

## Non-Classical Law of the Iterated Logarithm Behaviour for Trimmed Sums

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**Summary.** We study the law of the iterated logarithm for the partial sum of i.i.d. random variables when the  $r_n$  largest summands are excluded, where  $r_n = o(\log \log n)$ . This complements earlier work in which the case  $\log \log n = O(r_n)$  was considered. A law of the iterated logarithm is again seen to prevail for a wide class of distributions, but for reasons quite different from previously.

### 1. Introduction

Let  $X, X_1, X_2, \dots$  be a sequence of independent identically distributed random variables with common distribution function  $F$ . For  $x > 0$  define

$$G(x) = P(|X| > x), \quad K(x) = x^{-2} \int_{|y| \leq x} y^2 F(dy)$$

$$Q(x) = G(x) + K(x).$$

If we need to distinguish  $X$  from another random variable we will write  $F_X, G_X, K_X$  and  $Q_X$ .

Let  $(^{(1)}X_n, \dots, ^{(n)}X_n)$  be an arrangement of  $X_1, \dots, X_n$  in decreasing order of magnitude, i.e.  $|^{(1)}X_n| \geq \dots \geq |^{(n)}X_n|$ . We will assume throughout that the distribution function of  $X$  is continuous one effect of which is to make the ordering  $(^{(1)}X_n, \dots, ^{(n)}X_n)$  unique except on a null set. This assumption could be dispensed with but the ensuing technical details would only serve to obscure the main ideas. For  $r \geq 0$  an integer, define  $(^{(r)}S_n = ^{(r+1)}X_n + \dots + ^{(n)}X_n$ . We write  $S_n$  for  $(^{(0)}S_n$ . We will refer to  $(^{(r)}S_n$  as a trimmed sum.

The study of trimmed sums is motivated on the one hand by statistical considerations, (although it is perhaps more natural to consider trimming by the order statistics in this context) while on the other hand, probabilistically, by a desire to better understand partial sums of i.i.d. random variables and in particular to understand the role played by the summands of large modulus.

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This in turn leads to a deeper understanding of the classical limit theorems and puts them more sharply into perspective.

The present paper grew out of an attempt to answer some unresolved questions which arose in [3]. One of the main results in [3], Theorem 5.5, states that if the distribution of  $X$  satisfies

$$(1.1) \quad \limsup_{x \rightarrow \infty} G(x)/K(x) < \infty$$

and  $r_n$  is an increasing sequence of integers satisfying

$$(1.2) \quad \liminf_{n \rightarrow \infty} r_n / \log \log n > 0$$

$$(1.3) \quad \limsup_{n \rightarrow \infty} r_n n^{-1} < G(0),$$

then

$$(1.4) \quad 0 < \limsup_{n \rightarrow \infty} \frac{|{}^{(r_n)}S_n - nEX 1(|X| \leq b_n)|}{(n \log \log n b_n^2 K(b_n))^{1/2}} < \infty$$

where  $G(b_n) = r_n n^{-1}$ . Condition (1.1), first introduced by Feller [2], is equivalent to stochastic compactness of  $S_n$  and is discussed in detail in [4] where further references can be found. In particular (1.1) holds whenever  $X$  is in the domain of attraction of a stable law of index  $\alpha \in (0, 2]$ . The normalizer in (1.4) is the natural one to use for the Law of the Iterated Logarithm (L.I.L.) in that Pruitt [13] has shown that if  $X$  is symmetric, (1.1) holds,  $r_n \uparrow \infty$  and  $r_n n^{-1} \rightarrow 0$  then

$$(1.5) \quad \frac{{}^{(r_n)}S_n}{(n b_n^2 K(b_n))^{1/2}} \rightarrow N(0, 1)$$

where  $N(0, 1)$  is normal with mean 0 and variance 1. It is interesting to note that no symmetry assumption is needed for (1.4) to hold, but (1.5) may fail without it.

Results similar to (1.4), for other variants of the trimmed sum, have been discovered recently by several authors, see [5] and [6] for example. In each of these works it is also assumed that  $r_n$  satisfies (1.2). In light of (1.5) one might expect that (1.4) holds without this assumption. We will show that this is not the case although an L.I.L. result for  ${}^{(r_n)}S_n$  is still available but for entirely different reasons. In (1.4) the large values of  ${}^{(r_n)}S_n$  arise due to the cumulative effect of many summands as in the classical LIL, however when  $r_n = o(\log \log n)$  the large values of  ${}^{(r_n)}S_n$  are determined by a small number of large terms. For example, we will show that if  $r_n$  is an increasing sequence of integers tending to  $\infty$  such that

$$(1.6) \quad r_n (\log \log n)^{-1/2} \rightarrow 0$$

and if in addition to (1.1) the distribution of  $X$  satisfies

$$(1.7) \quad \liminf_{x \rightarrow \infty} G(x)/K(x) > 0$$

then the large values of  $(r_n^{-1})S_n$ , after centering, are due entirely to  $(r_n)X_n$  and further, that  $(r_n)X_n$  can be normalized to obtain a finite non-zero lim sup. That is, there exist  $\alpha_n, \delta_n$  such that

$$(1.8) \quad 0 < \limsup_{n \rightarrow \infty} |(r_n)X_n| \alpha_n^{-1} = \limsup_{n \rightarrow \infty} |(r_n^{-1})S_n - \delta_n| \alpha_n^{-1} < \infty$$

$$(1.9) \quad \limsup_{n \rightarrow \infty} |(r_n^{-1})X_n| \alpha_n^{-1} = \limsup_{n \rightarrow \infty} |(r_n)S_n - \delta_n| \alpha_n^{-1} = 0.$$

If instead of (1.6) we assume only that

$$(1.10) \quad r_n (\log \log n)^{-1} \rightarrow 0$$

then one can still find  $\alpha_n$  such that

$$(1.11) \quad 0 < \limsup_{n \rightarrow \infty} |(r_n)X_n| \alpha_n^{-1} < \infty$$

but now there may be other summands which are also comparable in size to  $\alpha_n$ . Nevertheless by controlling these terms we will show that under (1.1) and (1.7), there exists  $\gamma_n$  and  $\delta_n$  such that

$$(1.12) \quad 0 < \limsup_{n \rightarrow \infty} |(r_n^{-1})S_n - \delta_n| \gamma_n^{-1} < \infty.$$

The normalizer  $\gamma_n$  is given by  $N_n \alpha_n$  where  $N_n = [r_n^2 / \log \log n] + 1$  ( $[x]$  denotes the integer part of  $x$ ). The way in which this arises is that roughly speaking, the large values of  $(r_n^{-1})S_n$  occur because infinitely often there are  $N_n$  terms comparable in size to  $\alpha_n$  and these terms all have the same sign. This is quite different from the way the large values arise in the classical LIL, see Sect. 5 for a further discussion.

Condition (1.7) is equivalent, by a famous result of Lévy, to  $X$  not being in the domain of partial attraction of the normal law. Thus the class of distributions satisfying (1.1) and (1.7) is still quite large and includes all of those in the domain of attraction of a stable law of index  $\alpha \in (0, 2)$ . We should perhaps point out here that Maller [10], extending earlier work of Kesten [7] in the case  $r=0$ , has shown that the failure of (1.7) is necessary and sufficient for the existence of an increasing sequence  $\gamma_n$  such that (1.12) holds with  $\delta_n = \text{median}(S_n)$  and  $r_n$  a bounded sequence.

To illustrate the difference between the normalizers in (1.4) and (1.12), assume that  $X$  is symmetric stable of index  $\alpha \in (0, 2)$  and the scale parameter is chosen so that  $G(x) \sim x^{-\alpha}$ . Then the normalizer in (1.5) is given by  $n^{1/\alpha} (\alpha(2-\alpha)^{-1} r_n^{1-2/\alpha})^{1/2}$  and so in (1.4) it is  $n^{1/\alpha} (\alpha(2-\alpha)^{-1} r_n^{1-2/\alpha} \log \log n)^{1/2}$ . In (1.12) if we take for example  $r_n = [l_p n]$  for  $p \geq 3$ , where  $l_p n$  is the  $p^{\text{th}}$  iterate of the logarithm function, then

$$\limsup_{n \rightarrow \infty} |(r_n)X_n| \alpha_n^{-1} = \limsup_{n \rightarrow \infty} |(r_n^{-1})S_n| \alpha_n^{-1} = e^{2/\alpha}$$

where  $\alpha_n = n^{1/\alpha} (l_p n)^{-1/\alpha} \exp((l_2 n + \dots + l_{p-1} n)(\alpha r_n)^{-1})$ .

If (1.7) fails then the non-classical behaviour given by (1.12) need not hold. For example let  $X$  have bounded support, then it is easy to see that  $(r_n)S_n$  satisfies (1.4) no matter how slowly  $r_n$  increases to infinity, indeed (1.4) holds for  $r_n$  constant. In fact it can be shown that for any random variable  $X$  in the domain of attraction of the normal law, there exists an increasing sequence  $r_n$ , which depends on  $X$ , such that  $r_n = o(\log \log n)$  and (1.4) holds, c.f. [8]. On the other hand one can also construct examples of  $X$  in the domain of attraction of the normal law for which (1.12) holds provided  $r_n$  increases sufficiently slowly, this rate again depending on the distribution. Thus for distributions satisfying (1.1) and (1.7) there is a single level, namely  $\log \log n$ , which distinguishes between classical and non-classical LIL behaviour, while for distributions attracted to the normal such a cut-off, if it exists, seems to depend on the distribution.

**2. Preliminaries**

Our basic assumption on the underlying distribution will be

$$(2.1) \quad 0 < \liminf_{x \rightarrow \infty} G(x)/K(x) \leq \limsup_{x \rightarrow \infty} G(x)/K(x) < \infty$$

Hence for some  $\theta > 1$  and all  $x > 0$

$$(2.2) \quad G(x) \leq Q(x) \leq \theta G(x).$$

By (2.1) and Lemma 2.4 of Pruitt [12], there exists  $q > 0$  and  $x_0 > 0$  such that for all  $x \geq x_0$

$$(2.3) \quad x^q Q(x) \text{ is decreasing.}$$

On the other hand by Lemma 2.1 of [12]  $x^2 Q(x)$  is always increasing, thus for any  $\xi \in (0, 1)$  if  $\xi x \geq x_0$  then

$$(2.4) \quad \xi^2 \theta^{-1} G(\xi x) \leq G(x) \leq \theta \xi^q G(\xi x).$$

We will assume that  $r_n$  is a sequence of integers such that

$$(2.5) \quad r_n \text{ increases to } \infty, \quad r_n(l_2 n)^{-1} \rightarrow 0.$$

In order to describe the normalizing sequences  $\alpha_n$  and  $\gamma_n$  we must first introduce an auxilliary sequence. Thus let  $a_n$  be any sequence of positive reals satisfying the following conditions:

$$(2.6) \quad a_n \text{ is decreasing}$$

$$(2.7) \quad (ln)^{-2} \leq a_n \leq (ln)^{-1}$$

$$(2.8) \quad \sum_n a_n n^{-1} e^{\varepsilon r_n} \begin{cases} < \infty & \varepsilon < 0 \\ = \infty & \varepsilon \geq 0. \end{cases}$$

In the case that  $r_n$  satisfies  $\liminf r_n(l_p n)^{-1} > 0$  for some  $p \geq 2$ , one can easily check that  $a_n = ((l_1 n)(l_2 n) \dots (l_{p-1} n))^{-1}$  satisfies (2.6)–(2.8). The proof that such an  $a_n$  exists in general is not difficult but will be deferred to the appendix.

It is a simple consequence of the monotonicity of  $a_n$  and  $r_n$  that if  $b > 1$  then

$$(2.9) \quad \sum_k a_{\lfloor b^k \rfloor} e^{\varepsilon r_{\lfloor b^k \rfloor}} \begin{cases} < \infty & \text{if } \varepsilon < 0 \\ = \infty & \text{if } \varepsilon \geq 0 \end{cases}$$

and furthermore, again by monotonicity, if  $n_k \geq b^k$ , then for every  $\varepsilon < 0$

$$(2.10) \quad \sum_k a_{n_k} e^{\varepsilon r_{n_k}} < \infty.$$

Of course the sequence  $a_n$  depends on  $r_n$  but note that if  $a_n$  satisfies (2.6)–(2.8) then it satisfies (2.6)–(2.8) with  $r_n$  replaced by the sequence  $r_n + j$  for each fixed  $j$ . Also observe that if  $w_n$  is any sequence such that  $|r_n - w_n| = o(r_n)$  then (2.9) and (2.10) hold with  $w_n$  replacing  $r_n$  provided we exclude the case  $\varepsilon = 0$  in (2.9).

Now let

$$(2.11) \quad \beta_n = \exp((l a_n^{-1} - r_n l r_n + r_n) r_n^{-1}).$$

Thus

$$(2.12) \quad a_n = \exp(r_n - r_n l r_n - r_n l \beta_n).$$

For later reference note that by (2.5) and (2.7) for any  $p \in \mathbb{R}$

$$(2.13) \quad r_n^{p+1} (l_2 n)^{-p} \beta_n \geq (r_n / l_2 n)^p \exp((l_2 n / r_n)) \rightarrow \infty.$$

Set

$$(2.14) \quad N_n = \lceil r_n^2 / l_2 n \rceil + 1$$

and for  $\lambda > 0$  define

$$(2.15) \quad \alpha_n(\lambda) = \min \{x : G(x) = (\lambda n \beta_n)^{-1}\}$$

and let

$$(2.16) \quad \gamma_n(\lambda) = N_n \alpha_n(\lambda).$$

We will write  $\alpha_n$  for  $\alpha_n(1)$  and  $\gamma_n$  for  $\gamma_n(1)$ . The sequence  $\alpha_n$  will be used to normalize  ${}^{(r_n)}X_n$ , while  $\gamma_n$  will be used to normalize  ${}^{(r_n^{-1})}S_n$ . Note that by (2.5) and (2.13)  $n\beta_n \rightarrow \infty$ , so  $\alpha_n(\lambda)$  and  $\gamma_n(\lambda)$  both tend to infinity for every  $\lambda > 0$ . Thus by (2.4) if  $\lambda_1 < \lambda_2$  and  $n$  is sufficiently large

$$(2.17) \quad (\lambda_1 / \theta \lambda_2)^{q-1} \alpha_n(\lambda_2) \leq \alpha_n(\lambda_1) \leq (\theta \lambda_1 \lambda_2^{-1})^{1/2} \alpha_n(\lambda_2).$$

In the special case that  $X$  is in the domain of attraction of a stable law of index  $\alpha \in (0, 2)$ , then  $G(x)/K(x) \rightarrow (2 - \alpha)\alpha^{-1}$ . Thus using Lemma 2.4 of [12] instead of (2.4) one can improve (2.17) in this case to

$$(2.18) \quad \alpha_n(\lambda_1) \alpha_n(\lambda_2)^{-1} \rightarrow (\lambda_1 \lambda_2^{-1})^{1/\alpha}.$$

In many of our Borel-Cantelli arguments we will be using the same subsequence to sum along, so we will now describe this subsequence and also some of its properties that will be needed.

Let  $a > 1$  and set  $n_1 = [(a - 1)^{-1}] + 1$  and

$$(2.19) \quad n_{k+1} = \min \{n: r_n > r_{n_k} \text{ or } a_n < a_{n_k}/2\} \wedge [an_k].$$

We first note that for some  $b \in (1, a)$

$$(2.20) \quad n_k \geq b^k \quad \text{for all } k.$$

This is because for each given  $k$ , there are  $[k/3]$  values of  $j$  for which one of the following hold:

$$r_{n_j} > r_{n_{j-1}}, \quad a_{n_j} < a_{n_{j-1}}/2, \quad n_j = [an_{j-1}].$$

In the first case, since  $r_n$  is integer valued

$$\begin{aligned} [k/3] &\leq r_{n_k} \\ &\leq l_2 n_k \end{aligned}$$

for large  $k$  by (2.5). In the second case

$$\begin{aligned} 2^{[k/3]} a_{n_1}^{-1} &\leq a_{n_k}^{-1} \\ &\leq (ln_k)^2 \end{aligned}$$

by (2.7), while in the final case by the definition of  $n_1$ , it is not hard to see that for some  $c \in (1, a)$ , independent of  $k$ ,

$$n_k \geq c^{[k/3]}.$$

Consequently (2.20) holds, and also by (2.19)

$$(2.21) \quad n_{k+1} \leq an_k.$$

Set  $m_k = n_{k+1} - 1$ . Note that we trivially have

$$(2.22) \quad r_n \text{ is constant on } [n_k, m_k],$$

and since  $a_n$  is decreasing we see that

$$(2.23) \quad \beta_n \text{ and } \alpha_n(\lambda) \text{ are increasing on } [n_k, m_k].$$

Furthermore for some constant  $c > 0$  independent of  $k$

$$(2.24) \quad \beta_{n_k} \geq c \beta_{m_k}.$$

As a consequence of this and (2.21) we have by (2.4) that for some constant  $c > 0$  independent of  $k$  and  $\lambda$

$$(2.25) \quad \alpha_{n_k}(\lambda) \geq c \alpha_{m_k}(\lambda)$$

$$(2.26) \quad \gamma_{n_k}(\lambda) \geq c \gamma_{m_k}(\lambda).$$

*Remark.* Throughout we will use the letter  $c$  to denote a positive constant whose value may change from one useage to the next.

### 3. Probability Estimates

For  $b > 0$  and  $d > 0$  define

$$(3.1) \quad U_n(d) = \sum_{i=1}^n X_i 1(|X_i| \leq d)$$

$$(3.2) \quad J_n(b) = \sum_{i=1}^n 1(|X_i| > b).$$

In order to prove our main results we will need probability estimates on the size of  $J_n(b)$  and  $U_n(d)$ . Since we will be working outside the range for which the classical exponential bounds were designed (see p. 266 of [9]) we will use the following estimate which is an immediate consequence of Lemma 3.1 in [12].

**Lemma 3.1.** *For any  $v > 0, d > 0, s > 0$  and all  $n$*

$$(3.3) \quad P(|U_n(d) - EU_n(d)| \geq 2^{-1} v e^v n d K(d) + s d v^{-1}) \leq 2e^{-s}.$$

Given two sequences  $s_n$  and  $t_n$  we will write  $s_n \approx t_n$  if  $s_n t_n^{-1}$  and  $s_n^{-1} t_n$  are both bounded as  $n \rightarrow \infty$ .

**Lemma 3.2.** *There exist positive constants  $c_1$  and  $c_2$  such that for all  $r \geq 1$ , all  $n$  and all  $b \geq 0$*

$$(3.4) \quad c_1 r^{-1/2} \exp(r - rlr + rl(nG(b)) - 2nG(b)) \\ \leq P(J_n(b) \geq r) \leq c_2 r^{-1/2} \exp(r - rlr + rl(nG(b)) - (n-r)G(b))$$

*provided*

$$(3.5) \quad n > r^2$$

$$(3.6) \quad nG(b) < r/2.$$

*Proof.* For any  $b \geq 0, r \geq 1$  and  $n \geq r$

$$(3.7) \quad P(J_n(b) \geq r) = \sum_{j=r}^n \binom{n}{j} G(b)^j (1 - G(b))^{n-j}.$$

Set  $u_j = \binom{n}{j} G(b)^j (1 - G(b))^{n-j}$ . Then for  $r \leq j \leq n$

$$\frac{u_{j+1}}{u_j} = \frac{(n-j) G(b)}{(j+1)(1-G(b))} \leq \frac{nG(b)}{r(1-G(b))} \leq \frac{1}{2(1-G(b))}$$

by (3.6). Further since  $1 \leq r < n^{1/2}$ , we have by (3.6) that  $G(b) \leq (2n^{1/2})^{-1} \leq 2^{-(3/2)}$  and so  $u_{j+1} u_j^{-1} \leq 2^{-1/4}$ . Hence

$$(3.8) \quad \binom{n}{r} G(b)^r (1 - G(b))^{n-r} \leq P(J_n(b) \geq r) \leq c \binom{n}{r} G(b)^r (1 - G(b))^{n-r}$$

where the  $c$  is independent of  $n, r$  and  $b$ . Next by Stirling's formula there exist positive constants  $c_3$  and  $c_4$  such that for all  $r \geq 1$  and all  $n > r^2$

$$(3.9) \quad c_3 r^{-1/2} \left(\frac{n}{n-r}\right)^{n-r} \binom{n}{r} \leq \binom{n}{r} \leq c_4 r^{-1/2} \left(\frac{n}{n-r}\right)^{n-r} \left(\frac{n}{r}\right)^r.$$

Now it is a straightforward exercise to check that for all  $r$  and  $n$  satisfying  $1 \leq r^2 < n$ ,

$$(3.10) \quad e^{r-1} \leq \left(\frac{n}{n-r}\right)^{n-r} \leq e^r.$$

Also the elementary inequalities  $e^{-2x} \leq 1 - x \leq e^{-x}$  for  $0 \leq x \leq 1/2$ , give

$$(3.11) \quad \exp(-2nG(b)) \leq (1 - G(b))^{n-r} \leq \exp(-(n-r)G(b)).$$

Thus (3.4) follows from (3.8)–(3.11).  $\square$

**Corollary 3.3.** *For any sequence of integers  $s_n$  satisfying  $1 \leq s_n^2 < n$  and any sequence of real numbers  $b_n > 0$ , if  $nG(b_n) s_n^{-1} \rightarrow 0$  then*

$$(3.12) \quad P(J_n(b_n) \geq s_n) \rightarrow 0.$$

This also follows trivially from Markov's inequality. The following result is an easy consequence of a generalized Borel-Cantelli Lemma.

**Lemma 3.4.** *Assume  $B_k, C_k$  are two sequences of events such that  $B_k, k=1, 2, \dots$  are independent and for each  $k, B_k$  and  $C_k$  are independent. If  $\sum P(B_k) = \infty$  and  $P(C_k) \rightarrow 1$  then  $P(B_k C_k \text{ i.o.}) = 1$ .*

*Proof.* Let  $E_k = B_k C_k$ . Then  $P(E_k) = P(B_k) P(C_k) \sim P(B_k)$  and so  $\sum P(E_k) = \infty$ . If  $i < j$  then

$$P(E_i E_j) \leq P(B_i B_j) = P(B_i) P(B_j) \sim P(E_i) P(E_j)$$

as  $i \rightarrow \infty$ . From these two facts, it easily follows that

$$\limsup_{n \rightarrow \infty} \left( \sum_{i=1}^n \sum_{j=1}^n P(E_i E_j) \right) \left( \sum_{i=1}^n P(E_i) \right)^{-2} \leq 1.$$



The result now follows by P3 on page 317 of [14].  $\square$

We conclude this section with a simple Lemma which will prove useful later.

**Lemma 3.5.** For any  $x \geq 0$ ,  $\varepsilon \geq 0$  and  $N \geq 0$

$$(3.13) \quad [(N + \varepsilon)[x]] + 1 - [N[x]] \geq \varepsilon(1 + \varepsilon)^{-1} x$$

*Proof.* If  $\varepsilon = 0$  the result is trivial, thus we may assume  $\varepsilon > 0$ . If  $x \leq (1 + \varepsilon)^{-1} \varepsilon^{-1}$  then  $\text{RHS} \leq 1$  while  $\text{LHS} \geq 1$  for all  $x$ . If  $x > (1 + \varepsilon)^{-1} \varepsilon^{-1}$  then  $x - 1 > (1 + \varepsilon)^{-1} x$  and so

$$\begin{aligned} \text{LHS} &\geq (N + \varepsilon)[x] - N[x] \\ &\geq \varepsilon(x - 1) \\ &\geq \varepsilon(1 + \varepsilon)^{-1} x. \quad \square \end{aligned}$$

#### 4. Main Results

We begin this section by describing the growth of  $(r_n)X_n$ . The only consequences of (2.1) that will be used in this paper are (2.4), (2.17), (2.25) and (2.26). Since these are not needed in the proof of the following result, no restrictions need be placed on the distribution of  $X$ .

**Theorem 4.1.** Assume that  $r_n$  satisfies (2.5), then

$$(4.1) \quad \limsup_{n \rightarrow \infty} \frac{|(r_n)X_n|}{\alpha_n(\lambda)} \begin{cases} \leq 1 & \text{if } \lambda > 1 \\ \geq 1 & \text{if } \lambda < 1. \end{cases}$$

*Proof.* Given  $\lambda > 1$ , choose  $a \in (1, \lambda)$  and let  $n_k$  be defined by (2.19). Set  $m_k = n_{k+1} - 1$  and observe that by (2.13) and (2.21),  $m_k G(\alpha_{n_k}(\lambda)) r_{n_k}^{-1} \rightarrow 0$ . Thus for large  $k$  by (2.21), (2.23) and (3.4)

$$\begin{aligned} (4.2) \quad &P(|(r_n)X_n| > \alpha_n(\lambda) \text{ for some } n_k \leq n \leq m_k) \\ &= P(J_n(\alpha_n(\lambda)) \geq r_n \text{ for some } n_k \leq n \leq m_k) \\ &\leq P(J_{m_k}(\alpha_{n_k}(\lambda)) \geq r_{n_k}) \\ &\leq c_2 r_{n_k}^{-1/2} \exp(r_{n_k} - r_{n_k} l r_{n_k} + r_{n_k} l(m_k G(\alpha_{n_k}(\lambda)))) \\ &= c_2 r_{n_k}^{-1/2} \exp(r_{n_k} - r_{n_k} l r_{n_k} + r_{n_k} l(n_k G(\alpha_{n_k}(\lambda))) + r_{n_k} l(m_k n_k^{-1})) \\ &\leq c_2 r_{n_k}^{-1/2} \exp(r_{n_k} - r_{n_k} l r_{n_k} - r_{n_k} l \beta_{n_k} + r_{n_k} (la - l\lambda)) \\ &= c_2 r_{n_k}^{-1/2} a_{n_k} \exp((la - l\lambda) r_{n_k}) \end{aligned}$$

and this gives rise to a convergent series by (2.10) and (2.20) since  $a < \lambda$ . The upper bound now follows by the Borel-Cantelli Lemma.

Now fix  $\lambda < 1$  and choose  $D$ , an integer, large enough that

$$(4.3) \quad 1 - D^{-1} > \lambda$$

Set  $n_k = D^k$  and

$$A_k = \{J_{n_k}(\alpha_{n_k}(\lambda)) - J_{n_{k-1}}(\alpha_{n_k}(\lambda)) \geq r_{n_k}\}.$$

Now  $J_{n_k}(\alpha_{n_k}(\lambda)) - J_{n_{k-1}}(\alpha_{n_k}(\lambda))$  has the same distributions as  $J_{n_k - n_{k-1}}(\alpha_{n_k}(\lambda))$  and one can again easily check that the conditions of Lemma 3.2 are met, so

$$\begin{aligned} P(A_k) &\geq c_1 r_{n_k}^{-1/2} \exp(r_{n_k} - r_{n_k} l r_{n_k} + r_{n_k} l((n_k - n_{k-1}) G(\alpha_{n_k}(\lambda))) - 2(n_k - n_{k-1}) G(\alpha_{n_k}(\lambda))) \\ &\geq c_1 r_{n_k}^{-1/2} \exp(r_{n_k} - r_{n_k} l r_{n_k} - r_{n_k} l \beta_{n_k} + r_{n_k} (l(1 - D^1) - l\lambda) - 2(\lambda \beta_{n_k})^{-1}) \\ &= c_1 r_{n_k}^{-1/2} a_{n_k} \exp((l(1 - D^{-1}) - l\lambda) r_{n_k} - 2(\lambda \beta_{n_k})^{-1}) \end{aligned}$$

which gives rise to a divergent series by (2.9) and (4.3) since  $r_{n_k} \beta_{n_k} \rightarrow \infty$  by (2.13). Since  $A_k, k = 1, 2, \dots$  are independent events,  $P(A_k \text{ i.o.}) = 1$  from which the result follows.  $\square$

Of course in general  $\alpha_n(\lambda)$  and  $\alpha_n$  need not be comparable, for example when  $G$  is slowly varying, however in our case we have

**Corollary 4.2.** *Assume that  $r_n$  satisfies (2.5);*

(a) *if  $X$  satisfies (2.1) then*

$$(4.4) \quad 0 < \limsup_{n \rightarrow \infty} |^{(r_n)}X_n| \alpha_n^{-1} < \infty$$

(b) *if  $X$  is in the domain of attraction of a stable law of index  $\alpha \in (0, 2)$  then*

$$(4.5) \quad \limsup_{n \rightarrow \infty} |^{(r_n)}X_n| \alpha_n^{-1} = 1.$$

*Proof.* (a) follows from (2.17) and (4.1) while (b) follows from (2.18) and (4.1).  $\square$

*Remark.* It is easy to see that (4.4) holds more generally than under condition (2.1). What is needed for (4.4) is that  $\alpha_n(\lambda_1) \approx \alpha_n(\lambda_2)$  for some  $\lambda_1 < 1 < \lambda_2$ . This is true for example if there exists  $p > 0$  and a non-increasing function  $f$  such that  $x^p G(x) \approx f(x)$  as  $x \rightarrow \infty$ . In particular this is true for many random variables in the domain of attraction of the normal law.

As was pointed out earlier any sequence  $a_n$  satisfying (2.6)–(2.8), satisfies (2.6)–(2.8) with  $r_n$  replaced by  $r_n + j$ . Thus defining

$$(4.6) \quad \begin{aligned} \beta_n(j) &= \exp((l a_n^{-1} - (r_n + j) l(r_n + j) + (r_n + j))(r_n + j)^{-1}) \\ \alpha_n(\lambda, j) &= \min \{x: G(x) = (\lambda n \beta_n(j))^{-1}\} \end{aligned}$$

we have by Theorem 4.1 that for each  $j \geq 0$

$$(4.7) \quad \limsup_{n \rightarrow \infty} \frac{|^{(r_n + j)}X_n|}{\alpha_n(\lambda, j)} \begin{cases} \leq 1 & \text{if } \lambda > 1 \\ \geq 1 & \text{if } \lambda < 1. \end{cases}$$

In particular for each  $j \geq 0, |^{(r_n + j)}X_n| \leq \alpha_n(2, j)$  eventually. Our next aim, Lemma 4.4, is to show that this holds uniformly in  $j$  for  $0 \leq j \leq j_n$ , provided  $j_n = o(r_n)$ . Actually we will not prove quite this much, we will introduce a new sequence

$b_n(j)$ , which is more convenient to work with and show that  $|^{(r_n+j)}X_n| \leq b_n(j)$  for all  $0 \leq j \leq r_n$  eventually. To do this set

$$(4.8) \quad \delta_n = \exp\left(-\frac{l_2 n}{2r_n^2}\right)$$

and for  $0 \leq j \leq r_n$  define

$$(4.9) \quad b_n(j) = \min \{x: G(x) = (2n \delta_n^j \beta_n)^{-1}\}.$$

First note that by (2.5) and (2.7)

$$\begin{aligned} n \delta_n^{r_n} \beta_n &\geq n \exp\left(-\frac{l_2 n}{2r_n} + \frac{l_2 n - r_n l r_n + r_n}{r_n}\right) \\ &\geq n r_n^{-1} \\ &\rightarrow \infty. \end{aligned}$$

Thus

$$(4.10) \quad b_n(r_n) \rightarrow \infty$$

and since  $\delta_n < 1$  we trivially have

$$(4.11) \quad b_n(r_n) \leq b_n(j) < b_n(j-1) \leq \alpha_n(2)$$

for  $1 \leq j \leq r_n$ . Next observe that for  $0 \leq j \leq r_n$

$$\begin{aligned} l(\delta_n^j \beta_n \beta_n(j)^{-1}) &= -\frac{j l_2 n}{2r_n^2} + \frac{l a_n^{-1}}{r_n} - l r_n - \frac{l a_n^{-1}}{r_n + j} + l(r_n + j) \\ &\geq \frac{j l a_n^{-1}}{r_n(r_n + j)} - \frac{j l_2 n}{2r_n^2} \\ &\geq 0 \end{aligned}$$

by (2.7). Thus

$$(4.12) \quad \delta_n^j \beta_n \geq \beta_n(j), \quad 0 \leq j \leq r_n.$$

Later we will need to compare  $b_n(j)$  with  $b_n(1)$  and also  $b_n(1)$  with  $\alpha_n$ . To do this we let

$$(4.13) \quad \alpha = [2l(4\theta)] + 1$$

where  $\theta$  is given by (2.2). Thus by (2.14)

$$(4.14) \quad \delta_n^{-\alpha N_n} \geq 4\theta.$$

**Lemma 4.3.** *If  $n$  is sufficiently large, then for all integers  $k \geq 0$  satisfying  $\alpha k N_n + 1 \leq r_n$*

$$(4.15) \quad b_n(\alpha k N_n + 1) \leq 2^{-k} b_n(1).$$

*Proof.* If  $k = 0$  then the result is trivial, thus we may assume  $k \geq 1$ .

By (2.4) if  $2^{-k} b_n(1) \geq x_0$  then

$$\begin{aligned}
 (4.16) \quad G(2^{-k} b_n(1)) &\leq 4^k \theta G(b_n(1)) \\
 &\leq (4\theta)^k (2n \delta_n \beta_n)^{-1} \\
 &\leq (2n \delta^{\alpha k N_n + 1} \beta_n)^{-1} \\
 &= G(b_n(\alpha k N_n + 1)).
 \end{aligned}$$

Thus if we can show that  $2^{-k} b_n(1) \geq x_0$  holds whenever  $\alpha k N_n + 1 \leq r_n$ , provided  $n$  is sufficiently large, we will be done. Let

$$(4.17) \quad k_n = \max \{k : 2^{-k} b_n(1) \geq x_0\}.$$

Since  $b_n(1) \rightarrow \infty$  by (4.10) and (4.11), we must have  $k_n \rightarrow \infty$ . Suppose that  $\alpha k_n N_n + 1 \leq r_n$  infinitely often, then by (4.16) and (4.17) along some subsequence we have both

$$2^{-(k_n+1)} b_n(1) < x_0$$

and

$$\begin{aligned}
 2^{-k_n} b_n(1) &\geq b_n(\alpha k_n N_n + 1) \\
 &\geq b_n(r_n) \\
 &\rightarrow \infty
 \end{aligned}$$

by (4.10) and (4.11). This is a contradiction and so  $\alpha k_n N_n + 1 \geq r_n$  eventually which completes the proof.  $\square$

If  $r_n = o((l_2 n)^{1/2})$  then  $\delta_n \rightarrow 0$  and so for every  $\varepsilon > 0$  by (2.4)

$$\begin{aligned}
 G(\varepsilon \alpha_n) &\leq \varepsilon^{-2} \theta G(\alpha_n) \\
 &= \varepsilon^{-2} \theta (n \beta_n)^{-1} \\
 &\leq (2n \delta_n \beta_n)^{-1} \\
 &= G(b_n(1))
 \end{aligned}$$

provided  $n$  is sufficiently large. Thus if  $r_n = o((l_2 n)^{1/2})$  then

$$(4.18) \quad b_n(1) = o(\alpha_n).$$

**Lemma 4.4.** *Let  $j_n$  be any sequence of integers satisfying  $j_n r_n^{-1} \rightarrow 0$  and set*

$$(4.19) \quad A_n = \{ |r_n^{(j)} X_n| > b_n(j) \quad \text{for some } 0 \leq j \leq j_n \}$$

Then

$$(4.20) \quad P(A_n \text{ i.o.}) = 0.$$

*Proof.* Let  $a \in (1, 2)$  and define  $n_k$  by (2.19). Set  $m_k = n_{k+1} - 1$  and let

$$\begin{aligned}
 B_k &= \{A_n \text{ for some } n_k \leq n \leq m_k\}. \\
 \hat{b}_k(j) &= \min_{n_k \leq n \leq m_k} b_n(j).
 \end{aligned}$$

By (2.22), since we may assume  $j_n$  is nondecreasing, we have

$$(4.21) \quad P(B_k) \leq P(|^{(r_{m_k}+j)}X_{m_k}| > \hat{b}_k(j) \quad \text{for some } 0 \leq j \leq j_{m_k}) \\ \leq j_{m_k} \max_{0 \leq j \leq j_{m_k}} P(|^{(r_{m_k}+j)}X_{m_k}| > \hat{b}_k(j)).$$

We now wish to apply Lemma 3.2. By (2.5) it is clear that if  $k$  is sufficiently large then  $\max_{0 \leq j \leq j_{m_k}} (r_{m_k}+j)^2 < m_k$ . To check (3.6) first observe that by (4.12)

$$(4.22) \quad G(\hat{b}_k(j)) = \max_{n_k \leq n \leq m_k} G(b_n(j)) \\ = \max_{n_k \leq n \leq m_k} (2n \delta_n^j \beta_n)^{-1} \\ \leq \max_{n_k \leq n \leq m_k} (2n \beta_n(j))^{-1} \\ \leq (2n_k \beta_{n_k}(j))^{-1}$$

by (2.6) and (2.22). Thus for every  $0 \leq j \leq j_{m_k}$

$$(4.23) \quad m_k G(\hat{b}_k(j)) \leq a n_k G(\hat{b}_k(j)) \\ \leq a(2\beta_{n_k}(j))^{-1} \\ = (a/2) \exp\left(-\frac{la_{n_k}^{-1}}{r_{n_k}+j} + l(r_{n_k}+j) - 1\right) \\ \leq (a/2)(r_{n_k}+j) \exp\left(-\frac{la_{n_k}^{-1}}{r_{n_k}+j_{n_k}}\right) \\ \leq 1/2(r_{m_k}+j)$$

for large  $k$ , independent of  $j$ , by (2.5), (2.7) and (2.22). Consequently we can apply Lemma 3.2 to obtain

$$(4.24) \quad P(B_k) \leq j_{m_k} \max_{0 \leq j \leq j_{m_k}} c_2 (r_{m_k}+j)^{-1/2} \exp((r_{m_k}+j) \\ - (r_{m_k}+j) l(r_{m_k}+j) + (r_{m_k}+j) l(m_k G(\hat{b}_k(j))))$$

Using (2.22) and (4.22) the exponent above can be written as

$$(4.25) \quad (r_{m_k}+j) - (r_{m_k}+j) l(r_{m_k}+j) + (r_{m_k}+j) l(n_k G(\hat{b}_k(j))) + (r_{m_k}+j) l(m_k n_k^{-1}) \\ \leq (r_{n_k}+j) - (r_{n_k}+j) l(r_{n_k}+j) - (r_{n_k}+j) l\beta_{n_k}(j) + (r_{n_k}+j)(la - l2) \\ = la_{n_k} + (r_{n_k}+j)(la - l2)$$

by (4.6). Thus, since  $a < 2$

$$P(B_k) \leq c_2 j_{m_k} r_{m_k}^{-1/2} a_{n_k} e^{(la-l2)r_{n_k}}$$

and this gives rise to a convergent series by (2.10) since  $j_{m_k} r_{m_k}^{-1} = j_{m_k} r_{m_k}^{-1} \rightarrow 0$ . The result now follows from the Borel-Cantelli Lemma.  $\square$

In proving the lower bound in Theorem 4.8, we will, roughly speaking, need to ensure that infinitely often the  $N_n$  largest terms are all of size  $\alpha_n$  and further that all of these terms have the same sign. This will be formulated precisely, together with an additional requirement, in Lemma 4.6. To prove this we must first introduce some further notation.

For  $0 \leq y_1 \leq y_2$ , if  $G(y_1) > G(y_2)$  let  $X(y_1, y_2)$  be a random variable with distribution function  $F_{X(y_1, y_2)}$  given by

$$(4.26) \quad dF_{X(y_1, y_2)}(x) = 1(y_1 \leq |x| \leq y_2)(G(y_1) - G(y_2))^{-1} dF_X(x).$$

Thus  $X(y_1, y_2)$  is  $X$  conditioned to have absolute value between  $y_1$  and  $y_2$ . Note that

$$(4.27) \quad G_{X(y_1, y_2)}(x) = \begin{cases} 1 & \text{if } x \leq y_1 \\ (G(x) - G(y_2))(G(y_1) - G(y_2))^{-1} & \text{if } y_1 \leq x \leq y_2 \\ 0 & \text{if } x \geq y_2 \end{cases}$$

We will write  $X(y_1)$  for  $X(0, y_1)$ .

For  $r \geq 2$  and  $s \geq 0$  let  $H_{m, r+s, r-1}(y_1, y_2)$  denote the two-dimensional distribution function of  $(|^{(r+s)}X_m|, |^{(r-1)}X_m|)$ . Observe that this distribution assigns zero probability to the complement of the set  $\{(y_1, y_2): 0 \leq y_1 \leq y_2 \text{ and } G(y_1) > G(y_2)\}$ . The following proposition does not appear to be in the literature, but since variants of it are well known, (see for example Lemma 1.1 of [11]), we will not prove it here.

**Proposition 4.5.** *Let  $X_i(y_1)$ ,  $i = 1, 2, \dots$ , and  $X_j(y_1, y_2)$ ,  $j = 1, 2, \dots$ , be sequences of i.i.d. random variables with common distributions given by  $X(y_1)$  and  $X(y_1, y_2)$  respectively. Further assume that these sequences are independent. Then for all  $r \geq 2$ , all  $0 \leq s \leq u$ , all bounded Borel functions  $\phi_1: \mathbb{R}^{u-s} \rightarrow \mathbb{R}^1$ ,  $\phi_2: \mathbb{R}^s \rightarrow \mathbb{R}^1$ , and all Borel sets  $B \subseteq [0, \infty) \times [0, \infty)$ .*

$$(4.28) \quad E[\phi_1(^{(r+u)}X_m, \dots, ^{(r+s+1)}X_m) \phi_2(^{(r+s-1)}X_m, \dots, ^{(r)}X_m); \\ (|^{(r+s)}X_m|, |^{(r-1)}X_m|) \in B] \\ = \int_B E \phi_1(^{(u-s)}X_{m-r-s}(y_1), \dots, ^{(1)}X_{m-r-s}(y_1)) E \phi_2(^{(s)}X_s(y_1, y_2), \dots, \\ ^{(1)}X_s(y_1, y_2)) dH_{m, r+s, r-1}(y_1, y_2).$$

*Remarks.* 1. If  $s = 0$  or  $u = s$  we should explain what is meant by (4.28). If  $s = 0$  then  $\phi_2 \equiv 1$ , while if  $u = s$  then  $\phi_1 \equiv 1$ .

2. The more intuitive way of phrasing (4.28) is that the distribution of  $(^{(r+u)}X_m, \dots, ^{(r+s+1)}X_m, ^{(r+s-1)}X_m, \dots, ^{(r)}X_m)$  conditioned by  $|^{(r+s)}X_m| = y_1$ ,  $|^{(r-1)}X_m| = y_2$  is given by  $(^{(u-s)}X_{m-r-s}(y_1), \dots, ^{(1)}X_{m-r-s}(y_1), ^{(s)}X_s(y_1, y_2), \dots, ^{(1)}X_s(y_1, y_2))$ .

Set

$$(4.29) \quad p_n = [r_n^2 / l_2 n].$$

**Lemma 4.6.** For every integer  $N \geq 1$  there exists  $\lambda_2 \in (0, 1)$  such that for all  $\lambda_1 \in (0, \lambda_2)$  and all  $\varepsilon \in (0, 1)$

$$(4.30) \quad P(|^{(r_n+t_n)}X_n| \leq \alpha_n(\lambda_1), |^{(r_n+s_n)}X_n| \geq \alpha_n(\lambda_2), E_n \text{ i.o.}) = 1$$

where  $s_n = N p_n$ ,  $t_n = [(N + 2\varepsilon) p_n] + 1$  and  $E_n = E_n^+ \cup E_n^-$  where

$$(4.31) \quad E_n^+ = \left\{ \sum_{i=0}^{s_n-1} 1(^{(r_n+i)}X_n > 0) \geq (1 - \varepsilon) s_n \right\}$$

$$(4.32) \quad E_n^- = \left\{ \sum_{i=0}^{s_n-1} 1(^{(r_n+i)}X_n < 0) \geq (1 - \varepsilon) s_n \right\}.$$

*Remark.* If  $s_n = 0$ , then  $E_n$  is the whole space.

*Proof.* Fix  $N \geq 1$  and choose  $\lambda_2 \in (0, 1)$  so that

$$(4.33) \quad l(2\lambda_2) + 2N + 1 < 0$$

Let  $n_k = 2^k$ , and for notational convenience write  $r_k = r_{n_k}$ ,  $s_k = s_{n_k}$ ,  $t_k = t_{n_k}$ ,  $p_k = p_{n_k}$  and  $\alpha_k(\lambda) = \alpha_{n_k}(\lambda)$ . Define  $u_k = [(N + \varepsilon) p_k] + 1$  and  $v_k = [\varepsilon p_k]$ . Note that  $u_k > s_k$ . Let

$^{(r)}Z_k = r^{\text{th}}$  largest random variable in absolute value from amongst

$$X_{n_{k-1}+1}, \dots, X_{n_k}$$

$$A_k^+ = \{ ^{(r_k+i)}Z_k > 0 \quad \text{for all } 0 \leq i < s_k \}$$

$$A_k^- = \{ ^{(r_k+i)}Z_k < 0 \quad \text{for all } 0 \leq i < s_k \}$$

$$A_k = A_k^+ \cup A_k^-$$

$$B_k = \{ |^{(r_k+u_k)}Z_k| \leq \alpha_k(\lambda_1), |^{(r_k+s_k)}Z_k| \geq \alpha_k(\lambda_2) \}$$

$$C_k = \{ J_{n_{k-1}}(\alpha_k(\lambda_1)) \leq v_k \}.$$

If  $s_k = 0$  then  $A_k^+$  and  $A_k^-$  are the whole space. Observe that on the event  $A_k^+ B_k C_k$  since  $u_k + v_k \leq t_k$  we have

$$(4.34) \quad |^{(r_k+t_k)}X_{n_k}| \leq \alpha_k(\lambda_1)$$

$$(4.35) \quad |^{(r_k+s_k)}X_{n_k}| \geq \alpha_k(\lambda_2)$$

and

$$\begin{aligned} \sum_{i=0}^{s_k-1} 1(^{(r_k+i)}X_{n_k} > 0) &\geq s_k - v_k \\ &\geq (N - \varepsilon) p_k \\ &\geq (1 - \varepsilon) s_k. \end{aligned}$$

Similarly on the event  $A_k^- B_k C_k$  we have (4.34), (4.35) and

$$(4.36) \quad \sum_{i=0}^{s_k-1} 1(^{(r_k+i)}X_{n_k} < 0) \geq (1 - \varepsilon) s_k.$$

Hence to prove (4.30), it suffices to show  $P(A_k B_k C_k \text{ i.o.})=1$ . To do this we will use Lemma 3.4. First observe that  $A_k B_k, k=1, 2, \dots$  are independent, and for each  $k, C_k$  and  $A_k B_k$  are independent. Now setting  $v'_k = v_k + 1$  we have by (3.12)

$$1 - P(C_k) = P(J_{n_{k-1}}(\alpha_k(\lambda_1)) \geq v'_k) \rightarrow 0$$

since

$$\begin{aligned} n_{k-1} G(\alpha_k(\lambda_1))(v'_k)^{-1} &\leq (\lambda_1 \beta_k v'_k)^{-1} \\ &\leq (1 + \varepsilon)(\varepsilon \lambda_1 \beta_k r_k^2 / l_2 n_k)^{-1} \\ &\rightarrow 0 \end{aligned}$$

by (2.13) and (3.13). Next let  $m_k = n_k - n_{k-1} = 2^{k-1}$ , so  $\{^{(j)}Z_k: 1 \leq j \leq m_k\} = \{^{(j)}X_{m_k}: 1 \leq j \leq m_k\}$ . Thus to compute  $P(A_k B_k)$  we can use Proposition 4.5 with

$$\begin{aligned} \phi_1(x_1, \dots, x_{u_k - s_k}) &= 1(|x_1| \leq \alpha_k(\lambda_1)) \\ \phi_2(z_1, \dots, z_{s_k}) &= \begin{cases} \sum_{i=1}^{s_k} 1(z_i > 0) + \prod_{i=1}^{s_k} 1(z_i < 0) & \text{if } s_k \neq 0 \\ 1 & \text{if } s_k = 0 \end{cases} \end{aligned}$$

$$B = \{(y_1, y_2): \alpha_k(\lambda_2) \leq y_1 \leq y_2 < \infty\}.$$

Recall that  $u_k > s_k$  so there is no need to modify the definition of  $\phi_1$  to include the case  $u_k = s_k$ . Observe that if  $G(y_1) > G(y_2)$  then

$$\begin{aligned} E \phi_2(^{(s_k)}X_{s_k}(y_1, y_2), \dots, ^{(1)}X_{s_k}(y_1, y_2)) \\ = E \phi_2(X_1(y_1, y_2), \dots, X_{s_k}(y_1, y_2)) \\ \geq 2^{-s_k} \end{aligned}$$

while for any  $y_1 \geq \alpha_k(\lambda_2)$

$$\begin{aligned} E \phi_1(^{(u_k - s_k)}X_{m_k - r_k - s_k}(y_1), \dots, ^{(1)}X_{m_k - r_k - s_k}(y_1)) \\ = 1 - P(|^{(u_k - s_k)}X_{m_k - r_k - s_k}(y_1)| > \alpha_k(\lambda_1)) \\ \rightarrow 1 \end{aligned}$$

uniformly in  $y_1 \geq \alpha_k(\lambda_2)$  by (3.4), since by (4.27)

$$\begin{aligned} (m_k - r_k - s_k) G_{X(y_1)}(\alpha_k(\lambda_1))(u_k - s_k)^{-1} &\leq (m_k - r_k - s_k) G(\alpha_k(\lambda_1))(u_k - s_k)^{-1} \\ &\leq (\lambda_1 \beta_k (u_k - s_k))^{-1} \\ &\leq (1 + \varepsilon)(\varepsilon \lambda_1 \beta_k r_k^2 / l_2 n_k)^{-1} \\ &\rightarrow 0 \end{aligned}$$

by (2.13) and (3.13). Thus by (4.28), for large  $k$

$$P(A_k B_k) \geq 2^{-(s_k + 1)} P(|^{(r_k + s_k)}X_{m_k}| \geq \alpha_k(\lambda_2)),$$



and so to complete the proof we must show that this gives rise to a divergent series. Let  $w_n = r_n + s_n$  and write  $w_k = w_{n_k}$ . Note that  $|w_n - r_n| = o(r_n)$  by (2.5) thus it is easy to check that conditions (3.5) and (3.6) are met, so by (3.4)

$$(4.37) \quad P(|^{(w_k)}X_{m_k}| \geq \alpha_k(\lambda_2)) \geq c_1 w_k^{-1/2} \exp(w_k - w_k l w_k - w_k l \beta_k - w_k l(2\lambda_2) - (\lambda_2 \beta_k)^{-1}).$$

Now set  $\beta'_k = \exp((l a_n^{-1} - w_n l w_n + w_n) w_n^{-1})$  and  $\beta'_k = \beta'_{n_k}$ . Observe that by (2.5) and (2.7)

$$\begin{aligned} l\beta_k - l\beta'_k &= \frac{s_k l a_k^{-1}}{r_k(r_k + s_k)} + l(1 + s_k r_k^{-1}) \\ &\leq \frac{2s_k l_2 n_k}{r_k^2} + o(1) \\ &\leq 2N + o(1). \end{aligned}$$

Thus for large  $k$ ,  $l\beta_k - l\beta'_k \leq 2N + 1$  and so

$$\begin{aligned} P(A_k B_k) &\geq c_1 2^{-(s_k+1)} w_k^{-1/2} \exp(w_k - w_k l w_k - w_k l \beta'_k \\ &\quad - w_k(l(2\lambda_2) + 2N + 1) - (\lambda_2 \beta_k)^{-1}) \\ &= c_1 2^{-(s_k+1)} w_k^{-1/2} a_{n_k} \exp(-w_k(l(2\lambda_2) + 2N + 1) - (\lambda_2 \beta_k)^{-1}). \end{aligned}$$

Since  $|w_n - r_n| = s_n = o(r_n)$  we have  $s_k = o(w_k)$ , and further by (2.13) that  $\beta_k^{-1} = o(w_k)$ . Thus by (4.33) and the remarks following (2.10), the above give rise to a divergent series and the proof is complete.  $\square$

Fix  $p > 2$  and let

$$(4.38) \quad d_n = \min \{x : G(x) = (l_2 n/r_n)^p (n \beta_n)^{-1}\}.$$

One easily checks that  $G(d_n) \rightarrow 0$  and so  $d_n \rightarrow \infty$ . Let

$$(4.39) \quad j_n = \left[ 4p \left( \frac{r_n^2}{l_2 n} \right) l \left( \frac{l_2 n}{r_n} \right) \right] + 1.$$

Note that by (2.5)

$$(4.40) \quad j_n r_n^{-1} \rightarrow 0.$$

Let  $n_k$  defined by (2.19) with  $a = 2$  and set  $m_k = n_{k+1} - 1$ . Let

$$(4.41) \quad \hat{d}_k = \min_{n_k \leq n \leq m_k} d_n,$$

then

$$(4.42) \quad G(\hat{d}_k) \leq (l_2 m_k/r_{m_k})^p (n_k \beta_{n_k})^{-1}.$$

Observe that

$$\begin{aligned} G(b_{m_k}(j_{m_k})) &= (2m_k \delta_{m_k}^{j_{m_k}} \beta_{m_k})^{-1} \\ &= (2m_k \beta_{m_k})^{-1} \exp(j_{m_k} l_2 m_k (2r_{m_k}^2)^{-1}) \\ &\geq (l_2 m_k / r_{m_k})^{2p} (2m_k \beta_{m_k})^{-1} \\ &\geq c (l_2 m_k / r_{m_k})^{2p} (n_k \beta_{n_k})^{-1} \end{aligned}$$

by (2.21) and (2.24) where  $c$  is independent of  $k$ . Thus by (2.5) for large  $k$ ,  $G(b_{m_k}(j_{m_k})) \geq G(\hat{d}_k)$  and so

$$(4.43) \quad b_{m_k}(j_{m_k}) \leq \hat{d}_k.$$

Similarly one can show that for large  $n$

$$(4.44) \quad b_n(j_n) \leq d_n.$$

Now let

$$(4.45) \quad \check{d}_k = \max_{n_k \leq n \leq m_k} d_n$$

then by (2.5), (2.21) and (2.24) for some  $c > 0$ , independent of  $k$

$$\begin{aligned} G(\check{d}_k) &\geq (l_2 n_k / r_{n_k})^p (m_k \beta_{m_k})^{-1} \\ &\geq c (l_2 n_k / r_{n_k})^p (n_k \beta_{n_k})^{-1} \\ &= c G(d_{n_k}). \end{aligned}$$

Thus by (2.4) since  $d_n \rightarrow \infty$

$$(4.46) \quad \check{d}_k \leq c d_{n_k}$$

for large  $k$ , where  $c$  is independent of  $k$ . Also note that by (2.1) for any  $\varepsilon > 0$ ,  $x^{2+\varepsilon} G(x) \rightarrow \infty$  as  $x \rightarrow \infty$ , (this actually holds for  $\varepsilon = 0$  also). Thus by (2.7) and (2.11),  $n^{-s} \alpha_n \rightarrow \infty$  for all  $s < 1/2$ . In particular  $(r_n / l_2 n)^{p/2} \alpha_n \rightarrow \infty$  and so by (2.4) for large  $n$

$$(4.47) \quad d_n \leq \theta^{1/2} (r_n / l_2 n)^{p/2} \alpha_n.$$

From (2.5) it then easily follows that

$$(4.48) \quad d_n r_n = o(\gamma_n).$$

Note also that by (2.21), (2.22), (4.42) and (4.46) for some constant  $c$

$$(4.49) \quad \check{d}_k m_k G(\hat{d}_k) \leq c d_{n_k} (l_2 n_k / r_{n_k})^p \beta_{n_k}^{-1} = o(\gamma_{n_k})$$

by (2.13) and (4.48).

To simplify notation in the next Lemma, it is convenient to define

$$\bar{U}_n(d) = U_n(d) - EU_n(d).$$

**Lemma 4.7.**

$$(4.50) \quad \limsup_{n \rightarrow \infty} \frac{|U_n(d_n) - EU_n(d_n)|}{\gamma_n} = 0 \quad \text{a.s.}$$

*Proof.* Let  $n_k$  be defined by (2.19) with  $a = 2$  and set  $m_k = n_{k+1} - 1$ . By (2.21)–(2.23) it suffices to prove

$$\limsup_{k \rightarrow \infty} \max_{n_k \leq n \leq m_k} \frac{|\bar{U}_n(d_n)|}{\gamma_{n_k}} = 0 \quad \text{a.s.}$$

First observe that for  $n_k \leq n \leq m_k$

$$\begin{aligned} |\bar{U}_n(d_n)| &= |\bar{U}_n(\hat{d}_k) + \sum_{i=1}^n X_i \mathbf{1}(\hat{d}_k < |X_i| \leq d_n) - \sum_{i=1}^n EX_i \mathbf{1}(\hat{d}_k < |X_i| \leq d_n)| \\ &\leq |\bar{U}_n(\hat{d}_k)| + \check{d}_k \sum_{i=1}^{m_k} \mathbf{1}(\hat{d}_k < |X_i| \leq \check{d}_k) + \sum_{i=1}^{m_k} E|X_i| \mathbf{1}(\hat{d}_k < |X_i| \leq \check{d}_k) \end{aligned}$$

Thus

$$(4.51) \quad \begin{aligned} \max_{n_k \leq n \leq m_k} |\bar{U}_n(d_n)| &\leq \max_{n_k \leq n \leq m_k} |\bar{U}_n(\hat{d}_k)| + \check{d}_k \sum_1^{m_k} \mathbf{1}(\hat{d}_k < |X_i| \leq \check{d}_k) \\ &\quad + \sum_1^{m_k} E|X_i| \mathbf{1}(\hat{d}_k < |X_i| \leq \check{d}_k) \\ &= I + II + III. \end{aligned}$$

Now by (4.43) for large  $k$

$$\begin{aligned} II &\leq \check{d}_k J_{m_k}(\hat{d}_k) \\ &\leq \check{d}_k J_{m_k}(b_{m_k}(j_{m_k})) \end{aligned}$$

while by (4.19) and (4.40) for large  $k$  we have a.s.

$$J_{m_k}(b_{m_k}(j_{m_k})) < r_{m_k} + j_{m_k}.$$

Thus for large  $k$ , by (2.22), (4.40) and (4.46)

$$\begin{aligned} II &\leq 2\check{d}_k r_{m_k} \\ &\leq c d_{n_k} r_{n_k} \\ &= o(\gamma_{n_k}) \end{aligned}$$

by (4.48). Next

$$\begin{aligned} III &\leq \check{d}_k m_k G(\hat{d}_k) \\ &= o(\gamma_{n_k}) \end{aligned}$$

by (4.49). To deal with  $I$ , first observe that for  $n_k \leq n \leq m_k$  by Chebyshev and (2.2)

$$\begin{aligned} P(|\bar{U}_{m_k}(\hat{d}_k) - \bar{U}_n(\hat{d}_k)| > \varepsilon \gamma_{n_k}) &\leq (m_k - n) \hat{d}_k^2 K(\hat{d}_k) (\varepsilon \gamma_{n_k})^{-2} \\ &\leq \theta \hat{d}_k^2 m_k G(\hat{d}_k) (\varepsilon \gamma_{n_k})^{-2} \\ &\rightarrow 0 \end{aligned}$$

by (4.48) and (4.49), uniformly in  $n$ . Thus by Skorohod's Lemma (Breiman [1] p. 45) for large  $k$

$$P\left(\max_{n_k \leq n \leq m_k} |\bar{U}_n(\hat{d}_k)| > 2\varepsilon \gamma_{n_k}\right) \leq 2P(|\bar{U}_{m_k}(\hat{d}_k)| > \varepsilon \gamma_{n_k}).$$

We will now use Lemma 3.1 with  $s = 2l_2 n_k$  and  $v = l_2 n_k (2r_{n_k})^{-1}$ . Then

$$s \hat{d}_k v^{-1} \leq 4 d_{n_k} r_{n_k} = o(\gamma_{n_k})$$

by (4.48) while

$$\frac{1}{2} v e^v m_k \hat{d}_k K(\hat{d}_k) \leq \theta v e^v d_{n_k} n_k G(\hat{d}_k) = o(\gamma_{n_k})$$

by (4.49). Hence by (3.3) for any  $\varepsilon > 0$ , if  $k$  is sufficiently large

$$P(|\bar{U}_{m_k}(\hat{d}_k)| > \varepsilon \gamma_{n_k}) \leq 2 \exp(-2l_2 n_k)$$

which gives rise to a convergent series by (2.20). The result now follows from Borel-Cantelli.  $\square$

We now come to our main result describing the L.I.L. behaviour of  $(r_n^{-1})S_n$ .

**Theorem 4.8.** *Assume that  $r_n$  satisfies (2.5) and let  $\gamma_n$  and  $d_n$  be given by (2.16) and (4.38) respectively*

(a) *If (2.1) holds then*

$$(4.52) \quad 0 < \limsup_{n \rightarrow \infty} |(r_n^{-1})S_n - EU_n(d_n)| \gamma_n^{-1} < \infty$$

(b) *If (2.1) holds and in addition  $r_n = o((l_2 n)^{1/2})$  then*

$$(4.53) \quad 0 < \limsup_{n \rightarrow \infty} |(r_n)X_n| \gamma_n^{-1} = \limsup_{n \rightarrow \infty} |(r_n^{-1})S_n - EU_n(d_n)| \gamma_n^{-1} < \infty$$

Further

$$(4.54) \quad \limsup_{n \rightarrow \infty} |(r_n^{-1})X_n| \gamma_n^{-1} = 0$$

$$(4.55) \quad \limsup_{n \rightarrow \infty} |(r_n)S_n - EU_n(d_n)| \gamma_n^{-1} = 0$$

(c) *If  $r_n = o((l_2 n)^{1/2})$  and  $X$  is in the domain of attraction of a stable law of index  $\alpha \in (0, 2)$ , then (4.54) and (4.55) hold and (4.53) can be strengthened to*

$$(4.56) \quad \limsup_{n \rightarrow \infty} |(r_n)X_n| \gamma_n^{-1} = \limsup_{n \rightarrow \infty} |(r_n^{-1})S_n - EU_n(d_n)| \gamma_n^{-1} = 1.$$

*Proof.* First note that to prove (4.54), it suffices by (4.7) to show that  $\alpha_n(\lambda, 1) \alpha_n^{-1} \rightarrow 0$  for some  $\lambda > 1$ . To do this it suffices by (2.4) to show that  $G(\alpha_n(\lambda, 1))/G(\alpha_n) = \beta_n(\lambda \beta_n(1))^{-1} \rightarrow \infty$ . But by (2.7)

$$\begin{aligned} l(\beta_n \beta_n(1))^{-1} &\geq \frac{l a_n^{-1}}{r_n} - \frac{l a_n^{-1}}{r_n + 1} \\ &\geq \frac{l_2 n}{r_n(r_n + 1)} \\ &\rightarrow \infty \end{aligned}$$

if  $r_n = o((l_2 n)^{1/2})$ .

Next observe that (4.56) is an immediate consequence of (4.53) and (4.5). Further (4.53) follows from (4.55) and (4.4). Thus we only have to prove (4.52) and (4.55). We begin with the proof of (4.55) and the upper bound in (4.52), which will be proved simultaneously. Observe that to prove the upper bound in (4.52), it suffices by (4.4) to show that

$$(4.57) \quad \limsup_{n \rightarrow \infty} |^{(r_n)}S_n - EU_n(d_n)| \gamma_n^{-1} < \infty.$$

Fix  $n$ ; if  $|^{(r_n+1)}X_n| \leq d_n$  then

$$|^{(r_n)}S_n - U_n(d_n)| \leq d_n r_n$$

and so by (4.48) and (4.50)

$$(4.58) \quad |^{(r_n)}S_n - EU_n(d_n)| = o(\gamma_n).$$

If  $|^{(r_n+1)}X_n| > d_n$  then

$$\begin{aligned} (4.59) \quad |^{(r_n)}S_n - EU_n(d_n)| &\leq \left| \sum_{j=1}^{n-r_n} {}^{(r_n+j)}X_n \mathbf{1}(|^{(r_n+j)}X_n| > d_n) \right| \\ &\quad + |U_n(d_n) - EU_n(d_n)| \\ &= I + II. \end{aligned}$$

By (4.50),  $II = o(\gamma_n)$ , thus we have left to estimate  $I$ . Let  $j_n$  be as in (4.39), then by (4.20), (4.40) and (4.44) we have that  $|^{(r_n+j_n)}X_n| \leq d_n$  eventually. Thus for large  $n$ , using (4.11), (4.15) and (4.20) we have a.s.

$$\begin{aligned} (4.60) \quad I &\leq \sum_{j=1}^{j_n-1} |^{(r_n+j)}X_n| \\ &\leq \sum_{j=1}^{j_n-1} b_n(j) \\ &\leq \sum_{k=0}^{\lfloor j_n/\alpha N_n \rfloor} \sum_{j=k\alpha N_n+1}^{(k+1)\alpha N_n} b_n(j) \\ &\leq \sum_{k=0}^{\lfloor j_n/\alpha N_n \rfloor} \alpha N_n b_n(k\alpha N_n + 1) \\ &\leq \alpha N_n b_n(1) \sum_0^{\lfloor j_n/\alpha N_n \rfloor} 2^{-k} \\ &\leq 2\alpha N_n b_n(1). \end{aligned}$$

Thus by (4.59) and (4.60)

$$|^{(r_n)}S_n - EU_n(d_n)| \leq \gamma_n(2\alpha b_n(1) \alpha_n^{-1} + o(1))$$

Now by (2.17) and (4.11),  $b_n(1) \alpha_n^{-1} = O(1)$  which proves (4.57), while if  $r_n = o((l_2 n)^{1/2})$  then  $b_n(1) \alpha_n^{-1} \rightarrow 0$  by (4.18) which proves (4.55).

To prove the lower bound in (4.52) we begin by letting  $N=4$  in Lemma 4.6 and choosing  $\lambda_2 \in (0, 1)$  to satisfy (4.30). By (2.17) we have

$$(4.61) \quad \alpha_n(\lambda_2) \geq (\lambda_2/2\theta)^{q-1} \alpha_n(2).$$

Next choose  $M$  an integer, large enough that

$$(4.62) \quad 2^{-M} \leq (8\alpha)^{-1} (\lambda_2/2\theta)^{q-1}$$

and set  $\lambda_1 = (16\theta \alpha^2 M^2)^{-1} \lambda_2$ . Observe that  $\lambda_1 \in (0, \lambda_2)$  and by (2.17)

$$(4.63) \quad \alpha_n(\lambda_1) \leq (4\alpha M)^{-1} \alpha_n(\lambda_2).$$

Set

$$(4.64) \quad \varepsilon = (8(1 + (2\theta/\lambda_2)^{q-1}))^{-1}$$

in Lemma 4.6 and let

$$D_n = \{ |^{(r_n + t_n)}X_n| \leq \alpha_n(\lambda_1), |^{(r_n + s_n)}X_n| \geq \alpha_n(\lambda_2), E_n \}.$$

Thus  $P(D_n \text{ i.o.}) = 1$ . If  $\omega \in D_n$  and  $n$  is sufficiently large then  $|^{(r_n + s_n)}X_n| > d_n$  since  $\alpha_n(\lambda_2) > d_n$  for large  $n$  by (2.17) and (4.47). Thus for infinitely many  $n$ ,  $D_n$  occurs and

$$\begin{aligned} |^{(r_n - 1)}S_n - EU_n(d_n)| &\geq \left| \sum_{j=0}^{s_n} ^{(r_n + j)}X_n \right| - \sum_{j=s_n+1}^{n-r_n} |^{(r_n + j)}X_n| \mathbb{1}(|^{(r_n + j)}X_n| > d_n) \\ &\quad - |U_n(d_n) - EU_n(d_n)| \\ &= I - II - III. \end{aligned}$$

Now by (4.20) and (4.44) for large  $n$

$$\begin{aligned} II &\leq \sum_{j=s_n+1}^{j_n} |^{(r_n + j)}X_n| \\ &= \sum_{j=s_n+1}^{t_n-1} |^{(r_n + j)}X_n| + \sum_{j=t_n}^{\alpha MN_n} |^{(r_n + j)}X_n| + \sum_{\alpha MN_n+1}^{j_n} |^{(r_n + j)}X_n| \\ &= II_1 + II_2 + II_3. \end{aligned}$$

Now for large  $n$  by (4.1)

$$\begin{aligned} II_1 &\leq (t_n - s_n - 1) |^{(r_n + s_n + 1)} X_n| \\ &\leq 2\varepsilon p_n |^{(r_n)} X_n| \\ &\leq 2\varepsilon p_n \alpha_n(2) \\ &\leq (1/4) p_n \alpha_n(\lambda_2) \end{aligned}$$

by (4.61) and (4.64). Since  $\omega \in D_n$

$$\begin{aligned} II_2 &\leq \alpha M N_n \alpha_n(\lambda_1) \\ &\leq (1/4) N_n \alpha_n(\lambda_2) \end{aligned}$$

by (4.63). By (4.15) and (4.20)

$$\begin{aligned} III_3 &\leq \sum_{\alpha M N_n + 1}^{j_n} b_n(j) \\ &\leq \sum_{k=M}^{\lfloor j_n / \alpha N_n \rfloor} \sum_{k \alpha N_n + 1}^{(k+1) \alpha N_n} b_n(j) \\ &\leq \alpha N_n \sum_{k=M}^{\lfloor j_n / \alpha N_n \rfloor} b_n(k \alpha N_n + 1) \\ &\leq \alpha N_n \sum_{k=M}^{\lfloor j_n / \alpha N_n \rfloor} 2^{-k} b_n(1) \\ &\leq \alpha N_n 2^{-M+1} b_n(1) \\ &\leq \alpha N_n 2^{-M+1} \alpha_n(2) \\ &\leq (1/4) N_n \alpha_n(\lambda_2) \end{aligned}$$

by (4.11), (4.61) and (4.62). Thus for large  $n$ , if  $\omega \in D_n$

$$(4.65) \quad II \leq (3/4) N_n \alpha_n(\lambda_2).$$

Next for  $\omega \in D_n$ , if  $p_n = 0$  then

$$\begin{aligned} (4.66) \quad I &= |^{(r_n)} X_n| \\ &\geq \alpha_n(\lambda_2) \\ &= N_n \alpha_n(\lambda_2) \end{aligned}$$

while if  $p_n \neq 0$  then  $s_n \geq 4$  and so

$$I = \left| \sum_{j=0}^{s_n-1} {}^{(r_n+j)} X_n 1({}^{(r_n+j)} X_n > 0) + \sum_{j=0}^{s_n-1} {}^{(r_n+j)} X_n 1({}^{(r_n+j)} X_n < 0) + {}^{(r_n+s_n)} X_n \right|.$$

Now since  $\omega \in D_n$  we have that  $\omega \in E_n^+ \cup E_n^-$ . If  $\omega \in E_n^+$  let  $j_0 \in [0, s_n - 1]$  be such that  ${}^{(r_n + j_0)}X_n > 0$ . Then  ${}^{(r_n + j_0)}X_n + {}^{(r_n + s_n)}X_n \geq 0$  and so for large  $n$  by (4.1)

$$(4.67) \quad I \geq ((1 - \varepsilon) s_n - 1) \alpha_n(\lambda_2) - \varepsilon s_n \alpha_n(2).$$

If  $\omega \in E_n^-$  then the analogous argument shows that (4.67) still holds. Thus by (4.61) and (4.64)

$$(4.68) \quad I \geq ((3/4) s_n - 1) \alpha_n(\lambda_2).$$

Now if  $p_n \neq 0$  then  $(3/4) s_n - 1 \geq N_n$  thus combining this with (4.66) gives that for large  $n$ , if  $\omega \in D_n$

$$(4.69) \quad I \geq N_n \alpha_n(\lambda_2).$$

Thus by (4.50), (4.65) and (4.69), for infinitely many  $n$

$$\begin{aligned} |{}^{(r_n - 1)}S_n - EU_n(d_n)| &\geq (1/4 + o(1)) N_n \alpha_n(\lambda_2) \\ &\geq c N_n \alpha_n \end{aligned}$$

where  $c > 0$  by (2.17) and this completes the proof of the lower bound.  $\square$

As an example assume that  $X$  is symmetric stable of index  $\alpha \in (0, 2)$  with scale parameter chosen so that  $G(x) \sim x^{-\alpha}$ . If  $r_n = o(l_2 n)$  and  $\liminf r_n (l_p n)^{-1} > 0$  for some  $p \geq 3$ , then as mentioned in section 2 we may take  $a_n = ((l_n) \dots (l_{p-1} n))^{-1}$  and so

$$(4.70) \quad \gamma_n = N_n n^{1/\alpha} \exp((l_2 n + \dots + l_p n - r_n l r_n + r_n)(\alpha r_n)^{-1}).$$

In particular if  $r_n = [l_p n]$  for some  $p \geq 3$  then it's easy to see that

$$(4.71) \quad \gamma_n \sim e^{2/\alpha} n^{1/\alpha} (l_p n)^{-1/\alpha} \exp((l_2 n + \dots + l_{p-1} n)(\alpha r_n)^{-1})$$

As we remarked in the introduction, the assumption of continuity on the distribution of  $X$  is not needed. The general case can be dealt with using the techniques described in [3]. In particular take for the definition of  ${}^{(r)}S_n$  the one given in Sect. 6 of [3]. Next with  $\tilde{G}$  given by (6.1) of [3], let  $\tilde{\alpha}_n = \tilde{G}((n\beta_n)^{-1})$  and  $\tilde{d}_n = \tilde{G}((l_2 n/r_n)^p (n\beta_n)^{-1})$  where  $p > 2$ . Then Theorem 4.8 holds with  $\tilde{\alpha}_n$  and  $\tilde{d}_n$  replacing  $\alpha_n$  and  $d_n$  respectively. The proof follows along the lines given here but the technical details are made more complicated.

### 5. Classical and Non-Classical L.I.L. Behaviour

We would like to explain a little further the remarks made in the introduction about the different ways in which the large values arise in (1.4) and (1.12). For simplicity assume that  $X$  is symmetric, else what we are really talking about is fluctuations of  ${}^{(r_n)}S_n$  from some centering sequence. We also assume (2.1), so (1.4) and (1.12) both hold without need for centering.



If  $r_n(l_2 n)^{-1} \rightarrow 0$  let  $N_n$  and  $\alpha_n$  be given by (2.14) and (2.15) respectively. If  $r_n(l_2 n)^{-1} \rightarrow \infty$  let  $N_n = r_n$  and define  $\alpha_n$  by  $G(\alpha_n) = r_n n^{-1}$ . Then as we have seen, in the first case, the large values of  $^{(r_n)}X_n$  are comparable to  $\alpha_n$ , and the large values of  $^{(r_n^{-1})}S_n$  arise because infinitely often there are  $N_n$  terms comparable in size to  $\alpha_n$  and these terms have the same sign. If  $r_n(l_2 n)^{-1} \rightarrow \infty$ , then by (4.1) of [3], we can again show that the large values of  $^{(r_n)}X_n$  are comparable to  $\alpha_n$  and again there are  $N_n$  terms comparable in size to  $\alpha_n$ . However the correct normalization for  $^{(r_n^{-1})}S_n$  (or  $^{(r_n)}S_n$ ) in this case is not  $N_n \alpha_n = r_n \alpha_n$ , but  $(r_n l_2 n)^{1/2} \alpha_n$ . There are two things to notice about this. First, the minimal number of summands required to make  $^{(r_n^{-1})}S_n$  as large as  $(r_n l_2 n)^{1/2} \alpha_n$ , is greater than  $l_2 n$ , more precisely there exists a sequence  $s_n$  such that  $s_n(l_2 n)^{-1} \rightarrow \infty$  and  $(|^{(r_n)}X_n| + \dots + |^{(r_n + s_n)}X_n|) = o((r_n l_2 n)^{1/2} \alpha_n)$ . Secondly, since there are  $r_n$  terms of size  $\alpha_n$ , there needs to be a lot of cancellation amongst terms in order that  $(r_n l_2 n)^{1/2} \alpha_n$  be the correct normalizer for  $^{(r_n^{-1})}S_n$ . Both of these properties are typical of classical L.I.L. behaviour. For example if  $EX^2 < \infty$ , one can show that there exists a sequence  $s_n$ , depending on  $X$ , such that  $(l_2 n) = o(s_n)$  and  $(|^{(1)}X_n| + \dots + |^{(s_n)}X_n|) = o((nl_2 n)^{1/2})$ . Furthermore, despite the paradoxical sounding nature of the statement, there has to be a lot of cancellation in order for  $S_n$  to take values of order  $(nl_2 n)^{1/2}$ . One way of expressing this for example, is that if  $t_n$  is any sequence such that

$$\limsup_{n \rightarrow \infty} \left( \sum_{i=1}^{t_n} {}^{(i)}X_n \right) (nl_2 n)^{-1/2} > 0$$

then

$$\limsup_{n \rightarrow \infty} \left( \sum_{i=1}^{t_n} |{}^{(i)}X_n| \right) (nl_2 n)^{-1/2} = \infty.$$

The idea that classical L.I.L. behaviour is due to many moderate summands rather than a few large summands is a common (though often well hidden) theme; see Klass [15] for a nice discussion.

The borderline case  $r_n \approx l_2 n$  is not included in (1.12) but is included in (1.4). This might lead one to think of it as giving rise to classical L.I.L. behaviour. However it may be that the techniques used in this paper can be extended to cover this case. Notice that the two definitions of  $\alpha_n$  do agree up to constants when  $r_n \approx l_2 n$ , since we may take  $a_n = (ln)^{-1}$ . Thus the large values in this case may arise in both ways!

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### Appendix

Given a sequence of integers  $r_n$  increasing to infinity, we construct a sequence  $a_n$  satisfying (2.6)–(2.8). Let

$$\begin{aligned} n_1 &= \min \{n \geq 10 : r_n = 1\} \\ m_1 &= \max \{n : (ln) \leq (ln_1)^2\} \\ n_{k+1} &= \min \{n > m_k : r_n \neq r_{m_k}\} \\ m_{k+1} &= \max \{n : (ln) \leq (ln_{k+1})^2\}. \end{aligned}$$

Clearly  $n_k < m_k < n_{k+1}$  and since  $r_n$  is integer valued

$$(A1) \quad r_{n_k} \geq k.$$

Define

$$a_n = \begin{cases} (ln_k)^{-2} & n_k \leq n \leq m_k \\ (ln)^{-2} & m_k < n \leq n_{k+1}. \end{cases}$$

Clearly  $a_n$  satisfies (2.6) and (2.7). To check (2.8) first observe that

$$\begin{aligned} \sum_n a_n n^{-1} &\geq \sum_k \left( \sum_{n_k}^{m_k} a_n n^{-1} \right) \\ &= \sum_k (ln_k)^{-2} \sum_{n_k}^{m_k} n^{-1} \\ &\geq \sum_k (ln_k)^{-2} (l(m_k+1) - ln_k) \\ &\geq \sum_k (ln_k)^{-2} ((ln_k)^2 - (ln_k)) \end{aligned}$$

which diverges. Next let  $\varepsilon < 0$ , then by (A1)

$$\begin{aligned} \sum_n a_n n^{-1} e^{\varepsilon r_n} &\leq \sum_k \left( \sum_{n_k}^{m_k} a_n n^{-1} e^{\varepsilon r_n} + \sum_{m_k+1}^{n_{k+1}} a_n n^{-1} e^{\varepsilon r_n} \right) \\ &\leq \sum_k e^{\varepsilon k} \sum_{n_k}^{m_k} a_n n^{-1} + \sum_n (n(ln)^2)^{-1}. \end{aligned}$$

Now the latter series converges while

$$\begin{aligned} \sum_{n_k}^{m_k} a_n n^{-1} &= (ln_k)^{-2} \sum_{n_k}^{m_k} n^{-1} \\ &\leq (ln_k)^{-2} l m_k \\ &\leq 1, \end{aligned}$$

thus (2.8) holds.

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