

(origins or supply).

$$\sum_i x_{ij} = b_j, \quad j = 1, \dots, n$$

(destinations or demand), where  $x_{ij} \geq 0$  and  $\sum a_i = \sum b_j$ , which is called the *balance condition*. The assignment problem arises when  $m = n$  and all  $a_i$  and  $b_j$  are 1.

If all  $a_i$  and  $b_j$  in the transposed problem are integers, then there is an optimal solution for which all  $x_{ij}$  are integers (*Dantzig's theorem on integral solutions of the transport problem*).

In the assignment problem, for such a solution  $x_{ij}$  is either zero or one;  $x_{ij} = 1$  means that person  $i$  is assigned to job  $j$ ; the weight  $c_{ij}$  is the utility of person  $i$  assigned to job  $j$ .

The special structure of the transport problem and the assignment problem makes it possible to use algorithms that are more efficient than the **simplex method**. Some of these use the *Hungarian method* (see, e.g., [4], [5, Chapt. 7]), which is based on the König–Egervary theorem (see **König theorem**), the method of potentials (see [5], [1]), the *out-of-kilter algorithm* (see, e.g., [3]) or the transportation simplex method.

In turn, the transportation problem is a special case of the network optimization problem.

A totally different assignment problem is the **pole assignment problem** in control theory.

#### References

- [1] FRISCH, R.: 'La résolution des problèmes de programme linéaire par la méthode du potentiel logarithmique', *Cahier Sémin. Econom.* 4 (1956), 20–23.
- [2] GRÖTSCHEL, M., LOVÁSZ, L., AND SCHRIJVER, A.: *Geometric algorithms and combinatorial optimization*. Springer, 1987.
- [3] MURTZ, K.: *Linear and combinatorial programming*. Wiley, 1976.
- [4] PAPADIMITRIOU, C.H., AND STEIGLITZ, K.: *Combinatorial optimization*. Prentice-Hall, 1982.
- [5] YUDIN, D.B., AND GOL'SHTEIN, E.G.: *Linear programming*, Israel Program. Sci. Transl., 1965. (Translated from the Russian.)

*M. Hazewinkel*

MSC 1991: 90C05, 90B80, 90B06

**ASYMPTOTIC INVARIANT OF A GROUP** – A property of a **finitely-generated group**  $G$  which is a **quasi-isometry** invariant of the **metric space**  $(G, d_A)$ , where  $d_A$  is the **word metric** associated to a finite generating set  $A$  of  $G$  (cf. also **Quasi-isometric spaces**). This definition does not depend on the choice of the set  $A$ , since if  $B$  is another finite set of generators of  $G$ , then the metric spaces  $(G, d_A)$  and  $(G, d_B)$  are quasi-isometric.

The theory of asymptotic invariants of finitely-generated groups has been recently brought to the foreground by M. Gromov (see, in particular, [2] and [3]). As

Gromov says in [3, p. 8], 'one believes nowadays that the most essential invariants of an infinite group are asymptotic invariants'. For example, amenability (cf. **Invariant average**), hyperbolicity (in the sense of Gromov, cf. **Hyperbolic group**), the fact of being finitely presented (cf. **Finitely-presented group**), and the number of ends (cf. also **Absolute**) are all asymptotic invariants of finitely-generated groups. It is presently (1996) unknown whether the Kazhdan  $T$ -property is an asymptotic invariant. For an excellent survey on these matters, see [1].

A few examples of algebraic properties which are asymptotic invariants of finitely-generated groups are: being virtually nilpotent, being virtually Abelian, being virtually free.

#### References

- [1] GHYS, E.: 'Les groupes hyperboliques', *Astérisque* 189-190 (1990), 203–238, Sémin. Bourbaki Exp. 722.
- [2] GROMOV, M.: 'Hyperbolic groups', in S.M. GERSTEN (ed.): *Essays in Group Theory*, Vol. 8 of *MSRI Publ.*, Springer, 1987, pp. 75–263.
- [3] GROMOV, M.: 'Asymptotic invariants of infinite groups': *Proc. Symp. Sussex, 1991: II*, Vol. 182 of *London Math. Soc. Lecture Notes*, Cambridge Univ. Press, 1993, pp. 1–291.

*A. Papadopoulos*

MSC 1991: 20F32, 57M07

**ASYMPTOTIC OPTIMALITY** of estimating functions – Efficient estimation (cf. **Efficient estimator**) of parameters in stochastic models is most conveniently approached via properties of estimating functions, namely functions of the data and the parameter of interest, rather than estimators derived therefrom. For a detailed explanation see [3, Chapt. 1].

Let  $\{X_t: 0 \leq t \leq T\}$  be a **sample** in discrete or continuous time from a stochastic system taking values in an  $r$ -dimensional Euclidean space. The distribution of  $X_t$  depends on a parameter of interest  $\theta$  taking values in an open subset of a  $p$ -dimensional Euclidean space. The possible probability measures (cf. **Probability measure**) for  $X_t$  are  $\{P_\theta\}$ , a union of families of models.

Consider the class  $\mathcal{G}$  of zero-mean square-integrable estimating functions  $G_T = G_T(\{X_t: 0 \leq t \leq T\}, \theta)$ , which are vectors of dimension  $p$  and for which the matrices used below are non-singular.

Optimality in both the *fixed sample* and the *asymptotic* sense is considered. The former involves choice of an estimating function  $G_T$  to maximize, in the partial order of non-negative definite matrices, the information criterion

$$\mathcal{E}(G_T) := (\mathbb{E} \nabla G_T)' (\mathbb{E} G_T G_T')^{-1} (\mathbb{E} \nabla G_T),$$

which is a natural generalization of the **Fisher amount of information**. Here  $\nabla G$  is the  $(p \times p)$ -matrix of derivatives of the elements of  $G$  with respect to those of  $\theta$  and

prime denotes transposition. If  $\mathcal{H} \subset \mathcal{G}$  is a prespecified family of estimating functions, it is said that  $G_T^* \in \mathcal{H}$  is *fixed sample optimal* in  $\mathcal{H}$  if  $\mathcal{E}(G_T^*) - \mathcal{E}(G_T)$  is non-negative definite for all  $G_T \in \mathcal{H}$ ,  $\theta$  and  $\mathbb{P}_\theta$ . Then,  $G_T^*$  is the element of  $\mathcal{H}$  whose dispersion distance from the maximum information estimating function in  $\mathcal{G}$  (often the likelihood score) is least.

A focus on asymptotic properties can be made by confining attention to the subset  $\mathcal{M} \subset \mathcal{G}$  of estimating functions which are martingales (cf. **Martingale**). Here one considers  $T$  ranging over the positive real numbers and for  $\{G_T\} \in \mathcal{M}$  one writes  $\{\langle G \rangle_T\}$  for the *quadratic characteristic*, the predictable increasing process for which  $\{G_T G_T' - \langle G \rangle_T\}$  is a martingale. Also, write  $\{\bar{G}_T\}$  for the predictable process for which  $\{\nabla G_T - \bar{G}_T\}$  is a martingale. Then,  $G_T^* \in \mathcal{M}_1 \subset \mathcal{M}$  is *asymptotically optimal* in  $\mathcal{M}_1$  if  $\bar{\mathcal{E}}(G_T^*) - \bar{\mathcal{E}}(G_T)$  is almost surely non-negative definite for all  $G_T \in \mathcal{M}_1$ ,  $\theta$ ,  $\mathbb{P}_\theta$ , and  $T > 0$ ,

where

$$\bar{\mathcal{E}}(G_T) = \bar{G}_T' \langle G \rangle_T^{-1} \bar{G}_T.$$

Under suitable regularity conditions, asymptotically optimal estimating functions produce estimators for  $\theta$  which are consistent (cf. **Consistent estimator**), asymptotically unbiased (cf. **Unbiased estimator**) and asymptotically normally distributed (cf. **Normal distribution**) with minimum size asymptotic confidence zones (cf. **Confidence estimation**). For further details see [1], [2].

#### References

- [1] GODAMBE, V.P., AND HEYDE, C.C.: 'Quasi-likelihood and optimal estimation', *Internat. Statist. Rev.* **55** (1987), 231–244.
- [2] HEYDE, C.C.: *Quasi-likelihood and its application. A general approach to optimal parameter estimation*, Springer, 1997.
- [3] MCLEISH, D.L., AND SMALL, C.G.: *The theory and applications of statistical inference functions*, Lecture Notes in Statistics. Springer, 1988.

C.C. Heyde

MSC1991: 62F12, 62G20