

Upper Classes for the Increments of Fractional Wiener Processes [★]

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Summary. Let $(X(t), t \geq 0)$ be a centred Gaussian process with stationary increments and $EX^2(t) = C_0 t^{2\alpha}$ for some $C_0 > 0$, $0 < \alpha < 1$, and let $0 < a_t \leq t$ be a nondecreasing function of t with a_t/t nonincreasing. The asymptotic behaviour of several increment processes constructed from X and a_t is studied in terms of their upper classes.

1. Introduction

Let $(X(t), t \geq 0)$ be a centred Gaussian process with stationary increments, $X(0) = 0$ a.s. and define $\sigma^2(h) = EX^2(h) = E(X(t+h) - X(t))^2$. If $\sigma^2(h) = C_0 h^{2\alpha}$, $0 < \alpha < 1$ and $C_0 > 0$ then X is known as a fractional Wiener process of order α (FWP(α)). If $\alpha = 1/2$ and $C_0 = 1$ this is the standard Wiener process.

The purpose of this paper is the study of the asymptotic behaviour of the following increment processes: let a_t be a nondecreasing function of t with $0 < a_t \leq t$, and a_t/t nonincreasing. We define the following processes in terms of X and a_t :

$$Y_1(t) = \frac{X(t+a_t) - X(t)}{\sigma(a_t)}$$

$$Y_2(t) = \sup_{0 \leq s \leq t} Y_1(s)$$

$$Y_3(t) = \sup_{0 \leq u \leq a_t} \frac{X(t+u) - X(t)}{\sigma(a_t)}$$

$$Y_4(t) = \sup_{0 \leq s \leq t - a_t} Y_3(s)$$

$$Y_5(t) = \sup_{0 \leq s \leq t} Y_3(s)$$

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and the processes $Y_i^*(t), i = 1, \dots, 5$ defined using the absolute value of the increments. Thus, for example,

$$Y_1^*(t) = \frac{|X(t+a_t) - X(t)|}{\sigma(a_t)}$$

In [7] it was shown that if X is a FWP(α) then

$$\begin{aligned} \limsup_{T \rightarrow \infty} Y_1(T) \beta_T &= \limsup_{T \rightarrow \infty} Y_5^*(T) \beta_T \\ &= \limsup_{T \rightarrow \infty} \sup_{0 \leq t \leq T} \sup_{0 \leq s \leq a_T} \frac{|X(t+s) - X(t)|}{\sigma(a_T)} \beta_T = 1 \quad \text{a.s.} \end{aligned} \tag{1.1}$$

where $\beta_T = \left(2 \log \left(\frac{T \log T}{a_T}\right)\right)^{-1/2}$. For the Wiener process these results were obtained in 1979 by Csörgö and Révész [4]. Also, (1.1) implies the LIL for FWP proved by Orey [5]. Note that if $a_t = t, Y_4(t) = \sup_{0 \leq s \leq t} X(s)$.

We shall consider the asymptotic behaviour of these processes with respect to a nondecreasing function $\phi(t)$ in terms of their upper classes, which we define following Révész [13]:

Definition 1. The function ϕ belongs to the *upper-upper class* of the process $Z(\phi \in \text{UUC}(Z))$ if, with probability one, there exists a $t_0(\omega)$ such that $Z(t) < \phi(t)$ for all $t > t_0$.

Definition 2. The function ϕ belongs to the *upper-lower class* of the process $Z(\phi \in \text{ULC}(Z))$ if, with probability one, there exists a random sequence $0 < t_1 < t_2 < \dots$ with $t_i \rightarrow \infty$ as $i \rightarrow \infty$ such that $Z(t_i) \geq \phi(t_i)$ for $i \geq 1$.

In [6] the upper classes of the increments of the Wiener process were described by means of an integral test as follows.

Theorem A. Let $\alpha = 1/2$ and let Z be any of the processes $Y_i, Y_i^*, i = 1, \dots, 5$, then

$$\phi \in \text{UUC}(Z) \Leftrightarrow \int_1^\infty \frac{\phi(t)}{a(t)} \exp\left\{-\frac{\phi^2(t)}{2}\right\} dt < \infty.$$

The main result of this paper is the following:

Theorem 1. Let X be a FWP(α) and Z any of the processes $Y_i, Y_i^*, i = 1, \dots, 5$, then

$$\phi \in \text{UUC}(Z) \Leftrightarrow I_\alpha(\phi) = \int_1^\infty \frac{(\phi(t))^{(1/\alpha)-1}}{a(t)} \exp\left\{-\frac{\phi^2(t)}{2}\right\} dt < \infty.$$

This result includes Theorem A and Theorem 5 in [15] as special cases, and implies the first two inequalities in (1.1). It also extends previous works of the author on the increments of FWP [7, 8].

In the next section we give some preliminary results and in Sect. 3 the proof of Theorem 1. We shall assume without loss of generality that $C_0 = 1$. In what follows C denotes a positive constant which can take different values on each appearance and $\psi(x) = (1/\sqrt{2\pi x}) \exp\left(-\frac{x^2}{2}\right)$. We shall frequently use the well-known fact that if X is a standard Gaussian r.v. then $P(X > x) \leq \psi(x)$ for $x > 0$ and $\frac{P(X > x)}{\psi(x)} \rightarrow 1$ as $x \rightarrow \infty$.

2. Preliminary Results

In this section we give some results that will be used in the proof of Theorem 1. In particular, Lemma 5 is an asymptotic upper bound for the tail of the distribution of the oscillations of FWP which may be independent interest.

Lemma 1 (Qualls-Watanabe [11]). *Let $\{X(t), t \in \mathbb{R}^n\}$ be a continuous centred Gaussian random field with variance 1 satisfying*

$$E(X(p) - X(q))^2 \leq 2C_1 \|p - q\|^{2\alpha}$$

for all p, q in $D \subset \mathbb{R}^n$ with $\|p - q\| < \delta_1$, where $0 < \alpha \leq 1$ and $\|x\|$ is the Euclidean norm of $x \in \mathbb{R}^n$. Then

$$\limsup_{u \rightarrow \infty} \frac{P(Z(D) > u)}{\mu(D) \psi(u) (c(u))^n} \leq H_\alpha C_1^{n/2\alpha}$$

where $Z(D) = \sup\{X(p), p \in D\}$, D is an open bounded set with Lebesgue measure $\mu(D) = \mu(\bar{D})$, $c(u) = u^{1/\alpha}$ and $0 < H_\alpha < \infty$ is a constant which does not depend on u .

Let $(X(t), t \geq 0)$ be a FWP(α), we define the biparametric processes γ and Δ by

$$\gamma(t, t') = X(t) - X(t'); \quad \Delta(t, t') = \frac{\gamma(t, t')}{\sigma(t' - t)} \quad \text{for } t < t'.$$

Then $E\gamma(t, t') = E\Delta(t, t') = 0$ and $E\Delta^2(t, t') = 1$.

Lemma 2. *If $\mathbf{t} = (t, t')$ and $\mathbf{s} = (s, s')$ with $t < t'$ and $s < s'$, the incremental variance of Δ satisfies*

$$\bar{\sigma}^2(\mathbf{t}, \mathbf{s}) = E(\Delta(t, t') - \Delta(s, s'))^2 \leq \frac{4 \|\mathbf{t} - \mathbf{s}\|^{2\alpha}}{|t' - t|^\alpha |s' - s|^\alpha}$$

Proof.

$$\begin{aligned}
 \tilde{\sigma}^2(\mathbf{t}, \mathbf{s}) &= E \left(\frac{X(t') - X(t)}{\sigma(t' - t)} - \frac{X(s') - X(s)}{\sigma(s' - s)} \right)^2 \\
 &= \frac{E(X(t') - X(t) - X(s') + X(s))^2 - (\sigma(t' - t) - \sigma(s' - s))^2}{\sigma(t' - t) \sigma(s' - s)} \\
 &\leq \frac{2(\sigma^2(|t' - s'|) + \sigma^2(|t - s|))}{\sigma(t' - t) \sigma(s' - s)} \\
 &\leq \frac{4(|t' - s'|^2 + |t - s|^2)^\alpha}{|t' - t|^\alpha |s' - s|^\alpha} \\
 &= \frac{4 \|\mathbf{t} - \mathbf{s}\|^{2\alpha}}{|t' - t|^\alpha |s' - s|^\alpha} \quad \square
 \end{aligned}$$

We give now some results about the asymptotic distribution of the supremum of γ

Lemma 3. *Let $(X(t), t \geq 0)$ be a FWP(α) with $\alpha > 1/2$. Then for $h > 0$,*

$$\lim_{u \rightarrow \infty} \frac{P \left(\sup_{0 \leq t < t' \leq h} \gamma(t, t') \geq \sigma(h) u \right)}{P\{X(1) > u\}} = 1. \tag{2.1}$$

This lemma is a consequence of Theorem B below: let $\{X(t), t \in [0, 1]^n\}$ be a real separable centred Gaussian process with continuous covariance function, and put $\sigma^2(t) = EX^2(t)$. Suppose that there is a point τ in $[0, 1]^n$ such that $\sigma^2(t)$ has a unique maximum value at $t = \tau$, and put $\sigma^2 = \sigma^2(\tau)$. Define the metric $\|s - t\| = \max_i |s_i - t_i|$ where (s_i) and (t_i) are the real components of s and t . Suppose that there exist positive nondecreasing functions $q(t)$ and $g(t)$, $t > 0$, such that

$$\limsup_{\|t-s\| \rightarrow 0} \frac{E(X(s) - X(t))^2}{q^2(\|s-t\|)} < \infty \quad \limsup_{t,s \rightarrow \tau} \frac{E(X(s) - X(t))^2}{g^2(\|s-t\|)} < 1$$

and

$$\int_1^\infty q(e^{-y^2}) dy < \infty \quad \int_1^\infty g(e^{-y^2}) dy < \infty.$$

Define

$$Q(h) = q(h) + (2 + \sqrt{2}) \int_1^\infty q(h 2^{-y^2}) dy < \infty, \quad 0 < h \leq 1.$$

$$G(h) = g(h) + (2 + \sqrt{2}) \int_1^\infty g(h 2^{-y^2}) dy < \infty, \quad 0 < h \leq 1.$$

$$Q^{-1}(x) = \sup\{h: Q(h) \leq x\}$$

for $h > 0$, let $B(h) = \{t: \|t - \tau\| \leq h/2\}$ and define $\bar{\sigma}^2(h) = \max\{\sigma^2(t): t \in [0, 1]^n \cap B'(h)\}$ where A' is the complement of A . Then we have the following result:

Theorem B ([3], Theorem 2.1). *Suppose that there exist functions g and q satisfying the conditions stated above. If, for every $\varepsilon > 0$,*

$$\lim_{h \rightarrow 0} [Q^{-1}(G(h)/\varepsilon)]^{-n} \exp \left\{ -\frac{\varepsilon^2}{2\sigma^4} \left[\frac{\sigma^2 - \bar{\sigma}^2(h)}{G^2(h)} \right] \right\} = 0$$

then

$$\lim_{u \rightarrow \infty} P(\max_{[0, 1]^n} X(t) > u) / \psi(u/\sigma) = 1$$

To prove Lemma 3 it is enough to consider the case $h=1$ and prove that $E(\gamma(t, t') - \gamma(s, s'))^2 \leq 4(\|(t, t') - (s, s')\|^{2\alpha})$ (see Example 4.1 in [3]).

Lemma 4. *Let $\{X(t), t \geq 0\}$ be a FWP(α) with $\alpha < 1/2$ and $h > 0$. Then there exists a constant C_2 , which may depend on α but is independent of h , such that*

$$\limsup_{u \rightarrow \infty} \frac{P\left\{ \sup_{0 \leq t < t' \leq h} \gamma(t, t') \geq \sigma(h)u \right\}}{u^{(1/\alpha)-2} P\{X(1) > u\}} \leq C_2 \tag{2.2}$$

Proof. It is enough to consider the case $h=1$. We consider first the supremum of γ over the set $A = \{(t, t') : 0 \leq t < t' \leq 1, t' - t \leq 1/2\}$. For $\varepsilon > 0$ and $m = [2/\varepsilon]$ we have that

$$\begin{aligned} P\left\{ \sup_A \gamma(t, t') > u \right\} &\leq \sum_{j=0}^m P\left\{ \sup_{\substack{j\varepsilon/2 \leq t \leq (j+1)\varepsilon/2 \\ 0 < t' - t \leq 1/2}} \gamma(t, t') > u \right\} \\ &\leq (m+1) P\left\{ \sup_{0 < t' - t \leq (1+\varepsilon)/2} \gamma(t, t') > u \right\} \end{aligned} \tag{2.3}$$

and to obtain a bound for this probability we use Theorem 3.3 in [2] for the process γ over the square $R = \{0 \leq t \leq (1+\varepsilon)/2, 0 \leq t' \leq (1+\varepsilon)/2\}$. In our case $Q(x) = \mathcal{O}(x^\alpha)$ and $\sigma_R = \left(\frac{1+\varepsilon}{2}\right)^\alpha$, hence

$$\begin{aligned} (2.3) &\leq (m+1) \text{Const} (Q^{-1}(1/u))^{-2} \frac{\sigma_R}{u} \exp \left\{ -\frac{u^2}{2\sigma_R^2} \right\} \\ &\leq \text{Const} \frac{1}{\varepsilon} u^{(2/\alpha)-1} \exp \left\{ -\frac{u^2 2^{2\alpha}}{2(1+\varepsilon)^{2\alpha}} \right\} \end{aligned}$$

and if $\varepsilon < 1$ this is $\mathcal{O}(\psi(u))$ as $u \rightarrow \infty$.

It remains to consider the process γ over the set $E = \{(t, t') : 0 \leq t \leq 1/2, t + 1/2 < t' \leq 1\}$. To do this we cover E by squares of side η with η satisfying $uQ(\eta) \leq 1$, where Q is defined above. Since $Q(x) = \mathcal{O}(x^\alpha)$ in our case, we have $\eta = \mathcal{O}(u^{-1/\alpha})$. If u is large enough, the covering will be included in the set $S = \{(t, t') : 0 \leq t \leq 2/3, t + 1/3 \leq t' \leq 1\}$ and $\sigma^2 = \inf_S E(\gamma^2(t, t')) \geq 1/3^{2\alpha} > 1/3$. Therefore, the right hand

side of (3.12) in [2] is uniformly bounded for all $u > u_0$.

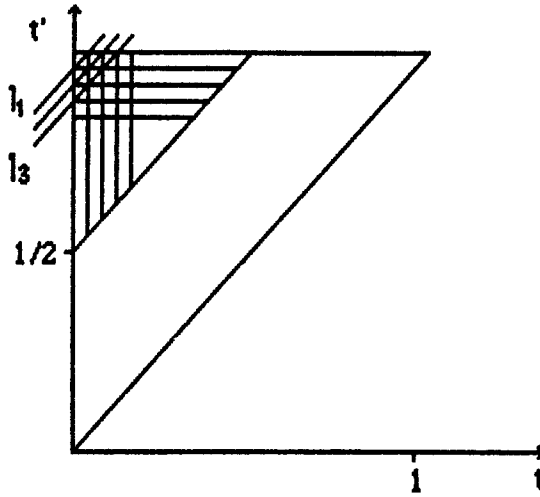


Fig. 1

We cover E by squares of sides parallel to the axes and length η starting at the right angled corner (see Fig. 1). Since X has stationary increments, the distribution of γ over squares of same size having diagonals over the same straight line is the same. Also, if we start counting from the right-angled corner, there are exactly i squares having diagonals over the straight line l_i of equation $t' = t + 1 - i\eta, i = 1, \dots, N$, where $N = [1/2\eta] + 1$.

Let $E_i(\eta) = \{(t, t') : 0 \leq t \leq \eta, 1 - (i - 1)\eta \leq t' \leq 1 - i\eta\}$. Then, using Corollary 3.1 in [2], with $\sigma_{E_i(\eta)} = (1 - (i - 1)\eta)^\alpha$, we have

$$\begin{aligned}
 P \left\{ \sup_E \gamma(t, t') > u \right\} &\leq \text{Const} \sum_{i=0}^N P \left\{ \sup_{E_i(\eta)} \gamma(t, t') > u \right\} \\
 &\leq \text{Const} \sum_{i=0}^N i \psi(u/\sigma_{E_i(\eta)}) \\
 &\leq \text{Const} \sum_{i=0}^N \frac{(i+1)(1+i\eta)^\alpha}{u} \exp \left\{ -\frac{u^2}{2(1-i\eta)^{2\alpha}} \right\} \\
 &\leq \text{Const} \int_0^{1/\eta} \frac{(y+1)(1-y\eta)^\alpha}{u} \exp \left\{ -\frac{u^2}{2(1-y\eta)^{2\alpha}} \right\} dy
 \end{aligned}$$

Using the method of Laplace to estimate this integral for large u , we get that it is asymptotically like $(\alpha\eta u^3)^{-1} \exp(-u^2/2)$ and since $\eta = \mathcal{O}(u^{-1/\alpha})$ we get the desired result. \square

Next we consider the supremum of the process η over parallelograms.

Lemma 5. Let $\{X(t), t \geq 0\}$ be a FWP(α), $0 < \alpha \leq 1$. Then

$$q(T, h, u) = P\left\{ \sup_{\substack{0 \leq t \leq T \\ 0 < t' - t \leq h}} (X(t') - X(t)) \geq \sigma(h) u \right\} \leq C_\alpha \frac{T}{h} u^{(1/\alpha)-1} e^{-u^2/2} \quad (2.4)$$

for $u \geq u_0$, where C_α is a constant which may depend on α but is independent of T and h .

Proof. The case $\alpha = 1/2$ was considered in [6] (see also [13]). Let $0 < \alpha < 1/2$, $0 < \varepsilon \leq 1$ and $m = [T/\varepsilon h]$, then

$$\begin{aligned} q(T, h, u) &\leq \sum_{j=0}^m P\left\{ \sup_{\substack{j\varepsilon h \leq t \leq (j+1)\varepsilon h \\ 0 < t' - t \leq h}} \gamma(t, t') \geq \sigma(h) u \right\} \\ &\leq (m+1) P\left\{ \sup_{0 \leq t < t' \leq h(1+\varepsilon)} \gamma(t, t') \geq \sigma(h) u \right\} \\ &\leq (m+1) P\left\{ \sup_{0 \leq t < t' \leq h(1+\varepsilon)} \frac{\gamma(t, t')}{\sigma(h(1+\varepsilon))} \geq \frac{u}{\sigma(1+\varepsilon)} \right\} \end{aligned}$$

and by Lemma 4, for u sufficiently large this is bounded by

$$\begin{aligned} &\leq \text{Const}(m+1) \left(\frac{u}{\sigma(1+\varepsilon)} \right)^{(1/\alpha)-2} P\left(X(1) \geq \frac{u}{\sigma(1+\varepsilon)} \right) \\ &\leq \text{Const} \frac{T}{\varepsilon h} u^{(1/\alpha)-3} \exp\left\{ -\frac{u^2}{2(1-\varepsilon)^{2\alpha}} \right\} \end{aligned}$$

and choosing $\varepsilon = 1/u^2$ one obtains (2.4) by means of a simple calculation.

For the case $1/2 < \alpha < 1$ the proof is divided in two steps. We consider first the process $\Delta(t, t')$ over the set $B = \{(t, t') : 0 \leq t \leq T, t + h(1+\varepsilon) \leq t' \leq t+h\}$ where $\varepsilon = 1/u^{1/\alpha}$. Lemma 2 shows that for $\mathbf{t} = (t, t')$ and $\mathbf{s} = (s, s')$ in B and u large we have

$$E(\Delta(t, t') - \Delta(s, s'))^2 \leq \frac{4 \|\mathbf{t} - \mathbf{s}\|^2}{h^{2\alpha}(1-\varepsilon)^{2\alpha}} \leq \frac{4^{1+\alpha}}{h^{2\alpha}} \|\mathbf{t} - \mathbf{s}\|^2$$

Therefore, by Lemma 1

$$\begin{aligned} P\left\{ \sup_B \gamma(t, t') > \sigma(h) u \right\} &\leq P\left\{ \sup_B \Delta(t, t') > u \right\} \\ &\leq C_3 \frac{T}{h} u^{(1/\alpha)-1} \exp\left(-\frac{u^2}{2} \right) \end{aligned} \quad (2.5)$$

for u large and some constant C_3 which may depend on α . We still have to consider the supremum of γ over the set $D = \{(t, t') : 0 \leq t \leq T, t < t' \leq t + h(1-\varepsilon)\}$.

Let $m = \left[\frac{Tu^{1/\alpha}}{h} \right]$, then

$$\begin{aligned} P\left\{ \sup_D \gamma(t, t') > \sigma(h) u \right\} &\leq \sum_{j=0}^m P\left\{ \sup_{\substack{j\varepsilon h \leq t \leq (j+1)\varepsilon h \\ 0 < t' - t \leq h(1-\varepsilon)}} \gamma(t, t') > \sigma(h) u \right\} \\ &\leq (m+1) P\left\{ \sup_{0 \leq t < t' \leq h} \gamma(t, t') > \sigma(h) u \right\} \end{aligned}$$

and by Lemma 3, for u sufficiently large this is bounded by

$$\begin{aligned} &\leq 2(m+1) P\{X(1) \geq u\} \\ &\leq \text{Const} \frac{T}{h} u^{(1/\alpha)-1} \exp\left\{-\frac{u^2}{2}\right\} \end{aligned} \tag{2.6}$$

Finally, combining (2.5) and (2.6) we obtain (2.4). \square

Lemma 6 (Berman, Plackett, Slepian [1, 9, 14]). *Let $\{X_i, i = 1, \dots, n\}$ be centred, stationary Gaussian r.v.'s with $EX_i^2 = 1$ for all i and $EX_i X_j = r_{ij}$. Let $I(c, +1) = [c, \infty)$ and $I(c, -1) = (-\infty, c)$. If $c_i \in \mathbb{R}, i = 1, \dots, n$ denote by F_i the event $\{X_i \in I(c_i, \varepsilon_i)\}$ where ε_i is either $+1$ or -1 . Let $K \subset \{1, \dots, n\}$, then:*

i) $P\{\bigcap_{i \in K} F_i\}$ is an increasing function of r_{ij} if $\varepsilon_i \varepsilon_j = +1$, otherwise it is decreasing.

ii) If $\{K_l, l = 1, \dots, s\}$ is a partition of K then

$$\left| P\{\bigcap_{i \in K} F_i\} - \prod_{l=1}^m P\{\bigcap_{i \in K_l} F_i\} \right| \leq \sum_{i \leq l < m \leq s} \sum_{j \in K_l} \sum_{j \in K_m} |r_{ij}| g(c_i, c_j; r_{ij}^*)$$

where $g(x, y; r)$ is the standard bivariate Gaussian density with correlation r and r_{ij}^* is a number between 0 and r_{ij} .

The proof of the following version of the Borel-Cantelli lemma can be found in [12].

Lemma 7. *Let $(G_n, n \geq 1)$ be a sequence of events. If*

$$\begin{aligned} i) \quad &\sum_{n=0}^{\infty} P(G_n) = \infty \\ ii) \quad &\liminf_{n \rightarrow \infty} \frac{\sum_{1 \leq i < k \leq n} [P(G_i \cap G_k) - P(G_i)P(G_k)]}{\left(\sum_1^n P(G_i)\right)^2} = 0 \end{aligned} \tag{2.7}$$

then $P(G_n \text{ i.o.}) = 1$.

The following lemma is a consequence of the proof of Lemma 2.3 in [10], taking $a = 1$.

Lemma 8 (Qualls-Watanabe). *Let $\{X(t), t \geq 0\}$ be a real centred Gaussian process with $EX^2(t) = 1$, which satisfies*

$$E(X(t) - X(s))^2 \geq C_4 |t - s|^{2\alpha}$$

for $0 \leq s, t \leq T, 0 < |t - s| < \delta_2$ where $0 < \alpha \leq 1$. Then

$$\liminf_{u \rightarrow \infty} \frac{P\{Z_u(T) > u\}}{T\psi(u)c(u)} \geq C_4^{1/2\alpha} H_\alpha$$

where $Z_u(T) = \max_{0 \leq k \leq m} X(k/c(u))$, $m = [Tc(u)]$, $c(u) = u^{1/\alpha}$ and $0 < H_\alpha < \infty$ is a constant which does not depend on T .

3. Proof of Theorem 1

A. $I_\alpha(\phi) < \infty \Rightarrow \phi \in UUC(Y_5^*)$

The proof is similar to the convergent half of Theorem 1 in [6] and some details will be omitted. We may assume without loss of generality, that $1 \leq \phi(t) \leq t^{1/4}$ for all large t . Define the increasing sequence T_k , $k \geq 0$ by

$$T_k = T_{k-1}(1 + \phi^{-2}(T_{k-1}))$$

where T_0 is chosen large enough so that $\phi(T_0) \geq 1$. Define

$$B_k = \left\{ \sup_{T_{k-1} \leq t \leq T_k} \sup_{0 \leq s \leq a(T_k)} (X(t+s) - X(t)) \geq \sigma(a(T_{k-1})) \phi(T_{k-1}) \right\}.$$

If T_0 is large we can use Lemma 5 to bound the probability of this event:

$$P(B_k) \leq C_\alpha \frac{T_k - T_{k-1}}{a(T_k)} \left(\frac{a(T_{k-1})}{a(T_k)} \right)^{1-\alpha} \phi^{\frac{1}{\alpha}-1}(T_{k-1}) \exp \left\{ -\frac{1}{2} \left(\frac{a(T_{k-1})}{a(T_k)} \right)^{2\alpha} \phi^2(T_{k-1}) \right\}$$

and since $a(T_k) \leq a(T_{k-1})(1 + \phi^{-2}(T_{k-1}))$ we obtain

$$\begin{aligned} P(B_k) &\leq \text{Const} \frac{T_k - T_{k-1}}{a(T_k)} \phi^{\frac{1}{\alpha}-1}(T_{k-1}) \exp \left\{ -\frac{\phi^2(T_{k-1})}{2} \right\} \\ &\leq \text{Const} \int_{T_{k-2}}^{T_{k-1}} \frac{(\phi(t))^{(1/\alpha)-1}}{a(t)} \exp \left\{ -\frac{\phi^2(t)}{2} \right\} dt \end{aligned}$$

Therefore $I_\alpha(\phi) < \infty \Rightarrow \sum_{n=0}^{\infty} P(B_n) < \infty$ and the Borel-Cantelli lemma implies by a standard argument that $\phi \in UUC(Y_5)$. A similar proof shows that $\phi \in UUC(Y_5^*)$.

B. $I_\alpha(\phi) = \infty \Rightarrow \phi \in ULC(Y_1)$

Let $\beta = \lim_{t \rightarrow \infty} \frac{a(t)}{t}$, if $\beta = 1$ then Theorem 5 in [15] implies the result. Assume $\beta < 1$ and define an increasing sequence $T_k \uparrow \infty$ by

$$0 < \frac{a(T_0)}{T_0} < \beta' = \frac{1 + \beta}{2}, \quad \phi(T_0) > 1$$

$$\xi(T_k) = T_k - a(T_k) = T_{k-1}, \quad k \geq 1.$$

We define also

$$\begin{aligned} \delta_k &= [\phi^{1/\alpha}(T_k)] & \delta'_k &= [\lambda \delta_k] & 0 < \lambda < 1; & k \geq 1 \\ \tau_{k,j} &= \frac{j a(T_{k-1})}{\delta_k} & t_{k,j} &= T_{k-1} + \tau_{k,j} & j = 1, \dots, \delta'_k; & k \geq 1 \\ G_k &= \left\{ \max_{1 \leq j \leq \delta'_k} Y_i(t_{k,j}) > \phi(T_k) \right\}. \end{aligned}$$

Suppose that s and t belong to the interval $[T_{k-1}, T_{k-1} + \lambda a(T_{k-1})]$ and $s \leq t$. We shall show that there exists a constant $D_\alpha > 0$ such that

$$E(Y_1(t) - Y_1(s))^2 \geq D_\alpha \frac{(t-s)^{2\alpha}}{a^{2\alpha}(T_{k-1})} \tag{2.8}$$

Start with

$$\begin{aligned} E(Y_1(t) - Y_1(s))^2 &= \frac{2\sigma(a_t)\sigma(a_s) + \sigma^2(t-s+a_t-a_s) + \sigma^2(t-s) - \sigma^2(a_t+t-s) - \sigma^2(a_s-t+s)}{\sigma(a_t)\sigma(a_s)} \end{aligned} \tag{2.9}$$

and consider

$$\sigma^2(a_t+t-s) + \sigma^2(a_s-t+s) = (a_t+t-s)^{2\alpha} + (a_s-(t-s))^{2\alpha}.$$

Using Taylor's theorem we get, for some $0 < \theta < (t-s)$, that this is

$$\begin{aligned} &= a_t^{2\alpha} + a_s^{2\alpha} + 2\alpha(a_t^{2\alpha-1} - a_s^{2\alpha-1})(t-s) + \alpha(2\alpha-1)(t-s)^2((a_t+\theta)^{2\alpha-2} \\ &\quad + (a_s-\theta)^{2\alpha-2}). \end{aligned} \tag{2.10}$$

If $0 < \alpha < 1/2$ this expression is bounded above by $a_t^{2\alpha} + a_s^{2\alpha}$ and

$$\begin{aligned} (2.9) &\geq \frac{(t-s+a_t-a_s)^{2\alpha} + (t-s)^{2\alpha} - (a_t^\alpha - a_s^\alpha)^2}{a_t^\alpha a_s^\alpha} \\ &\geq \frac{(t-s+a_t-a_s)^{2\alpha} + (t-s)^{2\alpha} - (a_t-a_s)^{2\alpha}}{a_t^\alpha a_s^\alpha} \\ &\geq \frac{(t-s)^{2\alpha}}{a_t^\alpha a_s^\alpha} \geq \frac{(t-s)^{2\alpha}}{a^{2\alpha}(T_{k-1})}. \end{aligned}$$

If $1/2 < \alpha < 1$ then $(a_t^{2\alpha-1} - a_s^{2\alpha-1}) \leq (a_t - a_s)^{2\alpha-1}$; also $\theta < (t-s) < \lambda a(T_{k-1}) \leq \lambda a_s \leq \lambda a_t$, and $a_t - a_s \leq t - s$, hence

$$\begin{aligned}
 (2.10) &\leq a_t^{2\alpha} + a_s^{2\alpha} + 2\alpha(t-s)^{2\alpha} + \alpha(2\alpha-1)(t-s)^{2\alpha} \left(\left(\frac{t-s}{a_t} \right)^{2-2\alpha} + \left(\frac{t-s}{(1-\lambda)a_s} \right)^{2-2\alpha} \right) \\
 &\leq a_t^{2\alpha} + a_s^{2\alpha} + (t-s)^{2\alpha} \left(2\alpha + \alpha(2\alpha-1) \left(\lambda^{2-2\alpha} + \left(\frac{\lambda}{1-\lambda} \right)^{2-2\alpha} \right) \right) \quad (2.11)
 \end{aligned}$$

On the other hand, since $1 < 2\alpha < 2$, we have $(t-s+a_t-a_s)^{2\alpha} \geq (t-s)^{2\alpha} + (a_t-a_s)^{2\alpha}$ and using this and (2.11) in (2.9) we get

$$\begin{aligned}
 (2.9) &\geq \frac{2a_t^\alpha a_s^\alpha + 2(t-s)^{2\alpha} + (a_t-a_s)^{2\alpha} - a_t^{2\alpha} - a_s^{2\alpha} - (t-s)^{2\alpha} (2\alpha + \alpha(2\alpha-1) (\lambda^{2-2\alpha} (1 + (1-\lambda)^{2\alpha-2})))}{a_t^\alpha a_s^\alpha}
 \end{aligned}$$

but $(a_t-a_s)^{2\alpha} \geq (a_t^\alpha - a_s^\alpha)^2 = a_t^{2\alpha} + a_s^{2\alpha} - 2a_t^\alpha a_s^\alpha$. Choosing $0 < \lambda < 1$ so that $D_\alpha = 2 - 2\alpha - \alpha(2\alpha-1)\lambda^{2-2\alpha}(1+(1-\lambda)^{2\alpha-2}) > 0$ we get

$$E(Y_1(t) - Y_1(s))^2 \geq D_\alpha \frac{(t-s)^{2\alpha}}{a^{2\alpha}(T_{k-1})}$$

and (2.8) is proved.

Now for $q \in [0, 1]$ define

$$V_k(q) = Y_1(T_{k-1} + q\lambda a(T_{k-1})).$$

Then (2.8) implies that for p and q in $[0, 1]$

$$E(V_k(p) - V_k(q))^2 \geq D_\alpha \frac{(|p-q|\lambda a(T_{k-1}))^{2\alpha}}{a^{2\alpha}(T_{k-1})} = D_\alpha \lambda^{2\alpha} |p-q|^{2\alpha}.$$

On the other hand

$$P(G_k) = P\left\{ \max_{1 \leq j \leq \delta'_k} V_k(j/\delta_k) > \phi(T_k) \right\}$$

and using Lemma 8 we see that if T_0 is sufficiently large there exists a positive constant C'_α such that for all k we have

$$\begin{aligned}
 P(G_k) &\geq C'_\alpha \delta_k \psi(\phi(T_k)) \\
 &\geq \text{Const } \phi^{\frac{1}{\alpha}-1}(T_k) \exp\left\{ -\frac{\phi^2(T_k)}{2} \right\} \\
 &\geq \text{Const } \frac{T_k - T_{k-1}}{a(T_k)} \phi^{\frac{1}{\alpha}-1}(T_k) \exp\left\{ -\frac{\phi^2(T_k)}{2} \right\} \\
 &\geq \int_{T_k}^{T_{k+1}} \frac{(\phi(t))^{(1/\alpha)-1}}{a(t)} \exp\left\{ -\frac{\phi^2(t)}{2} \right\} dt \quad (2.12)
 \end{aligned}$$

and $I_\alpha(\phi) = \infty \Rightarrow \sum_{n=0}^\infty P(G_n) = \infty$. To show that $P(G_k \text{ i.o.}) = 1$ we still have to prove (2.7). It is easy to see that

$$P(G_j \cap G_k) - P(G_j)P(G_k) = (P(G'_j \cap G'_k) - P(G'_j)P(G'_k))$$

where G' is the complement of G .

If $0 < \alpha < 1/2$ the variables $Y(t_{j,i})$ and $Y(t_{k,l})$ are negatively correlated if $j < k - 1$, and in that case by Lemma 6, $P(G'_j \cap G'_k) \leq P(G'_j)P(G'_k)$, which is sufficient for (2.7). If $1/2 < \alpha < 1$, using Lemma 6 we get

$$P(G_j \cap G_k) - P(G_j)P(G_k) \leq \sum_{l=0}^{\delta'_j} \sum_{m=0}^{\delta'_k} \rho(t_{j,l}; t_{k,m}) g(\phi(T_j), \phi(T_k); \rho^*(t_{j,l}; t_{k,m})) \quad (2.13)$$

where $\rho(t_{k,l}; t_{k,m}) = E(Y_1(t_{k,l}) Y_1(t_{j,m}))$. This correlation can be written as

$$\rho(t_{j,l}; t_{k,m}) = \frac{1}{2Q^\alpha R^\alpha} (P^{2\alpha} + H(Q, R)),$$

where $P = \sum_{i=j}^{k-1} a(T_i) + \frac{m}{\delta_k} a(T_{k-1}) - \frac{l}{\delta_j} a(T_{j-1}) - a(t_{j,l})$, $Q = a(t_{j,l})$, $R = a(t_{k,m})$ and $H(U, V) = (P + U + V)^{2\alpha} - (P + U)^{2\alpha} - (P + V)^{2\alpha}$. By Taylor's theorem

$$H(Q, R) = -P^{2\alpha} + 2\alpha(2\alpha - 1)QRP^{2(\alpha-1)} + S$$

where $S = \frac{2\alpha(2\alpha - 1)(2\alpha - 2)}{3!} ((Q + R)^3 (P + \theta R + \theta Q)^{2\alpha-3} - Q^3 (P + \theta Q)^{2\alpha-3} - R^3 (P + \theta R)^{2\alpha-3})$ for some $0 < \theta < 1$. It is easy to see that

$$S \leq \frac{2\alpha(2\alpha - 1)(2\alpha - 2)}{3!} Q^3 (P + \theta Q)^{2\alpha-3}$$

and

$$H(Q, R) \leq -P^{2\alpha} + 3\alpha(2\alpha - 1)QRP^{2\alpha-2}$$

whence

$$\begin{aligned} \rho(t_{j,l}; t_{k,m}) &\leq \text{Const} \frac{QRP^{2\alpha-2}}{Q^\alpha R^\alpha} \\ &\leq \text{Const} (a(t_{j,l}) a(t_{k,m}))^{1-\alpha} \left(\sum_{i=j+1}^{k-1} a(T_i) \right)^{2\alpha-2} \\ &\leq \text{Const} \left(\frac{a(T_j) a(T_k)}{\left(\sum_{i=j+1}^{k-1} a(T_i) \right)^2} \right)^{1-\alpha} \end{aligned}$$

since $\beta < 1$ we may assume, without loss of generality, that $a_1 < 1$, and then $T_k(1 - a_1) \leq T_{k-1}$ which implies $a(T_k) \leq (1 - a_1)^{-1} a(T_{k-1})$ and

$$\begin{aligned} \rho(t_{j,l}; t_{k,m}) &\leq \text{Const} \left(\frac{a(T_j)}{\sum_{i=j+1}^{k-1} a(T_i)} \right)^{1-\alpha} \\ &\leq \text{Const} (k-j-1)^{\alpha-1} \equiv \eta_{jk} \end{aligned}$$

as long as $k > j + 2$. We define

$$\rho_{jk} = \sup_{1 \leq l \leq \delta'_j} \sup_{1 \leq k \leq \delta'_k} \rho(t_{j,l}; t_{k,m}).$$

Using this we see that (2.13) is bounded by

$$\frac{\delta'_j \delta'_k \eta_{jk}}{2\pi(1 - \rho_{jk}^2)^{1/2}} \exp \left\{ - \frac{\phi^2(T_j) + \phi^2(T_k) - 2\hat{\rho}_{jk} \phi(T_j) \phi(T_k)}{2(1 - \hat{\rho}_{jk}^2)} \right\}$$

where $\hat{\rho}_{jk}$ is the value of $\rho(t_{j,l}; t_{k,m})$ which maximises the exponent. For our purpose it is sufficient to consider, for c fixed,

$$\begin{aligned} &\sum_{k=c}^n \sum_{j=1}^{k-1} [P(G_j \cap G_k) - P(G_j)P(G_k)] \\ &\leq \left(\sum_{k=c}^n \sum_{j=1}^{k-v_k} + \sum_{k=c}^n \sum_{j=k-v_k}^{k-1} \right) \frac{\delta'_j \delta'_k \eta_{jk}}{2\pi(1 - \rho_{jk}^2)^{1/2}} \exp \left\{ - \frac{\phi^2(T_j) + \phi^2(T_k) - 2\hat{\rho}_{jk} \phi(T_j) \phi(T_k)}{2(1 - \hat{\rho}_{jk}^2)} \right\} \end{aligned}$$

where $v_k = [(\phi(T_k))^{4/(1-\alpha)}]$. The first sum is bounded by

$$\sum_{k=c}^n \sum_{j=1}^{k-v_k} \frac{\eta_{jk} \phi(T_k) \phi(T_j)}{(1 - \rho_{jk}^2)^{1/2}} \delta_k \psi(\phi(T_k)) \delta_j \psi(\phi(T_j)) \exp \{ \phi^2(T_k) \eta_{jk} \}$$

and using (2.12) this is bounded by

$$\text{Const} \sum_{k=c}^n \sum_{j=1}^{k-v_k} \frac{\eta_{jk} \phi^2(T_k)}{(1 - \rho_{jk}^2)^{1/2}} \exp \{ \phi^2(T_k) \eta_{jk} \} P(G_k) P(G_j)$$

but $k - j \geq v_k$ implies $\eta_{jk} \phi^2(T_k) \leq \text{Const} v_k^{\alpha-1} \phi^2(T_k) \leq \text{Const} \phi^{-2}(T_k)$. Therefore, the sum above is bounded by

$$\text{Const } \phi^{-2}(T_c) \exp\{-\text{Const } \phi^{-2}(T_c)\} \sum_{k=c}^n \sum_{j=1}^{k-v_k} P(G_k) P(G_j)$$

and for c sufficiently large this is bounded by

$$\delta \left(\sum_{k=1}^n P(G_k) \right)^2.$$

The second sum is

$$\begin{aligned} & \sum_{k=c}^n \sum_{j=k-v_k}^{k-1} \frac{\delta_j \delta_k \eta_{jk}}{2\pi(1-\rho_{jk}^2)^{1/2}} \exp\left\{-\frac{\phi^2(T_j) + \phi^2(T_k) - 2\hat{\rho}_{jk} \phi(T_j) \phi(T_k)}{2(1-\hat{\rho}_{jk}^2)}\right\} \\ &= \sum_{k=c}^n \sum_{j=k-v_k}^{k-1} \frac{\eta_{jk} \phi^{\frac{1}{\alpha}}(T_k) \phi^{\frac{1}{\alpha}+1}(T_j) \psi(\phi(T_j))}{2\pi(1-\rho_{jk}^2)^{1/2}} \exp\left\{-\frac{(\phi(T_k) - \hat{\rho}_{jk} \phi(T_j))^2}{2(1-\hat{\rho}_{jk}^2)}\right\} \\ &\leq \text{Const} \sum_{k=c}^n \sum_{j=k-v_k}^{k-1} \phi^{\frac{1}{\alpha}+1}(T_k) \phi^{\frac{1}{\alpha}}(T_j) \psi(\phi(T_j)) \exp\left\{-\phi^2(T_k) \frac{(1-\rho)}{1+\rho}\right\} \end{aligned}$$

where ρ is the largest covariance among the terms considered. Let $B = \left(\frac{1-\rho}{1+\rho}\right)^2$

> 0 . Then if c is large enough this is

$$\begin{aligned} &\leq \text{Const} \sum_{k=c}^n v_k \phi^{\frac{1}{\alpha}+1}(T_k) \phi^{\frac{1}{\alpha}}(T_{k-v_k}) \psi(\phi(T_{k-v_k})) \exp\{-B\phi^2(T_k)\} \\ &\leq \text{Const} v_c \phi^{\frac{1}{\alpha}+1}(T_c) \exp\{-B\phi^2(T_c)\} \sum_{k=c-v_c}^n P(G_k) \\ &\leq \text{Const} \sum_{k=1}^n P(G_k). \end{aligned}$$

We have shown that given $\delta > 0$, for c large

$$\sum_{k=c}^n \sum_{j=j}^{k-1} P(G_j \cap G_k) - P(G_j) P(G_k) \leq \delta \left(\sum_{k=1}^n P(G_k) \right)^2 + \text{Const} \sum_{k=1}^n P(G_k)$$

and this is enough to prove (2.7). \square

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References

1. Berman, S.M.: Limit theorems for the maximum term in stationary sequences. *Ann. Math. Statist.* **35**, 502–516 (1964)
2. Berman, S.M.: An asymptotic bound for the distribution of the maximum of a Gaussian process. *Ann. Inst. Henri Poincaré* **21**, 47–57 (1985)
3. Berman, S.M.: The maximum of a Gaussian process with nonconstant variance. *Ann. Inst. Henri Poincaré* **21**, 383–391 (1985)

4. Csörgő, M., Révész, P.: How big are the increments of the Wiener process?. *Ann. Probab.* **7**, 731–743 (1979)
5. Orey, S.: Growth rate of certain Gaussian processes. In: *Proc. Sixth Berkeley, Symposium Math. Stat. Prob.*, vol. 2, pp. 443–451. Berkeley: University of California Press 1971
6. Ortega, J., Wschebor, M.: On the increments of the Wiener process. *Z. Wahrscheinlichkeitstheor. Verw. Geb.* **65**, 329–339 (1984)
7. Ortega, J.: On the size of the increments of non-stationary Gaussian processes. *Stochastic Proc. Appl.* **18**, 47–56 (1984)
8. Ortega, J.: Comportamiento asintótico de los incrementos de los procesos de Wiener fraccionarios. *Actas II Congreso Latinoamericano Prob. Est. Mat.*, pp. 195–215, Caracas: Equinoccio 1986
9. Plackett, R.L.: A reduction formula for normal multivariate integrals. *Biometrika* **41**, 351–360 (1954)
10. Qualls, C., Watanabe, H.: Asymptotic properties of Gaussian processes. *Ann. Math. Statist.* **43**, 580–596 (1972)
11. Qualls, C., Watanabe, H.: Asymptotic properties of Gaussian random fields. *Trans. Am. Math. Soc.* **177**, 155–171 (1973)
12. Rényi, A.: *Probability theory*. Amsterdam: North-Holland 1970
13. Révész, P.: On the increments of Wiener and related processes. *Ann. Probab.* **10**, 613–627 (1982)
14. Slepian, D.: The one-sided barrier problem for Gaussian noise. *Bell System Tech. J.* **41**, 463–501 (1962)
15. Watanabe, H.: An asymptotic property of Gaussian processes. *Trans. Am. Math. Soc.* **148**, 233–248 (1970)

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