

Limit Distribution of the Last Exit Time for Stationary Random Sequences

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Summary. For a strictly stationary random sequence $(X_i)_{i \geq 0}$ we find sufficient conditions such that the distribution of the last exit time $t_\beta = \max\{i: X_i > \beta i\}$ ($\beta > 0$) tends weakly to a nondegenerate limit distribution as $\beta \downarrow 0$.

1. Introduction and Result

Let $(X_i)_{i \geq 0}$ be a sequence of identically distributed random variables. Define the *last exit time* by $t_\beta = t_\beta(X) = \max\{i \geq 0: X_i > \beta i\}$, if such an i exists, and $t_\beta = 0$ else, $\beta > 0$. In [2] we proved that its limit distribution for $\beta \downarrow 0$ is highly connected with the limit distribution of the maximum $Z_n = \max_{i \leq n} X_i$. In this paper we derive limit distributions of t_β for strictly stationary sequences of not necessarily independent random variables. We give sufficient conditions in order that the limit of $P\{\beta t_\beta(X) \leq a_\beta x + b_\beta\}$ is the same as the limit of $P\{\beta t_\beta(X^*) \leq a_\beta x + b_\beta\}$, where X^* is a sequence of independent, identically distributed random variables $(X_i^*)_{i \geq 0}$ with the same marginal distribution as X_i :

$$P\{X_i \leq x\} = P\{X_i^* \leq x\} = F(x) \quad \text{for all } i \geq 0,$$

and a_β and b_β are norming values, depending on F and β (see [2]). According to the mentioned connection with extremal theory, we use some ideas from Leadbetter [4].

We assume that the sequence X is strictly stationary, which means that the finite-dimensional distributions

$$F_{i_1, \dots, i_n}(x_1, \dots, x_n) = P\{X_{i_1} \leq x_1, \dots, X_{i_n} \leq x_n\}$$

fulfill the condition

$$F_{i_1+l, \dots, i_n+l}(x_1, \dots, x_n) = F_{i_1, \dots, i_n}(x_1, \dots, x_n)$$

for all $l \geq 0$, $i_j \geq 0$, $n \geq 1$ and $x_j \in \mathbb{R}$. Therefore we have $F_i(x) = F(x)$ for all $i \geq 0$.

We also assume $EX_i^+ < \infty (\forall i \geq 0)$, which implies that t_β a.s. Thus t_β is well-defined.

In the following we set $\bar{F}(x) = 1 - F(x)$, $x_0 = \sup \{x : F(x) < 1\} \leq \infty$ and $H(x) = \int_x^{x_0} \bar{F}(y) dy$. The integral exists since $EX_i^+ < \infty$.

We prove the theorem under the conditions (A) and (B). Condition (A) restricts the dependence of events which are widely separated in time; it is some kind of a mixing condition. Condition (B) considers the dependence of events which are close in time. Condition (A) and (B) are fulfilled by a stationary Gaussian sequence with correlation function $r(n)$ where $r(n) \log n \rightarrow 0$ as $n \rightarrow \infty$. This verification is given in [3]. The result in this case which was proved in [1], is in a way the best possible, since if $r(n) \log n \rightarrow \gamma > 0$ as $n \rightarrow \infty$, the limit distribution of t_β is no more equal to $\exp(-e^{-x})$, which is the limit distribution of t_β for independent normal variables.

We formulate Condition (A) and (B) for any sequence u_N which will be later of the form $a_{1/N}x + b_{1/N} (N \rightarrow \infty)$.

Condition (A). Let for $n, K, m \in \mathbb{N}$ with $m < K$ and $N = nK$:

$$A_{l,n} = \left\{ X_{lK+i} \leq u_N + \frac{lK+i}{N}, i = 1, \dots, K-m \right\}, \quad l \geq 0,$$

$$\left| P \left(\bigcap_{k=0}^l A_{k,n} \right) - P \left(\bigcap_{k=0}^{l-1} A_{k,n} \right) P(A_{l,n}) \right| \leq \alpha_{n,m,K,l}.$$

Then we assume

$$\sum_{l=1}^{\infty} \alpha_{n,m,K,l} < \infty \quad \text{and} \quad \lim_{\substack{K \rightarrow \infty \\ m_K \rightarrow \infty}} \lim_{n \rightarrow \infty} \sum_{l=1}^{\infty} \alpha_{n,m_K,K,l} = 0,$$

where $m_K/K \rightarrow 0$ as $K \rightarrow \infty$.

Condition (B). Let n, K, m, N be as in Condition (A) and

$$\sum_{1 \leq i < j \leq K-m} P \left\{ X_i > u_N + \frac{lK+i}{N}, X_j > u_N + \frac{lK+j}{N} \right\} = \alpha_{n,m,K,l}^*.$$

Then we assume $\lim_{K \rightarrow \infty} \lim_{n \rightarrow \infty} \sum_{l=0}^{\infty} \alpha_{n,m,K,l}^* = 0$ for all sufficiently large m .

Instead of Condition (B) we could assume Condition (B')

$NH(u_N) = O(1)$ and

$$\lim_{K \rightarrow \infty} \lim_{n \rightarrow \infty} \sum_{l=0}^{\infty} \sum_{1 \leq i < j \leq K-m} \left| P \left\{ X_i > u_N + \frac{lK+i}{N}, X_j > u_N + \frac{lK+j}{N} \right\} - \bar{F} \left(u_N + \frac{lK+i}{N} \right) \bar{F} \left(u_N + \frac{lK+j}{N} \right) \right| = 0.$$

For

$$\begin{aligned} & \sum_{l=0}^{\infty} \sum_{1 \leq i < j \leq K-m} \bar{F}\left(u_N + \frac{lK+i}{N}\right) \bar{F}\left(u_N + \frac{lK+j}{N}\right) \\ & \leq \sum_{l=0}^{\infty} \left(\sum_{i=1}^K \bar{F}\left(u_N + \frac{lK+i}{N}\right) \right)^2 \\ & \leq \sum_{l=0}^{\infty} N^2 \left\{ H\left(u_N + \frac{lK}{N}\right) - H\left(u_N + \frac{lK+K}{N}\right) \right\}^2. \end{aligned}$$

Because of the monotony of the integrand

$$H\left(u_N + \frac{lK}{N}\right) - H\left(u_N + \frac{lK+K}{N}\right) \leq \int_{u_N}^{u_N + K/N} \bar{F}(y) dy \leq \frac{1}{n} \bar{F}(u_N),$$

the above sum is smaller than

$$N^2 \frac{1}{n} \bar{F}(u_N) \sum_{l=0}^{\infty} \left\{ H\left(u_N + \frac{l}{n}\right) - H\left(u_N + \frac{l+1}{n}\right) \right\} \leq K \bar{F}(u_N) N H(u_N) \rightarrow 0$$

as $n \rightarrow \infty$ and $K \rightarrow \infty$, for any m .

To simplify our formulation of the theorem we define

Condition (C). We assume that $F(x)$ is such that as $\beta \downarrow 0$ $P\{\beta t_{\beta}(X^*) \leq a_{\beta}x + b_{\beta}\}$ tends weakly to a nondegenerate limit distribution $\phi(x)$, where X^* is a sequence of independent, identically distributed random variables with distribution $F(x)$ and where $a_{\beta} > 0$ and b_{β} are norming values.

For the discussion of these limit distributions see [2].

Theorem. Let $(X_i)_{i \geq 0}$ be a stationary random sequence with $EX_i^+ < \infty$ and the marginal distribution $F(x)$, such that the conditions (A), (B) and (C) are fulfilled with $u_n = a_{1/n}x + b_{1/n}$, where $a_{1/n}$ and $b_{1/n}$ are the values of Condition (C), $\forall x \in \mathbb{R}$. Then

$$\lim_{\beta \downarrow 0} P\{\beta t_{\beta} \leq u_{[1/\beta]}\} = \phi(x)$$

for all points of continuity x of ϕ .

The proof of the theorem is given in the next section by the following steps:

- 1) we may restrict our calculation to sequences $\beta_n = n^{-1}$,
- 2) the probability of the event $\{t_{1/n} \leq nu_n\}$ is asymptotically equal to $P\left(\bigcap_{l=0}^{\infty} A_{l,n}\right)$, $A_{l,n}$ as in Condition (A),
- 3) $P\left(\bigcap_{l=0}^{\infty} A_{l,n}\right)$ is asymptotically equal to the product $\prod_{l=0}^{\infty} P(A_{l,n})$,
- 4) the last product tends to $\phi(x)$.

2. Proof of the Theorem

As mentioned we give the proof by showing the four steps.

Lemma 1. *If $u_n \rightarrow x_0$ and $\lim_{n \rightarrow \infty} P\{t_{1/n} \leq nu_n\} = \theta > 0$, then with $u^*(\beta) = u_n$ for $n \leq 1/\beta < n + 1$*

$$\lim_{\beta \downarrow 0} P\{\beta t_\beta \leq u^*(\beta)\} = \theta.$$

Proof. Since $P\{\beta t_\beta \leq u^*(\beta)\} = P\{X_i \leq \beta i, i > \beta^{-1} u^*(\beta)\}$ we have

$$\begin{aligned} P\{\beta t_\beta \leq u^*(\beta)\} &\leq P\{X_i \leq u^*(\beta) + \beta i, i \geq 1\} \\ &\geq P\{X_i \leq u^*(\beta) + \beta(i - 1), i \geq 1\} \end{aligned}$$

by using the stationarity of (X_i) and the monotony of the boundary. The difference between the two approximations tends to zero, since with the Bonferoni inequality

$$\begin{aligned} 0 &\leq P\{X_i \leq u^*(\beta) + \beta i, i \geq 1\} - P\{X_i \leq u^*(\beta) + \beta(i - 1), i \geq 1\} \\ &\leq \sum_{i=1}^{\infty} P\{u^*(\beta) + \beta(i - 1) < X_i \leq u^*(\beta) + \beta i\} \\ &= \sum_{i=1}^{\infty} \{\bar{F}(u^*(\beta) + \beta(i - 1)) - \bar{F}(u^*(\beta) + \beta i)\} = \bar{F}(u^*(\beta)) \rightarrow 0 \quad \text{as } \beta \downarrow 0. \end{aligned}$$

Let n be such that $n \leq 1/\beta < n + 1$. With the same argument we estimate

$$P\{X_i \leq u^*(\beta) + \beta i, i \geq 1\} \leq P\left\{X_i \leq u_n + \frac{i}{n}, i \geq 1\right\} = P\{t_{1/n} \leq nu_n\} + o(1)$$

and conversely

$$\begin{aligned} P\{X_i \leq u^*(\beta) + \beta i, i \geq 1\} &\geq P\left\{X_i \leq u_n + \frac{i}{n+1}, i \geq 1\right\} \\ &\geq P\left\{X_i \leq u_n + \frac{i}{n}, i \geq 1\right\} - \sum_{i=1}^{\infty} P\left\{u_n + \frac{i}{n+1} < X_i \leq u_n + \frac{i}{n}\right\} \\ &= P\{t_{1/n} \leq nu_n\} + o(1) \end{aligned}$$

since the sum is equal to

$$\begin{aligned} &\sum_{i=1}^{\infty} \left\{ \bar{F}\left(u_n + \frac{i}{n+1}\right) - \bar{F}\left(u_n + \frac{i}{n}\right) \right\} \\ &\leq (n+1) \int_{u_n}^{x_0} \bar{F}(y) dy - n \int_{u_n}^{x_0} \bar{F}(y) dy + \bar{F}(u_n) \\ &= H(u_n) + \bar{F}(u_n) = o(1) \quad \text{as } n \rightarrow \infty. \end{aligned}$$

Remark. For $u^*(\beta)$ we may obviously use any measurable function which is bounded by u_{n+1} and u_n on $[(n+1)^{-1}, n^{-1}]$.

Lemma 2. Let $NH(u_N) \rightarrow \theta > 0$ as $N \rightarrow \infty$, $A_{l,n}$ and n, m, K, N be as in Condition (A). Then

$$P \left\{ X_i \leq u_N + \frac{i}{N}, i \geq 1 \right\} - P \left(\bigcap_{l=0}^{\infty} A_{l,n} \right) = o(1) \quad \text{as } n \rightarrow \infty, K \rightarrow \infty.$$

Proof. Define $A_{l,n}^* = \left\{ X_{lK+i} \leq u_N + \frac{lK+i}{N}, i = K-m+1, \dots, K \right\}, l \geq 0$. Thus

$$P \left\{ X_i \leq u_N + \frac{i}{N}, i \geq 1 \right\} = P \left(\bigcap_{l=0}^{\infty} (A_{l,n} \cap A_{l,n}^*) \right).$$

Therefore we have to bound the difference

$$\begin{aligned} P \left(\bigcap_{l=0}^{\infty} A_{l,n} \right) - P \left(\bigcap_{l=0}^{\infty} (A_{l,n} \cap A_{l,n}^*) \right) &\leq \sum_{l=0}^{\infty} P(A_{l,n} \cap A_{l,n}^{*c}) \\ &\leq \sum_{l=0}^{\infty} P(A_{l,n}^{*c}) \leq \sum_{l=0}^{\infty} \sum_{i=K-m+1}^K \bar{F} \left(u_N + \frac{lK+i}{N} \right) \leq \sum_{l=0}^{\infty} m \bar{F} \left(u_N + \frac{l}{n} \right) \\ &\quad \text{by using the monotony of } \bar{F} \\ &\leq m \bar{F}(u_N) + mn \int_{u_N}^{x_0} \bar{F}(y) dy = o(1) + mnH(u_N) \\ &= o(1) + m \frac{\theta + o(1)}{K} = o(1) \quad \text{as } n \rightarrow \infty, K \rightarrow \infty. \end{aligned}$$

Lemma 3. Assume Condition (A), then as $n \rightarrow \infty, K \rightarrow \infty, m_K \rightarrow \infty$ with $m_K/K \rightarrow 0$

$$P \left(\bigcap_{l=0}^{\infty} A_{l,n} \right) - \prod_{l=0}^{\infty} P(A_{l,n}) = o(1).$$

Proof. Let L be a positive integer. Then by adding and subtracting we derive

$$\begin{aligned} \left| P \left(\bigcap_{l=0}^L A_{l,n} \right) - \prod_{l=0}^L P(A_{l,n}) \right| &= \left| P \left(\bigcap_{l=0}^L A_{l,n} \right) - P \left(\bigcap_{l=0}^{L-1} A_{l,n} \right) P(A_{L,n}) \right. \\ &\quad + P \left(\bigcap_{l=0}^{L-1} A_{l,n} \right) P(A_{L,n}) - P \left(\bigcap_{l=0}^{L-2} A_{l,n} \right) P(A_{L-1,n}) P(A_{L,n}) + \dots \\ &\quad \left. + P \left(\bigcap_{l=0}^1 A_{l,n} \right) P(A_{2,n}) P(A_{3,n}) \dots P(A_{L,n}) - \prod_{l=0}^L P(A_{l,n}) \right| \\ &\leq \sum_{l=1}^L \left| P \left(\bigcap_{l'=0}^l A_{l',n} \right) - P \left(\bigcap_{l'=0}^{l-1} A_{l',n} \right) P(A_{l,n}) \right| \\ &\leq \sum_{l=1}^L \alpha_{n,m,K,l} \leq \sum_{l=1}^{\infty} \alpha_{n,m,K,l} = o(1) \end{aligned}$$

by assumption for any L .

Lemma 4. *Assume Condition (B) and $NH(u_N) \rightarrow \theta > 0$. Then*

$$\lim_{K \rightarrow \infty} \lim_{n \rightarrow \infty} \prod_{l=0}^{\infty} P(A_{l,n}) = e^{-\theta}.$$

Proof. First we approximate $\sum_{l=0}^{\infty} (1 - P(A_{l,n}))$.

$$\begin{aligned} 1 - P(A_{l,n}) &= P \left\{ \exists i: X_i > u_N + \frac{lK+i}{N} \text{ with } 1 \leq i \leq K-m \right\} \\ &\leq \sum_{i=1}^{K-m} \bar{F} \left(u_N + \frac{lK+i}{N} \right) \leq N \int_0^{1/n} \bar{F} \left(u_N + \frac{l}{n} + y \right) dy \\ &= N \left\{ H \left(u_N + \frac{l}{n} \right) - H \left(u_N + \frac{l+1}{n} \right) \right\} \end{aligned}$$

by using again the monotony of \bar{F} . Therefore

$$\begin{aligned} \sum_{l=0}^{\infty} (1 - P(A_{l,n})) &\leq N \sum_{l=0}^{\infty} \left\{ H \left(u_N + \frac{l}{n} \right) - H \left(u_N + \frac{l+1}{n} \right) \right\} = NH(u_N) \rightarrow \theta \quad \text{as } n \rightarrow \infty. \end{aligned}$$

With the inequality of Bonferroni we approximate conversely

$$\begin{aligned} 1 - P(A_{l,n}) &\geq \sum_{i=1}^{K-m} \bar{F} \left(u_N + \frac{lK+i}{N} \right) - \sum_{1 \leq i < j \leq K-m} P \left\{ X_i > u_N + \frac{lK+i}{N}, X_j > u_N + \frac{lK+j}{N} \right\} \\ &\geq N \left\{ H \left(u_N + \frac{l}{n} \right) - H \left(u_N + \frac{l+1}{n} \right) \right\} - (m+1) \bar{F} \left(u_N + \frac{l}{n} \right) - \alpha_{n,m,K,l}^*. \end{aligned}$$

Hence by assumption

$$\sum_{l=0}^{\infty} (1 - P(A_{l,n})) \geq NH(u_N) + o(1) \quad \text{as } n \rightarrow \infty$$

since

$$\sum_{l=0}^{\infty} (m+1) \bar{F} \left(u_N + \frac{l}{n} \right) \leq o(1) + (m+1)nH(u_N) = o(1) \quad \text{as } n \rightarrow \infty, K \rightarrow \infty.$$

Finally we need to know that each term $P(A_{l,n}) \rightarrow 1$. This is established by using the monotony of the boundary

$$1 - P(A_{l,n}) \leq 1 - P(A_{0,n}) \leq \sum_{i=1}^{K-m} \bar{F} \left(u_N + \frac{i}{N} \right) \leq (K-m) \bar{F}(u_N) = o(1).$$

This implies now that

$$\prod_{l=0}^{\infty} P(A_{l,n}) = \exp \left\{ -(1 + o(1)) \sum_{l=0}^{\infty} (1 - P(A_{l,n})) \right\} \rightarrow e^{-\theta}$$

as $n \rightarrow \infty, K \rightarrow \infty$.

Proof of the Theorem. Let $\theta = -\log \phi(x) > 0$ for a point of continuity x of ϕ . Condition C implies $nH(u_n) \xrightarrow{n \rightarrow \infty} \theta$ (see [2]). By Lemma 1 it is sufficient to show that

$$P\{t_{1/n} \leq nu_n\} \rightarrow e^{-\theta} \text{ as } n \rightarrow \infty.$$

Lemma 2, 3 and 4 imply that $P\left\{X_i \leq u_N + \frac{i}{N}, i \geq 1\right\} \rightarrow e^{-\theta}$ as $n \rightarrow \infty$, $K \rightarrow \infty$ and $m_K \rightarrow \infty$ with $m_K/K \rightarrow 0$, $N = nK$. By the argument of the proof of Lemma 1 it follows that also $P\{t_{1/N} \leq Nu_N\} \rightarrow e^{-\theta}$. Therefore the proof is complete if we show that

$$P\left\{X_i \leq u_n + \frac{i}{n}, i \geq 1\right\} - P\left\{X_i \leq u_{rK} + \frac{i}{rK}, i \geq 1\right\} = o(1) \text{ as } n \rightarrow \infty,$$

with $rK \leq n < rK + K$, K any positive integer, $r \rightarrow \infty$. We estimate this difference with the argument of the proof of Lemma 1. The difference is bounded by the sum of

$$\left| P\left\{X_i \leq u_n + \frac{i}{n}, i \geq 1\right\} - P\left\{X_i \leq u_n + \frac{i}{rK}, i \geq 1\right\} \right|$$

and

$$\left| P\left\{X_i \leq u_n + \frac{i}{rK}, i \geq 1\right\} - P\left\{X_i \leq u_{rK} + \frac{i}{rK}, i \geq 1\right\} \right|.$$

The first term is bounded with the Bonferroni inequality by

$$\begin{aligned} \sum_{i=1}^{\infty} P\left\{u_n + \frac{i}{n} < X_i \leq u_n + \frac{i}{rK}\right\} &\leq \sum_{i=1}^{\infty} \left\{ \bar{F}\left(u_n + \frac{i}{n}\right) - \bar{F}\left(u_n + \frac{i}{rK}\right) \right\} \\ &\leq nH(u_n) - rKH(u_n) + \bar{F}(u_n) = o(1), \end{aligned}$$

since $nH(u_n) \rightarrow \theta$ and $rK/n \rightarrow 1$. For the second term we have to consider the two cases: $u_{rK} \leq u_n$ and $u_{rK} > u_n$. In the first case the term is bounded again with the same argument by

$$\sum_{i=1}^{\infty} P\left\{u_{rK} + \frac{i}{rK} < X_i \leq u_n + \frac{i}{rK}\right\} \leq rKH(u_{rK}) - rKH(u_n) + \bar{F}(u_n) = o(1)$$

Similar in the second case, the term is bounded by

$$rKH(u_n) - rKH(u_{rK}) + \bar{F}(u_{rK}) = o(1).$$

This finishes the proof of the Theorem.

Remark. As mentioned in [2] there is no essential restriction in considering the linear function as boundary instead of more general boundary functions $f(v)$ in the definition of the last exit time.

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