $q(s, x; t, y)$  is the value at  $y$  of the conditional density of  $X_t$  given  $X_s = x$ . Were  $\mu$  assumed absolutely continuous with respect to  $\lambda \times \lambda$ , then  $d\mu/d(\lambda \times \lambda)$  would be the joint density of  $X_a$  and  $X_b$ , and we could write down the corresponding conditional density. However, we are assuming only that  $\mu$  has  $\lambda$ -absolutely continuous marginals, and this by itself does not imply that  $\mu \ll \lambda \times \lambda$ . This last is indeed true, because  $\{X_t, a \leq t \leq b\}$  is Markovian. Fixing  $t \in (a, b)$ , we have already established the existence of joint densities for  $(X_a, X_t)$  and for  $(X_t, X_b)$ . It follows from the Chapman-Kolmogorov equation that  $X_a$  and  $X_b$  have a joint density, in other words, that  $\mu \ll \lambda \times \lambda$ . The argument of Doob used in the proof of Theorem 3.1 shows that there is a version  $r(x, y)$  of this joint density which is  $\Sigma \times \Sigma$ -measurable in  $(x, y)$ , and then  $q(a, x; b, y) = r(x, y)/f(x)$  is the value at  $y$  of the conditional density of  $X_b$  given  $X_a = x$ . The Chapman-Kolmogorov equation also shows that we may choose  $r(x, y)$ , hence  $q(a, x; b, y)$ , strictly positive. The conditional densities  $q(s, x; t, y)$  are now defined for all  $a \leq s < t \leq b$  and  $x, y$  in  $S$ . On the one hand, the value at  $y$  of the conditional density of  $X_t$  given  $X_s = x$  and  $X_n = z$  is given by  $p(s, x; t, y; u, z)$ . On the other hand, the fact that  $\{X_t, a \leq t \leq b\}$  is a Markov process with transition density  $q$  enables us to write this conditional density as the quotient of  $p(s, x) q(s, x; t, y) q(t, y; b, z)$  by  $\rho(s, x) q(s, x, u, y)$ . This establishes (4.2) and proves the lemma.

**Theorem 4.3.** *Assume the conditions on  $S, \Sigma$ , and  $\lambda$  given in the previous lemma. Suppose that  $p(s, x; t, y; u, z)$  is a strictly positive transition  $\lambda$ -density on  $[a, b]$  satisfying the following conditions:*

(i) *For each  $a \leq s < t < u \leq b$ , the map  $(x, y, z) \rightarrow p(s, x; t, y; u, z)$  is continuous on  $S \times S \times S$  and for each  $a < s_0 < t_0 < u_0 < v_0 < b$  is bounded uniformly in  $s \in [a, s_0]$ ,  $u \in [v_0, b]$ , and  $x, y, z$  in  $S$ .*

(ii) *For each  $t \in (a, b)$  and  $x, y, z$  in  $S$ ,*

$$\lim_{\substack{a' \downarrow a, b' \uparrow b}} p(a', x; t, y; b', z) = p(a, x; t, y; b, z)$$

*the limit approach being uniform for  $(s, y, z)$  in any compact subset of  $S \times S \times S$ .*

(iii) *There is a  $\eta > 0$  such that for each  $\delta > 0$  and  $x_0, z_0$  in  $S$ , and compact  $K$*

$$\lim_{u \rightarrow 0} \frac{1}{u} \max_1 \int_{R(y, \delta)} p(s, x; t, y; b, z_0) \lambda(dy) = 0,$$

$$\lim_{u \rightarrow 0} \frac{1}{u} \max_2 \int_{R(y, \delta)} p(a, x_0; s, y; t, x) \lambda(dy) = 0,$$

where  $\max_1$  is taken over  $a \leq s \leq a + \eta$ ,  $s < t < s + u$  and  $x \in K$ ,  $\max_2$  is taken over  $b - \eta \leq s \leq b$ ,  $t - u < s < t$ , and  $x \in K$ , while  $R(y, \delta)$  is the complement of the sphere of radius  $\delta$  centered at  $y$ .

Then there is a Markov transition  $\lambda$ -density  $q(s, x; t, y)$ ,  $a \leq s < t \leq b$ ,  $x, y$  in  $S$ , such that (4.2) holds for  $\lambda \times \lambda \times \lambda$ -almost all  $(x, y, z)$  in  $S \times S \times S$ , and all  $(x, y, z)$  if  $s > a$  and  $u < b$ .

*Proof.* Assume the hypotheses of the theorem. Let  $f$  and  $g$  be strictly positive  $\Sigma$ -measurable functions on  $S$  with  $\int f d\lambda = \int g d\lambda = 1$ . Let  $a_n \downarrow a < b \uparrow b$ . By virtue of the lemma, there is a Markov transition density  $q_n(s, x; t, y)$ ,  $a_n < s < t < b_n$  such that (4.2) holds. By (ii), we may assume that  $q_n(s, x; t, y; t, z)$  is continuous in  $(x, y, z)$  for each  $a_n \leq s < t < u \leq b_n$ . By theorem 3.2, for each  $n$  there is a measure  $\mu_n$  on  $\Sigma \times \Sigma$  whose marginals are given by the  $\lambda$ -densities  $f$  and  $g$  such that the reciprocal process  $\{X_t, a_n \leq t \leq b_n\}$  with transition density  $p(s, x; t, y; b, z)$  and endpoint measure  $\mu_n$  is Markov. Since the marginals of  $\mu_n$  do not depend on  $n$ ,  $\{\mu_n\}$  has a weakly convergent subsequence. Let  $\mu$  be its limit. The marginals of  $\mu$  are given by the densities  $f$  and  $g$ . We may assume without loss of generality that  $\{\mu_n\}$  itself converges. Consider now the process  $\{X_t, a \leq t \leq b\}$  determined by the reciprocal transition density  $p(s, x; t, y; u, t)$  and the endpoint measure  $\mu$ .

(a)  $\{X_t, a < t < b\}$  is Markov.

To prove (a), let  $a < s_1 < \dots < s_k < t < b$ . Choose  $n$  large enough so that  $a_n < s_1$ ,  $b_n < t$ . Let

$$\begin{aligned} \pi(x, x_1, \dots, x_k, y, z) &= p(a, x; s_1, x_1; b, z) p(s_1, x_1; s_2, x_2; b, z) \cdots \\ &\quad \cdots p(s_k, x_k; t, y; b, z), \end{aligned}$$

$$\begin{aligned} \sigma(x, x_1, \dots, x_k, z) &= p(a, x; s_1, x_1; b, z) p(s_1, x_1; s_2, x_2; b, z) \cdots \\ &\quad \cdots p(s_{k-1}, x_{k-1}; s_k, x_k; b, z) \end{aligned}$$

for  $x, x_1, \dots, x_k, y, z$  in  $S$ , and let  $\pi_n$  and  $\sigma_n$  be defined in exactly the same way, but with  $a_n$  and  $b_n$  replacing  $a$  and  $b$  respectively.  $(X_{s_1}, \dots, X_{s_k}, X_t)$  has a joint  $\lambda$ -density, and the value at  $y$  of the conditional density of  $X_t$  given  $X_{s_1} = x_1, \dots, X_{s_k} = x_k$  is easily seen to be equal to

$$\frac{\int \pi(x, x_1, \dots, x_k, y, z) d\mu(x, z)}{\int \sigma(x_1, x_1, \dots, x_k, z) d\mu(x, z)}. \quad (4.7)$$

It follows from conditions (i) and (ii) that for fixed  $x_1, \dots, x_k$  and  $y$   $\{\pi_n(x, x_1, \dots, x_k, y, z)\}$  is bounded uniformly in  $x, y$  and  $z$  and converges to  $\pi(x, x_1, \dots, x_k, y, z)$

uniformly for  $(x, y, z)$  in compact subsets of  $S \times S \times S$ . Therefore (4.7) is equal to

$$\frac{\lim_n \int \pi_n(x, x_1, \dots, x_k, y, z) d\mu_n(x, z)}{\lim_n \int \sigma_n(x, x_1, \dots, x_k, z) d\mu_n(x, z)} = \lim_n \frac{\int \pi_n(x, x_1, \dots, x_k, y, z) d\mu_n(x, z)}{\int \sigma_n(x, x_1, \dots, x_k, z) d\mu_n(x, z)}.$$

The expression whose limit is being taken is the value at  $y$  of the corresponding conditional density relative to the process  $\{X_t, a_n \leq t \leq b\}$  with transition density  $p(s, x; t, y; u, z), a_n \leq s < t < u \leq b_n$ , and endpoint measure  $\mu_n$ . This process is Markov, so the conditional density in question is independent of  $x_1, \dots, x_{k-1}$ . The same must be true of the limit (4.7). Since  $k$  and  $a < s_1 < \dots < s_k < t < b$  are arbitrary, this proves (a).

(b) *Almost all paths of  $\{X_t, a \leq t \leq b\}$  are continuous on  $[a, a + \eta]$  and  $[b - \eta, b]$ .*

First, fix  $x_0$  and  $z_0$  in  $S$ , and consider the reciprocal process  $\{X_t, a \leq t \leq b\}$  with transition density  $p(s, x; t, y; b, z)$  and endpoint measure equal to the point mass  $\delta_{(x_0, y_0)}$ . Then  $\{X_t, a \leq t \leq b\}$  is a Markov process with transition  $q$ -density  $p(s, x; t, y; b, z_0)$ , as is therefore  $\{X_t, a \leq t \leq a + \eta\}$ . The first condition of (iii) is guaranteed that this latter process has continuous paths by virtue of the corollary to Theorem 6.6 of [6].  $\{X_t, a \leq t \leq b\}$  with time reversed is also a Markov process, with transition density  $\rho(t, x; s, y) = p(a, x_0; s, y; t, x)$  for  $a \leq s < t \leq b, x, y \in S$ . The second condition of (iii) guarantees that  $\{X_t, b - \eta \leq t \leq b\}$  has continuous paths. Thus  $\{X_t, a \leq t \leq b\}$  has the desired continuity if the endpoint measure is a point mass, and it follows immediately that the same is true for any endpoint measure, in particular for the endpoint measure  $\mu$ . This proves (b).

(c)  *$\{X_t, a \leq t \leq b\}$  is a Markov process.*

First we show that  $\{X_t, a \leq t < b\}$  is a Markov process. For each  $a \leq c < d \leq b$  let  $\mathcal{F}_{[c, d]}$  be the  $\sigma$ -field generated by  $\{X_t; c \leq t \leq b\}$ . We need to show that  $E\{h(X_{t'}) | \mathcal{F}_{[0, t]}\} = E\{h(X_{t'}) | X_t\}$  for any  $a \leq t < t' \leq b$  and bounded  $\Sigma$ -measurable function  $h$ . Since  $\{X_t, a \leq t < b\}$  is a Markov process by (a), we know that  $E\{h(X_{t'}) | \mathcal{F}_0\} = E\{h(X_{t'}) | X_t\}$ , where  $\mathcal{F}_0$  is the smallest  $\sigma$ -field containing all the  $\sigma$ -fields  $\mathcal{F}_{[a + \frac{1}{n}, t]}$ ,  $n = 1, 2, \dots$ . Let  $Z$  be a bounded random variable measurable with respect to  $\mathcal{F}_0$ , let  $k$  be a bounded continuous function on  $S$  and let  $Y = k(X_a)Z$ . Then

$$\int Y h(X_{t'}) = \int k(X_a) Z h(X_{t'}) = \int \lim_n k(X_{a + \frac{1}{n}}) Z h(X_{t'}) = \lim_n \int k(X_{a + \frac{1}{n}}) Z h(X_{t'}),$$

the last two equalities holding by virtue of the continuity of  $k$ , the continuity (by (b)) at  $a$  of the paths of  $\{X_t, a \leq t \leq b\}$  and the bounded convergence theorem. But  $k(X_{a + \frac{1}{n}})Z$  is measurable with respect to  $\mathcal{F}_0$  for any  $n = 1, 2, \dots$ , so

$$\int k(X_{a + \frac{1}{n}}) Z h(X_{t'}) = \int k(X_{a + \frac{1}{n}}) Z E\{h(X_{t'}) | X_t\}.$$

Letting  $n \rightarrow \infty$ , and using again continuity and the bounded convergence theorem, we have  $\int Y h(X_{t'}) = \int Y E\{h(X_{t'}) | X_t\}$ . This is true for any  $Y$  of the form  $k(X_a)Z$ , with  $k$  continuous and  $Z$  bounded and  $\mathcal{F}_0$ -measurable. But bounded  $\mathcal{F}_{[a, t]}$ -measurable functions can be approximated by linear combinations of such  $Y$ 's, so the

last equality holds for arbitrary bounded  $\mathcal{F}_{[a,t]}$  measurable  $Y$ . It follows that  $E\{h(X_t)|\mathcal{F}_{[0,t]}\} = E\{h(X_t)|X_t\}$ . Reversal of the direction of time and use of the continuity of the paths of  $\{X_t, a \leq t \leq b\}$  at  $b$  show that  $\{X_t, a \leq t \leq b\}$  is Markovian, establishing (c).

Since  $\{X_t, a \leq t \leq b\}$  is a Markov process, and since the marginals  $f$  and  $g$  of the endpoint measure  $\mu$  are strictly positive, Lemma 4.2 applies to yield a Markov transition function  $g(s, x; t, y)$  for which (4.2) holds for  $\lambda \times \lambda \times \lambda$ -almost all  $(x, y, z)$  in  $S \times S \times S$ . If  $s > a$  and  $u < b$ , the joint density  $r(s, x; t, y)$  of  $(X_s, X_t)$  is continuous in  $(x, y)$  by virtue of assumption (i), (4.6), and the bounded convergence theorem, while the density  $\int p(a, w; s, x; b, y) d\mu(w, y)$  of  $X_s$  is continuous in  $x$  again by (ii) and the bounded convergence theorem. Thus  $q(s, x; t, y)$ , which is the quotient of  $r(s, x; t, y)$  and the (everywhere positive) density of  $x$ , is itself a continuous function of  $(x, y)$ , from which it follows that both sides of (4.2) are continuous functions of  $(x, y, z)$  if  $s > a$  and  $u < b$ , and (4.2) therefore holds for all  $(x, y, z)$  under this restriction. This completes the proof of the theorem.

We note that in keeping with the statement of the theorem, the reciprocal process  $\{X_t, 0 \leq t \leq 1\}$  studied by Slepian in [13] has a reciprocal transition density which satisfies (4.2) with  $q(s, x; t, y)$  being the transition density of the standard Brownian motion on  $[0, 1]$  and endpoint measure  $\mu$  the product measure whose factors are the one-dimensional Gaussian measures with zero mean and unit variance.

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