

# A Useful Estimate in the Multidimensional **Invariance** Principle

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Summary. An estimate of the convergence speed in the multidimensional invariance principle is obtained. Using this estimate, we can prove strong invariance principles for partial sums of independent not necessarily identically distributed multidimensional random vectors.

# 1. Introduction and Main Results

We denote by  $(\mathbb{R}^d, |\cdot|)$  the d-dimensional euclidean space. Let  $C_d[0, 1]$  be the space of all continuous  $\mathbb{R}^{d}$ -valued functions on [0,1] endowed with the supnorm  $\|\cdot\|$ . Log *t* stands for  $\log(\max(t, e), t \ge 0$ .

Let  $\xi_1, \ldots, \xi_n$ :  $\Omega \to \mathbb{R}^d$  be independent random vectors with zero means,  $\operatorname{cov}(\xi_k) = \sigma_k^2 I$ ,  $1 \le k \le n$  and  $\sum_{1}^n \sigma_k^2 = 1$ , where *I* denotes the *d*-dimensional unit

matrix.

Let  $S_{(n)}: \Omega \to C_d[0,1]$  be the partial sum process which is defined by

(1.1) 
$$S_{(n)}(t) := \sum_{1}^{m} \xi_{k} + \frac{t - t_{m}}{t_{m+1} - t_{m}} \xi_{m+1}, \quad t_{m} \leq t \leq t_{m+1},$$

 $0 \leq m < n$ , where  $t_m := \sum_{k=1}^{m} \sigma_k^2$ ,  $0 \leq m \leq n$ .  $\left( \sum_{k=1}^{0} \xi_k := \sum_{k=1}^{0} \sigma_k^2 := 0 \right)$ .

Denote by  $W_{\Gamma}|B(C_{d}[0,1])$  the Wiener-measure with covariance matrix  $\Gamma$ , i.e. the unique p-measure on the Borel- $\sigma$ -algebra of  $C_d[0,1]$  such that  $W \circ \pi_i$ =  $N(0, t\Gamma)$ ,  $0 \le t \le 1$  and  $\{\pi_t: t \in [0, 1]\}$  is a stochastic process with independent and stationary increments under  $W_{\Gamma}$ , where  $\pi_t$ ;  $0 \leq t \leq 1$  are the canonical projections (i.e.  $\pi_t(f) := f(t), f \in C_d[0, 1], 0 \le t \le 1$ ). Let  $W = W_t$ .

If  $Q_i | B(C_d[0,1])$ , i=1,2, are *p*-measures, we put for each  $\delta > 0$ :

(1.2) 
$$\lambda(Q_1, Q_2, \delta) := \sup \{Q_1(A) - Q_2(A^{\delta}) : A \subseteq C_d[0, 1] \text{ closed}\},\$$

where  $A^{\delta} := \{g \in C_d [0, 1] : \exists f \in A, \|g - f\| < \delta\}.$ 

Call  $\lambda(Q_1, Q_2, \delta)$  the " $\delta$ -distance" of  $Q_1$  and  $Q_2$ .

These  $\delta$ -distances have turned out to be extremely useful especially in treating the Prohorov-distance of *p*-measures (cf. [3, 4, 10, 11, 13]).

In this paper we prove an inequality for the  $\delta$ -distances of the partial sum process  $S_{(n)}$  from the Wiener-measure W (see Theorem 1). Our result concerns independent not necessarily identically distributed *d*-dimensional random vectors and does not require higher than second moments. Thus our inequality can be applied to rather different problems. Our inequality yields a multidimensional version of Prohorov's invariance principle for triangular arrays of rowwise independent r.v.'s fulfilling the Lindeberg-condition (see Corollary 4) as well as unimprovable estimates of the Prohorov-distance of the *p*-measures  $P \circ S_{(n)} | B(C_d[0,1])$  and  $W | B(C_d[0,1])$  (see Corollary 2); it can be used to obtain strong approximation results (see Theorems 2, 3, 5, 6) as well as functional (compact) laws of the iterated logarithm (see Corollaries 5/6).

Sahanenko (cf. [13]) obtained in the 1-dimensional case (i.e. d=1) the following estimate of the  $\delta$ -distances of  $P \circ S_{(m)}$  and W:

(1.3) 
$$\lambda(P \circ S_{(n)}, W, \delta) \leq c(s) \, \delta^{-s} K_{sn}, \quad s > 2,$$

where  $K_{sn} = \sum_{k=1}^{n} E[|\xi_k|^s]$ , c(s) is a positive constant depending on s only.

He used in his proof a refinement of the already rather complicated common probability space method of Komlós et al. (1976), which is restricted to the 1-dimensional case. The question remained open, whether (1.3) holds true in the multidimensional case (cf. [4], p. 64).

Using a recent result in connection with the multidimensional central limit theorem (cf. [8], Theorem 6), we now obtain by a comparatively simple construction method, which is related to the method of Csörgo and Révész (1975), Theorem 1 below enabling us to extend Sahanenko's result to the multidimensional case, if 2 < s < 4. At the same time, we are able to prove (1.3) for  $s \ge 4$  in the multidimensional case, if  $\delta$  is sufficiently large – in particular, if  $\delta \ge K_{sn}^{\gamma}$ , where  $\gamma < 1/(2s-4)$ . This shows that (1.3) can be obtained for any given s > 2 by a much easier construction method than that in [13], if  $\delta$  is large. On the other hand, it can be shown by similar arguments as in [2, Remark 3] that it is impossible to prove (1.3) by such a simple method if  $\delta$  is small and s > 4.

**Theorem 1.** Let  $\xi_1, \ldots, \xi_n: \Omega \to \mathbb{R}^d$  be independent random vectors with zero means,  $\operatorname{cov}(\xi_k) = \sigma_k^2 I$ ,  $1 \leq k \leq n$ , and  $\sum_{1}^{n} \sigma_k^2 = 1$ . Let  $S_{(n)}: \Omega \to C_d[0,1]$  be defined by (1.1). Let s > 2,  $0 < \gamma < 1/(2s-4)$ . Then we have for  $\delta \geq (K_{sn}(\delta) + \delta^{s-2} L_n(\delta))^{\gamma}$ :

(1.4) 
$$\lambda(P \circ S_{(n)}, W, c_1 \delta) \leq c_2(\delta^{-s} K_{sn}(\delta) + \delta^{-2} L_n(\delta)),$$

where  $K_{sn}(\delta) := \sum_{1}^{n} E[|\xi_k|^s \mathbbm{1}\{|\xi_k| \leq \delta\}], \ L_n(\delta) := \sum_{1}^{n} E[|\xi_k|^2 \mathbbm{1}\{|\xi_k| > \delta\}]. \ c_1 = c_1(\gamma, s, d)$ >0 and  $c_2 = c_2(\gamma, s, d) > 0$  are constants depending on  $\gamma$ , s and d only.

**Corollary 1.** Let  $S_{(n)}$ :  $\Omega \to C_d[0,1]$  be as in Theorem 1. Let 2 < s < 4. Then we have for  $\delta > 0$ :

(1.5) 
$$\lambda(P \circ S_{(n)}, W, c_3 \delta) \leq c_4(\delta^{-s} K_{sn}(\delta) + \delta^{-2} L_n(\delta)),$$

where  $c_3 = c_3(s, d)$ ,  $c_4 = c_4(s, d)$  are positive constants depending on s and d only.

Since  $\delta^{-s} K_{sn}(\delta) + \delta^{-2} L_n(\delta) \leq \delta^{-s} K_{sn}$ ,  $\delta > 0$ , we easily obtain from (1.5) a multidimensional version of (1.3), if 2 < s < 4. Furthermore, Theorem 1 implies the following unimprovable estimate of the Prohorov distance of  $P \circ S_{(n)}$  and W:

**Corollary 2.** Let  $S_{(n)}$ :  $\Omega \to C_d[0,1]$  be as in Theorem 1. Then we have for 2 < s < 5:

(1.6) 
$$\rho(P \circ S_{(n)}, W) \leq c_5 K_{sn}^{1/(s+1)},$$

where  $c_5 = c_5(s, d) > 0$ .

Thus we have obtained as a byproduct of (1.4) an improvement of the main result of Borovkov and Sahanenko (1980) (cf. [3], Theorem 4).

Recall that the Prohorov-distance of two *p*-measures  $Q_i|B(C_d[0,1]), i=1,2$ , is defined by

$$\rho(Q_1, Q_2) = \inf\{\delta > 0: \lambda(Q_1, Q_2, \delta) \leq \delta\}.$$

Furthermore, it easily follows from Theorem 1:

**Corollary 3.** Let  $S_{(n)}: \Omega \to C_d[0,1]$  be as in Theorem 1, let 2 < s < 4. Set  $\varepsilon := \varepsilon(s):=(s+1)/(s-2)$ . Suppose that  $L_n(\delta^s) \leq \delta^3$ . Then:  $\rho(P \circ S_{(n)}, W) \leq c_6 \delta$ , where  $c_6 = c_6(s,d)$  is a positive constant depending on s and d only.

An immediate consequence of Corollary 3 is the above announced invariance principle for triangular arrays of rowwise independent random vectors fulfilling the Lindeberg-condition.

**Corollary 4.** Let  $\{\xi_{nk}: 1 \leq k \leq k_n\}$ ,  $n \in \mathbb{N}$ , be a triangular array of rowwise independent random vectors with zero means  $\operatorname{cov}(\xi_{nk}) = \sigma_{nk}^2 \Gamma$ ,  $1 \leq k \leq k_n$  and  $\sum_{k=1}^{k_n} \sigma_{nk}^2 = 1$   $(n \in \mathbb{N})$ . Let  $S_{(n)}: \Omega \to C_d[0, 1]$  be defined by

$$S_{(n)}(t) := \sum_{k=1}^{m} \xi_{nk} + \frac{t - t_{nm}}{t_{nm+1} - t_{nm}} \xi_{nm+1}, \quad t_{nm} \leq k \leq t_{nm+1},$$

 $0 \leq m < k_n, \text{ where } t_{nm} := \sum_{k=1}^m \sigma_{nk}^2, 0 \leq m \leq k_n, n \in \mathbb{N}.$ Assume that for all  $\delta > 0$  the following is true:

$$L_n(\delta) := \sum_{k=1}^{k_n} E[|\xi_{nk}|^2 \, \mathbb{1}\{|\xi_{nk}| > \delta\}] \to 0 \quad \text{as } n \to \infty.$$

Then  $P \circ S_{(n)} | B(C_d[0,1])$  converges weakly to  $W_r | B(C_d[0,1])$ .

Theorem 1 is now used to obtain strong invariance principles for partial sums of independent not necessarily identically distributed d-dimensional random vectors.

We denote in the sequel by  $\mathscr{H}$  the set of all continuous functions  $H: [0, \infty) \to [0, \infty)$  such that  $t^{-2} H(t)$  is non-decreasing and  $t^{-4+r} H(t)$  is non-increasing for some r > 0.

**Theorem 2.** Let  $\{\xi_k\}$  be a sequence of independent random vectors with zero means and  $\operatorname{cov}(\xi_k) = \sigma_k^2 \Gamma$ ,  $k \in \mathbb{N}$ . Assume that the following holds true for some  $s \in (2, 4)$ :

(1.7) 
$$\sum_{1}^{\infty} a_{k}^{-s} E[|\xi_{k}|^{s} 1\{|\xi_{k}| \leq a_{k}\}] < \infty$$

and

(1.8) 
$$\sum_{1}^{\infty} a_k^{-2} E[|\xi_k|^2 1\{|\xi_k| > a_k\}] < \infty \quad \text{where } 0 < a_k \uparrow \infty.$$

Then one can construct a p-space  $(\Omega_0, \mathscr{A}_0, P_0)$  and two sequences of independent random vectors  $\{X_k\}$ ,  $\{Y_k\}$  with  $P_0 \circ X_k = P \circ \xi_k$ ,  $P_0 \circ Y_k = N(0, \sigma_k^2 \Gamma)$ ,  $k \in \mathbb{N}$ , such that the partial sums  $S_n := \sum_{i=1}^{n} X_k$ ,  $T_n := \sum_{i=1}^{n} Y_k$  fulfill:

$$S_n - T_n = o(a_n) \qquad a.s.$$

Remark. It is easy to see that the conditions (1.7) and (1.8) are fulfilled if

(1.10) 
$$\sum_{1}^{\infty} H(a_k)^{-1} E[H(|\xi_k|)] < \infty \quad \text{for some } H \in \mathscr{H}.$$

From Theorem 2 we obtain

**Theorem 3.** Let  $\{\xi_k\}$  be a sequence of independent random vectors with zero means and  $\operatorname{cov}(\xi_k) = \sigma_k^2 \Gamma$ ,  $k \in \mathbb{N}$ . Let  $f: [0, \infty) \to (0, \infty)$  be a non-decreasing function such that  $\sum_n \frac{1}{nf(n)} < \infty$ . Assume that  $\alpha_n := \sum_{1}^n E[H(|\xi_k|)] \to \infty$ , where  $H \in \mathcal{H}$ . Then a construction is possible such that

(1.11) 
$$S_n - T_n = o(H^{-1}(\alpha_n f(\alpha_n))) \quad a.s.$$

Let  $\{\xi_k\}$  be a sequence of independent random vectors such that  $0 < \lim_k E[H(|\xi_k|)] \le \lim_k E[H(|\xi_k|)] < \infty$ . According to (1.11), a construction is possible such that

(1.12) 
$$S_n - T_n = o(H^{-1}(nf(n)))$$
 a.s., if  $\sum_n \frac{1}{nf(n)} < \infty$ .

Since  $f(n)/\log n \to \infty$  as  $n \to \infty$  by Kronecker's Lemma, the convergence rate is considerably worse than that in the strong invariance principle for partial sums of i.i.d. random vectors. Recall that it was shown in [8], Theorem 2 for the

i.i.d. case that if  $H \in \mathcal{H}$  such that  $t^{-2}(\operatorname{LogLog} t)^{-1}H(t)$  is non-decreasing, a construction with error term  $o(H^{-1}(n))$  is possible. If  $H \in \mathcal{H}$  such that  $t^{-2}(\operatorname{LogLog} t)^{-1}H(t)$  is non-increasing it is still possible to obtain a construction with error term  $o\left(H^{-1}(n)\sqrt{\frac{\operatorname{LogLog} n}{h(n)}}\right)$ , where  $h(t):=t^{-2}H(t), t \ge 0$  (cf. [8], Theorem 3).

The following Theorem 4, however, shows that under the above assumptions no better convergence rate than (1.12) can be reached in general.

**Theorem 4.** Let  $o < f_n \uparrow \infty$  be such that  $\sum_n \frac{1}{nf_n} = \infty$ . Let  $H \in \mathscr{H}$ . There exists a sequence of independent r.v.'s  $\xi_k: \Omega \to \mathbb{R}$ ,  $k \in \mathbb{N}$  with zero means,  $E[\xi_k^2] = 1$  and  $E[H(|\xi_k|)] \leq 1 + H(1)$ ,  $k \in \mathbb{N}$ , such that for all sequences of independent r.v.'s  $\{X_k\}$  with  $P_0 \circ X_k = P \circ \xi_k$ ,  $k \in \mathbb{N}$ , and for all sequences of i.i.d. r.v.'s  $\{Y_k\}$  with  $P_0 \circ Y_1 = N(0, 1)$ , where  $(\Omega_0, \mathscr{A}_0, P_0)$  is an arbitrary p-space, the following holds true:

(1.13) 
$$\overline{\lim_{n}} \frac{|S_n - T_n|}{H^{-1}(nf_n)} = \infty \quad a.s.$$

Our next result implies the above announced functional law of the iterated logarithm. To simplify our notations, we set  $\text{Log}_2 t := \text{Log}(\text{Log} t), t \ge 0$ .

**Theorem 5.** Let  $\{\xi_k\}$  be a sequence of independent random vectors with zero means and  $\operatorname{cov}(\xi_k) = \sigma_k^2 \Gamma$ ,  $k \in \mathbb{N}$ . Assume  $B_n := \sum_{k=1}^{n} \sigma_k^2 \to \infty$ , (1.14)  $\sum_{k=1}^{\infty} (B_k \operatorname{Log}_2 B_k)^{-s/2} E[|\xi_k|^s \operatorname{1}\{|\xi_k| \le \sqrt{B_k \operatorname{Log}_2 B_k}\}] < \infty$  for some s > 2

and

(1.15) 
$$\sum_{1}^{\infty} (B_k \operatorname{Log}_2 B_k)^{-1} E[|\xi_k|^2 1\{|\xi_k| > \sqrt{B_k \operatorname{Log}_2 B_k}\}] < \infty.$$

Then a construction is possible such that

(1.16) 
$$S_n - T_n = o(\sqrt{B_n \log_2 B_n}) \quad a.s$$

(1.14) and (1.15) are fulfilled if

(1.17) 
$$\sum_{1}^{\infty} (B_k \operatorname{Log}_2 B_k)^{-s/2} E[|\xi_k|^s] < \infty \quad \text{for some } s > 2.$$

Let  $\mathscr{F}$  be the set of all absolutely continuous functions  $f: [0,1] \to \mathbb{R}^d$  such that f(0) = 0 and  $\int_0^1 |f'(t)|^2 dt \leq 1$ . Denote by  $\mathscr{F}_{\Gamma}$  the functions  $\{t \to \Gamma^{1/2} f(t), f \in \mathscr{F}\}$ .

Furthermore, we set:

$$S(t) := \sum_{1}^{m} \xi_{k} + \frac{t - B_{m}}{B_{m+1} - B_{m}} \xi_{m+1}, \qquad B_{m} \leq t \leq B_{m+1},$$
  
$$0 \leq m < \infty, \text{ where } B_{0} = \sum_{1}^{0} \xi_{k} = 0.$$

**Corollary 5.** Let  $\{\xi_k\}$  be as in Theorem 5. Let  $U_{(n)}: \Omega \to C_d[0,1]$  be defined by  $U_{(n)}(t):=(2B_n \log_2 B_n)^{-1/2} S(B_n t), \ 0 \le t \le 1$ . Suppose that (1.14) and (1.15) hold true,  $B_n \to \infty$  and

(1.18) 
$$\overline{\lim_{n}} \frac{B_{n+1}}{B_{n} \exp(\sqrt{\log_2 B_{n}})} < \infty.$$

With probability one,  $\{U_{(n)}: n \in \mathbb{N}\}$  is relatively compact in  $C_d[0,1]$  and the set of its limit points coincides with  $\mathscr{F}_{\Gamma}$ .

Corollary 5 implies immediately the following compact law of the iterated logarithm:

**Corollary 6.** Under the assumptions of Corollary 5 we have with probability one:  $\left\{ (2B_n \log_2 B_n)^{-1/2} \sum_{1}^{n} \xi_k; n \in \mathbb{N} \right\} \text{ is relatively compact in } \mathbb{R}^d \text{ and the set of its} \\ \text{limit points coincides with } \{\Gamma^{1/2} x; |x| \leq 1\}.$ 

Thus we have obtained a strong extension of [15], Theorem 1.2, which was proved in [15] under the more restrictive assumptions " $\overline{\lim}(B_{n+1}/B_n) < \infty$ " and

(1.17), if  $2 < s \le 4$ . Furthermore, we have shown that [15] Theorem 1.2 remains valid for any s > 4 and that the condition " $E[X_n^3] = 0$  ( $n \in \mathbb{N}$ )" can be omitted.

As a last application of Theorem 1 we present a sufficient condition which guarantees that in the strong invariance principle the convergence rate  $o(\sqrt{B_n})$  can be reached.

**Theorem 6.** Let  $\{\xi_k\}$  be a sequence of independent random vectors with zero means and  $\operatorname{cov}(\xi_k) = \sigma_k^2 \Gamma$ ,  $k \in \mathbb{N}$ .

Assume that  $B_n \to \infty$  and

(1.19) 
$$\sum_{k=1}^{\infty} B_k^{-s/2} E[|\xi_k|^s] < \infty \quad \text{for some } s > 2$$

Then a construction is possible such that

$$S_n - T_n = o(\sqrt{B_n})$$
 a.s.

Using Kronecker's lemma, we obtain from (1.19):

$$\sum_{1}^{n} E[|\xi_k|^s] = o(B_n^{s/2}) \quad \text{for some } s > 2,$$

i.e. Lyapunov's condition – a well known sufficient condition for the invariance principle (in distribution).

The paper is now organized as follows: In Sect. 2 we prove Theorem 1 and its corollaries. Some more or less known lemmas needed here are formulated and proved in Sect. 6. In Sect. 3 we infer from Theorem 1 the above strong invariance principles for partial sums of independent random vectors (i.e. Theorems 2, 3, 5 and 6). Theorem 4 is proved in Sect. 4, whereas the proof of Corollaries 5 and 6 is given in Sect. 5.

# 2. Proof of Theorem 1 and its Corollaries

Main tool of the proof is the following Theorem 7, which is related to [8], Theorem 7.

**Theorem 7.** Let  $\xi_1, \ldots, \xi_n: \Omega \to \mathbb{R}^d$  be independent random vectors with zero means. Assume  $\lambda_n/\Lambda_n \ge 1/2$ , where  $\lambda_n(\Lambda_n)$  is the smallest (largest) eigenvalue of  $\Gamma_n: = \operatorname{cov}(\xi_1 + \ldots + \xi_n)$ . Let  $3 \le \overline{s} < 4$ ,  $s > \overline{s}$ . There exist positive constants  $c_7 = c_7(\overline{s}, s, d)$ ,  $c_8 = c_8(\overline{s}, s, d)$  such that, if  $|\xi_k| \le c_7 \sqrt{\lambda_n \operatorname{Log} 1/\rho_{sn}}$  a.s.,  $1 \le k \le n$ , where  $\rho_{sn}:=\lambda_n^{-s/2}\sum_{j=1}^n E[|\xi_k|^s]$ , the following holds true:

One can construct a p-space  $(\Omega_0, \mathscr{A}_0, P_0)$  and random vectors  $S_n, T_n: \Omega_0 \to \mathbb{R}^d$ such that  $P_0 \circ S_n = P \circ \sum_{i=1}^{n} \xi_k$ ,  $P_0 \circ T_n = N(0, \Gamma_n)$ ,  $E[|S_n - T_n|^s] \leq c_8 \Lambda_n^{s/2} \rho_{sn}^{1+1/s}$  and  $E[|S_n - T_n|^2] \leq c_8 \Lambda_n \rho_{sn}^{1+\eta}$  for some  $\eta = \eta(\bar{s}, \bar{s}) > 0$ .

*Proof* (cf. [8], Theorem 7). We denote by  $c_9, \ldots, c_{14}$  constants depending on  $\overline{s}, \overline{s}$  and d only. Since by Lemma 1(b)

(2.1) 
$$\rho_n := \rho_{3n} \leq c_9 \cdot \rho_{\bar{s}n}^{1/(\bar{s}-2)} \leq c_{10} \cdot \rho_{sn}^{1/(s-2)},$$

we have

(2.2) 
$$\log 1/\rho_n \ge c_{11} \cdot \log 1/\rho_{sn}.$$

Setting  $c_7 := \left(\frac{c_{11}}{8d}\right)^{1/2} \cdot (4-\overline{s})(s+1)^{-2}$ , we infer from (2.2):

(2.3) 
$$|\xi_k| \leq \frac{4-\overline{s}}{\sqrt{8d}(s+1)^2} \sqrt{\lambda_n \log 1/\rho_n} \quad \text{a.s., } 1 \leq k \leq n.$$

By [8], Theorem 6 (applied with  $\varepsilon := (4 - \overline{s})/s$ ) we have:

(2.4) 
$$\lambda \left( P \circ \Gamma_n^{-1/2} \sum_{1}^n \xi_k, N(0, I), c_{12} \rho_n^{1-\varepsilon} \right) \leq c_{13} \rho_n^{s(s+1)}.$$

Thus we can obtain from the Strassen-Dudley theorem (cf. [7], Theorem 2) a *p*-space  $(\Omega_0, \mathscr{A}_0, P_0)$  and random vectors  $S_n, T_n: \Omega_0 \to \mathbb{R}^d$  with  $P_0 \circ S_n = P \circ \sum_{1}^{n} \xi_k$ ,  $P_0 \circ T_n = N(0, \Gamma_n)$ , such that

(2.5) 
$$P_0\{|S_n - T_n| > c_{12} \Lambda_n^{1/2} \rho_n^{1-\varepsilon}\} \leq c_{13} \rho_n^{s(s+1)}.$$

Since

$$\begin{split} & E[|S_n - T_n|^s] \\ & \leq x^s + E[|S_n - T_n|^{s+1}]^{s/(s+1)} P_0\{|S_n - T_n| > x\}^{1/(s+1)} \\ & \leq x^s + 2^{s-1} (E[|S_n|^{s+1}]^{s/(s+1)} + E[|T_n|^{s+1}]^{s/(s+1)}) P_0\{|S_n - T_n| > x\}^{1/(s+1)} \end{split}$$

for x > 0, we easily obtain from an obvious modification of [8], Lemma 3:

(2.6) 
$$E[|S_n - T_n|^s] \leq c_{14} A_n^{s/2} \rho_n^{s-(4-\bar{s})}.$$

Using (2.1), we immediately obtain  $E[|S_n - T_n|^s] \leq c_8 \Lambda_n^{s/2} \rho_{sn}^{1+1/s}$ . (Notice that  $(s + \overline{s} - 4)/(s - 2) \geq (s + 1)/s$ , since  $\overline{s} \geq 3$ .)

Finally

$$\begin{split} E[|S_n - T_n|^2] &\leq E[|S_n - T_n|^s]^{2/s} \leq c_{14}^{2/s} \Lambda_n \cdot \rho_n^{2-2(4-\bar{s})/s} \\ &\leq c_8 \Lambda_n \rho_{\bar{s}n}^{1+\eta}, \end{split}$$

where  $\eta := \frac{(s-2)(4-\bar{s})}{s(\bar{s}-2)}$ .

We now proceed to the proof of Theorem 1. Since  $\delta^{-s} K_{sn}(\delta) \leq \delta^{-s'} K_{s'n}(\delta)$ ,  $s \geq s'$ , it suffices to prove Theorem 1 for  $s \geq 3$ .

To simplify our notations, we set for  $1 \leq m \leq n$ :

$$\alpha_{m} := \sum_{k=1}^{m} \left( E\left[ |\xi_{k}|^{s} \, \mathbb{1}\left\{ |\xi_{k}| \leq \delta \right\} \right] + \delta^{s-2} E\left[ |\xi_{k}|^{2} \, \mathbb{1}\left\{ |\xi_{k}| > \delta \right\} \right] \right).$$

Furthermore, we define:

$$\bar{\xi_k} := \xi_k \, 1 \, \{ |\xi_k| \leq \delta \}, \qquad \tilde{\xi_k} := \bar{\xi_k} - E[\bar{\xi_k}], \qquad 1 \leq k \leq n.$$

We denote by  $c_{15}, \ldots, c_{33}$  positive constants depending on  $\gamma$ , s and d only.

(i) The purpose of this part of the proof is to show that there exist positive constants  $\bar{s} = \bar{s}(\gamma, s) \in [3, 4)$ ,  $\varepsilon = \varepsilon(\gamma, s) \in (0, 1]$  and  $c_{15}$  such that

(2.7) 
$$\delta^{-\bar{s}} K_{\bar{s}n}(\delta) \leq c_{15} (\delta^{-s} K_{sn}(\delta))^{\varepsilon}.$$

Since  $\sum_{1}^{n} E[|\bar{\xi}_{k}|^{2}] \leq d$ , we infer from Lemma 1(a):

(2.8) 
$$K_{\bar{s}n}(\delta) \leq c_{15} K_{sn}(\delta)^{(\bar{s}-2)/(s-2)} \quad \text{if } \bar{s} \leq s.$$

We choose  $\varepsilon$  and  $\overline{s}$  in a way such that

$$(2.9) \qquad \overline{s} - s \varepsilon \ge 0$$

and

(2.10) 
$$(\overline{s} - s \varepsilon) \gamma = (\overline{s} - 2)/(s - 2) - \varepsilon.$$

Since  $\gamma = \frac{1}{2}(1-\hat{\gamma})/(s-2)$ , where  $\hat{\gamma} > 0$ , (2.10) can be rewritten:

(2.11) 
$$\overline{s} = \frac{4}{\hat{\gamma}+1} + \left(s - \frac{4}{\hat{\gamma}+1}\right)\varepsilon.$$

It is now easy to see that a possible choice is:

$$\bar{s} = \frac{4+3\hat{\gamma}}{1+\hat{\gamma}}, \quad \varepsilon = \frac{3\hat{\gamma}}{(s-4)+s\hat{\gamma}} \quad (s \ge 4)$$

and (trivially)  $\bar{s} = s$ ,  $\varepsilon = 1$ , if s < 4.

Since  $\delta \ge (K_{sn}(\delta))^{\gamma}$ , we easily obtain (2.7) from (2.8), (2.9) and (2.10). Let now  $(\varepsilon, \vec{s})$  be fixed. We set for  $1 \le m \le n$ :

$$\beta_{m} := \sum_{k=1}^{m} \left( E\left[ |\xi_{k}|^{\bar{s}} \, \mathbb{1}\left\{ |\xi_{k}| \leq \delta \right\} \right] + \delta^{\bar{s}-2} E\left[ |\xi_{k}|^{2} \, \mathbb{1}\left\{ |\xi_{k}| > \delta \right\} \right] \right).$$

Using (2.7), we obtain

$$\delta^{-\bar{s}} \beta_n \leq c_{15} (\delta^{-s} K_{sn}(\delta))^{\varepsilon} + \delta^{-2} L_n(\delta) \leq (c_{15} + 1) 2^{1-\varepsilon} (\delta^{-s} \alpha_n)^{\varepsilon},$$

if we w.l.o.g. assume

$$(2.12) \qquad \qquad \alpha_n \leq e^{-2} \,\delta^s.$$

(In particular:  $\delta^{-2} L_n(\delta) \leq e^{-2}$ .) Thus we have

(2.13) 
$$\delta^{-\bar{s}}\beta_n \leq c_{16}(\delta^{-s}\alpha_n)^{\varepsilon}.$$

(ii) Let  $\tilde{S}_{(n)}: \Omega \to C_d[0,1]$  be defined by

(2.14) 
$$\tilde{S}_{(n)}(t) := \sum_{1}^{m} \tilde{\xi}_{k} + \frac{t - t_{m}}{t_{m+1} - t_{m}} \tilde{\xi}_{m+1}, \quad t_{m} \leq t \leq t_{m+1}, \quad 0 \leq m < n.$$

An application of Lemma 2 yields

(2.15) 
$$\lambda(P \circ S_{(n)}, P \circ \tilde{S}_{(n)}, 2\delta) \leq \delta^{-s} \alpha_n.$$

(iii) Let  $c_{17} > 0$  be a sufficiently large chosen constant, which is determined by (2.25).

We now define a finite sequence of non-negative integers  $0 = m_0 < ... < m_r$ = n by the following recursion:  $m_0 := 0$ . Let  $m_{j-1}$  be defined for a  $j \ge 1$ . Then we set

(2.16) 
$$m_j := \min\left\{k > m_{j-1}: c_{17} \,\delta^2 \left(\log \frac{\delta^s}{\alpha_k - \alpha_{m_{j-1}}}\right)^{-1} \leq t_k - t_{m_{j-1}}\right\} \wedge n,$$

where  $\min \phi := n + 1$ .

Put  $r := \min\{j \in \mathbb{N} : m_j = n\}$ .

Since

$$d\sigma_k^2 = E[|\xi_k|^2 1\{|\xi_k| \le \delta\}] + E[|\xi_k|^2 1\{|\xi_k| > \delta\}]$$
  
$$\le (\alpha_k - \alpha_{k-1})^{2/s} + \delta^{2-s}(\alpha_k - \alpha_{k-1}),$$

we obtain from (2.12):

(2.17) 
$$d\sigma_k^2 \leq 2(\alpha_k - \alpha_{k-1})^{2/s} \leq s \,\delta^2 \left( \log \frac{\delta^s}{\alpha_k - \alpha_{k-1}} \right)^{-1}, \quad 1 \leq k \leq n.$$

From the definition of the  $m_j$ 's and (2.17) it follows immediately

(2.18) 
$$t_{m_{j}} - t_{m_{j-1}} \leq c_{17} \,\delta^{2} \left( \log \frac{\delta^{s}}{\alpha_{m_{j}} - \alpha_{m_{j-1}}} \right)^{-1} + \sigma_{m_{j}}^{2}$$
$$\leq (c_{17} + s) \,\delta^{2} \left( \log \frac{\delta^{s}}{\alpha_{m_{j}} - \alpha_{m_{j-1}}} \right)^{-1}, \quad 1 \leq j \leq r$$

Furthermore, we infer from the definition of the  $m_j$ 's

(2.19) 
$$c_{17} \delta^2 \left( \log \frac{\delta^s}{\alpha_{m_j} - \alpha_{m_{j-1}}} \right)^{-1} \leq t_{m_j} - t_{m_{j-1}}, \quad 1 \leq j \leq r-1.$$

(iv) Let  $\lambda_j(\Lambda_j)$  be the smallest (largest) eigenvalue of

$$\begin{split} &\Gamma_{j} := \operatorname{cov}(\tilde{\xi}_{m_{j-1}+1} + \ldots + \tilde{\xi}_{m_{j}}), \quad \rho_{\bar{s}j} := \lambda_{j}^{-\bar{s}/2} \sum_{m_{j-1}+1}^{m_{j}} E[|\tilde{\xi}_{k}|^{\bar{s}}], \\ &\rho_{sj} := \lambda_{j}^{-s/2} \sum_{m_{j-1}+1}^{m_{j}} E[|\tilde{\xi}_{k}|^{\bar{s}}], \quad 1 \leq j \leq r. \end{split}$$

Since

$$\lambda_{j} \ge t_{m_{j}} - t_{m_{j-1}} - 2 \sum_{m_{j-1}+1}^{m_{j}} E[|\xi_{k}|^{2} 1\{|\xi_{k}| > \delta\}],$$

we obtain

$$(2.20) \quad t_{m_j} - t_{m_{j-1}} - \lambda_j \leq 2\delta^{2-s} (\alpha_{m_j} - \alpha_{m_{j-1}}) \leq 2\delta^2 \left( \log \frac{\delta^s}{\alpha_{m_j} - \alpha_{m_{j-1}}} \right)^{-1}, \quad 1 \leq j \leq r-1.$$

Thus we have by (2.19)

(2.21) 
$$\lambda_j \ge (c_{17} - 2) \, \delta^2 \left( \log \frac{\delta^s}{\alpha_{m_j} - \alpha_{m_{j-1}}} \right)^{-1}, \quad 1 \le j \le r - 1.$$

Since  $\Lambda_j \leq t_{m_j} - t_{m_{j-1}}$ , we infer from (2.18):

(2.22) 
$$\Lambda_j \leq (c_{17} + s) \,\delta^2 \left( \log \frac{\delta^s}{\alpha_{m_j} - \alpha_{m_{j-1}}} \right)^{-1}, \quad 1 \leq j \leq r.$$

Using the Hölder-inequality, we easily obtain:

$$E\left[|\tilde{\xi}_k|^s\right] \leq 2^s E\left[|\bar{\xi}_k|^s\right], \quad 1 \leq k \leq n.$$

Together with (2.21) this implies:

(2.23) 
$$\rho_{sj} \leq \frac{2^s}{(c_{17}-2)^{s/2}} \frac{\alpha_{m_j} - \alpha_{m_{j-1}}}{\delta^s} \left( \log \frac{\delta^s}{\alpha_{m_j} - \alpha_{m_{j-1}}} \right)^{s/2}, \quad 1 \leq j \leq r-1.$$

From (2.23) it easily follows if we choose  $c_{17}$  in a way such that  $c_{17} \ge c_{18} := 4s + 2$ 

(2.24) 
$$\log 1/\rho_{sj} \ge \frac{1}{2} \log \frac{\delta^s}{\alpha_{m_j} - \alpha_{m_{j-1}}}, \quad 1 \le j \le r-1.$$

We put  $c_{19} := c_7(\bar{s}, s, d)$ , where  $\bar{s} = \bar{s}(\gamma, s)$  is defined by (2.7). Recalling relations (2.21) and (2.24), it is straightforward to check that  $2\delta \le c_{19}\sqrt{\lambda_j \log 1/\rho_{sj}}$ ,  $1 \le j \le r-1$ , when setting

(2.25) 
$$c_{17} := \max(c_{18}, 10 c_{19}^{-2}).$$

Hence:

(2.26) 
$$|\xi_k| \leq c_{19} \sqrt{\lambda_j} \log 1/\rho_{sj}, \quad m_{j-1} < k \leq m_j, \quad 1 \leq j \leq r-1.$$

(v) We apply Theorem 7 and obtain a *p*-space  $(\Omega_1, \mathcal{A}_1, P_1)$  with independent random vectors  $(\tilde{U}_j, \tilde{V}_j), 1 \leq j \leq r-1$  such that

(2.27) 
$$P_1 \circ \tilde{U}_j = P \circ \sum_{m_j - 1 + 1}^{m_j} \tilde{\xi}_k, \quad P_1 \circ \tilde{V}_j = N(0, \Gamma_j),$$

(2.28) 
$$E[|\tilde{U}_j - \tilde{V}_j|^s] \leq c_{20}(\alpha_{m_j} - \alpha_{m_{j-1}})$$

and

(2.29) 
$$E[|\tilde{U}_j - \tilde{V}_j|^2] \leq c_{21} \cdot \Lambda_j \cdot \rho_{\tilde{s}j}^{1+\eta}, \quad 1 \leq j \leq r-1.$$

(Notice that  $\lambda_j \ge \frac{1}{2}\Lambda_j$ ,  $1 = j \le r - 1$ , since  $c_{17} \ge c_{18}$ . Use (2.21) and (2.22).) A similar argument as in (2.23) shows

$$\rho_{\bar{s}j} \leq \frac{2^{\bar{s}}}{(c_{17}-2)^{\bar{s}/2}} \cdot \frac{\beta_{m_j} - \beta_{m_{j-1}}}{\delta^{\bar{s}}} \left( \log \frac{\delta^s}{\alpha_{m_j} - \alpha_{m_{j-1}}} \right)^{\bar{s}/2}, \quad 1 \leq j \leq r-1.$$

Since  $\log 1/\rho_{\bar{s}j} \ge c_{22} \log 1/\rho_{sj}$  by Lemma 1, we infer from (2.22) and (2.24):

$$\begin{split} \Lambda_{j} \cdot \rho_{\bar{s}j}^{1+\eta} &\leq c_{23} \cdot \frac{\beta_{m_{j}} - \beta_{m_{j-1}}}{\delta^{\bar{s}-2}} \rho_{\bar{s}j}^{\eta} (\text{Log } 1/\rho_{\bar{s}j})^{\bar{s}/2} \\ &\leq c_{24} \cdot \frac{\beta_{m_{j}} - \beta_{m_{j-1}}}{\delta^{\bar{s}-2}}, \quad 1 \leq j \leq r-1. \end{split}$$

(Note that  $\rho_{\bar{s}j}$  remains bounded,  $1 \leq j \leq r-1$ . Use (2.12) and (2.13).) Therefore

(2.30) 
$$E[|\tilde{U}_j - \tilde{V}_j|^2] \leq c_{25} \cdot \delta^{2-\bar{s}} (\beta_{m_j} - \beta_{m_{j-1}}), \quad 1 \leq j \leq r-1.$$

We set  $V_j := (t_{m_j} - t_{m_{j-1}})^{1/2} \Gamma_j^{-1/2} \tilde{V}_j, \ 1 \le j \le r-1.$ Then  $P_1 \circ V_j = N(0, (t_{m_j} - t_{m_{j-1}}) I)$  and

(2.31) 
$$P_1 \circ (V_j - \tilde{V}_j) = N(0, ((t_{m_j} - t_{m_{j-1}})^{1/2} I - \Gamma_j^{1/2})^2), \quad 1 \leq j \leq r-1.$$

Denoting the largest eigenvalue of  $((t_{m_j} - t_{m_{j-1}})^{1/2} I - \Gamma_j^{1/2})^2$  by  $\mu_j$ , we obtain from the obvious inequality  $\mu_j \leq (t_{m_j} - t_{m_{j-1}} - \lambda_j)^2/(t_{m_j} - t_{m_{j-1}})$ ,  $1 \leq j \leq r-1$ , (2.19) and (2.20):

$$\mu_{j} \leq \frac{4}{c_{17}} \frac{(\alpha_{m_{i}} - \alpha_{m_{j-1}})^{2}}{\delta^{2s-2}} \log \frac{\delta^{s}}{\alpha_{m_{j}} - \alpha_{m_{j-1}}} \leq c_{26} \frac{\alpha_{m_{j}} - \alpha_{m_{j-1}}}{\delta^{s-2}}, \quad 1 \leq j \leq r-1.$$

Since  $E[|V_j - \tilde{V}_j|^2] \leq d\mu_j$ ,  $\delta^{2-s}(\alpha_{m_j} - \alpha_{m_{j-1}}) \leq \delta^{2-\bar{s}}(\beta_{m_j} - \beta_{m_{j-1}})$ ,  $1 \leq j \leq r-1$ , we infer from the above inequality and (2.30)

(2.32) 
$$E\left[|\tilde{U}_j - V_j|^2\right] \leq c_{27} \,\delta^{2-\bar{s}}(\beta_{m_j} - \beta_{m_{j-1}}), \quad 1 \leq j \leq r-1.$$

Since  $E[|V_j - \tilde{V}_j|^s] \leq c_{28} \cdot \mu_j^{s/2}$  by (2.31),

$$\delta^{2-s}(\alpha_{m_j}-\alpha_{m_{j-1}}) \leq e^{-2(s-2)/s}(\alpha_{m_j}-\alpha_{m_{j-1}})^{2/s}$$

by (2.12), we finally infer from (2.29) and the above estimate of  $\mu_j$ 

(2.33) 
$$E[|\tilde{U}_j - V_j|^s] \leq c_{29}(\alpha_{m_j} - \alpha_{m_{j-1}}), \quad 1 \leq j \leq r-1.$$

Using Lemma A.1., [1] we obtain a p-space  $(\Omega_2, \mathscr{A}_2, P_2)$  and two sequences of independent random vectors  $\tilde{X}_k$ ,  $1 \leq k \leq n$  and  $Y_k$ ,  $1 \leq k \leq n$  such that

$$(2.34) P_2 \circ \tilde{X}_k = P \circ \tilde{\xi}_k, P_2 \circ Y_k = N(0, \sigma_k^2 I), 1 \leq k \leq n$$

and

(2.35) 
$$P_2 \circ \left( \sum_{m_{j-1}+1}^{m_j} \tilde{X}_k, \sum_{m_{j-1}+1}^{m_j} Y_k \right) = P_1 \circ (\tilde{U}_j, V_j), \quad 1 \leq j \leq r-1.$$

We set  $\tilde{S}_m := \sum_{k=1}^{m} \tilde{X}_k, \ T_m := \sum_{k=1}^{m} Y_k, \ 1 \le m \le n.$ 

(vi) From Lemma 3, (2.35), (2.32) and (2.33) we obtain

$$P_{2}\left\{\max_{1\leq j\leq r-1}|\tilde{S}_{m_{j}}-T_{m_{j}}|\geq\delta\right\}=P_{1}\left\{\max_{1\leq j\leq r-1}\left|\sum_{k=1}^{j}(\tilde{U}_{k}-V_{k})\right|\geq\delta\right\}$$
$$\leq \bar{c}_{1}\left(\exp\left(-\frac{\bar{c}_{2}}{c_{27}}\frac{\delta^{\bar{s}}}{\beta_{n}}\right)+c_{29}\,\delta^{-s}\,\alpha_{n}\right).$$

Therefore by (2.13):

(2.36) 
$$P_{2}\{\max_{1 \leq j \leq r-1} |\tilde{S}_{m_{j}} - T_{m_{j}}| \geq \delta\} \leq c_{30} \, \delta^{-s} \, \alpha_{n}$$

(vii) If the constant  $c_{31}$  is chosen large enough, we infer from Lemma 3, and (2.22) for  $1 \le j \le r$ 

$$P_{2}\{\max_{m_{j-1}\leq k\leq m_{j}}|\tilde{S}_{k}-\tilde{S}_{m_{j-1}}|\geq c_{31}\cdot\delta\}\leq c_{32}\,\delta^{-s}(\alpha_{m_{j}}-\alpha_{m_{j-1}}).$$

(Notice that  $E[|\tilde{\xi}_k|^s] \leq 2^s E[|\bar{\xi}_k|^s], 1 \leq k \leq n$ .) Hence

(2.37) 
$$P_{2} \{ \max_{1 \leq j \leq r} \max_{m_{j-1} \leq k \leq m_{j}} |\tilde{S}_{k} - \tilde{S}_{m_{j-1}}| \geq c_{31} \delta \} \leq c_{32} \delta^{-s} \alpha_{n}$$

Using (2.18) instead of (2.22), we obtain similarly:

$$(2.38) P_2 \{ \max_{1 \le j \le r} \max_{m_{j-1} \le k \le m_j} |T_k - T_{m_{j-1}}| \ge c_{31} \, \delta \} \le c_{32} \, \delta^{-s} \, \alpha_n.$$

(viii) (2.36), (2.37) and (2.38) imply immediately:

(2.39) 
$$P_{2} \{ \max_{1 \le k \le n} |\tilde{S}_{k} - T_{k}| \ge (2c_{31} + 1)\delta \} \le (c_{30} + 2c_{32})\delta^{-s}\alpha_{n}$$

Therefore

(2.39') 
$$\lambda(P \circ \tilde{S}_{(n)}, P_2 \circ T_{(n)}, (2c_{31}+1)\delta) \leq (c_{30}+2c_{32})\delta^{-s}\alpha_n,$$

where  $T_{(n)}: \Omega_2 \to C_d[0,1]$  is defined by

$$T_{(n)}(t) := T_m + \frac{t - t_m}{t_{m+1} - t_m} Y_{m+1} \quad \text{for } t_m \le t \le t_{m+1},$$

 $0 \leq m < n$ .

To finish the proof is suffices to show

(2.40) 
$$\lambda(P_2 \circ T_n, W, \delta) \leq c_{33} \, \delta^{-s} \, \alpha_n.$$

Let  $w: \Omega \to C_d[0,1]$  be a Brownian motion, i.e.  $P \circ w = W$ . Let  $\hat{w}: \Omega \to C_d[0,1]$  be defined by

(2.41) 
$$\hat{w}(t) := w(t_m) + \frac{t - t_m}{t_{m+1} - t_m} (w(t_{m+1}) - w(t_m)) \quad \text{for } t_m \leq t \leq t_{m+1},$$

 $0 \leq m < n$ .

Since  $w = (w_1, ..., w_d)$ , where  $w_i: \Omega \to C[0, 1]$  is a 1-dimensional Brownian motion (i = 1, ..., d), we obtain from the definition of  $\hat{w}$ , ([6], Theorem 1.5.1) and (2.17)

$$P\{\sup_{\substack{0 \le t \le 1 \\ 1 \le k \le n}} |\hat{w}(t) - w(t)| \ge \delta\}$$
  
$$\leq P\{\max_{\substack{1 \le k \le n \\ t_{k-1} \le t \le t_{k}}} \sup_{|w(t) - w(t_{k-1})| \ge \delta/2}\}$$
  
$$\leq 2d \sum_{k=1}^{n} P\{|w_{1}(\sigma_{k}^{2})| \ge \delta/(2\sqrt{d})\} \le 4d \sum_{k=1}^{n} \exp\left(-\frac{\delta^{2}}{8\sigma_{k}^{2}d}\right) \le c_{33} \delta^{-s} \alpha_{n}.$$

This proves (2.40), since  $P \circ \hat{w} = P_2 \circ T_{(n)}$ .

Corollaries 1 and 2 are immediate consequences of Theorem 1 applied for  $\gamma = 1/s$  and  $\gamma = 1/(s+1)$ , respectively. (Notice that Corollary 1 is trivial for  $\delta < (K_{sn}(\delta) + \delta^{s-2} L_n(\delta))^{1/s}$ .) Corollary 3 follows from Corollary 1, using the obvious inequalities

$$K_{sn}(\delta) \leq d\delta^{(s-2)\varepsilon} + \delta^{s-2} L_n(\delta^{\varepsilon})$$

and

$$L_n(\delta) \leq L_n(\delta^{\epsilon}), \quad \delta \leq 1.$$

It suffices to prove Corollary 4 for  $\Gamma = I$ . Since the Prohorov distance metrizes the weak convergence, Corollary 4 is proved if we show that  $\rho(P \circ S_{(n)}, W) \to 0$  as  $n \to \infty$ .

Since  $\delta^{-1}L_n(\delta) \to 0$  for all  $\delta > 0$ , there exists a sequence  $\delta_n \downarrow 0$  such that  $\delta_n^{-3}L_n(\delta_n^3) \to 0$ . (Use Lemma 6, [12], Chap. X.) Hence,  $L_n(\delta_n^3) \leq \delta_n^3$  for sufficiently large *n*. From Corollary 3, applied with s = 7/2, we obtain:  $\rho(P \circ S_{(n)}, W) \leq \delta_n \downarrow 0$ .

### 3. Proof of Theorems 2, 3, 5 and 6

Using the well known Strassen-Dudley theorem (cf. [7], Theorem 2), we obtain from the proof of Theorem 1 the following proposition which improves ([11], Theorem 2).

**Proposition 1.** Let  $\xi_1, \ldots, \xi_n: \Omega \to \mathbb{R}^d$  be independent random vectors, defined on a *p*-space  $(\Omega, \mathscr{A}, P)$ , with zero means and  $\operatorname{cov}(\xi_k) = \sigma_k^2 I$ ,  $1 \leq k \leq n$ . Let s > 2,  $0 < \gamma < 1/(2s-4)$ . Let  $\delta/\sqrt{B_n} \geq [B_n^{-s/2}(K_{sn}(\delta) + \delta^{s-2}L_n(\delta))]^{\gamma}$ , where  $B_n:=\sum_{1}^n \sigma_k^2$ ,  $n \in \mathbb{N}$ . One can construct a *p*-space  $(\Omega_0, \mathscr{A}_0, P_0)$  and two finite sequences of independent random vectors  $X_1, \ldots, X_n$  and  $Y_1, \ldots, Y_n$  with  $P_0 \circ X_k = P \circ \xi_k$ ,  $P_0 \circ Y_k$ 

 $= N(0, \operatorname{cov}(\xi_k)), \quad 1 \leq k \leq n \quad \text{such that the partial sums } S_m := \sum_{1}^{m} X_k, \quad T_m := \sum_{1}^{m} Y_k, \quad 1 \leq m \leq n, \quad \text{fulfill:}$ 

$$P_0 \{ \max_{1 \le k \le n} |S_k - T_k| > c_{34} \, \delta \} \le c_{35} (\delta^{-s} K_{sn}(\delta) + \delta^{-2} L_n(\delta)),$$

where  $c_{34}$ ,  $c_{35} > 0$  are constants depending on  $\gamma$ , s and d only.

*Proof.* Using relations (2.15) and (2.39') of Theorem 1 (applied with  $\xi_1/\sqrt{B_n}, \dots, \xi_n/\sqrt{B_n}$  and  $\delta/\sqrt{B_n}$ , we obtain

(3.1) 
$$\lambda \left( P \circ S_{(n)}, P_2 \circ T_{(n)}, c_{34} \frac{\delta}{\sqrt{B_n}} \right) \leq c_{35} (\delta^{-s} K_{sn}(\delta) + \delta^{-2} L_n(\delta)),$$

where  $S_{(n)}: \Omega \to C_d[0,1]$  is the partial sum process of the random vectors  $\xi_k/\sqrt{B_n}$ ,  $1 \le k \le n$  and  $T_{(n)}: \Omega_2 \to C_d[0,1]$  is the partial sum process of a finite sequence of independent random vectors  $\eta_1, \ldots, \eta_n: \Omega_2 \to \mathbb{R}^d$  such that  $P_2 \circ \eta_k = N\left(0, \frac{\sigma_k^2}{B_n}I\right), 1 \le k \le n$ . The assertion follows by a straightforward application of [7], Theorem 2.

*Remarks* (a) By obvious modifications of the proof of Theorem 1 one can easily show that the above proposition remains valid, if the covariance matrices  $\Sigma_k := \operatorname{cov}(\xi_k), 1 \le k \le n$ , fulfill the condition

(3.2) 
$$\sigma_k^2 \mu_1 \leq \langle t, \Sigma_k t \rangle \leq \sigma_k^2 \mu_2 \quad \text{for } |t| = 1 \quad (1 \leq k \leq n)$$

where  $\mu_1, \mu_2$  are positive constants. We only have to replace the constants  $c_{34}$ ,  $c_{35}$  by constants depending on  $\mu_1, \mu_2, \gamma, s$  and d.

(b) Let 2 < s < 4. Then it is easy to see that for any given  $\delta > 0$  a construction is possible such that

(3.3) 
$$P_0 \{ \max_{1 \le k \le n} |S_k - T_k| > c_{36} \, \delta \} \le c_{37} (\delta^{-s} K_{sn}(\delta) + \delta^{-2} L_n(\delta)),$$

where  $c_{36}$ ,  $c_{37}$  are positive constants depending on s and d only.

(c) Since  $K_{sn}(\delta) + \delta^{s-2} L_n(\delta) \leq d\delta^{s-2} B_n$ , the condition

$$\delta \ge \sqrt{B_n} [B_n^{-s/2}(K_{sn}(\delta) + \delta^{s-2} L_n(\delta))^{\gamma} \text{ for some } \gamma < 1/(2s-4)"$$

is always fulfilled, if  $\delta \ge d^{1/(s-2)} \sqrt{B_n}$ .

Now we start to prove Theorem 2. W.l.o.g. we assume that  $\Gamma = I$ . Let  $0 < \tilde{a}_k \uparrow \infty$  such that

(3.4) 
$$\tilde{a}_k/a_k \to 0$$
 as  $k \to \infty$ 

and

(3.5) 
$$\sum_{k=1}^{\infty} \tilde{a}_{k}^{-s} (E[|\xi_{k}|^{s} 1\{|\xi_{k}| \leq \tilde{a}_{k}\}] + \tilde{a}_{k}^{s-2} E[|\xi_{k}|^{2} 1\{|\xi_{k}| > \tilde{a}_{k}\}]) < \infty.$$

(The existence of such a sequence follows easily from (1.7) and (1.8), since  $E[|\xi_k|^s 1\{|\xi_k| \le \delta\}] + \delta^{s-2} E[|\xi_k|^2 1\{|\xi_k| > \delta\}], \ \delta > 0$ , is non-decreasing.)

We define  $\{m_n : n \in \mathbb{N}\}$  by the following recursion:

$$m_0:=1, \quad m_n:=\min\{k: \tilde{a}_k \ge 2\,\tilde{a}_{m_{n-1}}\}, \quad n \ge 1.$$

By definition we have

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$$(3.6) \qquad \qquad \tilde{a}_{m_n-1} \leq 2 \tilde{a}_{m_n-1}, \qquad n \geq 1.$$

Applying the above remark (b) to  $\{\xi_k: m_{n-1} \leq k < m_n\}$   $(n \in \mathbb{N})$  we obtain a *p*-space  $(\Omega_0, \mathscr{A}_0, P_0)$  and two sequences of independent random vectors  $\{X_n\}, \{Y_n\}$  such that the following holds true:

$$(3.7) \quad P_{0} \{ \max_{m_{n-1} \leq k < m_{n}} |(S_{k} - S_{m_{n-1}-1}) - (T_{k} - T_{m_{n-1}-1})|| > c_{36} \tilde{a}_{m_{n-1}} \} \\ \leq c_{37} \tilde{a}_{m_{n-1}}^{-s} \sum_{m_{n-1}}^{m_{n-1}} (E[|\xi_{k}|^{s} 1\{|\xi_{k}| \leq \tilde{a}_{m_{n-1}}\}] + \tilde{a}_{m_{n-1}}^{s-2} E[|\xi_{k}|^{2} 1\{|\xi_{k}| > \tilde{a}_{m_{n-1}}\}]) \\ \leq 2^{s} c_{37} \sum_{m_{n-1}}^{m_{n-1}} \tilde{a}_{k}^{-s} (E[|\xi_{k}|^{s} 1\{|\xi_{k}| \leq \tilde{a}_{k}\}] + \tilde{a}_{k}^{s-2} E[|\xi_{k}|^{2} 1\{|\xi_{k}| \geq \tilde{a}_{k}\}]),$$

where  $S_0 := T_0 := 0$ .

Thus we obtain from the Borel-Cantelli lemma, (3.5) and the definition of the  $m_n$ 's that almost surely

$$\begin{split} |S_k - T_k| &\leq K(\omega) + \sum_{j=1}^n c_{36} \, \tilde{a}_{m_{j-1}} \leq K(\omega) + 2c_{36} \, \tilde{a}_{m_{n-1}} \\ &\leq K(\omega) + 2c_{36} \, \tilde{a}_k \quad \text{if } m_{n-1} \leq k < m_n. \end{split}$$

This proves the assertion (recall (3.4)).

Theorem 3 is an immediate consequence of the remark after Theorem 2 and Lemma 15, ([12], Chap. IX).

Theorems 5 and 6 are special cases of Theorem 2, if 2 < s < 4. Thus we have to prove these results only for  $s \ge 4$ . We will show that the above proof of Theorem 2 works in these cases, too. But we have to be a little more careful, since we can apply our proposition only for sufficiently large  $\delta$ , if  $s \ge 4$ .

The main problem is (3.7). Let  $a_n = \sqrt{B_n \log_2 B_n}$  (resp.  $a_n = \sqrt{B_n}$ ). It suffices to show that there exists a sequence  $\tilde{a}_n \uparrow \infty$  such that (3.4) and (3.5) as well as

(3.8) 
$$\tilde{a}_n \ge \sqrt{B_n} (B_n^{-s/2} (K_{sn}(\tilde{a}_n) + \tilde{a}_n^{s-2} L_n(\tilde{a}_n)))^{\gamma}$$

for some  $\gamma < 1/(2s-4)$  hold true.

It is easy to see that such a sequence can be found if

(3.9) 
$$[B_n^{-s/2}(K_{sn}(a_n) + a_n^{s-2}L_n(a_n))]^{\gamma} = o(a_n/\sqrt{B_n})$$

for some  $\gamma < 1/(2s-4)$ .

Using the same argument as in the remark (c) after the proposition, we immediately obtain (3.9) for  $a_n = \sqrt{B_n \log_2 B_n}$ , hence Theorem 5.

Since 
$$B_n^{-s/2}(K_{sn}(a_n) + a_n^{s-2}L_n(a_n)) \leq B_n^{-s/2} \sum_{1}^{n} E[|\xi_k|^s]$$
, we have by (1.19):  
 $B_n^{-s/2}(K_{sn}(a_n) + a_n^{s-2}L_n(a_n)) = o(n^0),$ 

therefore (3.9) for  $a_n = \sqrt{B_n}$ , hence Theorem 6.

## 4. Proof of Theorem 4

To simplify our notations, we set:  $G(x) := H^{-1}(x)$ , x > 0. Let  $1 + 2/f_1 < b_k \uparrow \infty$  such that

(4.1) 
$$\sum_{k=1}^{\infty} \frac{1}{k b_k f_k} = \infty$$

and

(4.2) 
$$0 \leq \frac{b_k^{1/2} G(kf_k)^2 - 1}{k b_k f_k} \leq \frac{1}{2}, \quad k \in \mathbb{N}.$$

Let  $\{\xi_k\}$  be a sequence of independent random variables with symmetric distributions such that

(4.3)  

$$P\{|\xi_{k}| = b_{k}^{1/4} G(kf_{k})\} = \frac{1}{k b_{k} f_{k}},$$

$$P\{|\xi_{k}| = 1\} = 1 - \frac{G(kf_{k})^{2}}{k b_{k}^{1/2} f_{k}} \text{ and }$$

$$P\{\xi_{k} = 0\} = \frac{b_{k}^{1/2} G(kf_{k})^{2} - 1}{k b_{k} f_{k}}, \quad k \in \mathbb{N}.$$

From (4.3) it follows immediately that  $E[\xi_k] = 0$ ,  $E[\xi_k^2] = 1$ ,  $k \in \mathbb{N}$ . Furthermore, we have

$$E[H(|\xi_k|)] \leq \frac{H(b_k^{1/4} G(kf_k))}{k b_k f_k} + H(1) \leq 1 + H(1), \quad k \in \mathbb{N},$$

since  $t^{-4}H(t)$  is non-increasing and H(0) = 0.

Using the obvious inequality:  $|X_n| \leq |S_n - T_n| + |S_{n-1} - T_{n-1}| + |Y_n|$ , we obtain:

(4.4) 
$$P_0\left\{\overline{\lim_n |S_n - T_n|}_{G(nf_n)} < \infty\right\} \leq P_0\left\{\overline{\lim_n |X_n|}_{G(nf_n)} < \infty\right\}$$

(Notice that  $P_0\left\{\overline{\lim_{n}} \frac{|Y_n|}{G(nf_n)} < \infty\right\} = 1$ , since  $\{Y_n\}$  is a sequence of independent N(0, 1)-distributed r.v.'s.)

From the Borel-Cantelli lemma, (4.1) and (4.3) we infer

$$P_0\left\{\overline{\lim_{n}} \frac{|X_n|}{G(nf_n)} < \infty\right\} = 0, \quad \text{hence the assertion by (4.4).}$$

# 5. Proof of Corollaries 5 and 6

We use the following proposition which is related to Strassen's functional law of the iterated logarithm for the Brownian motion (cf. [14] Theorem 1).

**Proposition 2.** Let  $\{\zeta(t): t \ge 0\}$  be the Brownian motion in  $\mathbb{R}^d$ . Let  $B_n \uparrow \infty$  such that

(5.1) 
$$\overline{\lim_{n}} \left( \frac{B_{n+1}}{B_n \exp(\sqrt{\log_2 B_n})} \right) < \infty.$$

Set  $\zeta_n(t) := (2B_n \log_2 B_n)^{-1/2} \zeta(B_n t), \ 0 \le t \le 1.$ 

With probability one,  $\{\zeta_n : n \in \mathbb{N}\}$  is relatively compact in  $C_d[0,1]$  and the set of its limit points coincides with  $\mathcal{F}$ .

*Proof* (cf. [14], Theorem 1). Since the first part of the assertion follows easily from [14], Corollary 1, we have only to show the second part.

We use similar arguments as in [14], Theorem 1 (cf. p. 214/215). It suffices to show that for given  $x \in \mathscr{F}$  and  $\varepsilon > 0$  the following holds true:

(5.2) 
$$P(\overline{\lim}_{n} \{ \|\zeta_{n} - x\| < \varepsilon \}) = 1.$$

Let *m* be a positive integer, let  $\delta > 0$ .

We denote by  $A_n$  the event

$$\left\{ \left| \zeta_n^{\kappa} \left( \frac{i}{m} \right) - \zeta_n^{\kappa} \left( \frac{i-1}{m} \right) - \left( x^{\kappa} \left( \frac{i}{m} \right) - x^{\kappa} \left( \frac{i-1}{m} \right) \right) \right| < \delta \quad \text{for } 2 \leq i \leq m, \quad 1 \leq \kappa \leq d \right\},$$

where  $\zeta_n^{\kappa}(t)(x^{\kappa}(t))$  denotes the  $\kappa$ -th coordinate of  $\zeta_n(t)(x(t)), 1 \leq \kappa \leq d$  ( $t \in [0, 1]$ ).

Using the same arguments as in [14], we obtain for sufficiently large n

(5.3) 
$$P(A_n) \ge \frac{c(\delta)}{\log B_n \sqrt{m \log \log B_n}},$$

where  $c(\delta) > 0$  depends only on  $\delta$ .

We define the subsequence  $\{n_i: j \ge 1\} \subseteq \mathbb{N}$  by the following recursion:

$$n_1 := 1, \quad n_j := \min\{k > n_{j-1} : B_k \ge m \cdot B_{n_{j-1}}\}, \quad j \ge 2.$$

By definition we have:

$$(5.4) B_{n_j} \ge m \cdot B_{n_{j-1}}, \quad j \ge 2.$$

Furthermore, we obtain from (5.1) and the definition of the  $n_i$ 's:

(5.5) 
$$B_{n_j} \leq K \cdot m \cdot B_{n_{j-1}} \exp(\sqrt{\operatorname{Log}_2 B_{n_{j-1}}}), \quad j \geq 2,$$

where K is a positive constant such that  $B_n \leq K \cdot B_{n-1} \exp(\sqrt{\log_2 B_{n-1}}), n \geq 2$ and  $\sigma_1^2 = B_1 \leq K$ .

This implies immediately:

(5.6) 
$$B_{n_j} \leq (K \cdot m)^j \exp(j \sqrt{\log_2 B_{n_j}})$$

From (5.6) we easily obtain:  $B_{n_j} \leq \exp(\bar{K}j\sqrt{\log j}), j \geq 1$ , where  $\bar{K} > 0$  is an appropriate constant.

Therefore by (5.3)

(5.7) 
$$\sum_{j=1}^{\infty} P(A_{n_j}) = \infty.$$

Since the events  $A_{n_j}$ ,  $j \ge 1$  are independent by (5.4), we infer from the Borel-Cantelli lemma

$$1 = P(\overline{\lim_{j}} A_{n_j}) = P(\overline{\lim_{n}} A_n).$$

We can now prove (5.2) by the same arguments as in [14], p. 215.

To prove Corollary 5, we use standard arguments. First, we assume w.l.o.g. that the sequence  $T_n$ ,  $n \ge 1$  in Theorem 5 fulfills

(5.8) 
$$T_n = \Gamma^{1/2} \zeta(B_n), \quad n \in \mathbb{N},$$

where  $\{\zeta(t): t \ge 0\}$  is a *d*-dimensional Brownian motion on  $(\Omega_0, \mathcal{A}_0, P_0)$ . This assumption can be justified by [6], Proposition 1.4.1 or [1], Lemma A.1.

Using the same arguments as in the proof of (2.40), we obtain from Theorem 5 and (5.8)

(5.9) 
$$\sup_{0 \le t \le 1} |\hat{U}_{(n)}(t) - \Gamma^{1/2} \zeta_n(t)| \to 0 \quad \text{a.s.},$$

where  $\{\zeta_n(t): 0 \leq t \leq 1\}$  is defined as in Proposition 2 and  $\{\hat{U}_{(n)}: n \in \mathbb{N}\}$  is defined as  $\{U_{(n)}: n \in \mathbb{N}\}$  with  $\xi_k$  replaced by  $X_k, k \in \mathbb{N}$ .

Since  $P_0 \circ (\hat{U}_{(n)})_{n \in \mathbb{N}} = P \circ (U_{(n)})_{n \in \mathbb{N}}$ , we easily obtain from (5.9) and Proposition 2 the assertion.

To prove Corollary 6, we remark that

and

{
$$(2B_n \log_2 B_n)^{-1/2} \sum_{1}^n \xi_k : n \in \mathbb{N}$$
} =  $\pi_1(\{U_{(n)} : n \in \mathbb{N}\})$ 

 $\{\Gamma^{1/2} x: |x| \leq 1\} = \pi_1(\mathscr{F}_{\Gamma}), \quad \text{where } \pi_1: \ C_d[0,1] \to \mathbb{R}^d$ 

is (trivially) continuous.

# 6. Lemmas

**Lemma 1.** Let  $\xi_1, \ldots, \xi_n: \Omega \to \mathbb{R}^d$  be random vectors such that  $E[|\xi_k|^s] < \infty$  for some s > 3  $(1 \le k \le n)$ . Set  $B_n := \sum_{k=1}^{n} E[|\xi_k|^2]$ . Then we have for any given  $\overline{s} \in [3, s]$ 

(a) 
$$B_n^{-\bar{s}/2} \sum_{1}^n E[|\xi_k|^{\bar{s}}] \leq (B_n^{-s/2} \sum_{1}^n E[|\xi_k|^s])^{(\bar{s}-2)/(s-2)}.$$

If we additionally assume that  $E[\xi_k] = 0$   $(1 \le k \le n)$ , we have furthermore

(b) 
$$\lambda_n^{-\bar{s}/2} \sum_{1}^n E[|\xi_k|^{\bar{s}}] \leq \left(\frac{A_n}{\lambda_n}\right)^{\bar{s}/2} d^{(s-\bar{s})/(s-2)} \left(\lambda_n^{-s/2} \sum_{1}^n E[|\xi_k|^s]\right)^{(\bar{s}-2)/(s-2)}$$

where  $\lambda_n(\Lambda_n)$  is the smallest (largest) eigenvalue of  $\Gamma_n := \operatorname{cov}(\xi_1 + \ldots + \xi_n)$ .

*Proof.* (a) follows from Lemma 2, [12] Chap. VI (applied to  $X_k = |\xi_k|$ ,  $1 \le k \le n$ ). Note that this lemma is formulated in [12] for r.v.'s with zero means. But it can be easily seen from the proof in [12] that this lemma is also applicable in our situation.

(b) follows from (a), since  $d\lambda_n \leq B_n \leq d\Lambda_n$ .

**Lemma 2.** Let  $\xi_1, \ldots, \xi_n: \Omega \to \mathbb{R}^d$  be integrable random vectors with zero means. Set for a fixed  $\delta > 0$ :

$$\bar{\xi}_k := \xi_k \, \mathbb{1} \, \{ |\xi_k| \leq \delta \}, \qquad \tilde{\xi}_k := \bar{\xi}_k - E \, [\bar{\xi}_k], \quad 1 \leq k \leq n.$$

Then we have:

$$P\left\{\max_{1 \le k \le n} \left| \sum_{m=1}^{k} \xi_{m} - \sum_{m=1}^{k} \tilde{\xi}_{m} \right| \ge 2\delta \right\} \le \delta^{-2} \sum_{k=1}^{n} E\left[ |\xi_{k}|^{2} 1\{|\xi_{k}| > \delta\} \right].$$

*Proof.* W.l.o.g. we assume  $\delta^{-2} \sum_{k=1}^{n} E[|\xi_k|^2 \mathbb{1}\{|\xi_k| > \delta\}] \leq 1$ . In this case we obtain:

$$\sum_{k=1}^{n} |E[\bar{\xi}_{k}]| \leq \sum_{k=1}^{n} E[|\xi_{k} - \bar{\xi}_{k}|] \leq \delta^{-1} \sum_{k=1}^{n} E[|\xi_{k}|^{2} \mathbb{1}\{|\xi_{k}| > \delta\}] \leq \delta,$$

hence

$$P\left\{\max_{1\leq k\leq n}\left|\sum_{m=1}^{k}\xi_{m}-\sum_{m=1}^{k}\tilde{\xi}_{m}\right|\geq 2\delta\right\}\leq P\left\{\max_{1\leq k\leq n}\left|\sum_{m=1}^{k}\xi_{m}-\sum_{m=1}^{k}\bar{\xi}_{m}\right|\geq \delta\right\}$$
$$\leq \sum_{k=1}^{n}P\left\{\xi_{k}+\bar{\xi}_{k}\right\}\leq \delta^{-2}\sum_{k=1}^{n}E\left[|\xi_{k}|^{2}1\left\{|\xi_{k}|>\delta\right\}\right].$$

**Lemma 3.** Let  $\xi_k: \Omega \to \mathbb{R}^d$ ,  $1 \leq k \leq n$ , be independent random vectors with zero means. Let s > 2. Let  $\Lambda_n$  be the largest eigenvalue of  $\Gamma_n := \operatorname{cov}(\xi_1 + \ldots + \xi_n)$ . Then we have for t > 0:

$$P\left\{\max_{1\leq k\leq n}\left|\sum_{m=1}^{k}\xi_{m}\right|\geq t\right\}\leq \overline{c}_{1}\left(\exp\left(-\frac{\overline{c}_{2}t^{2}}{\Lambda_{n}}\right)+t^{-s}\sum_{k=1}^{n}E\left[|\xi_{k}|^{s}\right]\right)$$

with positive constants  $\bar{c}_1, \bar{c}_2$  depending only on s and d.

*Proof.* W.l.o.g. we assume  $t \ge 3\sqrt{d}\sqrt{\Lambda_n}$ .

(If  $t < 3\sqrt{d}\sqrt{A_n}$ , the estimate is trivial, since in this case  $\exp\left(-\frac{\overline{c_2}t^2}{A_n}\right) \ge \exp(-9d\overline{c_2})$ .)

Let  $\xi_k = (\xi_{k,1}, \dots, \xi_{k,d}), 1 \leq k \leq n$ . Since we have

$$P\left\{\max_{1\leq k\leq n}\left|\sum_{m=1}^{k}\xi_{m}\right|\geq t\right\}\leq \sum_{i=1}^{d}P\left\{\max_{1\leq k\leq n}\left|\sum_{m=1}^{k}\xi_{m,i}\right|\geq t/\sqrt{d}\right\},$$

we obtain from Theorem 12 ([12], Chap. III)

$$P\left\{\max_{1\leq k\leq n}\left|\sum_{m=1}^{k}\xi_{m}\right|\geq t\right\}\leq 2\sum_{i=1}^{d}P\left\{\left|\sum_{m=1}^{n}\xi_{m,i}\right|\geq t/(2\sqrt{d})\right\}.$$

The assertion follows from the last inequality by an application of Corollary 4 ([9], p. 653).

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