

# The epidemiology of diabetes mellitus in the United Kingdom: The Yorkshire Regional Childhood Diabetes Register

A. Staines<sup>1</sup>, H. J. Bodansky<sup>2</sup>, H. E. B. Lilley<sup>1</sup>, C. Stephenson<sup>2</sup>, R. J. Q. McNally<sup>1</sup>, R. A. Cartwright<sup>1</sup>

<sup>1</sup> Leukaemia Research Fund Unit for Clinical Epidemiology, University of Leeds, Leeds, UK

<sup>2</sup> Academic Unit of General Medicine, University of Leeds, Leeds General Infirmary, Leeds, UK

**Summary.** A register of the incidence of Type 1 (insulin-dependent) diabetes mellitus in the Yorkshire region of the United Kingdom has been completed. A total of 1,490 subjects aged between 0 and 16 years were identified from 1978 to 1990, giving an incidence of 13.7 per 100,000 (ages 0–14) or 13.6 per 100,000 (ages 0–16), comparable to other recent studies in the United Kingdom. An age-period-cohort analysis shows evidence for a modest drift effect of 1.75% per year (95% confidence interval 0.28 to 3.25%). There is a marked epidemic pattern with peaks at 4-year intervals. The age-incidence curve is similar to that reported elsewhere,

having peaks in early childhood and puberty. Girls have an earlier pubertal peak than boys. There is substantial seasonal variation in incidence confined to those over 5 years of age. Ascertainment is believed to be very complete, and is estimated to be 97.6% (95% confidence interval 97.2% to 98.1%).

**Key words:** Type 1 (insulin-dependent) diabetes mellitus, epidemiology, aetiology, disease registry, childhood, ascertainment estimates.

Type 1 (insulin-dependent) diabetes mellitus is an increasingly common disorder in many European populations [1, 2]. Neither the causes of the condition, nor the causes of the increased incidence are known. Although a genetic predisposition is clearly of great importance, direct epidemiological evidence from case-control studies [3, 4] and indirect epidemiological evidence from migrant studies [5, 6] indicate that environmental agents are also very important.

The descriptive epidemiology of Type 1 diabetes is an important source of information for the framing of aetiological hypotheses. Several prominent features require explanation. There is great geographical variation in incidence, with as much as a 40-fold variation between countries; Finland and Sardinia have the highest incidence rates (over 30 per 100,000 per year); Japan, Africa and Korea have the lowest (0.3 to 2.5 per 100,000 per year) [7–13]. The age incidence curve typically shows a peak in early adolescence, and may have a secondary peak earlier in childhood. There is marked seasonality of onset in many (though not all) studies [12, 14–17], although this may relate more to infections precipitating a metabolic decompensation in pre-diabetic persons, than to any aetiological agent [8]. The incidence in a population may change rapidly during short periods of time, whether in the indigenous population [18], or in migrant populations [5].

A necessary first step in any series of investigations of the aetiology of Type 1 diabetes is to establish a good, well-validated register. This was the aim of the Yorkshire project.

## Subjects, materials and methods

### Subjects

A register of all diabetic children diagnosed while living in the Yorkshire Regional Health Authority (West Yorkshire, North Yorkshire and Humberside) in the United Kingdom was established in 1989. The Regional Health Authority had an average resident population of 840,000 persons under the age of 17 years during the period of this study.

Children newly diagnosed, satisfying standard criteria for the diagnosis of Type 1 diabetes, were recorded retrospectively back to 1978, and prospectively from 1989 onwards. All children diagnosed while aged under 17 years, and usually resident in the region at the time of diagnosis were eligible. Children with cystic fibrosis (five subjects), or who had undergone a pancreatectomy (one subject) were excluded.

Subjects were ascertained from three independent sources. First, all of the diabetologists and paediatricians in the region were approached, and all agreed to participate. Subjects were ascertained from their clinic records, and from the records of diabetic nurse specialists attached to the clinics. Secondly, all of the general practition-

**Table 1.** Numbers of cases identified through each method. Hospital indicates cases found in a search of hospital and clinic records; GP survey indicates cases found during a postal survey of general practitioners in Yorkshire; Tape indicates cases identified from a Hospital Activity Analysis computer tape

Hospital	Tape	GP survey	
		No	Yes
No	No	?	45
	Yes	118	108
Yes	No	44	59
	Yes	481	635

**Table 2.** Results of the best fitting log-linear model for estimating the completeness of ascertainment by time period

Estimated completeness of ascertainment			
Time period	Estimate	95% Confidence limits	
		Lower	Upper
1978–1980	98.8%	98.1%	99.6%
1981–1983	96.4%	92.9%	99.7%
1984–1986	98.8%	97.7%	99.8%
1987–1990	97.1%	95.6%	98.4%
1978–1990	97.7%	97.2%	98.1%

**Table 3.** Age-specific incidence of diabetes, per 100,000 persons per year, for males and females separately, and combined, in Yorkshire from 1978 to 1990, standardised to the 1981 census population

Age (years)	Male		Female		Combined	
	n	Rate	n	Rate	n	Rate
0–4	149	9.8	139	9.6	288	9.7
5–9	210	13.1	204	13.4	414	13.3
10–14	320	18.2	283	16.9	603	17.6
15–16	119	16.4	60	8.7	179	12.6

**Table 4.** Overall incidence per 100,000 persons per year reported from other studies in the United Kingdom

Study	Year of study	Age (years)	Incidence
Yorkshire <sup>a</sup>	1978–1990	0–16	13.6
Oxford 1 [29]	1985–1986	0–20	15.6
Oxford 2 [10]	1989–1990	0–14	16.4
British Isles [41]	1988	0–14	13.5
Scotland [28]	1977–1983	0–18	21.0

<sup>a</sup> Present study

**Table 5.** Standardised incidence ratios (SIR) by county of residence at diagnosis. The reference is the entire series

County	Cases	SIR	95% Confidence interval	p value
West Yorkshire	755	87.2	81.0–93.4	< 0.001
Humberside	438	121.2	109.8–132.5	< 0.001
North Yorkshire	291	113.3	100.2–126.3	0.04

Significance test for between-county heterogeneity:  $p < 0.0001$  ( $\chi^2 = 34.29$  on 2 *df*)

ers in the region were asked (by post) to provide a list of all known patients with diabetes, diagnosed under the age of 17 years, on their practice lists. One reminder letter was sent to those who did not reply within 4 weeks. Thirdly, a list of all admissions to hospitals in the Yorkshire region from 1969 to 1992, where a discharge diagnosis of

**Table 6.** Results of the age-period-cohort analysis. The best fitting model (with a Poisson error distribution, a logarithmic link function and the logarithm of the person-years at risk in each cell as an offset variable) had main effects for age, sex, county of residence, a drift term, a non-linear period term, and an age by sex interaction. The residual deviance after fitting was 96.44 on 83 *df*

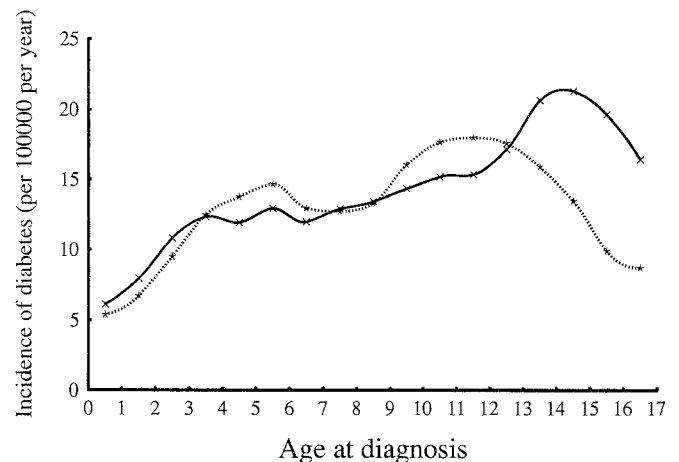
Effect	7.35 per 100,000 per year <sup>a</sup>	
	Rate Ratio	95% Confidence interval
Baseline rate	7.35 per 100,000 per year <sup>a</sup>	
Age group 1	1.00	– <sup>b</sup>
Age group 2	1.37	0.08, 1.74
Age group 3	1.82	1.45, 2.26
Age group 4	2.38	1.91, 2.95
West Yorkshire	1.00	– <sup>b</sup>
North Yorkshire	1.28	1.12, 1.47
Humberside	1.39	1.23, 1.56
Male	1.00	– <sup>b</sup>
Female	0.96	0.74, 1.25
Linear drift (per period)	1.06	1.01, 1.11
Non-linear period 1	1.00	– <sup>b</sup>
Non-linear period 2	0.84	0.73, 0.96
Non-linear period 3	0.95	0.83, 1.07
Non-linear period 4	1.00	– <sup>b</sup>
Age group 1. Female	1.00	– <sup>b</sup>
Age group 2. Female	1.10	0.78, 1.55
Age group 3. Female	1.15	0.83, 1.59
Age group 4. Female	0.67	0.49, 0.93

Age group 1, 0–4.25 years; 2, 4.2–8.5 years; 3, 8.5–12.75 years; 4, 12.75–under 17 years.

Period 1, 1 January 1978–31 March 1981; 2, 1 April 1981–30 June 1984; 3, 1 July 1984–30 September 1987; 4, 1 October 1987–31 December 1990

<sup>a</sup> Baseline is males aged 0 to 4.25 years, diagnosed between 1 January 1978 and 31 March 1981, and resident in West Yorkshire. All rate ratios are calculated relative to this baseline.

<sup>b</sup> These rate ratios are constrained to be equal to 1 by the model.



**Fig. 1.** Age-specific incidence of diabetes by single year of age from 1978 to 1990. Smoothed using a 3 year moving average. Males ×–, females \*–

diabetes (International Classification of Disease 250) was recorded for a person aged under 17 years, was obtained from the Yorkshire Regional Health Authority.

For each subject identified from any one of these sources the hospital notes were obtained. Following a structured abstraction form, basic demographic data were recorded from the notes. The child's name, sex, address at diagnosis, postcode at diagnosis, date of birth,

date of diagnosis, presenting symptoms, and family history of diabetes were noted. Where a set of notes recorded only follow-up details on a patient, the original notes from the time of diagnosis were obtained.

### Population estimates

Those subjects recorded as having been diagnosed between 1978 and 1990 were allocated to 1981 counties and wards using the frozen version of the central postcode directory [19]. Census small area statistics were extracted from the records of the 1971, 1981 and 1991 censuses held at the Manchester Computing Centre using the SASPAC package [20]. Incidence rates were calculated with locally written software [21].

Population figures were derived for each single year by interpolation from the 1971 census (adjusted to 1974 boundaries), the 1981 census and the 1991 census. For the age-period-cohort analysis linear interpolation within age and sex groups was used to estimate the population for four periods of 3.25 years each.

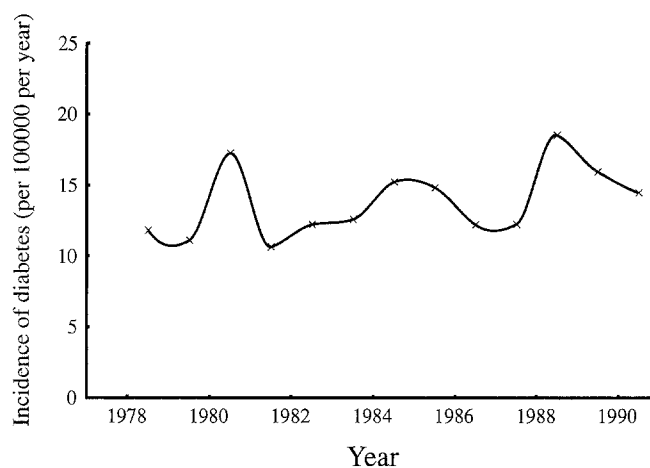
### Statistical analysis

Statistical analyses were performed using Genstat 5.1 [22]. Completeness of ascertainment was estimated using methods based on the capture-recapture technique, analysed using log-linear models [23]. This procedure enables one to test the independence of sources assumption of traditional methods. The age-period-cohort analysis used the method of Clayton and Schifflers [24, 25]. This approach addresses the fundamental problem of the non-identifiability of separate age, period and cohort effects in a novel way. Instead of simply fitting period or cohort effects, these terms are broken down into a linear drift term, and non-linear period or cohort effects. Drift is a linear increase with time, which may be due to period effects, cohort effects, or a combination of both. Due to the identifiability problem it cannot be directly attributed to one or the other. Tests for seasonal variation in incidence were carried out using the method of Walter and Elwood [26].

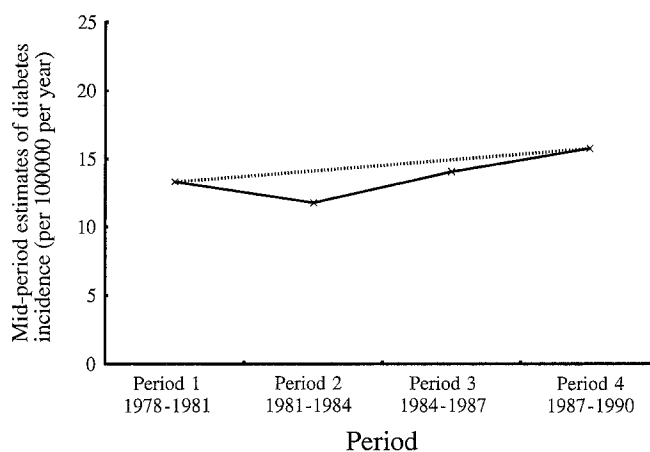
## Results

Between 1978 and 1990, 1,490 children were diagnosed as having Type 1 diabetes in Yorkshire. Of these 1,219 were identified through the hospital records, 847 through the postal survey of general practitioners, and 1,342 from the tape provided by the Regional Health Authority. A total of 635 subjects (43%) were identified by all three methods (Table 1). The response rate to the postal survey was 67%. The estimated completeness of ascertainment was 97.6% (95% confidence interval (CI) 97.2% to 98.1%). There was no evidence for any substantial variation in completeness of ascertainment over the period of the study (Table 2).

The incidence rate over the entire time period was 13.7 per 100,000 per year for ages 0 to 14 years, and 13.6 per 100,000 per year for ages 0 to 16 years (Table 3, Fig. 1). Results from other UK studies are shown in Table 4 for comparison. There are two peaks in incidence, at ages 4 to 6, and at ages 10 to 15. Girls have their peak incidence at age 10, earlier than boys, whose peak incidence is at age 14. The difference between the mean ages at onset for this peak (ages 8 to 16 years) was 0.64 of a year (95% CI 0.34 to 0.95 years). For children 7 years and under, the mean ages at onset for boys and girls were almost identical (dif-



**Fig. 2.** A plot of the annual incidence of diabetes (per 100,000 persons per year), directly standardised to the 1981 census population, in Yorkshire from 1978 to 1990 inclusive

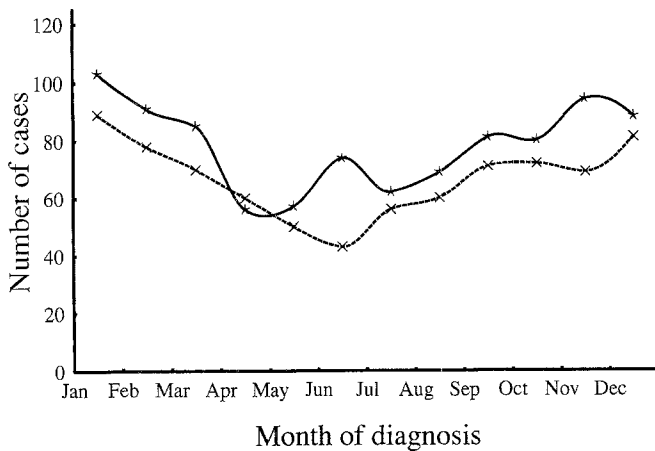


**Fig. 3.** The estimated change in diabetes incidence over the study period (i.e. the drift), and estimates of mid-period diabetes incidence from the ageperiod-cohort model. Drift effect  $\cdots$ , fitted incidence  $\times$ —

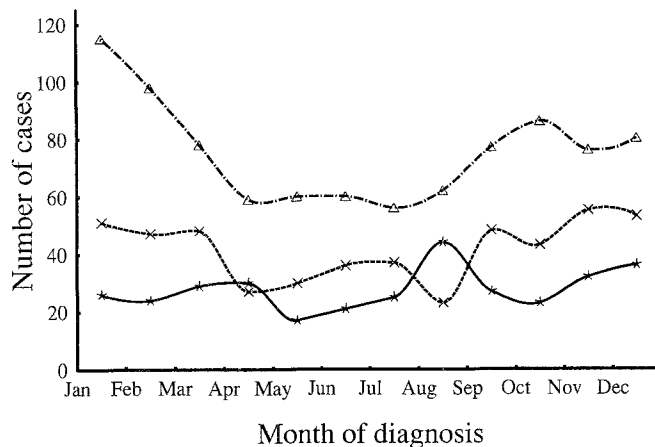
ference 0.14 years; 95% CI -0.4 to 0.2 years). There is a marked variation in incidence between the three counties, with West Yorkshire having the lowest incidence, and Humberside the highest (Table 5).

Figure 2 shows the changes in diabetes incidence during the 13 years from 1978 to 1990. For further analysis four equal age groups, four equal periods, and seven approximate birth cohorts were calculated from the individual records (Table 6). An age-period-cohort analysis was performed. The best fitting model was that which included main effects for age, sex and county, an age by sex interaction term, a drift term, and a non-linear period effect (Table 6). The regular drift effect suggests an increase in incidence from 1978 to 1990 of 1.75% a year (95% CI 0.28% to 3.25%). At this rate the incidence would double in 40 years. The predictions from this model for the middle of each period, and the drift effect are shown in Figure 3. Inspection of the incidence by single years suggests a cyclical pattern with a period of 4 years (Fig. 2).

Marked seasonal variation in incidence was identified with peak incidence in winter ( $\chi^2 = 48.24$ ;  $df = 2$ ;



**Fig. 4.** Diabetes incidence by month and sex showing the similar patterns of seasonal variation among males and females. Males \*—, females ×---



**Fig. 5.** Diabetes incidence by month and age, showing the marked variation in the seasonality of occurrence by age. 0–4 years \*—, 5–9 years ×---, 10 years and older Δ ···

$p < 0.0001$ ). There was no difference in the pattern of seasonal onset between boys and girls ( $\chi^2 = 23.73$ ;  $df = 2$ ;  $p < 0.00001$  for boys;  $\chi^2 = 25.34$ ;  $df = 2$ ;  $p < 0.00001$  for girls) (Fig. 4). There was no seasonal variation among children under 5 years old ( $\chi^2 = 3.21$ ;  $df = 2$ ;  $p = 0.2$ ), some seasonal variation among children aged between 5 and 9 ( $\chi^2 = 21.11$ ;  $df = 2$ ;  $p = 0.00002$ ), and very marked variation in incidence in the oldest children, 10 to 16 years old, ( $\chi^2 = 31.40$ ;  $df = 2$ ;  $p < 0.00001$ ) (Fig. 5).

## Discussion

Our results are broadly comparable to those from other European registries. Studies from Scotland [27, 28], which lies to the north of Yorkshire have reported substantially higher rates of diabetes than our study, while studies from Oxford [10, 29] which lies to the south have reported similar or slightly higher rates. This is consistent with a weak north-south gradient in incidence across the country, although more years of data from Oxford, and more recent data from Scotland would be required to confirm this.

The age-incidence curve is similar to that reported in other studies, both European [8, 30, 31] and Asian [12]. Boys have their peak incidence later than girls, presumably because they tend to enter puberty later than girls, and have a later pubertal growth spurt. This leads to an increased demand for insulin, and could cause a stressed pancreas to decompensate. There is modest evidence for a second peak in early childhood at 5 or 6 years of age.

The pattern of seasonality is similar to that reported in many other studies [8, 12, 14, 31–35] with excess incidence in the local winter. Some studies have not shown this [17, 36]. There is no seasonal variation in the youngest children, and the variation is most marked in the oldest children. This variation may reflect differing levels of exposure to infectious disease, leading to a different temporal pattern of metabolic stress triggering the clinical manifestations of ketoacidosis [8, 17]. Some studies have reported more seasonal variation in boys, than in girls, [8, 33] but we found no noticeable difference.

Our study confirms the general experience of registers in the United Kingdom such as those from Leicester [37] which showed a steady increase in incidence from 1950 to 1980; and with studies from Scotland [27, 28] which showed an increasing incidence from 1968 to 1983. Ascertainment in Leicester is believed to be very high [37], and while the methods used by Patterson et al. [28] have been criticised [2, 38], other studies from Scotland [39] tend to confirm their findings. Comparison of the two British national surveys of diabetes [40, 41] is difficult as the degree of ascertainment in the first study was unknown, and was probably low. Hence the apparent doubling in incidence between these two studies is probably artefactual.

Studies from continental Europe, reviewed by Bingley and Gale [2] suggested that an increasing incidence was a continent-wide event. Tuomilehto et al. [42] confirmed evidence of further increase from Finland, as did Jøner and Sjøvik [31] from Norway, and Nystrom et al. [43] from Sweden. Studies from the United States and Canada have shown either no evidence of an increase over time, or evidence for a very limited increase [15, 44]. The most extensive study yet conducted of temporal trends in incidence [1] confirms this pattern, significant increases from 2.6 to 12.2% per year, in Europe, Japan and New Zealand, but no consistent increase from North America. Our incidence is increasing more slowly than most other areas studied, at 1.75% per year.

Our temporal incidence curve was complex, with three clearly-defined peaks, at approximately 4-year intervals, in 1980, 1984–1985 and 1988. The pattern was similar for all age groups. Peaks in incidence in 1982 or 1983 and 1986 have been identified in several registries in Northern Europe [13, 18, 42, 43, 45, 46], suggesting that there may be large scale patterns in diabetes incidence.

The pattern of cyclical variation in temporal incidence found in our study and others is of great interest. Relatively few environmental variables fluctuate periodically with periods of more than 1 year. Clearly certain economic series behave in this way, and so social causes cannot be excluded. However, infectious diseases also behave in this way, often showing very marked periodicity [47]. One study [48] has linked cyclical variation in diabetes in-

vidence to cyclical variation in reported infections with the mumps virus, although other studies have not reported this [49]. Several other registries have published temporal incidence curves showing multiple peaks in incidence [1, 13, 18, 42, 43, 46], although this is not a universal feature [1]. Clearly short-term variations in incidence are a feature of diabetes epidemiology, a fact which complicates the interpretation of apparent epidemics of the disease.

More speculatively, given the limitations of the existing evidence, one could suggest that these peaks, occurring at different times in different places, may be markers for the continent-wide spread of some infectious agent. Epidemic waves with a very large scale have been identified for many infectious diseases, particularly influenza, plague and cholera [50, 51]. New methods are being developed to study the movements of infections over very large areas [52]. If large scale patterns of this kind can be identified, the implications for theories of the aetiology of diabetes would be substantial.

Further work on these questions will require international collaborative studies. These should be directed towards filling gaps in the areas covered by existing registries, and preparing for transnational analyses of disease variation, on a variety of spatial scales, and over time.

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Dr. A. Staines  
 LRF Unit for Clinical Epidemiology  
 17 Springfield Mount  
 Leeds LS2 9NG  
 UK