

On the Last Time and the Number of Boundary Crossings Related to the Strong Law of Large Numbers and the Law of the Iterated Logarithm

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1. Introduction and Summary

Let X_1, X_2, \dots be a sequence of i.i.d. random variables with $EX_1=0$, $S_n = X_1 + \dots + X_n$, and let $(b(n))_{n \geq 1}$ be a sequence of real numbers. Define

$$N(b(n)) = \sum_{n=1}^{\infty} I_{\{S_n \geq b(n)\}}, \quad (1.1 a)$$

$$\tilde{N}(b(n)) = \sum_{n=1}^{\infty} I_{\{|S_n| \geq b(n)\}}. \quad (1.1 b)$$

Kolmogorov's strong law of large numbers can be stated in terms of the random variable $\tilde{N}(\varepsilon n)$ as follows:

$$P[\tilde{N}(\varepsilon n) < \infty] = 1 \quad \text{for every } \varepsilon > 0. \quad (1.2)$$

To obtain a stronger result than Kolmogorov's strong law, Slivka and Severo [12] have discussed the moments of $\tilde{N}(\varepsilon n)$ and they proved that for $r \geq 1$,

$$E|X_1|^{r+1} < \infty \Leftrightarrow E\tilde{N}^r(\varepsilon n) < \infty \quad \text{for all } \varepsilon > 0. \quad (1.3)$$

Motivated by the Marcinkiewicz-Zygmund extension of Kolmogorov's strong law, Stratton [14] proved that for $r \geq 1$ and $\alpha > \frac{1}{2}$,

$$E|X_1|^{(r+1)\alpha} < \infty \Leftrightarrow E\tilde{N}^r(\varepsilon n^\alpha) < \infty \quad \text{for all } \varepsilon > 0. \quad (1.4)$$

There is in fact a misprint in [14] where the requirement $\alpha > 0$ should be replaced by $\alpha > \frac{1}{2}$.

An earlier result to strengthen almost sure convergence for normalized sample sums was due to Strassen [13] who, instead of $N(b(n))$ and $\tilde{N}(b(n))$, considered

$$L(b(n)) = \sup \{n \geq 1: S_n \geq b(n)\}, \quad (1.5 a)$$

$$\tilde{L}(b(n)) = \sup \{n \geq 1: |S_n| \geq b(n)\} \quad (\sup \emptyset = 0). \quad (1.5 b)$$

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In other words, Strassen considers the last exit time $\tilde{L}(b(n))$ while Slivka and Severo consider the number of exits $\tilde{N}(b(n))$. Again Kolmogorov's strong law can be restated in terms of the last exit time $\tilde{L}(\varepsilon n)$ as $P[\tilde{L}(\varepsilon n) < \infty] = 1$ for all $\varepsilon > 0$, and it follows easily from the results of Baum and Katz (see [1, 9, 13] or Section 4 below) that for $\alpha > \frac{1}{2}$ and $r > 0$,

$$E|X_1|^{(r+1)/\alpha} < \infty \Leftrightarrow E\tilde{L}^r(\varepsilon n^\alpha) < \infty \quad \text{for all } \varepsilon > 0. \quad (1.6)$$

An obvious connection between $\tilde{L}(b(n))$ and $\tilde{N}(b(n))$ is the following:

$$\tilde{N}(b(n)) \leq \tilde{L}(b(n)). \quad (1.7)$$

Likewise we have the inequality:

$$N(b(n)) \leq L(b(n)). \quad (1.8)$$

Making use of (1.6) and (1.7), together with Lemma 3 in Section 2, we can prove that in (1.4), the condition $r \geq 1$ as imposed by Stratton can be dropped. In fact we shall prove in Section 3 the following one-sided theorem involving $N(\varepsilon n^\alpha)$ which then immediately implies the corresponding two-sided result involving $\tilde{N}(\varepsilon n^\alpha)$.

Theorem 1. *Let X_1, X_2, \dots be i.i.d. random variables with $EX_1 = 0$, $E|X_1|^q < \infty$ for some $1 \leq q \leq 2$. Let $\alpha > 1/q$ and $r > 0$. Set $S_n = X_1 + \dots + X_n$ and define $N(\varepsilon n^\alpha)$ as in (1.1 a). Then*

$$\begin{aligned} E(X_1^+)^{(r+1)/\alpha} < \infty &\Leftrightarrow EN^r(\varepsilon n^\alpha) < \infty \quad \text{for all } \varepsilon > 0 \\ &\Leftrightarrow EN^r(\varepsilon n^\alpha) < \infty \quad \text{for some } \varepsilon > 0. \end{aligned}$$

We remark that for the one-sided result in Theorem 1, we require the two-sided moment condition $E|X_1|^q < \infty$ for some $1 \leq q \leq 2$. A counter-example to show that this condition cannot be dropped can be found in Section 2 of [2].

The law of the iterated logarithm can be formulated in terms of $\tilde{N}(\varepsilon(2n \log \log n)^{1/2})$ as follows: If $EX_1 = 0$ and $EX_1^2 = \sigma^2$, then

$$\begin{aligned} P[\tilde{N}(\varepsilon(2n \log \log n)^{1/2}) < \infty] &= 1 \quad \text{if } \varepsilon > \sigma, \\ &= 0 \quad \text{if } \varepsilon < \sigma. \end{aligned} \quad (1.9)$$

However, Slivka [11] showed that $E\tilde{N}^r(\varepsilon(2n \log \log n)^{1/2}) = \infty$ for all $r > 0$ and $\varepsilon > 0$. In [14], Stratton sharpened Slivka's result and found that if X_1 is symmetric with $EX_1 = 0$ and $EX_1^{2(m+1)} < \infty$ for some positive integer m , then

$$E\tilde{N}^r(\sigma(2(1+\delta)n \log n)^{1/2}) < \infty \quad \text{if } 1 \leq r < \min(m, 1+\delta); \quad (1.10a)$$

$$E\tilde{N}^r(\sigma(2(1+\delta)n \log n)^{1/2}) = \infty \quad \text{if } r > 1+\delta. \quad (1.10b)$$

There is in fact a misprint in Stratton's paper [14, p. 1012] where the moment condition $EX_1^{2m} < \infty$ should be changed to $EX_1^{2(m+1)} < \infty$ and the factor σ should be added to $(2(1+\delta)n \log n)^{1/2}$. Stratton's proof makes use of the Berry-Esseen bound, the form he quotes being the following result of Katz [7]:

$$E|X_1|^p < \infty \Rightarrow \sup_x |\Phi(x) - P[S_n \leq \sigma n^{1/2} x]| \leq C(p, \sigma^{-p} E|X_1|^p) n^{-(p-2)/2}, \quad (1.11)$$

where $C(p, \eta)$ is a universal constant depending only on p and η and Φ is the distribution function of the standard normal distribution. Now (1.11) is valid for $2 < p \leq 3$ and, as is well known (cf. [4, p. 53]), it cannot be extended to the case

$p > 3$. However, Stratton's proof deals with the case $p = 2(m + 1) \geq 4$ and therefore (1.11) cannot be applied as he stated. What should be applied instead is the following result due to Esseen [4, p. 73]: If $E|X_1|^k < \infty$ for some integer $k \geq 3$, then for $|x| \geq \{(1 + \delta)(k - 2) \log n\}^{1/2}$,

$$|\Phi(x) - P[S_n \leq \sigma n^{1/2} x]| \leq c(\delta, \beta_1, \dots, \beta_k) n^{-(k-2)/2} / (1 + |x|^k), \tag{1.12}$$

where δ is any fixed number with $0 < \delta < 1$ and $c(\delta, \beta_1, \dots, \beta_k)$ is a finite constant depending only on δ and the moments $\beta_i = E|X_1|^i, i = 1, \dots, k$.

If we apply (1.12) in place of (1.11), we can indeed prove Stratton's result. However, by applying a result of [2] instead, we can drop the assumption that X_1 is symmetric and weaken the moment condition $EX_1^{2(m+1)} < \infty$; also we shall no longer require m to be an integer and we can prove the one-sided theorem involving $N(\sigma(2(1 + \delta)n \log n)^{1/2})$ which then immediately implies the corresponding result for the number of exits \tilde{N} of the two-sided region studied by Stratton. Our result, which will be proved in Section 3, is the following theorem:

Theorem 2. *Let X_1, X_2, \dots be i.i.d. random variables such that $EX_1 = 0, \infty > EX_1^2 = \sigma^2 > 0$. Set $S_n = X_1 + \dots + X_n$ and define $N(\varepsilon(n \log n)^{1/2})$ as in (1.1 a) and $L(\varepsilon(n \log n)^{1/2})$ as in (1.5 a). Then for any $r > 0$, the following statements are equivalent:*

$$\int_{|X_1| > \varepsilon} X_1^{2(r+1)} (\log X_1)^{-(r+1)} dP < \infty. \tag{1.13}$$

$$EL(\varepsilon(n \log n)^{1/2}) < \infty \quad \text{for all } \varepsilon > (2r)^{1/2} \sigma. \tag{1.14}$$

$$EN^r(\varepsilon(n \log n)^{1/2}) < \infty \quad \text{for all } \varepsilon > (2r)^{1/2} \sigma. \tag{1.15}$$

$$EN^r(\varepsilon(n \log n)^{1/2}) < \infty \quad \text{for some } \varepsilon > 0. \tag{1.16}$$

$$EL(\varepsilon(n \log n)^{1/2}) < \infty \quad \text{for some } \varepsilon > 0. \tag{1.17}$$

As an easy corollary of Theorem 2, we obtain the following analogue for \tilde{L} and \tilde{N} :

Theorem 3. *Let X_1, X_2, \dots be i.i.d., $S_n = X_1 + \dots + X_n$ and define $\tilde{N}(\varepsilon(n \log n)^{1/2}), \tilde{L}(\varepsilon(n \log n)^{1/2})$ as in (1.1 b) and (1.5 b) respectively. Then for any $r > 0$, the following statements are equivalent:*

$$EX_1 = 0, \quad EX_1^2 = \sigma^2 \quad \text{and} \quad E|X_1|^{2(r+1)} (1 + \log^+ |X_1|)^{-(r+1)} < \infty. \tag{1.18}$$

$$E\tilde{L}(\varepsilon(n \log n)^{1/2}) < \infty \quad \text{for all } \varepsilon > (2r)^{1/2} \sigma. \tag{1.19}$$

$$E\tilde{N}^r(\varepsilon(n \log n)^{1/2}) < \infty \quad \text{for all } \varepsilon > (2r)^{1/2} \sigma. \tag{1.20}$$

$$E\tilde{N}^r(\varepsilon(n \log n)^{1/2}) < \infty \quad \text{for some } \varepsilon > 0. \tag{1.21}$$

$$E\tilde{L}(\varepsilon(n \log n)^{1/2}) < \infty \quad \text{for some } \varepsilon > 0. \tag{1.22}$$

The equivalence between (1.18) and (1.22) was first discovered in [9] where it was proved by applying Theorem 3 of [8]. This equivalence, as pointed out in [9], sharpens an earlier result of Strassen [13] who, by embedding the sample sum process in Brownian motion, has shown that if $EX_1 = 0, EX_1^2 = \sigma^2$ and

$E|X_1|^p < \infty$ for some $p > 2(r+1)$, then (1.19) holds. In [9], it is also proved that under the condition (1.18),

$$E\tilde{L}^r(\varepsilon(n \log n)^{1/2}) = \infty \quad \text{if } \varepsilon < (2r)^{1/2} \sigma, \quad (1.23)$$

a result first obtained by Strassen under the stronger moment condition $E|X_1|^p < \infty$ for some $p > 2(r+1)$. In view of (1.7), Stratton's result (1.10b) implies (1.23), but Stratton has to assume that X_1 is symmetric with $EX_1^{2(m+1)} < \infty$ for some integer $m > r$. In Section 3, we shall prove the following theorem which completely generalizes the results of Slivka and Stratton.

Theorem 4. *Let X_1, X_2, \dots be i.i.d. random variables, $S_n = X_1 + \dots + X_n$ and define $N(\varepsilon(n \log n)^{1/2})$ as in (1.1a). Let $r > 0$.*

(i) *If X_1 is symmetric, then*

$$EN^r(\varepsilon(n \log n)^{1/2}) = \infty \quad \text{for all } 0 < \varepsilon < (2r EX_1^2)^{1/2}. \quad (1.24)$$

(Since we do not assume any moment condition on X_1, EX_1^2 in (1.24) can be infinite.)

(ii) *If X_1 satisfies the moment condition (1.18), then (1.24) still holds.*

From Theorems 3 and 4, we see that for sample sums, Strassen's strengthening of the law of the iterated logarithm in terms of the finiteness of the r -th moment of the last exit time $\tilde{L}(\varepsilon(n \log n)^{1/2})$, a concept which he calls in [13] the r -quick lim sup, turns out to be equivalent to the finiteness of the r -th moment of the number of exits $\tilde{N}(\varepsilon(n \log n)^{1/2})$ considered by Stratton. Furthermore, in view of (1.6) and Theorem 1, the refinement of the strong law of large numbers by Severo, Slivka and Stratton, who consider the finiteness of r -th moment of $\tilde{N}(\varepsilon n^\alpha)$, again turns out to be equivalent to the notion of r -quick convergence studied by Strassen [13] and Lai [9], who consider the r -th moment of $\tilde{L}(\varepsilon n^\alpha)$ instead. Some statistical applications showing the usefulness of the concept of r -quick convergence can be found in [9] and the references thereof.

So far we have discussed boundaries of the form εn^α ($\alpha > 1/2$) and $\varepsilon(n \log n)^{1/2}$ to obtain the r -quick analogues of the Marcinkiewicz-Zygmund strong law and the law of the iterated logarithm. In Section 4, we shall extend our results to general boundaries and in this connection, we obtain a general form of the Hsu-Robbins-Erdős-Baum-Katz theorem on convergence rates for the tail probabilities of sample sums.

2. Some Preliminary Lemmas

Lemma 1. *Let X_1, X_2, \dots be i.i.d. random variables, $S_n = X_1 + \dots + X_n$ and let $(b(n))_{n \geq 1}$ be a nondecreasing sequence of nonnegative numbers. Define $N = N(b(n))$*

as in (1.1a) and let $N_m = \sum_{n=1}^m I_{[S_n \geq b(n)]}$. Assume that $EN^r < \infty$ for some $r > 0$.

(i) *If $r \geq 1$, then $\sum_{n=1}^{\infty} EN_n^{r-1} I_{[S_n \geq b(n)]} < \infty$.*

(ii) If $r > 0$ and r is not an integer, then $\sum_{n=1}^{\infty} n^{r-[r]-1} EN_n^{[r]} I_{[S_n \geq b(n)]} < \infty$.

Proof. (i) is an immediate consequence of the following relation:

$$N^r = N^{r-1} \sum_{n=1}^{\infty} I_{[S_n \geq b(n)]} \geq \sum_{n=1}^{\infty} N_n^{r-1} I_{[S_n \geq b(n)]}. \tag{2.1}$$

If $r > 1$, then (ii) follows directly from (i). Assume that $0 < r < 1$. It is easy to see that $(x+1)^r - x^r$ is decreasing in $x > 0$. Let $N_0 = 0$. Then

$$\begin{aligned} EN_m^r &= E \sum_{n=1}^m (N_n^r - N_{n-1}^r) = E \sum_{n=1}^m (N_n^r - N_{n-1}^r) I_{[S_n \geq b(n)]} \\ &\geq E \sum_{n=1}^m (n^r - (n-1)^r) I_{[S_n \geq b(n)]} \geq r \sum_{n=1}^m n^{r-1} EI_{[S_n \geq b(n)]}. \end{aligned}$$

Letting $m \rightarrow \infty$ above gives the desired conclusion.

Lemma 2. *With the same notations as in Lemma 1, assume that $EX_1 = 0$, $E|X_1|^q < \infty$ for some $1 \leq q \leq 2$ and $b^q(n) \geq n$ for all n . Then given any positive integer r and $0 < \delta < 1$, there exist positive constants α , C and integer m_0 depending only on r , δ , q and $E|X_1|^q$ such that $1 - \delta < \alpha < 1$ and for all $m \geq m_0$ and $\varepsilon > 0$,*

$$EN_m^r I_{[S_m \geq \varepsilon b(m)]} \geq Cm^r P[S_{[\alpha m]} \geq ((r+1)\delta + \varepsilon)b(m)]. \tag{2.2}$$

Proof. Let $1 - \delta < A < 1$ be a constant such that

$$4(1-A)E|X_1|^q \leq \delta^q. \tag{2.3}$$

Take any positive integer k and positive integers $1 \leq i_1 \leq \dots \leq i_k$ satisfying $i_k \geq Ai_{k+1}$. Then by the Markov inequality and an inequality due to von Bahr and Esseen [5],

$$\begin{aligned} P[S_{i_k} - S_{i_{k-1}} \geq -\delta b(i_k)] &\geq 1 - P[|S_{i_k} - S_{i_{k-1}}| > \delta b(i_k)] \\ &\geq 1 - 2(\delta b(i_k))^{-q} (i_k - i_{k-1}) E|X_1|^q \geq \frac{1}{2}. \end{aligned} \tag{2.4}$$

The last inequality above follows from (2.3) and the assumption that $b^q(n) \geq n$. Therefore

$$\begin{aligned} P\left(\bigcap_{j=1}^k [S_{i_j} \geq \varepsilon b(i_j)]\right) &\geq P\left(\bigcap_{j=1}^k [S_{i_j} \geq \varepsilon b(i_k)]\right) \\ &\geq P[S_{i_k} - S_{i_{k-1}} \geq -\delta b(i_k), S_{i_{k-1}} \geq (\delta + \varepsilon)b(i_k), S_{i_{k-2}} \geq \varepsilon b(i_k), \dots, S_{i_1} \geq \varepsilon b(i_k)] \\ &\geq \frac{1}{2} P\left(\bigcap_{j=1}^{k-1} [S_{i_j} \geq (\delta + \varepsilon)b(i_k)]\right) \\ &\geq \dots \geq \left(\frac{1}{2}\right)^{k-1} P[S_{i_1} \geq ((k-1)\delta + \varepsilon)b(i_k)]. \end{aligned} \tag{2.5}$$

Let $\alpha = (1+A)/2$ and $m_0 \geq 2/(1-A)$. Then for $m \geq m_0$, $[\alpha m] \geq Am$ and

$$\begin{aligned} EN_m^r I_{[S_m \geq \varepsilon b(m)]} &\geq E\left\{ I_{[S_{[\alpha m]} \geq \varepsilon b([\alpha m])]} \left(\sum_{i=[\alpha m]}^m I_{[S_i \geq \varepsilon b(i)]} \right)^r I_{[S_m \geq \varepsilon b(m)]} \right\} \\ &\geq ((1-\alpha)m)^r \left(\frac{1}{2}\right)^{r-1} P[S_{[\alpha m]} \geq ((r+1)\delta + \varepsilon)b(m)]. \end{aligned} \tag{2.6}$$

The last relation follows from (2.5) and the multinomial expansion of $\left(\sum_{i=[\alpha m]}^m I_{[S_i \geq \varepsilon b(i)]}\right)^r$ since there are at most r distinct factors in each of the $(m - [\alpha m])^r$ terms (not necessarily distinct) of the multinomial expansion.

Lemma 3. *With the same notations as in Lemma 1, if $EN^\gamma < \infty$ for some $0 < \gamma < 1$, then $\sum n^{\gamma-1} P[S_n \geq b(n)] < \infty$. Now assume that*

$$EN^r < \infty \quad \text{for some } r \geq 1. \quad (2.7)$$

(i) *If $EX_1 = 0$, $E|X_1|^q < \infty$ for some $1 \leq q \leq 2$ and $b^q(n) \geq n$ for all large n , then given any $\delta > 0$, there exists α such that $1 - \delta < \alpha < 1$ and*

$$\sum n^{r-1} P[S_{[\alpha n]} \geq ((r+1)\delta + 1)b(n)] < \infty. \quad (2.8)$$

(ii) *If X_1 is symmetric, then given any $0 < \alpha < 1$,*

$$\sum n^{r-1} P[S_{[\alpha n]} \geq b(n)] < \infty. \quad (2.9)$$

Proof. By Lemma 1 (ii), $EN^\gamma < \infty$ implies that $\sum n^{\gamma-1} P[S_n \geq b(n)] < \infty$ if $0 < \gamma < 1$. Now assume (2.7). To prove (i), since changing $b(n)$ for finitely many n 's does not change (2.7), we can without loss of generality assume that $b^q(n) \geq n$ for all n . If r is an integer, then (2.8) follows from Lemma 1 (i) and Lemma 2. If r is not an integer, then by Lemma 1 (ii), $\sum n^{r-[r]-1} EN_n^{[r]} I_{[S_n \geq b(n)]} < \infty$, so an application of Lemma 2 gives (2.8). To prove (ii), since X_1 and $-X_1$ have the same distribution, (2.7) implies that $E\tilde{N}^r < \infty$. An application of Lemmas 2 and 3 of [14] then proves (2.9) in a similar manner.

Lemma 4. *Suppose X_1, X_2, \dots are independent symmetric random variables and a_1, a_2, \dots are positive constants. Let $X'_n = X_n I_{[|X_n| \leq a_n]}$, $S'_n = X'_1 + \dots + X'_n$ and $S_n = X_1 + \dots + X_n$. Then for any $\varepsilon > 0$,*

$$P[S_n \geq \varepsilon] \geq \frac{1}{2} P[S'_n \geq \varepsilon]. \quad (2.10)$$

Proof. Let $X''_n = X_n - X'_n$, $S''_n = S_n - S'_n$. By symmetry, (X'_n, X''_n) and $(X'_n, -X''_n)$ have the same distribution. Hence by independence, (S'_n, S''_n) and $(S'_n, -S''_n)$ have the same distribution. Therefore

$$\begin{aligned} P[S'_n \geq \varepsilon] &\leq P[S'_n \geq \varepsilon, S''_n \geq 0] + P[S'_n \geq \varepsilon, S''_n \leq 0] \\ &= 2P[S'_n \geq \varepsilon, S''_n \geq 0] \leq 2P[S_n \geq \varepsilon]. \end{aligned}$$

3. Proof of Theorems

Proof of Theorem 1. Under the assumptions $EX_1 = 0$, $E|X_1|^q$ for some $1 \leq q \leq 2$ and $\alpha > 1/q$, it follows from the corollary to Theorem 1 of [2] (see also Theorem 5 of that paper) that the following statements are equivalent:

$$E(X_1^+)^{(r+1)/\alpha} < \infty; \quad (3.1)$$

$$EL(\varepsilon n^\alpha) < \infty \quad \text{for all } \varepsilon > 0; \quad (3.2)$$

$$\sum n^{r-1} P[S_n \geq \varepsilon n^\alpha] < \infty \quad \text{for some } \varepsilon > 0. \quad (3.3)$$

We shall now show that these statements are equivalent to each of the following statements:

$$EN^r(\varepsilon n^\alpha) < \infty \quad \text{for all } \varepsilon > 0; \quad (3.4)$$

$$EN^r(\varepsilon n^\alpha) < \infty \quad \text{for some } \varepsilon > 0. \quad (3.5)$$

Since $N(\varepsilon n^\alpha) \leq L(\varepsilon n^\alpha)$, (3.2) implies (3.4). Clearly (3.4) implies (3.5). Since $q\alpha > 1$, it follows from Lemma 3 that (3.5) implies (3.3). The proof of Theorem 1 is complete.

Proof of Theorem 2. By Theorem 4 of [2], (1.13) is equivalent to each of the following statements:

$$\sum n^{r-1} P[\sup_{k \geq n} (k \log k)^{-1/2} S_k \geq \varepsilon] < \infty \quad \text{for all } \varepsilon > (2r)^{1/2} \sigma; \quad (3.6)$$

$$\sum n^{r-1} P[S_n \geq \varepsilon (n \log n)^{1/2}] < \infty \quad \text{for some } \varepsilon > 0. \quad (3.7)$$

Clearly (3.6) \Leftrightarrow (1.14) \Rightarrow (1.17) \Rightarrow (3.7). It is obvious from (1.8) that (1.14) \Rightarrow (1.15). Since (1.15) \Rightarrow (1.16), it remains to prove (1.16) \Rightarrow (3.7). This implication follows from Lemma 3 since $\varepsilon^2 n \log n > n$ for all large n .

Proof of Theorem 4. To prove (i), since $0 < \varepsilon < (2r EX_1^2)^{1/2}$, we can choose $\alpha > 1$ and c such that

$$\alpha \varepsilon < (2r \text{Var } X_1 I_{\{|X_1| \leq c\}})^{1/2}. \quad (3.8)$$

Define

$$X'_n = X_n I_{\{|X_n| \leq c\}}, S'_n = X'_1 + \cdots + X'_n, \tau^2 = E(X'_1)^2. \quad (3.9)$$

Since X_1 is symmetric, $EX'_1 = 0$ and we obtain by Lemma 4 that

$$P[S_n \geq \alpha \varepsilon (n \log n)^{1/2}] \geq \frac{1}{2} P[S'_n \geq \alpha \varepsilon (n \log n)^{1/2}]. \quad (3.10)$$

Choose $k \geq 6$ such that $\alpha \varepsilon / \tau < (k-2)^{1/2}$. Then by an inequality due to Esseen [4, p. 75-76],

$$\begin{aligned} P[S'_n / (\tau n^{1/2}) \geq (\alpha \varepsilon / \tau) (\log n)^{1/2}] \\ \geq 1 - \Phi((\alpha \varepsilon / \tau) (\log n)^{1/2}) - c_1 n^{-1/2} \{1 + (\alpha \varepsilon / \tau)^3 (\log n)^{3/2}\} \\ \cdot \exp(-\alpha^2 \varepsilon^2 \log n / (2\tau^2)) - c_2 n^{-(k-2)/2} \\ = (c + o(1)) (\log n)^{-1/2} \exp(-\alpha^2 \varepsilon^2 \log n / (2\tau^2)), \end{aligned} \quad (3.11)$$

where c, c_1, c_2 are positive constants and Φ is the distribution function of the standard normal distribution. From (3.8) and (3.11), we obtain that

$$\sum n^{r-1} P[S'_n \geq \alpha \varepsilon (n \log n)^{1/2}] = \infty,$$

so by (3.10),

$$\sum n^{r-1} P[S_n \geq \alpha \varepsilon (n \log n)^{1/2}] = \infty. \quad (3.12)$$

If $0 < r < 1$, then by Lemma 3, (3.12) implies that $EN^r(\varepsilon (n \log n)^{1/2}) = \infty$. Now

assume $r \geq 1$ and take $1 < \beta < \alpha$. It is easy to see that (3.12) implies that

$$\sum n^{r-1} P[S_{[\beta^{-2}n]} \geq \varepsilon(n \log n)^{1/2}] = \infty. \quad (3.13)$$

Therefore by Lemma 3 (ii), $EN^r(\varepsilon(n \log n)^{1/2}) = \infty$. Hence we have proved part (i) of Theorem 4.

To prove part (ii) of the theorem, since X_1 need not be symmetric, instead of defining X'_n by (3.9), we define

$$\begin{aligned} X'_n &= X_n I_{[|X_n| \leq c]} - EX_1 I_{[|X_1| \leq c]}, & \tau^2 &= E(X'_1)^2, \\ X''_n &= X_n - X'_n, & S'_n &= X'_1 + \cdots + X'_n, & S''_n &= X''_1 + \cdots + X''_n, \end{aligned} \quad (3.14)$$

where we choose $\alpha > 1$ and c so large that (3.8) is satisfied and

$$2rE(X'_1)^2 < (\theta\varepsilon)^2 \quad \text{with } 1 < 1 + \theta < \alpha. \quad (3.15)$$

Take $\beta > 1$ such that $\beta + \theta < \alpha$. Then

$$\begin{aligned} \sum n^{r-1} P[S_n \geq \beta\varepsilon(n \log n)^{1/2}] + \sum n^{r-1} P[|S''_n| \geq \theta\varepsilon(n \log n)^{1/2}] \\ \geq \sum n^{r-1} P[S'_n \geq \alpha\varepsilon(n \log n)^{1/2}]. \end{aligned} \quad (3.16)$$

As before, making use of Esseen's inequality as in (3.11), we obtain that

$$\sum n^{r-1} P[S'_n \geq \alpha\varepsilon(n \log n)^{1/2}] = \infty. \quad (3.17)$$

In view of (3.15), it follows from Theorem 3 of [8] that

$$\sum n^{r-1} P[|S''_n| \geq \theta\varepsilon(n \log n)^{1/2}] < \infty. \quad (3.18)$$

From (3.16), (3.17) and (3.18), it follows that

$$\sum n^{r-1} P[S_n \geq \beta\varepsilon(n \log n)^{1/2}] = \infty. \quad (3.19)$$

Take $0 < \delta < 1$ such that $(1 - \delta)^{-1/2} \{(r + 1)\delta + 1\} < \beta$. It is easy to see from (3.19) that for all α with $1 - \delta < \alpha < 1$,

$$\sum n^{r-1} P[S_{[\alpha n]} \geq ((r + 1)\delta + 1)\varepsilon(n \log n)^{1/2}] = \infty. \quad (3.20)$$

Since $\varepsilon^2(n \log n) > n$ for all large n , it follows from Lemma 3(i) that

$$EN^r(\varepsilon(n \log n)^{1/2}) = \infty \quad (3.21)$$

when $r \geq 1$. When $0 < r < 1$, (3.19) implies (3.21) by Lemma 3.

4. General Upper-Class Boundaries and the Convergence Rate of Tail Probabilities for Sample Sums

In [13, pp. 339–340], Strassen has proved the following theorem: If $S_n = X_1 + \cdots + X_n$, where X_1, X_2, \dots are i.i.d. with $EX_1 = 0$, $EX_1^2 = \sigma^2$ and $E|X_1|^p < \infty$ for some $p > 2(r + 1)$, and if $t^{-1/2}b(t) \uparrow \infty$ while $t^{-q}b(t) \downarrow 0$ as $t \uparrow \infty$

for some $q > 1/2$, then the following two statements are equivalent:

$$\int_1^\infty t^{r-3/2} b(t) \exp(-b^2(t)/2\sigma^2 t) dt < \infty; \tag{4.1}$$

$$E\tilde{L}(b(n)) < \infty. \tag{4.2}$$

In particular, if for all large t ,

$$b(t) = \{2\sigma^2 t(r \log t + \frac{3}{2} \log_2 t + \log_3 t + \dots + \log_{k-1} t + \delta \log_k t)\}^{1/2},$$

where $\log_2 t = \log \log t$, etc., then

$$E\tilde{L}(b(n)) = \infty \text{ or } < \infty \text{ according as } \delta \leq 1 \text{ or } \delta > 1. \tag{4.3}$$

In what follows, we shall assume $EX_1 = 0$ and consider boundaries $b(t)$ which satisfy (4.1). We have seen in (1.6) that if $b(t) = t^\alpha$ ($\alpha > \frac{1}{2}$), then (4.2) holds if and only if $E|X_1|^{(r+1)/\alpha} < \infty$, i.e., $E(\Psi(|X_1|))^{r+1} < \infty$, where $\Psi(t) = t^{1/\alpha}$ is the inverse function of $b(t)$. Likewise Theorem 3 asserts that if $b(t) = \varepsilon(t \log t)^{1/2}$, $t \geq 1$, where $\varepsilon > (2r)^{1/2} \sigma$, then (4.2) holds if and only if $E(\Psi(|X_1|))^{r+1} < \infty$, where

$$\Psi(t) \sim t^2 / (2\varepsilon^2 \log t)$$

is the inverse function of $b(t)$. This observation suggests the general analogue in Theorems 5 and 6 below for the last exit time of the region bounded by general boundaries $b(t)$ and $-b(t)$.

In the proof of Theorems 1, 2 and 4 in Section 3, we have seen that the relation (4.2) is closely related to the convergence of a certain type of series. In [1], Baum and Katz studied some series of this type and proved that for $\alpha > \frac{1}{2}$ and $r > 0$, the following three statements are equivalent:

$$E|X_1|^{(r+1)/\alpha} < \infty \quad \text{and in the case } \alpha \leq 1, EX_1 = 0; \tag{4.4}$$

$$\sum n^{r-1} P[\sup_{k \geq n} k^{-\alpha} |S_k| \geq \varepsilon] < \infty \quad \text{for all } \varepsilon > 0; \tag{4.5}$$

$$\sum n^{r-1} P[|S_n| \geq \varepsilon n^\alpha] < \infty \quad \text{for some } \varepsilon > 0. \tag{4.6}$$

This result generalizes an earlier theorem of Hsu and Robbins [6] and Erdős [3] who consider the special case $\alpha = r = 1$, and the proof given by Baum and Katz follows closely that of Erdős. In [2], by using a different approach, Chow and Lai have obtained inequalities, called Paley-type inequalities, which relate the series in (4.5) or (4.6) with $E|X_1|^{(r+1)/\alpha}$ and thereby give another proof of the result of Baum and Katz. By an extension of the method of [2], we can prove the following generalization of the Hsu-Robbins-Erdős-Baum-Katz theorem.

Theorem 5. *Let X_1, X_2, \dots be i.i.d. random variables such that $EX_1 = 0$ and $E|X_1|^q < \infty$ for some $1 \leq q \leq 2$, and let $S_n = X_1 + \dots + X_n$. Let b be a real-valued function on $[1, \infty)$ satisfying the following condition:*

$$b(t) \text{ is ultimately nondecreasing, } \liminf_{t \rightarrow \infty} b(\delta t)/b(t) > 1 \text{ for all large } \delta, \text{ and} \\ \text{there exists } \beta > 1/q \text{ such that } \lim_{t \rightarrow \infty} t^{-\beta} b(t) = \infty. \tag{4.7}$$

Define $\Psi(x) = \inf\{t: b(t) > x\}$ for $x \geq 0$. Then for any $r > 0$, the following statements are equivalent:

$$E(\Psi(X_1^+))^{r+1} < \infty; \quad (4.8)$$

$$\sum n^{r-1} P[\max_{j \leq n} S_j \geq b(\varepsilon n)] < \infty \quad \text{for all } \varepsilon > 0; \quad (4.9)$$

$$\sum n^{r-1} P[S_n \geq \varepsilon b(\varepsilon n)] < \infty \quad \text{for some } \varepsilon > 0; \quad (4.10)$$

$$EL(b(n)) < \infty; \quad (4.11)$$

$$EN^r(b(n)) < \infty. \quad (4.12)$$

Proof. We shall prove (4.8) \Rightarrow (4.9) \Rightarrow (4.11) \Rightarrow (4.12) \Rightarrow (4.10) \Rightarrow (4.8). To prove (4.8) \Rightarrow (4.9), without loss of generality, we can assume that $E|X_1| \neq 0$. Let k be a positive integer such that $k(q\beta - 1) > r$. We note that

$$\begin{aligned} P[\max_{j \leq n} S_j \geq b(\varepsilon n)] &\leq P[\max_{j \leq n} X_j \geq b(\varepsilon n)/2k] \\ &\quad + P[\max_{j \leq n} S_j \geq b(\varepsilon n), \max_{j \leq n} X_j \leq b(\varepsilon n)/2k]. \end{aligned} \quad (4.13)$$

For large t , $x \geq b(t)$ implies that $\Psi(x) \geq t$. Therefore for all large n ,

$$P[\max_{j \leq n} X_j \geq b(\varepsilon n)/2k] \leq nP[X_1 \geq b(\varepsilon n)/2k] \leq nP[\Psi(2kX_1^+) \geq \varepsilon n]. \quad (4.14)$$

Define $\tau_1^{(n)} = \inf\{j \geq 1: S_j \geq b(\varepsilon n)/2k\}$, $\tau_2^{(n)} = \inf\{j \geq 1: S_{q^{(n)}+j} - S_{q^{(n)}} \geq b(\varepsilon n)/2k\}$, etc. Then

$$\begin{aligned} P[\max_{j \leq n} S_j \geq b(\varepsilon n), \max_{j \leq n} X_j \leq b(\varepsilon n)/2k] &\leq P[\tau_i^{(n)} \leq n \text{ for } i = 1, \dots, k] \\ &= P^k[\tau_1^{(n)} \leq n] = P^k[\max_{j \leq n} S_j \geq b(\varepsilon n)/2k] \\ &\leq \{(2k/b(\varepsilon n))^q E|S_n|^q\}^k, \quad \text{by the submartingale inequality} \\ &\leq \{(4k/b(\varepsilon n))^q n E|X_1|^q\}^k, \quad \text{by the Esseen-von Bahr inequality [5]} \\ &= o(n^{-k(q\beta-1)}). \end{aligned} \quad (4.15)$$

The last relation above follows from the assumption that $\lim_{n \rightarrow \infty} n^{-\beta} b(n) = \infty$. It is easy to see from (4.7) that

$$\limsup_{x \rightarrow \infty} \Psi(\eta x)/\Psi(x) < \infty \quad \text{for any } \eta > 1. \quad (4.16)$$

Hence (4.8) implies that $E(\Psi(2kX_1^+))^{r+1} < \infty$, and therefore

$$\sum n^r P[\Psi(2kX_1^+) \geq \varepsilon n] < \infty. \quad (4.17)$$

Since $k(q\beta - 1) > r$, it follows from (4.15) that

$$\sum n^{r-1} P[\max_{j \leq n} S_j \geq b(\varepsilon n), \max_{j \leq n} X_j \leq b(\varepsilon n)/2k] < \infty. \quad (4.18)$$

From (4.13), (4.14), (4.17) and (4.18), we obtain (4.9).

We shall now prove that (4.9) \Rightarrow (4.11). Without loss of generality, we can

assume that $b(t)$ is nondecreasing for $t \geq 1$. We need only note that

$$\begin{aligned} \sum_{i=1}^{\infty} 2^{ri} P[S_n \geq b(n) \text{ for some } n \geq 2^i] &\leq \sum_{i=1}^{\infty} 2^{ri} \sum_{j=i}^{\infty} P[\max_{n \leq 2^{j+1}} S_n \geq b(2^j)] \\ &\leq c \sum_{j=1}^{\infty} 2^{rj} P[\max_{n \leq 2^{j+1}} S_n \geq b(2^j)] \\ &\leq c_1 \sum_{j=0}^{\infty} \sum_{2^{j+1} \leq m < 2^{j+2}} m^{r-1} P[\max_{n \leq m} S_n \geq b(m/4)] \\ &= c_1 \sum_{m=2}^{\infty} m^{r-1} P[\max_{n \leq m} S_n \geq b(m/4)]. \end{aligned}$$

It is obvious from (1.8) that (4.11) \Rightarrow (4.12). To prove (4.12) \Rightarrow (4.10), we apply Lemma 3. If $0 < r < 1$, then the desired result is immediate from Lemma 3. Now assume that $r \geq 1$. By (4.7), $b^q(n) > n$ for all large n . Therefore if (4.12) holds, then by Lemma 3, there exists $\frac{1}{2} < \alpha < 1$ such that

$$\sum n^{r-1} P[S_{\lfloor \alpha n \rfloor} \geq (\frac{1}{2}(r+1)+1)b(n)] < \infty. \tag{4.19}$$

It is obvious that (4.19) \Rightarrow (4.10).

To prove (4.10) \Rightarrow (4.8), we follow the argument due to Erdős [3]. For $k=1, \dots, n$, let $A_k^{(n)} = [X_k \geq 2\epsilon b(\epsilon n)]$, $B_k^{(n)} = [|\sum_{1 \leq j \leq n, j \neq k} X_j| \leq \epsilon b(\epsilon n)]$. Since $\beta > 1/q$ and $E|X_1|^q < \infty$, $n^{-\beta} S_{n-1} \rightarrow 0$ a.s. by the Marcinkiewicz-Zygmund strong law of large numbers. Therefore by (4.7), $\lim_{n \rightarrow \infty} P(B_1^{(n)}) = 1$. Also for all large n , $P(A_1^{(n)}) \leq P[X_1 \geq n^{1/q}] = o(1/n)$ since $E|X_1|^q < \infty$. Therefore we can choose n_0 such that for $n \geq n_0$, $P(B_1^{(n)}) - nP(A_1^{(n)}) \leq 1/2$. Hence for $n \geq n_0$,

$$\begin{aligned} P[S_n \geq \epsilon b(\epsilon n)] &\geq \sum_{k=1}^n \{P(A_k^{(n)} \cap B_k^{(n)}) - P(A_k^{(n)} \cap (A_1^{(n)} \cup \dots \cup A_{k-1}^{(n)}))\} \\ &\geq \sum_{k=1}^n P(A_k^{(n)}) (P(B_1^{(n)}) - nP(A_1^{(n)})) \geq \frac{1}{2} nP(A_1^{(n)}). \end{aligned}$$

Hence (4.10) implies that

$$\sum n^r P[X_1 \geq 2\epsilon b(\epsilon n)] < \infty. \tag{4.20}$$

Since $\Psi(x) > t$ implies $x \geq b(t)$ for all large t , it is easy to see using (4.16) that (4.20) implies (4.8).

By a modification of the proof of Theorem 5, we obtain the following theorem which considers boundaries $b(t)$ satisfying weaker growth conditions than (4.7) but under stronger two-sided moment conditions on X_1 .

Theorem 6. *Let X_1, X_2, \dots be i.i.d. random variables such that $EX_1 = 0$ and $E|X_1|^{2+\theta} < \infty$ for some $\theta > 0$. Let $S_n = X_1 + \dots + X_n$, and let b be a real-valued function on $[1, \infty)$ satisfying the following condition:*

$$\begin{aligned} &b(t) \text{ is ultimately nondecreasing, } \liminf_{t \rightarrow \infty} b(\delta t)/b(t) > 1 \\ &\text{for all large } \delta \text{ and } \lim_{t \rightarrow \infty} (t \log t)^{-1/2} b(t) = \infty. \end{aligned} \tag{4.21}$$

Define $\Psi(x) = \inf\{t: b(t) > x\}$ for $x \geq 0$. Then for any $r > 0$, the statements (4.8), (4.9), (4.10), (4.11) and (4.12) are equivalent.

Proof. We first prove (4.8) \Rightarrow (4.9). Without loss of generality, we can assume that $0 < \theta < 1$. Choose a positive integer k such that $k\theta/2 > r$. Letting $\sigma^2 = EX_1^2$, we can without loss of generality assume that $\sigma \neq 0$. As in the proof of Theorem 5, we obtain that

$$P[\max_{j \leq n} S_j \geq b(\varepsilon n), \max_{j \leq n} X_j \leq b(\varepsilon n)/2k] \leq P^k[\max_{j \leq n} S_j \geq b(\varepsilon n)/2k]. \quad (4.22)$$

By the Lévy inequality (cf. [10, p. 248]),

$$\begin{aligned} P[\max_{j \leq n} S_j \geq b(\varepsilon n)/2k] &\leq 2P[S_n \geq b(\varepsilon n)/2k - (2n\sigma^2)^{1/2}] \\ &\leq 2P[S_n \geq b(\varepsilon n)/3k]. \end{aligned} \quad (4.23)$$

Using (1.11), we obtain that for all large n

$$\begin{aligned} P^k[S_n \geq b(\varepsilon n)/3k] &\leq \{1 - \Phi(b(\varepsilon n)/(3k\sigma n^{1/2})) + cn^{-\theta/2}\}^k \\ &\leq c_1 \exp(-b^2(\varepsilon n)/(18k\sigma^2 n)) + c_2 n^{-k\theta/2} \end{aligned} \quad (4.24)$$

where c , c_1 and c_2 are positive constants. Since $\lim_{n \rightarrow \infty} b^2(\varepsilon n)/(n \log n) = \infty$ and $k\theta/2 > r$, it follows from (4.23) and (4.24) that

$$\sum n^{r-1} P^k[\max_{j \leq n} S_j \geq b(\varepsilon n)/2k] < \infty. \quad (4.25)$$

Using (4.22) and (4.25), we can prove the implication (4.8) \Rightarrow (4.9) as in Theorem 5.

The proof of (4.9) \Rightarrow (4.11) in Theorem 5 carries over and since $\lim_{n \rightarrow \infty} b^2(n)/n = \infty$, so does the proof of (4.12) \Rightarrow (4.10). Since $EX_1^2 < \infty$, $(n \log n)^{-1/2} S_n \rightarrow 0$ a.s. and $P[X_1 \geq n^{1/2}] = o(1/n)$. Therefore we can again apply the argument due to Erdős as in Theorem 5 to prove (4.10) \Rightarrow (4.8).

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