

Lindeberg Functions and the Law of the Iterated Logarithm*

R.J. Tomkins

Department of Mathematics and Statistics, University of Regina,
Regina, Saskatchewan S4S0A2, Canada

Summary. For a sequence of independent random variables $\{X_n\}$ with zero means and finite variances, define $S_n = \sum_{j=1}^n X_j$, $s_n^2 = E(S_n^2)$ and $t_n^2 = 2 \log \log s_n^2$; assume $s_n \rightarrow \infty$. Kolmogorov's law of the iterated logarithm asserts that $\limsup_{n \rightarrow \infty} S_n / (s_n t_n) = 1$ a.s. if $t_n |X_n| \leq \varepsilon_n s_n$ for some real sequence $\varepsilon_n \rightarrow 0$. This paper will show that, under the weaker condition $t_n X_n / s_n \rightarrow 0$ a.s., the a.s. limiting value of $\limsup_{n \rightarrow \infty} S_n / (s_n t_n)$ depends on the limiting behaviour of the modified Lindeberg functions

$$s_n^{-2} \sum_{j=1}^n E(X_j^2 I(|X_j| \leq \varepsilon s_j t_j^{-1})), \quad \text{where } \varepsilon > 0.$$

1. Introduction

Consider a sequence X_1, X_2, \dots of independent random variables (r.v.) with $E(X_n) = 0$ and $E(X_n^2) < \infty$ for $n \geq 1$. Define $S_n = \sum_{j=1}^n X_j$, $s_n^2 = E(S_n^2)$ and $t_n = (2 \log \log s_n^2)^{1/2}$. Assume $s_n \rightarrow \infty$.

According to Kolmogorov's law of the iterated logarithm (LIL), if a positive sequence $\varepsilon_n \rightarrow 0$ exists such that

$$t_n |X_n| \leq \varepsilon_n s_n \text{ almost surely (a.s.) for all } n \geq 1, \quad (1.1)$$

then

$$A \equiv \limsup_{n \rightarrow \infty} S_n / (2 s_n^2 \log \log s_n^2)^{1/2} = 1 \quad \text{a.s.} \quad (1.2)$$

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This paper will investigate the value of λ when (1.1) is replaced by the weaker condition:

$$t_n X_n / s_n \rightarrow 0 \quad \text{a.s.} \tag{1.3}$$

or, equivalently (by the Borel Zero-One Law),

$$\sum_{n=1}^{\infty} P[|X_n| > \varepsilon s_n t_n^{-1}] < \infty \quad \text{for every } \varepsilon > 0. \tag{1.4}$$

Under the less restrictive assumption (1.3), λ need not be one. For example, suppose $P[X_n = \pm n] = (2n^2)^{-1}$ and $P[X_n = 0] = 1 - n^{-2}$ for $n \geq 1$. Then $s_n^2 = n$ and, for $\varepsilon > 0$,

$$\sum_{n=1}^{\infty} P[|X_n| > \varepsilon s_n t_n^{-1}] \leq \sum_{n=1}^{\infty} P[X_n \neq 0] = \sum_{n=1}^{\infty} n^{-2} < \infty,$$

establishing (1.4). But, since $\sum_{n=1}^{\infty} P[X_n \neq 0] < \infty$, $P[X_n \neq 0 \text{ infinitely often (i.o.)}] = 0$ by the Borel-Cantelli lemma. Consequently, $\sum_{n=1}^{\infty} X_n$ converges a.s. so that, trivially, $S_n / (s_n t_n) \rightarrow 0$ a.s.; i.e. $\lambda = 0$.

Theorem 1 of Teicher [6] implicitly suggests a relationship between the value of λ and the Lindeberg functions

$$L_n(x) = s_n^{-2} \sum_{j=1}^n E(X_j^2 I(|X_j| > x s_j)),$$

where $I(A)$ denotes the indicator function of the event A . Teicher's LIL asserts that $\lambda = 1$ if, for some $\delta > 0$, $\sum_{n=1}^{\infty} P[|X_n| > \delta s_n t_n] < \infty$ and, for every $\varepsilon > 0$,

$$\sum_{n=1}^{\infty} (s_n t_n)^{-2} E(X_n^2 I(\varepsilon s_n t_n^{-1} < |X_n| \leq \delta s_n t_n)) < \infty$$

and

$$\lim_{n \rightarrow \infty} L_n(\varepsilon t_n^{-1}) = 0. \tag{1.5}$$

Since (1.4) implies Teicher's first two hypotheses, (1.2) must hold when (1.3) and (1.5) hold.

Rather than deal with $L_n(x)$ or $L_n(x t_n^{-1})$, it will be more convenient in this paper to work with the functions

$$H_n(x) \equiv s_n^{-2} \sum_{j=1}^n E(X_j^2 I(|X_j| \leq x s_j t_j^{-1})). \tag{1.6}$$

Here is the main result of the paper.

Theorem 1.1. *Let X_1, X_2, \dots be independent r.v. with $E(X_n) = 0$ and $E(X_n^2) < \infty$ for all $n \geq 1$. For $n \geq 1$, define $S_n = \sum_{j=1}^n X_j$ and $s_n^2 = E(S_n^2)$. Suppose $s_n \rightarrow \infty$ and*

$(\log \log s_n^2)^{\frac{1}{2}} |X_n|/s_n \rightarrow 0$ a.s. Then

$$H_- \leq \limsup_{n \rightarrow \infty} S_n / (2s_n^2 \log \log s_n^2)^{\frac{1}{2}} \leq H_+ \quad \text{a.s.}, \tag{1.7}$$

where H_- and H_+ are numbers satisfying $0 \leq H_- \leq H_+ \leq 1$ and (cf. (1.6))

$$H_-^2 \equiv \liminf_{n \rightarrow \infty} H_n(x), \quad H_+^2 = \limsup_{n \rightarrow \infty} H_n(x); \tag{1.8}$$

the values H_- and H_+ are independent of x .

Theorem 1.1 will be proved in Sect. 3, following the establishment in Sect. 2 of several lemmas, some of which may be of interest in themselves. Section 4 will contain a number of examples which pertain to the main result.

2. Some Preliminary Results

The following lemma is known (cf. Egorov [2], p. 512).

Lemma 2.1. *Let $\{a_n(\varepsilon), n \geq 1\}$ be a sequence of non-negative functions defined for all $\varepsilon > 0$.*

(i) *If $\lim_{n \rightarrow \infty} a_n(\varepsilon) = 0$ for every $\varepsilon > 0$, then a sequence $\{\varepsilon_n\}$ exists such that $\varepsilon_n \downarrow 0$ and $a_n(\varepsilon_n) \rightarrow 0$ as $n \rightarrow \infty$.*

(ii) *If $\sum_{n=1}^{\infty} a_n(\varepsilon) < \infty$ for every $\varepsilon > 0$, then there exists a sequence $\{\varepsilon_n\}$ such that $\varepsilon_n \downarrow 0$ and $\sum_{n=1}^{\infty} a_n(\varepsilon_n) < \infty$.*

Lemma 2.2. *Let $\{a_n(\varepsilon)\}$ be a sequence of non-negative functions, defined for all $\varepsilon > 0$. Define $a^* = \liminf_{\varepsilon \downarrow 0} \liminf_{n \rightarrow \infty} a_n(\varepsilon)$. Then there exists a sequence $\{\varepsilon_n\}$ such that $\varepsilon_n \downarrow 0$ and $\liminf_{n \rightarrow \infty} a_n(\varepsilon_n) \geq a^*$. Moreover, if $a_n(\varepsilon)$ is a non-decreasing function of ε for each $n \geq 1$, then $\liminf_{n \rightarrow \infty} a_n(\delta_n) \leq a^*$ for every real sequence $\{\delta_n\}$ satisfying $\delta_n \downarrow 0$.*

Proof. Define $f_n(\varepsilon) = \inf_{m \geq n} a_m(\varepsilon)$ for $n \geq 1$ and let $a(\varepsilon) = \liminf_{n \rightarrow \infty} a_n(\varepsilon)$ where $\varepsilon > 0$. Then $a(\varepsilon) - f_n(\varepsilon) \downarrow 0$ as $n \rightarrow \infty$ for every $\varepsilon > 0$. By Lemma 2.1 (i), a sequence $\{\varepsilon_n\}$ exists such that $\varepsilon_n \downarrow 0$ and $\lim_{n \rightarrow \infty} [a(\varepsilon_n) - f_n(\varepsilon_n)] = 0$. However, $\liminf_{n \rightarrow \infty} a(\varepsilon_n) \geq a^*$, so $\lim_{n \rightarrow \infty} f_n(\varepsilon_n) \geq a^*$. However, $a_n(\varepsilon_n) \geq f_n(\varepsilon_n)$ so $\liminf_{n \rightarrow \infty} a_n(\varepsilon_n) \geq a^*$.

If $a_n(\varepsilon)$ is non-decreasing for each $n \geq 1$, then, for every $\varepsilon > 0$, $a_n(\delta_n) \leq a_n(\varepsilon)$ for all large n if $\delta_n \downarrow 0$. Hence $\liminf_{n \rightarrow \infty} a_n(\delta_n) \leq \liminf_{n \rightarrow \infty} a_n(\varepsilon)$ and the proof is completed by letting $\varepsilon \downarrow 0$. \square

Using Hartman's modification [3] of Kolmogorov's techniques, the final lemma will establish a refinement of Corollary 1 of Teicher [8] and of Lemma 1 of Tomkins [9].

Lemma 2.3. Let $(M_n, \mathcal{F}_n, n \geq 1)$ be a submartingale and let $\{\alpha_n\}$, $\{B_n\}$ and $\{c_n\}$ be positive real sequences. Suppose $B_n \uparrow \infty$ and define $\alpha = \limsup_{n \rightarrow \infty} \alpha_n$, $\gamma = \limsup_{n \rightarrow \infty} (\log \log B_n^2)^{\frac{1}{2}} c_n$, and $g(x) = x^{-2}(e^{-x} - 1 - x)$. Assume $\alpha < \infty$ and $\gamma < \infty$.

If positive numbers C , N and T exist such that

$$E \exp\{tM_n/(\alpha_n B_n)\} \leq C \exp\{t^2 g(tc_n)\} \tag{2.1}$$

whenever $n \geq N$ and $0 \leq tc_n < T$, then

$$\limsup_{n \rightarrow \infty} \frac{M_n}{(B_n^2 \log \log B_n^2)^{\frac{1}{2}}} \leq \alpha K_\gamma \quad \text{a.s.} \tag{2.2}$$

where

$$K_0 = 2^{\frac{1}{2}} \quad \text{and} \quad K_\gamma = \min_{0 < b \leq T\gamma^{-1}} (b^{-1} + bg(\gamma b)) \quad \text{for } \gamma > 0. \tag{2.3}$$

Proof. Fix b such that $0 < b \leq T/\gamma$ ($b > 0$ if $\gamma = 0$). Pick any $\alpha' > \alpha$ and let $\delta = \alpha'(b^{-1} + bg(\gamma b))$. Then choose $c > 1$ so close to 1 that $\delta' \equiv \delta c^{-2} > \alpha'(b^{-1} + bg(\gamma b))$; consequently $b\delta'/\alpha' - b^2 g(\gamma b) > 1$. Define $n_0 = 1$ and, for $k \geq 1$, $n_k = \min\{n | B_n \geq cB_{n_{k-1}}\}$, $M_k^* = \max_{n_{k-1} \leq n < n_k} M_n$ and $m_k = n_k - 1$. Let $b_n = (\log \log B_n^2)^{\frac{1}{2}}$ for $n \geq 1$.

Choose any $\gamma' > \gamma$ so close to γ that $\lambda \equiv b\delta'/\alpha' - b^2 g(\gamma' b) > 1$. Then, for all k so large that $m_k > N$, $\alpha_{m_k} < \alpha'$ and $b_{m_k} c_{m_k} < \gamma'$, it follows from Markov's inequality, Doob's inequality ([1], p. 314) and (2.1) that

$$\begin{aligned} P[M_k^* \geq \delta' B_{m_k} b_{m_k}] &\leq P[\exp\{b b_{m_k} M_k^*/(\alpha_{m_k} B_{m_k})\} \geq \exp\{b \delta' b_{m_k}^2/\alpha'\}] \\ &\leq C \exp\{-b \delta' b_{m_k}^2/\alpha' + b^2 b_{m_k}^2 g(b b_{m_k} c_{m_k})\} \\ &\leq C \exp\{-b \delta' b_{m_k}^2/\alpha' + b^2 b_{m_k}^2 g(\gamma' b)\} \\ &= C \exp\{-\lambda b_{m_k}^2\} = C (\log B_{m_k}^2)^{-\lambda} \leq C (\log B_{n_{k-1}}^2)^{-\lambda} = O((k-1)^{-\lambda}). \end{aligned}$$

Since $\lambda > 1$, $P[M_k^* \geq \delta' B_{m_k} b_{m_k} \text{ i.o.}] = 0$. But $B_{n_{k-1}} \leq B_{m_k} < cB_{n_{k-1}}$, so $b_{m_k} \sim b_{n_{k-1}}$ (i.e. $b_{m_k}/b_{n_{k-1}} \rightarrow 1$) as $k \rightarrow \infty$ and

$$\begin{aligned} P[M_n \geq \delta B_n b_n \text{ i.o.}] &\leq P[M_k^* \geq \delta B_{n_{k-1}} b_{n_{k-1}} \text{ i.o.}] \\ &\leq P[M_k^* \geq \delta c^{-1} B_{m_k} b_{n_{k-1}} \text{ i.o.}] \\ &\leq P[M_k^* \geq \delta' B_{m_k} b_{m_k} \text{ i.o.}] = 0. \end{aligned}$$

Therefore, $\limsup_{n \rightarrow \infty} M_n/(B_n b_n) \leq \delta = \alpha'(b^{-1} + bg(\gamma b))$ for all $\alpha' > \alpha$ and $0 < b \leq T/\gamma$ (where $T/\gamma \equiv \infty$ when $\gamma = 0$), proving (2.2). Note that, when $\gamma = 0$, $K_0 = \min_{b > 0} (b^{-1} + bg(0)) = 2^{\frac{1}{2}}$ as shown by Teicher [8]. \square

Remark. Lemma 2.3 remains valid under the less stringent hypothesis (cf. [1], p. 295) that $(e^{tM_n}, \mathcal{F}_n, n \geq 1)$ is a submartingale for every $t > 0$.

Corollary 2.4 (cf. Corollary 1 of [8]). Let $\{X_n\}$ be a sequence of independent r.v. with zero means and finite variances. Define $S_n = \sum_{j=1}^n X_j$ and suppose

$s_n^2 \equiv E(S_n^2) \rightarrow \infty$. If $X_n \leq c_n s_n$ for some positive sequence $\{c_n\}$ and all $n \geq 1$ and $a \equiv \limsup_{n \rightarrow \infty} c_n (\log \log s_n^2)^{\frac{1}{2}} < \infty$, then $\limsup_{n \rightarrow \infty} S_n / (s_n^2 \log \log s_n^2)^{\frac{1}{2}} \leq K_a$ a.s. where K_a is defined by (2.3).

Proof. By Lemma 1 (i) of Teicher [8], (2.1) holds with $M_n = S_n$, $\alpha_n = 1$, $B_n = s_n$, $N = C = 1$ and any $T > 0$. Since $\alpha = 1$ and $\gamma = a$ in this case, the desired result is clear from Lemma 2.3. \square

Remark. Corollary 2.4 shows that the hypothesis in Corollary 1 (i) of [8] that $E(X_n^2) = o(s_n^2)$ is unnecessary.

3. Proof of Theorem 1.1.

First, it will be shown that, for any pair $\varepsilon_2 > \varepsilon_1 > 0$,

$$\lim_{n \rightarrow \infty} (H_n(\varepsilon_2) - H_n(\varepsilon_1)) = 0; \tag{3.1}$$

this fact establishes the independence of H_- and H_+ from the value of x in (1.8). For $n \geq 1$, define the event $A_n = [\varepsilon_1 s_n < t_n |X_n| \leq \varepsilon_2 s_n]$ and the r.v. $Y_n = X_n I(A_n)$. Using (1.4),

$$\sum_{n=1}^{\infty} t_n^2 E(Y_n^2) / s_n^2 \leq \varepsilon_2^2 \sum_{n=1}^{\infty} P(A_n) < \infty,$$

so $(t_n/s_n)^2 \sum_{j=1}^n E(Y_j^2) \rightarrow 0$ by Kronecker's lemma. (3.1) is now clear because

$$H_n(\varepsilon_2) - H_n(\varepsilon_1) = s_n^{-2} \sum_{j=1}^n E(Y_j^2).$$

Turn now to the proof of (1.7). Let $\varepsilon > 0$. For $j \geq 1$, define $P_j = P[|X_j| > \varepsilon s_j t_j^{-1}]$. Then, using the Cauchy-Schwarz inequality twice,

$$\begin{aligned} \sum_{j=1}^n E(|X_j| I(|X_j| > \varepsilon s_j t_j^{-1})) &\leq \sum_{j=1}^n (E(X_j^2))^{\frac{1}{2}} P_j^{\frac{1}{2}} \\ &\leq \left(\left(\sum_{j=1}^n E(X_j^2) \right)^{\frac{1}{2}} \left(\sum_{j=1}^n P_j \right)^{\frac{1}{2}} \right) \leq s_n \left(\sum_{j=1}^n P_j \right)^{\frac{1}{2}}. \end{aligned}$$

Therefore, in view of (1.4),

$$(s_n t_n)^{-1} \sum_{j=1}^n E(|X_j| I(|X_j| > \varepsilon s_j t_j^{-1})) \rightarrow 0 \tag{3.2}$$

for every $\varepsilon > 0$.

Since (1.4) and (3.2) hold, Lemmas 2.1 and 2.2 ensure the existence of a sequence $\varepsilon_n \downarrow 0$ satisfying

$$\sum_{n=1}^{\infty} P[|X_n| > \varepsilon_n s_n t_n^{-1}] < \infty, \tag{3.3}$$

$$(s_n t_n)^{-1} \sum_{j=1}^n E(|X_j| I(|X_j| > \varepsilon_j s_j t_j^{-1})) \rightarrow 0, \tag{3.4}$$

and

$$\liminf_{n \rightarrow \infty} H_n(\varepsilon_n) = H_-^2. \tag{3.5}$$

(Actually, these lemmas yield possibly distinct sequences ε'_n , ε''_n and ε'''_n in (3.3), (3.4) and (3.5) respectively, but all three statements remain true using $\varepsilon_n = \max(\varepsilon'_n, \varepsilon''_n, \varepsilon'''_n)$.)

Now, for $n \geq 1$, define $X'_n = X_n I(|X_n| \leq \varepsilon_n s_n t_n^{-1})$,

$$S'_n = \sum_{j=1}^n X'_j, \quad (s'_n)^2 = \text{Var}(S'_n) \quad \text{and} \quad t'_n = (2 \log \log s'_n)^{\frac{1}{2}}.$$

Using (3.3) and the Borel-Cantelli lemma, $P[X_n \neq X'_n \text{ i.o.}] = 0$, so $\lim(S_n - S'_n)$ is finite a.s. and, hence, $(S_n - S'_n)/(s_n t_n) \rightarrow 0$ a.s. Furthermore, $E(S_n - S'_n)/(s_n t_n) \rightarrow 0$ by virtue of (3.4). Thus it will suffice to prove that

$$H_- \leq A' \equiv \limsup_{n \rightarrow \infty} (S'_n - E(S'_n))/(s_n t_n) \leq H_+ \quad \text{a.s.} \tag{3.6}$$

Because $E(X_j) = 0$,

$$\sum_{j=1}^n (E(X'_j))^2 \leq \sum_{j=1}^n (\varepsilon_j s_j t_j^{-1}) |E(X_j - X'_j)| \leq \varepsilon_1 s_n t_n^{-1} \sum_{j=1}^n E|X_j - X'_j|$$

so $s_n^{-2} \sum_{j=1}^n (E(X'_j))^2 \rightarrow 0$ by (3.4). This means that

$$(s'_n/s_n)^2 = s_n^{-2} \sum_{j=1}^n E(X_j^2 I(|X_j| \leq \varepsilon_j s_j t_j^{-1})) + o(1). \tag{3.7}$$

For $\varepsilon > 0$, choose N such that $\varepsilon_n < \varepsilon$ when $n \geq N$. If $n \geq N$,

$$\begin{aligned} H_n(\varepsilon_n) &\leq s_n^{-2} \sum_{j=1}^n E(X_j^2 I(|X_j| \leq \varepsilon_j s_j t_j^{-1})) \\ &\leq s_n^{-2} \sum_{j=1}^N E(X_j^2 I(\varepsilon < t_j |X_j|/s_j \leq \varepsilon_j)) + H_n(\varepsilon), \end{aligned}$$

so (3.5) and (3.7) yield

$$\liminf_{n \rightarrow \infty} s'_n/s_n \geq H_- \quad \text{and} \quad \limsup_{n \rightarrow \infty} (s'_n/s_n)^2 \leq \limsup_{n \rightarrow \infty} H_n(\varepsilon) \quad \text{for all } \varepsilon.$$

Letting $\varepsilon \downarrow 0$,

$$H_- \leq \liminf_{n \rightarrow \infty} s'_n/s_n \leq \limsup_{n \rightarrow \infty} s'_n/s_n \leq H_+. \tag{3.8}$$

If s'_n converges then $S'_n - E(S'_n)$ converges a.s. by the Kolmogorov Convergence Theorem (Loève [4], p. 248) so $A' = 0 = H_-$ a.s. trivially, since $s_n \rightarrow \infty$. So it may be assumed hereinafter that $s'_n \rightarrow \infty$.

Let $m > 1$ and $\lambda > H_+$. Define $Y_n = X'_n$ for $n \geq m$ and $Y_n = X_n I(|X_n| \leq \varepsilon_m s_n t_n^{-1})$ for $n < m$, and then let $R_n = \sum_{j=1}^n (Y_j - E(Y_j))$ and $r_n^2 = E(R_n^2)$. Since $S'_n - \sum_{j=1}^n Y_j$ depends only on X_1, X_2, \dots, X_m , clearly

$$(S'_n - E(S'_n) - R_n)/(s_n t_n) \rightarrow 0 \text{ a.s.}, \tag{3.9}$$

and $r_n/s'_n \rightarrow 1$. In view of (3.8), then, an integer N exists such that $r_n \leq \lambda s_n$ for $n \geq N$.

Now, $|Y_n - E(Y_n)| \leq 2\varepsilon_m s_n t_n^{-1}$, $n \geq 1$. Since $\varepsilon_m s_n t_n^{-1}$ is non-decreasing as $n \rightarrow \infty$, Lemma 1(i) of Teicher [8] implies

$$E \exp\{t R_n/r_n\} \leq \exp\{t^2 g(2t \varepsilon_m s_n/(r_n t_n))\}$$

for every $t > 0$ and $n \geq 1$, where $g(x) = x^{-2}(e^{-x} - 1 - x)$. Replacing t by $t r_n/(\lambda s_n)$ yields (2.1) for $n \geq N$, with $M_n = R_n$, $C = 1$, $\alpha_n = \lambda$, $B_n = s_n$, $c_n = 2\varepsilon_m/(\lambda t_n)$ and any $T > 0$. Since $\gamma = \gamma_m = \limsup_{n \rightarrow \infty} t_n c_n = 2\varepsilon_m/\lambda < \infty$, Lemma 2.3 implies

$$\limsup_{n \rightarrow \infty} R_n/(s_n t_n) \leq 2^{-\frac{1}{2}} \lambda K_{\gamma_m} \text{ a.s.}$$

for every $m > 1$ and $\lambda > H_+$. In view of (3.9), then,

$$\limsup_{n \rightarrow \infty} (S'_n - E(S'_n))/(s_n t_n) \leq 2^{-\frac{1}{2}} H_+ K_{\gamma_m} \text{ a.s.}$$

for every $m > 1$. The right-hand side of (3.6) now obtains by letting $m \rightarrow \infty$, which implies $\gamma_m \rightarrow 0$ and $K_{\gamma_m} \rightarrow 2^{\frac{3}{2}}$.

By Čebyšev's inequality and (3.8), $P[|S'_n - E(S'_n)| \geq \varepsilon s_n t_n] \leq (s'_n/(\varepsilon s_n t_n))^2 \rightarrow 0$ as $n \rightarrow \infty$ for every $\varepsilon > 0$. Hence $(S'_n - E(S'_n))/(s_n t_n) \rightarrow 0$ in probability, so $A' \geq 0$. Therefore, (3.6) is true when $H_- = 0$, so suppose $H_- > 0$. Since

$$|X'_n - E(X'_n)|/s'_n \leq 2\varepsilon_n s_n/(s'_n t_n) \equiv a_n \text{ and } a_n t'_n = O(\varepsilon_n t'_n t_n^{-1} H_-^{-1}) = O(\varepsilon_n),$$

Kolmogorov's LIL implies $\limsup_{n \rightarrow \infty} (S'_n - E(S'_n))/(s'_n t'_n) = 1$ a.s. Noting that $s_n H_-/2 < s'_n < 2s_n H_+$ for all large n when $H_- > 0$ (cf. (3.8)), clearly $t'_n \sim t_n$. Since $\liminf_{n \rightarrow \infty} s'_n/s_n \geq H_-$, the left side of (3.6) follows. \square

4. Some Illustrative Examples

The three examples to follow emphasize the importance of (1.3) in Theorem 1.1.

Example 4.1. Consider any numbers a, b such that $0 \leq a \leq b \leq 1$. An example will be constructed in which Theorem 1.1 applies with $H_- = a$ and $H_+ = b$. To this end, let $\{\delta_n\}$ be a sequence with $\delta_n = 0$ or 1 , $a^2 = \liminf_{n \rightarrow \infty} \sum_{j=1}^n \delta_j/n$ and $b^2 = \limsup_{n \rightarrow \infty} \sum_{j=1}^n \delta_j/n$. Let $\{Y_n\}, \{Z_n\}$ be independent sequences, independent of each

other, such that $P[Y_n = \pm 1] = 1/2$, $P[Z_n = \pm n] = 1/(2n^2)$ and $P[Z_n = 0] = 1 - n^{-2}$. Then let $X_n = \delta_n Y_n + (1 - \delta_n) Z_n$. Clearly $E(X_n) = 0$, $E(X_n^2) = 1$ so $s_n^2 = n$. Moreover, for $\varepsilon > 0$, $X_n I(|X_n| \leq \varepsilon s_n t_n^{-1}) = \delta_n Y_n$ for all n so large that $t_n < \varepsilon s_n < n t_n$, so $E(X_n^2 I(|X_n| \leq \varepsilon s_n t_n^{-1})) = \delta_n^2 = \delta_n$ for all such n . Hence

$$\lim_{n \rightarrow \infty} \left| H_n(\varepsilon) - \sum_{j=1}^n \delta_j/n \right| = 0, \quad \text{so } H_- = a, H_+ = b.$$

Note that (1.3) holds in this example, but (1.1) does not if $a < 1$. Indeed, if $a < 1$, then there is a subsequence $\{\delta_{n_k}\}$ such that $\delta_{n_k} = 0$ for all $k \geq 1$. But then

$$|X_{n_k}|/s_{n_k} = |Z_{n_k}|/n_k^{\frac{3}{2}} = n_k^{\frac{3}{2}}$$

on the event $[Z_{n_k} \neq 0]$, so (1.1) is false. But, for $\varepsilon > 0$,

$$\sum_{n=N}^{\infty} P[|X_n| \geq \varepsilon s_n t_n^{-1}] \leq \sum_{n=N}^{\infty} P[Z_n \neq 0] < \infty,$$

where $N = \min\{n | \varepsilon s_n > t_n\}$, so (1.3) holds.

Example 4.2. It will be shown that (3.1) may fail when (1.3) doesn't hold by means of a well-known example (see Marcinkiewicz and Zygmund [5] or Theorem 5 of Teicher [6]). Let Y_1, Y_2, \dots be independent r.v. with $P[Y_n = \pm 1] = 1/2$ for all $n \geq 1$. For $\lambda > 1$, let $\sigma_n^2 \equiv (\log n)^{-1} \exp(2\lambda n/\log n)$ and $X_n = \sigma_n Y_n$. Since $s_n^2 \sim (\log n) \sigma_n^2 / (2\lambda)$ and, hence, $t_n^2 \sim 2 \log n$ (cf. [5] p. 219) in this case, the (constant) sequence $t_n^2 X_n^2 / s_n^2 \rightarrow 4\lambda$. Hence (1.3) is false, $H_n(\varepsilon) \rightarrow 1$ if $\varepsilon^2 > 4\lambda$, and $H_n(\varepsilon) \rightarrow 0$ if $\varepsilon^2 < 4\lambda$.

Example 4.3. Suppose X_1, X_2, \dots are independent r.v. such that X_n is normal with mean zero and variance n^n . Then Hartman's LIL [3] implies (1.2). But $s_n^2 \sim n^n$ so, for every $\varepsilon > 0$,

$$P[|X_n| > \varepsilon s_n t_n^{-1}] = P[Z > \varepsilon s_n / (t_n n^{n/2})] \rightarrow P[Z > 0] = 1/2,$$

where Z is a standard normal r.v. Hence (1.4) fails in this case. This shows that (1.4) is not necessary for (1.2).

Teicher [7] showed that (1.2) always implies $\limsup_{n \rightarrow \infty} X_n / (s_n t_n) \leq 1$ a.s. In fact, the example above shows that Teicher's necessary condition is the best possible, in the sense that $\limsup_{n \rightarrow \infty} X_n / (s_n t_n) = 1$ a.s. above. It follows from the well-known fact that $P[Z > x] \sim e^{-x^2/2} / (2\pi x^2)^{\frac{1}{2}}$ as $x \rightarrow \infty$ that, for $\varepsilon > 0$,

$$P[|X_n| > \varepsilon s_n t_n] \sim (\pi \varepsilon^2 t_n^2 / 2)^{\frac{1}{2}} \exp\{-\varepsilon^2 s_n^2 \log \log s_n^2 / n^n\};$$

it follows readily that $\sum_{n=1}^{\infty} P[|X_n| > \varepsilon s_n t_n]$ converges if $\varepsilon > 1$ and diverges if $\varepsilon < 1$.

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