

Erratum to: The Asymptotic Distribution of Self-Normalized Triangular Arrays

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Abstract We correct and clarify some ambiguous statements in D. M. Mason (2005): The asymptotic distribution of self-normalized triangular arrays. *J. Theoret. Probab.*, **18**, 853–870.

Corrections and Clarifications of Mason (2005)

This note has two purposes. First is to correct some statements in the **Introduction and Statements of Results** of [4], and second is to provide the result given in Proposition [A] below, which clarifies a claim at the end of the proof of Theorem 2.

Our corrections are needed since it is not clear that (1.9) always implies (1.2). They are the following:

- (i) On page 855, line 10, change the “further shows” to “further shows that under the setup of Proposition [A] in this note” and on line 13 change “(1.4)” to “(1.4) nondegenerate”.
- (ii) On page 855, line 14, replace “or equivalently (1.9) holds” with “ $P(V > 0) = 1$ ”.
- (iii) On page 856, line 3, replace “Actually” with “Actually under the setup of Proposition [A] in this note”.

We remark in passing that the statements in [4] about triangular arrays of the form $X_{1,n}, \dots, X_{n,n}$, $n \geq 1$, are equally valid for triangular arrays of the form $X_{1,n_k}, \dots, X_{n_k,n_k}$, $k \geq 1$, where $\{n_k\}_{k \geq 1}$ is an infinite subsequence of the positive integers. Also we point out that everywhere *triangular array of infinitesimal independent*

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random variables should be changed to *infinitesimal triangular array of independent random variables*, as stated in Proposition [A]. The word *infinitesimal* was everywhere put in the wrong place.

The following Proposition [A] and Remark 1 justify the claim towards the end of the proof of Theorem 2 on page 868 that says, “This means that every subsequential distributional limit random variable T must be of the form (1.10).” They should have been included in an appendix in the original paper.

Proposition [A] *Let $\{n_k\}_{k \geq 1}$ be an infinite subsequence of the positive integers and $X_{1,n_k}, \dots, X_{n_k,n_k}$, $k \geq 1$, be an infinitesimal triangular array of independent random variables such that for each $k \geq 1$, $X_{1,n_k}, \dots, X_{n_k,n_k}$ are i.i.d. X_{1,n_k} . Assume that for a necessarily infinitely divisible random variable U ,*

$$\sum_{i=1}^{n_k} X_{i,n_k} \rightarrow_d U, \text{ as } k \rightarrow \infty. \quad (1.1)$$

Then

$$\left(\sum_{i=1}^{n_k} X_{i,n_k}, \sum_{i=1}^{n_k} X_{i,n_k}^2 \right) \rightarrow_d (U, V), \text{ as } k \rightarrow \infty, \quad (1.2)$$

where the two dimensional infinitely divisible random vector (U, V) in (1.2) has the representation:

$$(U, V) =_d (b + W + \tau Z, S + \tau^2), \quad (1.3)$$

with b and $\tau \geq 0$ being suitable constants,

$$\begin{aligned} W = & \int_0^1 \varphi_1(s) dN_1(s) + \int_1^\infty \varphi_1(s) d\{N_1(s) - s\} \\ & - \int_0^1 \varphi_2(s) dN_2(s) - \int_1^\infty \varphi_2(s) d\{N_2(s) - s\} \end{aligned} \quad (1.4)$$

and

$$S = \int_0^\infty \varphi_1^2(s) dN_1(s) + \int_0^\infty \varphi_2^2(s) dN_2(s), \quad (1.5)$$

with N_1 and N_2 being independent right continuous Poisson processes on $[0, \infty)$ with rate 1, Z being a standard normal random variable independent of N_1 and N_2 , and φ_1 and φ_2 being two left continuous, nonincreasing, nonnegative functions defined on $(0, \infty)$ satisfying for all $\delta > 0$,

$$\int_\delta^\infty \varphi_i^2(s) ds < \infty \text{ for } i = 1, 2. \quad (1.6)$$

Proof The proof that (1.1) implies (1.2) follows along very similar lines to that of Lemma 4 in [2]. To relieve the notational burden, in the following we shall write $n = n_k$. By parts (ii) and (iii) of Theorem 4.7 on page 61 of [1], the distributional convergence (1.1) implies that there exists a Lévy measure μ such that for every $\delta > 0$ such that $\mu\{-\delta, \delta\} = 0$,

$$\begin{aligned} w - \lim_{n \rightarrow \infty} \sum_{i=1}^n n \mathcal{L}(X_{i,n}) \mid (|x| > \delta) \\ = w - \lim_{n \rightarrow \infty} n \mathcal{L}(X_{1,n}) \mid (|x| > \delta) = \mu \mid (|x| > \delta); \end{aligned} \quad (1.7)$$

and for some a_δ

$$\lim_{n \rightarrow \infty} E S_{n,\delta} = \lim_{n \rightarrow \infty} (n E X_{1,n,\delta}) = a_\delta, \quad (1.8)$$

where

$$S_{n,\delta} := \sum_{i=1}^n X_{i,n} 1\{|X_{i,n}| \leq \delta\} =: \sum_{i=1}^n X_{i,n,\delta}. \quad (1.9)$$

(Note that we use here the notation of [1].) Now by part (i) of the same theorem, (1.1) also implies that for some $0 \leq \sigma^2 < \infty$,

$$\lim_{\delta \searrow 0} \left\{ \begin{array}{c} \limsup_{n \rightarrow \infty} \\ \liminf_{n \rightarrow \infty} \end{array} \right\} \sum_{i=1}^n E (X_{i,n,\delta} - E X_{i,n,\delta})^2 = \sigma^2. \quad (1.10)$$

Notice that

$$\sum_{i=1}^n E (X_{i,n,\delta} - E X_{i,n,\delta})^2 = n E X_{1,n,\delta}^2 - n^{-1} (n E X_{1,n,\delta})^2. \quad (1.11)$$

Further by (1.8) for every $\delta > 0$ such that $\mu\{-\delta, \delta\} = 0$, $n^{-1} (n E X_{1,n,\delta})^2 \rightarrow 0$, which by (1.10) and (1.11), implies

$$\lim_{\delta \searrow 0} \left\{ \begin{array}{c} \limsup_{n \rightarrow \infty} \\ \liminf_{n \rightarrow \infty} \end{array} \right\} E V_{n,\delta} = \sigma^2, \quad (1.12)$$

where

$$V_{n,\delta} := \sum_{i=1}^n X_{i,n}^2 1\{|X_{i,n}| \leq \delta\} = \sum_{i=1}^n X_{i,n,\delta}^2, \quad (1.13)$$

with

$$EV_{n,\delta} = nEX_{1,n,\delta}^2 = \int_{|x| \leq \delta} nx^2 d\mathcal{L}(X_{1,n}). \quad (1.14)$$

Now let $\delta_m, m \geq 1$, be a sequence of constants converging to zero such that $0 < \delta_{m+1} < \delta_m < \delta_0 = \delta$, and $\mu\{-\delta_m, \delta_m\} = 0, m \geq 0$. Then for each $m \geq 1$, by (1.7) and $\mu\{-\delta_m, \delta_m\} = \mu\{-\delta, \delta\} = 0$,

$$\begin{aligned} & \liminf_{n \rightarrow \infty} \int_{|x| \leq \delta_m} nx^2 d\mathcal{L}(X_{1,n}) + \int_{\delta_m < |x| \leq \delta} x^2 d\mu(x) \\ &= \liminf_{n \rightarrow \infty} \int_{|x| \leq \delta_m} nx^2 d\mathcal{L}(X_{1,n}) + \lim_{n \rightarrow \infty} \int_{\delta_m < |x| \leq \delta} nx^2 d\mathcal{L}(X_{1,n}) \\ &= \liminf_{n \rightarrow \infty} \int_{|x| \leq \delta} nx^2 d\mathcal{L}(X_{1,n}) \leq \limsup_{n \rightarrow \infty} \int_{|x| \leq \delta} nx^2 d\mathcal{L}(X_{1,n}) \\ &= \limsup_{n \rightarrow \infty} \int_{|x| \leq \delta_m} nx^2 d\mathcal{L}(X_{1,n}) + \lim_{n \rightarrow \infty} \int_{\delta_m < |x| \leq \delta} nx^2 d\mathcal{L}(X_{1,n}) \\ &= \limsup_{n \rightarrow \infty} \int_{|x| \leq \delta_m} nx^2 d\mathcal{L}(X_{1,n}) + \int_{\delta_m < |x| \leq \delta} x^2 d\mu(x). \end{aligned}$$

Now by letting $m \rightarrow \infty$, we see by (1.12) that

$$\lim_{n \rightarrow \infty} \int_{|x| \leq \delta} nx^2 d\mathcal{L}(X_{1,n}) = \sigma^2 + \int_{0 < |x| \leq \delta} x^2 d\mu(x) =: b_\delta. \quad (1.15)$$

Moreover, we get from (1.15) that for every $k > 2$,

$$\lim_{\delta \searrow 0} \limsup_{n \rightarrow \infty} \int_{|x| \leq \delta} n|x|^k d\mathcal{L}(X_{1,n}) = 0. \quad (1.16)$$

We now proceed as in the proof of Lemma 4 in [2]. We see using (1.16) that for any $\alpha, \beta \in \mathbb{R}$,

$$\lim_{\delta \searrow 0} \left\{ \limsup_{n \rightarrow \infty} \int_{|x| \leq \delta} E(\alpha(S_{n,\delta} - ES_{n,\delta}) + \beta(V_{n,\delta} - EV_{n,\delta}))^2 \right\} = \alpha^2 \sigma^2. \quad (1.17)$$

Write $\rho = \mu \circ T^{-1}$, where $T(x) = (x, x^2)$. Clearly by (1.7) for every $\delta > 0$ such that $\mu\{-\delta, \delta\} = 0$

$$w - \lim_{n \rightarrow \infty} n\mathcal{L}(X_{1,n}, X_{1,n}^2) |(\|x\| > \delta) = \rho |(\|x\| > \delta). \quad (1.18)$$

Furthermore, by (1.8) and (1.15) we have for every $\delta > 0$ such that $\mu\{-\delta, \delta\} = 0$

$$(ES_{n,\delta}, EV_{n,\delta}) \rightarrow (a_\delta, b_\delta).$$

Thus by the central limit theorem in \mathbb{R}^2 on pp. 67–68 of [1] and arguing just as in [2] we get that (1.2) holds with (U, V) having characteristic function $E \exp(sU + tV) =$

$$\exp \left\{ -\frac{\sigma^2 s^2}{2} + i(a_\delta s + \sigma^2 t) + \int (\exp(i(su + tu^2)) - 1 - i s u 1_{\{|u| \leq \delta\}}) d\mu(u) \right\}, \quad (1.19)$$

for any $\delta > 0$ such that $\mu\{-\delta, \delta\} = 0$. It can be shown using Proposition 5.7 in [3] that a pair of random variables (U, V) with this characteristic function has the distributional representation (1.3) where $\tau^2 = \sigma^2$ and b is a suitable constant. It is shown there how φ_1 and φ_2 are defined via the Lévy measure μ . \square

Remark 1 We note that if X is in the centered Feller class with a_n an appropriate sequence of norming constants and X_1, X_2, \dots , are i.i.d. X , then for every subsequence of $\{n\}$ there exists a further subsequence $\{n_k\}$ such that the triangular array $X_{i,n_k} = X_i/a_{n_k}$, $1 \leq i \leq n_k$, $k \geq 1$, satisfies (1.1), with U nondegenerate, and thus (1.2) and (1.19) hold, as was pointed out in [2]. Also we mention that it can be inferred using the Theorem in [5] that necessarily “ $P(V > 0) = 1$ ”.

Remark 2 A special case of the Proposition 1 implies that for any triangular array $X_{1,n_k}, \dots, X_{n_k,n_k}$, $k \geq 1$, satisfying its assumptions, and

$$\sum_{i=1}^{n_k} X_{i,n_k} \rightarrow_d N(0, \sigma^2), \quad \text{as } k \rightarrow \infty, \quad (1.20)$$

then

$$\sum_{i=1}^{n_k} X_{i,n_k}^2 \rightarrow_P \sigma^2, \quad \text{as } k \rightarrow \infty. \quad (1.21)$$

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