

Diabetes mellitus and risk of endometrial cancer: a meta-analysis

E. Friberg · N. Orsini · C. S. Mantzoros · A. Wolk

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Abstract

Aims/hypothesis Diabetes has been associated with a statistically significantly increased risk of endometrial cancer in most, but not all studies. To provide a quantitative assessment of the association between diabetes and risk of endometrial cancer, we conducted a meta-analysis of case-control studies and cohort studies.

Subjects and methods We identified studies by a literature search of PubMed and Embase through to January 2007 and by searching the reference lists of relevant articles. Summary relative risks (RRs) with 95% CIs were calculated using random-effects model.

Results The analysis of diabetes (largely type 2) and endometrial cancer is based on 16 studies (three cohort and 13 case-control studies), including 96,003 participants and 7,596 cases of endometrial cancer. Twelve of the studies showed a statistically significantly increased risk and four a non-significant increased risk of endometrial cancer. In our meta-analysis we found that diabetes was statistically significantly associated with an increased risk of endometrial cancer (summary RR 2.10, 95% CI 1.75–2.53). The risk estimates were somewhat stronger among case-control (RR 2.22, 95% CI 1.80–2.74) than among cohort studies (RR 1.62, 95% CI 1.21–2.16), stronger

among studies adjusting only for age (RR 2.74, 95% CI 1.87–4.00) compared with multivariate adjustment (RR 1.92, 95% CI 1.58–2.33) and slightly lower in studies performed in the USA than in those performed Europe. The analysis of type 1 diabetes and endometrial cancer was based on three studies and found a statistically significant positive association (summary RR 3.15, 95%CI 1.07–9.29). **Conclusions/interpretation** Results from the meta-analysis support a relationship between diabetes and increased risk of endometrial cancer.

Keywords Case-control · Cohort · Diabetes · Endometrial cancer · Meta-analysis

Abbreviation

RR relative risk

Introduction

The incidence of type 2 diabetes mellitus and endometrial cancer is increasing in westernised countries [1–3]. The major modifiable determinants of type 2 diabetes, hyperinsulinaemia and insulin resistance are obesity [4, 5] and low physical activity [6], both of which have also been shown to be risk factors for endometrial cancer [6–8]. A number of studies have reported a positive association between diabetes and incidence of [9–20] or mortality from [21] endometrial cancer. Many studies did not distinguish between type 1 and type 2 diabetes. However, on the basis of the relative prevalence of these two types, the vast majority of cases are likely to be type 2 diabetes. Although epidemiological studies of the relationship of diabetes with risk of endometrial cancer are not entirely consistent [22–26],

E. Friberg (✉) · N. Orsini · A. Wolk
Division of Nutritional Epidemiology,
The National Institute of Environmental Medicine,
Karolinska Institutet,
PO Box 210, 171 77 Stockholm, Sweden
e-mail: emilie.friberg@ki.se

C. S. Mantzoros
Division of Endocrinology, Diabetes and Metabolism,
Department of Medicine, Beth Israel Deaconess Medical Center,
Harvard Medical School,
Boston, MA, USA

Table 1 Characteristics of cohort studies of diabetes and endometrial cancer incidence and mortality

	Study population	Age range (years)	RR (95% CI) ^b	Case patients (<i>n</i>)	Controlled variables
Endometrial cancer incidence ^a					
Friberg, 2007 [9] Sweden (1998–2005)	Swedish Mammography Cohort Exposed group: 1,628 women with self-reported DM or DM from national inpatient register Comparison group: 35 145 women without self-reported DM or DM from national inpatient register	50–83	1.94 (1.23–3.08)	225	Age, BMI, total physical activity
Anderson, 2001 [24] USA (1986–1998)	Iowa Women's health study Exposed group: 1,325 women with self-reported DM Comparison group: 23 150 women without self-reported DM	55–69	1.43 (0.98–2.09)	346	Age, BMI, BMI ² , WHR, ovulatory span, gravidity, PMH, menstrual irregularities, hypertension
Terry, 1999 [25] Sweden (1967–1992)	Swedish Twin Registry Exposed group: 142 women with self-reported DM Comparison group: 10 012 women without self-reported DM	42–81	1.6 (0.2–11.3)	113	Age, physical activity, weight, parity
Endometrial cancer mortality ^a					
Folsom, 2004 [21] USA (1986–2000)	Iowa Women's health study Exposed group: 42 women with self-reported DM and an endometrial cancer diagnosis Comparison group: 373 women with self-reported DM and an endometrial cancer diagnosis	55–69	2.38 (1.05–5.37)	39	Age, extent of endometrial cancer at diagnosis
Coughlin, 2004 [26] USA (1982–1998)	US cohort: 588 321 women Exposed group: 33 women with self-reported DM Comparison group: 448 women without self-reported DM	≥30	1.33 (0.92–1.90)	33	Age, race, education, BMI, smoking, alcohol, red meat, citrus fruit and juice, vegetables, physical activity, PMH, parity, age at menarche, age at first live birth, menopausal status, OC (exclude hysterectomy)

International Classification of Diseases codes for endometrial cancer were not available except for Coughlin et al [26]: ICD9; 182–182.9.

Type of diabetes: largely type 2, unless otherwise specified.

DM, diabetes mellitus; OC, oral contraceptive use; HRT, hormone replacement therapy; PMH, postmenopausal hormone use

^a Authors, year [ref. no.], country (follow-up period)

^b The measure of RR is a rate ratio (hazard ratio) in all studies

most the studies are compatible with a positive association. However, because of the low precision of estimates in studies with a small sample size [12, 16, 25], reported risk estimates vary from 1.3 [23] to 7.8 [16].

We performed a meta-analysis to quantitatively summarise results from published cohort and case-control studies and to provide a more precise risk estimate for the association between diabetes and endometrial cancer.

Subjects and methods

Search strategy We identified studies by a literature search of the PubMed and Embase databases (from January 1966 to January 2007) with the following medical subject heading terms and/or text words: ‘diabetes mellitus’, ‘diabetes’, ‘endometrial cancer’, and ‘corpus uteri’. We also reviewed reference lists of the identified publications for additional pertinent studies. No language restrictions were imposed.

Inclusion and exclusion criteria The 33 studies considered for inclusion in this meta-analysis consisted of 13 cohort and 20 case-control studies on the association between diabetes and endometrial cancer incidence or mortality [9–41]. Studies were excluded if they provided only an effect estimate with no means to calculate a CI or if the estimates were not adjusted for age. We excluded two studies [28, 30] with unavailable effect estimates and three studies [29, 32, 33] that reported only crude data not adjusted for age. In the event of multiple publications from the same population or cohort, we included only data from the most recent report or, if the reports included exactly the same time frame, the publication with the most control for confounders. We excluded three candidate studies [27, 31, 41] because of overlapping publications.

We included 25 independent studies in this meta-analysis and excluded eight candidate studies.

Data extraction The data that we extracted included publication data (the first authors’ last names, year of publication and country of population studied), study design, number of exposed and unexposed subjects, follow-up period (for cohort studies), control source (in case-control studies), type of diabetes (type 1 or 2, when available), risk estimates with their corresponding CIs and variables controlled for by matching or in the multivariable model. From each study, we extracted the risk estimates that reflected the greatest degree of control for potential confounders.

Statistical analysis We divided epidemiological studies of the relationship between diabetes and endometrial cancer risk into three general types according to design: cohort studies (incidence and/or mortality rate ratio), case-control studies (odds ratio) and cohort studies with an external comparison group (standardised incidence and/or mortality ratio). We conducted separate meta-analyses of endometrial cancer incidence and mortality. The measure of effect of interest is the relative risk (RR). Because endometrial cancer is rare, the odds ratio in case-control studies and rate ratios in cohort studies yield similar estimates of RR

[42]. Cohort studies that reported standardised incidence: mortality ratio were analysed separately. Studies reporting an estimate for type 1 diabetes were analysed separately.

Summary RR estimates with their corresponding 95% CIs were derived with the method of DerSimonian and Laird [43] using the assumption of a random effects model that incorporated between-study variability. We calculated a pooled RR and its corresponding 95% CI. Statistical heterogeneity between studies was evaluated with Cochran’s Q test and the I^2 statistic [44]. Publication bias was assessed by constructing a funnel plot [45] and by Egger’s regression asymmetry test [46].

For cohort studies that reported incidence rate ratios and for case-control studies, we conducted subgroup meta-analyses to examine potential sources of heterogeneity, including study design, as well as type of control subjects in case-control studies. Statistical analyses were carried out with Stata, version 9.0 (Stata Corp., College Station, TX, USA). We considered p values that were less than 0.05 statistically significant. All statistical tests were two-sided.

Results

Study characteristics There were 25 independent studies that met the predefined inclusion criteria. Of these 25 studies, five were cohort studies that used incidence and/or mortality rate ratios as the measure of RR [9, 21, 24–26] (Table 1), 13 were case-control studies [10–20, 22, 23] (Table 2) and seven were cohort studies that used standardised incidence and/or mortality ratio as the measure of risk [34–40] (Table 3). Twelve studies were conducted in the USA, ten in Europe, one in South America and one in Asia. In the meta-analysis of diabetes (largely type 2) and endometrial cancer, we included three cohort studies that reported incidence rate ratios [9, 24, 25] and 13 case-control studies [10–20, 22, 23]. These 16 studies had 96,003 participants. The four cohort studies [34, 36–38] that reported standardised incidence ratios were analysed separately. For the meta-analysis of diabetes and endometrial cancer mortality, we included the two cohort studies that reported mortality rate ratio [21, 26]. These two studies enrolled a total of 896 participants. The four cohort studies [34, 36–38] that reported standardised incidence ratios and the three cohort studies [34, 39, 40] that reported standardised mortality ratios were analysed separately.

Four studies reported on type 1 diabetes and endometrial cancer, one case-control study [11], two studies providing standardised incidence ratio [34, 35] and one study

Table 2 Characteristics of case-control studies of diabetes and endometrial cancer

Study details ^a	Case patients (<i>n</i>)	Age range (years)	Control subjects, <i>n</i> , (selection methods)	Type of diabetes	RR (95% CI) ^b	Controlled variables
Weiss, 2006 [10] USA	1,281	45–74	1,779 (population, matched on age)	DM (self-reported)	1.58 (1.20–2.07)	Age, PMH, BMI, county, referent year, tumour aggressiveness
Weiderpass, 2000 [11] Sweden	709	50–74	3,368 (population, matched on age)	DM (self-reported)	1.7 (1.2–2.3)	Age, age at menarche, parity, age at last birth, age at menopause, smoking, OC, PMH, BMI
Salazar-Martínez, 2000 [12] Mexico	85	NA	668 (hospital, from primary health centre i.e. outpatient, matched on age)	Type 1 (self-reported DM with diagnosis at age <35 years and insulin treatment) DM (self-reported)	13.3 (3.1–56.4)	
Parazzini, 1999 [13] Italy	752	28–74	2,606 (hospital, admitted for acute, non-gynaecological, non-hormone related, non-neoplastic conditions)	DM (self-reported and medical records)	3.6 (1.7–7.4)	Age, anovulatory index, smoking, physical activity, menopausal status, hypertension, BMI
Shoff, 1998 [14] USA	723	40–79	2,291 (population controls, randomly selected from two sampling frames with individuals younger or older than 65 years)	DM (self-reported)	2.9 (2.2–3.9)	Age, calendar year, education, BMI, parity, OC, PMH, age at menopause, hypertension, smoking
Maatela, 1994 [15] Finland	1715	NA	1,715 (population controls, matched for age and place of residence)	DM (self-reported)	1.86 (1.37–2.52)	Age, BMI, smoking, PMH, parity, education
Inoue, 1994 [16] Japan	143	22–79	143 (hospital control who underwent hysterectomy due to benign gynaecological tumours, matched on year of admittance to hospital and age)	DM (medical registers)	4.0 (2.7–6.9)	Matched on age and place of residence
				DM (medical records)	7.75 (1.52–40.00)	Age, parity, cancer history, hypertension, obesity

Brinton, 1992 [17] USA	405	20–74	297 population controls random digit dialling for younger controls and + health care financing administration for older controls, older controls were matched on age, race and zip code)	DM (self-reported)	1.95 (1.1–3.6)	Age, education, number of births, weight, OC, PMH
Rubin, 1990 [22] USA	196	20–54	985 (population controls, matched for place of residence and age)	DM (self-reported)	1.8 (0.9–3.6)	Age
Lawrence, 1989 [18] USA	84	40–69	168 (population controls, matched on age and place of residence)	DM (self-reported)	4.14 (1.27–13.49)	Matched country, year of birth
Zemla, 1986 [19] Poland	173	24–77	346 (hospital controls not admitted for neoplastic disease and not related to the cases, matched on age)	DM (self-reported)	3.6 (1.4–9.5) ^c	Matched on age
O'Mara, 1985 [20] USA	479	30–89	2,420 (hospital controls, admitted for non-neoplastic disease)	DM (self-reported)	2.0 (1.3–3.0) ^c	Age
Kelsey, 1982 [23] USA	167	45–74	903 (hospital controls admitted to surgical services excluding gynaecology)	DM (self-reported)	1.3 (0.7–2.3)	Age, race, education, pregnancies, age at menopause, weight, tubes tied, vaginal hormone use, clots in veins, OC, PMH, menopausal status

International Classification of Diseases codes not available.

Type of diabetes: largely type 2, unless otherwise specified.

DM, diabetes mellitus; NA, data were not available; OC, oral contraceptive use; PMH, postmenopausal hormone use

^a Authors, year [ref. no.], country

^b Measure of relative risk is an odds ratio except for one (Maatela, 1994 [15]).

RR was calculated as a weighted average, from data presented by tumour aggressiveness in the article.

^c CIs were calculated from raw data reported in the article

Table 3 Characteristics of cohort studies of diabetes and endometrial cancer based on standardised incidence/mortality ratio

	Study population	SIR, RR (95% CI) ^b	Cases among DM patients (<i>n</i>)	Controlled variables	ICD for case status
Endometrial cancer incidence ^a					
Swerdlow, 2005 [34] UK (1972–2003)	Exposed group: 13 212 patients with insulin-treated DM Comparison group: UK population	Type 1: 1.20 (0.33–3.08)	4	Age, calendar year, country	ICD9: 179, 182
Zendehdel, 2003 [35] Sweden (1965–1999)	Exposed group: 14 323 hospitalised in Sweden for type 1 DM Comparison group: Swedish population	Type 2: 1.84 (0.95–3.22) Type 1: 2.7 (1.4–4.7)	12	Age, calendar year	NA
Weiderpass, 1997 [36] Sweden (1965–1989)	Exposed group: 70 110 women hospitalised for DM Comparison group: Swedish population	1.8 (1.6–2.0)	328	Age, calendar year	ICD7: 172
Wideroff, 1997 [37] Denmark (1977–1989)	Exposed group: 55010 women with a DM diagnosis in the Danish Central Hospital Discharge Register Control group: Danish population	1.4 (1.2–1.6)	231	Age, calendar year	ICD7: 172
Ragozzino, 1982 [38] USA (1945–1969)	Exposed group: 1135 DM patients. Control group: Rochester, NY, population	1.3 (0.3–3.8)	3	Age	NA
Endometrial cancer mortality ^a					
Swerdlow, 2005 [34] UK (1972–2003)	Exposed group: 13 212 patients with insulin-treated DM Comparison group: UK population	Type 1: 1.30 (0.03–726) Type 2: 2.13 (0.58–5.45)	1 4	Age, calendar year, country	ICD9: 179, 182
Verlato, 2003 [39] Italy (1987–1996)	Exposed group: 7148 type 2 DM patients. Comparison group: Verona population.	1.08 (0.54–1.93)	11	Age	ICD9: 179– 180, 182
Kessler, 1970 [40] USA (1941–1959)	Exposed group: 21 447 (male and female) DM patients in Boston, MA Comparison group: Massachusetts population	0.62 (0.38 1.0) ^c	19	Age, sex	NA

Type of diabetes—largely type 2, unless otherwise specified.

DM, diabetes mellitus; NA, data not available; SIR, standardised incidence ratio; ICD, International Classification of Diseases

^a Authors, year [ref. no.], country (follow-up period)

^b The measure of RR is a standardised incidence (or mortality) ratio

^c CIs were calculated from raw data reported in the article.

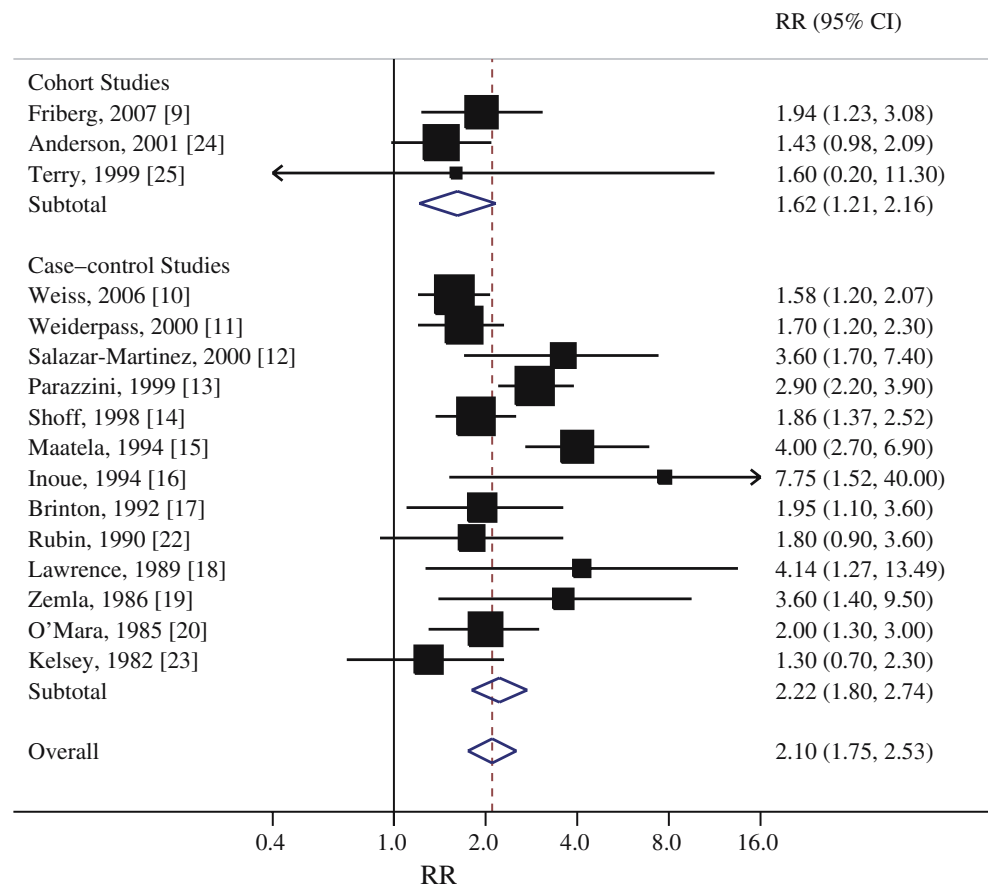
reporting on type 1 diabetes and endometrial cancer mortality providing a standardised mortality ratio [34].

Endometrial cancer Individual study results and the overall summary result for three cohort and 13 case-control studies of diabetes and endometrial cancer are shown in Fig. 1. Twelve of these 16 studies found a statistically significant positive association between diabetes and endometrial cancer (range of individual RRs 1.3–7.75; summary RR for all 16 studies 2.10, 95% CI 1.75–2.53). The heterogeneity among studies was $p=0.01$. In a sensitivity analysis in which one study at a time was excluded and the rest were analysed, we detected a statistically significant positive association between diabetes

and endometrial cancer (range of summary RRs 1.98–2.18 with the lower limit of the 95% CI never crossing 1.0).

We then conducted subgroup meta-analyses by study design, geographical area, control group (for case-control studies) and adjustments (including all the measured confounders vs adjusted for age only) (Table 4). The association between diabetes and endometrial cancer was somewhat stronger in Europe than in USA and also somewhat stronger among case-control studies. When adjustments are taken into account, the studies adjusting only for age reported a stronger association than those adjusting the RR with a model including all the measured confounders, indicating a presence of confounding.

Fig. 1 Association between diabetes (largely type 2) and endometrial cancer incidence in cohort and case-control studies. Studies are ordered by publication year and stratified on design. Squares, study-specific RR estimate (size of the square reflects the study-specific statistical weight, i.e. the inverse variance); horizontal lines, 95% CI; diamond, summary RR estimate and its corresponding 95% CI. All statistical tests were two-sided. Statistical heterogeneity between studies was assessed with Cochran's Q test. Test for heterogeneity among cohort studies: $Q=1.01$; $p=0.60$; $I^2=0.0\%$. Test for heterogeneity among case-control studies: $Q=28.34$; $p=0.01$; $I^2=57.7\%$. Test for heterogeneity between cohort and case-control studies: $Q=2.70$; $p=0.10$



Physical inactivity and body mass index are potentially the most important known confounders of the positive association between diabetes and endometrial cancer risk. When we restricted the meta-analysis to the two studies that controlled for these variables [9, 12], we found a positive association between diabetes and endometrial cancer (summary RR 2.47, 95% CI 1.37–4.45, test for heterogeneity $Q=1.95$, $p=0.16$, $I^2=48.8\%$). A positive association was observed between diabetes and endometrial cancer incidence in the cohort studies that reported standardised incidence ratios [34, 36–38] (Table 3) (summary RR 1.63, 95% CI 1.30–2.05, test for heterogeneity $Q=9.60$, $p=0.02$, $I^2=68.8\%$).

Endometrial cancer mortality Of the two cohort studies of diabetes and mortality from endometrial cancer [21, 26], one [21] reported a statistically significant positive association and one [26] observed a non-statistically significant positive association. When the studies were pooled, a positive, but non-significant association between diabetes and mortality from endometrial cancer was found (summary RR 1.58, 95% CI 0.94–2.66, test for heterogeneity $Q=1.63$, $p=0.20$, $I^2=38.7\%$). No association was observed between diabetes and endometrial cancer mortality in the cohort studies that reported standardised mortality ratios [34, 39, 40] (summary

RR 0.97, 95% CI 0.52–1.81, test for heterogeneity $Q=4.76$, $p=0.09$, $I^2=58.0\%$).

Type 1 diabetes Two cohorts providing standardised incidence ratios and one case-control study reported on the association between type 1 diabetes and endometrial cancer [11, 34, 35]. When the three studies were meta-analysed, a statistically significant positive association between type 1 diabetes and endometrial cancer was found (summary RR 3.15, 95%CI 1.07–9.29, test for heterogeneity $Q=6.66$, $p=0.04$, $I^2=70.0\%$).

Publication bias The funnel plot we constructed revealed no evidence for publication bias concerning diabetes and risk of endometrial cancer (data not shown). The p value for Egger's regression asymmetry test was 0.14, i.e. probability of publication bias was low.

Discussion

Findings from this meta-analysis show that women with diabetes (largely type 2 diabetes) may have an approximately twofold increased risk of developing endometrial

Table 4 Summary relative risk estimates and 95% CIs for case-control and cohort studies of the association between diabetes and endometrial cancer by study design, geographical area and adjustments

Subgroup	Studies (<i>n</i>)	Summary RR (95% CI)	Between studies			Between subgroups	
			<i>Q</i>	<i>p</i> value for heterogeneity	<i>I</i> ² statistic (%)	<i>Q</i>	<i>p</i> value for heterogeneity
Study design							
Cohort studies	3	1.62 (1.21–2.16)	1.01	0.60	0.0		
Case-control	13	2.22 (1.80–2.74)	28.34	0.01	57.7	2.74	0.10 ^a
Population-based	7	2.04 (1.58–2.63)	13.60	0.03	55.9		
Hospital-based	6	2.51 (1.78–3.56)	10.11	0.07	50.5	4.64	0.03 ^b
Geographical area							
USA	8	1.70 (1.47–1.98)	5.19	0.64	0.0		
Europe	6	2.51 (1.83–3.45)	12.16	0.03	59.0		
Other (Mexico, Japan)	2	4.10 (2.09–8.01)	0.70	0.40	0.0	14.00	0.00
Adjustments							
Multivariable	11	1.92 (1.58–2.33)	20.12	0.03	50.3		
Age	5	2.74 (1.87–4.00)	6.83	0.15	41.5	5.10	0.02

Diabetes, largely type 2

All statistical tests were two-sided.

^a Test for heterogeneity between case-control and cohort studies.

^b Test for heterogeneity between population-based and hospital-based case-control studies.

cancer compared with non-diabetic individuals. Our analyses on type 1 diabetes suggest that women with type 1 diabetes may have a threefold increased risk of endometrial cancer. However, our analysis does not support the notion that diabetes is associated with increased risk of endometrial cancer mortality.

All previous studies have consistently shown a positive association between diabetes and endometrial cancer incidence. However, heterogeneity exists due to differences in reported strength of the effect estimate. The summary RR was consistent, but slightly higher for case-control than cohort studies, as well as among European studies vs those done in the USA. The latter could reflect a residual confounding from BMI. The studies only adjusting for age also showed a slightly stronger effect estimate, probably due to lack of controlling for BMI. This finding was also consistent in cohort studies reporting a standardised incidence ratio.

Both our findings on endometrial cancer mortality in relation to diabetes (largely type 2), and our data on the risk of endometrial cancer in type 1 diabetes are more limited by uncertainty due to the smaller number of studies and to studies including small numbers of cases.

Our analysis must be interpreted in the context of the limitations of available data. Many studies did not distinguish between type 1 and type 2 diabetes, but given the relative prevalence of these two types, it is likely that the vast majority of cases are type 2 diabetes. Since diabetes is an underdiagnosed disease, some degree of misclassification of exposure to diabetes is probable, but

such non-differential misclassification would be expected to have attenuated the true relationship between diabetes and endometrial cancer. Moreover, some subgroup analyses were based on few studies and the results need to be interpreted with caution. As in all meta-analysis, the possibility of publication bias is of concern; however, a formal statistical test did not provide evidence for such bias. Type 2 diabetes and endometrial cancer share some risk factors such as obesity and low physical activity. Thus, the observed increased risk of endometrial cancer associated with diabetes may in part reflect confounding by these factors. Nevertheless, a statistically significant positive association remained when the analysis was limited to the two studies controlling for obesity and physical activity. That said, further residual confounding cannot be ruled out.

Several biological mechanisms have been proposed to potentially underlie the development of endometrial cancer in diabetic women. Hyperinsulinaemia is a common feature of diabetes, obesity and physical inactivity, and insulin has been shown to stimulate the growth of endometrial stromal cells by binding to insulin receptors on endometrial cells [47]. Epidemiological studies have observed an elevated risk of endometrial cancer in relation to high prediagnostic C-peptide concentrations indicating hyperinsulinaemia [48]. Long-term insulin therapy of patients with type 1 diabetes may explain the increased risk of endometrial cancer among diabetic women with type 1 diabetes [35]. Hyperinsulinaemia may also increase levels of bioactive oestrogens by decreasing concentrations of circulating sex hormone binding globulin [49, 50]. Oestrogens have been

shown to increase endometrial cancer risk by stimulating proliferation of endometrial cells [51], when unopposed by progesterone [52, 53]. Finally, hyperinsulinaemia leads to decreased levels of insulin-like growth factor binding protein 1. This in turn increases circulating free IGF-1, which stimulates endometrial cell proliferation [54–59]. Finally, the adipocyte-secreted hormone, adiponectin, is an endogenous insulin sensitiser lying upstream of all the above-mentioned hormonal factors and capable of affecting their circulatory levels [60]. Low levels of adiponectin predict not only diabetes but also endometrial cancer incidence [61] and can thus serve as the link between diabetes, hormonal abnormalities and endometrial cancer risk.

Our results have important clinical and public health implications, given the already high and continuously increasing prevalence of type 2 diabetes and endometrial cancer in western societies.

In summary, results from this meta-analysis support an association between diabetes and endometrial cancer. Future studies will need to clearly distinguish between type 1 and type 2 diabetes, as well as different types of treatment for diabetes. It would also be of interest to distinguish between specifically defined subtypes of endometrial cancer.

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Duality of interest The authors state that there are no dualities of interest to declare.

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