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**Probability** 

Theory Related Fields

## **Remarks on Limit Theorems for Nonlinear Functionals of Gaussian Sequences**

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**Summary.** Limit theorems for sums of nonlinear functionals of Gaussian sequences typically obtain as limit distribution that of a single term in an expansion given by Dobrushin [1] for a process subordinate to a Gaussian process. Here we show how one can obtain limit theorems of this type where the limit distribution is that of a full expansion of Dobrushin's type.

## Introduction

Let  $\{\xi_n\}$ , n = ..., -1, 0, 1, ... be a strictly stationary sequence. In dealing with limit laws one often defines the new sequences

$$S_n^N = A_N^{-1} \sum_{j \in B_n^N} \xi_j, \quad N = 1, 2, \dots,$$

where

$$B_n^N = \{j | nN \leq j < (n+1)N\}$$

and  $A_N$  is an appropriate norming constant. In investigating possible limit laws one is led to sequences  $\{\xi_n\}$  with the following property. The joint distribution of

is the same as that of

 $S_{n_1}^N,\ldots,S_{n_k}^N$ 

 $\xi_{n_1}, ..., \xi_{n_k}$ 

for all N = 1, 2, ... and integers  $n_1, ..., n_k$  with  $A_N = N^{\alpha}$  for some parameter  $\alpha > 0$ . The parameter  $\alpha$  is called the self-similarity parameter for the sequence  $\{\xi_n\}$  and the sequence itself is referred to as a self-similar process.

Dobrushin has characterized the self-similar processes subordinate to a Gaussian random spectral measure  $Z_G$  with spectral distribution G. For Borel

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sets A with  $G(A) < \infty$ , the random variables  $Z_G(A)$  have the following properties:

- (i) The random variables  $Z_G(A)$  are jointly Gaussian complex-valued.
- (ii)  $EZ_G(A) = 0$ ,  $EZ_G(A)\overline{Z_G(B)} = G(A \cap B)$ . (iii)  $\sum_{j=1}^n Z_G(A_j) = Z_G\left(\bigcup_{j=1}^n A_j\right)$  if the sets  $A_1, \dots, A_n$  are disjoint. (iv)  $Z_G(A) = \overline{Z_G(-A)}$ .

Dobrushin [1] has shown that if  $\{\xi_n\}$  has the form

$$\xi_{j} = \sum_{n=1}^{\infty} \frac{1}{n!} \int \exp\left\{ij \sum_{s=1}^{n} x_{s}\right\} \frac{\exp\left\{i \sum_{s=1}^{n} x_{s}\right\} - 1}{i\left(\sum_{s=1}^{n} x_{s}\right)}$$

$$f_{n}(x_{1}, \dots, x_{n}) Z_{G}(dx_{1}) \dots Z_{G}(dx_{n})$$
(1)

with

$$f_n(\lambda x_1, \dots, \lambda x_n) = \lambda^{1 - \frac{n\alpha}{2} - \beta} f_n(x_1, \dots, x_n)$$
$$G(\lambda A) = \lambda^{\alpha} G(A),$$

and

then  $\{\xi_n\}$  is a self-similar process with self-similarity parameter  $\beta$ . The integrals in the representation (1) are multiple Wiener-Ito integrals. One has to check whether the formula (1) is meaningful for a sequence  $f_n$  in the sense that the variance of (1) is finite. If we assume that the functions  $f_n$  are symmetric functions of the variables  $x_1, \ldots, x_n$  the variance of (1) is given by

$$\sum_{n=1}^{\infty} \frac{1}{n!} \int |f_n(x_1, \dots, x_n)|^2 K\left(\sum_{s=1}^n x_s\right) G(dx_1) \dots G(dx_n)$$

where K is the Fejer kernel

$$\mathbf{K}(\mu) = \frac{\sin^2 \frac{\mu}{2}}{\left(\frac{\mu}{2}\right)^2}.$$

The usual non Gaussian limit theorems for partial sums of nonlinear functionals of Gaussian sequences (see Dobrushin and Major [2]; Giraitis and Surgailis [5]; Major [3, 4], Taqqu [8, 9], and Rosenblatt [6, 7]) correspond to a single term in the expansion (1). In this paper, we will show that there are limit theorems in which the limiting distribution corresponds to a complete expansion of the type (1).

Let  $X_n$ ,  $n = ..., -1, 0, 1, ..., EX_n = 0$ ,  $EX_n^2 = 1$ , be a Gaussian stationary process with correlation function  $r(n) = EX_0X_n$  satisfying

$$r(n) \cong |n|^{-\alpha}, \quad 0 < \alpha < 1, \tag{2}$$

as  $|n| \rightarrow \infty$ . Let H be the spectral distribution function corresponding to (2) so that

$$r(n) = \int_{-\pi}^{\pi} e^{inx} H(dx).$$

The random spectral measure of the process  $\{X_n\}$  is  $Z_H$ . Let us consider the process  $\{Y_n\}$  subordinate to  $\{X_n\}$  given by

$$Y_{n} = \sum_{k > k_{0}} \frac{c_{k}}{k!} \int \exp\left[in(x_{1} + \dots + x_{k})\right] |x_{1}|^{\gamma_{k}} \dots |x_{k}|^{\gamma_{k}}$$
(3)  
$$Z_{H}(dx_{1}) \dots Z_{H}(dx_{k})$$

with

$$\gamma_k = \frac{1}{k} - \frac{\alpha}{2} - \frac{\beta}{k}, \quad k = 1, 2, \dots$$

and  $k_0$  the greatest integer less than or equal to  $2(1-\beta)/\alpha$ .

For the representation (3) to make sense we require that

$$\sum_{k>k_0} \frac{|c_k|^2}{k!} \int |x_1|^{2\gamma_k} \dots |x_k|^{2\gamma_k} H(dx_1) \dots H(dx_k) < \infty$$

$$\tag{4}$$

since it is the variance of (3). This requires at the very least that for any index k for which  $c_k \neq 0$  that

$$a_k = \int |x|^{2\gamma_k} H(dx) < \infty.$$

This will be finite if one has  $1 > \beta$ . Let us assume that  $2\gamma_k + \alpha = 2(1 - \beta)/k < 1/k$  so that

$$U_n^{(k)} = \int e^{in(x_1 + \dots + x_k)} \frac{e^{i(x_1 + \dots + x_k)} - 1}{i(x_1 + \dots + x_k)} |x_1|^{\gamma_k \dots} |x_k|^{\gamma_k}$$
$$|x_1|^{\frac{\alpha - 1}{2}} \dots |x_k|^{\frac{\alpha - 1}{2}} W(dx_1) \dots W(dx_k) D^{-k/2}$$

is well-defined where W is the random spectral measure of the white-noise process and  $D = 2\Gamma(\alpha) \cos\left(\frac{\pi}{2}\alpha\right)$ . This will happen if

 $\beta > \frac{1}{2}$ .

Notice that

$$V_{n} = \sum_{k > k_{0}} \frac{c_{k}}{k!} U_{n}^{(k)}$$
(5)

is well-defined if  $1 > \beta > \frac{1}{2}$  and

$$\sum_{k>k_0} \frac{|c_k|^2}{k!} \int \frac{\sin^2 \frac{1}{2} \left(\sum_{1}^k x_j\right)}{\left(\frac{1}{2} \sum_{1}^k x_j\right)^2} |x_1|^{2\gamma_k + \alpha - 1} \dots |x_k|^{2\gamma_k + \alpha - 1} dx_1 \dots dx_k D^{-k}$$

Notice that if (2) is satisfied, there is a constant

$$B > 2\Gamma(\alpha) \tag{6}$$

such that

$$|\int e^{ipx}H(dx)| < B(1+|p|)^{-a}$$

for all integral p. Let

$$A = 4B\left\{\Gamma(\alpha)\cos\left(\frac{\pi}{2}\alpha\right)\right\}^{-1}$$

We wish to show that under appropriate conditions, the partial sums of the process  $\{Y_n\}$  given by (3) will converge in distribution to the process  $\{V_n\}$  given by (5). More specifically, our object is to prove the following theorem.

**Theorem.** Let  $\{X_n\}$  be a stationary Gaussian process with mean zero and covariance function r(n) satisfying (2). If  $\{Y_n\}$  is the process (3) subordinate to  $\{X_n\}$ and satisfies

$$\sum_{k>k_0} \frac{|c_k|^2}{k!} A^k \Gamma\left(\frac{2}{k}(1-\beta)\right)^k < \infty$$
(7)

(see (8)) then the partial sums

$$N^{-\beta}\sum_{j\in B_n^N}Y_j, \quad n=1,\ldots,p,$$

converge in distribution to that of the random variables

$$V_n, n = 1, ..., p,$$

if  $1 > \beta > \frac{1}{2}$ . If  $2(1-\beta)/\alpha < 1$  or if

$$H(dx) = \left(\frac{c}{|x|^{1-\alpha}} + f(x)\right) dx, \quad 0 < \alpha < 1,$$

with the constant  $c = D^{-1}$  and f a function of bounded variation, the sum in (3), (5), and (7) can be taken from 1 to  $\infty$ .

The following simple lemma is helpful in the derivation.

**Lemma.** Under condition (2) and  $\gamma_k < 0$ ,

$$|\int e^{ipx} |x|^{2\gamma_{\kappa}} H(dx)| < B\left\{\frac{\Gamma(1-\alpha)}{\Gamma\left(1-\frac{2}{k}(1-\beta)\right)} + \frac{\Gamma\left(\frac{2}{k}(1-\beta)\right)}{\Gamma(\alpha)} \left(1 + \frac{\sin\pi\left(\alpha - \frac{2}{k}(1-\beta)\right)}{\sin\pi\alpha}\right)\right\}$$
(8)
$$\left\{\cos\left(\frac{\pi}{2}\eta\right)\right\}^{-1} (1+|p|)^{-2\gamma_{\kappa}-\alpha}.$$

In estimating (8) let us note that for k sufficiently large  $2\gamma_k = -\eta$  is negative. It is clearly of some interest to look at the Fourier coefficients of  $|x|^{-\eta}$ . Now

$$\int_{-\pi}^{\pi} e^{\ln x} |x|^{-\eta} dx = \int_{-\pi}^{\pi} \cos nx |x|^{-\eta} dx$$
$$= \int_{-\pi/2n}^{\pi/2n} \cos nx \left\{ \sum_{k=-n+1}^{n-1} (-1)^{k} \left| x + \frac{k\pi}{n} \right|^{-\eta} \right\} dx$$
$$+ (-1)^{n} \left[ \int_{\pi-\frac{\pi}{2n}}^{\pi} + \int_{-\pi}^{-\pi+\frac{\pi}{2n}} \right] \cos nx |x|^{-\eta} dx.$$
(9)

For n=0 we have

$$2\int_{0}^{\pi} x^{-\eta} dx = \frac{2}{1-\eta} \pi^{1-\eta}.$$

For  $n \neq 0$ , the first part of the expression on the right of (9) is

$$\int_{-\pi/2}^{\pi/2} \cos x \left\{ \sum_{k=-n+1}^{n-1} (-1)^k |x+k\pi|^{-\eta} \right\} dx n^{\eta-1}$$

and for large n this can be approximated by

$$\int_{-\pi/2}^{\pi/2} \cos x \sum_{k=-\infty}^{\infty} (-1)^k |x+k\pi|^{-\eta} dx n^{\eta-1} = \int_{-\infty}^{\infty} \cos x |x|^{-\eta} dx n^{\eta-1} = 2\Gamma(1-\eta) \cos\left(\frac{\pi}{2}(1-\eta)\right) n^{\eta-1}.$$
 (10)

If k is small and  $2\gamma_k$  is positive  $a_k$  is automatically finite. If  $2\gamma_k$  is negative we can still show that

$$\int |x|^{-\eta} H(dx) < \infty$$

Let

$$g_m(x) = \frac{1}{2\pi m} \int_{-\pi}^{\pi} \frac{\sin^2 \frac{m}{2} (x-y)}{\sin^2 \frac{1}{2} (x-y)} |y|^{-\eta} dy,$$

the Cesaro one sum of the truncated Fourier series of  $|x|^{-\eta}$ . Let the  $c_k$ 's be the Fourier coefficients of  $|x|^{-\eta}$ . Now

$$\int_{-\pi}^{\pi} g_m(x) dH(x) = \frac{1}{2\pi} \sum_{|k| \le m} c_k \left( 1 - \frac{|k|}{m} \right) r_{-k}$$
$$\leq \frac{1}{2\pi} \sum_k |c_k| |r_{-k}|.$$

Further  $g_m(x) \ge 0$  for all x and

$$\lim_{m\to\infty}g_m(x)=|x|^{-\eta}$$

for all  $x, 0 < |x| \le \pi$ . Therefore by Fatou's lemma

$$\int_{-\pi}^{\pi} |x|^{-\eta} dH(x) \leq \liminf_{m \to \infty} \int_{-\pi}^{\pi} g_m(x) dH(x)$$
$$\leq \frac{1}{2\pi} \sum |c_k| |r_{-k}|$$

and this last expression is finite by virtue of (2) and (10) since

$$\eta - 1 - \alpha = -\frac{2}{k}(1 - \beta) - 1.$$

We shall now show that

$$\int_{-\pi}^{\pi} e^{ipx} |x|^{-\eta} H(dx) = \frac{1}{2\pi} \sum_{j} c_{j} r_{p-j}.$$
(11)

Let us consider

$$\int_{-\pi}^{\pi} e^{ipx} g_m(x) dH(x) = \frac{1}{2\pi} \sum_{|k| \le m} c_k \left( 1 - \frac{|k|}{m} \right) r_{p-k}.$$
(12)

The equality (11) will follow from (12) on applying the Lebesgue convergence theorem with  $m \to \infty$ , if we can show that

$$g_m(x) \le K_1 + K_2 \left(\frac{1}{m^2} + x^2\right)^{-n/2}$$
(13)

for some constants  $K_1, K_2 > 0$ , First it is clear that for appropriate constants  $L_1, L_2 > 0$  one has

$$\frac{1}{2\pi m} \frac{\sin^2 \frac{m}{2}u}{\sin^2 \frac{1}{2}u} \leq L_1 + L_2 \frac{m}{\pi} \frac{1}{1 + m^2 u^2}$$

for  $|u| \leq \frac{\pi}{2}$ .

Now

$$\int_{-\infty}^{\infty} \frac{m}{\pi} \frac{1}{1 + m^2 (x - y)^2} |y|^{-\eta} dy = L_3 \operatorname{Re} \int_{-\infty}^{\infty} |\lambda|^{\eta - 1} e^{-\left|\frac{\lambda}{m}\right|} e^{iy\lambda} d\lambda$$

for an appropriate constant  $L_3 > 0$ . But

$$\operatorname{Re}\int_{0}^{\infty} \lambda^{\eta-1} e^{-\lambda \left(\frac{1}{m}-iy\right)} d\lambda = \operatorname{Re}\left(\frac{1}{m}-iy\right)^{-\eta} \Gamma(\eta)$$

and

$$\operatorname{Re}\left(\frac{1}{m}-iy\right)^{-\eta}=\cos\eta\,\theta\cdot\left(\frac{1}{m^2}+y^2\right)^{-\eta/2}$$

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with

$$\theta = \arccos\left(\frac{1}{m} \middle| \{m^{-2} + y^2\}^{\frac{1}{2}}\right).$$

This gives us the inequality (13) and consequently the relation (11). From (6) and (10) it follows that

$$\frac{1}{2\pi} \sum_{j} |c_{j}r_{p-j}| \leq \frac{B}{\pi} \Gamma(1-\eta) \cos\left(\frac{\pi}{2}(1-\eta)\right) \sum_{j} (1+|j|)^{\eta-1} (1+|p-j|)^{-\alpha}.$$
 (14)

The sum on the right side of inequality (14) can be partitioned into three parts

$$\left(\sum_{j=0}^{p} + \sum_{j<0} + \sum_{j>p}\right) (1+|j|)^{\eta-1} (1+|p-j|)^{-\alpha} = S_1 + S_2 + S_3.$$

An estimate for  $S_1$  is given by

$$\sum_{j=0}^{p} |j/p|^{\eta-1} \left| \frac{p-j}{p} \right|^{-\alpha} |p|^{\eta-1-\alpha}$$
$$\cong \int_{0}^{1} u^{\eta-1} (1-u)^{-\alpha} du |p|^{\eta-\alpha}$$
$$\equiv \frac{\Gamma(\eta)\Gamma(1-\alpha)}{\Gamma(\eta+1-\alpha)} |p|^{\eta-\alpha}.$$

if  $p \neq 0$ . A corresponding estimate for  $S_2$  is

$$\int_{-\infty}^{0} |u|^{\eta-1} |1-u|^{-\alpha} du|p|^{\eta-\alpha}$$
  
= 
$$\int_{0}^{1} |1-v|^{\eta-1} |v|^{1-\eta} |v|^{\alpha} |v|^{-2} dv|p|^{\eta-\alpha}$$
  
= 
$$\int_{0}^{1} |1-v|^{\eta-1} |v|^{\alpha-\eta-1} dv|p|^{\eta-\alpha}$$
  
= 
$$\frac{\Gamma(\eta)\Gamma(\alpha-\eta)}{\Gamma(\alpha)} |p|^{\eta-\alpha}$$

if  $p \neq 0$ . In a similar manner one obtains the estimate

$$\frac{\Gamma(\alpha-\eta)\Gamma(1-\alpha)}{\Gamma(1-\eta)}|p|^{\eta-\alpha}$$

for  $S_3$  if  $p \neq 0$ . These estimates imply that the right hand side of (14) is bounded by

$$\frac{B}{\pi}\Gamma(1-\eta)\cos\left(\frac{\pi}{2}(1-\eta)\right)\left\{\frac{\Gamma(\eta)\Gamma(1-\alpha)}{\Gamma(\eta+1-\alpha)} + \frac{\Gamma(\eta)\Gamma(\alpha-\eta)}{\Gamma(\alpha)} + \frac{\Gamma(\alpha-\eta)\Gamma(1-\alpha)}{\Gamma(1-\eta)}\right\}|p|^{\eta-\alpha}$$
(15)

if  $p \neq 0$ . The relation  $\Gamma(\eta)\Gamma(1-\eta) = \pi/\sin(\pi\eta)$  (recalling that  $\eta = \alpha - \frac{2}{k}(1-\beta)$ ) together with the bound (15) imply (8).

Let us now consider the proof of the theorem. The terms corresponding to different k in (3) are orthogonal to each other. The estimate of the lemma and (7) imply that  $Y_n$  is well-defined. Notice that

$$S_{n}^{N}(Y) = N^{-\beta} \sum_{j \in B_{n}^{N}} Y_{j}$$

$$= \sum_{k > k_{0}} \frac{c_{k}}{k!} \int e^{inN(x_{1} + \dots + x_{k})} \frac{e^{iN(x_{1} + \dots + x_{k})} - 1}{e^{i(x_{1} + \dots + x_{k})} - 1}$$

$$\prod_{s=1}^{k} (|x_{s}|^{y_{k}} Z_{G}(dx_{s}))$$

$$= \sum_{k > k_{0}} \frac{c_{k}}{k!} \int e^{in(\mu_{1} + \dots + \mu_{k})} \frac{e^{i(\mu_{1} + \dots + \mu_{k})} - 1}{\{e^{i\frac{1}{N}(\mu_{1} + \dots + \mu_{k})} - 1\}N}$$

$$Z_{kH_{N}}(d\mu_{1}) \dots Z_{kH_{N}}(d\mu_{k})$$

$${}_{k}H(dx) = |x|^{2\gamma_{k}} H(dx),$$

$$H_{n}(A) = N^{\tilde{k}_{k}} H(A^{1-1}A) = \tilde{x} - 2w + w^{-2}(1 - \theta)$$

with

$$_{k}H_{N}(A) = N^{\alpha_{k}}{}_{k}H(N^{-1}A), \quad \alpha_{k} = 2\gamma_{k} + \alpha = \frac{1}{k}(1-\beta).$$

The measures  ${}_{k}H_{N}(\cdot)$  tend locally weakly to the measures  ${}_{k}G(\cdot)$  where  ${}_{k}G(\cdot)$  has a spectral density

$$D^{-1}|x|^{\tilde{\alpha}_k-1}$$

with

$$D_k = 2\Gamma(\alpha)\cos\left(\frac{\alpha\,\pi}{2}\right).$$

Let

$$K_N(x_1, \dots, x_k) = \frac{e^{i(x_1 + \dots + x_k)} - 1}{\{e^{i\frac{1}{N}(x_1 + \dots + x_k)} - 1\}N}.$$

Just as in the case of the paper of Dobrushin and Major [2] one can show that the sequence of measures

$${}_{k}\mu_{N}(A) = \int_{A} \left| \sum_{p=1}^{l} \beta_{p} e^{in \, p(x_{1} + \dots + x_{k})} \right|^{2} |K_{N}(x_{1}, \dots, x_{k})|^{2} \\ {}_{k}H_{N}(dx_{1}) \dots {}_{k}H_{N}(dx_{k})$$
(16)

tends weakly to the finite measure

$${}_{k}\mu(A) = \int_{A} \left| \sum_{p=1}^{l} \beta_{p} e^{\operatorname{in} p(x_{1} + \dots + x_{k})} \right| p^{2} \left| K\left(\sum_{j=1}^{k} x_{j}\right) \right|^{2} G_{k}(dx_{1}) \dots G_{k}(dx_{k})$$
(17)

as  $N \rightarrow \infty$ ,  $k > k_0$ .

We shall now show that the distribution of

$$\sum_{k=k_{0}+1}^{m} \frac{c_{k}}{k!} \int e^{in(\mu_{1}+\ldots+\mu_{k})} K_{N}(\mu_{1},\ldots,\mu_{k})$$

$$Z_{kH_{N}}(d\mu_{1})\ldots Z_{kH_{N}}(d\mu_{k})$$
(18)

tends to that of

$$\sum_{k=k_{0}+1}^{m} \frac{c_{k}}{k!} \int e^{in(\mu_{1}+\ldots+\mu_{k})} K\left(\sum_{j=1}^{k} \mu_{j}\right)$$

$$\prod_{j=1}^{k} \left(\left|\mu_{j}\right|^{\gamma_{k}+\frac{\alpha-1}{2}} W(d\mu_{j})\right)$$
(19)

for each positive integer *m*. Let us now use the notation of Sect. 4 of the paper of Dobrushin [1]. Let  $h_k \in H^k_{G_k}$ ,  $k_0 < k \leq m$ . Then the distribution of

$$\sum_{k=k_0+1}^{m} \int h_k(x_1, \dots, x_k) Z_{kH_N}(dx_1) \dots Z_{kH_N}(dx_k)$$
(20)

tends to that of

$$\sum_{k=k_0+1}^{m} \int h_k(x_1, \dots, x_k) Z_{G_k}(dx_1) \dots Z_{G_k}(dx_k)$$
(21)

as  $N \to \infty$ . This follows since the integrals of (20) are polynomials in the random variables  $Z_{kH_N}(B)$  (the *B*s are the one-dimensional projections of the level sets of the functions  $h_k$ ). Since the joint distributions of the random variables  $Z_{kH_N}(B)$  tend to the joint distribution of the random variables  $Z_{G_k}(B)$ , it follows that the distribution of (20) tends to that of (21). Given any  $\varepsilon > 0$  there is an  $h_k \in H_{G_k}^k$  such that

$$\int_{\mathbb{R}^k} |K_N(x_1,\ldots,x_k) - h_k(x_1,\ldots,x_k)|^2 \qquad _k H_N(dx_1) \ldots {}_k H_N(dx_k) < \varepsilon,$$

 $k_0 < k \leq m$ , and

$$\int_{\mathbb{R}^k} \left| K\left( \sum_{j=1}^k x_j \right) - h_k(x_1, \dots, x_k) \right|^2 G_k(dx_1) \dots G_k(dx_k) < \varepsilon$$

The sets of inequalities follow from the convergence of the measures (16) to (17) and the convergence of  $K_N(x_1, \ldots, x_k)$  to  $K\left(\sum_{j=1}^k x_j\right)$  on every finite interval. The last comments imply that the distribution of (18) tends to the distribution of (19) as  $N \to \infty$  for each positive *m*.

The proof of the theorem will be complete if we can show that the variance of the remainder

k

$$\sum_{m=m}^{\infty} \frac{c_k}{k!} \int e^{in(\mu_1 + \dots + \mu_k)} K_N(\mu_1, \dots, \mu_k) Z_{\mu H_N}(d\mu_1) \dots Z_{\mu H_N}(d\mu_k)$$
(22)

can be made uniformly small as  $N \to \infty$  if m is fixed but sufficiently large. The variance of

is

If

$$\int e^{in(\mu_{1}+...+\mu_{k})} K_{N}(\mu_{1},...,\mu_{k}) Z_{kH_{N}}(d\mu_{1})...Z_{kH_{N}}(d\mu_{k})$$

$$\int |K_{N}(\mu_{1},...,\mu_{k})|^{2} {}_{k}H_{N}(d\mu_{1})...{}_{k}H_{N}(d\mu_{k}).$$
(23)

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 $\rho(j) = \int e^{ijx} |x|^{2\gamma_k} H(dx)$ 

then (23) by the argument on p. 34 of [2] can be rewritten

$$\frac{1}{N^{2-k\tilde{\alpha}}}\sum_{|j|< N} (N-|j|)\rho(j)^{k}.$$

Using the lemma, one can see that this is bounded in absolute value by

$$A^{k}\Gamma\left(\frac{2}{k}(1-\beta)\right)^{k}\sum_{|j|< N}\left(1-\frac{|j|}{N}\right)\left|\frac{j}{N}\right|^{-2(1-\beta)}N^{-1}.$$
(24)

The assumption (7) together with the estimate (24) implies that the variance of (22) is uniformly small as  $N \rightarrow \infty$  if m is sufficiently large.

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