

Variation in genes related to hepatic lipid metabolism and changes in waist circumference and body weight

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Abstract We analysed single nucleotide polymorphisms (SNPs) tagging the genetic variability of six candidate genes (*ATF6*, *FABP1*, *LPIN2*, *LPIN3*, *MLXIPL* and *MTTP*) involved in the regulation of hepatic lipid metabolism, an important regulatory site of energy balance for associations with body mass index (BMI) and changes in weight and waist circumference. We also investigated effect modification by sex and dietary intake. Data of 6,287 individuals participating in the European prospective investigation into cancer and nutrition were included in the analyses. Data on

weight and waist circumference were followed up for 6.9 ± 2.5 years. Association of 69 tagSNPs with baseline BMI and annual changes in weight as well as waist circumference were investigated using linear regression analysis. Interactions with sex, GI and intake of carbohydrates, fat as well as saturated, monounsaturated and polyunsaturated fatty acids were examined by including multiplicative SNP-covariate terms into the regression model. Neither baseline BMI nor annual weight or waist circumference changes were significantly associated with variation in the selected genes in the entire study population after correction for multiple testing. One SNP (rs1164) in *LPIN2* appeared to be significantly interacting with sex ($p = 0.0003$) and was associated with greater annual

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weight gain in men (56.8 ± 23.7 g/year per allele, $p = 0.02$) than in women (-25.5 ± 19.8 g/year per allele, $p = 0.2$). With respect to gene–nutrient interaction, we could not detect any significant interactions when accounting for multiple testing. Therefore, out of our six candidate genes, *LPIN2* may be considered as a candidate for further studies.

Keywords *LPIN2* · Obesity · Weight gain · Gene–diet interaction

Introduction

During the last decades, the prevalence of obesity has risen (Low et al. 2009) and a strong rise in prevalence of obesity with increasing age is observed (von Ruesten et al. 2011). Factors that influence weight gain during adulthood are incompletely understood, but it is believed that a person's susceptibility to gain weight is influenced by an interplay of many factors including lifestyle factors like physical activity and nutritional habits as well as genetic factors (WHO 2000).

From the physiological view, weight gain results from a positive energy balance, where excessive energy is stored in form of fat in the adipose tissue. The liver is the central site of fatty acid and lipid metabolism and plays a regulatory role in energy homeostasis. Impaired regulation of hepatic lipid metabolism may result in obesity (Guillou et al. 2008).

The activating transcription factor 6 (*ATF6*), liver fatty acid binding protein 1 (*FABP1*), lipin-2 (*LPIN2*), lipin-3 (*LPIN3*), carbohydrate response element binding protein (*MLXIPL*) and microsomal triglyceride transfer protein (*MTTP*) genes were considered to be candidate genes, and their function in relation to hepatic lipid metabolism is described as follows.

- *ATF6* is an endoplasmic reticulum (ER) membrane-bound transcription factor that is activated by ER stress.

ER stress can be caused among other stress factors by metabolic stress, like glucose deprivation (Salminen and Kaamiranta 2010; Zeng et al. 2004). Upon activation, *ATF6* regulates target genes for the unfolded protein response (Jäger et al. 2012) as well as several lipogenic genes (Zeng et al. 2004). It has been shown to antagonise the activity of sterol responsive element binding protein on cholesterol biosynthesis (Zeng et al. 2004).

- *FABP1* serves as an intracellular acceptor of long-chain fatty acids (McArthur et al. 1999) and is involved in the partitioning of fatty acids to specific lipid metabolic pathways (Veerkamp and van Moerkerk 1993; Atshaves et al. 2010). The *FABP1* knockout mouse was shown to be protected against diet-induced obesity (Newberry et al. 2006).
- The lipin proteins act as phosphatidate phosphatase-1 converting phosphatidate to diacylglycerol in triglyceride and phospholipid biosynthesis (Gropler et al. 2009). Lipin 2 mRNA and protein expression are induced by food deprivation (Gropler et al. 2009). Furthermore, lipin-2 and potentially lipin-3 function as transcriptional coactivators for peroxisome proliferator-activated receptor-response elements thereby promoting fatty acid oxidation (Donkor et al. 2009).
- The carbohydrate response element binding protein (*MLXIPL*) regulates transcriptionally key enzymes of glycolysis and lipogenesis in the liver in a glucose-dependent manner (Iizuka and Horikawa 2008; Poupeau and Postic 2011).
- *MTTP* is involved in the assembly and secretion of VLDLs in the liver (Gordon and Jamil 2000), and its expression is regulated by changes in the dietary intake of sugar and fat (Hussain et al. 2011).

Previous genetic studies have revealed associations of SNPs in these candidate genes anthropometric measures, i.e. body mass index (BMI) and fat mass (*ATF6* (Fougeray et al. 2011), *FABP1* (Weickert et al. 2007; Brouillette et al. 2004), *LPIN2* (Aulchenko et al. 2007), *MTTP* (Ledmyr et al. 2002; Bohme et al. 2008; Berthier et al. 2004; Rubin et al. 2006)) as well as with lipid levels [*ATF6* (Meex et al. 2009), *FABP1* (Fisher et al. 2007; Brouillette et al. 2004), *MLXIPL* (Talmud et al. 2009; Kooner et al. 2008; Willer et al. 2008), *MTTP* (Bohme et al. 2008; Berthier et al. 2004; Ledmyr et al. 2002)] and type two diabetes [*ATF6* (Meex et al. 2007; Thameem et al. 2006), *FABP1* (Mansego et al. 2012), *LPIN2* (Aulchenko et al. 2007; Parker et al. 2001)], *MTTP* (Rubin et al. 2006)].

Here, we aimed to study the association of variation in the selected candidate genes with changes in body weight and waist circumference. Since some previously reported associations of our candidate genes were sex specific

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(Bohme et al. 2008; Fisher et al. 2007), we further tested for interactions with sex. In addition, a major aim was to investigate possible interaction with dietary factors. As dietary effect modifiers, a quantitative (carbohydrate intake) and a qualitative measure [glycaemic index (GI)] of carbohydrate intake were chosen, since lipogenesis and lipolysis are mainly regulated by the intake of carbohydrates (Flatt 1995). Furthermore, interaction analyses were extended to dietary fat intake, since some of the genes have been shown to be sensitive to dietary fat with regard to regulation (Hussain et al. 2011) and interactions (Robitaille et al. 2004), respectively. Interaction and association analyses of the six selected candidate genes and baseline BMI, changes in body weight and waist circumference were carried out in a large prospective cohort from five European populations.

Methods

Study population

The study is based on several existing cohorts of the European prospective investigation into cancer and nutrition (EPIC) (Riboli and Kaaks 1997; Riboli et al. 2002) that participated in the DiOGenes project (www.diogenes-eu.org; Saris and Harper 2005) with the aim to study dietary components that influence weight gain in combination with genetic and behavioural factors. The EPIC study has been described in detail earlier (Riboli and Kaaks 1997; Riboli et al. 2002). Data from the cohorts in eight regions in five different European countries were included into the DiOGenes study: Copenhagen and Aarhus (Denmark), Potsdam (Germany), Florence (Italy), Bilthoven (combined EPIC centre from subcentres located in Doetinchem, Amsterdam and Maastricht; The Netherlands) and Norfolk (United Kingdom). Together, these cohorts comprise 146,543 individuals. Participants were considered eligible for the current analyses if they were younger than 60 years at baseline and younger than 65 years at follow-up, had blood samples available and had baseline information on diet, weight and height and follow-up information on weight. Inclusion criteria were also stable smoking habits, no diagnosis of cancer, cardiovascular disease and diabetes at baseline or during follow-up, and an annual weight change not higher than 5 kg/year. A total of 50,293 individuals were eligible. A subcohort was drawn randomly from the total eligible cohort as described previously (Du et al. 2011; Vimalaswaran et al. 2012; Fisher et al. 2012). The final subcohort including 7,061 individuals was used for the current analysis.

Anthropometrics, dietary intake and other measurements

Baseline weight, height and waist circumference were assessed by trained staff according to a predefined protocol (Rinaldi et al. 2006). During follow-up, participants in Doetinchem (Netherlands) and Norfolk (UK) were re-examined using the same protocol while in all other centres self-reported data was collected.

Validated country-specific food frequency questionnaires were completed by participants during the baseline assessment (Riboli and Kaaks 1997). Energy and nutrient intakes were calculated using country-specific food composition tables (Riboli and Kaaks 1997). Dietary GI was calculated using glucose as the reference food according to a methodology described elsewhere (van Bakel et al. 2009; Du et al. 2008). Data on lifestyle were collected using self-administered questionnaires at baseline and follow-up.

Selection and genotyping of candidate genes and SNPs

We have selected candidate genes based on their involvement in or regulation of lipid metabolism, expression in the liver, a potential nutrient-sensitive effect and previous reported genetic association with weight or body composition in human (Fougeray et al. 2011; Weickert et al. 2007; Aulchenko et al. 2007; Berthier et al. 2004; Bohme et al. 2008; Brouillette et al. 2004; Ledmyr et al. 2002; Rubin et al. 2006) or animal studies (He et al. 2009).

Tagging SNPs in the selected candidate genes were chosen based on data from the International HapMap project (The International HapMap Consortium 2005) using pairwise tagging and R^2 -threshold of 0.8 within the Tagger software (de Bakker et al. 2005) implemented in Haploview V3.3 (Barrett et al. 2005) to cover common genetic variation (MAF >0.05). SNPs were forced to be included when they had potential or proven functional implications or had been extensively studied in the past. High throughput genotyping of a total of 86 SNPs in the *ATF6*, *FABP1*, *LPIN2*, *LPIN3*, *MLXIPL* and *MTTP* genes was performed using an Illumina Beadstation Genotyping System along with other SNPs that served different genetic analyses within the DiOGenes project. The initial stage of quality control stratified by study centre was carried out jointly for all SNPs. Genotyping (i.e. individuals) passed the initial stage of quality control if their discordance rate of duplicate samples was $\leq 3\%$ and at least 95 % of SNPs per sample were genotyped. SNPs were excluded unless $\geq 95\%$ of samples were genotyped. Furthermore, SNPs were excluded for a particular country when the corresponding p value for Hardy–Weinberg equilibrium was below 0.001. Genotype data of 6,566 (93.0 %) individuals passed quality control. From the 86 SNPs that were

selected for this study, seven SNPs did not pass the initial stage of quality control (Online Resource 1-Table S1). Another ten SNPs had to be further excluded from the analyses because they only passed quality control in some, but not all, EPIC-centres (Online Resource 1-Table S1). Eventually, a total of 69 SNPs were included in the current analyses.

Linkage disequilibrium (LD)

LD between genotyped variants within the study population was calculated using Haploview V3.3 (Barrett et al. 2005). LD of genotyped LPIN2 variants and rs3745012 variant reported in literature (Aulchenko et al. 2007) was established using SNAP and the 1000GenomesPilot1 dataset (Johnson et al. 2008; Abecasis et al. 2010).

Statistical analyses

All analyses were carried out separately for each country, and combined effect estimates with corresponding *p* values were calculated using random effects meta-analysis. Due to differences in data collection of anthropometric measures and length of follow-up, data of the Dutch cohort was split into two study centres resulting in the final allocation of six study centres, namely Copenhagen and Århus (Denmark, DK-CopAa), Potsdam (Germany, GER-Pot), Florence (Italy, IT-Flo), Doetinchem (Netherlands, NL-Doe), Amsterdam and Maastricht (Netherlands, NL-AmMa), and Norfolk/England (United Kingdom, UK-Nor).

Linear regression models were applied assuming an additive genetic model. Analyses of baseline BMI were adjusted for baseline values of age, sex, height, smoking status, energy intake, physical activity and recruitment centre. Annual weight change was calculated by subtracting baseline body weight from follow-up body weight and dividing the difference in grams by the duration of follow-up in years (*y*). Analyses of annual weight change were adjusted for baseline values of age, sex, height, smoking status, weight, energy intake, physical activity and recruitment centre and follow-up time. Analyses of annual changes in waist circumference were adjusted for baseline age, sex, height, smoking status, waist circumference, energy intake, physical activity and recruitment centre and follow-up time. Interaction with sex was assessed by including the multiplicative interaction term of sex (indicator variable; 0/1) and each SNP in the model. When the interaction terms were statistically significant, sex-stratified analyses were carried out.

Interaction with carbohydrate intake was assessed by a multiplicative interaction term of the percentage of energy derived from carbohydrates and each SNP in a linear model

that was adjusted for percentages of energy derived from carbohydrates, protein and alcohol (all scaled by units corresponding to 5 % of total energy intake), the ratio of monounsaturated to saturated and the ratio of polyunsaturated to saturated fatty acids in addition to the basic set of covariates described above. Interactions with total fat intake were modelled in the same way with additional adjustment of energy derived from fat, protein, alcohol as well as the fatty acid ratios in addition to the basic set of covariates. Fat-subtype-based interaction analyses of saturated, monounsaturated and polyunsaturated fat, respectively, were modelled like total fat but additionally adjusted for the complementary fat intake, i.e., e.g. when saturated fat is used, it is also adjusted for the difference between total fat intake and saturated fat intake. Interactions between SNPs and GI were tested by adding the multiplicative interaction term of GI and each SNP as well as the GI main effect to the basic linear regression model for waist and weight gain. All statistical analyses were carried using Stata 9.2 for Windows (StataCorp LP, Texas, USA; www.stata.com). The forest plot was drawn using R version 2.15.1 (R Development Core Team 2010) and the 'meta' package (Schwarzer 2010).

The Bonferroni method was used with respect to each analysis case to correct for the multiple testing of 69 SNPs. Therefore, a *p* value lower than 0.0007 (0.05/69) was considered to be statistically significant. Power calculations were performed using the QUANTO software Version 1.2.4 (May 2009; Gauderman and Morrison 2006). The minimum detectable SNP effects with respect to different minor allele frequencies (MAFs) assuming a 80 % power rate and for a given significance level of $\alpha = 0.05$ were estimated to range from 40 g/year (MAF = 50 %) to 88 g/year (MAF = 5 %) for weight gain, from 0.18 kg/m² (MAF = 50 %) to 0.40 kg/m² (MAF = 5 %) for BMI and from 0.06 cm/y (MAF = 50 %) to 0.14 cm/y (MAF = 5 %) for waist circumference.

Results

Study population characteristics

The final study population included 6,287 individuals and consisted of 46 % men. The mean \pm SD baseline age and BMI were 48.0 \pm 7.2 year and 25.4 \pm 3.7 kg/m², respectively. Individuals had a mean follow-up time of 6.9 \pm 2.5 year, where they gained on average 235 \pm 784 g body weight per year (g/year) and 0.7 \pm 1.2 cm waist circumference per year (cm/year), respectively. Age and follow-up time were similar between men and women (Table 1). Annual changes in weight and waist circumference were slightly higher in women than in men (Table 1). A list of

Table 1 General and anthropometric characteristics as well as dietary intake of the study population

	Total 6,287	Men 2,891 (46 %)	Women 3,396 (54 %)
<i>N</i>			
Age (year)	48.0 ± 7.2	48.4 ± 7.0	47.8 ± 7.3
Follow-up time (year)	6.9 ± 2.5	6.6 ± 2.4	7.1 ± 2.5
Weight at baseline (kg)	73.2 ± 13.6	81.5 ± 11.3	66.1 ± 11.1
Annual weight change (g/y)	235 ± 784	232 ± 745	239 ± 815
Baseline waist circumference (cm)	84.9 ± 12.2	92.9 ± 9.3	78.3 ± 10.1
Annual waist change (cm/y)	0.7 ± 1.2	0.5 ± 1.1	0.8 ± 1.2
BMI at baseline (kg/m ²)	25.4 ± 3.7	26.1 ± 3.2	24.8 ± 3.9
BMI at follow-up (kg/m ²)	25.9 ± 3.9	26.5 ± 3.4	25.4 ± 4.2
Carbohydrate intake (% of total energy)	44.5 ± 6.5	43.6 ± 6.5	45.3 ± 6.5
Protein intake (% of total energy)	16.2 ± 2.7	15.8 ± 2.6	16.5 ± 2.8
Fat intake (% of total energy)	34.7 ± 5.4	34.7 ± 5.4	34.8 ± 5.7
Glycaemic index	56.5 ± 4.1	57.3 ± 4.1	55.7 ± 3.9

Values present *N* (%) or Mean ± SD

successfully genotyped SNPs and their MAFs can be found in Online Resource 1-Table S2.

In all genes, a number of nominally significant *p* values were observed, mainly in the interaction analyses and to a lesser extent in the main effects analyses (Table 2; detailed results in Online Resource 1-Table S3-11). Nominally significant main effects of *FABP1* rs894194 and *LPIN2* rs10438875 on baseline BMI as well as of *MLXIPL* rs6967028 and two SNPs in *MTTP* (rs1057613, rs3816873) on annual weight change were detected (Table 2).

Only the interaction of the *LPIN2* SNP rs1164 with sex in association with weight change was significant after adjustment for multiple testing ($\beta = -111.5 \pm 30.9$ g/year, $p = 0.0003$). For this SNP, sex-stratified analyses were carried out. As shown in Fig. 1, the presence of the minor allele of *LPIN2* rs1164 is estimated to lead to a weight gain of 56.8 ± 23.7 g/year per minor allele in men ($p = 0.02$) whereas in women, a small weight loss was observed ($\beta = -25.5 \pm 19.8$ g/year per minor allele, $p = 0.2$). The study-specific effect estimates in both men and women varied between studies (Fig. 1), but were consistently larger in men than in women throughout.

All tested gene–diet interactions were statistically not significant after correction for multiple testing. The lowest *p* values for gene–diet interactions were observed for firstly, an interaction of rs2919871 in *FABP1* and intake of

saturated fat in association with weight change ($\beta = 81.4 \pm 24.9$ g/year, $p = 0.001$) and secondly, an interaction of SNP rs3934427 in *LPIN2* with intake of carbohydrates in association with weight change ($\beta = 44.1 \pm 14.3$ g/year, $p = 0.002$). Both SNPs showed further nominally significant gene–diet interactions; the *FABP1* rs2919871 variant with total fat ($p = 0.04$) again in association with weight gain. The *LPIN2* rs3934427 SNP presented the same interaction with dietary carbohydrates in association with waist change ($p = 0.01$) as well as an interaction with the intake of fat in association with waist change ($p = 0.05$).

Discussion

For 69 SNPs located in the six candidate genes (*ATF6*, *FABP1*, *LPIN2*, *LPIN3*, *MLXIPL* and *MTTP*), no significant associations with baseline BMI, annual change in waist circumference or body weight as well as no interactions with dietary factors (carbohydrate intake, GI, total, saturated, monounsaturated and polyunsaturated fat) were detected following correction for multiple testing. Altogether, only the interaction of rs1164 in *LPIN2* with sex on annual weight change can be considered statistically significant taking multiple testing into account.

Strength and limitations

Our study investigated the interaction of common genetic variation in proteins involved in lipid uptake and metabolism of the liver and dietary factors (i.e. carbohydrates, fat and GI) in relation to weight development in five European populations. Large prospective cohort studies like this which offer detailed dietary records of their participants are still rare. Therefore, this study seems to be unique to investigate gene–nutrient interactions, but there are also important limitations. First, the study includes the possibility of some degree of under- and misreporting of weight and waist changes as well as of dietary intake. For all, but two centres, body weight and waist circumference were measured in the study centre by trained staff at baseline, while the follow-up measurement was self-reported. In the remaining two centres (UK-Nor, NL-Doe), body weight and waist circumference were also taken in the study centre during follow-up. Anthropometric measurements are usually underreported by self-reports (Park et al. 2011). Therefore, our data on weight gain based on measured baseline values and self-reported follow-up values would underestimate the true weight gain and could therefore attenuate the observed effect. Furthermore, the degree of underreporting might vary according to gender (Park et al. 2011), which could create an artificial interaction with sex.

Table 2 Nominally significant SNP main effect, SNP x diet interaction, and SNP x sex interaction of genetic variants in selected candidate genes for baseline BMI, and annual changes in waist circumference and body weight

Gene	SNP	SNP type ^d	Outcome	Estimate ^c	β (SE) ^{a,b}	<i>p</i> value ^a
ATF6	rs10918092	Intron	Waist change	SNP × carb	0.03 (0.02)	0.03
ATF6	rs12130299	Upstream	Weight change	SNP × mufa	−60.7 (25.7)	0.02
			Weight change	SNP × sfa	−55.8 (25.0)	0.03
			Weight change	SNP × fat	−27.1 (12.2)	0.03
ATF6	rs2499854	3' UTR	Waist change	SNP × pufa	0.13 (0.06)	0.03
ATF6	rs4657101	Intron	Waist change	SNP × sex	0.09 (0.04)	0.03
			Waist change	SNP × GI	−0.01 (0.01)	0.04
FABP1	rs2919871	Intron, nc	Waist change	SNP × sex	−0.08 (0.04)	0.04
			Weight change	SNP × sfa	81.4 (24.9)	0.001
			Weight change	SNP × fat	25.1 (12.3)	0.04
FABP1	rs2970902	Upstream	Weight change	SNP × sfa	69.1 (26.4)	0.009
			Weight change	SNP × fat	29.8 (11.5)	0.01
FABP1	rs2970903	Upstream	Weight change	SNP × GI	−11.1 (5.5)	0.04
FABP1	rs894194	Downstream	Baseline BMI	SNP	−0.16 (0.07)	0.03
LPIN2	rs10438875	Upstream	Baseline BMI	SNP	0.17 (0.09)	0.05
LPIN2	rs1164	3' UTR	Weight change	SNP × sex	−111.5 (30.9)	0.0003
LPIN2	rs11661932	Intron	Weight change	SNP × carb	−47.3 (22.5)	0.04
LPIN2	rs11665524	Intron	Waist change	SNP × pufa	−0.15 (0.08)	0.05
LPIN2	rs3934427	Intron, nc	Waist change	SNP × carb	0.05 (0.02)	0.01
			Waist change	SNP × fat	−0.07 (0.04)	0.05
			Weight change	SNP × carb	44.1 (14.3)	0.002
LPIN2	rs4797092	Intron	Weight change	SNP × GI	17.0 (7.2)	0.02
LPIN2	rs6506041	Upstream	Weight change	SNP × carb	−24.4 (10.0)	0.01
			Weight change	SNP × mufa	50.4 (25.4)	0.05
LPIN3	rs12625565	Missense (Q679H)	Waist change	SNP × sfa	0.09 (0.04)	0.02
LPIN3	rs16985665	Upstream	Weight change	SNP × pufa	102.7 (47.1)	0.03
LPIN3	rs4810312	Upstream	Waist change	SNP × mufa	0.07 (0.03)	0.04
			Weight change	SNP × mufa	65.5 (28.0)	0.02
			Weight change	SNP × fat	33.4 (15.0)	0.03
			Weight change	SNP × carb	−20.7 (9.4)	0.03
LPIN3	rs6029636	Upstream	Weight change	SNP × mufa	56.1 (26.4)	0.03
MLXIPL	rs11760752	Intron	Weight change	SNP × sex	65.5 (31.2)	0.04
MLXIPL	rs6967028	Upstream	Weight change	SNP	−50.5 (25.3)	0.05
MTTP	rs1057613	3' UTR	Weight change	SNP	28.8 (14.4)	0.04
MTTP	rs3816873	Missense (I128T)	Weight change	SNP	38.2 (16.5)	0.02

^a Overall estimate (β), standard error (SE) and *p* value from meta-analysis

^b For waist change in cm/year per minor allele, for weight change in g/year per minor allele, for baseline BMI in kg/m² per minor allele

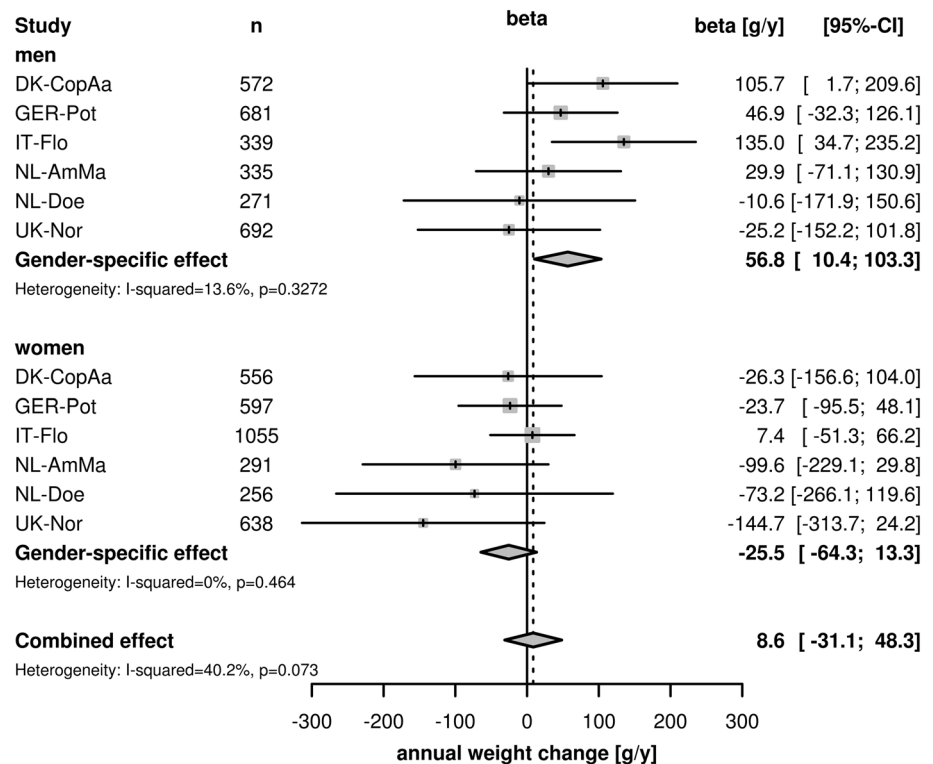
^c *carb* increment of 5 % energy from carbohydrates, *fat* increment of 5 % energy from fat, *sfa* increment of 5 % energy from saturated fatty acids, *mufa* increment of 5 % energy from monounsaturated fatty acids, *pufa* increment of 5 % energy from polysaturated fatty acids, *GI* glycaemic index

^d *nc* non-coding transcript

In the centre-specific analysis, the overall significant *LPIN2* rs1164–sex interaction was nominally significant only in two centres with self-reports (IT-Flo and NL-AmMa), but in none with measured follow-up data (UK-Nor, NL-Doe). However, the estimate for the interaction effect in the British centre with measured follow-up data

(UK-Nor -136.8 ± 107.1 g/year, $p = 0.20$) is of the same magnitude as in the two centres with self-reported follow-up data and significant interaction effect (It-Flo: -124.5 ± 59.9 g/year, $p = 0.04$; NL-AmMa: -175.7 ± 82.6 g/year, $p = 0.03$) suggesting that differences in data collection had no major influence on the observed effect.

Fig. 1 Forest plot of the random effect meta-analysis of beta estimates for the association of *LPIN2* rs1164 with annual weight change (g/y) stratified by sex



Second, while the genotype is a stable exposure, dietary intake may change over time, especially as a result of weight changes. We have analysed dietary intake only once at baseline and are therefore not able to take changes in dietary intake into account. Third, the selection of candidate genes was based on prior knowledge and financial resources and represents a small subset of the 'key players' within hepatic fat metabolism. We therefore have to acknowledge that, since the regulation of hepatic lipid metabolism is much more complex, other relevant genetic variants among these pathways might have been missed. Fourth, although our study is large, we cannot exclude the possibility that the lack of significant findings is due to a lack of power. Singular genetic effects in polygenic traits like weight gain are supposed to be small, and the power to detect an interaction with dietary intake is further limited by the measurement error in estimating dietary intake. Moreover, the Bonferroni method was used to correct for multiple testing. The Bonferroni method is frequently criticised for being relative strict and producing false negative results (Perneger 1998). However, since in most cases p values were just below the nominally significance threshold of 0.05, less conservative methods would most likely not have changed the overall result. However, since our study is based on meta-analyses of data from five different countries, a common type I error (rejection of null hypothesis that is actually true) seemed not very likely to have occurred. For that reason, nominally significant

results are discussed in the following section provided they have been reported in the literature previously.

Findings

The significant SNP rs1164–sex interaction resulted from a larger annual weight gain in men than in women. In the past, the rs3745012 *LPIN2* variant has been described to be associated with BMI, waist circumference and other anthropometric measurements in a Dutch study (Aulchenko et al. 2007). However, both variants are not in LD ($R^2 = 0.05$). All other variants in LD ($R^2 = 0.77$) with rs3745012 were not associated with any of the analysed traits. Therefore, we feel confident to exclude rs3745012 as a potential variant underlying the observed association.

The detected interaction effect led to weight gain in men, whereas in women, a small weight loss was observed. A recent GWAS meta-analysis of sex-specific genetic effects for height, weight, body mass index, waist circumference, hip circumference and waist-to-hip ratio did not find a sex-specific effect in *LPIN2* (Randall et al. 2013). Therefore, our finding needs to be treated very cautiously.

Lipin 2 catalyses as phosphatidate phosphatase a key step of lipid biosynthesis and regulates fatty acid oxidation as transcriptional coactivator of PPAR α as well as PPAR γ coactivator 1a (Reue and Dwyer 2009). Rare mutations in *LPIN2* cause the auto-inflammatory disorder Majeed syndrome suggesting a role in regulation of the inflammatory

response for *LPIN2* (Ferguson et al. 2005; Reue and Donkor 2007). It seems to be possible to influence body weight regulation via these mechanisms; however, as the causal variant is unknown, it is impossible to elucidate a biological mechanism for the observed association.

Two SNPs (rs1057613, rs3816873 (I128T), $R^2 = 0.41$) in *MTTP* showed a nominally significant association with annual changes in body weight. *MTTP* plays a central role in lipoprotein assembly (Hussain et al. 2003). The T128 allele of rs3816873 has been reported to lead to reduced function of *MTTP* (Ledmyr et al. 2006). The SNP is in complete LD with several promoter variants of the gene that have been shown to alter *MTTP* expression (Rubin et al. 2008; Ledmyr et al. 2004; Aminoff et al. 2010; Karpe et al. 1998). The rs3816873 polymorphism was reported to be associated with BMI and other anthropometric measures, but results were inconsistent (Rubin et al. 2006; Berthier et al. 2004; St-Pierre et al. 2002; Ledmyr et al. 2002). The so far largest study of *MTTP* variants on BMI in 7582 KORA participants found no association of BMI and waist circumference with the rs3816873 polymorphism, but an association with the H297Q (rs2306985) polymorphism in women (Bohme et al. 2008). In the present study, both variants were not associated with BMI, but the T128 allele was associated with increased weight gain ($\beta = 38.2 \pm 16.5$ g/year per minor allele, $p = 0.02$). Since it has been shown that the effect of *MTTP* on BMI was sex specific (Rubin et al. 2006; Berthier et al. 2004), sex-stratified analyses were carried out. Although no statistically significant interaction between sex and rs3816873 was detected, a nominally significant association of rs3816873 with increased annual weight gain was detected in men only (men: $\beta = 57.9 \pm 19.9$ g/year per minor allele, $p = 0.004$; women: $\beta = 17.8 \pm 17.6$ g/year per minor allele, $p = 0.31$). The so far contrasting findings might be explained by different nutritional influences. Here, we tested for interactions with the intake of carbohydrates, GI, fat and fat subtypes but did not detect any influence of those dietary factors on the association of SNPs in *MTTP* and weight change.

There have been previous reports on an association with BMI and waist circumference for the *FABPI* T94A variant (rs2241883) (Brouillette et al. 2004; Weickert et al. 2007). In the present study, we did not observe an association with BMI or changes in body weight or waist circumference of this variant.

The DiOGenes project also included an intervention study that investigated weight regain after an initial weight loss with dietary intervention differing in GI and protein intake (Larsen et al. 2012). This intervention study analysed the same genes as our study for main effects of weight, waist and fat mass regain after weight loss and interaction with protein content and GI of the dietary

intervention (Larsen et al. 2012). As in our study, no significant main or SNP–diet interaction effects were identified when correction for multiple testing was applied.

In summary, we found a significant interaction of rs1164 in *LPIN2* with sex in association with weight gain, leading to higher weight gain in of about 90 g/year in men than in women (men: 56.8 ± 23.7 g/year per minor allele; women: -25.5 ± 19.8 g/year per minor allele) than in women. We observed no significant association between SNPs located in *ATF6*, *FABP1*, *LPIN3*, *MTTP* and *MLXIPL* and annual changes in weight and waist circumference after correction for multiple testing. Furthermore, we found no convincing evidence for interactions of these candidate genes and either intake of carbohydrates, GI, total, saturated, mono-unsaturated and polyunsaturated fat. From our six selected candidate genes, *LPIN2* is the one that could be followed up in further studies with a special focus on sex-specific effects.

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Ethical standards The Sponsors of the DiOGenes Project are listed on the website of the project (<http://www.diogenes-eu.org/>). The funders had no role in study design, data collection and analysis, decision to publish, or preparation of the manuscript. Karina Meidtnier, Eva Fisher, Lars Ängquist, Claus Holst, Karani S Vimalaswaran, Jolanda MA Boer, Jytte Halkjær, Giovanna Masala, Jane N Østergaard, Lotte M Mortensen, Daphne L. van der A, Anne Tjønneland, Domenico Palli, Kim Overvad, Nicholas J Wareham, Ruth JF Loos, Thorkild IA Sørensen and Heiner Boeing declare that they have no conflict of interest. All procedures followed were in accordance with the ethical standards of the responsible committee on human experimentation (institutional and national) and with the Helsinki Declaration of 1975, as revised in 2000. Informed consent was obtained from all patients for being included in the study.

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