

On the Coupon Collector's Remainder Term

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Summary. Each element in a finite population π is assigned a “bonus value”, i.e. a real number. Elements are selected from π by simple random sampling with replacement and with equal draw probabilities. Each time we receive a “new” element, i.e. an element which has not been previously selected, we receive the corresponding bonus. Let W_n denote the bonus sum after n selections. It is well known that W_n is approximately normally distributed under mild conditions. We give a remainder term estimate of the Berry-Esseen type for this normal distribution approximation.

1. Introduction, Formulation of Main Results and Outline of Proofs

Consider a finite population $\pi = (1, 2, \dots, N)$, for which there is a “bonus value”, i.e. a real number, associated with each element. Let the bonus values be $\mathbf{a} = (a_1, a_2, \dots, a_N)$. In the sequel we will use the following notation.

$$m = \frac{1}{N} \sum_{s=1}^N a_s, \quad (1.1)$$

$$\sigma^2 = \frac{1}{N-1} \sum_{s=1}^N (a_s - m)^2, \quad (1.2)$$

$$\gamma^3 = \frac{1}{N-1} \sum_{s=1}^N |a_s - m|^3. \quad (1.3)$$

Elements are selected from π by simple random sampling with replacement and with equal draw probabilities. Each time we get a “new” element, i.e. an element which has not been previously selected, we receive the corresponding bonus. Let

$$W_n = \text{the bonus sum after } n \text{ selections.} \quad (1.4)$$

It is well known that W_n under mild conditions is approximately normally distributed, see e.g. Rosén (1969). Our main aim in this paper is to establish

the following remainder term estimate for this normal distribution approximation.

Theorem 1. *With assumptions and notation as above and with*

$$p = \frac{n}{N} \quad (1.5)$$

we have

$$\begin{aligned} \sup_{-\infty < x < \infty} \left| P(W_n \leq x) - \Phi \left(\frac{x - EW_n}{DW_n} \right) \right| &\leq (8 + 1.76 C_1) \left(\frac{\gamma}{\sigma} \right)^3 \frac{1}{\sqrt{N e^{-p}(1 - e^{-p})}} \\ &+ \min \left(\frac{13.4}{\sqrt{N e^{-p} \{1 - e^{-p}(1 + p)\}}}, 1.1 \left(\frac{m}{\sigma} \right)^2 p \right), \end{aligned} \quad (1.6)$$

where C_1 is the universal constant appearing in Theorem 2 below.

Next we present some basic steps and ingredients in the proof of Theorem 1. Let

$$Z_n = \text{the number of distinct elements from } \pi \text{ obtained in } n \text{ selections}, \quad (1.7)$$

and let further $J_1, J_2, \dots$ be the sequence of labels for the successive new elements obtained during the drawing procedure. Set

$$S_k = \sum_{j=1}^k a_{J_j}. \quad (1.8)$$

Then we have the following relation. Here and in the sequel $\mathcal{L}(X)$ denotes the distribution of the random variable X .

$$\mathcal{L}(W_n) = \mathcal{L}(S_{Z_n}). \quad (1.9)$$

The following proposition is an obvious consequence of the assumption of equal draw probabilities.

Proposition 1.

$$\mathcal{L}(S_{Z_n} | Z_n = k) = \mathcal{L}(S_k). \quad (1.10)$$

Note that the distribution of S_k is that of the sample sum in a size k sample drawn without replacement from π . We shall derive (1.6) by an appropriate application of (1.9), (1.10), the law of total probability and the following two estimates.

Theorem 2 (Höglund (1978)). *There is a universal constant C_1 such that*

$$\sup_x \left| P(S_k \leq x) - \Phi \left(\frac{x - ES_k}{DS_k} \right) \right| \leq C_1 \left(\frac{\gamma}{\sigma} \right)^3 \frac{1}{\sqrt{k \left(1 - \frac{k}{N} \right)}}. \quad (1.11)$$

Theorem 3 (Englund (1981)).

$$\frac{0.087}{(3vDZ_n)} \leq \sup_x \left| P(Z_n \leq x) - \Phi \left(\frac{x - EZ_n}{DZ_n} \right) \right| \leq \frac{10.4}{DZ_n}. \quad (1.12)$$

The course of the paper is as follows. In the next section we present some auxiliary results, and in Sect. 3 we complete the proof of Theorem 1.

To the best of our knowledge no numerical value for C_1 in Theorem 2 has been established. The method of Höglund seems to yield a very large value, but there is reason to believe that C_1 is of moderate size, i.e. about the same value as the constant in the traditional Berry-Esseen theorem for sums of independent random variables (where the current world record seems to be 0.7975, cf. van Beek (1972)).

We want to point out Theorem 4.10 in Lanke (1975), where by similar methods a remainder term estimate for the so-called Basu estimator was derived. In our notation the Basu estimator is W_n/Z_n .

We conclude this section with some remarks on the "strength" of the result in Theorem 1.

1. Note that if there are constants K_1 and K_2 such that

$$0 < K_1 \leq \frac{n}{N} \leq K_2 < \infty, \quad (1.13)$$

then the remainder term in (1.6) can be written in the form

$$C(K_1, K_2) \left(\frac{\gamma}{\sigma} \right)^3 \frac{1}{\sqrt{n}}, \quad (1.14)$$

where $C(K_1, K_2)$ is a quantity only depending on K_1 and K_2 . This is a remainder term similar to that of the classical Berry-Esseen theorem (cf. Feller (1971) p. 542).

2. If $m=0$, then the last term in (1.6) vanishes and the bound in (1.6) reduces to

$$C \left(\frac{\gamma}{\sigma} \right)^3 \frac{1}{\sqrt{N e^{-p}(1-e^{-p})}}. \quad (1.15)$$

Hence, if $m=0$ and $n/N \leq K_2 < \infty$ the bound in (1.15) is of the type $C(K_2)(\gamma/\sigma)^3/\sqrt{n}$, where $C(K_2)$ is a quantity only depending on K_2 . Note that this incorporates the case $n \leq N$.

3. We shall show below that the first term to the right in (1.6) is "necessary" by considering the following situation. Let the bonuses be $a_i=1$, $i=1, 2, \dots, N/2$, and $a_i=-1$, $i=N/2+1, \dots, N$ (we assume that N is even for simplicity). Then we have $m=0$, $\sigma \cong 1$ and $\gamma \cong 1$ and (1.6) reduces to

$$\frac{8 + 1.76 C_1}{\sqrt{N e^{-p}(1-e^{-p})}}. \quad (1.16)$$

By Čebyšev's inequality we have (note that $EW_n=0$, cf. Lemma 2.2)

$$P(-\sqrt{3} D W_n < W_n < \sqrt{3} D W_n) \geq \frac{2}{3}. \quad (1.17)$$

Since the bonuses are integer-valued, W_n is also integer-valued. The interval $(-\sqrt{3}DW_n, \sqrt{3}DW_n)$ contains at most $[2\sqrt{3}DW_n + 1]$ of the point masses in the distribution of W_n . Hence (1.17) yields that the largest point mass $p(n, N)$ satisfies

$$p(n, N) \geq \frac{\frac{2}{3}}{2\sqrt{3}DW_n + 1}. \quad (1.18)$$

From the subsequent Lemmas 2.2, 2.3 and 2.4 it can be deduced that

$$D^2 W_n = \sigma^2 N e^{-p}(1 - e^{-p})(1 + r(n, N)), \quad (1.19)$$

where $|r(n, N)| \leq 0.1$. Hence, since $\sigma \cong 1$, we have

$$p(n, N) \geq \frac{\frac{2}{3}}{2\sqrt{3.3}\sqrt{N e^{-p}(1 - e^{-p}) + 1}}. \quad (1.20)$$

Furthermore since $\Phi(x)$ is continuous we have

$$\sup_x \left| P(W_n \leq x) - \Phi\left(\frac{x - EW_n}{DW_n}\right) \right| \geq \frac{1}{2} p(n, N). \quad (1.21)$$

By combining (1.21) with (1.20) we see that the first term to the right in (1.6) is indeed “required”.

4. An interesting particular case of coupon collection is that of classical occupancy, i.e. when $a_i = 1$, $i = 1, 2, \dots, N$. By letting the bonus values be $a_i = 1 \pm \varepsilon$ where $\varepsilon \rightarrow 0$, it can fairly easily be shown that we do not “lose” the particular case of classical occupancy, i.e. Theorem 3 is in a sense “essentially” (apart from the numerical value of the constant) a particular case of Theorem 1.

5. As stated in 2 above, the term $1.1 (m/\sigma)^2 p$ has a reducing effect on the bound in (1.6) when $m = 0$, but it can also have a considerable reducing effect when $m \neq 0$. So e.g. if $m = \sigma = 1$ and $p \cong 1/\sqrt{n}$ the present bound (1.6) is of the order of magnitude $C/\sqrt{n}$, while omission of the term above would yield a bound of the order $C/n^{1/4}$.

2. Some Auxiliary Results

By well known results for sampling without replacement from a finite population we have the following result.

Lemma 2.1.

$$ES_k = km, \quad (2.1)$$

$$D^2 S_k = \sigma^2 k \left(1 - \frac{k}{N}\right). \quad (2.2)$$

From (1.9), (1.10), (2.1), (2.2) and well-known formulas for the moments of sums with randomly many terms the following formulas are readily derived.

Lemma 2.2.

$$EW_n = mEZ_n, \quad (2.3)$$

$$D^2 W_n = \sigma^2 N \frac{EZ_n}{N} \left(1 - \frac{EZ_n}{N}\right) + m^2 D^2 Z_n - \sigma^2 \frac{D^2 Z_n}{N}. \quad (2.4)$$

In proving Theorem 1 we can without loss of generality assume that

$$Ne^{-p}(1 - e^{-p}) \geq 32 \quad (2.5)$$

which can be seen by the following argument. By Jensen's inequality we have for $N \geq 2$ (the case $N=1$ is degenerate)

$$\left(\frac{\gamma}{\sigma}\right)^3 \geq \sqrt{\frac{N-1}{N}} \geq \frac{1}{\sqrt{2}}. \quad (2.6)$$

Hence Theorem 1 is trivial, in the sense that the right hand side of (1.6) exceeds 1, if (2.5) is not fulfilled. We will in the sequel use (2.5) without explicitly mentioning it each time. Next we state two results from England (1981). Note that (2.5) implies that $N \geq 128$.

Lemma 2.3. Define $r_1(n, N)$ by the relation

$$EZ_n = N(1 - e^{-p}) + r_1(n, N). \quad (2.7)$$

If $N \geq 100$, we have

$$0 \leq r_1(n, N) \leq 0.511 pe^{-p}. \quad (2.8)$$

Lemma 2.4. Define $r_2(n, N)$ by the relation

$$D^2 Z_n = Ne^{-p}(1 - e^{-p}(1 + p))(1 + r_2(n, N)). \quad (2.9)$$

If $N \geq 100$, we have

$$|r_2(n, N)| \leq \frac{6.13 pe^{-p}}{Ne^{-p}(1 - e^{-p}(1 + p))}. \quad (2.10)$$

Remark. By (2.10) and the elementary inequality $1 - e^{-p}(1 + p) \geq \frac{1}{2}(1 - e^{-p})^2$, $p \geq 0$, we obtain (using also (2.5) and $p \leq e^p - 1$, $p \geq 0$)

$$|r_2(n, N)| \leq \frac{6.13 pe^{-p}}{\frac{1}{2}Ne^{-p}(1 - e^{-p})^2} \leq 12.3 \cdot \frac{p}{e^p - 1} \cdot \frac{1}{Ne^{-p}(1 - e^{-p})} \leq 0.39. \quad \square \quad (2.11)$$

Lemma 2.5. Let X , Y , Z and Z' be random variables such that Z is independent of X , Z' is independent of Y and $\mathcal{L}(Z) = \mathcal{L}(Z')$. Then we have, for any real number a

$$\sup_x |P(aX + Z \leq x) - P(aY + Z' \leq x)| \leq \sup_x |P(X \leq x) - P(Y \leq x)|. \quad (2.12)$$

Proof. The case $a=0$ is trivial. For $a>0$ we have

$$\begin{aligned} |P(aX+Z \leq x) - P(aY+Z' \leq x)| &= \left| \int_{-\infty}^{\infty} (P(aX+z \leq x) - P(aY+z \leq x)) dF_Z(z) \right| \\ &\leq \int_{-\infty}^{\infty} \left| P\left(X \leq \frac{x-z}{a}\right) - P\left(Y \leq \frac{x-z}{a}\right) \right| dF_Z(z), \end{aligned} \quad (2.13)$$

which readily yields (2.12). The case $a<0$ can be treated quite similarly. $\square$

Next we introduce some additional notation. Let

$$\mathcal{K}(n, N) = \{k: |k - EZ_n| \leq (Ne^{-p}(1-e^{-p}))^{3/4}, \quad 1 \leq k \leq N\} \quad (2.14)$$

$$q(n, N) = [EZ_n], \quad (2.15)$$

where $[\cdot]$ denotes integer part. For notational convenience we suppress N and write $q(n)$ instead of $q(n, N)$.

Lemma 2.6. *If (2.5) holds, then*

$$q(n) \geq 0.96 N(1-e^{-p}) \quad \text{and} \quad N-q(n) \geq 0.99 Ne^{-p}. \quad (2.16)$$

Proof.

$$q(n) = N(1-e^{-p}) + (EZ_n - N(1-e^{-p})) + (q(n) - EZ_n). \quad (2.17)$$

By using (2.7), (2.8), (2.15) and (2.5) we get

$$\frac{q(n)}{N(1-e^{-p})} = 1 + \frac{r_1(n, N)}{N(1-e^{-p})} - \frac{EZ_n - q(n)}{N(1-e^{-p})} \geq 1 - \frac{1}{N(1-e^{-p})} \geq 1 - \frac{1}{32} \geq 0.96. \quad (2.18)$$

In the same manner we obtain

$$\frac{N-q(n)}{Ne^{-p}} = 1 - \frac{r_1(n, N)}{Ne^{-p}} + \frac{EZ_n - q(n)}{Ne^{-p}} \geq 1 - \frac{0.511 pe^{-p}}{Ne^{-p}} \geq 1 - \frac{0.511}{32e} \geq 0.99. \quad (2.19)$$

where we also used the inequality $pe^{-p} \leq 1/e$, $p \geq 0$. $\square$

Lemma 2.7. *If (2.5) holds and if $k \in \mathcal{K}(n, N)$, we have*

$$\frac{k}{N} \geq 0.57(1-e^{-p}) \quad \text{and} \quad 1 - \frac{k}{N} \geq 0.57e^{-p}. \quad (2.20)$$

Proof. We easily get from (2.14) that for $k \in \mathcal{K}(n, N)$ we have

$$\begin{aligned} \frac{k}{N} &\geq 1 - e^{-p} + \frac{r_1}{N} - \frac{(Ne^{-p}(1-e^{-p}))^{3/4}}{N} \geq (1-e^{-p}) \left(1 - \frac{e^{-p}}{\sqrt[4]{Ne^{-p}(1-e^{-p})}}\right) \\ &\geq (1-e^{-p}) \left(1 - \frac{1}{\sqrt[4]{32}}\right) \geq 0.57(1-e^{-p}). \end{aligned} \quad (2.21)$$

Similarly we get

$$\begin{aligned} 1 - \frac{k}{N} &\geq e^{-p} \left(1 - \frac{0.511 p e^{-p}}{N e^{-p}} - \frac{1 - e^{-p}}{4 \sqrt{N e^{-p} (1 - e^{-p})}} \right) \\ &\geq e^{-p} \left(1 - \frac{0.511}{32 e} - \frac{1}{4 \sqrt{32}} \right) \geq 0.57 e^{-p}. \quad \square \end{aligned} \quad (2.22)$$

Lemma 2.8. *If (2.5) holds, then*

$$\left| 1 - \frac{DS_k}{DS_{q(n)}} \right| \leq 1.06 \frac{1 + |k - EZ_n|}{N e^{-p} (1 - e^{-p})}. \quad (2.23)$$

Proof. By using (2.2) we get (also using $|1 - \sqrt{x}| \leq |1 - x|$, $x \geq 0$)

$$\begin{aligned} \left| 1 - \frac{DS_k}{DS_{q(n)}} \right| &= \left| 1 - \sqrt{\frac{k(N-k)}{q(n)(N-q(n))}} \right| \leq \left| 1 - \frac{\frac{k}{N} \left(1 - \frac{k}{N} \right)}{\frac{q(n)}{N} \left(1 - \frac{q(n)}{N} \right)} \right| \\ &= \frac{\left| \frac{q(n)}{N} - \frac{k}{N} \right| \left| 1 - \frac{q(n)+k}{N} \right|}{\frac{q(n)}{N} \left(1 - \frac{q(n)}{N} \right)} \leq \frac{1 + |k - EZ_n|}{N e^{-p} (1 - e^{-p})} \frac{N(1 - e^{-p})}{q(n)} \frac{N e^{-p}}{N - q(n)}. \end{aligned} \quad (2.24)$$

Application of (2.16) yields (2.23). $\square$

Lemma 2.9. *If $x > 0$ and $y > 0$, then we have for any real number z*

$$\left| \Phi\left(\frac{z}{x}\right) - \Phi\left(\frac{z}{y}\right) \right| \leq 1.25 \left| \frac{x}{y} - 1 \right|. \quad (2.25)$$

The lemma holds with the constant $1 + (2\pi e)^{-1/2} < 1.25$.

Proof. Since the left hand side of (2.25) is dominated by 1, we can without loss of generality assume that

$$\left| \frac{x}{y} - 1 \right| \leq \frac{1}{1.25}. \quad (2.26)$$

By Taylor expansion we obtain

$$\Delta = \left| \Phi\left(\frac{z}{x}\right) - \Phi\left(\frac{z}{y}\right) \right| = \left| \frac{z}{x} - \frac{z}{y} \right| \frac{1}{\sqrt{2\pi}} \exp\left(-\frac{1}{2} z^2 \left(\frac{1}{x} + \theta \left(\frac{1}{y} - \frac{1}{x}\right)\right)^2\right), \quad (2.27)$$

where $0 < \theta < 1$. By using (2.26) we obtain

$$\Delta \leq \frac{1}{\sqrt{2\pi}} \left| 1 - \frac{x}{y} \right| \cdot \frac{|z|}{x} \exp\left(-\frac{1}{2} \left(\frac{z}{x}\right)^2 \left(1 - \frac{1}{1.25}\right)^2\right), \quad (2.28)$$

which by using the simple inequality $|y| \cdot \exp(-\alpha^2 y^2) \leq 1/\alpha \sqrt{2e}$, $\alpha > 0$, yields (2.25). $\square$

3. Proof of Theorem 1

From (1.9) and (1.10) we obtain by conditioning on Z_n

$$\begin{aligned}
 & \left| P(W_n \leq x) - \Phi\left(\frac{x - EW_n}{DW_n}\right) \right| \leq P(Z_n \notin \mathcal{K}(n, N)) \\
 & + \left| \sum_{k \in \mathcal{K}(n, N)} P(S_k \leq x) P(Z_n = k) - \Phi\left(\frac{x - EW_n}{DW_n}\right) \right| \leq P(Z_n \notin \mathcal{K}(n, N)) \\
 & + \sup_{k \in \mathcal{K}(n, N)} \left| P(S_k \leq x) - \Phi\left(\frac{x - km}{DS_k}\right) \right| \\
 & + \left| \sum_{k \in \mathcal{K}(n, N)} P(Z_n = k) \cdot \left(\Phi\left(\frac{x - km}{DS_k}\right) - \Phi\left(\frac{x - km}{DS_{q(n)}}\right) \right) \right| \\
 & + P(Z_n \notin \mathcal{K}(n, N)) + \left| \sum_k \Phi\left(\frac{x - km}{DS_{q(n)}}\right) \cdot P(Z_n = k) - \Phi\left(\frac{x - mEZ_n}{\sqrt{D^2 S_{q(n)} + m^2 D^2 Z_n}}\right) \right| \\
 & + \left| \Phi\left(\frac{x - mEZ_n}{\sqrt{D^2 S_{q(n)} + m^2 D^2 Z_n}}\right) - \Phi\left(\frac{x - EW_n}{DW_n}\right) \right| \\
 & \leq 2P(Z_n \notin \mathcal{K}(n, N)) + \sup_{k \in \mathcal{K}(n, N)} \left| P(S_k \leq x) - \Phi\left(\frac{x - km}{DS_k}\right) \right| \\
 & + \sum_{k \in \mathcal{K}(n, N)} \left| \Phi\left(\frac{x - km}{DS_k}\right) - \Phi\left(\frac{x - km}{DS_{q(n)}}\right) \right| \cdot P(Z_n = k) \\
 & + \left| \sum_k \Phi\left(\frac{x - km}{DS_{q(n)}}\right) P(Z_n = k) - \Phi\left(\frac{x - mEZ_n}{\sqrt{D^2 S_{q(n)} + m^2 D^2 Z_n}}\right) \right| \\
 & + \left| \Phi\left(\frac{x - mEZ_n}{\sqrt{D^2 S_{q(n)} + m^2 D^2 Z_n}}\right) - \Phi\left(\frac{x - EW_n}{DW_n}\right) \right| \\
 & \triangleq A_1 + A_2(x) + A_3(x) + A_4(x) + A_5(x). \tag{3.1}
 \end{aligned}$$

We shall show that these five terms can be estimated as follows (C_1 denoting the universal constant in Theorem 2).

$$A_1 \leq \frac{2.78}{\sqrt{N e^{-p}(1 - e^{-p})}}, \tag{3.2}$$

$$A_2(x) \leq 1.76 C_1 \left(\frac{\gamma}{\sigma}\right)^3 \frac{1}{\sqrt{N e^{-p}(1 - e^{-p})}}, \tag{3.3}$$

$$A_3(x) \leq \frac{1.8}{\sqrt{N e^{-p}(1 - e^{-p})}}, \tag{3.4}$$

$$A_4(x) \leq \frac{13.4}{\sqrt{N e^{-p}(1 - e^{-p}(1 + p))}}, \tag{3.5}$$

$$A_4(x) \leq 1.1 \left(\frac{m}{\sigma}\right)^2 p, \tag{3.6}$$

$$A_5(x) \leq \frac{0.56}{\sqrt{N e^{-p}(1-e^{-p})}}. \quad (3.7)$$

The estimates (3.2)–(3.7) together with (3.1) and (2.6) prove Theorem 1 since $(2.78 + 1.8 + 0.56)\sqrt{2} \leq 8$. We now turn to the task of proving (3.2)–(3.7).

Proof of (3.2). By (2.14) and Čebyšev's inequality we obtain

$$P(Z_n \notin \mathcal{K}(n, N)) \leq \frac{D^2 Z_n}{(N e^{-p}(1-e^{-p}))^{3/2}} \leq \quad (3.8)$$

by using Lemma 2.4, (2.11) and the inequality $1 - e^{-p}(1+p) \leq 1 - e^{-p}$, $p \geq 0$,

$$\leq \frac{1.39 N e^{-p}(1 - e^{-p}(1+p))}{(N e^{-p}(1-e^{-p}))^{3/2}} \leq \frac{1.39}{\sqrt{N e^{-p}(1-e^{-p})}}, \quad (3.9)$$

proving (3.2) $\square$

Proof of (3.3). By Theorem 2 we have

$$\sup_{k \in \mathcal{K}(n, N)} \left| P(S_k \leq x) - \Phi\left(\frac{x - km}{DS_k}\right) \right| \leq C_1 \left(\frac{\gamma}{\sigma}\right)^3 \sup_{k \in \mathcal{K}(n, N)} \frac{1}{\sqrt{k\left(1 - \frac{k}{N}\right)}}. \quad (3.10)$$

By Lemma 2.7 we have

$$\sup_{k \in \mathcal{K}(n, N)} \frac{1}{\sqrt{k\left(1 - \frac{k}{N}\right)}} \leq \frac{1}{\sqrt{0.57 \times 0.57}} \times \frac{1}{\sqrt{N e^{-p}(1-e^{-p})}}, \quad (3.11)$$

which yields (3.3). $\square$

Proof of (3.4). Observing Lemma 2.9 and Lemma 2.8 we get for $k \in \mathcal{K}(n, N)$

$$\left| \Phi\left(\frac{x - km}{DS_k}\right) - \Phi\left(\frac{x - km}{DS_{q(n)}}\right) \right| \leq 1.25 \times 1.06 \frac{1 + |k - EZ_n|}{N e^{-p}(1-e^{-p})}, \quad (3.12)$$

yielding

$$A_3(x) \leq 1.325 \frac{1 + E|Z_n - EZ_n|}{N e^{-p}(1-e^{-p})} \leq \frac{1.325}{N e^{-p}(1-e^{-p})} + \frac{1.325 D Z_n}{N e^{-p}(1-e^{-p})}. \quad (3.13)$$

Simple use of (2.5), Lemma 2.4 and (2.11) as in (3.9) yields (3.4). $\square$

Proof of (3.5). By definition we have

$$A_4(x) = \left| \sum_{k=0}^{\infty} \Phi\left(\frac{x - km}{DS_{q(n)}}\right) P(Z_n = k) - \Phi\left(\frac{x - mEZ_n}{\sqrt{D^2 S_{q(n)} + m^2 D^2 Z_n}}\right) \right|. \quad (3.14)$$

Let U denote a standard normal random variable which is independent of Z_n . Then we have

$$\sum_{k=0}^{\infty} \Phi\left(\frac{x - km}{DS_{q(n)}}\right) P(Z_n = k) = P(mZ_n + UDS_{q(n)} \leq x). \quad (3.15)$$

If V is a normally distributed r.v. with $EV = EZ_n$ and $DV = DZ_n$, which is independent of U , we get

$$\Phi\left(\frac{x - mEZ_n}{\sqrt{D^2 S_{q(n)} + m^2 D^2 Z_n}}\right) = P(mV + UDS_{q(n)} \leq x). \quad (3.16)$$

By inserting (3.15) and (3.16) into (3.14) and using Lemma 2.5 we obtain (3.5) from Theorem 3 and Lemma 2.4 (see also the remark after Lemma 2.4). $\square$

Proof of (3.6). We easily get

$$\begin{aligned} A_4(x) \leq & \left| \sum_k \Phi\left(\frac{x - km}{DS_{q(n)}}\right) P(Z_n = k) - \Phi\left(\frac{x - mEZ_n}{DS_{q(n)}}\right) \right| \\ & + \left| \Phi\left(\frac{x - mEZ_n}{DS_{q(n)}}\right) - \Phi\left(\frac{x - mEZ_n}{\sqrt{D^2 S_{q(n)} + m^2 D^2 Z_n}}\right) \right| \triangleq A'_4(x) + A''_4(x). \end{aligned} \quad (3.17)$$

By Taylor expansion (where $0 < \theta_k < 1$) we have

$$\begin{aligned} \Phi\left(\frac{x - km}{DS_{q(n)}}\right) = & \Phi\left(\frac{x - mEZ_n}{DS_{q(n)}}\right) + \frac{m}{DS_{q(n)}} (EZ_n - k) \Phi'\left(\frac{x - mEZ_n}{DS_{q(n)}}\right) \\ & + \frac{m^2}{2D^2 S_{q(n)}} (EZ_n - k)^2 \Phi''\left(\frac{x - mEZ_n}{DS_{q(n)}} + \theta_k \left(\frac{x - km}{DS_{q(n)}} - \frac{x - mEZ_n}{DS_{q(n)}}\right)\right), \end{aligned} \quad (3.18)$$

which yields by Lemma 2.4, (2.11) and the inequality $|\Phi''(x)| \leq 1/\sqrt{2\pi e}$

$$\begin{aligned} A'_4(x) & \leq \frac{m^2}{2D^2 S_{q(n)}} D^2 Z_n \sup_x |\Phi''(x)| \\ & \leq \frac{1.39}{2\sqrt{2\pi e}} \left(\frac{m}{\sigma}\right)^2 \frac{1 - e^{-p}(1+p)}{1 - e^{-p}} \frac{Ne^{-p}}{N - q(n)} \frac{N(1 - e^{-p})}{q(n)}. \end{aligned} \quad (3.19)$$

By using the inequality $1 - e^{-p}(1+p) \leq p(1 - e^{-p})$ and (2.16) this yields

$$A'_4(x) \leq 0.18 \left(\frac{m}{\sigma}\right)^2 p. \quad (3.20)$$

In order to estimate $A''_4(x)$ we use Lemma 2.9 to get

$$A''_4(x) \leq 1.25 \left| \frac{\sqrt{D^2 S_{q(n)} + m^2 D^2 Z_n}}{DS_{q(n)}} - 1 \right| \leq 0.625 \frac{m^2 D^2 Z_n}{D^2 S_{q(n)}}, \quad (3.21)$$

where we used the inequality $|\sqrt{1+x} - 1| \leq \frac{1}{2}x$, $x \geq 0$ in the last step. Repeating the argument in (3.19) and (3.20) we see that

$$A''_4(x) \leq \frac{0.625 \times 1.39}{0.99 \times 0.96} \left(\frac{m}{\sigma}\right)^2 p \leq 0.92 \left(\frac{m}{\sigma}\right)^2 p. \quad (3.22)$$

By combining (3.22) and (3.20) we get (3.6). $\square$

Proof of (3.7). By using (2.4) we get

$$D^2 W_n = D^2 S_{q(n)} + m^2 D^2 Z_n + \sigma^2 \left(EZ_n \left(1 - \frac{EZ_n}{N} \right) - q(n) \left(1 - \frac{q(n)}{N} \right) - \frac{D^2 Z_n}{N} \right). \quad (3.23)$$

Hence we get

$$\begin{aligned} |D^2 W_n - (D^2 S_{q(n)} + m^2 D^2 Z_n)| &\leq \sigma^2 \left| EZ_n - q(n) + \frac{q^2(n) - E^2 Z_n}{N} \right| + \sigma^2 \frac{D^2 Z_n}{N} \\ &\leq \sigma^2 |EZ_n - q(n)| \cdot \left| 1 - \frac{q(n) + EZ_n}{N} \right| + 1.39 \sigma^2 \leq 2.39 \sigma^2, \end{aligned} \quad (3.24)$$

where we used Lemma 2.4, (2.11) and (2.15). Since $EW_n = mEZ_n$ (cf. (2.3)) we get by using Lemma 2.9

$$\begin{aligned} A_5(x) &= \left| \Phi \left(\frac{x - mEZ_n}{\sqrt{D^2 S_{q(n)} + m^2 D^2 Z_n}} \right) - \Phi \left(\frac{x - mEZ_n}{DW_n} \right) \right| \\ &\leq 1.25 \left| \frac{DW_n}{\sqrt{D^2 S_{q(n)} + m^2 D^2 Z_n}} - 1 \right|. \end{aligned} \quad (3.25)$$

By using the inequality $|x - 1| \leq |x^2 - 1|$, $x \geq 0$, we get from (3.25), (3.24) and (2.16)

$$A_5(x) \leq \frac{1.25 \times 2.39 \sigma^2}{D^2 S_{q(n)}} = \frac{1.25 \times 2.39}{N \frac{q(n)}{N} \left(1 - \frac{q(n)}{N} \right)} \leq \frac{3.15}{N e^{-p} (1 - e^{-p})}, \quad (3.26)$$

By using (2.5) we easily obtain (3.7) from (3.26). $\square$

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Note Added in Proof. Dr. Malcolm Quine of Univ. of Sydney informs me that he has obtained the preliminary estimate 48 of C_1 in Theorem 2.