

# Some Refinements in the Theory of Supercritical Multitype Markov Branching Processes

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## § 1. Introduction

Let  $\{Z(t); t \geq 0\}$  be a  $p$ -type ( $2 \leq p < \infty$ ) continuous time Markov branching process. The problem of studying the limit behavior of  $\eta \cdot Z(t)$  as  $t \rightarrow \infty$  where  $\eta$  is an arbitrary complex vector was attacked by the author in [2]. Kesten and Stigum [3] studied the discrete time case. The objectives of this paper are i) to refine some of these works and ii) to fill a gap mentioned in [2]. These are important in themselves as well as for their methodology which we believe will have applications elsewhere.

## § 2. The Statement of Results

We shall use the set up in [2]. Briefly, we consider a supercritical, positively regular and nonsingular multitype Markov branching process  $\{Z(t); t \geq 0\}$  with mean matrix  $M(t) = \exp(At)$  and finite second moments. Let the eigenvalues of  $A$  be arranged such that  $\lambda_1 > \operatorname{Re} \lambda_2 \geq \operatorname{Re} \lambda_3 \geq \dots \geq \operatorname{Re} \lambda_r$  where  $\lambda_1, \lambda_2, \dots, \lambda_r$  are the distinct eigenvalues and let  $u$  and  $v$  denote respectively the left and right eigenvectors of  $A$  with eigenvalue  $\lambda_1$  normalised such that  $u, v$  are both strictly positive and  $u \cdot v = 1$ . It is known that in the supercritical case, that is when  $\lambda_1 > 0$ ,  $Z(t) e^{-\lambda_1 t} \rightarrow uW$  almost surely as  $t \rightarrow \infty$  where  $W$  is a nonnegative random variable such that  $P(W > 0) > 0$  for any  $P$  with  $P(Z(0) \neq 0) > 0$ . If  $\eta$  is any complex vector such that  $\eta \cdot u \neq 0$  then it follows that  $\eta \cdot Z(t) e^{-\lambda_1 t} \rightarrow (\eta \cdot u)W$ . Thus the problem is easily solved in this case. When  $\eta \cdot u = 0$  we have to use different normalisation. In [2] we had established a trichotomy in the behavior of  $\eta \cdot Z(t)$ .

Let  $\{v_{jk}, k=1, 2, \dots, d_j, j=1, 2, \dots, r\}$  be the generalized eigenvectors of  $A$ . Here  $d_j$  is the algebraic multiplicity of the eigenvalue  $\lambda_j$  and  $\{v_{jk}\}$  satisfy

$$M(t) v_{jl} = e^{\lambda_j t} \sum_{k=1}^l v_{jk} \frac{t^{l-k}}{(l-k)!}, \quad 1 \leq l \leq d_j. \quad (1)$$

Given any vector  $\eta$  we can find unique constants  $c_{jk}$  such that

$$\eta = \sum_{j=1}^r \sum_{k=1}^{d_j} c_{jk} v_{jk}. \quad (2)$$

Define four objects connected with  $\eta$  as follows:

$$\begin{aligned} a(\eta) &= \sup \{ \operatorname{Re} \lambda_j; c_{jk} \neq 0 \text{ for some } k \} \\ \tilde{\gamma}(\eta) &= \sup \{ k: \text{there exists } j \text{ such that } \operatorname{Re} \lambda_j = a(\eta) \text{ and } c_{jk} \neq 0 \} \\ \gamma(\eta) &= \tilde{\gamma}(\eta) - 1 \\ I(\eta) &= \{ j: \operatorname{Re} \lambda_j = a(\eta), c_{j\tilde{\gamma}} \neq 0, c_{jk} = 0 \text{ for } k > \tilde{\gamma} \}. \end{aligned} \quad (3)$$

The trichotomy mentioned earlier is based on whether  $2a(\eta) >, =,$  or  $< \lambda_1$ . In this paper we shall refine the result proved in [2] for the case  $2a(\eta) > \lambda_1$  and also fill a gap in the case  $2a(\eta) = \lambda_1$ . The gap in question is simply to establish the following

**Theorem 1.** *Let  $2a(\eta) = \lambda_1$ . Then for each  $\varepsilon > 0$*

$$\lim_{s \rightarrow \infty} E_r \{ |\eta \cdot Z(s)|^2 e^{-\lambda_1 s} s^{-(2\gamma+1)}; |\eta \cdot Z(s)| > s^{2\gamma+1} e^{(\lambda_1 + \varepsilon)s} \} = 0 \tag{4}$$

where for any  $1 \leq r \leq p$ ,  $E_r$  stands for expectation when  $Z(0) = e_r = (\delta_{r1}, \delta_{r2}, \dots, \delta_{rp})$  where  $\delta_{rj}$ 's are Kronecker deltas.

We take this up in § 4. In [2] we established the following

**Theorem 2.** *Let  $2a(\eta) > \lambda_1$ . Then, there exist random variables  $Y_j$  for  $j \in I(\eta)$  such that if*

$$X(t) = \gamma! t^{-\gamma} e^{-at} \eta \cdot Z(t) - \sum_{j \in I(\eta)} Y_j e^{ib_j t} \tag{5}$$

then

$$E |X(t)|^2 = O(t^{-2}) \tag{6}$$

where  $b_j = \text{Im}(\lambda_j)$ .

It is immediate from this theorem and Borel-Cantelli that in the discrete time case

$$X(t) \rightarrow 0 \quad \text{a. s. as } t \rightarrow \infty. \tag{7}$$

However, the almost sure convergence is far from obvious in the continuous time case. We can, of course, establish the following

**Corollary 1.** *For each  $\delta > 0$  there exists a set  $A_\delta \in \mathbb{F}$  where  $(\Omega, \mathbb{F}, P)$  is our basic probability space such that  $P(A_\delta) = 1$  and*

$$\omega \in A_\delta \Rightarrow \lim_{n \rightarrow \infty} X(n\delta, \omega) = 0.$$

In general, a result of the above type does not imply the a. s. convergence of  $X(t)$ . For a counterexample see Kingman [4]. For a related difficulty see Athreya [1]. Our approach to establish (7) in the continuous time case uses the representation (2) of  $\eta$  in terms of generalized eigenvectors  $v_{ji}$ . The assertion (7) is an easy corollary to the following complete breakdown of the asymptotic behavior of  $\eta \cdot Z(t)$ .

**Theorem 3.** *Let  $\eta = \sum_{j=2}^r \sum_{k=1}^{d_j} v_{jk} c_{jk}$ . We can always write*

$$\begin{aligned} \eta \cdot Z(t) = & \sum_{(2 \text{ Re } \lambda_j > \lambda_1)} \sum_{k=1}^{d_j} c_{jk} \frac{e^{\lambda_j t} t^{(k-1)}}{(k-1)!} \bar{Y}_{jk}(t) \\ & + \sum_{(2 \text{ Re } \lambda_j = \lambda_1)} \sum_{k=1}^{d_j} c_{jk} e^{\lambda_1 t/2} t^{(k-\frac{1}{2})} \bar{Y}_{jk}(t) \\ & + \sum_{(2 \text{ Re } \lambda_j < \lambda_1)} \sum_{k=1}^{d_j} c_{jk} e^{\lambda_1 t/2} \bar{Y}_{jk}(t) \end{aligned} \tag{8}$$

where

$$\bar{Y}_{jk}(t) = \begin{cases} (k-1)! e^{-\lambda_j t} t^{-(k-1)} v_{jk} \cdot Z(t), & 2 \operatorname{Re} \lambda_j > \lambda \\ v_{jk} \cdot Z(t) e^{-\lambda_1 t/2} t^{-(k-\frac{1}{2})}, & \text{if } 2 \operatorname{Re} \lambda_j = \lambda_1 \\ v_{jk} \cdot Z(t) e^{-\lambda_1 t/2}, & \text{if } 2 \operatorname{Re} \lambda_j < \lambda_1. \end{cases}$$

Further, if  $2 \operatorname{Re} \lambda_j > \lambda_1$ , then for all  $k$ ,  $\bar{Y}_{jk}(t)$  converges in mean square and almost surely to a random variable  $Y_j$  independent of  $k$ ; if  $2 \operatorname{Re} \lambda_j \leq \lambda_1$ , then  $\bar{Y}_{jk}(t)$  converges in law to a mixture of normal distributions and  $\sup_t E |\bar{Y}_{jk}(t)|^2 < \infty$ . Finally if  $2\alpha > \lambda_1$  and  $2 \operatorname{Re} \lambda_j \leq \lambda_1$  then  $e^{-\alpha t} v_{jk} \cdot Z(t) \rightarrow 0$  almost surely and in mean square.

The case  $2a(\eta) > \lambda_1$  is discussed in Section 3.

### § 3. Two Basic Lemmas and Their Corollaries

**Lemma 1.** Let

$$w_{jl}(t) = e^{-\lambda_j t} t^{-(l-1)} \sum_{k=0}^{l-1} \frac{(-t)^k}{k!} v_{j(l-k)} \tag{9}$$

$$W_{jl}(t) = w_{jl}(t) \cdot Z(t).$$

Then for  $n \geq 1$

$$P_r \left( \sup_{n-1 \leq t \leq n} |W_{jl}(t)| > M \right) \leq K \frac{E_r |W_{jl}(n)|^2}{M^2} \tag{10}$$

where  $K$  is a constant independent of  $n$  and  $M$ .

**Lemma 2.**

$$E_r |W_{jl}(t)|^2 = \begin{cases} O(t^{-2(l-1)}) & \text{if } 2 \operatorname{Re} \lambda_j \geq \lambda_1, l \geq 2 \\ O(t^{-1}) & \text{if } 2 \operatorname{Re} \lambda_j = \lambda_1 \\ O(e^{(\lambda_1 - 2 \operatorname{Re} \lambda_j)t}) & \text{if } 2 \operatorname{Re} \lambda_j < \lambda_1. \end{cases} \tag{11}$$

We now get a few corollaries as easy consequences of the above two lemmas.

**Corollary 2.** Let  $2 \operatorname{Re} \lambda_j > \lambda_1$  and  $l \geq 2$ . Then  $W_{jl}(t) \rightarrow 0$  almost surely and in mean square as  $t \rightarrow \infty$ .

*Proof.* Use (10), (11) and Borel-Cantelli. q.e.d.

**Corollary 3.** Let  $2 \operatorname{Re} \lambda_j > \lambda_1$  and

$$Y_{jl}(t) = (l-1)! e^{-\lambda_j t} t^{-(l-1)} v_{jl} \cdot Z(t). \tag{12}$$

Then  $\lim_{t \rightarrow \infty} Y_{jl}(t)$  exists in mean square and almost surely and is independent of  $l$ .

*Proof.* For  $l=1$  the result follows from martingale arguments since  $v_{j1}$  being an eigenvector with eigenvalue  $\lambda_j$  makes  $Y_{j1}(t)$  a martingale and  $2 \operatorname{Re} \lambda_j > \lambda_1$  makes  $\sup_t E |Y_{j1}(t)|^2 < \infty$ .

By Corollary 2

$$\lim_{t \rightarrow \infty} W_{jl}(t) = 0 \quad \text{a.s. for } l \geq 2.$$

Note that

$$W_{jl}(t) = \sum_{k=0}^{l-1} \frac{(-1)^k}{k!(l-1-k)!} Y_{j,l-k}(t). \tag{13}$$

The result now follows by an induction on  $l$  and an use of the identity

$$\begin{aligned} \sum_{k=1}^{l-1} \frac{(-1)^k}{k!(l-1-k)!} &= \frac{1}{(l-1)!} \left[ \sum_{k=0}^{l-1} \frac{(l-1)! (-1)^k 1^{l-k}}{k!(l-k-1)!} - 1 \right] \\ &= \frac{1}{(l-1)!} [(-1+1)^{l-1} - 1] \\ &= -\frac{1}{(l-1)!}. \quad \text{q.e.d.} \end{aligned}$$

**Corollary 4.** Let  $2\text{Re } \lambda_j \leq \lambda_1$ . Let  $a$  and  $\gamma$  be constants and

$$\begin{aligned} \tilde{Y}_{j_l}(t) &= e^{-(a-\lambda_j)t} t^{-(\gamma-l+1)} Y_{j_l}(t) \\ \tilde{W}_{j_l}(t) &= e^{-(a-\lambda_j)t} t^{-(\gamma-l+1)} W_{j_l}(t). \end{aligned} \quad (14)$$

Then  $\tilde{W}_{j_l}(t)$  and  $\tilde{Y}_{j_l}(t)$  both tend to zero in mean square and almost surely provided either  $2a > \lambda_1$  or  $2a = \lambda_1$  and  $\gamma \geq l$ .

*Proof.* From Lemma 2 we get

$$E |\tilde{W}_{j_l}(t)|^2 \leq \text{const } e^{-(2a-\lambda_1)t} t^{-2(\gamma-l+1)}$$

and the mean square convergence follows. Also  $v_{j_l}$  being an eigenvector we know from [2] that

$$E |Y_{j_1}(t)|^2 \leq \text{const } e^{(\lambda_1 - 2\text{Re } \lambda_j)t}$$

and this clearly implies the mean square convergence of  $\tilde{Y}_{j_l}(t)$ . From (13) we get

$$\tilde{W}_{j_l}(t) = \sum_{k=0}^{l-1} \frac{(-1)^k}{k!(l-1-k)!} \tilde{Y}_{j, l-k}(t). \quad (15)$$

The mean square convergence of  $\tilde{Y}_{j_l}(t)$  follows by induction.

Next we turn to almost sure convergence. Since  $Y_{j_1}(t)$  is a martingale in  $t$

$$\begin{aligned} P \left\{ \sup_{n-1 \leq t \leq n} |\tilde{Y}_{j_1}(t)| > M \right\} &\leq P \left\{ \sup_{n-1 \leq t \leq n} |Y_{j_1}(t)| > M e^{(a-\text{Re } \lambda_j)(n-1)} (n-1)^{(\gamma-l+1)} \right\} \\ &\leq \text{const } \frac{E |Y_{j_1}(n)|^2}{M^2} e^{-2(a-\text{Re } \lambda_j)n} n^{-2(\gamma-l+1)}. \end{aligned}$$

Clearly,  $\sum_{n=1}^{\infty} e^{-\theta n} n^\delta < \infty$  for  $\theta > 0$ ,  $|\delta| < \infty$  or  $\theta = 0$ ,  $\delta > 1$ . The almost sure convergence of  $\tilde{Y}_{j_1}(t)$  follows by Borel-Cantelli. An exactly similar argument yields the almost sure convergence of  $\tilde{W}_{j_l}(t)$  for all  $l$ . The rest of the argument is the same as in Corollary 3 except that one uses (15) in place of (13). q.e.d.

It is easy to verify (7) from the definition of  $a(\eta)$ ,  $\gamma(\eta)$ ,  $I(\eta)$  and Corollaries 3 and 4.

Modulo a gap in the case  $2\text{Re } \lambda_j = \lambda_1$  it was established in [2] that  $\bar{Y}_{j_k}(t)$  as defined in Theorem 3 converges in law to a mixture of normal distributions when  $2\text{Re } \lambda_j \leq \lambda_1$ . Thus the proof of Theorem 3 is complete.

We shall finish this section with a proof of the basic lemmas.

*Proof of Lemma 1.* Let  $T = \inf_n \{t: n-1 \leq t \leq n, |W_{jl}(t)| > M\}$  and  $n$  if the set in braces is empty. Then  $T$  is a stopping time for our Markov process  $\{Z(t); t \geq 0\}$ . Hence if  $\mathbb{F}_T$  is the  $\sigma$ -field associated with  $T$  we have

$$E(W_{jl}(n) | \mathbb{F}_T) = e^{-\lambda_j n} n^{-(l-1)} \sum_{k=0}^{l-1} \frac{(-n)^k}{k!} M(n-T) v_{j,l-k} \cdot Z(T).$$

But on using (1) we get

$$\begin{aligned} \sum_{k=0}^{l-1} \frac{(-n)^k}{k!} M(n-T) v_{j,l-k} &= \sum_{k=0}^{l-1} \frac{(-n)^k}{k!} e^{\lambda_j(n-T)} \sum_{m=1}^{l-k} v_{j,m} \frac{(n-T)^{l-k-m}}{(l-k-m)!} \\ &= e^{\lambda_j(n-T)} \sum_{m=1}^l v_{j,m} \sum_{k=0}^{l-m} \frac{(n-T)^{l-k-m}}{(l-k-m)!} \frac{(-n)^k}{k!} \\ &= e^{\lambda_j(n-T)} \sum_{m=0}^{l-1} v_{j,l-m} \frac{(-T)^m}{m!}. \end{aligned}$$

Thus,

$$E(W_{jl}(n) | \mathbb{F}_T) = \left(\frac{T}{n}\right)^{l-1} W_{jl}(T). \tag{16}$$

Next

$$\begin{aligned} E|W_{jl}(n)|^2 &= E|W_{jl}(n) - W_{jl}(T) + W_{jl}(T)|^2 \\ &= E(W_{jl}(n) - W_{jl}(T)) \overline{W_{jl}(T)} + E \overline{(W_{jl}(n) - W_{jl}(T))} W_{jl}(T) \\ &\quad + E|W_{jl}(n) - W_{jl}(T)|^2 + E|W_{jl}(T)|^2. \end{aligned}$$

In view of (16)

$$\begin{aligned} E(W_{jl}(n) - W_{jl}(T)) \overline{W_{jl}(T)} &= E(\overline{W_{jl}(n) - W_{jl}(T)}) W_{jl}(T) \\ &= E\left(\left(\frac{T}{n}\right)^{l-1} - 1\right) |W_{jl}(T)|^2. \end{aligned}$$

Hence

$$E|W_{jl}(n)|^2 \geq E|W_{jl}(T)|^2 \left[ 2 \left\{ \left(\frac{T}{n}\right)^{l-1} - 1 \right\} + 1 \right].$$

Since  $n-1 \leq T \leq n$ ,

$$\begin{aligned} 2 \left\{ \left(\frac{T}{n}\right)^{l-1} - 1 \right\} + 1 &\geq 2 \left(\frac{n-1}{n}\right)^{l-1} - 1 \\ &\geq 2 \left(1 - \frac{1}{n}\right)^{p-1} - 1 \quad (l \leq p) \\ &\geq \frac{1}{2} \quad \text{for } n \text{ large enough.} \end{aligned}$$

For  $n$  large enough

$$\begin{aligned} P\left\{ \sup_{n-1 \leq t \leq n} |W_{jl}(t)| > M \right\} &\leq P\{W_{jl}(T) > M\} \\ &\leq \frac{E|W_{jl}(T)|^2}{M^2} \\ &\leq K \frac{E|W_{jl}(n)|^2}{M^2}, \end{aligned}$$

where  $K$  is any constant larger than 2. q.e.d.

*Proof of Lemma 2.* We get from [1] that

$$E_r |W_{ji}(t)|^2 = (M(t) \overline{w_{ji}(t)})^* D_r(0) (M(t) w_{ji}(t)) + \int_0^t (M(t-\tau) \overline{w_{ji}(t)})^* \left( \sum_{s=1}^p m_{rs}(\tau) B_s \right) (M(t-\tau) w_{ji}(t)) d\tau. \tag{17}$$

The calculation made in Lemma 1 tells us that

$$M(t-x) w_{ji}(t) = \left( \frac{x}{t} \right)^{l-1} w_{ji}(x) \quad \text{for } 0 \leq x \leq t.$$

Thus the first term on the right vanishes. The second term is of the form

$$t^{-2(l-1)} \int_0^t f(\tau) d\tau.$$

Now use Frobenius theory, the definition of  $w_{ji}(\tau)$ , and the relation between  $2\text{Re } \lambda_j$  and  $\lambda_1$  to estimate the growth rate of

$$\int_0^t |f(\tau)| d\tau. \quad \text{q.e.d.}$$

**§ 4. The Gap in the Case  $2a(\eta) = \lambda_1$**

We now tackle the gap mentioned in § 2 namely to establish Theorem 1. From the definition of  $a(\eta)$  and  $\gamma(\eta)$  it must be evident that

$$(M(s) \eta) e^{-as} s^{-(\gamma+\frac{1}{2})} \rightarrow 0 \quad \text{as } s \rightarrow \infty.$$

Thus (4) is equivalent to

$$\lim_{s \rightarrow \infty} E_r \{ |Y_r(s, \eta)|^2; |Y_r(s, \eta)| > c e^{\delta s} \} = 0 \tag{18}$$

for  $c > 0$  and  $\delta > 0$ , where

$$Y_r(s, \eta) = [\eta \cdot Z(s) - E_r(\eta \cdot Z(s))] e^{-as} s^{-(\gamma+\frac{1}{2})} = [\eta \cdot Z(s) - (M(s) \eta)_r] e^{-as} s^{-(\gamma+\frac{1}{2})}. \tag{19}$$

Using the fundamental property of branching processes, namely additivity, we get the identity (dropping the sub index  $r$ )

$$Y(s+1, \eta) = (s+1)^{-(\gamma+\frac{1}{2})} R(s, \eta) + \left( \frac{s}{s+1} \right)^{\gamma+\frac{1}{2}} e^{-a} Y(s, M \eta) \tag{20}$$

where

$$R(s, \eta) = \sum_{i=1}^p \frac{1}{e^{as}} \sum_{j=1}^{Z_i(s)} \frac{(\eta \cdot Z^{(ij)}(1) - (M \eta)_i)}{e^a}, \quad M = M(1),$$

and  $Z^{(ij)}(1)$  (for  $j=1, 2, \dots, Z_i(s)$ ) is the vector denoting the offspring population in one unit of time of the  $j$ -th particle among the  $Z_i(s)$  particles of type  $i$  at time  $s$ .

Now set

$$\begin{aligned} \bar{F}(s, c, \eta) &= E \{ |Y(s, \eta)|^2; |Y(s, \eta)| > c e^{\delta s} \} \\ F(s, c, \eta) &= s^{2\gamma+1} \bar{F}(s, c, \eta). \end{aligned} \tag{21}$$

Clearly,

$$\begin{aligned} \bar{F}(s+1, c, \eta) &\leq E\{|Y(s+1, \eta)|^2; |Y(s, M\eta)| > ce^a e^{\delta s}\} \\ &\quad + E\{|Y(s+1, \eta)|^2; |Y(s+1, \eta)| > ce^{\delta(s+1)}, |Y(s, M\eta)| < ce^a e^{\delta s}\} \\ &= I + II, \text{ say.} \end{aligned}$$

By the orthogonality of  $R(s, \eta)$  and the  $\sigma$ -field  $\mathbb{F}_s$ , we have

$$\begin{aligned} I &= (s+1)^{-(2\gamma+1)} E\{|R(s, \eta)|^2; |Y(s, M\eta)| > ce^a e^{\delta s}\} \\ &\quad + \left(\frac{s}{s+1}\right)^{2\gamma+1} \frac{1}{e^{2a}} E\{|Y(s, M\eta)|^2; |Y(s, M\eta)| > ce^a e^{\delta s}\} \end{aligned}$$

and

$$\begin{aligned} II &\leq 2E\left\{(s+1)^{-(2\gamma+1)} |R(s, \eta)|^2 + \left(\frac{s}{s+1}\right)^{2\gamma+1} \frac{1}{e^{2a}} |Y(s, \eta)|^2; \right. \\ &\quad \left. |Y(s+1, \eta)| > ce^{\delta(s+1)}, |Y(s, M\eta)| < ce^a e^{\delta s}\right\} \end{aligned}$$

using the trivial inequality  $(a+b)^2 \leq 2(a^2+b^2)$ .

This leads to the recurrence relation

$$F(s+1, c, \eta) \leq \rho F(s, ce^a, M\eta) + G(s, \eta, c) \tag{22}$$

where  $\rho = e^{2a}$ ,

$$\begin{aligned} G(s, \eta, c) &= G_1(s, \eta, c) + G_2(s, \eta, c) \\ G_1(s, \eta, c) &= 2E\{|R(s, \eta)|^2; A(s, \eta, c)\}, \quad A = A_1 \cup A_2, \\ A_1(s, \eta, c) &= \{|Y(s, M\eta)| > ce^a e^{\delta s}\} \\ A_2(s, \eta, c) &= \{|Y(s+1, \eta)| > ce^{\delta(s+1)}, |Y(s, M\eta)| \leq ce^a e^{\delta s}\} \\ G_2(s, \eta, c) &= 2E\{|Y(s, M\eta)|^2; A_2(s, \eta, c)\}. \end{aligned}$$

Iterating (22) yields

$$\begin{aligned} \frac{F(s+1, c, \eta)}{\rho^{s+1}} &\leq \frac{F(s, ce^a, M\eta)}{\rho^s} + \frac{G(s, \eta, c)}{\rho^s} \\ &\leq \sum_{r=0}^s \frac{G(s-r, M^r \eta, ce^{ar})}{\rho^{s-r}} \end{aligned}$$

which is the same as

$$\bar{F}(s+1, c, \eta) \leq \frac{\rho}{(s+1)^{2\gamma+1}} \sum_{r=0}^s \rho^r G(s-r, ce^{ar}, M^r \eta). \tag{23}$$

Let us first look at

$$\sum_{r=0}^s \rho^r G_1(s-r, ce^{ar}, M^r \eta).$$

By definition

$$\rho^r G_1(s-r, ce^{ar}, M^r \eta) = 2E\left\{\frac{|R(s-r, M^r \eta)|^2}{e^{2ar}}; A_{r,s}\right\}$$

where  $A_{r,s} = A(s-r, M^r \eta, ce^{ar})$ . We now make use of the definition of  $a(\eta)$ ,  $l(\eta)$  and  $\gamma(\eta)$ . If

$$l(r, \eta) = e^{-ar} r^\gamma \left[ M^r \eta - \frac{r^\gamma}{\gamma!} \sum_{j \in I(\eta)} \xi_j e^{\lambda_j r} \right]$$

then

$$l(r, \eta) = O\left(\frac{1}{r}\right), \quad \text{where } \xi_j = c_{j1} v_{j1}. \quad (24)$$

Next, by the orthogonality of  $R(s, \eta)$  and the  $\sigma$ -field  $\mathbb{F}_s$

$$E \left| \frac{R(s-r, M^r \eta)}{e^{ar} r^\gamma} - \sum_{j \in I(\eta)} e^{ib_j r} R(s-r, \xi_j) \right|^2 = E \left\{ \sum_i Z_i(s-r) e^{-\lambda_1(s-r)} \overline{l(r, \eta)^*} \pi_i l(r, \eta) \right\}$$

where  $\pi_i$  is the covariance matrix of the vector  $Z(1)$  with  $Z(0) = e_i$  (i.e., we start with one particle of type  $i$ ) and \* denotes transpose,

$$\leq \frac{\text{const}}{r^2} \quad \text{by (24).}$$

For any set  $A$

$$E \left\{ \left| \frac{R(s-r, M^r \eta)}{e^{ar} r^\gamma} \right|^2; A \right\} \leq 2E \left\{ \left| \sum_{j \in I(\eta)} R(s-r, \xi_j) e^{ib_j r} \right|^2; A \right\} + \frac{\text{const}}{r^2}.$$

Thus

$$\begin{aligned} & \sum_{r=1}^s \rho^r G_1(s-r, c e^{ar}, M^r \eta) \\ & \leq \text{const} \left( \sum_{r=1}^s r^{2\gamma} E \left\{ \left| \sum_{j \in I(\eta)} R(s-r, \xi_j) e^{ib_j r} \right|^2; A_{r,s} \right\} \right) + \text{const} \sum_{r=1}^s r^{2(\gamma-1)}. \end{aligned} \quad (25)$$

Clearly,

$$\frac{1}{(s+1)^{2\gamma+1}} \sum_{r=1}^s r^{2(\gamma-1)} \rightarrow 0 \quad \text{as } s \rightarrow \infty.$$

The first term on the right side of (25) is majorized using Minkowski's inequality by

$$\text{const} \sum_{j \in I(\eta)} \sum_{r=1}^s r^{2\gamma} E(|R(s-r, \xi_j)|^2; A_{r,s}).$$

We shall now show that for each  $j \in I(\eta)$

$$\frac{1}{(s+1)^{2\gamma+1}} \sum_{r=1}^s r^{2\gamma} E(|R(s-r, \xi_j)|^2; A_{r,s}) \rightarrow 0 \quad \text{as } s \rightarrow \infty. \quad (26)$$

This is, of course, implied by

$$\frac{1}{s} \sum_{r=1}^s E\{|R(s-r, \xi_j)|^2; A_{r,s}\} \rightarrow 0 \quad \text{as } s \rightarrow \infty. \quad (27)$$

To establish (27) break up  $r$  into two regions;  $r \leq s(1-\varepsilon)$  and  $r > s(1-\varepsilon)$ . Thus

$$\begin{aligned} \frac{1}{s} \sum_{r=1}^s E\{|R(s-r, \xi_j)|^2; A_{r,s}\} & \leq \frac{1}{s} \left[ \sum_{(1-\varepsilon)s < r \leq s} + \sum_{r \leq s(1-\varepsilon)} \right] \\ & \leq \text{const } \varepsilon + \frac{1}{s} \sum_{r \leq s(1-\varepsilon)} \end{aligned}$$



since  $\sup_s E |R(s, \xi_j)|^2 < \infty$  (use the fact  $2a = \lambda_1$  and  $\sup_s EZ_i(s) e^{-\lambda_1 s} < \infty$ ). Next,

$$\frac{1}{s} \sum_{r \leq s(1-\varepsilon)} = \frac{1}{s} \sum_{k \geq s\varepsilon} E(|R(k, \xi_j)|^2; A_{s-k, s}).$$

Since we can write

$$R(s, \xi_j) = \sum_{i=1}^p \frac{\sqrt{Z_i(s)}}{e^{as}} \frac{1}{\sqrt{Z_i(s)}} \sum_{j=1}^{Z_i(s)} \frac{(\xi \cdot Z^{(i,j)}(1) - \xi_i e^{\lambda})}{e^a}.$$

Now, by Renyi's generalization of the classical central limit theorem [5], we see that for each  $j$ ,  $R(s, \eta)$  converges in law to a random variable  $R$  (which is a mixture of normal distributions). Further  $E |R(s, \xi_j)|^2 \rightarrow ER^2$ . Thus the sequence  $|R(k, \xi_j)|^2$  for  $k=1, 2, \dots$  is uniformly integrable. Thus, if we show

$$\sup_{k \geq s\varepsilon} P(A_{s-k, s}) \rightarrow 0 \quad \text{as } s \rightarrow \infty. \tag{28}$$

(27) will follow since  $\varepsilon$  is arbitrary. Since  $A_{s-k, s} = A(k, M^{s-k}\eta, c e^{a(s-k)})$  and  $A(s, \eta, c) = A_1 \cup A_2$ ,

$$P(A_{s-k, s}) \leq P\{|Y(k, M^{s-k+1}\eta)| > c e^{a(s-k)} e^{\delta k}\} + P\{A_2(k, M^{s-k}\eta, c e^{a(s-k)})\} \tag{29}$$

$$= I' + II'.$$

By Chebychev's inequality

$$I' \leq \frac{\text{const}}{e^{2\delta k}} E \left\{ \left| Y \left( k, \frac{M^{s-k+1}\eta}{e^{a(s-k)}} \right) \right|^2 \right\}$$

$$\leq \frac{\text{const}}{e^{2\delta s\varepsilon}} \frac{1}{s} \left( \frac{1}{\varepsilon} \right)^{2\gamma+1} + \frac{(2s-k)^{2\gamma+1} - (s-k)^{2\gamma+1}}{s^{2\gamma+1}}$$

(the last step can be established using (17) in much the same way as Lemma 2)

$$\rightarrow 0 \quad \text{as } s \rightarrow \infty.$$

Turning to  $II'$  we notice first that

$$A_2(k, M^{s-k}\eta, c e^{a(s-k)}) \subset \{|R(k, M^{s-k}\eta)| > c e^{a(s-k)} e^{\delta k} [e^{\delta(k+1)^{\gamma+\frac{1}{2}}} - k^{\gamma+\frac{1}{2}}]\}.$$

Thus

$$II \leq \text{const} \cdot E \{|R(k, M^{s-k}\eta)|^2\} \{e^{2a(s-k)} e^{2\delta k} [e^{\delta(k+1)^{\gamma+\frac{1}{2}}} - k^{\gamma+\frac{1}{2}}]\}^{-2}$$

$$\leq \text{const} \cdot \frac{(s-k)^{2\gamma}}{e^{2\delta k} k^{2\gamma+1} \left[ e^{\delta \left(1 + \frac{1}{k}\right)^{\gamma+\frac{1}{2}}} - 1 \right]^{\frac{1}{2}}}$$

again using the fact

$$\sup_{s(1-\varepsilon) \leq k \leq s} \left( E \frac{|R(k, M^{s-k}\eta)|^2}{e^{2a(s-k)}(s-k)^{2\gamma}} \right) < \infty$$

$$\leq \text{const} \left( \frac{1}{1-\varepsilon} - 1 \right)^{2\gamma} \frac{1}{e^{2\delta s(1-\varepsilon)}} \frac{1}{(s(1-\varepsilon))^2} \quad (\text{for } s(1-\varepsilon) \leq k \leq s)$$

$$\rightarrow 0 \quad \text{as } s \rightarrow \infty.$$

This  $\Rightarrow$  (29)  $\Rightarrow$  (28)  $\Rightarrow$  (26). From (26) and (25) we conclude that

$$\frac{1}{(s+1)^{2\gamma+1}} \sum_{r=0}^s \rho^r G_1(s-r, c e^{ar}, M^r \eta) \rightarrow 0 \quad \text{as } s \rightarrow \infty. \quad (30)$$

To finish the proof we need to show

$$\frac{1}{(s+1)^{2\gamma+1}} \sum_{r=0}^s \rho^r G_2(s-r, c e^{ar}, M^r \eta) \rightarrow 0 \quad \text{as } s \rightarrow \infty. \quad (31)$$

Recall that

$$\begin{aligned} G_2(s, \eta, c) &= 2E \{ |Y(s, M\eta)|^2; |Y(s+1, \eta)| > c e^{\delta(s+1)}, |Y(s, M\eta)| \leq c e^{\delta s} \} \\ &= 2E \{ |Y(s, M\eta)|^2 \chi_{A_2(s, \eta, c)} \}, \end{aligned}$$

where  $\chi_A$  stands for the indicator function of the set  $A$

$$\leq 2E \{ |\tilde{Y}(s, M\eta)|^2 \chi_{\bar{A}_2(s, \eta, c)} \}$$

where  $\bar{A}_2(s, \eta, c) = \{ |R(s, \eta)| > c e^{\delta s} [e^{\delta(s+1)^{\gamma+\frac{1}{2}}} - s^{\gamma+\frac{1}{2}}] \}$ ,  $\tilde{Y} = Y$  on  $\{ |Y(s, M\eta)| \leq c e^{\delta s} \}$  and 0 otherwise

$$\leq 2E \{ |\tilde{Y}(s, M\eta)|^2 E(\chi_{\bar{A}_2(s, \eta, c)} | \mathbb{F}_s) \}$$

(by conditioning on the  $\sigma$ -field  $\mathbb{F}_s$ )

$$\leq \text{const } E \{ |\tilde{Y}(s, M\eta)|^2 E(|R(s, \eta)|^2 | \mathbb{F}_s) (f(s, c))^{-2} \}$$

where  $f(s, c) = c e^{\delta s} [e^{\delta(s+1)^{\gamma+\frac{1}{2}}} - s^{\gamma+\frac{1}{2}}]$ ,

$$\leq \text{const } (f(s))^{-2} E \left\{ |\tilde{Y}(s, M\eta)|^2 \sum_{i=1}^p Z_i(s) e^{-\lambda_1 s} \bar{\eta}^* \Pi_i \eta \right\}.$$

Thus,

$$\begin{aligned} & \rho^r G_2(s-r, c e^{ar}, M^r \eta) \\ & \leq \text{const } E \left\{ |\tilde{Y}(s-r, M^{r+1} \eta)|^2 \sum_{i=1}^p Z_i(s-r) e^{-\lambda_1(s-r)} \left( \frac{M^r \bar{\eta}}{e^{ar} r^\gamma} \right)^* \Pi_i \frac{M^r \eta}{e^{ar} r^\gamma} \right\} \\ & \quad \cdot r^{2\gamma} (f(s-r, c e^{ar}))^{-2} \\ & \leq \text{const } E \left\{ \left| \frac{\tilde{Y}(s-r, M^{r+1} \eta)}{c e^{ar} e^{\delta(s-r)}} \right|^2 \sum_{i=1}^p Z_i(s-r) e^{-\lambda_1(s-r)} \right\} \times (c e^{ar} e^{\delta(s-r)})^2 r^{2\gamma} \\ & \quad \times (c e^{ar} e^{\delta(s-r)})^2 \times [e^{\delta(s-r+1)^{\gamma+\frac{1}{2}}} - (s-r)^{\gamma+\frac{1}{2}}]^2 \\ & \leq \text{const} \left( \frac{r}{s-r} \right)^{2\gamma} \quad \text{if } (s-r) \geq 1 \\ & = r^{2\gamma} \quad \text{if } r=s. \end{aligned}$$

Again break up  $r$  into three regions  $r = s$ ,  $s(1 - \varepsilon) \leq r \leq s - 1$ , and  $r \leq s(1 - \varepsilon)$ .

$$\begin{aligned} & \frac{1}{(s+1)^{2\gamma+1}} \sum_{r=0}^s \rho^r G_2(s-r, c e^{ar}, M^r \eta) \\ & \leq \frac{1}{(s+1)^{2\gamma+1}} \left[ \sum_{r < s(1-\varepsilon)} + \sum_{s(1-\varepsilon) \leq r \leq s-1} + \sum_{r=s} \right] \\ & \leq \text{const} \left[ \left( \frac{1}{\varepsilon} - 1 \right)^{2\gamma} \frac{s(1-\varepsilon)}{s^{2\gamma+1}} + \frac{1}{s} \sum_{s(1-\varepsilon) \leq r \leq s-1} \frac{1}{s^{2\gamma}} \left( \frac{r}{s-r} \right)^{2\gamma} + \frac{1}{s} \right] \\ & \rightarrow 0 \quad \text{as } s \rightarrow \infty. \quad \text{q.e.d.} \end{aligned}$$

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