

A Ratio Ergodic Theorem for Increasing Additive Functionals

K. Bruce Erickson*

Department of Mathematics, University of Washington, Seattle, WA 98195, USA

Summary. Let B be a 1-dimensional Brownian motion. In this paper ratios of the form $A^+(t)/A^-(t)$, where A^+ is the $(0, \infty)$ -occupation time functional of B and A^- is a local time integral of an infinite (but locally finite) measure m with support in $(-\infty, 0]$, are studied. Conditions on m are given which ensure that such a ratio will be unbounded a.s. (or go to zero a.s.) as $t \rightarrow \infty$.

1. Introduction and Statement of Main Results

Let $\{B(t), t \geq 0\}$ be a standard Brownian motion on \mathbb{R} with $B(0)=0$ and let $L(t, x)$ be the local time functional

$$L(t, x) = \lim_{\varepsilon \rightarrow 0+} \text{meas.} \{s; s \leq t, x \leq B(s) \leq x + \varepsilon\} / \varepsilon.$$

Let m be a measure concentrated on $(-\infty, 0]$ and which satisfies

$$(1.1) \quad m(-\infty, 0] = \infty, \quad m\{I\} < \infty \quad \text{for bounded } I.$$

For $t \geq 0$ put

$$(1.2) \quad \begin{aligned} A^-(t) &= \int_{-\infty}^{0+} L(t, x) m\{dx\}, \\ A^+(t) &= \int_0^t I_{(0, \infty)}(B_s) ds = \int_0^\infty L(t, x) dx. \end{aligned}$$

(Note that $A^-(t) = \int_0^t m^\bullet(B_s) ds$ in the case that $m\{dx\} = m^\bullet(x) dx$.) The purpose of this paper is to determine when the ratio

$$K(t) = A^+(t)/A^-(t)$$

* The work was supported in part by a grant from the National Science Foundation.

is bounded or unbounded, a.s., as $t \rightarrow \infty$. More specifically let us write

$$k^* = \limsup_{t \rightarrow \infty} K(t) \quad \text{and} \quad k_* = \liminf_{t \rightarrow \infty} K(t).$$

By the 0-1 law k^* and k_* are constants a.s. which, as we will see, independently of each other (but for $k_* \leq k^*$), must be 0 or ∞ and our goal is to find criteria expressed as directly as possible in terms of m for deciding which of the two possibilities prevails. Note that if m were a finite measure, then we would immediately obtain from the ratio ergodic theorem, [5], p. 228, that $k^* = k_* = \infty$, so our results could be viewed as extensions of that theorem.

Our motivation for studying these limits stems in part from a recent (1982) paper by London, McKean, Rogers and Williams [7]. Let $A = A^+ - A^-$. A is a continuous additive functional which decreases when $B(t) \in \text{supp}(m)$ and increases (linearly) when $B(t) > 0$. Put $A^{-1}(t) = \inf\{s : A(s) \geq t\}$, $\inf \phi = \infty$, and $Y(t) = B(A^{-1}(t))$, $B(\infty) = \text{cemetery point } \delta$. Y is a Feller Brownian motion: a strong Markov process with state space $[0, \infty) \cup \delta$ which behaves like Brownian motion on $(0, \infty)$ (i.e. its local generator is $(1/2)d^2/dy^2$ there). The relationship of the measure m to the behavior of Y at the origin is the subject of [7]. The results of our paper pertain directly to the finiteness of the lifetime $\eta = \inf\{t : Y(t) = \delta\}$. As one may easily show

$$\eta = \infty \quad \text{if and only if} \quad a^* = \limsup_{t \rightarrow \infty} A(t) = +\infty,$$

and, under (1.1), $a^* = +\infty$ or $-\infty$ according as $k^* = \infty$ or 0. Note that in terms of the characteristics $(p_1, p_2, 0, p_4)$ of Y , $\eta = \infty$ if and only if p_1 , the killing rate at the origin, is 0. See [7], p. 44. We will not make any further reference to the process Y in the remainder of the paper.

Statement of Main Results. For any α put

$$D_\alpha = \inf\{t : B(t) = \alpha\}$$

and let

$$(1.3) \quad W = A^-(D_1) = \int_{-\infty}^{0+} L(D_1, x) m\{dx\}.$$

We will occasionally write f_t for $f(t)$ for functions on $[0, \infty)$.

Theorem 1. (i) $k^* = 0$ or ∞ according as

$$(1.4i) \quad \int_1^\infty \left(\int_0^x P(W > t) dt \right)^{-1} \frac{dx}{\sqrt{x}} < \infty \quad \text{or} \quad = \infty.$$

(ii) $k_* = \infty$ or 0 according as

$$(1.4ii) \quad \int_1^\infty P(W > t) \frac{dt}{\sqrt{t}} < \infty \quad \text{or} \quad = \infty.$$

(Note that the convergence of the integral (1.4ii) is equivalent to the finiteness of the moment $E\sqrt{W}$.)

For any positive decreasing function h on $(0, \infty)$, define

$$(1.5) \quad J^*(h) = \int_1^\infty \left(\int_0^x h(t) dt \right)^{-1} \frac{dx}{\sqrt{x}}$$

$$J_*(h) = \int_1^\infty h(x) \frac{dx}{\sqrt{x}}.$$

With m as in (1.1) put, for $x \geq 0$,

$$m_+ \{dx\} = m \{-dx\}, \quad m_+(x) = m_+[0, x] = m\{[-x, 0]\},$$

and define $\alpha(t)$ and $\beta(t)$ by

$$(1.6) \quad \alpha(t) = \inf \{a: a m_+(a) \geq t\} \vee 1,$$

$$(1.7) \quad \beta(t) = \sup \left\{ b: \frac{1}{b} \int_0^{b+} x^2 m_+ \{dx\} \leq t \right\} \vee 1.$$

It is useful to note that under (1.1)

$$(1.8) \quad \alpha(t) \uparrow \infty, \quad \beta(t) \uparrow \infty, \quad \alpha(t) \leq \beta(t), \quad \alpha(t) = o(t), \quad \text{as } t \rightarrow \infty.$$

We leave the easy verification to the reader. Note that α is also continuous.

Theorem 2. *We have the following implications:*

- (i) $J^*(1/\alpha) = \infty \Rightarrow k^* = \infty$.
- (ii) $J^*(1/\beta) < \infty \Rightarrow k^* = k_* = 0$.
- (iii) $J_*(1/\alpha) < \infty \Rightarrow k_* = k^* = \infty$.
- (iv) $J_*(1/\beta) = \infty \Rightarrow k_* = 0$.

Remark. It may be helpful to note that (1.8) implies $J^*(1/\alpha) \leq J^*(1/\beta)$, and $J_*(1/\beta) \leq J_*(1/\alpha)$, and that for any $0 < h \in \downarrow$ at most one of $J^*(h)$, $J_*(h)$ can be finite. This latter may be proved as in [2], p. 376.

Theorem 3. *If in addition to (1.1) we also have*

$$(1.9) \quad \limsup m_+(x)/m_+(2x) < 1,$$

then (i) $k^* = \infty \Leftrightarrow J^*(1/\alpha) = \infty$; (ii) $k_* = 0 \Leftrightarrow J_*(1/\alpha) = \infty$. (For a generalization see 5(b).)

Example. Interesting cases occur when m is near Lebesgue measure. Suppose for example $m \{dx\} = (\log^+ |x|)^r dx$ for $x \leq 0$. Then by Theorems 1 and 3

$$\limsup_{t \rightarrow \infty} \int_0^t I_{(0, \infty)}(B_s) ds \Big/ \int_0^t (\log^+ |B_u \wedge 0|)^r du = \infty \text{ or } 0$$

according as $r \leq 2$ or $r > 2$, and the $\liminf(\bullet) = 0$ or ∞ according as $r \geq -2$ or $r < -2$.

We prove Theorem 1 by a method similar, initially, to the proof of the ergodic theorem in [5]. By sampling A^+ and A_- at the successive passage

times (up and down) across a fixed interval, we find that $K(t)$ can be replaced by a ratio of two independent sums of positive independent random variables with infinite means. An application of some random walk methods in Erickson [2] and Kesten [6] completes the proof. The proof of Theorems 2 and 3 requires an asymptotic evaluation of $P(W > t)$. This is accomplished by conditioning the integral at (1.3) on $b = \min \{B_s; s \leq D_1\}$, applying Ray's representation of $L(D_1, x)$ in terms of BES(4) and estimating the integrals which develop. In the special case that m is regularly varying, application of a Tauberian theorem gives a more precise evaluation of $P(W > t)$, see § 5(a).

2. Proof of Theorem 1

Step 1. Let A_0^+ be occupation time of $[1, \infty)$: $A_0^+(t) = \int_0^t I_{[1, \infty)}(B_s) ds$. Then

$$k^*(k_*) = \limsup_{t \rightarrow \infty} (\inf) A_0^+(t) / A^-(t).$$

Proof. This follows immediately from

$$\lim_{t \rightarrow \infty} \int_0^t I_{(0, 1)}(B_s) ds / C(t) = 0 \quad \text{a.s.}$$

where C stands for any one of the functionals A^+, A^-, A_0^+ . See [5], p. 228.

Step 2. Define stopping times T_0, T_1, T_2, \dots , by $T_0 = 0, T_1 = D_1$, and

$$T_j = \begin{cases} \min \{t \geq T_{j-1}; B(t) = 0\} & \text{for } j \text{ even,} \\ \min \{t \geq T_{j-1}; B(t) = 1\} & \text{for } j \text{ odd,} \end{cases}$$

and for $n \geq 1$, let

$$(2.1) \quad \begin{aligned} W_n &= \int_{-\infty}^{0+} [L(T_{2n-1}, x) - L(T_{2n-2}, x)] m\{dx\} = A^-(T_{2n-1}) - A^-(T_{2n-2}), \\ V_n &= \int_1^\infty [L(T_{2n}, x) - L(T_{2n-1}, x)] dx = A_0^+(T_{2n}) - A_0^+(T_{2n-1}). \end{aligned}$$

Because A_0^+, A^- , increases only on intervals $[T_j, T_{j+1}]$ with j odd, j even, respectively, and is otherwise constant, we see immediately that for $n = 1, 2, \dots$,

$$(2.2) \quad \frac{V_1 + \dots + V_{n-1}}{W_1 + \dots + W_n} \leq \frac{A_0^+(t)}{A^-(t)} \leq \frac{V_1 + \dots + V_r}{W_1 + \dots + W_r}, \quad T_{2n-2} \leq t \leq T_{2n},$$

where $r = n - 1$ when $T_{2n-2} \leq t < T_{2n-1}$ and $r = n$ when $T_{2n-1} \leq t \leq T_{2n}$. Continuity of the sample paths of B and the strong Markov property imply

$$(2.3) \quad \begin{aligned} &\{V_n\} \text{ and } \{W_n\} \text{ are each sequences of independent,} \\ &\text{identically distributed, positive random variables} \\ &\text{and the two sequences are independent of each other.} \end{aligned}$$

In particular, the two sequences of partial sums $\{V_1 + \dots + V_{n-1}, n \geq 2\}$ and $\{V_2 + \dots + V_n, n \geq 2\}$ are identical in law and are independent of the sequence $\{W_n\}$. (From the point of view of the V 's the sequence $\{W_1 + \dots + W_n, n \geq 1\}$ may be regarded simply as a sequence of constants $\uparrow \infty$.) From these facts, (2.2), and Step 1, we conclude

$$(2.4) \quad k^*(k_*) = \limsup(\inf) \frac{V_1 + \dots + V_n}{W_1 + \dots + W_n} \quad \text{a.s.}$$

Now, by the method of problem 1, p. 230, in [5], it follows that $E[L(T_{2n-1}, x) - L(T_{2n-2}, x)] = EL(D_1, x) = 2$, for $x \leq 0$, and $E[L(T_{2n}, x) - L(T_{2n-1}, x)] = E[L(T_2, x) - L(D_1, x)] = 2$, for $x \geq 1$. Hence, see (2.1),

$$(2.5) \quad EV_1 = \infty \quad \text{and} \quad EW_1 = \infty.$$

Step 3. For positive random variables V and W define

$$J(V, W) = \int_1^\infty \left(x \int_0^x P(W > t) dt \right) P\{V \in dx\}.$$

Lemma A. *Let $\{V_n\}$ and $\{W_n\}$ be any two sequences of r.v.'s on the same probability space which satisfy (2.3) and (2.5). Then $J(V_1, W_1) + J(W_1, V_1) = \infty$ and the following implications hold*

$$(i) \quad J(V_1, W_1) = \infty \Rightarrow \limsup \frac{V_n}{W_1 + \dots + W_n} = \infty \quad \text{a.s.}$$

$$(ii) \quad J(V_1, W_1) < \infty \Rightarrow \limsup \frac{V_1 + \dots + V_n}{W_1 + \dots + W_n} = 0 \quad \text{a.s.}$$

Proof. That at most one of $J(V_1, W_1)$, $J(W_1, V_1)$ is finite can be proved with slight modification as in [2], pp. 375-6. From the estimates in [2], pp. 377-8 for any fixed $\varepsilon > 0$

$$(2.6) \quad \sum_{n=1}^\infty P(W_1 + \dots + W_n \leq \varepsilon x) \asymp x \int_0^x P(W_1 > t) dt, \quad x \rightarrow \infty,$$

(\asymp means the ratio of both sides is bounded away from 0 and ∞). It follows that $J(V_1, W_1) = \infty$ implies $\sum P(W_1 + \dots + W_n \leq \varepsilon V_n) = \infty$ and then, by Lemma 2 in [6], p. 1192, that for every $\varepsilon > 0$ $P(W_1 + \dots + W_n \leq \varepsilon V_n \text{ i.o.}) = 1$. Now suppose $J(V_1, W_1) < \infty$. Then by (2.6) and the Borel-Cantelli Lemma we have (*) $P(W_1 + \dots + W_n \leq V_n \text{ i.o.}) = 0$. Suppose, contrary to the conclusion of (ii), $P(V_1 + \dots + V_n \geq \varepsilon(W_1 + \dots + W_n) \text{ i.o.}) > 0$ for some $\varepsilon > 0$. This probability must be 1 and, as in [6], p. 1191, we get $P(V_n \geq \varepsilon \min_{j \leq n} (W_1 + \dots + W_j) \text{ i.o.}) = 1$ and then by Lemmas 3 and 4 of [6] we get $P(V_n \geq c(W_1 + \dots + W_n) \text{ i.o.}) = 1$ for any $c > 0$ which contradicts (*).

Step 4. It is now clear from Lemma A and (2.4) that $k^* = 0$ or ∞ according as $J(V_1, W_1)$ is finite or infinite, and, by interchanging the V 's and W 's in Lemma A, we also get $k^* = \infty$ or 0 according as $J(W_1, V_1)$ is finite or infinite. To complete

the proof we need to show that the integrals at (1.4i) and (1.4ii) are equivalent to $J(V_1, W_1)$ and $J(W_1, V_1)$, respectively, and to do this it suffices to show that

$$(2.7) \quad P\{V_1 \in dx\} \asymp x^{-3/2} dx \quad \text{as } x \rightarrow \infty.$$

This must be well known but we lack a ready reference so here is a quick proof. Let E_x denote expectation for paths starting at x and w_s^+ the shifted path: $B_t(w_s^+) = B_{t+s}(w)$. Then for any path w , $V_1 = A_0^+(s + D_0(w_s^+); w) - A_0^+(s; w) = A_0^+(D_0(w_s^+); w_s^+)$, $s = D_1(w)$, as everyone knows, hence, for any $\lambda \geq 0$,

$$E_0 \exp(-\lambda V_1) = E_1 \exp(-\lambda A_0^+(D_0)).$$

If $A_\varepsilon^+(t) = \int_{-\infty}^{\infty} L(t, x) \delta_\varepsilon(x) dx$ where $\delta_\varepsilon(x) = 1$ for $x \geq 1$, $\delta_\varepsilon(x) = \varepsilon$ for $x < 1$, then $A_\varepsilon^+(D_0)$ is the first passage time to 0 for the diffusion process on natural scale whose speed measure is $2\delta_\varepsilon(x) dx$. It follows that $g(x) = E_x \exp(-\lambda A_\varepsilon^+(D_0))$ satisfies $g(0) = 1$, g is bounded, g and g' are continuous, and $g''(x)/2 = \lambda \delta_\varepsilon(x) g(x)$, $x \neq 1$. Solving for g and setting $x = 1$, we obtain

$$(2.8) \quad E \exp(-\lambda V_1) = \lim_{\varepsilon \rightarrow 0^+} g(1) = \frac{1}{1 + \sqrt{2\lambda}}, \quad \lambda \geq 0.$$

The function $x \mapsto 2 \int_0^\infty e^{-t} P\{B(x) > t\} dt$ is a distribution function which also has the Laplace transform $(1 + \sqrt{2\lambda})^{-1}$. (To see this note that it is the distribution function of $T(\eta)$ where $\{T(t), t \geq 0\}$ is a stable process of index 1/2 and rate $\sqrt{2}$ and η is an independent $\text{Exp}(1)$ -distributed random variable.) It follows that

$$\begin{aligned} \frac{d}{dx} P(V_1 \leq x) &= \frac{d}{dx} 2 \int_0^\infty e^{-t} P\{B(x) > t\} dt \\ &= x^{-3/2} \int_0^\infty t e^{-t} e^{-t^2/2x} dt / \sqrt{2\pi}. \end{aligned}$$

This easily yields (2.7) and concludes the proof of Theorem 1.

3. Proof of Theorem 2

By Ray's Theorem, see Williams [9], p. 873, or Ray [8], for any fixed $b < 0$

$$\begin{aligned} &Law[\{L(D_1, x) : b \leq x \leq 0\} | \min_{s \leq D_1} B_s = b, B_0 = 0] \\ &= Law \left[\left\{ (1-x)^2 Z \left(\frac{1}{1-x} - \frac{1}{1-b} \right)^2 ; b \leq x \leq 0 \right\} \right], \end{aligned}$$

where Z is the radial part of a 4-dimensional Brownian motion starting at 0. (Note that our local time is twice the Ito-McKean local time.) Also

$$P[\min_{s \leq D_1} B_s \leq b] = \frac{1}{1-b}, \quad b \leq 0.$$

Hence, making the change of variable $x \rightarrow -x, b \rightarrow -b,$

$$P(W > t) = \int_0^\infty P \left[\int_0^{b^+} (1+x)^2 Z \left(\frac{1}{1+x} - \frac{1}{1+b} \right)^2 m_+ \{dx\} > t \right] \frac{db}{(1+b)^2}.$$

Lemma B. *With α as at (1.6), there is a constant $c_1, 0 < c_1 < \infty,$ such that*

$$(3.2) \quad P(W > t) \leq c_1/\alpha(t) \quad \text{for } t \geq 0.$$

Proof. Let B^4 be a 4-dimensional Brownian motion starting at the origin and put $H(s) = B^4(s) - sB^4(1),$ then the process

$$(3.3) \quad x \mapsto \sqrt{1+b} H \left(\frac{1+x}{1+b} \right), \quad -1 \leq x \leq b,$$

and the process

$$(3.4) \quad x \mapsto (1+x) B^4 \left(\frac{1}{1+x} - \frac{1}{1+b} \right), \quad -1 \leq x \leq b,$$

are law equivalent. To see this note that they are both 0 mean Gaussian processes in \mathbb{R}^4 with covariance $EU_i(x_1) U_j(x_2) = (1+x_1)(b-x_2)(1+b)^{-1} \delta_{ij},$ $x_1 < x_2,$ where U is either of the processes (3.3) or (3.4), δ_{ij} is the Kronecker delta, $i, j = 1, \dots, 4.$ From this equivalence we get

$$(3.5) \quad P(W > t) \leq \int_0^\infty P \left[N > \frac{t}{(1+b)m_+(b)} \right] \frac{db}{(1+b)^2},$$

where

$$\begin{aligned} N &= \max \left\{ \left\| H \left(\frac{1+x}{1+b} \right) \right\|^2; -1 \leq x \leq b \right\} \\ &= \max \{ \|H(s)\|^2; 0 \leq s \leq 1 \}. \end{aligned}$$

According to Fernique [4], for some $\varepsilon > 0$ we have $E \exp(\varepsilon N) = c_2 < \infty.$ Hence

$$(3.6) \quad P[N > u] \leq c_2 e^{-\varepsilon u}, \quad u \geq 0.$$

Put $c_3 = \varepsilon/(1 + 2m_+(1)).$ Then

$$P \left[N \geq \frac{t}{(1+b)m_+(b)} \right] \leq \begin{cases} c_2 e^{-c_3 t}, & 0 \leq b \leq 1, \\ c_2 e^{-\varepsilon t/2bm_+(b)}, & b \geq 1. \end{cases}$$

If $\alpha(t) > 2$ then for $1 \leq b \leq \alpha(t)/2,$

$$\begin{aligned} \exp(-\varepsilon/2bm_+(b)) &\leq \exp(-\varepsilon t/2bm_+(\alpha(t)/2)) \\ &\leq \exp(-\varepsilon \alpha(t)/4b). \end{aligned}$$

Going back to (3.5) with these bounds we get

$$(3.8) \quad \begin{aligned} P(W > t) &\leq \int_0^1 + \int_1^{\alpha/2} + \int_{\alpha/2}^\infty (\bullet) \\ &= 0(e^{-c_3 t}) + 0 \left(\int_1^{\alpha/2} e^{-\varepsilon \alpha(t)/4b} \frac{db}{b^2} \right) + 0 \left(\frac{1}{\alpha(t)} \right) \\ &= 0(e^{-c_3 t}) + 0 \left(\frac{1}{\alpha(t)} \right), \quad t \rightarrow \infty. \end{aligned}$$

From (1.8) we have $0(e^{-c_0 t})=0(1/t)=0(1/\alpha(t))$, as $t \rightarrow \infty$. With this (3.2) is established.

Lemma C. *With β as at (1.7), there is a constant $c_0 > 0$ such that*

$$(3.9) \quad P(W > t) \geq c_0/\beta(t), \quad \text{for } t \geq 0.$$

Proof. Using Brownian scaling we have for $b \geq 1$

$$\begin{aligned} & \min \left\{ Z \left(\frac{1}{1+x} - \frac{1}{1+b} \right)^2, 0 \leq x \leq b/2 \right\} \\ & =_d \min \left\{ \frac{1}{b} Z(br)^2; \frac{2}{2+b} - \frac{1}{1+b} \leq r \leq \frac{b}{1+b} \right\} \\ & \geq \inf \{ Z(s)^2; \frac{1}{6} \leq s < \infty \} / b = M/b \end{aligned}$$

where $=_d$ means equality in distribution. Since $Z(0)=0$ and since $\lim_{u \rightarrow \infty} Z(u) = \infty$ a.s. (four-dimensional Brownian motion is transient), it is clear that $P(M > \alpha) > 0$ for every $\alpha > 0$. Using these facts in (3.1) we obtain

$$\begin{aligned} P(W > t) & \geq \frac{1}{4} \int_1^\infty P \left[\int_0^{b/2} x^2 Z \left(\frac{1}{1+x} - \frac{1}{1+b} \right)^2 m_+ \{ dx \} > t \right] \frac{db}{b^2} \\ & \geq \frac{1}{4} \int_1^\infty P \left[M > \frac{bt}{\sigma(b/2)} \right] \frac{db}{b^2}, \end{aligned}$$

where $\sigma(x) = \int_0^x y^2 m_+ \{ dy \}$. With β as at (1.7) we have $(b/2)/\sigma(b/2) \leq 1/t$ whenever $b/2 \geq \beta(t) > 1$, hence

$$P \left[M > \frac{bt}{\sigma(b/2)} \right] \geq P[M > 2] > 0,$$

for $b \geq 2\beta(t)$. Consequently

$$P(W > t) > \frac{1}{4} \int_{2\beta(t)}^\infty (\bullet) \geq c_0/\beta(t),$$

for all $t \geq 0$ for some constant c_0 , $0 < c_0 \leq \frac{1}{16} P[M > 2]$.

The reader may easily complete the proof of Theorem 2 by substituting the bounds at (3.2) and (3.9) into the integrals which occur in Theorem 1 and reading off the implications.

Remark 1. It might be thought that one could get $c/\alpha(t)$ as a lower bound for $P(W > t)$ by staying with the process (3.3). Unfortunately $H(1/(1+b)) \rightarrow 0$ as $b \rightarrow \infty$ and $H(1)=0$, so there is trouble at both endpoints of $(-1, b)$ and this trouble neatly foils the attempt.

Remark 2. The assumption $m_+(\infty) = \infty$ is not necessary to get (3.9). It suffices to require only that $(1/b) \int_0^{b+} x^2 m_+ \{ dx \} \rightarrow \infty$ as $b \rightarrow \infty$.

Remark 3. Here is an example. Suppose

$$m\{dx\} = (|x| \log |x|)^{-1} dx, \quad x \leq -e.$$

Then $m_+(x) = \log \log x$, $\int_0^x y^2 m\{dy\} = \frac{1}{2}(x^2/\log x)(1 + o(1))$ for $x \rightarrow \infty$. So $\alpha(t) = (t/\log \log t)(1 + o(1))$, $\beta(t) = (2t \log t)(1 + o(1))$, and then

$$C_0 \frac{1}{t \log t} \leq P(W > t) \leq C_1 \frac{\log \log t}{t}, \quad t \geq 3.$$

It is not clear which bound is the best asymptotically, though one might suspect it is the lower one. See the last remarks in § 5(a).

4. Proof of Theorem 3

By (1.9) we can choose ε , $0 < \varepsilon < 1$ and $x_0 > 1$ and $x_0 > 1$ so that for $x \geq x_0$, $m_+(x)/m_+(2x) \leq 1 - \varepsilon$. Then for $b \geq 2x_0$

$$\begin{aligned} \varepsilon \frac{b^2}{4} m_+(b) &\leq \frac{b^2}{4} \left[m_+(b) - m_+\left(\frac{b}{2}\right) \right] \\ &\leq \int_0^{b+} x^2 m_+\{dx\} \leq b^2 m_+(b). \end{aligned}$$

From these inequalities and (1.6)-(1.7), it follows that $\alpha(t) \leq \beta(t) \leq \alpha(4t/\varepsilon)$ for all t sufficiently large. A simple scale change of variables now shows that the integrals $J^*(1/\beta)$ and $J_*(1/\beta)$ of Theorem 2 are equivalent to $J^*(1/\alpha)$ and $J_*(1/\alpha)$ respectively. (Note that the same argument shows that (1.9) can be replaced by $\limsup(m_+(x)/m_+(cx)) < 1$ for some $c > 1$.)

Remark. One should note that (1.9) does not imply regular variation. For example, the measure $m\{dx\} = \exp(|x|) dx$, $x \leq 0$, satisfies (1.9) but not (5.7) in the next section.

5. Miscellaneous Comments

(a) A more direct method is available for getting at the distribution of W in (1.3) which yields exact asymptotic estimates in special cases. Consider the process

$$(5.1) \quad \left\{ y_t \rightarrow A^-(D_y) = \int_{-\infty}^{0+} L(D_y; x) m\{dx\}, y \geq 0; P_0 \right\}.$$

The strong Markov property of B and the fact that $B(D_y) = y$ shows that (5.1) is a process with independent (but generally not stationary) increments. For $\lambda > 0$ write

$$(5.2) \quad u(y) = E_0 e^{-\lambda A^-(D_y)}.$$

Then for $h > 0$,

$$\begin{aligned} u(y+h) &= u(y) E_y e^{-\lambda A^-(D_{y+h})} \\ &= u(y) [E_y(e^{-\lambda A^-(D_{y+h})}; D_0 > D_{y+h}) + E_y(e^{-\lambda A^-(D_{y+h})}; D_0 \leq D_{y+h})] \\ &= u(y) \left[\frac{y}{y+h} + \frac{h}{y+h} u(y+h) \right], \end{aligned}$$

which, keeping in mind that (5.1) is continuous in probability, gives

$$(5.3) \quad u^+(y) = \lim_{h \rightarrow 0^+} (u(y+h) - u(y))/h = \frac{1}{y} u(y) (u(y) - 1),$$

for $y > 0$ and $u(0+) = u(0) = 1$. Solving (5.3) we arrive at the formula

$$(5.4) \quad E e^{-\lambda A^-(D_y)} = E \exp \left(-\lambda \int_{-\infty}^{0^+} L(D_y, x) m\{dx\} \right) = (j(\lambda; m) y + 1)^{-1}$$

where j does not depend on y and $j(0+; m) = 0$. The Brownian scaling property states that the path transformation $B(\bullet) \mapsto s B(\bullet/s^2)$ is P_0 -measure preserving for fixed $s > 0$. This and the definition of local time show that for any fixed $s > 0$,

$A^-(D_y)$ has the same distribution, under P_0 , as $s \int_{-\infty}^{0^+} L(D_{y/s}, x/s) m\{dx\}$. Consequently

$$(5.5) \quad E e^{-\lambda A^-(D_y)} = E \exp(-\lambda s m_+(s) A_s^-(D_{y/s}))$$

where $A_s^-(D_z) = \int_{-\infty}^{0^+} L(D_z, x) m_s\{dx\}$ and $m_s\{dx\} = m\{s dx\}/m_+(s)$. Combining (5.4) and (5.5) gives

$$(5.6) \quad j(\lambda; m) = j(\lambda s m_+(s); m_s)/s.$$

Now let us assume that m is regularly varying in the sense that for every $x \geq 0$, the

$$(5.7) \quad \lim_{s \rightarrow \infty} m_+(s x)/m_+(s) = \mu(x)$$

exists and is finite. Then, necessarily, $\mu(x) = x^q$ for some $q \geq 0$ (q is the ‘‘exponent’’) and, supposing $q > 0$,

$$\int_{-\infty}^{0^+} f(x) m_s\{dx\} \rightarrow q \int_{-\infty}^0 f(x) |x|^{q-1} dx, \quad s \rightarrow \infty,$$

for every continuous, compact f . Hence, since $x \mapsto L(D_y, x)$ has compact support a.s., $A_s^-(D_y) \rightarrow A_\mu^-(D_y)$ a.s. as $s \rightarrow \infty$ and then

$$\lim_{s \rightarrow \infty} j(\lambda; m_s) = j(\lambda; \mu)$$

uniformly on bounded intervals of $\lambda \geq 0$. Properties of regular variation, see, for example, Bojanic and Seneta (1971), imply that the function α defined at (1.6) is

also regularly varying (with exponent $(q + 1)^{-1}$) and that

$$\lambda s m_+(s) \rightarrow 1 \quad \text{as } \lambda \rightarrow 0+, \quad s = \alpha(1/\lambda).$$

Setting $s = \alpha(1/\lambda)$ in (5.6) we get

$$\begin{aligned} \int_0^\infty e^{-\lambda t} P(W > t) dt &= \frac{1 - E e^{-\lambda W}}{\lambda} = \frac{j(\lambda; m)/\lambda}{j(\lambda; m) + 1} \\ &= \frac{j(1 + o(1), m_s)}{o(1) + 1} \left(\frac{1}{\lambda \alpha(1/\lambda)} \right) \\ &= c(q) (\lambda \alpha(1/\lambda))^{-1} (1 + o(1)), \end{aligned}$$

as $\lambda \rightarrow 0+$, where $c(q) = E_0 \exp \left(-q \int_{-\infty}^0 L(D_1, x) |x|^{q-1} dx \right)$. Applying a Tauberian theorem, see Feller (1971), p. 446, now gives

$$(5.10) \quad P(W > t) = c(q) \Gamma \left(\frac{q}{q+1} \right)^{-1} \left(\frac{1}{\alpha(t)} \right) (1 + o(1))$$

as $t \rightarrow \infty$. If $q = 0$ about the best one can do with this argument is to show that $P(W > t) = o(1/\alpha(t))$. We omit the details. It should be pointed out, however, that in the case $q = 0$, even though (5.10) does not apply and even though Theorem 3 does not apply, it can still be shown, via Theorem 2 and properties of regularly varying functions, that $k^* = k_* = \infty$. This leads one to suspect that (1.9) in Theorem 3 is not necessary.

(b) If A^+ , the occupation time functional of $(0, \infty)$, is replaced by a more general functional say

$$A^+(t) = \int_0^\infty L(t, x) m^+ \{dx\}$$

(m^+ not to be confused with m_+ of §1), then the methods of this paper will yield results on the boundedness or unboundedness of $A^+(t)/A^-(t)$ (A^- as before). Unfortunately the analogue of Theorem 2 is rather unwieldy. Under a regularity condition such as (1.9), however, matters improve. Here is a sample. Write m^- for m . We assume $m^\pm(I) < \infty$ for bounded I and that $m^+(0, \infty) = m^-(-\infty, 0] = \infty$. Suppose also that for some $0 < \varepsilon < 1$

$$\begin{aligned} m^+(0, x]/m^+(0, 2x] &\leq 1 - \varepsilon \\ m^-[-x, 0]/m^-[-2x, 0] &\leq 1 - \varepsilon \end{aligned}$$

for all x sufficiently large. Let α^\pm be defined by $\alpha^- = \alpha$ at (1.6) and $\alpha^+(t) = \inf \{x: x m^+(0, x] \geq t\}$. Put $s(x) = \int_0^x \alpha^-(t)^{-1} dt$.

Theorem 4. *Under these assumptions $\limsup_{t \rightarrow \infty} A^+(t)/A^-(t) = \infty$ a.s. if and only if $\int [s(x) - x/\alpha^-(x)] s(x)^{-2} \alpha^+(x)^{-1} dx = \infty$.*

We omit the proof. As noted before our motivation for studying the particular case m^+ = Lebesgue measure on $(0, \infty)$ derives from the London et al. (1982) paper, but it also seems natural to compare an arbitrary additive functional with occupation time.

(c) A more interesting problem than the one discussed in (b) concerns the a.s. boundedness of $A_1(t)/A_2(t)$, $t \rightarrow \infty$, where A_i is now an arbitrary increasing continuous additive functional: $A_i(t) = \int_{-\infty}^{\infty} L(t, x) m_i\{dx\}$. I do not have a good conjecture but, as before, the ratio ergodic theorem does take care of the case in which at least one of m_1 or m_2 is a finite measure.

References

1. Bojanic, R., Seneta, E.: Slowly varying functions and asymptotic relations. *J. Math. Anal. Appl.* **34**, 302-315 (1971)
2. Erickson, K.B.: The strong law of large numbers when the mean is undefined. *Trans. Am. Math. Soc.* **185**, 371-381 (1973)
3. Feller, W.: An introduction to probability theory and its applications II. New York: Wiley 1971
4. Fernique, X.M.: Intégrabilité des vecteurs Gaussien. *C.R. Acad. Sci. Paris Ser. A*, **270**, 1698-1699 (1970)
5. Ito, K., McKean, H.P., Jr.: Diffusion processes and their sample paths. Berlin-Heidelberg-New York: Springer 1974
6. Kesten, H.: The limit points of a normalized random walk. *Ann. Math. Stat.* **41**, 1173-1205 (1970)
7. London, R.R., McKean, H.P., Rogers, L.C.G., Williams, D.: A martingale approach to some Wiener-Hopf problems, I. *Séminaire de Probabilités XVI. Lecture Notes in Math.* **920**, 41-67. Berlin-Heidelberg-New York: Springer 1982
8. Ray, D.: Sojourn times of diffusion process. III. *J. Math.* **7**, 615-630 (1963)
9. Williams, D.: Decomposing the Brownian path. *Bull. Am. Math. Soc.* **76**, 871-873 (1970)

Received July 15, 1984